

## HW1 DENSITY ESTIMATION

### 1 Problem 1: Bishop 2.7

|    |   |  |
|----|---|--|
| 1  | $\mu_{post} = \frac{a+m}{a+b+m+l}$  | posterior mean given by 2.20                           |
|    | $= \frac{a}{a+b+m+l} + \frac{m}{a+b+m+l}$ $= \frac{a}{(a+b+m+l)} \frac{(a+b)}{(a+b)} + \frac{m}{(a+b+m+l)} \frac{(m+l)}{(m+l)}$ $= \frac{(a+b)}{(a+b+m+l)} \frac{a}{(a+b)} + \frac{(m+l)}{(a+b+m+l)} \frac{m}{(m+l)}$ |  |
| 2  | let $\lambda = \frac{(a+b)}{(a+b+m+l)}$   | $a + b < a + b + m + l \implies 0 \leq \lambda \leq 1$ |
|    | it follows: $1 - \lambda = \frac{(a+b+m+l)}{(a+b+m+l)} - \frac{(a+b)}{(a+b+m+l)}$<br>$= \frac{m+l}{a+b+m+l}$  |  |
| 6  | $\mu_{post} = \lambda \frac{a}{(a+b)} + \frac{(m+l)}{(a+b+m+l)} \frac{m}{(m+l)}$  | substitute 5 into 4                                    |
| 9  | $\mu_{post} = \lambda \frac{a}{(a+b)} + (1 - \lambda) \frac{m}{(m+l)}$  | substitute 8 into 6                                    |
| 10 | $\mu_{prior} = \frac{a}{a+b}$   | prior mean given by 2.15                               |
| 11 | $\mu_{post} = \lambda * \mu_{prior} + (1 - \lambda) \frac{m}{(m+l)}$  | substitute 10 into 9                                   |
| 12 | $\mu_{ML} = \frac{m}{m+l}$  | maximum likelihood estimate given by 2.8               |
| 13 | $\mu_{post} = \lambda * \mu_{prior} + (1 - \lambda) \mu_{ML}$   | substitute 12 into 10                                  |

### 2 Problem 2: Bishop 2.12

The uniform distribution for a continuous variable  $x$  is defined by:

$$U(x|a, b) = \frac{1}{b - a}$$

for  $a \leq x \leq b$ .

This distribution is normalized, ie integrates to 1 as shown below

$$\begin{aligned}\int_a^b p(x) &= \int_a^b \frac{1}{b-a} \\ &= \frac{1}{b-a}(b-a) \\ &= 1\end{aligned}$$

The mean of the uniform distribution can be expressed as:

$$\mathbb{E}[U] = \frac{b+a}{2}$$

which is derived below

$$\begin{aligned}\mathbb{E}[f] &= \int p(x)f(x)dx \\ p(x) &= U(x) = \frac{1}{b-a} \\ f(x) &= x \\ \mathbb{E}[f] &= \int_a^b \frac{1}{b-a} x dx \\ &= \frac{1}{b-a} \int_a^b x dx \\ &= \frac{1}{b-a} \frac{b^2}{2} - \frac{a^2}{2} \\ &= \frac{1}{b-a} \frac{b^2-a^2}{2} \\ &= \frac{(b-a)(b+a)}{2(b-a)} \\ &= \frac{(b+a)}{2}\end{aligned}$$

The variance of the uniform distribution can be expressed as:

$$var[U] = \frac{(b-a)^2}{12}$$

which is derived below

$$\begin{aligned}
\text{var}[f] &= \mathbb{E}[(f(x) - \mathbb{E}[f(x)])^2] \\
&= \mathbb{E}\left[\left(x - \frac{(b+a)}{2}\right)^2\right] \\
&= \mathbb{E}\left[x^2 - 2x\frac{(b+a)}{2} + \frac{(b+a)^2}{4}\right] \\
&= \int_a^b \frac{1}{b-a} \left(x^2 - 2x\frac{(b+a)}{2} + \frac{(b+a)^2}{4}\right) \\
&= \frac{1}{b-a} \left(\int_a^b x^2 - ((b+a) \int_a^b x) + \int_a^b \frac{(b+a)^2}{4}\right) \\
&= \frac{1}{b-a} \left(\frac{b^3-a^3}{3} - \frac{(b+a)(b^2-a^2)}{2} + \frac{(b+a)^2(b-a)}{4}\right) \\
&= \frac{b^3-a^3}{3(b-a)} - \frac{(b+a)(b^2-a^2)}{2(b-a)} + \frac{(b+a)^2}{4} \\
&= \frac{(b^2+ab+a^2)(b-a)}{3(b-a)} - \frac{(b+a)(b+a)(b-a)}{2(b-a)} + \frac{(b+a)^2}{4} \\
&= \frac{(b^2+ab+a^2)}{3} - \frac{(b+a)^2}{2} + \frac{(b+a)^2}{4} \\
&= \frac{(4b^2+4ab+4a^2)}{12} - \frac{6(b+a)^2}{12} + \frac{3(b+a)^2}{12} \\
&= \frac{(4b^2+4ab+4a^2)}{12} - \frac{3(b+a)^2}{12} \\
&= \frac{(4b^2+4ab+4a^2)}{12} - \frac{3b^2+6ba+3a^2}{12} \\
&= \frac{(b^2-2ab+a^2)}{12} \\
&= \frac{(b-a)^2}{12}
\end{aligned}$$

### 3 Problem 3: Poisson

#### 3.1 3a

Using the definition of the Poisson distribution show that the sum of probabilities of all events is 1. (Hint: use the definition of  $e$  in terms of a sum).

By definition, the Poisson distribution is:

$$p(x|\lambda) = \frac{e^{-\lambda}\lambda^x}{x!}$$

The sum of probabilities for all  $x$  is then:

$$\sum_{x=0}^{\infty} \frac{e^{-\lambda} \lambda^x}{x!}$$

$$= \frac{1}{e^\lambda} \sum_{x=0}^{\infty} \frac{\lambda^x}{x!}$$

By definition of power series,

$$\sum_{x=0}^{\infty} \frac{\lambda^x}{x!} = e^\lambda$$

And as such

$$\sum_{x=0}^{\infty} \frac{e^{-\lambda} \lambda^x}{x!}$$

$$= \frac{1}{e^\lambda} e^\lambda = 1$$

### 3.2 3b

Derive the mean of the Poisson distribution.

The expected value of a (discrete) function given a distribution is given as:

$$\mathbb{E}[f] = \sum_{x=0}^{\infty} p(x) f(x)$$

For  $f(x) = x$  and the poisson distribution:

$$\mathbb{E}[f] = \frac{e^{-\lambda} \lambda^x}{x!}$$

The expected value is:

$$\mathbb{E}[f] = \sum_{x=0}^{\infty} \frac{e^{-\lambda} \lambda^x}{x!} x$$

Which can be manipulated to:

$$\begin{aligned}
 \mathbb{E}[f] &= \frac{1}{e^\lambda} \sum_{x=0}^{\infty} \frac{\lambda^x}{x!} x \\
 &= \frac{1}{e^\lambda} \sum_{x=0}^{\infty} \frac{\lambda^{x-1} \lambda x}{(x-1)! x} \\
 &= \frac{\lambda}{e^\lambda} \sum_{x=1}^{\infty} \frac{\lambda^{x-1}}{(x-1)!} \\
 &= \frac{\lambda}{e^\lambda} e^\lambda \\
 &= \lambda
 \end{aligned}$$

### 3.3 3c

$$\lambda_{ML} = \frac{1}{n} \sum_{i=1}^n x_i$$

### 3.4 3d

The posterior distribution poisson given a prior gamma is:

$$p_{post}(x|\lambda, a, b) = \frac{e^{-\lambda} \lambda^x}{x!} * \frac{\lambda^{a-1} e^{-\frac{\lambda}{b}}}{b^a \Gamma(a)}$$

$$= \frac{\lambda^{a-1+x} e^{-\frac{\lambda}{b} - \lambda}}{b^a \Gamma(a) x!}$$

Which is of the same form as the gamma distribution, except for the  $x!$  in the denominator which I have no idea what to do with.

### 3.5 3e

See figure 1.

### 3.6 3f

$$\lambda_{ML} = \frac{calls}{time}$$

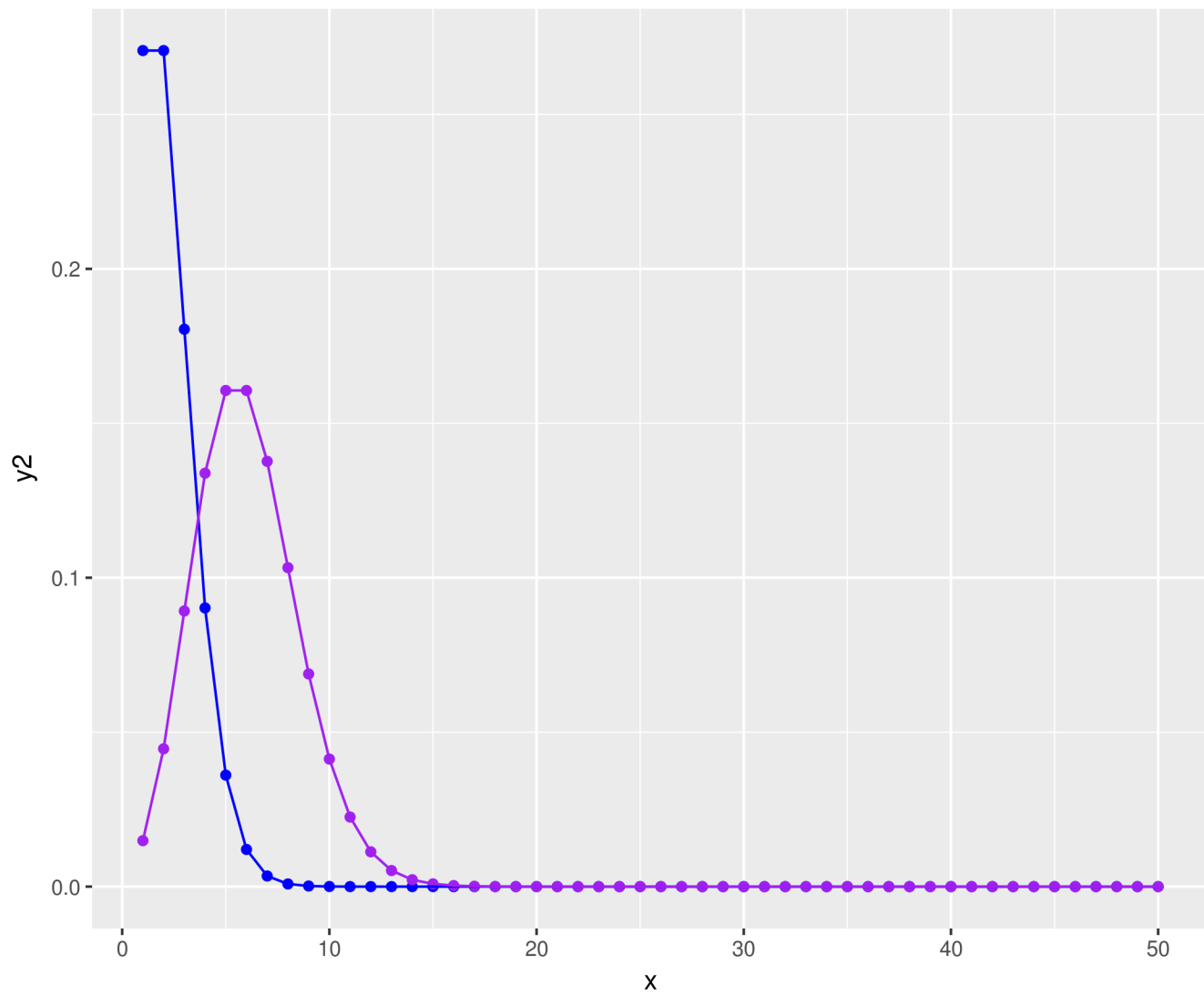


Figure 1: 3e: Combined plot showing difference between  $\lambda = 2$  and  $\lambda = 6$

### 3.7 3g

See figure 2.

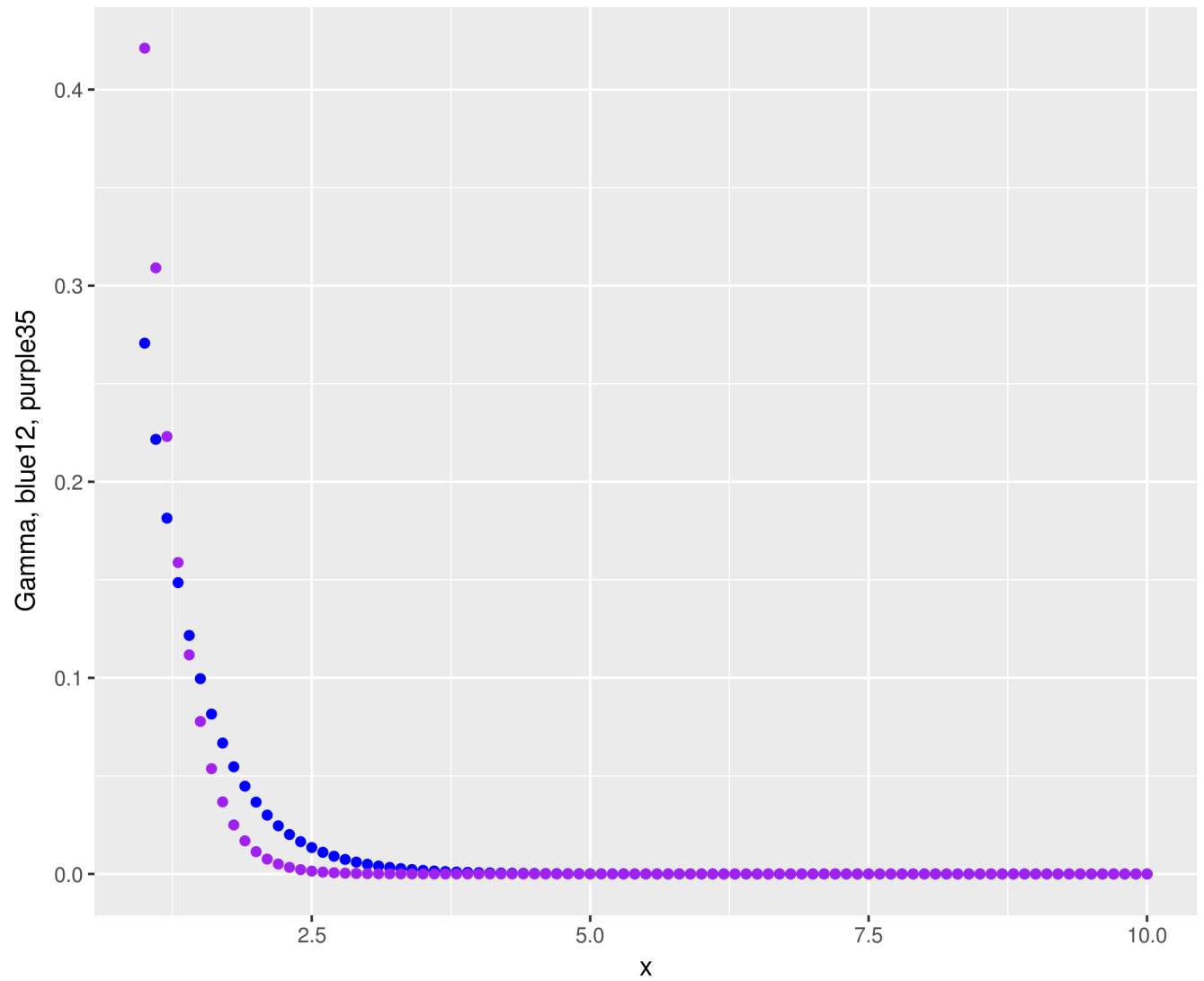


Figure 2: 3g. Plot of gamma distributions. Blue points show when  $a=1$   $b=2$ , and purple points show when  $a=3$  and  $b=5$ .

### 3.8 3h

## 4 Problem 4: Exponential Family

The Poisson distribution:

$$p(x|\lambda) = \frac{e^{-\lambda} \lambda^x}{x!}$$

Can be manipulated as:

$$\begin{aligned} p(x|\lambda) &= \frac{e^{-\lambda} \exp(\log(\lambda^x))}{x!} \\ &= \frac{e^{-\lambda} \exp(x \log(\lambda))}{x!} \end{aligned}$$

Which can be represented by the general form of the exponential family:

$$f(x|\eta) = \frac{1}{Z(\eta)} h(x) \exp[\eta^T t(x)]$$

where

$$\eta = \log(\lambda)$$

$$Z(\eta) = \frac{1}{e^{-\lambda}}$$

$$t(x) = x$$

$$h(x) = \frac{1}{x!}$$