

QUANTITATIVE ECONOMICS with Julia

Thomas Sargent and John Stachurski

December 16, 2016

CONTENTS

1 Programming in Julia	7
1.1 Setting up Your Julia Environment	7
1.2 An Introductory Example	22
1.3 Julia Essentials	30
1.4 Vectors, Arrays and Matrices	44
1.5 Types, Methods and Performance	58
1.6 Plotting in Julia	72
1.7 Useful Libraries	85
2 Introductory Applications	97
2.1 Linear Algebra	97
2.2 Finite Markov Chains	113
2.3 Orthogonal Projection and its Applications	134
2.4 Shortest Paths	144
2.5 The McCall Job Search Model	148
2.6 Schelling's Segregation Model	157
2.7 LLN and CLT	162
2.8 Linear State Space Models	175
2.9 A First Look at the Kalman Filter	198
2.10 Uncertainty Traps	210
2.11 A Simple Optimal Growth Model	217
2.12 LQ Dynamic Programming Problems	234
2.13 Discrete Dynamic Programming	260
2.14 Rational Expectations Equilibrium	273
2.15 An Introduction to Asset Pricing	282
2.16 The Permanent Income Model	297
3 Advanced Applications	313
3.1 Continuous State Markov Chains	313
3.2 The Lucas Asset Pricing Model	328
3.3 The Aiyagari Model	337
3.4 Modeling Career Choice	346
3.5 On-the-Job Search	355
3.6 Search with Offer Distribution Unknown	365
3.7 Optimal Savings	377

3.8 Robustness	390
3.9 Covariance Stationary Processes	413
3.10 Estimation of Spectra	429
3.11 Optimal Taxation in an LQ Economy	442
3.12 History Dependent Public Policies	459
3.13 Default Risk and Income Fluctuations	481
References	497

Note: You are currently viewing an automatically generated PDF version of our on-line lectures, which are located at

<http://quant-econ.net>

Please visit the website for more information on the aims and scope of the lectures and the two language options (Julia or Python). This PDF is generated from a set of

source files that are oriented towards the website and to HTML output. As a result, the presentation quality can be less consistent than the website.

CHAPTER
ONE

PROGRAMMING IN JULIA

This first part of the course provides a relatively fast-paced introduction to the Julia programming language

Setting up Your Julia Environment

Contents

- *Setting up Your Julia Environment*
 - *Overview*
 - *First Steps*
 - *Jupyter*
 - *QuantEcon*
 - *Exercises*

Overview

In this lecture we will cover how to get up and running with Julia

Topics:

1. Installation
2. Interactive Julia sessions
3. Running sample programs
4. Installation of libraries, including the Julia code that underpins these lectures

First Steps

Installation The first thing you will want to do is install Julia

The best option is probably to install the current release from the [download page](#)

- Read through any download and installation instructions specific to your OS on that page

- Unless you have good reason to do otherwise, choose the current release rather than nightly build and the platform specific binary rather than source

Assuming there were no problems, you should now be able to start Julia either by

- navigating to Julia through your menus or desktop icons (Windows, OSX), or
- opening a terminal and typing `julia` (Linux)

Either way you should now be looking at something like this (modulo your operating system — this is a Linux machine)

```

Terminal
+402 15:10
+402 15:11 (@godzilla3)

>>> julia
A fresh approach to technical computing
Documentation: http://docs.julialang.org
Type "?help" for help.

Version 0.4.2 (2015-12-06 21:47 UTC)
Official http://julialang.org release
x86_64-linux-gnu

julia>

```

The program that's running here is called the Julia REPL (Read Eval Print Loop) or Julia interpreter

Let's try some basic commands:

```

Terminal
+402 15:10
+402 15:11 (@godzilla3)

>>> julia
A fresh approach to technical computing
Documentation: http://docs.julialang.org
Type "?help" for help.

Version 0.4.2 (2015-12-06 21:47 UTC)
Official http://julialang.org release
x86_64-linux-gnu

julia> x = 10
10

julia> 2 * x
20

julia> 2x
20

julia>

```

The Julia interpreter has the kind of nice features you expect from a modern REPL

For example,

- Pushing the up arrow key retrieves the previously typed command
- If you type ? the prompt will change to help?> and give you access to online documentation

```

Terminal
A fresh approach to technical computing
Documentation: http://docs.julialang.org
Type "?help" for help.

Version 0.4.2 (2015-12-06 21:47 UTC)
Official http://julialang.org release
x86_64-linux-gnu

julia> x = 10
10

julia> 2 * x
20

julia> 2x
20

help?> linspace
search: linspace LinSpace

linspace(start, stop, n=100)

Construct a range of n linearly spaced elements from start to stop.

julia>

```

You can also type ; to get a shell prompt, at which you can enter shell commands

```

Terminal
Version 0.4.2 (2015-12-06 21:47 UTC)
Official http://julialang.org release
x86_64-linux-gnu

julia> x = 10
10

julia> 2 * x
20

julia> 2x
20

help?> linspace
search: linspace LinSpace

linspace(start, stop, n=100)

Construct a range of n linearly spaced elements from start to stop.

shell> ls
anaconda3 bin Downloads sync_dir temp_dir Webcam
backups Desktop Music temp versioned_dotfiles terminator.png

julia>

```

(Here ls is a UNIX style command that lists directory contents — your shell commands depend on your operating system)

From now on instead of showing terminal images we'll show interactions with the interpreter as follows

```
julia> x = 10
10

julia> 2 * x
20
```

Installing Packages Julia includes many useful tools in the base installation

However, you'll quickly find that you also have need for at least some of the many external Julia code libraries

Fortunately these are very easy to install using Julia's excellent package management system

For example, let's install `DataFrames`, which provides useful functions and data types for manipulating data sets

```
julia> Pkg.add("DataFrames")
```

Assuming you have a working Internet connection this should install the `DataFrames` package

If you now type `Pkg.status()` you'll see `DataFrames` and its version number

To pull the functionality from `DataFrames` into the current session we type using `DataFrames`

```
julia> using DataFrames
```

Now let's use one of its functions to create a data frame object (something like an R data frame, or a spreadsheet)

```
julia> df = DataFrame(x1=[1, 2], x2=["foo", "bar"])
2x2 DataFrame
| Row | x1 | x2    |
| ---- | --- | --- |
| 1   | 1  | "foo" |
| 2   | 2  | "bar" |
```

One quick point before we move on: Running

```
julia> Pkg.update()
```

will update your installed packages and also update local information on the set of available packages

Running Julia Scripts Let's say you have a file on your machine containing Julia code (a "script" file) and you want to run it

How would you do that?

Let's look at a "primitive" method for running Julia code from inside the Julia REPL

(This is actually the way that we run most of our Julia code day to day)

Suppose we have a Julia script called `test_script.jl` that we wish to run, with contents

```
# filename: test_script.jl

for i in 1:3
    println("i = $i")
end
```

Suppose that this file exists as a plain text file somewhere on your computer

If that somewhere happens to be the **present working directory** for your REPL then you are good to go

- The present working directory for your REPL can be found in a Julia session by typing `pwd()`

Then you can run it directly from within Julia by typing `include("test_script.jl")`

To try this,

1. find out what your present working directory by typing `pwd()`
2. copy the code above and paste it into a text editor (e.g., Notepad,TextEdit,TextMate)
3. Save it in the present working directory with name `test_script.jl`
4. Now go back to Julia and type `include("test_script.jl")`

Here's how it looks on our machine, when `test_script.jl` sits in directory `/home/john/temp`

(Paths to files will look different on different operating systems)

```
julia> pwd()
"/home/john/temp"

julia> include("test_script.jl")  # exists in /home/john/temp
i = 1
i = 2
i = 3
```

What if the file *isn't* in your present working directory?

Two simple options are

1. copy the file to your present working directory, or
2. change Julia's present working directory to the location of the script

For example:

```
julia> cd("/home/john/temp")  # paths look different on Windows or Mac
```

Now run using `include("test_script.jl")` as before

Editing Julia Scripts Hopefully you can now run Julia scripts

You also need to know how to edit them

Text Editors Some people like [IDEs](#) for coding but we usually use plain text editors

The beauty of text editors is that if you master one of them, you can use it for every coding task you come across, regardless of the language

At a minimum, a text editor for coding should provide

- syntax highlighting for the languages you want to work with
- automatic indentation
- efficient text manipulation (search and replace, copy and paste, etc.)

There are many text editors that speak Julia, and a lot of them are free

Suggestions:

- [Atom](#) is a popular open source next generation text editor
- [Sublime Text](#) is a modern, popular and highly regarded text editor with a relatively moderate learning curve (not free but trial period is unlimited)
- [Emacs](#) is a high quality free editor with a sharper learning curve

Finally, if you want an outstanding free text editor and don't mind a seemingly vertical learning curve plus long days of pain and suffering while all your neural pathways are rewired, try [Vim](#)

IDEs [IDEs](#) (Integrated Development Environments) combine an interpreter and text editing facilities in the one application

For Julia one nice option is [Juno](#)

Alternatively there's Jupyter, which is a little bit different again but has some great features that we now discuss

Jupyter

To work with Julia in a scientific context we need at a minimum

1. An environment for editing and running Julia code
2. The ability to generate figures and graphics

One option that provides these features is [Jupyter](#)

As a bonus, Jupyter also provides

- nicely formatted output in the browser, including tables, figures, animation, video, etc.
- The ability to mix in formatted text and mathematical expressions between cells
- Functions to generate PDF slides, static html, etc.

Whether you end up using Jupyter as your primary work environment or not, you'll find learning about it an excellent investment

Installing Jupyter There are two steps here:

1. Installing Jupyter itself
2. Installing IJulia, which serves as an interface between Jupyter notebooks and Julia

In fact you can get both by installing IJulia

However, if you have the bandwidth, we recommend that you

1. Do the two steps separately
2. In the first step, when installing Jupyter, do this by installing the larger package [Anaconda Python](#)

The advantage of this approach is that Anaconda gives you not just Jupyter by the whole scientific Python ecosystem

This includes things like plotting tools we'll make use of later

Installing Anaconda Installing Anaconda is straightforward: [download the binary](#) and follow the instructions

If you are asked during the installation process whether you'd like to make Anaconda your default Python installation, say yes — you can always remove it later

Otherwise you can accept all of the defaults

Note that the packages in Anaconda update regularly — you can keep up to date by typing `conda update anaconda` in a terminal

Installing IJulia Now open up a Julia terminal and type

```
julia> Pkg.add("IJulia")
```

Warning: The IJulia website states that if you get an error message you should force a rebuild with `Pkg.build("IJulia")`

If you have problems, consult [the installation instructions](#)

Other Requirements Since IJulia runs in the browser it might be a good time to update your browser

One good option is to install a free modern browser such as [Chrome](#) or [Firefox](#)

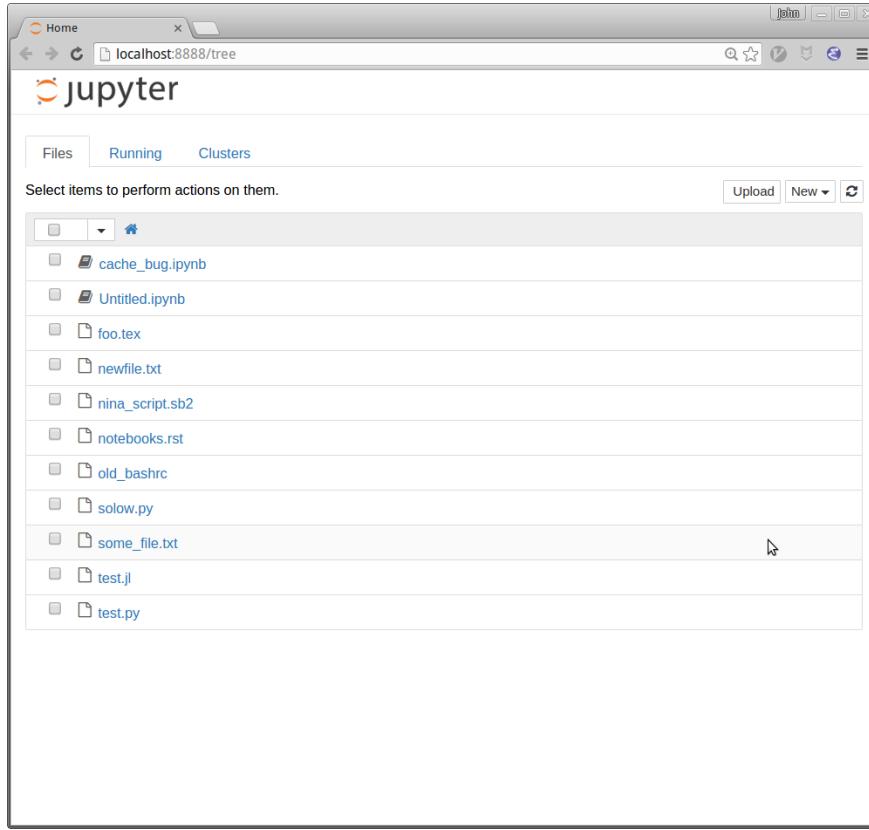
In our experience Chrome plays well with IJulia

Getting Starting Now either

1. search for and start the Jupyter notebook application on your machine or
2. open up a terminal (or `cmd` in Windows) and type

```
jupyter notebook
```

You should see something (not exactly) like this



The page you are looking at is called the “dashboard”

The address `localhost:8888/tree` you see in the image indicates that the browser is communicating with a Julia session via port 8888 of the local machine

If you click on “New” you should have the option to start a Julia notebook

Here’s what your Julia notebook should look like:

The notebook displays an *active cell*, into which you can type Julia commands

Notebook Basics Notice that in the previous figure the cell is surrounded by a green border

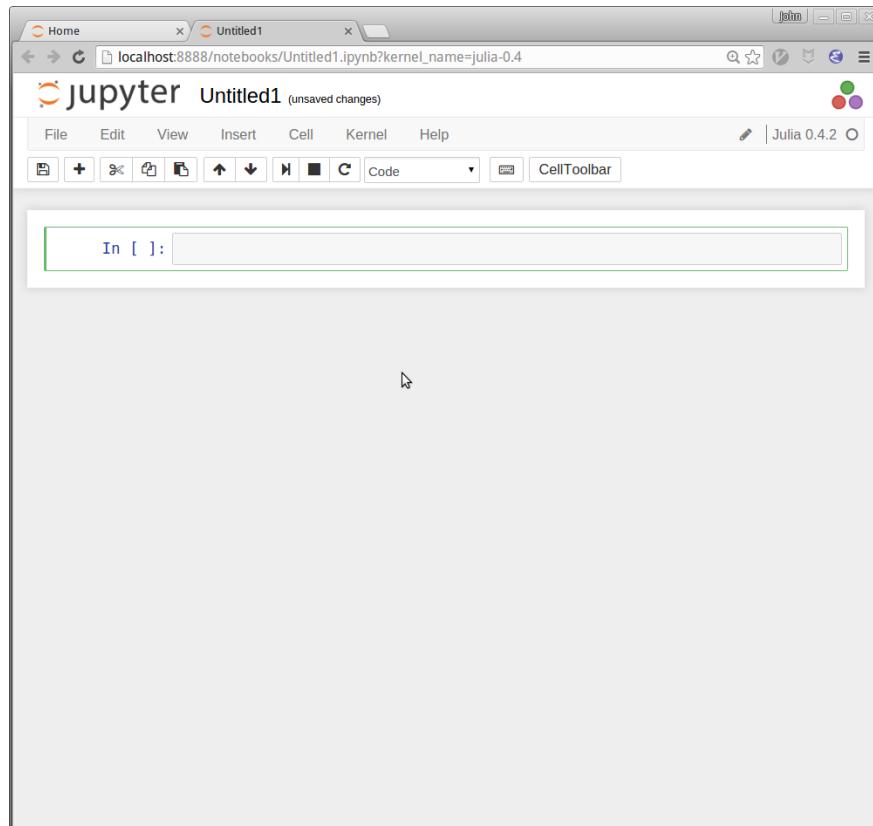
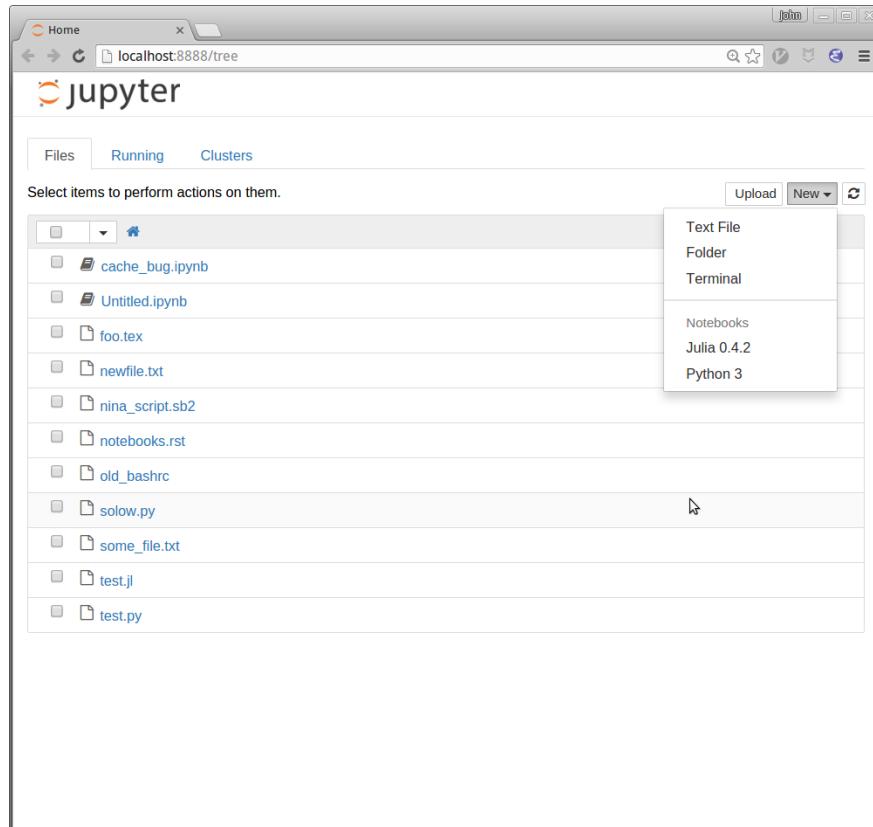
This means that the cell is in *edit mode*

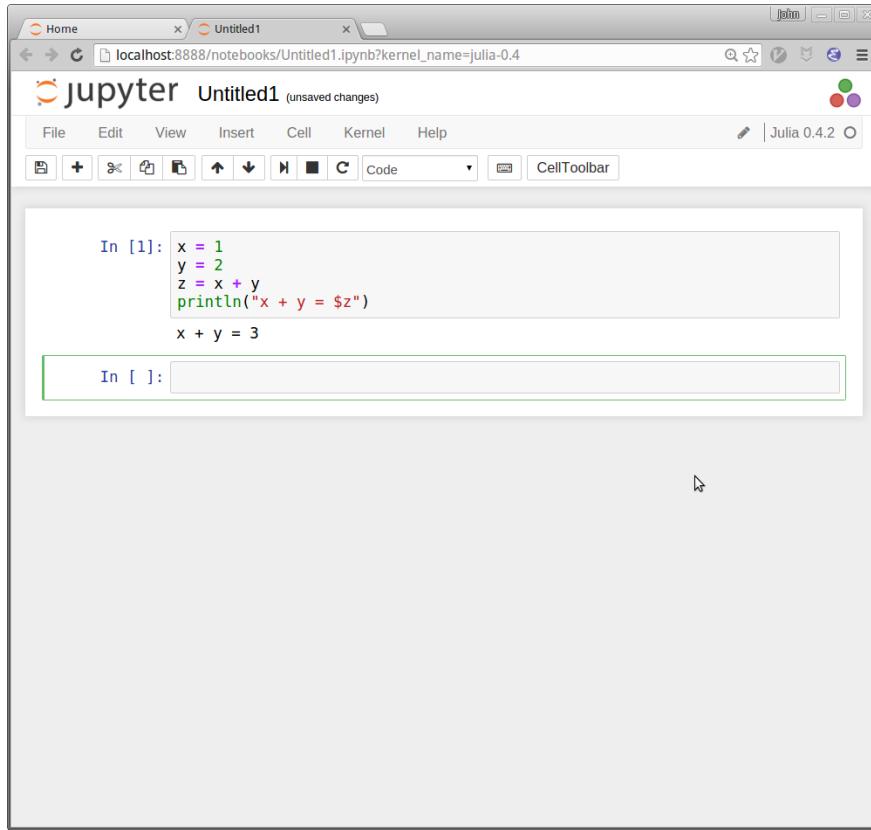
As a result, you can type in Julia code and it will appear in the cell

When you’re ready to execute these commands, hit `Shift-Enter` instead of the usual `Enter`

Modal Editing The next thing to understand about the Jupyter notebook is that it uses a *modal* editing system

This means that the effect of typing at the keyboard **depends on which mode you are in**





The two modes are

1. Edit mode
 - Indicated by a green border around one cell, as in the pictures above
 - Whatever you type appears as is in that cell
2. Command mode
 - The green border is replaced by a grey border
 - Key strokes are interpreted as commands — for example, typing *b* adds a new cell below the current one

(To learn about other commands available in command mode, go to “Keyboard Shortcuts” in the “Help” menu)

Switching modes

- To switch to command mode from edit mode, hit the `Esc` key
- To switch to edit mode from command mode, hit `Enter` or click in a cell

The modal behavior of the Jupyter notebook is a little tricky at first but very efficient when you get used to it

Plots Let's generate some plots

There are several options we'll discuss in detail later

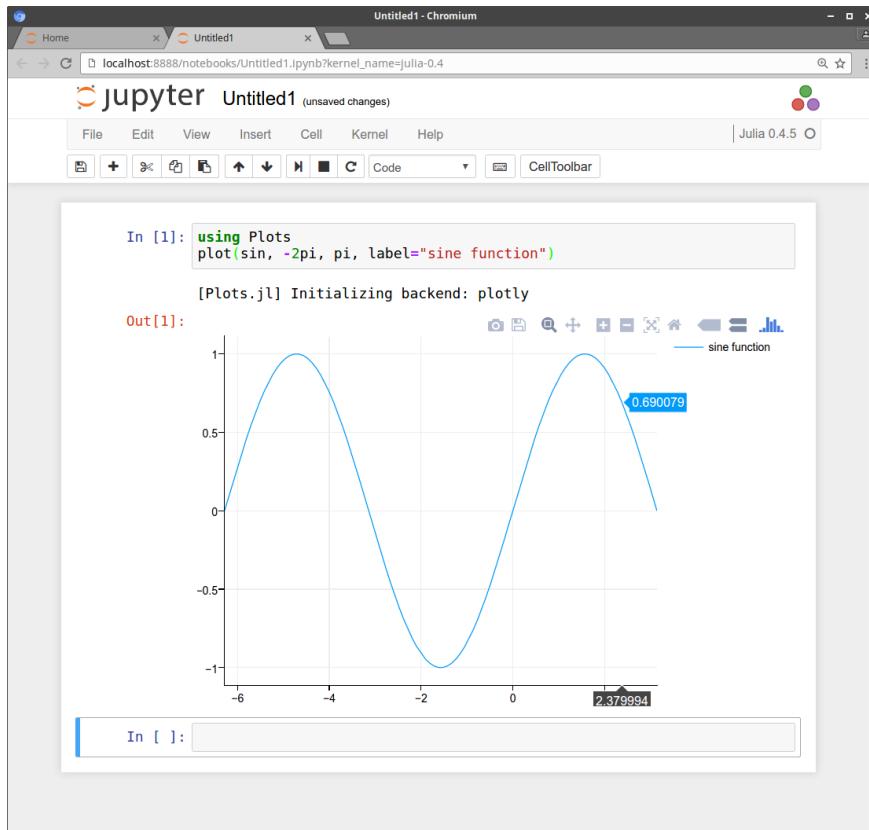
For now lets start with `Plots.jl`

```
julia> Pkg.add("Plots")
```

Now try copying the following into a notebook cell and hit Shift-Enter

```
using Plots
plot(sin, -2pi, pi, label="sine function")
```

This is what you should see:



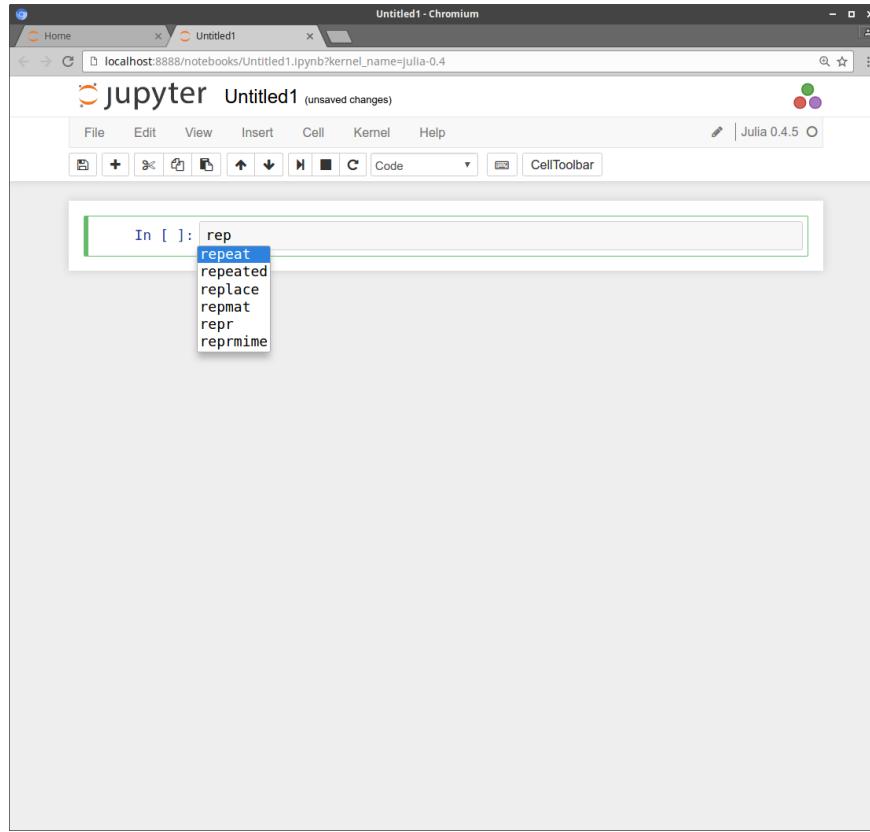
Working with the Notebook Let's go over some more Jupyter notebook features — enough so that we can press ahead with programming

Tab Completion A simple but useful feature of IJulia is tab completion

For example if you type `rep` and hit the tab key you'll get a list of all commands that start with `rep`

IJulia offers up the possible completions

This helps remind you of what's available and saves a bit of typing



On-Line Help To get help on the Julia function such as `repmat`, enter `?repmat`. Documentation should now appear in the browser

Other Content In addition to executing code, the Jupyter notebook allows you to embed text, equations, figures and even videos in the page

For example, here we enter a mixture of plain text and LaTeX instead of code

Next we `Esc` to enter command mode and then type `m` to indicate that we are writing [Markdown](#), a mark-up language similar to (but simpler than) LaTeX

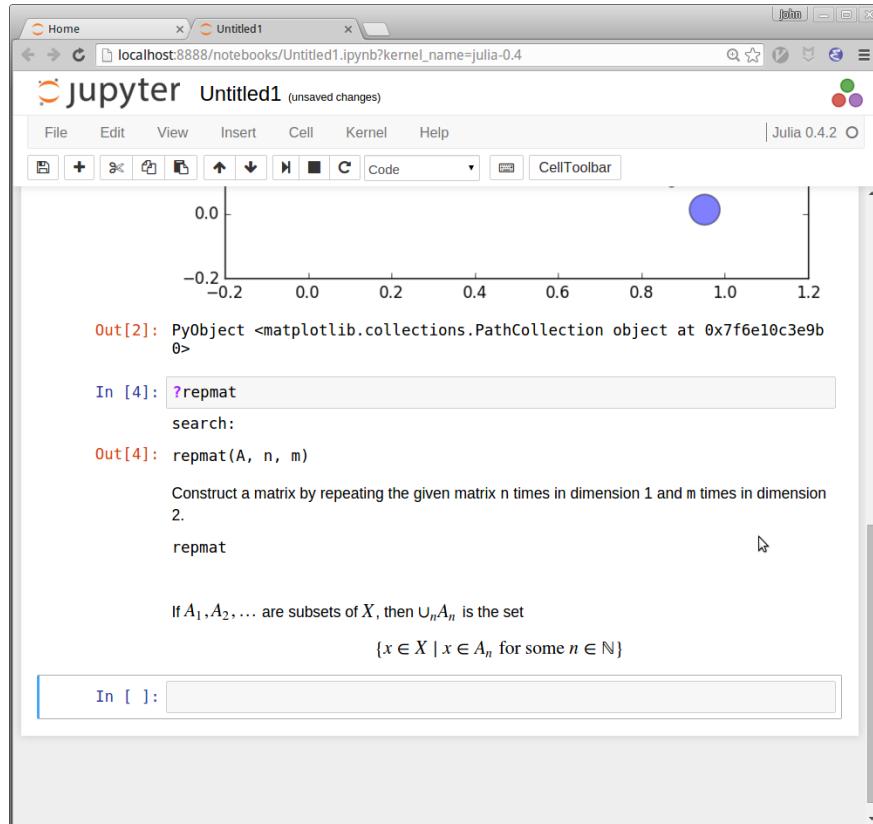
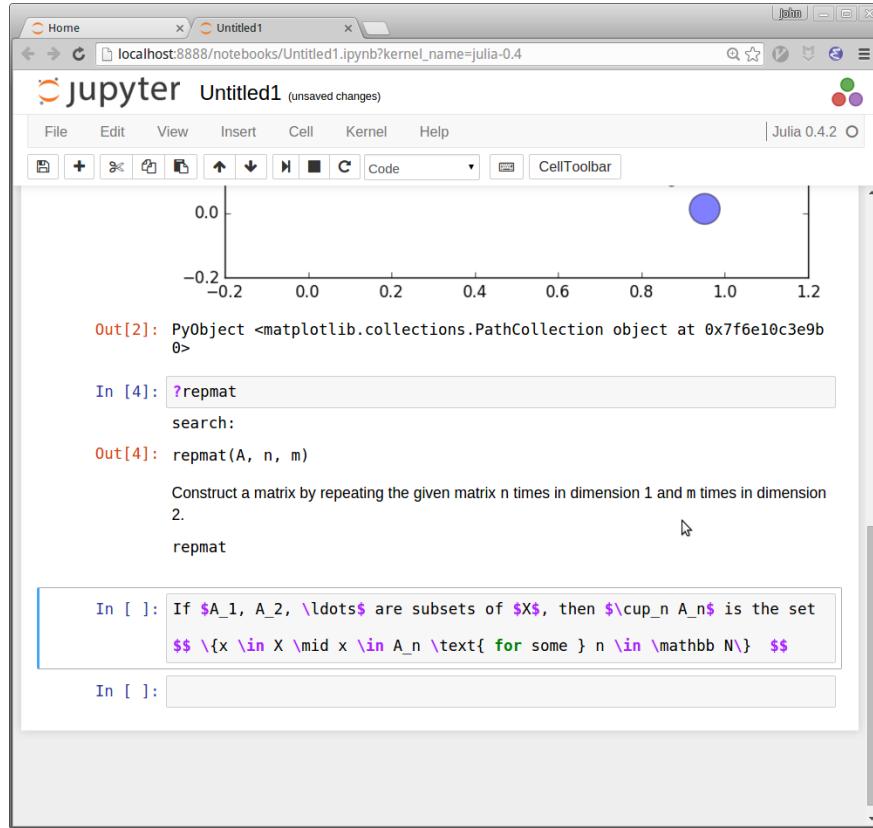
(You can also use your mouse to select Markdown from the Code drop-down box just below the list of menu items)

Now we `Shift + Enter` to produce this

Shell Commands You can execute shell commands (system commands) in IJulia by prepending a semicolon

For example, `;ls` will execute the UNIX style shell command `ls`, which — at least for UNIX style operating systems — lists the contents of the current working directory

These shell commands are handled by your default system shell and hence are platform specific



Working with Files To run an existing Julia file using the notebook we can either

1. copy and paste the contents into a cell in the notebook, or
2. use `include("filename")` in the same manner as for the Julia interpreter discussed above

More sophisticated methods for working with files are under active development and should be on-line soon

Sharing Notebooks Notebook files are just text files structured in [JSON](#) and typically ending with `.ipynb`

A notebook can easily be saved and shared between users — you just need to pass around the `ipynb` file

To open an existing `ipynb` file, import it from the dashboard (the first browser page that opens when you start Jupyter notebook) and run the cells or edit as discussed above

nbviewer The Jupyter organization has a site for sharing notebooks called [nbviewer](#)

The notebooks you see there are static HTML representations of notebooks

However, each notebook can be downloaded as an `ipynb` file by clicking on the download icon at the top right of its page

Once downloaded you can open it as a notebook, as we discussed just above

QuantEcon

[QuantEcon](#) is an organization that promotes development of open source code for economics and econometrics (feel free to get involved!)

QuantEcon.jl Among other things, QuantEcon supports [QuantEcon.jl](#), an open source code library for quantitative economic modeling in Julia

You can install this package through the usual Julia package manager:

```
julia> Pkg.add("QuantEcon")
```

For example, the following code creates a discrete approximation to an AR(1) process

```
julia> using QuantEcon: tauchen

julia> tauchen(4, 0.9, 1.0)
Discrete Markov Chain
stochastic matrix of type Array{Float64,2}:
4x4 Array{Float64,2}:
 0.945853   0.0541468   2.92863e-10   0.0
 0.00580845   0.974718   0.0194737   1.43534e-11
 1.43534e-11   0.0194737   0.974718   0.00580845
 2.08117e-27   2.92863e-10   0.0541468   0.945853
```

We'll learn more about the library as we go along

QuantEcon.applications The smaller Julia scripts we use in these lectures live in a separate repository called [QuantEcon.applications](#)

You can grab a copy of the files in this repo directly by downloading the zip file — try clicking the “Download ZIP” button on the [main page](#)

Alternatively, you can get a copy of the repo using [Git](#)

For more information see *Exercise 1*

Exercises

Exercise 1 If you haven't heard, [Git](#) is a *version control system* — a piece of software used to manage digital projects such as code libraries

In many cases the associated collections of files — called *repositories* — are stored on [GitHub](#)

[GitHub](#) is a wonderland of collaborative coding projects

Git is the underlying software used to manage these projects

Git is an extremely powerful tool for distributed collaboration — for example, we use it to share and synchronize all the source files for these lectures

There are two main flavors of Git

1. the plain vanilla command line version (which we recommend)
2. the point-and-click [GUI versions](#)

As an exercise, try getting a copy of [QuantEcon.applications](#) using Git

(Perhaps google for recent recommendations matching your operating system)

If you've installed the command line version, open up a terminal (`cmd` on Windows) and enter

```
git clone https://github.com/QuantEcon/QuantEcon.applications
```

This is just `git clone` in front of the URL for the repository

Even better, sign up to [GitHub](#) — it's free

Look into ‘forking’ GitHub repositories

(Loosely speaking, forking means making your own copy of a GitHub repository, stored on GitHub)

Try forking [QuantEcon.applications](#)

Now try cloning it to some local directory, making edits, adding and committing them, and pushing them back up to your forked GitHub repo

For reading on these and other topics, try

- [The official Git documentation](#)

- Reading through the docs on [GitHub](#)

An Introductory Example

Contents

- *An Introductory Example*
 - [Overview](#)
 - [Example: Plotting a White Noise Process](#)
 - [Exercises](#)
 - [Solutions](#)

Overview

We're now ready to start learning the Julia language itself

Level Our approach is aimed at those who already have at least some knowledge of programming — perhaps experience with Python, MATLAB, R, C or similar

In particular, we assume you have some familiarity with fundamental programming concepts such as

- variables
- loops
- conditionals (if/else)

If you have no such programming experience, then one option is to try Python first

Python is a great first language and there are many introductory treatments

Otherwise, just dive in and see how you go...

Approach In this lecture we will write and then pick apart small Julia programs

At this stage the objective is to introduce you to basic syntax and data structures

Deeper concepts—how things work—will be covered in later lectures

Since we are looking for simplicity the examples are a little contrived

Set Up We assume that you've worked your way through [our getting started lecture](#) already

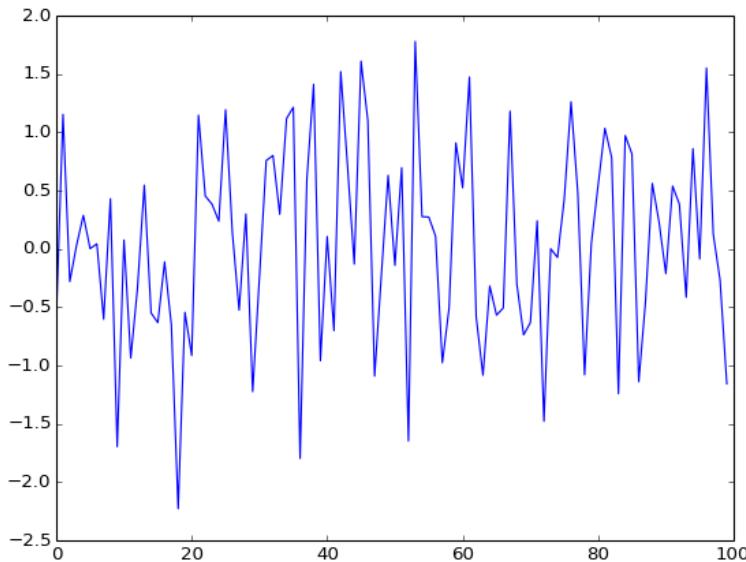
For this lecture, we recommend that you work in a Jupyter notebook, as described [here](#)

Other References The definitive reference is Julia's own documentation
 The manual is thoughtfully written but also quite dense (and somewhat evangelical)
 The presentation in this and our remaining lectures is more of a tutorial style based around examples

Example: Plotting a White Noise Process

To begin, let's suppose that we want to simulate and plot the white noise process $\epsilon_0, \epsilon_1, \dots, \epsilon_T$, where each draw ϵ_t is independent standard normal

In other words, we want to generate figures that look something like this:



This is straightforward using *Plots.jl*, which was discussed in our [set up](#) lecture

Fire up a *Jupyter notebook* and enter the following in a cell

```
using Plots
ts_length = 100
epsilon_values = randn(ts_length)
plot(epsilon_values, color="blue")
```

Let's break this down and see how it works

Importing Functions The effect of the statement `using Plots` is to make all the names exported by the `Plots` module available in the global scope

If you prefer to be more selective you can replace `using Plots` with `import Plots: plot`

Now only the `plot` function is accessible

Since our program uses only the plot function from this module, either would have worked in the previous example

Arrays The function call `epsilon_values = randn(ts_length)` creates one of the most fundamental Julia data types: an array

```
julia> typeof(epsilon_values)
Array{Float64,1}

julia> epsilon_values
100-element Array{Float64,1}:
 -0.908823
 -0.759142
 -1.42078
  0.792799
  0.577181
  1.74219
 -0.912529
  1.06259
  0.5766
 -0.0172788
 -0.591671
 -1.02792
 ...
 -1.29412
 -1.12475
  0.437858
 -0.709243
 -1.96053
  1.31092
  1.19819
  1.54028
 -0.246204
 -1.23305
 -1.16484
```

The information from `typeof()` tells us that `epsilon_values` is an array of 64 bit floating point values, of dimension 1

Julia arrays are quite flexible — they can store heterogeneous data for example

```
julia> x = [10, "foo", false]
3-element Array{Any,1}:
 10
 "foo"
 false
```

Notice now that the data type is recorded as `Any`, since the array contains mixed data

The first element of `x` is an integer

```
julia> typeof(x[1])
Int64
```

The second is a string

```
julia> typeof(x[2])
ASCIIString (constructor with 2 methods)
```

The third is the boolean value `false`

```
julia> typeof(x[3])
Bool
```

Notice from the above that

- array indices start at 1 (unlike Python, where arrays are zero-based)
- array elements are referenced using square brackets (unlike MATLAB and Fortran)

Julia contains many functions for acting on arrays — we'll review them later

For now here's several examples, applied to the same list `x = [10, "foo", false]`

```
julia> length(x)
3

julia> pop!(x)
false

julia> x
2-element Array{Any,1}:
 10
 "foo"

julia> push!(x, "bar")
3-element Array{Any,1}:
 10
 "foo"
 "bar"

julia> x
3-element Array{Any,1}:
 10
 "foo"
 "bar"
```

The first example just returns the length of the list

The second, `pop!()`, pops the last element off the list and returns it

In doing so it changes the list (by dropping the last element)

Because of this we call `pop!` a **mutating method**

It's conventional in Julia that mutating methods end in `!` to remind the user that the function has other effects beyond just returning a value

The function `push!()` is similar, expect that it appends its second argument to the array

For Loops Although there's no need in terms of what we wanted to achieve with our program, for the sake of learning syntax let's rewrite our program to use a for loop

```
using Plots
ts_length = 100
epsilon_values = Array(Float64, ts_length)
for i in 1:ts_length
    epsilon_values[i] = randn()
end
plot(epsilon_values, color="blue")
```

Here we first declared `epsilon_values` to be an empty array for storing 64 bit floating point numbers

The `for` loop then populates this array by successive calls to `randn()`

- Called without an argument, `randn()` returns a single float

Like all code blocks in Julia, the end of the for loop code block (which is just one line here) is indicated by the keyword `end`

The word `in` from the for loop can be replaced by symbol `=`

The expression `1:ts_length` creates an **iterator** that is looped over — in this case the integers from 1 to `ts_length`

Iterators are memory efficient because the elements are generated on the fly rather than stored in memory

In Julia you can also loop directly over arrays themselves, like so

```
words = ["foo", "bar"]
for word in words
    println("Hello $word")
end
```

The output is

```
Hello foo
Hello bar
```

While Loops The syntax for the while loop contains no surprises

```
using Plots
ts_length = 100
epsilon_values = Array(Float64, ts_length)
i = 1
while i <= ts_length
    epsilon_values[i] = randn()
    i = i + 1
end
plot(epsilon_values, color="blue")
```

The next example does the same thing with a condition and the `break` statement

```
using Plots
ts_length = 100
epsilon_values = Array(Float64, ts_length)
i = 1
while true
    epsilon_values[i] = randn()
    i = i + 1
    if i > ts_length
        break
    end
end
plot(epsilon_values, color="blue")
```

User-Defined Functions For the sake of the exercise, let's now go back to the `for` loop but re-structure our program so that generation of random variables takes place within a user-defined function

```
using Plots

function generate_data(n)
    epsilon_values = Array(Float64, n)
    for i = 1:n
        epsilon_values[i] = randn()
    end
    return epsilon_values
end

ts_length = 100
data = generate_data(ts_length)
plot(data, color="blue")
```

Here

- `function` is a Julia keyword that indicates the start of a function definition
- `generate_data` is an arbitrary name for the function
- `return` is a keyword indicating the return value

A Slightly More Useful Function Of course the function `generate_data` is completely contrived

We could just write the following and be done

```
ts_length = 100
data = randn(ts_length)
plot(data, color="blue")
```

Let's make a slightly more useful function

This function will be passed a choice of probability distribution and respond by plotting a histogram of observations

In doing so we'll make use of the Distributions package

```
julia> Pkg.add("Distributions")
```

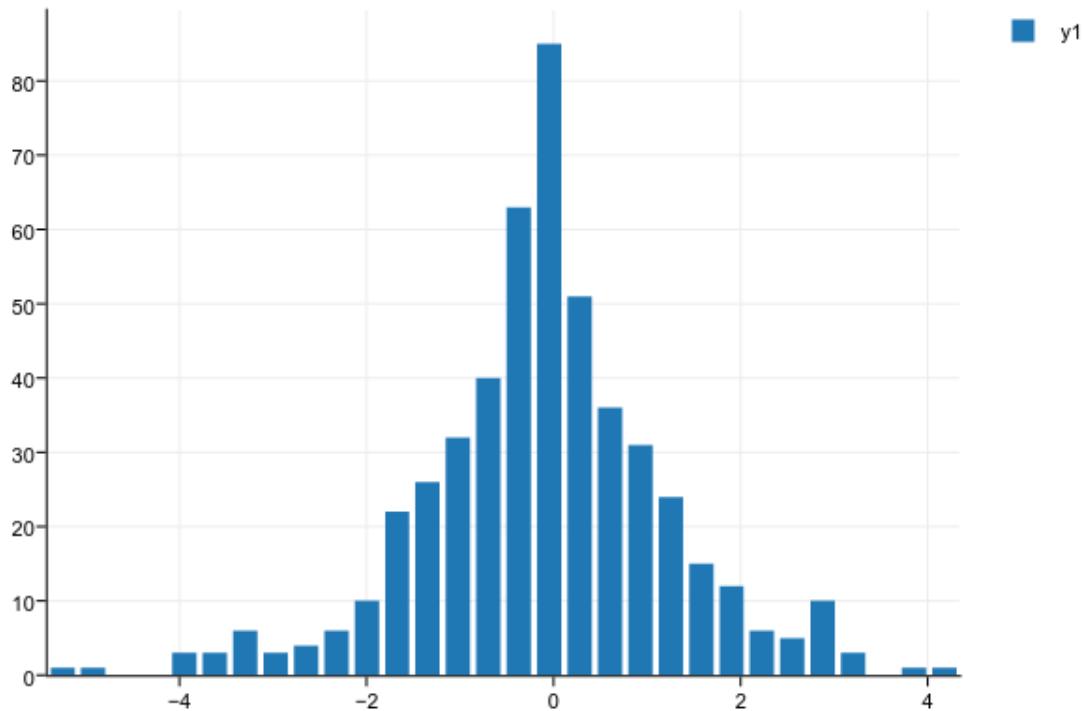
Here's the code

```
using Plots
using Distributions

function plot_histogram(distribution, n)
    epsilon_values = rand(distribution, n) # n draws from distribution
    histogram(epsilon_values)
end

lp = Laplace()
plot_histogram(lp, 500)
```

The resulting figure looks like this



Let's have a casual discussion of how all this works while leaving technical details for later in the lectures

First, `lp = Laplace()` creates an instance of a data type defined in the `Distributions` module that represents the Laplace distribution

The name `lp` is bound to this object

When we make the function call `plot_histogram(lp, 500)` the code in the body of the function `plot_histogram` is run with

- the name `distribution` bound to the same object as `lp`

- the name `n` bound to the integer 500

A Mystery Now consider the function call `rand(distribution, n)`

This looks like something of a mystery

The function `rand()` is defined in the base library such that `rand(n)` returns `n` uniform random variables on $[0, 1]$

```
julia> rand(3)
3-element Array{Float64,1}:
 0.856817
 0.981502
 0.510947
```

On the other hand, `distribution` points to a data type representing the Laplace distribution that has been defined in a third party package

So how can it be that `rand()` is able to take this kind of object as an argument and return the output that we want?

The answer in a nutshell is **multiple dispatch**

This refers to the idea that functions in Julia can have different behavior depending on the particular arguments that they're passed

Hence in Julia we can take an existing function and give it a new behavior by defining how it acts on a new type of object

The interpreter knows which function definition to apply in a given setting by looking at the types of the objects the function is called on

In Julia these alternative versions of a function are called **methods**

Exercises

Exercise 1 Recall that $n!$ is read as “ n factorial” and defined as $n! = n \times (n - 1) \times \cdots \times 2 \times 1$

In Julia you can compute this value with `factorial(n)`

Write your own version of this function, called `factorial2`, using a for loop

Exercise 2 The **binomial random variable** $Y \sim Bin(n, p)$ represents

- number of successes in n binary trials
- each trial succeeds with probability p

Using only `rand()` from the set of Julia's built in random number generators (not the Distributions package), write a function `binomial_rv` such that `binomial_rv(n, p)` generates one draw of Y

Hint: If U is uniform on $(0, 1)$ and $p \in (0, 1)$, then the expression $U < p$ evaluates to true with probability p

Exercise 3 Compute an approximation to π using Monte Carlo

For random number generation use only `rand()`

Your hints are as follows:

- If U is a bivariate uniform random variable on the unit square $(0, 1)^2$, then the probability that U lies in a subset B of $(0, 1)^2$ is equal to the area of B
- If U_1, \dots, U_n are iid copies of U , then, as n gets large, the fraction that falls in B converges to the probability of landing in B
- For a circle, area = $\pi * \text{radius}^2$

Exercise 4 Write a program that prints one realization of the following random device:

- Flip an unbiased coin 10 times
- If 3 consecutive heads occur one or more times within this sequence, pay one dollar
- If not, pay nothing

Once again use only `rand()` as your random number generator

Exercise 5 Simulate and plot the correlated time series

$$x_{t+1} = \alpha x_t + \epsilon_{t+1} \quad \text{where } x_0 = 0 \quad \text{and } t = 0, \dots, T$$

The sequence of shocks $\{\epsilon_t\}$ is assumed to be iid and standard normal

Set $T = 200$ and $\alpha = 0.9$

Exercise 6 Plot three simulated time series, one for each of the cases $\alpha = 0$, $\alpha = 0.8$ and $\alpha = 0.98$

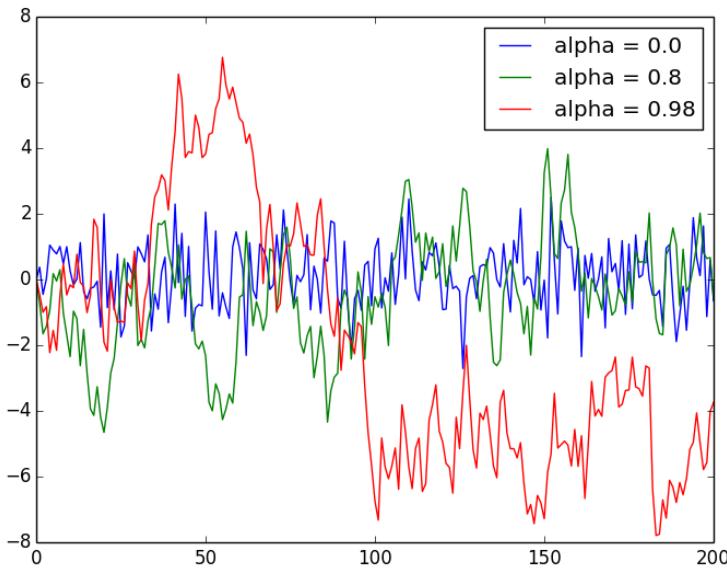
In particular, you should produce (modulo randomness) a figure that looks as follows

(The figure illustrates how time series with the same one-step-ahead conditional volatilities, as these three processes have, can have very different unconditional volatilities.)

Solutions

[Solution notebook](#)

Julia Essentials



Contents

- *Julia Essentials*
 - *Overview*
 - *Common Data Types*
 - *Input and Output*
 - *Iterating*
 - *Comparisons and Logical Operators*
 - *User Defined Functions*
 - *Exercises*
 - *Solutions*

Having covered a few examples, let's now turn to a more systematic exposition of the essential features of the language

Overview

Topics:

- Common data types
- Basic file I/O
- Iteration
- More on user-defined functions
- Comparisons and logic

Common Data Types

Like most languages, Julia language defines and provides functions for operating on standard data types such as

- integers
- floats
- strings
- arrays, etc...

Let's learn a bit more about them

Primitive Data Types A particularly simple data type is a Boolean value, which can be either `true` or `false`

```
julia> x = true
true

julia> typeof(x)
Bool

julia> y = 1 > 2  # Now y = false
false
```

Under addition, `true` is converted to 1 and `false` is converted to 0

```
julia> true + false
1

julia> sum([true, false, false, true])
2
```

The two most common data types used to represent numbers are integers and floats

(Computers distinguish between floats and integers because arithmetic is handled in a different way)

```
julia> typeof(1.0)
Float64

julia> typeof(1)
Int64
```

If you're running a 32 bit system you'll still see `Float64`, but you will see `Int32` instead of `Int64` (see [the section on Integer types](#) from the Julia manual)

Arithmetic operations are fairly standard

```
julia> x = 2; y = 1.0
1.0

julia> x * y
```

```
2.0

julia> x^2
4

julia> y / x
0.5
```

Although the * can be omitted for multiplication between variables and numeric literals

```
julia> 2x - 3y
1.0
```

Also, you can use function (instead of infix) notation if you so desire

```
julia> +(10, 20)
30

julia> *(10, 20)
200
```

Complex numbers are another primitive data type, with the imaginary part being specified by `im`

```
julia> x = 1 + 2im
1 + 2im

julia> y = 1 - 2im
1 - 2im

julia> x * y  # Complex multiplication
5 + 0im
```

There are several more primitive data types that we'll introduce as necessary

Strings A string is a data type for storing a sequence of characters

```
julia> x = "foobar"
"foobar"

julia> typeof(x)
ASCIIString (constructor with 2 methods)
```

You've already seen examples of Julia's simple string formatting operations

```
julia> x = 10; y = 20
20

julia> "x = $x"
"x = 10"

julia> "x + y = $(x + y)"
"x + y = 30"
```

To concatenate strings use *

```
julia> "foo" * "bar"
"foobar"
```

Julia provides many functions for working with strings

```
julia> s = "Charlie don't surf"
"Charlie don't surf"

julia> split(s)
3-element Array{SubString{ASCIIString},1}:
 "Charlie"
 "don't"
 "surf"

julia> replace(s, "surf", "ski")
"Charlie don't ski"

julia> split("fee,fi,fo", ",")
3-element Array{SubString{ASCIIString},1}:
 "fee"
 "fi"
 "fo"

julia> strip(" foobar ")  # Remove whitespace
"foobar"
```

Julia can also find and replace using regular expressions (see the documentation on regular expressions for more info)

```
julia> match(r"(\d+)", "Top 10")  # Find numerals in string
RegexMatch("10", 1="10")
```

Containers Julia has several basic types for storing collections of data

We have already discussed arrays

A related data type is **tuples**, which can act like “immutable” arrays

```
julia> x = ("foo", "bar")
("foo", "bar")

julia> typeof(x)
(ASCIIString, ASCIIString)
```

An immutable object is one that cannot be altered once it resides in memory

In particular, tuples do not support item assignment:

```
julia> x[1] = 42
ERROR: `setindex!` has no method matching setindex!(::(ASCIIString,ASCIIString), ::Int64, ::Int64)
```

This is similar to Python, as is the fact that the parenthesis can be omitted

```
julia> x = "foo", "bar"
("foo", "bar")
```

Another similarity with Python is tuple unpacking, which means that the following convenient syntax is valid

```
julia> x = ("foo", "bar")
("foo", "bar")

julia> word1, word2 = x
("foo", "bar")

julia> word1
"foo"

julia> word2
"bar"
```

Referencing Items The last element of a sequence type can be accessed with the keyword `end`

```
julia> x = [10, 20, 30, 40]
4-element Array{Int64,1}:
 10
 20
 30
 40

julia> x[end]
40

julia> x[end-1]
30
```

To access multiple elements of an array or tuple, you can use slice notation

```
julia> x[1:3]
3-element Array{Int64,1}:
 10
 20
 30

julia> x[2:end]
3-element Array{Int64,1}:
 20
 30
 40
```

The same slice notation works on strings

```
julia> "foobar"[3:end]
"obar"
```

Dictionaries Another container type worth mentioning is dictionaries

Dictionaries are like arrays except that the items are named instead of numbered

```
julia> d = Dict("name" => "Frodo", "age" => 33)
Dict{ASCIIString,Any} with 2 entries:
"name" => "Frodo"
"age"   => 33

julia> d["age"]
33
```

The strings `name` and `age` are called the **keys**

The objects that the keys are mapped to ("Frodo" and 33) are called the **values**

They can be accessed via `keys(d)` and `values(d)` respectively

Input and Output

Let's have a quick look at reading from and writing to text files

We'll start with writing

```
julia> f = open("newfile.txt", "w")  # "w" for writing
IOStream(<file newfile.txt>)

julia> write(f, "testing\n")        # \n for newline
7

julia> write(f, "more testing\n")
12

julia> close(f)
```

The effect of this is to create a file called `newfile.txt` in your present working directory with contents

```
testing
more testing
```

We can read the contents of `newline.txt` as follows

```
julia> f = open("newfile.txt", "r")  # Open for reading
IOStream(<file newfile.txt>)

julia> print(readall(f))
testing
more testing

julia> close(f)
```

Often when reading from a file we want to step through the lines of a file, performing an action on each one

There's a neat interface to this in Julia, which takes us to our next topic

Iterating

One of the most important tasks in computing is stepping through a sequence of data and performing a given action

Julia's provides neat, flexible tools for iteration as we now discuss

Iterables An iterable is something you can put on the right hand side of `for` and loop over

These include sequence data types like arrays

```
actions = ["surf", "ski"]
for action in actions
    println("Charlie don't $action")
end
```

They also include so-called **iterators**

You've already come across these types of objects

```
julia> for i in 1:3 print(i) end
123
```

If you ask for the keys of dictionary you get an iterator

```
julia> d = Dict("name" => "Frodo", "age" => 33)
Dict{ASCIIString,Any} with 2 entries:
  "name" => "Frodo"
  "age"   => 33

julia> keys(d)
Base.KeyIterator for a Dict{ASCIIString,Any} with 2 entries. Keys:
  "name"
  "age"
```

This makes sense, since the most common thing you want to do with keys is loop over them

The benefit of providing an iterator rather than an array, say, is that the former is more memory efficient

Should you need to transform an iterator into an array you can always use `collect()`

```
julia> collect(keys(d))
2-element Array{Any,1}:
  "name"
  "age"
```

Looping without Indices You can loop over sequences without explicit indexing, which often leads to neater code

For example compare

```
for x in x_values
    println(x * x)
end
```

with

```
for i in 1:length(x_values)
    println(x_values[i] * x_values[i])
end
```

Julia provides some functional-style helper functions (similar to Python) to facilitate looping without indices

One is `zip()`, which is used for stepping through pairs from two sequences

For example, try running the following code

```
countries = ("Japan", "Korea", "China")
cities = ("Tokyo", "Seoul", "Beijing")
for (country, city) in zip(countries, cities)
    println("The capital of $country is $city")
end
```

If we happen to need the index as well as the value, one option is to use `enumerate()`

The following snippet will give you the idea

```
countries = ("Japan", "Korea", "China")
cities = ("Tokyo", "Seoul", "Beijing")
for (i, country) in enumerate(countries)
    city = cities[i]
    println("The capital of $country is $city")
end
```

Comprehensions Comprehensions are an elegant tool for creating new arrays or dictionaries from iterables

Here's some examples

```
julia> doubles = [2i for i in 1:4]
4-element Array{Int64,1}:
 2
 4
 6
 8

julia> animals = ["dog", "cat", "bird"];    # semicolon suppresses output

julia> plurals = [animal * "s" for animal in animals]
3-element Array{ByteString,1}:
 "dogs"
 "cats"
 "birds"
```

```
julia> [i + j for i in 1:3, j in 4:6]
3x3 Array{Int64,2}:
 5  6  7
 6  7  8
 7  8  9

julia> [i + j + k for i in 1:3, j in 4:6, k in 7:9]
3x3x3 Array{Int64,3}:
[:, :, 1] =
 12 13 14
 13 14 15
 14 15 16

[:, :, 2] =
 13 14 15
 14 15 16
 15 16 17

[:, :, 3] =
 14 15 16
 15 16 17
 16 17 18
```

The same kind of expression works for dictionaries

```
julia> ["$i" => i for i in 1:3]
Dict{ASCIIString,Int64} with 3 entries:
"1" => 1
"2" => 2
"3" => 3
```

(This syntax is likely to change towards something like Dict("\$i" => i for i in 1:3) in future versions)

Comparisons and Logical Operators

Comparisons As we saw earlier, when testing for equality we use ==

```
julia> x = 1
1

julia> x == 2
false
```

For “not equal” use !=

```
julia> x != 3
true
```

We can chain inequalities:

```
julia> 1 < 2 < 3
true

julia> 1 <= 2 <= 3
true
```

In many languages you can use integers or other values when testing conditions but Julia is more fussy

```
julia> while 0 println("foo") end
ERROR: TypeError: non-boolean (Int64) used in boolean context
in anonymous at no file

julia> if 1 print("foo") end
ERROR: TypeError: non-boolean (Int64) used in boolean context
```

Combining Expressions Here are the standard logical connectives (conjunction, disjunction)

```
julia> true && false
false

julia> true || false
true
```

Remember

- P $\&\&$ Q is true if both are true, otherwise it's false
- P $\|$ Q is false if both are false, otherwise it's true

User Defined Functions

Let's talk a little more about user defined functions

User defined functions are important for improving the clarity of your code by

- separating different strands of logic
- facilitating code reuse (writing the same thing twice is always a bad idea)

Julia functions are convenient:

- Any number of functions can be defined in a given file
- Any “value” can be passed to a function as an argument, including other functions
- Functions can be (and often are) defined inside other functions
- A function can return any kind of value, including functions

We'll see many examples of these structures in the following lectures

For now let's just cover some of the different ways of defining functions

Return Statement In Julia, the `return` statement is optional, so that the following functions have identical behavior

```
function f1(a, b)
    return a * b
end

function f2(a, b)
    a * b
end
```

When no `return` statement is present, the last value obtained when executing the code block is returned

Although some prefer the second option, we often favor the former on the basis that explicit is better than implicit

A function can have arbitrarily many `return` statements, with execution terminating when the first `return` is hit

You can see this in action when experimenting with the following function

```
function foo(x)
    if x > 0
        return "positive"
    end
    return "nonpositive"
end
```

Other Syntax for Defining Functions For short function definitions Julia offers some attractive simplified syntax

First, when the function body is a simple expression, it can be defined without the `function` keyword or `end`

```
julia> f(x) = sin(1 / x)
f (generic function with 2 methods)
```

Let's check that it works

```
julia> f(1 / pi)
1.2246467991473532e-16
```

Julia also allows for you to define anonymous functions

For example, to define `f(x) = sin(1 / x)` you can use `x -> sin(1 / x)`

The difference is that the second function has no name bound to it

How can you use a function with no name?

Typically it's as an argument to another function

```
julia> map(x -> sin(1 / x), randn(3)) # Apply function to each element
3-element Array{Float64,1}:
```

```
0.744193
-0.370506
-0.458826
```

Optional and Keyword Arguments Function arguments can be given default values

```
function f(x, a=1)
    return exp(cos(a * x))
end
```

If the argument is not supplied the default value is substituted

```
julia> f(pi)
0.36787944117144233

julia> f(pi, 2)
2.718281828459045
```

Another option is to use **keyword** arguments

The difference between keyword and standard (positional) arguments is that they are parsed and bound by name rather than order in the function call

For example, in the call

```
simulate(param1, param2, max_iterations=100, error_tolerance=0.01)
```

the last two arguments are keyword arguments and their order is irrelevant (as long as they come after the positional arguments)

To define a function with keyword arguments you need to use ; like so

```
function simulate(param1, param2; max_iterations=100, error_tolerance=0.01)
    # Function body here
end
```

Exercises

Exercise 1 Part 1: Given two numeric arrays or tuples `x_vals` and `y_vals` of equal length, compute their inner product using `zip()`

Part 2: Using a comprehension, count the number of even numbers in 0,...,99

- Hint: `x % 2` returns 0 if `x` is even, 1 otherwise

Part 3: Using a comprehension, take `pairs = ((2, 5), (4, 2), (9, 8), (12, 10))` and count the number of pairs (`a, b`) such that both `a` and `b` are even

Exercise 2 Consider the polynomial

$$p(x) = a_0 + a_1x + a_2x^2 + \cdots a_nx^n = \sum_{i=0}^n a_i x^i \quad (1.1)$$

Using `enumerate()` in your loop, write a function `p` such that `p(x, coeff)` computes the value in (1.1) given a point `x` and an array of coefficients `coeff`

Exercise 3 Write a function that takes a string as an argument and returns the number of capital letters in the string

Hint: `uppercase("foo")` returns "FOO"

Exercise 4 Write a function that takes two sequences `seq_a` and `seq_b` as arguments and returns true if every element in `seq_a` is also an element of `seq_b`, else false

- By “sequence” we mean an array, tuple or string

Exercise 5 The Julia libraries include functions for interpolation and approximation

Nevertheless, let’s write our own function approximation routine as an exercise

In particular, write a function `linapprox` that takes as arguments

- A function `f` mapping some interval $[a, b]$ into \mathbb{R}
- two scalars `a` and `b` providing the limits of this interval
- An integer `n` determining the number of grid points
- A number `x` satisfying `a <= x <= b`

and returns the `piecewise linear interpolation` of `f` at `x`, based on `n` evenly spaced grid points `a = point[1] < point[2] < ... < point[n] = b`

Aim for clarity, not efficiency

Exercise 6 The following data lists US cities and their populations

```
new york: 8244910
los angeles: 3819702
chicago: 2707120
houston: 2145146
philadelphia: 1536471
phoenix: 1469471
san antonio: 1359758
san diego: 1326179
dallas: 1223229
```

Copy this text into a text file called `us_cities.txt` and save it in your present working directory

- That is, save it in the location Julia returns when you call `pwd()`

Write a program to calculate total population across these cities

Hints:

- If `f` is a file object then `eachline(f)` provides an iterable that steps you through the lines in the file

- `parse(Int, "100")` converts the string "100" into an integer

Solutions

[Solution notebook](#)

Vectors, Arrays and Matrices

Contents

- *Vectors, Arrays and Matrices*
 - *Overview*
 - *Array Basics*
 - *Operations on Arrays*
 - *Linear Algebra*
 - *Exercises*
 - *Solutions*

"Let's be clear: the work of science has nothing whatever to do with consensus. Consensus is the business of politics. Science, on the contrary, requires only one investigator who happens to be right, which means that he or she has results that are verifiable by reference to the real world. In science consensus is irrelevant. What is relevant is reproducible results." – Michael Crichton

Overview

In Julia, arrays are the most important data type for working with collections of numerical data

In this lecture we give more details on

- creating and manipulating Julia arrays
- fundamental array processing operations
- basic matrix algebra

Array Basics

Shape and Dimension We've already seen some Julia arrays in action

```
julia> a = [10, 20, 30]
3-element Array{Int64,1}:
 10
 20
 30
```

```
julia> a = ["foo", "bar", 10]
3-element Array{Any,1}:
 "foo"
 "bar"
 10
```

The REPL tells us that the arrays are of types `Array{Int64,1}` and `Array{Any,1}` respectively

Here `Int64` and `Any` are types for the elements inferred by the compiler

We'll talk more about types later on

The `1` in `Array{Int64,1}` and `Array{Any,1}` indicates that the array is one dimensional

This is the default for many Julia functions that create arrays

```
julia> typeof(randn(100))
Array{Float64,1}
```

To say that an array is one dimensional is to say that it is flat — neither a row nor a column vector

We can also confirm that `a` is flat using the `size()` or `ndims()` functions

```
julia> size(a)
(3,)

julia> ndims(a)
1
```

The syntax `(3,)` displays a tuple containing one element — the size along the one dimension that exists

Here's a function that creates a two-dimensional array

```
julia> eye(3)
3x3 Array{Float64,2}:
 1.0  0.0  0.0
 0.0  1.0  0.0
 0.0  0.0  1.0

julia> diagm([2, 4])
2x2 Array{Int64,2}:
 2  0
 0  4

julia> size(eye(3))
(3,3)
```

Array vs Vector vs Matrix In Julia, in addition to arrays you will see the types `Vector` and `Matrix`. However, these are just aliases for one- and two-dimensional arrays respectively

```
julia> Array{Int64, 1} == Vector{Int64}
true
```

```
julia> Array{Int64, 2} == Matrix{Int64}
true

julia> Array{Int64, 1} == Matrix{Int64}
false

julia> Array{Int64, 3} == Matrix{Int64}
false
```

In particular, a Vector in Julia is a flat array

Changing Dimensions The primary function for changing the dimension of an array is `reshape()`

```
julia> a = [10, 20, 30, 40]
4-element Array{Int64,1}:
 10
 20
 30
 40

julia> b = reshape(a, 2, 2)
2x2 Array{Int64,2}:
 10  30
 20  40

julia> b
2x2 Array{Int64,2}:
 10  30
 20  40
```

Notice that this function returns a “view” on the existing array

This means that changing the data in the new array will modify the data in the old one:

```
julia> b[1, 1] = 100 # Continuing the previous example
100

julia> b
2x2 Array{Int64,2}:
 100  30
 20   40

julia> a # First element has changed
4-element Array{Int64,1}:
 100
 20
 30
 40
```

To collapse an array along one dimension you can use `squeeze()`

```
julia> a = [1 2 3 4] # Two dimensional
1x4 Array{Int64,2}:
 1  2  3  4

julia> squeeze(a, 1)
4-element Array{Int64,1}:
 1
 2
 3
 4
```

The return value is an Array with the specified dimension “flattened”

Why Flat Arrays? As we’ve seen, in Julia we have both

- one-dimensional arrays (i.e., flat arrays)
- arrays of size $(1, n)$ or $(n, 1)$ that represent row and column vectors respectively

Why do we need both?

On one hand, dimension matters when we come to matrix algebra

- Multiplying by a row vector is different to multiplication by a column vector

On the other, we use arrays in many settings that don’t involve matrix algebra

In such cases, we don’t care about the distinction between row and column vectors

This is why many Julia functions return flat arrays by default

Creating Arrays

Functions that Return Arrays We’ve already seen some functions for creating arrays

```
julia> eye(2)
2x2 Array{Float64,2}:
 1.0  0.0
 0.0  1.0

julia> zeros(3)
3-element Array{Float64,1}:
 0.0
 0.0
 0.0
```

You can create an empty array using the `Array()` constructor

```
julia> x = Array(Float64, 2, 2)
2x2 Array{Float64,2}:
 0.0           2.82622e-316
 2.76235e-318 2.82622e-316
```

The printed values you see here are just garbage values
 (the existing contents of the allocated memory slots being interpreted as 64 bit floats)

Other important functions that return arrays are

```
julia> ones(2, 2)
2x2 Array{Float64,2}:
 1.0  1.0
 1.0  1.0

julia> fill("foo", 2, 2)
2x2 Array{ASCIIString,2}:
 "foo"  "foo"
 "foo"  "foo"
```

Manual Array Definitions As we've seen, you can create one dimensional arrays from manually specified data like so

```
julia> a = [10, 20, 30, 40]
4-element Array{Int64,1}:
 10
 20
 30
 40
```

In two dimensions we can proceed as follows

```
julia> a = [10 20 30 40]  # Two dimensional, shape is 1 x n
1x4 Array{Int64,2}:
 10  20  30  40

julia> ndims(a)
2

julia> a = [10 20; 30 40]  # 2 x 2
2x2 Array{Int64,2}:
 10  20
 30  40
```

You might then assume that `a = [10; 20; 30; 40]` creates a two dimensional column vector but unfortunately this isn't the case

```
julia> a = [10; 20; 30; 40]
4-element Array{Int64,1}:
 10
 20
 30
 40

julia> ndims(a)
1
```

Instead transpose the row vector

```
julia> a = [10 20 30 40]'  
4x1 Array{Int64,2}:  
10  
20  
30  
40  
  
julia> ndims(a)  
2
```

Array Indexing We've already seen the basics of array indexing

```
julia> a = collect(10:10:40)  
4-element Array{Int64,1}:  
10  
20  
30  
40  
  
julia> a[end-1]  
30  
  
julia> a[1:3]  
3-element Array{Int64,1}:  
10  
20  
30
```

For 2D arrays the index syntax is straightforward

```
julia> a = randn(2, 2)  
2x2 Array{Float64,2}:  
1.37556  0.924224  
1.52899  0.815694  
  
julia> a[1, 1]  
1.375559922478634  
  
julia> a[1, :]  # First row  
1x2 Array{Float64,2}:  
1.37556  0.924224  
  
julia> a[:, 1]  # First column  
2-element Array{Float64,1}:  
1.37556  
1.52899
```

Booleans can be used to extract elements

```
julia> a = randn(2, 2)  
2x2 Array{Float64,2}:  
-0.121311  0.654559
```

```
-0.297859  0.89208

julia> b = [true false; false true]
2x2 Array{Bool,2}:
 true  false
false  true

julia> a[b]
2-element Array{Float64,1}:
 -0.121311
 0.89208
```

This is useful for conditional extraction, as we'll see below

An aside: some or all elements of an array can be set equal to one number using slice notation

```
julia> a = Array(Float64, 4)
4-element Array{Float64,1}:
 1.30822e-282
 1.2732e-313
 4.48229e-316
 1.30824e-282

julia> a[2:end] = 42
42

julia> a
4-element Array{Float64,1}:
 1.30822e-282
 42.0
 42.0
 42.0
```

Passing Arrays As in Python, all arrays are passed by reference

What this means is that if `a` is an array and we set `b = a` then `a` and `b` point to exactly the same data

Hence any change in `b` is reflected in `a`

```
julia> a = ones(3)
3-element Array{Float64,1}:
 1.0
 1.0
 1.0

julia> b = a
3-element Array{Float64,1}:
 1.0
 1.0
 1.0

julia> b[3] = 44
```

```
44
```

```
julia> a
3-element Array{Float64,1}:
 1.0
 1.0
 44.0
```

If you are a MATLAB programmer perhaps you are recoiling in horror at this idea

But this is actually the more sensible default – after all, it's very inefficient to copy arrays unnecessarily

If you do need an actual copy in Julia, just use `copy()`

```
julia> a = ones(3)
3-element Array{Float64,1}:
 1.0
 1.0
 1.0

julia> b = copy(a)
3-element Array{Float64,1}:
 1.0
 1.0
 1.0

julia> b[3] = 44
44

julia> a
3-element Array{Float64,1}:
 1.0
 1.0
 1.0
```

Operations on Arrays

Array Methods Julia provides standard functions for acting on arrays, some of which we've already seen

```
julia> a = [-1, 0, 1]
3-element Array{Int64,1}:
 -1
  0
  1

julia> length(a)
3

julia> sum(a)
0
```

```

julia> mean(a)
0.0

julia> std(a)
1.0

julia> var(a)
1.0

julia> maximum(a)
1

julia> minimum(a)
-1

julia> b = sort(a, rev=true)    # Returns new array, original not modified
3-element Array{Int64,1}:
 1
 0
 -1

julia> b === a    # === tests if arrays are identical (i.e share same memory)
false

julia> b = sort!(a, rev=true)    # Returns *modified original* array
3-element Array{Int64,1}:
 1
 0
 -1

julia> b === a
true

```

Matrix Algebra For two dimensional arrays, `*` means matrix multiplication

```

julia> a = ones(1, 2)
1x2 Array{Float64,2}:
 1.0  1.0

julia> b = ones(2, 2)
2x2 Array{Float64,2}:
 1.0  1.0
 1.0  1.0

julia> a * b
1x2 Array{Float64,2}:
 2.0  2.0

julia> b * a'
2x1 Array{Float64,2}:
 2.0
 2.0

```

To solve the linear system $A \cdot X = B$ for X use $A \setminus B$

```
julia> A = [1 2; 2 3]
2x2 Array{Int64,2}:
 1  2
 2  3

julia> B = ones(2, 2)
2x2 Array{Float64,2}:
 1.0  1.0
 1.0  1.0

julia> A \ B
2x2 Array{Float64,2}:
 -1.0  -1.0
  1.0   1.0

julia> inv(A) * B
2x2 Array{Float64,2}:
 -1.0  -1.0
  1.0   1.0
```

Although the last two operations give the same result, the first one is numerically more stable and should be preferred in most cases

Multiplying two **one** dimensional vectors gives an error — which is reasonable since the meaning is ambiguous

```
julia> ones(2) * ones(2)
ERROR: `*` has no method matching *(::Array{Float64,1}, ::Array{Float64,1})
```

If you want an inner product in this setting use `dot()`

```
julia> dot(ones(2), ones(2))
2.0
```

Matrix multiplication using one dimensional vectors is a bit inconsistent — pre-multiplication by the matrix is OK, but post-multiplication gives an error

```
julia> b = ones(2, 2)
2x2 Array{Float64,2}:
 1.0  1.0
 1.0  1.0

julia> b * ones(2)
2-element Array{Float64,1}:
 2.0
 2.0

julia> ones(2) * b
ERROR: DimensionMismatch("*")
  in gemm_wrapper! at linalg/matmul.jl:275
  in * at linalg/matmul.jl:74
```

It's probably best to give your vectors dimension before you multiply them against matrices

Elementwise Operations

Algebraic Operations Suppose that we wish to multiply every element of matrix A with the corresponding element of matrix B

In that case we need to replace * (matrix multiplication) with .* (elementwise multiplication)

For example, compare

```
julia> ones(2, 2) * ones(2, 2)    # Matrix multiplication
2x2 Array{Float64,2}:
 2.0  2.0
 2.0  2.0

julia> ones(2, 2) .* ones(2, 2)    # Element by element multiplication
2x2 Array{Float64,2}:
 1.0  1.0
 1.0  1.0
```

This is a general principle: .x means apply operator x elementwise

```
julia> A = -ones(2, 2)
2x2 Array{Float64,2}:
 -1.0  -1.0
 -1.0  -1.0

julia> A.^2    # Square every element
2x2 Array{Float64,2}:
 1.0  1.0
 1.0  1.0
```

However in practice some operations are unambiguous and hence the . can be omitted

```
julia> ones(2, 2) + ones(2, 2)    # Same as ones(2, 2) .+ ones(2, 2)
2x2 Array{Float64,2}:
 2.0  2.0
 2.0  2.0
```

Scalar multiplication is similar

```
julia> A = ones(2, 2)
2x2 Array{Float64,2}:
 1.0  1.0
 1.0  1.0

julia> 2 * A    # Same as 2 .* A
2x2 Array{Float64,2}:
 2.0  2.0
 2.0  2.0
```

In fact you can omit the * altogether and just write 2A

Elementwise Comparisons Elementwise comparisons also use the .x style notation

```
julia> a = [10, 20, 30]
3-element Array{Int64,1}:
 10
 20
 30

julia> b = [-100, 0, 100]
3-element Array{Int64,1}:
 -100
   0
  100

julia> b .> a
3-element BitArray{1}:
 false
 false
 true

julia> a .== b
3-element BitArray{1}:
 false
 false
 false
```

We can also do comparisons against scalars with parallel syntax

```
julia> b
3-element Array{Int64,1}:
 -100
   0
  100

julia> b .> 1
3-element BitArray{1}:
 false
 false
 true
```

This is particularly useful for *conditional extraction* — extracting the elements of an array that satisfy a condition

```
julia> a = randn(4)
4-element Array{Float64,1}:
 0.0636526
 0.933701
 -0.734085
 0.531825

julia> a .< 0
4-element BitArray{1}:
 false
 false
 true
 false
```

```
julia> a[a .< 0]
1-element Array{Float64,1}:
 -0.734085
```

Vectorized Functions Julia provides standard mathematical functions such as `log`, `exp`, `sin`, etc.

```
julia> log(1.0)
0.0
```

By default, these functions act *elementwise* on arrays

```
julia> log(ones(4))
4-element Array{Float64,1}:
 0.0
 0.0
 0.0
 0.0
```

Functions that act elementwise on arrays in this manner are called **vectorized functions**

Note that we can get the same result as with a comprehension or more explicit loop

```
julia> [log(x) for x in ones(4)]
4-element Array{Float64,1}:
 0.0
 0.0
 0.0
 0.0
```

In Julia loops are typically fast and hence the need for vectorized functions is less intense than for some other high level languages

Nonetheless the syntax is convenient

Linear Algebra

Julia provides some a great deal of additional functionality related to linear operations

```
julia> A = [1 2; 3 4]
2x2 Array{Int64,2}:
 1  2
 3  4

julia> det(A)
-2.0

julia> trace(A)
5

julia> eigvals(A)
2-element Array{Float64,1}:
 -0.372281
```

```
5.37228
julia> rank(A)
2
```

For more details see the [linear algebra](#) section of the standard library

Exercises

Exercise 1 This exercise is on some matrix operations that arise in certain problems, including when dealing with linear stochastic difference equations

If you aren't familiar with all the terminology don't be concerned — you can skim read the background discussion and focus purely on the matrix exercise

With that said, consider the stochastic difference equation

$$X_{t+1} = AX_t + b + \Sigma W_{t+1} \quad (1.2)$$

Here

- X_t, b and X_{t+1} are $n \times 1$
- A is $n \times n$
- Σ is $n \times k$
- W_t is $k \times 1$ and $\{W_t\}$ is iid with zero mean and variance-covariance matrix equal to the identity matrix

Let S_t denote the $n \times n$ variance-covariance matrix of X_t

Using the rules for computing variances in matrix expressions, it can be shown from (1.2) that $\{S_t\}$ obeys

$$S_{t+1} = AS_t A' + \Sigma \Sigma' \quad (1.3)$$

It can be shown that, provided all eigenvalues of A lie within the unit circle, the sequence $\{S_t\}$ converges to a unique limit S

This is the **unconditional variance** or **asymptotic variance** of the stochastic difference equation

As an exercise, try writing a simple function that solves for the limit S by iterating on (1.3) given A and Σ

To test your solution, observe that the limit S is a solution to the matrix equation

$$S = ASA' + Q \quad \text{where} \quad Q := \Sigma \Sigma' \quad (1.4)$$

This kind of equation is known as a **discrete time Lyapunov equation**

The [QuantEcon package](#) provides a function called `solve_discrete_lyapunov` that implements a fast "doubling" algorithm to solve this equation

Test your iterative method against `solve_discrete_lyapunov` using matrices

$$A = \begin{bmatrix} 0.8 & -0.2 \\ -0.1 & 0.7 \end{bmatrix} \quad \Sigma = \begin{bmatrix} 0.5 & 0.4 \\ 0.4 & 0.6 \end{bmatrix}$$

Solutions

Solution notebook

Types, Methods and Performance

Contents

- *Types, Methods and Performance*
 - *Overview*
 - *Types*
 - *Defining Types and Methods*
 - *Writing Fast Code*
 - *Exercises*
 - *Solutions*

Overview

In this lecture we delve more deeply into the structure of Julia, and in particular into

- the concept of types
- building user defined types
- methods and multiple dispatch

These concepts relate to the way that Julia stores and acts on data

While they might be thought of as advanced topics, some understanding is necessary to

1. Read Julia code written by other programmers
2. Write “well organized” Julia code that’s easy to maintain and debug
3. Improve the speed at which your code runs

At the same time, don’t worry about following all the nuances on your first pass

If you return to these topics after doing some programming in Julia they will make more sense

Types

In Julia all objects (all “values” in memory) have a type, which can be queried using the `typeof()` function

```
julia> x = 42
42
```

```
julia> typeof(x)
Int64
```

Note here that the type resides with the object itself, not with the name `x`

The name `x` is just a symbol bound to an object of type `Int64`

Here we *rebind* it to another object, and now `typeof(x)` gives the type of that new object

```
julia> x = 42.0
42.0
```

```
julia> typeof(x)
Float64
```

Common Types We've already met many of the types defined in the core Julia language and its standard library

For numerical data, the most common types are integers and floats

For those working on a 64 bit machine, the default integers and floats are 64 bits, and are called `Int64` and `Float64` respectively (they would be `Int32` and `Float64` on a 32 bit machine)

There are many other important types, used for arrays, strings, iterators and so on

```
julia> typeof(1 + 1im)
Complex{Int64} (constructor with 1 method)

julia> typeof(linspace(0, 1, 100))
LinSpace{Float64}

julia> typeof(eye(2))
Array{Float64,2}

julia> typeof("foo")
ASCIIString

julia> typeof(1:10)
UnitRange{Int64}

julia> typeof('c') # Single character is a *Char*
Char
```

Type is important in Julia because it determines what operations will be performed on the data in a given situation

Moreover, if you try to perform an action that is unexpected for a given type the function call will usually fail

```
julia> 100 + "100"
ERROR: `+` has no method matching +(::Int64, ::ASCIIString)
```

Some languages will try to guess what the programmer wants here and return 200

Julia doesn't — in this sense, Julia is a “strongly typed” language

Type is important and it's up to the user to supply data in the correct form (as specified by type)

Methods and Multiple Dispatch Looking more closely at how this works brings us to a very important topic concerning Julia's data model — methods and multiple dispatch

Let's look again at the error message

```
julia> 100 + "100"
ERROR: `+` has no method matching +(::Int64, ::ASCIIString)
```

As discussed earlier, the operator `+` is just a function, and we can rewrite that call using functional notation to obtain exactly the same result

```
julia> +(100, "100")
ERROR: `+` has no method matching +(::Int64, ::ASCIIString)
```

Multiplication is similar

```
julia> 100 * "100"
ERROR: `*` has no method matching *(::Int64, ::ASCIIString)

julia> *(100, "100")
ERROR: `*` has no method matching *(::Int64, ::ASCIIString)
```

What the message tells us is that `*(a, b)` doesn't work when `a` is an integer and `b` is a string

In particular, the function `*` has no *matching method*

In essence, a **method** in Julia is a version of a function that acts on a particular tuple of data types

For example, if `a` and `b` are integers then a method for multiplying integers is invoked

```
julia> *(100, 100)
10000
```

On the other hand, if `a` and `b` are strings then a method for string concatenation is invoked

```
julia> *("foo", "bar")
"foobar"
```

In fact we can see the precise methods being invoked by applying `@which`

```
julia> @which *(100, 100)
*(x::Int64,y::Int64) at int.jl:47

julia> @which *("foo", "bar")
*(s1::AbstractString, ss::AbstractString...) at strings/basic.jl:50
```

We can see the same process with other functions and their methods

```
julia> isfinite(1.0) # Call isfinite on a float
true

julia> @which isfinite(1)
isfinite(x::Integer) at float.jl:311
```

```
julia> @which isfinite(1.0)
isfinite(x::AbstractFloat) at float.jl:309
```

Here `isfinite()` is a *function* with multiple *methods*

It has a method for acting on floating points and another method for acting on integers

In fact it has quite a few methods

```
julia> methods(isfinite)
# 11 methods for generic function "isfinite":
isfinite(x::Float16) at float16.jl:119
isfinite(x::BigFloat) at mpfr.jl:790
isfinite(x::AbstractFloat) at float.jl:309
isfinite(x::Integer) at float.jl:311
isfinite(::Irrational{sym}) at irrationals.jl:66
isfinite(x::Real) at float.jl:310
isfinite(z)::Complex{T<:Real}) at complex.jl:44
isfinite{T<:Number}(:AbstractArray{T<:Number,1}) at operators.jl:380
...
```

The particular method being invoked depends on the data type on which the function is called

We'll discuss some of the more complicated data types you see towards the end of this list as we go along

Abstract Types Looking at the list of methods above you can see references to types that we haven't met before, such as `Real` and `Number`

These two types are examples of what are known in Julia as **abstract types**

Abstract types serve a different purpose to **concrete types** such as `Int64` and `Float64`

To understand what that purpose is, imagine that you want to write a function with two methods, one to handle real numbers and the other for complex numbers

As we know, there are multiple types for real numbers, such as integers and floats

There are even multiple integer and float types, such as `Int32`, `Int64`, `Float32`, `Float64`, etc.

If we want to handle all of these types for real numbers in the same way, it's useful to have a "parent" type called `Real`

Rather than writing a separate method for each concrete type, we can just write one for the abstract `Real` type

In this way, the purpose of abstract types is to serve as a unifying parent type that concrete types can "inherit" from

Indeed, we can see that `Float64`, `Int64`, etc. are **subtypes** of `Real` as follows

```
julia> Float64 <: Real
true
```

```
julia> Int64 <: Real
true
```

On the other hand, 64 bit complex numbers are not reals

```
julia> Complex64 <: Real
false
```

They are, however, subtypes of Number

```
julia> Complex64 <: Number
true
```

Number in turn is a subtype of Any, which is a parent of all types

```
julia> Number <: Any
true
```

Type Hierarchy In fact the types form a hierarchy, with Any at the top of the tree and the concrete types at the bottom

Note that we never see *instances* of abstract types

For example, if x is a value then `typeof(x)` will never return an abstract type

The point of abstract types is simply to categorize the concrete types (as well as other abstract types that sit below them in the hierarchy)

On the other hand, we cannot subtype concrete types

While we can build subtypes of abstract types we cannot do the same for concrete types

Multiple Dispatch We can now be a little bit clearer about what happens when you call a function on given types

Suppose we execute the function call `f(a, b)` where a and b are of concrete types S and T respectively

The Julia interpreter first queries the types of a and b to obtain the tuple (S, T)

It then parses the list of methods belonging to f, searching for a match

If it finds a method matching (S, T) it calls that method

If not, it looks to see whether the pair (S, T) matches any method defined for *immediate parent types*

For example, if S is Float64 and T is Complex64 then the immediate parents are AbstractFloat and Number respectively

```
julia> super(Float64)
AbstractFloat

julia> super(Complex64)
Number
```

Hence the interpreter looks next for a method of the form `f(x::AbstractFloat, y::Number)`

If the interpreter can't find a match in immediate parents (supertypes) it proceeds up the tree, looking at the parents of the last type it checked at each iteration

- If it eventually finds a matching method it invokes that method
- If not, we get an error

This is the process that leads to the error that we saw above:

```
julia> *(100, "100")
ERROR: `*` has no method matching *(::Int64, ::ASCIIString)
```

The procedure of matching data to appropriate methods is called **dispatch**

Because the procedure starts from the concrete types and works upwards, dispatch always invokes the *most specific method* that is available

For example, if you have methods for function `f` that handle

1. (`Float64, Int64`) pairs
2. (`Number, Number`) pairs

and you call `f` with `f(0.5, 1)` then the first method will be invoked

This makes sense because (hopefully) the first method is designed to work well with exactly this kind of data

The second method is probably more of a “catch all” method that handles other data in a less optimal way

Defining Types and Methods

Let's look at defining our own methods and data types, including composite data types

User Defined Methods It's straightforward to add methods to existing functions or functions you've defined

In either case the process is the same:

- use the standard syntax to define a function of the same name
- but specify the data type for the method in the function signature

For example, we saw above that `+` is just a function with various methods

- recall that `a + b` and `+(a, b)` are equivalent

We saw also that the following call fails because it lacks a matching method

```
julia> +(100, "100")
ERROR: `+` has no method matching +(::Int64, ::ASCIIString)
```

This is sensible behavior, but if you want to change it by defining a method to handle the case in question there's nothing to stop you:

```
julia> importall Base.Operators

julia> +(x::Integer, y::ASCIIString) = x + parse(Int, y)
+ (generic function with 172 methods)

julia> +(100, "100")
200

julia> 100 + "100"
200
```

Here's another example, involving a user defined function

We begin with a file called `test.jl` in the present working directory with the following content

```
function f(x)
    println("Generic function invoked")
end

function f(x::Number)
    println("Number method invoked")
end

function f(x::Integer)
    println("Integer method invoked")
end
```

Clearly these methods do nothing more than tell you which method is being invoked

Let's now run this and see how it relates to our discussion of method dispatch above

```
julia> include("test.jl")
f (generic function with 3 methods)

julia> f(3)
Integer method invoked

julia> f(3.0)
Number method invoked

julia> f("foo")
Generic function invoked
```

Since 3 is an `Int64` and `Int64 <: Integer <: Number`, the call `f(3)` proceeds up the tree to `Integer` and invokes `f(x::Integer)`

On the other hand, 3.0 is a `Float64`, which is not a subtype of `Integer`

Hence the call `f(3.0)` continues up to `f(x::Number)`

Finally, `f("foo")` is handled by the generic function, since it is not a subtype of `Number`

User Defined Types Most languages have facilities for creating new data types and Julia is no exception

```
julia> type Foo end
julia> foo = Foo()
Foo()

julia> typeof(foo)
Foo (constructor with 1 method)
```

Let's make some observations about this code

First note that to create a new data type we use the keyword `type` followed by the name

- By convention, type names use CamelCase (e.g., `FloatingPoint`, `Array`, `AbstractArray`)

When a new data type is created in this way, the interpreter simultaneously creates a *default constructor* for the data type

This constructor is a function for generating new instances of the data type in question

It has the same name as the data type but uses function call notion — in this case `Foo()`

In the code above, `foo = Foo()` is a call to the default constructor

A new instance of type `Foo` is created and the name `foo` is bound to that instance

Now if we want to we can create methods that act on instances of `Foo`

Just for fun, let's define how to add one `Foo` to another

```
julia> +(x::Foo, y::Foo) = "twofoos"
+ (generic function with 126 methods)

julia> foo1, foo2 = Foo(), Foo()  # Create two Foos
(Foo(), Foo())

julia> +(foo1, foo2)
"twofoos"

julia> foo1 + foo2
"twofoos"
```

We can also create new functions to handle `Foo` data

```
julia> foofunc(x::Foo) = "onefoo"
foofunc (generic function with 1 method)

julia> foofunc(foo)
"onefoo"
```

This example isn't of much use but more useful examples follow

Composite Data Types Since the common primitive data types are already built in, most new user-defined data types are composite data types

Composite data types are data types that contain distinct fields of data as attributes

For example, let's say we are doing a lot of work with AR(1) processes, which are random sequences $\{X_t\}$ that follow a law of motion of the form

$$X_{t+1} = aX_t + b + \sigma W_{t+1} \quad (1.5)$$

Here a , b and σ are scalars and $\{W_t\}$ is an iid sequence of shocks with some given distribution ϕ

At times it might be convenient to take these primitives a , b , σ and ϕ and organize them into a single entity like so

```
type AR1
    a
    b
    sigma
    phi
end
```

For the distribution ϕ we'll assign a `Distribution` from the `Distributions` package

After reading in the `AR1` definition above we can do the following

```
julia> using Distributions

julia> m = AR1(0.9, 1, 1, Beta(5, 5))
AR1(0.9,1,1,Beta( alpha=5.0 beta=5.0 ))
```

In this call to the constructor we've created an instance of `AR1` and bound the name `m` to it

We can access the fields of `m` using their names and "dotted attribute" notation

```
julia> m.a
0.9

julia> m.b
1

julia> m.sigma
1

julia> m.phi
Beta( alpha=5.0 beta=5.0 )
```

For example, the attribute `m.phi` points to an instance of `Beta`, which is in turn a subtype of `Distribution` as defined in the `Distributions` package

```
julia> typeof(m.phi)
Beta (constructor with 3 methods)

julia> typeof(m.phi) <: Distribution
true
```

We can reach in to `m` and change this if we want to

```
julia> m.phi = Exponential(0.5)
Exponential( scale=0.5 )
```

In our type definition we can be explicit that we want phi to be a Distribution, and the other elements to be real scalars

```
type AR1
    a::Real
    b::Real
    sigma::Real
    phi::Distribution
end
```

(Before reading this in you might need to restart your REPL session in order to clear the old definition of AR1 from memory)

Now the constructor will complain if we try to use the wrong data type

```
julia> m = AR1(0.9, 1, "foo", Beta(5, 5))
ERROR: `convert` has no method matching convert(::Type{Real}, ::ASCIIString) in AR1 at no file
```

This is useful if we're going to have functions that act on instances of AR1

- e.g., simulate time series, compute variances, generate histograms, etc.

If those functions only work with AR1 instances built from the specified data types then it's probably best if we get an error as soon we try to make an instance that doesn't fit the pattern

Better to fail early rather than deeper into our code where errors are harder to debug

Type Parameters Consider the following output

```
julia> typeof([10, 20, 30])
Array{Int64,1}
```

Here Array is one of Julia's predefined types (Array <: DenseArray <: AbstractArray <: Any)

The Int64,1 in curly brackets are **type parameters**

In this case they are the element type and the dimension

Many other types have type parameters too

```
julia> typeof(1.0 + 1.0im)
Complex{Float64} (constructor with 1 method)

julia> typeof(1 + 1im)
Complex{Int64} (constructor with 1 method)
```

Types with parameters are therefore in fact an indexed family of types, one for each possible value of the parameter

Defining Parametric Types We can use parametric types in our own type definitions

Let's say we're defining a type called FooBar with attributes foo and bar

```
type FooBar
    foo
    bar
end
```

Suppose we now decide that we want foo and bar to have the same type, although we don't much care what that type is

We can achieve this with the syntax

```
type FooBar{T}
    foo::T
    bar::T
end
```

Now our constructor is happy provided that the arguments do in fact have the same type

```
julia> fb = FooBar(1.0, 2.0)
FooBar{Float64}(1.0, 2.0)

julia> fb = FooBar(1, 2)
FooBar{Int64}(1, 2)

julia> fb = FooBar(1, 2.0)
ERROR: `FooBar{T}` has no method matching FooBar{T}(::Int64, ::Float64)
```

Now let's say we want the data to be of the same type *and* that type must be a subtype of Number

We can achieve this as follows

```
type FooBar{T <: Number}
    foo::T
    bar::T
end
```

Let's try it

```
julia> fb = FooBar(1, 2)
FooBar{Int64}(1, 2)

julia> fb = FooBar("fee", "fi")
ERROR: `FooBar{T<:Number}` has no method matching FooBar{T<:Number}(::ASCIIString, ::ASCIIString)
```

In the second instance we get an error because ASCIIString is not a subtype of Number

Writing Fast Code

Let's briefly discuss how to write Julia code that executes quickly (for a given hardware configuration)

For now our focus is on generating more efficient machine code from essentially the same program (i.e., without parallelization or other more significant changes to the way the program runs)

Basic Concepts The benchmark for performance is well written *compiled* code, expressed in languages such as C and Fortran

This is because computer programs are essentially operations on data, and the details of the operations implemented by the CPU depend on the nature of the data

When code is written in a language like C and compiled, the compiler has access to sufficient information to build machine code that will organize the data optimally in memory and implement efficient operations as required for the task in hand

To approach this benchmark, Julia needs to know about the type of data it's processing as early as possible

An Example Consider the following function, which essentially does the same job as Julia's `sum()` function but acts only on floating point data

```
function sum_float_array(x::Array{Float64, 1})
    sum = 0.0
    for i in 1:length(x)
        sum += x[i]
    end
    return sum
end
```

Calls to this function run very quickly

```
julia> x = linspace(0, 1, 1e6)
linspace(0.0,1.0,1000000)

julia> x = collect(x); # Convert to array of Float64s

julia> typeof(x)
Array{Float64,1}

julia> @time sum_float_array(x)
0.001800 seconds (149 allocations: 10.167 KB)
499999.9999999796
```

One reason is that data types are fully specified

When Julia compiles this function via its just-in-time compiler, it knows that the data passed in as `x` will be an array of 64 bit floats

Hence it's known to the compiler that the relevant method for `+` is always addition of floating point numbers

Moreover, the data can be arranged into continuous 64 bit blocks of memory to simplify memory access

Finally, data types are stable — for example, the local variable `sum` starts off as a float and remains a float throughout

Type Inferences What happens if we don't supply type information?

Here's the same function minus the type annotation in the function signature

```
function sum_array(x)
    sum = 0.0
    for i in 1:length(x)
        sum += x[i]
    end
    return sum
end
```

When we run it with the same array of floating point numbers it executes at a similar speed as the function with type information

```
julia> @time sum_array(x)
0.001949 seconds (5 allocations: 176 bytes)
```

The reason is that when `sum_array()` is first called on a vector of a given data type, a newly compiled version of the function is produced to handle that type

In this case, since we're calling the function on a vector of floats, we get a compiled version of the function with essentially the same internal representation as `sum_float_array()`

Things get tougher for the interpreter when the data type within the array is imprecise

For example, the following snippet creates an array where the element type is `Any`

```
julia> x = Any[1/i for i in 1:1e6];
julia> eltype(x)
Any
```

Now summation is much slower and memory management is less efficient

```
julia> @time sum_array(x)
0.058874 seconds (1.00 M allocations: 15.259 MB, 41.67% gc time)
```

Summary and Tips To write efficient code use functions to segregate operations into logically distinct blocks

Data types will be determined at function boundaries

If types are not supplied then they will be inferred

If types are stable and can be inferred effectively your functions will run fast

Further Reading There are many other aspects to writing fast Julia code

A good next stop for further reading is the [relevant part](#) of the Julia documentation

Exercises

Exercise 1 Write a function with the signature `simulate(m::AR1, n::Integer, x0::Real)` that takes as arguments

- an instance `m` of `AR1`
- an integer `n`
- a real number `x0`

and returns an array containing a time series of length `n` generated according to (1.5) where

- the primitives of the AR(1) process are as specified in `m`
- the initial condition X_0 is set equal to `x0`

Here `AR1` is as defined above:

```
type AR1
    a::Real
    b::Real
    sigma::Real
    phi::Distribution
end
```

Hint: If `d` is an instance of `Distribution` then `rand(d)` generates one random draw from the distribution specified in `d`

Exercise 2 The term **universal function** is sometimes applied to functions which

- when called on a scalar return a scalar
- when called on an array of scalars return an array of the same length by acting elementwise on the scalars in the array

For example, `sin()` has this property in Julia

```
julia> sin(pi)
1.2246467991473532e-16

julia> sin([pi, 2pi])
2-element Array{Float64,1}:
 1.22465e-16
 -2.44929e-16
```

Write a universal function `f` such that

- `f(k)` returns a chi-squared random variable with `k` degrees of freedom when `k` is an integer
- `f(k_vec)` returns a vector where `f(k_vec)[i]` is chi-squared with `k_vec[i]` degrees of freedom

Hint: If we take `k` independent standard normals, square them all and sum we get a chi-squared with `k` degrees of freedom

Solutions

[Solution notebook](#)

Plotting in Julia

Contents

- *Plotting in Julia*
 - *Overview*
 - *PyPlot*
 - *PlotlyJS*
 - *Plots.jl*

Overview

Since it's inception, plotting in Julia has been a mix of happiness and frustration

Some initially promising libraries have stagnated, or failed to keep up with user needs

New packages have appeared to compete with them, but not all are fully featured

The good news is that the Julia community now has several very good options for plotting

In this lecture we'll try to save you some of our pain by focusing on what we believe are currently the best libraries

First we look at two high quality plotting packages that have proved useful to us in a range of applications

After that we turn to a relative newcomer called [Plots.jl](#)

The latter package takes a different – and intriguing – approach that combines and exploits the strengths of several existing plotting libraries

Below we assume that

- you've already read through [our getting started lecture](#)
- you are working in a Jupyter notebook, as described [here](#)

How to Read this Lecture If you want to get started quickly with relatively simple plots, you can skip straight to the [section on Plots.jl](#)

If you want a deeper understanding and more flexibility, continue from the next section and read on

Credits: Thanks to [@albep](#), [@vgregory757](#) and [@spencerlyon2](#) for help with the code examples below

PyPlot

Let's look at [PyPlot](#) first

PyPlot is a Julia front end to the excellent Python plotting library [Matplotlib](#)

Installing PyPlot One disadvantage of PyPlot is that it not only requires Python but also much of the scientific Python stack

Fortunately, installation of the latter has been greatly simplified by the excellent Anaconda Python distribution

Moreover, the tools that come with Anaconda (such as Jupyter) are too good to miss out on

So please go ahead and *install Anaconda* if you haven't yet

Next, start up Julia and type `Pkg.add("PyPlot")`

Usage There are two different interfaces to Matplotlib and hence to PyPlot

Let's look at them in turn

The Procedural API Matplotlib has a straightforward plotting API that essentially replicates the plotting routines in MATLAB

These plotting routines can be expressed in Julia with almost identical syntax

Here's an example

```
using PyPlot
x = linspace(0, 10, 200)
y = sin(x)
plot(x, y, "b-", linewidth=2)
```

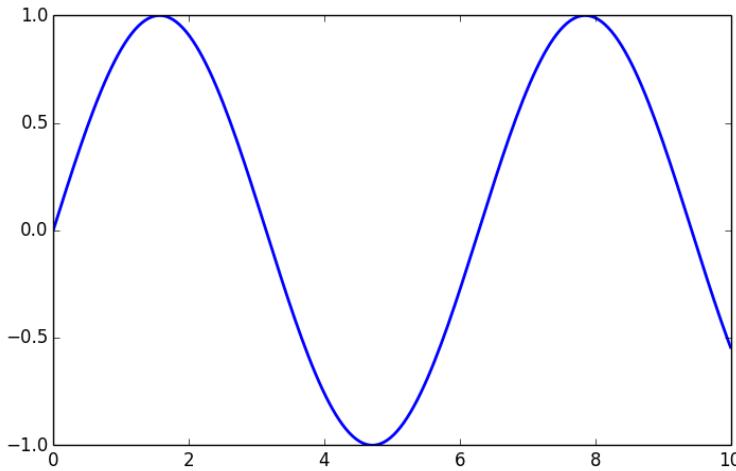
The resulting figure looks as follows

The Object Oriented API Matplotlib also has a more powerful and expressive object oriented API

Because Julia isn't object oriented in the same sense as Python, the syntax required to access this interface via PyPlot is a little awkward

Here's an example:

```
using PyPlot
x = linspace(0, 10, 200)
y = sin(x)
fig, ax = subplots()
ax[:plot](x, y, "b-", linewidth=2)
```



The resulting figure is the same

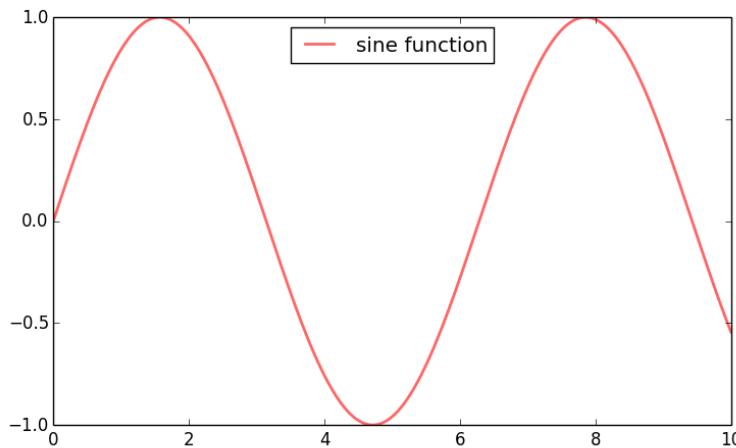
Here we get no particular benefit from switching APIs, while introducing a less attractive syntax

However, as plots get more complex, the more explicit syntax will give us greater control

Here's a similar plot with a bit more customization

```
using PyPlot
x = linspace(0, 10, 200)
y = sin(x)
fig, ax = subplots()
ax[:plot](x, y, "r-", linewidth=2, label="sine function", alpha=0.6)
ax[:legend](loc="upper center")
```

The resulting figure has a legend at the top center

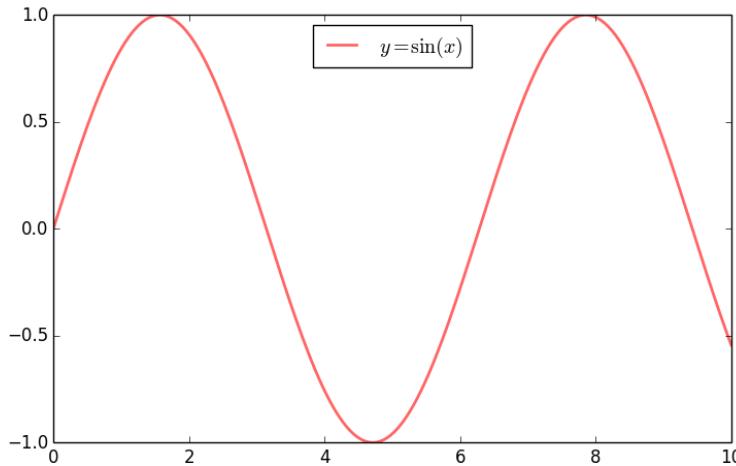


We can render the legend in LaTeX by changing the `ax[:plot]` line to

```
ax[:plot](x, y, "r-", linewidth=2, label=L"$y = \sin(x)$", alpha=0.6)
```

Note the L in front of the string to indicate LaTeX mark up

The result looks as follows



Multiple Plots on One Axis Here's another example, which helps illustrate how to put multiple plots on one figure

We use *Distributions.jl* to get the values of the densities given a randomly generated mean and standard deviation

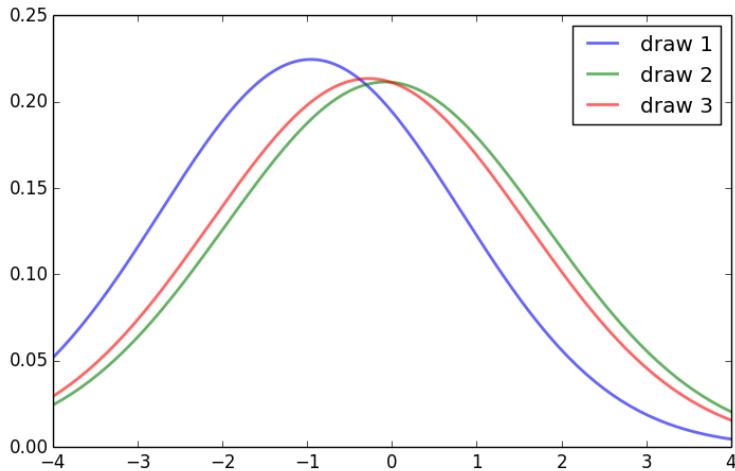
```
using PyPlot
using Distributions

u = Uniform()

fig, ax = subplots()
x = linspace(-4, 4, 150)
for i in 1:3
    # == Compute normal pdf from randomly generated mean and std ==
    m, s = rand(u) * 2 - 1, rand(u) + 1
    d = Normal(m, s)
    y = pdf(d, x)
    # == Plot current pdf ==
    ax[:plot](x, y, linewidth=2, alpha=0.6, label="draw $i")
end
ax[:legend]()
```

It generates the following plot

Subplots A figure containing n rows and m columns of subplots can be created by the call



```
fig, axes = subplots(num_rows, num_cols)
```

Here's an example that generates 6 normal distributions, takes 100 draws from each, and plots each of the resulting histograms

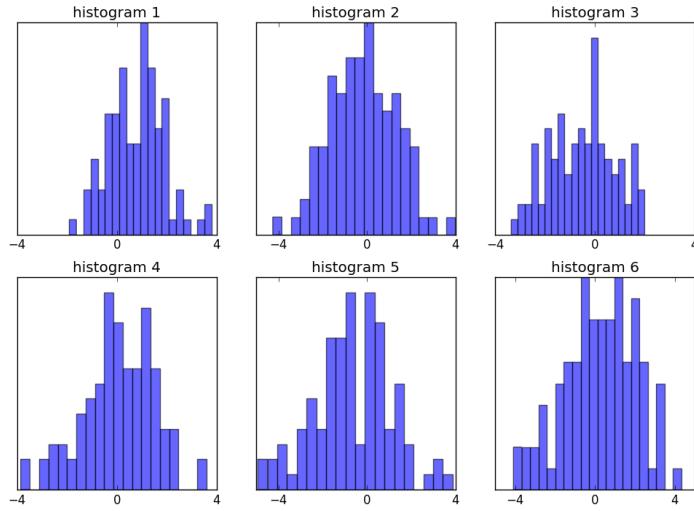
```
using PyPlot
using Distributions

u = Uniform()
num_rows, num_cols = 2, 3
fig, axes = subplots(num_rows, num_cols, figsize=(12, 8))
subplot_num = 0

for i in 1:num_rows
    for j in 1:num_cols
        ax = axes[i, j]
        subplot_num += 1
        # == Generate a normal sample with random mean and std ==
        m, s = rand(u) * 2 - 1, rand(u) + 1
        d = Normal(m, s)
        x = rand(d, 100)
        # == Histogram the sample ==
        ax[:hist](x, alpha=0.6, bins=20)
        ax[:set_title]("histogram $subplot_num")
        ax[:set_xticks]([-4, 0, 4])
        ax[:set_yticks]()
    end
end
```

The resulting figure is as follows

3D Plots Here's an example of how to create a 3D plot



```
using PyPlot
using Distributions
using QuantEcon: meshgrid

n = 50
x = linspace(-3, 3, n)
y = x

z = Array(Float64, n, n)
f(x, y) = cos(x^2 + y^2) / (1 + x^2 + y^2)
for i in 1:n
    for j in 1:n
        z[j, i] = f(x[i], y[j])
    end
end

fig = figure(figsize=(8,6))
ax = fig[:gca](projection="3d")
ax[:set_zlim](-0.5, 1.0)
xgrid, ygrid = meshgrid(x, y)
ax[:plot_surface](xgrid, ygrid, z, rstride=2, cstride=2, cmap=ColorMap("jet"), alpha=0.7, linewidth=0.25)
```

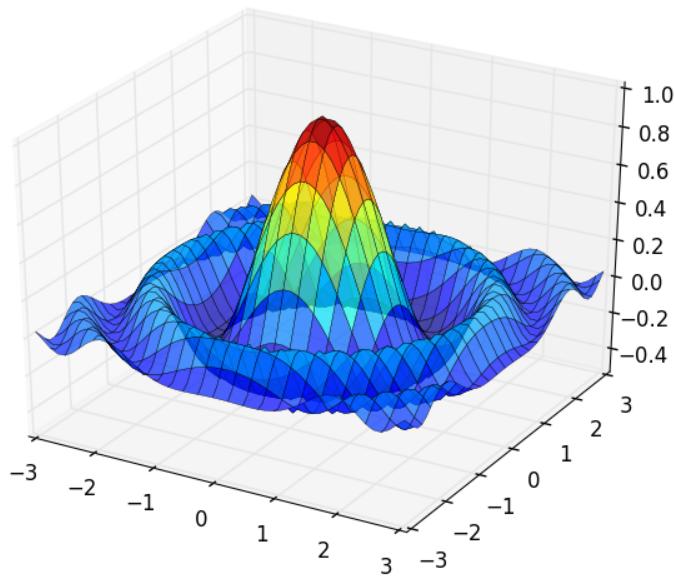
It creates this figure

PlotlyJS

Now let's turn to another plotting package — a promising new library called [PlotlyJS](#), authored by [Spencer Lyon](#)

PlotlyJS is a Julia interface to the [plotly.js visualization library](#)

It can be installed by typing `Pkg.add("PlotlyJS")` from within Julia



It has several advantages, one of which is beautiful interactive plots

While we won't treat the interface in great detail, we will frequently use PlotlyJS as a backend for Plots.jl

(More on this below)

Examples Let's look at some simple examples

Here's a version of the sine function plot you saw above

```
using PlotlyJS
x = linspace(0, 10, 200)
y = sin(x)
plot(scatter(x=x, y=y, marker_color="blue", line_width=2))
```

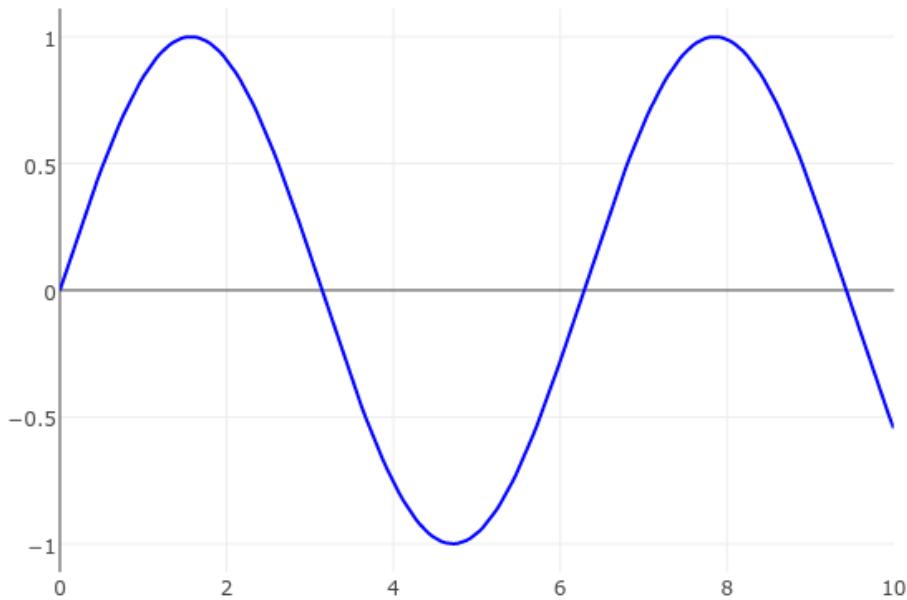
Here's the resulting figure:

Here's a replication of the *figure with multiple Gaussian densities*

```
using PlotlyJS
using Distributions

traces = GenericTrace[]
u = Uniform()

x = linspace(-4, 4, 150)
for i in 1:3
    # == Compute normal pdf from randomly generated mean and std == #
    traces
```



```
m, s = rand(u) * 2 - 1, rand(u) + 1
d = Normal(m, s)
y = pdf(d, x)
trace = scatter(x=x, y=y, name="draw $i")
push!(traces, trace)
end

plot(traces, Layout())
```

The output looks like this (modulo randomness):

Plots.jl

[Plots.jl](#) is another relative newcomer to the Julia plotting scene, authored by [Tom Breloff](#)

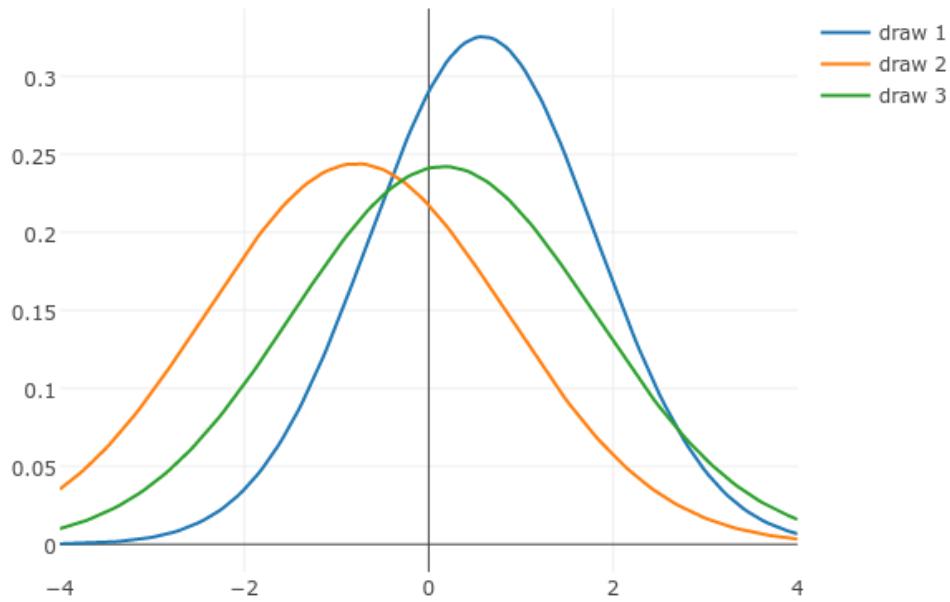
The approach of Plots.jl is to

1. provide a “frontend” plotting language
2. render the plots by using one of several existing plotting libraries as “backends”

In other words, Plots.jl plotting commands are translated internally to commands understood by a selected plotting library

Underlying libraries, or backends, can be swapped very easily

This is neat because each backend has a different look, as well as different capabilities



Also, Julia being Julia, it's quite possible that a given backend won't install or function on your machine at a given point in time

With Plots.jl, you can just change to another one

Simple Examples We produced some simple plots using Plots.jl back in our introductory Julia lecture

Here's another simple one:

```
using Plots
x = linspace(0, 10, 200)
y = sin(x)
plot(x, y, color=:blue, linewidth=2, label="sine")
```

On our machine this produces the following figure

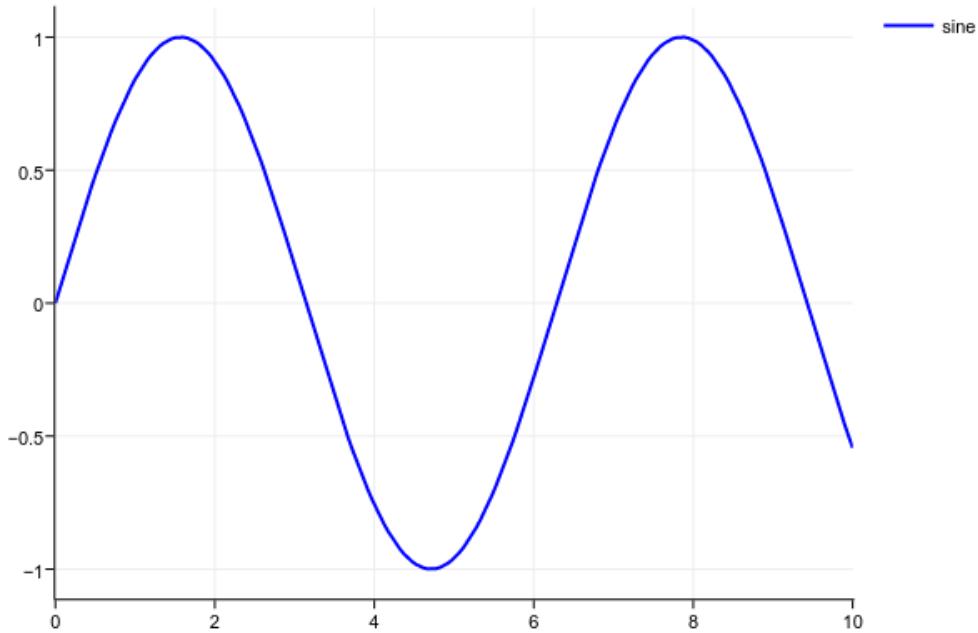
No backend was specified in the preceding code, and in this case it defaulted to PlotlyJS.jl

We can make this explicit by adding one extra line

```
using Plots
plotlyjs()      # specify backend
x = linspace(0, 10, 200)
y = sin(x)
plot(x, y, color=:blue, linewidth=2, label="sine")
```

To switch your backend to PyPlot, change `plotlyjs()` to `pyplot()`

Your figure should now look more like the plots produced by PyPlot



Here's a slightly more complex plot using Plots.jl with PyPlot backend

```
using Plots
using LaTeXStrings      # Install this package
pyplot()
x = linspace(0, 10, 100)
plot(x,
    sin,
    color=:red,
    lw=2,
    yticks=-1:1:1,
    title="sine function",
    label=L"$y = \sin(x)$",  # L for LaTeX string
    alpha=0.6)
```

Here's the figure it produces:

Use `legend=:none` if you want no legend on the plot

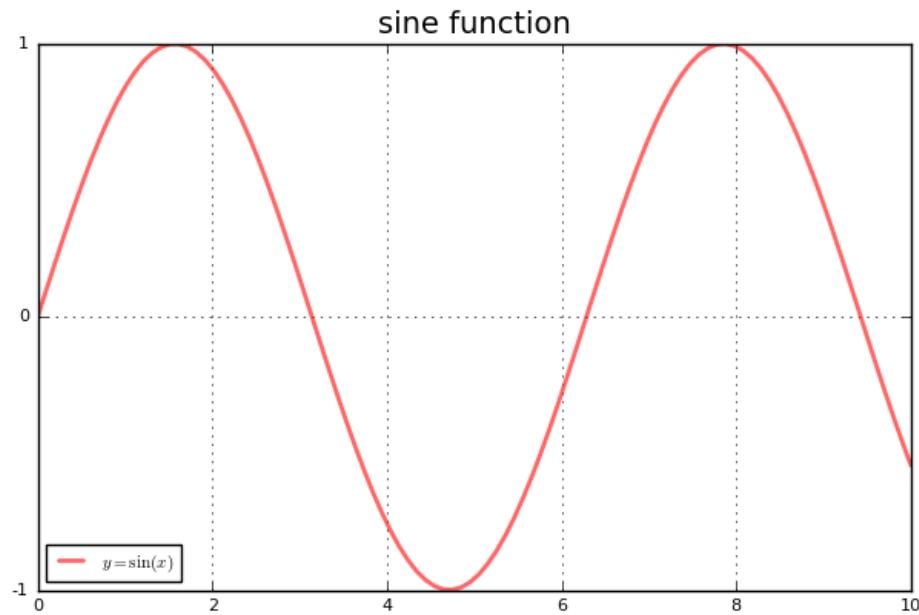
Notice that in the preceding code example, the second argument to `plot()` is a function rather than an array of data points

This is valid syntax, as is

```
plot(sin, 0, 10) # Plot the sine function from 0 to 10
```

Plots.jl accommodates these useful variations in syntax by exploiting multiple dispatch

Multiple Plots on One Axis Next, let's replicate the *figure with multiple Gaussian densities*



```
using Distributions
using Plots
plotlyjs()

x = linspace(-4, 4, 150)
y_vals = []
labels = []
for i = 1:3
    m, s = 2*(rand() - 0.5), rand() + 1
    d = Normal(m, s)
    push!(y_vals, pdf(d, x))
    l = string("mu = ", round(m, 2))
    push!(labels, l)
end

plot(x, y_vals, linewidth=2, alpha=0.6, label=labels')
```

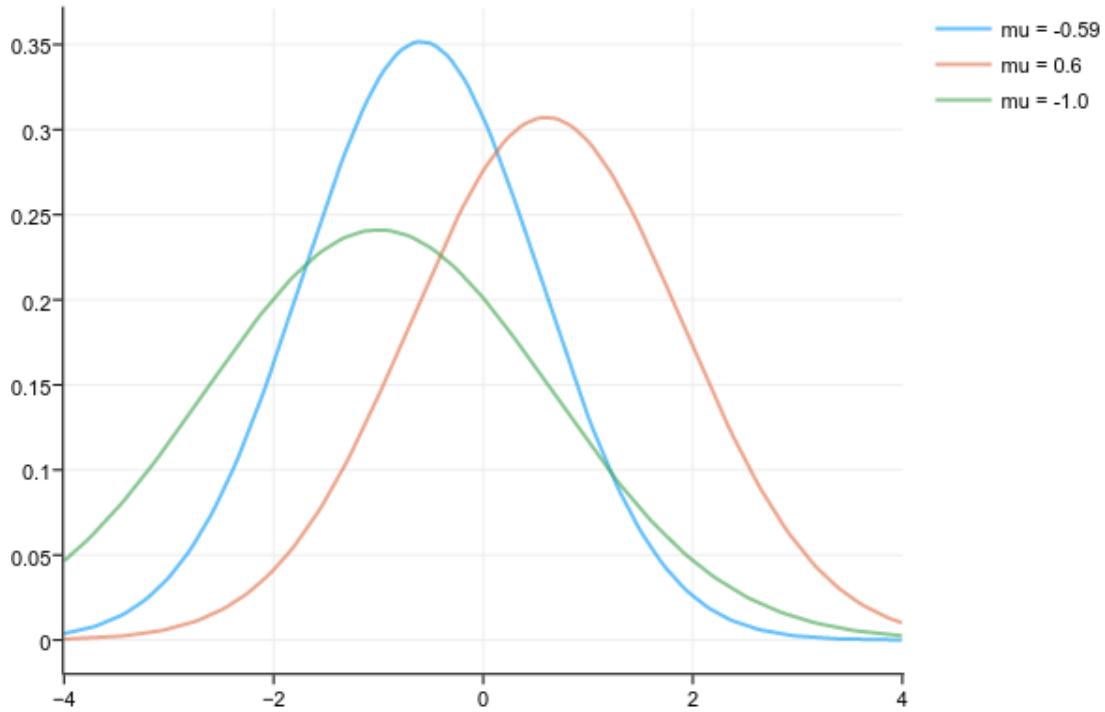
Notice that the *labels* vector is transposed because *Plots.jl* will interpret each row as a separate label
Also, when you have multiple y-series, *Plots.jl* can accept one x-values vector and apply it to each y-series

Here's the resulting figure:

Subplots Let's replicate the subplots figure shown above

```
using Distributions
using Plots
using LaTeXStrings
pyplot()

draws = []
```



```

titles = []
for i = 1:6
    m, s = 2*(rand() - 0.5), rand() + 1
    d = Normal(m, s)
    push!(draws, rand(d, 100))
    t = string(L"\mu = $", round(m, 2), L", \sigma = $", round(s, 2))
    push!(titles, t)
end

histogram(draws,
          layout=6,
          title=titles',
          legend=:none,
          titlefont=font(9),
          bins=20)

```

Notice that the font and bins settings get applied to each subplot

Here's the resulting figure:

When you want to pass individual arguments to subplots, you can use a row vector of arguments

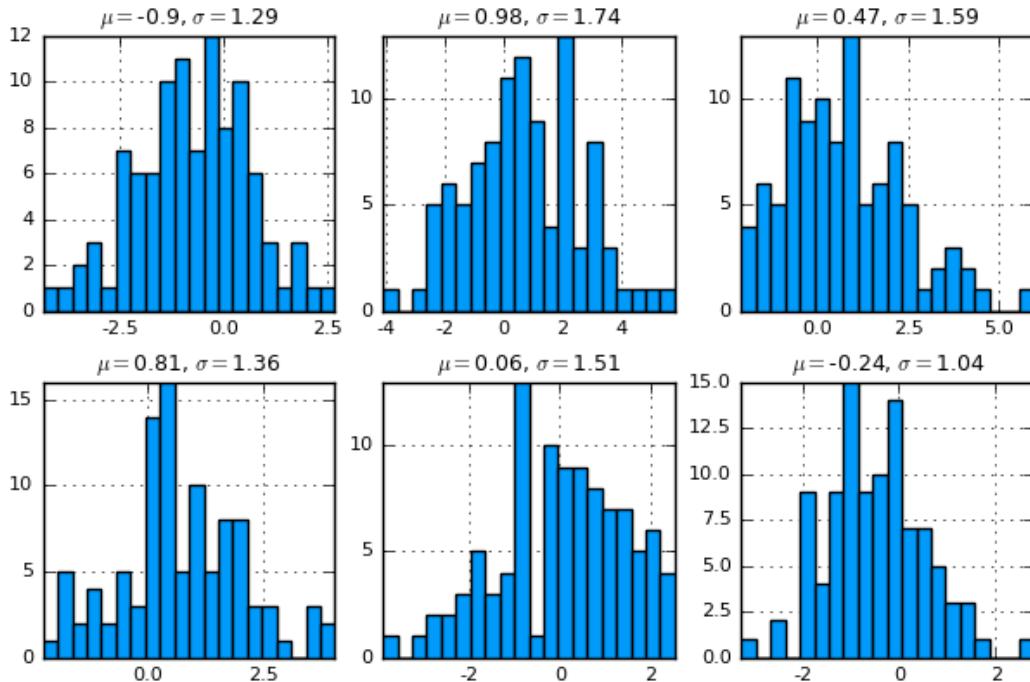
- For example, in the preceding code, `titles'` is a 1×6 row vector

Here's another example of this, with a row vector of different colors for the histograms

```

using Distributions
using Plots
using LaTeXStrings
pyplot()

```



```

draws = []
titles = []
for i = 1:6
    m, s = 2*(rand() - 0.5), rand() + 1
    d = Normal(m, s)
    push!(draws, rand(d, 100))
    t = string(L"\mu = $", round(m, 2), L", \sigma = $", round(s, 2))
    push!(titles, t)
end

histogram(draws,
    layout=6,
    title=titles',
    legend=:none,
    titlefont=font(9),
    color=[:red :blue :yellow :green :black :purple],
    bins=20)

```

The result is a bit garish but hopefully the message is clear

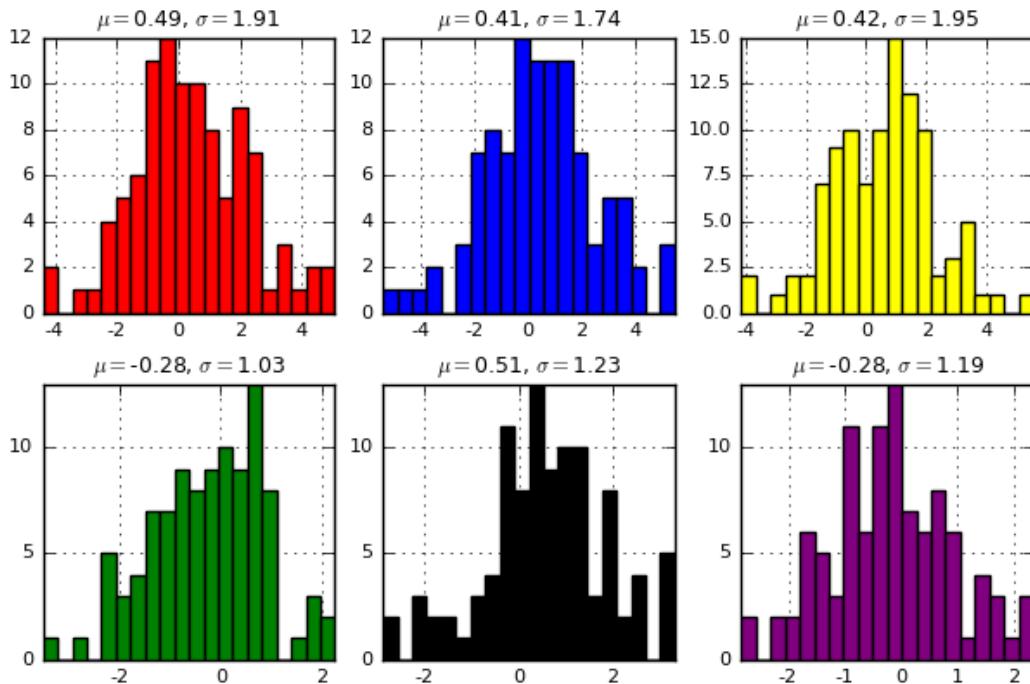
3D Plots Here's a sample 3D plot

```

using Plots
plotlyjs()

n = 50
x = linspace(-3, 3, n)
y = x

```



```

z = Array(Float64, n, n)
f(x, y) = cos(x^2 + y^2) / (1 + x^2 + y^2)
for i in 1:n
    for j in 1:n
        z[j, i] = f(x[i], y[j])
    end
end

surface(x, y, z)

```

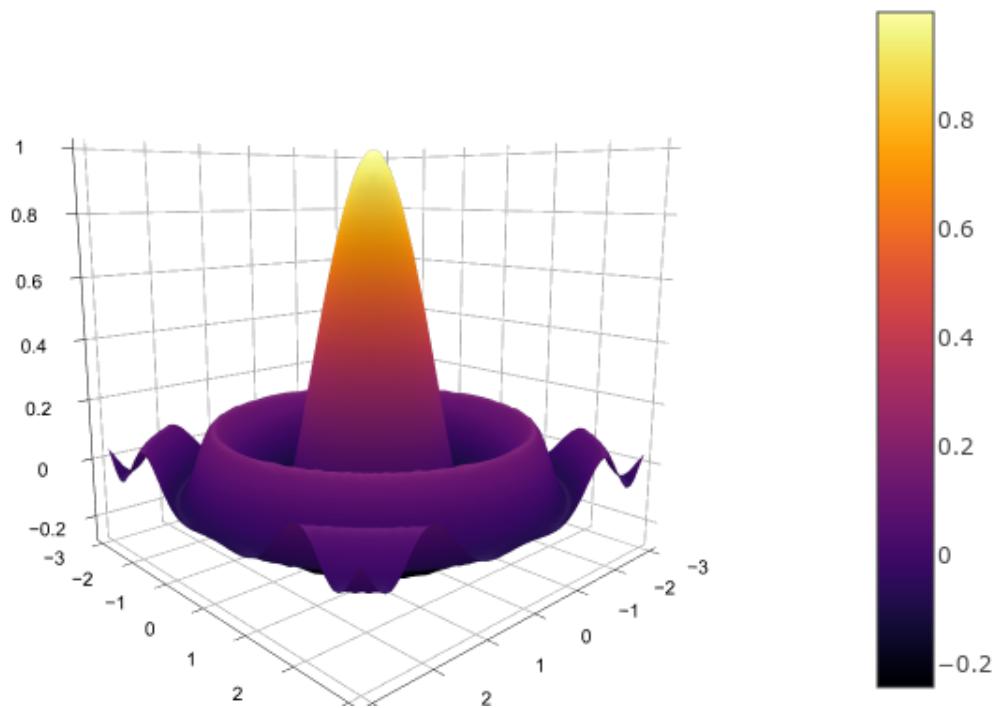
The resulting figure looks like this:

Further Reading Hopefully this tutorial has given you some ideas on how to get started with Plots.jl

We'll see more examples of this package in action through the lectures

Additional information can be found in the [official documentation](#)

Useful Libraries



Contents

- *Useful Libraries*
 - *Overview*
 - *Distributions*
 - *Working with Data*
 - *Interpolation*
 - *Optimization, Roots and Fixed Points*
 - *Others Topics*
 - *Further Reading*

Overview

While Julia lacks the massive scientific ecosystem of Python, it has successfully attracted a small army of enthusiastic and talented developers

As a result, its package system is moving towards a critical mass of useful, well written libraries

In addition, a major advantage of Julia libraries is that, because Julia itself is sufficiently fast, there is less need to mix in low level languages like C and Fortran

As a result, most Julia libraries are written exclusively in Julia

Not only does this make the libraries more portable, it makes them much easier to dive into, read, learn from and modify

In this lecture we introduce a few of the Julia libraries that we've found particularly useful for quantitative work in economics

Credits: Thanks to [@cc7768](#), [@vgregory757](#) and [@spencerlyon2](#) for keeping us up to date with current best practice

Distributions

Functions for manipulating probability distributions and generating random variables are supplied by the excellent [Distributions.jl](#) package

We'll restrict ourselves to a few simple examples (the package itself has [detailed documentation](#))

- `d = Normal(m, s)` creates a normal distribution with mean m and standard deviation s
 - defaults are $m = 0$ and $s = 1$
- `d = Uniform(a, b)` creates a uniform distribution on interval $[a, b]$
 - defaults are $a = 0$ and $b = 1$
- `d = Binomial(n, p)` creates a binomial over n trials with success probability p
 - defaults are $n = 1$ and $p = 0.5$

[Distributions.jl](#) defines various methods for acting on these instances in order to obtain

- random draws
- evaluations of pdfs (densities), cdfs (distribution functions), quantiles, etc.
- mean, variance, kurtosis, etc.

For example,

- To generate k draws from the instance `d` use `rand(d, k)`
- To obtain the mean of the distribution use `mean(d)`
- To evaluate the probability density function of `d` at `x` use `pdf(d, x)`

Further details on the interface can be found [here](#)

Several multivariate distributions are also implemented

Working with Data

A useful package for working with data is [DataFrames](#)

The most important data type provided is a `DataFrame`, a two dimensional array for storing heterogeneous data

Although data can be heterogeneous within a `DataFrame`, the contents of the columns must be homogeneous

This is analogous to a `data.frame` in R, a `DataFrame` in Pandas (Python) or, more loosely, a spreadsheet in Excel

The `DataFrames` package also supplies a `DataArray` type, which is like a one dimensional `DataFrame`

In terms of working with data, the advantage of a `DataArray` over a standard numerical array is that it can handle missing values

Here's an example

```
julia> using DataFrames

julia> commodities = ["crude", "gas", "gold", "silver"]
4-element Array{ASCIIString,1}:
 "crude"
 "gas"
 "gold"
 "silver"

julia> last = @data([4.2, 11.3, 12.1, NA]) # Create DataArray
4-element DataArray{Float64,1}:
 4.2
 11.3
 12.1
 NA

julia> df = DataFrame(commod = commodities, price = last)
4x2 DataFrame
|-----|-----|-----|
| Row # | commod   | price  |
| 1     | "crude"  | 4.2    |
| 2     | "gas"     | 11.3   |
| 3     | "gold"    | 12.1   |
| 4     | "silver"  | NA     |
```

Columns of the `DataFrame` can be accessed by name

```
julia> df[:price]
4-element DataArray{Float64,1}:
 4.2
 11.3
 12.1
 NA

julia> df[:commod]
4-element DataArray{ASCIIString,1}:
 "crude"
 "gas"
 "gold"
 "silver"
```

The `DataFrames` package provides a number of methods for acting on `DataFrames`

A simple one is `describe()`

```
julia> describe(df)
commod
Length 4
Type ASCIIString
NAs 0
NA% 0.0%
Unique 4

price
Min 4.2
1st Qu. 7.75
Median 11.3
Mean 9.200000000000001
3rd Qu. 11.7
Max 12.1
NAs 1
NA% 25.0%
```

There are also functions for splitting, merging and other data munging operations

Data can be read from and written to CSV files using syntax `df = readtable("data_file.csv")` and `writetable("data_file.csv", df)` respectively

Other packages for working with data can be found at [JuliaStats](#) and [JuliaQuant](#)

Interpolation

In economics we often wish to interpolate discrete data (i.e., build continuous functions that join discrete sequences of points)

We also need such representations to be fast and efficient

The package we usually turn to for this purpose is [Interpolations.jl](#)

One downside of [Interpolations.jl](#) is that the code to set up simple interpolation objects is relatively verbose

The upside is that the routines have excellent performance

The package is also well written and well maintained

Univariate Interpolation

Let's start with the univariate case

We begin by creating some data points, using a sine function

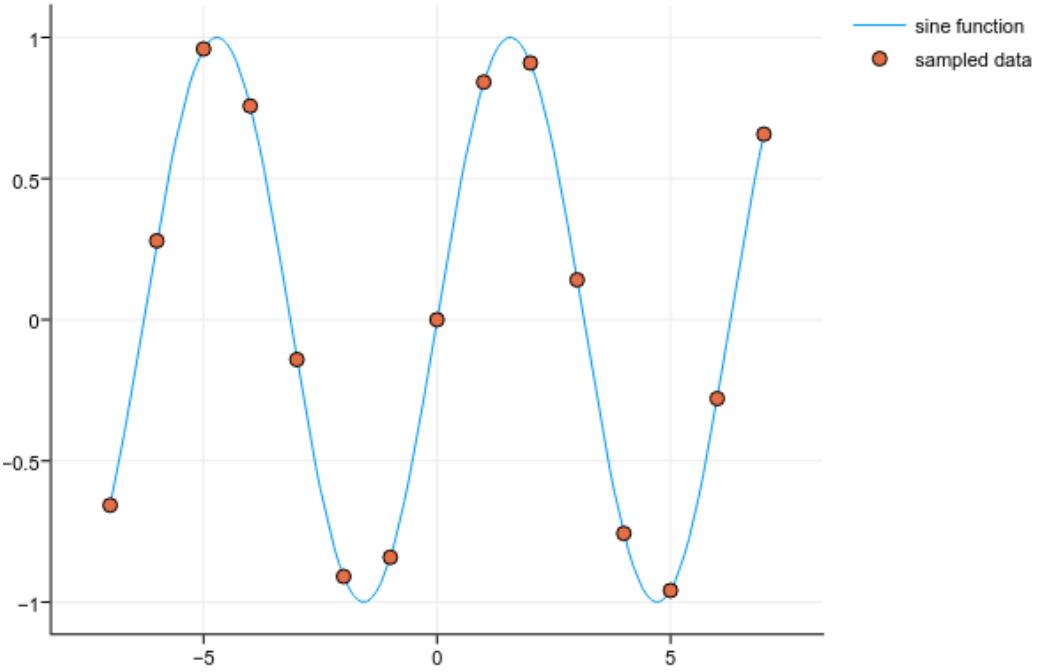
```
using Interpolations
using Plots
plotlyjs()

x = -7:7      # x points, coarse grid
y = sin(x)    # corresponding y points

xf = -7:0.1:7 # fine grid
```

```
plot(xf, sin(xf), label="sine function")
scatter!(x, y, label="sampled data", markersize=4)
```

Here's the resulting figure



Now let's interpolate the sampled data points using piecewise constant, piecewise linear and cubic interpolation

```
itp_const = scale(interpolate(y, BSpline(Constant()), OnGrid()), x)
itp_linear = scale(interpolate(y, BSpline(Linear()), OnGrid()), x)
itp_cubic = scale(interpolate(y, BSpline(Cubic(Line())), OnGrid()), x)
```

When we want to evaluate them at points in their domain (i.e., between `min(x)` and `max(x)`) we can do so as follows

```
julia> itp_cubic[0.3]
0.29400097760820687
```

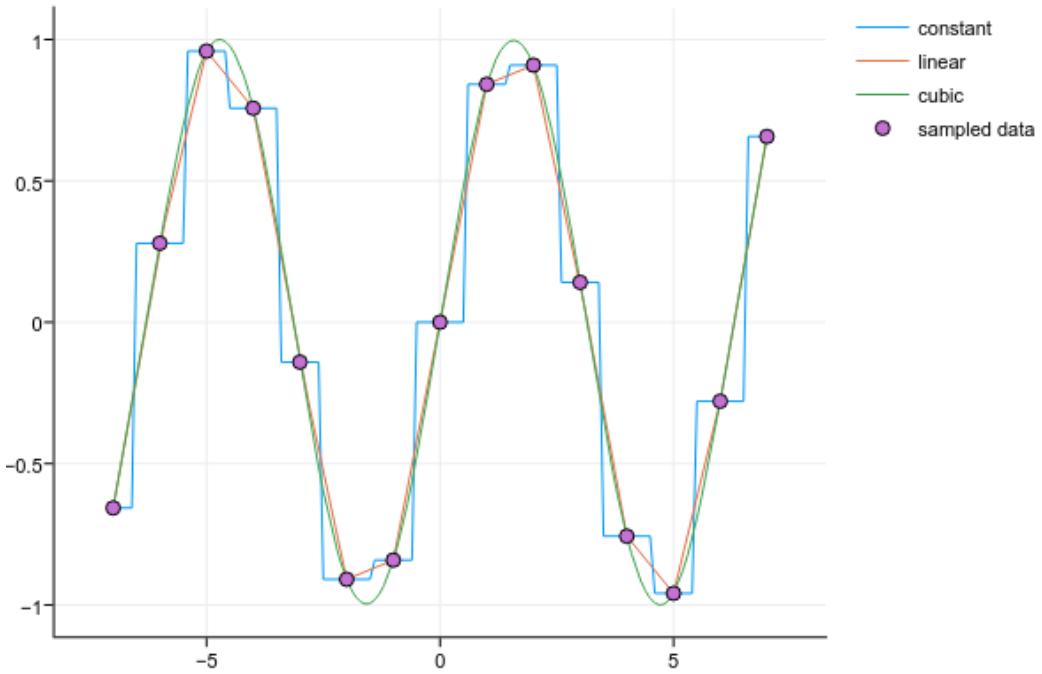
Note the use of square brackets, rather than parentheses!

Let's plot these functions

```
xf = -7:0.1:7
y_const = [itp_const[x] for x in xf]
y_linear = [itp_linear[x] for x in xf]
y_cubic = [itp_cubic[x] for x in xf]

plot(xf, [y_const y_linear y_cubic], label=["constant" "linear" "cubic"])
scatter!(x, y, label="sampled data", markersize=4)
```

Here's the figure we obtain



Univariate with Irregular Grid Here's an example with an irregular grid

```
using Interpolations
using Plots
plotlyjs()

x = log(linspace(1, exp(4), 10)) + 1 # Uneven grid
y = log(x) # Corresponding y points

itp_const = interpolate((x, ), y, Gridded(Constant()))
itp_linear = interpolate((x, ), y, Gridded(Linear()))

xf = log(linspace(1, exp(4), 100)) + 1
y_const = [itp_const[x] for x in xf]
y_linear = [itp_linear[x] for x in xf]
y_true = [log(x) for x in xf]

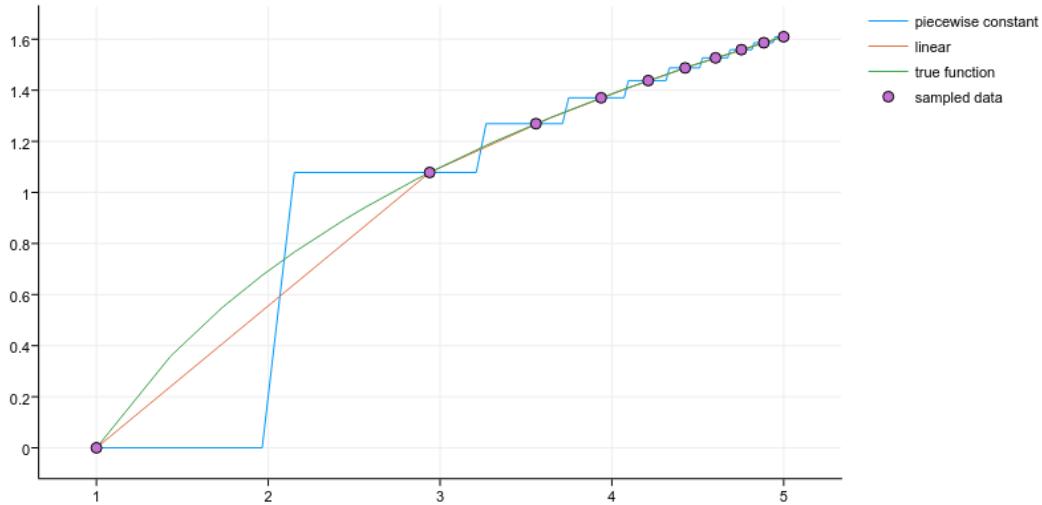
labels = ["piecewise constant" "linear" "true function"]
plot(xf, [y_const y_linear y_true], label=labels)
scatter!(x, y, label="sampled data", markersize=4, size=(800, 400))
```

The figure looks as follows

Multivariate Interpolation We can also interpolate in higher dimensions

The following example gives one illustration

```
using Interpolations
using Plots
plotlyjs()
```



```

using QuantEcon: gridmake

n = 5
x = linspace(-3, 3, n)
y = copy(x)

z = Array(Float64, n, n)
f(x, y) = cos(x^2 + y^2) / (1 + x^2 + y^2)
for i in 1:n
    for j in 1:n
        z[j, i] = f(x[i], y[j])
    end
end

itp = interpolate((x, y), z, Gridded(Linear()));

nf = 50
xf = linspace(-3, 3, nf)
yf = copy(xf)

zf = Array(Float64, nf, nf)
ztrue = Array(Float64, nf, nf)
for i in 1:nf
    for j in 1:nf
        zf[j, i] = itp[xf[i], yf[j]]
        ztrue[j, i] = f(xf[i], yf[j])
    end
end

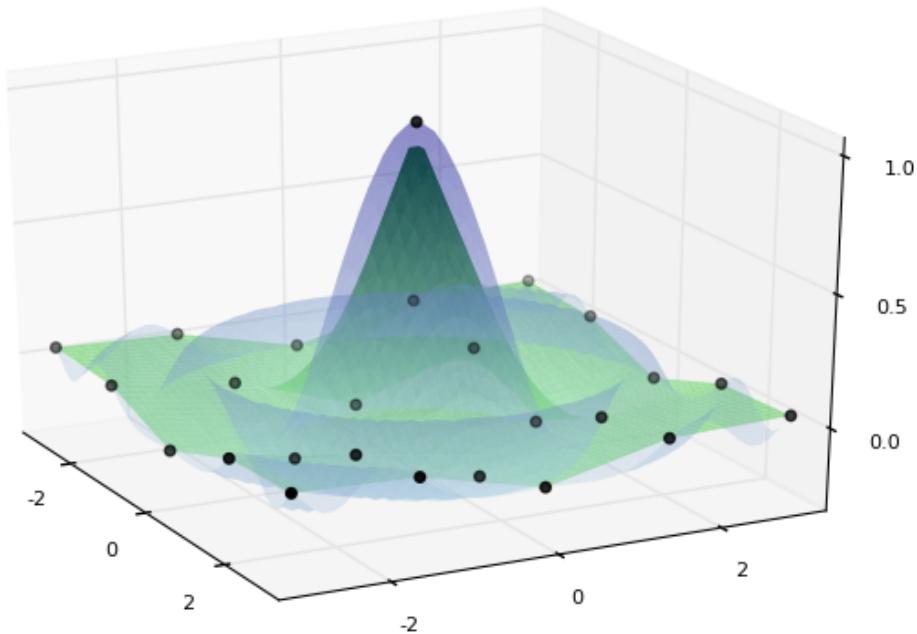
grid = gridmake(x, y)
z = reshape(z, n*n, 1)

pyplot()
surface(xf, yf, zf, color=:greens, falpha=0.7, cbar=false)
surface!(xf, yf, ztrue, fcolor=:blues, falpha=0.25, cbar=false)

```

```
scatter!(grid[:, 1], grid[:, 2], vec(z), legend=:none, color=:black, markersize=4)
```

This code produces the following figure



The original function is in blue, while the linear interpolant is shown in green

Optimization, Roots and Fixed Points

Let's look briefly at the optimization and root finding algorithms

Roots A root of a real function f on $[a, b]$ is an $x \in [a, b]$ such that $f(x) = 0$

For example, if we plot the function

$$f(x) = \sin(4(x - 1/4)) + x + x^{20} - 1 \quad (1.6)$$

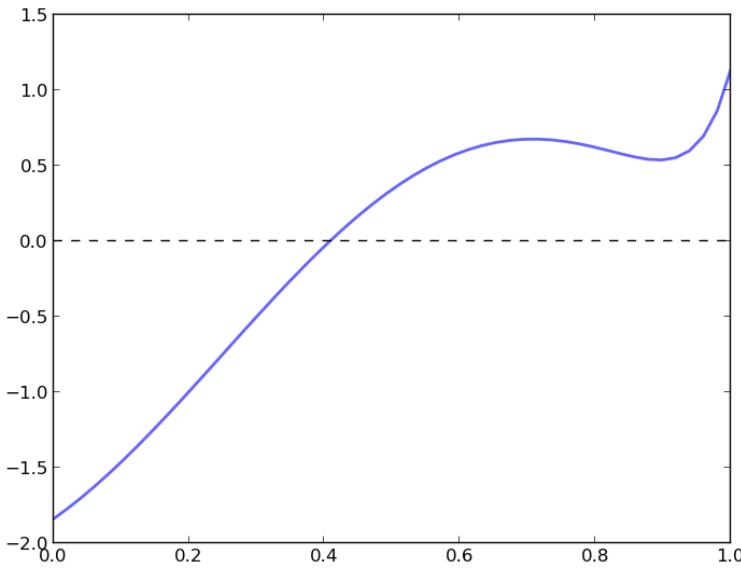
with $x \in [0, 1]$ we get

The unique root is approximately 0.408

One common root-finding algorithm is the Newton-Raphson method

This is implemented as `newton()` in the `Roots` package and is called with the function and an initial guess

```
julia> using Roots
julia> f(x) = sin(4 * (x - 1/4)) + x + x^20 - 1
```



```
f (generic function with 1 method)
```

```
julia> newton(f, 0.2)
0.40829350427936706
```

The Newton-Raphson method uses local slope information, which can lead to failure of convergence for some initial conditions

```
julia> newton(f, 0.7)
-1.0022469256696989
```

For this reason most modern solvers use more robust “hybrid methods”, as does Roots’s `fzero()` function

```
julia> fzero(f, 0, 1)
0.40829350427936706
```

Optimization For constrained, univariate minimization a useful option is `optimize()` from the `Optim` package

This function defaults to a robust hybrid optimization routine called Brent’s method

```
julia> using Optim

julia> optimize(x -> x^2, -1.0, 1.0)
Results of Optimization Algorithm
* Algorithm: Brent's Method
* Search Interval: [-1.000000, 1.000000]
* Minimizer: -0.000000
* Minimum: 0.000000
```

```
* Iterations: 5
* Convergence: max(|x - x_upper|, |x - x_lower|) <= 2*(1.5e-08*|x|+2.2e-16): true
* Objective Function Calls: 6
```

For other optimization routines, including least squares and multivariate optimization, see [the documentation](#)

A number of alternative packages for optimization can be found at [JuliaOpt](#)

Others Topics

Numerical Integration The base library contains a function called `quadgk()` that performs Gaussian quadrature

```
julia> quadgk(x -> cos(x), -2pi, 2pi)
(5.644749237155177e-15, 4.696156369056425e-22)
```

This is an adaptive Gauss-Kronrod integration technique that's relatively accurate for smooth functions

However, its adaptive implementation makes it slow and not well suited to inner loops

For this kind of integration you can use the quadrature routines from QuantEcon

```
julia> using QuantEcon

julia> nodes, weights = qnwlege(65, -2pi, 2pi);

julia> integral = do_quad(x -> cos(x), nodes, weights)
-2.912600716165059e-15
```

Let's time the two implementations

```
julia> @time quadgk(x -> cos(x), -2pi, 2pi)
elapsed time: 2.732162971 seconds (984420160 bytes allocated, 40.55% gc time)

julia> @time do_quad(x -> cos(x), nodes, weights)
elapsed time: 0.002805691 seconds (1424 bytes allocated)
```

We get similar accuracy with a speed up factor approaching three orders of magnitude

More numerical integration (and differentiation) routines can be found in the package [Calculus](#)

Linear Algebra The standard library contains many useful routines for linear algebra, in addition to standard functions such as `det()`, `inv()`, `eye()`, etc.

Routines are available for

- Cholesky factorization
- LU decomposition
- Singular value decomposition,

- Schur factorization, etc.

See [here](#) for further details

Further Reading

The full set of libraries available under the Julia packaging system can be browsed at pkg.julialang.org

CHAPTER
TWO

INTRODUCTORY APPLICATIONS

This section of the course contains intermediate and foundational applications.

Linear Algebra

Contents

- *Linear Algebra*
 - *Overview*
 - *Vectors*
 - *Matrices*
 - *Solving Systems of Equations*
 - *Eigenvalues and Eigenvectors*
 - *Further Topics*
 - *Exercises*
 - *Solutions*

Overview

Linear algebra is one of the most useful branches of applied mathematics for economists to invest in

For example, many applied problems in economics and finance require the solution of a linear system of equations, such as

$$\begin{aligned}y_1 &= ax_1 + bx_2 \\y_2 &= cx_1 + dx_2\end{aligned}$$

or, more generally,

$$\begin{aligned}y_1 &= a_{11}x_1 + a_{12}x_2 + \cdots + a_{1k}x_k \\&\vdots \\y_n &= a_{n1}x_1 + a_{n2}x_2 + \cdots + a_{nk}x_k\end{aligned}\tag{2.1}$$

The objective here is to solve for the “unknowns” x_1, \dots, x_k given a_{11}, \dots, a_{nk} and y_1, \dots, y_n

When considering such problems, it is essential that we first consider at least some of the following questions

- Does a solution actually exist?
- Are there in fact many solutions, and if so how should we interpret them?
- If no solution exists, is there a best “approximate” solution?
- If a solution exists, how should we compute it?

These are the kinds of topics addressed by linear algebra

In this lecture we will cover the basics of linear and matrix algebra, treating both theory and computation

We admit some overlap with [this lecture](#), where operations on Julia arrays were first explained

Note that this lecture is more theoretical than most, and contains background material that will be used in applications as we go along

Vectors

A *vector* of length n is just a sequence (or array, or tuple) of n numbers, which we write as $x = (x_1, \dots, x_n)$ or $x = [x_1, \dots, x_n]$

We will write these sequences either horizontally or vertically as we please

(Later, when we wish to perform certain matrix operations, it will become necessary to distinguish between the two)

The set of all n -vectors is denoted by \mathbb{R}^n

For example, \mathbb{R}^2 is the plane, and a vector in \mathbb{R}^2 is just a point in the plane

Traditionally, vectors are represented visually as arrows from the origin to the point

The following figure represents three vectors in this manner

If you’re interested, the Julia code for producing this figure is [here](#)

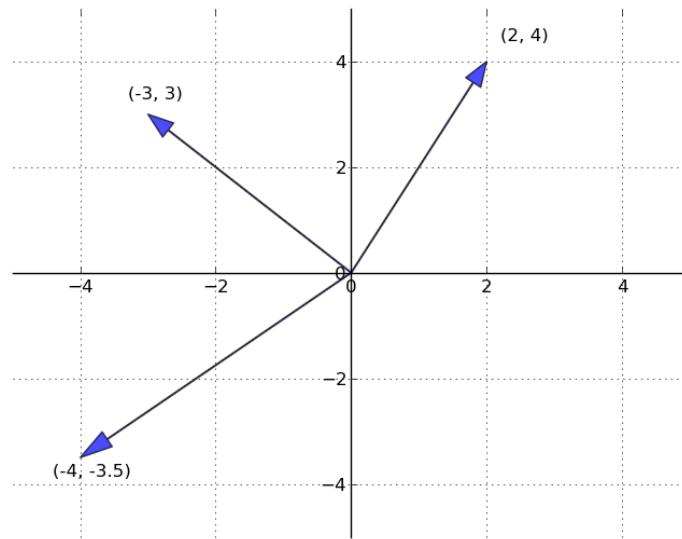
Vector Operations The two most common operators for vectors are addition and scalar multiplication, which we now describe

As a matter of definition, when we add two vectors, we add them element by element

$$x + y = \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_n \end{bmatrix} + \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix} := \begin{bmatrix} x_1 + y_1 \\ x_2 + y_2 \\ \vdots \\ x_n + y_n \end{bmatrix}$$

Scalar multiplication is an operation that takes a number γ and a vector x and produces

$$\gamma x := \begin{bmatrix} \gamma x_1 \\ \gamma x_2 \\ \vdots \\ \gamma x_n \end{bmatrix}$$



Scalar multiplication is illustrated in the next figure

In Julia, a vector can be represented as a one dimensional *Array*

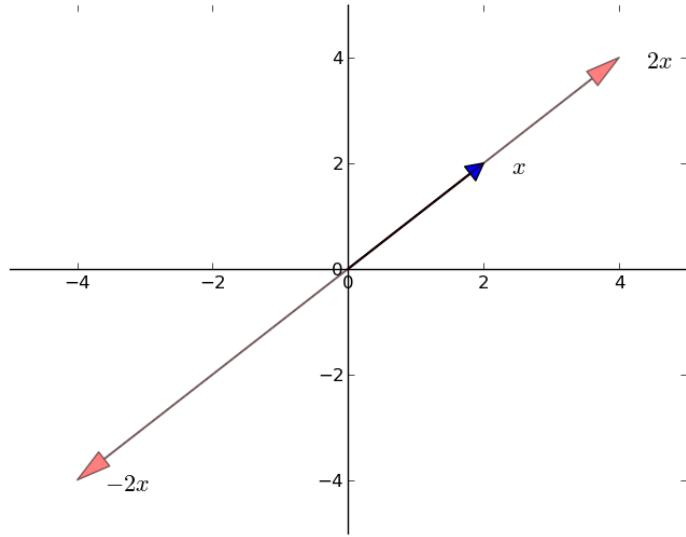
Julia *Arrays* allow us to express scalar multiplication and addition with a very natural syntax

```
julia> x = ones(3)
3-element Array{Float64,1}:
 1.0
 1.0
 1.0

julia> y = [2, 4, 6]
3-element Array{Int64,1}:
 2
 4
 6

julia> x + y
3-element Array{Float64,1}:
 3.0
 5.0
 7.0

julia> 4x  # equivalent to 4 * x and 4 .* x
3-element Array{Float64,1}:
 4.0
 4.0
 4.0
```



Inner Product and Norm The *inner product* of vectors $x, y \in \mathbb{R}^n$ is defined as

$$x'y := \sum_{i=1}^n x_i y_i$$

Two vectors are called *orthogonal* if their inner product is zero

The *norm* of a vector x represents its “length” (i.e., its distance from the zero vector) and is defined as

$$\|x\| := \sqrt{x'x} := \left(\sum_{i=1}^n x_i^2 \right)^{1/2}$$

The expression $\|x - y\|$ is thought of as the distance between x and y

Continuing on from the previous example, the inner product and norm can be computed as follows

```
julia> dot(x, y)          # Inner product of x and y
12.0

julia> sum(x .* y)        # Gives the same result
12.0

julia> norm(x)            # Norm of x
1.7320508075688772

julia> sqrt(sum(x.^2))    # Gives the same result
1.7320508075688772
```

Span Given a set of vectors $A := \{a_1, \dots, a_k\}$ in \mathbb{R}^n , it’s natural to think about the new vectors we can create by performing linear operations

New vectors created in this manner are called *linear combinations* of A

In particular, $y \in \mathbb{R}^n$ is a linear combination of $A := \{a_1, \dots, a_k\}$ if

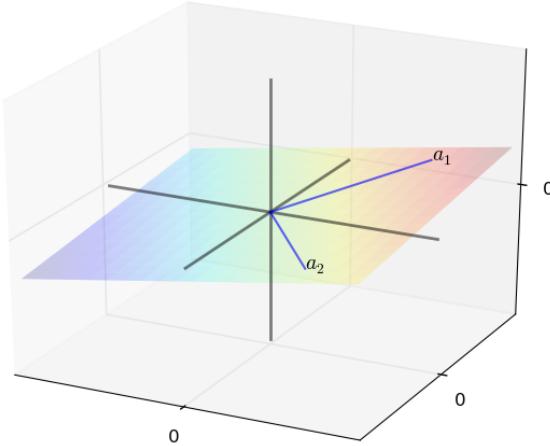
$$y = \beta_1 a_1 + \dots + \beta_k a_k \text{ for some scalars } \beta_1, \dots, \beta_k$$

In this context, the values β_1, \dots, β_k are called the *coefficients* of the linear combination

The set of linear combinations of A is called the *span* of A

The next figure shows the span of $A = \{a_1, a_2\}$ in \mathbb{R}^3

The span is a 2 dimensional plane passing through these two points and the origin



The code for producing this figure can be found [here](#)

Examples If A contains only one vector $a_1 \in \mathbb{R}^2$, then its span is just the scalar multiples of a_1 , which is the unique line passing through both a_1 and the origin

If $A = \{e_1, e_2, e_3\}$ consists of the *canonical basis vectors* of \mathbb{R}^3 , that is

$$e_1 := \begin{bmatrix} 1 \\ 0 \\ 0 \end{bmatrix}, \quad e_2 := \begin{bmatrix} 0 \\ 1 \\ 0 \end{bmatrix}, \quad e_3 := \begin{bmatrix} 0 \\ 0 \\ 1 \end{bmatrix}$$

then the span of A is all of \mathbb{R}^3 , because, for any $x = (x_1, x_2, x_3) \in \mathbb{R}^3$, we can write

$$x = x_1 e_1 + x_2 e_2 + x_3 e_3$$

Now consider $A_0 = \{e_1, e_2, e_1 + e_2\}$

If $y = (y_1, y_2, y_3)$ is any linear combination of these vectors, then $y_3 = 0$ (check it)

Hence A_0 fails to span all of \mathbb{R}^3

Linear Independence As we'll see, it's often desirable to find families of vectors with relatively large span, so that many vectors can be described by linear operators on a few vectors

The condition we need for a set of vectors to have a large span is what's called linear independence

In particular, a collection of vectors $A := \{a_1, \dots, a_k\}$ in \mathbb{R}^n is said to be

- *linearly dependent* if some strict subset of A has the same span as A
- *linearly independent* if it is not linearly dependent

Put differently, a set of vectors is linearly independent if no vector is redundant to the span, and linearly dependent otherwise

To illustrate the idea, recall *the figure* that showed the span of vectors $\{a_1, a_2\}$ in \mathbb{R}^3 as a plane through the origin

If we take a third vector a_3 and form the set $\{a_1, a_2, a_3\}$, this set will be

- linearly dependent if a_3 lies in the plane
- linearly independent otherwise

As another illustration of the concept, since \mathbb{R}^n can be spanned by n vectors (see the discussion of canonical basis vectors above), any collection of $m > n$ vectors in \mathbb{R}^n must be linearly dependent

The following statements are equivalent to linear independence of $A := \{a_1, \dots, a_k\} \subset \mathbb{R}^n$

1. No vector in A can be formed as a linear combination of the other elements
2. If $\beta_1 a_1 + \dots + \beta_k a_k = 0$ for scalars β_1, \dots, β_k , then $\beta_1 = \dots = \beta_k = 0$

(The zero in the first expression is the origin of \mathbb{R}^n)

Unique Representations Another nice thing about sets of linearly independent vectors is that each element in the span has a unique representation as a linear combination of these vectors

In other words, if $A := \{a_1, \dots, a_k\} \subset \mathbb{R}^n$ is linearly independent and

$$y = \beta_1 a_1 + \dots + \beta_k a_k$$

then no other coefficient sequence $\gamma_1, \dots, \gamma_k$ will produce the same vector y

Indeed, if we also have $y = \gamma_1 a_1 + \dots + \gamma_k a_k$, then

$$(\beta_1 - \gamma_1) a_1 + \dots + (\beta_k - \gamma_k) a_k = 0$$

Linear independence now implies $\gamma_i = \beta_i$ for all i

Matrices

Matrices are a neat way of organizing data for use in linear operations

An $n \times k$ matrix is a rectangular array A of numbers with n rows and k columns:

$$A = \begin{bmatrix} a_{11} & a_{12} & \cdots & a_{1k} \\ a_{21} & a_{22} & \cdots & a_{2k} \\ \vdots & \vdots & & \vdots \\ a_{n1} & a_{n2} & \cdots & a_{nk} \end{bmatrix}$$

Often, the numbers in the matrix represent coefficients in a system of linear equations, as discussed at the start of this lecture

For obvious reasons, the matrix A is also called a vector if either $n = 1$ or $k = 1$

In the former case, A is called a *row vector*, while in the latter it is called a *column vector*

If $n = k$, then A is called *square*

The matrix formed by replacing a_{ij} by a_{ji} for every i and j is called the *transpose* of A , and denoted A' or A^\top

If $A = A'$, then A is called *symmetric*

For a square matrix A , the i elements of the form a_{ii} for $i = 1, \dots, n$ are called the *principal diagonal*
 A is called *diagonal* if the only nonzero entries are on the principal diagonal

If, in addition to being diagonal, each element along the principal diagonal is equal to 1, then A is called the *identity matrix*, and denoted by I

Matrix Operations Just as was the case for vectors, a number of algebraic operations are defined for matrices

Scalar multiplication and addition are immediate generalizations of the vector case:

$$\gamma A = \gamma \begin{bmatrix} a_{11} & \cdots & a_{1k} \\ \vdots & \vdots & \vdots \\ a_{n1} & \cdots & a_{nk} \end{bmatrix} := \begin{bmatrix} \gamma a_{11} & \cdots & \gamma a_{1k} \\ \vdots & \vdots & \vdots \\ \gamma a_{n1} & \cdots & \gamma a_{nk} \end{bmatrix}$$

and

$$A + B = \begin{bmatrix} a_{11} & \cdots & a_{1k} \\ \vdots & \vdots & \vdots \\ a_{n1} & \cdots & a_{nk} \end{bmatrix} + \begin{bmatrix} b_{11} & \cdots & b_{1k} \\ \vdots & \vdots & \vdots \\ b_{n1} & \cdots & b_{nk} \end{bmatrix} := \begin{bmatrix} a_{11} + b_{11} & \cdots & a_{1k} + b_{1k} \\ \vdots & \vdots & \vdots \\ a_{n1} + b_{n1} & \cdots & a_{nk} + b_{nk} \end{bmatrix}$$

In the latter case, the matrices must have the same shape in order for the definition to make sense

We also have a convention for *multiplying* two matrices

The rule for matrix multiplication generalizes the idea of inner products discussed above, and is designed to make multiplication play well with basic linear operations

If A and B are two matrices, then their product AB is formed by taking as its i, j -th element the inner product of the i -th row of A and the j -th column of B

There are many tutorials to help you visualize this operation, such as [this one](#), or the discussion on the [Wikipedia page](#)

If A is $n \times k$ and B is $j \times m$, then to multiply A and B we require $k = j$, and the resulting matrix AB is $n \times m$

As perhaps the most important special case, consider multiplying $n \times k$ matrix A and $k \times 1$ column vector x

According to the preceding rule, this gives us an $n \times 1$ column vector

$$Ax = \begin{bmatrix} a_{11} & \cdots & a_{1k} \\ \vdots & \vdots & \vdots \\ a_{n1} & \cdots & a_{nk} \end{bmatrix} \begin{bmatrix} x_1 \\ \vdots \\ x_k \end{bmatrix} := \begin{bmatrix} a_{11}x_1 + \cdots + a_{1k}x_k \\ \vdots \\ a_{n1}x_1 + \cdots + a_{nk}x_k \end{bmatrix} \quad (2.2)$$

Note: AB and BA are not generally the same thing

Another important special case is the identity matrix

You should check that if A is $n \times k$ and I is the $k \times k$ identity matrix, then $AI = A$

If I is the $n \times n$ identity matrix, then $IA = A$

Matrices in Julia Julia arrays are also used as matrices, and have fast, efficient functions and methods for all the standard matrix operations

You can create them as follows

```
julia> A = [1 2
            3 4]
2x2 Array{Int64,2}:
 1  2
 3  4

julia> typeof(A)
Array{Int64,2}

julia> size(A)
(2,2)
```

The `size` function returns a tuple giving the number of rows and columns

To get the transpose of A , use `transpose(A)` or, more simply, A'

There are many convenient functions for creating common matrices (matrices of zeros, ones, etc.) — see [here](#)

Since operations are performed elementwise by default, scalar multiplication and addition have very natural syntax

```
julia> A = eye(3)
3x3 Array{Float64,2}:
 1.0  0.0  0.0
```

```

0.0 1.0 0.0
0.0 0.0 1.0

julia> B = ones(3, 3)
3x3 Array{Float64, 2}:
 1.0 1.0 1.0
 1.0 1.0 1.0
 1.0 1.0 1.0

julia> 2A
3x3 Array{Float64, 2}:
 2.0 0.0 0.0
 0.0 2.0 0.0
 0.0 0.0 2.0

julia> A + B
3x3 Array{Float64, 2}:
 2.0 1.0 1.0
 1.0 2.0 1.0
 1.0 1.0 2.0

```

To multiply matrices we use the `*` operator

In particular, `A * B` is matrix multiplication, whereas `A .* B` is element by element multiplication

Matrices as Maps Each $n \times k$ matrix A can be identified with a function $f(x) = Ax$ that maps $x \in \mathbb{R}^k$ into $y = Ax \in \mathbb{R}^n$

These kinds of functions have a special property: they are *linear*

A function $f: \mathbb{R}^k \rightarrow \mathbb{R}^n$ is called *linear* if, for all $x, y \in \mathbb{R}^k$ and all scalars α, β , we have

$$f(\alpha x + \beta y) = \alpha f(x) + \beta f(y)$$

You can check that this holds for the function $f(x) = Ax + b$ when b is the zero vector, and fails when b is nonzero

In fact, it's known that f is linear if and *only if* there exists a matrix A such that $f(x) = Ax$ for all x .

Solving Systems of Equations

Recall again the system of equations (2.1)

If we compare (2.1) and (2.2), we see that (2.1) can now be written more conveniently as

$$y = Ax \tag{2.3}$$

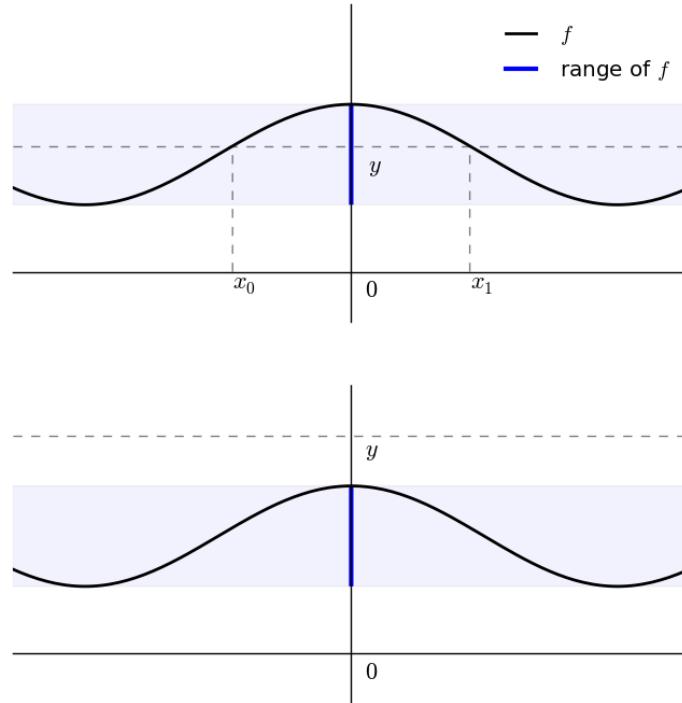
The problem we face is to determine a vector $x \in \mathbb{R}^k$ that solves (2.3), taking y and A as given

This is a special case of a more general problem: Find an x such that $y = f(x)$

Given an arbitrary function f and a y , is there always an x such that $y = f(x)$?

If so, is it always unique?

The answer to both these questions is negative, as the next figure shows



In the first plot there are multiple solutions, as the function is not one-to-one, while in the second there are no solutions, since y lies outside the range of f

Can we impose conditions on A in (2.3) that rule out these problems?

In this context, the most important thing to recognize about the expression Ax is that it corresponds to a linear combination of the columns of A

In particular, if a_1, \dots, a_k are the columns of A , then

$$Ax = x_1 a_1 + \cdots + x_k a_k$$

Hence the range of $f(x) = Ax$ is exactly the span of the columns of A

We want the range to be large, so that it contains arbitrary y

As you might recall, the condition that we want for the span to be large is *linear independence*

A happy fact is that linear independence of the columns of A also gives us uniqueness

Indeed, it follows from our *earlier discussion* that if $\{a_1, \dots, a_k\}$ are linearly independent and $y = Ax = x_1 a_1 + \cdots + x_k a_k$, then no $z \neq x$ satisfies $y = Az$

The $n \times n$ Case Let's discuss some more details, starting with the case where A is $n \times n$

This is the familiar case where the number of unknowns equals the number of equations

For arbitrary $y \in \mathbb{R}^n$, we hope to find a unique $x \in \mathbb{R}^n$ such that $y = Ax$

In view of the observations immediately above, if the columns of A are linearly independent, then their span, and hence the range of $f(x) = Ax$, is all of \mathbb{R}^n

Hence there always exists an x such that $y = Ax$

Moreover, the solution is unique

In particular, the following are equivalent

1. The columns of A are linearly independent
2. For any $y \in \mathbb{R}^n$, the equation $y = Ax$ has a unique solution

The property of having linearly independent columns is sometimes expressed as having *full column rank*

Inverse Matrices Can we give some sort of expression for the solution?

If y and A are scalar with $A \neq 0$, then the solution is $x = A^{-1}y$

A similar expression is available in the matrix case

In particular, if square matrix A has full column rank, then it possesses a multiplicative *inverse matrix* A^{-1} , with the property that $AA^{-1} = A^{-1}A = I$

As a consequence, if we pre-multiply both sides of $y = Ax$ by A^{-1} , we get $x = A^{-1}y$

This is the solution that we're looking for

Determinants Another quick comment about square matrices is that to every such matrix we assign a unique number called the *determinant* of the matrix — you can find the expression for it [here](#)

If the determinant of A is not zero, then we say that A is *nonsingular*

Perhaps the most important fact about determinants is that A is nonsingular if and only if A is of full column rank

This gives us a useful one-number summary of whether or not a square matrix can be inverted

More Rows than Columns This is the $n \times k$ case with $n > k$

This case is very important in many settings, not least in the setting of linear regression (where n is the number of observations, and k is the number of explanatory variables)

Given arbitrary $y \in \mathbb{R}^n$, we seek an $x \in \mathbb{R}^k$ such that $y = Ax$

In this setting, existence of a solution is highly unlikely

Without much loss of generality, let's go over the intuition focusing on the case where the columns of A are linearly independent

It follows that the span of the columns of A is a k -dimensional subspace of \mathbb{R}^n

This span is very "unlikely" to contain arbitrary $y \in \mathbb{R}^n$

To see why, recall the *figure above*, where $k = 2$ and $n = 3$

Imagine an arbitrarily chosen $y \in \mathbb{R}^3$, located somewhere in that three dimensional space

What's the likelihood that y lies in the span of $\{a_1, a_2\}$ (i.e., the two dimensional plane through these points)?

In a sense it must be very small, since this plane has zero "thickness"

As a result, in the $n > k$ case we usually give up on existence

However, we can still seek a best approximation, for example an x that makes the distance $\|y - Ax\|$ as small as possible

To solve this problem, one can use either calculus or the theory of orthogonal projections

The solution is known to be $\hat{x} = (A'A)^{-1}A'y$ — see for example chapter 3 of these notes

More Columns than Rows This is the $n \times k$ case with $n < k$, so there are fewer equations than unknowns

In this case there are either no solutions or infinitely many — in other words, uniqueness never holds

For example, consider the case where $k = 3$ and $n = 2$

Thus, the columns of A consists of 3 vectors in \mathbb{R}^2

This set can never be linearly independent, since it is possible to find two vectors that span \mathbb{R}^2

(For example, use the canonical basis vectors)

It follows that one column is a linear combination of the other two

For example, let's say that $a_1 = \alpha a_2 + \beta a_3$

Then if $y = Ax = x_1 a_1 + x_2 a_2 + x_3 a_3$, we can also write

$$y = x_1(\alpha a_2 + \beta a_3) + x_2 a_2 + x_3 a_3 = (x_1\alpha + x_2)a_2 + (x_1\beta + x_3)a_3$$

In other words, uniqueness fails

Linear Equations with Julia Here's an illustration of how to solve linear equations with Julia's built-in linear algebra facilities

```
julia> A = [1.0 2.0; 3.0 4.0];
julia> y = ones(2, 1); # A column vector
```

```
julia> det(A)
-2.0

julia> A_inv = inv(A)
2x2 Array{Float64,2}:
 -2.0   1.0
  1.5  -0.5

julia> x = A_inv * y  # solution
2x1 Array{Float64,2}:
 -1.0
  1.0

julia> A * x  # should equal y (a vector of ones)
2x1 Array{Float64,2}:
 1.0
 1.0

julia> A\y  # produces the same solution
2x1 Array{Float64,2}:
 -1.0
  1.0
```

Observe how we can solve for $x = A^{-1}y$ by either via `inv(A) * y`, or using `A \ y`

The latter method is preferred because it automatically selects the best algorithm for the problem based on the values of A and y

If A is not square then $A \ y$ returns the least squares solution $\hat{x} = (A'A)^{-1}A'y$

Eigenvalues and Eigenvectors

Let A be an $n \times n$ square matrix

If λ is scalar and v is a non-zero vector in \mathbb{R}^n such that

$$Av = \lambda v$$

then we say that λ is an *eigenvalue* of A , and v is an *eigenvector*

Thus, an eigenvector of A is a vector such that when the map $f(x) = Ax$ is applied, v is merely scaled

The next figure shows two eigenvectors (blue arrows) and their images under A (red arrows)

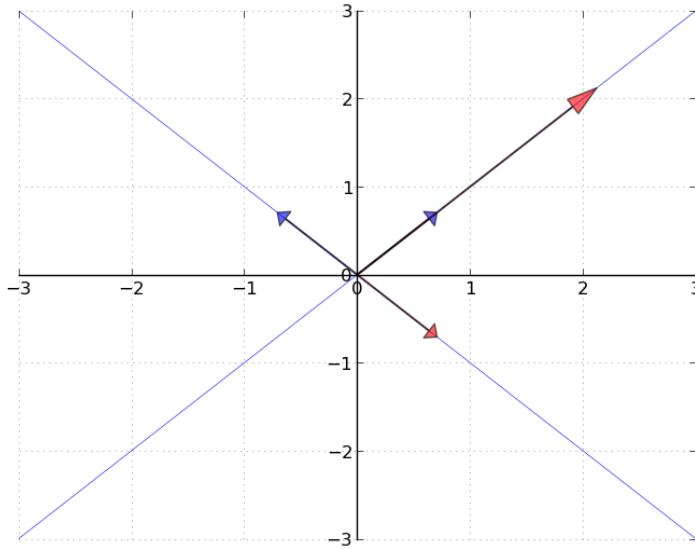
As expected, the image Av of each v is just a scaled version of the original

The eigenvalue equation is equivalent to $(A - \lambda I)v = 0$, and this has a nonzero solution v only when the columns of $A - \lambda I$ are linearly dependent

This in turn is equivalent to stating that the determinant is zero

Hence to find all eigenvalues, we can look for λ such that the determinant of $A - \lambda I$ is zero

This problem can be expressed as one of solving for the roots of a polynomial in λ of degree n



This in turn implies the existence of n solutions in the complex plane, although some might be repeated

Some nice facts about the eigenvalues of a square matrix A are as follows

1. The determinant of A equals the product of the eigenvalues
2. The trace of A (the sum of the elements on the principal diagonal) equals the sum of the eigenvalues
3. If A is symmetric, then all of its eigenvalues are real
4. If A is invertible and $\lambda_1, \dots, \lambda_n$ are its eigenvalues, then the eigenvalues of A^{-1} are $1/\lambda_1, \dots, 1/\lambda_n$

A corollary of the first statement is that a matrix is invertible if and only if all its eigenvalues are nonzero

Using Julia, we can solve for the eigenvalues and eigenvectors of a matrix as follows

```
julia> A = [1.0 2.0; 2.0 1.0];

julia> evals, evecs = eig(A);

julia> evals
2-element Array{Float64,1}:
 -1.0
  3.0

julia> evecs
2x2 Array{Float64,2}:
 -0.707107  0.707107
  0.707107  0.707107
```

Note that the *columns* of `evecs` are the eigenvectors

Since any scalar multiple of an eigenvector is an eigenvector with the same eigenvalue (check it), the `eig` routine normalizes the length of each eigenvector to one

Generalized Eigenvalues It is sometimes useful to consider the *generalized eigenvalue problem*, which, for given matrices A and B , seeks generalized eigenvalues λ and eigenvectors v such that

$$Av = \lambda Bv$$

This can be solved in Julia via `eig(A, B)`

Of course, if B is square and invertible, then we can treat the generalized eigenvalue problem as an ordinary eigenvalue problem $B^{-1}Av = \lambda v$, but this is not always the case

Further Topics

We round out our discussion by briefly mentioning several other important topics

Series Expansions Recall the usual summation formula for a geometric progression, which states that if $|a| < 1$, then $\sum_{k=0}^{\infty} a^k = (1 - a)^{-1}$

A generalization of this idea exists in the matrix setting

Matrix Norms Let A be a square matrix, and let

$$\|A\| := \max_{\|x\|=1} \|Ax\|$$

The norms on the right-hand side are ordinary vector norms, while the norm on the left-hand side is a *matrix norm* — in this case, the so-called *spectral norm*

For example, for a square matrix S , the condition $\|S\| < 1$ means that S is *contractive*, in the sense that it pulls all vectors towards the origin ¹

Neumann's Theorem Let A be a square matrix and let $A^k := AA^{k-1}$ with $A^1 := A$

In other words, A^k is the k -th power of A

Neumann's theorem states the following: If $\|A^k\| < 1$ for some $k \in \mathbb{N}$, then $I - A$ is invertible, and

$$(I - A)^{-1} = \sum_{k=0}^{\infty} A^k \tag{2.4}$$

¹ Suppose that $\|S\| < 1$. Take any nonzero vector x , and let $r := \|x\|$. We have $\|Sx\| = r\|S(x/r)\| \leq r\|S\| < r = \|x\|$. Hence every point is pulled towards the origin.

Spectral Radius A result known as Gelfand's formula tells us that, for any square matrix A ,

$$\rho(A) = \lim_{k \rightarrow \infty} \|A^k\|^{1/k}$$

Here $\rho(A)$ is the *spectral radius*, defined as $\max_i |\lambda_i|$, where $\{\lambda_i\}_i$ is the set of eigenvalues of A

As a consequence of Gelfand's formula, if all eigenvalues are strictly less than one in modulus, there exists a k with $\|A^k\| < 1$

In which case (2.4) is valid

Positive Definite Matrices Let A be a symmetric $n \times n$ matrix

We say that A is

1. *positive definite* if $x'Ax > 0$ for every $x \in \mathbb{R}^n \setminus \{0\}$
2. *positive semi-definite* or *nonnegative definite* if $x'Ax \geq 0$ for every $x \in \mathbb{R}^n$

Analogous definitions exist for negative definite and negative semi-definite matrices

It is notable that if A is positive definite, then all of its eigenvalues are strictly positive, and hence A is invertible (with positive definite inverse)

Differentiating Linear and Quadratic forms The following formulas are useful in many economic contexts. Let

- z, x and a all be $n \times 1$ vectors
- A be an $n \times n$ matrix
- B be an $m \times n$ matrix and y be an $m \times 1$ vector

Then

1. $\frac{\partial a'x}{\partial x} = a$
2. $\frac{\partial Ax}{\partial x} = A'$
3. $\frac{\partial x'Ax}{\partial x} = (A + A')x$
4. $\frac{\partial y'Bz}{\partial y} = Bz$
5. $\frac{\partial y'Bz}{\partial B} = yz'$

Exercise 1 below asks you to apply these formulas

Further Reading The documentation of the linear algebra features built into Julia can be found [here](#)

Chapters 2 and 3 of the [Econometric Theory](#) contains a discussion of linear algebra along the same lines as above, with solved exercises

If you don't mind a slightly abstract approach, a nice intermediate-level text on linear algebra is [\[Janich94\]](#)

Exercises

Exercise 1 Let x be a given $n \times 1$ vector and consider the problem

$$v(x) = \max_{y,u} \{-y'Py - u'Qu\}$$

subject to the linear constraint

$$y = Ax + Bu$$

Here

- P is an $n \times n$ matrix and Q is an $m \times m$ matrix
- A is an $n \times n$ matrix and B is an $n \times m$ matrix
- both P and Q are symmetric and positive semidefinite

(What must the dimensions of y and u be to make this a well-posed problem?)

One way to solve the problem is to form the Lagrangian

$$\mathcal{L} = -y'Py - u'Qu + \lambda' [Ax + Bu - y]$$

where λ is an $n \times 1$ vector of Lagrange multipliers

Try applying the formulas given above for differentiating quadratic and linear forms to obtain the first-order conditions for maximizing \mathcal{L} with respect to y, u and minimizing it with respect to λ

Show that these conditions imply that

1. $\lambda = -2Py$
2. The optimizing choice of u satisfies $u = -(Q + B'PB)^{-1}B'PAx$
3. The function v satisfies $v(x) = -x'\tilde{P}x$ where $\tilde{P} = A'PA - A'PB(Q + B'PB)^{-1}B'PA$

As we will see, in economic contexts Lagrange multipliers often are shadow prices

Note: If we don't care about the Lagrange multipliers, we can substitute the constraint into the objective function, and then just maximize $-(Ax + Bu)'P(Ax + Bu) - u'Qu$ with respect to u . You can verify that this leads to the same maximizer.

Solutions

[Solution notebook](#)

Finite Markov Chains

Contents

- *Finite Markov Chains*
 - *Overview*
 - *Definitions*
 - *Simulation*
 - *Marginal Distributions*
 - *Irreducibility and Aperiodicity*
 - *Stationary Distributions*
 - *Ergodicity*
 - *Computing Expectations*
 - *Exercises*
 - *Solutions*

Overview

Markov chains are one of the most useful classes of stochastic processes, being

- simple, flexible and supported by many elegant theoretical results
- valuable for building intuition about random dynamic models
- central to quantitative modeling in their own right

You will find them in many of the workhorse models of economics and finance

In this lecture we review some of the theory of Markov chains

We will also introduce some of the high quality routines for working with Markov chains available in [QuantEcon](#)

Prerequisite knowledge is basic probability and linear algebra

Definitions

The following concepts are fundamental

Stochastic Matrices A **stochastic matrix** (or **Markov matrix**) is an $n \times n$ square matrix P such that

1. each element of P is nonnegative, and
2. each row of P sums to one

Each row of P can be regarded as a probability mass function over n possible outcomes

It is too not difficult to check ¹ that if P is a stochastic matrix, then so is the k -th power P^k for all $k \in \mathbb{N}$

¹ Hint: First show that if P and Q are stochastic matrices then so is their product — to check the row sums, try postmultiplying by a column vector of ones. Finally, argue that P^n is a stochastic matrix using induction.

Markov Chains There is a close connection between stochastic matrices and Markov chains

To begin, let S be a finite set with n elements $\{x_1, \dots, x_n\}$

The set S is called the **state space** and x_1, \dots, x_n are the **state values**

A **Markov chain** $\{X_t\}$ on S is a sequence of random variables on S that have the **Markov property**

This means that, for any date t and any state $y \in S$,

$$\mathbb{P}\{X_{t+1} = y | X_t\} = \mathbb{P}\{X_{t+1} = y | X_t, X_{t-1}, \dots\} \quad (2.5)$$

In other words, knowing the current state is enough to know probabilities for future states

In particular, the dynamics of a Markov chain are fully determined by the set of values

$$P(x, y) := \mathbb{P}\{X_{t+1} = y | X_t = x\} \quad (x, y \in S) \quad (2.6)$$

By construction,

- $P(x, y)$ is the probability of going from x to y in one unit of time (one step)
- $P(x, \cdot)$ is the conditional distribution of X_{t+1} given $X_t = x$

We can view P as a stochastic matrix where

$$P_{ij} = P(x_i, x_j) \quad 1 \leq i, j \leq n$$

Going the other way, if we take a stochastic matrix P , we can generate a Markov chain $\{X_t\}$ as follows:

- draw X_0 from some specified distribution
- for each $t = 0, 1, \dots$, draw X_{t+1} from $P(X_t, \cdot)$

By construction, the resulting process satisfies (2.6)

Example 1 Consider a worker who, at any given time t , is either unemployed (state 1) or employed (state 2)

Suppose that, over a one month period,

1. An employed worker loses her job and becomes unemployed with probability $\beta \in (0, 1)$
2. An unemployed worker finds a job with probability $\alpha \in (0, 1)$

In terms of a Markov model, we have

- $S = \{1, 2\}$
- $P(1, 2) = \alpha$ and $P(2, 1) = \beta$

We can write out the transition probabilities in matrix form as

$$P = \begin{pmatrix} 1 - \alpha & \alpha \\ \beta & 1 - \beta \end{pmatrix}$$

Once we have the values α and β , we can address a range of questions, such as

- What is the average duration of unemployment?
- Over the long-run, what fraction of time does a worker find herself unemployed?
- Conditional on employment, what is the probability of becoming unemployed at least once over the next 12 months?

We'll cover such applications below

Example 2 Using US unemployment data, Hamilton [Ham05] estimated the stochastic matrix

$$P = \begin{pmatrix} 0.971 & 0.029 & 0 \\ 0.145 & 0.778 & 0.077 \\ 0 & 0.508 & 0.492 \end{pmatrix}$$

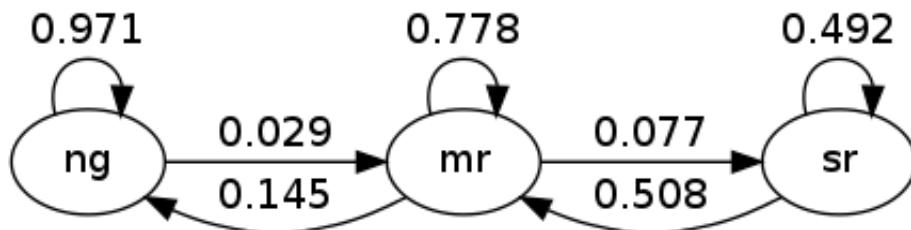
where

- the frequency is monthly
- the first state represents “normal growth”
- the second state represents “mild recession”
- the third state represents “severe recession”

For example, the matrix tells us that when the state is normal growth, the state will again be normal growth next month with probability 0.97

In general, large values on the main diagonal indicate persistence in the process $\{X_t\}$

This Markov process can also be represented as a directed graph, with edges labeled by transition probabilities



Here “ng” is normal growth, “mr” is mild recession, etc.

Simulation

One natural way to answer questions about Markov chains is to simulate them

(To approximate the probability of event E , we can simulate many times and count the fraction of times that E occurs)

Nice functionality for simulating Markov chains exists in [QuantEcon](#)

- Efficient, bundled with lots of other useful routines for handling Markov chains

However, it's also a good exercise to roll our own routines — let's do that first and then come back to the methods in QuantEcon

In this exercises we'll take the state space to be $S = 1, \dots, n$

Rolling our own To simulate a Markov chain, we need its stochastic matrix P and either an initial state or a probability distribution ψ for initial state to be drawn from

The Markov chain is then constructed as discussed above. To repeat:

1. At time $t = 0$, the X_0 is set to some fixed state or chosen from ψ
2. At each subsequent time t , the new state X_{t+1} is drawn from $P(X_t, \cdot)$

In order to implement this simulation procedure, we need a method for generating draws from a discrete distributions

For this task we'll use `DiscreteRV` from QuantEcon

```
julia> using QuantEcon

julia> psi = [0.1, 0.9];      # Probabilities over sample space {1, 2}

julia> d = DiscreteRV(psi);

julia> draw(d, 5)           # Generate 5 independent draws from psi
5-element Array{Int64,1}:
 1
 2
 2
 1
 2
```

We'll write our code as a function that takes the following three arguments

- A stochastic matrix P
- An initial state `init`
- A positive integer `sample_size` representing the length of the time series the function should return

```
using QuantEcon

function mc_sample_path(P; init=1, sample_size=1000)
    X = Array(Int64, sample_size) # allocate memory
    X[1] = init
    # === convert each row of P into a distribution === #
    n = size(P)[1]
    P_dist = [DiscreteRV(vec(P[i,:])) for i in 1:n]

    # === generate the sample path === #
    for t in 1:(sample_size - 1)
        X[t+1] = draw(P_dist[X[t]])
    end
end
```

```

    return X
end

```

Let's see how it works using the small matrix

$$P := \begin{pmatrix} 0.4 & 0.6 \\ 0.2 & 0.8 \end{pmatrix} \quad (2.7)$$

As we'll see later, for a long series drawn from P , the fraction of the sample that takes value 1 will be about 0.25

If you run the following code you should get roughly that answer

```

julia> P = [0.4 0.6; 0.2 0.8]
2x2 Array{Float64,2}:
 0.4  0.6
 0.2  0.8

julia> X = mc_sample_path(P, sample_size=100000);

julia> println(mean(X .== 1))
0.25171

```

Using QuantEcon's Routines As discussed above, QuantEcon has routines for handling Markov chains, including simulation

Here's an illustration using the same P as the preceding example

```

julia> using QuantEcon

julia> P = [0.4 0.6; 0.2 0.8];

julia> mc = MarkovChain(P)
Discrete Markov Chain
stochastic matrix of type Array{Float64,2}:
2x2 Array{Float64,2}:
 0.4  0.6
 0.2  0.8

julia> X = simulate(mc, 100000);

julia> mean(X .== 1)  # Should be close to 0.25
0.25031

```

Adding state values and initial conditions If we wish to, we can provide a specification of state values to `MarkovChain`

These state values can be integers, floats, or even strings

The following code illustrates

```
julia> mc = MarkovChain(P, ["employed", "unemployed"])
Discrete Markov Chain
stochastic matrix of type Array{Float64,2}:
2x2 Array{Float64,2}:
 0.4  0.6
 0.2  0.8

julia> simulate(mc, 4, init=1)  # Start at state 1 (employed)
4-element Array{String,1}:
 "employed"
 "employed"
 "unemployed"
 "unemployed"

julia> simulate(mc, 4, init=2)  # Start at state 2 (unemployed)
4-element Array{String,1}:
 "unemployed"
 "unemployed"
 "unemployed"
 "unemployed"

julia> simulate(mc, 4)  # Start with randomly chosen initial condition
4-element Array{ASCIIString,1}:
 "employed"
 "unemployed"
 "unemployed"
 "unemployed"
```

If we want indices rather than state values we can use

```
julia> simulate_indices(mc, 4)
4-element Array{Int64,1}:
 2
 2
 2
 1
```

Marginal Distributions

Suppose that

1. $\{X_t\}$ is a Markov chain with stochastic matrix P
2. the distribution of X_t is known to be ψ_t

What then is the distribution of X_{t+1} , or, more generally, of X_{t+m} ?

Solution Let ψ_t be the distribution of X_t for $t = 0, 1, 2, \dots$

Our first aim is to find ψ_{t+1} given ψ_t and P

To begin, pick any $y \in S$

Using the law of total probability, we can decompose the probability that $X_{t+1} = y$ as follows:

$$\mathbb{P}\{X_{t+1} = y\} = \sum_{x \in S} \mathbb{P}\{X_{t+1} = y | X_t = x\} \cdot \mathbb{P}\{X_t = x\}$$

In words, to get the probability of being at y tomorrow, we account for all ways this can happen and sum their probabilities

Rewriting this statement in terms of marginal and conditional probabilities gives

$$\psi_{t+1}(y) = \sum_{x \in S} P(x, y) \psi_t(x)$$

There are n such equations, one for each $y \in S$

If we think of ψ_{t+1} and ψ_t as *row vectors* (as is traditional in this literature), these n equations are summarized by the matrix expression

$$\psi_{t+1} = \psi_t P \tag{2.8}$$

In other words, to move the distribution forward one unit of time, we postmultiply by P

By repeating this m times we move forward m steps into the future

Hence, iterating on (2.8), the expression $\psi_{t+m} = \psi_t P^m$ is also valid — here P^m is the m -th power of P . As a special case, we see that if ψ_0 is the initial distribution from which X_0 is drawn, then $\psi_0 P^m$ is the distribution of X_m .

This is very important, so let's repeat it

$$X_0 \sim \psi_0 \implies X_m \sim \psi_0 P^m \tag{2.9}$$

and, more generally,

$$X_t \sim \psi_t \implies X_{t+m} \sim \psi_t P^m \tag{2.10}$$

Multiple Step Transition Probabilities We know that the probability of transitioning from x to y in one step is $P(x, y)$

It turns out that the probability of transitioning from x to y in m steps is $P^m(x, y)$, the (x, y) -th element of the m -th power of P

To see why, consider again (2.10), but now with ψ_t putting all probability on state x

- 1 in the x -th position and zero elsewhere

Inserting this into (2.10), we see that, conditional on $X_t = x$, the distribution of X_{t+m} is the x -th row of P^m

In particular

$$\mathbb{P}\{X_{t+m} = y\} = P^m(x, y) = (x, y)\text{-th element of } P^m$$

Example: Probability of Recession Recall the stochastic matrix P for recession and growth *considered above*

Suppose that the current state is unknown — perhaps statistics are available only at the *end* of the current month

We estimate the probability that the economy is in state x to be $\psi(x)$

The probability of being in recession (either mild or severe) in 6 months time is given by the inner product

$$\psi P^6 \cdot \begin{pmatrix} 0 \\ 1 \\ 1 \end{pmatrix}$$

Example 2: Cross-Sectional Distributions The marginal distributions we have been studying can be viewed either as probabilities or as cross-sectional frequencies in large samples

To illustrate, recall our model of employment / unemployment dynamics for a given worker *discussed above*

Consider a large (i.e., tending to infinite) population of workers, each of whose lifetime experiences are described by the specified dynamics, independently of one another

Let ψ be the current *cross-sectional* distribution over $\{1, 2\}$

- For example, $\psi(1)$ is the unemployment rate

The cross-sectional distribution records the fractions of workers employed and unemployed at a given moment

The same distribution also describes the fractions of a particular worker's career spent being employed and unemployed, respectively

Irreducibility and Aperiodicity

Irreducibility and aperiodicity are central concepts of modern Markov chain theory

Let's see what they're about

Irreducibility Let P be a fixed stochastic matrix

Two states x and y are said to **communicate** with each other if there exist positive integers j and k such that

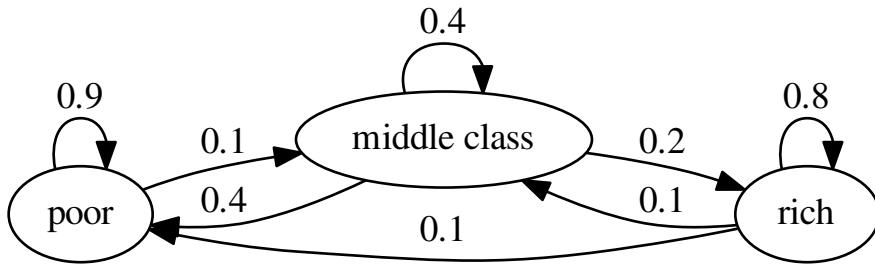
$$P^j(x, y) > 0 \quad \text{and} \quad P^k(y, x) > 0$$

In view of our discussion *above*, this means precisely that

- state x can be reached eventually from state y , and
- state y can be reached eventually from state x

The stochastic matrix P is called **irreducible** if all states communicate; that is, if x and y communicate for all (x, y) in $S \times S$

For example, consider the following transition probabilities for wealth of a fictitious set of households



We can translate this into a stochastic matrix, putting zeros where there's no edge between nodes

$$P := \begin{pmatrix} 0.9 & 0.1 & 0 \\ 0.4 & 0.4 & 0.2 \\ 0.1 & 0.1 & 0.8 \end{pmatrix}$$

It's clear from the graph that this stochastic matrix is irreducible: we can reach any state from any other state eventually

We can also test this using QuantEcon's `MarkovChain` class

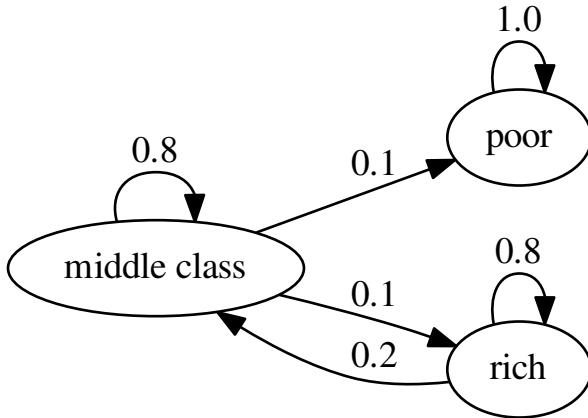
```
julia> using QuantEcon

julia> P = [0.9 0.1 0.0; 0.4 0.4 0.2; 0.1 0.1 0.8];

julia> mc = MarkovChain(P)
Discrete Markov Chain
stochastic matrix of type Array{Float64,2}:
3x3 Array{Float64,2}:
 0.9  0.1  0.0
 0.4  0.4  0.2
 0.1  0.1  0.8

julia> is_irreducible(mc)
true
```

Here's a more pessimistic scenario, where the poor are poor forever



This stochastic matrix is not irreducible, since, for example, *rich* is not accessible from *poor*. Let's confirm this

```

julia> using QuantEcon

julia> P = [1.0 0.0 0.0; 0.1 0.8 0.1; 0.0 0.2 0.8];

julia> mc = MarkovChain(P);

julia> is_irreducible(mc)
false
  
```

We can also determine the “communication classes”

```

julia> communication_classes(mc)
2-element Array{Array{Int64,1},1}:
 [1]
 [2,3]
  
```

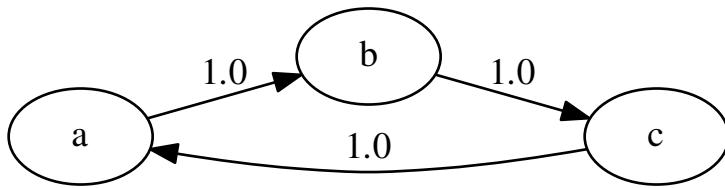
It might be clear to you already that irreducibility is going to be important in terms of long run outcomes

For example, poverty is a life sentence in the second graph but not the first

We'll come back to this a bit later

Aperiodicity Loosely speaking, a Markov chain is called periodic if it cycles in a predictable way, and aperiodic otherwise

Here's a trivial example with three states



The chain cycles with period 3:

```

julia> using QuantEcon
julia> P = [0 1 0; 0 0 1; 1 0 0];
julia> mc = MarkovChain(P);
julia> period(mc)
3
  
```

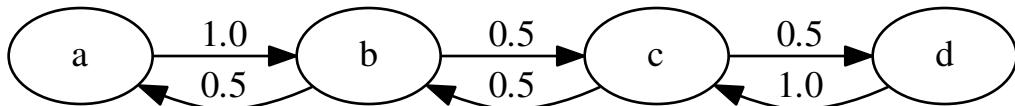
More formally, the **period** of a state x is the greatest common divisor of the set of integers

$$D(x) := \{j \geq 1 : P^j(x, x) > 0\}$$

In the last example, $D(x) = \{3, 6, 9, \dots\}$ for every state x , so the period is 3

A stochastic matrix is called **aperiodic** if the period of every state is 1, and **periodic** otherwise

For example, the stochastic matrix associated with the transition probabilities below is periodic because, for example, state a has period 2



We can confirm that the stochastic matrix is periodic as follows

```

julia> P = zeros(4, 4);
julia> P[1,2] = 1;
julia> P[2, 1] = P[2, 3] = 0.5;
julia> P[3, 2] = P[3, 4] = 0.5;
  
```

```
julia> P[4, 3] = 1;
julia> mc = MarkovChain(P);
julia> period(mc)
2
julia> is_aperiodic(mc)
false
```

Stationary Distributions

As seen in (2.8), we can shift probabilities forward one unit of time via postmultiplication by P . Some distributions are invariant under this updating process — for example,

```
julia> P = [.4 .6; .2 .8];
julia> psi = [0.25, 0.75];
julia> psi' * P
1x2 Array{Float64,2}:
 0.25  0.75
```

Such distributions are called **stationary**, or **invariant**. Formally, a distribution ψ^* on S is called **stationary** for P if $\psi^* = \psi^*P$.

From this equality we immediately get $\psi^* = \psi^*P^t$ for all t .

This tells us an important fact: If the distribution of X_0 is a stationary distribution, then X_t will have this same distribution for all t .

Hence stationary distributions have a natural interpretation as stochastic steady states — we'll discuss this more in just a moment.

Mathematically, a stationary distribution is a fixed point of P when P is thought of as the map $\psi \mapsto \psi P$ from (row) vectors to (row) vectors.

Theorem Every stochastic matrix P has at least one stationary distribution.

(We are assuming here that the state space S is finite; if not more assumptions are required.)

For a proof of this result you can apply Brouwer's fixed point theorem, or see [EDTC](#), theorem 4.3.5.

There may in fact be many stationary distributions corresponding to a given stochastic matrix P .

- For example, if P is the identity matrix, then all distributions are stationary.

Since stationary distributions are long run equilibria, to get uniqueness we require that initial conditions are not infinitely persistent.

Infinite persistence of initial conditions occurs if certain regions of the state space cannot be accessed from other regions, which is the opposite of irreducibility.

This gives some intuition for the following fundamental theorem **Theorem**. If P is both aperiodic and irreducible, then

1. P has exactly one stationary distribution ψ^*
2. For any initial distribution ψ_0 , we have $\|\psi_0 P^t - \psi^*\| \rightarrow 0$ as $t \rightarrow \infty$

For a proof, see, for example, theorem 5.2 of [Haggstrom02]

(Note that part 1 of the theorem requires only irreducibility, whereas part 2 requires both irreducibility and aperiodicity)

A stochastic matrix satisfying the conditions of the theorem is sometimes called **uniformly ergodic**

One easy sufficient condition for aperiodicity and irreducibility is that every element of P is strictly positive

- Try to convince yourself of this

Example Recall our model of employment / unemployment dynamics for a given worker *discussed above*

Assuming $\alpha \in (0, 1)$ and $\beta \in (0, 1)$, the uniform ergodicity condition is satisfied

Let $\psi^* = (p, 1 - p)$ be the stationary distribution, so that p corresponds to unemployment (state 1)

Using $\psi^* = \psi^* P$ and a bit of algebra yields

$$p = \frac{\beta}{\alpha + \beta}$$

This is, in some sense, a steady state probability of unemployment — more on interpretation below

Not surprisingly it tends to zero as $\beta \rightarrow 0$, and to one as $\alpha \rightarrow 0$

Calculating Stationary Distributions As discussed above, a given Markov matrix P can have many stationary distributions

That is, there can be many row vectors ψ such that $\psi = \psi P$

In fact if P has two distinct stationary distributions ψ_1, ψ_2 then it has infinitely many, since in this case, as you can verify,

$$\psi_3 := \lambda\psi_1 + (1 - \lambda)\psi_2$$

is a stationary distribution for P for any $\lambda \in [0, 1]$

If we restrict attention to the case where only one stationary distribution exists, one option for finding it is to try to solve the linear system $\psi(I_n - P) = 0$ for ψ , where I_n is the $n \times n$ identity

But the zero vector solves this equation

Hence we need to impose the restriction that the solution must be a probability distribution

A suitable algorithm is implemented in QuantEcon — the next code block illustrates

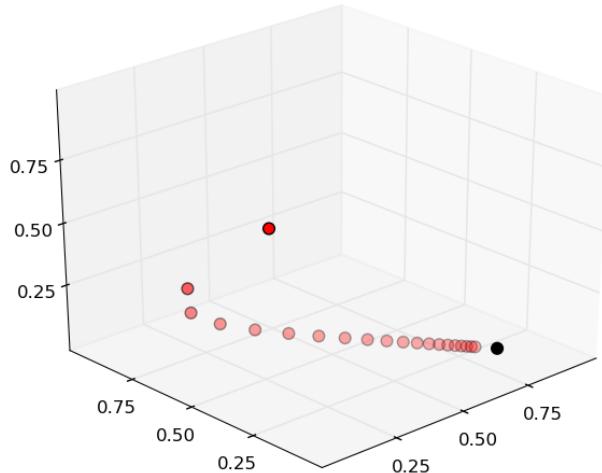
```
julia> P = [.4 .6; .2 .8];
julia> mc = MarkovChain(P);
julia> stationary_distributions(mc)
1-element Array{Array{Float64,1},1}:
 [0.25, 0.7499999999999999]
```

The stationary distribution is unique

Convergence to Stationarity Part 2 of the Markov chain convergence theorem *stated above* tells us that the distribution of X_t converges to the stationary distribution regardless of where we start off

This adds considerable weight to our interpretation of ψ^* as a stochastic steady state

The convergence in the theorem is illustrated in the next figure



Here

- P is the stochastic matrix for recession and growth *considered above*
- The highest red dot is an arbitrarily chosen initial probability distribution ψ , represented as a vector in \mathbb{R}^3
- The other red dots are the distributions ψP^t for $t = 1, 2, \dots$
- The black dot is ψ^*

The code for the figure [can be found](#) in the QuantEcon applications library — you might like to try experimenting with different initial conditions

Ergodicity

Under irreducibility, yet another important result obtains: For all $x \in S$,

$$\frac{1}{n} \sum_{t=1}^n \mathbf{1}\{X_t = x\} \rightarrow \psi^*(x) \quad \text{as } n \rightarrow \infty \quad (2.11)$$

Here

- $\mathbf{1}\{X_t = x\} = 1$ if $X_t = x$ and zero otherwise
- convergence is with probability one
- the result does not depend on the distribution (or value) of X_0

The result tells us that the fraction of time the chain spends at state x converges to $\psi^*(x)$ as time goes to infinity. This gives us another way to interpret the stationary distribution — provided that the convergence result in (2.11) is valid.

The convergence in (2.11) is a special case of a law of large numbers result for Markov chains — see [EDTC](#), section 4.3.4 for some additional information.

Example Recall our cross-sectional interpretation of the employment / unemployment model discussed above

Assume that $\alpha \in (0, 1)$ and $\beta \in (0, 1)$, so that irreducibility and aperiodicity both hold.

We saw that the stationary distribution is $(p, 1 - p)$, where

$$p = \frac{\beta}{\alpha + \beta}$$

In the cross-sectional interpretation, this is the fraction of people unemployed.

In view of our latest (ergodicity) result, it is also the fraction of time that a worker can expect to spend unemployed.

Thus, in the long-run, cross-sectional averages for a population and time-series averages for a given person coincide.

This is one interpretation of the notion of ergodicity.

Computing Expectations

We are interested in computing expectations of the form

$$\mathbb{E}[h(X_t)] \quad (2.12)$$

and conditional expectations such as

$$\mathbb{E}[h(X_{t+k}) \mid X_t = x] \quad (2.13)$$

where

- $\{X_t\}$ is a Markov chain generated by $n \times n$ stochastic matrix P
- h is a given function, which, in expressions involving matrix algebra, we'll think of as the column vector

$$h = \begin{pmatrix} h(x_1) \\ \vdots \\ h(x_n) \end{pmatrix}$$

The unconditional expectation (2.12) is easy: We just sum over the distribution of X_t to get

$$\mathbb{E}[h(X_t)] = \sum_{x \in S} (\psi P^t)(x)h(x)$$

Here ψ is the distribution of X_0

Since ψ and hence ψP^t are row vectors, we can also write this as

$$\mathbb{E}[h(X_t)] = \psi P^t h$$

For the conditional expectation (2.13), we need to sum over the conditional distribution of X_{t+k} given $X_t = x$

We already know that this is $P^k(x, \cdot)$, so

$$\mathbb{E}[h(X_{t+k}) \mid X_t = x] = (P^k h)(x) \quad (2.14)$$

The vector $P^k h$ stores the conditional expectation $\mathbb{E}[h(X_{t+k}) \mid X_t = x]$ over all x

Expectations of Geometric Sums Sometimes we also want to compute expectations of a geometric sum, such as $\sum_t \beta^t h(X_t)$

In view of the preceding discussion, this is

$$\mathbb{E} \left[\sum_{j=0}^{\infty} \beta^j h(X_{t+j}) \mid X_t = x \right] = [(I - \beta P)^{-1} h](x)$$

where

$$(I - \beta P)^{-1} = I + \beta P + \beta^2 P^2 + \dots$$

Premultiplication by $(I - \beta P)^{-1}$ amounts to “applying the **resolvent operator**”

Exercises

Exercise 1 According to the discussion *immediately above*, if a worker's employment dynamics obey the stochastic matrix

$$P = \begin{pmatrix} 1 - \alpha & \alpha \\ \beta & 1 - \beta \end{pmatrix}$$

with $\alpha \in (0, 1)$ and $\beta \in (0, 1)$, then, in the long-run, the fraction of time spent unemployed will be

$$p := \frac{\beta}{\alpha + \beta}$$

In other words, if $\{X_t\}$ represents the Markov chain for employment, then $\bar{X}_n \rightarrow p$ as $n \rightarrow \infty$, where

$$\bar{X}_n := \frac{1}{n} \sum_{t=1}^n \mathbf{1}\{X_t = 1\}$$

Your exercise is to illustrate this convergence

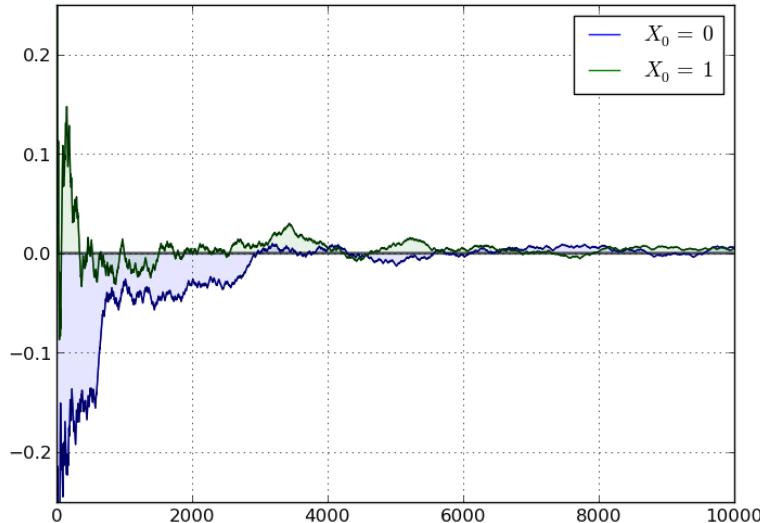
First,

- generate one simulated time series $\{X_t\}$ of length 10,000, starting at $X_0 = 1$
- plot $\bar{X}_n - p$ against n , where p is as defined above

Second, repeat the first step, but this time taking $X_0 = 2$

In both cases, set $\alpha = \beta = 0.1$

The result should look something like the following — modulo randomness, of course



(You don't need to add the fancy touches to the graph—see the solution if you're interested)

Exercise 2 A topic of interest for economics and many other disciplines is *ranking*

Let's now consider one of the most practical and important ranking problems — the rank assigned to web pages by search engines

(Although the problem is motivated from outside of economics, there is in fact a deep connection between search ranking systems and prices in certain competitive equilibria — see [DLP13])

To understand the issue, consider the set of results returned by a query to a web search engine

For the user, it is desirable to

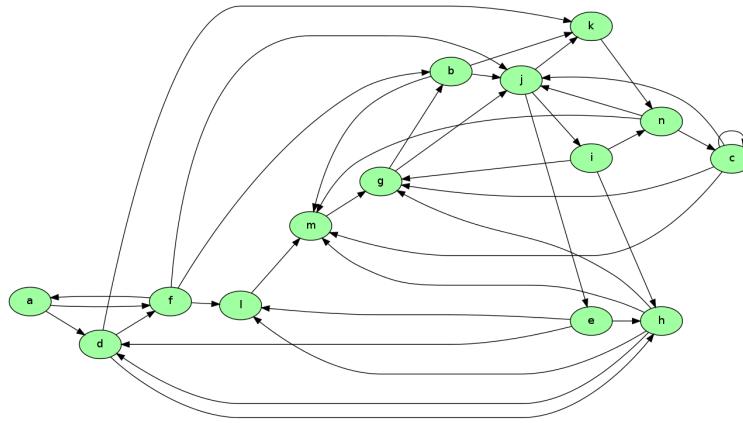
1. receive a large set of accurate matches

2. have the matches returned in order, where the order corresponds to some measure of “importance”

Ranking according to a measure of importance is the problem we now consider

The methodology developed to solve this problem by Google founders Larry Page and Sergey Brin is known as [PageRank](#)

To illustrate the idea, consider the following diagram



Imagine that this is a miniature version of the WWW, with

- each node representing a web page
- each arrow representing the existence of a link from one page to another

Now let's think about which pages are likely to be important, in the sense of being valuable to a search engine user

One possible criterion for importance of a page is the number of inbound links — an indication of popularity

By this measure, m and j are the most important pages, with 5 inbound links each

However, what if the pages linking to m , say, are not themselves important?

Thinking this way, it seems appropriate to weight the inbound nodes by relative importance

The PageRank algorithm does precisely this

A slightly simplified presentation that captures the basic idea is as follows

Letting j be (the integer index of) a typical page and r_j be its ranking, we set

$$r_j = \sum_{i \in L_j} \frac{r_i}{\ell_i}$$

where

- ℓ_i is the total number of outbound links from i
- L_j is the set of all pages i such that i has a link to j

This is a measure of the number of inbound links, weighted by their own ranking (and normalized by $1/\ell_i$)

There is, however, another interpretation, and it brings us back to Markov chains

Let P be the matrix given by $P(i, j) = \mathbf{1}\{i \rightarrow j\}/\ell_i$ where $\mathbf{1}\{i \rightarrow j\} = 1$ if i has a link to j and zero otherwise

The matrix P is a stochastic matrix provided that each page has at least one link

With this definition of P we have

$$r_j = \sum_{i \in L_j} \frac{r_i}{\ell_i} = \sum_{\text{all } i} \mathbf{1}\{i \rightarrow j\} \frac{r_i}{\ell_i} = \sum_{\text{all } i} P(i, j) r_i$$

Writing r for the row vector of rankings, this becomes $r = rP$

Hence r is the stationary distribution of the stochastic matrix P

Let's think of $P(i, j)$ as the probability of "moving" from page i to page j

The value $P(i, j)$ has the interpretation

- $P(i, j) = 1/k$ if i has k outbound links, and j is one of them
- $P(i, j) = 0$ if i has no direct link to j

Thus, motion from page to page is that of a web surfer who moves from one page to another by randomly clicking on one of the links on that page

Here "random" means that each link is selected with equal probability

Since r is the stationary distribution of P , assuming that the uniform ergodicity condition is valid, we can interpret r_j as the fraction of time that a (very persistent) random surfer spends at page j

Your exercise is to apply this ranking algorithm to the graph pictured above, and return the list of pages ordered by rank

The data for this graph is in the `web_graph_data.txt` file from the [main repository](#) — you can also view it [here](#)

There is a total of 14 nodes (i.e., web pages), the first named `a` and the last named `n`

A typical line from the file has the form

```
d -> h;
```

This should be interpreted as meaning that there exists a link from `d` to `h`

To parse this file and extract the relevant information, you can use [regular expressions](#)

The following code snippet provides a hint as to how you can go about this

```
julia> matchall(r"\w", "x +++ y ***** z")
3-element Array{SubString{UTF8String},1}:
 "x"
 "y"
 "z"
```

```
julia> matchall(r"\w", "a ^ b && \$\$ c")
3-element Array{SubString{UTF8String},1}:
 "a"
 "b"
 "c"
```

When you solve for the ranking, you will find that the highest ranked node is in fact g , while the lowest is a

Exercise 3 In numerical work it is sometimes convenient to replace a continuous model with a discrete one

In particular, Markov chains are routinely generated as discrete approximations to AR(1) processes of the form

$$y_{t+1} = \rho y_t + u_{t+1}$$

Here u_t is assumed to be iid and $N(0, \sigma_u^2)$

The variance of the stationary probability distribution of $\{y_t\}$ is

$$\sigma_y^2 := \frac{\sigma_u^2}{1 - \rho^2}$$

Tauchen's method [Tau86] is the most common method for approximating this continuous state process with a finite state Markov chain

A routine for this already exists in QuantEcon.jl but let's write our own version as an exercise

As a first step we choose

- n , the number of states for the discrete approximation
- m , an integer that parameterizes the width of the state space

Next we create a state space $\{x_0, \dots, x_{n-1}\} \subset \mathbb{R}$ and a stochastic $n \times n$ matrix P such that

- $x_0 = -m \sigma_y$
- $x_{n-1} = m \sigma_y$
- $x_{i+1} = x_i + s$ where $s = (x_{n-1} - x_0)/(n - 1)$

Let F be the cumulative distribution function of the normal distribution $N(0, \sigma_u^2)$

The values $P(x_i, x_j)$ are computed to approximate the AR(1) process — omitting the derivation, the rules are as follows:

1. If $j = 0$, then set

$$P(x_i, x_0) = P(x_i, x_0) = F(x_0 - \rho x_i + s/2)$$

2. If $j = n - 1$, then set

$$P(x_i, x_{n-1}) = P(x_i, x_{n-1}) = 1 - F(x_{n-1} - \rho x_i - s/2)$$

3. Otherwise, set

$$P(x_i, x_j) = F(x_j - \rho x_i + s/2) - F(x_j - \rho x_i - s/2)$$

The exercise is to write a function `approx_markov(rho, sigma_u, m=3, n=7)` that returns $\{x_0, \dots, x_{n-1}\} \subset \mathbb{R}$ and $n \times n$ matrix P as described above

- Even better, write a function that returns an instance of QuantEcon.jl's `MarkovChain` type

Solutions

[Solution notebook](#)

Orthogonal Projection and its Applications

Contents

- Orthogonal Projection and its Applications
 - [Overview](#)
 - [Key Definitions](#)
 - [The Orthogonal Projection Theorem](#)
 - [Orthonormal Bases](#)
 - [Projection Using Matrix Algebra](#)
 - [Least Squares Regression](#)
 - [Orthogonalization and Decomposition](#)
 - [Exercises](#)
 - [Solutions](#)

Overview

Orthogonal projection is a cornerstone of vector space methods, with many diverse applications

This include, but are not limited to,

- Least squares and linear regression
- Conditional expectation
- Gram–Schmidt orthogonalization
- QR decomposition
- Orthogonal polynomials
- etc

In this lecture we focus on

- key results
- standard applications such as least squares regression

Further Reading For background and foundational concepts, see our lecture on linear algebra

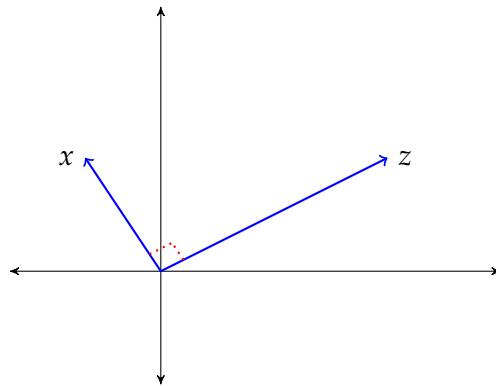
For more proofs and greater theoretical detail, see [A Primer in Econometric Theory](#)

For a complete set of proofs in a general setting, see, for example, [\[Rom05\]](#)

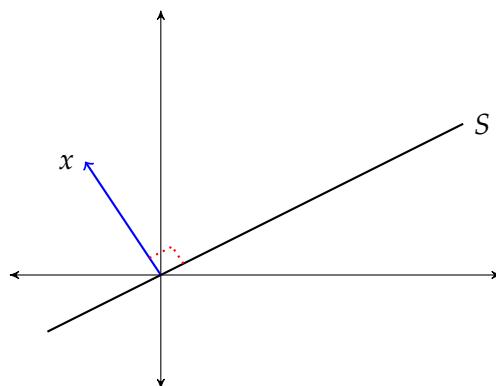
For an advanced treatment of projection in the context of least squares prediction, see [this book chapter](#)

Key Definitions

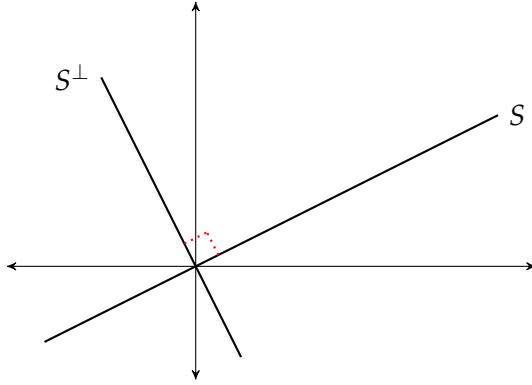
If $x, z \in \mathbb{R}^n$ and $\langle x, z \rangle = 0$ then x and z are said to be **orthogonal**, and we write $x \perp z$



Given $S \subset \mathbb{R}^n$, we call $x \in \mathbb{R}^n$ **orthogonal to S** if $x \perp z$ for all $z \in S$, and write $x \perp S$



The **orthogonal complement** of linear subspace S is the set $S^\perp := \{x \in \mathbb{R}^n : x \perp S\}$



Note that S^\perp is always a linear subspace of \mathbb{R}^n

To see this, fix $x, y \in S^\perp$ and $\alpha, \beta \in \mathbb{R}$

Observe that if $z \in S$, then

$$\langle \alpha x + \beta y, z \rangle = \alpha \langle x, z \rangle + \beta \langle y, z \rangle = \alpha \times 0 + \beta \times 0 = 0$$

Hence $\alpha x + \beta y \in S^\perp$, as was to be shown

A set of vectors $\{x_1, \dots, x_k\} \subset \mathbb{R}^n$ is called an **orthogonal set** if $x_i \perp x_j$ whenever $i \neq j$

If $\{x_1, \dots, x_k\}$ is an orthogonal set, then the **Pythagorean Law** states that

$$\|x_1 + \dots + x_k\|^2 = \|x_1\|^2 + \dots + \|x_k\|^2$$

In the case of $k = 2$ this is easy to see, since orthogonality gives

$$\|x_1 + x_2\|^2 = \langle x_1 + x_2, x_1 + x_2 \rangle = \langle x_1, x_1 \rangle + 2\langle x_2, x_1 \rangle + \langle x_2, x_2 \rangle = \|x_1\|^2 + \|x_2\|^2$$

Linear Independence vs Orthogonality If $X \subset \mathbb{R}^n$ is an orthogonal set and $0 \notin X$, then X is linearly independent

Proving this is a nice exercise

While the converse is not true, a kind of partial converse holds, as we'll see below

The Orthogonal Projection Theorem

The problem considered by the orthogonal projection theorem (OPT) is to find the closest approximation to an arbitrary vector from within a given linear subspace

The theorem, stated below, tells us that this problem always has a unique solution, and provides a very useful characterization

Theorem (OPT) Given $y \in \mathbb{R}^n$ and linear subspace $S \subset \mathbb{R}^n$, there exists a unique solution to the minimization problem

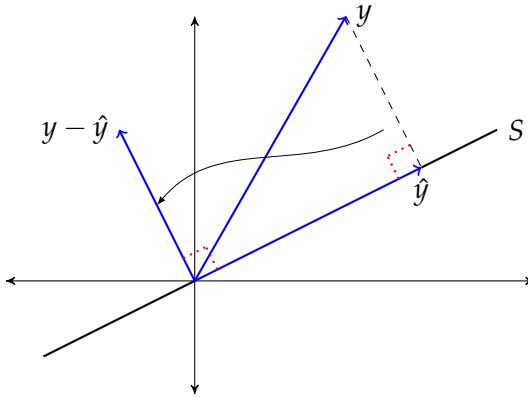
$$\hat{y} := \arg \min_{z \in S} \|y - z\|$$

Moreover, the solution \hat{y} is the unique vector in \mathbb{R}^n such that

- $\hat{y} \in S$
- $y - \hat{y} \perp S$

The vector \hat{y} is called the **orthogonal projection** of y onto S

The next figure provides some intuition



We'll omit the full proof but let's at least cover sufficiency of the conditions

To this end, let $y \in \mathbb{R}^n$ and let S be a linear subspace of \mathbb{R}^n

Let \hat{y} be a vector in \mathbb{R}^n such that $\hat{y} \in S$ and $y - \hat{y} \perp S$

Letting z be any other point in S and using the fact that S is a linear subspace, we have

$$\|y - z\|^2 = \|(y - \hat{y}) + (\hat{y} - z)\|^2 = \|y - \hat{y}\|^2 + \|\hat{y} - z\|^2$$

Hence $\|y - z\| \geq \|y - \hat{y}\|$, which completes the proof

Orthogonal Projection as a Mapping Holding S fixed, we have a functional relationship

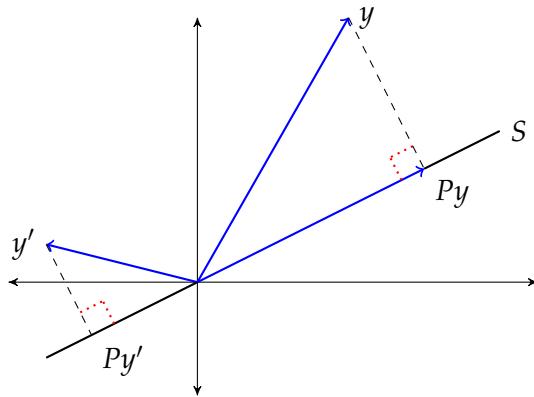
$$y \mapsto \text{its orthogonal projection } \hat{y} \in S$$

By the OPT, this is a well-defined mapping from \mathbb{R}^n to \mathbb{R}^n

In what follows it is denoted by P

- Py represents the projection \hat{y}
- We write $P = \text{proj } S$

The operator P is called the **orthogonal projection mapping onto S**



It is immediate from the OPT that, for any $y \in \mathbb{R}^n$,

1. $Py \in S$ and
2. $y - Py \perp S$

From this we can deduce additional properties, such as

1. $\|y\|^2 = \|Py\|^2 + \|y - Py\|^2$ and
2. $\|Py\| \leq \|y\|$

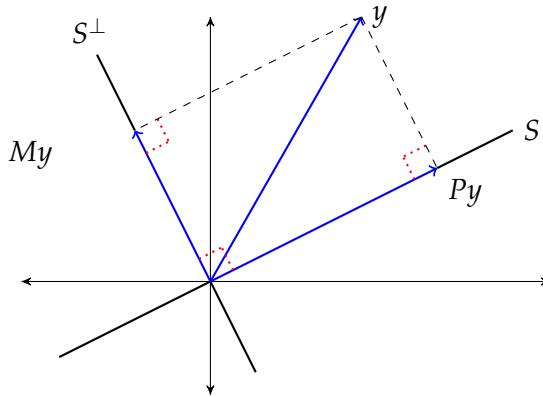
For example, to prove 1, observe that $y = Py + y - Py$ and apply the Pythagorean law

The Residual Projection Here's another version of the OPT

Theorem. If S is a linear subspace of \mathbb{R}^n , $P = \text{proj } S$ and $M = \text{proj } S^\perp$, then

$$Py \perp My \quad \text{and} \quad y = Py + My \quad \text{for all } y \in \mathbb{R}^n$$

The next figure illustrates



Orthonormal Bases

An orthogonal set $O \subset \mathbb{R}^n$ is called an **orthonormal set** if $\|u\| = 1$ for all $u \in O$

Let S be a linear subspace of \mathbb{R}^n and let $O \subset S$

If O is orthonormal and $\text{span } O = S$, then O is called an **orthonormal basis** of S

It is, necessarily, a basis of S (being independent by orthogonality and the fact that no element is the zero vector)

One example of an orthonormal set is the canonical basis $\{e_1, \dots, e_n\}$, which forms an orthonormal basis of \mathbb{R}^n

If $\{u_1, \dots, u_k\}$ is an orthonormal basis of linear subspace S , then we have

$$x = \sum_{i=1}^k \langle x, u_i \rangle u_i \quad \text{for all } x \in S$$

To see this, observe that since $x \in \text{span}\{u_1, \dots, u_k\}$, we can find scalars $\alpha_1, \dots, \alpha_k$ s.t.

$$x = \sum_{j=1}^k \alpha_j u_j \tag{2.15}$$

Taking the inner product with respect to u_i gives

$$\langle x, u_i \rangle = \sum_{j=1}^k \alpha_j \langle u_j, u_i \rangle = \alpha_i$$

Combining this result with (2.15) verifies the claim

Projection onto an Orthonormal Basis When we have an orthonormal basis for the subspace we are projecting onto, computing the projection is straightforward:

Theorem If $\{u_1, \dots, u_k\}$ is an orthonormal basis for S , then

$$Py = \sum_{i=1}^k \langle y, u_i \rangle u_i, \quad \forall y \in \mathbb{R}^n \tag{2.16}$$

Proof: Fix $y \in \mathbb{R}^n$ and let Py be as defined in (2.16)

Clearly, $Py \in S$

We claim that $y - Py \perp S$ also holds

It suffices to show that $y - Py \perp$ any basis element (why?)

This is true because

$$\left\langle y - \sum_{i=1}^k \langle y, u_i \rangle u_i, u_j \right\rangle = \langle y, u_j \rangle - \sum_{i=1}^k \langle y, u_i \rangle \langle u_i, u_j \rangle = 0$$

Projection Using Matrix Algebra

It is not too difficult to show that if S is any linear subspace of \mathbb{R}^n and $P = \text{proj } S$, then P is a linear function from \mathbb{R}^n to \mathbb{R}^n

It follows that $P = \text{proj } S$ can be represented as a matrix

Below we use P for both the orthogonal projection mapping and the matrix that represents it

But what does the matrix look like?

Theorem. If $P = \text{proj } S$ and the columns of $n \times k$ matrix X form a basis of S , then

$$P = X(X'X)^{-1}X'$$

Proof: Given arbitrary $y \in \mathbb{R}^n$ and $P = X(X'X)^{-1}X'$, our claim is that

1. $Py \in S$, and

2. $y - Py \perp S$

Here 1 is true because

$$Py = X(X'X)^{-1}X'y = Xa \quad \text{when } a := (X'X)^{-1}X'y$$

An expression of the form Xa is precisely a linear combination of the columns of X , and hence an element of S

On the other hand, 2 is equivalent to the statement

$$y - X(X'X)^{-1}X'y \perp Xb \quad \text{for all } b \in \mathbb{R}^k$$

This is true: If $b \in \mathbb{R}^k$, then

$$(Xb)'[y - X(X'X)^{-1}X'y] = b'[X'y - X'y] = 0$$

The proof is now complete

It is common in applications to start with $n \times k$ matrix X with linearly independent columns and let

$$S := \text{span } X := \text{span}\{\text{col}_1 X, \dots, \text{col}_k X\}$$

Then the columns of X form a basis of S

From the preceding theorem, $P = X(X'X)^{-1}X'$ projects onto S

In this context, $P = \text{proj } S$ is often called the **projection matrix**

- The matrix $M = I - P$ satisfies $M = \text{proj}(S^\perp)$ and is sometimes called the **annihilator**

As a further illustration of the last result, suppose that U is $n \times k$ with orthonormal columns

Let $u_i := \text{col } U_i$ for each i , let $S := \text{span } U$ and let $y \in \mathbb{R}^n$

We know that the projection of y onto S is

$$Py = U(U'U)^{-1}U'y$$

Since U has orthonormal columns, we have $U'U = I$

Hence

$$Py = UU'y = \sum_{i=1}^k \langle u_i, y \rangle u_i$$

We have recovered our earlier result about projecting onto the span of an orthonormal basis

Application: Overdetermined Systems of Equations Consider linear system $Xb = y$ where $y \in \mathbb{R}^n$ and X is $n \times k$ with linearly independent columns

Given X and y , we seek $b \in \mathbb{R}^k$ satisfying this equation

If $n > k$ (more equations than unknowns), then the system is said to be **overdetermined**

Intuitively, we may not be able find a b that satisfies all n equations

The best approach here is to

- Accept that an exact solution may not exist
- Look instead for an approximate solution

By approximate solution, we mean a $b \in \mathbb{R}^k$ such that Xb is as close to y as possible

The next theorem shows that the solution is well defined and unique

The proof is based around the OPT

Theorem The unique minimizer of $\|y - Xb\|$ over $b \in \mathbb{R}^K$ is

$$\hat{\beta} := (X'X)^{-1}X'y$$

Proof: Note that

$$X\hat{\beta} = X(X'X)^{-1}X'y = Py$$

Since Py is the orthogonal projection onto $\text{span}(X)$ we have

$$\|y - Py\| \leq \|y - z\| \text{ for any } z \in \text{span}(X)$$

In other words,

$$\|y - X\hat{\beta}\| \leq \|y - Xb\| \text{ for any } b \in \mathbb{R}^K$$

This is what we aimed to show

Least Squares Regression

Let's apply the theory of orthogonal projection to least squares regression

This approach provides insight on many fundamental geometric and theoretical properties of linear regression

We treat only some of the main ideas

The Setting Here's one way to introduce linear regression

Given pairs $(x, y) \in \mathbb{R}^K \times \mathbb{R}$, consider choosing $f: \mathbb{R}^K \rightarrow \mathbb{R}$ to minimize the **risk**

$$R(f) := \mathbb{E} [(y - f(x))^2]$$

If probabilities and hence \mathbb{E} are unknown, we cannot solve this problem directly

However, if a sample is available, we can estimate the risk with the **empirical risk**:

$$\min_{f \in \mathcal{F}} \frac{1}{N} \sum_{n=1}^N (y_n - f(x_n))^2$$

Minimizing this expression is called **empirical risk minimization**

The set \mathcal{F} is sometimes called the hypothesis space

The theory of statistical learning tells us we should take it to be relatively simple to prevent overfitting

If we let \mathcal{F} be the class of linear functions and drop the constant $1/N$, the problem is

$$\min_{b \in \mathbb{R}^K} \sum_{n=1}^N (y_n - b' x_n)^2$$

This is the **linear least squares problem**

Solution To switch to matrix notation, define

$$y := \begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_N \end{pmatrix}, \quad x_n := \begin{pmatrix} x_{n1} \\ x_{n2} \\ \vdots \\ x_{nK} \end{pmatrix} = n\text{-th obs on all regressors}$$

and

We assume throughout that $N > K$ and X is full column rank

If you work through the algebra, you will be able to verify that $\|y - Xb\|^2 = \sum_{n=1}^N (y_n - b' x_n)^2$

Since increasing transforms don't affect minimizers we have

$$\arg \min_{b \in \mathbb{R}^K} \sum_{n=1}^N (y_n - b' x_n)^2 = \arg \min_{b \in \mathbb{R}^K} \|y - Xb\|^2$$

By our results above on overdetermined systems, the solution is

$$\hat{\beta} := (X' X)^{-1} X' y$$

Let P and M be the projection and annihilator associated with X :

$$P := X(X' X)^{-1} X' \quad \text{and} \quad M := I - P$$

The **vector of fitted values** is

$$\hat{y} := X\hat{\beta} = Py$$

The **vector of residuals** is

$$\hat{u} := y - \hat{y} = y - Py = My$$

Here are some more standard definitions:

- The **total sum of squares** is $\|y\|^2$

- The **sum of squared residuals** is $:= \|\hat{u}\|^2$
- The **explained sum of squares** is $:= \|\hat{y}\|^2$

It's well known that $TSS = ESS + SSR$ always holds

We can prove this easily using the OPT

From the OPT we have $y = \hat{y} + \hat{u}$ and $\hat{u} \perp \hat{y}$

Applying the Pythagorean law completes the proof

Many other standards results about least squares regression follow easily from the OPT

Orthogonalization and Decomposition

Let's return to the connection between linear independence and orthogonality touched on above

The main result of interest is a famous algorithm for generating orthonormal sets from linearly independent sets

The next section gives details

Gram-Schmidt Orthogonalization Theorem For each linearly independent set $\{x_1, \dots, x_k\} \subset \mathbb{R}^n$, there exists an orthonormal set $\{u_1, \dots, u_k\}$ with

$$\text{span}\{x_1, \dots, x_i\} = \text{span}\{u_1, \dots, u_i\} \quad \text{for } i = 1, \dots, k$$

Construction uses the **Gram-Schmidt orthogonalization** procedure

One version of this procedure is as follows: For $i = 1, \dots, k$, set

- $S_i := \text{span}\{x_1, \dots, x_i\}$ and $M_i := \text{proj } S_i^\perp$
- $v_i := M_{i-1}x_i$ where M_0 is the identity mapping
- $u_i := v_i / \|v_i\|$

The sequence u_1, \dots, u_k has the stated properties

In the exercises below you are asked to implement this algorithm and test it using projection

QR Decomposition Here's a well known result that uses the preceding algorithm to produce a useful decomposition

Theorem If X is $n \times k$ with linearly independent columns, then there exists a factorization $X = QR$ where

- R is $k \times k$, upper triangular and nonsingular
- Q is $n \times k$, with orthonormal columns

Proof sketch: Let

- $x_j := \text{col}_j(X)$

- $\{u_1, \dots, u_k\}$ be orthonormal with same span as $\{x_1, \dots, x_k\}$ (to be constructed using Gram-Schmidt)
- Q be formed from cols u_i

Since $x_j \in \text{span}\{u_1, \dots, u_j\}$, we have

$$x_j = \sum_{i=1}^j \langle u_i, x_j \rangle u_i \quad \text{for } j = 1, \dots, k$$

Some rearranging gives $X = QR$

Linear Regression via QR Decomposition For X and y as above we have $\hat{\beta} = (X'X)^{-1}X'y$

Using the QR decomposition $X = QR$ gives

$$\begin{aligned} \hat{\beta} &= (R'Q'QR)^{-1}R'Q'y \\ &= (R'R)^{-1}R'Q'y \\ &= R^{-1}(R')^{-1}R'Q'y = R^{-1}Q'y \end{aligned}$$

Numerical routines would in this case use the alternative form $R\hat{\beta} = Q'y$ and back substitution

Exercises

Exercise 1 Show that, for any linear subspace $S \subset \mathbb{R}^n$, we have $S \cap S^\perp = \{0\}$

Exercise 2 Let $P = X(X'X)^{-1}X'$ and let $M = I - P$. Show that P and M are both idempotent and symmetric. Can you give any intuition as to why they should be idempotent?

Solutions

[Solution notebook](#)

Shortest Paths

Contents

- *Shortest Paths*
 - [Overview](#)
 - [Outline of the Problem](#)
 - [Finding Least-Cost Paths](#)
 - [Solving for \$J\$](#)
 - [Exercises](#)
 - [Solutions](#)

Overview

The shortest path problem is a classic problem in mathematics and computer science with applications in

- Economics (sequential decision making, analysis of social networks, etc.)
- Operations research and transportation
- Robotics and artificial intelligence
- Telecommunication network design and routing
- etc., etc.

Variations of the methods we discuss in this lecture are used millions of times every day, in applications such as

- Google Maps
- routing packets on the internet

For us, the shortest path problem also provides a nice introduction to the logic of **dynamic programming**

Dynamic programming is an extremely powerful optimization technique that we apply in many lectures on this site

Outline of the Problem

The shortest path problem is one of finding how to traverse a graph from one specified node to another at minimum cost

Consider the following graph

We wish to travel from node (vertex) A to node G at minimum cost

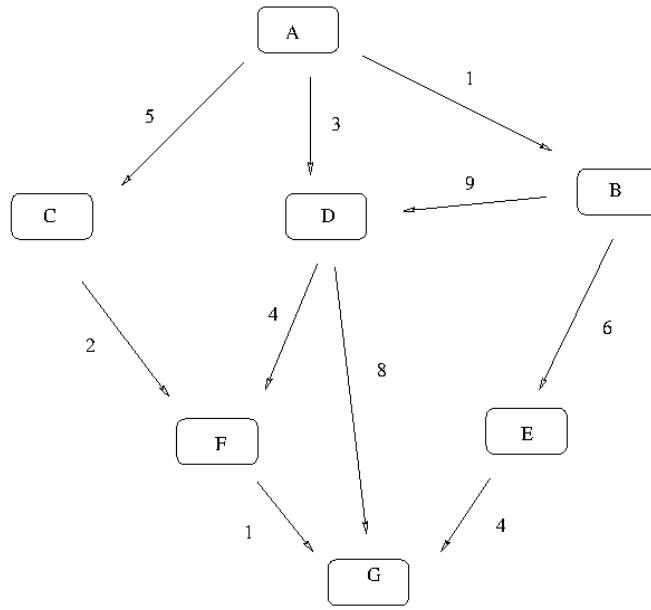
- Arrows (edges) indicate the movements we can take
- Numbers next to edges indicate the cost of traveling that edge

Possible interpretations of the graph include

- Minimum cost for supplier to reach a destination
- Routing of packets on the internet (minimize time)
- Etc., etc.

For this simple graph, a quick scan of the edges shows that the optimal paths are

- A, C, F, G at cost 8
- A, D, F, G at cost 8



Finding Least-Cost Paths

For large graphs we need a systematic solution

Let $J(v)$ denote the minimum cost-to-go from node v , understood as the total cost from v if we take the best route

Suppose that we know $J(v)$ for each node v , as shown below for the graph from the preceding example

Note that $J(G) = 0$

Intuitively, the best path can now be found as follows

- Start at A
- From node v , move to any node that solves

$$\min_{w \in F_v} \{c(v, w) + J(w)\} \quad (2.17)$$

where

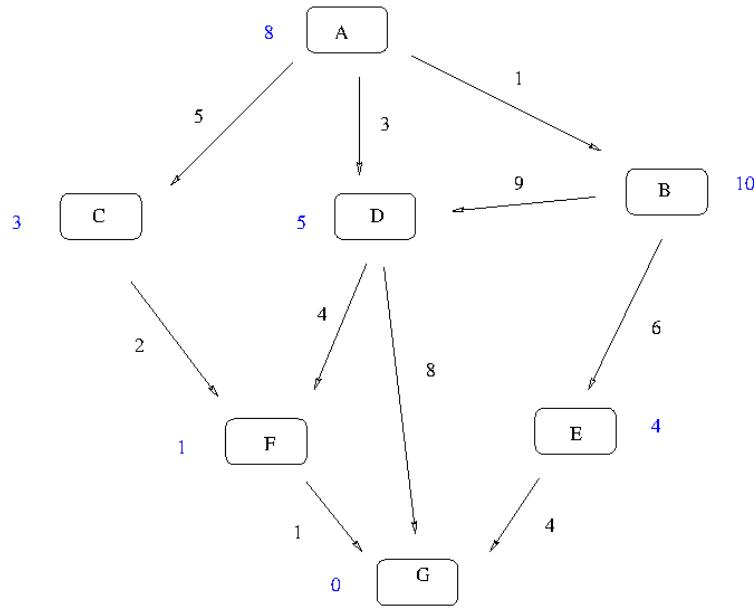
- F_v is the set of nodes that can be reached from v in one step
- $c(v, w)$ is the cost of traveling from v to w

Hence, if we know the function J , then finding the best path is almost trivial

But how to find J ?

Some thought will convince you that, for every node v , the function J satisfies

$$J(v) = \min_{w \in F_v} \{c(v, w) + J(w)\} \quad (2.18)$$



This is known as the Bellman equation

- That is, J is the solution to the Bellman equation
- There are algorithms for computing the minimum cost-to-go function J

Solving for J

The standard algorithm for finding J is to start with

$$J_0(v) = M \text{ if } v \neq \text{destination, else } J_0(v) = 0 \quad (2.19)$$

where M is some large number

Now we use the following algorithm

1. Set $n = 0$
2. Set $J_{n+1}(v) = \min_{w \in F_v} \{c(v, w) + J_n(w)\}$ for all v
3. If J_{n+1} and J_n are not equal then increment n , go to 2

In general, this sequence converges to J —the proof is omitted

Exercises

Exercise 1 Use the algorithm given above to find the optimal path (and its cost) for this graph

Here the line `node0, node1 0.04, node8 11.11, node14 72.21` means that from node0 we can go to

- node1 at cost 0.04
- node8 at cost 11.11
- node14 at cost 72.21

and so on

According to our calculations, the optimal path and its cost are like this

Your code should replicate this result

Solutions

[Solution notebook](#)

The McCall Job Search Model

Contents

- *The McCall Job Search Model*
 - [Overview](#)
 - [The Model](#)
 - [Solving the Model using Dynamic Programming](#)
 - [Implementation](#)
 - [The Reservation Wage](#)
 - [Exercises](#)
 - [Solutions](#)

Overview

The McCall search model [\[McC70\]](#) helped transform economists' way of thinking about labor markets

It did this by casting

- the loss of a job as a capital loss, and
- a spell of unemployment as an *investment* in searching for an acceptable job

To solve the model, we follow McCall in using dynamic programming

Dynamic programming was discussed previously in the [lecture on shortest paths](#)

The McCall model is a nice vehicle for readers to start to make themselves more comfortable with this approach to optimization

(More extensive and more formal treatments of dynamic programming are given in later lectures)

The Model

The model concerns the life of an infinitely lived worker and

- the opportunities he or she (let's say he to save one symbol) has to work at different wages
- exogenous events that destroy his current job
- his decision making process while unemployed

It is assumed that the worker lives forever

He can be in one of two states: employed or unemployed

He wants to maximize

$$\mathbb{E} \sum_{t=0}^{\infty} \beta^t u(y_t) \quad (2.20)$$

which represents the expected value of the discounted utility of his income

The constant β lies in $(0, 1)$ and is called a **discount factor**

The smaller is β , the more the worker discounts future utility relative to current utility

The variable y_t is

- his wage w_t when employed
- unemployment compensation c when unemployed

The function u is a utility function satisfying $u' > 0$ and $u'' < 0$

Timing and Decisions Let's consider what happens at the start of a given period (e.g., a month, if the timing of the model is monthly)

If currently *employed*, the worker consumes his wage w , receiving utility $u(w)$

If currently *unemployed*, he

- receives and consumes unemployment compensation c
- receives an offer to start work *next period* at a wage w' drawn from a known distribution p

He can either accept or reject the offer

If he accepts the offer, he enters next period employed with wage w'

If he rejects the offer, he enters next period unemployed

(Note that we do not allow for job search while employed—this topic is taken up in a [later lecture](#))

Job Termination When employed, he faces a constant probability α of becoming unemployed at the end of the period

Solving the Model using Dynamic Programming

As promised, we shall solve the McCall search model using dynamic programming

Dynamic programming is an ingenious method for solving a problem that starts by

1. assuming that you know the answer,
2. writing down some natural conditions that the answer must satisfy, then
3. solving those conditions to find the answer

So here goes

Let

- $V(w)$ be the total lifetime *value* accruing to a worker who enters the current period employed with wage w
- U be the total lifetime value accruing to a worker who is unemployed this period

Here *value* means the value of the objective function (2.20) when the worker makes optimal decisions now and at all future points in time

Suppose for now that the worker can calculate the function V and the constant U and use them in his decision making

In this case, a little thought will convince you that V and U should satisfy

$$V(w) = u(w) + \beta[(1 - \alpha)V(w) + \alpha U] \quad (2.21)$$

and

$$U = u(c) + \beta \sum_{w'} \max \{U, V(w')\} p(w') \quad (2.22)$$

The sum is over all possible wage values, which we assume for convenience is finite

Let's interpret these two equations in light of the fact that today's tomorrow is tomorrow's today

- The left hand sides of equations (2.21) and (2.22) are the values of a worker in a particular situation *today*
- The right hand sides of the equations are the discounted (by β) expected values of the possible situations that worker can be in *tomorrow*
- But *tomorrow* the worker can be in only one of the situations whose values *today* are on the left sides of our two equations

Equation (2.22) incorporates the fact that a currently unemployed worker will maximize his own welfare

In particular, if his next period wage offer is w' , he will choose to remain unemployed unless $U < V(w')$

Equations (2.21) and (2.22) are called *Bellman equations* after the mathematician Richard Bellman

It turns out that equations (2.21) and (2.22) provide enough information to solve out for both V and U

Before discussing this, however, let's make a small extension to the model

Stochastic Offers Let's suppose now that unemployed workers don't always receive job offers

Instead, let's suppose that unemployed workers only receive an offer with probability γ

If our worker does receive an offer, the wage offer is drawn from p as before

He either accepts or rejects the offer

Otherwise the model is the same

With some thought, you will be able to convince yourself that V and U should now satisfy

$$V(w) = u(w) + \beta[(1 - \alpha)V(w) + \alpha U] \quad (2.23)$$

and

$$U = u(c) + \beta(1 - \gamma)U + \beta\gamma \sum_{w'} \max\{U, V(w')\} p(w') \quad (2.24)$$

Solving the Bellman Equations The Bellman equations are nonlinear in U and V , and hence not trivial to solve

One way to solve them is to

1. make guesses for U and V
2. plug these guesses into the right hand sides of (2.23) and (2.24)
3. update the left hand sides from this rule and then repeat

In other words, we are iterating using the rules

$$V_{n+1}(w) = u(w) + \beta[(1 - \alpha)V_n(w) + \alpha U_n] \quad (2.25)$$

and

$$U_{n+1} = u(c) + \beta(1 - \gamma)U_n + \beta\gamma \sum_{w'} \max\{U_n, V_n(w')\} p(w') \quad (2.26)$$

starting from some initial conditions U_0, V_0

This procedure is called *iterating on the Bellman equations*

It turns out that these iterations are guaranteed to converge to the V and U that solve (2.23) and (2.24)

We discuss the theory behind this property extensively in later lectures (see, e.g., the discussion in [this lecture](#))

For now let's try implementing the iteration scheme to see what the solutions look like

Implementation

Code to iterate on the Bellman equations can be found in [mccall_bellman_iteration.jl](#) from the [applications repository](#)

We repeat it here for convenience

In the code you'll see that we use a type to store the various parameters and other objects associated with a given model

This helps to tidy up the code and provides an object that's easy to pass to functions

The default utility function is a CRRA utility function

```
#=
Implements iteration on the Bellman equations to solve the McCall growth model

=#
using Distributions

# A default utility function

function u(c, sigma)
    if c > 0
        return (c^(1 - sigma) - 1) / (1 - sigma)
    else
        return -10e6
    end
end

# default wage vector with probabilities

const n = 60                                # n possible outcomes for wage
const default_w_vec = linspace(10, 20, n)      # wages between 10 and 20
const a, b = 600, 400                          # shape parameters
const dist = BetaBinomial(n-1, a, b)
const default_p_vec = pdf(dist)

type McCallModel
    alpha::Float64          # Job separation rate
    beta::Float64            # Discount rate
    gamma::Float64           # Job offer rate
    c::Float64                # Unemployment compensation
    sigma::Float64            # Utility parameter
    w_vec::Vector{Float64} # Possible wage values
    p_vec::Vector{Float64} # Probabilities over w_vec

    function McCallModel(alpha=0.2,
                        beta=0.98,
                        gamma=0.7,
                        c=6.0,
                        sigma=2.0,
                        w_vec=default_w_vec,
                        p_vec=default_p_vec)

        return new(alpha, beta, gamma, c, sigma, w_vec, p_vec)
    end
end
```

```

"""
A function to update the Bellman equations. Note that V_new is modified in
place (i.e., modified by this function). The new value of U is returned.

"""

function update_bellman!(mcm, V, V_new, U)
    # Simplify notation
    alpha, beta, sigma, c, gamma = mcm.alpha, mcm.beta, mcm.sigma, mcm.c, mcm.gamma

    for (w_idx, w) in enumerate(mcm.w_vec)
        # w_idx indexes the vector of possible wages
        V_new[w_idx] = u(w, sigma) + beta * ((1 - alpha) * V[w_idx] + alpha * U)
    end

    U_new = u(c, sigma) + beta * (1 - gamma) * U +
            beta * gamma * sum(max(U, V) .* mcm.p_vec)

    return U_new
end

function solve_mccall_model(mcm; tol::Float64=1e-5, max_iter::Int=2000)

    V = ones(length(mcm.w_vec))      # Initial guess of V
    V_new = similar(V)              # To store updates to V
    U = 1.0                         # Initial guess of U
    i = 0
    error = tol + 1

    while error > tol && i < max_iter
        U_new = update_bellman!(mcm, V, V_new, U)
        error_1 = maximum(abs(V_new - V))
        error_2 = abs(U_new - U)
        error = max(error_1, error_2)
        V[:] = V_new
        U = U_new
        i += 1
    end

    return V, U
end

```

The approach is to iterate until successive iterates are closer together than some small tolerance level

We then return the current iterate as an approximate solution

Let's plot the approximate solutions U and V to see what they look like

We'll use the default parameterizations found in the code above

```
#=
Generate plots of value of employment and unemployment in the McCall model.

#=

using Plots, LaTeXStrings
pyplot()

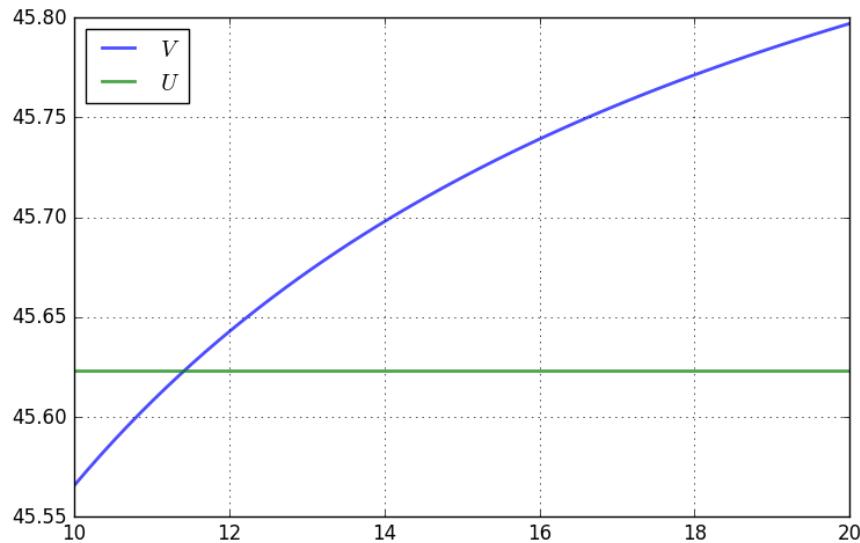
include("mccall_bellman_iteration.jl")
include("compute_reservation_wage.jl")

mcm = McCallModel()
V, U = solve_mccall_model(mcm)
U_vec = U .* ones(length(mcm.w_vec))

plot(mcm.w_vec,
      [V U_vec],
      lw=2,
      alpha=0.7,
      label=[L"\$V\$" L"\$U\$"])

```

Here's the plot this code produces



The value V is increasing because higher w generates a higher wage flow conditional on staying employed

The Reservation Wage

Once V and U are known, the agent can use them to make decisions in the face of a given wage offer

If $V(w) > U$, then working at wage w is preferred to unemployment

If $V(w) < U$, then remaining unemployed will generate greater lifetime value

Suppose in particular that V crosses U (as it does in the preceding figure)

Then, since V is increasing, there is a unique smallest w in the set of possible wages such that $V(w) \geq U$

We denote this wage \bar{w} and call it the **reservation wage**

Optimal behavior for the worker is characterized by \bar{w}

- if the wage offer w in hand is greater than or equal to \bar{w} , then the worker accepts
- if the wage offer w in hand is less than \bar{w} , then the worker rejects

We've written a function called `compute_reservation_wage` that takes an instance of a McCall model and returns the reservation wage associated with a given model

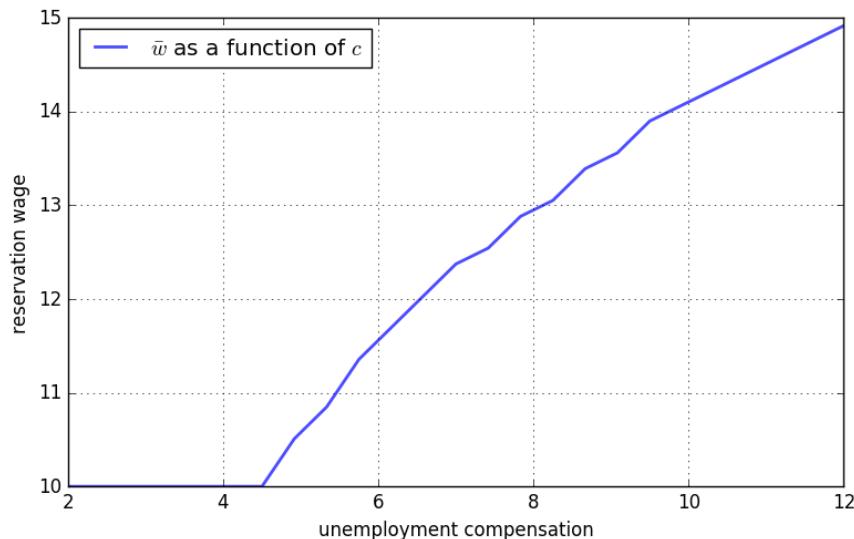
If $V(w) < U$ for all w , then the function returns `np.inf`

Below you'll be asked to try to produce a version of this function as an exercise

For now let's use it to look at how the reservation wage varies with parameters

The Reservation Wage and Unemployment Compensation First, let's look at how \bar{w} varies with unemployment compensation

In the figure below, we use the default parameters in the `McCallModel` type, apart from c (which takes the values given on the horizontal axis)



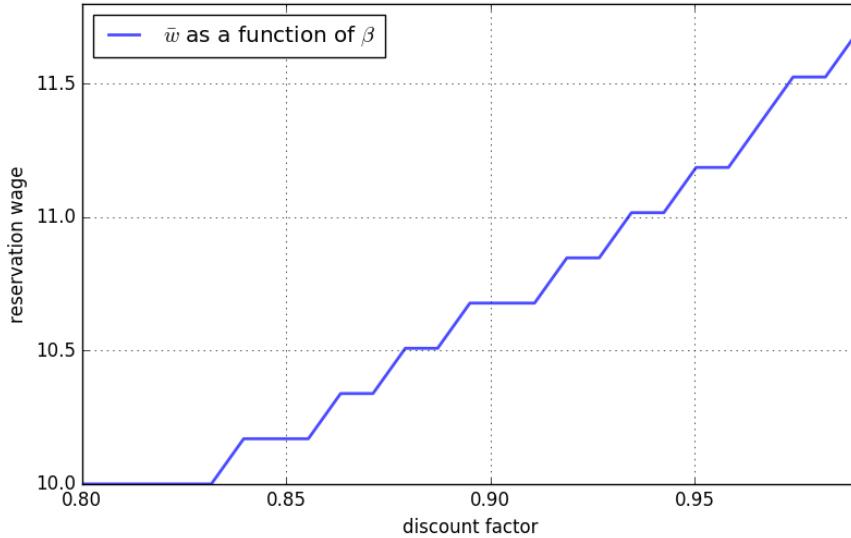
As expected, higher unemployment compensation causes the worker to hold out for higher wages

In effect, the cost of continuing job search is reduced

(Code to reproduce the figure can be found in [this directory](#))

The Reservation Wage and Discounting Next let's investigate how \bar{w} varies with the discount rate

The next figure plots the reservation wage associated with different values of β



Again, the results are intuitive: More patient workers will hold out for higher wages

(Again, code to reproduce the figure can be found in [this directory](#))

The Reservation Wage and Job Destruction Finally, let's look at how \bar{w} varies with the job separation rate α

Higher α translates to a greater chance that a worker will face termination in each period once employed

Once more, the results are in line with our intuition

If the separation rate is high, then the benefit of holding out for a higher wage falls

Hence the reservation wage is lower

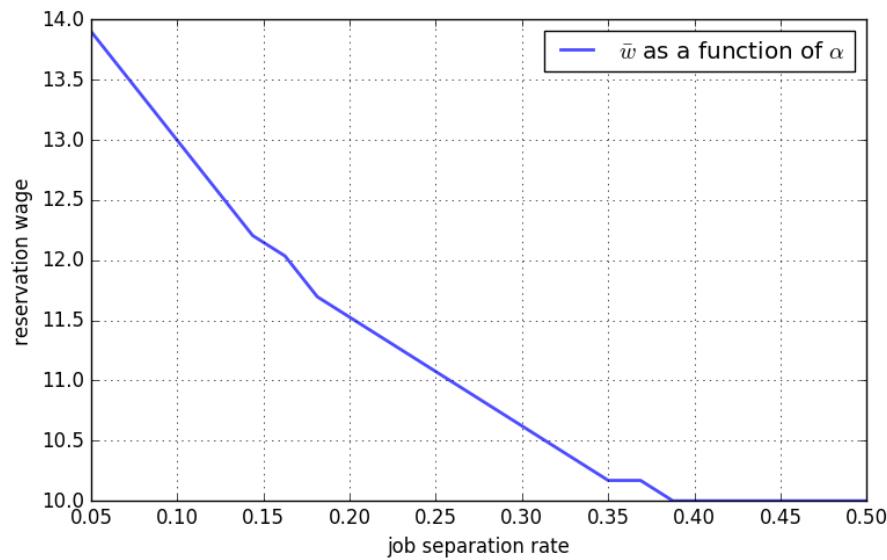
Exercises

Exercise 1 In the preceding discussion we computed the reservation wage for various instances of the McCall model

Try implementing your own function that accomplishes this task

Its input should be an instance of *McCallModel* as defined in `mccall_bellman_iteration.jl` and its output should be the corresponding reservation wage

In doing so, you can make use of



- the logic for computing the reservation wage discussed above
- the code for computing value functions in `mccall_bellman_iteration.jl`

Exercise 2 Use your function from Exercise 1 to plot \bar{w} against the job offer rate γ

Interpret your results

Solutions

Solution notebook

Schelling's Segregation Model

Contents

- *Schelling's Segregation Model*
 - *Outline*
 - *The Model*
 - *Results*
 - *Exercises*
 - *Solutions*

Outline

In 1969, Thomas C. Schelling developed a simple but striking model of racial segregation [Sch69]

His model studies the dynamics of racially mixed neighborhoods

Like much of Schelling's work, the model shows how local interactions can lead to surprising aggregate structure

In particular, it shows that relatively mild preference for neighbors of similar race can lead in aggregate to the collapse of mixed neighborhoods, and high levels of segregation

In recognition of this and other research, Schelling was awarded the 2005 Nobel Prize in Economic Sciences (joint with Robert Aumann)

In this lecture we (in fact you) will build and run a version of Schelling's model

The Model

We will cover a variation of Schelling's model that is easy to program and captures the main idea

Set Up Suppose we have two types of people: orange people and green people

For the purpose of this lecture, we will assume there are 250 of each type

These agents all live on a single unit square

The location of an agent is just a point (x, y) , where $0 < x, y < 1$

Preferences We will say that an agent is *happy* if half or more of her 10 nearest neighbors are of the same type

Here 'nearest' is in terms of Euclidean distance

An agent who is not happy is called *unhappy*

An important point here is that agents are not averse to living in mixed areas

They are perfectly happy if half their neighbors are of the other color

Behavior Initially, agents are mixed together (integrated)

In particular, the initial location of each agent is an independent draw from a bivariate uniform distribution on $S = (0, 1)^2$

Now, cycling through the set of all agents, each agent is now given the chance to stay or move

We assume that each agent will stay put if they are happy and move if unhappy

The algorithm for moving is as follows

1. Draw a random location in S
2. If happy at new location, move there
3. Else, go to step 1

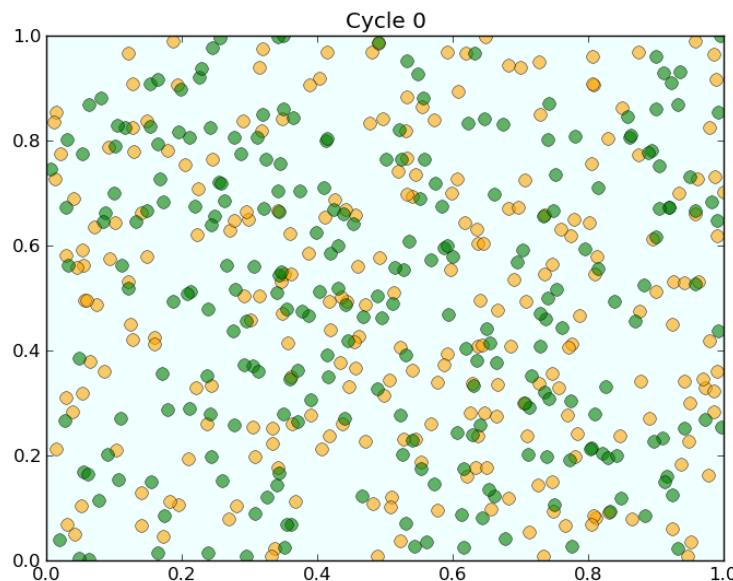
In this way, we cycle continuously through the agents, moving as required

We continue to cycle until no one wishes to move

Results

Let's have a look at the results we got when we coded and ran this model

As discussed above, agents are initially mixed randomly together



But after several cycles they become segregated into distinct regions

In this instance, the program terminated after 4 cycles through the set of agents, indicating that all agents had reached a state of happiness

What is striking about the pictures is how rapidly racial integration breaks down

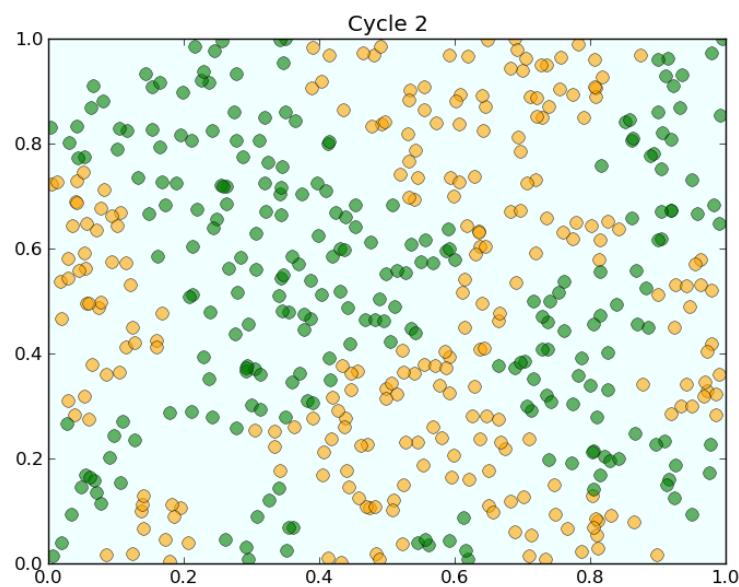
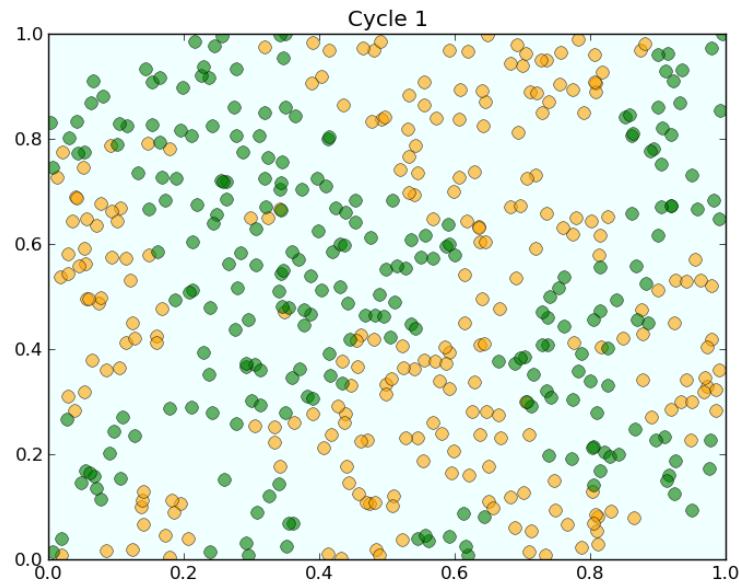
This is despite the fact that people in the model don't actually mind living mixed with the other type

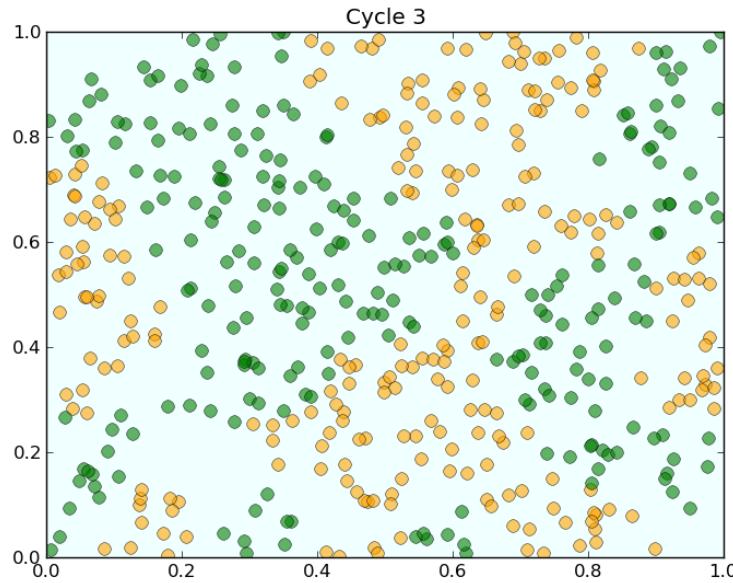
Even with these preferences, the outcome is a high degree of segregation

Exercises

Rather than show you the program that generated these figures, we'll now ask you to write your own version

You can see our program at the end, when you look at the solution





Exercise 1 Implement and run this simulation for yourself

Consider the following structure for your program

Agents are modeled as objects

(Have a look at *this lecture* if you've forgotten how to build your own objects)

Here's an indication of how they might look

```
* Data:
    * type (green or orange)
    * location

* Methods:
    * Determine whether happy or not given locations of other agents
    * If not happy, move
        * find a new location where happy
```

And here's some pseudocode for the main loop

```
while agents are still moving
    for agent in agents
        give agent the opportunity to move
    end
end
```

Use 250 agents of each type

Solutions

[Solution notebook](#)

LLN and CLT

Contents

- *LLN and CLT*
 - *Overview*
 - *Relationships*
 - *LLN*
 - *CLT*
 - *Exercises*
 - *Solutions*

Overview

This lecture illustrates two of the most important theorems of probability and statistics: The law of large numbers (LLN) and the central limit theorem (CLT)

These beautiful theorems lie behind many of the most fundamental results in econometrics and quantitative economic modeling

The lecture is based around simulations that show the LLN and CLT in action

We also demonstrate how the LLN and CLT break down when the assumptions they are based on do not hold

In addition, we examine several useful extensions of the classical theorems, such as

- The delta method, for smooth functions of random variables
- The multivariate case

Some of these extensions are presented as exercises

Relationships

The CLT refines the LLN

The LLN gives conditions under which sample moments converge to population moments as sample size increases

The CLT provides information about the rate at which sample moments converge to population moments as sample size increases

LLN

We begin with the law of large numbers, which tells us when sample averages will converge to their population means

The Classical LLN The classical law of large numbers concerns independent and identically distributed (IID) random variables

Here is the strongest version of the classical LLN, known as *Kolmogorov's strong law*

Let X_1, \dots, X_n be independent and identically distributed scalar random variables, with common distribution F

When it exists, let μ denote the common mean of this sample:

$$\mu := \mathbb{E}X = \int xF(dx)$$

In addition, let

$$\bar{X}_n := \frac{1}{n} \sum_{i=1}^n X_i$$

Kolmogorov's strong law states that, if $\mathbb{E}|X|$ is finite, then

$$\mathbb{P}\{\bar{X}_n \rightarrow \mu \text{ as } n \rightarrow \infty\} = 1 \quad (2.27)$$

What does this last expression mean?

Let's think about it from a simulation perspective, imagining for a moment that our computer can generate perfect random samples (which of course **it can't**)

Let's also imagine that we can generate infinite sequences, so that the statement $\bar{X}_n \rightarrow \mu$ can be evaluated

In this setting, (2.27) should be interpreted as meaning that the probability of the computer producing a sequence where $\bar{X}_n \rightarrow \mu$ fails to occur is zero

Proof The proof of Kolmogorov's strong law is nontrivial – see, for example, theorem 8.3.5 of [Dud02]

On the other hand, we can prove a weaker version of the LLN very easily and still get most of the intuition

The version we prove is as follows: If X_1, \dots, X_n is IID with $\mathbb{E}X_i^2 < \infty$, then, for any $\epsilon > 0$, we have

$$\mathbb{P}\{|\bar{X}_n - \mu| \geq \epsilon\} \rightarrow 0 \quad \text{as } n \rightarrow \infty \quad (2.28)$$

(This version is weaker because we claim only **convergence in probability** rather than **almost sure convergence**, and assume a finite second moment)

To see that this is so, fix $\epsilon > 0$, and let σ^2 be the variance of each X_i

Recall the [Chebyshev inequality](#), which tells us that

$$\mathbb{P}\{|\bar{X}_n - \mu| \geq \epsilon\} \leq \frac{\mathbb{E}[(\bar{X}_n - \mu)^2]}{\epsilon^2} \quad (2.29)$$

Now observe that

$$\begin{aligned} \mathbb{E}[(\bar{X}_n - \mu)^2] &= \mathbb{E}\left\{\left[\frac{1}{n} \sum_{i=1}^n (X_i - \mu)\right]^2\right\} \\ &= \frac{1}{n^2} \sum_{i=1}^n \sum_{j=1}^n \mathbb{E}(X_i - \mu)(X_j - \mu) \\ &= \frac{1}{n^2} \sum_{i=1}^n \mathbb{E}(X_i - \mu)^2 \\ &= \frac{\sigma^2}{n} \end{aligned}$$

Here the crucial step is at the third equality, which follows from independence

Independence means that if $i \neq j$, then the covariance term $\mathbb{E}(X_i - \mu)(X_j - \mu)$ drops out

As a result, $n^2 - n$ terms vanish, leading us to a final expression that goes to zero in n

Combining our last result with (2.29), we come to the estimate

$$\mathbb{P}\{|\bar{X}_n - \mu| \geq \epsilon\} \leq \frac{\sigma^2}{n\epsilon^2} \quad (2.30)$$

The claim in (2.28) is now clear

Of course, if the sequence X_1, \dots, X_n is correlated, then the cross-product terms $\mathbb{E}(X_i - \mu)(X_j - \mu)$ are not necessarily zero

While this doesn't mean that the same line of argument is impossible, it does mean that if we want a similar result then the covariances should be "almost zero" for "most" of these terms

In a long sequence, this would be true if, for example, $\mathbb{E}(X_i - \mu)(X_j - \mu)$ approached zero when the difference between i and j became large

In other words, the LLN can still work if the sequence X_1, \dots, X_n has a kind of "asymptotic independence", in the sense that correlation falls to zero as variables become further apart in the sequence

This idea is very important in time series analysis, and we'll come across it again soon enough

Illustration Let's now illustrate the classical IID law of large numbers using simulation

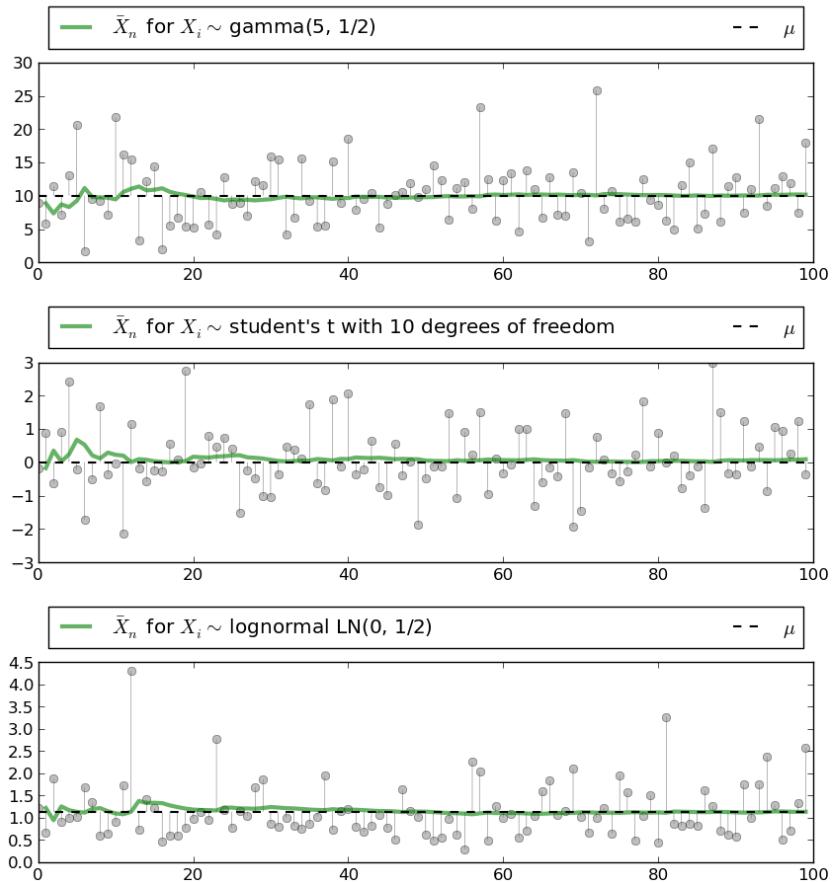
In particular, we aim to generate some sequences of IID random variables and plot the evolution of \bar{X}_n as n increases

Below is a figure that does just this (as usual, you can click on it to expand it)

It shows IID observations from three different distributions and plots \bar{X}_n against n in each case

The dots represent the underlying observations X_i for $i = 1, \dots, 100$

In each of the three cases, convergence of \bar{X}_n to μ occurs as predicted



The figure was produced by `illuminates_lln.jl`, which is shown below (and can be found in the `lln_clt` directory of the [applications repository](#))

The three distributions are chosen at random from a selection stored in the dictionary `distributions`

```
#=
Visual illustration of the law of large numbers.

@author : Spencer Lyon <spencer.lyon@nyu.edu>
          Victoria Gregory <victoria.gregory@nyu.edu>

References
-----
Based off the original python file illuminates_lln.py
=#

```

```

using Plots
pyplot()
using Distributions
using LaTeXStrings

n = 100
srand(42) # reproducible results

# == Arbitrary collection of distributions == #
distributions = Dict("student's t with 10 degrees of freedom" => TDist(10),
                      "beta(2, 2)" => Beta(2.0, 2.0),
                      "lognormal LN(0, 1/2)" => LogNormal(0.5),
                      "gamma(5, 1/2)" => Gamma(5.0, 2.0),
                      "poisson(4)" => Poisson(4),
                      "exponential with lambda = 1" => Exponential(1))

num_plots = 3
dist_data = zeros(num_plots, n)
sample_means = []
dist_means = []
titles = []
for i = 1:num_plots
    dist_names = collect(keys(distributions))
    # == Choose a randomly selected distribution == #
    name = dist_names[rand(1:length(dist_names))]
    dist = pop!(distributions, name)

    # == Generate n draws from the distribution == #
    data = rand(dist, n)

    # == Compute sample mean at each n == #
    sample_mean = Array(Float64, n)
    for j=1:n
        sample_mean[j] = mean(data[1:j])
    end

    m = mean(dist)

    dist_data[i, :] = data'
    push!(sample_means, sample_mean)
    push!(dist_means, m*ones(n))
    push!(titles, name)
end

# == Plot == #
N = repmat(reshape(repmat(linspace(1, n, n), 1, num_plots)', 1, n*num_plots), 2, 1)
heights = [zeros(1,n*num_plots); reshape(dist_data, 1, n*num_plots)]
plot(N, heights, layout=(3, 1), label="", color=:grey, alpha=0.5)
plot!(1:n, dist_data', layout=(3, 1), color=:grey, markershape=:circle,
      alpha=0.5, label="", linewidth=0)
plot!(1:n, sample_means, linewidth=3, alpha=0.6, color=:green, legend=:topleft,
      layout=(3, 1), label=[LaTeXString("\$\\bar{X}_n\$") " " ""])

```

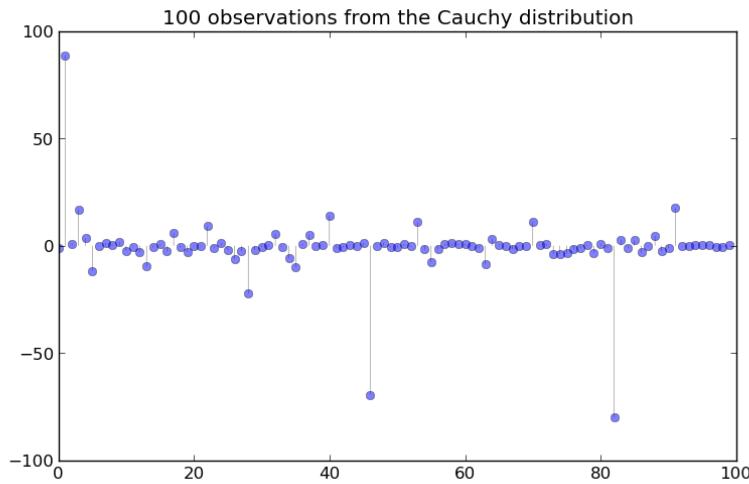
```
plot!(1:n, dist_means, color=:black, linewidth=1.5, layout=(3, 1),
      linestyle=:dash, grid=false, label=[LaTeXString("\$\mu\$") " " ""])
plot!(title="titles")
```

Infinite Mean What happens if the condition $\mathbb{E}|X| < \infty$ in the statement of the LLN is not satisfied?

This might be the case if the underlying distribution is heavy tailed — the best known example is the Cauchy distribution, which has density

$$f(x) = \frac{1}{\pi(1+x^2)} \quad (x \in \mathbb{R})$$

The next figure shows 100 independent draws from this distribution



Notice how extreme observations are far more prevalent here than the previous figure

Let's now have a look at the behavior of the sample mean

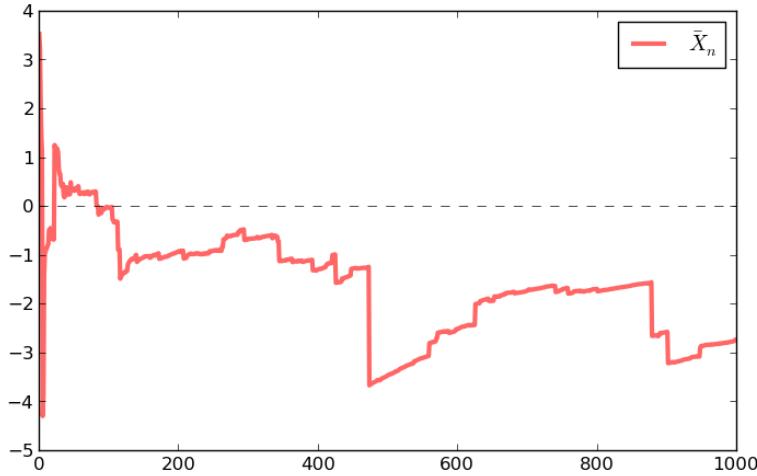
Here we've increased n to 1000, but the sequence still shows no sign of converging

Will convergence become visible if we take n even larger?

The answer is no

To see this, recall that the characteristic function of the Cauchy distribution is

$$\phi(t) = \mathbb{E}e^{itX} = \int e^{itx} f(x) dx = e^{-|t|} \tag{2.31}$$



Using independence, the characteristic function of the sample mean becomes

$$\begin{aligned}\mathbb{E}e^{it\bar{X}_n} &= \mathbb{E} \exp \left\{ i \frac{t}{n} \sum_{j=1}^n X_j \right\} \\ &= \mathbb{E} \prod_{j=1}^n \exp \left\{ i \frac{t}{n} X_j \right\} \\ &= \prod_{j=1}^n \mathbb{E} \exp \left\{ i \frac{t}{n} X_j \right\} = [\phi(t/n)]^n\end{aligned}$$

In view of (2.31), this is just $e^{-|t|}$

Thus, in the case of the Cauchy distribution, the sample mean itself has the very same Cauchy distribution, regardless of n

In particular, the sequence \bar{X}_n does not converge to a point

CLT

Next we turn to the central limit theorem, which tells us about the distribution of the deviation between sample averages and population means

Statement of the Theorem The central limit theorem is one of the most remarkable results in all of mathematics

In the classical IID setting, it tells us the following: If the sequence X_1, \dots, X_n is IID, with common mean μ and common variance $\sigma^2 \in (0, \infty)$, then

$$\sqrt{n}(\bar{X}_n - \mu) \xrightarrow{d} N(0, \sigma^2) \quad \text{as } n \rightarrow \infty \quad (2.32)$$

Here $\xrightarrow{d} N(0, \sigma^2)$ indicates convergence in distribution to a centered (i.e., zero mean) normal with standard deviation σ

Intuition The striking implication of the CLT is that for **any** distribution with finite second moment, the simple operation of adding independent copies **always** leads to a Gaussian curve

A relatively simple proof of the central limit theorem can be obtained by working with characteristic functions (see, e.g., theorem 9.5.6 of [Dud02])

The proof is elegant but almost anticlimactic, and it provides surprisingly little intuition

In fact all of the proofs of the CLT that we know are similar in this respect

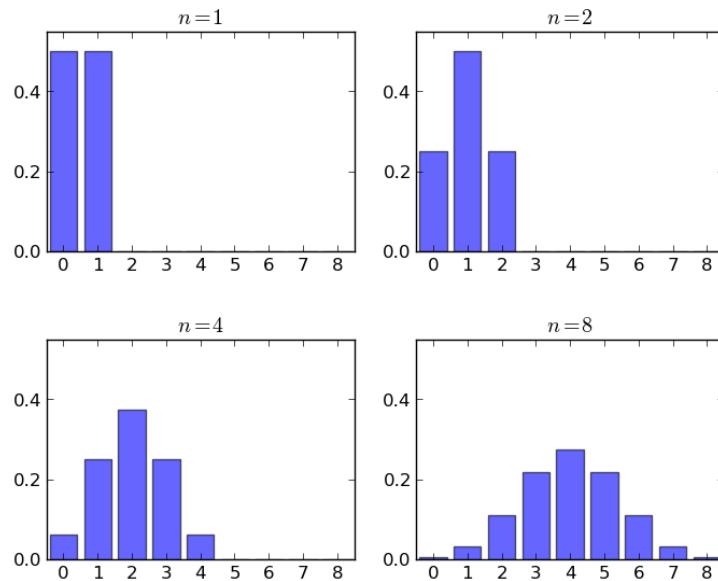
Why does adding independent copies produce a bell-shaped distribution?

Part of the answer can be obtained by investigating addition of independent Bernoulli random variables

In particular, let X_i be binary, with $\mathbb{P}\{X_i = 0\} = \mathbb{P}\{X_i = 1\} = 0.5$, and let X_1, \dots, X_n be independent

Think of $X_i = 1$ as a “success”, so that $Y_n = \sum_{i=1}^n X_i$ is the number of successes in n trials

The next figure plots the probability mass function of Y_n for $n = 1, 2, 4, 8$



When $n = 1$, the distribution is flat — one success or no successes have the same probability

When $n = 2$ we can either have 0, 1 or 2 successes

Notice the peak in probability mass at the mid-point $k = 1$

The reason is that there are more ways to get 1 success (“fail then succeed” or “succeed then fail”) than to get zero or two successes

Moreover, the two trials are independent, so the outcomes “fail then succeed” and “succeed then fail” are just as likely as the outcomes “fail then fail” and “succeed then succeed”

(If there was positive correlation, say, then “succeed then fail” would be less likely than “succeed then succeed”)

Here, already we have the essence of the CLT: addition under independence leads probability mass to pile up in the middle and thin out at the tails

For $n = 4$ and $n = 8$ we again get a peak at the “middle” value (halfway between the minimum and the maximum possible value)

The intuition is the same — there are simply more ways to get these middle outcomes

If we continue, the bell-shaped curve becomes ever more pronounced

We are witnessing the binomial approximation of the normal distribution

Simulation 1 Since the CLT seems almost magical, running simulations that verify its implications is one good way to build intuition

To this end, we now perform the following simulation

1. Choose an arbitrary distribution F for the underlying observations X_i
2. Generate independent draws of $Y_n := \sqrt{n}(\bar{X}_n - \mu)$
3. Use these draws to compute some measure of their distribution — such as a histogram
4. Compare the latter to $N(0, \sigma^2)$

Here's some code that does exactly this for the exponential distribution $F(x) = 1 - e^{-\lambda x}$

(Please experiment with other choices of F , but remember that, to conform with the conditions of the CLT, the distribution must have finite second moment)

```
#=
Visual illustration of the central limit theorem

@author : Spencer Lyon <spencer.lyon@nyu.edu>
          Victoria Gregory <victoria.gregory@nyu.edu>

References
-----

Based off the original python file illustrates_clt.py
=#
using Plots
pyplot()
using Distributions
using LaTeXStrings

# == Set parameters == #
srand(42) # reproducible results
n = 250    # Choice of n
k = 100000 # Number of draws of Y_n
dist = Exponential(1./2.) # Exponential distribution, lambda = 1/2
mu, s = mean(dist), std(dist)
```

```
# == Draw underlying RVs. Each row contains a draw of  $X_1, \dots, X_n$  ==
data = rand(dist, (k, n))

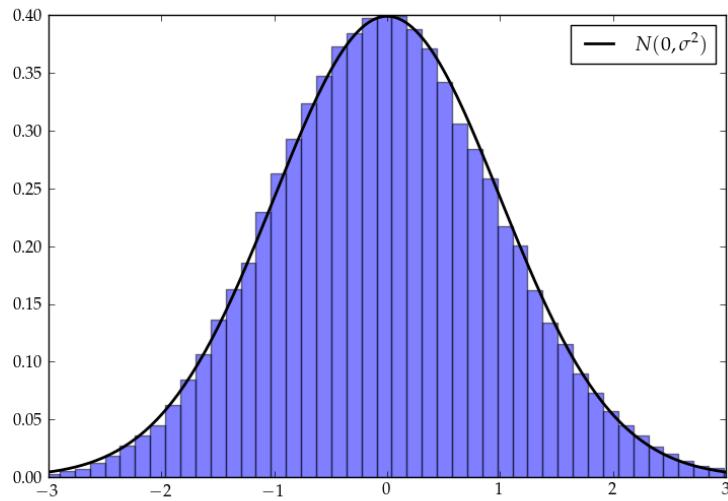
# == Compute mean of each row, producing k draws of  $\bar{X}_n$  ==
sample_means = mean(data, 2)

# == Generate observations of  $Y_n$  ==
Y = sqrt(n) * (sample_means .- mu)

# == Plot ==
xmin, xmax = -3 * s, 3 * s
histogram(Y, nbins=60, alpha=0.5, xlims=(xmin, xmax),
           norm=true, label="")
xgrid = linspace(xmin, xmax, 200)
plot!(xgrid, pdf(Normal(0.0, s), xgrid), color=:black,
      linewidth=2, label=LaTeXString("\$N(0, \sigma^2)\$"),
      legendfont=font(12))
```

The file is `illustiates_clt.jl`, from the [QuantEcon.applications](#) repo

The program produces figures such as the one below



The fit to the normal density is already tight, and can be further improved by increasing n

You can also experiment with other specifications of F

Simulation 2 Our next simulation is somewhat like the first, except that we aim to track the distribution of $Y_n := \sqrt{n}(\bar{X}_n - \mu)$ as n increases

In the simulation we'll be working with random variables having $\mu = 0$

Thus, when $n = 1$, we have $Y_1 = X_1$, so the first distribution is just the distribution of the underlying random variable

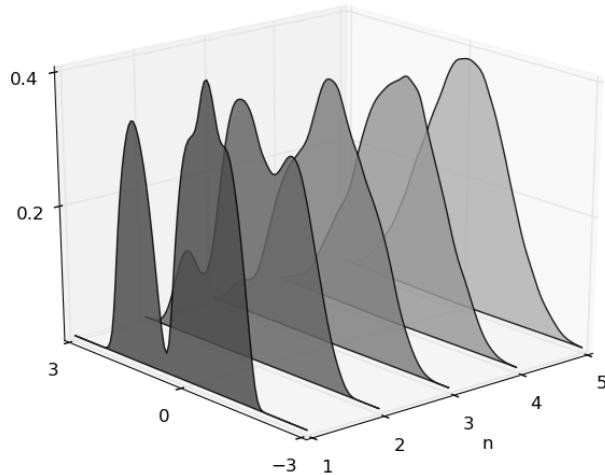
For $n = 2$, the distribution of Y_2 is that of $(X_1 + X_2)/\sqrt{2}$, and so on

What we expect is that, regardless of the distribution of the underlying random variable, the distribution of Y_n will smooth out into a bell shaped curve

The next figure shows this process for $X_i \sim f$, where f was specified as the convex combination of three different beta densities

(Taking a convex combination is an easy way to produce an irregular shape for f)

In the figure, the closest density is that of Y_1 , while the furthest is that of Y_5



As expected, the distribution smooths out into a bell curve as n increases

The figure is generated by file `lln_clt/clt3d.jl`, which is available from the [applications repository](#)

We leave you to investigate its contents if you wish to know more

If you run the file from the ordinary Julia or IJulia shell, the figure should pop up in a window that you can rotate with your mouse, giving different views on the density sequence

The Multivariate Case The law of large numbers and central limit theorem work just as nicely in multidimensional settings

To state the results, let's recall some elementary facts about random vectors

A random vector \mathbf{X} is just a sequence of k random variables (X_1, \dots, X_k)

Each realization of \mathbf{X} is an element of \mathbb{R}^k

A collection of random vectors $\mathbf{X}_1, \dots, \mathbf{X}_n$ is called independent if, given any n vectors $\mathbf{x}_1, \dots, \mathbf{x}_n$ in \mathbb{R}^k , we have

$$\mathbb{P}\{\mathbf{X}_1 \leq \mathbf{x}_1, \dots, \mathbf{X}_n \leq \mathbf{x}_n\} = \mathbb{P}\{\mathbf{X}_1 \leq \mathbf{x}_1\} \times \dots \times \mathbb{P}\{\mathbf{X}_n \leq \mathbf{x}_n\}$$

(The vector inequality $\mathbf{X} \leq \mathbf{x}$ means that $X_j \leq x_j$ for $j = 1, \dots, k$)

Let $\mu_j := \mathbb{E}[X_j]$ for all $j = 1, \dots, k$

The expectation $\mathbb{E}[\mathbf{X}]$ of \mathbf{X} is defined to be the vector of expectations:

$$\mathbb{E}[\mathbf{X}] := \begin{pmatrix} \mathbb{E}[X_1] \\ \mathbb{E}[X_2] \\ \vdots \\ \mathbb{E}[X_k] \end{pmatrix} = \begin{pmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_k \end{pmatrix} =: \boldsymbol{\mu}$$

The *variance-covariance matrix* of random vector \mathbf{X} is defined as

$$\text{Var}[\mathbf{X}] := \mathbb{E}[(\mathbf{X} - \boldsymbol{\mu})(\mathbf{X} - \boldsymbol{\mu})']$$

Expanding this out, we get

$$\text{Var}[\mathbf{X}] = \begin{pmatrix} \mathbb{E}[(X_1 - \mu_1)(X_1 - \mu_1)] & \cdots & \mathbb{E}[(X_1 - \mu_1)(X_k - \mu_k)] \\ \mathbb{E}[(X_2 - \mu_2)(X_1 - \mu_1)] & \cdots & \mathbb{E}[(X_2 - \mu_2)(X_k - \mu_k)] \\ \vdots & \vdots & \vdots \\ \mathbb{E}[(X_k - \mu_k)(X_1 - \mu_1)] & \cdots & \mathbb{E}[(X_k - \mu_k)(X_k - \mu_k)] \end{pmatrix}$$

The j, k -th term is the scalar covariance between X_j and X_k

With this notation we can proceed to the multivariate LLN and CLT

Let $\mathbf{X}_1, \dots, \mathbf{X}_n$ be a sequence of independent and identically distributed random vectors, each one taking values in \mathbb{R}^k

Let $\boldsymbol{\mu}$ be the vector $\mathbb{E}[\mathbf{X}_i]$, and let Σ be the variance-covariance matrix of \mathbf{X}_i

Interpreting vector addition and scalar multiplication in the usual way (i.e., pointwise), let

$$\bar{\mathbf{X}}_n := \frac{1}{n} \sum_{i=1}^n \mathbf{X}_i$$

In this setting, the LLN tells us that

$$\mathbb{P}\{\bar{\mathbf{X}}_n \rightarrow \boldsymbol{\mu} \text{ as } n \rightarrow \infty\} = 1 \quad (2.33)$$

Here $\bar{\mathbf{X}}_n \rightarrow \boldsymbol{\mu}$ means that $\|\bar{\mathbf{X}}_n - \boldsymbol{\mu}\| \rightarrow 0$, where $\|\cdot\|$ is the standard Euclidean norm

The CLT tells us that, provided Σ is finite,

$$\sqrt{n}(\bar{\mathbf{X}}_n - \boldsymbol{\mu}) \xrightarrow{d} N(\mathbf{0}, \Sigma) \quad \text{as } n \rightarrow \infty \quad (2.34)$$

Exercises

Exercise 1 One very useful consequence of the central limit theorem is as follows

Assume the conditions of the CLT as *stated above*

If $g: \mathbb{R} \rightarrow \mathbb{R}$ is differentiable at μ and $g'(\mu) \neq 0$, then

$$\sqrt{n}\{g(\bar{X}_n) - g(\mu)\} \xrightarrow{d} N(0, g'(\mu)^2\sigma^2) \quad \text{as } n \rightarrow \infty \quad (2.35)$$

This theorem is used frequently in statistics to obtain the asymptotic distribution of estimators — many of which can be expressed as functions of sample means

(These kinds of results are often said to use the “delta method”)

The proof is based on a Taylor expansion of g around the point μ

Taking the result as given, let the distribution F of each X_i be uniform on $[0, \pi/2]$ and let $g(x) = \sin(x)$

Derive the asymptotic distribution of $\sqrt{n}\{g(\bar{X}_n) - g(\mu)\}$ and illustrate convergence in the same spirit as the program `illustrate_clt.jl` discussed above

What happens when you replace $[0, \pi/2]$ with $[0, \pi]$?

What is the source of the problem?

Exercise 2 Here’s a result that’s often used in developing statistical tests, and is connected to the multivariate central limit theorem

If you study econometric theory, you will see this result used again and again

Assume the setting of the multivariate CLT *discussed above*, so that

1. $\mathbf{X}_1, \dots, \mathbf{X}_n$ is a sequence of IID random vectors, each taking values in \mathbb{R}^k
2. $\boldsymbol{\mu} := \mathbb{E}[\mathbf{X}_i]$, and Σ is the variance-covariance matrix of \mathbf{X}_i
3. The convergence

$$\sqrt{n}(\bar{\mathbf{X}}_n - \boldsymbol{\mu}) \xrightarrow{d} N(\mathbf{0}, \Sigma) \quad (2.36)$$

is valid

In a statistical setting, one often wants the right hand side to be **standard** normal, so that confidence intervals are easily computed

This normalization can be achieved on the basis of three observations

First, if \mathbf{X} is a random vector in \mathbb{R}^k and \mathbf{A} is constant and $k \times k$, then

$$\text{Var}[\mathbf{AX}] = \mathbf{A} \text{Var}[\mathbf{X}] \mathbf{A}'$$

Second, by the **continuous mapping theorem**, if $\mathbf{Z}_n \xrightarrow{d} \mathbf{Z}$ in \mathbb{R}^k and \mathbf{A} is constant and $k \times k$, then

$$\mathbf{AZ}_n \xrightarrow{d} \mathbf{AZ}$$

Third, if \mathbf{S} is a $k \times k$ symmetric positive definite matrix, then there exists a symmetric positive definite matrix \mathbf{Q} , called the inverse square root of \mathbf{S} , such that

$$\mathbf{Q}\mathbf{S}\mathbf{Q}' = \mathbf{I}$$

Here \mathbf{I} is the $k \times k$ identity matrix

Putting these things together, your first exercise is to show that if \mathbf{Q} is the inverse square root of Σ , then

$$\mathbf{Z}_n := \sqrt{n}\mathbf{Q}(\bar{\mathbf{X}}_n - \boldsymbol{\mu}) \xrightarrow{d} \mathbf{Z} \sim N(\mathbf{0}, \mathbf{I})$$

Applying the continuous mapping theorem one more time tells us that

$$\|\mathbf{Z}_n\|^2 \xrightarrow{d} \|\mathbf{Z}\|^2$$

Given the distribution of \mathbf{Z} , we conclude that

$$n\|\mathbf{Q}(\bar{\mathbf{X}}_n - \boldsymbol{\mu})\|^2 \xrightarrow{d} \chi^2(k) \quad (2.37)$$

where $\chi^2(k)$ is the chi-squared distribution with k degrees of freedom

(Recall that k is the dimension of \mathbf{X}_i , the underlying random vectors)

Your second exercise is to illustrate the convergence in (2.37) with a simulation

In doing so, let

$$\mathbf{x}_i := \begin{pmatrix} W_i \\ U_i + W_i \end{pmatrix}$$

where

- each W_i is an IID draw from the uniform distribution on $[-1, 1]$
- each U_i is an IID draw from the uniform distribution on $[-2, 2]$
- U_i and W_i are independent of each other

Hints:

1. `sqrtm(A)` computes the square root of A . You still need to invert it
2. You should be able to work out Σ from the preceding information

Solutions

[Solution notebook](#)

Linear State Space Models

Contents

- *Linear State Space Models*
 - *Overview*
 - *The Linear State Space Model*
 - *Distributions and Moments*
 - *Stationarity and Ergodicity*
 - *Noisy Observations*
 - *Prediction*
 - *Code*
 - *Exercises*
 - *Solutions*

“We may regard the present state of the universe as the effect of its past and the cause of its future” – Marquis de Laplace

Overview

This lecture introduces the linear state space dynamic system

Easy to use and carries a powerful theory of prediction

A workhorse with many applications

- representing dynamics of higher-order linear systems
- predicting the position of a system j steps into the future
- predicting a geometric sum of future values of a variable like
 - non financial income
 - dividends on a stock
 - the money supply
 - a government deficit or surplus
 - etc., etc., ...
- key ingredient of useful models
 - Friedman’s permanent income model of consumption smoothing
 - Barro’s model of smoothing total tax collections
 - Rational expectations version of Cagan’s model of hyperinflation
 - Sargent and Wallace’s “unpleasant monetarist arithmetic”
 - etc., etc., ...

The Linear State Space Model

Objects in play

- An $n \times 1$ vector x_t denoting the **state** at time $t = 0, 1, 2, \dots$
- An iid sequence of $m \times 1$ random vectors $w_t \sim N(0, I)$
- A $k \times 1$ vector y_t of **observations** at time $t = 0, 1, 2, \dots$
- An $n \times n$ matrix A called the **transition matrix**
- An $n \times m$ matrix C called the **volatility matrix**
- A $k \times n$ matrix G sometimes called the **output matrix**

Here is the linear state-space system

$$\begin{aligned} x_{t+1} &= Ax_t + Cw_{t+1} \\ y_t &= Gx_t \\ x_0 &\sim N(\mu_0, \Sigma_0) \end{aligned} \tag{2.38}$$

Primitives The primitives of the model are

1. the matrices A, C, G
2. shock distribution, which we have specialized to $N(0, I)$
3. the distribution of the initial condition x_0 , which we have set to $N(\mu_0, \Sigma_0)$

Given A, C, G and draws of x_0 and w_1, w_2, \dots , the model (2.38) pins down the values of the sequences $\{x_t\}$ and $\{y_t\}$

Even without these draws, the primitives 1–3 pin down the *probability distributions* of $\{x_t\}$ and $\{y_t\}$

Later we'll see how to compute these distributions and their moments

Martingale difference shocks We've made the common assumption that the shocks are independent standardized normal vectors

But some of what we say will go through under the assumption that $\{w_{t+1}\}$ is a **martingale difference sequence**

A martingale difference sequence is a sequence that is zero mean when conditioned on past information

In the present case, since $\{x_t\}$ is our state sequence, this means that it satisfies

$$\mathbb{E}[w_{t+1}|x_t, x_{t-1}, \dots] = 0$$

This is a weaker condition than that $\{w_t\}$ is iid with $w_{t+1} \sim N(0, I)$

Examples By appropriate choice of the primitives, a variety of dynamics can be represented in terms of the linear state space model

The following examples help to highlight this point

They also illustrate the wise dictum *finding the state is an art*

Second-order difference equation Let $\{y_t\}$ be a deterministic sequence that satisfies

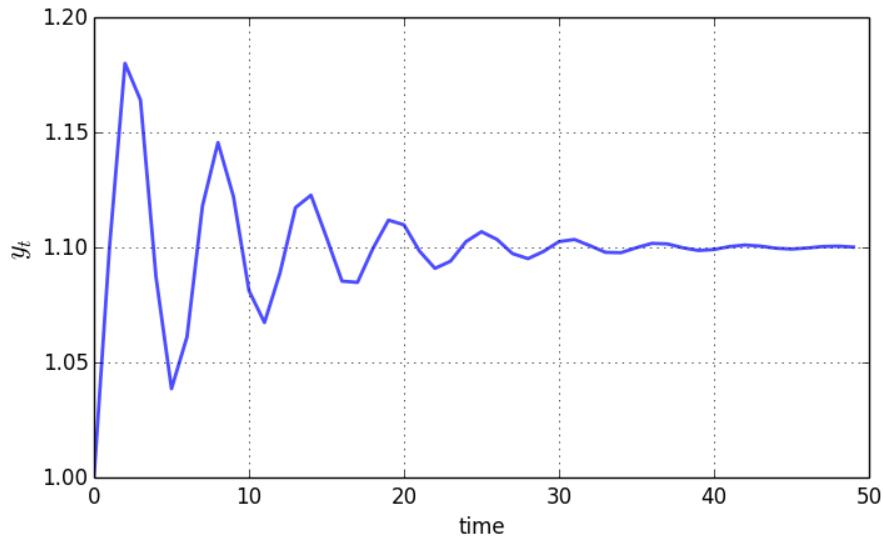
$$y_{t+1} = \phi_0 + \phi_1 y_t + \phi_2 y_{t-1} \quad \text{s.t. } y_0, y_{-1} \text{ given} \quad (2.39)$$

To map (2.39) into our state space system (2.38), we set

$$x_t = \begin{bmatrix} 1 \\ y_t \\ y_{t-1} \end{bmatrix} \quad A = \begin{bmatrix} 1 & 0 & 0 \\ \phi_0 & \phi_1 & \phi_2 \\ 0 & 1 & 0 \end{bmatrix} \quad C = \begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix} \quad G = [0 \ 1 \ 0]$$

You can confirm that under these definitions, (2.38) and (2.39) agree

The next figure shows dynamics of this process when $\phi_0 = 1.1, \phi_1 = 0.8, \phi_2 = -0.8, y_0 = y_{-1} = 1$



Later you'll be asked to recreate this figure

Univariate Autoregressive Processes We can use (2.38) to represent the model

$$y_{t+1} = \phi_1 y_t + \phi_2 y_{t-1} + \phi_3 y_{t-2} + \phi_4 y_{t-3} + \sigma w_{t+1} \quad (2.40)$$

where $\{w_t\}$ is iid and standard normal

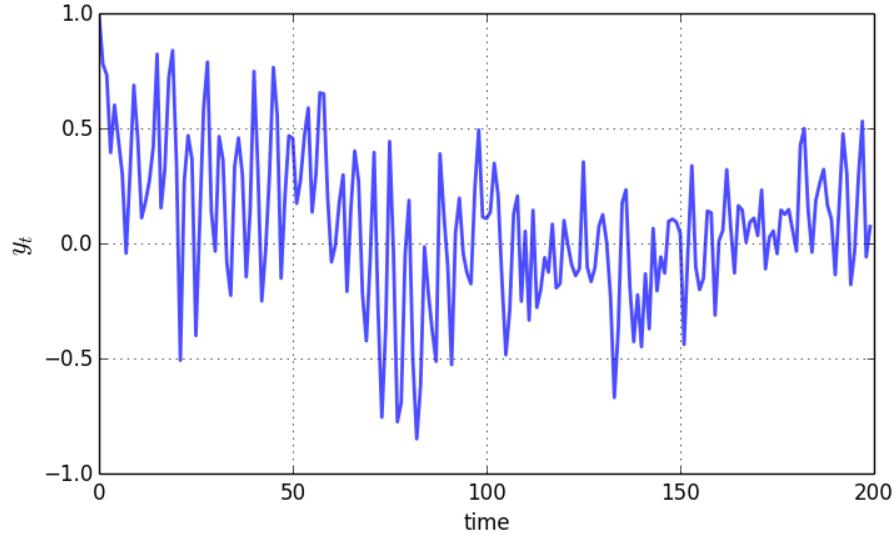
To put this in the linear state space format we take $x_t = [y_t \ y_{t-1} \ y_{t-2} \ y_{t-3}]'$ and

$$A = \begin{bmatrix} \phi_1 & \phi_2 & \phi_3 & \phi_4 \\ 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \end{bmatrix} \quad C = \begin{bmatrix} \sigma \\ 0 \\ 0 \\ 0 \end{bmatrix} \quad G = [1 \ 0 \ 0 \ 0]$$

The matrix A has the form of the *companion matrix* to the vector $[\phi_1 \ \phi_2 \ \phi_3 \ \phi_4]$.

The next figure shows dynamics of this process when

$$\phi_1 = 0.5, \phi_2 = -0.2, \phi_3 = 0, \phi_4 = 0.5, \sigma = 0.2, y_0 = y_{-1} = y_{-2} = y_{-3} = 1$$



Vector Autoregressions Now suppose that

- y_t is a $k \times 1$ vector
- ϕ_j is a $k \times k$ matrix and
- w_t is $k \times 1$

Then (2.40) is termed a *vector autoregression*

To map this into (2.38), we set

$$x_t = \begin{bmatrix} y_t \\ y_{t-1} \\ y_{t-2} \\ y_{t-3} \end{bmatrix} \quad A = \begin{bmatrix} \phi_1 & \phi_2 & \phi_3 & \phi_4 \\ I & 0 & 0 & 0 \\ 0 & I & 0 & 0 \\ 0 & 0 & I & 0 \end{bmatrix} \quad C = \begin{bmatrix} \sigma \\ 0 \\ 0 \\ 0 \end{bmatrix} \quad G = [I \ 0 \ 0 \ 0]$$

where I is the $k \times k$ identity matrix and σ is a $k \times k$ matrix

Seasonals We can use (2.38) to represent

1. the *deterministic seasonal* $y_t = y_{t-4}$
2. the *indeterministic seasonal* $y_t = \phi_4 y_{t-4} + w_t$

In fact both are special cases of (2.40)

With the deterministic seasonal, the transition matrix becomes

$$A = \begin{bmatrix} 0 & 0 & 0 & 1 \\ 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \end{bmatrix}$$

It is easy to check that $A^4 = I$, which implies that x_t is strictly periodic with period 4:¹

$$x_{t+4} = x_t$$

Such an x_t process can be used to model deterministic seasonals in quarterly time series.

The *indeterministic* seasonal produces recurrent, but aperiodic, seasonal fluctuations.

Time Trends The model $y_t = at + b$ is known as a *linear time trend*

We can represent this model in the linear state space form by taking

$$A = \begin{bmatrix} 1 & 1 \\ 0 & 1 \end{bmatrix} \quad C = \begin{bmatrix} 0 \\ 0 \end{bmatrix} \quad G = [a \ b] \quad (2.41)$$

and starting at initial condition $x_0 = [0 \ 1]'$

In fact it's possible to use the state-space system to represent polynomial trends of any order

For instance, let

$$x_0 = \begin{bmatrix} 0 \\ 0 \\ 1 \end{bmatrix} \quad A = \begin{bmatrix} 1 & 1 & 0 \\ 0 & 1 & 1 \\ 0 & 0 & 1 \end{bmatrix} \quad C = \begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix}$$

It follows that

$$A^t = \begin{bmatrix} 1 & t & t(t-1)/2 \\ 0 & 1 & t \\ 0 & 0 & 1 \end{bmatrix}$$

Then $x'_t = [t(t-1)/2 \ t \ 1]$, so that x_t contains linear and quadratic time trends

As a variation on the linear time trend model, consider $y_t = at + b$

To modify (2.41) accordingly, we set

$$A = \begin{bmatrix} 1 & 1 \\ 0 & 1 \end{bmatrix} \quad C = \begin{bmatrix} 0 \\ 0 \end{bmatrix} \quad G = [a \ b] \quad (2.42)$$

¹ The eigenvalues of A are $(1, -1, i, -i)$.

Moving Average Representations A nonrecursive expression for x_t as a function of $x_0, w_1, w_2, \dots, w_t$ can be found by using (2.38) repeatedly to obtain

$$\begin{aligned} x_t &= Ax_{t-1} + Cw_t \\ &= A^2x_{t-2} + ACw_{t-1} + Cw_t \\ &\quad \vdots \\ &= \sum_{j=0}^{t-1} A^j C w_{t-j} + A^t x_0 \end{aligned} \tag{2.43}$$

Representation (2.43) is a *moving average* representation

It expresses $\{x_t\}$ as a linear function of

1. current and past values of the process $\{w_t\}$ and
2. the initial condition x_0

As an example of a moving average representation, let the model be

$$A = \begin{bmatrix} 1 & 1 \\ 0 & 1 \end{bmatrix} \quad C = \begin{bmatrix} 1 \\ 0 \end{bmatrix}$$

You will be able to show that $A^t = \begin{bmatrix} 1 & t \\ 0 & 1 \end{bmatrix}$ and $A^j C = [1 \ 0]'$

Substituting into the moving average representation (2.43), we obtain

$$x_{1t} = \sum_{j=0}^{t-1} w_{t-j} + [1 \ t] x_0$$

where x_{1t} is the first entry of x_t

The first term on the right is a cumulated sum of martingale differences, and is therefore a *martingale*

The second term is a translated linear function of time

For this reason, x_{1t} is called a *martingale with drift*

Distributions and Moments

Unconditional Moments Using (2.38), it's easy to obtain expressions for the (unconditional) means of x_t and y_t

We'll explain what *unconditional* and *conditional* mean soon

Letting $\mu_t := \mathbb{E}[x_t]$ and using linearity of expectations, we find that

$$\mu_{t+1} = A\mu_t \quad \text{with } \mu_0 \text{ given} \tag{2.44}$$

Here μ_0 is a primitive given in (2.38)

The variance-covariance matrix of x_t is $\Sigma_t := \mathbb{E}[(x_t - \mu_t)(x_t - \mu_t)']$

Using $x_{t+1} - \mu_{t+1} = A(x_t - \mu_t) + Cw_{t+1}$, we can determine this matrix recursively via

$$\Sigma_{t+1} = A\Sigma_t A' + CC' \quad \text{with } \Sigma_0 \text{ given} \quad (2.45)$$

As with μ_0 , the matrix Σ_0 is a primitive given in (2.38)

As a matter of terminology, we will sometimes call

- μ_t the *unconditional mean* of x_t
- Σ_t the *unconditional variance-covariance matrix* of x_t

This is to distinguish μ_t and Σ_t from related objects that use conditioning information, to be defined below

However, you should be aware that these “unconditional” moments do depend on the initial distribution $N(\mu_0, \Sigma_0)$

Moments of the Observations Using linearity of expectations again we have

$$\mathbb{E}[y_t] = \mathbb{E}[Gx_t] = G\mu_t \quad (2.46)$$

The variance-covariance matrix of y_t is easily shown to be

$$\text{Var}[y_t] = \text{Var}[Gx_t] = G\Sigma_t G' \quad (2.47)$$

Distributions In general, knowing the mean and variance-covariance matrix of a random vector is not quite as good as knowing the full distribution

However, there are some situations where these moments alone tell us all we need to know

One such situation is when the vector in question is Gaussian (i.e., normally distributed)

This is the case here, given

1. our Gaussian assumptions on the primitives
2. the fact that normality is preserved under linear operations

In fact, it's well-known that

$$u \sim N(\bar{u}, S) \quad \text{and} \quad v = a + Bu \implies v \sim N(a + B\bar{u}, BSB') \quad (2.48)$$

In particular, given our Gaussian assumptions on the primitives and the linearity of (2.38) we can see immediately that both x_t and y_t are Gaussian for all $t \geq 0$ ²

Since x_t is Gaussian, to find the distribution, all we need to do is find its mean and variance-covariance matrix

But in fact we've already done this, in (2.44) and (2.45)

² The correct way to argue this is by induction. Suppose that x_t is Gaussian. Then (2.38) and (2.48) imply that x_{t+1} is Gaussian. Since x_0 is assumed to be Gaussian, it follows that every x_t is Gaussian. Evidently this implies that each y_t is Gaussian.

Letting μ_t and Σ_t be as defined by these equations, we have

$$x_t \sim N(\mu_t, \Sigma_t) \quad (2.49)$$

By similar reasoning combined with (2.46) and (2.47),

$$y_t \sim N(G\mu_t, G\Sigma_t G') \quad (2.50)$$

Ensemble Interpretations How should we interpret the distributions defined by (2.49)–(2.50)?

Intuitively, the probabilities in a distribution correspond to relative frequencies in a large population drawn from that distribution

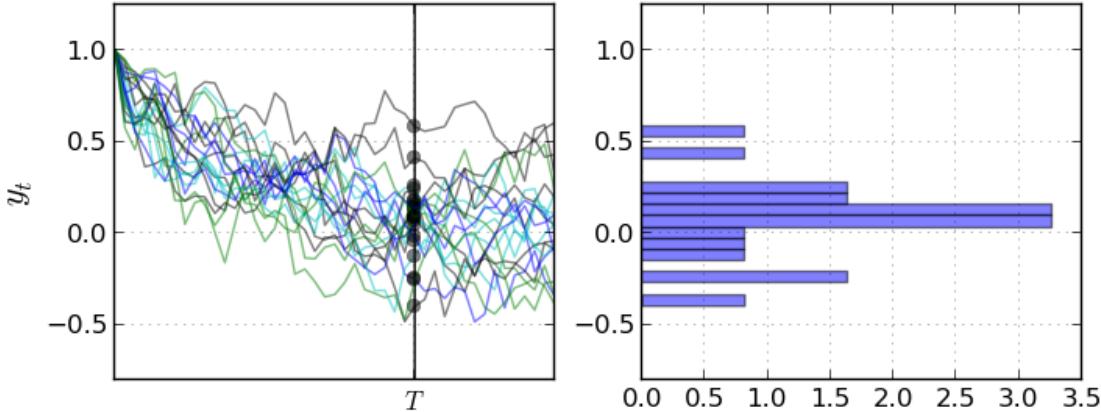
Let's apply this idea to our setting, focusing on the distribution of y_T for fixed T

We can generate independent draws of y_T by repeatedly simulating the evolution of the system up to time T , using an independent set of shocks each time

The next figure shows 20 simulations, producing 20 time series for $\{y_t\}$, and hence 20 draws of y_T

The system in question is the univariate autoregressive model (2.40)

The values of y_T are represented by black dots in the left-hand figure



In the right-hand figure, these values are converted into a rotated histogram that shows relative frequencies from our sample of 20 y_T 's

(The parameters and source code for the figures can be found in file `linear_models/paths_and_hist.py` from the [applications repository](#))

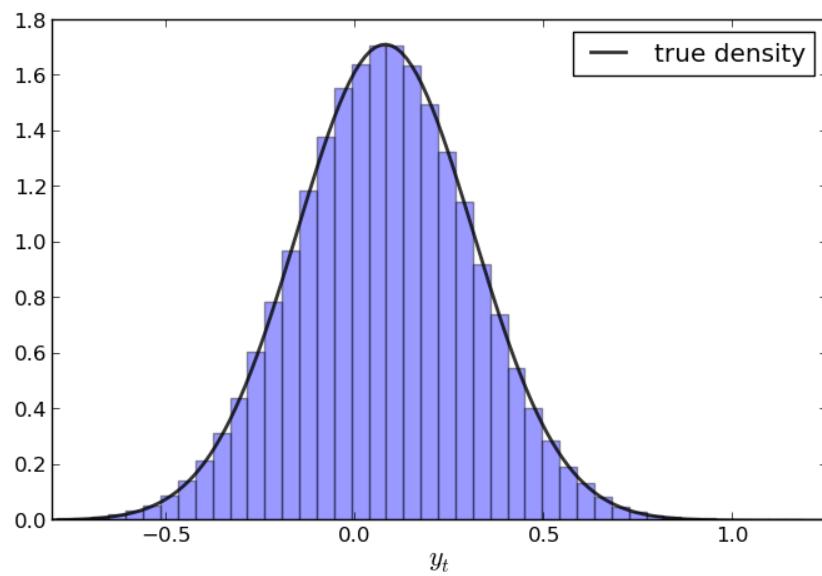
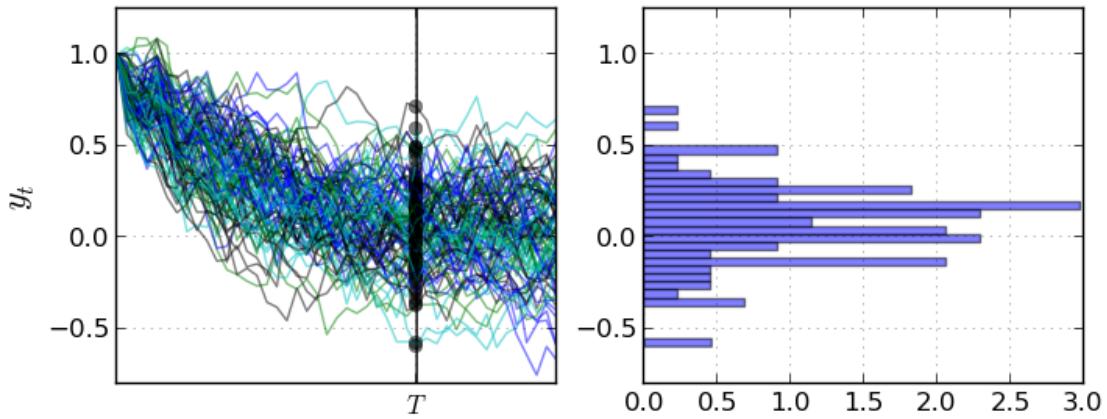
Here is another figure, this time with 100 observations

Let's now try with 500,000 observations, showing only the histogram (without rotation)

The black line is the density of y_T calculated analytically, using (2.50)

The histogram and analytical distribution are close, as expected

By looking at the figures and experimenting with parameters, you will gain a feel for how the distribution depends on the model primitives *listed above*



Ensemble means In the preceding figure we recovered the distribution of y_T by

1. generating I sample paths (i.e., time series) where I is a large number
2. recording each observation y_T^i
3. histogramming this sample

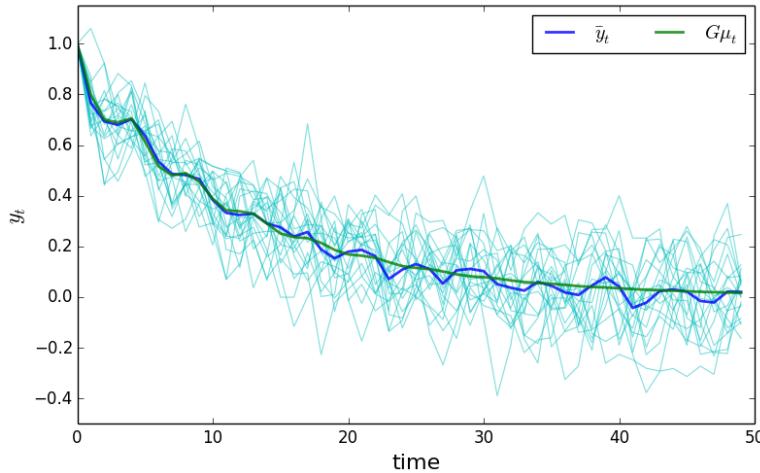
Just as the histogram corresponds to the distribution, the *ensemble* or *cross-sectional average*

$$\bar{y}_T := \frac{1}{I} \sum_{i=1}^I y_T^i$$

approximates the expectation $\mathbb{E}[y_T] = G\mu_T$ (as implied by the law of large numbers)

Here's a simulation comparing the ensemble averages and population means at time points $t = 0, \dots, 50$

The parameters are the same as for the preceding figures, and the sample size is relatively small ($I = 20$)



The ensemble mean for x_t is

$$\bar{x}_T := \frac{1}{I} \sum_{i=1}^I x_T^i \rightarrow \mu_T \quad (I \rightarrow \infty)$$

The limit μ_T can be thought of as a "population average"

(By *population average* we mean the average for an infinite ($I = \infty$) number of sample x_T 's)

Another application of the law of large numbers assures us that

$$\frac{1}{I} \sum_{i=1}^I (x_T^i - \bar{x}_T)(x_T^i - \bar{x}_T)' \rightarrow \Sigma_T \quad (I \rightarrow \infty)$$

Joint Distributions In the preceding discussion we looked at the distributions of x_t and y_t in isolation

This gives us useful information, but doesn't allow us to answer questions like

- what's the probability that $x_t \geq 0$ for all t ?
- what's the probability that the process $\{y_t\}$ exceeds some value a before falling below b ?
- etc., etc.

Such questions concern the *joint distributions* of these sequences

To compute the joint distribution of x_0, x_1, \dots, x_T , recall that joint and conditional densities are linked by the rule

$$p(x, y) = p(y | x)p(x) \quad (\text{joint} = \text{conditional} \times \text{marginal})$$

From this rule we get $p(x_0, x_1) = p(x_1 | x_0)p(x_0)$

The Markov property $p(x_t | x_{t-1}, \dots, x_0) = p(x_t | x_{t-1})$ and repeated applications of the preceding rule lead us to

$$p(x_0, x_1, \dots, x_T) = p(x_0) \prod_{t=0}^{T-1} p(x_{t+1} | x_t)$$

The marginal $p(x_0)$ is just the primitive $N(\mu_0, \Sigma_0)$

In view of (2.38), the conditional densities are

$$p(x_{t+1} | x_t) = N(Ax_t, CC')$$

Autocovariance functions An important object related to the joint distribution is the *autocovariance function*

$$\Sigma_{t+j,t} := \mathbb{E} [(x_{t+j} - \mu_{t+j})(x_t - \mu_t)'] \quad (2.51)$$

Elementary calculations show that

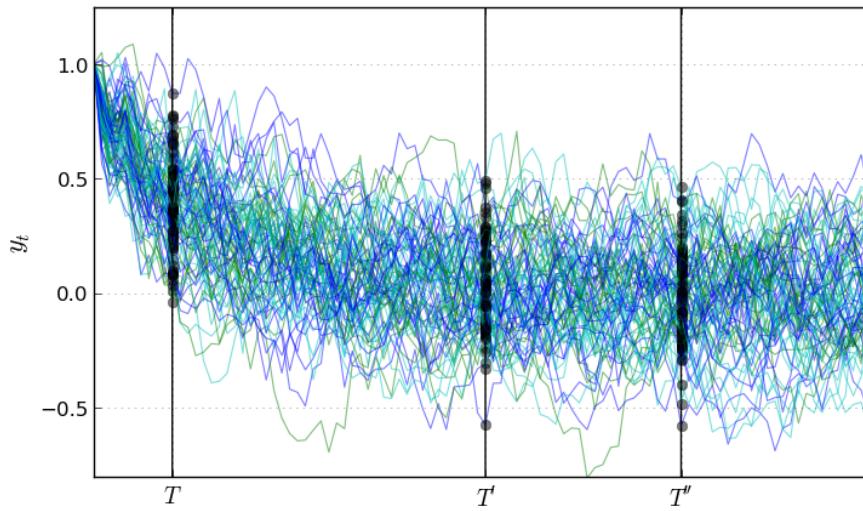
$$\Sigma_{t+j,t} = A^j \Sigma_t \quad (2.52)$$

Notice that $\Sigma_{t+j,t}$ in general depends on both j , the gap between the two dates, and t , the earlier date

Stationarity and Ergodicity

Stationarity and ergodicity are two properties that, when they hold, greatly aid analysis of linear state space models

Let's start with the intuition



Visualizing Stability Let's look at some more time series from the same model that we analyzed above

This picture shows cross-sectional distributions for y_t at times T, T', T''

Note how the time series "settle down" in the sense that the distributions at T' and T'' are relatively similar to each other — but unlike the distribution at T

Apparently, the distributions of y_t converge to a fixed long-run distribution as $t \rightarrow \infty$

When such a distribution exists it is called a *stationary distribution*

Stationary Distributions In our setting, a distribution ψ_∞ is said to be *stationary* for x_t if

$$x_t \sim \psi_\infty \quad \text{and} \quad x_{t+1} = Ax_t + Cw_{t+1} \implies x_{t+1} \sim \psi_\infty$$

Since

1. in the present case all distributions are Gaussian
2. a Gaussian distribution is pinned down by its mean and variance-covariance matrix

we can restate the definition as follows: ψ_∞ is stationary for x_t if

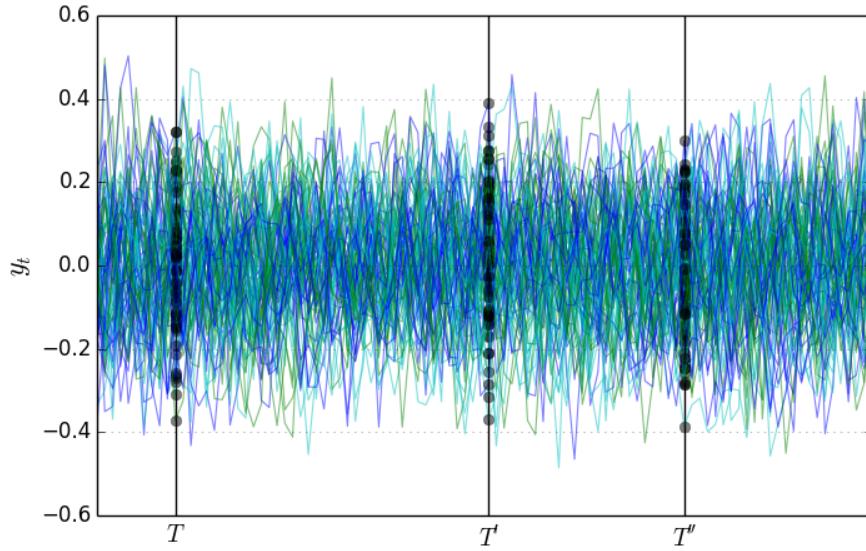
$$\psi_\infty = N(\mu_\infty, \Sigma_\infty)$$

where μ_∞ and Σ_∞ are fixed points of (2.44) and (2.45) respectively

Covariance Stationary Processes Let's see what happens to the preceding figure if we start x_0 at the stationary distribution

Now the differences in the observed distributions at T, T' and T'' come entirely from random fluctuations due to the finite sample size

By



- our choosing $x_0 \sim N(\mu_\infty, \Sigma_\infty)$
- the definitions of μ_∞ and Σ_∞ as fixed points of (2.44) and (2.45) respectively

we've ensured that

$$\mu_t = \mu_\infty \quad \text{and} \quad \Sigma_t = \Sigma_\infty \quad \text{for all } t$$

Moreover, in view of (2.52), the autocovariance function takes the form $\Sigma_{t+j,t} = A^j \Sigma_\infty$, which depends on j but not on t

This motivates the following definition

A process $\{x_t\}$ is said to be *covariance stationary* if

- both μ_t and Σ_t are constant in t
- $\Sigma_{t+j,t}$ depends on the time gap j but not on time t

In our setting, $\{x_t\}$ will be covariance stationary if μ_0, Σ_0, A, C assume values that imply that none of $\mu_t, \Sigma_t, \Sigma_{t+j,t}$ depends on t

Conditions for Stationarity

The globally stable case The difference equation $\mu_{t+1} = A\mu_t$ is known to have *unique* fixed point $\mu_\infty = 0$ if all eigenvalues of A have moduli strictly less than unity

That is, if `all(abs(eigvals(A)) .< 1) == true`

The difference equation (2.45) also has a unique fixed point in this case, and, moreover

$$\mu_t \rightarrow \mu_\infty = 0 \quad \text{and} \quad \Sigma_t \rightarrow \Sigma_\infty \quad \text{as} \quad t \rightarrow \infty$$

regardless of the initial conditions μ_0 and Σ_0

This is the *globally stable case* — see these notes for more a theoretical treatment

However, global stability is more than we need for stationary solutions, and often more than we want

To illustrate, consider *our second order difference equation example*

Here the state is $x_t = [1 \ y_t \ y_{t-1}]'$

Because of the constant first component in the state vector, we will never have $\mu_t \rightarrow 0$

How can we find stationary solutions that respect a constant state component?

Processes with a constant state component To investigate such a process, suppose that A and C take the form

$$A = \begin{bmatrix} A_1 & a \\ 0 & 1 \end{bmatrix} \quad C = \begin{bmatrix} C_1 \\ 0 \end{bmatrix}$$

where

- A_1 is an $(n - 1) \times (n - 1)$ matrix
- a is an $(n - 1) \times 1$ column vector

Let $x_t = [x'_{1t} \ 1]'$ where x_{1t} is $(n - 1) \times 1$

It follows that

$$x_{1,t+1} = A_1 x_{1t} + a + C_1 w_{t+1}$$

Let $\mu_{1t} = \mathbb{E}[x_{1t}]$ and take expectations on both sides of this expression to get

$$\mu_{1,t+1} = A_1 \mu_{1,t} + a \tag{2.53}$$

Assume now that the moduli of the eigenvalues of A_1 are all strictly less than one

Then (2.53) has a unique stationary solution, namely,

$$\mu_{1\infty} = (I - A_1)^{-1} a$$

The stationary value of μ_t itself is then $\mu_\infty := [\mu'_{1\infty} \ 1]'$

The stationary values of Σ_t and $\Sigma_{t+j,t}$ satisfy

$$\begin{aligned} \Sigma_\infty &= A \Sigma_\infty A' + C C' \\ \Sigma_{t+j,t} &= A^j \Sigma_\infty \end{aligned} \tag{2.54}$$

Notice that here $\Sigma_{t+j,t}$ depends on the time gap j but not on calendar time t

In conclusion, if

- $x_0 \sim N(\mu_\infty, \Sigma_\infty)$ and
- the moduli of the eigenvalues of A_1 are all strictly less than unity

then the $\{x_t\}$ process is covariance stationary, with constant state component

Note: If the eigenvalues of A_1 are less than unity in modulus, then (a) starting from any initial value, the mean and variance-covariance matrix both converge to their stationary values; and (b) iterations on (2.45) converge to the fixed point of the *discrete Lyapunov equation* in the first line of (2.54)

Ergodicity Let's suppose that we're working with a covariance stationary process

In this case we know that the ensemble mean will converge to μ_∞ as the sample size T approaches infinity

Averages over time Ensemble averages across simulations are interesting theoretically, but in real life we usually observe only a *single* realization $\{x_t, y_t\}_{t=0}^T$

So now let's take a single realization and form the time series averages

$$\bar{x} := \frac{1}{T} \sum_{t=1}^T x_t \quad \text{and} \quad \bar{y} := \frac{1}{T} \sum_{t=1}^T y_t$$

Do these time series averages converge to something interpretable in terms of our basic state-space representation?

The answer depends on something called *ergodicity*

Ergodicity is the property that time series and ensemble averages coincide

More formally, ergodicity implies that time series sample averages converge to their expectation under the stationary distribution

In particular,

- $\frac{1}{T} \sum_{t=1}^T x_t \rightarrow \mu_\infty$
- $\frac{1}{T} \sum_{t=1}^T (x_t - \bar{x}_T)(x_t - \bar{x}_T)' \rightarrow \Sigma_\infty$
- $\frac{1}{T} \sum_{t=1}^T (x_{t+j} - \bar{x}_T)(x_t - \bar{x}_T)' \rightarrow A^j \Sigma_\infty$

In our linear Gaussian setting, any covariance stationary process is also ergodic

Noisy Observations

In some settings the observation equation $y_t = Gx_t$ is modified to include an error term

Often this error term represents the idea that the true state can only be observed imperfectly

To include an error term in the observation we introduce

- An iid sequence of $\ell \times 1$ random vectors $v_t \sim N(0, I)$
- A $k \times \ell$ matrix H

and extend the linear state-space system to

$$\begin{aligned}x_{t+1} &= Ax_t + Cw_{t+1} \\y_t &= Gx_t + Hv_t \\x_0 &\sim N(\mu_0, \Sigma_0)\end{aligned}\tag{2.55}$$

The sequence $\{v_t\}$ is assumed to be independent of $\{w_t\}$

The process $\{x_t\}$ is not modified by noise in the observation equation and its moments, distributions and stability properties remain the same

The unconditional moments of y_t from (2.46) and (2.47) now become

$$\mathbb{E}[y_t] = \mathbb{E}[Gx_t + Hv_t] = G\mu_t\tag{2.56}$$

The variance-covariance matrix of y_t is easily shown to be

$$\text{Var}[y_t] = \text{Var}[Gx_t + Hv_t] = G\Sigma_t G' + HH'\tag{2.57}$$

The distribution of y_t is therefore

$$y_t \sim N(G\mu_t, G\Sigma_t G' + HH')$$

Prediction

The theory of prediction for linear state space systems is elegant and simple

Forecasting Formulas – Conditional Means The natural way to predict variables is to use conditional distributions

For example, the optimal forecast of x_{t+1} given information known at time t is

$$\mathbb{E}_t[x_{t+1}] := \mathbb{E}[x_{t+1} | x_t, x_{t-1}, \dots, x_0] = Ax_t$$

The right-hand side follows from $x_{t+1} = Ax_t + Cw_{t+1}$ and the fact that w_{t+1} is zero mean and independent of x_t, x_{t-1}, \dots, x_0

That $\mathbb{E}_t[x_{t+1}] = \mathbb{E}[x_{t+1} | x_t]$ is an implication of $\{x_t\}$ having the *Markov property*

The one-step-ahead forecast error is

$$x_{t+1} - \mathbb{E}_t[x_{t+1}] = Cw_{t+1}$$

The covariance matrix of the forecast error is

$$\mathbb{E}[(x_{t+1} - \mathbb{E}_t[x_{t+1}])(x_{t+1} - \mathbb{E}_t[x_{t+1}])'] = CC'$$

More generally, we'd like to compute the j -step ahead forecasts $\mathbb{E}_t[x_{t+j}]$ and $\mathbb{E}_t[y_{t+j}]$

With a bit of algebra we obtain

$$x_{t+j} = A^j x_t + A^{j-1} C w_{t+1} + A^{j-2} C w_{t+2} + \cdots + A^0 C w_{t+j}$$

In view of the iid property, current and past state values provide no information about future values of the shock

Hence $\mathbb{E}_t[w_{t+k}] = \mathbb{E}[w_{t+k}] = 0$

It now follows from linearity of expectations that the j -step ahead forecast of x is

$$\mathbb{E}_t[x_{t+j}] = A^j x_t$$

The j -step ahead forecast of y is therefore

$$\mathbb{E}_t[y_{t+j}] = \mathbb{E}_t[Gx_{t+j} + Hv_{t+j}] = GA^j x_t$$

Covariance of Prediction Errors It is useful to obtain the covariance matrix of the vector of j -step-ahead prediction errors

$$x_{t+j} - \mathbb{E}_t[x_{t+j}] = \sum_{s=0}^{j-1} A^s C w_{t-s+j} \quad (2.58)$$

Evidently,

$$V_j := \mathbb{E}_t[(x_{t+j} - \mathbb{E}_t[x_{t+j}]) (x_{t+j} - \mathbb{E}_t[x_{t+j}])'] = \sum_{k=0}^{j-1} A^k C C' A^k' \quad (2.59)$$

V_j defined in (2.59) can be calculated recursively via $V_1 = CC'$ and

$$V_j = CC' + AV_{j-1}A', \quad j \geq 2 \quad (2.60)$$

V_j is the *conditional covariance matrix* of the errors in forecasting x_{t+j} , conditioned on time t information x_t

Under particular conditions, V_j converges to

$$V_\infty = CC' + AV_\infty A' \quad (2.61)$$

Equation (2.61) is an example of a *discrete Lyapunov equation* in the covariance matrix V_∞

A sufficient condition for V_j to converge is that the eigenvalues of A be strictly less than one in modulus.

Weaker sufficient conditions for convergence associate eigenvalues equaling or exceeding one in modulus with elements of C that equal 0

Forecasts of Geometric Sums In several contexts, we want to compute forecasts of geometric sums of future random variables governed by the linear state-space system (2.38)

We want the following objects

- Forecast of a geometric sum of future x 's, or $\mathbb{E}_t \left[\sum_{j=0}^{\infty} \beta^j x_{t+j} \right]$
- Forecast of a geometric sum of future y 's, or $\mathbb{E}_t \left[\sum_{j=0}^{\infty} \beta^j y_{t+j} \right]$

These objects are important components of some famous and interesting dynamic models

For example,

- if $\{y_t\}$ is a stream of dividends, then $\mathbb{E} \left[\sum_{j=0}^{\infty} \beta^j y_{t+j} | x_t \right]$ is a model of a stock price
- if $\{y_t\}$ is the money supply, then $\mathbb{E} \left[\sum_{j=0}^{\infty} \beta^j y_{t+j} | x_t \right]$ is a model of the price level

Formulas Fortunately, it is easy to use a little matrix algebra to compute these objects

Suppose that every eigenvalue of A has modulus strictly less than $\frac{1}{\beta}$

It then follows that $I + \beta A + \beta^2 A^2 + \dots = [I - \beta A]^{-1}$

This leads to our formulas:

- Forecast of a geometric sum of future x 's

$$\mathbb{E}_t \left[\sum_{j=0}^{\infty} \beta^j x_{t+j} \right] = [I + \beta A + \beta^2 A^2 + \dots] x_t = [I - \beta A]^{-1} x_t$$

- Forecast of a geometric sum of future y 's

$$\mathbb{E}_t \left[\sum_{j=0}^{\infty} \beta^j y_{t+j} \right] = G[I + \beta A + \beta^2 A^2 + \dots] x_t = G[I - \beta A]^{-1} x_t$$

Code

Our preceding simulations and calculations are based on code in the file `lss.py` from the `QuantEcon` package

The code implements a type for handling linear state space models (simulations, calculating moments, etc.)

We repeat it here for convenience

```
#=
Computes quantities related to the Gaussian linear state space model

x_{t+1} = A x_t + C w_{t+1}

y_t = G x_t

The shocks {w_t} are iid and N(0, I)

@author : Spencer Lyon <spencer.lyon@nyu.edu>
@date : 2014-07-28

References
-----
```

```

TODO: Come back and update to match `LinearStateSpace` type from py side
TODO: Add docstrings

http://quant-econ.net/jl/linear_models.html

=#
import Distributions: MultivariateNormal, rand
import Base: ==

#=

    numpy allows its multivariate_normal function to have a matrix of
    zeros for the covariance matrix; Stats.jl doesn't. This type just
    gives a `rand` method when we pass in a matrix of zeros for Sigma_0
    so the rest of the api can work, unaffected

    The behavior of `rand` is to just pass back the mean vector when
    the covariance matrix is zero.

=#
type FakeMVTNorm{T <: Real}
    mu_0::Array{T}
    Sigma_0::Array{T}
end

==(f1::FakeMVTNorm, f2::FakeMVTNorm) =
    (f1.mu_0 == f2.mu_0) && (f1.Sigma_0 == f2.Sigma_0)

Base.rand{T}(d::FakeMVTNorm{T}) = copy(d.mu_0)
"""

A type that describes the Gaussian Linear State Space Model
of the form:

     $x_{t+1} = A x_t + C w_{t+1}$ 

     $y_t = G x_t$ 

where  $\{w_t\}$  and  $\{v_t\}$  are independent and standard normal with dimensions
 $k$  and  $l$  respectively. The initial conditions are  $\mu_0$  and  $\Sigma_0$  for  $x_0 \sim N(\mu_0, \Sigma_0)$ . When  $\Sigma_0=0$ , the draw of  $x_0$  is exactly  $\mu_0$ .

#### Fields

- `A::Matrix` Part of the state transition equation. It should be `n x n`
- `C::Matrix` Part of the state transition equation. It should be `n x m`
- `G::Matrix` Part of the observation equation. It should be `k x n`
- `k::Int` Dimension
- `n::Int` Dimension
- `m::Int` Dimension
- `mu_0::Vector` This is the mean of initial draw and is of length `n`
- `Sigma_0::Matrix` This is the variance of the initial draw and is `n x n` and
    also should be positive definite and symmetric

"""

type LSS

```

```

A::Matrix
C::Matrix
G::Matrix
k::Int
n::Int
m::Int
mu_0::Vector
Sigma_0::Matrix
dist::Union{MultivariateNormal, FakeMVTNorm}
end

function LSS(A::ScalarOrArray, C::ScalarOrArray, G::ScalarOrArray,
            mu_0::ScalarOrArray,
            Sigma_0::Matrix=zeros(size(G, 2), size(G, 2)))
    k = size(G, 1)
    n = size(G, 2)
    m = size(C, 2)

    # coerce shapes
    A = reshape(vcat(A), n, n)
    C = reshape(vcat(C), n, m)
    G = reshape(vcat(G), k, n)

    mu_0 = reshape([mu_0;], n)

    # define distribution
    if all(Sigma_0 .== 0.0)    # no variance -- no distribution
        dist = FakeMVTNorm(mu_0, Sigma_0)
    else
        dist = MultivariateNormal(mu_0, Sigma_0)
    end
    LSS(A, C, G, k, n, m, mu_0, Sigma_0, dist)
end

# make kwarg version
function LSS(A::ScalarOrArray, C::ScalarOrArray, G::ScalarOrArray;
            mu_0::Vector=zeros(size(G, 2)),
            Sigma_0::Matrix=zeros(size(G, 2), size(G, 2)))
    return LSS(A, C, G, mu_0, Sigma_0)
end

function simulate(lss::LSS, ts_length=100)
    x = Array(Float64, lss.n, ts_length)
    x[:, 1] = rand(lss.dist)
    w = randn(lss.m, ts_length - 1)
    for t=1:ts_length-1
        x[:, t+1] = lss.A * x[:, t] .+ lss.C * w[:, t]
    end
    y = lss.G * x

    return x, y
end

```

```

end

"""
Simulate num_reps observations of x_T and y_T given x_0 ~ N(mu_0, Sigma_0).

#### Arguments

- `lss::LSS` An instance of the Gaussian linear state space model.
- `t::Int = 10` The period that we want to replicate values for.
- `num_reps::Int = 100` The number of replications we want

#### Returns

- `x::Matrix` An n x num_reps matrix, where the j-th column is the j-th
  observation of x_T
- `y::Matrix` An k x num_reps matrix, where the j-th column is the j-th
  observation of y_T

"""

function replicate(lss::LSS, t::Integer, num_reps::Integer=100)
    x = Array(Float64, lss.n, num_reps)
    for j=1:num_reps
        x_t, _ = simulate(lss, t+1)
        x[:, j] = x_t[:, end]
    end

    y = lss.G * x
    return x, y
end

replicate(lss::LSS; t::Integer=10, num_reps::Integer=100) =
    replicate(lss, t, num_reps)

"""

Create a generator to calculate the population mean and
variance-convariance matrix for both x_t and y_t, starting at
the initial condition (self.mu_0, self.Sigma_0). Each iteration
produces a 4-tuple of items (mu_x, mu_y, Sigma_x, Sigma_y) for
the next period.

#### Arguments

- `lss::LSS` An instance of the Gaussian linear state space model

"""

function moment_sequence(lss::LSS)
    A, C, G = lss.A, lss.C, lss.G
    mu_x, Sigma_x = copy(lss.mu_0), copy(lss.Sigma_0)
    while true
        mu_y, Sigma_y = G * mu_x, G * Sigma_x * G'
        produce((mu_x, mu_y, Sigma_x, Sigma_y))

        # Update moments of x

```

```

    mu_x = A * mu_x
    Sigma_x = A * Sigma_x * A' + C * C'
end
end

"""
Compute the moments of the stationary distributions of x_t and
y_t if possible. Computation is by iteration, starting from the
initial conditions lss.mu_0 and lss.Sigma_0

#### Arguments

- `lss::LSS` An instance of the Guassian linear state space model
- `;max_iter::Int = 200` The maximum number of iterations allowed
- `;tol::Float64 = 1e-5` The tolerance level one wishes to achieve

#### Returns

- `mu_x::Vector` Represents the stationary mean of x_t
- `mu_y::Vector` Represents the stationary mean of y_t
- `Sigma_x::Matrix` Represents the var-cov matrix
- `Sigma_y::Matrix` Represents the var-cov matrix

"""

function stationary_distributions(lss::LSS; max_iter=200, tol=1e-5)
    # Initialize iteration
    m = @task moment_sequence(lss)
    mu_x, mu_y, Sigma_x, Sigma_y = consume(m)

    i = 0
    err = tol + 1.

    while err > tol
        if i > max_iter
            error("Convergence failed after $i iterations")
        else
            i += 1
            mu_x1, mu_y, Sigma_x1, Sigma_y = consume(m)
            err_mu = Base.maxabs(mu_x1 - mu_x)
            err_Sigma = Base.maxabs(Sigma_x1 - Sigma_x)
            err = max(err_Sigma, err_mu)
            mu_x, Sigma_x = mu_x1, Sigma_x1
        end
    end

    return mu_x, mu_y, Sigma_x, Sigma_y
end

function geometric_sums(lss::LSS, bet, x_t)
    I = eye(lss.n)
    S_x = (I - bet .* lss.A) \ x_t
    S_y = lss.G * S_x

```

```

    return S_x, S_y
end

```

Hopefully the code is relatively self explanatory and adequately documented

Examples of usage are given in the solutions to the exercises

Exercises

Exercise 1 Replicate *this figure* using the `LinearStateSpace` type from `lss.py`

Exercise 2 Replicate *this figure* modulo randomness using the same type

Exercise 3 Replicate *this figure* modulo randomness using the same type

The state space model and parameters are the same as for the preceding exercise

Exercise 4 Replicate *this figure* modulo randomness using the same type

The state space model and parameters are the same as for the preceding exercise, except that the initial condition is the stationary distribution

Hint: You can use the `stationary_distributions` method to get the initial conditions

The number of sample paths is 80, and the time horizon in the figure is 100

Producing the vertical bars and dots is optional, but if you wish to try, the bars are at dates 10, 50 and 75

Solutions

Solution notebook

A First Look at the Kalman Filter

Contents

- *A First Look at the Kalman Filter*
 - *Overview*
 - *The Basic Idea*
 - *Convergence*
 - *Implementation*
 - *Exercises*
 - *Solutions*

Overview

This lecture provides a simple and intuitive introduction to the Kalman filter, for those who either

- have heard of the Kalman filter but don't know how it works, or
- know the Kalman filter equations, but don't know where they come from

For additional (more advanced) reading on the Kalman filter, see

- [\[LS12\]](#), section 2.7.
- [\[AM05\]](#)

The last reference gives a particularly clear and comprehensive treatment of the Kalman filter

Required knowledge: Familiarity with matrix manipulations, multivariate normal distributions, covariance matrices, etc.

The Basic Idea

The Kalman filter has many applications in economics, but for now let's pretend that we are rocket scientists

A missile has been launched from country Y and our mission is to track it

Let $x \in \mathbb{R}^2$ denote the current location of the missile—a pair indicating latitude-longitude coordinates on a map

At the present moment in time, the precise location x is unknown, but we do have some beliefs about x

One way to summarize our knowledge is a point prediction \hat{x}

- But what if the President wants to know the probability that the missile is currently over the Sea of Japan?
- Better to summarize our initial beliefs with a bivariate probability density p
 - $\int_E p(x)dx$ indicates the probability that we attach to the missile being in region E

The density p is called our *prior* for the random variable x

To keep things tractable, we will always assume that our prior is Gaussian. In particular, we take

$$p = N(\hat{x}, \Sigma) \quad (2.62)$$

where \hat{x} is the mean of the distribution and Σ is a 2×2 covariance matrix. In our simulations, we will suppose that

$$\hat{x} = \begin{pmatrix} 0.2 \\ -0.2 \end{pmatrix}, \quad \Sigma = \begin{pmatrix} 0.4 & 0.3 \\ 0.3 & 0.45 \end{pmatrix} \quad (2.63)$$

This density $p(x)$ is shown below as a contour map, with the center of the red ellipse being equal to \hat{x}

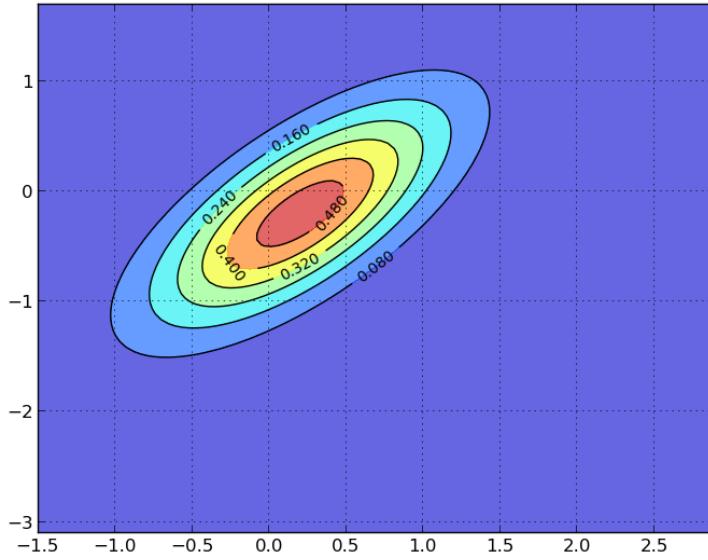


Fig. 2.1: Prior density (Click this or any other figure to enlarge.)

The Filtering Step We are now presented with some good news and some bad news

The good news is that the missile has been located by our sensors, which report that the current location is $y = (2.3, -1.9)$

The next figure shows the original prior $p(x)$ and the new reported location y

The bad news is that our sensors are imprecise.

In particular, we should interpret the output of our sensor not as $y = x$, but rather as

$$y = Gx + v, \quad \text{where} \quad v \sim N(0, R) \quad (2.64)$$

Here G and R are 2×2 matrices with R positive definite. Both are assumed known, and the noise term v is assumed to be independent of x

How then should we combine our prior $p(x) = N(\hat{x}, \Sigma)$ and this new information y to improve our understanding of the location of the missile?

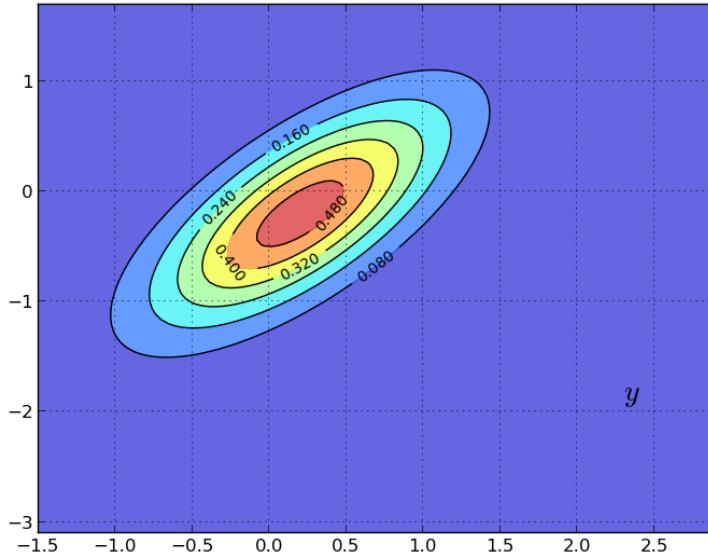
As you may have guessed, the answer is to use Bayes' theorem, which tells us we should update our prior $p(x)$ to $p(x | y)$ via

$$p(x | y) = \frac{p(y | x) p(x)}{p(y)}$$

where $p(y) = \int p(y | x) p(x) dx$

In solving for $p(x | y)$, we observe that

- $p(x) = N(\hat{x}, \Sigma)$
- In view of (2.64), the conditional density $p(y | x)$ is $N(Gx, R)$



- $p(y)$ does not depend on x , and enters into the calculations only as a normalizing constant

Because we are in a linear and Gaussian framework, the updated density can be computed by calculating population linear regressions.

In particular, the solution is known¹ to be

$$p(x|y) = N(\hat{x}^F, \Sigma^F)$$

where

$$\hat{x}^F := \hat{x} + \Sigma G' (G \Sigma G' + R)^{-1} (y - G\hat{x}) \quad \text{and} \quad \Sigma^F := \Sigma - \Sigma G' (G \Sigma G' + R)^{-1} G \Sigma \quad (2.65)$$

Here $\Sigma G' (G \Sigma G' + R)^{-1}$ is the matrix of population regression coefficients of the hidden object $x - \hat{x}$ on the surprise $y - G\hat{x}$

This new density $p(x|y) = N(\hat{x}^F, \Sigma^F)$ is shown in the next figure via contour lines and the color map

The original density is left in as contour lines for comparison

Our new density twists the prior $p(x)$ in a direction determined by the new information $y - G\hat{x}$

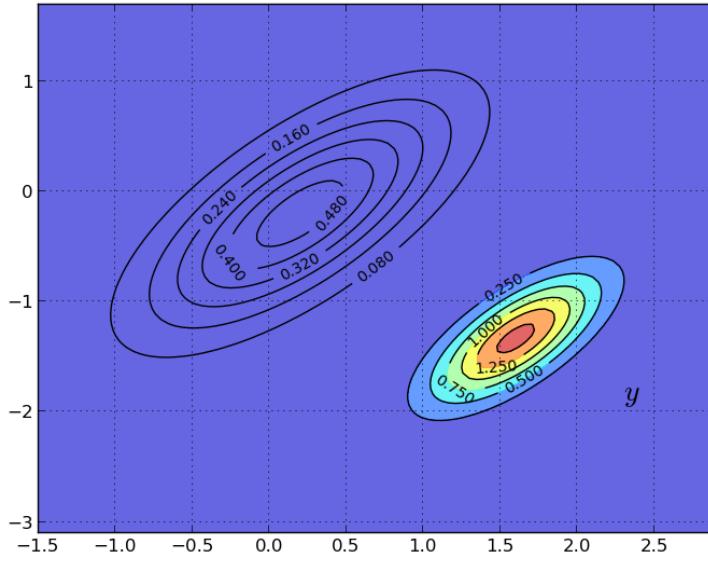
In generating the figure, we set G to the identity matrix and $R = 0.5\Sigma$ for Σ defined in (2.63)

(The code for generating this and the preceding figures can be found in the file `gaussian_contours.jl` from the `QuantEcon.applications` package

The Forecast Step

What have we achieved so far?

¹ See, for example, page 93 of [Bis06]. To get from his expressions to the ones used above, you will also need to apply the [Woodbury matrix identity](#).



We have obtained probabilities for the current location of the state (missile) given prior and current information

This is called “filtering” rather than forecasting, because we are filtering out noise rather than looking into the future

- $p(x|y) = N(\hat{x}^F, \Sigma^F)$ is called the *filtering distribution*

But now let’s suppose that we are given another task: To predict the location of the missile after one unit of time (whatever that may be) has elapsed

To do this we need a model of how the state evolves

Let’s suppose that we have one, and that it’s linear and Gaussian: In particular,

$$x_{t+1} = Ax_t + w_{t+1}, \quad \text{where } w_t \sim N(0, Q) \quad (2.66)$$

Our aim is to combine this law of motion and our current distribution $p(x|y) = N(\hat{x}^F, \Sigma^F)$ to come up with a new *predictive* distribution for the location one unit of time hence

In view of (2.66), all we have to do is introduce a random vector $x^F \sim N(\hat{x}^F, \Sigma^F)$ and work out the distribution of $Ax^F + w$ where w is independent of x^F and has distribution $N(0, Q)$

Since linear combinations of Gaussians are Gaussian, $Ax^F + w$ is Gaussian

Elementary calculations and the expressions in (2.65) tell us that

$$\mathbb{E}[Ax^F + w] = A\mathbb{E}x^F + \mathbb{E}w = A\hat{x}^F = A\hat{x} + A\Sigma G'(G\Sigma G' + R)^{-1}(y - G\hat{x})$$

and

$$\text{Var}[Ax^F + w] = A \text{Var}[x^F] A' + Q = A\Sigma^F A' + Q = A\Sigma A' - A\Sigma G'(G\Sigma G' + R)^{-1}G\Sigma A' + Q$$

The matrix $A\Sigma G'(G\Sigma G' + R)^{-1}$ is often written as K_Σ and called the *Kalman gain*

- the subscript Σ has been added to remind us that K_Σ depends on Σ , but not y or \hat{x}

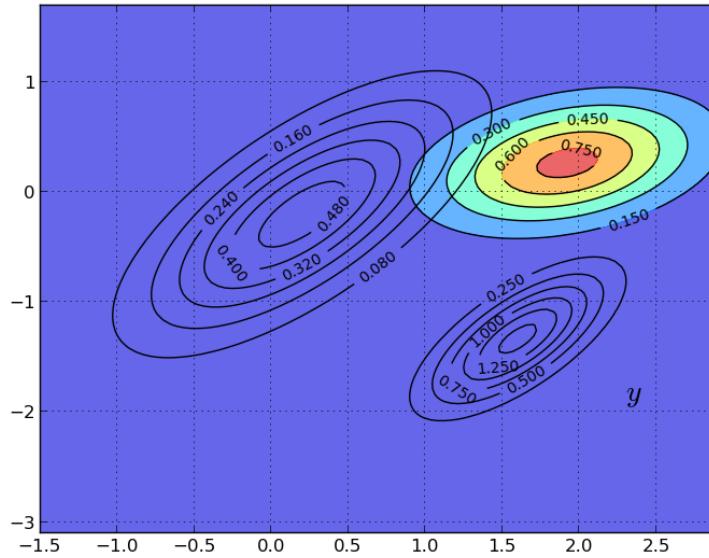
Using this notation, we can summarize our results as follows: Our updated prediction is the density $N(\hat{x}_{new}, \Sigma_{new})$ where

$$\begin{aligned}\hat{x}_{new} &:= A\hat{x} + K_\Sigma(y - G\hat{x}) \\ \Sigma_{new} &:= A\Sigma A' - K_\Sigma G\Sigma A' + Q\end{aligned}\tag{2.67}$$

- The density $p_{new}(x) = N(\hat{x}_{new}, \Sigma_{new})$ is called the *predictive distribution*

The predictive distribution is the new density shown in the following figure, where the update has used parameters

$$A = \begin{pmatrix} 1.2 & 0.0 \\ 0.0 & -0.2 \end{pmatrix}, \quad Q = 0.3 * \Sigma$$



The Recursive Procedure Let's look back at what we've done.

We started the current period with a prior $p(x)$ for the location x of the missile

We then used the current measurement y to update to $p(x | y)$

Finally, we used the law of motion (2.66) for $\{x_t\}$ to update to $p_{new}(x)$

If we now step into the next period, we are ready to go round again, taking $p_{new}(x)$ as the current prior

Swapping notation $p_t(x)$ for $p(x)$ and $p_{t+1}(x)$ for $p_{new}(x)$, the full recursive procedure is:

1. Start the current period with prior $p_t(x) = N(\hat{x}_t, \Sigma_t)$
2. Observe current measurement y_t

3. Compute the filtering distribution $p_t(x | y) = N(\hat{x}_t^F, \Sigma_t^F)$ from $p_t(x)$ and y_t , applying Bayes rule and the conditional distribution (2.64)
4. Compute the predictive distribution $p_{t+1}(x) = N(\hat{x}_{t+1}, \Sigma_{t+1})$ from the filtering distribution and (2.66)
5. Increment t by one and go to step 1

Repeating (2.67), the dynamics for \hat{x}_t and Σ_t are as follows

$$\begin{aligned}\hat{x}_{t+1} &= A\hat{x}_t + K_{\Sigma_t}(y_t - G\hat{x}_t) \\ \Sigma_{t+1} &= A\Sigma_t A' - K_{\Sigma_t}G\Sigma_t A' + Q\end{aligned}\tag{2.68}$$

These are the standard dynamic equations for the Kalman filter. See, for example, [LS12], page 58.

Convergence

The matrix Σ_t is a measure of the uncertainty of our prediction \hat{x}_t of x_t

Apart from special cases, this uncertainty will never be fully resolved, regardless of how much time elapses

One reason is that our prediction \hat{x}_t is made based on information available at $t - 1$, not t

Even if we know the precise value of x_{t-1} (which we don't), the transition equation (2.66) implies that $x_t = Ax_{t-1} + w_t$

Since the shock w_t is not observable at $t - 1$, any time $t - 1$ prediction of x_t will incur some error (unless w_t is degenerate)

However, it is certainly possible that Σ_t converges to a constant matrix as $t \rightarrow \infty$

To study this topic, let's expand the second equation in (2.68):

$$\Sigma_{t+1} = A\Sigma_t A' - A\Sigma_t G'(G\Sigma_t G' + R)^{-1}G\Sigma_t A' + Q\tag{2.69}$$

This is a nonlinear difference equation in Σ_t

A fixed point of (2.69) is a constant matrix Σ such that

$$\Sigma = A\Sigma A' - A\Sigma G'(G\Sigma G' + R)^{-1}G\Sigma A' + Q\tag{2.70}$$

Equation (2.69) is known as a discrete time Riccati difference equation

Equation (2.70) is known as a **discrete time algebraic Riccati equation**

Conditions under which a fixed point exists and the sequence $\{\Sigma_t\}$ converges to it are discussed in [AHMS96] and [AM05], chapter 4

One sufficient (but not necessary) condition is that all the eigenvalues λ_i of A satisfy $|\lambda_i| < 1$ (cf. e.g., [AM05], p. 77)

(This strong condition assures that the unconditional distribution of x_t converges as $t \rightarrow +\infty$)

In this case, for any initial choice of Σ_0 that is both nonnegative and symmetric, the sequence $\{\Sigma_t\}$ in (2.69) converges to a nonnegative symmetric matrix Σ that solves (2.70)

Implementation

The type `Kalman` from the `QuantEcon.jl` package implements the Kalman filter

- Instance data:
 - The parameters A, G, Q, R of a given model
 - the moments (\hat{x}_t, Σ_t) of the current prior
- The main methods are:
 - `prior_to_filtered`, which updates (\hat{x}_t, Σ_t) to $(\hat{x}_t^F, \Sigma_t^F)$
 - `filtered_to_forecast`, which updates the filtering distribution to the predictive distribution – which becomes the new prior $(\hat{x}_{t+1}, \Sigma_{t+1})$
 - `update`, which combines the last two methods
 - a `stationary_values`, which computes the solution to (2.70) and the corresponding (stationary) Kalman gain

You can view the program [on GitHub](#) but we repeat it here for convenience

```
#=
Implements the Kalman filter for a linear Gaussian state space model.

@author : Spencer Lyon <spencer.lyon@nyu.edu>

@date: 2014-07-29

References
-----

http://quant-econ.net/jl/kalman.html

TODO: Do docstrings here after implementing LinerStateSpace
=#

type Kalman
    A
    G
    Q
    R
    k
    n
    cur_x_hat
    cur_sigma
end

# Initializes current mean and cov to zeros
function Kalman(A, G, Q, R)
    k = size(G, 1)
    n = size(G, 2)
```

```

xhat = n == 1 ? zero(eltype(A)) : zeros(n)
Sigma = n == 1 ? zero(eltype(A)) : zeros(n, n)
return Kalman(A, G, Q, R, k, n, xhat, Sigma)
end

function set_state!(k::Kalman, x_hat, Sigma)
    k.cur_x_hat = x_hat
    k.cur_sigma = Sigma
    Void
end

"""
Updates the moments (cur_x_hat, cur_sigma) of the time t prior to the
time t filtering distribution, using current measurement y_t.
The updates are according to

$$\begin{aligned} \hat{x}_F &= \hat{x}_t + \Sigma G' (G \Sigma G' + R)^{-1} \\ &\quad (y - G \hat{x}_t) \\ \Sigma^F &= \Sigma - \Sigma G' (G \Sigma G' + R)^{-1} G \Sigma \end{aligned}$$

"""

#### Arguments
- `k::Kalman` An instance of the Kalman filter
- `y` The current measurement

function prior_to_filtered!(k::Kalman, y)
    # simplify notation
    G, R = k.G, k.R
    x_hat, Sigma = k.cur_x_hat, k.cur_sigma

    # and then update
    if k.k > 1
        reshape(y, k.k, 1)
    end
    A = Sigma * G'
    B = G * Sigma' * G' + R
    M = A * inv(B)
    k.cur_x_hat = x_hat + M * (y - G * x_hat)
    k.cur_sigma = Sigma - M * G * Sigma
    Void
end

"""
Updates the moments of the time t filtering distribution to the
moments of the predictive distribution, which becomes the time
t+1 prior

#### Arguments
- `k::Kalman` An instance of the Kalman filter
"""

```

```

function filtered_to_forecast!(k::Kalman)
    # simplify notation
    A, Q = k.A, k.Q
    x_hat, Sigma = k.cur_x_hat, k.cur_sigma

    # and then update
    k.cur_x_hat = A * x_hat
    k.cur_sigma = A * Sigma * A' + Q
    Void
end

"""
Updates cur_x_hat and cur_sigma given array `y` of length `k`. The full
update, from one period to the next

#### Arguments

- `k::Kalman` An instance of the Kalman filter
- `y` An array representing the current measurement

"""

function update!(k::Kalman, y)
    prior_to_filtered!(k, y)
    filtered_to_forecast!(k)
    Void
end

function stationary_values(k::Kalman)
    # simplify notation
    A, Q, G, R = k.A, k.Q, k.G, k.R

    # solve Riccati equation, obtain Kalman gain
    Sigma_inf = solve_discrete_riccati(A', G', Q, R)
    K_inf = A * Sigma_inf * G' * inv(G * Sigma_inf * G' + R)
    return Sigma_inf, K_inf
end

```

Exercises

Exercise 1 Consider the following simple application of the Kalman filter, loosely based on [LS12], section 2.9.2

Suppose that

- all variables are scalars
- the hidden state $\{x_t\}$ is in fact constant, equal to some $\theta \in \mathbb{R}$ unknown to the modeler

State dynamics are therefore given by (2.66) with $A = 1$, $Q = 0$ and $x_0 = \theta$

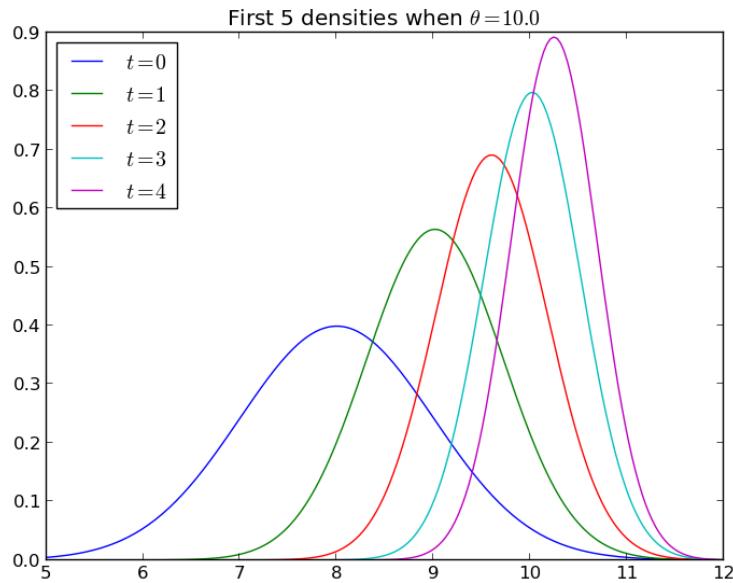
The measurement equation is $y_t = \theta + v_t$ where v_t is $N(0, 1)$ and iid

The task of this exercise is to simulate the model and, using the code from `kalman.jl`, plot the first five predictive densities $p_t(x) = N(\hat{x}_t, \Sigma_t)$

As shown in [LS12], sections 2.9.1–2.9.2, these distributions asymptotically put all mass on the unknown value θ

In the simulation, take $\theta = 10$, $\hat{x}_0 = 8$ and $\Sigma_0 = 1$

Your figure should – modulo randomness – look something like this



Exercise 2 The preceding figure gives some support to the idea that probability mass converges to θ

To get a better idea, choose a small $\epsilon > 0$ and calculate

$$z_t := 1 - \int_{\theta-\epsilon}^{\theta+\epsilon} p_t(x) dx$$

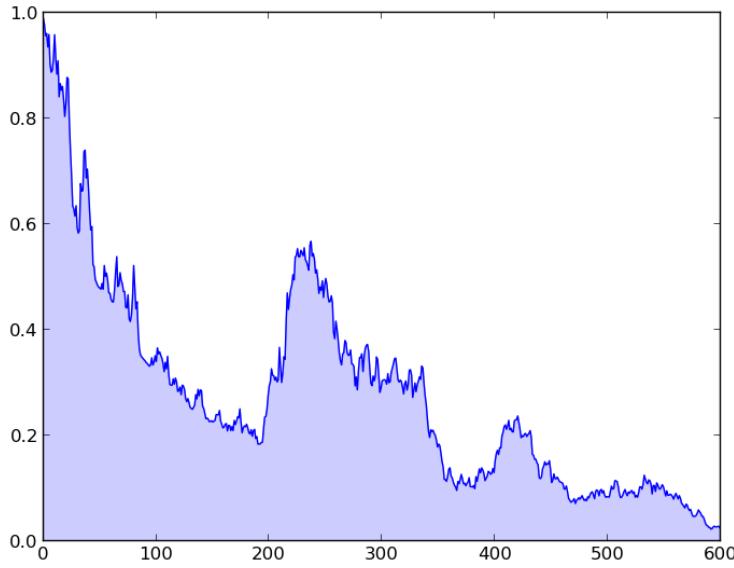
for $t = 0, 1, 2, \dots, T$

Plot z_t against T , setting $\epsilon = 0.1$ and $T = 600$

Your figure should show error erratically declining something like this

Exercise 3 As discussed above, if the shock sequence $\{w_t\}$ is not degenerate, then it is not in general possible to predict x_t without error at time $t - 1$ (and this would be the case even if we could observe x_{t-1})

Let's now compare the prediction \hat{x}_t made by the Kalman filter against a competitor who is allowed to observe x_{t-1}



This competitor will use the conditional expectation $\mathbb{E}[x_t | x_{t-1}]$, which in this case is Ax_{t-1}

The conditional expectation is known to be the optimal prediction method in terms of minimizing mean squared error

(More precisely, the minimizer of $\mathbb{E} \|x_t - g(x_{t-1})\|^2$ with respect to g is $g^*(x_{t-1}) := \mathbb{E}[x_t | x_{t-1}]$)

Thus we are comparing the Kalman filter against a competitor who has more information (in the sense of being able to observe the latent state) and behaves optimally in terms of minimizing squared error

Our horse race will be assessed in terms of squared error

In particular, your task is to generate a graph plotting observations of both $\|x_t - Ax_{t-1}\|^2$ and $\|x_t - \hat{x}_t\|^2$ against t for $t = 1, \dots, 50$

For the parameters, set $G = I$, $R = 0.5I$ and $Q = 0.3I$, where I is the 2×2 identity

Set

$$A = \begin{pmatrix} 0.5 & 0.4 \\ 0.6 & 0.3 \end{pmatrix}$$

To initialize the prior density, set

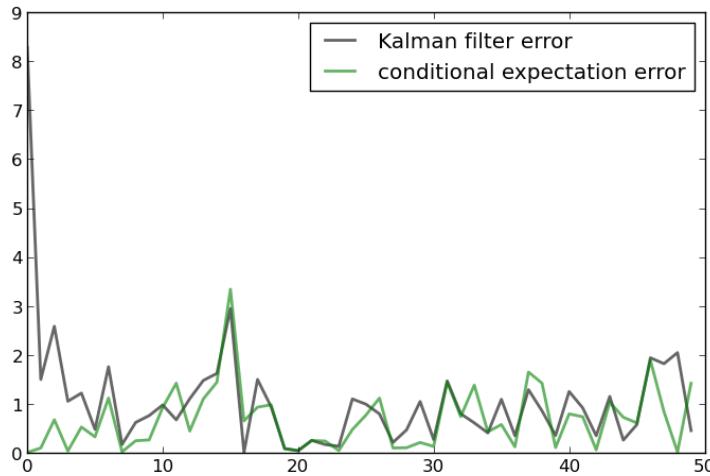
$$\Sigma_0 = \begin{pmatrix} 0.9 & 0.3 \\ 0.3 & 0.9 \end{pmatrix}$$

and $\hat{x}_0 = (8, 8)$

Finally, set $x_0 = (0, 0)$

You should end up with a figure similar to the following (modulo randomness)

Observe how, after an initial learning period, the Kalman filter performs quite well, even relative to the competitor who predicts optimally with knowledge of the latent state



Exercise 4 Try varying the coefficient 0.3 in $Q = 0.3I$ up and down

Observe how the diagonal values in the stationary solution Σ (see (2.70)) increase and decrease in line with this coefficient

The interpretation is that more randomness in the law of motion for x_t causes more (permanent) uncertainty in prediction

Solutions

[Solution notebook](#)

Uncertainty Traps

Overview

In this lecture we study a simplified version of an uncertainty traps model of Fajgelbaum, Schaal and Taschereau-Dumouchel [FSTD15]

The model features self-reinforcing uncertainty that has big impacts on economic activity

In the model,

- Fundamentals vary stochastically and are not fully observable
- At any moment there are both active and inactive entrepreneurs; only active entrepreneurs produce
- Agents – active and inactive entrepreneurs – have beliefs about the fundamentals expressed as probability distributions
- Greater uncertainty means greater dispersions of these distributions

- Entrepreneurs are risk averse and hence less inclined to be active when uncertainty is high
- The output of active entrepreneurs is observable, supplying a noisy signal that helps everyone inside the model infer fundamentals
- Entrepreneurs update their beliefs about fundamentals using Bayes' Law, implemented via [Kalman filtering](#)

Uncertainty traps emerge because:

- High uncertainty discourages entrepreneurs from becoming active
- A low level of participation – i.e., a smaller number of active entrepreneurs – diminishes the flow of information about fundamentals
- Less information translates to higher uncertainty, further discouraging entrepreneurs from choosing to be active, and so on

Uncertainty traps stem from a positive externality: high aggregate economic activity levels generates valuable information

The Model

The original model described in [\[FSTD15\]](#) has many interesting moving parts

Here we examine a simplified version that nonetheless captures many of the key ideas

Fundamentals The evolution of the fundamental process $\{\theta_t\}$ is given by

$$\theta_{t+1} = \rho\theta_t + \sigma_\theta w_{t+1}$$

where

- $\sigma_\theta > 0$ and $0 < \rho < 1$
- $\{w_t\}$ is IID and standard normal

The random variable θ_t is not observable at any time

Output There is a total \bar{M} of risk averse entrepreneurs

Output of the m -th entrepreneur, conditional on being active in the market at time t , is equal to

$$x_m = \theta + \epsilon_m \quad \text{where} \quad \epsilon_m \sim N(0, \gamma_x^{-1}) \quad (2.71)$$

Here the time subscript has been dropped to simplify notation

The inverse of the shock variance, γ_x , is called the shock's **precision**

The higher is the precision, the more informative x_m is about the fundamental

Output shocks are independent across time and firms

Information and Beliefs All entrepreneurs start with identical beliefs about θ_0

Signals are publicly observable and hence all agents have identical beliefs always

Dropping time subscripts, beliefs for current θ are represented by the normal distribution $N(\mu, \gamma^{-1})$

Here γ is the precision of beliefs; its inverse is the degree of uncertainty

These parameters are updated by Kalman filtering

Let

- $\mathbb{M} \subset \{1, \dots, \bar{M}\}$ denote the set of currently active firms
- $M := |\mathbb{M}|$ denote the number of currently active firms
- X be the average output $\frac{1}{M} \sum_{m \in \mathbb{M}} x_m$ of the active firms

With this notation and primes for next period values, we can write the updating of the mean and precision via

$$\mu' = \rho \frac{\gamma \mu + M \gamma_x X}{\gamma + M \gamma_x} \quad (2.72)$$

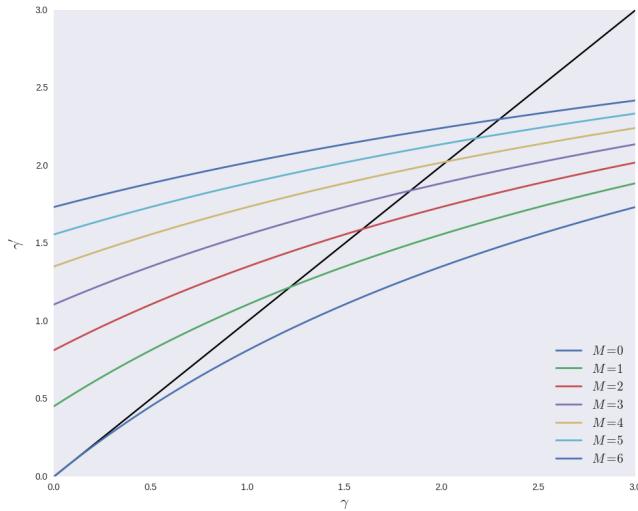
$$\gamma' = \left(\frac{\rho^2}{\gamma + M \gamma_x} + \sigma_\theta^2 \right)^{-1} \quad (2.73)$$

These are standard Kalman filtering results applied to the current setting

Exercise 1 provides more details on how (2.72) and (2.73) are derived, and then asks you to fill in remaining steps

The next figure plots the law of motion for the precision in (2.73) as a 45 degree diagram, with one curve for each $M \in \{0, \dots, 6\}$

The other parameter values are $\rho = 0.99, \gamma_x = 0.5, \sigma_\theta = 0.5$



Points where the curves hit the 45 degree lines are long run steady states for precision for different values of M

Thus, if one of these values for M remains fixed, a corresponding steady state is the equilibrium level of precision

- high values of M correspond to greater information about the fundamental, and hence more precision in steady state
- low values of M correspond to less information and more uncertainty in steady state

In practice, as we'll see, the number of active firms fluctuates stochastically

Participation Omitting time subscripts once more, entrepreneurs enter the market in the current period if

$$\mathbb{E}[u(x_m - F_m)] > c \quad (2.74)$$

Here

- the mathematical expectation of x_m is based on (2.71) and beliefs $N(\mu, \gamma^{-1})$ for θ
- F_m is a stochastic but previsible fixed cost, independent across time and firms
- c is a constant reflecting opportunity costs

The statement that F_m is previsible means that it is realized at the start of the period and treated as a constant in (2.74)

The utility function has the constant absolute risk aversion form

$$u(x) = \frac{1}{a} (1 - \exp(-ax)) \quad (2.75)$$

where a is a positive parameter

Combining (2.74) and (2.75), entrepreneur m participates in the market (or is said to be active) when

$$\frac{1}{a} \{1 - \mathbb{E}[\exp(-a(\theta + \epsilon_m - F_m))] \} > c$$

Using standard formulas for expectations of lognormal random variables, this is equivalent to the condition

$$\psi(\mu, \gamma, F_m) := \frac{1}{a} \left(1 - \exp \left(-a\mu + aF_m + \frac{a^2 \left(\frac{1}{\gamma} + \frac{1}{\gamma_x} \right)}{2} \right) \right) - c > 0 \quad (2.76)$$

Implementation

We want to simulate this economy

As a first step, let's put together a class that bundles

- the parameters, the current value of θ and the current values of the two belief parameters μ and γ

- methods to update θ , μ and γ , as well as to determine the number of active firms and their outputs

The updating methods follow the laws of motion for θ , μ and γ given above

The method to evaluate the number of active firms generates F_1, \dots, F_M and tests condition (2.76) for each firm

The function *UncertaintyTrapEcon* encodes as default values the parameters we'll use in the simulations below

Here's the code, which can also be obtained from [GitHub](#)

```

type UncertaintyTrapEcon
    a::Float64          # Risk aversion
    gx::Float64          # Production shock precision
    rho::Float64          # Correlation coefficient for theta
    sig_theta::Float64    # Std dev of theta shock
    num_firms::Int        # Number of firms
    sig_F::Float64        # Std dev of fixed costs
    c::Float64            # External opportunity cost
    mu::Float64           # Initial value for mu
    gamma::Float64         # Initial value for gamma
    theta::Float64         # Initial value for theta
    sd_x::Float64          # standard deviation of shock
end

function UncertaintyTrapEcon(;a::Real=1.5, gx::Real=0.5, rho::Real=0.99,
                           sig_theta::Real=0.5, num_firms::Int=100,
                           sig_F::Real=1.5, c::Real=-420, mu_init::Real=0,
                           gamma_init::Real=4, theta_init::Real=0)
    sd_x = sqrt(a/gx)
    UncertaintyTrapEcon(a, gx, rho, sig_theta, num_firms, sig_F, c, mu_init,
                        gamma_init, theta_init, sd_x)
end

function psi(uc::UncertaintyTrapEcon, F)
    temp1 = -uc.a * (uc.mu - F)
    temp2 = 0.5 * uc.a^2 * (1/uc.gamma + 1/uc.gx)
    return (1/uc.a) * (1 - exp(temp1 + temp2)) - uc.c
end

"""
Update beliefs (mu, gamma) based on aggregates X and M.
"""
function update_beliefs!(uc::UncertaintyTrapEcon, X, M)
    # Simplify names
    gx, rho, sig_theta = uc.gx, uc.rho, uc.sig_theta

    # Update mu
    temp1 = rho * (uc.gamma*uc.mu + M*gx*X)
    temp2 = uc.gamma + M*gx
    uc.mu = temp1 / temp2

```

```

# Update gamma
uc.gamma = 1 / (rho^2 / (uc.gamma + M * gx) + sig_theta^2)
end

update_theta!(uc::UncertaintyTrapEcon, w) =
    (uc.theta = uc.rho*uc.theta + uc.sig_theta*w)

"""
Generate aggregates based on current beliefs (mu, gamma). This
is a simulation step that depends on the draws for F.
"""

function gen_aggregates(uc::UncertaintyTrapEcon)
    F_vals = uc.sig_F * randn(uc.num_firms)

    M = sum(psi(uc, F_vals) .> 0)::Int # Counts number of active firms
    if M > 0
        x_vals = uc.theta + uc.sd_x * randn(M)
        X = mean(x_vals)
    else
        X = 0.0
    end
    return X, M
end

```

In the results below we use this code to simulate time series for the major variables

Results

Let's look first at the dynamics of μ , which the agents use to track θ



We see that μ tracks θ well when there are sufficient firms in the market

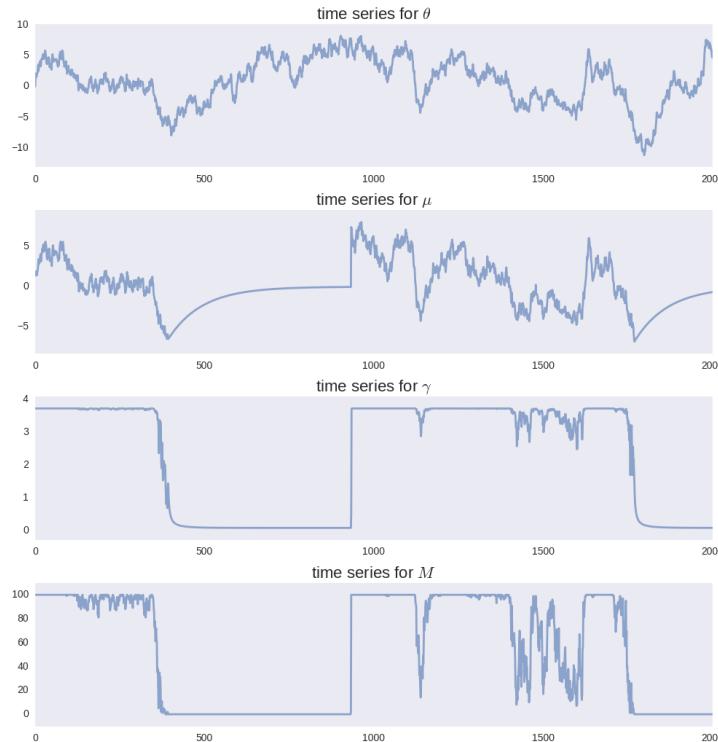
However, there are times when μ tracks θ poorly due to insufficient information

These are episodes where the uncertainty traps take hold

During these episodes

- precision is low and uncertainty is high
- few firms are in the market

To get a clearer idea of the dynamics, let's look at all the main time series at once, for a given set of shocks



Notice how the traps only take hold after a sequence of bad draws for the fundamental

Thus, the model gives us a *propagation mechanism* that maps bad random draws into long downturns in economic activity

Exercises

Exercise 1 Fill in the details behind (2.72) and (2.73) based on the following standard result (see, e.g., p. 24 of [YS05])

Fact Let $\mathbf{x} = (x_1, \dots, x_M)$ be a vector of IID draws from common distribution $N(\theta, 1/\gamma_x)$ and let \bar{x} be the sample mean. If γ_x is known and the prior for θ is $N(\mu, 1/\gamma)$, then the posterior distribution of θ given \mathbf{x} is

$$\pi(\theta | \mathbf{x}) = N(\mu_0, 1/\gamma_0)$$

where

$$\mu_0 = \frac{\mu\gamma + M\bar{x}\gamma_x}{\gamma + M\gamma_x} \quad \text{and} \quad \gamma_0 = \gamma + M\gamma_x$$

Exercise 2 Modulo randomness, replicate the simulation figures shown above

- Use the parameter values listed as defaults in the function *UncertaintyTrapEcon*

Solutions

[Solution notebook](#)

A Simple Optimal Growth Model

Contents

- *A Simple Optimal Growth Model*
 - [Overview](#)
 - [The Model](#)
 - [Dynamic Programming](#)
 - [Computation](#)
 - [Writing Reusable Code](#)
 - [Exercises](#)
 - [Solutions](#)

Overview

In this lecture we're going to study a simple optimal growth model with one agent

The model is a version of the standard one sector infinite horizon growth model studied in

- [\[SLP89\]](#), chapter 2
- [\[LS12\]](#), section 3.1
- [EDTC](#), chapter 1
- [\[Sun96\]](#), chapter 12

The model is intentionally simplistic — for now we favor ease of exposition over realism

The technique we use to solve the model is dynamic programming, which is

- pervasive in economics, finance and many other fields
- general and powerful, yielding both intuition and practical computational methods

Our treatment of dynamic programming follows on from earlier, less technical treatments in our lectures on [shortest paths](#) and [job search](#)

The Model

Consider an agent who owns at time t capital stock $k_t \in \mathbb{R}_+ := [0, \infty)$ and produces output

$$y_t := f(k_t) \in \mathbb{R}_+$$

This output can either be consumed or saved as capital for next period

For simplicity we assume that depreciation is total, so that next period capital is just output minus consumption:

$$k_{t+1} = y_t - c_t \quad (2.77)$$

Taking k_0 as given, we suppose that the agent wishes to maximize

$$\sum_{t=0}^{\infty} \beta^t u(c_t) \quad (2.78)$$

where u is a given utility function and $\beta \in (0, 1)$ is a discount factor

More precisely, the agent wishes to select a path c_0, c_1, c_2, \dots for consumption that is

1. nonnegative
2. feasible, in the sense that the capital path $\{k_t\}$ determined by $\{c_t\}$, k_0 and (2.77) is always nonnegative
3. optimal, in the sense that it maximizes (2.78) relative to all other feasible consumption sequences

In the present context

- k is called the *state* variable — it summarizes the “state of the world” at the start of each period
- c is called the *control* variable — a value chosen by the agent each period after observing the state

A well-known result from dynamic programming theory (cf., e.g., [SLP89], section 4.1) states that, for kind of this problem, any optimal consumption sequence $\{c_t\}$ is *Markov*

That is, there exists a function σ such that

$$c_t = \sigma(k_t) \quad \text{for all } t$$

In other words, the current control is a fixed (i.e., time homogeneous) function of the current state

The Policy Function Approach As it turns out, we are better off seeking the function σ directly, rather than the optimal consumption sequence

The main reason is that the functional approach — seeking the optimal policy — translates directly over to the stochastic case, whereas the sequential approach does not

For this model, we will say that function σ mapping \mathbb{R}_+ into \mathbb{R}_+ is a *feasible consumption policy* if it satisfies

$$\sigma(k) \leq f(k) \quad \text{for all } k \in \mathbb{R}_+ \quad (2.79)$$

The set of all such policies will be denoted by Σ

Using this notation, the agent's decision problem can be rewritten as

$$\max_{\sigma \in \Sigma} \left\{ \sum_{t=0}^{\infty} \beta^t u(\sigma(k_t)) \right\} \quad (2.80)$$

where the sequence $\{k_t\}$ in (2.80) is given by

$$k_{t+1} = f(k_t) - \sigma(k_t), \quad k_0 \text{ given} \quad (2.81)$$

In the next section we discuss how to solve this problem for the maximizing σ

Dynamic Programming

We will solve for the optimal policy using [dynamic programming](#)

The first step is to define the *policy value function* v_σ associated with a given policy σ , which is

$$v_\sigma(k_0) := \sum_{t=0}^{\infty} \beta^t u(\sigma(k_t)) \quad (2.82)$$

when $\{k_t\}$ is given by (2.81)

Evidently $v_\sigma(k_0)$ is the total present value of discounted utility associated with following policy σ forever, given initial capital k_0

The *value function* for this optimization problem is then defined as

$$v^*(k_0) := \sup_{\sigma \in \Sigma} v_\sigma(k_0) \quad (2.83)$$

The value function gives the maximal value that can be obtained from state k_0 , after considering all feasible policies

A policy $\sigma \in \Sigma$ is called *optimal* if it attains the supremum in (2.83) for all $k_0 \in \mathbb{R}_+$

The *Bellman equation* for this problem takes the form

$$v^*(k) = \max_{0 \leq c \leq f(k)} \{u(c) + \beta v^*(f(k) - c)\} \quad \text{for all } k \in \mathbb{R}_+ \quad (2.84)$$

It states that maximal value from a given state can be obtained by trading off

- current reward from a given action (in this case utility from current consumption) vs
- the discounted future value of the state resulting from that action

(If the intuition behind the Bellman equation is not clear to you, try working through *this lecture*)

As a matter of notation, given a continuous function w on \mathbb{R}_+ , we say that policy $\sigma \in \Sigma$ is *w-greedy* if $\sigma(k)$ is a solution to

$$\max_{0 \leq c \leq f(k)} \{u(c) + \beta w(f(k) - c)\} \quad (2.85)$$

for every $k \in \mathbb{R}_+$

Theoretical Results As with most optimization problems, conditions for existence of a solution typically require some form of continuity and compactness

In addition, some restrictions are needed to ensure that the sum of discounted utility is always finite

For example, if we are prepared to assume that f and u are continuous and u is bounded, then

1. The value function v^* is finite, bounded, continuous and satisfies the Bellman equation
2. At least one optimal policy exists
3. A policy is optimal if and only if it is v^* -greedy

(For a proof see, for example, proposition 10.1.13 of [EDTC](#))

In view of these results, to find an optimal policy, one option — perhaps the most common — is to

1. compute v^*
2. solve for a v^* -greedy policy

The advantage is that, once we get to the second step, we are solving a one-dimensional optimization problem — the problem on the right-hand side of (2.84)

This is much easier than an infinite-dimensional optimization problem, which is what we started out with

(An infinite sequence $\{c_t\}$ is a point in an infinite-dimensional space)

In fact step 2 is almost trivial once v^* is obtained

For this reason, most of our focus is on the first step — how to obtain the value function

Value Function Iteration The value function v^* can be obtained by an iterative technique:

- Start with a guess — some initial function w
- successively improve it

The improvement step involves applying an operator (i.e., a map from functions to functions) called the *Bellman operator*

The Bellman operator for this problem is denoted T and sends w into Tw via

$$Tw(k) := \max_{0 \leq c \leq f(k)} \{u(c) + \beta w(f(k) - c)\} \quad (2.86)$$

Now let w be any continuous bounded function

Iteratively applying T from initial condition w produces a sequence of functions $w, Tw, T(Tw) = T^2w, \dots$ that converges uniformly to v^*

For a proof see, for example, lemma 10.1.20 of [EDTC](#)

This convergence will be prominent in our numerical experiments

Unbounded Utility The theoretical results stated above assume that the utility function is bounded

In practice economists often work with unbounded utility functions

For utility functions that are bounded below (but possibly unbounded above), a clean and comprehensive theory now exists

(Section 12.2 of [EDTC](#) provides one exposition)

For utility functions that are unbounded both below and above the situation is more complicated

For recent work on deterministic problems, see, for example, [\[Kam12\]](#) or [\[MdRV10\]](#)

In this lecture we will use both bounded and unbounded utility functions without dwelling on the theory

Computation

Let's now look at computing the value function and the optimal policy

Fitted Value Iteration The first step is to compute the value function by iterating with the Bellman operator

In theory, the algorithm is as follows

1. Begin with a function w — an initial condition
2. Solving [\(2.86\)](#), obtain the function Tw
3. Unless some stopping condition is satisfied, set $w = Tw$ and go to step 2

However, there is a problem we must confront before we implement this procedure: The iterates can neither be calculated exactly nor stored on a computer

To see the issue, consider [\(2.86\)](#)

Even if w is a known function, unless Tw can be shown to have some special structure, the only way to store this function is to record the value $Tw(k)$ for every $k \in \mathbb{R}_+$

Clearly this is impossible

What we will do instead is use *fitted value function iteration*

The procedure is to record the value of the function Tw at only finitely many “grid” points $\{k_1, \dots, k_I\} \subset \mathbb{R}_+$, and reconstruct it from this information when required

More precisely, the algorithm will be

1. Begin with an array of values $\{w_1, \dots, w_I\}$, typically representing the values of some initial function w on the grid points $\{k_1, \dots, k_I\}$
2. build a function \hat{w} on the state space \mathbb{R}_+ by interpolating the points $\{w_1, \dots, w_I\}$
3. By repeatedly solving [\(2.86\)](#), obtain and record the value $T\hat{w}(k_i)$ on each grid point k_i

4. Unless some stopping condition is satisfied, set $\{w_1, \dots, w_I\} = \{T\hat{w}(k_1), \dots, T\hat{w}(k_I)\}$ and go to step 2

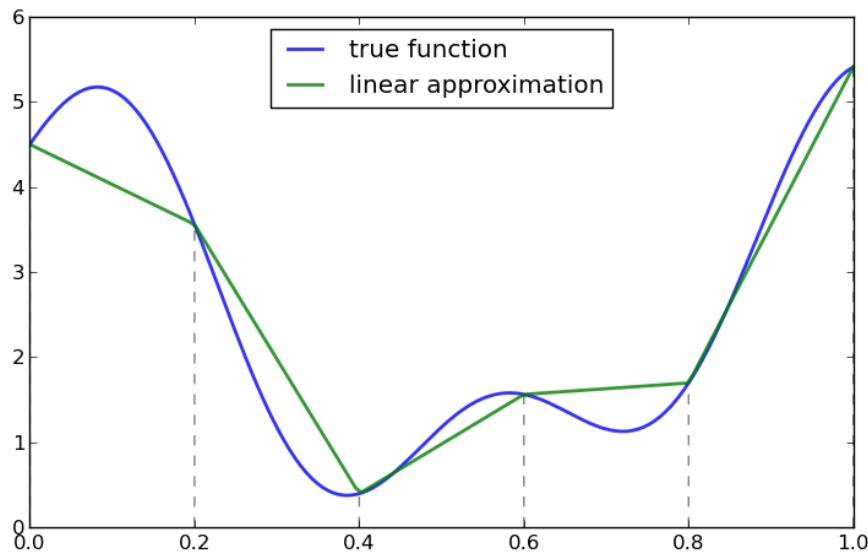
How should we go about step 2?

This is a problem of function approximation, and there are many ways to approach it

What's important here is that the function approximation scheme must not only produce a good approximation to $T\hat{w}$, but also combine well with the broader iteration algorithm described above

One good choice from both respects is continuous piecewise linear interpolation (see this paper for further discussion)

The next figure illustrates piecewise linear interpolation of an arbitrary function on grid points $0, 0.2, 0.4, \dots, 1$



Another advantage of piecewise linear interpolation is that it preserves useful shape properties such as monotonicity and concavity / convexity

A First Pass Implementation Let's now look at an implementation of fitted value function iteration using Julia

In the example below,

- $f(k) = k^\alpha$ with $\alpha = 0.65$
- $u(c) = \ln c$ and $\beta = 0.95$

As is well-known (see [LS12], section 3.1.2), for this particular problem an exact analytical solution is available, with

$$v^*(k) = c_1 + c_2 \ln k \quad (2.87)$$

for

$$c_1 := \frac{\ln(1 - \alpha\beta)}{1 - \beta} + \frac{\ln(\alpha\beta)\alpha\beta}{(1 - \alpha\beta)(1 - \beta)} \quad \text{and} \quad c_2 := \frac{\alpha}{1 - \alpha\beta}$$

At this stage, our only aim is to see if we can replicate this solution numerically, using fitted value function iteration

Here's a first-pass solution, the details of which are explained *below*

The code can be found in file `optgrowth/optgrowth_v0.jl` from the `QuantEcon.applications` repository

We repeat it here for convenience

```
#=
A first pass at solving the optimal growth problem via value function
iteration. A more general version is provided in optgrowth.py.

@author : Spencer Lyon <spencer.lyon@nyu.edu>
          Victoria Gregory <victoria.gregory@nyu.edu>
=#

using Optim: optimize
using Interpolations
using Plots
pyplot()

## Primitives and grid
alpha = 0.65
bet = 0.95
grid_max = 2
grid_size = 150
grid = 1e-6:(grid_max-1e-6)/(grid_size-1):grid_max

## Exact solution
ab = alpha * bet
c1 = (log(1 - ab) + log(ab) * ab / (1 - ab)) / (1 - bet)
c2 = alpha / (1 - ab)
v_star(k) = c1 .+ c2 .* log(k)

function bellman_operator(grid, w)
    Aw = scale(interpolate(w, BSpline(Linear()), OnGrid()), grid)

    Tw = zeros(w)

    for (i, k) in enumerate(grid)
        objective(c) = - log(c) - bet * Aw[k^alpha - c]
        res = optimize(objective, 1e-6, k^alpha)
        Tw[i] = - objective(res.minimum)
    end
    return Tw
end
```

```

function main(n::Int=35)
    w_init = 5 .* log(grid) .- 25 # An initial condition -- fairly arbitrary
    w = copy(w_init)

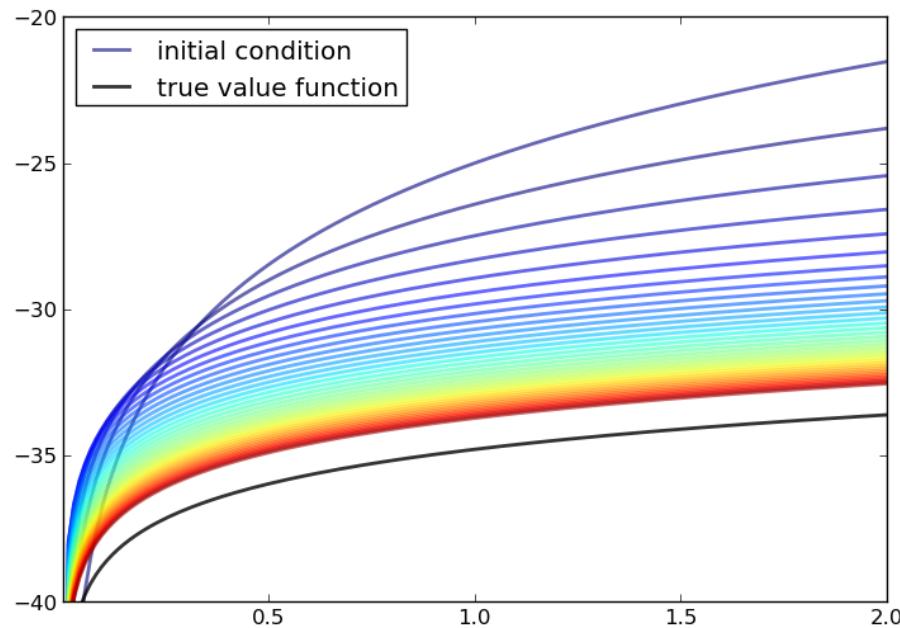
    ws = []
    colors = []
    for i=1:n
        w = bellman_operator(grid, w)
        push!(ws, w)
        push!(colors, RGBA(0, 0, 0, i/n))
    end

    p = plot(grid, w_init, color=:green, linewidth=2, alpha=0.6,
              label="initial condition")
    plot!(grid, ws, color=colors', label="", linewidth=2)
    plot!(grid, v_star(grid), color=:blue, linewidth=2, alpha=0.8,
          label="true value function")
    plot!(ylims=(-40, -20), xlims=(minimum(grid), maximum(grid)))

    return p
end

```

Running the code produces the following figure



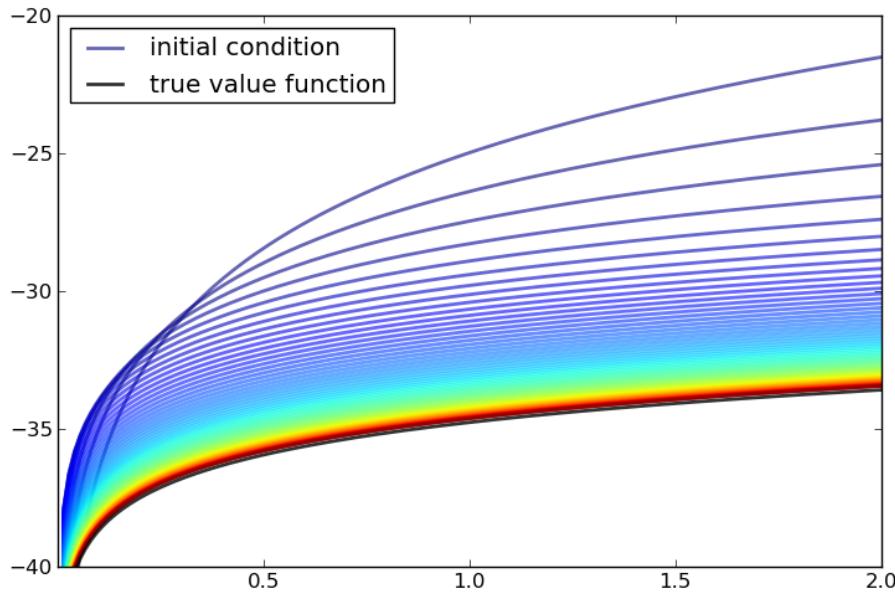
The curves in this picture represent

1. the first 36 functions generated by the fitted value function iteration algorithm described above, with hotter colors given to higher iterates

2. the true value function as specified in (2.87), drawn in black

The sequence of iterates converges towards v^*

If we increase n and run again we see further improvement — the next figure shows $n = 75$



Incidentally, it is true that knowledge of the functional form of v^* for this model has influenced our choice of the initial condition

```
w = 5 * log(grid) - 25
```

In more realistic problems such information is not available, and convergence will probably take longer

Comments on the Code The function `bellman_operator` implements steps 2–3 of the fitted value function algorithm discussed *above*

Linear interpolation is performed by the `getindex` method on the `CoordInterpGrid` from `Grid.jl`

The numerical solver `optimize` from `Optim.jl` minimizes its objective, so we use the identity $\max_x f(x) = -\min_x -f(x)$ to solve (2.86)

Notice that we wrap the code used to generate the figure in a function named `main`. This allows us to import the functionality of `optgrowth_v0.jl` into a Julia session, without necessarily generating the figure.

The Policy Function To compute an approximate optimal policy, we run the *fitted value function algorithm* until approximate convergence

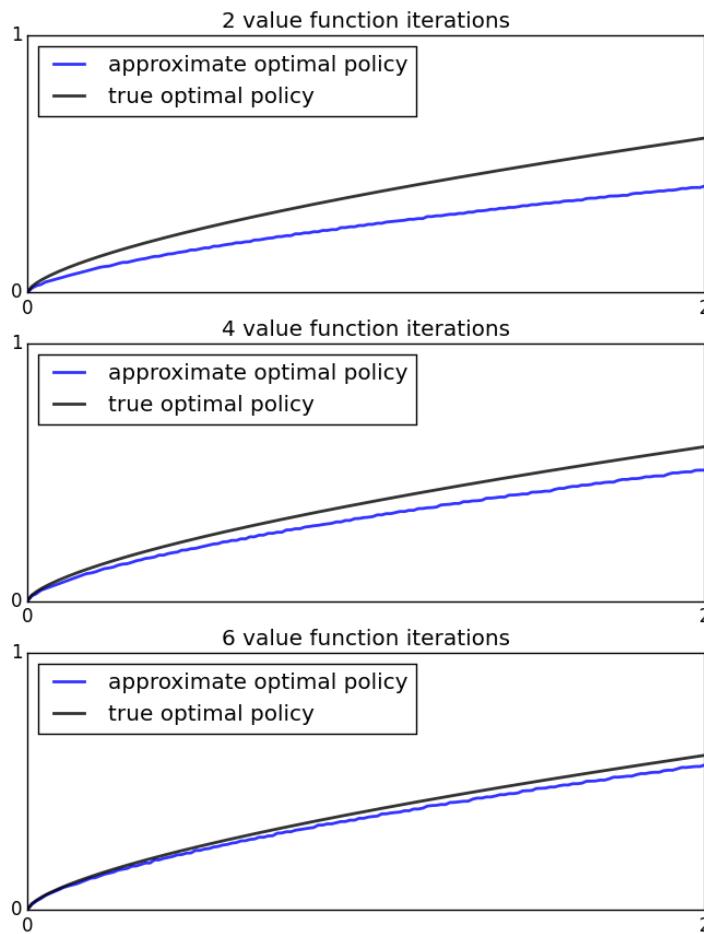
Taking the function so produced as an approximation to v^* , we then compute the (approximate) v^* -greedy policy

For this particular problem, the optimal consumption policy has the known analytical solution $\sigma(k) = (1 - \alpha\beta)k^\alpha$

The next figure compares the numerical solution to this exact solution

In the three figures, the approximation to v^* is obtained by running the loop in the *fitted value function algorithm* 2, 4 and 6 times respectively

Even with as few as 6 iterates, the numerical result is quite close to the true policy



Exercise 1 asks you to reproduce this figure

Writing Reusable Code

The title of this section might sound uninteresting and a departure from our topic, but it's equally important if not more so

It's understandable that many economists never consider the basic principles of software development, preoccupied as they are with the applied aspects of trying to implement their projects

However, in programming as in many things, success tends to find those who focus on what is important, not just what is urgent

The Danger of Copy and Paste For computing the value function of the particular growth model studied above, the code we have already written (in file `optgrowth_v0.jl`, shown *here*) is perfectly fine

However, suppose that we now want to solve a different growth model, with different technology and preferences

Probably we want to keep our existing code, so let's follow our first instinct and

1. copy the contents of `optgrowth_v0.jl` to a new file
2. then make the necessary changes

Now let's suppose that we repeat this process again and again over several years, so we now have many similar files

(And perhaps we're doing similar things with other projects, leading to hundreds of specialized and partially related Julia files lying around our file system)

There are several potential problems here

Problem 1 First, if we now realize we've been making some small but fundamental error with our dynamic programming all this time, we have to modify all these different files

And now we realize that we don't quite remember which files they were, and where exactly we put them...

So we fix all the ones we can find — spending a few hours in the process, since each implementation is slightly different and takes time to remember — and leave the ones we can't

Now, 6 weeks later, we need to use one of these files

But is file X one that we have fixed, or is it not?

In this way, our code base becomes a mess, with related functionality scattered across many files, and errors creeping into them

Problem 2 A second issue here is that since all these files are specialized and might not be used again, there's little incentive to invest in writing them cleanly and efficiently

DRY The preceding discussion leads us to one of the most fundamental principles of code development: **don't repeat yourself**

To the extent that it's practical,

- always strive to write code that is abstract and generic in order to facilitate reuse
- try to ensure that each distinct logical concept is repeated in your code base as few times as possible

To this end, we are now going to rewrite our solution to the optimal growth problem given in `optgrowth_v0.jl` (shown *above*) with the intention of producing a more generic version

While some aspects of this exercise might seem like overkill, the principles are important, and easy to illustrate in the context of the current problem

Implementation 2 In writing our second implementation, we want our function `bellman_operator` to be able to handle a wider class of models

In particular, we don't want model specifics hardwired into this function

Instead, we would like `bellman_operator` to act in conjunction with a more general description of a model (technology, preferences, etc.)

To do so it's convenient to wrap the model description up in a type and add the Bellman operator as a method

This idea is implemented in the code below, in file `optgrowth.jl` from the `QuantEcon.applications` repo

```
#=
Solving the optimal growth problem via value function iteration.

@author : Spencer Lyon <spencer.lyon@nyu.edu>

@date : 2014-07-05

References
-----

Simple port of the file quantecon.models.optgrowth

http://quant-econ.net/jl/dp_intro.html
=#

#=

This type defines the primitives representing the growth model. The
default values are

f(k) = k**alpha, i.e., Cobb-Douglas production function
u(c) = ln(c), i.e., log utility

See the constructor below for details
=#
using Interpolations
```

```

using Optim

"""
Neoclassical growth model

##### Fields

- `f::Function` : Production function
- `bet::Real` : Discount factor in (0, 1)
- `u::Function` : Utility function
- `grid_max::Int` : Maximum for grid over savings values
- `grid_size::Int` : Number of points in grid for savings values
- `grid::LinSpace{Float64}` : The grid for savings values

"""

type GrowthModel
    f::Function
    bet::Float64
    u::Function
    grid_max::Int
    grid_size::Int
    grid::LinSpace{Float64}
end

default_f(k) = k^0.65
default_u(c) = log(c)

"""
Constructor of `GrowthModel`

##### Arguments

- `f::Function(k->k^0.65)` : Production function
- `bet::Real(0.95)` : Discount factor in (0, 1)
- `u::Function(log)` : Utility function
- `grid_max::Int(2)` : Maximum for grid over savings values
- `grid_size::Int(150)` : Number of points in grid for savings values

"""

function GrowthModel(f=default_f, bet=0.95, u=default_u, grid_max=2,
                      grid_size=150)
    grid = linspace(1e-6, grid_max, grid_size)
    return GrowthModel(f, bet, u, grid_max, grid_size, grid)
end

"""
Apply the Bellman operator for a given model and initial value.

##### Arguments

- `g::GrowthModel` : Instance of `GrowthModel`
- `w::Vector` : Current guess for the value function

```

```

- `out::Vector` : Storage for output.
- `;ret_policy::Bool(false)` : Toggles return of value or policy functions

##### Returns

None, `out` is updated in place. If `ret_policy == true` out is filled with the
policy function, otherwise the value function is stored in `out`.

"""

function bellman_operator!(g::GrowthModel, w::Vector, out::Vector;
                           ret_policy::Bool=false)
    # Apply linear interpolation to w
    Aw = scale(interpolate(w, BSpline(Linear()), OnGrid()), g.grid)

    for (i, k) in enumerate(g.grid)
        objective(c) = - g.u(c) - g.bet * Aw[g.f(k) - c]
        res = optimize(objective, 1e-6, g.f(k))
        c_star = res.minimum

        if ret_policy
            # set the policy equal to the optimal c
            out[i] = c_star
        else
            # set Tw[i] equal to max_c { u(c) + beta w(f(k_i) - c)}
            out[i] = - objective(c_star)
        end
    end

    return out
end

function bellman_operator(g::GrowthModel, w::Vector;
                           ret_policy::Bool=false)
    out = similar(w)
    bellman_operator!(g, w, out, ret_policy=ret_policy)
end

"""

Extract the greedy policy (policy function) of the model.

##### Arguments

- `g::GrowthModel` : Instance of `GrowthModel`
- `w::Vector` : Current guess for the value function
- `out::Vector` : Storage for output

##### Returns

None, `out` is updated in place to hold the policy function

"""

function get_greedy!(g::GrowthModel, w::Vector, out::Vector)
    bellman_operator!(g, w, out, ret_policy=true)

```

```
end

get_greedy(g::GrowthModel, w::Vector) = bellman_operator(g, w, ret_policy=true)
```

Of course we could omit the type structure and just pass data to `bellman_operator` and `compute_greedy` as a list of separate arguments

For example

```
Tw = bellman_operator(f, beta, u, grid_max, grid_size, w)
```

This approach is also fine, and many prefer it

However, we find that the type structure is more convenient and a bit less error prone because once an instance is created we can call the methods repeatedly without having to specify a lot of arguments

Iteration The next thing we need to do is implement iteration of the Bellman operator

Since iteratively applying an operator is something we'll do a lot of, let's write this as generic, reusable code

Our code is written in the file `compute_fp.jl` from the [main repository](#), and displayed below

```
#=

Compute the fixed point of a given operator T, starting from
specified initial condition v.

@author : Spencer Lyon <spencer.lyon@nyu.edu>

@date: 2014-07-05

References
-----

http://quant-econ.net/jl/dp_intro.html
=#

"""

Repeatedly apply a function to search for a fixed point

Approximates `T^∞ v`, where `T` is an operator (function) and `v` is an initial
guess for the fixed point. Will terminate either when `T^{k+1}(v) - T^k v <
err_tol` or `max_iter` iterations has been exceeded.

Provided that `T` is a contraction mapping or similar, the return value will
be an approximation to the fixed point of `T`.

##### Arguments

* `T`: A function representing the operator `T`
* `v::TV`: The initial condition. An object of type `TV`
```

```

* `;err_tol(1e-3)` : Stopping tolerance for iterations
* `;max_iter(50)` : Maximum number of iterations
* `;verbose(2)` : Level of feedback (0 for no output, 1 for warnings only, 2
    for warning and convergence messages during iteration)
* `;print_skip(10)` : if `verbose == 2` , how many iterations to apply between
    print messages

##### Returns
---

* `::TV` : The fixed point of the operator `T` . Has type `TV`


##### Example

```julia
using QuantEcon
T(x, μ) = 4.0 * μ * x * (1.0 - x)
x_star = compute_fixed_point(x->T(x, 0.3), 0.4) # (4μ - 1)/(4μ)
```

"""

function compute_fixed_point{TV}(T::Function,
                                v::TV;
                                err_tol=1e-4,
                                max_iter=100,
                                verbose=2,
                                print_skip=10)

    if !(verbose in (0, 1, 2))
        throw(ArgumentError("verbose should be 0, 1 or 2"))
    end

    iterate = 0
    err = err_tol + 1
    while iterate < max_iter && err > err_tol
        new_v = T(v)::TV
        iterate += 1
        err = Base.maxabs(new_v - v)
        if verbose == 2
            if iterate % print_skip == 0
                println("Compute iterate $iterate with error $err")
            end
        end
        v = new_v
    end

    if verbose >= 1
        if iterate == max_iter
            warn("max_iter attained in compute_fixed_point")
        elseif verbose == 2
            println("Converged in $iterate steps")
        end
    end
end

```

```

    return v
end

```

As currently written, the code continues iteration until one of two stopping conditions holds

1. Successive iterates become sufficiently close together, in the sense that the maximum deviation between them falls below `error_tol`
2. The number of iterations exceeds `max_iter`

Examples of usage for all the code above can be found in the solutions to the exercises

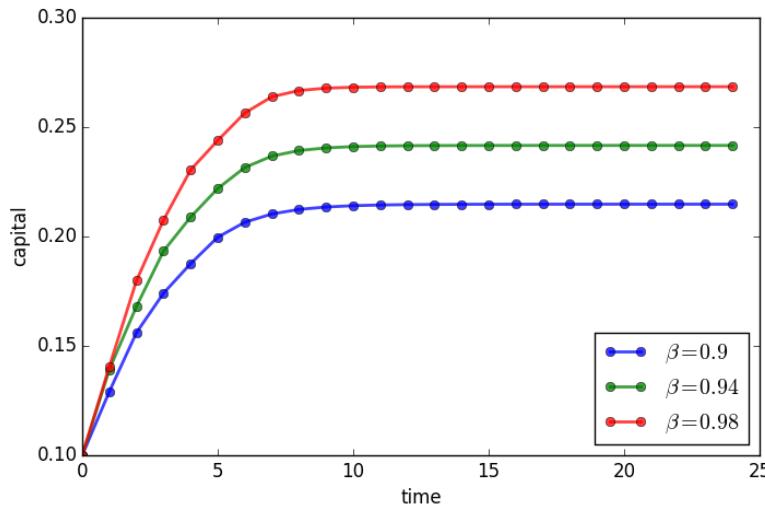
Exercises

Exercise 1 Replicate the optimal policy figure *shown above*

Use the same parameters and initial condition found in `optgrowth_v0.jl`

Exercise 2 Once an optimal consumption policy σ is given, the dynamics for the capital stock follows (2.81)

The next figure shows the first 25 elements of this sequence for three different discount factors (and hence three different policies)



In each sequence, the initial condition is $k_0 = 0.1$

The discount factors are `discount_factors = (0.9, 0.94, 0.98)`

Otherwise, the parameters and primitives are the same as found in `optgrowth_v0.jl`

Replicate the figure

Solutions

[Solution notebook](#)

LQ Dynamic Programming Problems

Contents

- *LQ Dynamic Programming Problems*
 - [Overview](#)
 - [Introduction](#)
 - [Optimality – Finite Horizon](#)
 - [Extensions and Comments](#)
 - [Implementation](#)
 - [Further Applications](#)
 - [Exercises](#)
 - [Solutions](#)

Overview

Linear quadratic (LQ) control refers to a class of dynamic optimization problems that have found applications in almost every scientific field

This lecture provides an introduction to LQ control and its economic applications

As we will see, LQ systems have a simple structure that makes them an excellent workhorse for a wide variety of economic problems

Moreover, while the linear-quadratic structure is restrictive, it is in fact far more flexible than it may appear initially

These themes appear repeatedly below

Mathematically, LQ control problems are closely related to [the Kalman filter](#), although we won't pursue the deeper connections in this lecture

In reading what follows, it will be useful to have some familiarity with

- matrix manipulations
- vectors of random variables
- dynamic programming and the Bellman equation (see for example [this lecture](#) and [this lecture](#))

For additional reading on LQ control, see, for example,

- [\[LS12\]](#), chapter 5
- [\[HS08\]](#), chapter 4

- [HLL96], section 3.5

In order to focus on computation, we leave longer proofs to these sources (while trying to provide as much intuition as possible)

Introduction

The “linear” part of LQ is a linear law of motion for the state, while the “quadratic” part refers to preferences

Let’s begin with the former, move on to the latter, and then put them together into an optimization problem

The Law of Motion Let x_t be a vector describing the state of some economic system

Suppose that x_t follows a linear law of motion given by

$$x_{t+1} = Ax_t + Bu_t + Cw_{t+1}, \quad t = 0, 1, 2, \dots \quad (2.88)$$

Here

- u_t is a “control” vector, incorporating choices available to a decision maker confronting the current state x_t
- $\{w_t\}$ is an uncorrelated zero mean shock process satisfying $\mathbb{E}w_tw_t' = I$, where the right-hand side is the identity matrix

Regarding the dimensions

- x_t is $n \times 1$, A is $n \times n$
- u_t is $k \times 1$, B is $n \times k$
- w_t is $j \times 1$, C is $n \times j$

Example 1 Consider a household budget constraint given by

$$a_{t+1} + c_t = (1 + r)a_t + y_t$$

Here a_t is assets, r is a fixed interest rate, c_t is current consumption, and y_t is current non-financial income

If we suppose that $\{y_t\}$ is serially uncorrelated and $N(0, \sigma^2)$, then, taking $\{w_t\}$ to be standard normal, we can write the system as

$$a_{t+1} = (1 + r)a_t - c_t + \sigma w_{t+1}$$

This is clearly a special case of (2.88), with assets being the state and consumption being the control

Example 2 One unrealistic feature of the previous model is that non-financial income has a zero mean and is often negative

This can easily be overcome by adding a sufficiently large mean

Hence in this example we take $y_t = \sigma w_{t+1} + \mu$ for some positive real number μ

Another alteration that's useful to introduce (we'll see why soon) is to change the control variable from consumption to the deviation of consumption from some "ideal" quantity \bar{c}

(Most parameterizations will be such that \bar{c} is large relative to the amount of consumption that is attainable in each period, and hence the household wants to increase consumption)

For this reason, we now take our control to be $u_t := c_t - \bar{c}$

In terms of these variables, the budget constraint $a_{t+1} = (1+r)a_t - c_t + y_t$ becomes

$$a_{t+1} = (1+r)a_t - u_t - \bar{c} + \sigma w_{t+1} + \mu \quad (2.89)$$

How can we write this new system in the form of equation (2.88)?

If, as in the previous example, we take a_t as the state, then we run into a problem: the law of motion contains some constant terms on the right-hand side

This means that we are dealing with an *affine* function, not a linear one (recall *this discussion*)

Fortunately, we can easily circumvent this problem by adding an extra state variable

In particular, if we write

$$\begin{pmatrix} a_{t+1} \\ 1 \end{pmatrix} = \begin{pmatrix} 1+r & -\bar{c} + \mu \\ 0 & 1 \end{pmatrix} \begin{pmatrix} a_t \\ 1 \end{pmatrix} + \begin{pmatrix} -1 \\ 0 \end{pmatrix} u_t + \begin{pmatrix} \sigma \\ 0 \end{pmatrix} w_{t+1} \quad (2.90)$$

then the first row is equivalent to (2.89)

Moreover, the model is now linear, and can be written in the form of (2.88) by setting

$$x_t := \begin{pmatrix} a_t \\ 1 \end{pmatrix}, \quad A := \begin{pmatrix} 1+r & -\bar{c} + \mu \\ 0 & 1 \end{pmatrix}, \quad B := \begin{pmatrix} -1 \\ 0 \end{pmatrix}, \quad C := \begin{pmatrix} \sigma \\ 0 \end{pmatrix} \quad (2.91)$$

In effect, we've bought ourselves linearity by adding another state

Preferences In the LQ model, the aim is to minimize a flow of losses, where time- t loss is given by the quadratic expression

$$x_t' R x_t + u_t' Q u_t \quad (2.92)$$

Here

- R is assumed to be $n \times n$, symmetric and nonnegative definite
- Q is assumed to be $k \times k$, symmetric and positive definite

Note: In fact, for many economic problems, the definiteness conditions on R and Q can be relaxed. It is sufficient that certain submatrices of R and Q be nonnegative definite. See [HS08] for details

Example 1 A very simple example that satisfies these assumptions is to take R and Q to be identity matrices, so that current loss is

$$x_t' I x_t + u_t' I u_t = \|x_t\|^2 + \|u_t\|^2$$

Thus, for both the state and the control, loss is measured as squared distance from the origin

(In fact the general case (2.92) can also be understood in this way, but with R and Q identifying other – non-Euclidean – notions of “distance” from the zero vector)

Intuitively, we can often think of the state x_t as representing deviation from a target, such as

- deviation of inflation from some target level
- deviation of a firm’s capital stock from some desired quantity

The aim is to put the state close to the target, while using controls parsimoniously

Example 2 In the household problem *studied above*, setting $R = 0$ and $Q = 1$ yields preferences

$$x_t' R x_t + u_t' Q u_t = u_t^2 = (c_t - \bar{c})^2$$

Under this specification, the household’s current loss is the squared deviation of consumption from the ideal level \bar{c}

Optimality – Finite Horizon

Let’s now be precise about the optimization problem we wish to consider, and look at how to solve it

The Objective We will begin with the finite horizon case, with terminal time $T \in \mathbb{N}$

In this case, the aim is to choose a sequence of controls $\{u_0, \dots, u_{T-1}\}$ to minimize the objective

$$\mathbb{E} \left\{ \sum_{t=0}^{T-1} \beta^t (x_t' R x_t + u_t' Q u_t) + \beta^T x_T' R_f x_T \right\} \quad (2.93)$$

subject to the law of motion (2.88) and initial state x_0

The new objects introduced here are β and the matrix R_f

The scalar β is the discount factor, while $x' R_f x$ gives terminal loss associated with state x

Comments:

- We assume R_f to be $n \times n$, symmetric and nonnegative definite
- We allow $\beta = 1$, and hence include the undiscounted case
- x_0 may itself be random, in which case we require it to be independent of the shock sequence w_1, \dots, w_T

Information There's one constraint we've neglected to mention so far, which is that the decision maker who solves this LQ problem knows only the present and the past, not the future

To clarify this point, consider the sequence of controls $\{u_0, \dots, u_{T-1}\}$

When choosing these controls, the decision maker is permitted to take into account the effects of the shocks $\{w_1, \dots, w_T\}$ on the system

However, it is typically assumed — and will be assumed here — that the time- t control u_t can be made with knowledge of past and present shocks only

The fancy [measure-theoretic](#) way of saying this is that u_t must be measurable with respect to the σ -algebra generated by $x_0, w_1, w_2, \dots, w_t$

This is in fact equivalent to stating that u_t can be written in the form $u_t = g_t(x_0, w_1, w_2, \dots, w_t)$ for some Borel measurable function g_t

(Just about every function that's useful for applications is Borel measurable, so, for the purposes of intuition, you can read that last phrase as "for some function g_t ")

Now note that x_t will ultimately depend on the realizations of $x_0, w_1, w_2, \dots, w_t$

In fact it turns out that x_t summarizes all the information about these historical shocks that the decision maker needs to set controls optimally

More precisely, it can be shown that any optimal control u_t can always be written as a function of the current state alone

Hence in what follows we restrict attention to control policies (i.e., functions) of the form $u_t = g_t(x_t)$

Actually, the preceding discussion applies to all standard dynamic programming problems

What's special about the LQ case is that — as we shall soon see — the optimal u_t turns out to be a linear function of x_t

Solution To solve the finite horizon LQ problem we can use a dynamic programming strategy based on backwards induction that is conceptually similar to the approach adopted in [this lecture](#)

For reasons that will soon become clear, we first introduce the notation $J_T(x) = x' R_f x$

Now consider the problem of the decision maker in the second to last period

In particular, let the time be $T - 1$, and suppose that the state is x_{T-1}

The decision maker must trade off current and (discounted) final losses, and hence solves

$$\min_u \{x'_{T-1} Rx_{T-1} + u' Qu + \beta \mathbb{E} J_T(Ax_{T-1} + Bu + Cw_T)\}$$

At this stage, it is convenient to define the function

$$J_{T-1}(x) = \min_u \{x' Rx + u' Qu + \beta \mathbb{E} J_T(Ax + Bu + Cw_T)\} \quad (2.94)$$

The function J_{T-1} will be called the $T - 1$ value function, and $J_{T-1}(x)$ can be thought of as representing total "loss-to-go" from state x at time $T - 1$ when the decision maker behaves optimally

Now let's step back to $T - 2$

For a decision maker at $T - 2$, the value $J_{T-1}(x)$ plays a role analogous to that played by the terminal loss $J_T(x) = x' R_f x$ for the decision maker at $T - 1$

That is, $J_{T-1}(x)$ summarizes the future loss associated with moving to state x

The decision maker chooses her control u to trade off current loss against future loss, where

- the next period state is $x_{T-1} = Ax_{T-2} + Bu + Cw_{T-1}$, and hence depends on the choice of current control
- the “cost” of landing in state x_{T-1} is $J_{T-1}(x_{T-1})$

Her problem is therefore

$$\min_u \{x'_{T-2} Rx_{T-2} + u' Qu + \beta \mathbb{E} J_{T-1}(Ax_{T-2} + Bu + Cw_{T-1})\}$$

Letting

$$J_{T-2}(x) = \min_u \{x' Rx + u' Qu + \beta \mathbb{E} J_{T-1}(Ax + Bu + Cw_{T-1})\}$$

the pattern for backwards induction is now clear

In particular, we define a sequence of value functions $\{J_0, \dots, J_T\}$ via

$$J_{t-1}(x) = \min_u \{x' Rx + u' Qu + \beta \mathbb{E} J_t(Ax + Bu + Cw_t)\} \quad \text{and} \quad J_T(x) = x' R_f x$$

The first equality is the Bellman equation from dynamic programming theory specialized to the finite horizon LQ problem

Now that we have $\{J_0, \dots, J_T\}$, we can obtain the optimal controls

As a first step, let's find out what the value functions look like

It turns out that every J_t has the form $J_t(x) = x' P_t x + d_t$ where P_t is a $n \times n$ matrix and d_t is a constant

We can show this by induction, starting from $P_T := R_f$ and $d_T = 0$

Using this notation, (2.94) becomes

$$J_{T-1}(x) = \min_u \{x' Rx + u' Qu + \beta \mathbb{E} (Ax + Bu + Cw_T)' P_T (Ax + Bu + Cw_T)\} \quad (2.95)$$

To obtain the minimizer, we can take the derivative of the r.h.s. with respect to u and set it equal to zero

Applying the relevant rules of *matrix calculus*, this gives

$$u = -(Q + \beta B' P_T B)^{-1} \beta B' P_T A x \quad (2.96)$$

Plugging this back into (2.95) and rearranging yields

$$J_{T-1}(x) = x' P_{T-1} x + d_{T-1}$$

where

$$P_{T-1} = R - \beta^2 A' P_T B (Q + \beta B' P_T B)^{-1} B' P_T A + \beta A' P_T A \quad (2.97)$$

and

$$d_{T-1} := \beta \operatorname{trace}(C' P_T C) \quad (2.98)$$

(The algebra is a good exercise — we'll leave it up to you)

If we continue working backwards in this manner, it soon becomes clear that $J_t(x) = x' P_t x + d_t$ as claimed, where $\{P_t\}$ and $\{d_t\}$ satisfy the recursions

$$P_{t-1} = R - \beta^2 A' P_t B (Q + \beta B' P_t B)^{-1} B' P_t A + \beta A' P_t A \quad \text{with} \quad P_T = R_f \quad (2.99)$$

and

$$d_{t-1} = \beta(d_t + \operatorname{trace}(C' P_t C)) \quad \text{with} \quad d_T = 0 \quad (2.100)$$

Recalling (2.96), the minimizers from these backward steps are

$$u_t = -F_t x_t \quad \text{where} \quad F_t := (Q + \beta B' P_{t+1} B)^{-1} \beta B' P_{t+1} A \quad (2.101)$$

These are the linear optimal control policies we *discussed above*

In particular, the sequence of controls given by (2.101) and (2.88) solves our finite horizon LQ problem

Rephrasing this more precisely, the sequence u_0, \dots, u_{T-1} given by

$$u_t = -F_t x_t \quad \text{with} \quad x_{t+1} = (A - BF_t)x_t + Cw_{t+1} \quad (2.102)$$

for $t = 0, \dots, T-1$ attains the minimum of (2.93) subject to our constraints

An Application Early Keynesian models assumed that households have a constant marginal propensity to consume from current income

Data contradicted the constancy of the marginal propensity to consume

In response, Milton Friedman, Franco Modigliani and others built models based on a consumer's preference for an intertemporally smooth consumption stream

(See, for example, [Fri56] or [MB54])

One property of those models is that households purchase and sell financial assets to make consumption streams smoother than income streams

The household savings problem *outlined above* captures these ideas

The optimization problem for the household is to choose a consumption sequence in order to minimize

$$\mathbb{E} \left\{ \sum_{t=0}^{T-1} \beta^t (c_t - \bar{c})^2 + \beta^T q a_T^2 \right\} \quad (2.103)$$

subject to the sequence of budget constraints $a_{t+1} = (1+r)a_t - c_t + y_t$, $t \geq 0$

Here q is a large positive constant, the role of which is to induce the consumer to target zero debt at the end of her life

(Without such a constraint, the optimal choice is to choose $c_t = \bar{c}$ in each period, letting assets adjust accordingly)

As before we set $y_t = \sigma w_{t+1} + \mu$ and $u_t := c_t - \bar{c}$, after which the constraint can be written as in (2.89)

We saw how this constraint could be manipulated into the LQ formulation $x_{t+1} = Ax_t + Bu_t + Cw_{t+1}$ by setting $x_t = (a_t \ 1)'$ and using the definitions in (2.91)

To match with this state and control, the objective function (2.103) can be written in the form of (2.93) by choosing

$$Q := 1, \quad R := \begin{pmatrix} 0 & 0 \\ 0 & 0 \end{pmatrix}, \quad \text{and} \quad R_f := \begin{pmatrix} q & 0 \\ 0 & 0 \end{pmatrix}$$

Now that the problem is expressed in LQ form, we can proceed to the solution by applying (2.99) and (2.101)

After generating shocks w_1, \dots, w_T , the dynamics for assets and consumption can be simulated via (2.102)

We provide code for all these operations below

The following figure was computed using this code, with $r = 0.05, \beta = 1/(1+r), \bar{c} = 2, \mu = 1, \sigma = 0.25, T = 45$ and $q = 10^6$

The shocks $\{w_t\}$ were taken to be iid and standard normal

The top panel shows the time path of consumption c_t and income y_t in the simulation

As anticipated by the discussion on consumption smoothing, the time path of consumption is much smoother than that for income

(But note that consumption becomes more irregular towards the end of life, when the zero final asset requirement impinges more on consumption choices)

The second panel in the figure shows that the time path of assets a_t is closely correlated with cumulative unanticipated income, where the latter is defined as

$$z_t := \sum_{j=0}^t \sigma w_j$$

A key message is that unanticipated windfall gains are saved rather than consumed, while unanticipated negative shocks are met by reducing assets

(Again, this relationship breaks down towards the end of life due to the zero final asset requirement)

These results are relatively robust to changes in parameters

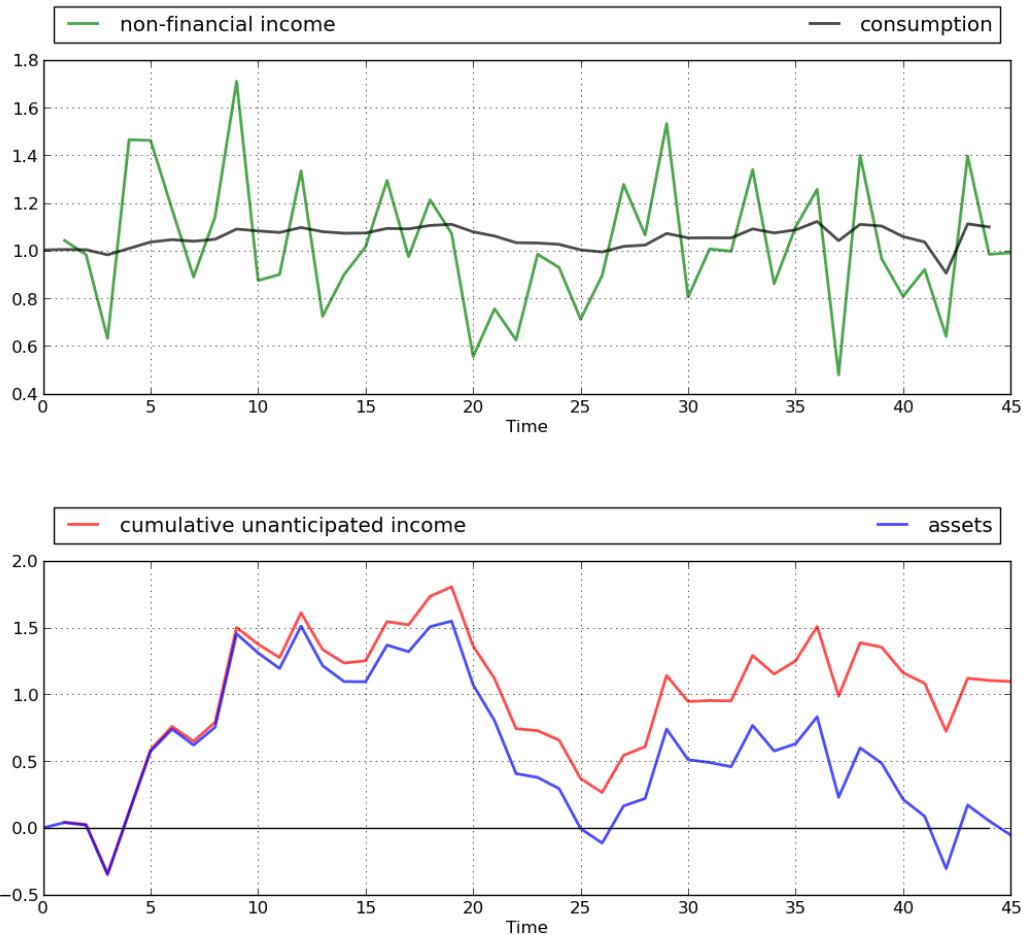
For example, let's increase β from $1/(1+r) \approx 0.952$ to 0.96 while keeping other parameters fixed

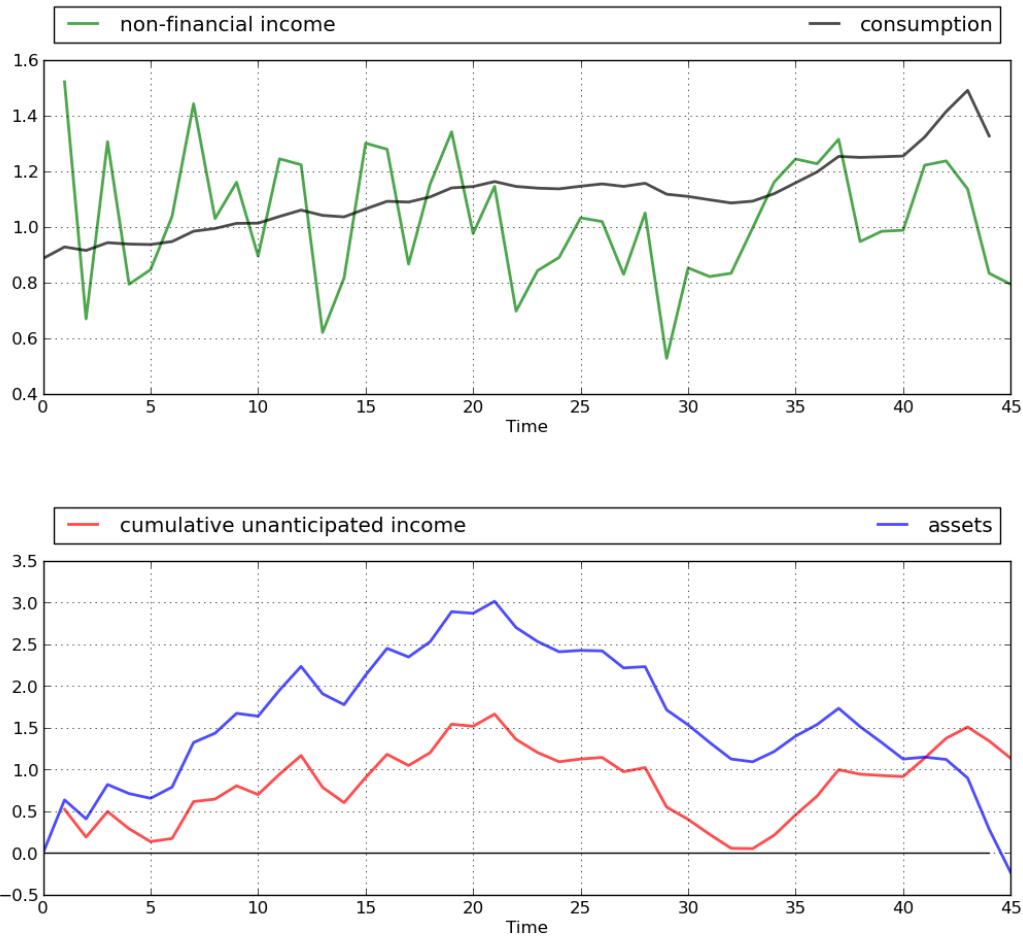
This consumer is slightly more patient than the last one, and hence puts relatively more weight on later consumption values

A simulation is shown below

We now have a slowly rising consumption stream and a hump-shaped build up of assets in the middle periods to fund rising consumption

However, the essential features are the same: consumption is smooth relative to income, and assets are strongly positively correlated with cumulative unanticipated income





Extensions and Comments

Let's now consider a number of standard extensions to the LQ problem treated above

Time-Varying Parameters In some settings it can be desirable to allow A, B, C, R and Q to depend on t

For the sake of simplicity, we've chosen not to treat this extension in our implementation given below

However, the loss of generality is not as large as you might first imagine

In fact, we can tackle many models with time-varying parameters by suitable choice of state variables

One illustration is given *below*

For further examples and a more systematic treatment, see [HS13], section 2.4

Adding a Cross-Product Term In some LQ problems, preferences include a cross-product term $u_t' N x_t$, so that the objective function becomes

$$\mathbb{E} \left\{ \sum_{t=0}^{T-1} \beta^t (x_t' R x_t + u_t' Q u_t + 2u_t' N x_t) + \beta^T x_T' R_f x_T \right\} \quad (2.104)$$

Our results extend to this case in a straightforward way

The sequence $\{P_t\}$ from (2.99) becomes

$$P_{t-1} = R - (\beta B' P_t A + N)' (Q + \beta B' P_t B)^{-1} (\beta B' P_t A + N) + \beta A' P_t A \quad \text{with} \quad P_T = R_f \quad (2.105)$$

The policies in (2.101) are modified to

$$u_t = -F_t x_t \quad \text{where} \quad F_t := (Q + \beta B' P_{t+1} B)^{-1} (\beta B' P_{t+1} A + N) \quad (2.106)$$

The sequence $\{d_t\}$ is unchanged from (2.100)

We leave interested readers to confirm these results (the calculations are long but not overly difficult)

Infinite Horizon Finally, we consider the infinite horizon case, with *cross-product term*, unchanged dynamics and objective function given by

$$\mathbb{E} \left\{ \sum_{t=0}^{\infty} \beta^t (x_t' R x_t + u_t' Q u_t + 2u_t' N x_t) \right\} \quad (2.107)$$

In the infinite horizon case, optimal policies can depend on time only if time itself is a component of the state vector x_t

In other words, there exists a fixed matrix F such that $u_t = -F x_t$ for all t

That decision rules are constant over time is intuitive — after all, the decision maker faces the same infinite horizon at every stage, with only the current state changing

Not surprisingly, P and d are also constant

The stationary matrix P is the solution to the discrete time algebraic Riccati equation

$$P = R - (\beta B'PA + N)'(Q + \beta B'PB)^{-1}(\beta B'PA + N) + \beta A'PA \quad (2.108)$$

Equation (2.108) is also called the *LQ Bellman equation*, and the map that sends a given P into the right-hand side of (2.108) is called the *LQ Bellman operator*

The stationary optimal policy for this model is

$$u = -Fx \quad \text{where} \quad F = (Q + \beta B'PB)^{-1}(\beta B'PA + N) \quad (2.109)$$

The sequence $\{d_t\}$ from (2.100) is replaced by the constant value

$$d := \text{trace}(C'PC) \frac{\beta}{1 - \beta} \quad (2.110)$$

The state evolves according to the time-homogeneous process $x_{t+1} = (A - BF)x_t + Cw_{t+1}$

An example infinite horizon problem is treated *below*

Certainty Equivalence Linear quadratic control problems of the class discussed above have the property of *certainty equivalence*

By this we mean that the optimal policy F is not affected by the parameters in C , which specify the shock process

This can be confirmed by inspecting (2.109) or (2.106)

It follows that we can ignore uncertainty when solving for optimal behavior, and plug it back in when examining optimal state dynamics

Implementation

We have put together some code for solving finite and infinite horizon linear quadratic control problems

The code can be found in the file `lqcontrol.jl` from the `QuantEcon.jl` package

You can view the program [on GitHub](#) but we repeat it here for convenience

```
#=
Provides a type called LQ for solving linear quadratic control
problems.

@author : Spencer Lyon <spencer.lyon@nyu.edu>
@author : Zac Cranko <zaccranko@gmail.com>

@date : 2014-07-05
```

References

```
http://quant-econ.net/jl/lqcontrol.html
=#
"""
Linear quadratic optimal control of either infinite or finite horizon
```

The infinite horizon problem can be written

```
min E sum_{t=0}^{infty} beta^t r(x_t, u_t)
```

with

```
r(x_t, u_t) := x_t' R x_t + u_t' Q u_t + 2 u_t' N x_t
```

The finite horizon form is

```
min E sum_{t=0}^{T-1} beta^t r(x_t, u_t) + beta^T x_T' R_f x_T
```

Both are minimized subject to the law of motion

```
x_{t+1} = A x_t + B u_t + C w_{t+1}
```

Here x is $n \times 1$, u is $k \times 1$, w is $j \times 1$ and the matrices are conformable for these dimensions. The sequence $\{w_t\}$ is assumed to be white noise, with zero mean and $E w_t w_t' = I$, the $j \times j$ identity.

For this model, the time t value (i.e., cost-to-go) function V_t takes the form

```
x' P_T x + d_T
```

and the optimal policy is of the form $u_T = -F_T x_T$. In the infinite horizon case, V , P , d and F are all stationary.

Fields

- `Q::ScalarOrArray` : $k \times k$ payoff coefficient for control variable u . Must be symmetric and nonnegative definite
- `R::ScalarOrArray` : $n \times n$ payoff coefficient matrix for state variable x . Must be symmetric and nonnegative definite
- `A::ScalarOrArray` : $n \times n$ coefficient on state in state transition
- `B::ScalarOrArray` : $n \times k$ coefficient on control in state transition
- `C::ScalarOrArray` : $n \times j$ coefficient on random shock in state transition
- `N::ScalarOrArray` : $k \times n$ cross product in payoff equation
- `bet::Real` : Discount factor in $[0, 1]$
- `capT::Union{Int, Void}` : Terminal period in finite horizon problem
- `rf::ScalarOrArray` : $n \times n$ terminal payoff in finite horizon problem. Must be symmetric and nonnegative definite
- `P::ScalarOrArray` : $n \times n$ matrix in value function representation $V(x) = x'Px + d$

```

- `d::Real` : Constant in value function representation
- `F::ScalarOrArray` : Policy rule that specifies optimal control in each period

"""

type LQ
    Q::ScalarOrArray
    R::ScalarOrArray
    A::ScalarOrArray
    B::ScalarOrArray
    C::ScalarOrArray
    N::ScalarOrArray
    bet::Real
    capT::Union{Int, Void} # terminal period
    rf::ScalarOrArray
    P::ScalarOrArray
    d::Real
    F::ScalarOrArray # policy rule
end

"""

Main constructor for LQ type

Specifies default arguments for all fields not part of the payoff function or
transition equation.

##### Arguments

- `Q::ScalarOrArray` : k x k payoff coefficient for control variable u. Must be
symmetric and nonnegative definite
- `R::ScalarOrArray` : n x n payoff coefficient matrix for state variable x.
Must be symmetric and nonnegative definite
- `A::ScalarOrArray` : n x n coefficient on state in state transition
- `B::ScalarOrArray` : n x k coefficient on control in state transition
- `C::ScalarOrArray(zeros(size(R, 1)))` : n x j coefficient on random shock in
state transition
- `N::ScalarOrArray(zeros(size(B, 1), size(A, 2)))` : k x n cross product in
payoff equation
- `bet::Real(1.0)` : Discount factor in [0, 1]
- `capT::Union{Int, Void}(Void)` : Terminal period in finite horizon
problem
- `rf::ScalarOrArray(fill(NaN, size(R)...))` : n x n terminal payoff in finite
horizon problem. Must be symmetric and nonnegative definite.

"""

function LQ(Q::ScalarOrArray,
            R::ScalarOrArray,
            A::ScalarOrArray,
            B::ScalarOrArray,
            C::ScalarOrArray=zeros(size(R, 1)),
            N::ScalarOrArray=zero(B'A);
            bet::ScalarOrArray=1.0,

```

```

capT::Union{Int,Void}=nothing,
rf::ScalarOrArray=fill(NaN, size(R)...))

k = size(Q, 1)
n = size(R, 1)
F = k==n==1 ? zero(Float64) : zeros(Float64, k, n)
P = copy(rf)
d = 0.0

LQ(Q, R, A, B, C, N, bet, capT, rf, P, d, F)
end

"""
Update `P` and `d` from the value function representation in finite horizon case

##### Arguments

- `lq::LQ` : instance of `LQ` type

##### Returns

- `P::ScalarOrArray` : n x n matrix in value function representation
V(x) = x'Px + d
- `d::Real` : Constant in value function representation

##### Notes

This function updates the `P` and `d` fields on the `lq` instance in addition to
returning them

"""

function update_values!(lq::LQ)
    # Simplify notation
    Q, R, A, B, N, C, P, d = lq.Q, lq.R, lq.A, lq.B, lq.N, lq.C, lq.P, lq.d

    # Some useful matrices
    s1 = Q + lq.bet * (B'P*B)
    s2 = lq.bet * (B'P*A) + N
    s3 = lq.bet * (A'P*A)

    # Compute F as (Q + B'PB)^{-1} (beta B'PA)
    lq.F = s1 \ s2

    # Shift P back in time one step
    new_P = R - s2'lq.F + s3

    # Recalling that trace(AB) = trace(BA)
    new_d = lq.bet * (d + trace(P * C * C'))

    # Set new state
    lq.P, lq.d = new_P, new_d
end

```

```
"""
Computes value and policy functions in infinite horizon model

##### Arguments

- `lq::LQ` : instance of `LQ` type

##### Returns

- `P::ScalarOrArray` : n x n matrix in value function representation
V(x) = x'Px + d
- `d::Real` : Constant in value function representation
- `F::ScalarOrArray` : Policy rule that specifies optimal control in each period

##### Notes

This function updates the `P`, `d`, and `F` fields on the `lq` instance in
addition to returning them

"""

function stationary_values!(lq::LQ)
    # simplify notation
    Q, R, A, B, N, C = lq.Q, lq.R, lq.A, lq.B, lq.N, lq.C

    # solve Riccati equation, obtain P
    A0, B0 = sqrt(lq.bet) * A, sqrt(lq.bet) * B
    P = solve_discrete_riccati(A0, B0, R, Q, N)

    # Compute F
    s1 = Q + lq.bet * (B' * P * B)
    s2 = lq.bet * (B' * P * A) + N
    F = s1 \ s2

    # Compute d
    d = lq.bet * trace(P * C * C') / (1 - lq.bet)

    # Bind states
    lq.P, lq.F, lq.d = P, F, d
end

"""

Non-mutating routine for solving for `P`, `d`, and `F` in infinite horizon model

See docstring for stationary_values! for more explanation
"""

function stationary_values(lq::LQ)
    _lq = LQ(copy(lq.Q),
              copy(lq.R),
              copy(lq.A),
              copy(lq.B),
              copy(lq.C),
              copy(lq.N),
              bet=copy(lq.bet),
              
```

```

    capT=lq.capT,
    rf=copy(lq.rf))

stationary_values!(_lq)
return _lq.P, _lq.F, _lq.d
end

"""
Private method implementing `compute_sequence` when state is a scalar
"""
function _compute_sequence{T}(lq::LQ, x0::T, policies)
    capT = length(policies)

    x_path = Array(T, capT+1)
    u_path = Array(T, capT)

    x_path[1] = x0
    u_path[1] = -(first(policies)*x0)
    w_path = lq.C * randn(capT+1)

    for t = 2:capT
        f = policies[t]
        x_path[t] = lq.A*x_path[t-1] + lq.B*u_path[t-1] + w_path[t]
        u_path[t] = -(f*x_path[t])
    end
    x_path[end] = lq.A*x_path[capT] + lq.B*u_path[capT] + w_path[end]

    x_path, u_path, w_path
end

"""
Private method implementing `compute_sequence` when state is a scalar
"""
function _compute_sequence{T}(lq::LQ, x0::Vector{T}, policies)
    # Ensure correct dimensionality
    n, j, k = size(lq.C, 1), size(lq.C, 2), size(lq.B, 2)
    capT = length(policies)

    A, B, C = lq.A, reshape(lq.B, n, k), reshape(lq.C, n, j)

    x_path = Array(T, n, capT+1)
    u_path = Array(T, k, capT)
    w_path = C*randn(j, capT+1)

    x_path[:, 1] = x0
    u_path[:, 1] = -(first(policies)*x0)

    for t = 2:capT
        f = policies[t]
        x_path[:, t] = A*x_path[:, t-1] + B*u_path[:, t-1] + w_path[:, t]
        u_path[:, t] = -(f*x_path[:, t])
    end
    x_path[:, end] = A*x_path[:, capT] + B*u_path[:, capT] + w_path[:, end]

```

```

    x_path, u_path, w_path
end

"""

Compute and return the optimal state and control sequence, assuming innovation N(0,1)

##### Arguments

- `lq::LQ` : instance of `LQ` type
- `x0::ScalarOrArray` : initial state
- `ts_length::Integer(100)` : maximum number of periods for which to return
process. If `lq` instance is finite horizon type, the sequences are returned
only for `min(ts_length, lq.capT)`

##### Returns

- `x_path::Matrix{Float64}` : An n x T+1 matrix, where the t-th column
represents `x_t`
- `u_path::Matrix{Float64}` : A k x T matrix, where the t-th column represents
`u_t`
- `w_path::Matrix{Float64}` : A n x T+1 matrix, where the t-th column represents
`lq.C*N(0,1)`

"""

function compute_sequence(lq::LQ, x0::ScalarOrArray, ts_length::Integer=100)

    # Compute and record the sequence of policies
    if isa(lq.capT, Void)
        stationary_values!(lq)
        policies = fill(lq.F, ts_length)
    else
        capT = min(ts_length, lq.capT)
        policies = Array(typeof(lq.F), capT)
        for t = capT:-1:1
            update_values!(lq)
            policies[t] = lq.F
        end
    end

    _compute_sequence(lq, x0, policies)
end

```

In the module, the various updating, simulation and fixed point methods are wrapped in a type called LQ, which includes

- Instance data:
 - The required parameters Q, R, A, B and optional parameters C, β, T, R_f, N specifying a given LQ model
 - * set T and R_f to None in the infinite horizon case
 - * set $C = \text{None}$ (or zero) in the deterministic case

- the value function and policy data
 - * d_t, P_t, F_t in the finite horizon case
 - * d, P, F in the infinite horizon case
- Methods:
 - `update_values` — shifts d_t, P_t, F_t to their $t - 1$ values via (2.99), (2.100) and (2.101)
 - `stationary_values` — computes P, d, F in the infinite horizon case
 - `compute_sequence` — simulates the dynamics of x_t, u_t, w_t given x_0 and assuming standard normal shocks

An example of usage is given in `lq_permanent_1.jl` from the [applications repository](#), the contents of which are shown below

This program can be used to replicate the figures shown in our *section on the permanent income model*

(Some of the plotting techniques are rather fancy and you can ignore those details if you wish)

```
using QuantEcon
using Plots
plotlyjs(size = (900, 900))

# == Model parameters ==
r = 0.05
bet = 1 / (1 + r)
T = 45
c_bar = 2.0
sigma = 0.25
mu = 1.0
q = 1e6

# == Formulate as an LQ problem ==
Q = 1.0
R = zeros(2, 2)
Rf = zeros(2, 2); Rf[1, 1] = q
A = [1.0+r -c_bar+mu;
      0.0 1.0]
B = [-1.0; 0.0]
C = [sigma; 0.0]

# == Compute solutions and simulate ==
lq = LQ(Q, R, A, B, C; bet = bet, capT = T, rf = Rf)
x0 = [0.0; 1.0]
xp, up, wp = compute_sequence(lq, x0)

# == Convert back to assets, consumption and income ==
assets = vec(xp[1, :])           # a_t
c = vec(up + c_bar)             # c_t
income = vec(wp[1, 2:end] + mu) # y_t

# == Plot results ==

```

```
plot(Vector[assets, c, zeros(T + 1), income, cumsum(income - mu)],
  lab = ["assets" "consumption" "" "non-financial income" "cumulative unanticipated income"],
  color = [:blue :green :black :orange :red],
  width = 3, xaxis = ("Time"), layout = (2, 1),
  bottom_margin = 20mm, show = true)
```

Further Applications

Application 1: Age-Dependent Income Process Previously we studied a permanent income model that generated consumption smoothing

One unrealistic feature of that model is the assumption that the mean of the random income process does not depend on the consumer's age

A more realistic income profile is one that rises in early working life, peaks towards the middle and maybe declines toward end of working life, and falls more during retirement

In this section, we will model this rise and fall as a symmetric inverted "U" using a polynomial in age

As before, the consumer seeks to minimize

$$\mathbb{E} \left\{ \sum_{t=0}^{T-1} \beta^t (c_t - \bar{c})^2 + \beta^T q a_T^2 \right\} \quad (2.111)$$

subject to $a_{t+1} = (1+r)a_t - c_t + y_t$, $t \geq 0$

For income we now take $y_t = p(t) + \sigma w_{t+1}$ where $p(t) := m_0 + m_1 t + m_2 t^2$

(In the next section we employ some tricks to implement a more sophisticated model)

The coefficients m_0, m_1, m_2 are chosen such that $p(0) = 0$, $p(T/2) = \mu$, and $p(T) = 0$

You can confirm that the specification $m_0 = 0, m_1 = T\mu/(T/2)^2, m_2 = -\mu/(T/2)^2$ satisfies these constraints

To put this into an LQ setting, consider the budget constraint, which becomes

$$a_{t+1} = (1+r)a_t - u_t - \bar{c} + m_1 t + m_2 t^2 + \sigma w_{t+1} \quad (2.112)$$

The fact that a_{t+1} is a linear function of $(a_t, 1, t, t^2)$ suggests taking these four variables as the state vector x_t

Once a good choice of state and control (recall $u_t = c_t - \bar{c}$) has been made, the remaining specifications fall into place relatively easily

Thus, for the dynamics we set

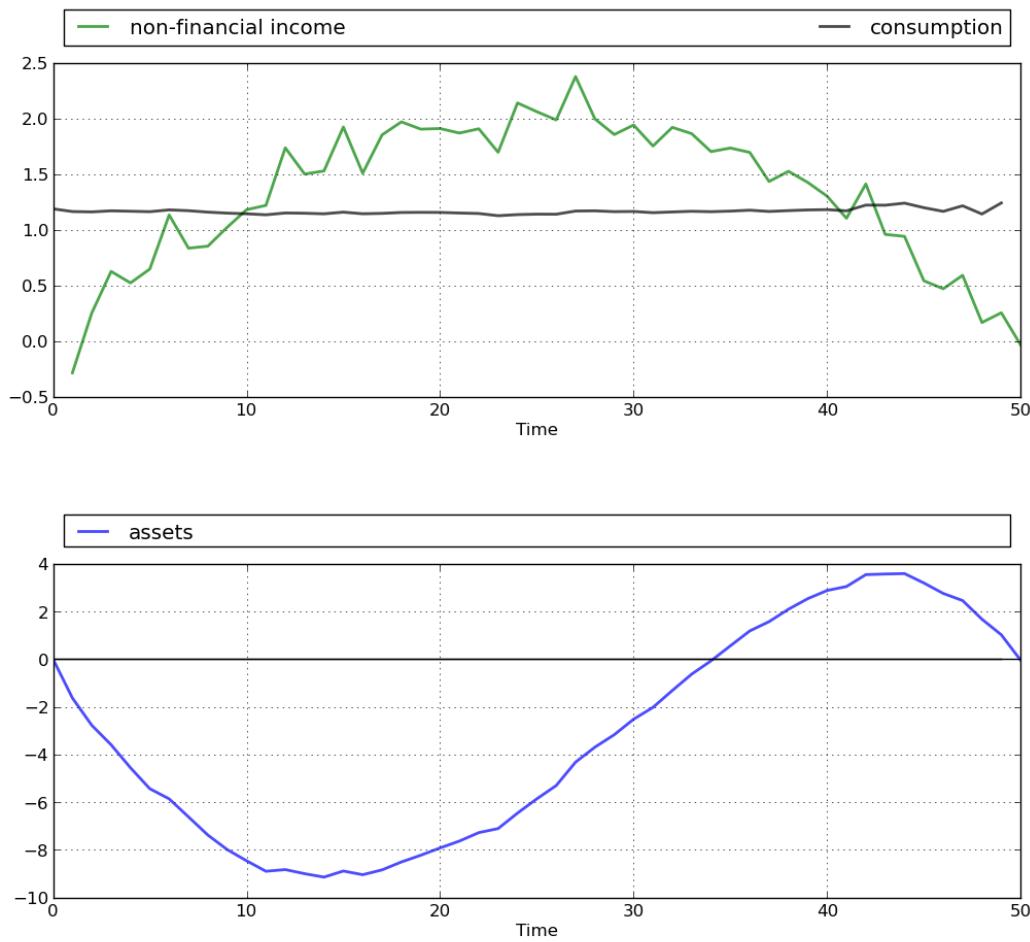
$$x_t := \begin{pmatrix} a_t \\ 1 \\ t \\ t^2 \end{pmatrix}, \quad A := \begin{pmatrix} 1+r & -\bar{c} & m_1 & m_2 \\ 0 & 1 & 0 & 0 \\ 0 & 1 & 1 & 0 \\ 0 & 1 & 2 & 1 \end{pmatrix}, \quad B := \begin{pmatrix} -1 \\ 0 \\ 0 \\ 0 \end{pmatrix}, \quad C := \begin{pmatrix} \sigma \\ 0 \\ 0 \\ 0 \end{pmatrix} \quad (2.113)$$

If you expand the expression $x_{t+1} = Ax_t + Bu_t + Cw_{t+1}$ using this specification, you will find that assets follow (2.112) as desired, and that the other state variables also update appropriately

To implement preference specification (2.111) we take

$$Q := 1, \quad R := \begin{pmatrix} 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \end{pmatrix} \quad \text{and} \quad R_f := \begin{pmatrix} q & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \end{pmatrix} \quad (2.114)$$

The next figure shows a simulation of consumption and assets computed using the compute_sequence method of lqcontrol.jl with initial assets set to zero



Once again, smooth consumption is a dominant feature of the sample paths

The asset path exhibits dynamics consistent with standard life cycle theory

Exercise 1 gives the full set of parameters used here and asks you to replicate the figure

Application 2: A Permanent Income Model with Retirement In the *previous application*, we generated income dynamics with an inverted U shape using polynomials, and placed them in an LQ framework

It is arguably the case that this income process still contains unrealistic features

A more common earning profile is where

1. income grows over working life, fluctuating around an increasing trend, with growth flattening off in later years
2. retirement follows, with lower but relatively stable (non-financial) income

Letting K be the retirement date, we can express these income dynamics by

$$y_t = \begin{cases} p(t) + \sigma w_{t+1} & \text{if } t \leq K \\ s & \text{otherwise} \end{cases} \quad (2.115)$$

Here

- $p(t) := m_1 t + m_2 t^2$ with the coefficients m_1, m_2 chosen such that $p(K) = \mu$ and $p(0) = p(2K) = 0$
- s is retirement income

We suppose that preferences are unchanged and given by (2.103)

The budget constraint is also unchanged and given by $a_{t+1} = (1+r)a_t - c_t + y_t$

Our aim is to solve this problem and simulate paths using the LQ techniques described in this lecture

In fact this is a nontrivial problem, as the kink in the dynamics (2.115) at K makes it very difficult to express the law of motion as a fixed-coefficient linear system

However, we can still use our LQ methods here by suitably linking two component LQ problems

These two LQ problems describe the consumer's behavior during her working life (`lq_working`) and retirement (`lq_retired`)

(This is possible because in the two separate periods of life, the respective income processes [polynomial trend and constant] each fit the LQ framework)

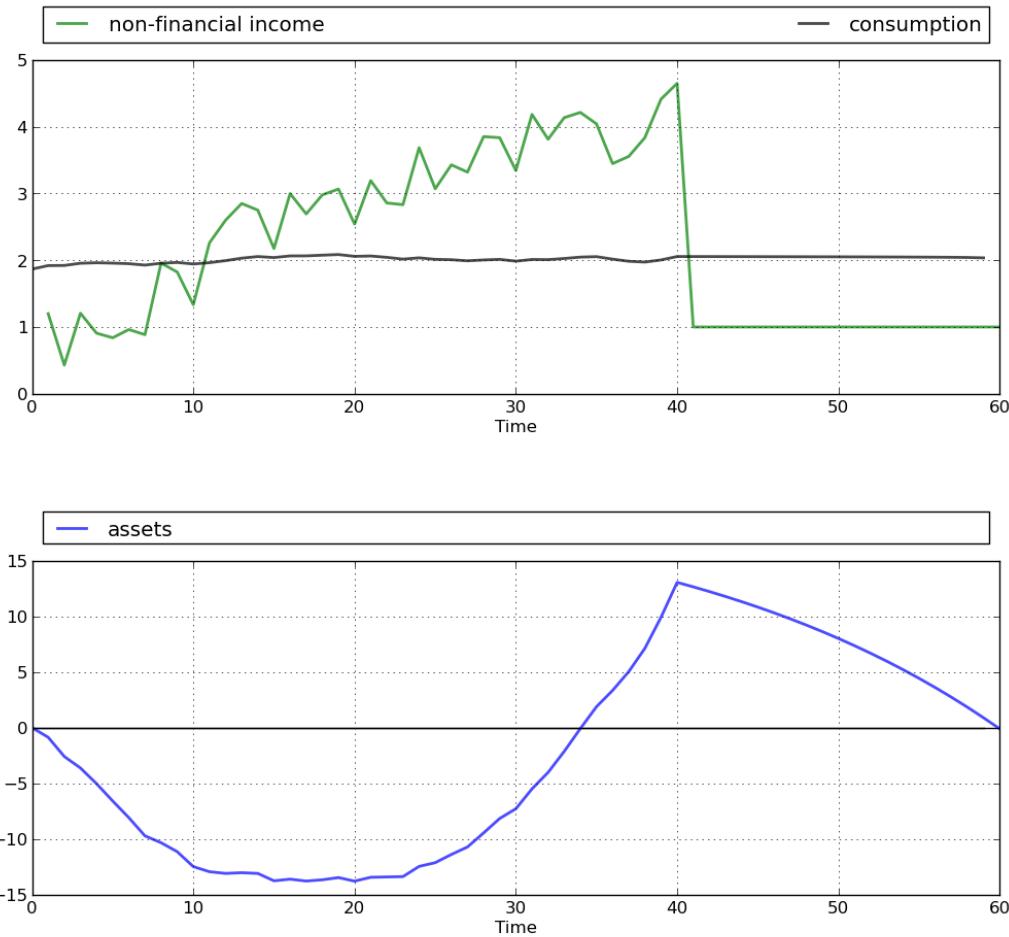
The basic idea is that although the whole problem is not a single time-invariant LQ problem, it is still a dynamic programming problem, and hence we can use appropriate Bellman equations at every stage

Based on this logic, we can

1. solve `lq_retired` by the usual backwards induction procedure, iterating back to the start of retirement
2. take the start-of-retirement value function generated by this process, and use it as the terminal condition R_f to feed into the `lq_working` specification
3. solve `lq_working` by backwards induction from this choice of R_f , iterating back to the start of working life

This process gives the entire life-time sequence of value functions and optimal policies

The next figure shows one simulation based on this procedure



The full set of parameters used in the simulation is discussed in *Exercise 2*, where you are asked to replicate the figure

Once again, the dominant feature observable in the simulation is consumption smoothing

The asset path fits well with standard life cycle theory, with dissaving early in life followed by later saving

Assets peak at retirement and subsequently decline

Application 3: Monopoly with Adjustment Costs Consider a monopolist facing stochastic inverse demand function

$$p_t = a_0 - a_1 q_t + d_t$$

Here q_t is output, and the demand shock d_t follows

$$d_{t+1} = \rho d_t + \sigma w_{t+1}$$

where $\{w_t\}$ is iid and standard normal

The monopolist maximizes the expected discounted sum of present and future profits

$$\mathbb{E} \left\{ \sum_{t=0}^{\infty} \beta^t \pi_t \right\} \quad \text{where} \quad \pi_t := p_t q_t - c q_t - \gamma (q_{t+1} - q_t)^2 \quad (2.116)$$

Here

- $\gamma (q_{t+1} - q_t)^2$ represents adjustment costs
- c is average cost of production

This can be formulated as an LQ problem and then solved and simulated, but first let's study the problem and try to get some intuition

One way to start thinking about the problem is to consider what would happen if $\gamma = 0$

Without adjustment costs there is no intertemporal trade-off, so the monopolist will choose output to maximize current profit in each period

It's not difficult to show that profit-maximizing output is

$$\bar{q}_t := \frac{a_0 - c + d_t}{2a_1}$$

In light of this discussion, what we might expect for general γ is that

- if γ is close to zero, then q_t will track the time path of \bar{q}_t relatively closely
- if γ is larger, then q_t will be smoother than \bar{q}_t , as the monopolist seeks to avoid adjustment costs

This intuition turns out to be correct

The following figures show simulations produced by solving the corresponding LQ problem

The only difference in parameters across the figures is the size of γ

To produce these figures we converted the monopolist problem into an LQ problem

The key to this conversion is to choose the right state — which can be a bit of an art

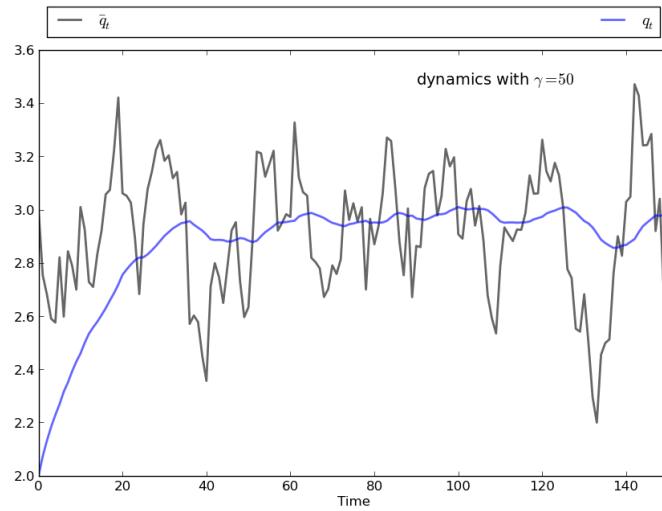
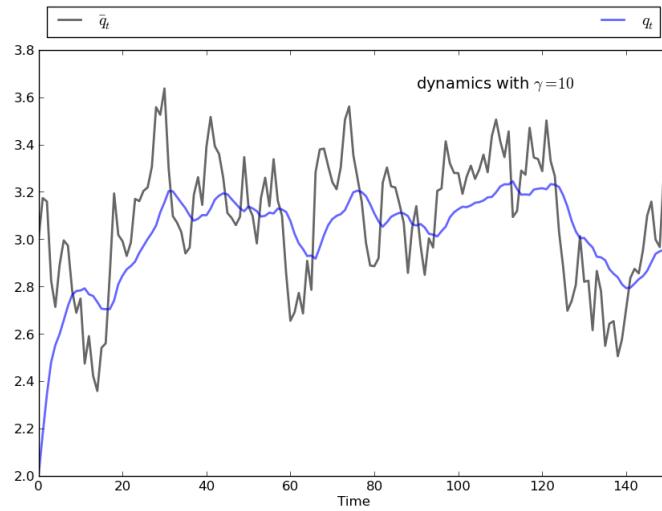
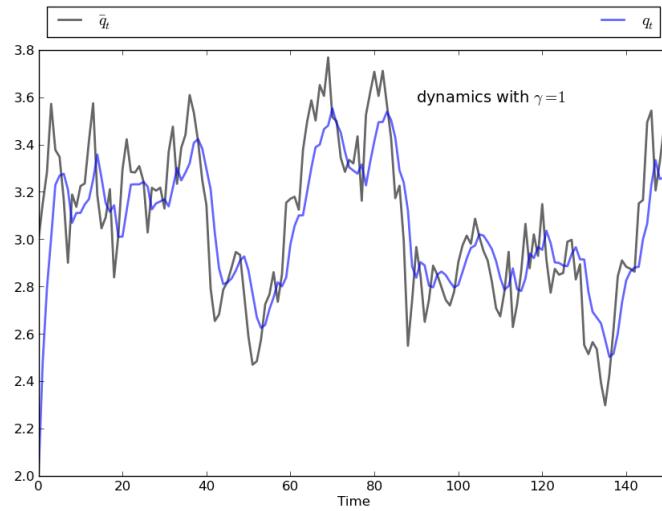
Here we take $x_t = (\bar{q}_t \ q_t \ 1)'$, while the control is chosen as $u_t = q_{t+1} - q_t$

We also manipulated the profit function slightly

In (2.116), current profits are $\pi_t := p_t q_t - c q_t - \gamma (q_{t+1} - q_t)^2$

Let's now replace π_t in (2.116) with $\hat{\pi}_t := \pi_t - a_1 \bar{q}_t^2$

This makes no difference to the solution, since $a_1 \bar{q}_t^2$ does not depend on the controls



(In fact we are just adding a constant term to (2.116), and optimizers are not affected by constant terms)

The reason for making this substitution is that, as you will be able to verify, $\hat{\pi}_t$ reduces to the simple quadratic

$$\hat{\pi}_t = -a_1(q_t - \bar{q}_t)^2 - \gamma u_t^2$$

After negation to convert to a minimization problem, the objective becomes

$$\min \mathbb{E} \sum_{t=0}^{\infty} \beta^t \{ a_1(q_t - \bar{q}_t)^2 + \gamma u_t^2 \} \quad (2.117)$$

It's now relatively straightforward to find R and Q such that (2.117) can be written as (2.107)

Furthermore, the matrices A , B and C from (2.88) can be found by writing down the dynamics of each element of the state

Exercise 3 asks you to complete this process, and reproduce the preceding figures

Exercises

Exercise 1 Replicate the figure with polynomial income *shown above*

The parameters are $r = 0.05$, $\beta = 1/(1+r)$, $\bar{c} = 1.5$, $\mu = 2$, $\sigma = 0.15$, $T = 50$ and $q = 10^4$

Exercise 2 Replicate the figure on work and retirement *shown above*

The parameters are $r = 0.05$, $\beta = 1/(1+r)$, $\bar{c} = 4$, $\mu = 4$, $\sigma = 0.35$, $K = 40$, $T = 60$, $s = 1$ and $q = 10^4$

To understand the overall procedure, carefully read the section containing that figure

Some hints are as follows:

First, in order to make our approach work, we must ensure that both LQ problems have the same state variables and control

As with previous applications, the control can be set to $u_t = c_t - \bar{c}$

For `lq_working`, x_t , A , B , C can be chosen as in (2.113)

- Recall that m_1, m_2 are chosen so that $p(K) = \mu$ and $p(2K) = 0$

For `lq_retired`, use the same definition of x_t and u_t , but modify A , B , C to correspond to constant income $y_t = s$

For `lq_retired`, set preferences as in (2.114)

For `lq_working`, preferences are the same, except that R_f should be replaced by the final value function that emerges from iterating `lq_retired` back to the start of retirement

With some careful footwork, the simulation can be generated by patching together the simulations from these two separate models

Exercise 3 Reproduce the figures from the monopolist application *given above*

For parameters, use $a_0 = 5, a_1 = 0.5, \sigma = 0.15, \rho = 0.9, \beta = 0.95$ and $c = 2$, while γ varies between 1 and 50 (see figures)

Solutions

[Solution notebook](#)

Discrete Dynamic Programming

Contents

- *Discrete Dynamic Programming*
 - *Overview*
 - *Discrete DPs*
 - *Solving Discrete DPs*
 - *Example: A Growth Model*
 - *Exercises*
 - *Solutions*
 - *Appendix: Algorithms*

Overview

In this lecture we discuss a family of dynamic programming problems with the following features:

1. a discrete state space and discrete choices (actions)
2. an infinite horizon
3. discounted rewards
4. Markov state transitions

We call such problems discrete dynamic programs, or discrete DPs

Discrete DPs are the workhorses in much of modern quantitative economics, including

- monetary economics
- search and labor economics
- household savings and consumption theory
- investment theory
- asset pricing
- industrial organization, etc.

When a given model is not inherently discrete, it is common to replace it with a discretized version in order to use discrete DP techniques

This lecture covers

- the theory of dynamic programming in a discrete setting, plus examples and applications
- a powerful set of routines for solving discrete DPs from the [QuantEcon code library](#)

How to Read this Lecture We have used dynamic programming in a number of earlier lectures, such as

- The [shortest path lecture](#)
- The [McCall growth model lecture](#)
- The [optimal growth lecture](#)

Here we shift to a more systematic and theoretical treatment, including algorithms and implementation

The code discussed below was authored primarily by [Daisuke Oyama](#)

References For background reading on dynamic programming and additional applications, see, for example,

- [\[LS12\]](#)
- [\[HLL96\]](#), section 3.5
- [\[Put05\]](#)
- [\[SLP89\]](#)
- [\[Rus96\]](#)
- [\[MF02\]](#)
- [EDTC](#), chapter 5

Discrete DPs

Loosely speaking, a discrete DP is a maximization problem with an objective function of the form

$$\mathbb{E} \sum_{t=0}^{\infty} \beta^t r(s_t, a_t) \quad (2.118)$$

where

- s_t is the state variable
- a_t is the action
- β is a discount factor
- $r(s_t, a_t)$ is interpreted as a current reward when the state is s_t and the action chosen is a_t

Each pair (s_t, a_t) pins down transition probabilities $Q(s_t, a_t, s_{t+1})$ for the next period state s_{t+1}

Thus, actions influence not only current rewards but also the future time path of the state

The essence of dynamic programming problems is to trade off current rewards vs favorable positioning of the future state (modulo randomness)

Examples:

- consuming today vs saving and accumulating assets
- accepting a job offer today vs seeking a better one in the future
- exercising an option now vs waiting

Policies The most fruitful way to think about solutions to discrete DP problems is to compare *policies*

In general, a policy is a randomized map from past actions and states to current action

In the setting formalized below, it suffices to consider so-called *stationary Markov policies*, which consider only the current state

In particular, a stationary Markov policy is a map σ from states to actions

- $a_t = \sigma(s_t)$ indicates that a_t is the action to be taken in state s_t

It is known that, for any arbitrary policy, there exists a stationary Markov policy that dominates it at least weakly

- See section 5.5 of [Put05] for discussion and proofs

In what follows, stationary Markov policies are referred to simply as policies

The aim is to find an optimal policy, in the sense of one that maximizes (2.118)

Let's now step through these ideas more carefully

Formal definition Formally, a discrete dynamic program consists of the following components:

1. A finite set of *states* $S = \{0, \dots, n - 1\}$
2. A finite set of *feasible actions* $A(s)$ for each state $s \in S$, and a corresponding set

$$SA := \{(s, a) \mid s \in S, a \in A(s)\}$$

of *feasible state-action pairs*

3. A *reward function* $r: SA \rightarrow \mathbb{R}$
4. A *transition probability function* $Q: SA \rightarrow \Delta(S)$, where $\Delta(S)$ is the set of probability distributions over S
5. A *discount factor* $\beta \in [0, 1)$

We also use the notation $A := \bigcup_{s \in S} A(s) = \{0, \dots, m - 1\}$ and call this set the *action space*

A *policy* is a function $\sigma: S \rightarrow A$

A policy is called *feasible* if it satisfies $\sigma(s) \in A(s)$ for all $s \in S$

Denote the set of all feasible policies by Σ

If a decision maker uses a policy $\sigma \in \Sigma$, then

- the current reward at time t is $r(s_t, \sigma(s_t))$
- the probability that $s_{t+1} = s'$ is $Q(s_t, \sigma(s_t), s')$

For each $\sigma \in \Sigma$, define

- r_σ by $r_\sigma(s) := r(s, \sigma(s))$
- Q_σ by $Q_\sigma(s, s') := Q(s, \sigma(s), s')$

Notice that Q_σ is a *stochastic matrix* on S

It gives transition probabilities of the *controlled chain* when we follow policy σ

If we think of r_σ as a column vector, then so is $Q_\sigma^t r_\sigma$, and the s -th row of the latter has the interpretation

$$(Q_\sigma^t r_\sigma)(s) = \mathbb{E}[r(s_t, \sigma(s_t)) \mid s_0 = s] \quad \text{when } \{s_t\} \sim Q_\sigma \quad (2.119)$$

Comments

- $\{s_t\} \sim Q_\sigma$ means that the state is generated by stochastic matrix Q_σ
- See *this discussion* on computing expectations of Markov chains for an explanation of the expression in (2.119)

Notice that we're not really distinguishing between functions from S to \mathbb{R} and vectors in \mathbb{R}^n

This is natural because they are in one to one correspondence

Value and Optimality Let $v_\sigma(s)$ denote the discounted sum of expected reward flows from policy σ when the initial state is s

To calculate this quantity we pass the expectation through the sum in (2.118) and use (2.119) to get

$$v_\sigma(s) = \sum_{t=0}^{\infty} \beta^t (Q_\sigma^t r_\sigma)(s) \quad (s \in S)$$

This function is called the *policy value function* for the policy σ

The *optimal value function*, or simply *value function*, is the function $v^*: S \rightarrow \mathbb{R}$ defined by

$$v^*(s) = \max_{\sigma \in \Sigma} v_\sigma(s) \quad (s \in S)$$

(We can use \max rather than \sup here because the domain is a finite set)

A policy $\sigma \in \Sigma$ is called *optimal* if $v_\sigma(s) = v^*(s)$ for all $s \in S$

Given any $w: S \rightarrow \mathbb{R}$, a policy $\sigma \in \Sigma$ is called w -greedy if

$$\sigma(s) \in \arg \max_{a \in A(s)} \left\{ r(s, a) + \beta \sum_{s' \in S} w(s') Q(s, a, s') \right\} \quad (s \in S)$$

As discussed in detail below, optimal policies are precisely those that are v^* -greedy

Two Operators It is useful to define the following operators:

- The *Bellman operator* $T: \mathbb{R}^S \rightarrow \mathbb{R}^S$ is defined by

$$(Tv)(s) = \max_{a \in A(s)} \left\{ r(s, a) + \beta \sum_{s' \in S} v(s') Q(s, a, s') \right\} \quad (s \in S)$$

- For any policy function $\sigma \in \Sigma$, the operator $T_\sigma: \mathbb{R}^S \rightarrow \mathbb{R}^S$ is defined by

$$(T_\sigma v)(s) = r(s, \sigma(s)) + \beta \sum_{s' \in S} v(s') Q(s, \sigma(s), s') \quad (s \in S)$$

This can be written more succinctly in operator notation as

$$T_\sigma v = r_\sigma + \beta Q_\sigma v$$

The two operators are both monotone

- $v \leq w$ implies $Tv \leq Tw$ pointwise on S , and similarly for T_σ

They are also contraction mappings with modulus β

- $\|Tv - Tw\| \leq \beta \|v - w\|$ and similarly for T_σ , where $\|\cdot\|$ is the max norm

For any policy σ , its value v_σ is the unique fixed point of T_σ

For proofs of these results and those in the next section, see, for example, [EDTC](#), chapter 10

The Bellman Equation and the Principle of Optimality The main principle of the theory of dynamic programming is that

- the optimal value function v^* is a unique solution to the *Bellman equation*,

$$v(s) = \max_{a \in A(s)} \left\{ r(s, a) + \beta \sum_{s' \in S} v(s') Q(s, a, s') \right\} \quad (s \in S),$$

or in other words, v^* is the unique fixed point of T , and

- σ^* is an optimal policy function if and only if it is v^* -greedy

By the definition of greedy policies given above, this means that

$$\sigma^*(s) \in \arg \max_{a \in A(s)} \left\{ r(s, a) + \beta \sum_{s' \in S} v^*(s') Q(s, a, s') \right\} \quad (s \in S)$$

Solving Discrete DPs

Now that the theory has been set out, let's turn to solution methods

Code for solving discrete DPs is available in `ddp.jl` from the `QuantEcon.jl` code library

It implements the three most important solution methods for discrete dynamic programs, namely

- value function iteration
- policy function iteration
- modified policy function iteration

Let's briefly review these algorithms and their implementation

Value Function Iteration Perhaps the most familiar method for solving all manner of dynamic programs is value function iteration

This algorithm uses the fact that the Bellman operator T is a contraction mapping with fixed point v^*

Hence, iterative application of T to any initial function $v^0: S \rightarrow \mathbb{R}$ converges to v^*

The details of the algorithm can be found in *the appendix*

Policy Function Iteration This routine, also known as Howard's policy improvement algorithm, exploits more closely the particular structure of a discrete DP problem

Each iteration consists of

1. A policy evaluation step that computes the value v_σ of a policy σ by solving the linear equation $v = T_\sigma v$
2. A policy improvement step that computes a v_σ -greedy policy

In the current setting policy iteration computes an exact optimal policy in finitely many iterations

- See theorem 10.2.6 of `EDTC` for a proof

The details of the algorithm can be found in *the appendix*

Modified Policy Function Iteration Modified policy iteration replaces the policy evaluation step in policy iteration with "partial policy evaluation"

The latter computes an approximation to the value of a policy σ by iterating T_σ for a specified number of times

This approach can be useful when the state space is very large and the linear system in the policy evaluation step of policy iteration is correspondingly difficult to solve

The details of the algorithm can be found in *the appendix*

Example: A Growth Model

Let's consider a simple consumption-saving model

A single household either consumes or stores its own output of a single consumption good

The household starts each period with current stock s

Next, the household chooses a quantity a to store and consumes $c = s - a$

- Storage is limited by a global upper bound M
- Flow utility is $u(c) = c^\alpha$

Output is drawn from a discrete uniform distribution on $\{0, \dots, B\}$

The next period stock is therefore

$$s' = a + U \quad \text{where} \quad U \sim U[0, \dots, B]$$

The discount factor is $\beta \in [0, 1)$

Discrete DP Representation We want to represent this model in the format of a discrete dynamic program

To this end, we take

- the state variable to be the stock s
- the state space to be $S = \{0, \dots, M + B\}$
 - hence $n = M + B + 1$
- the action to be the storage quantity a
- the set of feasible actions at s to be $A(s) = \{0, \dots, \min\{s, M\}\}$
 - hence $A = \{0, \dots, M\}$ and $m = M + 1$
- the reward function to be $r(s, a) = u(s - a)$
- the transition probabilities to be

$$Q(s, a, s') := \begin{cases} \frac{1}{B+1} & \text{if } a \leq s' \leq a + B \\ 0 & \text{otherwise} \end{cases} \quad (2.120)$$

Defining a DiscreteDP Instance This information will be used to create an instance of *DiscreteDP* by passing the following information

1. An $n \times m$ reward array R
2. An $n \times m \times n$ transition probability array Q
3. A discount factor β

For R we set $R[s, a] = u(s - a)$ if $a \leq s$ and $-\infty$ otherwise

For Q we follow the rule in (2.120)

Note:

- The feasibility constraint is embedded into R by setting $R[s, a] = -\infty$ for $a \notin A(s)$
- Probability distributions for (s, a) with $a \notin A(s)$ can be arbitrary

A simple type that sets up these objects for us in the current application can be found in the QuantEcon.applications repository

For convenience let's repeat it here:

```
type SimpleOG
    B :: Int64
    M :: Int64
    alpha :: Float64
    beta :: Float64
    R :: Array{Float64}
    Q :: Array{Float64}
end

function SimpleOG(;B=10, M=5, alpha=0.5, beta=0.9)

    u(c) = c^alpha
    n = B + M + 1
    m = M + 1

    R = Array(Float64,n,m)
    Q = zeros(Float64,n,m,n)

    for a in 0:M
        Q[:, a + 1, (a:(a + B)) + 1] = 1 / (B + 1)
        for s in 0:(B + M)
            R[s + 1, a + 1] = a <= s ? u(s - a) : -Inf
        end
    end

    return SimpleOG(B, M, alpha, beta, R, Q)
end
```

Let's run this code and create an instance of SimpleOG

```
julia> include("finite_dp_og_example.jl")
SimpleOG

julia> g = SimpleOG();
```

Instances of DiscreteDP are created using the signature DiscreteDP(R, Q, beta)

Let's create an instance using the objects stored in g

```
julia> using QuantEcon
```

```
julia> ddp = DiscreteDP(g.R, g.Q, g.beta);
```

Now that we have an instance `ddp` of `DiscreteDP` we can solve it as follows

```
julia> results = solve(ddp, PFI);
```

Let's see what we've got here

```
julia> fieldnames(results)
5-element Array{Symbol,1}:
 :v
 :Tv
 :num_iter
 :sigma
 :mc
```

The most important attributes are `v`, the value function, and `sigma`, the optimal policy

```
julia> results.v
16-element Array{Float64,1}:
 19.0174
 20.0174
 20.4316
 20.7495
 21.0408
 21.3087
 21.5448
 21.7693
 21.9827
 22.1882
 22.3845
 22.5781
 22.7611
 22.9438
 23.1153
 23.2776
```

```
julia> results.sigma - 1
16-element Array{Int64,1}:
 0
 0
 0
 0
 1
 1
 1
 2
 2
 3
 3
 4
 5
 5
 5
```

| |
|---|
| 5 |
|---|

Here 1 is subtracted from `results.sigma` because we added 1 to each state and action to create valid indices

Since we've used policy iteration, these results will be exact unless we hit the iteration bound `max_iter`

Let's make sure this didn't happen

| |
|---|
| <code>julia> results.num_iter</code> |
| 3 |

In this case we converged in only 3 iterations

Another interesting object is `results.mc`, which is the controlled chain defined by Q_{σ^*} , where σ^* is the optimal policy

In other words, it gives the dynamics of the state when the agent follows the optimal policy

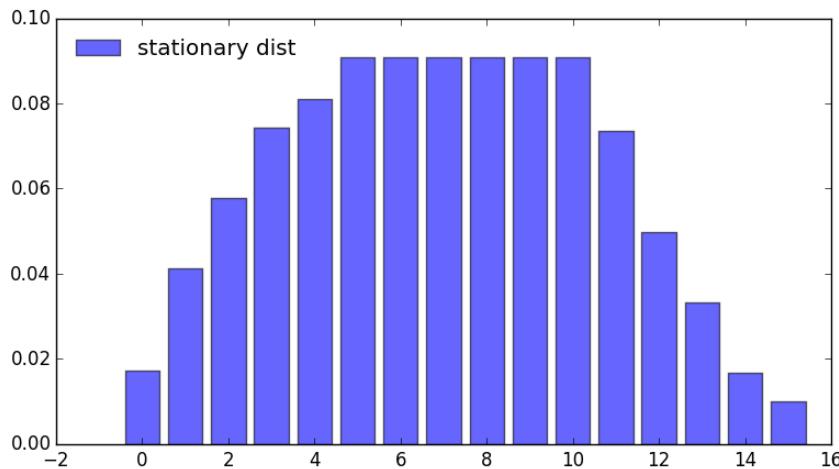
Since this object is an instance of *MarkovChain* from `QuantEcon.jl` (see [this lecture](#) for more discussion), we can easily simulate it, compute its stationary distribution and so on

| |
|---|
| <code>julia> stationary_distributions(results.mc) [1]</code> |
| 16-element Array{Float64,1}: |
| 0.0173219 |
| 0.0412106 |
| 0.0577396 |
| 0.0742685 |
| 0.0809582 |
| 0.0909091 |
| 0.0909091 |
| 0.0909091 |
| 0.0909091 |
| 0.0909091 |
| 0.0735872 |
| 0.0496985 |
| 0.0331695 |
| 0.0166406 |
| 0.00995086 |

Here's the same information in a bar graph

What happens if the agent is more patient?

| |
|--|
| <code>julia> g_2 = SimpleOG(beta=0.99);</code> |
| <code>julia> ddp_2 = DiscreteDP(g_2.R, g_2.Q, g_2.beta);</code> |
| <code>julia> results_2 = solve(ddp_2, PFI);</code> |
| <code>julia> stationary_distributions(results_2.mc) [1]</code> |
| 16-element Array{Float64,1}: |
| 0.00546913 |



```

0.0232134
0.0314779
0.0480068
0.0562713
0.0909091
0.0909091
0.0909091
0.0909091
0.0909091
0.0909091
0.08544
0.0676957
0.0594312
0.0429023
0.0346378

```

If we look at the bar graph we can see the rightward shift in probability mass

State-Action Pair Formulation The DiscreteDP type in fact provides a second interface to setting up an instance

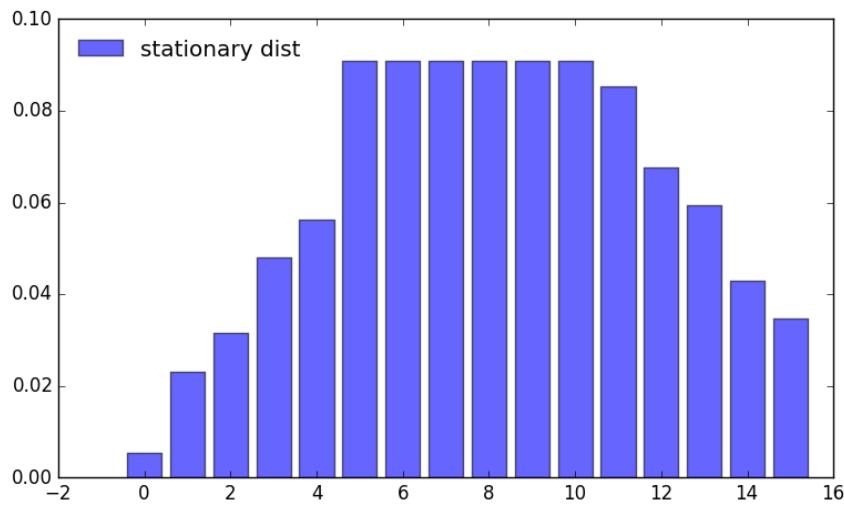
One of the advantages of this alternative set up is that it permits use of a sparse matrix for Q

(An example of using sparse matrices is given in the exercise solution notebook below)

The call signature of the second formulation is DiscreteDP(R, Q, beta, s_indices, a_indices) where

- s_indices and a_indices are arrays of equal length L enumerating all feasible state-action pairs
- R is an array of length L giving corresponding rewards
- Q is an L x n transition probability array

Here's how we could set up these objects for the preceding example



```

using QuantEcon

B = 10
M = 5
alpha = 0.5
beta = 0.9
u(c) = c^alpha
n = B + M + 1
m = M + 1

s_indices = Int64[]
a_indices = Int64[]
Q = Array(Float64, 0, n)
R = Float64[]

b = 1.0 / (B + 1)

for s in 0:(M + B)
    for a in 0:min(M, s)
        s_indices = [s_indices; s + 1]
        a_indices = [a_indices; a + 1]
        q = zeros(Float64, 1, n)
        q[(a + 1):((a + B) + 1)] = b
        Q = [Q; q]
        R = [R; u(s-a)]
    end
end

ddp = DiscreteDP(R, Q, beta, s_indices, a_indices);
results = solve(ddp, PFI)

```

Exercises

In the deterministic optimal growth [dynamic programming lecture](#), we solved a *benchmark model* that has an analytical solution to check we could replicate it numerically

The exercise is to replicate this solution using `DiscreteDP`

Solutions

[Solution notebook](#)

Appendix: Algorithms

This appendix covers the details of the solution algorithms implemented for `DiscreteDP`

We will make use of the following notions of approximate optimality:

- For $\varepsilon > 0$, v is called an ε -approximation of v^* if $\|v - v^*\| < \varepsilon$
- A policy $\sigma \in \Sigma$ is called ε -optimal if v_σ is an ε -approximation of v^*

Value Iteration The `DiscreteDP` value iteration method implements value function iteration as follows

1. Choose any $v^0 \in \mathbb{R}^n$, and specify $\varepsilon > 0$; set $i = 0$
2. Compute $v^{i+1} = T v^i$
3. If $\|v^{i+1} - v^i\| < [(1 - \beta) / (2\beta)]\varepsilon$, then go to step 4; otherwise, set $i = i + 1$ and go to step 2
4. Compute a v^{i+1} -greedy policy σ , and return v^{i+1} and σ

Given $\varepsilon > 0$, the value iteration algorithm

- terminates in a finite number of iterations
- returns an $\varepsilon/2$ -approximation of the optimal value function and an ε -optimal policy function (unless `iter_max` is reached)

(While not explicit, in the actual implementation each algorithm is terminated if the number of iterations reaches `iter_max`)

Policy Iteration The `DiscreteDP` policy iteration method runs as follows

1. Choose any $v^0 \in \mathbb{R}^n$ and compute a v^0 -greedy policy σ^0 ; set $i = 0$
2. Compute the value v_{σ^i} by solving the equation $v = T_{\sigma^i} v$
3. Compute a v_{σ^i} -greedy policy σ^{i+1} ; let $\sigma^{i+1} = \sigma^i$ if possible
4. If $\sigma^{i+1} = \sigma^i$, then return v_{σ^i} and σ^{i+1} ; otherwise, set $i = i + 1$ and go to step 2

The policy iteration algorithm terminates in a finite number of iterations

It returns an optimal value function and an optimal policy function (unless `iter_max` is reached)

Modified Policy Iteration The `DiscreteDP` modified policy iteration method runs as follows:

1. Choose any $v^0 \in \mathbb{R}^n$, and specify $\varepsilon > 0$ and $k \geq 0$; set $i = 0$
2. Compute a v^i -greedy policy σ^{i+1} ; let $\sigma^{i+1} = \sigma^i$ if possible (for $i \geq 1$)
3. Compute $u = T v^i (= T_{\sigma^{i+1}} v^i)$. If $\text{span}(u - v^i) < [(1 - \beta)/\beta]\varepsilon$, then go to step 5; otherwise go to step 4
 - Span is defined by $\text{span}(z) = \max(z) - \min(z)$
4. Compute $v^{i+1} = (T_{\sigma^{i+1}})^k u (= (T_{\sigma^{i+1}})^{k+1} v^i)$; set $i = i + 1$ and go to step 2
5. Return $v = u + [\beta/(1 - \beta)][(\min(u - v^i) + \max(u - v^i))/2]\mathbf{1}$ and σ_{i+1}

Given $\varepsilon > 0$, provided that v^0 is such that $T v^0 \geq v^0$, the modified policy iteration algorithm terminates in a finite number of iterations

It returns an $\varepsilon/2$ -approximation of the optimal value function and an ε -optimal policy function (unless `iter_max` is reached).

See also the documentation for `DiscreteDP`

Rational Expectations Equilibrium

Contents

- Rational Expectations Equilibrium
 - Overview
 - Defining Rational Expectations Equilibrium
 - Computation of an Equilibrium
 - Exercises
 - Solutions

“If you’re so smart, why aren’t you rich?”

Overview

This lecture introduces the concept of *rational expectations equilibrium*

To illustrate it, we describe a linear quadratic version of a famous and important model due to Lucas and Prescott [LP71]

This 1971 paper is one of a small number of research articles that kicked off the *rational expectations revolution*

We follow Lucas and Prescott by employing a setting that is readily “Bellmanized” (i.e., capable of being formulated in terms of dynamic programming problems)

Because we use linear quadratic setups for demand and costs, we can adapt the LQ programming techniques described in [this lecture](#)

We will learn about how a representative agent’s problem differs from a planner’s, and how a planning problem can be used to compute rational expectations quantities

We will also learn about how a rational expectations equilibrium can be characterized as a **fixed point** of a mapping from a *perceived law of motion* to an *actual law of motion*

Equality between a perceived and an actual law of motion for endogenous market-wide objects captures in a nutshell what the rational expectations equilibrium concept is all about

Finally, we will learn about the important “Big K , little k ” trick, a modeling device widely used in macroeconomics

Except that for us

- Instead of “Big K ” it will be “Big Y ”
- Instead of “little k ” it will be “little y ”

The Big Y , little y trick This widely used method applies in contexts in which a “representative firm” or agent is a “price taker” operating within a competitive equilibrium

We want to impose that

- The representative firm or individual takes *aggregate Y* as given when it chooses individual y , but ...
- At the end of the day, $Y = y$, so that the representative firm is indeed representative

The Big Y , little y trick accomplishes these two goals by

- Taking Y as beyond control when posing the choice problem of who chooses y ; but ...
- Imposing $Y = y$ after having solved the individual’s optimization problem

Please watch for how this strategy is applied as the lecture unfolds

We begin by applying the Big Y , little y trick in a very simple static context

A simple static example of the Big Y , little y trick Consider a static model in which a collection of n firms produce a homogeneous good that is sold in a competitive market

Each of these n firms sells output y

The price p of the good lies on an inverse demand curve

$$p = a_0 - a_1 Y \quad (2.121)$$

where

- $a_i > 0$ for $i = 0, 1$

- $Y = ny$ is the market-wide level of output

Each firm has total cost function

$$c(y) = c_1y + 0.5c_2y^2, \quad c_i > 0 \text{ for } i = 1, 2$$

The profits of a representative firm are $py - c(y)$

Using (2.121), we can express the problem of the representative firm as

$$\max_y [(a_0 - a_1Y)y - c_1y - 0.5c_2y^2] \quad (2.122)$$

In posing problem (2.122), we want the firm to be a *price taker*

We do that by regarding p and therefore Y as exogenous to the firm

The essence of the Big Y , little y trick is *not* to set $Y = ny$ before taking the first-order condition with respect to y in problem (2.122)

This assures that the firm is a price taker

The first order condition for problem (2.122) is

$$a_0 - a_1Y - c_1 - c_2y = 0 \quad (2.123)$$

At this point, *but not before*, we substitute $Y = ny$ into (2.123) to obtain the following linear equation

$$a_0 - c_1 - (a_1 + n^{-1}c_2)Y = 0 \quad (2.124)$$

to be solved for the competitive equilibrium market wide output Y

After solving for Y , we can compute the competitive equilibrium price p from the inverse demand curve (2.121)

Further Reading References for this lecture include

- [LP71]
- [Sar87], chapter XIV
- [LS12], chapter 7

Defining Rational Expectations Equilibrium

Our first illustration of a rational expectations equilibrium involves a market with n firms, each of which seeks to maximize the discounted present value of profits in the face of adjustment costs

The adjustment costs induce the firms to make gradual adjustments, which in turn requires consideration of future prices

Individual firms understand that, via the inverse demand curve, the price is determined by the amounts supplied by other firms

Hence each firm wants to forecast future total industry supplies

In our context, a forecast is generated by a belief about the law of motion for the aggregate state
Rational expectations equilibrium prevails when this belief coincides with the actual law of motion generated by production choices induced by this belief

We formulate a rational expectations equilibrium in terms of a fixed point of an operator that maps beliefs into optimal beliefs

Competitive Equilibrium with Adjustment Costs To illustrate, consider a collection of n firms producing a homogeneous good that is sold in a competitive market.

Each of these n firms sells output y_t

The price p_t of the good lies on the inverse demand curve

$$p_t = a_0 - a_1 Y_t \quad (2.125)$$

where

- $a_i > 0$ for $i = 0, 1$
- $Y_t = ny_t$ is the market-wide level of output

The Firm's Problem Each firm is a price taker

While it faces no uncertainty, it does face adjustment costs

In particular, it chooses a production plan to maximize

$$\sum_{t=0}^{\infty} \beta^t r_t \quad (2.126)$$

where

$$r_t := p_t y_t - \frac{\gamma(y_{t+1} - y_t)^2}{2}, \quad y_0 \text{ given} \quad (2.127)$$

Regarding the parameters,

- $\beta \in (0, 1)$ is a discount factor
- $\gamma > 0$ measures the cost of adjusting the rate of output

Regarding timing, the firm observes p_t and y_t when it chooses y_{t+1} at time t

To state the firm's optimization problem completely requires that we specify dynamics for all state variables

This includes ones that the firm cares about but does not control like p_t

We turn to this problem now

Prices and Aggregate Output In view of (2.125), the firm's incentive to forecast the market price translates into an incentive to forecast aggregate output Y_t

Aggregate output depends on the choices of other firms

We assume that n is such a large number that the output of any single firm has a negligible effect on aggregate output

That justifies firms in regarding their forecasts of aggregate output as being unaffected by their own output decisions

The Firm's Beliefs We suppose the firm believes that market-wide output Y_t follows the law of motion

$$Y_{t+1} = H(Y_t) \quad (2.128)$$

where Y_0 is a known initial condition

The *belief function* H is an equilibrium object, and hence remains to be determined

Optimal Behavior Given Beliefs For now let's fix a particular belief H in (2.128) and investigate the firm's response to it

Let v be the optimal value function for the firm's problem given H

The value function satisfies the Bellman equation

$$v(y, Y) = \max_{y'} \left\{ a_0 y - a_1 y Y - \frac{\gamma(y' - y)^2}{2} + \beta v(y', H(Y)) \right\} \quad (2.129)$$

Let's denote the firm's optimal policy function by h , so that

$$y_{t+1} = h(y_t, Y_t) \quad (2.130)$$

where

$$h(y, Y) := \arg \max_{y'} \left\{ a_0 y - a_1 y Y - \frac{\gamma(y' - y)^2}{2} + \beta v(y', H(Y)) \right\} \quad (2.131)$$

Evidently v and h both depend on H

First-Order Characterization of h In what follows it will be helpful to have a second characterization of h , based on first order conditions

The first-order necessary condition for choosing y' is

$$-\gamma(y' - y) + \beta v_y(y', H(Y)) = 0 \quad (2.132)$$

An important useful envelope result of Benveniste-Scheinkman [BS79] implies that to differentiate v with respect to y we can naively differentiate the right side of (2.129), giving

$$v_y(y, Y) = a_0 - a_1 Y + \gamma(y' - y)$$

Substituting this equation into (2.132) gives the *Euler equation*

$$-\gamma(y_{t+1} - y_t) + \beta[a_0 - a_1 Y_{t+1} + \gamma(y_{t+2} - y_{t+1})] = 0 \quad (2.133)$$

The firm optimally sets an output path that satisfies (2.133), taking (2.128) as given, and subject to

- the initial conditions for (y_0, Y_0)
- the terminal condition $\lim_{t \rightarrow \infty} \beta^t y_t v_y(y_t, Y_t) = 0$

This last condition is called the *transversality condition*, and acts as a first-order necessary condition “at infinity”

The firm’s decision rule solves the difference equation (2.133) subject to the given initial condition y_0 and the transversality condition

Note that solving the Bellman equation (2.129) for v and then h in (2.131) yields a decision rule that automatically imposes both the Euler equation (2.133) and the transversality condition

The Actual Law of Motion for $\{Y_t\}$ As we’ve seen, a given belief translates into a particular decision rule h

Recalling that $Y_t = ny_t$, the *actual law of motion* for market-wide output is then

$$Y_{t+1} = nh(Y_t/n, Y_t) \quad (2.134)$$

Thus, when firms believe that the law of motion for market-wide output is (2.128), their optimizing behavior makes the actual law of motion be (2.134)

Definition of Rational Expectations Equilibrium A *rational expectations equilibrium* or *recursive competitive equilibrium* of the model with adjustment costs is a decision rule h and an aggregate law of motion H such that

1. Given belief H , the map h is the firm’s optimal policy function
2. The law of motion H satisfies $H(Y) = nh(Y/n, Y)$ for all Y

Thus, a rational expectations equilibrium equates the perceived and actual laws of motion (2.128) and (2.134)

Fixed point characterization As we’ve seen, the firm’s optimum problem induces a mapping Φ from a perceived law of motion H for market-wide output to an actual law of motion $\Phi(H)$

The mapping Φ is the composition of two operations, taking a perceived law of motion into a decision rule via (2.129)–(2.131), and a decision rule into an actual law via (2.134)

The H component of a rational expectations equilibrium is a fixed point of Φ

Computation of an Equilibrium

Now let’s consider the problem of computing the rational expectations equilibrium

Misbehavior of Φ Readers accustomed to dynamic programming arguments might try to address this problem by choosing some guess H_0 for the aggregate law of motion and then iterating with Φ

Unfortunately, the mapping Φ is not a contraction

In particular, there is no guarantee that direct iterations on Φ converge¹

Fortunately, there is another method that works here

The method exploits a general connection between equilibrium and Pareto optimality expressed in the fundamental theorems of welfare economics (see, e.g, [MCWG95])

Lucas and Prescott [LP71] used this method to construct a rational expectations equilibrium

The details follow

A Planning Problem Approach Our plan of attack is to match the Euler equations of the market problem with those for a single-agent choice problem

As we'll see, this planning problem can be solved by LQ control (linear regulator)

The optimal quantities from the planning problem are rational expectations equilibrium quantities

The rational expectations equilibrium price can be obtained as a shadow price in the planning problem

For convenience, in this section we set $n = 1$

We first compute a sum of consumer and producer surplus at time t

$$s(Y_t, Y_{t+1}) := \int_0^{Y_t} (a_0 - a_1 x) dx - \frac{\gamma(Y_{t+1} - Y_t)^2}{2} \quad (2.135)$$

The first term is the area under the demand curve, while the second measures the social costs of changing output

The *planning problem* is to choose a production plan $\{Y_t\}$ to maximize

$$\sum_{t=0}^{\infty} \beta^t s(Y_t, Y_{t+1})$$

subject to an initial condition for Y_0

Solution of the Planning Problem Evaluating the integral in (2.135) yields the quadratic form $a_0 Y_t - a_1 Y_t^2 / 2$

As a result, the Bellman equation for the planning problem is

$$V(Y) = \max_{Y'} \left\{ a_0 Y - \frac{a_1}{2} Y^2 - \frac{\gamma(Y' - Y)^2}{2} + \beta V(Y') \right\} \quad (2.136)$$

¹ A literature that studies whether models populated with agents who learn can converge to rational expectations equilibria features iterations on a modification of the mapping Φ that can be approximated as $\gamma\Phi + (1 - \gamma)I$. Here I is the identity operator and $\gamma \in (0, 1)$ is a *relaxation parameter*. See [MS89] and [EH01] for statements and applications of this approach to establish conditions under which collections of adaptive agents who use least squares learning converge to a rational expectations equilibrium.

The associated first order condition is

$$-\gamma(Y' - Y) + \beta V'(Y') = 0 \quad (2.137)$$

Applying the same Benveniste-Scheinkman formula gives

$$V'(Y) = a_0 - a_1 Y + \gamma(Y' - Y)$$

Substituting this into equation (2.137) and rearranging leads to the Euler equation

$$\beta a_0 + \gamma Y_t - [\beta a_1 + \gamma(1 + \beta)]Y_{t+1} + \gamma\beta Y_{t+2} = 0 \quad (2.138)$$

The Key Insight Return to equation (2.133) and set $y_t = Y_t$ for all t

(Recall that for this section we've set $n = 1$ to simplify the calculations)

A small amount of algebra will convince you that when $y_t = Y_t$, equations (2.138) and (2.133) are identical

Thus, the Euler equation for the planning problem matches the second-order difference equation that we derived by

1. finding the Euler equation of the representative firm and
2. substituting into it the expression $Y_t = ny_t$ that "makes the representative firm be representative"

If it is appropriate to apply the same terminal conditions for these two difference equations, which it is, then we have verified that a solution of the planning problem is also a rational expectations equilibrium quantity sequence

It follows that for this example we can compute equilibrium quantities by forming the optimal linear regulator problem corresponding to the Bellman equation (2.136)

The optimal policy function for the planning problem is the aggregate law of motion H that the representative firm faces within a rational expectations equilibrium.

Structure of the Law of Motion As you are asked to show in the exercises, the fact that the planner's problem is an LQ problem implies an optimal policy — and hence aggregate law of motion — taking the form

$$Y_{t+1} = \kappa_0 + \kappa_1 Y_t \quad (2.139)$$

for some parameter pair κ_0, κ_1

Now that we know the aggregate law of motion is linear, we can see from the firm's Bellman equation (2.129) that the firm's problem can also be framed as an LQ problem

As you're asked to show in the exercises, the LQ formulation of the firm's problem implies a law of motion that looks as follows

$$y_{t+1} = h_0 + h_1 y_t + h_2 Y_t \quad (2.140)$$

Hence a rational expectations equilibrium will be defined by the parameters $(\kappa_0, \kappa_1, h_0, h_1, h_2)$ in (2.139)–(2.140)

Exercises

Exercise 1 Consider the firm problem *described above*

Let the firm's belief function H be as given in (2.139)

Formulate the firm's problem as a discounted optimal linear regulator problem, being careful to describe all of the objects needed

Use the type LQ from the `QuantEcon.jl` package to solve the firm's problem for the following parameter values:

$$a_0 = 100, a_1 = 0.05, \beta = 0.95, \gamma = 10, \kappa_0 = 95.5, \kappa_1 = 0.95$$

Express the solution of the firm's problem in the form (2.140) and give the values for each h_j

If there were n identical competitive firms all behaving according to (2.140), what would (2.140) imply for the *actual* law of motion (2.128) for market supply

Exercise 2 Consider the following κ_0, κ_1 pairs as candidates for the aggregate law of motion component of a rational expectations equilibrium (see (2.139))

Extending the program that you wrote for exercise 1, determine which if any satisfy *the definition* of a rational expectations equilibrium

- (94.0886298678, 0.923409232937)
- (93.2119845412, 0.984323478873)
- (95.0818452486, 0.952459076301)

Describe an iterative algorithm that uses the program that you wrote for exercise 1 to compute a rational expectations equilibrium

(You are not being asked actually to use the algorithm you are suggesting)

Exercise 3 Recall the planner's problem *described above*

1. Formulate the planner's problem as an LQ problem
2. Solve it using the same parameter values in exercise 1
 - $a_0 = 100, a_1 = 0.05, \beta = 0.95, \gamma = 10$
3. Represent the solution in the form $Y_{t+1} = \kappa_0 + \kappa_1 Y_t$
4. Compare your answer with the results from exercise 2

Exercise 4 A monopolist faces the industry demand curve (2.125) and chooses $\{Y_t\}$ to maximize $\sum_{t=0}^{\infty} \beta^t r_t$ where

$$r_t = p_t Y_t - \frac{\gamma(Y_{t+1} - Y_t)^2}{2}$$

Formulate this problem as an LQ problem

Compute the optimal policy using the same parameters as the previous exercise

In particular, solve for the parameters in

$$Y_{t+1} = m_0 + m_1 Y_t$$

Compare your results with the previous exercise. Comment.

Solutions

[Solution notebook](#)

An Introduction to Asset Pricing

Contents

- *An Introduction to Asset Pricing*
 - [Overview](#)
 - [Pricing Models](#)
 - [Prices in the Risk Neutral Case](#)
 - [Asset Prices under Risk Aversion](#)
 - [Implementation](#)
 - [Exercises](#)
 - [Solutions](#)

“A little knowledge of geometric series goes a long way” – Robert E. Lucas, Jr.

“Asset pricing is all about covariances” – Lars Peter Hansen

Overview

An asset is a claim on one or more future payoffs

The spot price of an asset depends primarily on

- the anticipated dynamics for the stream of income accruing to the owners
- attitudes to risk
- rates of time preference

In this lecture we consider some standard pricing models and dividend stream specifications

We study how prices and dividend-price ratios respond in these different scenarios

We also look at creating and pricing *derivative* assets by repackaging income streams

Key tools for the lecture are

- formulas for predicting future values of functions of a Markov state
- a formula for predicting the discounted sum of future values of a Markov state

Pricing Models

In what follows let $\{d_t\}_{t \geq 0}$ be a stream of dividends

- A time- t **cum-dividend** asset is a claim to the stream d_t, d_{t+1}, \dots
- A time- t **ex-dividend** asset is a claim to the stream d_{t+1}, d_{t+2}, \dots

Let's look at some equations that we expect to hold for prices of assets under cum-dividend and ex-dividend contracts respectively

Risk Neutral Pricing Our first scenario is risk-neutral pricing

Let $\beta = 1/(1 + \rho)$ be an intertemporal discount factor, where ρ is the rate at which agents discount the future

The basic risk-neutral asset pricing equation for pricing one unit of a cum-dividend asset is

$$p_t = d_t + \beta \mathbb{E}_t[p_{t+1}] \quad (2.141)$$

This is a simple "cost equals expected benefit" relationship

Here $\mathbb{E}_t[y]$ denotes the best forecast of y , conditioned on information available at time t

For an ex-dividend asset, the basic risk-neutral asset pricing equation is

$$p_t = \beta \mathbb{E}_t[d_{t+1} + p_{t+1}] \quad (2.142)$$

Pricing with Random Discount Factor What happens if for some reason traders discount pay-outs differently depending on the state of the world?

Michael Harrison and David Kreps [HK79] and Lars Peter Hansen and Scott Richard [HR87] showed that in quite general settings the price of an ex-dividend asset obeys

$$p_t = \mathbb{E}_t[m_{t+1}(d_{t+1} + p_{t+1})] \quad (2.143)$$

for some **stochastic discount factor** m_{t+1}

The fixed discount factor β in (2.142) has been replaced by the random variable m_{t+1}

The way anticipated future payoffs are evaluated can now depend on various random outcomes

One example of this idea is that assets that tend to have good payoffs in bad states of the world might be regarded as more valuable

This is because they pay well when the funds are more urgently needed

We give examples of how the stochastic discount factor has been modeled below

Asset Pricing and Covariances Recall that, from the definition of a conditional covariance $\text{cov}_t(x_{t+1}, y_{t+1})$, we have

$$\mathbb{E}_t(x_{t+1}y_{t+1}) = \text{cov}_t(x_{t+1}, y_{t+1}) + \mathbb{E}_tx_{t+1}\mathbb{E}_ty_{t+1} \quad (2.144)$$

If we apply this definition to the asset pricing equation (2.143) we obtain

$$p_t = \mathbb{E}_tm_{t+1}\mathbb{E}_t(d_{t+1} + p_{t+1}) + \text{cov}_t(m_{t+1}, d_{t+1} + p_{t+1}) \quad (2.145)$$

It is useful to regard equation (2.145) as a generalization of equation (2.142)

- In equation (2.142), the stochastic discount factor $m_{t+1} = \beta$, a constant
- In equation (2.142), the covariance term $\text{cov}_t(m_{t+1}, d_{t+1} + p_{t+1})$ is zero because $m_{t+1} = \beta$

Equation (2.145) asserts that the covariance of the stochastic discount factor with the one period payout $d_{t+1} + p_{t+1}$ is an important determinant of the price p_t

We give examples of some models of stochastic discount factors that have been proposed later in this lecture and also in a [later lecture](#)

The Price-Dividend Ratio Aside from prices, another quantity of interest is the **price-dividend ratio** $v_t := p_t/d_t$

Let's write down some expressions that this ratio should satisfy

For the case of an ex-dividend contract, we can divide both sides of (2.143) by d_t to get

$$v_t = \mathbb{E}_t \left[m_{t+1} \frac{d_{t+1}}{d_t} (1 + v_{t+1}) \right] \quad (2.146)$$

For the cum-dividend case, the corresponding expression is

$$v_t = 1 + \mathbb{E}_t \left[m_{t+1} \frac{d_{t+1}}{d_t} v_{t+1} \right] \quad (2.147)$$

Below we'll discuss the implications of these equations

Prices in the Risk Neutral Case

What can we say about price dynamics on the basis of the models described above?

The answer to this question depends on

1. the process we specify for dividends
2. the stochastic discount factor and how it correlates with dividends

For now let's focus on the risk neutral case, where the stochastic discount factor is constant, and study how prices depend on the dividend process

Example 1: Constant dividends The simplest case is risk neutral pricing in the face of a constant, non-random dividend stream $d_t = d > 0$

Removing the expectation from (2.141) and iterating forward gives

$$\begin{aligned} p_t &= d + \beta p_{t+1} \\ &= d + \beta(d + \beta p_{t+2}) \\ &\vdots \\ &= d + \beta d + \beta^2 d + \cdots + \beta^{k-1} d + \beta^k p_{t+k} \end{aligned}$$

Unless prices explode in the future, this sequence converges to

$$\bar{p} := \frac{1}{1-\beta} d \quad (2.148)$$

This price is the equilibrium price in the constant dividend case

Indeed, simple algebra shows that setting $p_t = \bar{p}$ for all t satisfies the equilibrium condition $p_t = d + \beta p_{t+1}$

The ex-dividend equilibrium price is $(1 - \beta)^{-1} \beta d$

Example 2: Dividends with deterministic growth paths Consider a growing, non-random dividend process $d_{t+1} = g d_t$ where $0 < g\beta < 1$

While prices are not usually constant when dividends grow over time, the price dividend-ratio might be

If we guess this, substituting $v_t = v$ into (2.147) as well as our other assumptions, we get $v = 1 + \beta g v$

Since $\beta g < 1$, we have a unique positive solution for the cum-dividend case:

$$v = \frac{1}{1 - \beta g}$$

The cum-dividend price is then

$$p_t = \frac{1}{1 - \beta g} d_t \quad (2.149)$$

In view of (2.146), the ex-dividend formulas are

$$v = \frac{\beta g}{1 - \beta g} \quad \text{and} \quad p_t = \frac{\beta g}{1 - \beta g} d_t$$

If, in this example, we take $g = 1 + \kappa$ and let $\rho := 1/\beta - 1$, then the ex-dividend price becomes

$$p_t = \frac{1 + \kappa}{\rho - \kappa} d_t$$

This is called the *Gordon formula*

Example 3: Markov growth, risk neutral pricing Next we consider a dividend process

$$d_{t+1} = g_{t+1} d_t \quad (2.150)$$

The stochastic growth factor $\{g_t\}$ is given by

$$g_t = g(X_t), \quad t = 1, 2, \dots$$

where

1. $\{X_t\}$ is a finite Markov chain with state space S and transition probabilities

$$P(x, y) := \mathbb{P}\{X_{t+1} = y \mid X_t = x\} \quad (x, y \in S)$$

1. g is a given function on S taking positive values

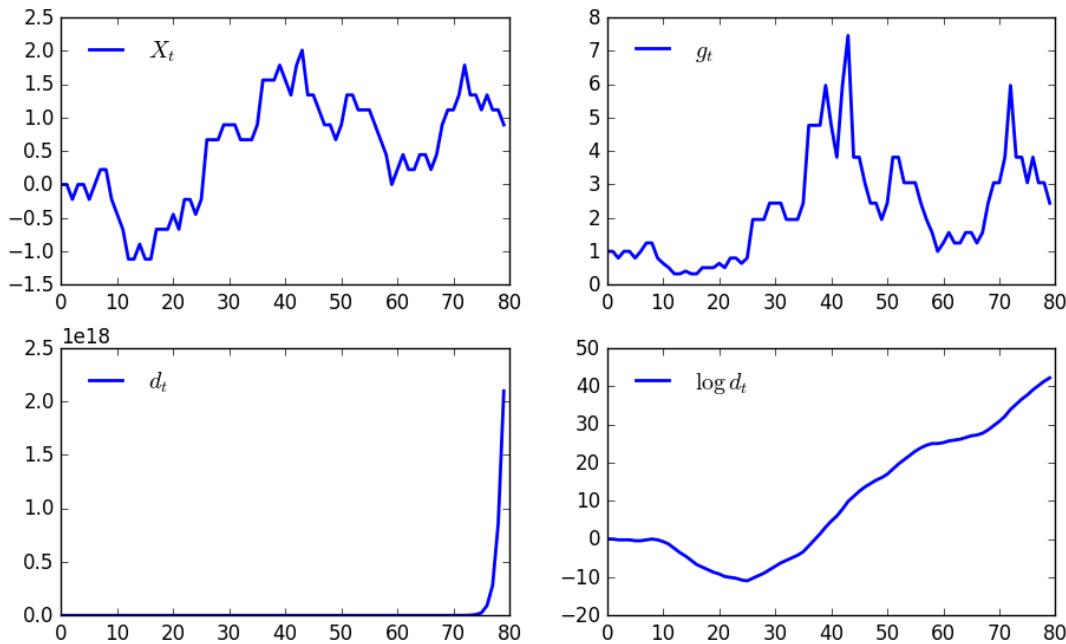
You can think of

- S as n possible “states of the world” and X_t as the current state
- g as a function that maps a given state X_t into a growth factor $g_t = g(X_t)$ for the endowment
- $\ln g_t = \ln(d_{t+1}/d_t)$ is the growth rate of dividends

(For a refresher on notation and theory for finite Markov chains see [this lecture](#))

The next figure shows a simulation, where

- $\{X_t\}$ evolves as a discretized AR1 process produced using *Tauchen's method*
- $g_t = \exp(X_t)$, so that $\ln g_t = X_t$ is the growth rate



The code can be found [here](#)

Pricing To obtain asset prices in this setting, let's adapt our analysis from the case of deterministic growth

In that case we found that v is constant

This encourages us to guess that, in the current case, v_t is constant given the state X_t

In other words, we are looking for a fixed function v such that the price-dividend ratio satisfies $v_t = v(X_t)$

Starting with the cum-dividend case, we can substitute this guess into (2.147) to get

$$v(X_t) = 1 + \beta \mathbb{E}_t[g(X_{t+1})v(X_{t+1})]$$

If we condition on $X_t = x$, this becomes

$$v(x) = 1 + \beta \sum_{y \in S} g(y)v(y)P(x, y)$$

or

$$v(x) = 1 + \beta \sum_{y \in S} K(x, y)v(y) \quad \text{where } K(x, y) := g(y)P(x, y) \quad (2.151)$$

Suppose that there are n possible states x_1, \dots, x_n

We can then think of (2.151) as n stacked equations, one for each state, and write it in matrix form as

$$v = \mathbb{1} + \beta Kv \quad (2.152)$$

Here

- v is understood to be the column vector $(v(x_1), \dots, v(x_n))'$
- K is the matrix $(K(x_i, x_j))_{1 \leq i, j \leq n}$
- $\mathbb{1}$ is a column vector of ones

When does (2.152) have a unique solution?

From the *Neumann series lemma* and Gelfand's formula, this will be the case if βK has spectral radius strictly less than one

In other words, we require that the eigenvalues of K be strictly less than β^{-1} in modulus

The solution is then

$$v = (I - \beta K)^{-1}\mathbb{1}$$

Similar reasoning in the ex-dividend case yields

$$v = (I - \beta K)^{-1}\beta K\mathbb{1} \quad (2.153)$$

Code Let's calculate and plot the price-dividend ratio for the ex-dividend case at a set of parameters

As before, we'll generate $\{X_t\}$ as a *discretized AR1 process* and set $g_t = \exp(X_t)$

Here's the code, including a test of the spectral radius condition

```
#=
```

Plot the price-dividend ratio in the risk neutral case, for the Markov asset pricing lecture.

```
=#
```

```
using QuantEcon
```

```
using Plots
```

```
using LaTeXStrings
```

```
pyplot()
```

```
n = 25 # size of state space
```

```
beta = 0.9
```

```
mc = tauchen(n, 0.96, 0.02)
```

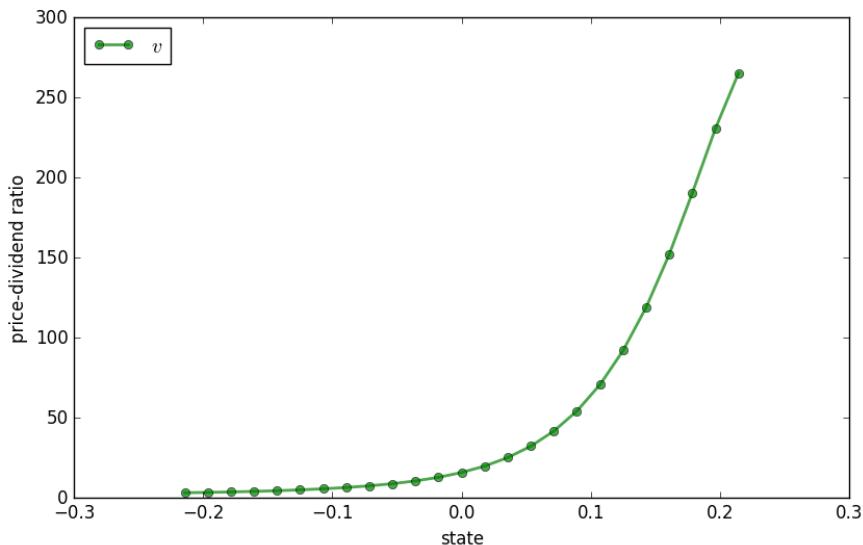
```
K = mc.p .* reshape(exp(mc.state_values), 1, n)
```

```
I = eye(n)
```

```
v = (I - beta * K) \ (beta * K * ones(n, 1))
```

```
plot(mc.state_values,
      v,
      lw=2,
      ylabel="price-dividend ratio",
      xlabel="state",
      alpha=0.7,
      label=L"\$v\$")
```

Here's the figure it produces



Why does the price-dividend ratio increase with the state?

The reason is that this Markov process is positively correlated, so high current states suggest high

future states

Moreover, dividend growth is increasing in the state

Anticipation of high future dividend growth leads to a high price-dividend ratio

Asset Prices under Risk Aversion

Now let's turn to the case where agents are risk averse

We'll price several distinct assets, including

- The price of an endowment stream
- A consol (a type of bond issued by the UK government in the 19th century)
- Call options on a consol

Pricing a Lucas tree Let's start with a version of the celebrated asset pricing model of Robert E. Lucas, Jr. [\[Luc78\]](#)

As in [\[Luc78\]](#), suppose that the stochastic discount factor takes the form

$$m_{t+1} = \beta \frac{u'(c_{t+1})}{u'(c_t)} \quad (2.154)$$

where u is a concave utility function and c_t is time t consumption of a representative consumer

(A derivation of this expression is given in a [later lecture](#))

Assume the existence of an endowment that follows [\(2.150\)](#)

The asset being priced is a claim on the endowment process

Following [\[Luc78\]](#), suppose further that in equilibrium, consumption is equal to the endowment, so that $d_t = c_t$ for all t

For utility, we'll assume the **constant relative risk aversion** (CRRA) specification

$$u(c) = \frac{c^{1-\gamma}}{1-\gamma} \text{ with } \gamma > 0 \quad (2.155)$$

When $\gamma = 1$ we let $u(c) = \ln c$

Inserting the CRRA specification into [\(2.154\)](#) and using $c_t = d_t$ gives

$$m_{t+1} = \beta \left(\frac{c_{t+1}}{c_t} \right)^{-\gamma} = \beta g_{t+1}^{-\gamma} \quad (2.156)$$

Substituting this into [\(2.146\)](#) gives the ex-dividend price-dividend ratio formula

$$v(X_t) = \beta \mathbb{E}_t \left[g(X_{t+1})^{1-\gamma} (1 + v(X_{t+1})) \right]$$

Conditioning on $X_t = x$, we can write this as

$$v(x) = \beta \sum_{y \in S} g(y)^{1-\gamma} (1 + v(y)) P(x, y)$$

If we let

$$J(x, y) := g(y)^{1-\gamma} P(x, y)$$

then we can rewrite in vector form as

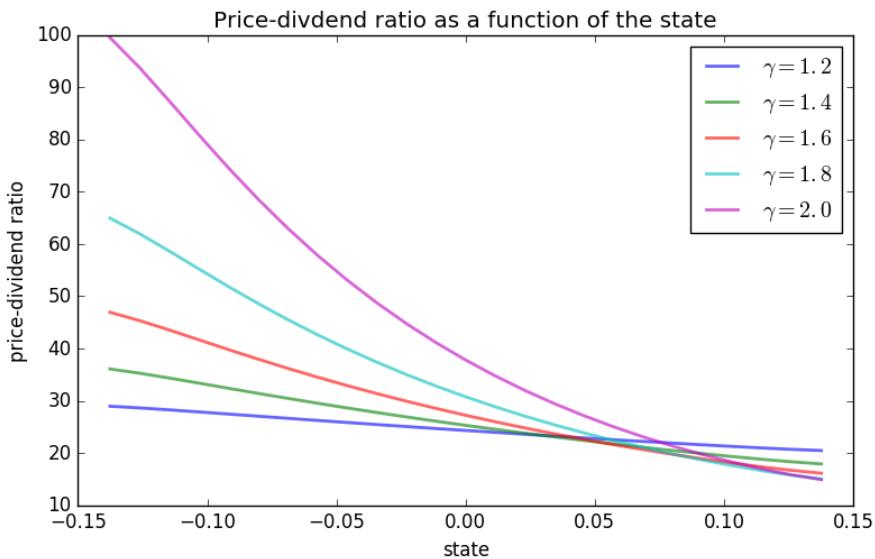
$$v = \beta J(\mathbb{1} + v)$$

Assuming that the spectral radius of J is strictly less than β^{-1} , this equation has the unique solution

$$v = (I - \beta J)^{-1} \beta J \mathbb{1} \quad (2.157)$$

Here's a plot of v as a function of the state for several values of γ , with a positively correlated Markov process and $g(x) = \exp(x)$

The code with all details can be found [here](#)



Notice that v is decreasing in each case

This is because, with a positively correlated state process, higher states suggest higher future consumption growth

In the stochastic discount factor (2.156), higher growth decreases the discount factor, lowering the weight placed on future returns

Special cases In the special case $\gamma = 1$, we have $J = P$

Recalling that $P^i \mathbb{1} = \mathbb{1}$ for all i and applying Neumann's geometric series lemma, we are led to

$$v = \beta(I - \beta P)^{-1} \mathbb{1} = \beta \sum_{i=0}^{\infty} \beta^i P^i \mathbb{1} = \beta \frac{1}{1 - \beta} \mathbb{1}$$

Thus, with log preferences, the price-dividend ratio for a Lucas tree is constant

Alternatively, if $\gamma = 0$, then $J = K$ and we recover the risk neutral solution (2.153)

This is as expected, since $\gamma = 0$ implies $u(c) = c$ (and hence agents are risk neutral)

A Risk-Free Consol Consider the same pure exchange representative agent economy

A risk-free consol promises to pay a constant amount $\zeta > 0$ each period

Recycling notation, let p_t now be the price of an ex-coupon claim to the consol

An ex-coupon claim to the consol entitles the owner at the end of period t to

- ζ in period $t + 1$, plus
- the right to sell the claim for p_{t+1} next period

The price satisfies (2.143) with $d_t = \zeta$, or

$$p_t = \mathbb{E}_t [m_{t+1}(\zeta + p_{t+1})]$$

We maintain the stochastic discount factor (2.156), so this becomes

$$p_t = \mathbb{E}_t [\beta g_{t+1}^{-\gamma} (\zeta + p_{t+1})] \quad (2.158)$$

Guessing a solution of the form $p_t = p(X_t)$ and conditioning on $X_t = x$, we get

$$p(x) = \beta \sum_{y \in S} g(y)^{-\gamma} (\zeta + p(y)) P(x, y)$$

Letting $M(x, y) = P(x, y)g(y)^{-\gamma}$ and rewriting in vector notation yields the solution

$$p = (I - \beta M)^{-1} \beta M \zeta \mathbf{1} \quad (2.159)$$

Pricing an Option to Purchase the Consol Let's now price options of varying maturity that give the right to purchase a consol at a price p_S

An infinite horizon call option We want to price an infinite horizon option to purchase a consol at a price p_S

The option entitles the owner at the beginning of a period either to

1. purchase the bond at price p_S now, or
2. Not to exercise the option now but to retain the right to exercise it later

Thus, the owner either *exercises* the option now, or chooses *not to exercise* and wait until next period

This is termed an infinite-horizon *call option* with *strike price* p_S

The owner of the option is entitled to purchase the consol at the price p_S at the beginning of any period, after the coupon has been paid to the previous owner of the bond

The fundamentals of the economy are identical with the one above, including the stochastic discount factor and the process for consumption

Let $w(X_t, p_S)$ be the value of the option when the time t growth state is known to be X_t but *before* the owner has decided whether or not to exercise the option at time t (i.e., today)

Recalling that $p(X_t)$ is the value of the consol when the initial growth state is X_t , the value of the option satisfies

$$w(X_t, p_S) = \max \left\{ \beta \mathbb{E}_t \frac{u'(c_{t+1})}{u'(c_t)} w(X_{t+1}, p_S), p(X_t) - p_S \right\}$$

The first term on the right is the value of waiting, while the second is the value of exercising now

We can also write this as

$$w(x, p_S) = \max \left\{ \beta \sum_{y \in S} P(x, y) g(y)^{-\gamma} w(y, p_S), p(x) - p_S \right\} \quad (2.160)$$

With $M(x, y) = P(x, y)g(y)^{-\gamma}$ and w as the vector of values $(w(x_i), p_S)_{i=1}^n$, we can express (2.160) as the nonlinear vector equation

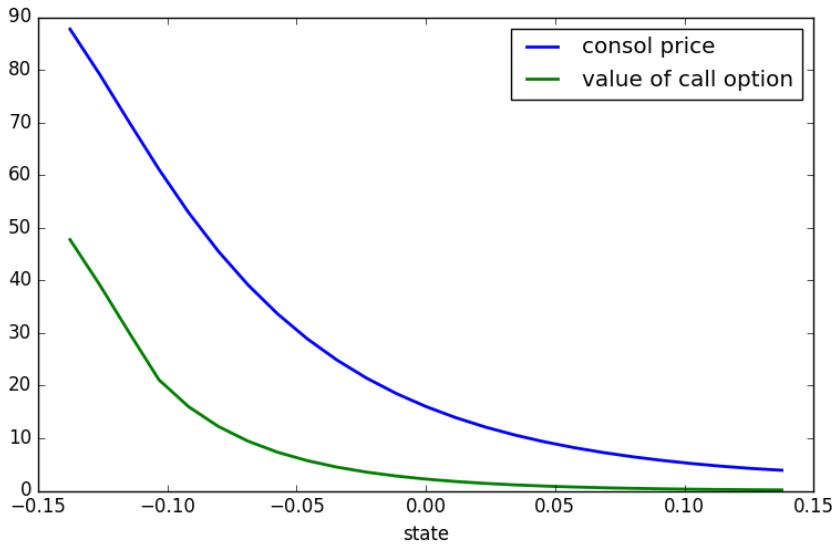
$$w = \max\{\beta Mw, p - p_S \mathbf{1}\} \quad (2.161)$$

To solve (2.161), form the operator T mapping vector w into vector Tw via

$$Tw = \max\{\beta Mw, p - p_S \mathbf{1}\}$$

Start at some initial w and iterate to convergence with T

Here's a plot of w compared to the consol price when $P_S = 40$



The code with all details can be found [here](#)

In large states the value of the option is close to zero

This is despite the fact the Markov chain is irreducible and low states — where the consol prices is high — will eventually be visited

The reason is that $\beta = 0.9$, so the future is discounted relatively rapidly

Risk Free Rates Let's look at risk free interest rates over different periods

The one-period risk-free interest rate As before, the stochastic discount factor is $m_{t+1} = \beta g_{t+1}^{-\gamma}$

It follows that the reciprocal R_t^{-1} of the gross risk-free interest rate R_t in state x is

$$\mathbb{E}_t m_{t+1} = \beta \sum_{y \in S} P(x, y) g(y)^{-\gamma}$$

We can write this as

$$m_1 = \beta M \mathbf{1}$$

where the i -th element of m_1 is the reciprocal of the one-period gross risk-free interest rate in state x_i

Other terms Let m_j be an $n \times 1$ vector whose i th component is the reciprocal of the j -period gross risk-free interest rate in state x_i

Then $m_1 = \beta M$, and $m_{j+1} = M m_j$ for $j \geq 1$

Implementation

The file `asset_pricing.py` from `QuantEcon.applications` provides some functions for computing prices of the Lucas tree, consol and call option described above

Its contents are as follows

```
#=
Filename: asset_pricing.jl
@authors: Spencer Lyon, Tom Sargent, John Stachurski
Computes asset prices with a Lucas style discount factor when the endowment obeys geometric growth driven by a finite state Markov chain. That is,
.. math::
d_{\{t+1\}} = g(X_{\{t+1\}}) d_t
where
* :math:`\{X_t\}` is a finite Markov chain with transition matrix P.
* :math:`g` is a given positive-valued function
```

```

References
-----
http://quant-econ.net/py/markov_asset.html

#=

using QuantEcon

# A default Markov chain for the state process
rho = 0.9
sigma = 0.02
n = 25
default_mc = tauchen(n, rho, sigma)

type AssetPriceModel
    beta :: Float64      # Discount factor
    gamma :: Float64     # Coefficient of risk aversion
    mc :: MarkovChain   # State process
    n :: Int             # Number of states
    g :: Function        # Function mapping states into growth rates
end

function AssetPriceModel(;beta=0.96, gamma=2.0, mc=default_mc, g=exp)
    n = size(mc.p)[1]
    return AssetPriceModel(beta, gamma, mc, n, g)
end

"""
Stability test for a given matrix Q.
"""
function test_stability(ap::AssetPriceModel, Q::Matrix)
    sr = maximum(abs(eigvals(Q)))
    if sr >= 1 / ap.beta
        msg = "Spectral radius condition failed with radius = $sr"
        throw(ArgumentError(msg))
    end
end

"""
Computes the price-dividend ratio of the Lucas tree.
"""

function tree_price(ap::AssetPriceModel)
    # == Simplify names, set up matrices == #
    beta, gamma, P, y = ap.beta, ap.gamma, ap.mc.p, ap.mc.state_values
    y = reshape(y, 1, ap.n)
    J = P .* ap.g(y).^ (1 - gamma)

    # == Make sure that a unique solution exists == #
    test_stability(ap, J)

```

```

# == Compute v == #
I = eye(ap.n)
Ones = ones(ap.n)
v = (I - beta * J) \ (beta * J * Ones)

return v
end

"""

Computes price of a consol bond with payoff zeta

"""

function consol_price(ap::AssetPriceModel, zeta::Float64)
    # == Simplify names, set up matrices == #
    beta, gamma, P, y = ap.beta, ap.gamma, ap.mc.p, ap.mc.state_values
    y = reshape(y, 1, ap.n)
    M = P .* ap.g(y).^( -gamma)

    # == Make sure that a unique solution exists == #
    test_stability(ap, M)

    # == Compute price == #
    I = eye(ap.n)
    Ones = ones(ap.n)
    p = (I - beta * M) \ (beta * zeta * M * Ones)

    return p
end

"""

Computes price of a perpetual call option on a consol bond.

"""

function call_option(ap::AssetPriceModel, zeta::Float64, p_s::Float64, epsilon=1e-7)

    # == Simplify names, set up matrices == #
    beta, gamma, P, y = ap.beta, ap.gamma, ap.mc.p, ap.mc.state_values
    y = reshape(y, 1, ap.n)
    M = P .* ap.g(y).^( -gamma)

    # == Make sure that a unique console price exists == #
    test_stability(ap, M)

    # == Compute option price == #
    p = consol_price(ap, zeta)
    w = zeros(ap.n, 1)
    error = epsilon + 1
    while (error > epsilon)
        # == Maximize across columns == #
        w_new = max(beta * M * w, p - p_s)
        # == Find maximal difference of each component and update == #
        error = norm(w - w_new)
        w = w_new
    end
end

```

```

    error = maximum(abs(w-w_new))
    w = w_new
end

return w
end

```

Exercise 1 asks you to make use of this code

Exercises

Exercise 1 Consider the following primitives

```

n = 5
P = 0.0125 .* ones(n, n)
P .+= diagm(0.95 - 0.0125 .* ones(5))
s = [1.05, 1.025, 1.0, 0.975, 0.95]
gamm = 2.0
bet = 0.94
zet = 1.0

```

Let g be defined by $g(x) = x$ (that is, g is the identity map)

Compute the price of the Lucas tree

Do the same for

- the price of the risk-free console when $\zeta = 1$
- the call option on the console when $\zeta = 1$ and $p_S = 150.0$

Exercise 2 Let's consider finite horizon call options, which are more common than the infinite horizon variety

Finite horizon options obey functional equations closely related to (2.160)

A k period option expires after k periods

If we view today as date zero, a k period option gives the owner the right to exercise the option to purchase the risk-free consol at the strike price p_S at dates $0, 1, \dots, k-1$

The option expires at time k

Thus, for $k = 1, 2, \dots$, let $w(x, k)$ be the value of a k -period option

It obeys

$$w(x, k) = \max \left\{ \beta \sum_{y \in S} P(x, y) g(y)^{-\gamma} w(y, k-1), p(x) - p_S \right\}$$

where $w(x, 0) = 0$ for all x

We can express the preceding as the sequence of nonlinear vector equations

$$w_k = \max\{\beta M w_{k-1}, p - p_S \mathbf{1}\} \quad k = 1, 2, \dots \quad \text{with } w_0 = 0$$

Write a function that computes w_k for any given k

Compute the value of the option with $k = 5$ and $k=25$ using parameter values as in Exercise 1

Is one higher than the other? Can you give intuition?

Solutions

[Solution notebook](#)

The Permanent Income Model

Contents

- *The Permanent Income Model*
 - [Overview](#)
 - [The Savings Problem](#)
 - [Alternative Representations](#)
 - [Two Classic Examples](#)
 - [Further Reading](#)
 - [Appendix: The Euler Equation](#)

Overview

This lecture describes a rational expectations version of the famous permanent income model of Friedman [[Fri56](#)]

Hall cast Friedman's model within a linear-quadratic setting [[Hal78](#)]

Like Hall, we formulate an infinite-horizon linear-quadratic savings problem

We use the model as a vehicle for illustrating

- alternative formulations of the *state* of a dynamic system
- the idea of *cointegration*
- impulse response functions
- the idea that changes in consumption are useful as predictors of movements in income

Background readings on the linear-quadratic-Gaussian permanent income model are Robert Hall's [[Hal78](#)] and chapter 2 of [[LS12](#)]

The Savings Problem

In this section we state and solve the savings and consumption problem faced by the consumer

Preliminaries The discussion below requires a casual familiarity with [martingales](#)

A discrete time martingale is a stochastic process (i.e., a sequence of random variables) $\{X_t\}$ with finite mean and satisfying

$$\mathbb{E}_t[X_{t+1}] = X_t, \quad t = 0, 1, 2, \dots$$

Here $\mathbb{E}_t := \mathbb{E}[\cdot | \mathcal{F}_t]$ is a mathematical expectation conditional on the time t information set \mathcal{F}_t

The latter is just a collection of random variables that the modeler declares to be visible at t

- When not explicitly defined, it is usually understood that $\mathcal{F}_t = \{X_t, X_{t-1}, \dots, X_0\}$

Martingales have the feature that the history of past outcomes provides no predictive power for changes between current and future outcomes

For example, the current wealth of a gambler engaged in a “fair game” has this property

One common class of martingales is the family of *random walks*

A *random walk* is a stochastic process $\{X_t\}$ that satisfies

$$X_{t+1} = X_t + w_{t+1}$$

for some iid zero mean *innovation* sequence $\{w_t\}$

Evidently X_t can also be expressed as

$$X_t = \sum_{j=1}^t w_j + X_0$$

Not every martingale arises as a random walk (see, for example, [Wald’s martingale](#))

The Decision Problem A consumer has preferences over consumption streams that are ordered by the utility functional

$$\mathbb{E}_0 \left[\sum_{t=0}^{\infty} \beta^t u(c_t) \right] \tag{2.162}$$

where

- \mathbb{E}_t is the mathematical expectation conditioned on the consumer’s time t information
- c_t is time t consumption
- u is a strictly concave one-period utility function
- $\beta \in (0, 1)$ is a discount factor

The consumer maximizes (2.162) by choosing a consumption, borrowing plan $\{c_t, b_{t+1}\}_{t=0}^{\infty}$ subject to the sequence of budget constraints

$$c_t + b_t = \frac{1}{1+r} b_{t+1} + y_t \quad t \geq 0 \tag{2.163}$$

Here

- y_t is an exogenous endowment process

- $r > 0$ is the risk-free interest rate
- b_t is one-period risk-free debt maturing at t

The consumer also faces initial conditions b_0 and y_0 , which can be fixed or random

Assumptions For the remainder of this lecture, we follow Friedman and Hall in assuming that $(1+r)^{-1} = \beta$

Regarding the endowment process, we assume it has the state-space representation

$$z_{t+1} = Az_t + Cw_{t+1} \quad (2.164)$$

$$y_t = Uz_t \quad (2.165)$$

where

- $\{w_t\}$ is an iid vector process with $\mathbb{E}w_t = 0$ and $\mathbb{E}w_tw_t' = I$
- the spectral radius of A satisfies $\rho(A) < 1/\beta$
- U is a selection vector that pins down y_t as a particular linear combination of the elements of z_t .

The restriction on $\rho(A)$ prevents income from growing so fast that some discounted geometric sums of some infinite sequences below become infinite

Regarding preferences, we assume the quadratic utility function

$$u(c_t) = -(c_t - \gamma)^2$$

where γ is a bliss level of consumption

Note: Along with this quadratic utility specification, we allow consumption to be negative. However, by choosing parameters appropriately, we can make the probability that the model generates negative consumption paths as low as desired.

Finally, we impose the *no Ponzi scheme* condition

$$\mathbb{E}_0 \left[\sum_{t=0}^{\infty} \beta^t b_t^2 \right] < \infty \quad (2.166)$$

This condition rules out an always-borrow scheme that would allow the household to enjoy bliss consumption forever

First Order Conditions First-order conditions for maximizing (2.162) subject to (2.163) are

$$\mathbb{E}_t[u'(c_{t+1})] = u'(c_t), \quad t = 0, 1, \dots \quad (2.167)$$

These equations are also known as the *Euler equations* for the model

If you're not sure where they come from, you can find a proof sketch in the *appendix*

With our quadratic preference specification, (2.167) has the striking implication that consumption follows a martingale:

$$\mathbb{E}_t[c_{t+1}] = c_t \quad (2.168)$$

(In fact quadratic preferences are *necessary* for this conclusion ¹)

One way to interpret (2.168) is that consumption will only change when “new information” about permanent income is revealed

These ideas will be clarified below

The Optimal Decision Rule Now let’s deduce the optimal decision rule ²

Note: One way to solve the consumer’s problem is to apply *dynamic programming* as in [this lecture](#). We do this later. But first we use an alternative approach that is revealing and shows the work that dynamic programming does for us automatically

In doing so, we need to combine

1. the optimality condition (2.168)
2. the period-by-period budget constraint (2.163), and
3. the boundary condition (2.166)

To accomplish this, observe first that (2.166) implies $\lim_{t \rightarrow \infty} \beta^t b_{t+1} = 0$

Using this restriction on the debt path and solving (2.163) forward yields

$$b_t = \sum_{j=0}^{\infty} \beta^j (y_{t+j} - c_{t+j}) \quad (2.169)$$

Take conditional expectations on both sides of (2.169) and use the martingale property of consumption and the *law of iterated expectations* to deduce

$$b_t = \sum_{j=0}^{\infty} \beta^j \mathbb{E}_t[y_{t+j}] - \frac{c_t}{1-\beta} \quad (2.170)$$

Expressed in terms of c_t we get

$$c_t = (1-\beta) \left[\sum_{j=0}^{\infty} \beta^j \mathbb{E}_t[y_{t+j}] - b_t \right] = \frac{r}{1+r} \left[\sum_{j=0}^{\infty} \beta^j \mathbb{E}_t[y_{t+j}] - b_t \right] \quad (2.171)$$

where the last equality uses $(1+r)\beta = 1$

These last two equations assert that consumption equals *economic income*

¹ A linear marginal utility is essential for deriving (2.168) from (2.167). Suppose instead that we had imposed the following more standard assumptions on the utility function: $u'(c) > 0, u''(c) < 0, u'''(c) > 0$ and required that $c \geq 0$. The Euler equation remains (2.167). But the fact that $u''' < 0$ implies via Jensen’s inequality that $\mathbb{E}_t[u'(c_{t+1})] > u'(\mathbb{E}_t[c_{t+1}])$. This inequality together with (2.167) implies that $\mathbb{E}_t[c_{t+1}] > c_t$ (consumption is said to be a ‘submartingale’), so that consumption stochastically diverges to $+\infty$. The consumer’s savings also diverge to $+\infty$.

² An optimal decision rule is a map from current state into current actions—in this case, consumption

- *financial wealth* equals $-b_t$
- *non-financial wealth* equals $\sum_{j=0}^{\infty} \beta^j \mathbb{E}_t[y_{t+j}]$
- A *marginal propensity to consume out of wealth* equals the interest factor $\frac{r}{1+r}$
- *economic income* equals
 - a constant marginal propensity to consume times the sum of nonfinancial wealth and financial wealth
 - the amount the household can consume while leaving its wealth intact

Reacting to the state The *state* vector confronting the household at t is $[b_t \ z_t]$

Here

- z_t is an *exogenous* component, unaffected by household behavior
- b_t is an *endogenous* component (since it depends on the decision rule)

Note that z_t contains all variables useful for forecasting the household's future endowment

It seems likely that current decisions c_t and b_{t+1} should be expressible as functions of z_t and b_t

This is indeed the case

In fact, from *this discussion* we see that

$$\sum_{j=0}^{\infty} \beta^j \mathbb{E}_t[y_{t+j}] = \mathbb{E}_t \left[\sum_{j=0}^{\infty} \beta^j y_{t+j} \right] = U(I - \beta A)^{-1} z_t$$

Combining this with (2.171) gives

$$c_t = \frac{r}{1+r} [U(I - \beta A)^{-1} z_t - b_t] \quad (2.172)$$

Using this equality to eliminate c_t in the budget constraint (2.163) gives

$$\begin{aligned} b_{t+1} &= (1+r)(b_t + c_t - y_t) \\ &= (1+r)b_t + r[U(I - \beta A)^{-1} z_t - b_t] - (1+r)Uz_t \\ &= b_t + U[r(I - \beta A)^{-1} - (1+r)I]z_t \\ &= b_t + U(I - \beta A)^{-1}(A - I)z_t \end{aligned}$$

To get from the second last to the last expression in this chain of equalities is not trivial

Try using the fact that $(1+r)\beta = 1$ and $(I - \beta A)^{-1} = \sum_{j=0}^{\infty} \beta^j A^j$

We've now successfully written c_t and b_{t+1} as functions of b_t and z_t

A State-Space Representation We can summarize our dynamics in the form of a linear state-space system governing consumption, debt and income:

$$z_{t+1} = Az_t + Cw_{t+1} \quad (2.173)$$

$$b_{t+1} = b_t + U[(I - \beta A)^{-1}(A - I)]z_t \quad (2.174)$$

$$y_t = Uz_t \quad (2.175)$$

$$c_t = (1 - \beta)[U(I - \beta A)^{-1}z_t - b_t] \quad (2.176)$$

To write this more succinctly, let

$$x_t = \begin{bmatrix} z_t \\ b_t \end{bmatrix}, \quad \tilde{A} = \begin{bmatrix} A & 0 \\ U(I - \beta A)^{-1}(A - I) & 1 \end{bmatrix}, \quad \tilde{C} = \begin{bmatrix} C \\ 0 \end{bmatrix}$$

and

$$\tilde{U} = \begin{bmatrix} U & 0 \\ (1 - \beta)U(I - \beta A)^{-1} & -(1 - \beta) \end{bmatrix}, \quad \tilde{y}_t = \begin{bmatrix} y_t \\ b_t \end{bmatrix}$$

Then we can express equation (2.173) as

$$x_{t+1} = \tilde{A}x_t + \tilde{C}w_{t+1} \quad (2.177)$$

$$\tilde{y}_t = \tilde{U}x_t \quad (2.178)$$

We can use the following formulas from [state-space representation](#) to compute population mean $\mu_t = \mathbb{E}x_t$ and covariance $\Sigma_t := \mathbb{E}[(x_t - \mu_t)(x_t - \mu_t)']$

$$\mu_{t+1} = \tilde{A}\mu_t \quad \text{with } \mu_0 \text{ given} \quad (2.179)$$

$$\Sigma_{t+1} = \tilde{A}\Sigma_t\tilde{A}' + \tilde{C}\tilde{C}' \quad \text{with } \Sigma_0 \text{ given} \quad (2.180)$$

We can then compute the mean and covariance of \tilde{y}_t from

$$\mu_{y,t} = \tilde{U}\mu_t\Sigma_{y,t} = \tilde{U}\Sigma_t\tilde{U}' \quad (2.181)$$

A Simple Example with iid Income To gain some preliminary intuition on the implications of (2.173), let's look at a highly stylized example where income is just iid

(Later examples will investigate more realistic income streams)

In particular, let $\{w_t\}_{t=1}^\infty$ be iid and scalar standard normal, and let

$$z_t = \begin{bmatrix} z_t^1 \\ 1 \end{bmatrix}, \quad A = \begin{bmatrix} 0 & 0 \\ 0 & 1 \end{bmatrix}, \quad U = [1 \quad \mu], \quad C = \begin{bmatrix} \sigma \\ 0 \end{bmatrix}$$

Finally, let $b_0 = z_0^1 = 0$

Under these assumptions we have $y_t = \mu + \sigma w_t \sim N(\mu, \sigma^2)$

Further, if you work through the state space representation, you will see that

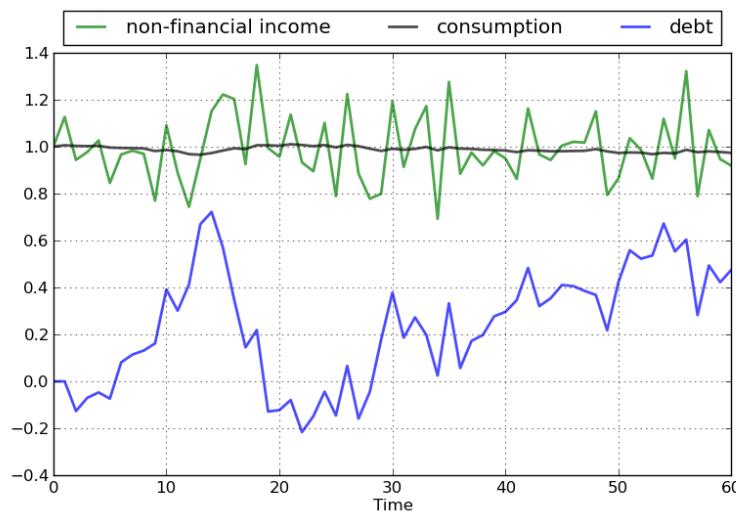
$$b_t = -\sigma \sum_{j=1}^{t-1} w_j$$

$$c_t = \mu + (1 - \beta)\sigma \sum_{j=1}^t w_j$$

Thus income is iid and debt and consumption are both Gaussian random walks

Defining assets as $-b_t$, we see that assets are just the cumulative sum of unanticipated income prior to the present date

The next figure shows a typical realization with $r = 0.05$, $\mu = 1$ and $\sigma = 0.15$



Observe that consumption is considerably smoother than income

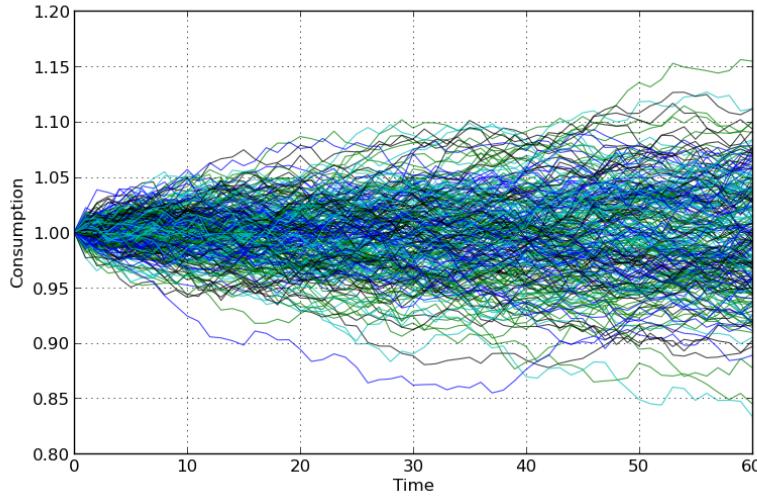
The figure below shows the consumption paths of 250 consumers with independent income streams

The code for these figures can be found in [perm_inc_figs.jl](#)

Alternative Representations

In this section we shed more light on the evolution of savings, debt and consumption by representing their dynamics in several different ways

Hall's Representation Hall [Hal78] suggests a sharp way to summarize the implications of LQ permanent income theory



First, to represent the solution for b_t , shift (2.171) forward one period and eliminate b_{t+1} by using (2.163) to obtain

$$c_{t+1} = (1 - \beta) \sum_{j=0}^{\infty} \beta^j \mathbb{E}_{t+1}[y_{t+j+1}] - (1 - \beta) [\beta^{-1}(c_t + b_t - y_t)]$$

If we add and subtract $\beta^{-1}(1 - \beta) \sum_{j=0}^{\infty} \beta^j \mathbb{E}_t y_{t+j}$ from the right side of the preceding equation and rearrange, we obtain

$$c_{t+1} - c_t = (1 - \beta) \sum_{j=0}^{\infty} \beta^j \{ \mathbb{E}_{t+1}[y_{t+j+1}] - \mathbb{E}_t[y_{t+j+1}] \} \quad (2.182)$$

The right side is the time $t + 1$ *innovation to the expected present value* of the endowment process $\{y_t\}$

We can represent the optimal decision rule for c_t, b_{t+1} in the form of (2.182) and (2.170), which is repeated here:

$$b_t = \sum_{j=0}^{\infty} \beta^j \mathbb{E}_t[y_{t+j}] - \frac{1}{1 - \beta} c_t \quad (2.183)$$

Equation (2.183) asserts that the household's debt due at t equals the expected present value of its endowment minus the expected present value of its consumption stream

A high debt thus indicates a large expected present value of surpluses $y_t - c_t$

Recalling again our discussion on *forecasting geometric sums*, we have

$$\begin{aligned} \mathbb{E}_t \sum_{j=0}^{\infty} \beta^j y_{t+j} &= U(I - \beta A)^{-1} z_t \\ \mathbb{E}_{t+1} \sum_{j=0}^{\infty} \beta^j y_{t+j+1} &= U(I - \beta A)^{-1} z_{t+1} \\ \mathbb{E}_t \sum_{j=0}^{\infty} \beta^j y_{t+j+1} &= U(I - \beta A)^{-1} A z_t \end{aligned}$$

Using these formulas together with (2.164) and substituting into (2.182) and (2.183) gives the following representation for the consumer's optimum decision rule:

$$c_{t+1} = c_t + (1 - \beta)U(I - \beta A)^{-1}Cw_{t+1} \quad (2.184)$$

$$b_t = U(I - \beta A)^{-1}z_t - \frac{1}{1 - \beta}c_t \quad (2.185)$$

$$y_t = Uz_t \quad (2.186)$$

$$z_{t+1} = Az_t + Cw_{t+1} \quad (2.187)$$

Representation (2.184) makes clear that

- The state can be taken as (c_t, z_t)
 - The endogenous part is c_t and the exogenous part is z_t
 - Debt b_t has disappeared as a component of the state because it is encoded in c_t
- Consumption is a random walk with innovation $(1 - \beta)U(I - \beta A)^{-1}Cw_{t+1}$
 - This is a more explicit representation of the martingale result in (2.168)

Cointegration Representation (2.184) reveals that the joint process $\{c_t, b_t\}$ possesses the property that Engle and Granger [EG87] called **cointegration**

Cointegration is a tool that allows us to apply powerful results from the theory of stationary processes to (certain transformations of) nonstationary models

To clarify cointegration in the present context, suppose that z_t is asymptotically stationary⁴

Despite this, both c_t and b_t will be non-stationary because they have unit roots (see (2.173) for b_t)

Nevertheless, there is a linear combination of c_t, b_t that is asymptotically stationary

In particular, from the second equality in (2.184) we have

$$(1 - \beta)b_t + c_t = (1 - \beta)U(I - \beta A)^{-1}z_t \quad (2.188)$$

Hence the linear combination $(1 - \beta)b_t + c_t$ is asymptotically stationary

Accordingly, Granger and Engle would call $[(1 - \beta) \ 1]$ a *cointegrating vector* for the state

When applied to the nonstationary vector process $[b_t \ c_t]'$, it yields a process that is asymptotically stationary

Equation (2.188) can be arranged to take the form

$$(1 - \beta)b_t + c_t = (1 - \beta)\mathbb{E}_t \sum_{j=0}^{\infty} \beta^j y_{t+j}, \quad (2.189)$$

Equation (2.189) asserts that the *cointegrating residual* on the left side equals the conditional expectation of the geometric sum of future incomes on the right⁶

⁴ This would be the case if, for example, the *spectral radius* of A is strictly less than one

⁶ See [JYC88], [LL01], [LL04] for interesting applications of related ideas.

Cross-Sectional Implications Consider again (2.184), this time in light of our discussion of distribution dynamics in the [lecture on linear systems](#)

The dynamics of c_t are given by

$$c_{t+1} = c_t + (1 - \beta)U(I - \beta A)^{-1}Cw_{t+1} \quad (2.190)$$

or

$$c_t = c_0 + \sum_{j=1}^t \hat{w}_j \quad \text{for } \hat{w}_{t+1} := (1 - \beta)U(I - \beta A)^{-1}Cw_{t+1}$$

The unit root affecting c_t causes the time t variance of c_t to grow linearly with t

In particular, since $\{\hat{w}_t\}$ is iid, we have

$$\text{Var}[c_t] = \text{Var}[c_0] + t \hat{\sigma}^2 \quad (2.191)$$

when

$$\hat{\sigma}^2 := (1 - \beta)^2 U(I - \beta A)^{-1} C C' (I - \beta A')^{-1} U'$$

Assuming that $\hat{\sigma} > 0$, this means that $\{c_t\}$ has no asymptotic distribution

Let's consider what this means for a cross-section of ex ante identical households born at time 0

Let the distribution of c_0 represent the cross-section of initial consumption values

Equation (2.191) tells us that the distribution of c_t spreads out over time at a rate proportional to t

A number of different studies have investigated this prediction (see, e.g., [\[DP94\]](#), [\[STY04\]](#))

Impulse Response Functions Impulse response functions measure the change in a dynamic system subject to a given impulse (i.e., temporary shock)

The impulse response function of $\{c_t\}$ to the innovation $\{w_t\}$ is a box

In particular, the response of c_{t+j} to a unit increase in the innovation w_{t+1} is $(1 - \beta)U(I - \beta A)^{-1}C$ for all $j \geq 1$

Moving Average Representation It's useful to express the innovation to the expected present value of the endowment process in terms of a moving average representation for income y_t

The endowment process defined by (2.164) has the moving average representation

$$y_{t+1} = d(L)w_{t+1} \quad (2.192)$$

where

- $d(L) = \sum_{j=0}^{\infty} d_j L^j$ for some sequence d_j , where L is the lag operator ³
- at time t , the household has an information set ⁵ $w^t = [w_t, w_{t-1}, \dots]$

³ Representation (2.164) implies that $d(L) = U(I - AL)^{-1}C$.

⁵ A moving average representation for a process y_t is said to be *fundamental* if the linear space spanned by y^t is equal to the linear space spanned by w^t . A time-invariant innovations representation, attained via the Kalman filter, is by construction fundamental.

Notice that

$$y_{t+j} - \mathbb{E}_t[y_{t+j}] = d_0 w_{t+j} + d_1 w_{t+j-1} + \cdots + d_{j-1} w_{t+1}$$

It follows that

$$\mathbb{E}_{t+1}[y_{t+j}] - \mathbb{E}_t[y_{t+j}] = d_{j-1} w_{t+1} \quad (2.193)$$

Using (2.193) in (2.182) gives

$$c_{t+1} - c_t = (1 - \beta)d(\beta)w_{t+1} \quad (2.194)$$

The object $d(\beta)$ is the *present value of the moving average coefficients* in the representation for the endowment process y_t

Two Classic Examples

We illustrate some of the preceding ideas with the following two examples

In both examples, the endowment follows the process $y_t = z_{1t} + z_{2t}$ where

$$\begin{bmatrix} z_{1t+1} \\ z_{2t+1} \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ 0 & 0 \end{bmatrix} \begin{bmatrix} z_{1t} \\ z_{2t} \end{bmatrix} + \begin{bmatrix} \sigma_1 & 0 \\ 0 & \sigma_2 \end{bmatrix} \begin{bmatrix} w_{1t+1} \\ w_{2t+1} \end{bmatrix}$$

Here

- w_{t+1} is an iid 2×1 process distributed as $N(0, I)$
- z_{1t} is a permanent component of y_t
- z_{2t} is a purely transitory component

Example 1 Assume as before that the consumer observes the state z_t at time t

In view of (2.184) we have

$$c_{t+1} - c_t = \sigma_1 w_{1t+1} + (1 - \beta)\sigma_2 w_{2t+1} \quad (2.195)$$

Formula (2.195) shows how an increment $\sigma_1 w_{1t+1}$ to the permanent component of income z_{1t+1} leads to

- a permanent one-for-one increase in consumption and
- no increase in savings $-b_{t+1}$

But the purely transitory component of income $\sigma_2 w_{2t+1}$ leads to a permanent increment in consumption by a fraction $1 - \beta$ of transitory income

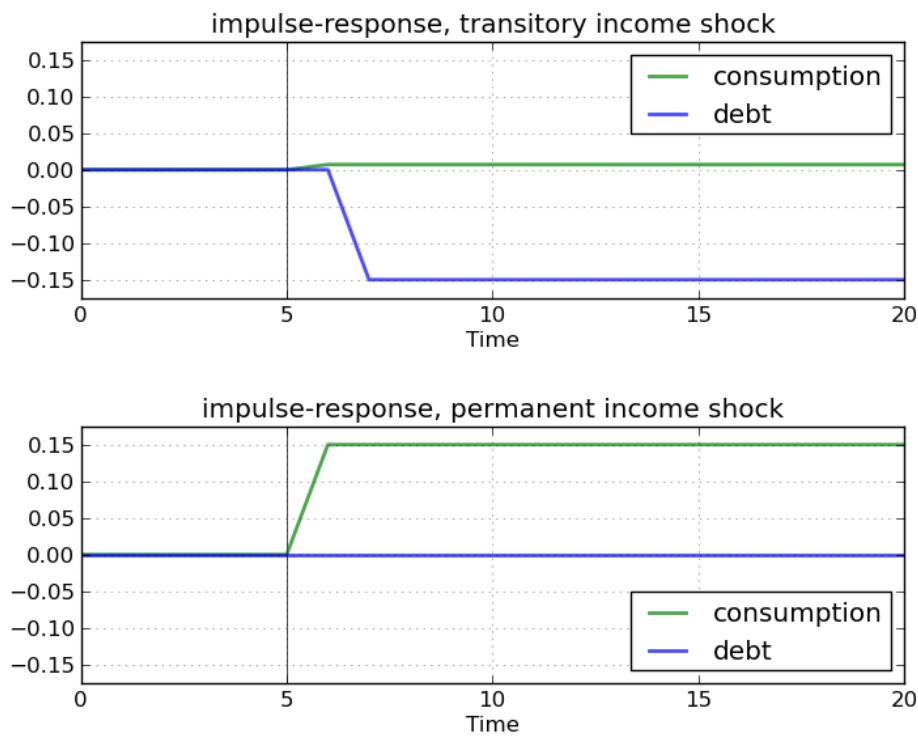
The remaining fraction β is saved, leading to a permanent increment in $-b_{t+1}$

Application of the formula for debt in (2.173) to this example shows that

$$b_{t+1} - b_t = -z_{2t} = -\sigma_2 w_{2t} \quad (2.196)$$

This confirms that none of $\sigma_1 w_{1t}$ is saved, while all of $\sigma_2 w_{2t}$ is saved

The next figure illustrates these very different reactions to transitory and permanent income shocks using impulse-response functions



The code for generating this figure is in file `perm_income/perm_inc_ir.jl` from the [applications repository](#), as shown below

```
#=
@author : Spencer Lyon
          Victoria Gregory

@date: 07/09/2014
=#

using Plots
pyplot()

const r = 0.05
const beta = 1.0 / (1.0 + r)
const T = 20 # Time horizon
const S = 5 # Impulse date
const sigma1 = 0.15
const sigma2 = 0.15

function time_path(permanent=false)
    w1 = zeros(T+1)
    w2 = zeros(T+1)
    b = zeros(T+1)
    c = zeros(T+1)
```

```

if permanent === false
    w2[S+2] = 1.0
else
    w1[S+2] = 1.0
end

for t=2:T
    b[t+1] = b[t] - sigma2 * w2[t]
    c[t+1] = c[t] + sigma1 * w1[t+1] + (1 - beta) * sigma2 * w2[t+1]
end

return b, c
end

function main()
L = 0.175

b1, c1 = time_path(false)
b2, c2 = time_path(true)
p = plot(0:T, [c1 c2 b1 b2], layout=(2, 1),
         color=[:green :green :blue :blue],
         label=["consumption" "consumption" "debt" "debt"])
t = ["impulse-response, transitory income shock"
     "impulse-response, permanent income shock"]
plot!(title=t', xlabel="Time", ylims=(-L, L), legend=[:topright :bottomright])
vline!([S S], color=:black, layout=(2, 1), label="")

return p
end

```

Example 2 Assume now that at time t the consumer observes y_t , and its history up to t , but not z_t

Under this assumption, it is appropriate to use an *innovation representation* to form A, C, U in (2.184)

The discussion in sections 2.9.1 and 2.11.3 of [LS12] shows that the pertinent state space representation for y_t is

$$\begin{bmatrix} y_{t+1} \\ a_{t+1} \end{bmatrix} = \begin{bmatrix} 1 & -(1-K) \\ 0 & 0 \end{bmatrix} \begin{bmatrix} y_t \\ a_t \end{bmatrix} + \begin{bmatrix} 1 \\ 1 \end{bmatrix} a_{t+1}$$

$$y_t = [1 \ 0] \begin{bmatrix} y_t \\ a_t \end{bmatrix}$$

where

- $K :=$ the stationary Kalman gain
- $a_t := y_t - E[y_t | y_{t-1}, \dots, y_0]$

In the same discussion in [LS12] it is shown that $K \in [0, 1]$ and that K increases as σ_1 / σ_2 does

In other words, as the ratio of the standard deviation of the permanent shock to that of the transitory shock increases

Applying formulas (2.184) implies

$$c_{t+1} - c_t = [1 - \beta(1 - K)]a_{t+1} \quad (2.197)$$

where the endowment process can now be represented in terms of the univariate innovation to y_t as

$$y_{t+1} - y_t = a_{t+1} - (1 - K)a_t \quad (2.198)$$

Equation (2.198) indicates that the consumer regards

- fraction K of an innovation a_{t+1} to y_{t+1} as *permanent*
- fraction $1 - K$ as purely transitory

The consumer permanently increases his consumption by the full amount of his estimate of the permanent part of a_{t+1} , but by only $(1 - \beta)$ times his estimate of the purely transitory part of a_{t+1}

Therefore, in total he permanently increments his consumption by a fraction $K + (1 - \beta)(1 - K) = 1 - \beta(1 - K)$ of a_{t+1}

He saves the remaining fraction $\beta(1 - K)$

According to equation (2.198), the first difference of income is a first-order moving average

Equation (2.197) asserts that the first difference of consumption is iid

Application of formula to this example shows that

$$b_{t+1} - b_t = (K - 1)a_t \quad (2.199)$$

This indicates how the fraction K of the innovation to y_t that is regarded as permanent influences the fraction of the innovation that is saved

Further Reading

The model described above significantly changed how economists think about consumption

At the same time, it's generally recognized that Hall's version of the permanent income hypothesis fails to capture all aspects of the consumption/savings data

For example, liquidity constraints and buffer stock savings appear to be important

Further discussion can be found in, e.g., [HM82], [Par99], [Dea91], [Car01]

Appendix: The Euler Equation

Where does the first order condition (2.167) come from?

Here we'll give a proof for the two period case, which is representative of the general argument

The finite horizon equivalent of the no-Ponzi condition is that the agent cannot end her life in debt, so $b_2 = 0$

From the budget constraint (2.163) we then have

$$c_0 = \frac{b_1}{1+r} - b_0 + y_0 \quad \text{and} \quad c_1 = y_1 - b_1$$

Here b_0 and y_0 are given constants

Subsituting these constraints into our two period objective $u(c_0) + \beta \mathbb{E}_0[u(c_1)]$ gives

$$\max_{b_1} \left\{ u \left(\frac{b_1}{R} - b_0 + y_0 \right) + \beta \mathbb{E}_0[u(y_1 - b_1)] \right\}$$

You will be able to verify that the first order condition is

$$u'(c_0) = \beta R \mathbb{E}_0[u'(c_1)]$$

Using $\beta R = 1$ gives (2.167) in the two period case

The proof for the general case is similar

ADVANCED APPLICATIONS

This advanced section of the course contains more complex applications, and can be read selectively, according to your interests

Continuous State Markov Chains

Contents

- *Continuous State Markov Chains*
 - *Overview*
 - *The Density Case*
 - *Beyond Densities*
 - *Stability*
 - *Exercises*
 - *Solutions*
 - *Appendix*

Overview

In a [previous lecture](#) we learned about finite Markov chains, a relatively elementary class of stochastic dynamic models

The present lecture extends this analysis to continuous (i.e., uncountable) state Markov chains

Most stochastic dynamic models studied by economists either fit directly into this class or can be represented as continuous state Markov chains after minor modifications

In this lecture, our focus will be on continuous Markov models that

- evolve in discrete time
- are often nonlinear

The fact that we accommodate nonlinear models here is significant, because linear stochastic models have their own highly developed tool set, as we'll see [later on](#)

The question that interests us most is: Given a particular stochastic dynamic model, how will the state of the system evolve over time?

In particular,

- What happens to the distribution of the state variables?
- Is there anything we can say about the “average behavior” of these variables?
- Is there a notion of “steady state” or “long run equilibrium” that’s applicable to the model?
 - If so, how can we compute it?

Answering these questions will lead us to revisit many of the topics that occupied us in the finite state case, such as simulation, distribution dynamics, stability, ergodicity, etc.

Note: For some people, the term “Markov chain” always refers to a process with a finite or discrete state space. We follow the mainstream mathematical literature (e.g., [MT09]) in using the term to refer to any discrete **time** Markov process

The Density Case

You are probably aware that some distributions can be represented by densities and some cannot (For example, distributions on the real numbers \mathbb{R} that put positive probability on individual points have no density representation)

We are going to start our analysis by looking at Markov chains where the one step transition probabilities have density representations

The benefit is that the density case offers a very direct parallel to the finite case in terms of notation and intuition

Once we’ve built some intuition we’ll cover the general case

Definitions and Basic Properties In our [lecture on finite Markov chains](#), we studied discrete time Markov chains that evolve on a finite state space S

In this setting, the dynamics of the model are described by a stochastic matrix — a nonnegative square matrix $P = P[i, j]$ such that each row $P[i, \cdot]$ sums to one

The interpretation of P is that $P[i, j]$ represents the probability of transitioning from state i to state j in one unit of time

In symbols,

$$\mathbb{P}\{X_{t+1} = j \mid X_t = i\} = P[i, j]$$

Equivalently,

- P can be thought of as a family of distributions $P[i, \cdot]$, one for each $i \in S$
- $P[i, \cdot]$ is the distribution of X_{t+1} given $X_t = i$

(As you probably recall, when using Julia arrays, $P[i, \cdot]$ is expressed as $P[i, :]$)

In this section, we'll allow S to be a subset of \mathbb{R} , such as

- \mathbb{R} itself
- the positive reals $(0, \infty)$
- a bounded interval (a, b)

The family of discrete distributions $P[i, \cdot]$ will be replaced by a family of densities $p(x, \cdot)$, one for each $x \in S$

Analogous to the finite state case, $p(x, \cdot)$ is to be understood as the distribution (density) of X_{t+1} given $X_t = x$

More formally, a *stochastic kernel on S* is a function $p: S \times S \rightarrow \mathbb{R}$ with the property that

1. $p(x, y) \geq 0$ for all $x, y \in S$
2. $\int p(x, y) dy = 1$ for all $x \in S$

(Integrals are over the whole space unless otherwise specified)

For example, let $S = \mathbb{R}$ and consider the particular stochastic kernel p_w defined by

$$p_w(x, y) := \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{(y-x)^2}{2}\right\} \quad (3.1)$$

What kind of model does p_w represent?

The answer is, the (normally distributed) random walk

$$X_{t+1} = X_t + \xi_{t+1} \quad \text{where} \quad \{\xi_t\} \stackrel{\text{IID}}{\sim} N(0, 1) \quad (3.2)$$

To see this, let's find the stochastic kernel p corresponding to (3.2)

Recall that $p(x, \cdot)$ represents the distribution of X_{t+1} given $X_t = x$

Letting $X_t = x$ in (3.2) and considering the distribution of X_{t+1} , we see that $p(x, \cdot) = N(x, 1)$

In other words, p is exactly p_w , as defined in (3.1)

Connection to Stochastic Difference Equations In the previous section, we made the connection between stochastic difference equation (3.2) and stochastic kernel (3.1)

In economics and time series analysis we meet stochastic difference equations of all different shapes and sizes

It will be useful for us if we have some systematic methods for converting stochastic difference equations into stochastic kernels

To this end, consider the generic (scalar) stochastic difference equation given by

$$X_{t+1} = \mu(X_t) + \sigma(X_t) \xi_{t+1} \quad (3.3)$$

Here we assume that

- $\{\xi_t\} \stackrel{\text{IID}}{\sim} \phi$, where ϕ is a given density on \mathbb{R}
- μ and σ are given functions on S , with $\sigma(x) > 0$ for all x

Example 1: The random walk (3.2) is a special case of (3.3), with $\mu(x) = x$ and $\sigma(x) = 1$

Example 2: Consider the ARCH model

$$X_{t+1} = \alpha X_t + \sigma_t \xi_{t+1}, \quad \sigma_t^2 = \beta + \gamma X_t^2, \quad \beta, \gamma > 0$$

Alternatively, we can write the model as

$$X_{t+1} = \alpha X_t + (\beta + \gamma X_t^2)^{1/2} \xi_{t+1} \quad (3.4)$$

This is a special case of (3.3) with $\mu(x) = \alpha x$ and $\sigma(x) = (\beta + \gamma x^2)^{1/2}$ **Example 3:** With stochastic production and a constant savings rate, the one-sector neoclassical growth model leads to a law of motion for capital per worker such as

$$k_{t+1} = s A_{t+1} f(k_t) + (1 - \delta) k_t \quad (3.5)$$

Here

- s is the rate of savings
- A_{t+1} is a production shock
 - The $t + 1$ subscript indicates that A_{t+1} is not visible at time t
- δ is a depreciation rate
- $f: \mathbb{R}_+ \rightarrow \mathbb{R}_+$ is a production function satisfying $f(k) > 0$ whenever $k > 0$

(The fixed savings rate can be rationalized as the optimal policy for a particular set of technologies and preferences (see [LS12], section 3.1.2), although we omit the details here)

Equation (3.5) is a special case of (3.3) with $\mu(x) = (1 - \delta)x$ and $\sigma(x) = sf(x)$

Now let's obtain the stochastic kernel corresponding to the generic model (3.3)

To find it, note first that if U is a random variable with density f_U , and $V = a + bU$ for some constants a, b with $b > 0$, then the density of V is given by

$$f_V(v) = \frac{1}{b} f_U \left(\frac{v - a}{b} \right) \quad (3.6)$$

(The proof is *below*. For a multidimensional version see EDTC, theorem 8.1.3)

Taking (3.6) as given for the moment, we can obtain the stochastic kernel p for (3.3) by recalling that $p(x, \cdot)$ is the conditional density of X_{t+1} given $X_t = x$

In the present case, this is equivalent to stating that $p(x, \cdot)$ is the density of $Y := \mu(x) + \sigma(x) \xi_{t+1}$ when $\xi_{t+1} \sim \phi$

Hence, by (3.6),

$$p(x, y) = \frac{1}{\sigma(x)} \phi \left(\frac{y - \mu(x)}{\sigma(x)} \right) \quad (3.7)$$

For example, the growth model in (3.5) has stochastic kernel

$$p(x, y) = \frac{1}{sf(x)} \phi \left(\frac{y - (1 - \delta)x}{sf(x)} \right) \quad (3.8)$$

where ϕ is the density of A_{t+1}

(Regarding the state space S for this model, a natural choice is $(0, \infty)$ — in which case $\sigma(x) = sf(x)$ is strictly positive for all s as required)

Distribution Dynamics In this section of our lecture on **finite** Markov chains, we asked the following question: If

1. $\{X_t\}$ is a Markov chain with stochastic matrix P
2. the distribution of X_t is known to be ψ_t

then what is the distribution of X_{t+1} ?

Letting ψ_{t+1} denote the distribution of X_{t+1} , the answer we gave was that

$$\psi_{t+1}[j] = \sum_{i \in S} P[i, j] \psi_t[i]$$

This intuitive equality states that the probability of being at j tomorrow is the probability of visiting i today and then going on to j , summed over all possible i

In the density case, we just replace the sum with an integral and probability mass functions with densities, yielding

$$\psi_{t+1}(y) = \int p(x, y) \psi_t(x) dx, \quad \forall y \in S \quad (3.9)$$

It is convenient to think of this updating process in terms of an operator

(An operator is just a function, but the term is usually reserved for a function that sends functions into functions)

Let \mathcal{D} be the set of all densities on S , and let P be the operator from \mathcal{D} to itself that takes density ψ and sends it into new density ψP , where the latter is defined by

$$(\psi P)(y) = \int p(x, y) \psi(x) dx \quad (3.10)$$

This operator is usually called the *Markov operator* corresponding to p

Note: Unlike most operators, we write P to the right of its argument, instead of to the left (i.e., ψP instead of $P\psi$). This is a common convention, with the intention being to maintain the parallel with the finite case — see [here](#)

With this notation, we can write (3.9) more succinctly as $\psi_{t+1}(y) = (\psi_t P)(y)$ for all y , or, dropping the y and letting “=” indicate equality of functions,

$$\psi_{t+1} = \psi_t P \quad (3.11)$$

Equation (3.11) tells us that if we specify a distribution for ψ_0 , then the entire sequence of future distributions can be obtained by iterating with P

It's interesting to note that (3.11) is a deterministic difference equation

Thus, by converting a stochastic difference equation such as (3.3) into a stochastic kernel p and hence an operator P , we convert a stochastic difference equation into a deterministic one (albeit in a much higher dimensional space)

Note: Some people might be aware that discrete Markov chains are in fact a special case of the continuous Markov chains we have just described. The reason is that probability mass functions are densities with respect to the [counting measure](#).

Computation To learn about the dynamics of a given process, it's useful to compute and study the sequences of densities generated by the model

One way to do this is to try to implement the iteration described by (3.10) and (3.11) using numerical integration

However, to produce ψP from ψ via (3.10), you would need to integrate at every y , and there is a continuum of such y

Another possibility is to discretize the model, but this introduces errors of unknown size

A nicer alternative in the present setting is to combine simulation with an elegant estimator called the *look ahead* estimator

Let's go over the ideas with reference to the growth model *discussed above*, the dynamics of which we repeat here for convenience:

$$k_{t+1} = sA_{t+1}f(k_t) + (1 - \delta)k_t \quad (3.12)$$

Our aim is to compute the sequence $\{\psi_t\}$ associated with this model and fixed initial condition ψ_0

To approximate ψ_t by simulation, recall that, by definition, ψ_t is the density of k_t given $k_0 \sim \psi_0$

If we wish to generate observations of this random variable, all we need to do is

1. draw k_0 from the specified initial condition ψ_0
2. draw the shocks A_1, \dots, A_t from their specified density ϕ
3. compute k_t iteratively via (3.12)

If we repeat this n times, we get n independent observations k_t^1, \dots, k_t^n

With these draws in hand, the next step is to generate some kind of representation of their distribution ψ_t

A naive approach would be to use a histogram, or perhaps a [smoothed histogram](#) using the `kde` function from [KernelDensity.jl](#)

However, in the present setting there is a much better way to do this, based on the look-ahead estimator

With this estimator, to construct an estimate of ψ_t , we actually generate n observations of k_{t-1} , rather than k_t

Now we take these n observations $k_{t-1}^1, \dots, k_{t-1}^n$ and form the estimate

$$\psi_t^n(y) = \frac{1}{n} \sum_{i=1}^n p(k_{t-1}^i, y) \quad (3.13)$$

where p is the growth model stochastic kernel in (3.8)

What is the justification for this slightly surprising estimator?

The idea is that, by the strong *law of large numbers*,

$$\frac{1}{n} \sum_{i=1}^n p(k_{t-1}^i, y) \rightarrow \mathbb{E} p(k_{t-1}^i, y) = \int p(x, y) \psi_{t-1}(x) dx = \psi_t(y)$$

with probability one as $n \rightarrow \infty$

Here the first equality is by the definition of ψ_{t-1} , and the second is by (3.9)

We have just shown that our estimator $\psi_t^n(y)$ in (3.13) converges almost surely to $\psi_t(y)$, which is just what we want to compute

In fact much stronger convergence results are true (see, for example, this paper)

Implementation A type called LAE for estimating densities by this technique can be found in QuantEcon

We repeat it here for convenience

```
#=
Computes a sequence of marginal densities for a continuous state space
Markov chain :math:`X_t` where the transition probabilities can be represented
as densities. The estimate of the marginal density of X_t is

1/n sum_{i=0}^n p(X_{t-1})^i, y)

This is a density in y.

@author : Spencer Lyon <spencer.lyon@nyu.edu>

@date: 2014-08-01

References
-----
http://quant-econ.net/jl/stationary_densities.html
=#

"""
A look ahead estimator associated with a given stochastic kernel p and a vector
```

```

of observations X.

##### Fields

- `p::Function`: The stochastic kernel. Signature is `p(x, y)` and it should be
vectorized in both inputs
- `X::Matrix`: A vector containing observations. Note that this can be passed as
any kind of `AbstractArray` and will be coerced into an `n x 1` vector.

"""

type LAE
    p::Function
    X::Matrix

    function LAE(p::Function, X::AbstractArray)
        n = length(X)
        new(p, reshape(X, n, 1))
    end
end

"""

A vectorized function that returns the value of the look ahead estimate at the
values in the array y.

##### Arguments

- `l::LAE`: Instance of `LAE` type
- `y::Array`: Array that becomes the `y` in `l.p(l.x, y)`

##### Returns

- `psi_vals::Vector`: Density at `(x, y)`

"""

function lae_est{T}(l::LAE, y::AbstractArray{T})
    k = length(y)
    v = l.p(l.X, reshape(y, 1, k))
    psi_vals = mean(v, 1)
    return squeeze(psi_vals, 1)
end

```

This function returns the right-hand side of (3.13) using

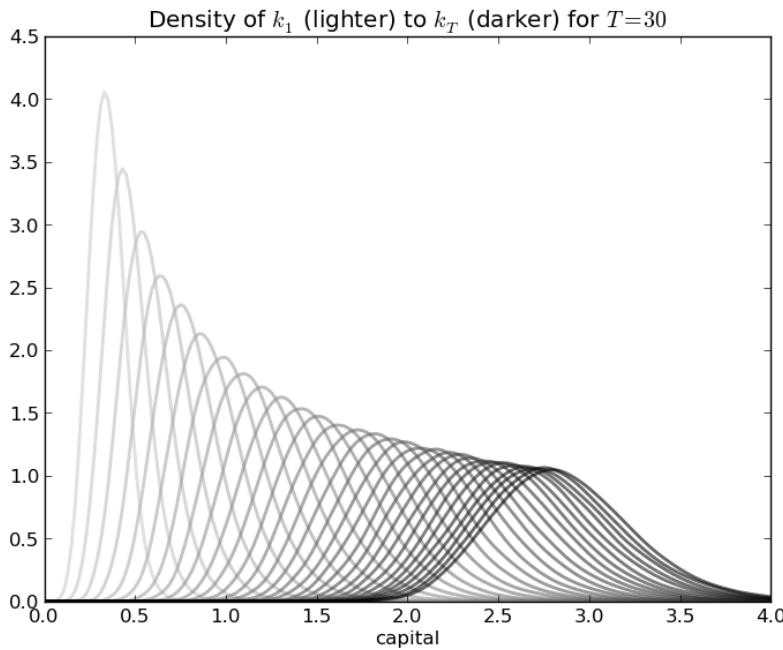
- an object of type LAE that stores the stochastic kernel and the observations
- the value y as its second argument

The function is vectorized, in the sense that if ψ is such an instance and y is an array, then the call $\psi(y)$ acts elementwise

(This is the reason that we reshaped X and y inside the type — to make vectorization work)

Example An example of usage for the stochastic growth model *described above* can be found in `stationary_densities/stochasticgrowth.py`

When run, the code produces a figure like this



The figure shows part of the density sequence $\{\psi_t\}$, with each density computed via the look ahead estimator

Notice that the sequence of densities shown in the figure seems to be converging — more on this in just a moment

Another quick comment is that each of these distributions could be interpreted as a cross sectional distribution (recall *this discussion*)

Beyond Densities

Up until now, we have focused exclusively on continuous state Markov chains where all conditional distributions $p(x, \cdot)$ are densities

As discussed above, not all distributions can be represented as densities

If the conditional distribution of X_{t+1} given $X_t = x$ **cannot** be represented as a density for some $x \in S$, then we need a slightly different theory

The ultimate option is to switch from densities to **probability measures**, but not all readers will be familiar with measure theory

We can, however, construct a fairly general theory using distribution functions

Example and Definitions To illustrate the issues, recall that Hopenhayn and Rogerson [HR93] study a model of firm dynamics where individual firm productivity follows the exogenous process

$$X_{t+1} = a + \rho X_t + \xi_{t+1}, \quad \text{where } \{\xi_t\} \stackrel{\text{IID}}{\sim} N(0, \sigma^2)$$

As is, this fits into the density case we treated above

However, the authors wanted this process to take values in $[0, 1]$, so they added boundaries at the end points 0 and 1

One way to write this is

$$X_{t+1} = h(a + \rho X_t + \xi_{t+1}) \quad \text{where } h(x) := x \mathbf{1}\{0 \leq x \leq 1\} + (1-x) \mathbf{1}\{x > 1\}$$

If you think about it, you will see that for any given $x \in [0, 1]$, the conditional distribution of X_{t+1} given $X_t = x$ puts positive probability mass on 0 and 1

Hence it cannot be represented as a density

What we can do instead is use cumulative distribution functions (cdfs)

To this end, set

$$G(x, y) := \mathbb{P}\{h(a + \rho x + \xi_{t+1}) \leq y\} \quad (0 \leq x, y \leq 1)$$

This family of cdfs $G(x, \cdot)$ plays a role analogous to the stochastic kernel in the density case

The distribution dynamics in (3.9) are then replaced by

$$F_{t+1}(y) = \int G(x, y) F_t(dx) \tag{3.14}$$

Here F_t and F_{t+1} are cdfs representing the distribution of the current state and next period state

The intuition behind (3.14) is essentially the same as for (3.9)

Computation If you wish to compute these cdfs, you cannot use the look-ahead estimator as before

Indeed, you should not use any density estimator, since the objects you are estimating/computing are not densities

One good option is simulation as before, combined with the *empirical distribution function*

Stability

In our lecture on finite Markov chains we also studied stationarity, stability and ergodicity

Here we will cover the same topics for the continuous case

We will, however, treat only the density case (as in *this section*), where the stochastic kernel is a family of densities

The general case is relatively similar — references are given below

Theoretical Results Analogous to *the finite case*, given a stochastic kernel p and corresponding Markov operator as defined in (3.10), a density ψ^* on S is called *stationary* for P if it is a fixed point of the operator P

In other words,

$$\psi^*(y) = \int p(x, y)\psi^*(x) dx, \quad \forall y \in S \quad (3.15)$$

As with the finite case, if ψ^* is stationary for P , and the distribution of X_0 is ψ^* , then, in view of (3.11), X_t will have this same distribution for all t

Hence ψ^* is the stochastic equivalent of a steady state

In the finite case, we learned that at least one stationary distribution exists, although there may be many

When the state space is infinite, the situation is more complicated

Even existence can fail very easily

For example, the random walk model has no stationary density (see, e.g., **EDTC**, p. 210)

However, there are well-known conditions under which a stationary density ψ^* exists

With additional conditions, we can also get a unique stationary density ($\psi \in \mathcal{D}$ and $\psi = \psi P \implies \psi = \psi^*$), and also global convergence in the sense that

$$\forall \psi \in \mathcal{D}, \quad \psi P^t \rightarrow \psi^* \quad \text{as } t \rightarrow \infty \quad (3.16)$$

This combination of existence, uniqueness and global convergence in the sense of (3.16) is often referred to as *global stability*

Under very similar conditions, we get *ergodicity*, which means that

$$\frac{1}{n} \sum_{t=1}^n h(X_t) \rightarrow \int h(x)\psi^*(x)dx \quad \text{as } n \rightarrow \infty \quad (3.17)$$

for any (measurable) function $h: S \rightarrow \mathbb{R}$ such that the right-hand side is finite

Note that the convergence in (3.17) does not depend on the distribution (or value) of X_0

This is actually very important for simulation — it means we can learn about ψ^* (i.e., approximate the right hand side of (3.17) via the left hand side) without requiring any special knowledge about what to do with X_0

So what are these conditions we require to get global stability and ergodicity?

In essence, it must be the case that

1. Probability mass does not drift off to the “edges” of the state space
2. Sufficient “mixing” obtains

For one such set of conditions see theorem 8.2.14 of **EDTC**

In addition

- [SLP89] contains a classic (but slightly outdated) treatment of these topics

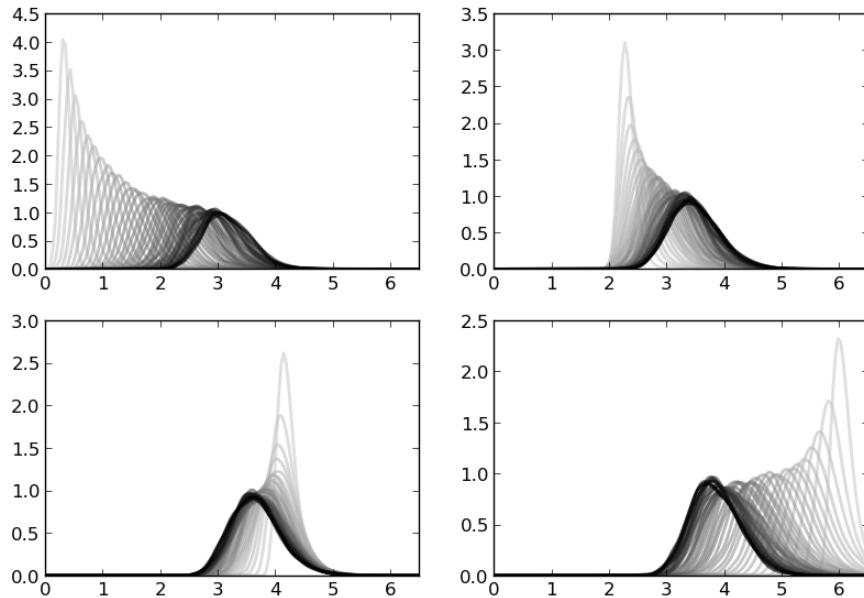
- From the mathematical literature, [LM94] and [MT09] give outstanding in depth treatments
- Section 8.1.2 of [EDTC](#) provides detailed intuition, and section 8.3 gives additional references
- [EDTC](#), section 11.3.4 provides a specific treatment for the growth model we considered in this lecture

An Example of Stability As stated above, the *growth model treated here* is stable under mild conditions on the primitives

- See [EDTC](#), section 11.3.4 for more details

We can see this stability in action — in particular, the convergence in (3.16) — by simulating the path of densities from various initial conditions

Here is such a figure



All sequences are converging towards the same limit, regardless of their initial condition

The details regarding initial conditions and so on are given in *this exercise*, where you are asked to replicate the figure

Computing Stationary Densities In the preceding figure, each sequence of densities is converging towards the unique stationary density ψ^*

Even from this figure we can get a fair idea what ψ^* looks like, and where its mass is located

However, there is a much more direct way to estimate the stationary density, and it involves only a slight modification of the look ahead estimator

Let's say that we have a model of the form (3.3) that is stable and ergodic

Let p be the corresponding stochastic kernel, as given in (3.7)

To approximate the stationary density ψ^* , we can simply generate a long time series X_0, X_1, \dots, X_n and estimate ψ^* via

$$\psi_n^*(y) = \frac{1}{n} \sum_{t=1}^n p(X_t, y) \quad (3.18)$$

This is essentially the same as the look ahead estimator (3.13), except that now the observations we generate are a single time series, rather than a cross section

The justification for (3.18) is that, with probability one as $n \rightarrow \infty$,

$$\frac{1}{n} \sum_{t=1}^n p(X_t, y) \rightarrow \int p(x, y) \psi^*(x) dx = \psi^*(y)$$

where the convergence is by (3.17) and the equality on the right is by (3.15)

The right hand side is exactly what we want to compute

On top of this asymptotic result, it turns out that the rate of convergence for the look ahead estimator is very good

The first exercise helps illustrate this point

Exercises

Exercise 1 Consider the simple threshold autoregressive model

$$X_{t+1} = \theta |X_t| + (1 - \theta^2)^{1/2} \xi_{t+1} \quad \text{where } \{\xi_t\} \stackrel{\text{IID}}{\sim} N(0, 1) \quad (3.19)$$

This is one of those rare nonlinear stochastic models where an analytical expression for the stationary density is available

In particular, provided that $|\theta| < 1$, there is a unique stationary density ψ^* given by

$$\psi^*(y) = 2\phi(y) \Phi \left[\frac{\theta y}{(1 - \theta^2)^{1/2}} \right] \quad (3.20)$$

Here ϕ is the standard normal density and Φ is the standard normal cdf

As an exercise, compute the look ahead estimate of ψ^* , as defined in (3.18), and compare it with ψ^* in (3.20) to see whether they are indeed close for large n

In doing so, set $\theta = 0.8$ and $n = 500$

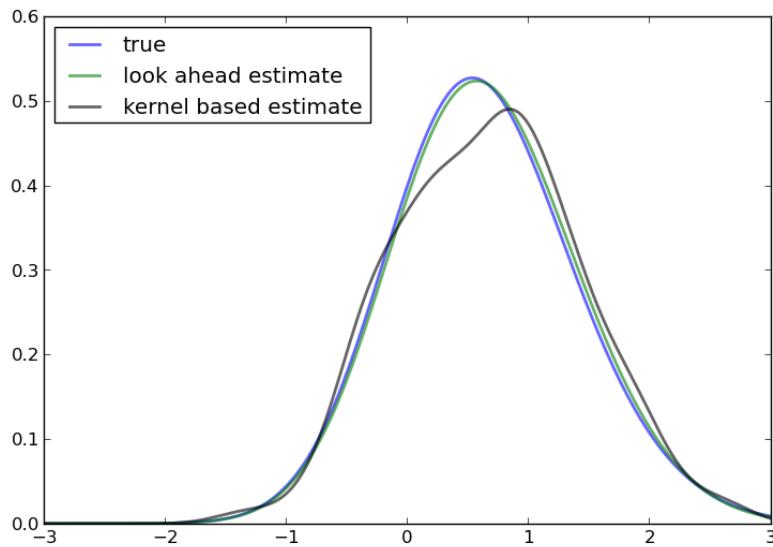
The next figure shows the result of such a computation

The additional density (black line) is a [nonparametric kernel density estimate](#), added to the solution for illustration

(You can try to replicate it before looking at the solution if you want to)

As you can see, the look ahead estimator is a much tighter fit than the kernel density estimator

If you repeat the simulation you will see that this is consistently the case



Exercise 2 Replicate the figure on global convergence shown above

The densities come from the stochastic growth model treated *at the start of the lecture*

Begin with the code found in `stationary_densities/stochasticgrowth.py`

Use the same parameters

For the four initial distributions, use the beta distribution and shift the random draws as shown below

```
psi_0 = Beta(5.0, 5.0) # Initial distribution
n = 1000
# .... more setup

for i=1:4
    # .... some code
    rand_draws = (rand(psi_0, n) .+ 2.5i) ./ 2
```

Exercise 3 A common way to compare distributions visually is with `boxplots`

To illustrate, let's generate three artificial data sets and compare them with a boxplot

```
using Plots
pyplot()
using LaTeXStrings
using StatPlots      # needed for box plot support

n = 500
x = randn(n) # N(0, 1)
x = exp(x) # Map x to lognormal
y = randn(n) + 2.0 # N(2, 1)
z = randn(n) + 4.0 # N(4, 1)
```

```

data = vcat(x, y, z)
l = [LaTeXString("\$X\$") LaTeXString("\$Y\$") LaTeXString("\$Z\$")]
xlabels = reshape(repmat(l, n), n*3, 1)

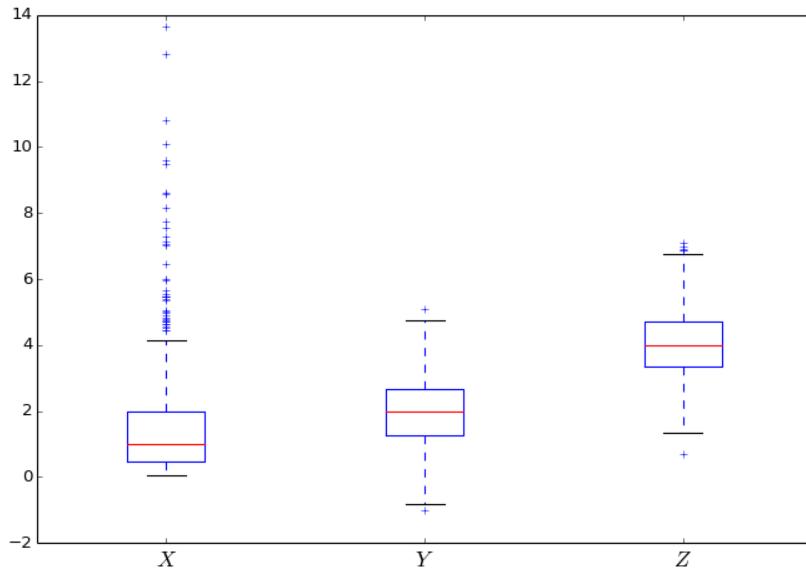
boxplot(xlabels, data, label="", ylims=(-2, 14))

```

The three data sets are

$$\{X_1, \dots, X_n\} \sim LN(0, 1), \quad \{Y_1, \dots, Y_n\} \sim N(2, 1), \quad \text{and} \quad \{Z_1, \dots, Z_n\} \sim N(4, 1),$$

The figure looks as follows



Each data set is represented by a box, where the top and bottom of the box are the third and first quartiles of the data, and the red line in the center is the median

The boxes give some indication as to

- the location of probability mass for each sample
- whether the distribution is right-skewed (as is the lognormal distribution), etc

Now let's put these ideas to use in a simulation

Consider the threshold autoregressive model in (3.19)

We know that the distribution of X_t will converge to (3.20) whenever $|\theta| < 1$

Let's observe this convergence from different initial conditions using boxplots

In particular, the exercise is to generate J boxplot figures, one for each initial condition X_0 in

```
initial_conditions = linspace(8, 0, J)
```

For each X_0 in this set,

1. Generate k time series of length n , each starting at X_0 and obeying (3.19)
2. Create a boxplot representing n distributions, where the t -th distribution shows the k observations of X_t

Use $\theta = 0.9, n = 20, k = 5000, J = 8$

Solutions

[Solution notebook](#)

Appendix

Here's the proof of (3.6)

Let F_U and F_V be the cumulative distributions of U and V respectively

By the definition of V , we have $F_V(v) = \mathbb{P}\{a + bU \leq v\} = \mathbb{P}\{U \leq (v - a)/b\}$

In other words, $F_V(v) = F_U((v - a)/b)$

Differentiating with respect to v yields (3.6)

The Lucas Asset Pricing Model

Contents

- *The Lucas Asset Pricing Model*
 - *Overview*
 - *The Lucas Model*
 - *Exercises*
 - *Solutions*

Overview

As stated in an [earlier lecture](#), an asset is a claim on a stream of prospective payments

What is the correct price to pay for such a claim?

The elegant asset pricing model of Lucas [[Luc78](#)] attempts to answer this question in an equilibrium setting with risk averse agents

While we mentioned some consequences of Lucas' model *earlier*, it is now time to work through the model more carefully, and try to understand where the fundamental asset pricing equation comes from

A side benefit of studying Lucas' model is that it provides a beautiful illustration of model building in general and equilibrium pricing in competitive models in particular

The Lucas Model

Lucas studied a pure exchange economy with a representative consumer (or household), where

- *Pure exchange* means that all endowments are exogenous
- *Representative consumer* means that either
 - there is a single consumer (sometimes also referred to as a household), or
 - all consumers have identical endowments and preferences

Either way, the assumption of a representative agent means that prices adjust to eradicate desires to trade

This makes it very easy to compute competitive equilibrium prices

Basic Setup

Let's review the set up

Assets There is a single “productive unit” that costlessly generates a sequence of consumption goods $\{y_t\}_{t=0}^{\infty}$

Another way to view $\{y_t\}_{t=0}^{\infty}$ is as a *consumption endowment* for this economy

We will assume that this endowment is Markovian, following the exogenous process

$$y_{t+1} = G(y_t, \xi_{t+1})$$

Here $\{\xi_t\}$ is an iid shock sequence with known distribution ϕ and $y_t \geq 0$

An asset is a claim on all or part of this endowment stream

The consumption goods $\{y_t\}_{t=0}^{\infty}$ are nonstorable, so holding assets is the only way to transfer wealth into the future

For the purposes of intuition, it's common to think of the productive unit as a “tree” that produces fruit

Based on this idea, a “Lucas tree” is a claim on the consumption endowment

Consumers A representative consumer ranks consumption streams $\{c_t\}$ according to the time separable utility functional

$$\mathbb{E} \sum_{t=0}^{\infty} \beta^t u(c_t) \tag{3.21}$$

Here

- $\beta \in (0, 1)$ is a fixed discount factor
- u is a strictly increasing, strictly concave, continuously differentiable period utility function
- \mathbb{E} is a mathematical expectation

Pricing a Lucas Tree What is an appropriate price for a claim on the consumption endowment?

We'll price an *ex dividend* claim, meaning that

- the seller retains this period's dividend
- the buyer pays p_t today to purchase a claim on
 - y_{t+1} and
 - the right to sell the claim tomorrow at price p_{t+1}

Since this is a competitive model, the first step is to pin down consumer behavior, taking prices as given

Next we'll impose equilibrium constraints and try to back out prices

In the consumer problem, the consumer's control variable is the share π_t of the claim held in each period

Thus, the consumer problem is to maximize (3.21) subject to

$$c_t + \pi_{t+1}p_t \leq \pi_t y_t + \pi_t p_t$$

along with $c_t \geq 0$ and $0 \leq \pi_t \leq 1$ at each t

The decision to hold share π_t is actually made at time $t - 1$

But this value is inherited as a state variable at time t , which explains the choice of subscript

The dynamic program We can write the consumer problem as a dynamic programming problem

Our first observation is that prices depend on current information, and current information is really just the endowment process up until the current period

In fact the endowment process is Markovian, so that the only relevant information is the current state $y \in \mathbb{R}_+$ (dropping the time subscript)

This leads us to guess an equilibrium where price is a function p of y

Remarks on the solution method

- Since this is a competitive (read: price taking) model, the consumer will take this function p as given
- In this way we determine consumer behavior given p and then use equilibrium conditions to recover p
- This is the standard way to solve competitive equilibrium models

Using the assumption that price is a given function p of y , we write the value function and constraint as

$$v(\pi, y) = \max_{c, \pi'} \left\{ u(c) + \beta \int v(\pi', G(y, z)) \phi(dz) \right\}$$

subject to

$$c + \pi' p(y) \leq \pi y + \pi p(y) \tag{3.22}$$

We can invoke the fact that utility is increasing to claim equality in (3.22) and hence eliminate the constraint, obtaining

$$v(\pi, y) = \max_{\pi'} \left\{ u[\pi(y + p(y)) - \pi' p(y)] + \beta \int v(\pi', G(y, z)) \phi(dz) \right\} \quad (3.23)$$

The solution to this dynamic programming problem is an optimal policy expressing either π' or c as a function of the state (π, y)

- Each one determines the other, since $c(\pi, y) = \pi(y + p(y)) - \pi'(\pi, y)p(y)$

Next steps What we need to do now is determine equilibrium prices

It seems that to obtain these, we will have to

1. Solve this two dimensional dynamic programming problem for the optimal policy
2. Impose equilibrium constraints
3. Solve out for the price function $p(y)$ directly

However, as Lucas showed, there is a related but more straightforward way to do this

Equilibrium constraints Since the consumption good is not storable, in equilibrium we must have $c_t = y_t$ for all t

In addition, since there is one representative consumer (alternatively, since all consumers are identical), there should be no trade in equilibrium

In particular, the representative consumer owns the whole tree in every period, so $\pi_t = 1$ for all t

Prices must adjust to satisfy these two constraints

The equilibrium price function Now observe that the first order condition for (3.23) can be written as

$$u'(c)p(y) = \beta \int v'_1(\pi', G(y, z)) \phi(dz)$$

where v'_1 is the derivative of v with respect to its first argument

To obtain v'_1 we can simply differentiate the right hand side of (3.23) with respect to π , yielding

$$v'_1(\pi, y) = u'(c)(y + p(y))$$

Next we impose the equilibrium constraints while combining the last two equations to get

$$p(y) = \beta \int \frac{u'[G(y, z)]}{u'(y)} [G(y, z) + p(G(y, z))] \phi(dz) \quad (3.24)$$

In sequential rather than functional notation, we can also write this as

$$p_t = \mathbb{E}_t \left[\beta \frac{u'(c_{t+1})}{u'(c_t)} (c_{t+1} + p_{t+1}) \right] \quad (3.25)$$

This is the famous consumption-based asset pricing equation

Before discussing it further we want to solve out for prices

Solving the Model Equation (3.24) is a *functional equation* in the unknown function p

The solution is an equilibrium price function p^*

Let's look at how to obtain it

Setting up the problem Instead of solving for it directly we'll follow Lucas' indirect approach, first setting

$$f(y) := u'(y)p(y) \quad (3.26)$$

so that (3.24) becomes

$$f(y) = h(y) + \beta \int f[G(y, z)]\phi(dz) \quad (3.27)$$

Here $h(y) := \beta \int u'[G(y, z)]G(y, z)\phi(dz)$ is a function that depends only on the primitives

Equation (3.27) is a functional equation in f

The plan is to solve out for f and convert back to p via (3.26)

To solve (3.27) we'll use a standard method: convert it to a fixed point problem

First we introduce the operator T mapping f into Tf as defined by

$$(Tf)(y) = h(y) + \beta \int f[G(y, z)]\phi(dz) \quad (3.28)$$

The reason we do this is that a solution to (3.27) now corresponds to a function f^* satisfying $(Tf^*)(y) = f^*(y)$ for all y

In other words, a solution is a *fixed point* of T

This means that we can use fixed point theory to obtain and compute the solution

A little fixed point theory Let $cb\mathbb{R}_+$ be the set of continuous bounded functions $f: \mathbb{R}_+ \rightarrow \mathbb{R}_+$

We now show that

1. T has exactly one fixed point f^* in $cb\mathbb{R}_+$
2. For any $f \in cb\mathbb{R}_+$, the sequence $T^k f$ converges uniformly to f^*

(Note: If you find the mathematics heavy going you can take 1–2 as given and skip to the *next section*)

Recall the [Banach contraction mapping theorem](#)

It tells us that the previous statements will be true if we can find an $\alpha < 1$ such that

$$\|Tf - Tg\| \leq \alpha \|f - g\|, \quad \forall f, g \in cb\mathbb{R}_+ \quad (3.29)$$

Here $\|h\| := \sup_{x \in \mathbb{R}_+} |h(x)|$

To see that (3.29) is valid, pick any $f, g \in cb\mathbb{R}_+$ and any $y \in \mathbb{R}_+$

Observe that, since integrals get larger when absolute values are moved to the inside,

$$\begin{aligned}|Tf(y) - Tg(y)| &= \left| \beta \int f[G(y, z)]\phi(dz) - \beta \int g[G(y, z)]\phi(dz) \right| \\&\leq \beta \int |f[G(y, z)] - g[G(y, z)]| \phi(dz) \\&\leq \beta \int \|f - g\| \phi(dz) \\&= \beta \|f - g\|\end{aligned}$$

Since the right hand side is an upper bound, taking the sup over all y on the left hand side gives (3.29) with $\alpha := \beta$

Computation – An Example The preceding discussion tells that we can compute f^* by picking any arbitrary $f \in cb\mathbb{R}_+$ and then iterating with T

The equilibrium price function p^* can then be recovered by $p^*(y) = f^*(y)/u'(y)$

Let's try this when $\ln y_{t+1} = \alpha \ln y_t + \sigma \epsilon_{t+1}$ where $\{\epsilon_t\}$ is iid and standard normal

Utility will take the isoelastic form $u(c) = c^{1-\gamma}/(1-\gamma)$, where $\gamma > 0$ is the coefficient of relative risk aversion

Some code to implement the iterative computational procedure can be found in `lucastree.jl` from the `QuantEcon.applications` repo

We repeat it here for convenience

```
#=
Solves the price function for the Lucas tree in a continuous state
setting, using piecewise linear approximation for the sequence of
candidate price functions. The consumption endowment follows the
log linear AR(1) process

log y' = alpha log y + sigma epsilon

where y' is a next period y and epsilon is an iid standard normal
shock. Hence

y' = y^alpha * xi    where xi = e^(sigma * epsilon)

The distribution phi of xi is

phi = LN(0, sigma^2) where LN means lognormal

@authors : Spencer Lyon <spencer.lyon@nyu.edu>, John Stachurski

References
-----
http://quant-econ.net/jl/markov_asset.html
```

```

#=

using QuantEcon
using Distributions
using Interpolations


"""
A function that takes two arrays and returns a function that approximates the
data using continuous piecewise linear interpolation.

"""

function lin_interp(x_vals::Vector{Float64}, y_vals::Vector{Float64})
    # == linear interpolation inside grid ==
    w = interpolate((x_vals,), y_vals, Gridded(Linear()))
    # == constant values outside grid ==
    w = extrapolate(w, Interpolations.Flat())
    return w
end


"""

The Lucas asset pricing model --- parameters and grid data
"""

type LucasTree
    gamma::Real          # coefficient of risk aversion
    beta::Real            # Discount factor in (0, 1)
    alpha::Real           # Correlation coefficient in the shock process
    sigma::Real           # Volatility of shock process
    phi::Distribution     # Distribution for shock process
    grid::Vector          # Grid of points on which to evaluate prices.
    shocks::Vector        # Draws of the shock
    h::Vector              # The h function represented as a vector
end


"""

Constructor for the Lucas asset pricing model
"""

function LucasTree(;gamma=2.0,
                    beta=0.95,
                    alpha=0.9,
                    sigma=0.1,
                    grid_size=100)

    phi = LogNormal(0.0, sigma)
    shocks = rand(phi, 500)

    # == build a grid with mass around stationary distribution ==
    ssd = sigma / sqrt(1 - alpha^2)
    grid_min, grid_max = exp(-4 * ssd), exp(4 * ssd)

```

```

grid = collect(linspace(grid_min, grid_max, grid_size))

# == set h(y) = beta * int u'(G(y,z)) G(y,z) phi(dz) == #
h = similar(grid)
for (i, y) in enumerate(grid)
    h[i] = beta * mean((y^alpha .* shocks).^(1 - gamma))
end

return LucasTree(gamma,
                 beta,
                 alpha,
                 sigma,
                 phi,
                 grid,
                 shocks,
                 h)
end

"""
The approximate Lucas operator, which computes and returns updated function
Tf on the grid points.
"""
function lucas_operator(lt::LucasTree, f::Vector{Float64})

    # == unpack names == #
    grid, alpha, beta, h = lt.grid, lt.alpha, lt.beta, lt.h
    z = lt.shocks

    Tf = similar(f)
    Af = lin_interp(grid, f)

    for (i, y) in enumerate(grid)
        Tf[i] = h[i] + beta * mean(Af[y^alpha .* z])
    end
    return Tf
end

"""
Compute the equilibrium price function associated with Lucas tree `lt`
"""
function compute_lt_price(lt::LucasTree, max_iter=500)

    # == Simplify names == #
    grid = lt.grid
    alpha, beta, gamma = lt.alpha, lt.beta, lt.gamma

    # == Create suitable initial vector to iterate from == #
    f_init = zeros(grid)

    func(f_vec) = lucas_operator(lt, f_vec)
    f = compute_fixed_point(func, f_init);

```

```

    max_iter=max_iter,
    err_tol=1e-4,
    verbose=false)

#  $p(y) = f(y) * y^\gamma$ 
price = f .* grid.^gamma)

return price
end

```

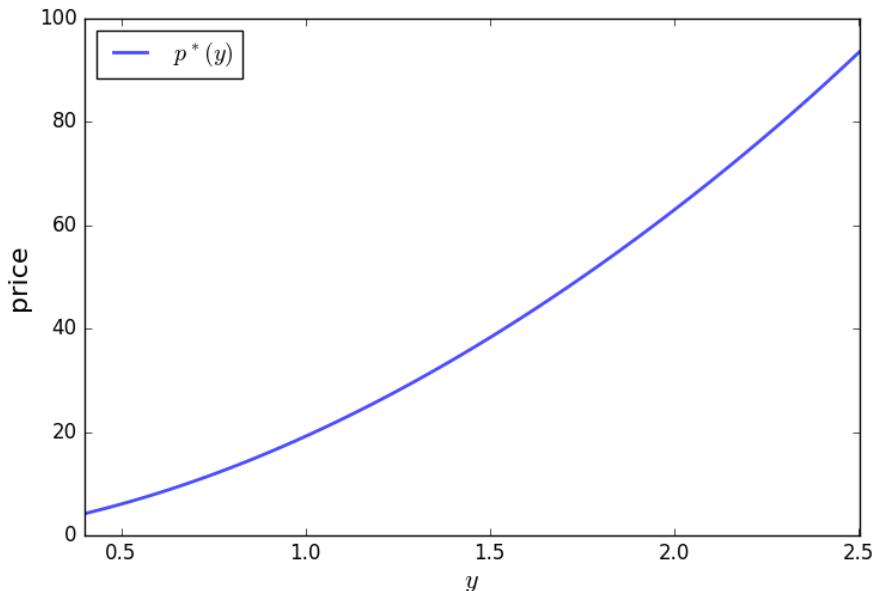
An example of usage is given in the docstring and repeated here

```

tree = LucasTree(2, 0.95, 0.90, 0.1)
grid, price_vals = compute_lt_price(tree)

```

Here's the resulting price function



The price is increasing, even if we remove all serial correlation from the endowment process

The reason is that a larger current endowment reduces current marginal utility

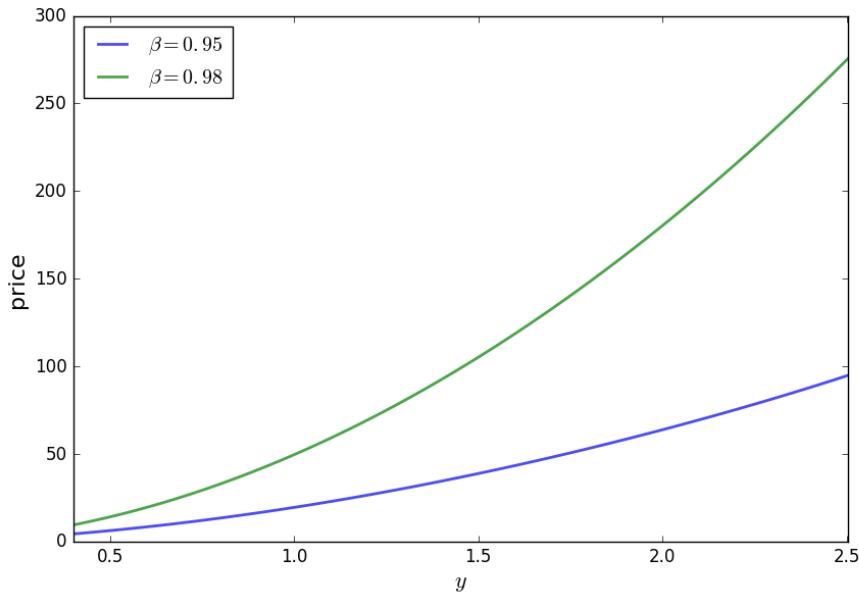
The price must therefore rise to induce the household to consume the entire endowment (and hence satisfy the resource constraint)

What happens with a more patient consumer?

Here the blue line corresponds to the previous parameters and the green line is price when $\beta = 0.98$

We see that when consumers are more patient the asset becomes more valuable, and the price of the Lucas tree shifts up

Exercise 1 asks you to replicate this figure



Exercises

Exercise 1 Replicate *the figure* to show how discount rates affect prices

Solutions

[Solution notebook](#)

The Aiyagari Model

Overview

In this lecture we describe the structure of a class of models that build on work by Truman Bewley [[Bew77](#)]

We begin by discussing an example of a Bewley model due to Rao Aiyagari

The model features

- Heterogeneous agents
- A single exogenous vehicle for borrowing and lending
- Limits on amounts individual agents may borrow

The Aiyagari model has been used to investigate many topics, including

- precautionary savings and the effect of liquidity constraints [[Aiy94](#)]

- risk sharing and asset pricing [HL96]
- the shape of the wealth distribution [BBZ15]
- etc., etc., etc.

References The primary reference for this lecture is [Aiy94]

A textbook treatment is available in chapter 18 of [LS12]

A continuous time version of the model by SeHyoun Ahn and Benjamin Moll can be found [here](#)

The Economy

Households Infinitely lived households / consumers face idiosyncratic income shocks

A unit interval of *ex ante* identical households face a common borrowing constraint

The savings problem faced by a typical household is

$$\max \mathbb{E} \sum_{t=0}^{\infty} \beta^t u(c_t)$$

subject to

$$a_{t+1} + c_t \leq wz_t + (1+r)a_t \quad c_t \geq 0, \quad \text{and} \quad a_t \geq -B$$

where

- c_t is current consumption
- a_t is assets
- z_t is an exogenous component of labor income capturing stochastic unemployment risk, etc.
- w is a wage rate
- r is a net interest rate
- B is the borrowing constraint

The exogenous process $\{z_t\}$ follows a finite state Markov chain with given stochastic matrix P

The wage and interest rate are fixed over time

In this simple version of the model, households supply labor inelastically because they do not value leisure

Firms

Firms produce output by hiring capital and labor

Firms act competitively and face constant returns to scale

Since returns to scale are constant the number of firms does not matter

Hence we can consider a single (but nonetheless competitive) representative firm

The firm's output is

$$Y_t = AK_t^\alpha N^{1-\alpha}$$

where

- A and α are parameters with $A > 0$ and $\alpha \in (0, 1)$
- K_t is aggregate capital
- N is total labor supply (which is constant in this simple version of the model)

The firm's problem is

$$\max_{K,N} \left\{ AK_t^\alpha N^{1-\alpha} - (r + \delta)K - wN \right\}$$

The parameter δ is the depreciation rate

From the first-order condition with respect to capital, the firm's inverse demand for capital is

$$r = A\alpha \left(\frac{N}{K} \right)^{1-\alpha} - \delta \quad (3.30)$$

Using this expression and the firm's first-order condition for labor, we can pin down the equilibrium wage rate as a function of r as

$$w(r) = A(1 - \alpha)(A\alpha/(r + \delta))^{\alpha/(1-\alpha)} \quad (3.31)$$

Equilibrium We construct a *stationary rational expectations equilibrium* (SREE)

In such an equilibrium

- prices induce behavior that generates aggregate quantities consistent with the prices
- aggregate quantities and prices are constant over time

In more detail, an SREE lists a set of prices, savings and production policies such that

- households want to choose the specified savings policies taking the prices as given
- firms maximize profits taking the same prices as given
- the resulting aggregate quantities are consistent with the prices; in particular, the demand for capital equals the supply
- aggregate quantities (defined as cross-sectional averages) are constant

In practice, once parameter values are set, we can check for an SREE by the following steps

1. pick a proposed quantity K for aggregate capital
2. determine corresponding prices, with interest rate r determined by (3.30) and a wage rate $w(r)$ as given in (3.31)
3. determine the common optimal savings policy of the households given these prices
4. compute aggregate capital as the mean of steady state capital given this savings policy

If this final quantity agrees with K then we have a SREE

Code

Let's look at how we might compute such an equilibrium in practice

To solve the household's dynamic programming problem we'll use the `DiscreteDP` type from `QuantEcon.jl`

Our first task is the least exciting one: write code that maps parameters for a household problem into the `R` and `Q` matrices needed to generate an instance of `DiscreteDP`

Below is a piece of boilerplate code that does just this

It comes from the file `aiyagari_household.jl` from the `QuantEcon.applications` repository

In reading the code, the following information will be helpful

- `R` needs to be a matrix where `R[s, a]` is the reward at state `s` under action `a`
- `Q` needs to be a three dimensional array where `Q[s, a, s']` is the probability of transitioning to state `s'` when the current state is `s` and the current action is `a`

(For a detailed discussion of `DiscreteDP` see [this lecture](#))

Here we take the state to be $s_t := (a_t, z_t)$, where a_t is assets and z_t is the shock

The action is the choice of next period asset level a_{t+1}

The type also includes a default set of parameters that we'll adopt unless otherwise specified

```
#=
Filename: aiyagari_household.jl
Author: Victoria Gregory
Date: 8/29/2016

This file defines the Household type (and its constructor)
for setting up an Aiyagari household problem.
=#

using QuantEcon

"""
Stores all the parameters that define the household's
problem.

##### Fields

- `r::Float64` : interest rate
- `w::Float64` : wage
- `beta::Float64` : discount factor
- `z_chain::MarkovChain` : MarkovChain for income
- `a_min::Float64` : minimum on asset grid
- `a_max::Float64` : maximum on asset grid
- `a_size::Int64` : number of points on asset grid
- `z_size::Int64` : number of points on income grid
- `n::Int64` : number of points in state space: (a, z)
- `s_vals::Array{Float64}` : stores all the possible (a, z) combinations
```

```

- `s_i_vals::Array{Int64}` : stores indices of all the possible (a, z) combinations
- `R::Array{Float64}` : reward array
- `Q::Array{Float64}` : transition probability array
"""

type Household
    r::Float64
    w::Float64
    beta::Float64
    z_chain::MarkovChain{Float64,Array{Float64,2},Array{Float64,1}}
    a_min::Float64
    a_max::Float64
    a_size::Int64
    a_vals::Vector{Float64}
    z_size::Int64
    n::Int64
    s_vals::Array{Float64}
    s_i_vals::Array{Int64}
    R::Array{Float64}
    Q::Array{Float64}
end

"""

Constructor for `Household`


##### Arguments
- `r::Float64(0.01)` : interest rate
- `w::Float64(1.0)` : wage
- `beta::Float64(0.96)` : discount factor
- `z_chain::MarkovChain` : MarkovChain for income
- `a_min::Float64(1e-10)` : minimum on asset grid
- `a_max::Float64(18.0)` : maximum on asset grid
- `a_size::Int64(200)` : number of points on asset grid

"""

function Household(;r::Float64=0.01, w::Float64=1.0, beta::Float64=0.96,
                  z_chain::MarkovChain{Float64,Array{Float64,2},Array{Float64,1}}=
                  MarkovChain([0.9 0.1; 0.1 0.9], [0.1; 1.0]), a_min::Float64=1e-10,
                  a_max::Float64=18.0, a_size::Int64=200)

    # set up grids
    a_vals = linspace(a_min, a_max, a_size)
    z_size = length(z_chain.state_values)
    n = a_size*z_size
    s_vals = gridmake(a_vals, z_chain.state_values)
    s_i_vals = gridmake(1:a_size, 1:z_size)

    # set up Q
    Q = zeros(Float64, n, a_size, n)
    for next_s_i in 1:n
        for a_i in 1:a_size
            for s_i in 1:n
                z_i = s_i_vals[s_i, 2]
                next_z_i = s_i_vals[next_s_i, 2]

```

```

        next_a_i = s_i_vals[next_s_i, 1]
        if next_a_i == a_i
            Q[s_i, a_i, next_s_i] = z_chain.p[z_i, next_z_i]
        end
    end
end

# placeholder for R
R = fill(-Inf, n, a_size)
h = Household(r, w, beta, z_chain, a_min, a_max, a_size,
              a_vals, z_size, n, s_vals, s_i_vals, R, Q)

setup_R!(h, r, w)

return h

end

"""
Update the reward array of a Household object, given
a new interest rate and wage.

##### Arguments
- `h::Household` : instance of Household type
- `r::Float64(0.01)` : interest rate
- `w::Float64(1.0)` : wage
"""
function setup_R!(h::Household, r::Float64, w::Float64)

    # set up R
    R = h.R
    for new_a_i in 1:h.a_size
        a_new = h.a_vals[new_a_i]
        for s_i in 1:h.n
            a = h.s_vals[s_i, 1]
            z = h.s_vals[s_i, 2]
            c = w * z + (1 + r) * a - a_new
            if c > 0
                R[s_i, new_a_i] = log(c)
            end
        end
    end

    h.r = r
    h.w = w
    h.R = R
    h
end

```

In the following examples our import statements assume that this code is stored as `aiyagari_household.jl` in the present working directory

(Or you can copy it into a Jupyter notebook cell and delete the corresponding import statement)

As a first example of what we can do, let's compute and plot an optimal accumulation policy at fixed prices

```
#=
Filename: aiyagari_compute_policy.jl
Author: Victoria Gregory
Date: 8/29/2016

Computes and plots the optimal policy of a household
from the Aiyagari model, given prices.
=#

using QuantEcon
using LaTeXStrings
using Plots
pyplot()
include("aiyagari_household.jl")

# Example prices
r = 0.03
w = 0.956

# Create an instance of Household
am = Household(a_max=20.0, r=r, w=w)

# Use the instance to build a discrete dynamic program
am_ddp = DiscreteDP(am.R, am.Q, am.beta)

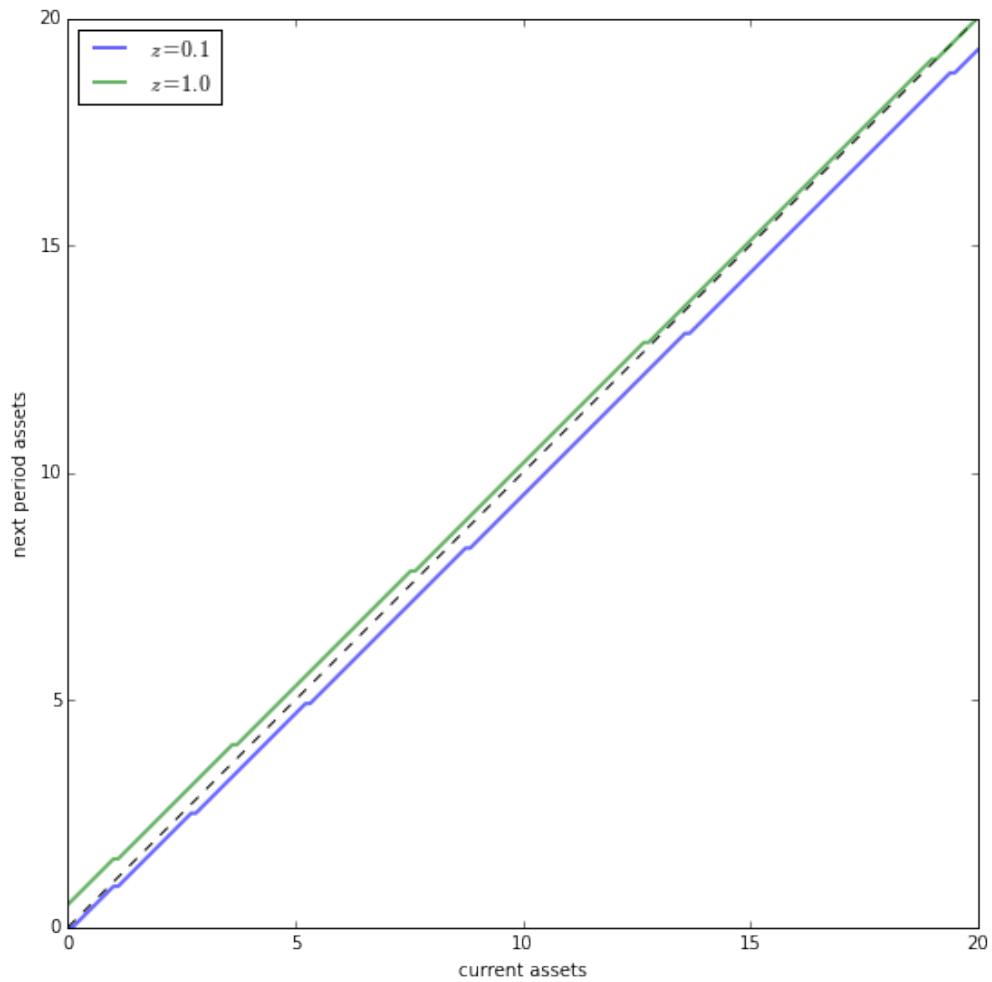
# Solve using policy function iteration
results = solve(am_ddp, PFI)

# Simplify names
z_size, a_size = am.z_size, am.a_size
z_vals, a_vals = am.z_chain.state_values, am.a_vals
n = am.n

# Get all optimal actions across the set of
# a indices with z fixed in each column
a_star = Array(Float64, a_size, z_size)
s_i_vals = gridmake(1:a_size, 1:z_size)
for s_i in 1:n
    a_i = s_i_vals[s_i, 1]
    z_i = s_i_vals[s_i, 2]
    a_star[a_i, z_i] = a_vals[results.sigma[s_i]]
end

labels = [string(L"$z = $", z_vals[1]); string(L"$z = $", z_vals[2])]
plot(a_vals, a_star, label=labels', lw=2, alpha=0.6)
plot!(a_vals, a_vals, label="", color=:black, linestyle=:dash)
plot!(xlabel="current assets", ylabel="next period assets", grid=false)
```

Here's the output



The plot shows asset accumulation policies at different values of the exogenous state

Now we want to calculate the equilibrium

Let's do this visually as a first pass

The following code draws aggregate supply and demand curves

The intersection gives equilibrium interest rates and capital

```
#=
Filename: aiyagari_compute_equilibrium.jl
Author: Victoria Gregory
Date: 8/30/2016

Draws the aggregate supply and demand curves for
the Aiyagari model.

#=

using QuantEcon
include("aiyagari_household.jl")
using Plots
pyplot()

# Firms' parameters
A = 1
N = 1
alpha = 0.33
beta = 0.96
delta = 0.05

"""
Compute wage rate given an interest rate, r
"""

function r_to_w(r::Float64)
    return A * (1 - alpha) * (A * alpha / (r + delta)) ^ (alpha / (1 - alpha))
end

"""
Inverse demand curve for capital. The interest rate
associated with a given demand for capital K.
"""

function rd(K::Float64)
    return A * alpha * (N / K) ^ (1 - alpha) - delta
end

"""
Map prices to the induced level of capital stock.

#### Arguments
- `am::Household` : Household instance for problem we want to solve
- `r::Float64` : interest rate

#### Returns
- The implied level of aggregate capital

```

```
"""
function prices_to_capital_stock(am::Household, r::Float64)

    # Set up problem
    w = r_to_w(r)
    setup_R!(am, r, w)
    aiyagari_ddp = DiscreteDP(am.R, am.Q, am.beta)

    # Compute the optimal policy
    results = solve(aiyagari_ddp, PFI)

    # Compute the stationary distribution
    stationary_probs = stationary_distributions(results.mc)[:, 1][1]

    # Return K
    return sum(am.s_vals[:, 1] .* stationary_probs)
end

# Create an instance of Household
z_chain = MarkovChain([0.67 0.33; 0.33 0.67], [0.5, 1.5])
am = Household(z_chain=z_chain, beta=beta, a_max=20.0)

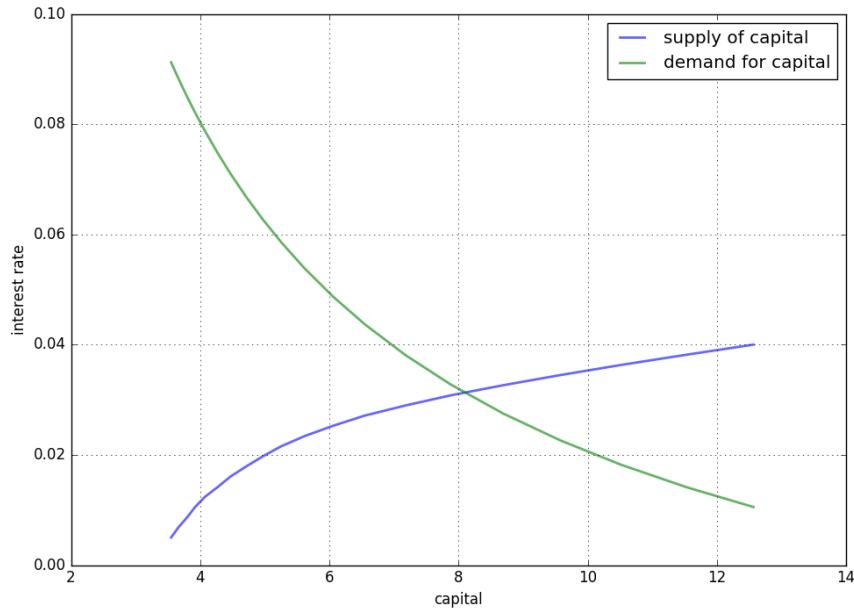
# Create a grid of r values at which to compute demand and supply of capital
num_points = 20
r_vals = linspace(0.02, 1/beta - 1, num_points)

# Compute supply of capital
k_vals = Array(Float64, num_points, 1)
for i in 1:num_points
    k_vals[i] = prices_to_capital_stock(am, r_vals[i])
end

# Plot against demand for capital by firms
demand = [rd(k) for k in k_vals]
labels = ["demand for capital"; "supply of capital"]
plot(k_vals, [demand r_vals], label=labels', lw=2, alpha=0.6)
plot!(xlabel="capital", ylabel="interest rate")
```

Here's the corresponding plot

Modeling Career Choice



Contents

- *Modeling Career Choice*
 - *Overview*
 - *Model*
 - *Implementation: career.jl*
 - *Exercises*
 - *Solutions*

Overview

Next we study a computational problem concerning career and job choices. The model is originally due to Derek Neal [Nea99] and this exposition draws on the presentation in [LS12], section 6.5.

Model features

- career and job within career both chosen to maximize expected discounted wage flow
- infinite horizon dynamic programming with two states variables

Model

In what follows we distinguish between a career and a job, where

- a *career* is understood to be a general field encompassing many possible jobs, and

- a *job* is understood to be a position with a particular firm

For workers, wages can be decomposed into the contribution of job and career

- $w_t = \theta_t + \epsilon_t$, where
 - θ_t is contribution of career at time t
 - ϵ_t is contribution of job at time t

At the start of time t , a worker has the following options

- retain a current (career, job) pair (θ_t, ϵ_t) — referred to hereafter as “stay put”
- retain a current career θ_t but redraw a job ϵ_t — referred to hereafter as “new job”
- redraw both a career θ_t and a job ϵ_t — referred to hereafter as “new life”

Draws of θ and ϵ are independent of each other and past values, with

- $\theta_t \sim F$
- $\epsilon_t \sim G$

Notice that the worker does not have the option to retain a job but redraw a career — starting a new career always requires starting a new job

A young worker aims to maximize the expected sum of discounted wages

$$\mathbb{E} \sum_{t=0}^{\infty} \beta^t w_t \quad (3.32)$$

subject to the choice restrictions specified above

Let $V(\theta, \epsilon)$ denote the value function, which is the maximum of (3.32) over all feasible (career, job) policies, given the initial state (θ, ϵ)

The value function obeys

$$V(\theta, \epsilon) = \max\{I, II, III\},$$

where

$$\begin{aligned} I &= \theta + \epsilon + \beta V(\theta, \epsilon) \\ II &= \theta + \int \epsilon' G(d\epsilon') + \beta \int V(\theta, \epsilon') G(d\epsilon') \\ III &= \int \theta' F(d\theta') + \int \epsilon' G(d\epsilon') + \beta \int \int V(\theta', \epsilon') G(d\epsilon') F(d\theta') \end{aligned} \quad (3.33)$$

Evidently I , II and III correspond to “stay put”, “new job” and “new life”, respectively

Parameterization As in [LS12], section 6.5, we will focus on a discrete version of the model, parameterized as follows:

- both θ and ϵ take values in the set `linspace(0, B, N)` — an even grid of N points between 0 and B inclusive
- $N = 50$

- $B = 5$
- $\beta = 0.95$

The distributions F and G are discrete distributions generating draws from the grid points `linspace(0, B, N)`

A very useful family of discrete distributions is the Beta-binomial family, with probability mass function

$$p(k | n, a, b) = \binom{n}{k} \frac{B(k + a, n - k + b)}{B(a, b)}, \quad k = 0, \dots, n$$

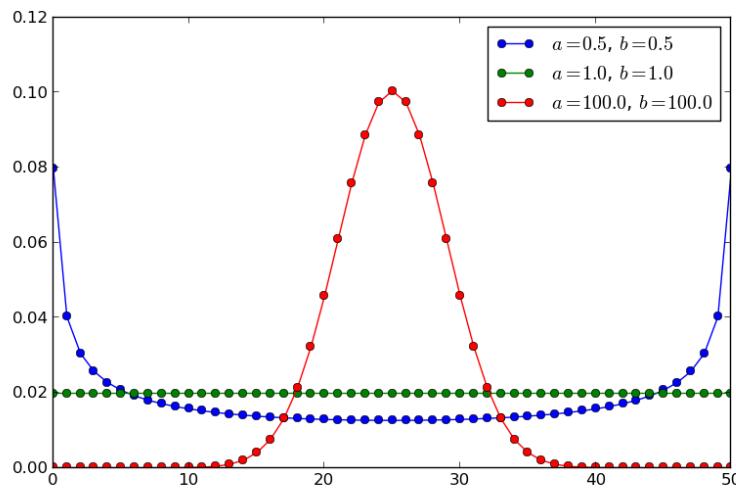
Interpretation:

- draw q from a Beta distribution with shape parameters (a, b)
- run n independent binary trials, each with success probability q
- $p(k | n, a, b)$ is the probability of k successes in these n trials

Nice properties:

- very flexible class of distributions, including uniform, symmetric unimodal, etc.
- only three parameters

Here's a figure showing the effect of different shape parameters when $n = 50$



The code that generated this figure can be found [here](#)

Implementation: `career.jl`

The `QuantEcon.applications` repo provides some code for solving the DP problem described above. See in particular [this file](#), which is repeated here for convenience

```

#=

A type to solve the career / job choice model due to Derek Neal.

@author : Spencer Lyon <spencer.lyon@nyu.edu>

@date: 2014-08-05

References
-----

http://quant-econ.net/jl/career.html

[Neal1999] Neal, D. (1999). The Complexity of Job Mobility among Young Men,
Journal of Labor Economics, 17(2), 237-261.

#


"""

Career/job choice model fo Derek Neal (1999)

##### Fields

- `beta::Real` : Discount factor in (0, 1)
- `N::Int` : Number of possible realizations of both epsilon and theta
- `B::Real` : upper bound for both epsilon and theta
- `theta::AbstractVector` : A grid of values on [0, B]
- `epsilon::AbstractVector` : A grid of values on [0, B]
- `F_probs::AbstractVector` : The pdf of each value associated with of F
- `G_probs::AbstractVector` : The pdf of each value associated with of G
- `F_mean::Real` : The mean of the distribution F
- `G_mean::Real` : The mean of the distribution G

"""

type CareerWorkerProblem
    beta::Real
    N::Int
    B::Real
    theta::AbstractVector
    epsilon::AbstractVector
    F_probs::AbstractVector
    G_probs::AbstractVector
    F_mean::Real
    G_mean::Real
end

"""

Constructor with default values for `CareerWorkerProblem`


##### Arguments

- `beta::Real(0.95)` : Discount factor in (0, 1)
- `B::Real(5.0)` : upper bound for both epsilon and theta
- `N::Real(50)` : Number of possible realizations of both epsilon and theta
- `F_a::Real(1), F_b::Real(1)` : Parameters of the distribution F

```

```

- `G_a::Real(1), G_b::Real(1)` : Parameters of the distribution F

##### Notes

There is also a version of this function that accepts keyword arguments for
each parameter
"""

function CareerWorkerProblem(beta::Real=0.95, B::Real=5.0, N::Real=50,
                             F_a::Real=1, F_b::Real=1, G_a::Real=1,
                             G_b::Real=1)
    theta = linspace(0, B, N)
    epsilon = copy(theta)
    F_probs::Vector{Float64} = pdf(BetaBinomial(N-1, F_a, F_b))
    G_probs::Vector{Float64} = pdf(BetaBinomial(N-1, G_a, G_b))
    F_mean = sum(theta .* F_probs)
    G_mean = sum(epsilon .* G_probs)
    CareerWorkerProblem(beta, N, B, theta, epsilon, F_probs, G_probs,
                        F_mean, G_mean)
end

# create kwarg version
function CareerWorkerProblem(;beta::Real=0.95, B::Real=5.0, N::Real=50,
                           F_a::Real=1, F_b::Real=1, G_a::Real=1,
                           G_b::Real=1)
    CareerWorkerProblem(beta, B, N, F_a, F_b, G_a, G_b)
end

"""

Apply the Bellman operator for a given model and initial value.

##### Arguments

- `cp::CareerWorkerProblem` : Instance of `CareerWorkerProblem`
- `v::Matrix` : Current guess for the value function
- `out::Matrix` : Storage for output
- `;ret_policy::Bool(false)` : Toggles return of value or policy functions

##### Returns

None, `out` is updated in place. If `ret_policy == true` `out` is filled with the
policy function, otherwise the value function is stored in `out`.

"""

function bellman_operator!(cp::CareerWorkerProblem, v::Array, out::Array;
                          ret_policy=false)
    # new life. This is a function of the distribution parameters and is
    # always constant. No need to recompute it in the loop
    v3 = (cp.G_mean + cp.F_mean + cp.beta .*
          cp.F_probs' * v * cp.G_probs)[1] # don't need 1 element array

    for j=1:cp.N
        for i=1:cp.N
            # stay put

```

```

v1 = cp.theta[i] + cp.epsilon[j] + cp.beta * v[i, j]

# new job
v2 = (cp.theta[i] .+ cp.G_mean .+ cp.beta .* v[i, :]*cp.G_probs)[1] # don't need a single element array

if ret_policy
    if v1 > max(v2, v3)
        action = 1
    elseif v2 > max(v1, v3)
        action = 2
    else
        action = 3
    end
    out[i, j] = action
else
    out[i, j] = max(v1, v2, v3)
end
end
end

function bellman_operator(cp::CareerWorkerProblem, v::Array; ret_policy=false)
    out = similar(v)
    bellman_operator!(cp, v, out, ret_policy=ret_policy)
    return out
end

"""
Extract the greedy policy (policy function) of the model.

##### Arguments

- `cp::CareerWorkerProblem` : Instance of `CareerWorkerProblem`
- `v::Matrix` : Current guess for the value function
- `out::Matrix` : Storage for output

##### Returns

None, `out` is updated in place to hold the policy function

"""

function get_greedy!(cp::CareerWorkerProblem, v::Array, out::Array)
    bellman_operator!(cp, v, out, ret_policy=true)
end

function get_greedy(cp::CareerWorkerProblem, v::Array)
    bellman_operator(cp, v, ret_policy=true)
end

```

The code defines

- a type `CareerWorkerProblem` that
 - encapsulates all the details of a particular parameterization
 - implement the Bellman operator T

In this model, T is defined by $Tv(\theta, \epsilon) = \max\{I, II, III\}$, where I , II and III are as given in (3.33), replacing V with v

The default probability distributions in `CareerWorkerProblem` correspond to discrete uniform distributions (see *the Beta-binomial figure*)

In fact all our default settings correspond to the version studied in [LS12], section 6.5.

Hence we can reproduce figures 6.5.1 and 6.5.2 shown there, which exhibit the value function and optimal policy respectively

Here's the value function

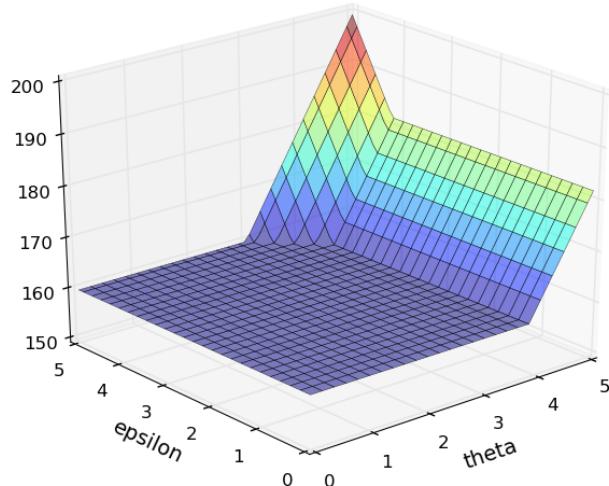


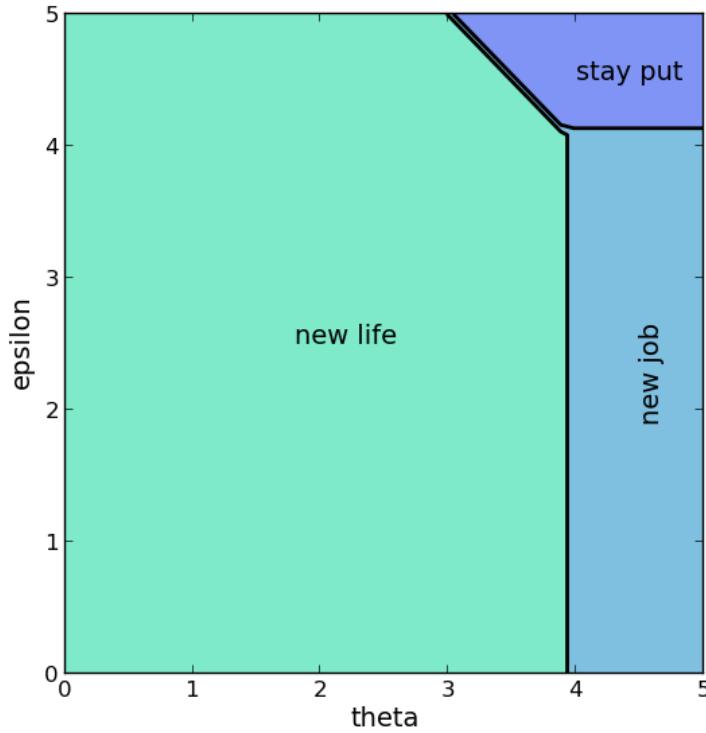
Fig. 3.1: Value function with uniform probabilities

The code used to produce this plot was `career/career_vf_plot.py`

The optimal policy can be represented as follows (see *Exercise 3* for code)

Interpretation:

- If both job and career are poor or mediocre, the worker will experiment with new job and new career
- If career is sufficiently good, the worker will hold it and experiment with new jobs until a sufficiently good one is found
- If both job and career are good, the worker will stay put



Notice that the worker will always hold on to a sufficiently good career, but not necessarily hold on to even the best paying job

The reason is that high lifetime wages require both variables to be large, and the worker cannot change careers without changing jobs

- Sometimes a good job must be sacrificed in order to change to a better career

Exercises

Exercise 1 Using the default parameterization in the type `CareerWorkerProblem`, generate and plot typical sample paths for θ and ϵ when the worker follows the optimal policy

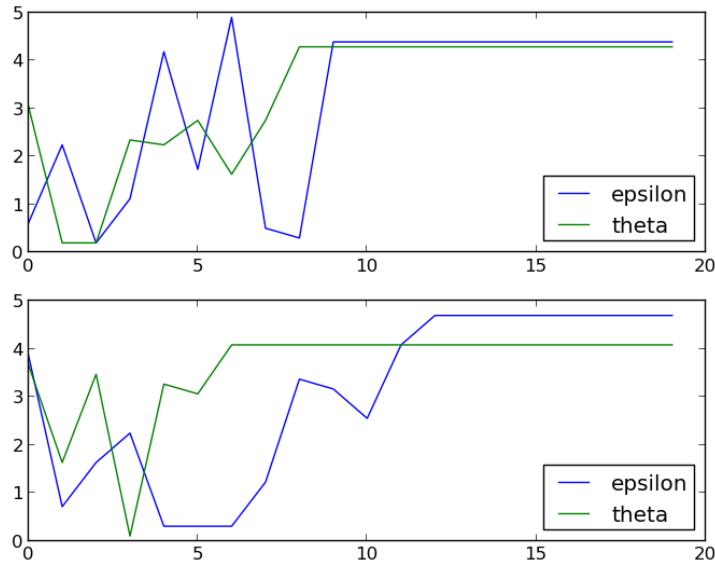
In particular, modulo randomness, reproduce the following figure (where the horizontal axis represents time)

Hint: To generate the draws from the distributions F and G , use the type `DiscreteRV`

Exercise 2 Let's now consider how long it takes for the worker to settle down to a permanent job, given a starting point of $(\theta, \epsilon) = (0, 0)$

In other words, we want to study the distribution of the random variable

$T^* :=$ the first point in time from which the worker's job no longer changes



Evidently, the worker's job becomes permanent if and only if (θ_t, ϵ_t) enters the "stay put" region of (θ, ϵ) space

Letting S denote this region, T^* can be expressed as the first passage time to S under the optimal policy:

$$T^* := \inf\{t \geq 0 \mid (\theta_t, \epsilon_t) \in S\}$$

Collect 25,000 draws of this random variable and compute the median (which should be about 7)

Repeat the exercise with $\beta = 0.99$ and interpret the change

Exercise 3 As best you can, reproduce *the figure showing the optimal policy*

Hint: The `get_greedy()` function returns a representation of the optimal policy where values 1, 2 and 3 correspond to "stay put", "new job" and "new life" respectively. Use this and `contourf` from `PyPlot.jl` to produce the different shadings.

Now set `G_a = G_b = 100` and generate a new figure with these parameters. Interpret.

Solutions

[Solution notebook](#)

On-the-Job Search

Contents

- *On-the-Job Search*
 - *Overview*
 - *Model*
 - *Implementation*
 - *Solving for Policies*
 - *Exercises*
 - *Solutions*

Overview

In this section we solve a simple on-the-job search model

- based on [LS12], exercise 6.18
- see also [add Jovanovic reference]

Model features

- job-specific human capital accumulation combined with on-the-job search
- infinite horizon dynamic programming with one state variable and two controls

Model

Let

- x_t denote the time- t job-specific human capital of a worker employed at a given firm
- w_t denote current wages

Let $w_t = x_t(1 - s_t - \phi_t)$, where

- ϕ_t is investment in job-specific human capital for the current role
- s_t is search effort, devoted to obtaining new offers from other firms.

For as long as the worker remains in the current job, evolution of $\{x_t\}$ is given by $x_{t+1} = G(x_t, \phi_t)$

When search effort at t is s_t , the worker receives a new job offer with probability $\pi(s_t) \in [0, 1]$

Value of offer is U_{t+1} , where $\{U_t\}$ is iid with common distribution F

Worker has the right to reject the current offer and continue with existing job.

In particular, $x_{t+1} = U_{t+1}$ if accepts and $x_{t+1} = G(x_t, \phi_t)$ if rejects

Letting $b_{t+1} \in \{0, 1\}$ be binary with $b_{t+1} = 1$ indicating an offer, we can write

$$x_{t+1} = (1 - b_{t+1})G(x_t, \phi_t) + b_{t+1} \max\{G(x_t, \phi_t), U_{t+1}\} \quad (3.34)$$

Agent's objective: maximize expected discounted sum of wages via controls $\{s_t\}$ and $\{\phi_t\}$

Taking the expectation of $V(x_{t+1})$ and using (3.34), the Bellman equation for this problem can be written as

$$V(x) = \max_{s+\phi \leq 1} \left\{ x(1-s-\phi) + \beta(1-\pi(s))V[G(x, \phi)] + \beta\pi(s) \int V[G(x, \phi) \vee u]F(du) \right\}. \quad (3.35)$$

Here nonnegativity of s and ϕ is understood, while $a \vee b := \max\{a, b\}$

Parameterization In the implementation below, we will focus on the parameterization

$$G(x, \phi) = A(x\phi)^\alpha, \quad \pi(s) = \sqrt{s} \quad \text{and} \quad F = \text{Beta}(2, 2)$$

with default parameter values

- $A = 1.4$
- $\alpha = 0.6$
- $\beta = 0.96$

The Beta(2,2) distribution is supported on $(0, 1)$. It has a unimodal, symmetric density peaked at 0.5.

Back-of-the-Envelope Calculations Before we solve the model, let's make some quick calculations that provide intuition on what the solution should look like.

To begin, observe that the worker has two instruments to build capital and hence wages:

1. invest in capital specific to the current job via ϕ
2. search for a new job with better job-specific capital match via s

Since wages are $x(1-s-\phi)$, marginal cost of investment via either ϕ or s is identical

Our risk neutral worker should focus on whatever instrument has the highest expected return

The relative expected return will depend on x

For example, suppose first that $x = 0.05$

- If $s = 1$ and $\phi = 0$, then since $G(x, \phi) = 0$, taking expectations of (3.34) gives expected next period capital equal to $\pi(s)\mathbb{E}U = \mathbb{E}U = 0.5$
- If $s = 0$ and $\phi = 1$, then next period capital is $G(x, \phi) = G(0.05, 1) \approx 0.23$

Both rates of return are good, but the return from search is better

Next suppose that $x = 0.4$

- If $s = 1$ and $\phi = 0$, then expected next period capital is again 0.5
- If $s = 0$ and $\phi = 1$, then $G(x, \phi) = G(0.4, 1) \approx 0.8$

Return from investment via ϕ dominates expected return from search

Combining these observations gives us two informal predictions:

1. At any given state x , the two controls ϕ and s will function primarily as substitutes — worker will focus on whichever instrument has the higher expected return
2. For sufficiently small x , search will be preferable to investment in job-specific human capital. For larger x , the reverse will be true

Now let's turn to implementation, and see if we can match our predictions.

Implementation

The QuantEcon package provides some code for solving the DP problem described above

See in particular `jv.jl`, which is repeated here for convenience

```
#=
@author : Spencer Lyon <spencer.lyon@nyu.edu>
@date: 2014-06-27

References
-----

Simple port of the file quantecon.models.jv

http://quant-econ.net/jl/jv.html
=#

using Distributions
using Interpolations

# NOTE: only brute-force approach is available in bellman operator. Waiting on a simple constrained opt

"""
A Jovanovic-type model of employment with on-the-job search.

The value function is given by

\\[V(x) = \\max_{\\phi, s} w(x, \\phi, s)\\]

for

w(x, phi, s) := x(1 - phi - s) + beta (1 - pi(s)) V(G(x, phi)) +
beta pi(s) E V[ max(G(x, phi), U)

where

* `x` : human capital
* `s` : search effort
* `phi` : investment in human capital
* `pi(s)` : probability of new offer given search level s
* `x(1 - \phi - s)` : wage
* `G(x, \phi)` : new human capital when current job retained

```

```

* `U` : Random variable with distribution F -- new draw of human capital

##### Fields

- `A::Real` : Parameter in human capital transition function
- `alpha::Real` : Parameter in human capital transition function
- `bet::Real` : Discount factor in (0, 1)
- `x_grid::Vector` : Grid for potential levels of x
- `G::Function` : Transition `function` for human capital
- `pi_func::Function` : `function` mapping search effort to the probability of getting a new job offer
- `F::UnivariateDistribution` : A univariate distribution from which the value of new job offers is drawn
- `quad_nodes::Vector` : Quadrature nodes for integrating over phi
- `quad_weights::Vector` : Quadrature weights for integrating over phi
- `epsilon::Float64` : A small number, used in optimization routine

"""
type JvWorker
    A::Real
    alpha::Real
    bet::Real
    x_grid::Vector{Float64}
    G::Function
    pi_func::Function
    F::UnivariateDistribution
    quad_nodes::Vector
    quad_weights::Vector
    epsilon::Float64
end

"""
Constructor with default values for `JvWorker`

##### Arguments

- `A::Real(1.4)` : Parameter in human capital transition function
- `alpha::Real(0.6)` : Parameter in human capital transition function
- `bet::Real(0.96)` : Discount factor in (0, 1)
- `grid_size::Int(50)` : Number of points in discrete grid for `x`
- `epsilon::Float(1e-4)` : A small number, used in optimization routine

##### Notes

There is also a version of this function that accepts keyword arguments for
each parameter

"""

function JvWorker(A=1.4, alpha=0.6, bet=0.96, grid_size=50, epsilon=1e-4)
    G(x, phi) = A .* (x .* phi).^alpha
    pi_func = sqrt
    F = Beta(2, 2)

    # integration bounds
    a, b, = quantile(F, 0.005), quantile(F, 0.995)

```

```

# quadrature nodes/weights
nodes, weights = qnwlege(21, a, b)

# Set up grid over the state space for DP
# Max of grid is the max of a large quantile value for F and the
# fixed point y = G(y, 1).
grid_max = max(A^(1.0 / (1.0 - alpha)), quantile(F, 1 - epsilon))

# range for linspace(epsilon, grid_max, grid_size). Needed for
# CoordInterpGrid below
x_grid = collect(linspace(epsilon, grid_max, grid_size))

JvWorker(A, alpha, bet, x_grid, G, pi_func, F, nodes, weights, epsilon)
end

# make kwarg version
JvWorker(;A=1.4, alpha=0.6, bet=0.96, grid_size=50, epsilon=1e-4) = JvWorker(A, alpha, bet, grid_size, epsilon)

"""
Apply the Bellman operator for a given model and initial value, returning only the value function

##### Arguments

- `jv::JvWorker` : Instance of `JvWorker`
- `V::Vector` : Current guess for the value function
- `new_V::Vector` : Storage for updated value function

##### Returns

None, `new_V` is updated in place with the value function.

##### Notes

Currently, only the brute-force approach is available. We are waiting on a simple constrained optimizer

"""

function bellman_operator!(jv::JvWorker, V::Vector, new_V::Vector)

    # simplify notation
    G, pi_func, F, bet, epsilon = jv.G, jv.pi_func, jv.F, jv.bet, jv.epsilon
    nodes, weights = jv.quad_nodes, jv.quad_weights

    # prepare interpoland of value function
    Vf = extrapolate(interpolate((jv.x_grid, ), V, Gridded(Linear())), Flat())

    # instantiate the linesearch variables
    max_val = -1.0
    cur_val = 0.0
    max_s = 1.0
    max_phi = 1.0
    search_grid = linspace(epsilon, 1.0, 15)

    for (i, x) in enumerate(jv.x_grid)

```

```

function w(z)
    s, phi = z
    function h(u)
        out = similar(u)
        for j in 1:length(u)
            out[j] = Vf[max(G(x, phi), u[j])] * pdf(F, u[j])
        end
        out
    end
    integral = do_quad(h, nodes, weights)
    q = pi_func(s) * integral + (1.0 - pi_func(s)) * Vf[G(x, phi)]

    return -x * (1.0 - phi - s) - bet * q
end

for s in search_grid
    for phi in search_grid
        if s + phi <= 1.0
            cur_val = -w((s, phi))
        else
            cur_val = -1.0
        end
        if cur_val > max_val
            max_val, max_s, max_phi = cur_val, s, phi
        end
    end
end

new_V[i] = max_val
end
end
"""

```

Apply the Bellman operator for a given model and initial value, returning policies

Arguments

- `jv::JvWorker` : Instance of `JvWorker`
- `V::Vector` : Current guess for the value function
- `out::Tuple{Vector, Vector}` : Storage for the two policy rules

Returns

None, `out` is updated in place with the two policy functions.

Notes

Currently, only the brute-force approach is available. We are waiting on a simple constrained optimizer

"""

```
function bellman_operator!(jv::JvWorker, V::Vector, out::Tuple{Vector, Vector})
```

```
# simplify notation
```

```

G, pi_func, F, bet, epsilon = jv.G, jv.pi_func, jv.F, jv.bet, jv.epsilon
nodes, weights = jv.quad_nodes, jv.quad_weights

# prepare interpoland of value function
Vf = extrapolate(interpolate((jv.x_grid, ), V, Gridded(Linear())), Flat())

# instantiate variables
s_policy, phi_policy = out[1], out[2]

# instantiate the linesearch variables
max_val = -1.0
cur_val = 0.0
max_s = 1.0
max_phi = 1.0
search_grid = linspace(epsilon, 1.0, 15)

for (i, x) in enumerate(jv.x_grid)

    function w(z)
        s, phi = z
        function h(u)
            out = similar(u)
            for j in 1:length(u)
                out[j] = Vf[max(G(x, phi), u[j])] * pdf(F, u[j])
            end
            out
        end
        integral = do_quad(h, nodes, weights)
        q = pi_func(s) * integral + (1.0 - pi_func(s)) * Vf[G(x, phi)]

        return - x * (1.0 - phi - s) - bet * q
    end

    for s in search_grid
        for phi in search_grid
            if s + phi <= 1.0
                cur_val = -w((s, phi))
            else
                cur_val = -1.0
            end
            if cur_val > max_val
                max_val, max_s, max_phi = cur_val, s, phi
            end
        end
    end

    s_policy[i], phi_policy[i] = max_s, max_phi
end
end

function bellman_operator(jv::JvWorker, V::Vector; ret_policies=false)
    if ret_policies
        out = (similar(V), similar(V))
    end
end

```

```

    else
        out = similar(V)
    end
    bellman_operator!(jv, V, out)
    return out
end

```

The code is written to be relatively generic—and hence reusable

- For example, we use generic $G(x, \phi)$ instead of specific $A(x\phi)^\alpha$

Regarding the imports

- `fixed_quad` is a simple non-adaptive integration routine
- `fmin_slsqp` is a minimization routine that permits inequality constraints

Next we build a type called `JvWorker` that

- packages all the parameters and other basic attributes of a given model
- Implements the method `bellman_operator` for value function iteration

The `bellman_operator` method takes a candidate value function V and updates it to TV via

$$TV(x) = - \min_{s+\phi \leq 1} w(s, \phi)$$

where

$$w(s, \phi) := - \left\{ x(1 - s - \phi) + \beta(1 - \pi(s))V[G(x, \phi)] + \beta\pi(s) \int V[G(x, \phi) \vee u]F(du) \right\} \quad (3.36)$$

Here we are minimizing instead of maximizing to fit with SciPy's optimization routines

When we represent V , it will be with a Julia array `V` giving values on grid `x_grid`

But to evaluate the right-hand side of (3.36), we need a function, so we replace the arrays `V` and `x_grid` with a function `Vf` that gives linear interpolation of `V` on `x_grid`

Hence in the preliminaries of `bellman_operator`

- from the array `V` we define a linear interpolation `Vf` of its values
 - `c1` is used to implement the constraint $s + \phi \leq 1$
 - `c2` is used to implement $s \geq \epsilon$, a numerically stable alternative to the true constraint $s \geq 0$
 - `c3` does the same for ϕ

Inside the `for` loop, for each `x` in the grid over the state space, we set up the function $w(z) = w(s, \phi)$ defined in (3.36).

The function is minimized over all feasible (s, ϕ) pairs, either by

- a relatively sophisticated solver from SciPy called `fmin_slsqp`, or
- brute force search over a grid

The former is much faster, but convergence to the global optimum is not guaranteed. Grid search is a simple way to check results

Solving for Policies

Let's plot the optimal policies and see what they look like

The code is in a file `jv/jv_test.jl` from the [applications repository](#) and looks as follows

```
V = compute_fixed_point(f, v_init, max_iter=200)

s_policy, phi_policy = bellman_operator(wp, V, ret_policies=true)

# === plot solution === #

plot(wp.x_grid, [phi_policy s_policy V],
      title = ["phi policy" "s policy" "value function"],
      color = [:orange :blue :green], width = 3,
      xaxis = ("x", (0.0, maximum(wp.x_grid))),
      yaxis = ((-0.1, 1.1)),
      layout = (3,1), bottom_margin = 20mm, show = true)
```

It produces the following figure

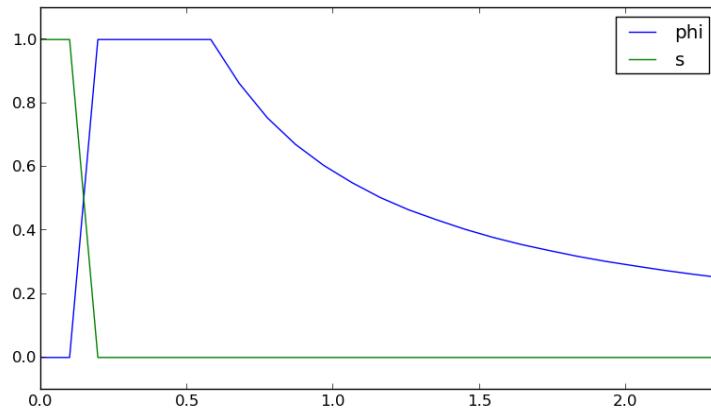


Fig. 3.2: Optimal policies

The horizontal axis is the state x , while the vertical axis gives $s(x)$ and $\phi(x)$

Overall, the policies match well with our predictions from *section*.

- Worker switches from one investment strategy to the other depending on relative return
- For low values of x , the best option is to search for a new job
- Once x is larger, worker does better by investing in human capital specific to the current position

Exercises

Exercise 1 Let's look at the dynamics for the state process $\{x_t\}$ associated with these policies.

The dynamics are given by (3.34) when ϕ_t and s_t are chosen according to the optimal policies, and $\mathbb{P}\{b_{t+1} = 1\} = \pi(s_t)$.

Since the dynamics are random, analysis is a bit subtle

One way to do it is to plot, for each x in a relatively fine grid called `plot_grid`, a large number K of realizations of x_{t+1} given $x_t = x$. Plot this with one dot for each realization, in the form of a 45 degree diagram. Set:

```
K = 50
plot_grid_max, plot_grid_size = 1.2, 100
plot_grid = linspace(0, plot_grid_max, plot_grid_size)
fig, ax = subplots()
ax[:set_xlim](0, plot_grid_max)
ax[:set_ylim](0, plot_grid_max)
```

By examining the plot, argue that under the optimal policies, the state x_t will converge to a constant value \bar{x} close to unity

Argue that at the steady state, $s_t \approx 0$ and $\phi_t \approx 0.6$.

Exercise 2 In the preceding exercise we found that s_t converges to zero and ϕ_t converges to about 0.6

Since these results were calculated at a value of β close to one, let's compare them to the best choice for an *infinitely* patient worker.

Intuitively, an infinitely patient worker would like to maximize steady state wages, which are a function of steady state capital.

You can take it as given—it's certainly true—that the infinitely patient worker does not search in the long run (i.e., $s_t = 0$ for large t)

Thus, given ϕ , steady state capital is the positive fixed point $x^*(\phi)$ of the map $x \mapsto G(x, \phi)$.

Steady state wages can be written as $w^*(\phi) = x^*(\phi)(1 - \phi)$

Graph $w^*(\phi)$ with respect to ϕ , and examine the best choice of ϕ

Can you give a rough interpretation for the value that you see?

Solutions

Solution notebook

Search with Offer Distribution Unknown

Contents

- *Search with Offer Distribution Unknown*
 - *Overview*
 - *Model*
 - *Take 1: Solution by VFI*
 - *Take 2: A More Efficient Method*
 - *Exercises*
 - *Solutions*

Overview

In this lecture we consider an extension of the job search model developed by John J. McCall [[McC70](#)]

In the McCall model, an unemployed worker decides when to accept a permanent position at a specified wage, given

- his or her discount rate
- the level of unemployment compensation
- the distribution from which wage offers are drawn

In the version considered below, the wage distribution is unknown and must be learned

- Based on the presentation in [[LS12](#)], section 6.6

Model features

- Infinite horizon dynamic programming with two states and one binary control
- Bayesian updating to learn the unknown distribution

Model

Let's first recall the basic McCall model [[McC70](#)] and then add the variation we want to consider

The Basic McCall Model Consider an unemployed worker who is presented in each period with a permanent job offer at wage w_t

At time t , our worker has two choices

1. Accept the offer and work permanently at constant wage w_t
2. Reject the offer, receive unemployment compensation c , and reconsider next period

The wage sequence $\{w_t\}$ is iid and generated from known density h

The worker aims to maximize the expected discounted sum of earnings $\mathbb{E} \sum_{t=0}^{\infty} \beta^t y_t$

Trade-off:

- Waiting too long for a good offer is costly, since the future is discounted
- Accepting too early is costly, since better offers will arrive with probability one

Let $V(w)$ denote the maximal expected discounted sum of earnings that can be obtained by an unemployed worker who starts with wage offer w in hand

The function V satisfies the recursion

$$V(w) = \max \left\{ \frac{w}{1-\beta}, c + \beta \int V(w') h(w') dw' \right\} \quad (3.37)$$

where the two terms on the r.h.s. are the respective payoffs from accepting and rejecting the current offer w

The optimal policy is a map from states into actions, and hence a binary function of w

Not surprisingly, it turns out to have the form $\mathbf{1}\{w \geq \bar{w}\}$, where

- \bar{w} is a constant depending on (β, h, c) called the *reservation wage*
- $\mathbf{1}\{w \geq \bar{w}\}$ is an indicator function returning 1 if $w \geq \bar{w}$ and 0 otherwise
- 1 indicates “accept” and 0 indicates “reject”

For further details see [LS12], section 6.3

Offer Distribution Unknown Now let's extend the model by considering the variation presented in [LS12], section 6.6

The model is as above, apart from the fact that

- the density h is unknown
- the worker learns about h by starting with a prior and updating based on wage offers that he/she observes

The worker knows there are two possible distributions F and G — with densities f and g

At the start of time, “nature” selects h to be either f or g — the wage distribution from which the entire sequence $\{w_t\}$ will be drawn

This choice is not observed by the worker, who puts prior probability π_0 on f being chosen

Update rule: worker's time t estimate of the distribution is $\pi_t f + (1 - \pi_t)g$, where π_t updates via

$$\pi_{t+1} = \frac{\pi_t f(w_{t+1})}{\pi_t f(w_{t+1}) + (1 - \pi_t)g(w_{t+1})} \quad (3.38)$$

This last expression follows from Bayes' rule, which tells us that

$$\mathbb{P}\{h = f \mid W = w\} = \frac{\mathbb{P}\{W = w \mid h = f\} \mathbb{P}\{h = f\}}{\mathbb{P}\{W = w\}} \quad \text{and} \quad \mathbb{P}\{W = w\} = \sum_{\psi \in \{f,g\}} \mathbb{P}\{W = w \mid h = \psi\} \mathbb{P}\{h = \psi\}$$

The fact that (3.38) is recursive allows us to progress to a recursive solution method

Letting

$$h_\pi(w) := \pi f(w) + (1 - \pi)g(w) \quad \text{and} \quad q(w, \pi) := \frac{\pi f(w)}{\pi f(w) + (1 - \pi)g(w)}$$

we can express the value function for the unemployed worker recursively as follows

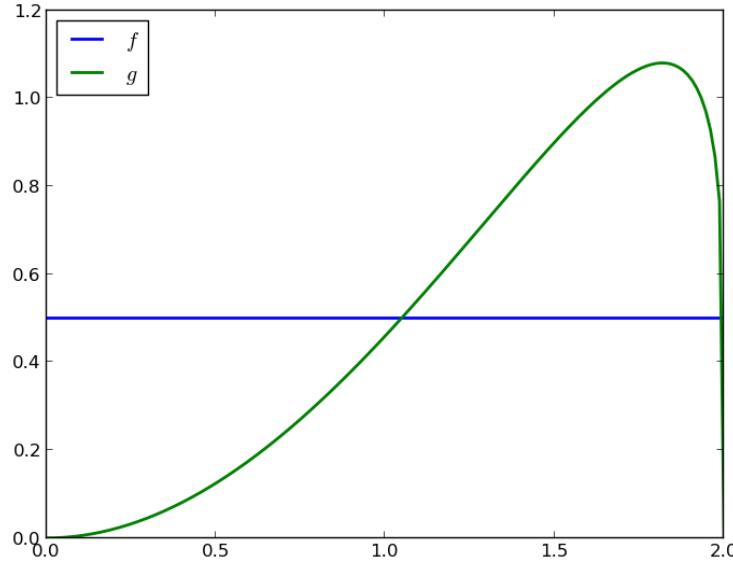
$$V(w, \pi) = \max \left\{ \frac{w}{1 - \beta}, c + \beta \int V(w', \pi') h_\pi(w') dw' \right\} \quad \text{where} \quad \pi' = q(w', \pi) \quad (3.39)$$

Notice that the current guess π is a state variable, since it affects the worker's perception of probabilities for future rewards

Parameterization Following section 6.6 of [LS12], our baseline parameterization will be

- f is Beta(1, 1) scaled (i.e., draws are multiplied by) some factor w_m
- g is Beta(3, 1.2) scaled (i.e., draws are multiplied by) the same factor w_m
- $\beta = 0.95$ and $c = 0.6$

With $w_m = 2$, the densities f and g have the following shape



Looking Forward What kind of optimal policy might result from (3.39) and the parameterization specified above?

Intuitively, if we accept at w_a and $w_a \leq w_b$, then — all other things being given — we should also accept at w_b

This suggests a policy of accepting whenever w exceeds some threshold value \bar{w}

But \bar{w} should depend on π — in fact it should be decreasing in π because

- f is a less attractive offer distribution than g
- larger π means more weight on f and less on g

Thus larger π depresses the worker's assessment of her future prospects, and relatively low current offers become more attractive

Summary: We conjecture that the optimal policy is of the form $\mathbb{1}\{w \geq \bar{w}(\pi)\}$ for some decreasing function \bar{w}

Take 1: Solution by VFI

Let's set about solving the model and see how our results match with our intuition

We begin by solving via value function iteration (VFI), which is natural but ultimately turns out to be second best

VFI is implemented in the file `odu/odu.jl` contained in the [QuantEcon.applications](#) repo

The code is as follows

```
#=
Solves the "Offer Distribution Unknown" Model by value function
iteration and a second faster method discussed in the corresponding
quantecon lecture.

@author : Spencer Lyon <spencer.lyon@nyu.edu>

@date: 2014-08-14

References
-----

http://quant-econ.net/jl/odu.html

=#
using Distributions
"""

Unemployment/search problem where offer distribution is unknown

##### Fields

- `bet::Real` : Discount factor on (0, 1)
- `c::Real` : Unemployment compensation
- `F::Distribution` : Offer distribution `F`
- `G::Distribution` : Offer distribution `G`
- `f::Function` : The pdf of `F`
- `g::Function` : The pdf of `G`
- `n_w::Int` : Number of points on the grid for w
- `w_max::Real` : Maximum wage offer
- `w_grid::AbstractVector` : Grid of wage offers w
- `n_pi::Int` : Number of points on grid for pi
```

```

- `pi_min::Real` : Minimum of pi grid
- `pi_max::Real` : Maximum of pi grid
- `pi_grid::AbstractVector` : Grid of probabilities pi
- `quad_nodes::Vector` : Notes for quadrature ofer offers
- `quad_weights::Vector` : Weights for quadrature ofer offers

"""
type SearchProblem
    bet::Real
    c::Real
    F::Distribution
    G::Distribution
    f::Function
    g::Function
    n_w::Int
    w_max::Real
    w_grid::AbstractVector
    n_pi::Int
    pi_min::Real
    pi_max::Real
    pi_grid::AbstractVector
    quad_nodes::Vector
    quad_weights::Vector
end

"""
Constructor for `SearchProblem` with default values

##### Arguments

- `bet::Real(0.95)` : Discount factor in (0, 1)
- `c::Real(0.6)` : Unemployment compensation
- `F_a::Real(1), F_b::Real(1)` : Parameters of `Beta` distribution for `F`
- `G_a::Real(3), G_b::Real(1.2)` : Parameters of `Beta` distribution for `G`
- `w_max::Real(2)` : Maximum of wage offer grid
- `w_grid_size::Int(40)` : Number of points in wage offer grid
- `pi_grid_size::Int(40)` : Number of points in probability grid

##### Notes

There is also a version of this function that accepts keyword arguments for
each parameter

"""
function SearchProblem(bet=0.95, c=0.6, F_a=1, F_b=1, G_a=3, G_b=1.2,
                      w_max=2, w_grid_size=40, pi_grid_size=40)

    F = Beta(F_a, F_b)
    G = Beta(G_a, G_b)

    # NOTE: the x./w_max)./w_max in these functions makes our dist match
    #       the scipy one with scale=w_max given
    f(x) = pdf(F, x./w_max)./w_max

```

```

g(x) = pdf(G, x./w_max)./w_max

pi_min = 1e-3 # avoids instability
pi_max = 1 - pi_min

w_grid = linspace(0, w_max, w_grid_size)
pi_grid = linspace(pi_min, pi_max, pi_grid_size)

nodes, weights = qnwlege(21, 0.0, w_max)

SearchProblem(bet, c, F, G, f, g,
              w_grid_size, w_max, w_grid,
              pi_grid_size, pi_min, pi_max, pi_grid, nodes, weights)
end

# make kwarg version
function SearchProblem(;bet=0.95, c=0.6, F_a=1, F_b=1, G_a=3, G_b=1.2,
                      w_max=2, w_grid_size=40, pi_grid_size=40)
    SearchProblem(bet, c, F_a, F_b, G_a, G_b, w_max, w_grid_size,
                  pi_grid_size)
end

function q(sp::SearchProblem, w, pi_val)
    new_pi = 1.0 ./ (1 + ((1 - pi_val) .* sp.g(w)) ./ (pi_val .* sp.f(w)))

    # Return new_pi when in [pi_min, pi_max] and else end points
    return clamp(new_pi, sp.pi_min, sp.pi_max)
end

"""
Apply the Bellman operator for a given model and initial value.

#### Arguments

- `sp::SearchProblem` : Instance of `SearchProblem`
- `v::Matrix` : Current guess for the value function
- `out::Matrix` : Storage for output.
- `;ret_policy::Bool(false)` : Toggles return of value or policy functions

#### Returns

None, `out` is updated in place. If `ret_policy == true` `out` is filled with the
policy function, otherwise the value function is stored in `out`.

"""

function bellman_operator!(sp::SearchProblem, v::Matrix, out::Matrix;
                           ret_policy::Bool=false)
    # Simplify names
    f, g, bet, c = sp.f, sp.g, sp.bet, sp.c
    nodes, weights = sp.quad_nodes, sp.quad_weights

    vf = extrapolate(interpolate((sp.w_grid, sp.pi_grid), v,
                                 Gridded(Linear())), Flat())

```

```

# set up quadrature nodes/weights
# q_nodes, q_weights = qnwlege(21, 0.0, sp.w_max)

for w_i=1:sp.n_w
    w = sp.w_grid[w_i]

    # calculate v1
    v1 = w / (1 - bet)

    for pi_j=1:sp.n_pi
        _pi = sp.pi_grid[pi_j]

        # calculate v2
        function integrand(m)
            quad_out = similar(m)
            for i=1:length(m)
                mm = m[i]
                quad_out[i] = vf[mm, q(sp, mm, _pi)] * (_pi*f(mm) +
                                                (1-_pi)*g(mm))
            end
            return quad_out
        end
        integral = do_quad(integrand, nodes, weights)
        # integral = do_quad(integrand, q_nodes, q_weights)
        v2 = c + bet * integral

        # return policy if asked for, otherwise return max of values
        out[w_i, pi_j] = ret_policy ? v1 > v2 : max(v1, v2)
    end
end
return out
end

function bellman_operator(sp::SearchProblem, v::Matrix;
                           ret_policy::Bool=false)
    out_type = ret_policy ? Bool : Float64
    out = Array(out_type, sp.n_w, sp.n_pi)
    bellman_operator!(sp, v, out, ret_policy=ret_policy)
end

"""
Extract the greedy policy (policy function) of the model.

##### Arguments

- `sp::SearchProblem` : Instance of `SearchProblem`
- `v::Matrix` : Current guess for the value function
- `out::Matrix` : Storage for output

##### Returns

None, `out` is updated in place to hold the policy function

```

```

"""
function get_greedy!(sp::SearchProblem, v::Matrix, out::Matrix)
    bellman_operator!(sp, v, out, ret_policy=true)
end

get_greedy(sp::SearchProblem, v::Matrix) = bellman_operator(sp, v,
                                                               ret_policy=true)

"""

Updates the reservation wage function guess phi via the operator Q.

##### Arguments

- `sp::SearchProblem` : Instance of `SearchProblem`
- `phi::Vector` : Current guess for phi
- `out::Vector` : Storage for output

##### Returns

None, `out` is updated in place to hold the updated levels of phi
"""

function res_wage_operator!(sp::SearchProblem, phi::Vector, out::Vector)
    # Simplify name
    f, g, bet, c = sp.f, sp.g, sp.bet, sp.c

    # Construct interpolator over pi_grid, given phi
    phi_f = extrapolate(interpolate((sp.pi_grid, ), phi, Gridded(Linear())),
                         Flat())

    # set up quadrature nodes/weights
    q_nodes, q_weights = qnwlege(7, 0.0, sp.w_max)

    for (i, _pi) in enumerate(sp.pi_grid)
        integrand(x) = max(x, phi_f[q(sp, x, _pi)]).*( _pi*f(x) + (1-_pi)*g(x))
        integral = do_quad(integrand, q_nodes, q_weights)
        out[i] = (1 - bet)*c + bet*integral
    end
end

"""

Updates the reservation wage function guess phi via the operator Q.

See the documentation for the mutating method of this function for more details
on arguments
"""

function res_wage_operator(sp::SearchProblem, phi::Vector)
    out = similar(phi)
    res_wage_operator!(sp, phi, out)
    return out
end

```

The type `SearchProblem` is used to store parameters and methods needed to compute optimal

actions

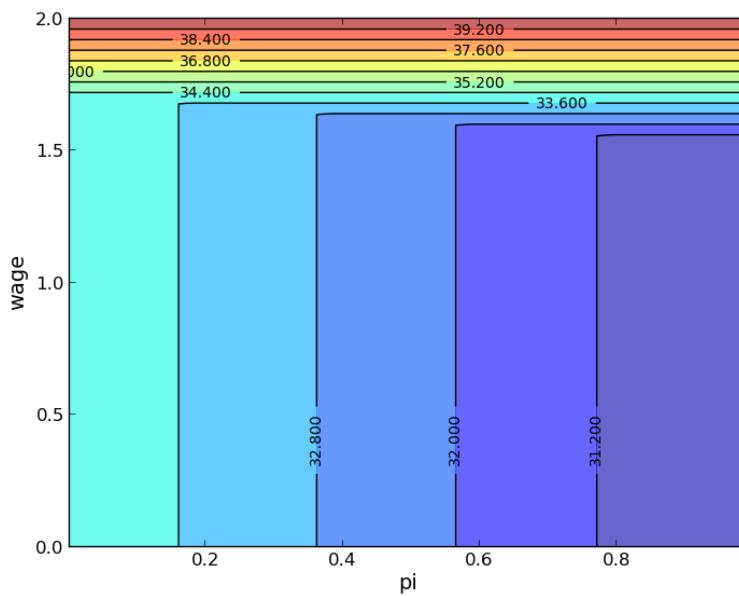
The Bellman operator is implemented as the method `bellman_operator()`, while `get_greedy()` computes an approximate optimal policy from a guess v of the value function

We will omit a detailed discussion of the code because there is a more efficient solution method

These ideas are implemented in the `res_wage_operator` method

Before explaining it let's look quickly at solutions computed from value function iteration

Here's the value function:



The optimal policy:

Code for producing these figures can be found in file `odu/odu_vfi_plots.jl` from the [applications repository](#)

The code takes several minutes to run

The results fit well with our intuition from section *looking forward*

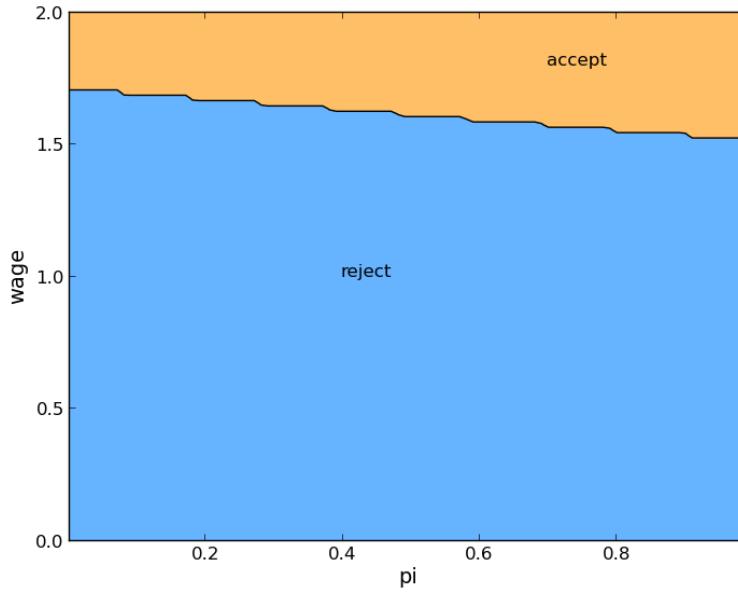
- The black line in the figure above corresponds to the function $\bar{w}(\pi)$ introduced there
- decreasing as expected

Take 2: A More Efficient Method

Our implementation of VFI can be optimized to some degree,

But instead of pursuing that, let's consider another method to solve for the optimal policy

Uses iteration with an operator having the same contraction rate as the Bellman operator, but



- one dimensional rather than two dimensional
- no maximization step

As a consequence, the algorithm is orders of magnitude faster than VFI

This section illustrates the point that when it comes to programming, a bit of mathematical analysis goes a long way

Another Functional Equation To begin, note that when $w = \bar{w}(\pi)$, the worker is indifferent between accepting and rejecting

Hence the two choices on the right-hand side of (3.39) have equal value:

$$\frac{\bar{w}(\pi)}{1 - \beta} = c + \beta \int V(w', \pi') h_\pi(w') dw' \quad (3.40)$$

Together, (3.39) and (3.40) give

$$V(w, \pi) = \max \left\{ \frac{w}{1 - \beta}, \frac{\bar{w}(\pi)}{1 - \beta} \right\} \quad (3.41)$$

Combining (3.40) and (3.41), we obtain

$$\frac{\bar{w}(\pi)}{1 - \beta} = c + \beta \int \max \left\{ \frac{w'}{1 - \beta}, \frac{\bar{w}(\pi')}{1 - \beta} \right\} h_\pi(w') dw'$$

Multiplying by $1 - \beta$, substituting in $\pi' = q(w', \pi)$ and using \circ for composition of functions yields

$$\bar{w}(\pi) = (1 - \beta)c + \beta \int \max \{ w', \bar{w} \circ q(w', \pi) \} h_\pi(w') dw' \quad (3.42)$$

Equation (3.42) can be understood as a functional equation, where \bar{w} is the unknown function

- Let's call it the *reservation wage functional equation* (RWFE)
- The solution \bar{w} to the RWFE is the object that we wish to compute

Solving the RWFE To solve the RWFE, we will first show that its solution is the fixed point of a contraction mapping

To this end, let

- $b[0, 1]$ be the bounded real-valued functions on $[0, 1]$
- $\|\psi\| := \sup_{x \in [0, 1]} |\psi(x)|$

Consider the operator Q mapping $\psi \in b[0, 1]$ into $Q\psi \in b[0, 1]$ via

$$(Q\psi)(\pi) = (1 - \beta)c + \beta \int \max \{w', \psi \circ q(w', \pi)\} h_\pi(w') dw' \quad (3.43)$$

Comparing (3.42) and (3.43), we see that the set of fixed points of Q exactly coincides with the set of solutions to the RWFE

- If $Q\bar{w} = \bar{w}$ then \bar{w} solves (3.42) and vice versa

Moreover, for any $\psi, \phi \in b[0, 1]$, basic algebra and the triangle inequality for integrals tells us that

$$|(Q\psi)(\pi) - (Q\phi)(\pi)| \leq \beta \int |\max \{w', \psi \circ q(w', \pi)\} - \max \{w', \phi \circ q(w', \pi)\}| h_\pi(w') dw' \quad (3.44)$$

Working case by case, it is easy to check that for real numbers a, b, c we always have

$$|\max\{a, b\} - \max\{a, c\}| \leq |b - c| \quad (3.45)$$

Combining (3.44) and (3.45) yields

$$|(Q\psi)(\pi) - (Q\phi)(\pi)| \leq \beta \int |\psi \circ q(w', \pi) - \phi \circ q(w', \pi)| h_\pi(w') dw' \leq \beta \|\psi - \phi\| \quad (3.46)$$

Taking the supremum over π now gives us

$$\|Q\psi - Q\phi\| \leq \beta \|\psi - \phi\| \quad (3.47)$$

In other words, Q is a contraction of modulus β on the complete metric space $(b[0, 1], \|\cdot\|)$

Hence

- A unique solution \bar{w} to the RWFE exists in $b[0, 1]$
- $Q^k \psi \rightarrow \bar{w}$ uniformly as $k \rightarrow \infty$, for any $\psi \in b[0, 1]$

Implementation These ideas are implemented in the `res_wage_operator` method from `odu.jl` as shown above

The method corresponds to action of the operator Q

The following exercise asks you to exploit these facts to compute an approximation to \bar{w}

Exercises

Exercise 1 Use the default parameters and the `res_wage_operator` method to compute an optimal policy

Your result should coincide closely with the figure for the optimal policy *shown above*

Try experimenting with different parameters, and confirm that the change in the optimal policy coincides with your intuition

Solutions

[Solution notebook](#)

Optimal Savings

Contents

- *Optimal Savings*
 - [Overview](#)
 - [The Optimal Savings Problem](#)
 - [Computation](#)
 - [Exercises](#)
 - [Solutions](#)

Overview

Next we study the standard optimal savings problem for an infinitely lived consumer—the “common ancestor” described in [\[LS12\]](#), section 1.3

- Also known as the income fluctuation problem
- An important sub-problem for many representative macroeconomic models
 - [\[Aiy94\]](#)
 - [\[Hug93\]](#)
 - etc.
- Useful references include [\[Dea91\]](#), [\[DH10\]](#), [\[Kuh13\]](#), [\[Rab02\]](#), [\[Rei09\]](#) and [\[SE77\]](#)

Our presentation of the model will be relatively brief

- For further details on economic intuition, implication and models, see [\[LS12\]](#)
- Proofs of all mathematical results stated below can be found in this paper

In this lecture we will explore an alternative to value function iteration (VFI) called *policy function iteration* (PFI)

- Based on the Euler equation, and not to be confused with Howard's policy iteration algorithm
- Globally convergent under mild assumptions, even when utility is unbounded (both above and below)
- Numerically, turns out to be faster and more efficient than VFI for this model

Model features

- Infinite horizon dynamic programming with two states and one control

The Optimal Savings Problem

Consider a household that chooses a state-contingent consumption plan $\{c_t\}_{t \geq 0}$ to maximize

$$\mathbb{E} \sum_{t=0}^{\infty} \beta^t u(c_t)$$

subject to

$$c_t + a_{t+1} \leq Ra_t + z_t, \quad c_t \geq 0, \quad a_t \geq -b \quad t = 0, 1, \dots \quad (3.48)$$

Here

- $\beta \in (0, 1)$ is the discount factor
- a_t is asset holdings at time t , with ad-hoc borrowing constraint $a_t \geq -b$
- c_t is consumption
- z_t is non-capital income (wages, unemployment compensation, etc.)
- $R := 1 + r$, where $r > 0$ is the interest rate on savings

Assumptions

1. $\{z_t\}$ is a finite Markov process with Markov matrix Π taking values in Z
2. $|Z| < \infty$ and $Z \subset (0, \infty)$
3. $r > 0$ and $\beta R < 1$
4. u is smooth, strictly increasing and strictly concave with $\lim_{c \rightarrow 0} u'(c) = \infty$ and $\lim_{c \rightarrow \infty} u'(c) = 0$

The asset space is $[-b, \infty)$ and the state is the pair $(a, z) \in S := [-b, \infty) \times Z$

A *feasible consumption path* from $(a, z) \in S$ is a consumption sequence $\{c_t\}$ such that $\{c_t\}$ and its induced asset path $\{a_t\}$ satisfy

1. $(a_0, z_0) = (a, z)$
2. the feasibility constraints in (3.48), and

3. measurability of c_t w.r.t. the filtration generated by $\{z_1, \dots, z_t\}$

The meaning of the third point is just that consumption at time t can only be a function of outcomes that have already been observed

The *value function* $V: S \rightarrow \mathbb{R}$ is defined by

$$V(a, z) := \sup \mathbb{E} \left\{ \sum_{t=0}^{\infty} \beta^t u(c_t) \right\} \quad (3.49)$$

where the supremum is over all feasible consumption paths from (a, z) .

An *optimal consumption path* from (a, z) is a feasible consumption path from (a, z) that attains the supremum in (3.49)

Given our assumptions, it is known that

1. For each $(a, z) \in S$, a unique optimal consumption path from (a, z) exists
2. This path is the unique feasible path from (a, z) satisfying the Euler equality

$$u'(c_t) = \max \{ \beta R \mathbb{E}_t [u'(c_{t+1})], u'(Ra_t + z_t + b) \} \quad (3.50)$$

and the transversality condition

$$\lim_{t \rightarrow \infty} \beta^t \mathbb{E} [u'(c_t) a_{t+1}] = 0. \quad (3.51)$$

Moreover, there exists an *optimal consumption function* $c^*: S \rightarrow [0, \infty)$ such that the path from (a, z) generated by

$$(a_0, z_0) = (a, z), \quad z_{t+1} \sim \Pi(z_t, dy), \quad c_t = c^*(a_t, z_t) \quad \text{and} \quad a_{t+1} = Ra_t + z_t - c_t$$

satisfies both (3.50) and (3.51), and hence is the unique optimal path from (a, z)

In summary, to solve the optimization problem, we need to compute c^*

Computation

There are two standard ways to solve for c^*

1. Value function iteration (VFI)
2. Policy function iteration (PFI) using the Euler equality

Policy function iteration

We can rewrite (3.50) to make it a statement about functions rather than random variables

In particular, consider the functional equation

$$u' \circ c(a, z) = \max \left\{ \gamma \int u' \circ c \{ Ra + z - c(a, z), \dot{z} \} \Pi(z, d\dot{z}), u'(Ra + z + b) \right\} \quad (3.52)$$

where $\gamma := \beta R$ and $u' \circ c(s) := u'(c(s))$

Equation (3.52) is a functional equation in c

In order to identify a solution, let \mathcal{C} be the set of candidate consumption functions $c: S \rightarrow \mathbb{R}$ such that

- each $c \in \mathcal{C}$ is continuous and (weakly) increasing
- $\min Z \leq c(a, z) \leq Ra + z + b$ for all $(a, z) \in S$

In addition, let $K: \mathcal{C} \rightarrow \mathcal{C}$ be defined as follows:

For given $c \in \mathcal{C}$, the value $Kc(a, z)$ is the unique $t \in J(a, z)$ that solves

$$u'(t) = \max \left\{ \gamma \int u' \circ c \{Ra + z - t, \dot{z}\} \Pi(z, d\dot{z}), u'(Ra + z + b) \right\} \quad (3.53)$$

where

$$J(a, z) := \{t \in \mathbb{R} : \min Z \leq t \leq Ra + z + b\} \quad (3.54)$$

We refer to K as Coleman's policy function operator [Col90]

It is known that

- K is a contraction mapping on \mathcal{C} under the metric

$$\rho(c, d) := \|u' \circ c - u' \circ d\| := \sup_{s \in S} |u'(c(s)) - u'(d(s))| \quad (c, d \in \mathcal{C})$$

- The metric ρ is complete on \mathcal{C}
- Convergence in ρ implies uniform convergence on compacts

In consequence, K has a unique fixed point $c^* \in \mathcal{C}$ and $K^n c \rightarrow c^*$ as $n \rightarrow \infty$ for any $c \in \mathcal{C}$

By the definition of K , the fixed points of K in \mathcal{C} coincide with the solutions to (3.52) in \mathcal{C}

In particular, it can be shown that the path $\{c_t\}$ generated from $(a_0, z_0) \in S$ using policy function c^* is the unique optimal path from $(a_0, z_0) \in S$

TL;DR The unique optimal policy can be computed by picking any $c \in \mathcal{C}$ and iterating with the operator K defined in (3.53)

Value function iteration

The Bellman operator for this problem is given by

$$Tv(a, z) = \max_{0 \leq c \leq Ra + z + b} \left\{ u(c) + \beta \int v(Ra + z - c, \dot{z}) \Pi(z, d\dot{z}) \right\} \quad (3.55)$$

We have to be careful with VFI (i.e., iterating with T) in this setting because u is not assumed to be bounded

- In fact typically unbounded both above and below — e.g. $u(c) = \log c$
- In which case, the standard DP theory does not apply
- $T^n v$ is not guaranteed to converge to the value function for arbitrary continuous bounded v

Nonetheless, we can always try the strategy "iterate and hope"

- In this case we can check the outcome by comparing with PFI
- The latter is known to converge, as described above

Implementation The code in `ifp.jl` from `QuantEcon.applications` provides implementations of both VFI and PFI

The code is repeated here and a description and clarifications are given below

```
#=
Tools for solving the standard optimal savings / income fluctuation
problem for an infinitely lived consumer facing an exogenous income
process that evolves according to a Markov chain.

@author : Spencer Lyon <spencer.lyon@nyu.edu>

@date: 2014-08-18

References
-----
http://quant-econ.net/jl/ifp.html

=#
using Interpolations
using Optim

# utility and marginal utility functions
u(x) = log(x)
du(x) = 1 ./ x

"""
Income fluctuation problem

##### Fields

- `r::Float64` : Strictly positive interest rate
- `R::Float64` : The interest rate plus 1 (strictly greater than 1)
- `bet::Float64` : Discount rate in (0, 1)
- `b::Float64` : The borrowing constraint
- `Pi::Matrix{Float64}` : Transition matrix for `z`
- `z_vals::Vector{Float64}` : Levels of productivity
- `asset_grid::LinSpace{Float64}` : Grid of asset values

"""
type ConsumerProblem
    r::Float64
    R::Float64
    bet::Float64
    b::Float64
    Pi::Matrix{Float64}
    z_vals::Vector{Float64}
    asset_grid::LinSpace{Float64}
end

function ConsumerProblem(;r=0.01, bet=0.96, Pi=[0.6 0.4; 0.05 0.95],
                      z_vals=[0.5, 1.0], b=0.0, grid_max=16, grid_size=50)
    R = 1 + r
```

```

asset_grid = linspace(-b, grid_max, grid_size)

ConsumerProblem(r, R, bet, b, Pi, z_vals, asset_grid)
end

"""
Given a matrix of size `length(cp.asset_grid), length(cp.z_vals)` , construct
an interpolation object that does linear interpolation in the asset dimension
and has a lookup table in the z dimension
"""
function Interpolations.interpolate(cp::ConsumerProblem, x::AbstractMatrix)
    sz = (length(cp.asset_grid), length(cp.z_vals))
    if size(x) != sz
        msg = "x must have dimensions $(sz)"
        throw(DimensionMismatch(msg))
    end

    itp = interpolate(x, (BSpline(Linear()), NoInterp()), OnGrid())
    scale(itp, cp.asset_grid, 1:sz[2])
end

"""
Apply the Bellman operator for a given model and initial value.

##### Arguments

- `cp::ConsumerProblem` : Instance of `ConsumerProblem`
- `v::Matrix` : Current guess for the value function
- `out::Matrix` : Storage for output
- `;ret_policy::Bool(false)` : Toggles return of value or policy functions

##### Returns

None, `out` is updated in place. If `ret_policy == true` `out` is filled with the
policy function, otherwise the value function is stored in `out`.

"""

function bellman_operator!(cp::ConsumerProblem, V::Matrix, out::Matrix;
                           ret_policy::Bool=false)
    # simplify names, set up arrays
    R, Pi, bet, b = cp.R, cp.Pi, cp.bet, cp.b
    asset_grid, z_vals = cp.asset_grid, cp.z_vals

    z_idx = 1:length(z_vals)
    vf = interpolate(cp, V)

    # compute lower bound for optimization
    opt_lb = 1e-8

    # solve for RHS of Bellman equation
    for (i_z, z) in enumerate(z_vals)
        for (i_a, a) in enumerate(asset_grid)

```

```

        function obj(c)
            y = 0.0
            for j in z_idx
                y += vf[R*a+z-c, j] * Pi[i_z, j]
            end
            return -u(c) - bet * y
        end

        res = optimize(obj, opt_lb, R.*a.+z.+b)
        c_star = res.minimum

        if ret_policy
            out[i_a, i_z] = c_star
        else
            out[i_a, i_z] = - obj(c_star)
        end

    end
    out
end

bellman_operator(cp::ConsumerProblem, V::Matrix; ret_policy=false) =
    bellman_operator!(cp, V, similar(V); ret_policy=ret_policy)

"""
Extract the greedy policy (policy function) of the model.

##### Arguments

- `cp::CareerWorkerProblem` : Instance of `CareerWorkerProblem`
- `v::Matrix` : Current guess for the value function
- `out::Matrix` : Storage for output

##### Returns

None, `out` is updated in place to hold the policy function

"""
get_greedy!(cp::ConsumerProblem, V::Matrix, out::Matrix) =
    bellman_operator!(cp, V, out, ret_policy=true)

get_greedy(cp::ConsumerProblem, V::Matrix) =
    bellman_operator(cp, V, ret_policy=true)

"""
The approximate Coleman operator.

Iteration with this operator corresponds to policy function
iteration. Computes and returns the updated consumption policy
c. The array c is replaced with a function cf that implements
univariate linear interpolation over the asset grid for each
possible value of z.

```

```

##### Arguments

- `cp::CareerWorkerProblem` : Instance of `CareerWorkerProblem`
- `c::Matrix` : Current guess for the policy function
- `out::Matrix` : Storage for output

##### Returns

None, `out` is updated in place to hold the policy function

"""

function coleman_operator!(cp::ConsumerProblem, c::Matrix, out::Matrix)
    # simplify names, set up arrays
    R, Pi, bet, b = cp.R, cp.Pi, cp.bet, cp.b
    asset_grid, z_vals = cp.asset_grid, cp.z_vals
    z_size = length(z_vals)
    gam = R * bet
    vals = Array(Float64, z_size)

    cf = interpolate(cp, c)

    # linear interpolation to get consumption function. Updates vals inplace
    cf!(a, vals) = map!(i->cf[a, i], vals, 1:z_size)

    # compute lower_bound for optimization
    opt_lb = 1e-8

    for (i_z, z) in enumerate(z_vals)
        for (i_a, a) in enumerate(asset_grid)
            function h(t)
                cf!(R*a+z-t, vals) # update vals
                expectation = dot(du(vals), vec(Pi[i_z, :]))
                return abs(du(t) - max(gam * expectation, du(R*a+z+b)))
            end
            opt_ub = R*a + z + b # addresses issue #8 on github
            res = optimize(h, min(opt_lb, opt_ub - 1e-2), opt_ub,
                           method=Optim.Brent())
            out[i_a, i_z] = res.minimum
        end
    end
    out
end

"""

Apply the Coleman operator for a given model and initial value

See the specific methods of the mutating version of this function for more
details on arguments
"""

coleman_operator(cp::ConsumerProblem, c::Matrix) =
    coleman_operator!(cp, c, similar(c))

```

```

function init_values(cp::ConsumerProblem)
    # simplify names, set up arrays
    R, bet, b = cp.R, cp.bet, cp.b
    asset_grid, z_vals = cp.asset_grid, cp.z_vals
    shape = length(asset_grid), length(z_vals)
    V, c = Array(Float64, shape...), Array(Float64, shape...)

    # Populate V and c
    for (i_z, z) in enumerate(z_vals)
        for (i_a, a) in enumerate(asset_grid)
            c_max = R*a + z + b
            c[i_a, i_z] = c_max
            V[i_a, i_z] = u(c_max) ./ (1 - bet)
        end
    end

    return V, c
end

```

The code contains a type called `ConsumerProblem` that

- stores all the relevant parameters of a given model
- defines methods
 - `bellman_operator`, which implements the Bellman operator T specified above
 - `coleman_operator`, which implements the Coleman operator K specified above
 - `initialize`, which generates suitable initial conditions for iteration

The methods `bellman_operator` and `coleman_operator` both use linear interpolation along the asset grid to approximate the value and consumption functions

The following exercises walk you through several applications where policy functions are computed

In exercise 1 you will see that while VFI and PFI produce similar results, the latter is much faster

- Because we are exploiting analytically derived first order conditions

Another benefit of working in policy function space rather than value function space is that value functions typically have more curvature

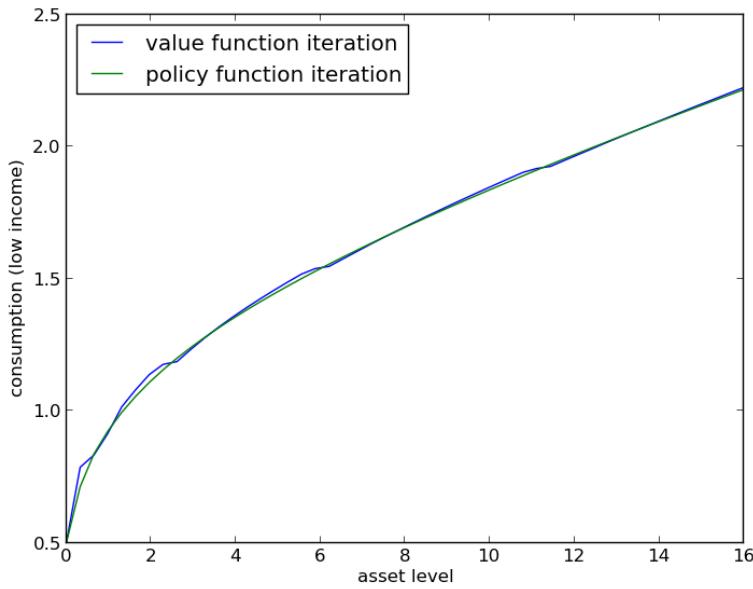
- Makes them harder to approximate numerically

Exercises

Exercise 1 The first exercise is to replicate the following figure, which compares PFI and VFI as solution methods

The figure shows consumption policies computed by iteration of K and T respectively

- In the case of iteration with T , the final value function is used to compute the observed policy



Consumption is shown as a function of assets with income z held fixed at its smallest value

The following details are needed to replicate the figure

- The parameters are the default parameters in the definition of `consumerProblem`
- The initial conditions are the default ones from `initialize(cp)`
- Both operators are iterated 80 times

When you run your code you will observe that iteration with K is faster than iteration with T

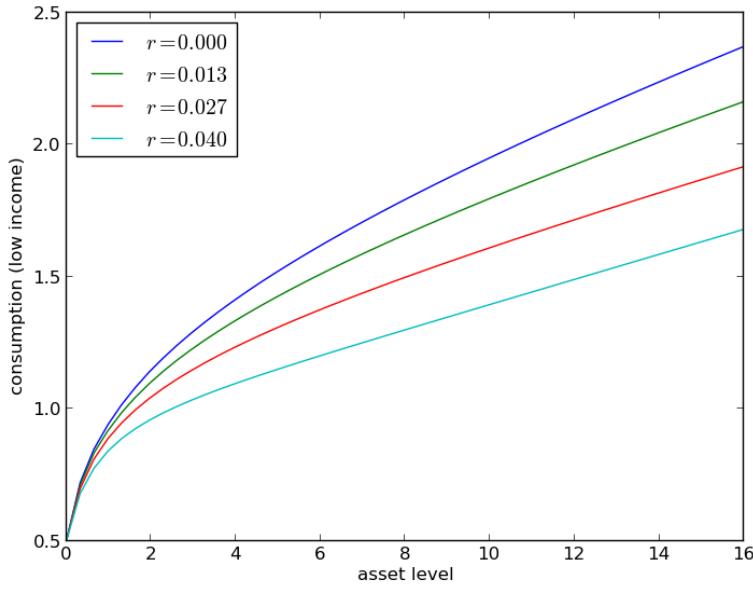
In the Julia console, a comparison of the operators can be made as follows

```
julia> using QuantEcon
julia> cp = ConsumerProblem();
julia> v, c, = initialize(cp);
julia> @time bellman_operator(cp, v);
elapsed time: 0.095017748 seconds (24212168 bytes allocated, 30.48% gc time)
julia> @time coleman_operator(cp, c);
elapsed time: 0.0696242 seconds (23937576 bytes allocated)
```

Exercise 2 Next let's consider how the interest rate affects consumption

Reproduce the following figure, which shows (approximately) optimal consumption policies for different interest rates

- Other than r , all parameters are at their default values



- r steps through $\text{linspace}(0, 0.04, 4)$
- Consumption is plotted against assets for income shock fixed at the smallest value

The figure shows that higher interest rates boost savings and hence suppress consumption

Exercise 3 Now let's consider the long run asset levels held by households

We'll take $r = 0.03$ and otherwise use default parameters

The following figure is a 45 degree diagram showing the law of motion for assets when consumption is optimal

The green line and blue line represent the function

$$a' = h(a, z) := Ra + z - c^*(a, z)$$

when income z takes its high and low values respectively

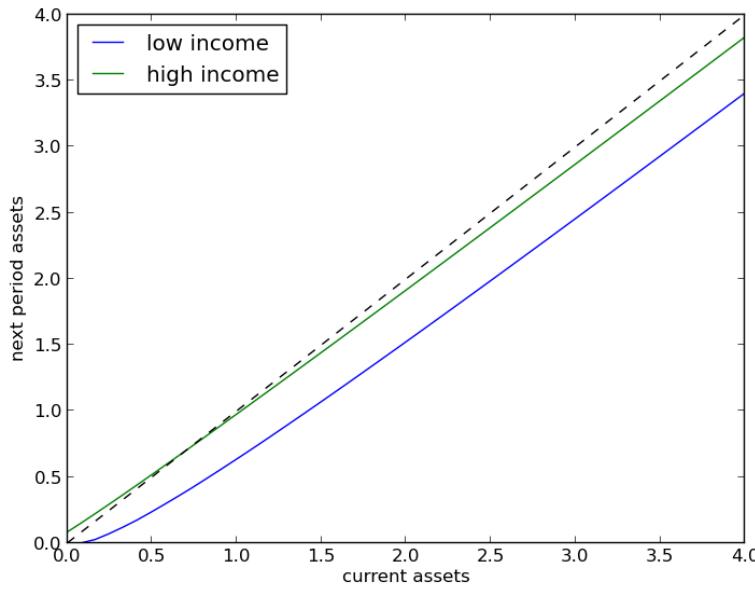
The dashed line is the 45 degree line

We can see from the figure that the dynamics will be stable — assets do not diverge

In fact there is a unique stationary distribution of assets that we can calculate by simulation

- Can be proved via theorem 2 of [HP92]
- Represents the long run dispersion of assets across households when households have idiosyncratic shocks

Ergodicity is valid here, so stationary probabilities can be calculated by averaging over a single long time series



- Hence to approximate the stationary distribution we can simulate a long time series for assets and histogram, as in the following figure

Your task is to replicate the figure

- Parameters are as discussed above
- The histogram in the figure used a single time series $\{a_t\}$ of length 500,000
- Given the length of this time series, the initial condition (a_0, z_0) will not matter
- You might find it helpful to use the `MarkovChain` class from `quantecon`

Exercise 4 Following on from exercises 2 and 3, let's look at how savings and aggregate asset holdings vary with the interest rate

- Note: [LS12] section 18.6 can be consulted for more background on the topic treated in this exercise

For a given parameterization of the model, the mean of the stationary distribution can be interpreted as aggregate capital in an economy with a unit mass of *ex-ante* identical households facing idiosyncratic shocks

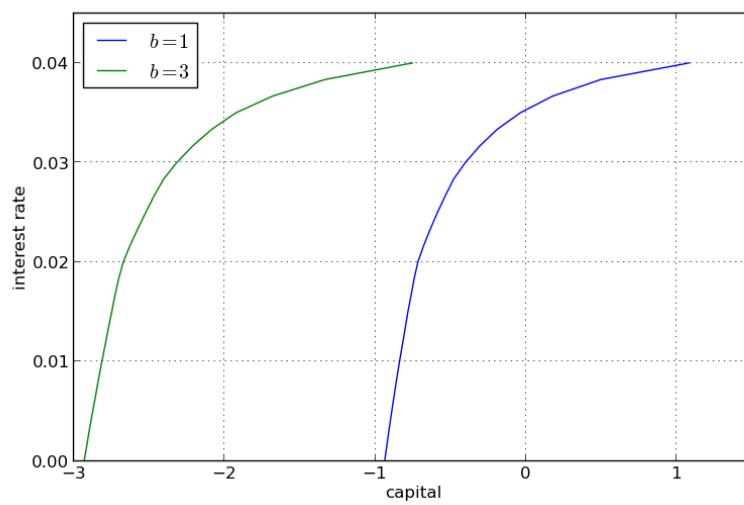
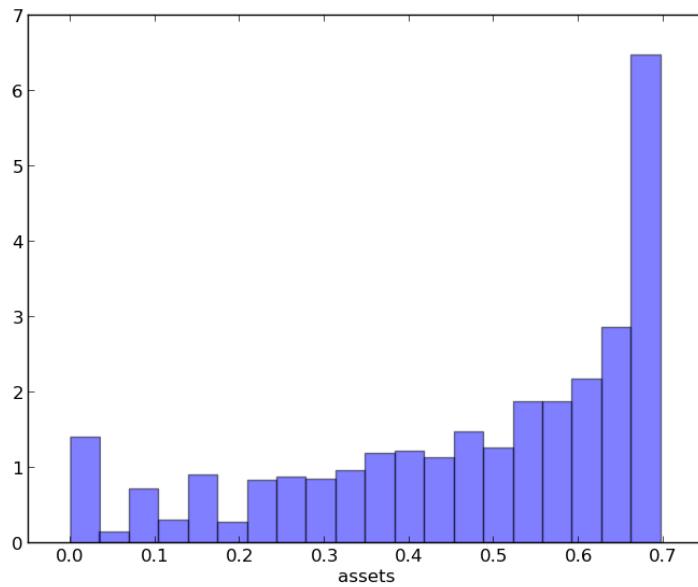
Let's look at how this measure of aggregate capital varies with the interest rate and borrowing constraint

The next figure plots aggregate capital against the interest rate for b in $(1, 3)$

As is traditional, the price (interest rate) is on the vertical axis

The horizontal axis is aggregate capital computed as the mean of the stationary distribution

Exercise 4 is to replicate the figure, making use of code from previous exercises



Try to explain why the measure of aggregate capital is equal to $-b$ when $r = 0$ for both cases shown here

Solutions

[Solution notebook](#)

Robustness

Contents

- *Robustness*
 - *Overview*
 - *The Model*
 - *Constructing More Robust Policies*
 - *Robustness as Outcome of a Two-Person Zero-Sum Game*
 - *The Stochastic Case*
 - *Implementation*
 - *Application*
 - *Appendix*

Overview

This lecture modifies a Bellman equation to express a decision maker's doubts about transition dynamics

His specification doubts make the decision maker want a *robust* decision rule

Robust means insensitive to misspecification of transition dynamics

The decision maker has a single *approximating model*

He calls it *approximating* to acknowledge that he doesn't completely trust it

He fears that outcomes will actually be determined by another model that he cannot describe explicitly

All that he knows is that the actual data-generating model is in some (uncountable) set of models that surrounds his approximating model

He quantifies the discrepancy between his approximating model and the genuine data-generating model by using a quantity called *entropy*

(We'll explain what entropy means below)

He wants a decision rule that will work well enough no matter which of those other models actually governs outcomes

This is what it means for his decision rule to be “robust to misspecification of an approximating model”

This may sound like too much to ask for, but . . .

. . . a *secret weapon* is available to design robust decision rules

The secret weapon is max-min control theory

A value-maximizing decision maker enlists the aid of an (imaginary) value-minimizing model chooser to construct *bounds* on the value attained by a given decision rule under different models of the transition dynamics

The original decision maker uses those bounds to construct a decision rule with an assured performance level, no matter which model actually governs outcomes

Note: In reading this lecture, please don’t think that our decision maker is paranoid when he conducts a worst-case analysis. By designing a rule that works well against a worst-case, his intention is to construct a rule that will work well across a *set* of models.

Sets of Models Imply Sets Of Values Our “robust” decision maker wants to know how well a given rule will work when he does not *know* a single transition law . . .

. . . he wants to know *sets* of values that will be attained by a given decision rule F under a *set* of transition laws

Ultimately, he wants to design a decision rule F that shapes these *sets* of values in ways that he prefers

With this in mind, consider the following graph, which relates to a particular decision problem to be explained below

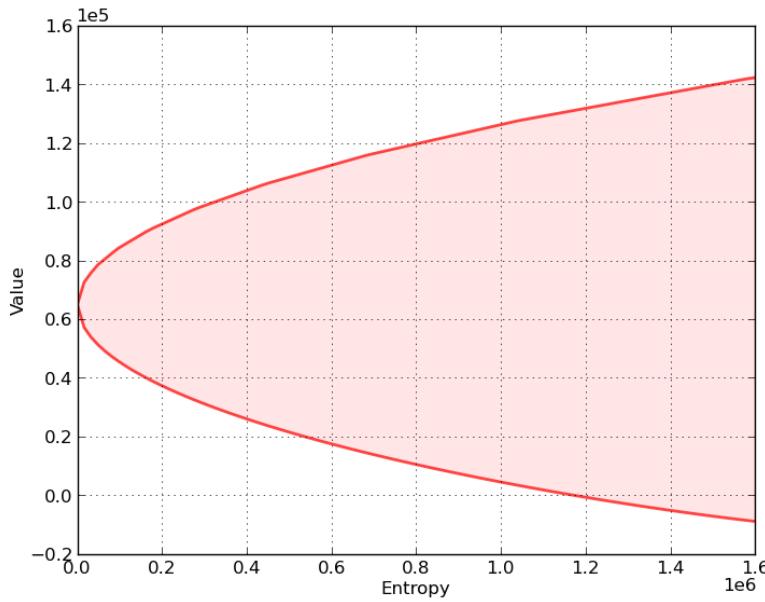
The figure shows a *value-entropy correspondence* for a particular decision rule F

The shaded set is the graph of the correspondence, which maps entropy to a set of values associated with a set of models that surround the decision maker’s approximating model

Here

- *Value* refers to a sum of discounted rewards obtained by applying the decision rule F when the state starts at some fixed initial state x_0
- *Entropy* is a nonnegative number that measures the size of a set of models surrounding the decision maker’s approximating model
 - Entropy is zero when the set includes only the approximating model, indicating that the decision maker completely trusts the approximating model
 - Entropy is bigger, and the set of surrounding models is bigger, the less the decision maker trusts the approximating model

The shaded region indicates that for **all** models having entropy less than or equal to the number on the horizontal axis, the value obtained will be somewhere within the indicated set of values



Now let's compare sets of values associated with two different decision rules, F_r and F_b

In the next figure,

- The red set shows the value-entropy correspondence for decision rule F_r
- The blue set shows the value-entropy correspondence for decision rule F_b

The blue correspondence is skinnier than the red correspondence

This conveys the sense in which the decision rule F_b is *more robust* than the decision rule F_r

- *more robust* means that the set of values is less sensitive to *increasing misspecification* as measured by entropy

Notice that the less robust rule F_r promises higher values for small misspecifications (small entropy)

(But it is more fragile in the sense that it is more sensitive to perturbations of the approximating model)

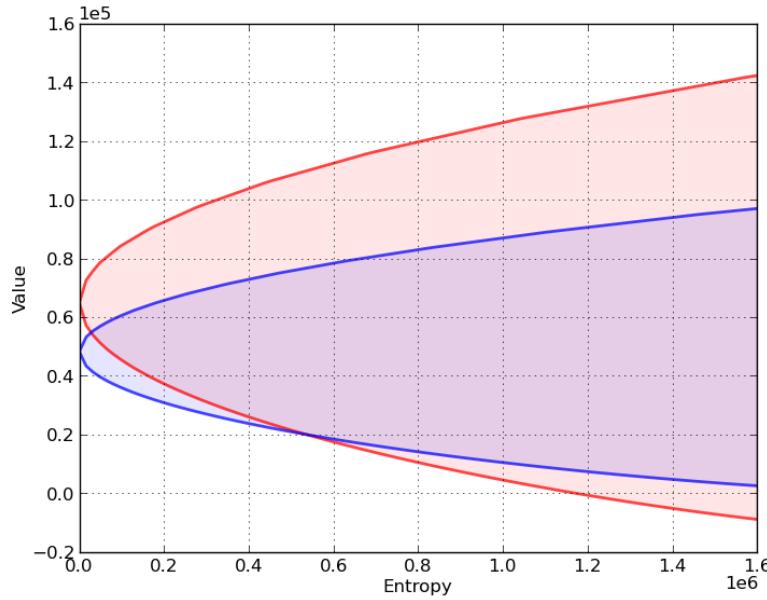
Below we'll explain in detail how to construct these sets of values for a given F , but for now ...

Here is a hint about the *secret weapons* we'll use to construct these sets

- We'll use some min problems to construct the lower bounds
- We'll use some max problems to construct the upper bounds

We will also describe how to choose F to shape the sets of values

This will involve crafting a *skinnier* set at the cost of a lower *level* (at least for low values of entropy)



Inspiring Video If you want to understand more about why one serious quantitative researcher is interested in this approach, we recommend Lars Peter Hansen's Nobel lecture

Other References Our discussion in this lecture is based on

- [HS00]
- [HS08]

The Model

For simplicity, we present ideas in the context of a class of problems with linear transition laws and quadratic objective functions

To fit in with our earlier lecture on LQ control, we will treat loss minimization rather than value maximization

To begin, recall the *infinite horizon LQ problem*, where an agent chooses a sequence of controls $\{u_t\}$ to minimize

$$\sum_{t=0}^{\infty} \beta^t \{ x_t' R x_t + u_t' Q u_t \} \quad (3.56)$$

subject to the linear law of motion

$$x_{t+1} = Ax_t + Bu_t + Cw_{t+1}, \quad t = 0, 1, 2, \dots \quad (3.57)$$

As before,

- x_t is $n \times 1$, A is $n \times n$
- u_t is $k \times 1$, B is $n \times k$

- w_t is $j \times 1$, C is $n \times j$
- R is $n \times n$ and Q is $k \times k$

Here x_t is the state, u_t is the control, and w_t is a shock vector.

For now we take $\{w_t\} := \{w_t\}_{t=1}^\infty$ to be deterministic — a single fixed sequence

We also allow for *model uncertainty* on the part of the agent solving this optimization problem

In particular, the agent takes $w_t = 0$ for all $t \geq 0$ as a benchmark model, but admits the possibility that this model might be wrong

As a consequence, she also considers a set of alternative models expressed in terms of sequences $\{w_t\}$ that are “close” to the zero sequence

She seeks a policy that will do well enough for a set of alternative models whose members are pinned down by sequences $\{w_t\}$

Soon we'll quantify the quality of a model specification in terms of the maximal size of the expression $\sum_{t=0}^{\infty} \beta^{t+1} w'_{t+1} w_{t+1}$

Constructing More Robust Policies

If our agent takes $\{w_t\}$ as a given deterministic sequence, then, drawing on intuition from earlier lectures on dynamic programming, we can anticipate Bellman equations such as

$$J_{t-1}(x) = \min_u \{x'Rx + u'Qu + \beta J_t(Ax + Bu + Cw_t)\}$$

(Here J depends on t because the sequence $\{w_t\}$ is not recursive)

Our tool for studying robustness is to construct a rule that works well even if an adverse sequence $\{w_t\}$ occurs

In our framework, “adverse” means “loss increasing”

As we'll see, this will eventually lead us to construct the Bellman equation

$$J(x) = \min_u \max_w \{x'Rx + u'Qu + \beta [J(Ax + Bu + Cw) - \theta w'w]\} \quad (3.58)$$

Notice that we've added the penalty term $-\theta w'w$

Since $w'w = \|w\|^2$, this term becomes influential when w moves away from the origin

The penalty parameter θ controls how much we penalize the maximizing agent for “harming” the minimizing agent

By raising θ more and more, we more and more limit the ability of maximizing agent to distort outcomes relative to the approximating model

So bigger θ is implicitly associated with smaller distortion sequences $\{w_t\}$

Analyzing the Bellman equation So what does J in (3.58) look like?

As with the ordinary LQ control model, J takes the form $J(x) = x'Px$ for some symmetric positive definite matrix P

One of our main tasks will be to analyze and compute the matrix P

Related tasks will be to study associated feedback rules for u_t and w_{t+1}

First, using *matrix calculus*, you will be able to verify that

$$\begin{aligned} \max_w \{ (Ax + Bu + Cw)'P(Ax + Bu + Cw) - \theta w'w \} \\ = (Ax + Bu)'D(P)(Ax + Bu) \end{aligned} \quad (3.59)$$

where

$$D(P) := P + PC(\theta I - C'PC)^{-1}C'P \quad (3.60)$$

and I is a $j \times j$ identity matrix. Substituting this expression for the maximum into (3.58) yields

$$x'Px = \min_u \{ x'Rx + u'Qu + \beta (Ax + Bu)'D(P)(Ax + Bu) \} \quad (3.61)$$

Using similar mathematics, the solution to this minimization problem is $u = -Fx$ where $F := (Q + \beta B'D(P)B)^{-1}\beta B'D(P)A$

Substituting this minimizer back into (3.61) and working through the algebra gives $x'Px = x'B(D(P))x$ for all x , or, equivalently,

$$P = B(D(P))$$

where D is the operator defined in (3.60) and

$$B(P) := R - \beta^2 A'PB(Q + \beta B'PB)^{-1}B'PA + \beta A'PA$$

The operator B is the standard (i.e., non-robust) LQ Bellman operator, and $P = B(P)$ is the standard matrix Riccati equation coming from the Bellman equation — see *this discussion*

Under some regularity conditions (see [HS08]), the operator $B \circ D$ has a unique positive definite fixed point, which we denote below by \hat{P}

A robust policy, indexed by θ , is $u = -\hat{F}x$ where

$$\hat{F} := (Q + \beta B'D(\hat{P})B)^{-1}\beta B'D(\hat{P})A \quad (3.62)$$

We also define

$$\hat{K} := (\theta I - C'\hat{P}C)^{-1}C'\hat{P}(A - B\hat{F}) \quad (3.63)$$

The interpretation of \hat{K} is that $w_{t+1} = \hat{K}x_t$ on the worst-case path of $\{x_t\}$, in the sense that this vector is the maximizer of (3.59) evaluated at the fixed rule $u = -\hat{F}x$

Note that $\hat{P}, \hat{F}, \hat{K}$ are all determined by the primitives and θ

Note also that if θ is very large, then D is approximately equal to the identity mapping

Hence, when θ is large, \hat{P} and \hat{F} are approximately equal to their standard LQ values

Furthermore, when θ is large, \hat{K} is approximately equal to zero

Conversely, smaller θ is associated with greater fear of model misspecification, and greater concern for robustness

Robustness as Outcome of a Two-Person Zero-Sum Game

What we have done above can be interpreted in terms of a two-person zero-sum game in which \hat{F}, \hat{K} are Nash equilibrium objects

Agent 1 is our original agent, who seeks to minimize loss in the LQ program while admitting the possibility of misspecification

Agent 2 is an imaginary malevolent player

Agent 2's malevolence helps the original agent to compute bounds on his value function across a set of models

We begin with agent 2's problem

Agent 2's Problem Agent 2

1. knows a fixed policy F specifying the behavior of agent 1, in the sense that $u_t = -Fx_t$ for all t
2. responds by choosing a shock sequence $\{w_t\}$ from a set of paths sufficiently close to the benchmark sequence $\{0, 0, 0, \dots\}$

A natural way to say "sufficiently close to the zero sequence" is to restrict the summed inner product $\sum_{t=1}^{\infty} w_t' w_t$ to be small

However, to obtain a time-invariant recursive formulation, it turns out to be convenient to restrict a discounted inner product

$$\sum_{t=1}^{\infty} \beta^t w_t' w_t \leq \eta \quad (3.64)$$

Now let F be a fixed policy, and let $J_F(x_0, \mathbf{w})$ be the present-value cost of that policy given sequence $\mathbf{w} := \{w_t\}$ and initial condition $x_0 \in \mathbb{R}^n$

Substituting $-Fx_t$ for u_t in (3.56), this value can be written as

$$J_F(x_0, \mathbf{w}) := \sum_{t=0}^{\infty} \beta^t x_t' (R + F' Q F) x_t \quad (3.65)$$

where

$$x_{t+1} = (A - BF)x_t + Cw_{t+1} \quad (3.66)$$

and the initial condition x_0 is as specified in the left side of (3.65)

Agent 2 chooses \mathbf{w} to maximize agent 1's loss $J_F(x_0, \mathbf{w})$ subject to (3.64)

Using a Lagrangian formulation, we can express this problem as

$$\max_{\mathbf{w}} \sum_{t=0}^{\infty} \beta^t \{ x_t' (R + F' Q F) x_t - \beta \theta (w_{t+1}' w_{t+1} - \eta) \}$$

where $\{x_t\}$ satisfied (3.66) and θ is a Lagrange multiplier on constraint (3.64)

For the moment, let's take θ as fixed, allowing us to drop the constant $\beta\theta\eta$ term in the objective function, and hence write the problem as

$$\max_{\mathbf{w}} \sum_{t=0}^{\infty} \beta^t \{ x_t' (R + F' QF) x_t - \beta\theta w_{t+1}' w_{t+1} \}$$

or, equivalently,

$$\min_{\mathbf{w}} \sum_{t=0}^{\infty} \beta^t \{ -x_t' (R + F' QF) x_t + \beta\theta w_{t+1}' w_{t+1} \} \quad (3.67)$$

subject to (3.66)

What's striking about this optimization problem is that it is once again an LQ discounted dynamic programming problem, with $\mathbf{w} = \{w_t\}$ as the sequence of controls

The expression for the optimal policy can be found by applying the usual LQ formula (*see here*)

We denote it by $K(F, \theta)$, with the interpretation $w_{t+1} = K(F, \theta)x_t$

The remaining step for agent 2's problem is to set θ to enforce the constraint (3.64), which can be done by choosing $\theta = \theta_\eta$ such that

$$\beta \sum_{t=0}^{\infty} \beta^t x_t' K(F, \theta_\eta)' K(F, \theta_\eta) x_t = \eta \quad (3.68)$$

Here x_t is given by (3.66) — which in this case becomes $x_{t+1} = (A - BF + CK(F, \theta))x_t$

Using Agent 2's Problem to Construct Bounds on the Value Sets

The Lower Bound Define the minimized object on the right side of problem (3.67) as $R_\theta(x_0, F)$.

Because “minimizers minimize” we have

$$R_\theta(x_0, F) \leq \sum_{t=0}^{\infty} \beta^t \{ -x_t' (R + F' QF) x_t \} + \beta\theta \sum_{t=0}^{\infty} \beta^t w_{t+1}' w_{t+1},$$

where $x_{t+1} = (A - BF + CK(F, \theta))x_t$ and x_0 is a given initial condition.

This inequality in turn implies the inequality

$$R_\theta(x_0, F) - \theta \text{ ent} \leq \sum_{t=0}^{\infty} \beta^t \{ -x_t' (R + F' QF) x_t \} \quad (3.69)$$

where

$$\text{ent} := \beta \sum_{t=0}^{\infty} \beta^t w_{t+1}' w_{t+1}$$

The left side of inequality (3.69) is a straight line with slope $-\theta$

Technically, it is a “separating hyperplane”

At a particular value of entropy, the line is tangent to the lower bound of values as a function of entropy

In particular, the lower bound on the left side of (3.69) is attained when

$$\text{ent} = \beta \sum_{t=0}^{\infty} \beta^t x_t' K(F, \theta)' K(F, \theta) x_t \quad (3.70)$$

To construct the *lower bound* on the set of values associated with all perturbations \mathbf{w} satisfying the entropy constraint (3.64) at a given entropy level, we proceed as follows:

- For a given θ , solve the minimization problem (3.67)
- Compute the minimizer $R_\theta(x_0, F)$ and the associated entropy using (3.70)
- Compute the lower bound on the value function $R_\theta(x_0, F) - \theta \text{ent}$ and plot it against ent
- Repeat the preceding three steps for a range of values of θ to trace out the lower bound

Note: This procedure sweeps out a set of separating hyperplanes indexed by different values for the Lagrange multiplier θ

The Upper Bound To construct an *upper bound* we use a very similar procedure

We simply replace the *minimization* problem (3.67) with the *maximization* problem

$$V_{\tilde{\theta}}(x_0, F) = \max_{\mathbf{w}} \sum_{t=0}^{\infty} \beta^t \left\{ -x_t'(R + F'QF)x_t - \beta\tilde{\theta}w_{t+1}'w_{t+1} \right\} \quad (3.71)$$

where now $\tilde{\theta} > 0$ penalizes the choice of \mathbf{w} with larger entropy.

(Notice that $\tilde{\theta} = -\theta$ in problem (3.67))

Because “maximizers maximize” we have

$$V_{\tilde{\theta}}(x_0, F) \geq \sum_{t=0}^{\infty} \beta^t \left\{ -x_t'(R + F'QF)x_t \right\} - \beta\tilde{\theta} \sum_{t=0}^{\infty} \beta^t w_{t+1}' w_{t+1}$$

which in turn implies the inequality

$$V_{\tilde{\theta}}(x_0, F) + \tilde{\theta} \text{ent} \geq \sum_{t=0}^{\infty} \beta^t \left\{ -x_t'(R + F'QF)x_t \right\} \quad (3.72)$$

where

$$\text{ent} \equiv \beta \sum_{t=0}^{\infty} \beta^t w_{t+1}' w_{t+1}$$

The left side of inequality (3.72) is a straight line with slope $\tilde{\theta}$

The upper bound on the left side of (3.72) is attained when

$$\text{ent} = \beta \sum_{t=0}^{\infty} \beta^t x_t' K(F, \tilde{\theta})' K(F, \tilde{\theta}) x_t \quad (3.73)$$

To construct the *upper bound* on the set of values associated all perturbations \mathbf{w} with a given entropy we proceed much as we did for the lower bound

- For a given $\tilde{\theta}$, solve the maximization problem (3.71)
- Compute the maximizer $V_{\tilde{\theta}}(x_0, F)$ and the associated entropy using (3.73)
- Compute the upper bound on the value function $V_{\tilde{\theta}}(x_0, F) + \tilde{\theta} \text{ent}$ and plot it against ent
- Repeat the preceding three steps for a range of values of $\tilde{\theta}$ to trace out the upper bound

Reshaping the set of values Now in the interest of *reshaping* these sets of values by choosing F , we turn to agent 1's problem

Agent 1's Problem Now we turn to agent 1, who solves

$$\min_{\{u_t\}} \sum_{t=0}^{\infty} \beta^t \{ x_t' Rx_t + u_t' Qu_t - \beta \theta w_{t+1}' w_{t+1} \} \quad (3.74)$$

where $\{w_{t+1}\}$ satisfies $w_{t+1} = Kx_t$

In other words, agent 1 minimizes

$$\sum_{t=0}^{\infty} \beta^t \{ x_t' (R - \beta \theta K' K) x_t + u_t' Qu_t \} \quad (3.75)$$

subject to

$$x_{t+1} = (A + CK)x_t + Bu_t \quad (3.76)$$

Once again, the expression for the optimal policy can be found *here* — we denote it by \tilde{F}

Nash Equilibrium Clearly the \tilde{F} we have obtained depends on K , which, in agent 2's problem, depended on an initial policy F

Holding all other parameters fixed, we can represent this relationship as a mapping Φ , where

$$\tilde{F} = \Phi(K(F, \theta))$$

The map $F \mapsto \Phi(K(F, \theta))$ corresponds to a situation in which

1. agent 1 uses an arbitrary initial policy F
2. agent 2 best responds to agent 1 by choosing $K(F, \theta)$
3. agent 1 best responds to agent 2 by choosing $\tilde{F} = \Phi(K(F, \theta))$

As you may have already guessed, the robust policy \hat{F} defined in (3.62) is a fixed point of the mapping Φ

In particular, for any given θ ,

1. $K(\hat{F}, \theta) = \hat{K}$, where \hat{K} is as given in (3.63)
2. $\Phi(\hat{K}) = \hat{F}$

A sketch of the proof is given in *the appendix*

The Stochastic Case

Now we turn to the stochastic case, where the sequence $\{w_t\}$ is treated as an iid sequence of random vectors

In this setting, we suppose that our agent is uncertain about the *conditional probability distribution* of w_{t+1}

The agent takes the standard normal distribution $N(0, I)$ as the baseline conditional distribution, while admitting the possibility that other “nearby” distributions prevail

These alternative conditional distributions of w_{t+1} might depend nonlinearly on the history $x_s, s \leq t$

To implement this idea, we need a notion of what it means for one distribution to be near another one

Here we adopt a very useful measure of closeness for distributions known as the *relative entropy*, or **Kullback-Leibler divergence**

For densities p, q , the Kullback-Leibler divergence of q from p is defined as

$$D_{KL}(p, q) := \int \ln \left[\frac{p(x)}{q(x)} \right] p(x) dx$$

Using this notation, we replace (3.58) with the stochastic analogue

$$J(x) = \min_u \max_{\psi \in \mathcal{P}} \left\{ x' Rx + u' Qu + \beta \left[\int J(Ax + Bu + Cw) \psi(dw) - \theta D_{KL}(\psi, \phi) \right] \right\} \quad (3.77)$$

Here \mathcal{P} represents the set of all densities on \mathbb{R}^n and ϕ is the benchmark distribution $N(0, I)$

The distribution ϕ is chosen as the least desirable conditional distribution in terms of next period outcomes, while taking into account the penalty term $\theta D_{KL}(\psi, \phi)$

This penalty term plays a role analogous to the one played by the deterministic penalty $\theta w' w$ in (3.58), since it discourages large deviations from the benchmark

Solving the Model The maximization problem in (3.77) appears highly nontrivial — after all, we are maximizing over an infinite dimensional space consisting of the entire set of densities

However, it turns out that the solution is tractable, and in fact also falls within the class of normal distributions

First, we note that J has the form $J(x) = x' Px + d$ for some positive definite matrix P and constant real number d

Moreover, it turns out that if $(I - \theta^{-1} C' P C)^{-1}$ is nonsingular, then

$$\begin{aligned} \max_{\psi \in \mathcal{P}} & \left\{ \int (Ax + Bu + Cw)' P (Ax + Bu + Cw) \psi(dw) - \theta D_{KL}(\psi, \phi) \right\} \\ &= (Ax + Bu)' \mathcal{D}(P)(Ax + Bu) + \kappa(\theta, P) \end{aligned} \quad (3.78)$$

where

$$\kappa(\theta, P) := \theta \ln[\det(I - \theta^{-1}C'PC)^{-1}]$$

and the maximizer is the Gaussian distribution

$$\psi = N\left((\theta I - C'PC)^{-1}C'P(Ax + Bu), (I - \theta^{-1}C'PC)^{-1}\right) \quad (3.79)$$

Substituting the expression for the maximum into Bellman equation (3.77) and using $J(x) = x'Px + d$ gives

$$x'Px + d = \min_u \{x'Rx + u'Qu + \beta(Ax + Bu)'D(P)(Ax + Bu) + \beta[d + \kappa(\theta, P)]\} \quad (3.80)$$

Since constant terms do not affect minimizers, the solution is the same as (3.61), leading to

$$x'Px + d = x'\mathcal{B}(D(P))x + \beta[d + \kappa(\theta, P)]$$

To solve this Bellman equation, we take \hat{P} to be the positive definite fixed point of $\mathcal{B} \circ D$

In addition, we take \hat{d} as the real number solving $d = \beta[d + \kappa(\theta, P)]$, which is

$$\hat{d} := \frac{\beta}{1 - \beta}\kappa(\theta, P) \quad (3.81)$$

The robust policy in this stochastic case is the minimizer in (3.80), which is once again $u = -\hat{F}x$ for \hat{F} given by (3.62)

Substituting the robust policy into (3.79) we obtain the worst case shock distribution:

$$w_{t+1} \sim N(\hat{K}x_t, (I - \theta^{-1}C'\hat{P}C)^{-1})$$

where \hat{K} is given by (3.63)

Note that the mean of the worst-case shock distribution is equal to the same worst-case w_{t+1} as in the earlier deterministic setting

Computing Other Quantities Before turning to implementation, we briefly outline how to compute several other quantities of interest

Worst-Case Value of a Policy One thing we will be interested in doing is holding a policy fixed and computing the discounted loss associated with that policy

So let F be a given policy and let $J_F(x)$ be the associated loss, which, by analogy with (3.77), satisfies

$$J_F(x) = \max_{\psi \in \mathcal{P}} \left\{ x'(R + F'QF)x + \beta \left[\int J_F((A - BF)x + Cw) \psi(dw) - \theta D_{KL}(\psi, \phi) \right] \right\}$$

Writing $J_F(x) = x'P_Fx + d_F$ and applying the same argument used to derive (3.78) we get

$$x'P_Fx + d_F = x'(R + F'QF)x + \beta [x'(A - BF)'D(P_F)(A - BF)x + d_F + \kappa(\theta, P_F)]$$

To solve this we take P_F to be the fixed point

$$P_F = R + F'QF + \beta(A - BF)'D(P_F)(A - BF)$$

and

$$d_F := \frac{\beta}{1-\beta} \kappa(\theta, P_F) = \frac{\beta}{1-\beta} \theta \ln[\det(I - \theta^{-1} C' P_F C)^{-1}] \quad (3.82)$$

If you skip ahead to *the appendix*, you will be able to verify that $-P_F$ is the solution to the Bellman equation in agent 2's problem *discussed above* — we use this in our computations

Implementation

The QuantEcon.jl package provides a type called `RBLQ` for implementation of robust LQ optimal control

Here's the relevant code, from file `robustlq.jl`

```
#=
Provides a type called RBLQ for solving robust linear quadratic control
problems.

@author : Spencer Lyon <spencer.lyon@nyu.edu>

@date : 2014-08-19

References
-----

http://quant-econ.net/jl/robustness.html
=#

"""
Represents infinite horizon robust LQ control problems of the form

    min_{u_t} sum_t beta^t {x_t' R x_t + u_t' Q u_t }

subject to

    x_{t+1} = A x_t + B u_t + C w_{t+1}

and with model misspecification parameter theta.

##### Fields

- `Q::Matrix{Float64}` : The cost(payoff) matrix for the controls. See above
for more. `Q` should be k x k and symmetric and positive definite
- `R::Matrix{Float64}` : The cost(payoff) matrix for the state. See above for
more. `R` should be n x n and symmetric and non-negative definite
- `A::Matrix{Float64}` : The matrix that corresponds with the state in the
state space system. `A` should be n x n
- `B::Matrix{Float64}` : The matrix that corresponds with the control in the
state space system. `B` should be n x k
- `C::Matrix{Float64}` : The matrix that corresponds with the random process in
the state space system. `C` should be n x j
- `beta::Real` : The discount factor in the robust control problem
- `theta::Real` The robustness factor in the robust control problem
```

```

- `k, n, j::Int` : Dimensions of input matrices

"""

type RBLQ
    A::Matrix
    B::Matrix
    C::Matrix
    Q::Matrix
    R::Matrix
    k::Int
    n::Int
    j::Int
    bet::Real
    theta::Real
end

function RBLQ(Q::ScalarOrArray, R::ScalarOrArray, A::ScalarOrArray,
              B::ScalarOrArray, C::ScalarOrArray, bet::Real, theta::Real)
    k = size(Q, 1)
    n = size(R, 1)
    j = size(C, 2)

    # coerce sizes
    A = reshape([A], n, n)
    B = reshape([B], n, k)
    C = reshape([C], n, j)
    R = reshape([R], n, n)
    Q = reshape([Q], k, k)
    RBLQ(A, B, C, Q, R, k, n, j, bet, theta)
end

"""

The D operator, mapping P into

D(P) := P + PC(theta I - C'PC)^{-1} C'P.

##### Arguments

- `rlq::RBLQ` : Instance of `RBLQ` type
- `P::Matrix{Float64}` : `size` is n x n

##### Returns

- `dP::Matrix{Float64}` : The matrix P after applying the D operator

"""

function d_operator(rlq::RBLQ, P::Matrix)
    C, theta, I = rlq.C, rlq.theta, eye(rlq.j)
    S1 = P*C
    dP = P + S1*((theta.*I - C'*S1) \ (S1'))

    return dP

```

```

end

"""
The D operator, mapping P into

    B(P) := R - beta^2 A'PB(Q + beta B'PB)^{-1}B'PA + beta A'PA

and also returning

    F := (Q + beta B'PB)^{-1} beta B'PA

##### Arguments

- `rlq::RBLQ`: Instance of `RBLQ` type
- `P::Matrix{Float64}` : `size` is n x n

##### Returns

- `F::Matrix{Float64}` : The F matrix as defined above
- `new_p::Matrix{Float64}` : The matrix P after applying the B operator

"""

function b_operator(rlq::RBLQ, P::Matrix)
    A, B, Q, R, bet = rlq.A, rlq.B, rlq.Q, rlq.R, rlq.bet

    S1 = Q + bet.*B'*P*B
    S2 = bet.*B'*P*A
    S3 = bet.*A'*P*A

    F = S1 \ S2
    new_P = R - S2'*F + S3

    return F, new_P
end

"""
Solves the robust control problem.

The algorithm here tricks the problem into a stacked LQ problem, as described in
chapter 2 of Hansen- Sargent's text "Robustness." The optimal control with
observed state is

    u_t = - F x_t

And the value function is -x'Px

##### Arguments

- `rlq::RBLQ`: Instance of `RBLQ` type

##### Returns

```

```

- `F::Matrix{Float64}` : The optimal control matrix from above
- `P::Matrix{Float64}` : The positive semi-definite matrix defining the value
function
- `K::Matrix{Float64}` : the worst-case shock matrix `K`, where
`w_{t+1} = K x_t` is the worst case shock

"""

function robust_rule(rlq::RBLQ)
    A, B, C, Q, R = rlq.A, rlq.B, rlq.C, rlq.Q, rlq.R
    bet, theta, k, j = rlq.bet, rlq.theta, rlq.k, rlq.j

    # Set up LQ version
    I = eye(j)
    Z = zeros(k, j)
    Ba = [B C]
    Qa = [Q Z
          Z' -bet.*I.*theta]
    lq = LQ(Qa, R, A, Ba, bet=bet)

    # Solve and convert back to robust problem
    P, f, d = stationary_values(lq)
    F = f[1:k, :]
    K = -f[k+1:end, :]

    return F, K, P
end

"""

Solve the robust LQ problem

A simple algorithm for computing the robust policy F and the
corresponding value function P, based around straightforward
iteration with the robust Bellman operator. This function is
easier to understand but one or two orders of magnitude slower
than self.robust_rule(). For more information see the docstring
of that method.

##### Arguments

- `rlq::RBLQ`: Instance of `RBLQ` type
- `P_init::Matrix{Float64}(zeros(rlq.n, rlq.n))` : The initial guess for the
value function matrix
- `;max_iter::Int(80)` : Maximum number of iterations that are allowed
- `;tol::Real(1e-8)` The tolerance for convergence

##### Returns

- `F::Matrix{Float64}` : The optimal control matrix from above
- `P::Matrix{Float64}` : The positive semi-definite matrix defining the value
function
- `K::Matrix{Float64}` : the worst-case shock matrix `K`, where
`w_{t+1} = K x_t` is the worst case shock

```

```

"""
function robust_rule_simple(rlq::RBLQ,
                           P::Matrix=zeros(Float64, rlq.n, rlq.n);
                           max_iter=80,
                           tol=1e-8)
    # Simplify notation
    A, B, C, Q, R = rlq.A, rlq.B, rlq.C, rlq.Q, rlq.R
    bet, theta, k, j = rlq.bet, rlq.theta, rlq.k, rlq.j
    iterate, e = 0, tol + 1.0

    F = similar(P) # instantiate so available after loop

    while iterate <= max_iter && e > tol
        F, new_P = b_operator(rlq, d_operator(rlq, P))
        e = sqrt(sum((new_P - P).^2))
        iterate += 1
        copy!(P, new_P)
    end

    if iterate >= max_iter
        warn("Maximum iterations in robust_rule_simple")
    end

    I = eye(j)
    K = (theta.*I - C'*P*C)\(C'*P)*(A - B*F)

    return F, K, P
end

"""

Compute agent 2's best cost-minimizing response `K`, given `F`.

##### Arguments

- `rlq::RBLQ`: Instance of `RBLQ` type
- `F::Matrix{Float64}`: A k x n array representing agent 1's policy

##### Returns

- `K::Matrix{Float64}` : Agent's best cost minimizing response corresponding to `F`
- `P::Matrix{Float64}` : The value function corresponding to `F`

"""

function F_to_K(rlq::RBLQ, F::Matrix)
    # simplify notation
    R, Q, A, B, C = rlq.R, rlq.Q, rlq.A, rlq.B, rlq.C
    bet, theta = rlq.bet, rlq.theta

    # set up lg
    Q2 = bet * theta
    R2 = -R - F'*Q*F
    A2 = A - B*F

```

```

B2 = C
lq = LQ(Q2, R2, A2, B2, bet=bet)

neg_P, neg_K, d = stationary_values(lq)

return -neg_K, -neg_P
end

"""
Compute agent 1's best cost-minimizing response `K`, given `F`.

##### Arguments

- `rlq::RBLQ`: Instance of `RBLQ` type
- `K::Matrix{Float64}`: A k x n array representing the worst case matrix

##### Returns

- `F::Matrix{Float64}` : Agent's best cost minimizing response corresponding to `K`
- `P::Matrix{Float64}` : The value function corresponding to `K`

"""
function K_to_F(rlq::RBLQ, K::Matrix)
    R, Q, A, B, C = rlq.R, rlq.Q, rlq.A, rlq.B, rlq.C
    bet, theta = rlq.bet, rlq.theta

    A1, B1, Q1, R1 = A+C*K, B, Q, R-bet*theta.*K'*K
    lq = LQ(Q1, R1, A1, B1, bet=bet)

    P, F, d = stationary_values(lq)

    return F, P
end

"""
Given `K` and `F`, compute the value of deterministic entropy, which is sum_t beta^t x_t' K' K x_t with x_{t+1} = (A - BF + CK) x_t.

##### Arguments

- `rlq::RBLQ`: Instance of `RBLQ` type
- `F::Matrix{Float64}` The policy function, a k x n array
- `K::Matrix{Float64}` The worst case matrix, a j x n array
- `x0::Vector{Float64}` : The initial condition for state

##### Returns

- `e::Float64` The deterministic entropy

"""
function compute_deterministic_entropy(rlq::RBLQ, F, K, x0)
    B, C, bet = rlq.B, rlq.C, rlq.bet

```

```

H0 = K'*K
C0 = zeros(Float64, rlq.n, 1)
A0 = A - B*F + C*K
    return var_quadratic_sum(A0, C0, H0, bet, x0)
end

"""
Given a fixed policy `F`, with the interpretation  $u = -F x$ , this function
computes the matrix  $P_F$  and constant  $d_F$  associated with discounted cost  $J_F(x) =$ 
 $x' P_F x + d_F$ .

##### Arguments

- `rlq::RBLQ`: Instance of `RBLQ` type
- `F::Matrix{Float64}` : The policy function, a  $k \times n$  array

##### Returns

- `P_F::Matrix{Float64}` : Matrix for discounted cost
- `d_F::Float64` : Constant for discounted cost
- `K_F::Matrix{Float64}` : Worst case policy
- `O_F::Matrix{Float64}` : Matrix for discounted entropy
- `o_F::Float64` : Constant for discounted entropy

"""

function evaluate_F(rlq::RBLQ, F::Matrix)
    R, Q, A, B, C = rlq.R, rlq.Q, rlq.A, rlq.B, rlq.C
    bet, theta, j = rlq.bet, rlq.theta, rlq.j

    # Solve for policies and costs using agent 2's problem
    K_F, P_F = F_to_K(rlq, F)
    I = eye(j)
    H = inv(I - C'*P_F*C./theta)
    d_F = log(det(H))

    # compute O_F and o_F
    sig = -1.0 / theta
    A0 = sqrt(bet) .* (A - B*F + C*K_F)
    O_F = solve_discrete_lyapunov(A0', bet*K_F'*K_F)
    ho = (trace(H - 1) - d_F) / 2.0
    tr = trace(O_F*C*H*C')
    o_F = (ho + bet*tr) / (1 - bet)

    return K_F, P_F, d_F, O_F, o_F
end

```

Here is a brief description of the methods of the type

- `d_operator()` and `b_operator()` implement \mathcal{D} and \mathcal{B} respectively
- `robust_rule()` and `robust_rule_simple()` both solve for the triple $\hat{F}, \hat{K}, \hat{P}$, as described in equations (3.62) – (3.63) and the surrounding discussion
 - `robust_rule()` is more efficient

- `robust_rule_simple()` is more transparent and easier to follow
- `K_to_F()` and `F_to_K()` solve the decision problems of *agent 1* and *agent 2* respectively
- `compute_deterministic_entropy()` computes the left-hand side of (3.68)
- `evaluate_F()` computes the loss and entropy associated with a given policy — see *this discussion*

Application

Let us consider a monopolist similar to *this one*, but now facing model uncertainty

The inverse demand function is $p_t = a_0 - a_1 y_t + d_t$

where

$$d_{t+1} = \rho d_t + \sigma_d w_{t+1}, \quad \{w_t\} \stackrel{\text{iid}}{\sim} N(0, 1)$$

and all parameters are strictly positive

The period return function for the monopolist is

$$r_t = p_t y_t - \gamma \frac{(y_{t+1} - y_t)^2}{2} - c y_t$$

Its objective is to maximize expected discounted profits, or, equivalently, to minimize $\mathbb{E} \sum_{t=0}^{\infty} \beta^t (-r_t)$

To form a linear regulator problem, we take the state and control to be

$$x_t = \begin{bmatrix} 1 \\ y_t \\ d_t \end{bmatrix} \quad \text{and} \quad u_t = y_{t+1} - y_t$$

Setting $b := (a_0 - c)/2$ we define

$$R = - \begin{bmatrix} 0 & b & 0 \\ b & -a_1 & 1/2 \\ 0 & 1/2 & 0 \end{bmatrix} \quad \text{and} \quad Q = \gamma/2$$

For the transition matrices we set

$$A = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & \rho \end{bmatrix}, \quad B = \begin{bmatrix} 0 \\ 1 \\ 0 \end{bmatrix}, \quad C = \begin{bmatrix} 0 \\ 0 \\ \sigma_d \end{bmatrix}$$

Our aim is to compute the value-entropy correspondences *shown above*

The parameters are

$$a_0 = 100, a_1 = 0.5, \rho = 0.9, \sigma_d = 0.05, \beta = 0.95, c = 2, \gamma = 50.0$$

The standard normal distribution for w_t is understood as the agent's baseline, with uncertainty parameterized by θ

We compute value-entropy correspondences for two policies

1. The no concern for robustness policy F_0 , which is the ordinary LQ loss minimizer
2. A “moderate” concern for robustness policy F_b , with $\theta = 0.02$

The code for producing the graph shown above, with blue being for the robust policy, is given in `robustness/robust_monopolist.jl`

We repeat it here for convenience

```
#=
The robust control problem for a monopolist with adjustment costs. The
inverse demand curve is:

p_t = a_0 - a_1 y_t + d_t

where d_{t+1} = \rho d_t + \sigma_d w_{t+1} for w_t ~ N(0, 1) and iid.
The period return function for the monopolist is

r_t = p_t y_t - \gamma (y_{t+1} - y_t)^2 / 2 - c y_t

The objective of the firm is E_t \sum_{t=0}^{\infty} \beta^t r_t

For the linear regulator, we take the state and control to be

x_t = (1, y_t, d_t) and u_t = y_{t+1} - y_t

@author : Spencer Lyon <spencer.lyon@nyu.edu>

@date : 2014-07-05

References
-----

Simple port of the file examples/robust_monopolist.py

http://quant-econ.net/robustness.html#application

=#
using QuantEcon
using Plots
pyplot()
using Grid

# model parameters
a_0      = 100
a_1      = 0.5
rho      = 0.9
sigma_d = 0.05
bet      = 0.95
c        = 2
gam     = 50.0
theta   = 0.002
ac      = (a_0 - c) / 2.0
```

```

# Define LQ matrices
R = [0 ac 0
      ac -a_1 0.5
      0. 0.5 0]
R = -R # For minimization
Q = [gam / 2.0] '
A = [1. 0. 0.
      0. 1. 0.
      0. 0. rho]
B = [0. 1. 0.] '
C = [0. 0. sigma_d] '

## Functions

function evaluate_policy(theta, F)
    rlq = RBLQ(Q, R, A, B, C, bet, theta)
    K_F, P_F, d_F, O_F, o_F = evaluate_F(rlq, F)
    x0 = [1.0 0.0 0.0] '
    value = - x0'*P_F*x0 - d_F
    entropy = x0'*O_F*x0 + o_F
    return value[1], entropy[1] # return scalars
end

function value_and_entropy(emax, F, bw, grid_size=1000)
    if lowercase(bw) == "worst"
        thetas = 1 ./ linspace(1e-8, 1000, grid_size)
    else
        thetas = -1 ./ linspace(1e-8, 1000, grid_size)
    end

    data = Array(Float64, grid_size, 2)

    for (i, theta) in enumerate(thetas)
        data[i, :] = collect(evaluate_policy(theta, F))
        if data[i, 2] >= emax # stop at this entropy level
            data = data[1:i, :]
            break
        end
    end
    return data
end

## Main

# compute optimal rule
optimal_lq = LQ(Q, R, A, B, C, zero(B'A), bet)
Po, Fo, Do = stationary_values(optimal_lq)

# compute robust rule for our theta
baseline_robust = RBLQ(Q, R, A, B, C, bet, theta)
Fb, Kb, Pb = robust_rule(baseline_robust)

```

```

# Check the positive definiteness of worst-case covariance matrix to
# ensure that theta exceeds the breakdown point
test_matrix = eye(size(Pb, 1)) - (C' * Pb * C ./ theta)[1]
eigenvals, eigenvecs = eig(test_matrix)
@assert all(eigenvals .>= 0)

emax = 1.6e6

# compute values and entropies
optimal_best_case = value_and_entropy(emax, Fo, "best")
robust_best_case = value_and_entropy(emax, Fb, "best")
optimal_worst_case = value_and_entropy(emax, Fo, "worst")
robust_worst_case = value_and_entropy(emax, Fb, "worst")

# we reverse order of "worst_case"s so values are ascending
data_pairs = ((optimal_best_case, optimal_worst_case),
               (robust_best_case, robust_worst_case))

egrid = linspace(0, emax, 100)
egrid_data = Array{Float64}[]
for data_pair in data_pairs
    for data in data_pair
        x, y = data[:, 2], data[:, 1]
        curve(z) = InterpIrregular(x, y, BCnearest, InterpLinear)[z]
        push!(egrid_data, curve(egrid))
    end
end
plot(egrid, egrid_data, color=[:red :red :blue])
plot!(egrid, egrid_data[1], fillrange=egrid_data[2],
      fillcolor=:red, fillalpha=0.1, color=:red, legend=:none)
plot!(egrid, egrid_data[3], fillrange=egrid_data[4],
      fillcolor=:blue, fillalpha=0.1, color=:blue, legend=:none)
plot!(xlabel="Entropy", ylabel="Value")

```

Here's another such figure, with $\theta = 0.002$ instead of 0.02

Can you explain the different shape of the value-entropy correspondence for the robust policy?

Appendix

We sketch the proof only of the first claim in *this section*, which is that, for any given θ , $K(\hat{F}, \theta) = \hat{K}$, where \hat{K} is as given in (3.63)

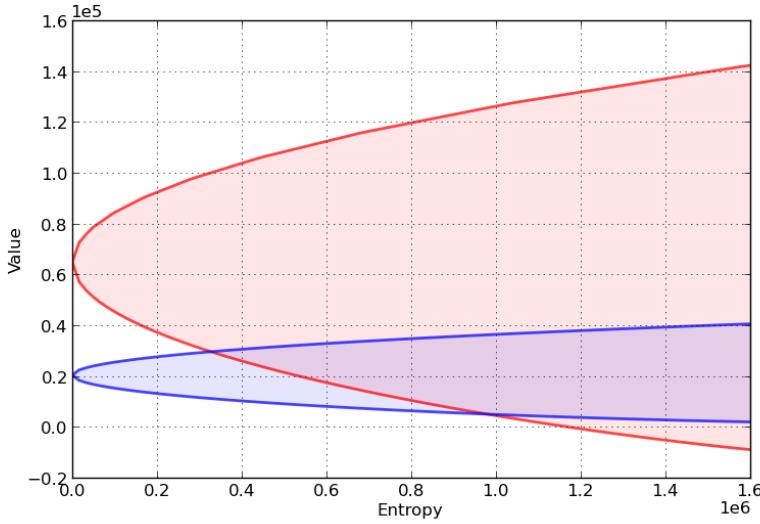
This is the content of the next lemma

Lemma. If \hat{P} is the fixed point of the map $\mathcal{B} \circ \mathcal{D}$ and \hat{F} is the robust policy as given in (3.62), then

$$K(\hat{F}, \theta) = (\theta I - C' \hat{P} C)^{-1} C' \hat{P} (A - B \hat{F}) \quad (3.83)$$

Proof: As a first step, observe that when $F = \hat{F}$, the Bellman equation associated with the LQ problem (3.66) – (3.67) is

$$\tilde{P} = -R - \hat{F}' Q \hat{F} - \beta^2 (A - B \hat{F})' \tilde{P} C (\beta \theta I + \beta C' \tilde{P} C)^{-1} C' \tilde{P} (A - B \hat{F}) + \beta (A - B \hat{F})' \tilde{P} (A - B \hat{F}) \quad (3.84)$$



(revisit *this discussion* if you don't know where (3.84) comes from) and the optimal policy is

$$w_{t+1} = -\beta(\beta\theta I + \beta C' \tilde{P} C)^{-1} C' \tilde{P} (A - B \hat{F}) x_t$$

Suppose for a moment that $-\hat{P}$ solves the Bellman equation (3.84)

In this case the policy becomes

$$w_{t+1} = (\theta I - C' \hat{P} C)^{-1} C' \hat{P} (A - B \hat{F}) x_t$$

which is exactly the claim in (3.83)

Hence it remains only to show that $-\hat{P}$ solves (3.84), or, in other words,

$$\hat{P} = R + \hat{F}' Q \hat{F} + \beta(A - B \hat{F})' \hat{P} C (\theta I - C' \hat{P} C)^{-1} C' \hat{P} (A - B \hat{F}) + \beta(A - B \hat{F})' \hat{P} (A - B \hat{F})$$

Using the definition of \mathcal{D} , we can rewrite the right-hand side more simply as

$$R + \hat{F}' Q \hat{F} + \beta(A - B \hat{F})' \mathcal{D}(\hat{P})(A - B \hat{F})$$

Although it involves a substantial amount of algebra, it can be shown that the latter is just \hat{P}

(Hint: Use the fact that $\hat{P} = \mathcal{B}(\mathcal{D}(\hat{P}))$)

Covariance Stationary Processes

Contents

- Covariance Stationary Processes
 - Overview
 - Introduction
 - Spectral Analysis
 - Implementation

Overview

In this lecture we study covariance stationary linear stochastic processes, a class of models routinely used to study economic and financial time series

This class has the advantage of being

1. simple enough to be described by an elegant and comprehensive theory
2. relatively broad in terms of the kinds of dynamics it can represent

We consider these models in both the time and frequency domain

ARMA Processes We will focus much of our attention on linear covariance stationary models with a finite number of parameters

In particular, we will study stationary ARMA processes, which form a cornerstone of the standard theory of time series analysis

It's well known that every ARMA processes can be represented in [linear state space](#) form

However, ARMA have some important structure that makes it valuable to study them separately

Spectral Analysis Analysis in the frequency domain is also called spectral analysis

In essence, spectral analysis provides an alternative representation of the autocovariance of a covariance stationary process

Having a second representation of this important object

- shines new light on the dynamics of the process in question
- allows for a simpler, more tractable representation in certain important cases

The famous *Fourier transform* and its inverse are used to map between the two representations

Other Reading For supplementary reading, see

- [\[LS12\]](#), chapter 2
- [\[Sar87\]](#), chapter 11
- John Cochrane's notes on time series analysis, chapter 8
- [\[Shi95\]](#), chapter 6
- [\[CC08\]](#), all

Introduction

Consider a sequence of random variables $\{X_t\}$ indexed by $t \in \mathbb{Z}$ and taking values in \mathbb{R}

Thus, $\{X_t\}$ begins in the infinite past and extends to the infinite future — a convenient and standard assumption

As in other fields, successful economic modeling typically requires identifying some deep structure in this process that is relatively constant over time

If such structure can be found, then each new observation X_t, X_{t+1}, \dots provides additional information about it — which is how we learn from data

For this reason, we will focus in what follows on processes that are *stationary* — or become so after some transformation (differencing, cointegration, etc.)

Definitions A real-valued stochastic process $\{X_t\}$ is called *covariance stationary* if

1. Its mean $\mu := \mathbb{E}X_t$ does not depend on t
2. For all k in \mathbb{Z} , the k -th autocovariance $\gamma(k) := \mathbb{E}(X_t - \mu)(X_{t+k} - \mu)$ is finite and depends only on k

The function $\gamma: \mathbb{Z} \rightarrow \mathbb{R}$ is called the *autocovariance function* of the process

Throughout this lecture, we will work exclusively with zero-mean (i.e., $\mu = 0$) covariance stationary processes

The zero-mean assumption costs nothing in terms of generality, since working with non-zero-mean processes involves no more than adding a constant

Example 1: White Noise Perhaps the simplest class of covariance stationary processes is the white noise processes

A process $\{\epsilon_t\}$ is called a *white noise process* if

1. $\mathbb{E}\epsilon_t = 0$
2. $\gamma(k) = \sigma^2 \mathbf{1}\{k = 0\}$ for some $\sigma > 0$

(Here $\mathbf{1}\{k = 0\}$ is defined to be 1 if $k = 0$ and zero otherwise)

Example 2: General Linear Processes From the simple building block provided by white noise, we can construct a very flexible family of covariance stationary processes — the *general linear processes*

$$X_t = \sum_{j=0}^{\infty} \psi_j \epsilon_{t-j}, \quad t \in \mathbb{Z} \tag{3.85}$$

where

- $\{\epsilon_t\}$ is white noise
- $\{\psi_t\}$ is a square summable sequence in \mathbb{R} (that is, $\sum_{t=0}^{\infty} \psi_t^2 < \infty$)

The sequence $\{\psi_t\}$ is often called a *linear filter*

With some manipulations it is possible to confirm that the autocovariance function for (3.85) is

$$\gamma(k) = \sigma^2 \sum_{j=0}^{\infty} \psi_j \psi_{j+k} \tag{3.86}$$

By the [Cauchy-Schwartz inequality](#) one can show that the last expression is finite. Clearly it does not depend on t

Wold's Decomposition Remarkably, the class of general linear processes goes a long way towards describing the entire class of zero-mean covariance stationary processes

In particular, [Wold's theorem](#) states that every zero-mean covariance stationary process $\{X_t\}$ can be written as

$$X_t = \sum_{j=0}^{\infty} \psi_j \epsilon_{t-j} + \eta_t$$

where

- $\{\epsilon_t\}$ is white noise
- $\{\psi_t\}$ is square summable
- η_t can be expressed as a linear function of X_{t-1}, X_{t-2}, \dots and is perfectly predictable over arbitrarily long horizons

For intuition and further discussion, see [\[Sar87\]](#), p. 286

AR and MA General linear processes are a very broad class of processes, and it often pays to specialize to those for which there exists a representation having only finitely many parameters

(In fact, experience shows that models with a relatively small number of parameters typically perform better than larger models, especially for forecasting)

One very simple example of such a model is the AR(1) process

$$X_t = \phi X_{t-1} + \epsilon_t \quad \text{where } |\phi| < 1 \quad \text{and } \{\epsilon_t\} \text{ is white noise} \quad (3.87)$$

By direct substitution, it is easy to verify that $X_t = \sum_{j=0}^{\infty} \phi^j \epsilon_{t-j}$

Hence $\{X_t\}$ is a general linear process

Applying (3.86) to the previous expression for X_t , we get the AR(1) autocovariance function

$$\gamma(k) = \phi^k \frac{\sigma^2}{1 - \phi^2}, \quad k = 0, 1, \dots \quad (3.88)$$

The next figure plots this function for $\phi = 0.8$ and $\phi = -0.8$ with $\sigma = 1$

Another very simple process is the MA(1) process

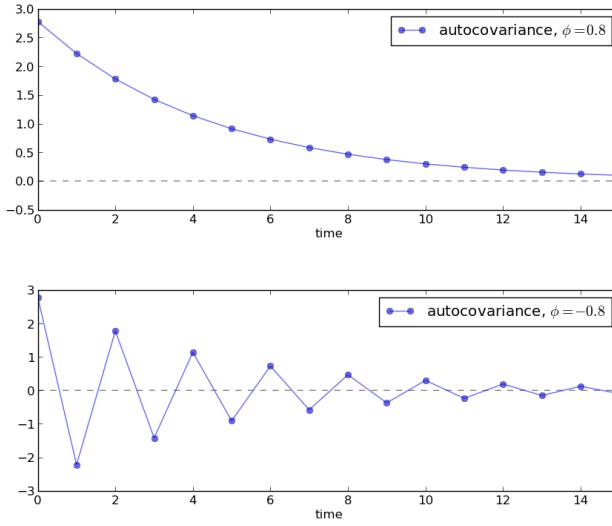
$$X_t = \epsilon_t + \theta \epsilon_{t-1}$$

You will be able to verify that

$$\gamma(0) = \sigma^2(1 + \theta^2), \quad \gamma(1) = \sigma^2 \theta, \quad \text{and} \quad \gamma(k) = 0 \quad \forall k > 1$$

The AR(1) can be generalized to an AR(p) and likewise for the MA(1)

Putting all of this together, we get the



ARMA Processes A stochastic process $\{X_t\}$ is called an *autoregressive moving average process*, or ARMA(p, q), if it can be written as

$$X_t = \phi_1 X_{t-1} + \cdots + \phi_p X_{t-p} + \epsilon_t + \theta_1 \epsilon_{t-1} + \cdots + \theta_q \epsilon_{t-q} \quad (3.89)$$

where $\{\epsilon_t\}$ is white noise

There is an alternative notation for ARMA processes in common use, based around the *lag operator* L

Def. Given arbitrary variable Y_t , let $L^k Y_t := Y_{t-k}$

It turns out that

- lag operators can lead to very succinct expressions for linear stochastic processes
- algebraic manipulations treating the lag operator as an ordinary scalar often are legitimate

Using L , we can rewrite (3.89) as

$$L^0 X_t - \phi_1 L^1 X_t - \cdots - \phi_p L^p X_t = L^0 \epsilon_t + \theta_1 L^1 \epsilon_t + \cdots + \theta_q L^q \epsilon_t \quad (3.90)$$

If we let $\phi(z)$ and $\theta(z)$ be the polynomials

$$\phi(z) := 1 - \phi_1 z - \cdots - \phi_p z^p \quad \text{and} \quad \theta(z) := 1 + \theta_1 z + \cdots + \theta_q z^q \quad (3.91)$$

then (3.90) simplifies further to

$$\phi(L) X_t = \theta(L) \epsilon_t \quad (3.92)$$

In what follows we **always assume** that the roots of the polynomial $\phi(z)$ lie outside the unit circle in the complex plane

This condition is sufficient to guarantee that the ARMA(p, q) process is covariance stationary

In fact it implies that the process falls within the class of general linear processes *described above*

That is, given an ARMA(p, q) process $\{X_t\}$ satisfying the unit circle condition, there exists a square summable sequence $\{\psi_t\}$ with $X_t = \sum_{j=0}^{\infty} \psi_j \epsilon_{t-j}$ for all t

The sequence $\{\psi_t\}$ can be obtained by a recursive procedure outlined on page 79 of [CC08]

In this context, the function $t \mapsto \psi_t$ is often called the *impulse response function*

Spectral Analysis

Autocovariance functions provide a great deal of information about covariance stationary processes

In fact, for zero-mean Gaussian processes, the autocovariance function characterizes the entire joint distribution

Even for non-Gaussian processes, it provides a significant amount of information

It turns out that there is an alternative representation of the autocovariance function of a covariance stationary process, called the *spectral density*

At times, the spectral density is easier to derive, easier to manipulate and provides additional intuition

Complex Numbers Before discussing the spectral density, we invite you to recall the main properties of complex numbers (or *skip to the next section*)

It can be helpful to remember that, in a formal sense, complex numbers are just points $(x, y) \in \mathbb{R}^2$ endowed with a specific notion of multiplication

When (x, y) is regarded as a complex number, x is called the *real part* and y is called the *imaginary part*

The *modulus* or *absolute value* of a complex number $z = (x, y)$ is just its Euclidean norm in \mathbb{R}^2 , but is usually written as $|z|$ instead of $\|z\|$

The product of two complex numbers (x, y) and (u, v) is defined to be $(xu - vy, xv + yu)$, while addition is standard pointwise vector addition

When endowed with these notions of multiplication and addition, the set of complex numbers forms a *field* — addition and multiplication play well together, just as they do in \mathbb{R}

The complex number (x, y) is often written as $x + iy$, where i is called the *imaginary unit*, and is understood to obey $i^2 = -1$

The $x + iy$ notation can be thought of as an easy way to remember the definition of multiplication given above, because, proceeding naively,

$$(x + iy)(u + iv) = xu - vy + i(xv + yu)$$

Converted back to our first notation, this becomes $(xu - vy, xv + yu)$, which is the same as the product of (x, y) and (u, v) from our previous definition

Complex numbers are also sometimes expressed in their polar form $re^{i\omega}$, which should be interpreted as

$$re^{i\omega} := r(\cos(\omega) + i \sin(\omega))$$

Spectral Densities Let $\{X_t\}$ be a covariance stationary process with autocovariance function γ satisfying $\sum_k \gamma(k)^2 < \infty$

The *spectral density* f of $\{X_t\}$ is defined as the [discrete time Fourier transform](#) of its autocovariance function γ

$$f(\omega) := \sum_{k \in \mathbb{Z}} \gamma(k) e^{-i\omega k}, \quad \omega \in \mathbb{R}$$

(Some authors normalize the expression on the right by constants such as $1/\pi$ — the chosen convention makes little difference provided you are consistent)

Using the fact that γ is *even*, in the sense that $\gamma(t) = \gamma(-t)$ for all t , you should be able to show that

$$f(\omega) = \gamma(0) + 2 \sum_{k \geq 1} \gamma(k) \cos(\omega k) \tag{3.93}$$

It is not difficult to confirm that f is

- real-valued
- even ($f(\omega) = f(-\omega)$), and
- 2π -periodic, in the sense that $f(2\pi + \omega) = f(\omega)$ for all ω

It follows that the values of f on $[0, \pi]$ determine the values of f on all of \mathbb{R} — the proof is an exercise

For this reason it is standard to plot the spectral density only on the interval $[0, \pi]$

Example 1: White Noise Consider a white noise process $\{\epsilon_t\}$ with standard deviation σ

It is simple to check that in this case we have $f(\omega) = \sigma^2$. In particular, f is a constant function

As we will see, this can be interpreted as meaning that “all frequencies are equally present”

(White light has this property when frequency refers to the visible spectrum, a connection that provides the origins of the term “white noise”)

Example 2: AR and MA and ARMA It is an exercise to show that the MA(1) process $X_t = \theta \epsilon_{t-1} + \epsilon_t$ has spectral density

$$f(\omega) = \sigma^2 (1 + 2\theta \cos(\omega) + \theta^2) \tag{3.94}$$

With a bit more effort, it’s possible to show (see, e.g., p. 261 of [\[Sar87\]](#)) that the spectral density of the AR(1) process $X_t = \phi X_{t-1} + \epsilon_t$ is

$$f(\omega) = \frac{\sigma^2}{1 - 2\phi \cos(\omega) + \phi^2} \tag{3.95}$$

More generally, it can be shown that the spectral density of the ARMA process (3.89) is

$$f(\omega) = \left| \frac{\theta(e^{i\omega})}{\phi(e^{i\omega})} \right|^2 \sigma^2 \quad (3.96)$$

where

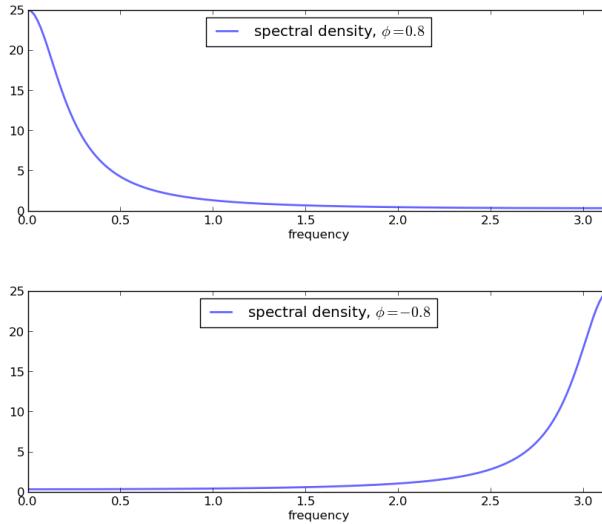
- σ is the standard deviation of the white noise process $\{\epsilon_t\}$
- the polynomials $\phi(\cdot)$ and $\theta(\cdot)$ are as defined in (3.91)

The derivation of (3.96) uses the fact that convolutions become products under Fourier transformations

The proof is elegant and can be found in many places — see, for example, [Sar87], chapter 11, section 4

It's a nice exercise to verify that (3.94) and (3.95) are indeed special cases of (3.96)

Interpreting the Spectral Density Plotting (3.95) reveals the shape of the spectral density for the AR(1) model when ϕ takes the values 0.8 and -0.8 respectively



These spectral densities correspond to the autocovariance functions for the AR(1) process *shown above*

Informally, we think of the spectral density as being large at those $\omega \in [0, \pi]$ such that the autocovariance function exhibits significant cycles at this “frequency”

To see the idea, let's consider why, in the lower panel of the preceding figure, the spectral density for the case $\phi = -0.8$ is large at $\omega = \pi$

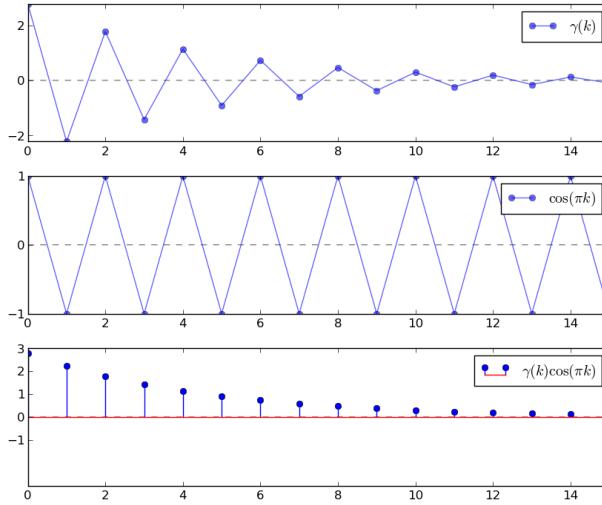
Recall that the spectral density can be expressed as

$$f(\omega) = \gamma(0) + 2 \sum_{k \geq 1} \gamma(k) \cos(\omega k) = \gamma(0) + 2 \sum_{k \geq 1} (-0.8)^k \cos(\omega k) \quad (3.97)$$

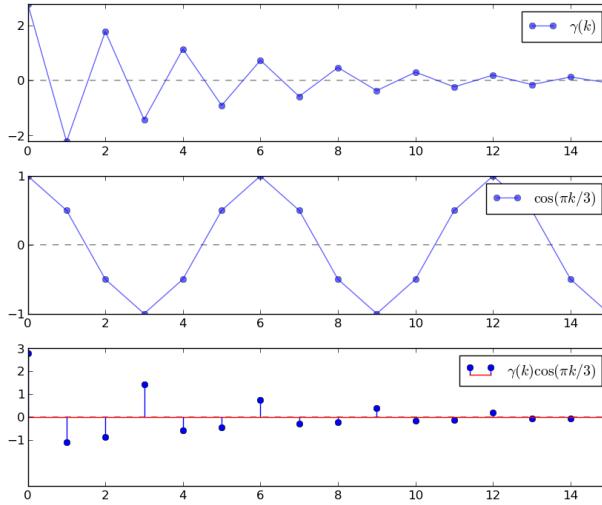
When we evaluate this at $\omega = \pi$, we get a large number because $\cos(\pi k)$ is large and positive when $(-0.8)^k$ is positive, and large in absolute value and negative when $(-0.8)^k$ is negative

Hence the product is always large and positive, and hence the sum of the products on the right-hand side of (3.97) is large

These ideas are illustrated in the next figure, which has k on the horizontal axis (click to enlarge)



On the other hand, if we evaluate $f(\omega)$ at $\omega = \pi/3$, then the cycles are not matched, the sequence $\gamma(k) \cos(\omega k)$ contains both positive and negative terms, and hence the sum of these terms is much smaller



In summary, the spectral density is large at frequencies ω where the autocovariance function ex-

hibits cycles

Inverting the Transformation We have just seen that the spectral density is useful in the sense that it provides a frequency-based perspective on the autocovariance structure of a covariance stationary process

Another reason that the spectral density is useful is that it can be “inverted” to recover the autocovariance function via the *inverse Fourier transform*

In particular, for all $k \in \mathbb{Z}$, we have

$$\gamma(k) = \frac{1}{2\pi} \int_{-\pi}^{\pi} f(\omega) e^{i\omega k} d\omega \quad (3.98)$$

This is convenient in situations where the spectral density is easier to calculate and manipulate than the autocovariance function

(For example, the expression (3.96) for the ARMA spectral density is much easier to work with than the expression for the ARMA autocovariance)

Mathematical Theory This section is loosely based on [Sar87], p. 249-253, and included for those who

- would like a bit more insight into spectral densities
- and have at least some background in Hilbert space theory

Others should feel free to skip to the *next section* — none of this material is necessary to progress to computation

Recall that every **separable** Hilbert space H has a countable orthonormal basis $\{h_k\}$

The nice thing about such a basis is that every $f \in H$ satisfies

$$f = \sum_k \alpha_k h_k \quad \text{where} \quad \alpha_k := \langle f, h_k \rangle \quad (3.99)$$

where $\langle \cdot, \cdot \rangle$ denotes the inner product in H

Thus, f can be represented to any degree of precision by linearly combining basis vectors

The scalar sequence $\alpha = \{\alpha_k\}$ is called the *Fourier coefficients* of f , and satisfies $\sum_k |\alpha_k|^2 < \infty$

In other words, α is in ℓ_2 , the set of square summable sequences

Consider an operator T that maps $\alpha \in \ell_2$ into its expansion $\sum_k \alpha_k h_k \in H$

The Fourier coefficients of $T\alpha$ are just $\alpha = \{\alpha_k\}$, as you can verify by confirming that $\langle T\alpha, h_k \rangle = \alpha_k$

Using elementary results from Hilbert space theory, it can be shown that

- T is one-to-one — if α and β are distinct in ℓ_2 , then so are their expansions in H
- T is onto — if $f \in H$ then its preimage in ℓ_2 is the sequence α given by $\alpha_k = \langle f, h_k \rangle$
- T is a linear isometry — in particular $\langle \alpha, \beta \rangle = \langle T\alpha, T\beta \rangle$

Summarizing these results, we say that any separable Hilbert space is isometrically isomorphic to ℓ_2

In essence, this says that each separable Hilbert space we consider is just a different way of looking at the fundamental space ℓ_2

With this in mind, let's specialize to a setting where

- $\gamma \in \ell_2$ is the autocovariance function of a covariance stationary process, and f is the spectral density
- $H = L_2$, where L_2 is the set of square summable functions on the interval $[-\pi, \pi]$, with inner product $\langle g, h \rangle = \int_{-\pi}^{\pi} g(\omega)h(\omega)d\omega$
- $\{h_k\}$ = the orthonormal basis for L_2 given by the set of trigonometric functions

$$h_k(\omega) = \frac{e^{i\omega k}}{\sqrt{2\pi}}, \quad k \in \mathbb{Z}, \quad \omega \in [-\pi, \pi]$$

Using the definition of T from above and the fact that f is even, we now have

$$T\gamma = \sum_{k \in \mathbb{Z}} \gamma(k) \frac{e^{i\omega k}}{\sqrt{2\pi}} = \frac{1}{\sqrt{2\pi}} f(\omega) \quad (3.100)$$

In other words, apart from a scalar multiple, the spectral density is just a transformation of $\gamma \in \ell_2$ under a certain linear isometry — a different way to view γ

In particular, it is an expansion of the autocovariance function with respect to the trigonometric basis functions in L_2

As discussed above, the Fourier coefficients of $T\gamma$ are given by the sequence γ , and, in particular, $\gamma(k) = \langle T\gamma, h_k \rangle$

Transforming this inner product into its integral expression and using (3.100) gives (3.98), justifying our earlier expression for the inverse transform

Implementation

Most code for working with covariance stationary models deals with ARMA models

Julia code for studying ARMA models can be found in the `DSP.jl` package

Since this code doesn't quite cover our needs — particularly vis-a-vis spectral analysis — we've put together the module `arma.jl`, which is part of `QuantEcon.jl` package.

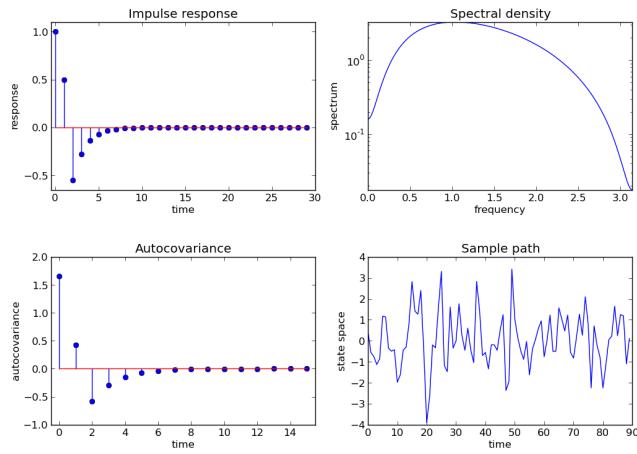
The module provides functions for mapping ARMA(p, q) models into their

1. impulse response function
2. simulated time series
3. autocovariance function
4. spectral density

In addition to individual plots of these entities, we provide functionality to generate 2x2 plots containing all this information

In other words, we want to replicate the plots on pages 68–69 of [LS12]

Here's an example corresponding to the model $X_t = 0.5X_{t-1} + \epsilon_t - 0.8\epsilon_{t-2}$



Code For interest's sake, `arma.jl` is printed below

```
#=
@authors: John Stachurski
>Date: Thu Aug 21 11:09:30 EST 2014

Provides functions for working with and visualizing scalar ARMA processes.
Ported from Python module quantecon.arma, which was written by Doc-Jin Jang,
Jerry Choi, Thomas Sargent and John Stachurski

References
-----
http://quant-econ.net/jl/arma.html

=#
"""
Represents a scalar ARMA(p, q) process

If phi and theta are scalars, then the model is
understood to be

    X_t = phi X_{t-1} + epsilon_t + theta epsilon_{t-1}

where epsilon_t is a white noise process with standard
deviation sigma.
```

If phi and theta are arrays or sequences,
then the interpretation is the ARMA(p, q) model

```
X_t = phi_1 X_{t-1} + ... + phi_p X_{t-p} +
epsilon_t + theta_1 epsilon_{t-1} + ... +
theta_q epsilon_{t-q}
```

where

```
* phi = (phi_1, phi_2, ..., phi_p)
* theta = (theta_1, theta_2, ..., theta_q)
* sigma is a scalar, the standard deviation of the white noise
```

Fields

- `phi::Vector` : AR parameters phi_1, ..., phi_p
- `theta::Vector` : MA parameters theta_1, ..., theta_q
- `p::Integer` : Number of AR coefficients
- `q::Integer` : Number of MA coefficients
- `sigma::Real` : Standard deviation of white noise
- `ma_poly::Vector` : MA polynomial --- filtering representation
- `ar_poly::Vector` : AR polynomial --- filtering representation

Examples

```
```julia
using QuantEcon
phi = 0.5
theta = [0.0, -0.8]
sigma = 1.0
lp = ARMA(phi, theta, sigma)
require(joinpath(dirname(@__FILE__), "..", "examples", "arma_plots.jl"))
quad_plot(lp)
```
```
type ARMA
 phi::Vector # AR parameters phi_1, ..., phi_p
 theta::Vector # MA parameters theta_1, ..., theta_q
 p::Integer # Number of AR coefficients
 q::Integer # Number of MA coefficients
 sigma::Real # Variance of white noise
 ma_poly::Vector # MA polynomial --- filtering representation
 ar_poly::Vector # AR polynomial --- filtering representation
end

constructors to coerce phi/theta to vectors
ARMA(phi::Real, theta::Real, sigma::Real) = ARMA([phi;], [theta;], sigma)
ARMA(phi::Real, theta::Vector, sigma::Real) = ARMA([phi;], theta, sigma)
ARMA(phi::Vector, theta::Real, sigma::Real) = ARMA(phi, [theta;], sigma)

function ARMA(phi::AbstractVector, theta::AbstractVector=[0.0], sigma::Real=1.0)
 # == Record dimensions == #
 p = length(phi)
```

```

q = length(theta)

== Build filtering representation of polynomials ==
ma_poly = [1.0; theta]
ar_poly = [1.0; -phi]
return ARMA(phi, theta, p, q, sigma, ma_poly, ar_poly)
end

"""

Compute the spectral density function.

The spectral density is the discrete time Fourier transform of the
autocovariance function. In particular,

f(w) = sum_k gamma(k) exp(-ikw)

where gamma is the autocovariance function and the sum is over
the set of all integers.

Arguments

- `arma::ARMA`: Instance of `ARMA` type
- `;two_pi::Bool(true)": Compute the spectral density function over [0, pi] if
 false and [0, 2 pi] otherwise.
- `;res(1200)": If `res` is a scalar then the spectral density is computed at
 `res` frequencies evenly spaced around the unit circle, but if `res` is an array
 then the function computes the response at the frequencies given by the array

Returns

- `w::Vector{Float64}`: The normalized frequencies at which h was computed, in
 radians/sample
- `spect::Vector{Float64}` : The frequency response
"""

function spectral_density(arma::ARMA; res=1200, two_pi::Bool=true)
 # Compute the spectral density associated with ARMA process arma
 wmax = two_pi ? 2pi : pi
 w = linspace(0, wmax, res)
 tf = TFFilter(reverse(arma.ma_poly), reverse(arma.ar_poly))
 h = freqz(tf, w)
 spect = arma.sigma^2 * abs(h).^2
 return w, spect
end

"""

Compute the autocovariance function from the ARMA parameters
over the integers range(num_autocov) using the spectral density
and the inverse Fourier transform.

Arguments

- `arma::ARMA`: Instance of `ARMA` type
- `;num_autocov::Integer(16)": The number of autocovariances to calculate

```

```

"""
function autocovariance(arma::ARMA; num_autocov::Integer=16)
 # Compute the autocovariance function associated with ARMA process arma
 # Computation is via the spectral density and inverse FFT
 (w, spect) = spectral_density(arma)
 acov = real(Base.ifft(spect))
 # num_autocov should be <= len(acov) / 2
 return acov[1:num_autocov]
end

"""
Get the impulse response corresponding to our model.

Arguments

- `arma::ARMA`: Instance of `ARMA` type
- `;impulse_length::Integer(30)`: Length of horizon for calculating impulse
 response. Must be at least as long as the `p` fields of `arma`

#####
>Returns

- `psi::Vector{Float64}`: `psi[j]` is the response at lag j of the impulse
 response. We take psi[1] as unity.

"""
function impulse_response(arma::ARMA; impulse_length=30)
 # Compute the impulse response function associated with ARMA process arma
 err_msg = "Impulse length must be greater than number of AR coefficients"
 @assert impulse_length >= arma.p err_msg
 # == Pad theta with zeros at the end == #
 theta = [arma.theta; zeros(impulse_length - arma.q)]
 psi_zero = 1.0
 psi = Array(Float64, impulse_length)
 for j = 1:impulse_length
 psi[j] = theta[j]
 for i = 1:min(j, arma.p)
 psi[j] += arma.phi[i] * (j-i > 0 ? psi[j-i] : psi_zero)
 end
 end
 return [psi_zero; psi[1:end-1]]
end

"""
Compute a simulated sample path assuming Gaussian shocks.

Arguments

- `arma::ARMA`: Instance of `ARMA` type
- `;ts_length::Integer(90)`: Length of simulation
- `;impulse_length::Integer(30)`: Horizon for calculating impulse response
 (see also docstring for `impulse_response`)

```

```
Returns

- `X::Vector{Float64}`: Simulation of the ARMA model `arma`

"""

function simulation(arma::ARMA; ts_length=90, impulse_length=30)
 # Simulate the ARMA process arma assuing Gaussian shocks
 J = impulse_length
 T = ts_length
 psi = impulse_response(arma, impulse_length=impulse_length)
 epsilon = arma.sigma * randn(T + J)
 X = Array(Float64, T)
 for t=1:T
 X[t] = dot(epsilon[t:J+t-1], psi)
 end
 return X
end
```

Here's an example of usage

```
julia> using QuantEcon

julia> using QuantEcon

julia> phi = 0.5;

julia> theta = [0, -0.8];

julia> lp = ARMA(phi, theta);

julia> QuantEcon.quad_plot(lp)
```

**Explanation** The call

```
lp = ARMA(phi, theta, sigma)
```

creates an instance `lp` that represents the ARMA( $p, q$ ) model

$$X_t = \phi_1 X_{t-1} + \dots + \phi_p X_{t-p} + \epsilon_t + \theta_1 \epsilon_{t-1} + \dots + \theta_q \epsilon_{t-q}$$

If `phi` and `theta` are arrays or sequences, then the interpretation will be

- `phi` holds the vector of parameters  $(\phi_1, \phi_2, \dots, \phi_p)$
- `theta` holds the vector of parameters  $(\theta_1, \theta_2, \dots, \theta_q)$

The parameter `sigma` is always a scalar, the standard deviation of the white noise

We also permit `phi` and `theta` to be scalars, in which case the model will be interpreted as

$$X_t = \phi X_{t-1} + \epsilon_t + \theta \epsilon_{t-1}$$

The two numerical packages most useful for working with ARMA models are `DSP.jl` and the `fft` routine in Julia

**Computing the Autocovariance Function** As discussed above, for ARMA processes the spectral density has a *simple representation* that is relatively easy to calculate

Given this fact, the easiest way to obtain the autocovariance function is to recover it from the spectral density via the inverse Fourier transform

Here we use Julia's Fourier transform routine `fft`, which wraps a standard C-based package called FFTW

A look at [the fft documentation](#) shows that the inverse transform `ifft` takes a given sequence  $A_0, A_1, \dots, A_{n-1}$  and returns the sequence  $a_0, a_1, \dots, a_{n-1}$  defined by

$$a_k = \frac{1}{n} \sum_{t=0}^{n-1} A_t e^{ik2\pi t/n}$$

Thus, if we set  $A_t = f(\omega_t)$ , where  $f$  is the spectral density and  $\omega_t := 2\pi t/n$ , then

$$a_k = \frac{1}{n} \sum_{t=0}^{n-1} f(\omega_t) e^{i\omega_t k} = \frac{1}{2\pi} \frac{2\pi}{n} \sum_{t=0}^{n-1} f(\omega_t) e^{i\omega_t k}, \quad \omega_t := 2\pi t/n$$

For  $n$  sufficiently large, we then have

$$a_k \approx \frac{1}{2\pi} \int_0^{2\pi} f(\omega) e^{i\omega k} d\omega = \frac{1}{2\pi} \int_{-\pi}^{\pi} f(\omega) e^{i\omega k} d\omega$$

(You can check the last equality)

In view of (3.98) we have now shown that, for  $n$  sufficiently large,  $a_k \approx \gamma(k)$  — which is exactly what we want to compute

## Estimation of Spectra

### Contents

- *Estimation of Spectra*
  - *Overview*
  - *Periodograms*
  - *Smoothing*
  - *Exercises*
  - *Solutions*

### Overview

In a *previous lecture* we covered some fundamental properties of covariance stationary linear stochastic processes

One objective for that lecture was to introduce spectral densities — a standard and very useful technique for analyzing such processes

In this lecture we turn to the problem of estimating spectral densities and other related quantities from data

Estimates of the spectral density are computed using what is known as a periodogram — which in turn is computed via the famous [fast Fourier transform](#)

Once the basic technique has been explained, we will apply it to the analysis of several key macroeconomic time series

For supplementary reading, see [\[Sar87\]](#) or [\[CC08\]](#).

### Periodograms

Recall that the spectral density  $f$  of a covariance stationary process with autocorrelation function  $\gamma$  can be written as

$$f(\omega) = \gamma(0) + 2 \sum_{k \geq 1} \gamma(k) \cos(\omega k), \quad \omega \in \mathbb{R}$$

Now consider the problem of estimating the spectral density of a given time series, when  $\gamma$  is unknown

In particular, let  $X_0, \dots, X_{n-1}$  be  $n$  consecutive observations of a single time series that is assumed to be covariance stationary

The most common estimator of the spectral density of this process is the *periodogram* of  $X_0, \dots, X_{n-1}$ , which is defined as

$$I(\omega) := \frac{1}{n} \left| \sum_{t=0}^{n-1} X_t e^{it\omega} \right|^2, \quad \omega \in \mathbb{R} \tag{3.101}$$

(Recall that  $|z|$  denotes the modulus of complex number  $z$ )

Alternatively,  $I(\omega)$  can be expressed as

$$I(\omega) = \frac{1}{n} \left\{ \left[ \sum_{t=0}^{n-1} X_t \cos(\omega t) \right]^2 + \left[ \sum_{t=0}^{n-1} X_t \sin(\omega t) \right]^2 \right\}$$

It is straightforward to show that the function  $I$  is even and  $2\pi$ -periodic (i.e.,  $I(\omega) = I(-\omega)$  and  $I(\omega + 2\pi) = I(\omega)$  for all  $\omega \in \mathbb{R}$ )

From these two results, you will be able to verify that the values of  $I$  on  $[0, \pi]$  determine the values of  $I$  on all of  $\mathbb{R}$

The next section helps to explain the connection between the periodogram and the spectral density

**Interpretation** To interpret the periodogram, it is convenient to focus on its values at the *Fourier frequencies*

$$\omega_j := \frac{2\pi j}{n}, \quad j = 0, \dots, n-1$$

In what sense is  $I(\omega_j)$  an estimate of  $f(\omega_j)$ ?

The answer is straightforward, although it does involve some algebra

With a bit of effort one can show that, for any integer  $j > 0$ ,

$$\sum_{t=0}^{n-1} e^{it\omega_j} = \sum_{t=0}^{n-1} \exp \left\{ i2\pi j \frac{t}{n} \right\} = 0$$

Letting  $\bar{X}$  denote the sample mean  $n^{-1} \sum_{t=0}^{n-1} X_t$ , we then have

$$nI(\omega_j) = \left| \sum_{t=0}^{n-1} (X_t - \bar{X}) e^{it\omega_j} \right|^2 = \sum_{t=0}^{n-1} (X_t - \bar{X}) e^{it\omega_j} \sum_{r=0}^{n-1} (X_r - \bar{X}) e^{-ir\omega_j}$$

By carefully working through the sums, one can transform this to

$$nI(\omega_j) = \sum_{t=0}^{n-1} (X_t - \bar{X})^2 + 2 \sum_{k=1}^{n-1} \sum_{t=k}^{n-1} (X_t - \bar{X})(X_{t-k} - \bar{X}) \cos(\omega_j k)$$

Now let

$$\hat{\gamma}(k) := \frac{1}{n} \sum_{t=k}^{n-1} (X_t - \bar{X})(X_{t-k} - \bar{X}), \quad k = 0, 1, \dots, n-1$$

This is the sample autocovariance function, the natural “plug-in estimator” of the *autocovariance function*  $\gamma$

(“Plug-in estimator” is an informal term for an estimator found by replacing expectations with sample means)

With this notation, we can now write

$$I(\omega_j) = \hat{\gamma}(0) + 2 \sum_{k=1}^{n-1} \hat{\gamma}(k) \cos(\omega_j k)$$

Recalling our expression for  $f$  given *above*, we see that  $I(\omega_j)$  is just a sample analog of  $f(\omega_j)$

**Calculation** Let’s now consider how to compute the periodogram as defined in (3.101)

There are already functions available that will do this for us — an example is `periodogram` in the `DSP.jl` package

However, it is very simple to replicate their results, and this will give us a platform to make useful extensions

The most common way to calculate the periodogram is via the discrete Fourier transform, which in turn is implemented through the `fast Fourier transform` algorithm

In general, given a sequence  $a_0, \dots, a_{n-1}$ , the discrete Fourier transform computes the sequence

$$A_j := \sum_{t=0}^{n-1} a_t \exp \left\{ i2\pi \frac{tj}{n} \right\}, \quad j = 0, \dots, n-1$$

With  $a_0, \dots, a_{n-1}$  stored in Julia array `a`, the function call `fft(a)` returns the values  $A_0, \dots, A_{n-1}$  as a Julia array

It follows that, when the data  $X_0, \dots, X_{n-1}$  is stored in array  $\mathbf{X}$ , the values  $I(\omega_j)$  at the Fourier frequencies, which are given by

$$\frac{1}{n} \left| \sum_{t=0}^{n-1} X_t \exp \left\{ i2\pi \frac{tj}{n} \right\} \right|^2, \quad j = 0, \dots, n-1$$

can be computed by `abs(fft(X)) .^ 2 / length(X)`

Note: The Julia function `abs` acts elementwise, and correctly handles complex numbers (by computing their modulus, which is exactly what we need)

Here's a function that puts all this together

```
function periodogram(x::Array):
 n = length(x)
 I_w = abs(fft(x)) .^ 2 / n
 w = 2pi * [0:n-1] ./ n # Fourier frequencies
 w, I_w = w[1:round(Int, n/2)], I_w[1:round(Int, n/2)] # Truncate to interval [0, pi]
 return w, I_w
end
```

Let's generate some data for this function using the ARMA type from QuantEcon

(See the *lecture on linear processes* for details on this class)

Here's a code snippet that, once the preceding code has been run, generates data from the process

$$X_t = 0.5X_{t-1} + \epsilon_t - 0.8\epsilon_{t-2} \quad (3.102)$$

where  $\{\epsilon_t\}$  is white noise with unit variance, and compares the periodogram to the actual spectral density

```
import PyPlot: plt
import QuantEcon: ARMA

n = 40 # Data size
phi, theta = 0.5, [0, -0.8] # AR and MA parameters
lp = ARMA(phi, theta)
X = simulation(lp, ts_length=n)

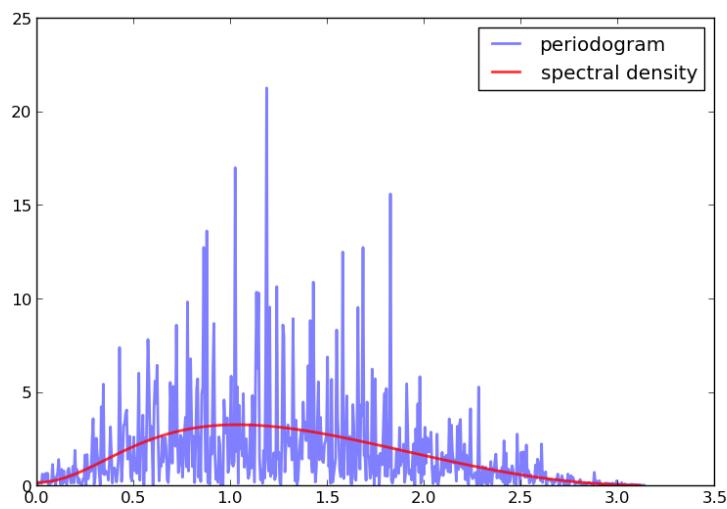
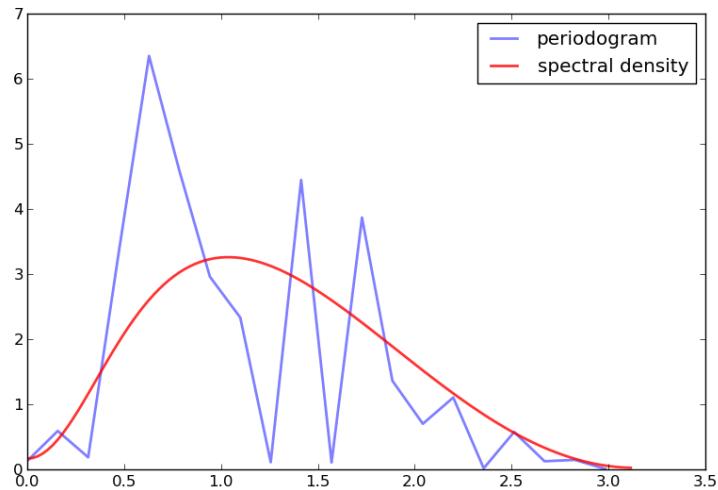
fig, ax = plt.subplots()
x, y = periodogram(X)
ax[:plot](x, y, "b-", lw=2, alpha=0.5, label="periodogram")
x_sd, y_sd = spectral_density(lp, two_pi=False, resolution=120)
ax[:plot](x_sd, y_sd, "r-", lw=2, alpha=0.8, label="spectral density")
ax[:legend]()
plt.show()
```

Running this should produce a figure similar to this one

This estimate looks rather disappointing, but the data size is only 40, so perhaps it's not surprising that the estimate is poor

However, if we try again with  $n = 1200$  the outcome is not much better

The periodogram is far too irregular relative to the underlying spectral density



This brings us to our next topic

### Smoothing

There are two related issues here

One is that, given the way the fast Fourier transform is implemented, the number of points  $\omega$  at which  $I(\omega)$  is estimated increases in line with the amount of data

In other words, although we have more data, we are also using it to estimate more values

A second issue is that densities of all types are fundamentally hard to estimate without parametric assumptions

Typically, nonparametric estimation of densities requires some degree of smoothing

The standard way that smoothing is applied to periodograms is by taking local averages

In other words, the value  $I(\omega_j)$  is replaced with a weighted average of the adjacent values

$$I(\omega_{j-p}), I(\omega_{j-p+1}), \dots, I(\omega_j), \dots, I(\omega_{j+p})$$

This weighted average can be written as

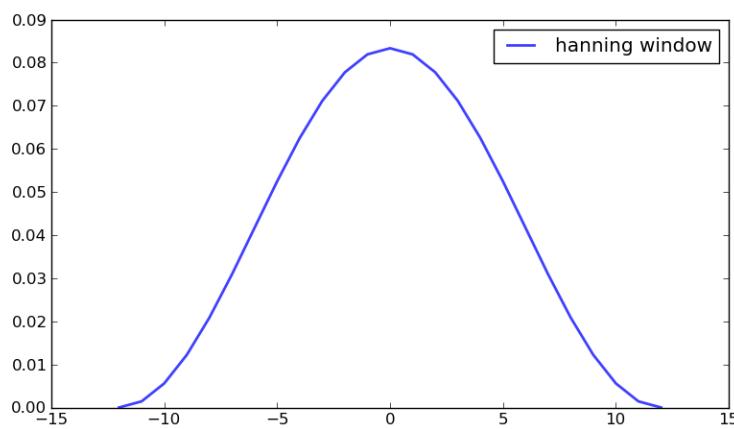
$$I_S(\omega_j) := \sum_{\ell=-p}^p w(\ell) I(\omega_{j+\ell}) \quad (3.103)$$

where the weights  $w(-p), \dots, w(p)$  are a sequence of  $2p + 1$  nonnegative values summing to one

In generally, larger values of  $p$  indicate more smoothing — more on this below

The next figure shows the kind of sequence typically used

Note the smaller weights towards the edges and larger weights in the center, so that more distant values from  $I(\omega_j)$  have less weight than closer ones in the sum (3.103)



**Estimation with Smoothing** Our next step is to provide code that will not only estimate the periodogram but also provide smoothing as required

Such functions have been written in `estspec.jl` and are available via `QuantEcon.jl`

The file `estspec.jl` are printed below

```
#=
Functions for working with periodograms of scalar data.

@author : Spencer Lyon <spencer.lyon@nyu.edu>

@date : 2014-08-21

References

http://quant-econ.net/jl/estspec.html

=#
import DSP

"""
Smooth the data in x using convolution with a window of requested size and type.

Arguments

- `x::Array`: An array containing the data to smooth
- `window_len::Int(7)`: An odd integer giving the length of the window
- `window::AbstractString("hanning")`: A string giving the window type. Possible values are `flat`, `hanning`, `hamming`, `bartlett`, or `blackman`

Returns

- `out::Array`: The array of smoothed data
"""
function smooth(x::Array, window_len::Int, window::AbstractString="hanning")
 if length(x) < window_len
 throw(ArgumentError("Input vector length must be >= window length"))
 end

 if window_len < 3
 throw(ArgumentError("Window length must be at least 3."))
 end

 if iseven(window_len)
 window_len += 1
 println("Window length must be odd, reset to $window_len")
 end

 windows = Dict("hanning" => DSP.hanning,
 "hamming" => DSP.hamming,
 "bartlett" => DSP.bartlett,
 "blackman" => DSP.blackman,
```

```

 "flat" => DSP.rect # moving average
)

Reflect x around x[0] and x[-1] prior to convolution
k = round(Int, window_len / 2)
xb = x[1:k] # First k elements
xt = x[end-k+1:end] # Last k elements
s = [reverse(xb); x; reverse(xt)]

=== Select window values ===
if !haskey(windows, window)
 msg = "Unrecognized window type '$window'"
 print(msg * " Defaulting to hanning")
 window = "hanning"
end

w = windows[window](window_len)

return conv(w ./ sum(w), s)[window_len+1:end-window_len]
end

"Version of `smooth` where `window_len` and `window` are keyword arguments"
function smooth(x::Array; window_len::Int=7, window::AbstractString="hanning")
 smooth(x, window_len, window)
end

function periodogram(x::Vector)
 n = length(x)
 I_w = abs(fft(x)).^2 ./ n
 w = 2pi * (0:n-1) ./ n # Fourier frequencies

 # int rounds to nearest integer. We want to round up or take 1/2 + 1 to
 # make sure we get the whole interval from [0, pi]
 ind = iseven(n) ? round(Int, n / 2 + 1) : ceil(Int, n / 2)
 w, I_w = w[1:ind], I_w[1:ind]
 return w, I_w
end

function periodogram(x::Vector, window::AbstractString, window_len::Int=7)
 w, I_w = periodogram(x)
 I_w = smooth(I_w, window_len=window_len, window=window)
 return w, I_w
end

"""
Computes the periodogram

I(w) = (1 / n) | sum_{t=0}^{n-1} x_t e^{itw} |^2
at the Fourier frequencies w_j := 2 pi j / n, j = 0, ..., n - 1, using the fast
Fourier transform. Only the frequencies w_j in [0, pi] and corresponding values
I(w_j) are returned. If a window type is given then smoothing is performed.

```

```

Arguments

- `x::Array`: An array containing the data to smooth
- `window_len::Int(7)`: An odd integer giving the length of the window
- `window::AbstractString("hanning")`: A string giving the window type. Possible values
are `flat`, `hanning`, `hamming`, `bartlett`, or `blackman`

Returns

- `w::Array{Float64}`: Fourier frequencies at which the periodogram is evaluated
- `I_w::Array{Float64}`: The periodogram at frequencies `w`

"""
periodogram

"""

Compute periodogram from data `x`, using prewhitening, smoothing and recoloring.
The data is fitted to an AR(1) model for prewhitening, and the residuals are
used to compute a first-pass periodogram with smoothing. The fitted
coefficients are then used for recoloring.

Arguments

- `x::Array`: An array containing the data to smooth
- `window_len::Int(7)`: An odd integer giving the length of the window
- `window::AbstractString("hanning")`: A string giving the window type. Possible values
are `flat`, `hanning`, `hamming`, `bartlett`, or `blackman`

Returns

- `w::Array{Float64}`: Fourier frequencies at which the periodogram is evaluated
- `I_w::Array{Float64}`: The periodogram at frequencies `w`

"""

function ar_periodogram(x::Array, window::AbstractString="hanning", window_len::Int=7)
 # run regression
 x_current, x_lagged = x[2:end], x[1:end-1] # x_t and x_{t-1}
 coefs = collect(linreg(x_lagged, x_current))

 # get estimated values and compute residual
 est = [ones(x_lagged) x_lagged] * coefs
 e_hat = x_current - est

 phi = coefs[2]

 # compute periodogram on residuals
 w, I_w = periodogram(e_hat, window, window_len)

 # recolor and return
 I_w = I_w ./ abs(1 - phi .* exp(im.*w)).^2

 return w, I_w
end

```

The listing displays three functions, `smooth()`, `periodogram()`, `ar_periodogram()`. We will discuss the first two here and the third one *below*

The `periodogram()` function returns a periodogram, optionally smoothed via the `smooth()` function

Regarding the `smooth()` function, since smoothing adds a nontrivial amount of computation, we have applied a fairly terse array-centric method based around `conv`

Readers are left to either explore or simply use this code according to their interests

The next three figures each show smoothed and unsmoothed periodograms, as well as the true spectral density

(The model is the same as before — see equation (3.102) — and there are 400 observations)

From top figure to bottom, the window length is varied from small to large

In looking at the figure, we can see that for this model and data size, the window length chosen in the middle figure provides the best fit

Relative to this value, the first window length provides insufficient smoothing, while the third gives too much smoothing

Of course in real estimation problems the true spectral density is not visible and the choice of appropriate smoothing will have to be made based on judgement/priors or some other theory

**Pre-Filtering and Smoothing** In the code listing *above* we showed three functions from the file `estspec.jl`

The third function in the file (`ar_periodogram()`) adds a pre-processing step to periodogram smoothing

First we describe the basic idea, and after that we give the code

The essential idea is to

1. Transform the data in order to make estimation of the spectral density more efficient
2. Compute the periodogram associated with the transformed data
3. Reverse the effect of the transformation on the periodogram, so that it now estimates the spectral density of the original process

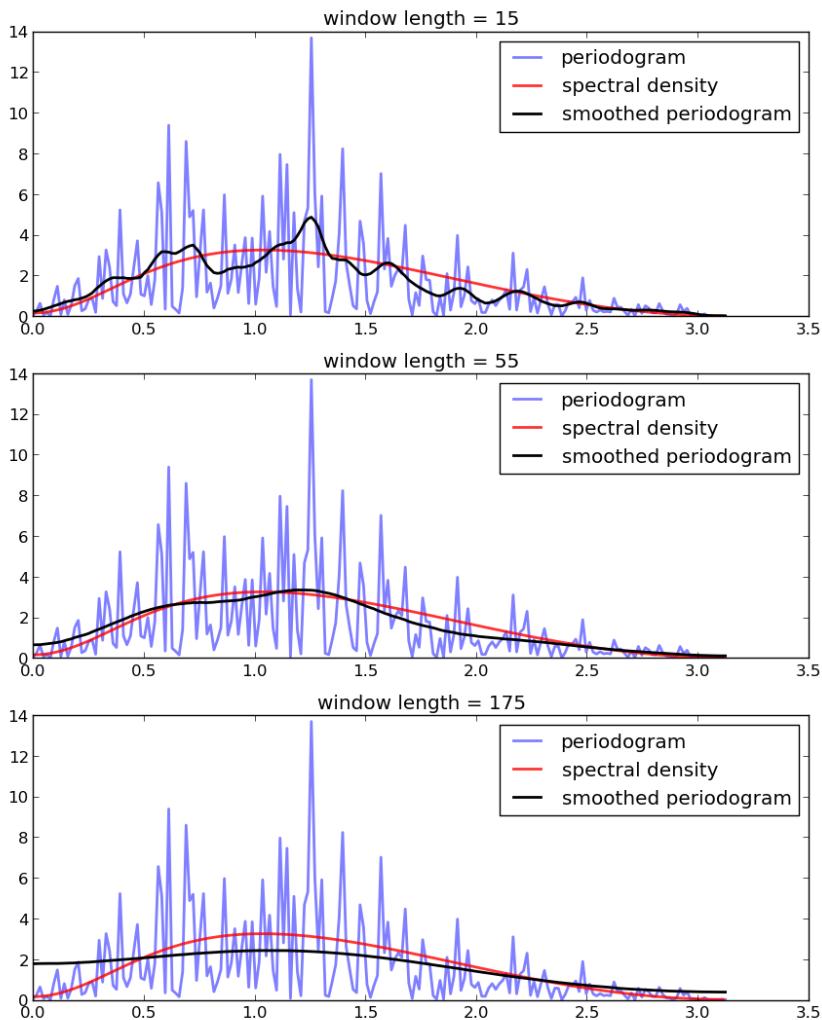
Step 1 is called *pre-filtering* or *pre-whitening*, while step 3 is called *recoloring*

The first step is called pre-whitening because the transformation is usually designed to turn the data into something closer to white noise

Why would this be desirable in terms of spectral density estimation?

The reason is that we are smoothing our estimated periodogram based on estimated values at nearby points — recall (3.103)

The underlying assumption that makes this a good idea is that the true spectral density is relatively regular — the value of  $I(\omega)$  is close to that of  $I(\omega')$  when  $\omega$  is close to  $\omega'$



This will not be true in all cases, but it is certainly true for white noise

For white noise,  $I$  is as regular as possible — *it is a constant function*

In this case, values of  $I(\omega')$  at points  $\omega'$  near to  $\omega$  provided the maximum possible amount of information about the value  $I(\omega)$

Another way to put this is that if  $I$  is relatively constant, then we can use a large amount of smoothing without introducing too much bias

**The AR(1) Setting** Let's examine this idea more carefully in a particular setting — where the data is assumed to be AR(1)

(More general ARMA settings can be handled using similar techniques to those described below)

Suppose in particular that  $\{X_t\}$  is covariance stationary and AR(1), with

$$X_{t+1} = \mu + \phi X_t + \epsilon_{t+1} \quad (3.104)$$

where  $\mu$  and  $\phi \in (-1, 1)$  are unknown parameters and  $\{\epsilon_t\}$  is white noise

It follows that if we regress  $X_{t+1}$  on  $X_t$  and an intercept, the residuals will approximate white noise

Let

- $g$  be the spectral density of  $\{\epsilon_t\}$  — a constant function, as discussed above
- $I_0$  be the periodogram estimated from the residuals — an estimate of  $g$
- $f$  be the spectral density of  $\{X_t\}$  — the object we are trying to estimate

In view of *an earlier result* we obtained while discussing ARMA processes,  $f$  and  $g$  are related by

$$f(\omega) = \left| \frac{1}{1 - \phi e^{i\omega}} \right|^2 g(\omega) \quad (3.105)$$

This suggests that the recoloring step, which constructs an estimate  $I$  of  $f$  from  $I_0$ , should set

$$I(\omega) = \left| \frac{1}{1 - \hat{\phi} e^{i\omega}} \right|^2 I_0(\omega)$$

where  $\hat{\phi}$  is the OLS estimate of  $\phi$

The code for `ar_periodogram()` — the third function in `estspec.jl` — does exactly this. (See the code [here](#))

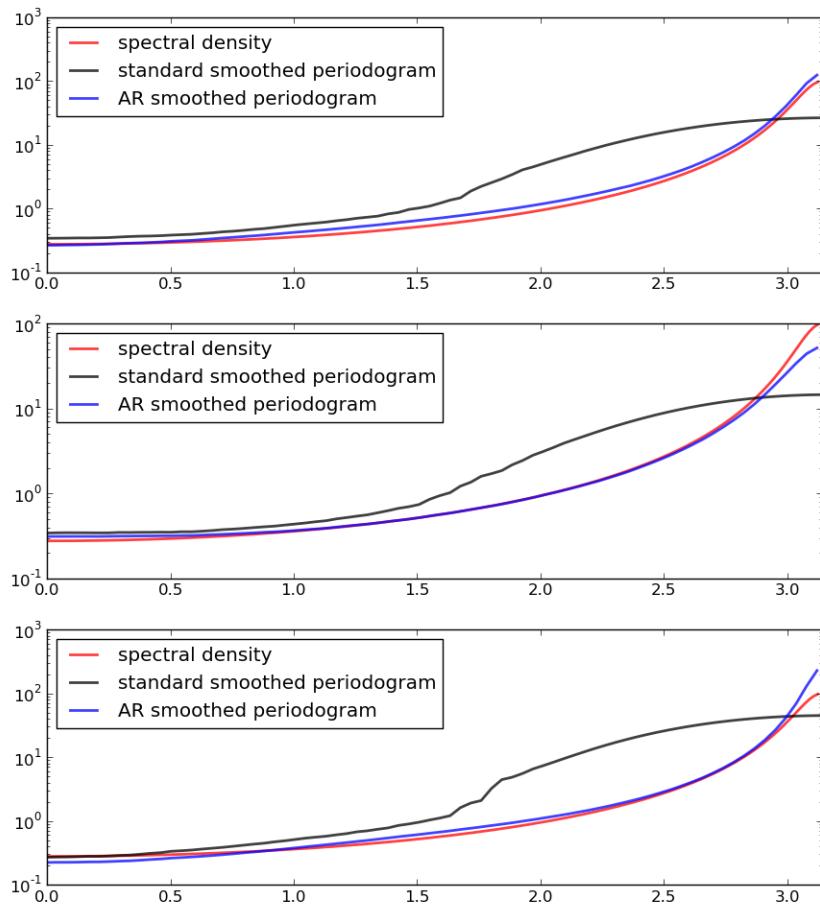
The next figure shows realizations of the two kinds of smoothed periodograms

1. “standard smoothed periodogram”, the ordinary smoothed periodogram, and
2. “AR smoothed periodogram”, the pre-whitened and recolored one generated by `ar_periodogram()`

The periodograms are calculated from time series drawn from (3.104) with  $\mu = 0$  and  $\phi = -0.9$

Each time series is of length 150

The difference between the three subfigures is just randomness — each one uses a different draw of the time series



In all cases, periodograms are fit with the “hamming” window and window length of 65

Overall, the fit of the AR smoothed periodogram is much better, in the sense of being closer to the true spectral density

### Exercises

**Exercise 1** Replicate *this figure* (modulo randomness)

The model is as in equation (3.102) and there are 400 observations

For the smoothed periodogram, the window type is “hamming”

**Exercise 2** Replicate *this figure* (modulo randomness)

The model is as in equation (3.104), with  $\mu = 0$ ,  $\phi = -0.9$  and 150 observations in each time series

All periodograms are fit with the “hamming” window and window length of 65

**Exercise 3** To be written. The exercise will be to use the code from *this lecture* to download FRED data and generate periodograms for different kinds of macroeconomic data.

## Solutions

[Solution notebook](#)

# Optimal Taxation in an LQ Economy

## Contents

- *Optimal Taxation in an LQ Economy*
  - *Overview*
  - *The Ramsey Problem*
  - *Implementation*
  - *Examples*
  - *Exercises*
  - *Solutions*

## Overview

In this lecture we study optimal fiscal policy in a linear quadratic setting

We slightly modify a well-known model of Robert Lucas and Nancy Stokey [LS83] so that convenient formulas for solving linear-quadratic models can be applied to simplify the calculations

The economy consists of a representative household and a benevolent government

The government finances an exogenous stream of government purchases with state-contingent loans and a linear tax on labor income

A linear tax is sometimes called a flat-rate tax

The household maximizes utility by choosing paths for consumption and labor, taking prices and the government’s tax rate and borrowing plans as given

Maximum attainable utility for the household depends on the government’s tax and borrowing plans

The *Ramsey problem* [Ram27] is to choose tax and borrowing plans that maximize the household's welfare, taking the household's optimizing behavior as given

There is a large number of competitive equilibria indexed by different government fiscal policies

The Ramsey planner chooses the best competitive equilibrium

We want to study the dynamics of tax rates, tax revenues, government debt under a Ramsey plan

Because the Lucas and Stokey model features state-contingent government debt, the government debt dynamics differ substantially from those in a model of Robert Barro [Bar79]

The treatment given here closely follows this manuscript, prepared by Thomas J. Sargent and Francois R. Velde

We cover only the key features of the problem in this lecture, leaving you to refer to that source for additional results and intuition

### Model Features

- Linear quadratic (LQ) model
- Representative household
- Stochastic dynamic programming over an infinite horizon
- Distortionary taxation

### The Ramsey Problem

We begin by outlining the key assumptions regarding technology, households and the government sector

**Technology** Labor can be converted one-for-one into a single, non-storable consumption good

In the usual spirit of the LQ model, the amount of labor supplied in each period is unrestricted

This is unrealistic, but helpful when it comes to solving the model

Realistic labor supply can be induced by suitable parameter values

**Households** Consider a representative household who chooses a path  $\{\ell_t, c_t\}$  for labor and consumption to maximize

$$-\mathbb{E} \frac{1}{2} \sum_{t=0}^{\infty} \beta^t [(c_t - b_t)^2 + \ell_t^2] \quad (3.106)$$

subject to the budget constraint

$$\mathbb{E} \sum_{t=0}^{\infty} \beta^t p_t^0 [d_t + (1 - \tau_t)\ell_t + s_t - c_t] = 0 \quad (3.107)$$

Here

- $\beta$  is a discount factor in  $(0, 1)$
- $p_t^0$  is a scaled Arrow-Debreu price at time 0 of history contingent goods at time  $t + j$
- $b_t$  is a stochastic preference parameter
- $d_t$  is an endowment process
- $\tau_t$  is a flat tax rate on labor income
- $s_t$  is a promised time- $t$  coupon payment on debt issued by the government

The scaled Arrow-Debreu price  $p_t^0$  is related to the unscaled Arrow-Debreu price as follows.

If we let  $\pi_t^0(x^t)$  denote the probability (density) of a history  $x^t = [x_t, x_{t-1}, \dots, x_0]$  of the state  $x^t$ , then the Arrow-Debreu time 0 price of a claim on one unit of consumption at date  $t$ , history  $x^t$  would be

$$\frac{\beta^t p_t^0}{\pi_t^0(x^t)}$$

Thus, our scaled Arrow-Debreu price is the ordinary Arrow-Debreu price multiplied by the discount factor  $\beta^t$  and divided by an appropriate probability.

The budget constraint (3.107) requires that the present value of consumption be restricted to equal the present value of endowments, labor income and coupon payments on bond holdings

**Government** The government imposes a linear tax on labor income, fully committing to a stochastic path of tax rates at time zero

The government also issues state-contingent debt

Given government tax and borrowing plans, we can construct a competitive equilibrium with distorting government taxes

Among all such competitive equilibria, the Ramsey plan is the one that maximizes the welfare of the representative consumer

**Exogenous Variables** Endowments, government expenditure, the preference shock process  $b_t$ , and promised coupon payments on initial government debt  $s_t$  are all exogenous, and given by

- $d_t = S_d x_t$
- $g_t = S_g x_t$
- $b_t = S_b x_t$
- $s_t = S_s x_t$

The matrices  $S_d, S_g, S_b, S_s$  are primitives and  $\{x_t\}$  is an exogenous stochastic process taking values in  $\mathbb{R}^k$

We consider two specifications for  $\{x_t\}$

1. Discrete case:  $\{x_t\}$  is a discrete state Markov chain with transition matrix  $P$
2. VAR case:  $\{x_t\}$  obeys  $x_{t+1} = Ax_t + Cw_{t+1}$  where  $\{w_t\}$  is independent zero mean Gaussian with identify covariance matrix

**Feasibility** The period-by-period feasibility restriction for this economy is

$$c_t + g_t = d_t + \ell_t \quad (3.108)$$

A labor-consumption process  $\{\ell_t, c_t\}$  is called *feasible* if (3.108) holds for all  $t$

**Government budget constraint** Where  $p_t^0$  is again a scaled Arrow-Debreu price, the time zero government budget constraint is

$$\mathbb{E} \sum_{t=0}^{\infty} \beta^t p_t^0 (s_t + g_t - \tau_t \ell_t) = 0 \quad (3.109)$$

**Equilibrium** An *equilibrium* is a feasible allocation  $\{\ell_t, c_t\}$ , a sequence of prices  $\{p_t^0\}$ , and a tax system  $\{\tau_t\}$  such that

1. The allocation  $\{\ell_t, c_t\}$  is optimal for the household given  $\{p_t^0\}$  and  $\{\tau_t\}$
2. The government's budget constraint (3.109) is satisfied

The *Ramsey problem* is to choose the equilibrium  $\{\ell_t, c_t, \tau_t, p_t^0\}$  that maximizes the household's welfare

If  $\{\ell_t, c_t, \tau_t, p_t^0\}$  solves the Ramsey problem, then  $\{\tau_t\}$  is called the *Ramsey plan*

The solution procedure we adopt is

1. Use the first-order conditions from the household problem to pin down prices and allocations given  $\{\tau_t\}$
2. Use these expressions to rewrite the government budget constraint (3.109) in terms of exogenous variables and allocations
3. Maximize the household's objective function (3.106) subject to the constraint constructed in step 2 and the feasibility constraint (3.108)

The solution to this maximization problem pins down all quantities of interest

**Solution** Step one is to obtain the first-conditions for the household's problem, taking taxes and prices as given

Letting  $\mu$  be the Lagrange multiplier on (3.107), the first-order conditions are  $p_t^0 = (c_t - b_t)/\mu$  and  $\ell_t = (c_t - b_t)(1 - \tau_t)$

Rearranging and normalizing at  $\mu = b_0 - c_0$ , we can write these conditions as

$$p_t^0 = \frac{b_t - c_t}{b_0 - c_0} \quad \text{and} \quad \tau_t = 1 - \frac{\ell_t}{b_t - c_t} \quad (3.110)$$

Substituting (3.110) into the government's budget constraint (3.109) yields

$$\mathbb{E} \sum_{t=0}^{\infty} \beta^t [(b_t - c_t)(s_t + g_t - \ell_t) + \ell_t^2] = 0 \quad (3.111)$$

The Ramsey problem now amounts to maximizing (3.106) subject to (3.111) and (3.108)

The associated Lagrangian is

$$\mathcal{L} = \mathbb{E} \sum_{t=0}^{\infty} \beta^t \left\{ -\frac{1}{2} [(c_t - b_t)^2 + \ell_t^2] + \lambda [(b_t - c_t)(\ell_t - s_t - g_t) - \ell_t^2] + \mu_t [d_t + \ell_t - c_t - g_t] \right\} \quad (3.112)$$

The first order conditions associated with  $c_t$  and  $\ell_t$  are

$$-(c_t - b_t) + \lambda[-\ell_t + (g_t + s_t)] = \mu_t$$

and

$$\ell_t - \lambda[(b_t - c_t) - 2\ell_t] = \mu_t$$

Combining these last two equalities with (3.108) and working through the algebra, one can show that

$$\ell_t = \bar{\ell}_t - \nu m_t \quad \text{and} \quad c_t = \bar{c}_t - \nu m_t \quad (3.113)$$

where

- $\nu := \lambda / (1 + 2\lambda)$
- $\bar{\ell}_t := (b_t - d_t + g_t) / 2$
- $\bar{c}_t := (b_t + d_t - g_t) / 2$
- $m_t := (b_t - d_t - s_t) / 2$

Apart from  $\nu$ , all of these quantities are expressed in terms of exogenous variables

To solve for  $\nu$ , we can use the government's budget constraint again

The term inside the brackets in (3.111) is  $(b_t - c_t)(s_t + g_t) - (b_t - c_t)\ell_t + \ell_t^2$

Using (3.113), the definitions above and the fact that  $\bar{\ell} = b - \bar{c}$ , this term can be rewritten as

$$(b_t - \bar{c}_t)(g_t + s_t) + 2m_t^2(\nu^2 - \nu)$$

Reinserting into (3.111), we get

$$\mathbb{E} \left\{ \sum_{t=0}^{\infty} \beta^t (b_t - \bar{c}_t)(g_t + s_t) \right\} + (\nu^2 - \nu) \mathbb{E} \left\{ \sum_{t=0}^{\infty} \beta^t 2m_t^2 \right\} = 0 \quad (3.114)$$

Although it might not be clear yet, we are nearly there because:

- The two expectations terms in (3.114) can be solved for in terms of model primitives
- This in turn allows us to solve for the Lagrange multiplier  $\nu$
- With  $\nu$  in hand, we can go back and solve for the allocations via (3.113)
- Once we have the allocations, prices and the tax system can be derived from (3.110)

**Computing the Quadratic Term** Let's consider how to obtain the term  $v$  in (3.114)

If we can compute the two expected geometric sums

$$b_0 := \mathbb{E} \left\{ \sum_{t=0}^{\infty} \beta^t (b_t - \bar{c}_t)(g_t + s_t) \right\} \quad \text{and} \quad a_0 := \mathbb{E} \left\{ \sum_{t=0}^{\infty} \beta^t 2m_t^2 \right\} \quad (3.115)$$

then the problem reduces to solving

$$b_0 + a_0(v^2 - v) = 0$$

for  $v$

Provided that  $4b_0 < a_0$ , there is a unique solution  $v \in (0, 1/2)$ , and a unique corresponding  $\lambda > 0$

Let's work out how to compute mathematical expectations in (3.115)

For the first one, the random variable  $(b_t - \bar{c}_t)(g_t + s_t)$  inside the summation can be expressed as

$$\frac{1}{2} x_t' (S_b - S_d + S_g)' (S_g + S_s) x_t$$

For the second expectation in (3.115), the random variable  $2m_t^2$  can be written as

$$\frac{1}{2} x_t' (S_b - S_d - S_s)' (S_b - S_d - S_s) x_t$$

It follows that both objects of interest are special cases of the expression

$$q(x_0) = \mathbb{E} \sum_{t=0}^{\infty} \beta^t x_t' H x_t \quad (3.116)$$

where  $H$  is a matrix conformable to  $x_t$  and  $x_t'$  is the transpose of column vector  $x_t$

Suppose first that  $\{x_t\}$  is the Gaussian VAR described *above*

In this case, the formula for computing  $q(x_0)$  is known to be  $q(x_0) = x_0' Q x_0 + v$ , where

- $Q$  is the solution to  $Q = H + \beta A' Q A$ , and
- $v = \text{trace}(C' Q C) \beta / (1 - \beta)$

The first equation is known as a discrete Lyapunov equation, and can be solved using [this function](#)

**Finite state Markov case** Next suppose that  $\{x_t\}$  is the discrete Markov process described *above*

Suppose further that each  $x_t$  takes values in the state space  $\{x^1, \dots, x^N\} \subset \mathbb{R}^k$

Let  $h: \mathbb{R}^k \rightarrow \mathbb{R}$  be a given function, and suppose that we wish to evaluate

$$q(x_0) = \mathbb{E} \sum_{t=0}^{\infty} \beta^t h(x_t) \quad \text{given} \quad x_0 = x^j$$

For example, in the discussion above,  $h(x_t) = x_t' H x_t$

It is legitimate to pass the expectation through the sum, leading to

$$q(x_0) = \sum_{t=0}^{\infty} \beta^t (P^t h)[j] \quad (3.117)$$

Here

- $P^t$  is the  $t$ -th power of the transition matrix  $P$
- $h$  is, with some abuse of notation, the vector  $(h(x^1), \dots, h(x^N))$
- $(P^t h)[j]$  indicates the  $j$ -th element of  $P^t h$

It can be shown that (3.117) is in fact equal to the  $j$ -th element of the vector  $(I - \beta P)^{-1} h$

This last fact is applied in the calculations below

**Other Variables** We are interested in tracking several other variables besides the ones described above.

To prepare the way for this, we define

$$p_{t+j}^t = \frac{b_{t+j} - c_{t+j}}{b_t - c_t}$$

as the scaled Arrow-Debreu time  $t$  price of a history contingent claim on one unit of consumption at time  $t + j$

These are prices that would prevail at time  $t$  if market were reopened at time  $t$

These prices are constituents of the present value of government obligations outstanding at time  $t$ , which can be expressed as

$$B_t := \mathbb{E}_t \sum_{j=0}^{\infty} \beta^j p_{t+j}^t (\tau_{t+j} \ell_{t+j} - g_{t+j}) \quad (3.118)$$

Using our expression for prices and the Ramsey plan, we can also write  $B_t$  as

$$B_t = \mathbb{E}_t \sum_{j=0}^{\infty} \beta^j \frac{(b_{t+j} - c_{t+j})(\ell_{t+j} - g_{t+j}) - \ell_{t+j}^2}{b_t - c_t}$$

This version is more convenient for computation

Using the equation

$$p_{t+j}^t = p_{t+1}^t p_{t+j}^{t+1}$$

it is possible to verify that (3.118) implies that

$$B_t = (\tau_t \ell_t - g_t) + E_t \sum_{j=1}^{\infty} p_{t+j}^t (\tau_{t+j} \ell_{t+j} - g_{t+j})$$

and

$$B_t = (\tau_t \ell_t - g_t) + \beta E_t p_{t+1}^t B_{t+1} \quad (3.119)$$

Define

$$R_t^{-1} := \mathbb{E}_t \beta^j p_{t+1}^t \quad (3.120)$$

$R_t$  is the gross 1-period risk-free rate for loans between  $t$  and  $t + 1$

**A martingale** We now want to study the following two objects, namely,

$$\pi_{t+1} := B_{t+1} - R_t[B_t - (\tau_t \ell_t - g_t)]$$

and the cumulation of  $\pi_t$

$$\Pi_t := \sum_{s=0}^t \pi_s$$

The term  $\pi_{t+1}$  is the difference between two quantities:

- $B_{t+1}$ , the value of government debt at the start of period  $t + 1$
- $R_t[B_t + g_t - \tau_t]$ , which is what the government would have owed at the beginning of period  $t + 1$  if it had simply borrowed at the one-period risk-free rate rather than selling state-contingent securities.

Thus,  $\pi_{t+1}$  is the excess payout on the actual portfolio of state contingent government debt relative to an alternative portfolio sufficient to finance  $B_t + g_t - \tau_t \ell_t$  and consisting entirely of risk-free one-period bonds.

Use expressions (3.119) and (3.120) to obtain

$$\pi_{t+1} = B_{t+1} - \frac{1}{\beta E_t p_{t+1}^t} [\beta E_t p_{t+1}^t B_{t+1}]$$

or

$$\pi_{t+1} = B_{t+1} - \tilde{E}_t B_{t+1} \quad (3.121)$$

where  $\tilde{E}_t$  is the conditional mathematical expectation taken with respect to a one-step transition density that has been formed by multiplying the original transition density with the likelihood ratio

$$m_{t+1}^t = \frac{p_{t+1}^t}{E_t p_{t+1}^t}$$

It follows from equation (3.121) that

$$\tilde{E}_t \pi_{t+1} = \tilde{E}_t B_{t+1} - \tilde{E}_t B_{t+1} = 0$$

which asserts that  $\{\pi_{t+1}\}$  is a martingale difference sequence under the distorted probability measure, and that  $\{\Pi_t\}$  is a martingale under the distorted probability measure.

In the tax-smoothing model of Robert Barro [Bar79], government debt is a random walk. In the current model, government debt  $\{B_t\}$  is not a random walk, but the “excess payoff”  $\{\Pi_t\}$  on it is.

### Implementation

The following code provides functions for

1. Solving for the Ramsey plan given a specification of the economy
2. Simulating the dynamics of the major variables

The file is `lqramsey/lqramsey.jl` from the [applications repository](#)

Description and clarifications are given below

```
#=
```

*This module provides code to compute Ramsey equilibria in a LQ economy with distortionary taxation. The program computes allocations (consumption, leisure), tax rates, revenues, the net present value of the debt and other related quantities.*

*Functions for plotting the results are also provided below.*

*@author : Spencer Lyon <spencer.lyon@nyu.edu>*

*@date: 2014-08-21*

*References*

-----

*Simple port of the file examples/lqramsey.py*

*http://quant-econ.net/lqramsey.html*

```
=#
using QuantEcon
using PlotlyJS

abstract AbstractStochProcess

type ContStochProcess <: AbstractStochProcess
 A::Matrix
 C::Matrix
end

type DiscreteStochProcess <: AbstractStochProcess
 P::Matrix
 x_vals::Array
end

type Economy{SP <: AbstractStochProcess}
 bet::Real
 Sg::Matrix
 Sd::Matrix
 Sb::Matrix
 Ss::Matrix
 is_discrete::Bool
 proc::SP
end

type Path
 g
 d
```

```

b
s
c
l
p
tau
rvn
B
R
pi
Pi
xi
end

function compute_exog_sequences(econ::Economy, x)
 # Compute exogenous variable sequences
 Sg, Sd, Sb, Ss = econ.Sg, econ.Sd, econ.Sb, econ.Ss
 g, d, b, s = [squeeze(S * x, 1) for S in (Sg, Sd, Sb, Ss)]

 #= Solve for Lagrange multiplier in the govt budget constraint
 In fact we solve for nu = lambda / (1 + 2*lambda). Here nu is the
 solution to a quadratic equation a(nu^2 - nu) + b = 0 where
 a and b are expected discounted sums of quadratic forms of the state. =#
 Sm = Sb - Sd - Ss

 return g, d, b, s, Sm
end

function compute_allocation(econ::Economy, Sm, nu, x, b)
 Sg, Sd, Sb, Ss = econ.Sg, econ.Sd, econ.Sb, econ.Ss

 # Solve for the allocation given nu and x
 Sc = 0.5 .* (Sb + Sd - Sg - nu .* Sm)
 Sl = 0.5 .* (Sb - Sd + Sg - nu .* Sm)
 c = squeeze(Sc * x, 1)
 l = squeeze(Sl * x, 1)
 p = squeeze((Sb - Sc) * x, 1) # Price without normalization
 tau = 1 .- l ./ (b .- c)
 rvn = l .* tau

 return Sc, Sl, c, l, p, tau, rvn
end

function compute_nu(a0, b0)
 disc = a0^2 - 4a0*b0

 if disc >= 0
 nu = 0.5 * (a0 - sqrt(disc)) / a0
 else
 println("There is no Ramsey equilibrium for these parameters.")
 end

```

```

 error("Government spending (economy.g) too low")
 end

 # Test that the Lagrange multiplier has the right sign
 if nu * (0.5 - nu) < 0
 print("Negative multiplier on the government budget constraint.")
 error("Government spending (economy.g) too low")
 end

 return nu
end

function compute_Pi(B, R, rvn, g, xi)
 pi = B[2:end] - R[1:end-1] .* B[1:end-1] - rvn[1:end-1] + g[1:end-1]
 Pi = cumsum(pi .* xi)
 return pi, Pi
end

function compute_paths(econ::Economy{DiscreteStochProcess}, T)
 # simplify notation
 bet, Sg, Sd, Sb, Ss = econ.bet, econ.Sg, econ.Sd, econ.Sb, econ.Ss
 P, x_vals = econ.proc.P, econ.proc.x_vals

 state = mc_sample_path(P, 1, T)
 x = x_vals[:, state]

 # Compute exogenous sequence
 g, d, b, s, Sm = compute_exog_sequences(econ, x)

 # compute a0, b0
 ns = size(P, 1)
 F = eye(ns) - bet.*P
 a0 = (F \ ((Sm * x_vals)'.^2))[1] ./ 2
 H = ((Sb - Sd + Sg) * x_vals) .* ((Sg - Ss)*x_vals)
 b0 = (F \ H')[1] ./ 2

 # compute lagrange multiplier
 nu = compute_nu(a0, b0)

 # Solve for the allocation given nu and x
 Sc, Sl, c, l, p, tau, rvn = compute_allocation(econ, Sm, nu, x, b)

 # compute remaining variables
 H = ((Sb - Sc)*x_vals) .* ((Sl - Sg)*x_vals) - (Sl*x_vals).^2
 temp = squeeze(F*H', 2)
 B = temp[state] ./ p
 H = squeeze(P[state, :] * ((Sb - Sc)*x_vals)', 2)
 R = p ./ (bet .* H)
 temp = squeeze(P[state, :] *((Sb - Sc) * x_vals)', 2)
 xi = p[2:end] ./ temp[1:end-1]

```

```

compute pi
pi, Pi = compute_Pi(B, R, rvn, g, xi)

Path(g, d, b, s, c, l, p, tau, rvn, B, R, pi, Pi, xi)
end

function compute_paths(econ::Economy{ContStochProcess}, T)
 # simplify notation
 bet, Sg, Sd, Sb, Ss = econ.bet, econ.Sg, econ.Sd, econ.Sb, econ.Ss
 A, C = econ.proc.A, econ.proc.C

 # Generate an initial condition x0 satisfying x0 = A x0
 nx, nx = size(A)
 x0 = nullspace((eye(nx) - A))
 x0 = x0[end] < 0 ? -x0 : x0
 x0 = x0 ./ x0[end]
 x0 = squeeze(x0, 2)

 # Generate a time series x of length T starting from x0
 nx, nw = size(C)
 x = zeros(nx, T)
 w = randn(nw, T)
 x[:, 1] = x0
 for t=2:T
 x[:, t] = A *x[:, t-1] + C * w[:, t]
 end

 # Compute exogenous sequence
 g, d, b, s, Sm = compute_exog_sequences(econ, x)

 # compute a0 and b0
 H = Sm' Sm
 a0 = 0.5 * var_quadratic_sum(A, C, H, bet, x0)
 H = (Sb - Sd + Sg)*(Sg + Ss)
 b0 = 0.5 * var_quadratic_sum(A, C, H, bet, x0)

 # compute lagrange multiplier
 nu = compute_nu(a0, b0)

 # Solve for the allocation given nu and x
 Sc, Sl, c, l, p, tau, rvn = compute_allocation(econ, Sm, nu, x, b)

 # compute remaining variables
 H = Sl' Sl - (Sb - Sc)'*(Sl - Sg)
 L = Array(Float64, T)
 for t=1:T
 L[t] = var_quadratic_sum(A, C, H, bet, x[:, t])
 end
 B = L ./ p
 Rinv = squeeze(bet .* (Sb - Sc)*A*x, 1) ./ p
 R = 1 ./ Rinv
 AF1 = (Sb - Sc) * x[:, 2:end]

```

```

AF2 = (Sb - Sc) * A * x[:, 1:end-1]
xi = AF1 ./ AF2
xi = squeeze(xi, 1)

compute pi
pi, Pi = compute_Pi(B, R, rvn, g, xi)

Path(g, d, b, s, c, l, p, tau, rvn, B, R, pi, Pi, xi)
end

function gen_fig_1(path::Path)
T = length(path.c)

tr1, tr2, tr4 = GenericTrace[], GenericTrace[], GenericTrace[]

Plot consumption, govt expenditure and revenue
push!(tr1, scatter(; y = path.rvn, legendgroup = "rev", marker_color = "blue", name = L"\tau_t \ell_t"))
push!(tr1, scatter(; y = path.g, legendgroup = "gov", marker_color = "red", name = L"g_t"))
push!(tr1, scatter(; y = path.c, marker_color = "green", name = L"c_t"))

Plot govt expenditure and debt
push!(tr2, scatter(; x = 1:T, y = path.rvn, legendgroup = "rev", marker_color = "blue", showlegend = false))
push!(tr2, scatter(; x = 1:T, y = path.g, legendgroup = "gov", marker_color = "red", showlegend = false))
push!(tr2, scatter(; x = 1:T-1, y = path.B[2:end], marker_color = "orange", name = L"B_{t+1}"))

Plot risk free return
tr3 = scatter(; x = 1:T, y = path.R - 1, marker_color = "pink", name = L"R_{t-1}")

Plot revenue, expenditure and risk free rate
push!(tr4, scatter(; x = 1:T, y = path.rvn, legendgroup = "rev", marker_color = "blue", showlegend = false))
push!(tr4, scatter(; x = 1:T, y = path.g, legendgroup = "gov", marker_color = "red", showlegend = false))
push!(tr4, scatter(; x = 1:T-1, y = path.pi, marker_color = "violet", name = L"\pi_{t+1}"))

p1 = plot(tr1)
p2 = plot(tr2)
p3 = plot(tr3, Layout(; xaxis_title = "Time"))
p4 = plot(tr4, Layout(; xaxis_title = "Time"))
p = [p1 p2; p3 p4]
relayout!(p, height = 900)

p
end

function gen_fig_2(path::Path)
T = length(path.c)

Plot adjustment factor
p1 = plot(scatter(; x = 2:T, y = path.xi, name = L"\xi_t"))

Plot adjusted cumulative return
p2 = plot(scatter(; x = 2:T, y = path.Pi, name = L"\Pi_t"), Layout(; xaxis_title = "Time"))

```

```

p = [p1; p2]
relayout!(p, height = 600)

p
end

```

**Comments on the Code** The function `var_quadratic_sum` from `QuantEcon.jl` is for computing the value of (3.116) when the exogenous process  $\{x_t\}$  is of the VAR type described *above*

This code defines two Types: `Economy` and `Path`

The first is used to collect all the parameters and primitives of a given LQ economy, while the second collects output of the computations

### Examples

Let's look at two examples of usage

**The Continuous Case** Our first example adopts the VAR specification described *above*

Regarding the primitives, we set

- $\beta = 1/1.05$
- $b_t = 2.135$  and  $s_t = d_t = 0$  for all  $t$

Government spending evolves according to

$$g_{t+1} - \mu_g = \rho(g_t - \mu_g) + C_g w_{g,t+1}$$

with  $\rho = 0.7$ ,  $\mu_g = 0.35$  and  $C_g = \mu_g \sqrt{1 - \rho^2}/10$

Here's the code, from file `lqramsey/lqramsey_ar1.jl`

```

#=

Example 1: Govt spending is AR(1) and state is (g, 1).

@author : Spencer Lyon <spencer.lyon@nyu.edu>

@date: 2014-08-21

References

Simple port of the file examples/lqramsey_ar1.py

http://quant-econ.net/lqramsey.html

=#
include("lqramsey.jl")

```

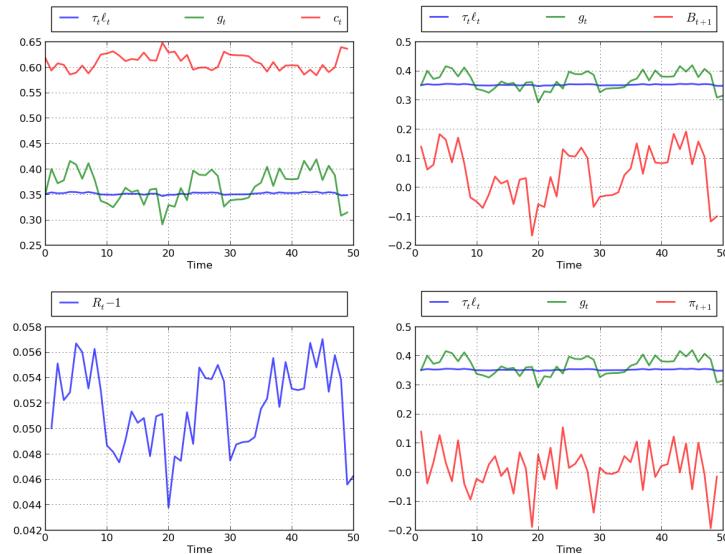
```
== Parameters ==
bet = 1 / 1.05
rho, mg = .7, .35
A = eye(2)
A = [rho mg*(1 - rho); 0.0 1.0]
C = [sqrt(1 - rho^2)*mg/10 0.0] '
Sg = [1.0 0.0]
Sd = [0.0 0.0]
Sb = [0 2.135]
Ss = [0.0 0.0]
discrete = false
proc = ContStochProcess(A, C)

econ = Economy(bet, Sg, Sd, Sb, Ss, discrete, proc)
T = 50

path = compute_paths(econ, T)

gen_fig_1(path)
```

Running the program produces the figure

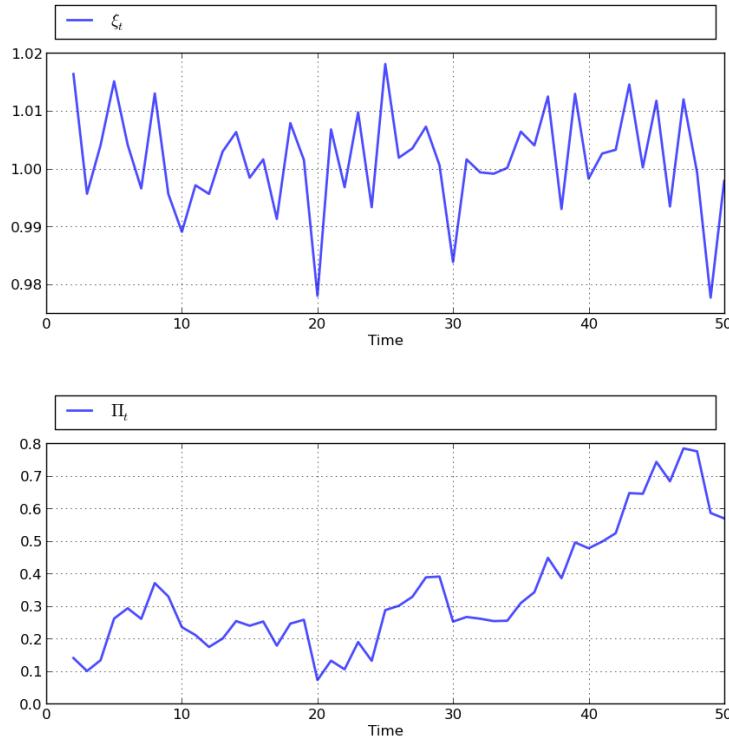


The legends on the figures indicate the variables being tracked

Most obvious from the figure is tax smoothing in the sense that tax revenue is much less variable than government expenditure

After running the code above, if you then execute `gen_fig_2(path)` from your Julia console you will produce the figure

See the original manuscript for comments and interpretation



**The Discrete Case** Our second example adopts a discrete Markov specification for the exogenous process

Here's the code, from file `lqramsey/lqramsey_discrete.jl`

```
#=
Example 2: LQ Ramsey model with discrete exogenous process.

@author : Spencer Lyon <spencer.lyon@nyu.edu>
@date: 2014-08-21

References

Simple port of the file examples/lqramsey_discrete.py

http://quant-econ.net/lqramsey.html

=#
include("lqramsey.jl")

Parameters
bet = 1 / 1.05
P = [0.8 0.2 0.0
 0.0 0.5 0.5
 0.0 0.0 1.0]
```

```

Possible states of the world
Each column is a state of the world. The rows are [g d b s 1]
x_vals = [0.5 0.5 0.25
 0.0 0.0 0.0
 2.2 2.2 2.2
 0.0 0.0 0.0
 1.0 1.0 1.0]
Sg = [1.0 0.0 0.0 0.0 0.0]
Sd = [0.0 1.0 0.0 0.0 0.0]
Sb = [0.0 0.0 1.0 0.0 0.0]
Ss = [0.0 0.0 0.0 1.0 0.0]
discrete = true
proc = DiscreteStochProcess(P, x_vals)

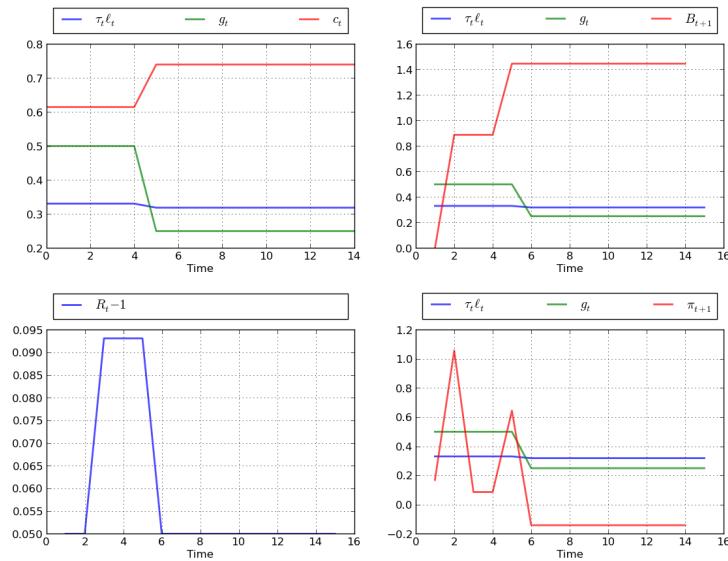
econ = Economy(bet, Sg, Sd, Sb, Ss, discrete, proc)
T = 15

path = compute_paths(econ, T)

gen_fig_1(path)

```

The call `gen_fig_1(path)` generates the figure



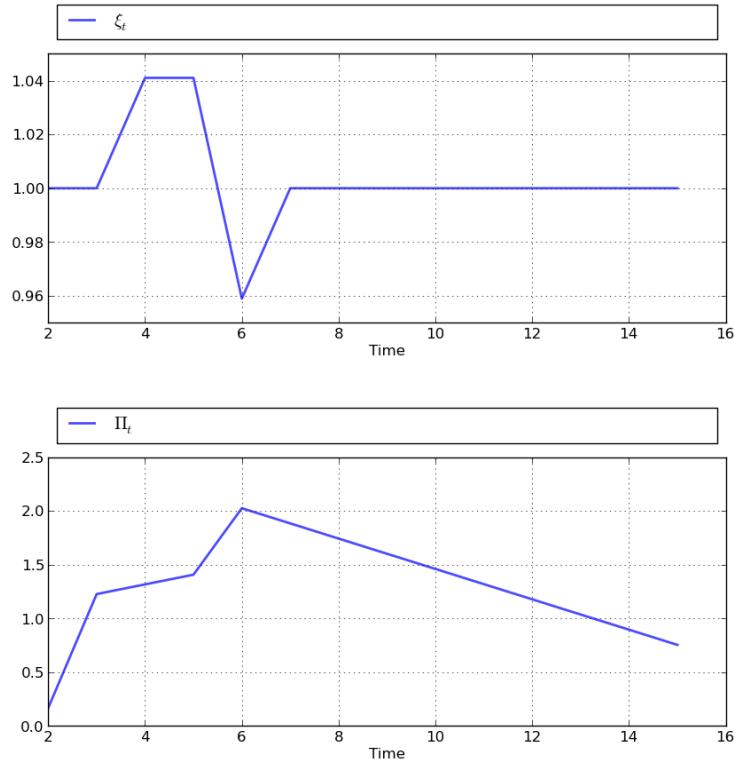
while `gen_fig_2(path)` generates

See the original manuscript for comments and interpretation

### Exercises

**Exercise 1** Modify the VAR example given above, setting

$$g_{t+1} - \mu_g = \rho(g_{t-3} - \mu_g) + C_g w_{g,t+1}$$



with  $\rho = 0.95$  and  $C_g = 0.7\sqrt{1 - \rho^2}$

Produce the corresponding figures

### Solutions

[Solution notebook](#)

## History Dependent Public Policies

## Contents

- *History Dependent Public Policies*
  - *Overview*
  - *Two Sources of History Dependence*
  - *Competitive equilibrium*
  - *Ramsey Problem*
  - *Two Subproblems*
  - *Time Inconsistency*
  - *Credible Policy*
  - *Concluding remarks*

### Overview

This lecture describes history-dependent public policies and some of their representations

History dependent policies are decision rules that depend on the entire past history of the state variables

History dependent policies naturally emerge in [Ramsey problems](#)

A Ramsey planner (typically interpreted as a government) devises a plan of actions at time  $t = 0$  to follow at all future dates and for all contingencies

In order to make a plan, he takes as given Euler equations expressing private agents' first-order necessary conditions

He also takes into account that his *future* actions affect earlier decisions by private agents, an avenue opened up by the maintained assumption of *rational expectations*

Another setting in which history dependent policies naturally emerge is where instead of a Ramsey planner there is a *sequence* of government administrators whose time  $t$  member takes as given the policies used by its successors

We study these ideas in the context of a model in which a benevolent tax authority is forced

- to raise a prescribed present value of revenues
- to do so by imposing a distorting flat rate tax on the output of a competitive representative firm

The firm faces costs of adjustment and lives within a competitive equilibrium, which in turn imposes restrictions on the tax authority<sup>1</sup>

**References** The presentation below is based on a recent paper by Evans and Sargent [[ES13](#)]

Regarding techniques, we will make use of the methods described in

1. the [linear regulator lecture](#)
2. the [solving LQ dynamic Stackelberg problems lecture](#)

<sup>1</sup> We could also call a competitive equilibrium a rational expectations equilibrium.

## Two Sources of History Dependence

We compare two timing protocols

1. An infinitely lived benevolent tax authority solves a Ramsey problem
2. There is a sequence of tax authorities, each choosing only a time  $t$  tax rate

Under both timing protocols, optimal tax policies are *history-dependent*

But history dependence captures different economic forces across the two timing protocols

In the first timing protocol, history dependence expresses the *time-inconsistency of the Ramsey plan*

In the second timing protocol, history dependence reflects the unfolding of constraints that assure that a time  $t$  government administrator wants to confirm the representative firm's expectations about government actions

We describe recursive representations of history-dependent tax policies under both timing protocols

**Ramsey Timing Protocol** The first timing protocol models a policy maker who can be said to 'commit', choosing a sequence of tax rates once-and-for-all at time 0

**Sequence of Governments Timing Protocol** For the second timing protocol we use the notion of a *sustainable plan* proposed in [CK90], also referred to as a *credible public policy* in [Sto89]

A key idea here is that history-dependent policies can be arranged so that, when regarded as a representative firm's forecasting functions, they confront policy makers with incentives to confirm them

We follow Chang [Cha98] in expressing such history-dependent plans recursively

Credibility considerations contribute an additional auxiliary state variable in the form of a promised value to the planner

It expresses how decisions must unfold to give the government the incentive to confirm private sector expectations when the government chooses sequentially

---

**Note:** We occasionally hear confusion about the consequences of recursive representations of government policies under our two timing protocols. It is incorrect to regard a recursive representation of the Ramsey plan as in any way 'solving a time-inconsistency problem'. On the contrary, the evolution of the auxiliary state variable that augments the authentic ones under our first timing protocol ought to be viewed as *expressing* the time-inconsistency of a Ramsey plan. Despite that, in literatures about practical monetary policy one sometimes hears interpretations that sell Ramsey plans in settings where our sequential timing protocol is the one that more accurately characterizes decision making. Please beware of discussions that toss around claims about credibility if you don't also see recursive representations of policies with the complete list of state variables appearing in our [Cha98]-like analysis that we present *below*.

---

### Competitive equilibrium

A representative competitive firm sells output  $q_t$  at price  $p_t$  when market-wide output is  $Q_t$

The market as a whole faces a downward sloping inverse demand function

$$p_t = A_0 - A_1 Q_t, \quad A_0 > 0, A_1 > 0 \quad (3.122)$$

The representative firm

- has given initial condition  $q_0$
- endures quadratic adjustment costs  $\frac{d}{2}(q_{t+1} - q_t)^2$
- pays a flat rate tax  $\tau_t$  per unit of output
- treats  $\{p_t, \tau_t\}_{t=0}^{\infty}$  as exogenous
- chooses  $\{q_{t+1}\}_{t=0}^{\infty}$  to maximize

$$\sum_{t=0}^{\infty} \beta^t \left\{ p_t q_t - \frac{d}{2} (q_{t+1} - q_t)^2 - \tau_t q_t \right\} \quad (3.123)$$

Let  $u_t := q_{t+1} - q_t$  be the firm's 'control variable' at time  $t$

First-order conditions for the representative firm's problem are

$$u_t = \frac{\beta}{d} p_{t+1} + \beta u_{t+1} - \frac{\beta}{d} \tau_{t+1}, \quad t = 0, 1, \dots \quad (3.124)$$

To compute a competitive equilibrium, it is appropriate to take (3.124), eliminate  $p_t$  in favor of  $Q_t$  by using (3.122), and then set  $q_t = Q_t$

This last step *makes the representative firm be representative*<sup>2</sup>

We arrive at

$$u_t = \frac{\beta}{d} (A_0 - A_1 Q_{t+1}) + \beta u_{t+1} - \frac{\beta}{d} \tau_{t+1} \quad (3.125)$$

$$Q_{t+1} = Q_t + u_t \quad (3.126)$$

**Notation:** For any scalar  $x_t$ , let  $\vec{x} = \{x_t\}_{t=0}^{\infty}$

Given a tax sequence  $\{\tau_{t+1}\}_{t=0}^{\infty}$ , a **competitive equilibrium** is a price sequence  $\vec{p}$  and an output sequence  $\vec{Q}$  that satisfy (3.122), (3.125), and (3.126)

For any sequence  $\vec{x} = \{x_t\}_{t=0}^{\infty}$ , the sequence  $\vec{x}_1 := \{x_t\}_{t=1}^{\infty}$  is called the **continuation sequence** or simply the **continuation**

Note that a competitive equilibrium consists of a first period value  $u_0 = Q_1 - Q_0$  and a continuation competitive equilibrium with initial condition  $Q_1$

Also, a continuation of a competitive equilibrium is a competitive equilibrium

Following the lead of [Cha98], we shall make extensive use of the following property:

---

<sup>2</sup> It is important not to set  $q_t = Q_t$  prematurely. To make the firm a price taker, this equality should be imposed *after* and not *before* solving the firm's optimization problem.

- A continuation  $\vec{\tau}_1 = \{\tau_t\}_{t=1}^\infty$  of a tax policy  $\vec{\tau}$  influences  $u_0$  via (3.125) entirely through its impact on  $u_1$

A continuation competitive equilibrium can be indexed by a  $u_1$  that satisfies (3.125)

In the spirit of [KP80], we shall use  $u_{t+1}$  to describe what we shall call a **promised marginal value** that a competitive equilibrium offers to a representative firm<sup>3</sup>

Define  $Q^t := [Q_0, \dots, Q_t]$

A **history-dependent tax policy** is a sequence of functions  $\{\sigma_t\}_{t=0}^\infty$  with  $\sigma_t$  mapping  $Q^t$  into a choice of  $\tau_{t+1}$

Below, we shall

- Study history-dependent tax policies that either solve a Ramsey plan or are credible
- Describe recursive representations of both types of history-dependent policies

### Ramsey Problem

The planner's objective is cast in terms of consumer surplus net of the firm's adjustment costs

Consumer surplus is

$$\int_0^Q (A_0 - A_1 x) dx = A_0 Q - \frac{A_1}{2} Q^2$$

Hence the planner's one-period return function is

$$A_0 Q_t - \frac{A_1}{2} Q_t^2 - \frac{d}{2} u_t^2 \quad (3.127)$$

At time  $t = 0$ , a Ramsey planner faces the intertemporal budget constraint

$$\sum_{t=1}^{\infty} \beta^t \tau_t Q_t = G_0 \quad (3.128)$$

Note that (3.128) forbids taxation of initial output  $Q_0$

The **Ramsey problem** is to choose a tax sequence  $\vec{\tau}_1$  and a competitive equilibrium outcome  $(\vec{Q}, \vec{u})$  that maximize

$$\sum_{t=0}^{\infty} \beta^t \left\{ A_0 Q_t - \frac{A_1}{2} Q_t^2 - \frac{d}{2} u_t^2 \right\} \quad (3.129)$$

subject to (3.128)

Thus, the Ramsey timing protocol is:

1. At time 0, knowing  $(Q_0, G_0)$ , the Ramsey planner chooses  $\{\tau_{t+1}\}_{t=0}^\infty$
2. Given  $(Q_0, \{\tau_{t+1}\}_{t=0}^\infty)$ , a competitive equilibrium outcome  $\{u_t, Q_{t+1}\}_{t=0}^\infty$  emerges

<sup>3</sup> We could instead, perhaps with more accuracy, define a promised marginal value as  $\beta(A_0 - A_1 Q_{t+1}) - \beta \tau_{t+1} + u_{t+1}/\beta$ , since this is the object to which the firm's first-order condition instructs it to equate to the marginal cost  $d u_t$  of  $u_t = q_{t+1} - q_t$ . This choice would align better with how Chang [Cha98] chose to express his competitive equilibrium recursively. But given  $(u_t, Q_t)$ , the representative firm knows  $(Q_{t+1}, \tau_{t+1})$ , so it is adequate to take  $u_{t+1}$  as the intermediate variable that summarizes how  $\vec{\tau}_{t+1}$  affects the firm's choice of  $u_t$ .

**Note:** In bringing out the timing protocol associated with a Ramsey plan, we run head on into a set of issues analyzed by Bassetto [Bas05]. This is because our definition of the Ramsey timing protocol doesn't completely describe all conceivable actions by the government and firms as time unfolds. For example, the definition is silent about how the government would respond if firms, for some unspecified reason, were to choose to deviate from the competitive equilibrium associated with the Ramsey plan, possibly prompting violation of government budget balance. This is an example of the issues raised by [Bas05], who identifies a class of government policy problems whose proper formulation requires supplying a complete and coherent description of all actors' behavior across all possible histories. Implicitly, we are assuming that a more complete description of a government strategy could be specified that (a) agrees with ours along the Ramsey outcome, and (b) suffices uniquely to implement the Ramsey plan by deterring firms from taking actions that deviate from the Ramsey outcome path.

**Computing a Ramsey Plan** The planner chooses  $\{u_t\}_{t=0}^{\infty}, \{\tau_t\}_{t=1}^{\infty}$  to maximize (3.129) subject to (3.125), (3.126), and (3.128)

To formulate this problem as a Lagrangian, attach a Lagrange multiplier  $\mu$  to the budget constraint (3.128)

Then the planner chooses  $\{u_t\}_{t=0}^{\infty}, \{\tau_t\}_{t=1}^{\infty}$  to maximize and the Lagrange multiplier  $\mu$  to minimize

$$\sum_{t=0}^{\infty} \beta^t (A_0 Q_t - \frac{A_1}{2} Q_t^2 - \frac{d}{2} u_t^2) + \mu \left[ \sum_{t=0}^{\infty} \beta^t \tau_t Q_t - G_0 - \tau_0 Q_0 \right] \quad (3.130)$$

subject to and (3.125) and (3.126)

The Ramsey problem is a special case of the linear quadratic dynamic Stackelberg problem analyzed in this lecture

The key implementability conditions are (3.125) for  $t \geq 0$

Holding fixed  $\mu$  and  $G_0$ , the Lagrangian for the planning problem can be abbreviated as

$$\max_{\{u_t, \tau_{t+1}\}} \sum_{t=0}^{\infty} \beta^t \left\{ A_0 Q_t - \frac{A_1}{2} Q_t^2 - \frac{d}{2} u_t^2 + \mu \tau_t Q_t \right\}$$

Define

$$z_t := \begin{bmatrix} 1 \\ Q_t \\ \tau_t \end{bmatrix} \quad \text{and} \quad y_t := \begin{bmatrix} z_t \\ u_t \end{bmatrix} = \begin{bmatrix} 1 \\ Q_t \\ \tau_t \\ u_t \end{bmatrix}$$

Here the elements of  $z_t$  are natural state variables and  $u_t$  is a forward looking variable that we treat as a state variable for  $t \geq 1$

But  $u_0$  is a choice variable for the Ramsey planner.

We include  $\tau_t$  as a state variable for bookkeeping purposes: it helps to map the problem into a linear regulator problem with no cross products between states and controls

However, it will be a redundant state variable in the sense that the optimal tax  $\tau_{t+1}$  will not depend on  $\tau_t$

The government chooses  $\tau_{t+1}$  at time  $t$  as a function of the time  $t$  state

Thus, we can rewrite the Ramsey problem as

$$\max_{\{y_t, \tau_{t+1}\}} - \sum_{t=0}^{\infty} \beta^t y_t' R y_t \quad (3.131)$$

subject to  $z_0$  given and the law of motion

$$y_{t+1} = A y_t + B \tau_{t+1} \quad (3.132)$$

where

$$R = \begin{bmatrix} 0 & -\frac{A_0}{2} & 0 & 0 \\ -\frac{A_0}{2} & \frac{A_1}{2} & \frac{-\mu}{2} & 0 \\ 0 & \frac{-\mu}{2} & 0 & 0 \\ 0 & 0 & 0 & \frac{d}{2} \end{bmatrix}, \quad A = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 1 \\ 0 & 0 & 0 & 0 \\ -\frac{A_0}{d} & \frac{A_1}{d} & 0 & \frac{A_1}{d} + \frac{1}{\beta} \end{bmatrix}, \quad B = \begin{bmatrix} 0 \\ 0 \\ 1 \\ \frac{1}{d} \end{bmatrix}$$

### Two Subproblems

Working backwards, we first present the Bellman equation for the value function that takes both  $z_t$  and  $u_t$  as given. Then we present a value function that takes only  $z_0$  as given and is the indirect utility function that arises from choosing  $u_0$  optimally.

Let  $v(Q_t, \tau_t, u_t)$  be the optimum value function for the time  $t \geq 1$  government administrator facing state  $Q_t, \tau_t, u_t$ .

Let  $w(Q_0)$  be the value of the Ramsey plan starting from  $Q_0$

**Subproblem 1** Here the Bellman equation is

$$v(Q_t, \tau_t, u_t) = \max_{\tau_{t+1}} \left\{ A_0 Q_t - \frac{A_1}{2} Q_t^2 - \frac{d}{2} u_t^2 + \mu \tau_t Q_t + \beta v(Q_{t+1}, \tau_{t+1}, u_{t+1}) \right\}$$

where the maximization is subject to the constraints

$$Q_{t+1} = Q_t + u_t$$

and

$$u_{t+1} = -\frac{A_0}{d} + \frac{A_1}{d} Q_t + \frac{A_1}{d} + \frac{1}{\beta} u_t + \frac{1}{d} \tau_{t+1}$$

Here we regard  $u_t$  as a state

**Subproblem 2** The subproblem 2 Bellman equation is

$$w(z_0) = \max_{u_0} v(Q_0, 0, u_0)$$

**Details** Define the state vector to be

$$y_t = \begin{bmatrix} 1 \\ Q_t \\ \tau_t \\ u_t \end{bmatrix} = \begin{bmatrix} z_t \\ u_t \end{bmatrix},$$

where  $z_t = [1 \ Q_t \ \tau_t]'$  are authentic state variables and  $u_t$  is a variable whose time 0 value is a 'jump' variable but whose values for dates  $t \geq 1$  will become state variables that encode history dependence in the Ramsey plan

$$v(y_t) = \max_{\tau_{t+1}} \{-y_t' R y_t + \beta v(y_{t+1})\} \quad (3.133)$$

where the maximization is subject to the constraint

$$y_{t+1} = A y_t + B \tau_{t+1}$$

and where

$$R = \begin{bmatrix} 0 & -\frac{A_0}{2} & 0 & 0 \\ -\frac{A_0}{2} & \frac{A_1}{2} & \frac{-\mu}{2} & 0 \\ 0 & \frac{-\mu}{2} & 0 & 0 \\ 0 & 0 & 0 & \frac{d}{2} \end{bmatrix}, \quad A = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 1 \\ 0 & 0 & 0 & 0 \\ -\frac{A_0}{d} & \frac{A_1}{d} & 0 & \frac{A_1}{d} + \frac{1}{\beta} \end{bmatrix}, \text{ and } B = \begin{bmatrix} 0 \\ 0 \\ 1 \\ \frac{1}{d} \end{bmatrix}.$$

Functional equation (3.133) has solution

$$v(y_t) = -y_t' P y_t$$

where

- $P$  solves the algebraic matrix Riccati equation  $P = R + \beta A' P A - \beta A' P B (B' P B)^{-1} B' P A$
- the optimal policy function is given by  $\tau_{t+1} = -F y_t$  for  $F = (B' P B)^{-1} B' P A$

Now we turn to subproblem 1.

Evidently the optimal choice of  $u_0$  satisfies  $\frac{\partial v}{\partial u_0} = 0$

If we partition  $P$  as

$$P = \begin{bmatrix} P_{11} & P_{12} \\ P_{21} & P_{22} \end{bmatrix}$$

then we have

$$0 = \frac{\partial}{\partial u_0} (z_0' P_{11} z_0 + z_0' P_{12} u_0 + u_0' P_{21} z_0 + u_0' P_{22} u_0) = P_{12}' z_0 + P_{21} z_0 + 2P_{22} u_0$$

which implies

$$u_0 = -P_{22}^{-1} P_{21} z_0 \quad (3.134)$$

Thus, the Ramsey plan is

$$\tau_{t+1} = -F \begin{bmatrix} z_t \\ u_t \end{bmatrix} \quad \text{and} \quad \begin{bmatrix} z_{t+1} \\ u_{t+1} \end{bmatrix} = (A - BF) \begin{bmatrix} z_t \\ u_t \end{bmatrix}$$

with initial state  $[z_0 \ -P_{22}^{-1} P_{21} z_0]'$

**Recursive Representation** An outcome of the preceding results is that the Ramsey plan can be represented recursively as the choice of an initial marginal utility (or rate of growth of output) according to a function

$$u_0 = v(Q_0|\mu) \quad (3.135)$$

that obeys (3.134) and the following updating equations for  $t \geq 0$ :

$$\tau_{t+1} = \tau(Q_t, u_t|\mu) \quad (3.136)$$

$$Q_{t+1} = Q_t + u_t \quad (3.137)$$

$$u_{t+1} = u(Q_t, u_t|\mu) \quad (3.138)$$

We have conditioned the functions  $v$ ,  $\tau$ , and  $u$  by  $\mu$  to emphasize how the dependence of  $F$  on  $G_0$  appears indirectly through the Lagrange multiplier  $\mu$

**An Example Calculation** We'll discuss how to compute  $\mu$  below but first consider the following numerical example

We take the parameter set  $[A_0, A_1, d, \beta, Q_0] = [100, .05, .2, .95, 100]$  and compute the Ramsey plan with the following piece of code

```
#=
In the following, ``uhat`` and ``tauhat`` are what the planner would
choose if he could reset at time t , ``uhatdif`` and ``tauhatdif`` are
the difference between those and what the planner is constrained to
choose. The variable ``mu`` is the Lagrange multiplier associated with
the constraint at time t .
For more complete description of inputs and outputs see the website.
@author : Spencer Lyon <spencer.lyon@nyu.edu>
 Victoria Gregory <victoria.gregory@nyu.edu>
@date: 2014-08-21
References

Simple port of the file examples/evans_sargent.py
http://quant-econ.net/hist_dep_policies.html
=#
using QuantEcon
using Optim
using Plots
using LaTeXStrings
pyplot()
```

```

type HistDepRamsey
 # These are the parameters of the economy
 A0::Real
 A1::Real
 d::Real
 Q0::Real
 tau0::Real
 mu0::Real
 bet::Real

 # These are the LQ fields and stationary values
 R::Matrix
 A::Matrix
 B::Array
 Q::Real
 P::Matrix
 F::Matrix
 lq::LQ
end

type RamseyPath
 y::Matrix
 uhat::Vector
 uhatdif::Vector
 tauhat::Vector
 tauhatdif::Vector
 mu::Vector
 G::Vector
 GPay::Vector
end

function HistDepRamsey(A0, A1, d, Q0, tau0, mu, bet)
 # Create Matrices for solving Ramsey problem
 R = [0.0 -A0/2 0.0 0.0
 -A0/2 A1/2 -mu/2 0.0
 0.0 -mu/2 0.0 0.0
 0.0 0.0 0.0 d/2]

 A = [1.0 0.0 0.0 0.0
 0.0 1.0 0.0 1.0
 0.0 0.0 0.0 0.0
 -A0/d A1/d 0.0 A1/d+1.0/bet]

 B = [0.0; 0.0; 1.0; 1.0/d]

 Q = 0.0

 # Use LQ to solve the Ramsey Problem.
 lq = LQ(Q, -R, A, B, bet=bet)

 P, F, _d = stationary_values(lq)

```

```

 HistDepRamsey(A0, A1, d, Q0, tau0, mu0, bet, R, A, B, Q, P, F, lq)
end

function compute_G(hdr::HistDepRamsey, mu)
 # simplify notation
 Q0, tau0, P, F, d, A, B = hdr.Q0, hdr.tau0, hdr.P, hdr.F, hdr.d, hdr.A, hdr.B
 bet = hdr.bet

 # Need y_0 to compute government tax revenue.
 u0 = compute_u0(hdr, P)
 y0 = vcat([1.0 Q0 tau0]', u0)

 # Define A_F and S matrixies
 AF = A - B * F
 S = [0.0 1.0 0.0 0]' * [0.0 0.0 1.0 0]

 # Solves equation (25)
 Omega = solve_discrete_lyapunov(sqrt(bet) .* AF', bet .* AF' * S * AF)
 T0 = y0' * Omega * y0

 return T0[1], A, B, F, P
end

function compute_u0(hdr::HistDepRamsey, P::Matrix)
 # simplify notation
 Q0, tau0 = hdr.Q0, hdr.tau0

 P21 = P[4, 1:3]
 P22 = P[4, 4]
 z0 = [1.0 Q0 tau0]'
 u0 = -P22^(-1) .* P21*(z0)

 return u0[1]
end

function init_path(hdr::HistDepRamsey, mu0, T::Int=20)
 # Construct starting values for the path of the Ramsey economy
 G0, A, B, F, P = compute_G(hdr, mu0)

 # Compute the optimal u0
 u0 = compute_u0(hdr, P)

 # Initialize vectors
 y = Array(Float64, 4, T)
 uhat = Array(Float64, T)
 uhatdif = Array(Float64, T)
 tauhat = Array(Float64, T)
 tauhatdif = Array(Float64, T-1)
 mu = Array(Float64, T)
 G = Array(Float64, T)

```

```

GPay = Array(Float64, T)

Initial conditions
G[1] = G0
mu[1] = mu0
uhatdif[1] = 0.0
uhat[1] = u0
y[:, 1] = vcat([1.0 hdr.Q0 hdr.tau0]', u0)

return RamseyPath(y, uhat, uhatdif, tauhat, tauhatdif, mu, G, GPay)
end

function compute_ramsey_path!(hdr::HistDepRamsey, rp::RamseyPath)
 # simplify notation
 y, uhat, uhatdif, tauhat, = rp.y, rp.uhat, rp.uhatdif, rp.tauhat
 tauhatdif, mu, G, GPay = rp.tauhatdif, rp.mu, rp.G, rp.GPay
 bet = hdr.bet

 G0, A, B, F, P = compute_G(hdr, mu[1])

 for t=2:T
 # iterate government policy
 y[:, t] = (A - B * F) * y[:, t-1]

 # update G
 G[t] = (G[t-1] - bet*y[2, t]*y[3, t])/bet
 GPay[t] = bet.*y[2, t]*y[3, t]

 #=
 Compute the mu if the government were able to reset its plan
 ff is the tax revenues the government would receive if they reset the
 plan with Lagrange multiplier mu minus current G
 =
 ff(mu) = abs(compute_G(hdr, mu)[1]-G[t])
 # find ff = 0
 mu[t] = optimize(ff, mu[t-1]-1e4, mu[t-1]+1e4).minimum
 temp, Atemp, Btemp, Ftemp, Ptemp = compute_G(hdr, mu[t])

 # Compute alternative decisions
 P21temp = Ptemp[4, 1:3]
 P22temp = P[4, 4]
 uhat[t] = (-P22temp^(-1) .* P21temp * y[1:3, t])[1]

 yhat = (Atemp-Btemp * Ftemp) * [y[1:3, t-1]; uhat[t-1]]
 tauhat[t] = yhat[3]
 tauhatdif[t-1] = tauhat[t] - y[3, t]
 uhatdif[t] = uhat[t] - y[3, t]
 end

 return rp

```

```

end

function plot1(rp::RamseyPath)
 tt = 1:length(rp.mu) # tt is used to make the plot time index correct.
 y = rp.y

 ylabels = [L"Q" L"τ" L"u"]
 y_vals = [squeeze(y[2, :], 1) squeeze(y[3, :], 1) squeeze(y[4, :], 1)]
 p = plot(tt, y_vals, color=:blue,
 label=["output" "tax rate" "first difference in output"],
 lw=2, alpha=0.7, ylabel=ylabels, layout=(3,1),
 xlims=(0, 15), xlabel="" "time", legend=:topright,
 xticks=0:5:15)
 return p
end

function plot2(rp::RamseyPath)
 y, uhatdif, tauhatdif, mu = rp.y, rp.uhatdif, rp.tauhatdif, rp.mu
 G, GPay = rp.G, rp.GPay
 T = length(rp.mu)
 tt = 1:T # tt is used to make the plot time index correct.
 tt2 = 0:T-1
 tauhatdif = [NaN; tauhatdif]

 x_vals = [tt2 tt tt tt]
 y_vals = [tauhatdif uhatdif mu G]
 ylabels = [L"$\Delta\tau$ L"Δu" L"μ" L"G"]
 labels = ["time inconsistency differential for tax rate" "time inconsistency differential for u"]
 p = plot(x_vals, y_vals, ylabel=ylabels, label=labels,
 layout=(4, 1), xlims=(-0.5, 15), lw=2, alpha=0.7,
 legend=:topright, color=:blue, xlabel="" "time"))
 return p
end

Primitives
T = 20
A0 = 100.0
A1 = 0.05
d = 0.20
bet = 0.95

Initial conditions
mu0 = 0.0025
Q0 = 1000.0
tau0 = 0.0

Solve Ramsey problem and compute path
hdr = HistDepRamsey(A0, A1, d, Q0, tau0, mu0, bet)
rp = init_path(hdr, mu0, T)
compute_ramsey_path!(hdr, rp) # updates rp in place
plot1(rp)

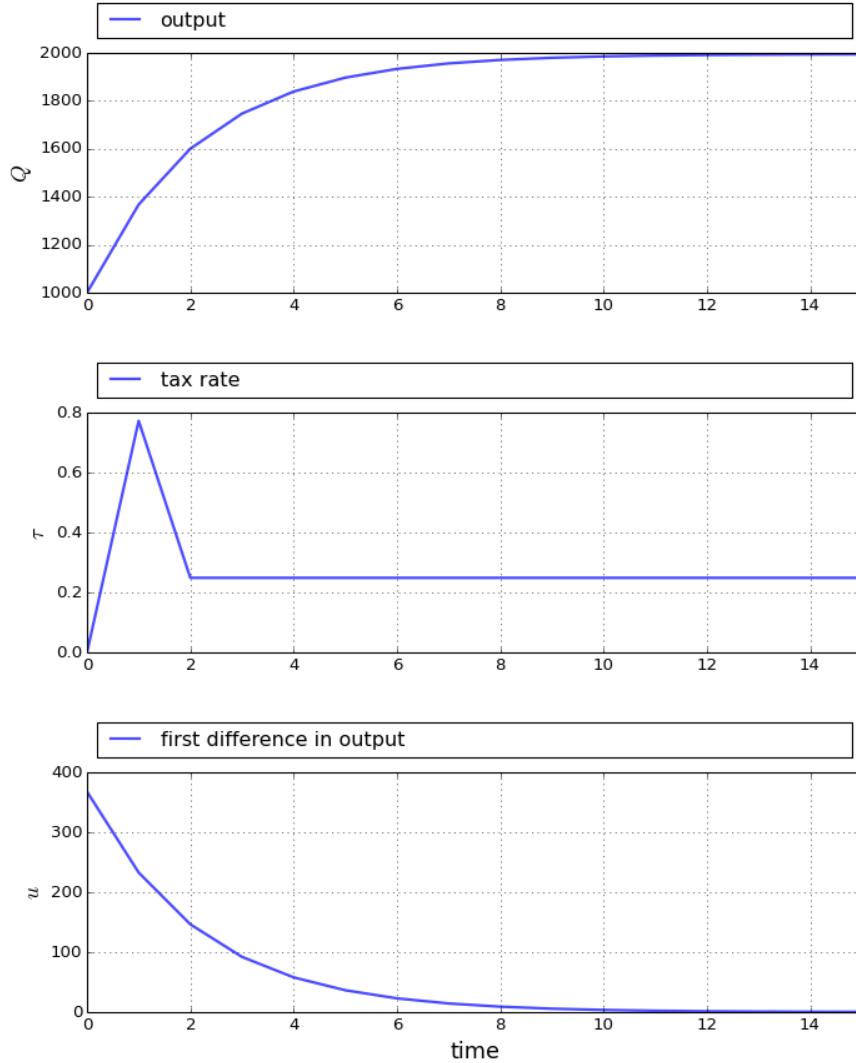
```

```
plot2(rp)
```

The program can also be [found](#) in the QuantEcon GitHub repository

It computes a number of sequences besides the Ramsey plan, some of which have already been discussed, while others will be described below

The next figure uses the program to compute and show the Ramsey plan for  $\tau$  and the Ramsey outcome for  $(Q_t, u_t)$



From top to bottom, the panels show  $Q_t$ ,  $\tau_t$  and  $u_t := Q_{t+1} - Q_t$  over  $t = 0, \dots, 15$

The optimal decision rule is <sup>4</sup>

$$\tau_{t+1} = -248.0624 - 0.1242Q_t - 0.3347u_t \quad (3.139)$$

Notice how the Ramsey plan calls for a high tax at  $t = 1$  followed by a perpetual stream of lower taxes

Taxing heavily at first, less later expresses time-inconsistency of the optimal plan for  $\{\tau_{t+1}\}_{t=0}^{\infty}$

We'll characterize this formally after first discussing how to compute  $\mu$ .

**Computing  $\mu$**  Define the selector vectors  $e_{\tau} = [0 \ 0 \ 1 \ 0]'$  and  $e_Q = [0 \ 1 \ 0 \ 0]'$  and express  $\tau_t = e_{\tau}'y_t$  and  $Q_t = e_Q'y_t$

Evidently  $Q_t\tau_t = y_t'e_Qe_{\tau}'y_t = y_t'Sy_t$  where  $S := e_Qe_{\tau}'$

We want to compute

$$T_0 = \sum_{t=1}^{\infty} \beta^t \tau_t Q_t = \tau_1 Q_1 + \beta T_1$$

where  $T_1 = \sum_{t=2}^{\infty} \beta^{t-1} Q_t \tau_t$

The present values  $T_0$  and  $T_1$  are connected by

$$T_0 = \beta y_0' A_F' S A_F y_0 + \beta T_1$$

Guess a solution that takes the form  $T_t = y_t'\Omega y_t$ , then find an  $\Omega$  that satisfies

$$\Omega = \beta A_F' S A_F + \beta A_F' \Omega A_F \quad (3.140)$$

Equation (3.140) is a discrete Lyapunov equation that can be solved for  $\Omega$  using QuantEcon's `solve_discrete_lyapunov` function

The matrix  $F$  and therefore the matrix  $A_F = A - BF$  depend on  $\mu$

To find a  $\mu$  that guarantees that  $T_0 = G_0$  we proceed as follows:

1. Guess an initial  $\mu$ , compute a tentative Ramsey plan and the implied  $T_0 = y_0'\Omega(\mu)y_0$
2. If  $T_0 > G_0$ , lower  $\mu$ ; otherwise, raise  $\mu$
3. Continue iterating on step 3 until  $T_0 = G_0$

### Time Inconsistency

Recall that the Ramsey planner chooses  $\{u_t\}_{t=0}^{\infty}, \{\tau_t\}_{t=1}^{\infty}$  to maximize

$$\sum_{t=0}^{\infty} \beta^t \left\{ A_0 Q_t - \frac{A_1}{2} Q_t^2 - \frac{d}{2} u_t^2 \right\}$$

subject to (3.125), (3.126), and (3.128)

---

<sup>4</sup> As promised,  $\tau_t$  does not appear in the Ramsey planner's decision rule for  $\tau_{t+1}$ .

We express the outcome that a Ramsey plan is time-inconsistent the following way

**Proposition.** A continuation of a Ramsey plan is not a Ramsey plan

Let

$$w(Q_0, u_0 | \mu_0) = \sum_{t=0}^{\infty} \beta^t \left\{ A_0 Q_t - \frac{A_1}{2} Q_t^2 - \frac{d}{2} u_t^2 \right\} \quad (3.141)$$

where

- $\{Q_t, u_t\}_{t=0}^{\infty}$  are evaluated under the Ramsey plan whose recursive representation is given by (3.136), (3.137), (3.138)
- $\mu_0$  is the value of the Lagrange multiplier that assures budget balance, computed as described above

Evidently, these continuation values satisfy the recursion

$$w(Q_t, u_t | \mu_0) = A_0 Q_t - \frac{A_1}{2} Q_t^2 - \frac{d}{2} u_t^2 + \beta w(Q_{t+1}, u_{t+1} | \mu_0) \quad (3.142)$$

for all  $t \geq 0$ , where  $Q_{t+1} = Q_t + u_t$

Under the timing protocol affiliated with the Ramsey plan, the planner is committed to the outcome of iterations on (3.136), (3.137), (3.138)

In particular, when time  $t$  comes, the Ramsey planner is committed to the value of  $u_t$  implied by the Ramsey plan and receives continuation value  $w(Q_t, u_t, \mu_0)$

That the Ramsey plan is time-inconsistent can be seen by subjecting it to the following ‘revolutionary’ test

First, define continuation revenues  $G_t$  that the government raises along the original Ramsey outcome by

$$G_t = \beta^{-t} (G_0 - \sum_{s=1}^t \beta^s \tau_s Q_s) \quad (3.143)$$

where  $\{\tau_t, Q_t\}_{t=0}^{\infty}$  is the original Ramsey outcome <sup>5</sup>

Then at time  $t \geq 1$ ,

1. take  $(Q_t, G_t)$  inherited from the original Ramsey plan as initial conditions
2. invite a brand new Ramsey planner to compute a new Ramsey plan, solving for a new  $u_t$ , to be called  $\check{u}_t$ , and for a new  $\mu$ , to be called  $\check{\mu}_t$

The revised Lagrange multiplier  $\check{\mu}_t$  is chosen so that, under the new Ramsey plan, the government is able to raise enough continuation revenues  $G_t$  given by (3.143)

Would this new Ramsey plan be a continuation of the original plan?

The answer is no because along a Ramsey plan, for  $t \geq 1$ , in general it is true that

$$w(Q_t, v(Q_t | \check{\mu}) | \check{\mu}) > w(Q_t, u_t | \mu_0) \quad (3.144)$$

---

<sup>5</sup> The continuation revenues  $G_t$  are the time  $t$  present value of revenues that must be raised to satisfy the original time 0 government intertemporal budget constraint, taking into account the revenues already raised from  $s = 1, \dots, t$  under the original Ramsey plan.

Inequality (3.144) expresses a continuation Ramsey planner's incentive to deviate from a time 0 Ramsey plan by

1. resetting  $u_t$  according to (3.135)
2. adjusting the Lagrange multiplier on the continuation appropriately to account for tax revenues already collected<sup>6</sup>

Inequality (3.144) expresses the time-inconsistency of a Ramsey plan

**A Simulation** To bring out the time inconsistency of the Ramsey plan, we compare

- the time  $t$  values of  $\tau_{t+1}$  under the original Ramsey plan with
- the value  $\check{\tau}_{t+1}$  associated with a new Ramsey plan begun at time  $t$  with initial conditions  $(Q_t, G_t)$  generated by following the *original* Ramsey plan

Here again  $G_t := \beta^{-t}(G_0 - \sum_{s=1}^t \beta^s \tau_s Q_s)$

The difference  $\Delta\tau_t := \check{\tau}_t - \tau_t$  is shown in the top panel of the following figure

In the second panel we compare the time  $t$  outcome for  $u_t$  under the original Ramsey plan with the time  $t$  value of this new Ramsey problem starting from  $(Q_t, G_t)$

To compute  $u_t$  under the new Ramsey plan, we use the following version of formula (3.134):

$$\check{u}_t = -P_{22}^{-1}(\check{\mu}_t) P_{21}(\check{\mu}_t) z_t$$

Here  $z_t$  is evaluated along the Ramsey outcome path, where we have included  $\check{\mu}_t$  to emphasize the dependence of  $P$  on the Lagrange multiplier  $\mu_0$ <sup>7</sup>

To compute  $u_t$  along the Ramsey path, we just iterate the recursion starting (??) from the initial  $Q_0$  with  $u_0$  being given by formula (3.134)

Thus the second panel indicates how far the reinitialized value  $\check{u}_t$  value departs from the time  $t$  outcome along the Ramsey plan

Note that the restarted plan raises the time  $t + 1$  tax and consequently lowers the time  $t$  value of  $u_t$

Associated with the new Ramsey plan at  $t$  is a value of the Lagrange multiplier on the continuation government budget constraint

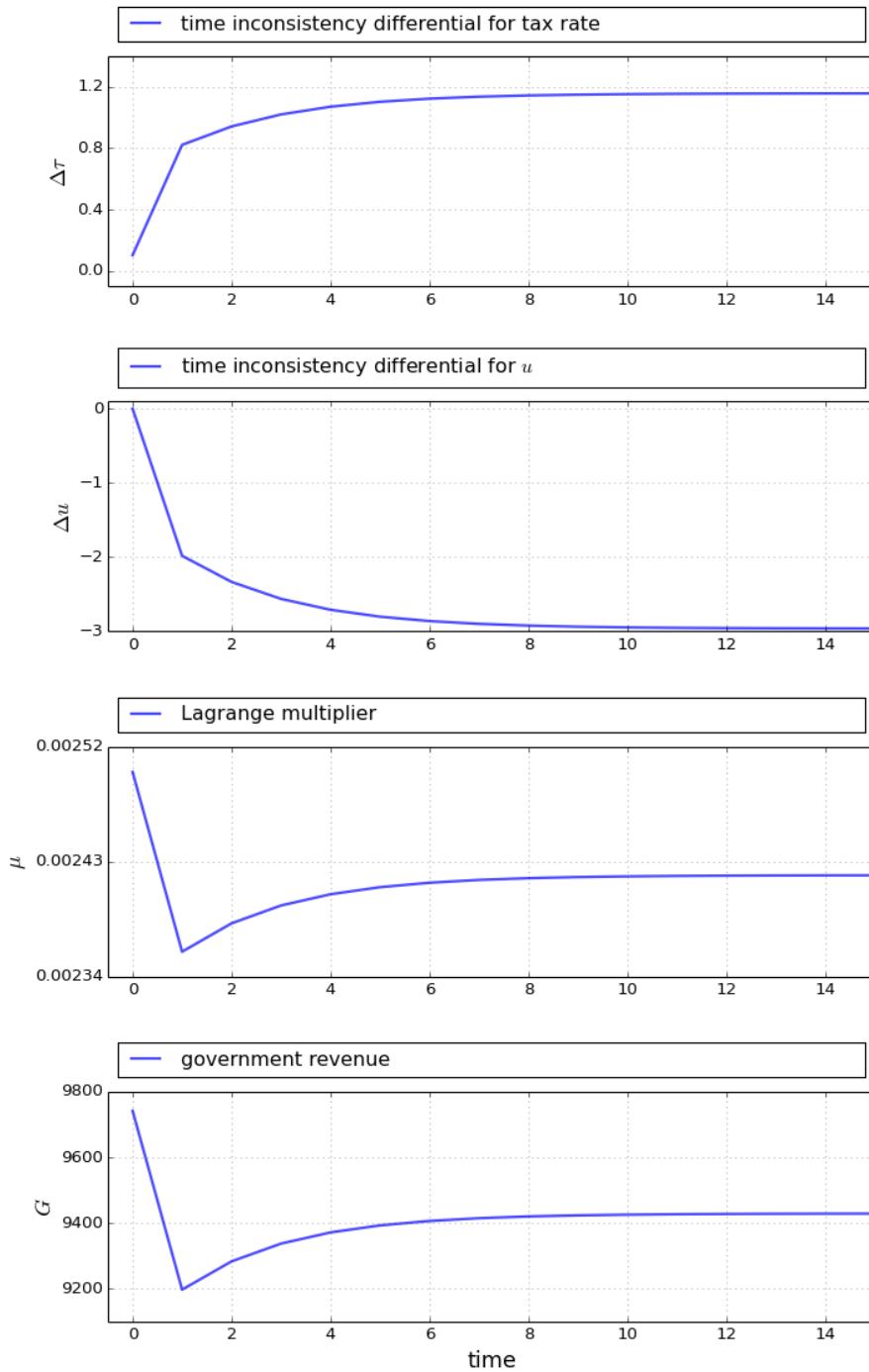
This is the third panel of the figure

The fourth panel plots the required continuation revenues  $G_t$  implied by the original Ramsey plan

These figures help us understand the time inconsistency of the Ramsey plan

<sup>6</sup> For example, let the Ramsey plan yield time 1 revenues  $Q_1 \tau_1$ . Then at time 1, a continuation Ramsey planner would want to raise continuation revenues, expressed in units of time 1 goods, of  $\tilde{G}_1 := \frac{G - \beta Q_1 \tau_1}{\beta}$ . To finance the remainder revenues, the continuation Ramsey planner would find a continuation Lagrange multiplier  $\mu$  by applying the three-step procedure from the previous section to revenue requirements  $\tilde{G}_1$ .

<sup>7</sup> It can be verified that this formula puts non-zero weight only on the components 1 and  $Q_t$  of  $z_t$ .



**Further Intuition** One feature to note is the large difference between  $\tau_{t+1}$  and  $\tau_{t+1}$  in the top panel of the figure

If the government is able to reset to a new Ramsey plan at time  $t$ , it chooses a significantly higher tax rate than if it were required to maintain the original Ramsey plan

The intuition here is that the government is required to finance a given present value of expenditures with distorting taxes  $\tau$

The quadratic adjustment costs prevent firms from reacting strongly to variations in the tax rate for next period, which tilts a time  $t$  Ramsey planner toward using time  $t + 1$  taxes

As was noted before, this is evident in *the first figure*, where the government taxes the next period heavily and then falls back to a constant tax from then on

This can also be seen in the third panel of *the second figure*, where the government pays off a significant portion of the debt using the first period tax rate

The similarities between the graphs in the last two panels of *the second figure* reveals that there is a one-to-one mapping between  $G$  and  $\mu$

The Ramsey plan can then only be time consistent if  $G_t$  remains constant over time, which will not be true in general

### Credible Policy

We express the theme of this section in the following: In general, a continuation of a Ramsey plan is not a Ramsey plan

This is sometimes summarized by saying that a Ramsey plan is not *credible*

On the other hand, a continuation of a credible plan is a credible plan

The literature on a credible public policy ([\[CK90\]](#) and [\[Sto89\]](#)) arranges strategies and incentives so that public policies can be implemented by a *sequence* of government decision makers instead of a single Ramsey planner who chooses an entire sequence of history-dependent actions once and for all at time  $t = 0$

Here we confine ourselves to sketching how recursive methods can be used to characterize credible policies in our model

A key reference on these topics is [\[Cha98\]](#)

A credibility problem arises because we assume that the timing of decisions differs from those for a Ramsey problem

A **sequential timing protocol** is a protocol such that

1. At each  $t \geq 0$ , given  $Q_t$  and expectations about a continuation tax policy  $\{\tau_{s+1}\}_{s=t}^{\infty}$  and a continuation price sequence  $\{p_{s+1}\}_{s=t}^{\infty}$ , the representative firm chooses  $u_t$
2. At each  $t$ , given  $(Q_t, u_t)$ , a government chooses  $\tau_{t+1}$

Item (2) captures that taxes are now set sequentially, the time  $t + 1$  tax being set *after* the government has observed  $u_t$

Of course, the representative firm sets  $u_t$  in light of its expectations of how the government will ultimately choose to set future taxes

A credible tax plan  $\{\tau_{s+1}\}_{s=t}^{\infty}$

- is anticipated by the representative firm, and
- is one that a time  $t$  government chooses to confirm

We use the following recursion, closely related to but different from (3.142), to define the continuation value function for the government:

$$J_t = A_0 Q_t - \frac{A_1}{2} Q_t^2 - \frac{d}{2} u_t^2 + \beta J_{t+1}(\tau_{t+1}, G_{t+1}) \quad (3.145)$$

This differs from (3.142) because

- continuation values are now allowed to depend explicitly on values of the choice  $\tau_{t+1}$ , and
- continuation government revenue to be raised  $G_{t+1}$  need not be ones called for by the prevailing government policy

Thus, deviations from that policy are allowed, an alteration that recognizes that  $\tau_t$  is chosen sequentially

Express the government budget constraint as requiring that  $G_0$  solves the difference equation

$$G_t = \beta \tau_{t+1} Q_{t+1} + \beta G_{t+1}, \quad t \geq 0 \quad (3.146)$$

subject to the terminal condition  $\lim_{t \rightarrow +\infty} \beta^t G_t = 0$

Because the government is choosing sequentially, it is convenient to

- take  $G_t$  as a state variable at  $t$  and
- to regard the time  $t$  government as choosing  $(\tau_{t+1}, G_{t+1})$  subject to constraint (3.146)

To express the notion of a credible government plan concisely, we expand the strategy space by also adding  $J_t$  itself as a state variable and allowing policies to take the following recursive forms<sup>8</sup>

Regard  $J_0$  as an discounted present value promised to the Ramsey planner and take it as an initial condition.

Then after choosing  $u_0$  according to

$$u_0 = v(Q_0, G_0, J_0), \quad (3.147)$$

choose subsequent taxes, outputs, and continuation values according to recursions that can be represented as

$$\hat{\tau}_{t+1} = \tau(Q_t, u_t, G_t, J_t) \quad (3.148)$$

$$u_{t+1} = \xi(Q_t, u_t, G_t, J_t, \tau_{t+1}) \quad (3.149)$$

$$G_{t+1} = \beta^{-1} G_t - \tau_{t+1} Q_{t+1} \quad (3.150)$$

<sup>8</sup> This choice is the key to what [LS12] call ‘dynamic programming squared’.

$$J_{t+1}(\tau_{t+1}, G_{t+1}) = \nu(Q_t, u_t, G_{t+1}, J_t, \tau_{t+1}) \quad (3.151)$$

Here

- $\hat{\tau}_{t+1}$  is the time  $t + 1$  government action called for by the plan, while
- $\tau_{t+1}$  is possibly some one-time deviation that the time  $t + 1$  government contemplates and
- $G_{t+1}$  is the associated continuation tax collections

The plan is said to be **credible** if, for each  $t$  and each state  $(Q_t, u_t, G_t, J_t)$ , the plan satisfies the incentive constraint

$$J_t = A_0 Q_t - \frac{A_1}{2} Q_t^2 - \frac{d}{2} u_t^2 + \beta J_{t+1}(\hat{\tau}_{t+1}, \hat{G}_{t+1}) \quad (3.152)$$

$$\geq A_0 Q_t - \frac{A_1}{2} Q_t^2 - \frac{d}{2} u_t^2 + \beta J_{t+1}(\tau_{t+1}, G_{t+1}) \quad (3.153)$$

for all tax rates  $\tau_{t+1} \in \mathbb{R}$  available to the government

$$\text{Here } \hat{G}_{t+1} = \frac{G_t - \hat{\tau}_{t+1} Q_{t+1}}{\beta}$$

- Inequality expresses that continuation values adjust to deviations in ways that discourage the government from deviating from the prescribed  $\hat{\tau}_{t+1}$
- Inequality (3.152) indicates that *two* continuation values  $J_{t+1}$  contribute to sustaining time  $t$  promised value  $J_t$ 
  - $J_{t+1}(\hat{\tau}_{t+1}, \hat{G}_{t+1})$  is the continuation value when the government chooses to confirm the private sector's expectation, formed according to the decision rule (3.148)<sup>9</sup>
  - $J_{t+1}(\tau_{t+1}, G_{t+1})$  tells the continuation consequences should the government disappoint the private sector's expectations

The internal structure of a credible plan deters deviations from it

That (3.152) maps *two* continuation values  $J_{t+1}(\tau_{t+1}, G_{t+1})$  and  $J_{t+1}(\hat{\tau}_{t+1}, \hat{G}_{t+1})$  into one promised value  $J_t$  reflects how a credible plan arranges a system of private sector expectations that induces the government to choose to confirm them

Chang [Cha98] builds on how inequality (3.152) maps two continuation values into one

**Remark** Let  $\mathcal{J}$  be the set of values associated with credible plans

Every value  $J \in \mathcal{J}$  can be attained by a credible plan that has a recursive representation of form form (3.148), (3.149), (3.150)

The set of values can be computed as the largest fixed point of an operator that maps sets of candidate values into sets of values

Given a value within this set, it is possible to construct a government strategy of the recursive form (3.148), (3.149), (3.150) that attains that value

In many cases, there is a **set** of values and associated credible plans

---

<sup>9</sup> Note the double role played by (3.148): as decision rule for the government and as the private sector's rule for forecasting government actions.

In those cases where the Ramsey outcome is credible, a multiplicity of credible plans is a key part of the story because, as we have seen earlier, a continuation of a Ramsey plan is not a Ramsey plan

For it to be credible, a Ramsey outcome must be supported by a worse outcome associated with another plan, the prospect of reversion to which sustains the Ramsey outcome

### Concluding remarks

The term ‘optimal policy’, which pervades an important applied monetary economics literature, means different things under different timing protocols

Under the ‘static’ Ramsey timing protocol (i.e., choose a sequence once-and-for-all), we obtain a unique plan

Here the phrase ‘optimal policy’ seems to fit well, since the Ramsey planner optimally reaps early benefits from influencing the private sector’s beliefs about the government’s later actions

When we adopt the sequential timing protocol associated with credible public policies, ‘optimal policy’ is a more ambiguous description

There is a multiplicity of credible plans

True, the theory explains how it is optimal for the government to confirm the private sector’s expectations about its actions along a credible plan

But some credible plans have very bad outcomes

These bad outcomes are central to the theory because it is the presence of bad credible plans that makes possible better ones by sustaining the low continuation values that appear in the second line of incentive constraint (3.152)

Recently, many have taken for granted that ‘optimal policy’ means ‘follow the Ramsey plan’<sup>10</sup>

In pursuit of more attractive ways to describe a Ramsey plan when policy making is in practice done sequentially, some writers have repackaged a Ramsey plan in the following way

- Take a Ramsey *outcome* - a sequence of endogenous variables under a Ramsey plan - and reinterpret it (or perhaps only a subset of its variables) as a *target path* of relationships among outcome variables to be assigned to a sequence of policy makers<sup>11</sup>
- If appropriate (infinite dimensional) invertibility conditions are satisfied, it can happen that following the Ramsey plan is the *only* way to hit the target path<sup>12</sup>
- The spirit of this work is to say, “in a democracy we are obliged to live with the sequential timing protocol, so let’s constrain policy makers’ objectives in ways that will force them to follow a Ramsey plan in spite of their benevolence”<sup>13</sup>
- By this slight of hand, we acquire a theory of an *optimal outcome target path*

<sup>10</sup> It is possible to read [Woo03] and [GW10] as making some carefully qualified statements of this type. Some of the qualifications can be interpreted as advice ‘eventually’ to follow a tail of a Ramsey plan.

<sup>11</sup> In our model, the Ramsey outcome would be a path  $(\vec{p}, \vec{Q})$ .

<sup>12</sup> See [GW10].

<sup>13</sup> Sometimes the analysis is framed in terms of following the Ramsey plan only from some future date  $T$  onwards.

This ‘invertibility’ argument leaves open two important loose ends:

1. implementation, and
2. time consistency

As for (1), repackaging a Ramsey plan (or the tail of a Ramsey plan) as a target outcome sequence does not confront the delicate issue of *how* that target path is to be implemented <sup>14</sup>

As for (2), it is an interesting question whether the ‘invertibility’ logic can repackage and conceal a Ramsey plan well enough to make policy makers forget or ignore the benevolent intentions that give rise to the time inconsistency of a Ramsey plan in the first place

To attain such an optimal output path, policy makers must forget their benevolent intentions because there will inevitably occur temptations to deviate from that target path, and the implied relationship among variables like inflation, output, and interest rates along it

**Remark** The continuation of such an optimal target path is not an optimal target path

## Default Risk and Income Fluctuations

### Contents

- *Default Risk and Income Fluctuations*
  - Overview
  - Structure
  - Equilibrium
  - Computation
  - Results
  - Exercises
  - Solutions

### Overview

This lecture computes versions of Arellano’s [Are08] model of sovereign default

The model describes interactions among default risk, output, and an equilibrium interest rate that includes a premium for endogenous default risk

The decision maker is a government of a small open economy that borrows from risk-neutral foreign creditors

The foreign lenders must be compensated for default risk

The government borrows and lends abroad in order to smooth the consumption of its citizens

The government repays its debt only if it wants to, but declining to pay has adverse consequences

---

<sup>14</sup> See [Bas05] and [ACK10].

The interest rate on government debt adjusts in response to the state-dependent default probability chosen by government

The model yields outcomes that help interpret sovereign default experiences, including

- countercyclical interest rates on sovereign debt
- countercyclical trade balances
- high volatility of consumption relative to output

Notably, long recessions caused by bad draws in the income process increase the government's incentive to default

This can lead to

- spikes in interest rates
- temporary losses of access to international credit markets
- large drops in output, consumption, and welfare
- large capital outflows during recessions

Such dynamics are consistent with experiences of many countries

### Structure

In this section we describe the main features of the model

**Output, Consumption and Debt** A small open economy is endowed with an exogenous stochastically fluctuating potential output stream  $\{y_t\}$

Potential output is realized only in periods in which the government honors its sovereign debt

The output good can be traded or consumed

The sequence  $\{y_t\}$  is described by a Markov process with stochastic density kernel  $p(y, y')$

Households within the country are identical and rank stochastic consumption streams according to

$$\mathbb{E} \sum_{t=0}^{\infty} \beta^t u(c_t) \quad (3.154)$$

Here

- $0 < \beta < 1$  is a time discount factor
- $u$  is an increasing and strictly concave utility function

Consumption sequences enjoyed by households are affected by the government's decision to borrow or lend internationally

The government is benevolent in the sense that its aim is to maximize (3.154)

The government is the only domestic actor with access to foreign credit

Because household are averse to consumption fluctuations, the government will try to smooth consumption by borrowing from (and lending to) foreign creditors

**Asset Markets** The only credit instrument available to the government is a one-period bond traded in international credit markets

The bond market has the following features

- The bond matures in one period and is not state contingent
- A purchase of a bond with face value  $B'$  is a claim to  $B'$  units of the consumption good next period
- To purchase  $B'$  next period costs  $qB'$  now, or, what is equivalent
- For selling  $-B'$  units of next period goods the seller earns  $-qB'$  of today's goods
  - if  $B' < 0$ , then  $-qB'$  units of the good are received in the current period, for a promise to repay  $-B'$  units next period
  - there is an equilibrium price function  $q(B', y)$  that makes  $q$  depend on both  $B'$  and  $y$

Earnings on the government portfolio are distributed (or, if negative, taxed) lump sum to households

When the government is not excluded from financial markets, the one-period national budget constraint is

$$c = y + B - q(B', y)B' \quad (3.155)$$

Here and below, a prime denotes a next period value or a claim maturing next period

To rule out Ponzi schemes, we also require that  $B \geq -Z$  in every period

- $Z$  is chosen to be sufficiently large that the constraint never binds in equilibrium

**Financial Markets** Foreign creditors

- are risk neutral
- know the domestic output stochastic process  $\{y_t\}$  and observe  $y_t, y_{t-1}, \dots$ , at time  $t$
- can borrow or lend without limit in an international credit market at a constant international interest rate  $r$
- receive full payment if the government chooses to pay
- receive zero if the government defaults on its one-period debt due

When a government is expected to default next period with probability  $\delta$ , the expected value of a promise to pay one unit of consumption next period is  $1 - \delta$ .

Therefore, the discounted expected value of a promise to pay  $B$  next period is

$$q = \frac{1 - \delta}{1 + r} \quad (3.156)$$

Next we turn to how the government in effect chooses the default probability  $\delta$

**Government's decisions** At each point in time  $t$ , the government chooses between

1. defaulting
2. meeting its current obligations and purchasing or selling an optimal quantity of one-period sovereign debt

Defaulting means declining to repay all of its current obligations

If the government defaults in the current period, then consumption equals current output

But a sovereign default has two consequences:

1. Output immediately falls from  $y$  to  $h(y)$ , where  $0 \leq h(y) \leq y$ 
  - it returns to  $y$  only after the country regains access to international credit markets
2. The country loses access to foreign credit markets

**Reentering international credit market** While in a state of default, the economy regains access to foreign credit in each subsequent period with probability  $\theta$

### Equilibrium

Informally, an equilibrium is a sequence of interest rates on its sovereign debt, a stochastic sequence of government default decisions and an implied flow of household consumption such that

1. Consumption and assets satisfy the national budget constraint
2. The government maximizes household utility taking into account
  - the resource constraint
  - the effect of its choices on the price of bonds
  - consequences of defaulting now for future net output and future borrowing and lending opportunities
3. The interest rate on the government's debt includes a risk-premium sufficient to make foreign creditors expect on average to earn the constant risk-free international interest rate

To express these ideas more precisely, consider first the choices of the government, which

1. enters a period with initial assets  $B$ , or what is the same thing, initial debt to be repaid now of  $-B$
2. observes current output  $y$ , and
3. chooses either
  - (a) to default, or
  - (b) to pay  $-B$  and set next period's debt due to  $-B'$

In a recursive formulation,

- state variables for the government comprise the pair  $(B, y)$

- $v(B, y)$  is the optimum value of the government's problem when at the beginning of a period it faces the choice of whether to honor or default
- $v_c(B, y)$  is the value of choosing to pay obligations falling due
- $v_d(y)$  is the value of choosing to default

$v_d(y)$  does not depend on  $B$  because, when access to credit is eventually regained, net foreign assets equal 0

Expressed recursively, the value of defaulting is

$$v_d(y) = u(h(y)) + \beta \int \{ \theta v(0, y') + (1 - \theta)v_d(y') \} p(y, y') dy'$$

The value of paying is

$$v_c(B, y) = \max_{B' \geq -Z} \left\{ u(y - q(B', y)B' + B) + \beta \int v(B', y') p(y, y') dy' \right\}$$

The three value functions are linked by

$$v(B, y) = \max\{v_c(B, y), v_d(y)\}$$

The government chooses to default when

$$v_c(B, y) < v_d(y)$$

and hence given  $B'$  the probability of default next period is

$$\delta(B', y) := \int \mathbb{1}\{v_c(B', y') < v_d(y')\} p(y, y') dy' \quad (3.157)$$

Given zero profits for foreign creditors in equilibrium, we can combine (3.156) and (3.157) to pin down the bond price function:

$$q(B', y) = \frac{1 - \delta(B', y)}{1 + r} \quad (3.158)$$

**Definition of equilibrium** An *equilibrium* is

- a pricing function  $q(B', y)$ ,
- a triple of value functions  $(v_c(B, y), v_d(y), v(B, y))$ ,
- a decision rule telling the government when to default and when to pay as a function of the state  $(B, y)$ , and
- an asset accumulation rule that, conditional on choosing not to default, maps  $(B, y)$  into  $B'$

such that

- The three Bellman equations for  $(v_c(B, y), v_d(y), v(B, y))$  are satisfied
- Given the price function  $q(B', y)$ , the default decision rule and the asset accumulation decision rule attain the optimal value function  $v(B, y)$ , and
- The price function  $q(B', y)$  satisfies equation (3.158)

### Computation

Let's now compute an equilibrium of Arellano's model

The equilibrium objects are the value function  $v(B, y)$ , the associated default decision rule, and the pricing function  $q(B', y)$

We'll use our code to replicate Arellano's results

After that we'll perform some additional simulations

The majority of the code below was written by Chase Coleman

It uses a slightly modified version of the algorithm recommended by Arellano

- The appendix to [Are08] recommends value function iteration until convergence, updating the price, and then repeating
- Instead, we update the bond price at every value function iteration step

The second approach is faster and the two different procedures deliver very similar results

Here is a more detailed description of our algorithm:

1. Guess a value function  $v(B, y)$  and price function  $q(B', y)$
2. At each pair  $(B, y)$ ,
  - update the value of defaulting  $v_d(y)$
  - update the value of continuing  $v_c(B, y)$
3. Update the value function  $v(B, y)$ , the default rule, the implied ex ante default probability, and the price function
4. Check for convergence. If converged, stop. If not, go to step 2.

We use simple discretization on a grid of asset holdings and income levels

The output process is discretized using Tauchen's quadrature method

The code can be found in the file `arellano_vfi.jl` from the `QuantEcon.applications` package but we repeat it here for convenience

(Results and discussion follow the code)

```
using QuantEcon: tauchen, MarkovChain, simulate

Define the main Arellano Economy type

"""
Arellano 2008 deals with a small open economy whose government
invests in foreign assets in order to smooth the consumption of
domestic households. Domestic households receive a stochastic
path of income.
"""


```

```

Fields
* `β::Real`: Time discounting parameter
* `γ::Real`: Risk aversion parameter
* `r::Real`: World interest rate
* `ρ::Real`: Autoregressive coefficient on income process
* `η::Real`: Standard deviation of noise in income process
* `θ::Real`: Probability of re-entering the world financial sector after default
* `ny::Int`: Number of points to use in approximation of income process
* `nB::Int`: Number of points to use in approximation of asset holdings
* `ygrid::Vector{Float64}`: This is the grid used to approximate income process
* `ydefgrid::Vector{Float64}`: When in default get less income than process
 would otherwise dictate
* `Bgrid::Vector{Float64}`: This is grid used to approximate choices of asset
 holdings
* `Π::Array{Float64, 2}`: Transition probabilities between income levels
* `vf::Array{Float64, 2}`: Place to hold value function
* `vd::Array{Float64, 2}`: Place to hold value function when in default
* `vc::Array{Float64, 2}`: Place to hold value function when choosing to
 continue
* `policy::Array{Float64, 2}`: Place to hold asset policy function
* `q::Array{Float64, 2}`: Place to hold prices at different pairs of (y, B')
* `defprob::Array{Float64, 2}`: Place to hold the default probabilities for
 pairs of (y, B')
"""

immutable ArellanoEconomy
 # Model Parameters
 β::Float64
 γ::Float64
 r::Float64
 ρ::Float64
 η::Float64
 θ::Float64

 # Grid Parameters
 ny::Int
 nB::Int
 ygrid::Array{Float64, 1}
 ydefgrid::Array{Float64, 1}
 Bgrid::Array{Float64, 1}
 Π::Array{Float64, 2}

 # Value function
 vf::Array{Float64, 2}
 vd::Array{Float64, 2}
 vc::Array{Float64, 2}
 policy::Array{Float64, 2}
 q::Array{Float64, 2}
 defprob::Array{Float64, 2}
end

"""
This is the default constructor for building an economy as presented
in Arellano 2008.

```

```

Arguments
* `β::Real(0.953)`: Time discounting parameter
* `γ::Real(2.0)`: Risk aversion parameter
* `r::Real(0.017)`: World interest rate
* `ρ::Real(0.945)`: Autoregressive coefficient on income process
* `η::Real(0.025)`: Standard deviation of noise in income process
* `θ::Real(0.282)`: Probability of re-entering the world financial sector
 after default
* `ny::Int(21)`: Number of points to use in approximation of income process
* `nB::Int(251)`: Number of points to use in approximation of asset holdings
"""

function ArellanoEconomy(;β=.953, γ=2., r=0.017, ρ=0.945, η=0.025, θ=0.282,
 ny=21, nB=251)

 # Create grids
 Bgrid = collect(linspace(-.4, .4, nB))
 mc = tauchen(ny, ρ, η)
 Π = mc.p
 ygrid = exp(mc.state_values)
 ydefgrid = min(.969 * mean(ygrid), ygrid)

 # Define value functions (Notice ordered different than Python to take
 # advantage of column major layout of Julia)
 vf = zeros(nB, ny)
 vd = zeros(1, ny)
 vc = zeros(nB, ny)
 policy = Array(Int, nB, ny)
 q = ones(nB, ny) .* (1 / (1 + r))
 defprob = Array(Float64, nB, ny)

 return ArellanoEconomy(β, γ, r, ρ, η, θ, ny, nB, ygrid, ydefgrid, Bgrid, Π,
 vf, vd, vc, policy, q, defprob)
end

u(ae::ArellanoEconomy, c) = c^(1 - ae.γ) / (1 - ae.γ)
_unpack(ae::ArellanoEconomy) =
 ae.β, ae.γ, ae.r, ae.ρ, ae.η, ae.θ, ae.ny, ae.nB
_unpackgrids(ae::ArellanoEconomy) =
 ae.ygrid, ae.ydefgrid, ae.Bgrid, ae.Π, ae.vf, ae.vd, ae.vc, ae.policy, ae.q, ae.defprob

Write the value function iteration

"""

This function performs the one step update of the value function for the
Arellano model-- Using current value functions and their expected value,
it updates the value function at every state by solving for the optimal
choice of savings

Arguments

* `ae::ArellanoEconomy`: This is the economy we would like to update the

```

```

 value functions for
* `EV::Matrix<Float64>`: Expected value function at each state
* `EVd::Matrix<Float64>`: Expected value function of default at each state
* `EVc::Matrix<Float64>`: Expected value function of continuing at each state

Notes

* This function updates value functions and policy functions in place.
"""

function one_step_update!(ae::ArellanoEconomy, EV::Matrix{Float64},
 EVd::Matrix{Float64}, EVc::Matrix{Float64})

 # Unpack stuff
 β, γ, r, ρ, η, θ, ny, nB = _unpack(ae)
 ygrid, ydefgrid, Bgrid, Π, vf, vd, vc, policy, q, defprob = _unpackgrids(ae)
 zero_ind = searchsortedfirst(Bgrid, 0.)

 for iy=1:ny
 y = ae.ygrid[iy]
 ydef = ae.ydefgrid[iy]

 # Value of being in default with income y
 defval = u(ae, ydef) + β*(θ*EVc[zero_ind, iy] + (1-θ)*EVd[1, iy])
 ae.vd[1, iy] = defval

 for ib=1:nB
 B = ae.Bgrid[ib]

 current_max = -1e14
 pol_ind = 0
 for ib_next=1:nB
 c = max(y - ae.q[ib_next, iy]*Bgrid[ib_next] + B, 1e-14)
 m = u(ae, c) + β * EV[ib_next, iy]

 if m > current_max
 current_max = m
 pol_ind = ib_next
 end
 end

 ae.vc[ib, iy] = current_max
 ae.policy[ib, iy] = pol_ind
 ae.vf[ib, iy] = defval > current_max ? defval : current_max
 end
 end

 Void
end

"""
This function takes the Arellano economy and its value functions and

```

```

policy functions and then updates the prices for each (y, B') pair

Arguments

* `ae::ArellanoEconomy` : This is the economy we would like to update the
 prices for

Notes

* This function updates the prices and default probabilities in place
"""

function compute_prices!(ae::ArellanoEconomy)
 # Unpack parameters
 β, γ, r, ρ, η, θ, ny, nB = _unpack(ae)

 # Create default values with a matching size
 vd_compat = repmat(ae.vd, nB)
 default_states = vd_compat .> ae.vc

 # Update default probabilities and prices
 copy!(ae.defprob, default_states * ae.Π')
 copy!(ae.q, (1 - ae.defprob) / (1 + r))

 Void
end

"""
This performs value function iteration and stores all of the data inside
the ArellanoEconomy type.

Arguments

* `ae::ArellanoEconomy` : This is the economy we would like to solve
* `tol::Float64(1e-8)` : Level of tolerance we would like to achieve
* `maxit::Int(10000)` : Maximum number of iterations

Notes

* This updates all value functions, policy functions, and prices in place.

"""

function vfi!(ae::ArellanoEconomy; tol=1e-8, maxit=10000)

 # Unpack stuff
 β, γ, r, ρ, η, θ, ny, nB = _unpack(ae)
 ygrid, ydefgrid, Bgrid, Π, vf, vd, vc, policy, q, defprob = _unpackgrids(ae)
 Πt = Π'

 # Iteration stuff
 it = 0
 dist = 10.

 # Allocate memory for update

```

```

V_upd = zeros(ae.vf)

while dist > tol && it < maxit
 it += 1

 # Compute expectations for this iterations
 # (We need Π' because of order value function dimensions)
 copy!(V_upd, ae.vf)
 EV = ae.vf * Π_t
 EVd = ae.vd * Π_t
 EVc = ae.vc * Π_t

 # Update Value Function
 one_step_update!(ae, EV, EVd, EVc)

 # Update prices
 compute_prices!(ae)

 dist = maxabs(V_upd - ae.vf)

 if it%25 == 0
 println("Finished iteration $(it) with dist of $(dist)")
 end
 end

 Void
end

"""
This function simulates the Arellano economy

Arguments

* `ae::ArellanoEconomy`: This is the economy we would like to solve
* `capT::Int`: Number of periods to simulate
* `y_init::Float64(mean(ae.ygrid))`: The level of income we would like to
 start with
* `B_init::Float64(mean(ae.Bgrid))`: The level of asset holdings we would like
 to start with

Returns

* `B_sim_val::Vector{Float64}`: Simulated values of assets
* `y_sim_val::Vector{Float64}`: Simulated values of income
* `q_sim_val::Vector{Float64}`: Simulated values of prices
* `default_status::Vector{Float64}`: Simulated default status
 (true if in default)

Notes

* This updates all value functions, policy functions, and prices in place.

"""

```

```

function QuantEcon.simulate(ae::ArellanoEconomy, capT::Int=5000;
 y_init=mean(ae.ygrid), B_init=mean(ae.Bgrid))

 # Get initial indices
 zero_index = searchsortedfirst(ae.Bgrid, 0.)
 y_init_ind = searchsortedfirst(ae.ygrid, y_init)
 B_init_ind = searchsortedfirst(ae.Bgrid, B_init)

 # Create a QE MarkovChain
 mc = MarkovChain(ae.Π)
 y_sim_indices = simulate(mc, capT+1; init=y_init_ind)

 # Allocate and Fill output
 y_sim_val = Array(Float64, capT+1)
 B_sim_val, q_sim_val = similar(y_sim_val), similar(y_sim_val)
 B_sim_indices = Array(Int, capT+1)
 default_status = fill(false, capT+1)
 B_sim_indices[1], default_status[1] = B_init_ind, false
 y_sim_val[1], B_sim_val[1] = ae.ygrid[y_init_ind], ae.Bgrid[B_init_ind]

 for t=1:capT
 # Get today's indexes
 yi, Bi = y_sim_indices[t], B_sim_indices[t]
 defstat = default_status[t]

 # If you are not in default
 if !defstat
 default_today = ae.vc[Bi, yi] < ae.vd[yi] ? true: false

 if default_today
 # Default values
 default_status[t] = true
 default_status[t+1] = true
 y_sim_val[t] = ae.ydefgrid[y_sim_indices[t]]
 B_sim_indices[t+1] = zero_index
 B_sim_val[t+1] = 0.
 q_sim_val[t] = ae.q[zero_index, y_sim_indices[t]]
 else
 default_status[t] = false
 y_sim_val[t] = ae.ygrid[y_sim_indices[t]]
 B_sim_indices[t+1] = ae.policy[Bi, yi]
 B_sim_val[t+1] = ae.Bgrid[B_sim_indices[t+1]]
 q_sim_val[t] = ae.q[B_sim_indices[t+1], y_sim_indices[t]]
 end

 # If you are in default
 else
 B_sim_indices[t+1] = zero_index
 B_sim_val[t+1] = 0.
 y_sim_val[t] = ae.ydefgrid[y_sim_indices[t]]
 q_sim_val[t] = ae.q[zero_index, y_sim_indices[t]]

 # With probability θ exit default status
 end
 end
end

```

```

 if rand() < ae.θ
 default_status[t+1] = false
 else
 default_status[t+1] = true
 end
end

return (y_sim_val[1:capT], B_sim_val[1:capT], q_sim_val[1:capT],
 default_status[1:capT])
end

```

## Results

Let's start by trying to replicate the results obtained in [Are08]

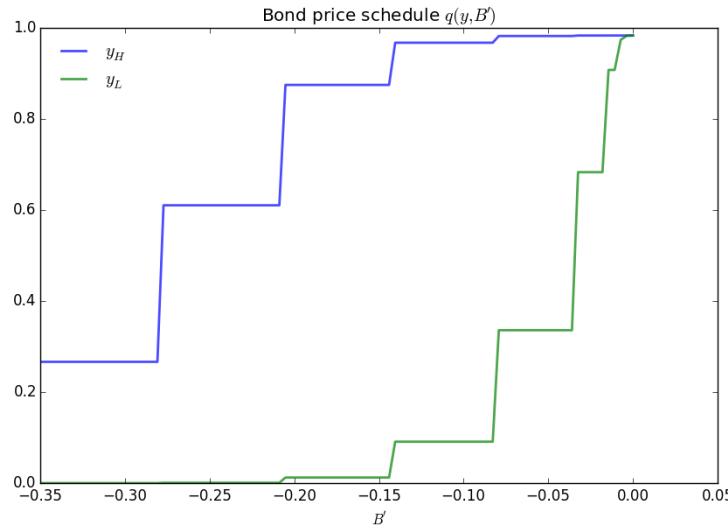
In what follows, all results are computed using Arellano's parameter values

The values can be seen in the function *ArellanoEconomy* shown above

- For example,  $r=0.017$  matches the average quarterly rate on a 5 year US treasury over the period 1983–2001

Details on how to compute the figures are reported as solutions to the exercises

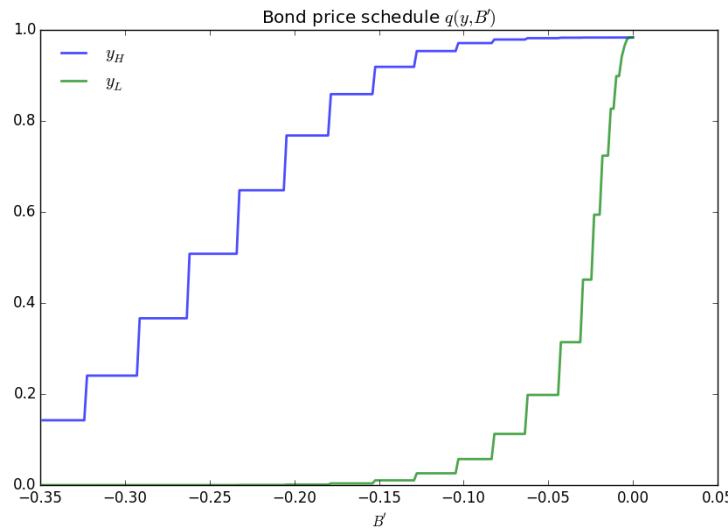
The first figure shows the bond price schedule and replicates Figure 3 of Arellano, where  $y_L$  and  $y_H$  are particular below average and above average values of output  $y$



- $y_L$  is 5% below the mean of the  $y$  grid values
- $y_H$  is 5% above the mean of the  $y$  grid values

The grid used to compute this figure was relatively coarse ( $ny, nB = 21, 251$ ) in order to match Arrelano's findings

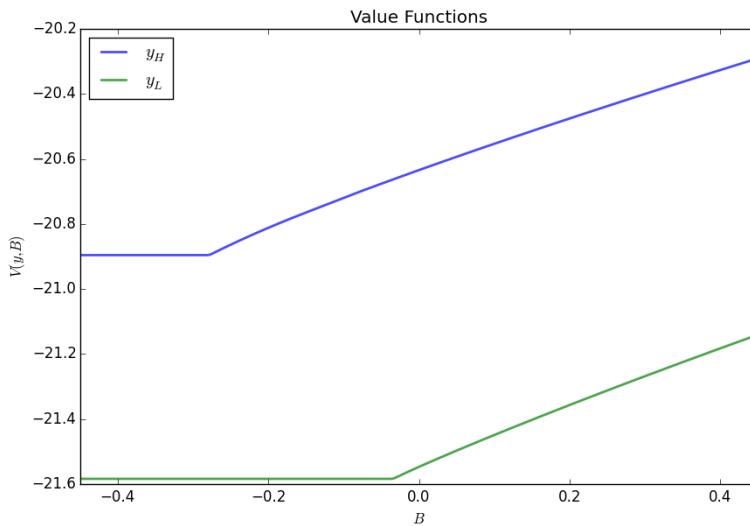
Here's the same relationships computed on a finer grid ( $ny, nB = 51, 551$ )



In either case, the figure shows that

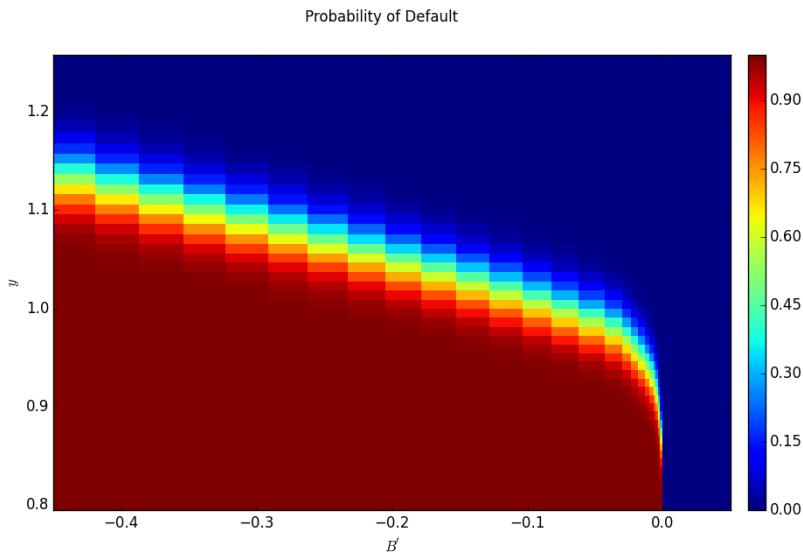
- Higher levels of debt (larger  $-B'$ ) induce larger discounts on the face value, which correspond to higher interest rates
- Lower income also causes more discounting, as foreign creditors anticipate greater likelihood of default

The next figure plots value functions and replicates the right hand panel of Figure 4 of [Are08]



We can use the results of the computation to study the default probability  $\delta(B', y)$  defined in (3.157)

The next plot shows these default probabilities over  $(B', y)$  as a heat map



As anticipated, the probability that the government chooses to default in the following period increases with indebtedness and falls with income

Next let's run a time series simulation of  $\{y_t\}$ ,  $\{B_t\}$  and  $q(B_{t+1}, y_t)$

The grey vertical bars correspond to periods when the economy is excluded from financial markets because of a past default

One notable feature of the simulated data is the nonlinear response of interest rates

Periods of relative stability are followed by sharp spikes in the discount rate on government debt

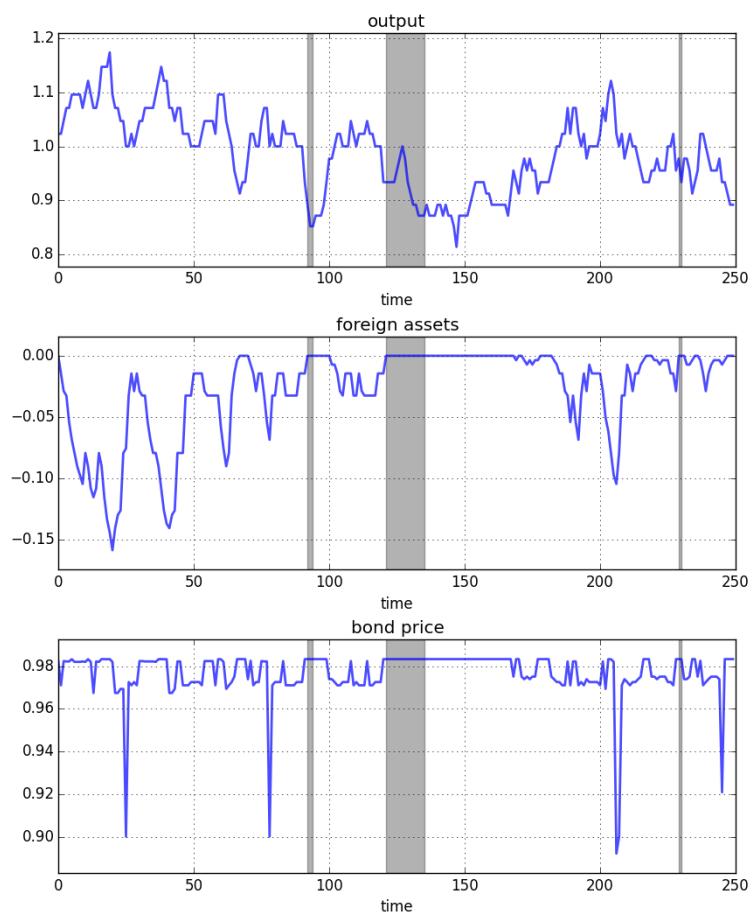
### Exercises

**Exercise 1** To the extent that you can, replicate the figures shown above

- Use the parameter values listed as defaults in the function *ArellanoEconomy*
- The time series will of course vary depending on the shock draws

### Solutions

Solution notebook



## REFERENCES

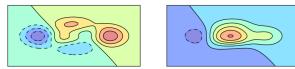
- [Aiy94] S Rao Aiyagari. Uninsured Idiosyncratic Risk and Aggregate Saving. *The Quarterly Journal of Economics*, 109(3):659–684, 1994.
- [AM05] D. B. O. Anderson and J. B. Moore. *Optimal Filtering*. Dover Publications, 2005.
- [AHMS96] E. W. Anderson, L. P. Hansen, E. R. McGrattan, and T. J. Sargent. Mechanics of Forming and Estimating Dynamic Linear Economies. In *Handbook of Computational Economics*. Elsevier, vol 1 edition, 1996.
- [Are08] Cristina Arellano. Default risk and income fluctuations in emerging economies. *The American Economic Review*, pages 690–712, 2008.
- [ACK10] Andrew Atkeson, Varadarajan V Chari, and Patrick J Kehoe. Sophisticated monetary policies\*. *The Quarterly journal of economics*, 125(1):47–89, 2010.
- [Bar79] Robert J Barro. On the Determination of the Public Debt. *Journal of Political Economy*, 87(5):940–971, 1979.
- [Bas05] Marco Bassetto. Equilibrium and government commitment. *Journal of Economic Theory*, 124(1):79–105, 2005.
- [BBZ15] Jess Benhabib, Alberto Bisin, and Shenghao Zhu. The wealth distribution in beasley economies with capital income risk. *Journal of Economic Theory*, 159:489–515, 2015.
- [BS79] L M Benveniste and J A Scheinkman. On the Differentiability of the Value Function in Dynamic Models of Economics. *Econometrica*, 47(3):727–732, 1979.
- [Bew77] Truman Bewley. The permanent income hypothesis: a theoretical formulation. *Journal of Economic Theory*, 16(2):252–292, 1977.
- [Bis06] C. M. Bishop. *Pattern Recognition and Machine Learning*. Springer, 2006.
- [Car01] Christopher D Carroll. A Theory of the Consumption Function, with and without Liquidity Constraints. *Journal of Economic Perspectives*, 15(3):23–45, 2001.
- [Cha98] Roberto Chang. Credible monetary policy in an infinite horizon model: recursive approaches. *Journal of Economic Theory*, 81(2):431–461, 1998.
- [CK90] Varadarajan V Chari and Patrick J Kehoe. Sustainable plans. *Journal of Political Economy*, pages 783–802, 1990.

- [Col90] Wilbur John Coleman. Solving the Stochastic Growth Model by Policy-Function Iteration. *Journal of Business & Economic Statistics*, 8(1):27–29, 1990.
- [CC08] J. D. Cryer and K-S. Chan. *Time Series Analysis*. Springer, 2nd edition edition, 2008.
- [Dea91] Angus Deaton. Saving and Liquidity Constraints. *Econometrica*, 59(5):1221–1248, 1991.
- [DP94] Angus Deaton and Christina Paxson. Intertemporal Choice and Inequality. *Journal of Political Economy*, 102(3):437–467, 1994.
- [DH10] Wouter J Den Haan. Comparison of solutions to the incomplete markets model with aggregate uncertainty. *Journal of Economic Dynamics and Control*, 34(1):4–27, 2010.
- [DLP13] Y E Du, Ehud Lehrer, and A D Y Pauzner. Competitive economy as a ranking device over networks. submitted, 2013.
- [Dud02] R M Dudley. *Real Analysis and Probability*. Cambridge Studies in Advanced Mathematics. Cambridge University Press, 2002.
- [EG87] Robert F Engle and Clive W J Granger. Co-integration and Error Correction: Representation, Estimation, and Testing. *Econometrica*, 55(2):251–276, 1987.
- [ES13] David Evans and Thomas J Sargent. *History dependent public policies*. Oxford University Press, 2013.
- [EH01] G W Evans and S Honkapohja. *Learning and Expectations in Macroeconomics*. Frontiers of Economic Research. Princeton University Press, 2001.
- [FSTD15] Pablo Fajgelbaum, Edouard Schaal, and Mathieu Taschereau-Dumouchel. Uncertainty traps. Technical Report, National Bureau of Economic Research, 2015.
- [Fri56] M. Friedman. *A Theory of the Consumption Function*. Princeton University Press, 1956.
- [GW10] Marc P Giannoni and Michael Woodford. Optimal target criteria for stabilization policy. Technical Report, National Bureau of Economic Research, 2010.
- [Hal78] Robert E Hall. Stochastic Implications of the Life Cycle-Permanent Income Hypothesis: Theory and Evidence. *Journal of Political Economy*, 86(6):971–987, 1978.
- [HM82] Robert E Hall and Frederic S Mishkin. The Sensitivity of Consumption to Transitory Income: Estimates from Panel Data on Households. *National Bureau of Economic Research Working Paper Series*, 1982.
- [Ham05] James D Hamilton. What's real about the business cycle? *Federal Reserve Bank of St. Louis Review*, pages 435–452, 2005.
- [HS08] L P Hansen and T J Sargent. *Robustness*. Princeton University Press, 2008.
- [HS13] L P Hansen and T J Sargent. *Recursive Models of Dynamic Linear Economies*. The Gorman Lectures in Economics. Princeton University Press, 2013.
- [HR87] Lars Peter Hansen and Scott F Richard. The Role of Conditioning Information in Deducing Testable. *Econometrica*, 55(3):587–613, May 1987.

- [HS00] Lars Peter Hansen and Thomas J Sargent. Wanting robustness in macroeconomics. *Manuscript, Department of Economics, Stanford University.*, 2000.
- [HK79] J. Michael Harrison and David M. Kreps. Martingales and arbitrage in multi-period securities markets. *Journal of Economic Theory*, 20(3):381–408, June 1979.
- [HL96] John Heaton and Deborah J Lucas. Evaluating the effects of incomplete markets on risk sharing and asset pricing. *Journal of Political Economy*, pages 443–487, 1996.
- [HLL96] O Hernandez-Lerma and J B Lasserre. *Discrete-Time Markov Control Processes: Basic Optimality Criteria*. number Vol 1 in Applications of Mathematics Stochastic Modelling and Applied Probability. Springer, 1996.
- [HP92] Hugo A Hopenhayn and Edward C Prescott. Stochastic Monotonicity and Stationary Distributions for Dynamic Economies. *Econometrica*, 60(6):1387–1406, 1992.
- [HR93] Hugo A Hopenhayn and Richard Rogerson. Job Turnover and Policy Evaluation: A General Equilibrium Analysis. *Journal of Political Economy*, 101(5):915–938, 1993.
- [Hug93] Mark Huggett. The risk-free rate in heterogeneous-agent incomplete-insurance economies. *Journal of Economic Dynamics and Control*, 17(5-6):953–969, 1993.
- [Haggstrom02] Olle Häggström. *Finite Markov chains and algorithmic applications*. volume 52. Cambridge University Press, 2002.
- [JYC88] Robert J. Shiller John Y. Campbell. The Dividend-Price Ratio and Expectations of Future Dividends and Discount Factors. *Review of Financial Studies*, 1(3):195–228, 1988.
- [Janich94] K Jänich. *Linear Algebra*. Springer Undergraduate Texts in Mathematics and Technology. Springer, 1994.
- [Kam12] Takashi Kamihigashi. Elementary results on solutions to the bellman equation of dynamic programming: existence, uniqueness, and convergence. Technical Report, Kobe University, 2012.
- [Kuh13] Moritz Kuhn. Recursive Equilibria In An Aiyagari-Style Economy With Permanent Income Shocks. *International Economic Review*, 54:807–835, 2013.
- [KP80] Finn E Kydland and Edward C Prescott. Dynamic optimal taxation, rational expectations and optimal control. *Journal of Economic Dynamics and Control*, 2:79–91, 1980.
- [LM94] A Lasota and M C MacKey. *Chaos, Fractals, and Noise: Stochastic Aspects of Dynamics*. Applied Mathematical Sciences. Springer-Verlag, 1994.
- [LL01] Martin Lettau and Sydney Ludvigson. Consumption, Aggregate Wealth, and Expected Stock Returns. *Journal of Finance*, 56(3):815–849, 06 2001.
- [LL04] Martin Lettau and Sydney C. Ludvigson. Understanding Trend and Cycle in Asset Values: Reevaluating the Wealth Effect on Consumption. *American Economic Review*, 94(1):276–299, March 2004.
- [LS12] L Ljungqvist and T J Sargent. *Recursive Macroeconomic Theory*. MIT Press, 3 edition, 2012.

- [Luc78] Robert E Lucas, Jr. Asset prices in an exchange economy. *Econometrica: Journal of the Econometric Society*, 46(6):1429–1445, 1978.
- [LP71] Robert E Lucas, Jr and Edward C Prescott. Investment under uncertainty. *Econometrica: Journal of the Econometric Society*, pages 659–681, 1971.
- [LS83] Robert E Lucas, Jr and Nancy L Stokey. Optimal Fiscal and Monetary Policy in an Economy without Capital. *Journal of monetary Economics*, 12(3):55–93, 1983.
- [MS89] Albert Marcet and Thomas J Sargent. Convergence of Least-Squares Learning in Environments with Hidden State Variables and Private Information. *Journal of Political Economy*, 97(6):1306–1322, 1989.
- [MdRV10] V Filipe Martins-da-Rocha and Yiannis Vailakis. Existence and Uniqueness of a Fixed Point for Local Contractions. *Econometrica*, 78(3):1127–1141, 2010.
- [MCWG95] A Mas-Colell, M D Whinston, and J R Green. *Microeconomic Theory*. volume 1. Oxford University Press, 1995.
- [McC70] J J McCall. Economics of Information and Job Search. *The Quarterly Journal of Economics*, 84(1):113–126, 1970.
- [MT09] S P Meyn and R L Tweedie. *Markov Chains and Stochastic Stability*. Cambridge University Press, 2009.
- [MF02] Mario J Miranda and P L Fackler. *Applied Computational Economics and Finance*. Cambridge: MIT Press, 2002.
- [MB54] F. Modigliani and R. Brumberg. Utility analysis and the consumption function: An interpretation of cross-section data. In K.K Kurihara, editor, *Post-Keynesian Economics*. 1954.
- [Nea99] Derek Neal. The Complexity of Job Mobility among Young Men. *Journal of Labor Economics*, 17(2):237–261, 1999.
- [Par99] Jonathan A Parker. The Reaction of Household Consumption to Predictable Changes in Social Security Taxes. *American Economic Review*, 89(4):959–973, 1999.
- [Put05] Martin L Puterman. *Markov decision processes: discrete stochastic dynamic programming*. John Wiley & Sons, 2005.
- [Rab02] Guillaume Rabault. When do borrowing constraints bind? Some new results on the income fluctuation problem. *Journal of Economic Dynamics and Control*, 26(2):217–245, 2002.
- [Ram27] F. P. Ramsey. A Contribution to the theory of taxation. *Economic Journal*, 37(145):47–61, 1927.
- [Rei09] Michael Reiter. Solving heterogeneous-agent models by projection and perturbation. *Journal of Economic Dynamics and Control*, 33(3):649–665, 2009.
- [Rom05] Steven Roman. *Advanced linear algebra*. volume 3. Springer, 2005.
- [Rus96] John Rust. Numerical dynamic programming in economics. *Handbook of computational economics*, 1:619–729, 1996.
- [Sar87] T J Sargent. *Macroeconomic Theory*. Academic Press, 2nd edition, 1987.

- [SE77] Jack Schechtman and Vera L S Escudero. Some results on “an income fluctuation problem”. *Journal of Economic Theory*, 16(2):151–166, 1977.
- [Sch69] Thomas C Schelling. Models of Segregation. *American Economic Review*, 59(2):488–493, 1969.
- [Shi95] A N Shiriaev. *Probability*. Graduate texts in mathematics. Springer. Springer, 2nd edition, 1995.
- [SLP89] N L Stokey, R E Lucas, and E C Prescott. *Recursive Methods in Economic Dynamics*. Harvard University Press, 1989.
- [Sto89] Nancy L Stokey. Reputation and time consistency. *The American Economic Review*, pages 134–139, 1989.
- [STY04] Kjetil Storesletten, Christopher I Telmer, and Amir Yaron. Consumption and risk sharing over the life cycle. *Journal of Monetary Economics*, 51(3):609–633, 2004.
- [Sun96] R K Sundaram. *A First Course in Optimization Theory*. Cambridge University Press, 1996.
- [Tau86] George Tauchen. Finite state markov-chain approximations to univariate and vector autoregressions. *Economics Letters*, 20(2):177–181, 1986.
- [Woo03] Michael Woodford. *Interest and Prices: Foundations of a Theory of Monetary Policy*. Princeton University Press, 2003.
- [YS05] G Alastair Young and Richard L Smith. *Essentials of statistical inference*. Cambridge University Press, 2005.



Acknowledgements: These lectures have benefitted greatly from comments and suggestion from our colleagues, students and friends. Special thanks go to Anmol Bhandari, Jeong-Hun Choi, Chase Coleman, David Evans, Chenghan Hou, Doc-Jin Jang, Spencer Lyon, Qingyin Ma, Matthew McKay, Tomohito Okabe, Alex Olssen, Nathan Palmer and Yixiao Zhou.

## INDEX

### A

A Simple Optimal Growth Model, 217  
An Introduction to Asset Pricing, 282  
An Introduction to Job Search, 148  
AR, 419  
ARMA, 417, 419  
ARMA Processes, 414

### B

Bellman Equation, 390

### C

Central Limit Theorem, 162, 168  
Intuition, 169  
Multivariate Case, 172  
CLT, 162  
Complex Numbers, 418  
Continuous State Markov Chains, 313  
Covariance Stationary, 415  
Covariance Stationary Processes, 413  
    AR, 416  
    MA, 416

### D

Discrete Dynamic Programming, 260  
Dynamic Programming, 217, 219  
    Computation, 221  
    Shortest Paths, 144  
    Theory, 220  
    Unbounded Utility, 221  
    Value Function Iteration, 220, 221

### E

Eigenvalues, 97, 109  
Eigenvectors, 97, 109  
Ergodicity, 114, 128

### F

Finite Markov Asset Pricing  
    Lucas Tree, 289  
Finite Markov Chains, 113–115  
    Stochastic Matrices, 114  
Fixed Point Theory, 332

### G

General Linear Processes, 415

### H

History Dependent Public Policies, 459, 460  
    Competitive Equilibrium, 462  
    Ramsey Timing, 461  
    Sequence of Governments Timing, 461  
    Timing Protocols, 461

### I

Irreducibility and Aperiodicity, 114, 121

### K

Kalman Filter, 198  
    Programming Implementation, 205  
    Recursive Procedure, 203

### L

Law of Large Numbers, 162, 163  
    Illustration, 164  
    Multivariate Case, 172  
    Proof, 163

Linear Algebra, 97

    Differentiating Linear and Quadratic Forms, 112  
    Eigenvalues, 109  
    Eigenvectors, 109  
    Matrices, 102  
    Matrix Norms, 111  
    Neumann's Theorem, 111

- Positive Definite Matrices, 112
- Series Expansions, 111
- Spectral Radius, 112
- Vectors, 98
- Linear State Space Models, 175
  - Distributions, 181, 182
  - Ergodicity, 186
  - Martingale Difference Shocks, 177
  - Moments, 181
  - Moving Average Representations, 181
  - Prediction, 191
  - Seasonals, 179
  - Stationarity, 186
  - Time Trends, 180
  - Univariate Autoregressive Processes, 178
  - Vector Autoregressions, 179
- LLN, 162
- LQ Control, 234
  - Infinite Horizon, 244
  - Optimality (Finite Horizon), 237
- Lucas Model, 329
  - Assets, 329
  - Computation, 333
  - Consumers, 329
  - Dynamic Program, 330
  - Equilibrium Constraints, 331
  - Equilibrium Price Function, 331
  - Pricing, 330
  - Solving, 332
- M
  - MA, 419
  - Marginal Distributions, 114, 119
  - Markov Asset Pricing
    - Overview, 282
  - Markov Chains, 115
    - Calculating Stationary Distributions, 126
    - Continuous State, 313
    - Convergence to Stationarity, 127
    - Cross-Sectional Distributions, 121
    - Ergodicity, 128
    - Forecasting Future Values, 128
    - Future Probabilities, 121
    - Irreducibility, Aperiodicity, 121
    - Marginal Distributions, 119
    - Simulation, 116
    - Stationary Distributions, 125
  - Matrix
- Determinants, 107
- Inverse, 107
- Maps, 105
- Operations, 103
- Solving Systems of Equations, 105
- McCall Model, 366
- Modeling
  - Career Choice, 346
- Models
  - Linear State Space, 177
  - Lucas Asset Pricing, 328
  - Markov Asset Pricing, 282
  - McCall, 366
  - On-the-Job Search, 355
  - Permanent Income, 297
  - Pricing, 283
  - Schelling's Segregation Model, 157
- N
- Neumann's Theorem, 111
- Nonparametric Estimation, 434
- O
- On-the-Job Search, 355, 356
  - Model, 356
  - Model Features, 356
  - Parameterization, 357
  - Programming Implementation, 358
  - Solving for Policies, 364
- Optimal Growth
  - Model, 218
  - Policy Function, 225
  - Policy Function Approach, 218
- Optimal Savings, 377
  - Computation, 379
  - Problem, 378
  - Programming Implementation, 381
- Optimal Taxation in an LQ Economy, 442
- Orthogonal Projection, 134
- P
- Periodograms, 429, 430
  - Computation, 431
  - Interpretation, 430
- Permanent Income Model, 297
  - Hall's Representation, 303
  - Savings Problem, 297
- Positive Definite Matrices, 112

Pricing Models, 282, 283  
 Risk Aversion, 283  
 Risk Neutral, 283  
 Programming  
 Dangers, 227  
 Iteration, 231  
 Writing Reusable Code, 227

## R

Ramsey Problem, 460, 463  
 Computing, 464  
 Credible Policy, 477  
 Optimal Taxation, 442  
 Recursive Representation, 467  
 Time Inconsistency, 473  
 Two Subproblems, 465  
 Rational Expectations Equilibrium, 273  
 Competitive Equilibrium (w. Adjustment Costs), 276  
 Computation, 278  
 Definition, 275  
 Planning Problem Approach, 279  
 Robustness, 390

## S

Schelling Segregation Model, 157  
 Smoothing, 429, 434  
 Spectra, 429  
 Estimation, 429  
 Spectra, Estimation  
 AR(1) Setting, 440  
 Fast Fourier Transform, 430  
 Pre-Filtering, 438  
 Smoothing, 434, 435, 438  
 Spectral Analysis, 413, 414, 418  
 Spectral Densities, 419  
 Spectral Density, 420  
 interpretation, 420  
 Inverting the Transformation, 422  
 Mathematical Theory, 422  
 Spectral Radius, 112  
 Stationary Distributions, 114, 125  
 Stochastic Matrices, 114

## U

Unbounded Utility, 221

## V

Value Function Iteration, 220

Vectors, 97, 98  
 Inner Product, 99  
 Linear Independence, 102  
 Norm, 99  
 Operations, 98  
 Span, 100

## W

White Noise, 415, 419  
 Wold's Decomposition, 416