

Complete Proofs

1 Highest Posterior Density (HPD) Intervals

We report the complete proofs of Theorems 1-3.

Theorem 1. Given a prior distribution $\text{Beta}(a, b)$ for μ and a Knowledge Graph (KG) correctness annotation process, where $0 < \tau_S < n_S$, the $1 - \alpha$ Highest Posterior Density (HPD) interval is the smallest (l, u) interval satisfying the condition $F(u) - F(l) = 1 - \alpha$.

PROOF. When $0 < \tau_S < n_S$, the beta posterior $\text{Beta}(a + \tau_S, b + n_S - \tau_S)$ is unimodal and continuous over the $[0, 1]$ interval, guaranteeing the existence of a Credible Interval (CrI). The method of Lagrange multipliers allows us to find the smallest interval satisfying $F(u) - F(l) = 1 - \alpha$. That is, we minimize the following Lagrangian function:

$$\mathcal{L} = (u - l) + \lambda \left(\int_l^u f(\mu | G_S) d\mu - (1 - \alpha) \right) \quad (1)$$

where λ is the Lagrange multiplier.

The first-order partial derivatives of \mathcal{L} in ∂l and ∂u lead to:

$$\begin{aligned} \frac{\partial \mathcal{L}}{\partial l} &= -1 - \lambda f(l | G_S) = 0 \Rightarrow f(l | G_S) = -\frac{1}{\lambda}; \\ \frac{\partial \mathcal{L}}{\partial u} &= 1 + \lambda f(u | G_S) = 0 \Rightarrow f(u | G_S) = -\frac{1}{\lambda}. \end{aligned}$$

Since $f(\cdot | G_S) > 0$, it follows that $\lambda < 0$. To confirm that this yields a minimum, we compute the second-order partial derivatives:

$$\begin{aligned} \frac{\partial^2 \mathcal{L}}{(\partial l)^2} &= -\lambda \frac{\partial f(l | G_S)}{\partial l}; \\ \frac{\partial^2 \mathcal{L}}{(\partial u)^2} &= \lambda \frac{\partial f(u | G_S)}{\partial u}; \\ \frac{\partial^2 \mathcal{L}}{(\partial l)(\partial u)} &= \frac{\partial^2 \mathcal{L}}{(\partial u)(\partial l)} = 0. \end{aligned}$$

The posterior distribution is unimodal, hence $\partial f(l | G_S) / \partial l > 0$, leading to $\partial^2 \mathcal{L} / (\partial l)^2 > 0$. Similarly, $\partial f(u | G_S) / \partial u < 0$ implies $\partial^2 \mathcal{L} / (\partial u)^2 > 0$. Therefore, the Hessian matrix is positive definite, confirming that the interval minimizing the Lagrangian in Eq. (1) is the smallest satisfying $F(u) - F(l) = 1 - \alpha$. \square

Theorem 2. Under the assumptions of Theorem 1, the $1 - \alpha$ HPD interval is unique.

PROOF. Ab absurdo, let (l^*, u^*) represent the HPD interval, and assume there exists another interval (l', u') such that:

- (1) $|u' - l'| = |u^* - l^*|$;
- (2) $\int_{l'}^{u'} f(\mu | G_S) d\mu = 1 - \alpha$.

Denote the mode of the posterior distribution as ω . Since the posterior is unimodal, Th. 1 ensures that the HPD interval contains the mode, meaning $l^* < \omega < u^*$. Following the same reasoning, the alternative interval (l', u') must also satisfy $l' < \omega < u'$.

Given that $(l', u') \neq (l^*, u^*)$ but $|u' - l'| = |u^* - l^*|$, there are two possible cases:

- (i) $l' < l^*$ and $u' < u^*$;
- (ii) $l' > l^*$ and $u' > u^*$.

We prove case (i); the proof for case (ii) is analogous.

Since both intervals satisfy the condition $F(u) - F(l) = 1 - \alpha$

$$\int_{l'}^{u'} f(\mu | G_S) d\mu = \int_{l^*}^{u^*} f(\mu | G_S) d\mu$$

Under case (i), we can rewrite the left-hand side as:

$$\int_{l'}^{u'} f(\mu | G_S) d\mu = \int_{l'}^{l^*} f(\mu | G_S) d\mu + \int_{l^*}^{u'} f(\mu | G_S) d\mu$$

and the right-hand side as:

$$\int_{l^*}^{u^*} f(\mu | G_S) d\mu = \int_{l^*}^{u'} f(\mu | G_S) d\mu + \int_{u'}^{u^*} f(\mu | G_S) d\mu$$

Eliding the common term $\int_{l^*}^{u'} f(\mu | G_S) d\mu$ leads to:

$$\int_{l'}^{l^*} f(\mu | G_S) d\mu = \int_{u'}^{u^*} f(\mu | G_S) d\mu$$

that is a contradiction. Indeed, the posterior distribution is unimodal; the Probability Density Function (PDF) $f(\mu | G_S)$ is increasing before the mode and decreasing after it. Therefore, $\int_{l'}^{l^*} f(\mu | G_S) d\mu < \int_{u'}^{u^*} f(\mu | G_S) d\mu$, as l' lies outside the HPD region while u' lies inside. Despite l' and u' are equidistant from l^* and u^* , respectively, the densities of (l', l^*) and (u', u^*) differ due to their positions relative to the HPD region. Therefore, no other interval can have the same properties as the HPD interval, proving its uniqueness. \square

Theorem 3. When the beta posterior for μ is unimodal and symmetric, the $1 - \alpha$ HPD interval is equivalent to the $1 - \alpha$ Equal-Tailed (ET) interval.

PROOF. Consider a unimodal and symmetric beta posterior. Let ω represent the mode of this posterior. By the symmetry property, for any $\delta > 0$, we have $f(\omega - \delta | G_S) = f(\omega + \delta | G_S)$. Therefore, the $1 - \alpha$ HPD interval is symmetric around ω – i.e., the lower and upper bounds, l and u , are equidistant from ω .

Let us rewrite $\int_l^u f(\mu | G_S) d\mu = 1 - \alpha$ as $\int_l^\omega f(\mu | G_S) d\mu + \int_\omega^u f(\mu | G_S) d\mu = 1 - \alpha$. Due to the symmetry of the posterior:

$$\int_l^\omega f(\mu | G_S) d\mu = \int_\omega^u f(\mu | G_S) d\mu = \frac{1 - \alpha}{2}$$

By construction, it follows that:

$$\int_0^l f(\mu | G_S) d\mu = \int_u^1 f(\mu | G_S) d\mu = \frac{\alpha}{2}$$

corresponding to the Cumulative Distribution Function (CDF) of the $\alpha/2$ and $1 - \alpha/2$ quantiles, respectively. Thus, the $1 - \alpha$ HPD interval coincides with the $1 - \alpha$ ET interval. \square