## Team 2

I pledge my Honor that I have abided by the Stevens Honor System.

Kush Parmar, Jo-Anne Rivera, Simerjeet Mudhar

Query 2

1. 
$$\sum_{i=1}^{n} e_{i}^{2} = \sum_{i=1}^{n} (y_{i} - \widehat{y}_{i})^{2} = \sum_{i=1}^{n} [y_{i} - (\beta_{0} + \beta_{1} x_{i})]^{2} =$$

Knowing that  $\beta_0$  and  $\beta_1$  represent the constants of the regression model we can take derivation in regards to them.

Derivation for  $\beta_0$ 

$$\begin{split} &\frac{\partial S}{\partial \beta_0} \left[ \sum_{i=1}^n \left[ y_i - (\beta_0 + \beta_1 x_i) \right]^2 \right] = 0 \rightarrow \text{chain rule} \quad y_i - (\beta_0 + \beta_1 x_i) = u \rightarrow \\ &\frac{\partial S}{\partial u} \left( 2u \right) \rightarrow \frac{\partial u}{\partial \beta_0} \left( 2(y_i - \beta_0 - \beta_1 x_i) \right) \rightarrow 2(\frac{\partial u}{\partial \beta_0} y_i - \frac{\partial u}{\partial \beta_0} \beta_0 - \frac{\partial u}{\partial \beta_0} \beta_1 x_i)) \\ &= 2(0 - 1 - 0) = -2 \\ &\frac{\partial S}{\partial \beta_0} \left[ \sum_{i=1}^n -2 \left[ y_i - \beta_0 - \beta_1 x_i \right] \right] = 0 \rightarrow -2 \sum_{i=1}^n y_i + 2 \sum_{i=1}^n \beta_0 + 2 \sum_{i=1}^n \beta_1 x_i = 0 \rightarrow \\ &- \sum_{i=1}^n y_i + \sum_{i=1}^n \beta_0 + \sum_{i=1}^n \beta_1 x_i = 0 \rightarrow \sum_{i=1}^n \beta_0 + \sum_{i=1}^n \beta_1 x_i = \sum_{i=1}^n y_i \rightarrow \\ &n\beta_0 + \sum_{i=1}^n \beta_1 x_i = \sum_{i=1}^n y_i \rightarrow \beta_0 = \overline{y} - \beta_1 \overline{x} \end{split}$$

Derivation for  $\beta_1$ 

$$\frac{\partial S}{\partial \beta_{1}} \left[ \sum_{i=1}^{n} \left[ y_{i} - (\beta_{0} + \beta_{1} x_{i}) \right]^{2} \right] = 0 \rightarrow \text{chain rule } y_{i} - (\beta_{0} + \beta_{1} x_{i}) = u \rightarrow \frac{\partial S}{\partial u} \left( 2u \right) \rightarrow \frac{\partial u}{\partial \beta_{1}} \left( 2(y_{i} - \beta_{0} - \beta_{1} x_{i}) \right) \rightarrow 2(\frac{\partial u}{\partial \beta_{1}} y_{i} - \frac{\partial u}{\partial \beta_{1}} \beta_{0} - \frac{\partial u}{\partial \beta_{1}} \beta_{1} x_{i}))$$

$$= 2(0 - 0 - x_{i}) = -2x_{i}$$

$$\frac{\partial S}{\partial \beta_{1}} \left[ \sum_{i=1}^{n} -2 \left[ y_{i} - \beta_{0} - \beta_{1} x_{i} \right) (x_{i}) \right] = 0 \rightarrow -2 \sum_{i=1}^{n} y_{i} x_{i} + 2 \sum_{i=1}^{n} \beta_{0} x_{i} + 2 \sum_{i=1}^{n} \beta_{1} x_{i}^{2} = 0$$

$$\beta_{0} \sum_{i=1}^{n} x_{i} + \beta_{1} \sum_{i=1}^{n} x_{i}^{2} = \sum_{i=1}^{n} y_{i} x_{i} \text{ One possible answer is } \beta_{1} \sum_{i=1}^{n} x_{i}^{2} = \sum_{i=1}^{n} y_{i} x_{i} - \beta_{0} \sum_{i=1}^{n} x_{i}$$

Alternatively substitute the equation found for  $\beta_0$ 

$$(\overline{y} - \beta_1 \overline{x}) n \overline{x} + \beta_1 \sum_{i=1}^n x_i^2 = \sum_{i=1}^n y_i x_i \rightarrow \text{Simplifies to } \beta_1 = \frac{\sum_{i=1}^n y_i x_i - n \overline{y}(\overline{x})}{\sum_{i=1}^n x_i^2 - n(\overline{x})^2}$$

2. Our derivatives in number one were set to zero, hence we can see our equations come out to he:

$$\beta_0 = \overline{y} - \beta_1 \overline{x}$$

$$\beta_{1} n (\overline{x})^{2} = n(\overline{y})(\overline{x}) - \beta_{0} n (\overline{x}) \rightarrow \beta_{1} (\overline{x})^{2} + \beta_{0} (\overline{x}) = (\overline{y})(\overline{x}) \rightarrow (\overline{y})(\overline{x}) - \beta_{1} (\overline{x})^{2} = \beta_{0} (\overline{x}) \rightarrow (\overline{y}) - \beta_{1} (\overline{x}) = \beta_{0}$$

3. 
$$\beta_{1} = \frac{\begin{vmatrix} \frac{n}{n} & \Sigma y_{i} \\ \sum x_{i} & \sum x_{i} y_{i} \end{vmatrix}}{\begin{vmatrix} n & \Sigma x_{i} \\ \sum x_{i} & \sum x_{i}^{2} \end{vmatrix}} = \frac{n \sum x_{i} y_{i} - \sum x_{i} y_{i}}{n \sum x_{i}^{2} - \sum x_{i}^{2}} = \frac{\sum x_{i} y_{i} (n-1)}{\sum x_{i}^{2} (n-1)} = \frac{\sum x_{i} y_{i}}{\sum x_{i}^{2}} = \frac{\sum y_{i}}{\sum x_{i}}$$

Since 
$$\beta_0 = \overline{y} - \beta_1(\overline{x})$$
 and  $\overline{x} = \frac{1}{n}\sum x_i$ ,  $\overline{y} = \frac{1}{n}\sum y_i$ :

$$\beta_0 = \overline{y} - \frac{\sum y_i}{\sum x_i} (\overline{x}) = \overline{y} - \frac{\sum y_i}{\sum x_i} (\frac{1}{n} \sum x_i) = \overline{y} - \sum y_i (\frac{1}{n}) = \overline{y} - \overline{y} = 0$$

$$(\beta_0, \beta_1) = (0, \frac{\Sigma y_i}{\Sigma x_i})$$

4

$$\begin{vmatrix} \frac{\partial S^2}{\partial \beta_0^2} SSE & \frac{\partial S^2}{\partial \beta_0 \partial \beta_1} SSE \\ \frac{\partial S^2}{\partial \beta_0 \partial \beta_1} SSE & \frac{\partial S^2}{\partial \beta_1^2} SSE \end{vmatrix} = \begin{vmatrix} n & 0 \\ 0 & \sum\limits_{i=1}^n x_i^2 \end{vmatrix}$$

$$\frac{\partial S}{\partial \beta_0^2} (n \beta_0 + n \beta_1 \overline{x} - n \overline{y})$$

$$(\frac{\partial S}{\partial \beta_0} n \beta_0 + \frac{\partial S}{\partial \beta_0} n \beta \overline{x} - \frac{\partial S}{\partial \beta_0} n \overline{y})$$

$$(n 1 + 0 - 0) = n$$

$$\frac{\partial S}{\partial \beta_0 \partial \beta_1} (\overline{y} - \beta_1 \overline{x}) n \overline{x} + \beta_1 \sum_{i=1}^n x_i^2 - \sum_{i=1}^n y_i x_i 
\frac{\partial S}{\partial \beta_0 \partial \beta_1} [(\overline{y} - \beta_1 \overline{x}) n \overline{x} + n \beta_1 (\overline{x})^2 - n (\overline{x} * \overline{y})] 
[((\overline{y} - \beta_1 \overline{x}) n \overline{x}) \frac{\partial S}{\partial \beta_0} + (n \beta_1 (\overline{x})^2) \frac{\partial S}{\partial \beta_0} - (n (\overline{x} * \overline{y})) \frac{\partial S}{\partial \beta_0}] = [0 + 0 - 0] = 0$$

$$\frac{\partial S}{\partial \beta_{1}^{2}} (\overline{y} - \beta_{1} \overline{x}) n \overline{x} + \beta_{1} \sum_{i=1}^{n} x_{i}^{2} - \sum_{i=1}^{n} y_{i} x_{i}$$

$$\frac{\partial S}{\partial \beta_{1}} (\overline{y} - \beta_{1} \overline{x}) n \overline{x} + (\frac{\partial S}{\partial \beta_{1}}) \beta_{1} \sum_{i=1}^{n} x_{i}^{2} - \frac{\partial S}{\partial \beta_{1}} \sum_{i=1}^{n} y_{i} x_{i}$$

$$\frac{\partial S}{\partial \beta_{1}} (\overline{y} - \beta_{1} \overline{x}) n \overline{x} + (\frac{\partial S}{\partial \beta_{1}}) \beta_{1} \sum_{i=1}^{n} x_{i}^{2} - \frac{\partial S}{\partial \beta_{1}} \sum_{i=1}^{n} y_{i} x_{i} = (0 + \sum_{i=1}^{n} x_{i}^{2} - 0) = \sum_{i=1}^{n} x_{i}^{2}$$

5. 
$$\begin{vmatrix} n & 0 \\ 0 & \sum_{i=1}^{n} x_i^2 \end{vmatrix} \rightarrow \text{critical point - } (\beta_0 = 0, \beta_1 = \frac{\sum y_i}{\sum x_i})$$

The Hessian matrix is symmetric. Since n represents the number of values in a given dataset as seen with  $\sum_{i=1}^{n} x_i^2$  n is a non-zero, positive value. The given Hessian matrix is in its reduced form. Since the pivots are positive, we can conclude that this is a positive definite matrix.

6. According to the multivariable second derivative test, critical point is a local minimum if:

$$\frac{\partial S}{\partial \beta_0^2}(\beta_0, \beta_1) > 0 \text{ and } \frac{\partial S}{\partial \beta_0^2}(\beta_0, \beta_1) \cdot \frac{\partial S}{\partial \beta_1^2}(\beta_0, \beta_1) - \frac{\partial S}{\partial \beta_0}(\beta_0, \beta_1) > 0$$

$$n > 0$$
 and  $n \cdot \sum_{i=1}^{n} x_i^2 - 0 > 0$ 

Both of these are true, due to how n is a positive value since it represents the number of values in a dataset. Moreover  $(n \cdot \sum_{i=1}^{n} x_i^2 - 0)$  represents a positive value since the

$$\sum_{i=1}^{n} x_{i}^{2}$$
 summation will always result in a positive value since  $x_{i}$  is squared.

Also, according to the slides, upon proving that the Hassian matrix is positive definite, the critical point is proven to be the unique minimizer of SSE.