

Particle Filters

M. Sami Fadali
Professor of EE
University of Nevada

Outline

- Monte Carlo integration.
- Particle filter.
- Importance sampling.
- Degeneracy
- Resampling
- Example.

Monte Carlo Integration

- Numerical integration using randomly generated points.
- No “curse of dimensionality”
- Use to calculate moments (integrals)
- Integral of $g(x)$, continuous $[a, b]$

$$\begin{aligned}\int_a^b g(x)dx &= (b-a) \int_a^b \frac{1}{b-a} g(x)dx \\ &= (b-a)E\{g(X)\}, X \sim U[a, b]\end{aligned}$$

MC Integration Procedure

- Generate $X_i \sim U[a, b], i = 1, \dots, N$
- Calculate $Y_i = g(X_i), i = 1, \dots, N$
- Calculate the estimate of $E\{g(X)\}$

$$\bar{Y} = \frac{1}{N} \sum_{i=1}^N Y_i$$

- Calculate the estimate of $\int_a^b g(x)dx$
 $(b-a)\bar{Y}$

Properties of \bar{Y}

- Use $X_i \sim U[a, b], i = 1, \dots, N$
- Unbiased

$$E\{\bar{Y}\} = \frac{1}{N} \sum_{i=1}^N E\{g(X_i)\} = E\{g(X)\}$$

- Consistent (shown earlier)

$$\text{Var}\{\bar{Y}\} = \frac{1}{N} \text{Var}\{g(X)\}$$

Example (Hogg et al., p. 289)

- Estimate π using Monte Carlo integration
- Use $g(X) = 4\sqrt{1-x^2}$

$$\int_0^1 g(x) dx = \left[2x\sqrt{1-x^2} + 2\sin^{-1}(x) \right]_0^1 = \pi$$

- Generate $X_i \sim U[0, 1], i = 1, \dots, N$
- Calculate $Y_i = 4\sqrt{1-X_i^2}, i = 1, \dots, N$
- \bar{Y} = consistent estimate of π (estimate improves as N increases)

MC Integration

- Factorization
 $g(x) = f(x)p(x), x \in [a, b]$
 $p(x)$ = pdf (integral is unity, nonnegative)
- Generate $X_i \sim p(x), i = 1, \dots, N$
- Calculate $Y_i = f(X_i), i = 1, \dots, N$
- Calculate the estimate of $E\{f(X)\} = \int_a^b g(x) dx$

$$\bar{Y} = \frac{1}{N} \sum_{i=1}^N Y_i$$

Importance Sampling

- Not always possible to generate samples from $p(x)$
- Generate samples from a similar pdf $q(x)$
 $p(x) > 0 \Rightarrow q(x) > 0, \quad \forall x \in \mathcal{R}^n$
(same support)

$$\begin{aligned} \int_a^b g(x) dx &= \int_a^b f(x)p(x) dx \\ &= \int_a^b f(x) \frac{p(x)}{q(x)} q(x) dx \end{aligned}$$

Use MC Integration

- Generate $X_i \sim q(x), i = 1, \dots, N$
- Calculate $Y_i = f(X_i)W(X_i), i = 1, \dots, N$

$$W(X_i) = \frac{p(X_i)}{q(X_i)} / \sum_{i=1}^N \frac{p(X_i)}{q(X_i)}$$

$$\bar{Y} = \frac{1}{N} \sum_{i=1}^N Y_i = \frac{1}{N} \sum_{i=1}^N f(X_i)W(X_i)$$

•

•9

Relation to Estimation

- Fundamental theorem of estimation theory
- Minimum mean-square error estimator:

$$\hat{x}(k) = E\{x(k)|z^*(k)\}$$

$$= \int x(k)p_{x|z}(x(k)|z^*(k))dx(k)$$

$$z^*(k) = \text{col}\{z(i), i = 0, 1, \dots, k\}.$$
- $z^*(k)$ = all measurements up to and including time k

•

•10

Particle Filter

- View random samples as “particles”
- Can handle non-Gaussian statistics and multimodal distributions.
- Approximate a continuous pdf with appropriately weighted random samples.
- A posteriori conditional density

$$p(x_k|z_k^*) \approx \sum_{i=1}^n W_k^{(i)} \delta(x_k - x_k^{(i)}), \sum_{i=1}^n W_k^{(i)} = 1$$

•

•11

Approximation of PDF Using Weighted Random Samples

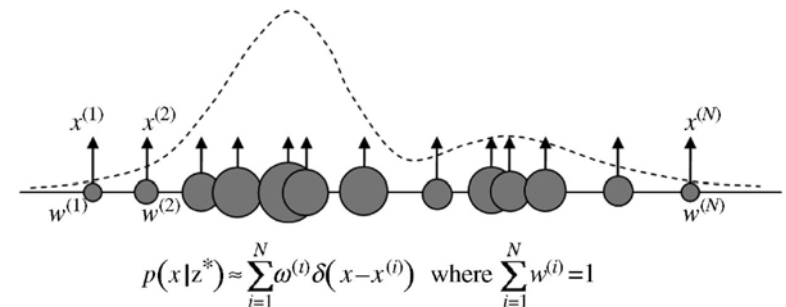


Figure 7.10
© John Wiley & Sons, Inc. All rights reserved.

•

•12

Model

- Nonlinear process model

$$\mathbf{x}_{k+1} = \phi(\mathbf{x}_k, \mathbf{w}_k)$$
- Measurement model

$$\mathbf{z}_k = \mathbf{h}(\mathbf{x}_k, \mathbf{v}_k)$$
- Noise processes are not necessarily Gaussian.
- \mathbf{z}_k^* = measurements up to and including time k

• 13

Recursive Bayesian Filter

- Predictor

$$p(\mathbf{x}_{k+1} | \mathbf{z}_k^*) = \int p(\mathbf{x}_{k+1} | \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{z}_k^*) d\mathbf{x}_k$$
- Corrector

$$p(\mathbf{x}_k | \mathbf{z}_k^*) = \kappa_k p(\mathbf{z}_k | \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{z}_{k-1}^*)$$
- Initial condition $p(\mathbf{x}_0)$ for recursion

$$p(\mathbf{x}_k | \mathbf{z}_k^*) = \kappa_k p(\mathbf{z}_k | \mathbf{x}_k) \int p(\mathbf{x}_k | \mathbf{x}_{k-1}) p(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*) d\mathbf{x}_{k-1}$$

• 14

Corrector Modified

- Rewrite using Markov property

$$p(\mathbf{x}_k | \mathbf{z}_k^*) = \kappa_k p(\mathbf{z}_k | \mathbf{x}_k) \int p(\mathbf{x}_k | \mathbf{x}_{k-1}) p(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*) d\mathbf{x}_{k-1}$$

$$p(\mathbf{x}_k | \mathbf{z}_{k-1}^*) = \int p(\mathbf{x}_k, \mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*) d\mathbf{x}_{k-1}$$

$$p(\mathbf{x}_k | \mathbf{x}_{k-1}, \mathbf{z}_{k-1}^*) p(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*) = \frac{p(\mathbf{x}_k, \mathbf{x}_{k-1}, \mathbf{z}_{k-1}^*) p(\mathbf{x}_{k-1}, \mathbf{z}_{k-1}^*)}{p(\mathbf{x}_{k-1}, \mathbf{z}_{k-1}^*) p(\mathbf{z}_{k-1}^*)} = p(\mathbf{x}_k, \mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*)$$

• 15

Introduce Importance Density

$$\begin{aligned}
 p(\mathbf{x}_k | \mathbf{z}_k^*) &= \kappa_k p(\mathbf{z}_k | \mathbf{x}_k) \int p(\mathbf{x}_k | \mathbf{x}_{k-1}) p(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*) d\mathbf{x}_{k-1} \\
 &\propto p(\mathbf{z}_k | \mathbf{x}_k) \int p(\mathbf{x}_k | \mathbf{x}_{k-1}) \frac{p(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*)}{q(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*)} q(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*) d\mathbf{x}_{k-1}
 \end{aligned}$$

• 16

Choose Importance Density

- Choose to allow the factorization

$$q(\mathbf{x}_k | \mathbf{z}_k^*) = q(\mathbf{x}_k | \mathbf{x}_{k-1}, \mathbf{z}_k^*) \cdot q(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*)$$

- Observe that

$$q(\mathbf{x}_k | \mathbf{x}_{k-1}, \mathbf{z}_k^*) = q(\mathbf{x}_k | \mathbf{x}_{k-1}, \mathbf{z}_k)$$

- Substitute

$$\begin{aligned} q(\mathbf{x}_k | \mathbf{z}_k^*) &= q(\mathbf{x}_k | \mathbf{x}_{k-1}, \mathbf{z}_k) \cdot q(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*) \\ \frac{p(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*)}{q(\mathbf{x}_k | \mathbf{z}_k^*)} q(\mathbf{x}_k | \mathbf{z}_k^*) &= \frac{p(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*)}{q(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*)} \cdot \frac{q(\mathbf{x}_k | \mathbf{x}_{k-1}, \mathbf{z}_k) \cdot q(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*)}{q(\mathbf{x}_k | \mathbf{x}_{k-1}, \mathbf{z}_k)} \end{aligned}$$

• 17

Back to Integral

$$p(\mathbf{x}_k | \mathbf{z}_k^*) = \kappa_k p(\mathbf{z}_k | \mathbf{x}_k) \int p(\mathbf{x}_k | \mathbf{x}_{k-1}) p(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*) d\mathbf{x}_{k-1}$$

$$\propto p(\mathbf{z}_k | \mathbf{x}_k) \int p(\mathbf{x}_k | \mathbf{x}_{k-1}) \tilde{w}_{k-1} q(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*) d\mathbf{x}_{k-1}$$

$$\tilde{w}_{k-1} = \frac{p(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*)}{q(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*)}$$

$$\begin{aligned} p(\mathbf{x}_k | \mathbf{z}_k^*) &\propto \sum_{i=1}^n \tilde{w}_k^{(i)} \delta(\mathbf{x}_k - \mathbf{x}_k^{(i)}) \\ &= \sum_{i=1}^n \frac{p(\mathbf{z}_k | \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{x}_{k-1})}{q(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*)} \tilde{w}_{k-1}^{(i)} \delta(\mathbf{x}_k - \mathbf{x}_k^{(i)}) \end{aligned}$$

• 18

Weights

$$\tilde{w}_{k-1} \propto \frac{p(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*)}{q(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*)}$$

$$\tilde{w}_k^{(i)} = \frac{p(\mathbf{z}_k | \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{x}_{k-1})}{q(\mathbf{x}_{k-1} | \mathbf{z}_{k-1}^*)} \tilde{w}_{k-1}^{(i)}$$

Monte Carlo Estimate (normalize)

$$W_k^{(j)} = \frac{\tilde{w}_k^{(j)}}{\sum_{i=1}^N \tilde{w}_{k-1}^{(i)}}$$

• 19

Substitute

$$p(\mathbf{x}_k | \mathbf{z}_k^*) = \sum_{i=1}^n W_k^{(i)} \delta(\mathbf{x}_k - \mathbf{x}_k^{(i)})$$

$$W_k^{(i)} = W_{k-1}^{(i)} \frac{p(\mathbf{z}_k | \mathbf{x}_k^{(i)}) p(\mathbf{x}_k | \mathbf{x}_{k-1}^{(i)})}{q(\mathbf{x}_k | \mathbf{x}_{k-1}^{(i)}, \mathbf{z}_k)}$$

• 20

Choice of Importance Density

- Critical step in particle filter design.
- Optimal importance density: minimizes the variance of importance weights conditioned upon $\mathbf{x}_{k-1}^{(i)}$ and \mathbf{z}_k

$$q(\mathbf{x}_k | \mathbf{x}_{k-1}^{(i)}, \mathbf{z}_k) = p(\mathbf{x}_k | \mathbf{x}_{k-1}^{(i)}, \mathbf{z}_k)$$

•

•21

Choice of Importance Density

- Popular choice: bootstrap filter

$$q(\mathbf{x}_k | \mathbf{x}_{k-1}^{(i)}, \mathbf{z}_k) = p(\mathbf{x}_k | \mathbf{x}_{k-1}^{(i)})$$

$$W_k^{(i)} \propto W_{k-1}^{(i)} p(\mathbf{z}_k | \mathbf{x}_k^{(i)})$$

- Assume that the weights are resampled every cycle such that they become uniformly distributed.

$$W_k^{(i)} = p(\mathbf{z}_k | \mathbf{x}_k^{(i)})$$

•

•22

Importance Sampling

- Unlike Monte Carlo and ensemble filters, does not use a simple distribution to generate the random samples.
- Generate points that fit a general pdf.
- Importance Density: proposal density that can easily generate samples.
- Algorithm used: Sequential Importance Sampling (SIS)

•

•23

Sequential Importance Sampling

- Weight recursion

$$W_k^{(i)} = W_{k-1}^{(i)} \frac{p(\mathbf{z}_k | \mathbf{x}_k^{(i)}) p(\mathbf{x}_k | \mathbf{x}_{k-1}^{(i)})}{q(\mathbf{x}_k | \mathbf{x}_{k-1}^{(i)}, \mathbf{z}_k)}$$

$$p(\mathbf{z}_k | \mathbf{x}_k^{(i)}) = \text{likelihood function}$$

$$p(\mathbf{x}_k | \mathbf{x}_{k-1}^{(i)}) = \text{transition prior}$$

$$q(\mathbf{x}_k | \mathbf{x}_{k-1}^{(i)}, \mathbf{z}_k) = \text{importance density}$$

•

•24

Estimates

- Based on the approximate discretize pdf
- State Estimate

$$\hat{\mathbf{x}}_k^+ = \sum_{i=1}^n W_k^{(i)} \mathbf{x}_k^{(i)}$$

- Error Covariance

$$P_k^+ = \sum_{i=1}^n W_k^{(i)} (\mathbf{x}_k^{(i)} - \hat{\mathbf{x}}_k^+) (\mathbf{x}_k^{(i)} - \hat{\mathbf{x}}_k^+)^T$$

• 25

SIS Algorithm

- Input $\{\mathbf{x}_k^{(i-1)}, W_k^{(i-1)}, i = 1, \dots, N\}, \mathbf{z}_k$

- Output $\{\mathbf{x}_k^{(i)}, W_k^{(i)}, i = 1, \dots, N\}$

$Sum_w = 0$

for $i = 1:N$

Draw $\mathbf{x}_k^{(i)}$ from $q(\mathbf{x}_k | \mathbf{x}_{k-1}^{(i)}, \mathbf{z}_k)$

$W_k^{(i)} = W_{k-1}^{(i)} p(\mathbf{z}_k | \mathbf{x}_k^{(i)}) p(\mathbf{x}_k | \mathbf{x}_{k-1}^{(i)}) / q(\mathbf{x}_k | \mathbf{x}_{k-1}^{(i)}, \mathbf{z}_k)$

$Sum_w = Sum_w + W_k^{(i)}$

end

for $i = 1:N$

$W_k^{(i)} = W_k^{(i)} / Sum_w$

end

• 26

Degeneracy Phenomenon

- Normalized weights tend to concentrate into one particle after a number of recursions (all other particles degenerate)
- Measure of degeneracy N_{eff}

$$\hat{N}_{eff} = 1 / \sum_{i=1}^n (w_k^{(i)})^2, 1 \leq \hat{N}_{eff} \leq N$$

$$\hat{N}_{eff} = N, w_k^{(i)} = 1/N, i = 1, 2, \dots, N$$

$$\hat{N}_{eff} = 1, w_k^{(j)} = 1,$$

$$w_k^{(i)} = 0, i \neq j, i = 1, \dots, j-1, j+1, \dots, N$$

• 27

Degeneracy After 2 Cycles

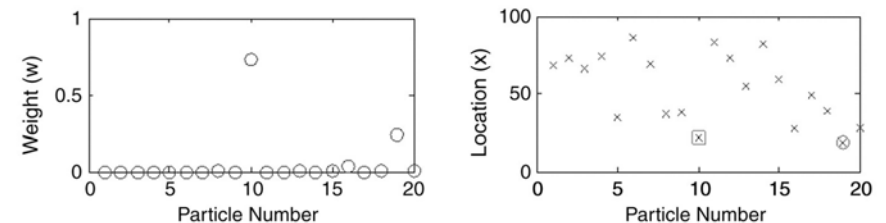


Figure 7.11
© John Wiley & Sons, Inc. All rights reserved.

• 28

Resampling

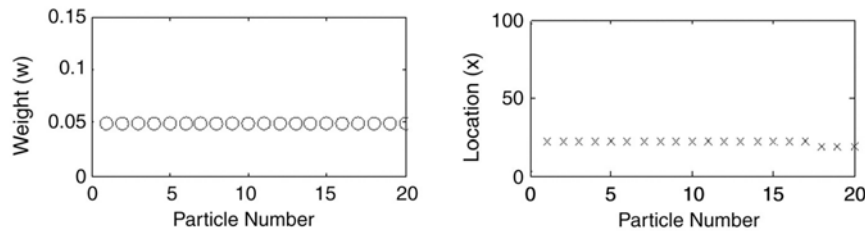


Figure 7.12
© John Wiley & Sons, Inc. All rights reserved.

• 29

Resampling

- Required if $N_{eff} < N_{thresh}$
- Particles get more meaningful weights (uniformly distributed weights)
- Eliminates samples with low weights and produces a set with uniform weights.

•

• 30

Resampling Algorithms Idea

- Generate a random number $u_1 \sim U[0,1]$
- Add the weights until their sum exceeds u_1 at which point choose the corresponding particle \mathbf{x}_k^i as a resampled point \mathbf{x}_k^{j*} and save the index i^j of its parent
- Can show that the resampled particles' pdf converges to $p(\mathbf{x}_k|\mathbf{z}_k)$

•

• 31

Resampling Algorithm

- Input $\{\mathbf{x}_k^{(i)}, w_k^{(i)}, i = 1, \dots, N\}$
 - Output $\{\mathbf{x}_k^{(j)*}, w_k^{(j)}, i^j, i = 1, \dots, N\}$
- ```

 $c_1 = w_k^{(1)}$
for $i = 2:N$
 $c_i = c_{i-1} + w_k^{(i)}$
end
Draw a starting point $u_1 \sim U[0, N^{-1}]$
for $j = 1:N$
 $u_j = u_1 + N^{-1}(j - 1)$
While $u_j > c_i$
 $i = i + 1$
end
 $\mathbf{x}_k^{(j)*} = \mathbf{x}_k^{(i)}$
 $w_k^{(j)} = N^{-1}$
 $i^j = i$
end

```

•

• 32



## Particle Filter Algorithm (Simon)

- Generate  $N$  initial particles  $\mathbf{x}_0^{+(i)}, i = 1, \dots, N$ , using the pdf  $p(\mathbf{x}^0)$
- Predictor step to obtain
 
$$\mathbf{x}_k^{-(i)} = \phi(\mathbf{x}_{k-1}^{+(i)}, \mathbf{w}_{k-1}^{(i)}), i = 1, \dots, N$$
- Compute the weights  $W_k^{(i)}$  then scale them so that their sum is unity.
- Corrector step:
 
$$\mathbf{x}_k^{+(i)}, i = 1, \dots, N, \quad \mathbf{x}_k^{+(i)} = \frac{1}{N} \sum_{i=1}^N \mathbf{x}_k^{+(i)}$$

• 33

## SIR Particle Filter Algorithm

- Input  $\{\mathbf{x}_{k-1}^{(i)}, i = 1, \dots, N\}, \mathbf{z}_k$
  - Output  $\{\mathbf{x}_k^{(i)}, i = 1, \dots, N\}$
- ```

Sum_w = 0
for i = 1:N
    Draw  $\mathbf{x}_k^{(i)} \sim p(\mathbf{x}_k | \mathbf{x}_{k-1}^{(i)})$ 
    Calculate  $W_k^{(i)} = W_{k-1}^{(i)} p(\mathbf{z}_k | \mathbf{x}_k^{(i)})$ 
    Sum_w = Sum_w + W_k^{(i)}
End
    for j = 1:N
         $W_k^{(j)} = W_k^{(j)} / \text{Sum}_w$ 
    end
     $\mathbf{x}_k^{(j)*} = \mathbf{x}_k^{(j)}$ 
Resample using SIR
    
```

• 34

References

- R. G. Brown and P. Y. C. Hwang, *Introduction to Random Signals and Applied Kalman Filtering*, 4ed, J. Wiley, NY, 2012.
- R. V. Hogg, J. W. McKean, A. T. Craig, *Introduction to Mathematical Statistics*, 6th Ed., Pearson/Prentice-Hall, Upper Saddle River, NJ, 2005.
- B. Ristic, S. Arulampalam, N. Gordon, *Beyond the Kalman Filter: Particle Filters for Tracking Applications*, Artech House, Boston, 2004.

• 35