# CS559 Lecture 5: Density Estimation (3)

Reading: Bishop Book, Chapter 2

#### **Outline**

#### **Outline:**

- Density estimation:
  - Maximum likelihood (ML)
  - Bayesian parameter estimates
  - MAP
- Bernoulli distribution.
- Binomial distribution
- Multinomial distribution
- Normal distribution
- Exponential family

#### **Exponential family:**

 all probability mass / density functions that can be written in the exponential normal form

$$f(\mathbf{x} \mid \mathbf{\eta}) = \frac{1}{Z(\mathbf{\eta})} h(\mathbf{x}) \exp \left[ \mathbf{\eta}^T t(\mathbf{x}) \right]$$

- η a vector of natural (or canonical) parameters
- $t(\mathbf{x})$  a function referred to as a sufficient statistic
- $h(\mathbf{x})$  a function of x (it is less important)
- $Z(\eta)$  a normalization constant (a partition function)  $Z(\eta) = \int h(\mathbf{x}) \exp \left\{ \eta^T t(\mathbf{x}) \right\} d\mathbf{x}$

Other common form:

$$f(\mathbf{x} \mid \mathbf{\eta}) = h(\mathbf{x}) \exp \left[ \mathbf{\eta}^T t(\mathbf{x}) - A(\mathbf{\eta}) \right]$$
 log  $Z(\mathbf{\eta}) = A(\mathbf{\eta})$ 

#### Bernoulli distribution

$$p(x \mid \pi) = \pi^{x} (1 - \pi)^{1 - x}$$

$$= \exp\left\{\log\left(\frac{\pi}{1 - \pi}\right)x + \log(1 - \pi)\right\}$$

$$= \exp\left\{\log(1 - \pi)\right\} \exp\left\{\log\left(\frac{\pi}{1 - \pi}\right)x\right\}$$

#### Exponential family

$$f(\mathbf{x} \mid \mathbf{\eta}) = \frac{1}{Z(\mathbf{\eta})} h(\mathbf{x}) \exp \left[ \mathbf{\eta}^T t(\mathbf{x}) \right]$$

#### Parameters

$$\eta = ?$$
  $t(\mathbf{x}) = ?$ 

$$Z(\mathbf{\eta}) = ?$$
  $h(\mathbf{x}) = ?$ 

#### Bernoulli distribution

$$p(x \mid \pi) = \pi^{x} (1 - \pi)^{1 - x}$$

$$= \exp\left\{\log\left(\frac{\pi}{1 - \pi}\right)x + \log(1 - \pi)\right\}$$

$$= \exp\left\{\log(1 - \pi)\right\} \exp\left\{\log\left(\frac{\pi}{1 - \pi}\right)x\right\}$$

Exponential family

$$f(\mathbf{x} \mid \mathbf{\eta}) = \frac{1}{Z(\mathbf{\eta})} h(\mathbf{x}) \exp \left[ \mathbf{\eta}^T t(\mathbf{x}) \right]$$

Parameters

$$\mathbf{\eta} = \log \frac{\pi}{1 - \pi} \quad \text{(note } \pi = \frac{1}{1 + e^{-\eta}} \text{)}$$

$$t(\mathbf{x}) = x$$

$$Z(\mathbf{\eta}) = \frac{1}{1 - \pi} = 1 + e^{\eta}$$

$$h(\mathbf{x}) = 1$$

Univariate Gaussian distribution

$$p(x \mid \mu, \sigma) = \frac{1}{\sigma \sqrt{2\pi}} \exp\left[-\frac{1}{2\sigma^2} (x - \mu)^2\right]$$
$$= \frac{1}{2\pi} \exp\left(-\frac{\mu}{2\sigma^2} - \log \sigma\right) \exp\left\{\frac{\mu}{\sigma^2} x - \frac{1}{2\sigma^2} x^2\right\}$$

• Exponential family  $f(\mathbf{x} \mid \mathbf{\eta}) = \frac{1}{Z(\mathbf{n})} h(x) \exp \left[ \eta^T t(x) \right]$ 

Parameters

$$\mathbf{\eta} = ? \qquad \qquad t(\mathbf{x}) = ?$$

$$Z(\mathbf{\eta}) = ?$$
  $h(\mathbf{x}) = ?$ 

Univariate Gaussian distribution

$$p(x \mid \mu, \sigma) = \frac{1}{\sigma \sqrt{2\pi}} \exp\left[-\frac{1}{2\sigma^2} (x - \mu)^2\right]$$
$$= \frac{1}{2\pi} \exp\left(-\frac{\mu}{2\sigma^2} - \log \sigma\right) \exp\left\{\frac{\mu}{\sigma^2} x - \frac{1}{2\sigma^2} x^2\right\}$$

- Exponential family  $f(\mathbf{x} \mid \mathbf{\eta}) = \frac{1}{Z(\mathbf{\eta})} h(x) \exp \left[ \eta^T t(x) \right]$
- Parameters

$$\mathbf{\eta} = \begin{bmatrix} \mu / 2\sigma^2 \\ -1 / 2\sigma^2 \end{bmatrix} \qquad t(\mathbf{x}) = \begin{bmatrix} x \\ x^2 \end{bmatrix}$$

$$Z(\mathbf{\eta}) = \exp\left\{\frac{\mu}{2\sigma^2} + \log\sigma\right\} = \exp\left\{-\frac{\eta_1^2}{4\eta_2} - \frac{1}{2}\log(-2\eta_2)\right\}$$

$$h(\mathbf{x}) = 1/\sqrt{2\pi}$$

For iid samples, the likelihood of data is

$$P(D \mid \mathbf{\eta}) = \prod_{i=1}^{n} p(\mathbf{x}_{i} \mid \mathbf{\eta}) = \prod_{i=1}^{n} h(\mathbf{x}_{i}) \exp \left[\mathbf{\eta}^{T} t(\mathbf{x}_{i}) - A(\mathbf{\eta})\right]$$
$$= \left[\prod_{i=1}^{n} h(\mathbf{x}_{i})\right] \exp \left[\sum_{i=1}^{n} \mathbf{\eta}^{T} t(\mathbf{x}_{i}) - A(\mathbf{\eta})\right]$$
$$= \left[\prod_{i=1}^{n} h(\mathbf{x}_{i})\right] \exp \left[\mathbf{\eta}^{T} \left(\sum_{i=1}^{n} t(\mathbf{x}_{i})\right) - nA(\mathbf{\eta})\right]$$

#### Important:

 the dimensionality of the sufficient statistic remains the same for different sample sizes (that is, different number of examples in D)

The log likelihood of data is

$$l(D, \mathbf{\eta}) = \log \left[ \prod_{i=1}^{n} h(\mathbf{x}_{i}) \right] \exp \left[ \mathbf{\eta}^{T} \left( \sum_{i=1}^{n} t(\mathbf{x}_{i}) \right) - nA(\mathbf{\eta}) \right]$$
$$= \log \left[ \prod_{i=1}^{n} h(\mathbf{x}_{i}) \right] + \left[ \mathbf{\eta}^{T} \left( \sum_{i=1}^{n} t(\mathbf{x}_{i}) \right) - nA(\mathbf{\eta}) \right]$$

Optimizing the loglikelihood

$$\nabla_{\mathbf{\eta}} l(D, \mathbf{\eta}) = \left(\sum_{i=1}^{n} t(\mathbf{x}_{i})\right) - n \nabla_{\mathbf{\eta}} A(\mathbf{\eta}) = \mathbf{0}$$

For the ML estimate it must hold

$$\nabla_{\mathbf{\eta}} A(\mathbf{\eta}) = \frac{1}{n} \left( \sum_{i=1}^{n} t(\mathbf{x}_i) \right)$$

Rewritting the gradient:

$$\nabla_{\eta} A(\eta) = \nabla_{\eta} \log Z(\eta) = \nabla_{\eta} \log \int h(\mathbf{x}) \exp \left\{ \eta^{T} t(\mathbf{x}) \right\} d\mathbf{x}$$

$$\nabla_{\eta} A(\eta) = \frac{\int t(\mathbf{x}) h(\mathbf{x}) \exp \left\{ \eta^{T} t(\mathbf{x}) \right\} d\mathbf{x}}{\int h(\mathbf{x}) \exp \left\{ \eta^{T} t(\mathbf{x}) \right\} d\mathbf{x}}$$

$$\nabla_{\eta} A(\eta) = \int t(\mathbf{x}) h(\mathbf{x}) \exp \left\{ \eta^{T} t(\mathbf{x}) - A(\eta) \right\} d\mathbf{x}$$

$$\nabla_{\eta} A(\eta) = E(t(\mathbf{x}))$$

- Result:  $E(t(\mathbf{x})) = \frac{1}{n} \left( \sum_{i=1}^{n} t(\mathbf{x}_i) \right)$
- For the ML estimate the parameters η should be adjusted such that the expectation of the statistic t(x) is equal to the observed sample statistics

#### Moments of the distribution

- For the exponential family
  - The k-th moment of the statistic corresponds to the k-th derivative of  $A(\eta)$
  - If x is a component of t(x) then we get the moments of the distribution by differentiating its corresponding natural parameter
- Example: Bernoulli  $p(x \mid \pi) = \exp\left\{\log\left(\frac{\pi}{1-\pi}\right)x + \log(1-\pi)\right\}$   $A(\eta) = \log\frac{1}{1-\pi} = \log(1+e^{\eta})$
- Derivatives:

$$\frac{\partial A(\mathbf{\eta})}{\partial \eta} = \frac{\partial}{\partial \eta} \log(1 + e^{\eta}) = \frac{e^{\eta}}{(1 + e^{\eta})} = \frac{1}{(1 + e^{-\eta})} = \pi$$
$$\frac{\partial A(\mathbf{\eta})}{\partial \eta^2} = \frac{\partial}{\partial \eta} \frac{1}{(1 + e^{-\eta})} = \pi (1 - \pi)$$

## **Conjugate priors**

For any member of the exponential family

$$f(\mathbf{x} \mid \mathbf{\eta}) = \frac{1}{Z(\mathbf{\eta})} h(\mathbf{x}) \exp \left[ \mathbf{\eta}^T \mathbf{t}(\mathbf{x}) \right]$$

there exists a prior:

$$p(\mathbf{\eta} \mid \mathbf{\chi}, \mathbf{v}) = u(\mathbf{\chi}, \mathbf{v}) g(\mathbf{\eta})^{\mathbf{v}} \exp \left[ \mathbf{v} \, \mathbf{\eta}^{\mathrm{T}} \mathbf{\chi} \right]$$

Such that for n examples, the posterior is

$$p(\mathbf{\eta} \mid D, \mathbf{\chi}, \mathbf{\nu}) \propto g(\mathbf{\eta})^{\nu+n} \exp \left[ \mathbf{\eta}^T \left( \left[ \sum_{i=1}^n \mathbf{t}(x_i) \right] + \nu \mathbf{\chi} \right) \right]$$

Note that:

$$P(D \mid \mathbf{\eta}) = \left(\frac{1}{Z(\mathbf{\eta})}\right)^n \left[\prod_{i=1}^n h(\mathbf{x}_i)\right] \exp\left[\mathbf{\eta}^T \left(\sum_{i=1}^n t(\mathbf{x}_i)\right)\right]$$

## **Conjugate priors**

For any member of the exponential family

$$f(\mathbf{x} \mid \mathbf{\eta}) = \frac{1}{Z(\mathbf{\eta})} h(\mathbf{x}) \exp \left[ \mathbf{\eta}^T \mathbf{t}(\mathbf{x}) \right]$$

there exists a prior:

$$p(\mathbf{\eta} \mid \mathbf{\chi}, \mathbf{v}) = u(\mathbf{\chi}, \mathbf{v}) g(\mathbf{\eta})^{\mathbf{v}} \exp \left[ \mathbf{v} \, \mathbf{\eta}^{\mathrm{T}} \mathbf{\chi} \right]$$

Such that for n examples, the posterior is

$$p(\mathbf{\eta} \mid D, \mathbf{\chi}, \mathbf{\nu}) \propto g(\mathbf{\eta})^{\nu+n} \exp \left[ \mathbf{\eta}^T \left( \left[ \sum_{i=1}^n \mathbf{t}(x_i) \right] + \nu \mathbf{\chi} \right) \right]$$

Prior corresponds to V observations with value  $\chi$ .

$$P(D \mid \mathbf{\eta}) = \left\lfloor \frac{1}{Z(\mathbf{\eta})} \right\rfloor \left\lfloor \prod_{i=1}^{n} h(\mathbf{x}_i) \right\rfloor \exp \left\lfloor \mathbf{\eta}^{T} \left\lfloor \sum_{i=1}^{n} t(\mathbf{x}_i) \right\rfloor \right\rfloor$$

## **Nonparametric Methods**

#### Parametric distribution models are:

- restricted to specific forms, which may not always be suitable;
- Example: modelling a multimodal distribution with a single, unimodal model.

#### Nonparametric approaches:

 make few assumptions about the overall shape of the distribution being modelled.

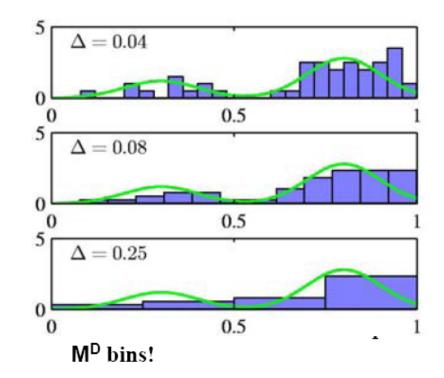
## **Nonparametric Methods**

#### **Histogram methods:**

partition the data space into distinct bins with widths  $\Delta_i$  and count the number of observations,  $n_i$ , in each bin.

$$p_i = \frac{n_i}{N\Delta_i}$$

- Often, the same width is used for all bins,  $\Delta_i = \Delta$ .
- $\Delta$  acts as a smoothing parameter.



## **Nonparametric Methods**

 Assume observations drawn from a density p(x) and consider a small region R containing x such that

$$P = \int_{R} p(x) dx$$

 The probability that K out of N observations lie inside R is Bin(K,N,P) and if N is large

$$K \cong NP$$

If the volume of R, V, is sufficiently small, p(x) is approximately constant over R and

$$P \cong p(x)V$$

Thus

$$p(x) = \frac{P}{V}$$

$$p(x) = \frac{K}{NV}$$

## Nonparametric Methods: kernel methods

#### **Kernel Density Estimation:**

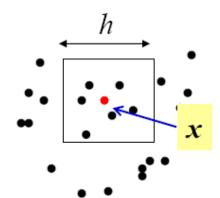
Fix V, estimate K from the data. Let R be a hypercube centred on **x** and define the kernel function (Parzen window)

$$k\left(\frac{x-x_n}{h}\right) = \begin{cases} 1 & |(x_i-x_{ni})|/h \le 1/2 \\ 0 & otherwise \end{cases} i = 1, \dots D$$

- It follows that

• and hence 
$$K = \sum_{n=1}^{N} k \left( \frac{x - x_n}{h} \right)$$

$$p(x) = \frac{1}{N} \sum_{n=1}^{N} \frac{1}{h^{D}} k \left( \frac{x - x_{n}}{h} \right)$$



## Nonparametric Methods: smooth kernels

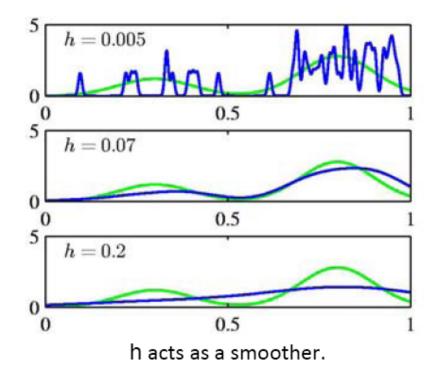
To avoid discontinuities in p(x) because of sharp boundaries use a smooth kernel, e.g. a Gaussian

$$p(\mathbf{x}) = rac{1}{N} \sum_{n=1}^{N} rac{1}{(2\pi h^2)^{D/2}} \ \exp\left\{-rac{\|\mathbf{x} - \mathbf{x}_n\|^2}{2h^2}
ight\}$$

· Any kernel such that

$$k(\mathbf{u}) \geqslant 0,$$

$$\int k(\mathbf{u}) d\mathbf{u} = 1$$

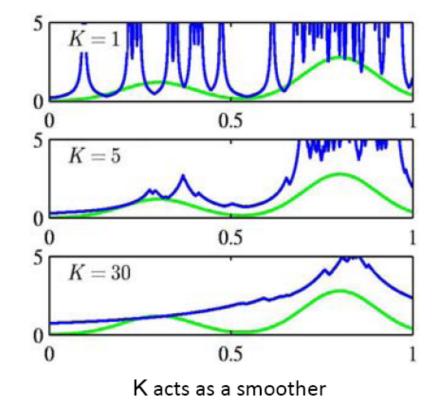


## Nonparametric Methods: kNN estimation

## **Nearest Neighbour Density Estimation:**

fix K, estimate V from the data. Consider a hyper-sphere centred on X and let it grow to a volume, V\*, that includes K of the given N data points. Then

$$p(\mathbf{x}) \simeq \frac{K}{NV^{\star}}.$$



## Nonparametric vs Parametric Methods

#### Nonparametric models:

- More flexibility no density model is needed
- But require storing the entire dataset
- and the computation is performed with all data examples.

#### Parametric models:

- Once fitted, only parameters need to be stored
- They are much more efficient in terms of computation
- But the model needs to be picked in advance

## **K-Nearest-Neighbours for Classification**

• Given a data set with  $N_k$  data points from class  $C_k$  and  $\sum_k N_k = N$ , we have

$$p(\mathbf{x}) = \frac{K}{NV}$$

and correspondingly

$$p(\mathbf{x}|\mathcal{C}_k) = rac{K_k}{N_k V}.$$

• Since  $p(C_k) = N_k/N$ , Bayes' theorem gives

$$p(\mathcal{C}_k|\mathbf{x}) = rac{p(\mathbf{x}|\mathcal{C}_k)p(\mathcal{C}_k)}{p(\mathbf{x})} = rac{K_k}{K}.$$

## **K-Nearest-Neighbours for Classification**

