

# Statistical Machine Learning

4주차

담당: 18기 신인수

# 3주차 과제 우수자

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김재훈님😊

코드를화면공유해주시고, 간단히설명해주세요!

리마인드: **매주 수요일 22시까지** 과제를 제출해주셔야 합니다! (github)

과제제출시, 파일명은 기존.ipynb 파일이름을 유지하되, '\_이름'만  
추가해주세요 😊 (ML\_week4\_HW\_방서연)

## 0. Conditional Expectation

### 1. Regression

### 2. Linear Regression

### 3. Regularization

## 4. Regression Diagnostics

# 0. Conditional Expectation

# Conditional Expectation

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## Conditional Distribution (조건부 분포)

- Recap) 조건부 확률:  $B$ 라는 조건이 주어졌을 때  $A$ 의 확률

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

- $Y$ 라는 조건이 주어졌을 때,  $X$ 의 분포

$$f(X|Y) = \frac{f(x, y)}{f(y)} = \frac{P(X \cap Y)}{P(Y)}$$

# Conditional Expectation

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## Conditional Expectation (조건부 기댓값)

- $Y$ 라는 조건이 주어졌을 때,  $X$ 의 기댓값

$$E(X|Y) = \int_{-\infty}^{\infty} x f(x|y) dx, X \text{ is continuous random variable}$$

$$E(X|Y) = \sum_x x f(x|y), X \text{ is discrete random variable}$$

## Conditional Variance (조건부 분산)

- $Y$ 라는 조건이 주어졌을 때,  $X$ 의 분산 (역시 편차 제곱의 평균)

$$\text{Var}(X|Y) = E[(X - E(X|Y))^2 | Y]$$

# Conditional Expectation

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## Properties

$$E(X) = E[E(X|Y)]$$

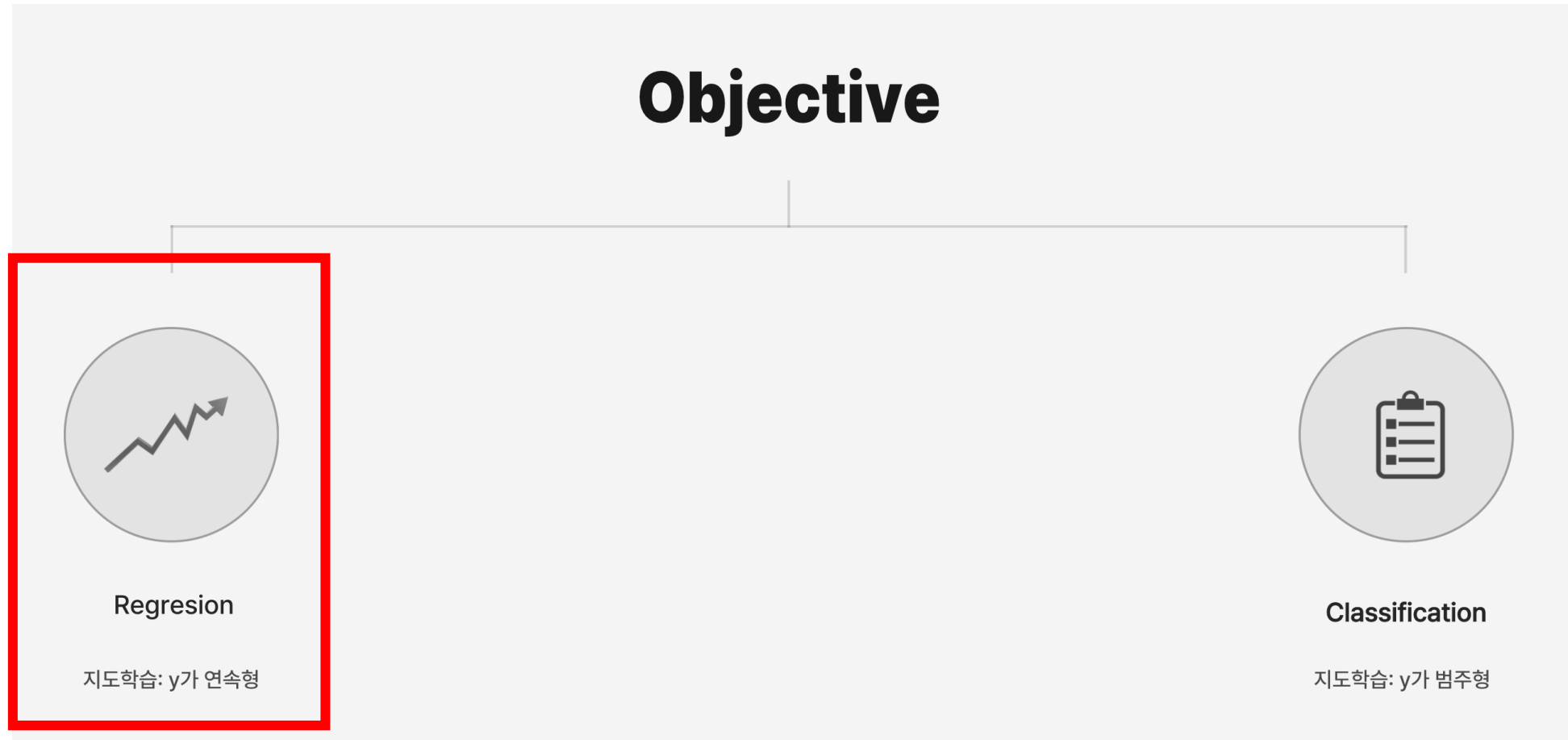
$$\text{Var}(X) = \text{Var}[E(X|Y)] + E[\text{Var}(X|Y)]$$

# 1. Regression



# Supervised Learning

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# Regression

Galton의 height 데이터.

X : 아버지의 키

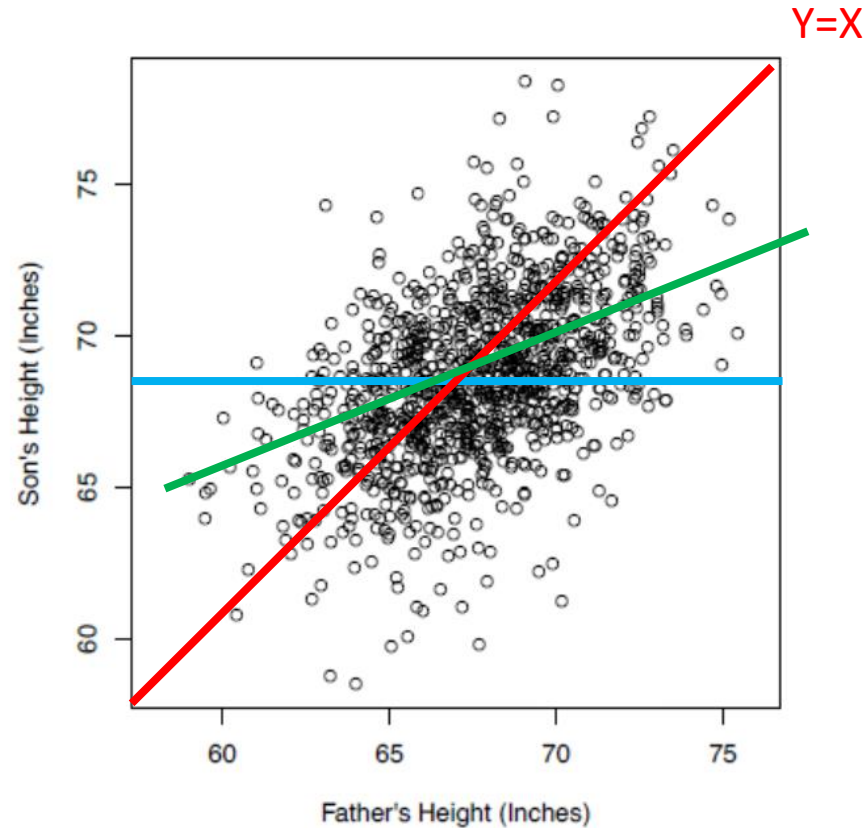
Y : 아들의 키

두 변수의 관계를

$$Y = aX + b$$

로 표현하려고 할 때, a와 b를  
어떻게 정할 것인가?

→ 원칙이 필요하다



이런 식으로 아무렇게나  
그어버리면 곤란하다.

여전히 우리는 조건부  
기댓값을 모델링하고 싶다.

# Regression

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## Recap: Bias and Variance tradeoff & MSE

$$\begin{aligned}MSE &= E[y - \hat{f}(x)]^2 \\&= E[f(x) + \epsilon - \hat{f}(x)]^2 \\&= E[f(x) - \hat{f}(x)]^2 + E[\epsilon]^2 + 2E[\epsilon(f(x) - \hat{f}(x))] \\&= [Var(\hat{f}(x)) + Bias(\hat{f}(x))^2] + Var(\epsilon) \\&= Reducible Error + Irreducible Error\end{aligned}$$

- 원칙1: **Bias = 0으로 만들고(unbiased estimator)**, 분산이 최소인 직선을(추정량) 구하겠다 → Linear Regression (1장)
- 원칙2: Bias가 있더라도 전체 MSE를 최소화시키는 직선을(추정량) 구하겠다 → LASSO, Ridge 등의 regularization (2장)

# Regression

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어떤 원칙으로 직선을 만들게 되는가?

$$SST = SSR + SSE$$

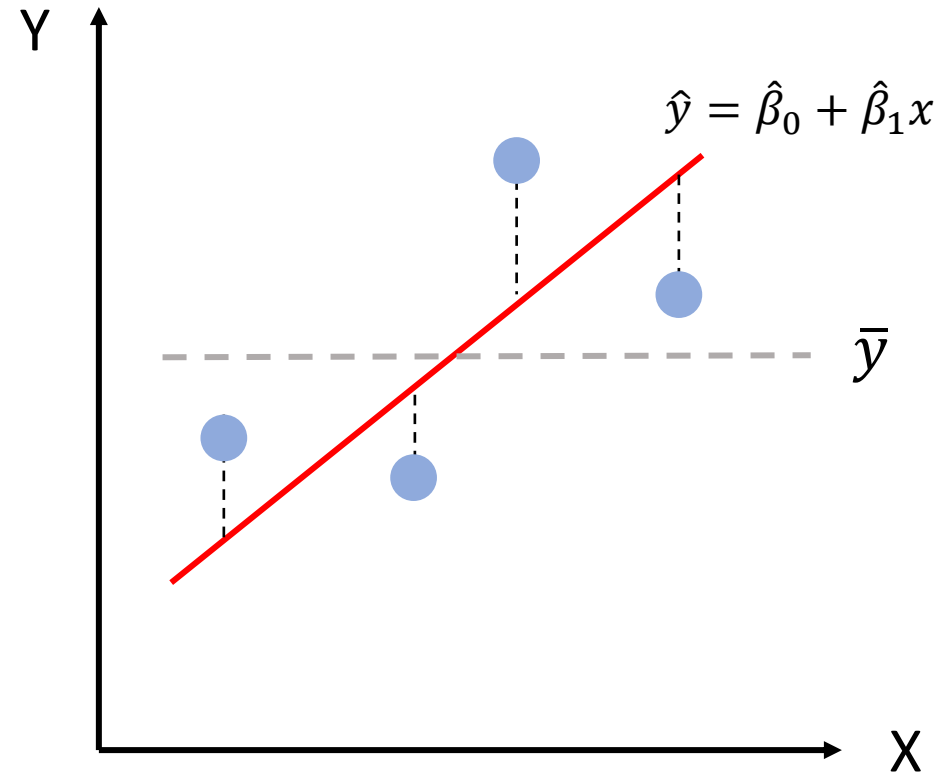
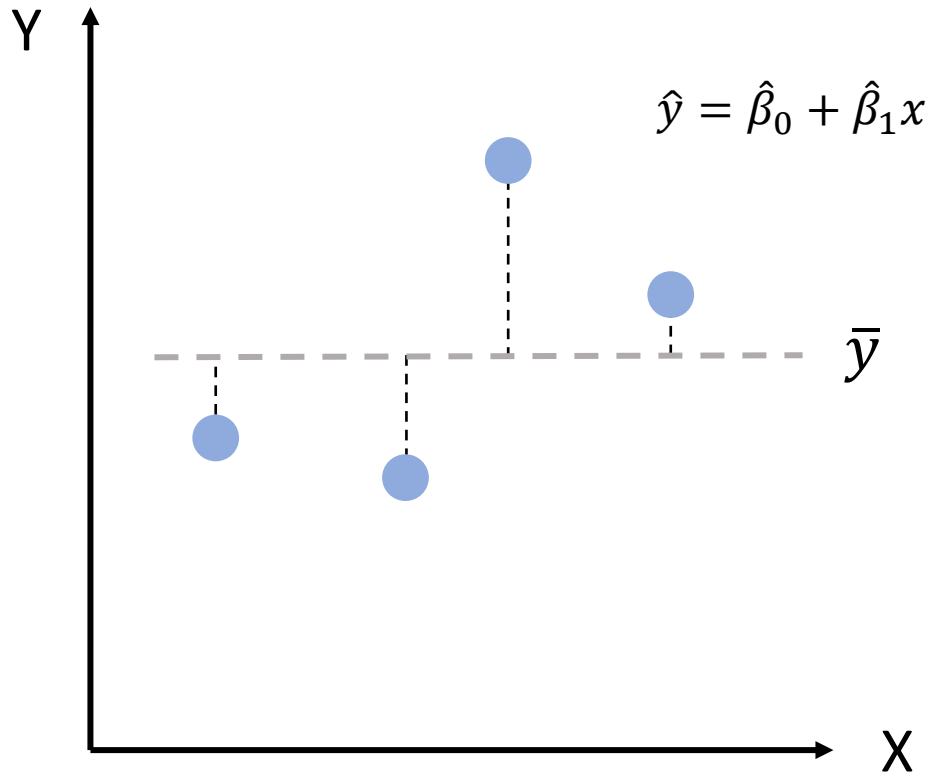
Sum of squares total = sum of squares regression + sum of squares error

$$\sum_{i=1}^n (y_i - \bar{y})^2 = \sum_{i=1}^n (\hat{y}_i - \bar{y})^2 + \sum_{i=1}^n (y_i - \hat{y}_i)^2$$

데이터의 분산 = 회귀모형이 설명 가능한 부분 + 불가능한 부분

→ SSE를 최소화하는 직선 찾아라!

# Regression



→ SSE를 최소화하는 직선 찾기 → **LSE (Least Squares Estimate)**

## 2. Linear Regression

# Linearity & Linear Model

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- Linearity?

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \cdots + \beta_p X_{pi} + \epsilon_i$$

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 X_{1i} X_{2i} + \epsilon_i$$

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 X_i^2 + \cdots + \beta_p X_i^p + \epsilon_i$$

# Linear Regression

예전에 사용하던 슬라이드  
중요한 부분 짚어가면서 배워봅시다!

$$r = \underline{f(x)} + \varepsilon$$

Estimator this one directly!  $= g(x|w) = w_1x_1 + \dots + w_dx_d + w_0 = \underline{w^T x} + w_0$

Assume as Linear model

[Assumptions of error]: Normality & Homoscedasticity & independent

→ 회귀진단에 이용!

$$\varepsilon \sim N(0, \sigma^2)$$

$$r \sim N(g(x|w), \sigma^2)$$

$$\log \prod_{t=1}^N p(r^t | x^t) = \log \prod_t \frac{1}{\sqrt{2\pi}\sigma} \exp \left[ -\frac{(r^t - g(x^t|w))^2}{2\sigma^2} \right] \rightarrow \text{Maximize!}$$

→ 갑자기 MLE??? → 다음 슬라이드

**Minimize** : Loss function  $E(w|x) = \frac{1}{2} \sum_{t=1}^N [(r^t - g(x^t|w))]^2$

→ MSE 최소화  $\frac{1}{n} \sum_{i=1}^n (y_i - \hat{y}_i)^2$





# Linear Regression

- 분명 MLE와 LSE는 원칙이 다르다.
  - MLE: likelihood을 최대로 만드는 추정량 (estimator)
  - LSE: MSE, 혹은 SSE를 최소로 만드는 추정량
- 그러나 오차의 정규성, 등분산성, 독립성을 만족할 때 두 추정량은 같다.

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon} \quad \boldsymbol{\epsilon} \sim \mathcal{N}(0, \sigma^2 \mathbf{I})$$

MLE

$$L(\boldsymbol{\beta}, \sigma^2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(y_i - \mathbf{x}_i^T \boldsymbol{\beta})^2}{2\sigma^2}\right)$$

$$\text{Maximize } \ell(\boldsymbol{\beta}, \sigma^2) = -\frac{n}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mathbf{x}_i^T \boldsymbol{\beta})^2$$

$$\text{Minimize } -\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mathbf{x}_i^T \boldsymbol{\beta})^2$$

$$\text{Equivalent to } S(\boldsymbol{\beta}) = (\mathbf{y} - \mathbf{X}\boldsymbol{\beta})^T (\mathbf{y} - \mathbf{X}\boldsymbol{\beta})$$

LSE

$$\text{Minimize } S(\boldsymbol{\beta}) = (\mathbf{y} - \mathbf{X}\boldsymbol{\beta})^T (\mathbf{y} - \mathbf{X}\boldsymbol{\beta})$$

$$S(\boldsymbol{\beta}) = \mathbf{y}^T \mathbf{y} - 2\mathbf{y}^T \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\beta}^T \mathbf{X}^T \mathbf{X} \boldsymbol{\beta}$$

$$\frac{\partial S(\boldsymbol{\beta})}{\partial \boldsymbol{\beta}} = -2\mathbf{X}^T \mathbf{y} + 2\mathbf{X}^T \mathbf{X} \boldsymbol{\beta}$$

$$\text{Let } -2\mathbf{X}^T \mathbf{y} + 2\mathbf{X}^T \mathbf{X} \boldsymbol{\beta} = 0 \rightarrow \mathbf{X}^T \mathbf{X} \boldsymbol{\beta} = \mathbf{X}^T \mathbf{y}$$

$$\boldsymbol{\beta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

$$\text{Therefore } \hat{\boldsymbol{\beta}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

# Gauss-Markov Theorem

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$\hat{\beta}$ 가 LSE이면서 비편향 추정량(unbiased estimator)일때, 임의의 비편향추정량  $\tilde{\beta}$ 에 대해 다음을 만족

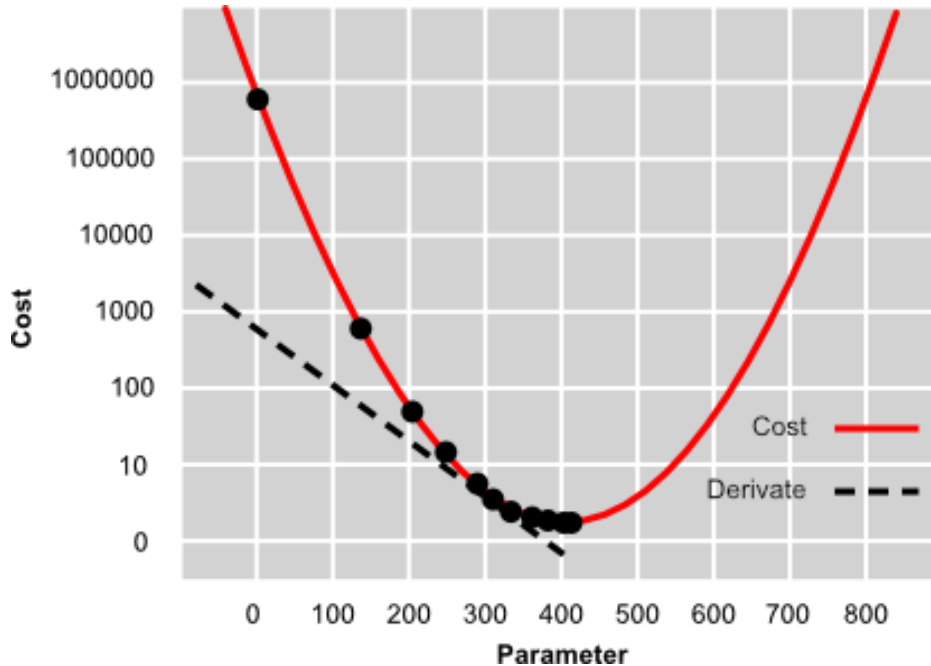
$$Var[\hat{\beta}|X] \leq Var[\tilde{\beta}|X]$$

따라서 LSE가 BLUE (best linear unbiased estimator)를 만족!

# Gradient Descent

**Minimize** : Loss function  $E(w|x) = \frac{1}{N} \sum_{t=1}^N [(r^t - g(x^t|w))^2] = \text{MSE}(\text{Mean Squared Error})$

$$w^* = \operatorname{argmin}_w E(w|x) \quad w_{j+1} \leftarrow w_j - \eta \frac{\partial E}{\partial w_j} \quad \text{iteratively}$$



- 컴퓨터가 어떻게 MSE를 최소화 하는 추정량을 찾는가 → Gradient Descent
- MSE가 2차함수이기 때문에 금방 찾을 수 있다.
- 다른 Loss 함수를 사용하는 경우도 사용 가능
  - Ex) MAE, RMSE, Huber 등등

# Least Square Estimation

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**Minimize** : Loss function  $E(w|x) = \frac{1}{2} \sum_{t=1}^N [(r^t - g(x^t|w))^2]$

\*\*  $g(x^t|w) = w_1 x^t + w_0$  : 1<sup>st</sup> order

$$w^* = \operatorname{argmin}_w E(w|x) \rightarrow \frac{\partial E}{\partial w_1} = 0 \ \& \ \frac{\partial E}{\partial w_0} = 0$$

$$A = \begin{bmatrix} N & \sum_t x^t \\ \sum_t x^t & \sum_t (x^t)^2 \end{bmatrix} \quad w = \begin{bmatrix} w_0 \\ w_1 \end{bmatrix} \quad y = \begin{bmatrix} \sum_t y \\ \sum_t r^t x^t \end{bmatrix}$$

$$w^* = A^{-1}y$$

# Multivariate Regression Multiple Regression

$$r^t = g(\mathbf{x}^t | w_0, w_1, \dots, w_d) + \epsilon = w_0 + w_1 x_1^t + w_2 x_2^t + \dots + w_d x_d^t + \epsilon$$

$$E(w_0, w_1, \dots, w_d | \mathcal{X}) = \frac{1}{2} \sum_t (r^t - w_0 - w_1 x_1^t - w_2 x_2^t - \dots - w_d x_d^t)^2$$

$$\begin{aligned} \sum_t r^t &= Nw_0 + w_1 \sum_t x_1^t + w_2 \sum_t x_2^t + \dots + w_d \sum_t x_d^t \\ \sum_t x_1^t r^t &= w_0 \sum_t x_1^t + w_1 \sum_t (x_1^t)^2 + w_2 \sum_t x_1^t x_2^t + \dots + w_d \sum_t x_1^t x_d^t \\ \sum_t x_2^t r^t &= w_0 \sum_t x_2^t + w_1 \sum_t x_1^t x_2^t + w_2 \sum_t (x_2^t)^2 + \dots + w_d \sum_t x_2^t x_d^t \\ &\vdots \\ \sum_t x_d^t r^t &= w_0 \sum_t x_d^t + w_1 \sum_t x_d^t x_1^t + w_2 \sum_t x_d^t x_2^t + \dots + w_d \sum_t (x_d^t)^2 \end{aligned}$$

$$\mathbf{X} = \begin{bmatrix} 1 & x_1^1 & x_2^1 & \dots & x_d^1 \\ 1 & x_1^2 & x_2^2 & \dots & x_d^2 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & x_1^N & x_2^N & \dots & x_d^N \end{bmatrix}, \mathbf{w} = \begin{bmatrix} w_0 \\ w_1 \\ \vdots \\ w_d \end{bmatrix}, \mathbf{r} = \begin{bmatrix} r^1 \\ r^2 \\ \vdots \\ r^N \end{bmatrix}$$

Then the normal equations can be written as

$$\mathbf{X}^T \mathbf{X} \mathbf{w} = \mathbf{X}^T \mathbf{r}$$

$$\mathbf{w} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{r}$$

$$\mathbf{w}^* = \mathbf{A}^{-1} \mathbf{y}$$

# ※주의

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Multiple Linear Regression

종속 변수 1개

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \cdots + \beta_p x_p + \epsilon$$

Multivariate Linear Regression

종속변수 여러 개

$$Y_1 = \beta_{01} + \beta_{11}X_1 + \beta_{21}X_2 + \cdots + \beta_{p1}X_p + \epsilon_1$$

$$Y_2 = \beta_{02} + \beta_{12}X_1 + \beta_{22}X_2 + \cdots + \beta_{p2}X_p + \epsilon_2$$

$$Y_3 = \beta_{03} + \beta_{13}X_1 + \beta_{23}X_2 + \cdots + \beta_{p3}X_p + \epsilon_3$$

# Polynomial Regression

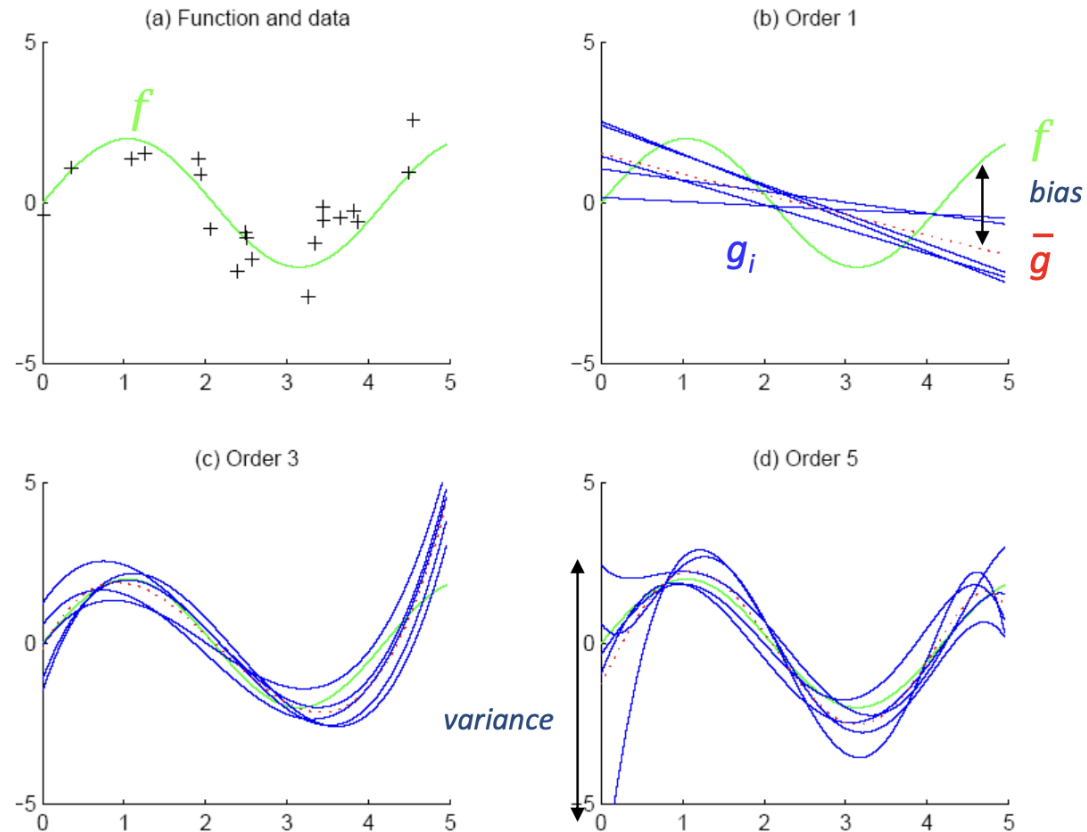
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$$g(x^t | w_k, \dots, w_2, w_1, w_0) = w_k (x^t)^k + \dots + w_2 (x^t)^2 + w_1 x^t + w_0$$

$$A = \begin{bmatrix} N & \sum_t x^t & \sum_t (x^t)^2 & \dots & \sum_t (x^t)^k \\ \sum_t x^t & \sum_t (x^t)^2 & \sum_t (x^t)^3 & \dots & \sum_t (x^t)^{k+1} \\ \vdots & & & & \\ \sum_t (x^t)^k & \sum_t (x^t)^{k+1} & \sum_t (x^t)^{k+2} & \dots & \sum_t (x^t)^{2k} \end{bmatrix}$$
$$w = \begin{bmatrix} w_0 \\ w_1 \\ w_2 \\ \vdots \\ w_k \end{bmatrix}, \quad y = \begin{bmatrix} \sum_t r^t \\ \sum_t r^t x^t \\ \sum_t r^t (x^t)^2 \\ \vdots \\ \sum_t r^t (x^t)^k \end{bmatrix}$$

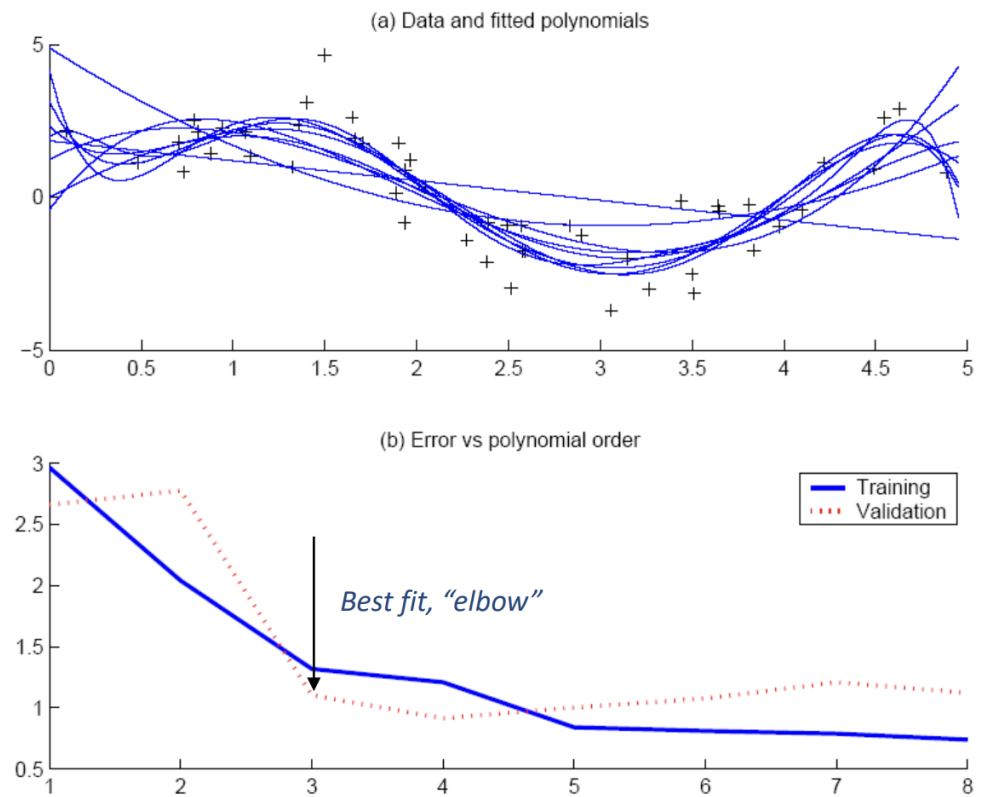
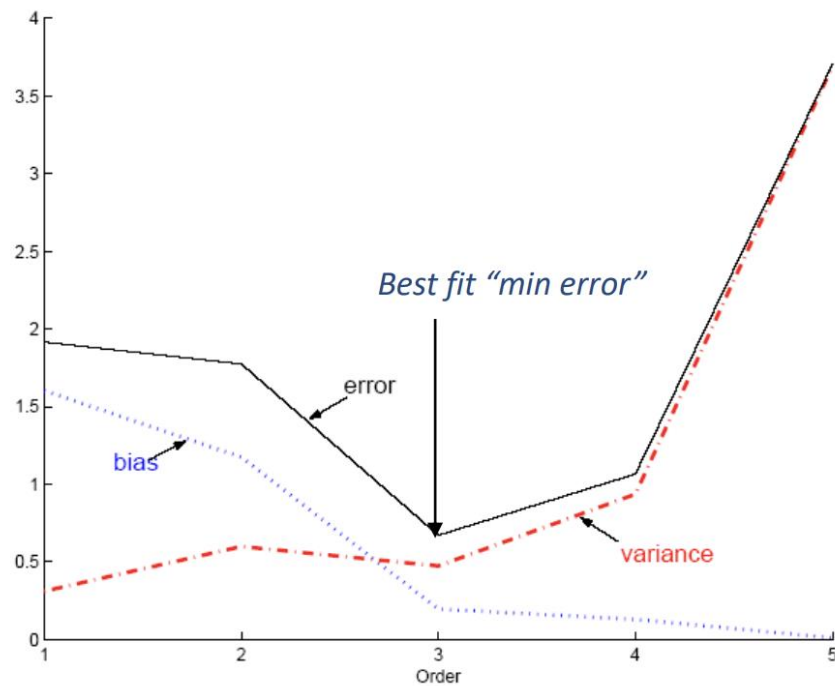
$$w^* = A^{-1}y$$

# Polynomial Regression





# Model Selection



# Cross Validation

To estimate **generalization error**, we need data unseen during training.

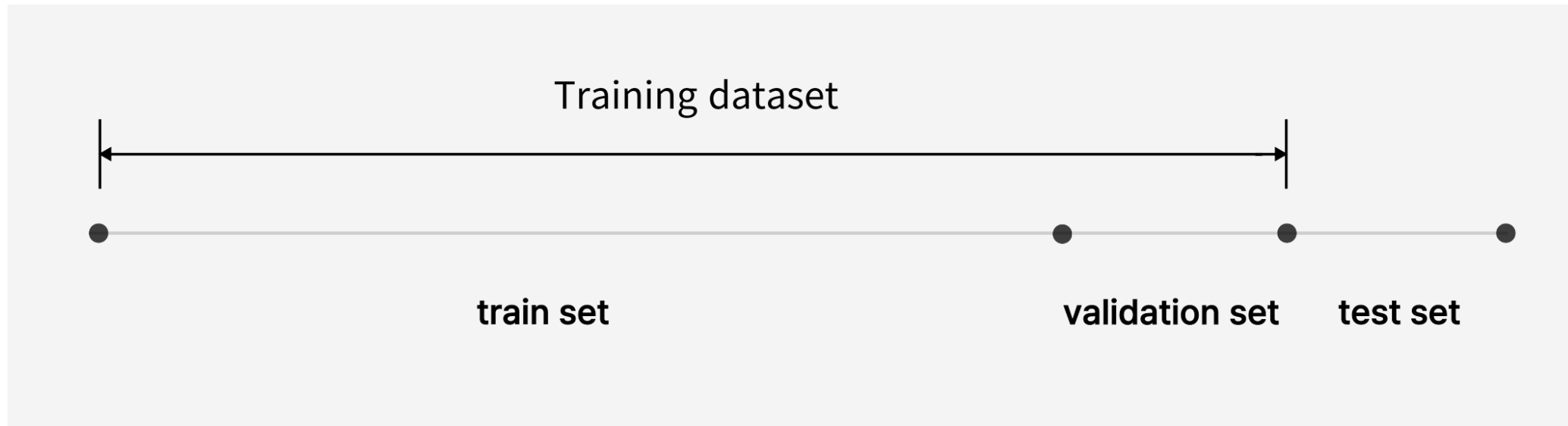
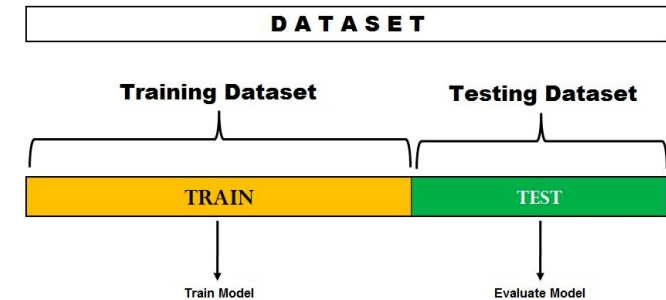
We split the data as

- Training set (50%)
- Validation set (25%)
- Test (publication) set (25%)

= Test Error 추정

Measure generalization accuracy by testing on data unused during training

Hold out



# Regularization

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## Penalize complex models

-  $E' = \text{error on data} + \lambda * \text{model complexity}$

\* If  $\lambda$  increases, variance decreases, but bias increases

→ Penalty term이 생기는 순간  
unbiased 포기 (Gauss-Markov Theorem)

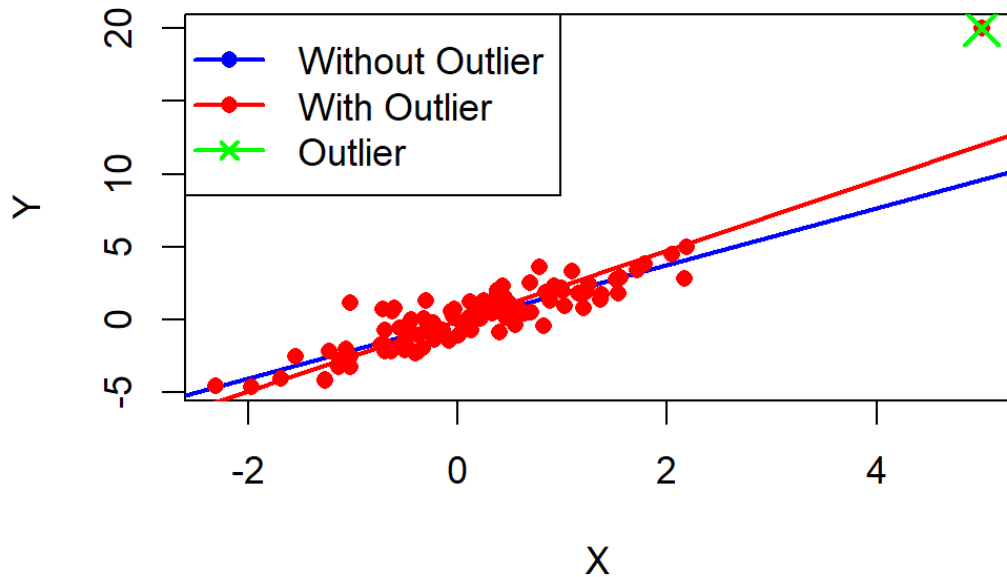
In regression...

Regularization (L2): 
$$E(\mathbf{w} | \mathcal{X}) = \frac{1}{2} \sum_{t=1}^N [r^t - g(x^t | \mathbf{w})]^2 + \lambda \sum_i w_i^2$$

Loss Function인 MSE에 Penalty Term  
추가 → 그런데 이게 왜 필요?  
→ 다음 슬라이드

# Linear Regression: Limitations

Linear Regression with and without an Outlier



- Linear Regression은 outlier에 취약!
- 이를 완화하기 위해 Regularization 사용
  - Ridge, LASSO

### 3. Regularization

# Distance

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- Distance measure

$$d(\mathbf{u}, \mathbf{v}) = (\sum |u_i - v_i|^2)^{\frac{1}{2}} = ||\mathbf{u} - \mathbf{v}||_2 \quad \text{Euclidean (L2 norm)}$$

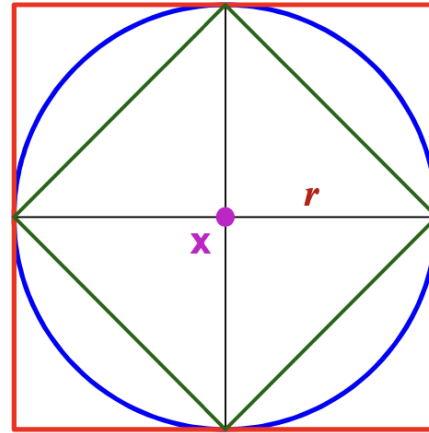
$$d(\mathbf{u}, \mathbf{v}) = \sum |u_i - v_i| = ||\mathbf{u} - \mathbf{v}||_1 \quad \text{Manhattan (L1 norm)}$$

$$d(\mathbf{u}, \mathbf{v}) = (\sum |u_i - v_i|^p)^{\frac{1}{p}} = ||\mathbf{u} - \mathbf{v}||_p \quad \text{Minkowski (Lp norm)}$$

$$d(\mathbf{u}, \mathbf{v}) = \sqrt{(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})} \quad \text{Mahalanobis Distance}$$

# Distance

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- **Green:** All points  $y$  at distance  $L_1(x, y) = r$  from point  $x$
- **Blue:** All points  $y$  at distance  $L_2(x, y) = r$  from point  $x$
- **Red:** All points  $y$  at distance  $L_\infty(x, y) = r$  from point  $x$

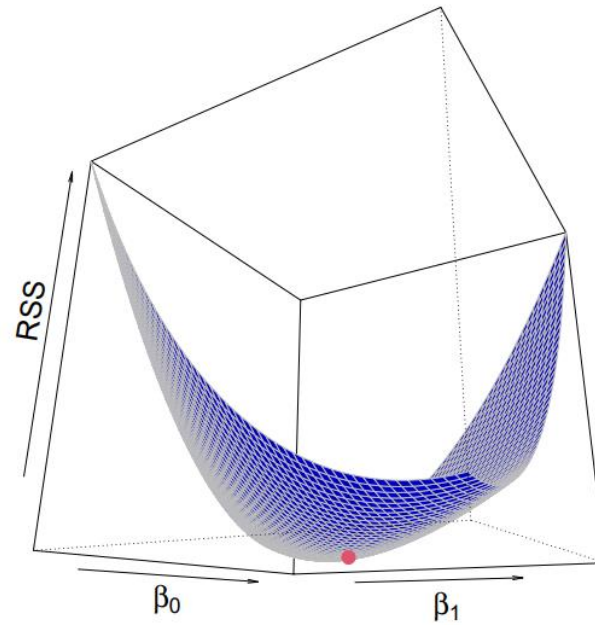
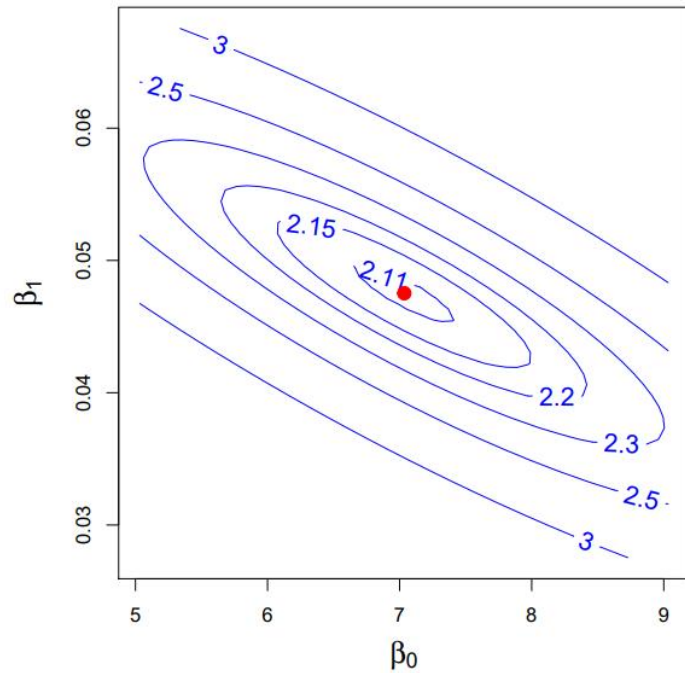
$$\|\mathbf{x}\|_1 = \sum_{i=1}^n |x_i|$$

$$\|\mathbf{x}\|_2 = \sqrt{\sum_{i=1}^n x_i^2}$$

$$\|\mathbf{x}\|_\infty = \max_{i=1, \dots, n} |x_i|$$

$$\|\mathbf{x}\|_p = (\sum_{i=1}^n |x_i|^p)^{1/p}$$

# SSE와 추정량의 contour

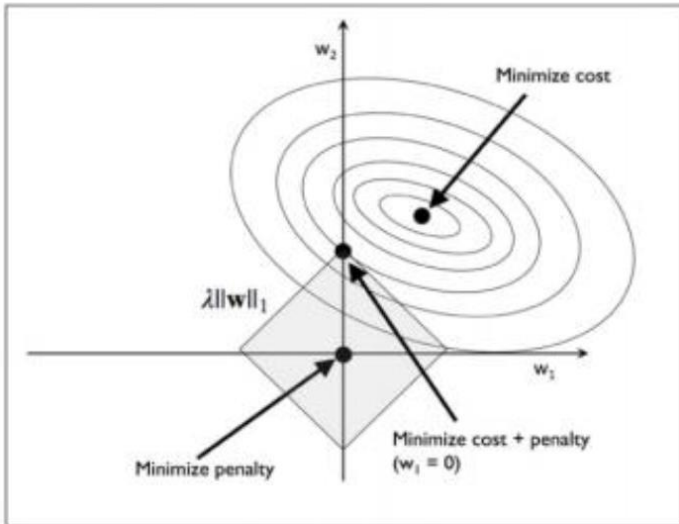


- 실제로 시뮬레이션을 해보면 이차 곡면이 나온다.
- 붉은 점이 LSE
- $RSS = \text{residual sum of square} = SSE$

→ 다음 슬라이드에 LASSO, Ridge regression 설명을 위함!



# Lasso Regression



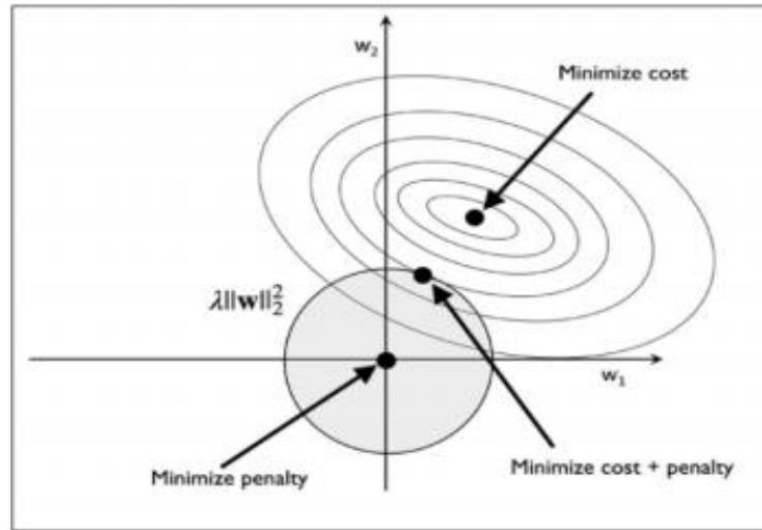
- LASSO (Least Absolute Shrinkage and Selection Operator)

$$(\hat{\beta}^{\lambda,1} =) \hat{\beta}_{LASSO} = \underset{\beta}{\operatorname{argmin}} (\mathbf{Y} - \mathbf{X}\beta)^T (\mathbf{Y} - \mathbf{X}\beta) + \lambda \|\beta\|_1$$

where  $\|\beta\|_1 = \sum_j^p |\beta_j|$   $\rightarrow$  Lagrangian Form

$$\Leftrightarrow \hat{\beta}_{LASSO} = \underset{\beta}{\operatorname{argmin}} (\mathbf{Y} - \mathbf{X}\beta)^T (\mathbf{Y} - \mathbf{X}\beta), \text{ with constraint } \sum_{i=1}^p |\beta_i| \leq t$$

# Ridge Regression

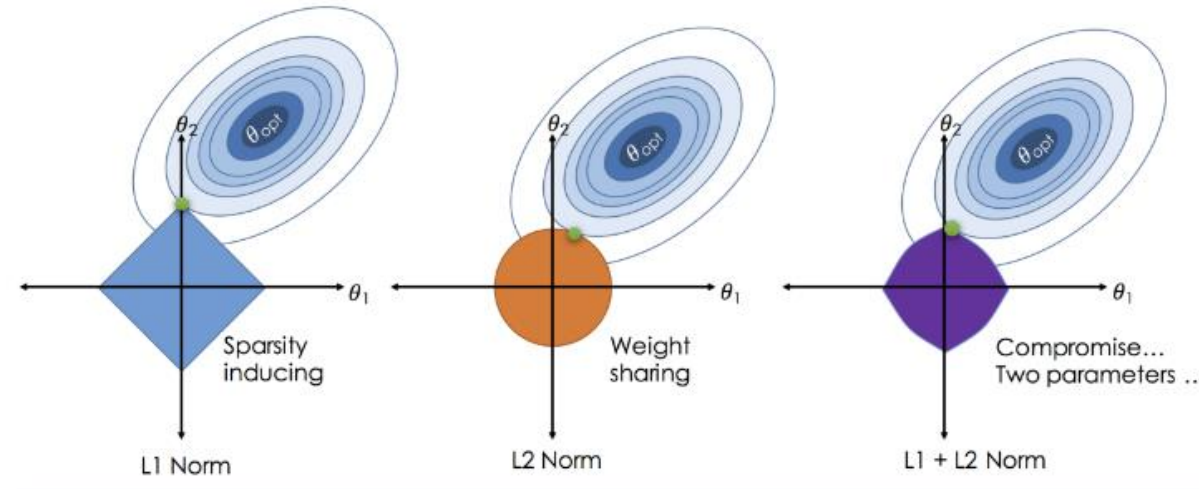


Ridge Regression solves → Lagrangian Form

$$\min_{\beta} (Y - X\beta)^T (Y - X\beta) + \lambda \|\beta\|_2^2 \quad (L2 \text{ penalty})$$

$$\Leftrightarrow \hat{\beta}_{Ridge} = \underset{\beta}{\operatorname{argmin}} (Y - X\beta)^T (Y - X\beta), \text{ with constraint } \sum_{i=1}^p \beta_i^2 \leq t$$

# Elastic-Net Regression



$$\frac{\sum_{i=1}^n (y_i - x_i^J \hat{\beta})^2}{2n} + \lambda \left( \frac{1-\alpha}{2} \sum_{j=1}^m \hat{\beta}_j^2 + \alpha \sum_{j=1}^m |\hat{\beta}_j| \right)$$

## 4. Regression Diagnostics

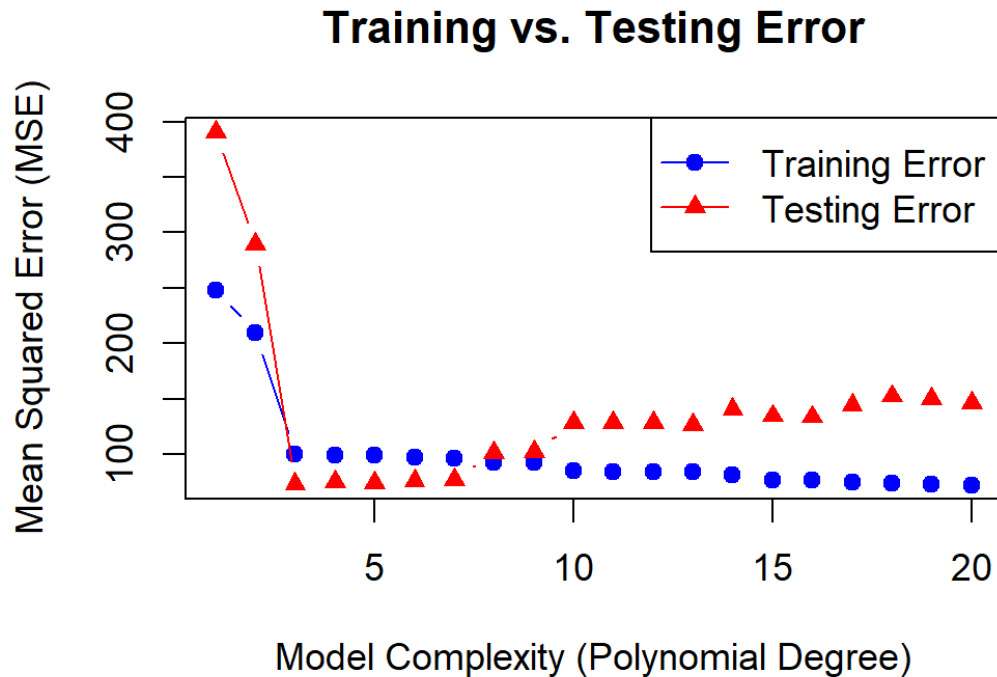
# Regression Diagnostics (회귀 진단)

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## 종류

1. Optimism based method → AIC, BIC
2. Test error based → Cross Validation, Bootstrap
3. Multicollinearity
4. Residual plot: 오차의 iid 가정 체크 (독립, 등분산, 정규성)
5. Goodness of Fit Test

# Optimism of Training Error



- 우리는 test error을 최소화 하고 싶다!
- 그러나 우리가 알고 있 수 있는 것은 training error 뿐
- 그러나 training error은 test error보다 낮다.
- Train error: 단조 감소
- Test error: 감소 → 증가

$$op = Err_{in} - \overline{err}$$

- $op$ : optimism
- $Err_{in}$ : in-sample error
- $\overline{err}$ : training error

# Optimism of Training Error

1. Optimism based method → AIC, BIC  
AIC와 BIC는 optimism (op)를 추정  
이 둘을 최소화하는 모델 찾기

$$AIC = 2k - 2 \ln(L)$$

k = number of estimated parameters  
L = maximum likelihood of model

$$BIC = \ln(n)k - 2 \ln(L)$$

k = number of estimated parameters  
L = maximum likelihood of model  
n = sample size

2. Test error based → Cross Validation, Bootstrap

두 방법은 test error를 직접 추정

Cross validation: sample size N, fold K일 때  
 $\kappa: \{1, \dots, N\} \rightarrow \{1, \dots, K\}$ : indexing function  
 $\hat{f}^{-k}(x)$ : fitted function with  $k$ th chunk left out

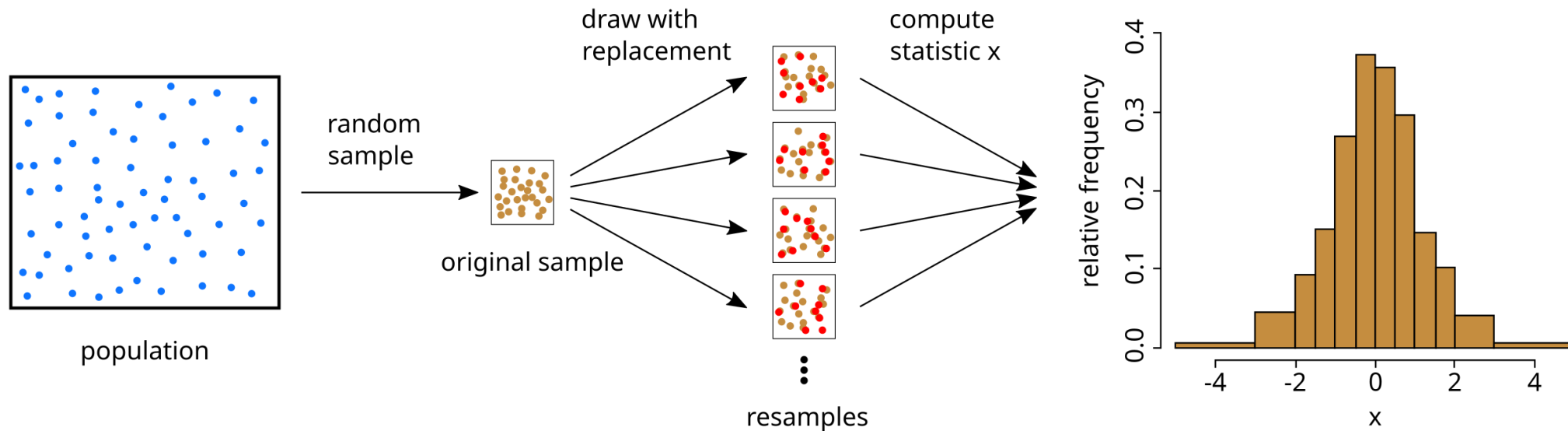
$$\widehat{test\ error} = CV(\hat{f}) = \frac{1}{N} \sum_{i=1}^N L(y_i, \hat{f}^{-\kappa(i)}(x_i))$$

Bootstrapping: B 가 replication 수 일때  
 $\hat{f}^{*b}(x_i)$ : predicted value of  $b$ th fitted bootstrap dataset

$$\widehat{test\ error} = \frac{1}{B} \frac{1}{N} \sum_{b=1}^B \sum_{i=1}^N L(y_i, \hat{f}^{*b}(x_i))$$



# Bootstrapping



Sample size가  $N$ , replication 수를  $B$  라고 하자.

- sample에서  $N$ 개를 복원 추출 (중복된 값 허용)
- $B$ 번 반복
- $B$ 개의 sample을 통해 통계량 계산 → bootstrap mean, bootstrap variance, bootstrap sd 등등



# Multicollinearity (다중공선성)

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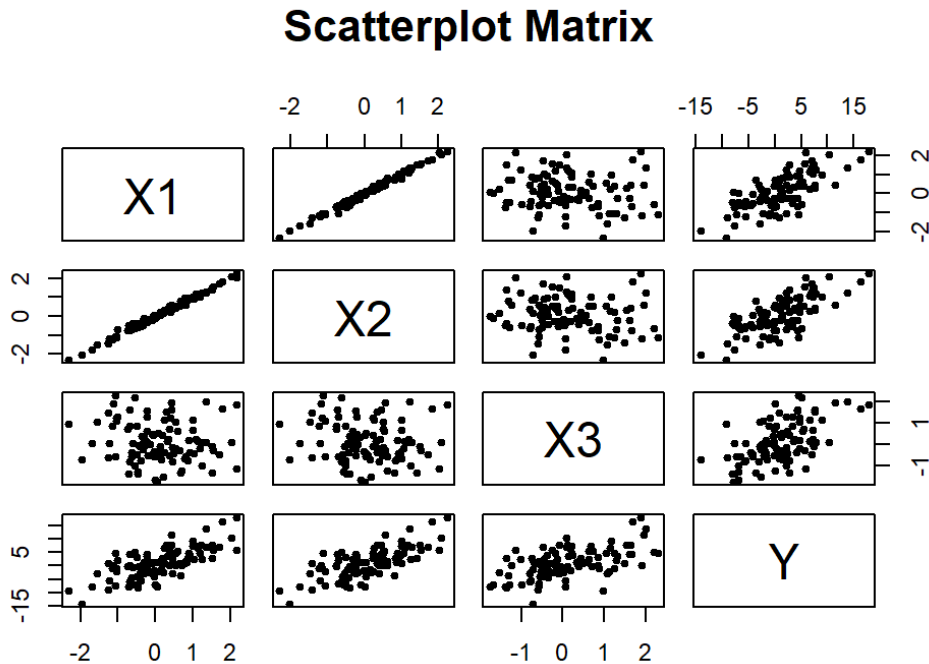
종속변수  $y$ , 독립변수를  $x_1, x_2$ 라고 하자

- $y$ 와  $x_1$ 의 correlation이 높으면 (좋다/나쁘다)
- $x_1$ 과  $x_2$ 의 correlation이 높으면 (좋다/나쁘다)

# Multicollinearity

## 파악하는 방법

1. Scatterplot matrix: X1과 X2 다중공선성



2. VIF (variance inflation factor)

$$VIF_j = \frac{1}{1 - R_j^2}$$

$R_j^2$ :  $R^2$  calculated by  $j$ th predictor regressed on other predictors  
→ 보통 cutoff 3 또는 5로 잡는다

→  $R^2$ 에 대해서는 뒤에 등장

```
> print(vif_values)
```

X1	X2	X3
89.523514	89.447291	1.017576

# Multicollinearity

---

처리하는 방법

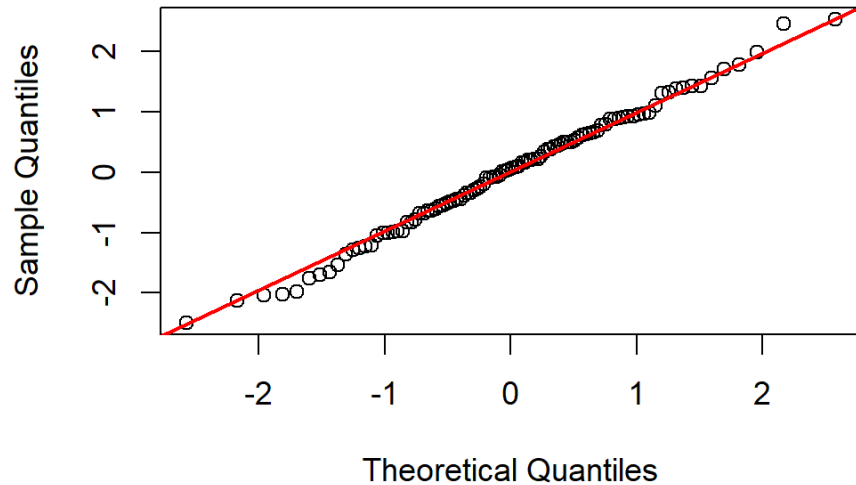
1. 변수 제거 ( $x_2$  제거)
2. Ridge regression, LASSO regression → robust to multicollinearity
3. 파생변수로 만들기

# Residual Plot

## Residual QQ Plot

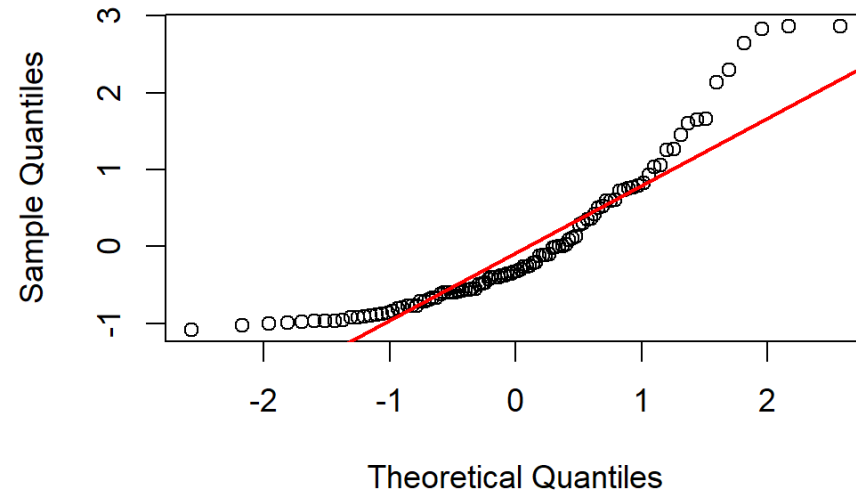
오차의 정규성 체크: 잔차의 qqplot이 직선을 따르면 정규성 만족

Normal Q-Q Plot



정규 가정 만족

Normal Q-Q Plot



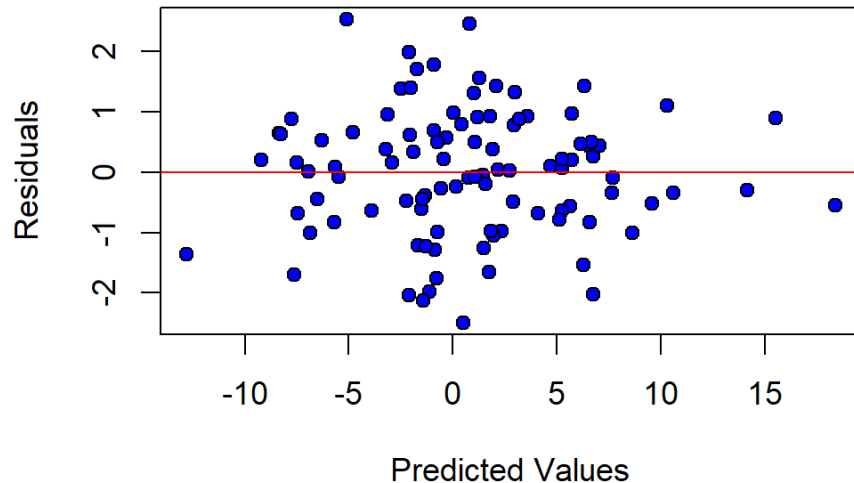
정규 가정 불만족

# Residual Plot

## Residual vs Predicted Values Plot

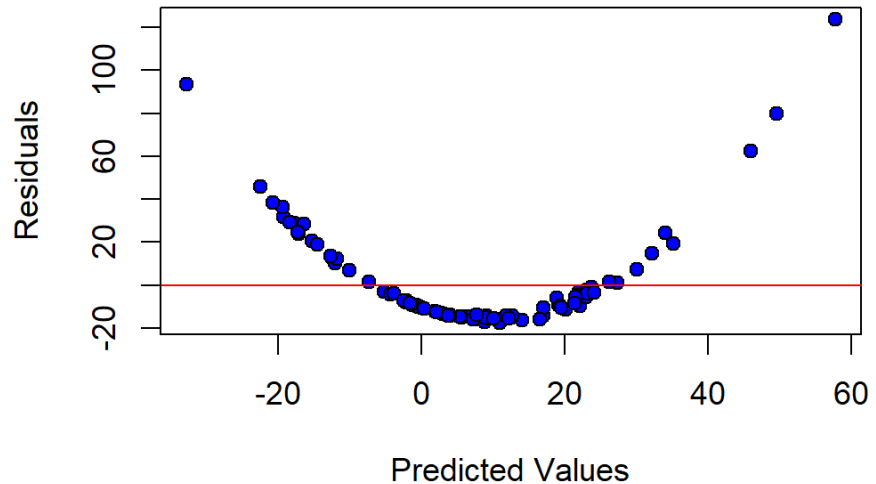
오차의 독립성 체크: predicted value vs residual 산점도 패턴 없어야 만족

Residuals vs. Predicted Values (No Pattern)



독립 가정 만족

Residuals vs. Predicted Values (Quadratic Pattern)

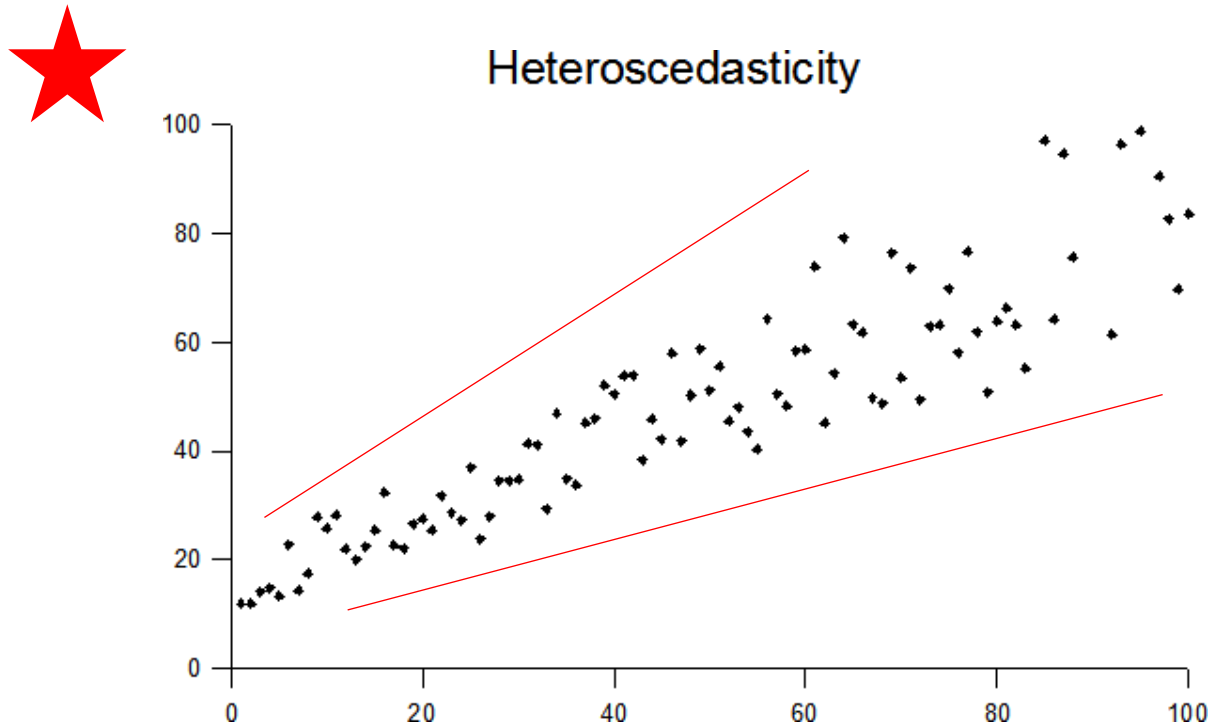


독립 가정 불만족

# Residual Plot

## Residual vs Predicted Values Plot

오차의 등분산성 체크: predicted value vs residual 퍼져있는 정도가 일정하면 만족



점점 퍼지는 중  
→ 오차 등분산 불만족

해결 방법:  
Variance stabilizing transformation(VST) →  
분산을 줄이는 변수변환 이용

$X \rightarrow \sqrt{X}, \log X, \arcsin X$  등으로 변환

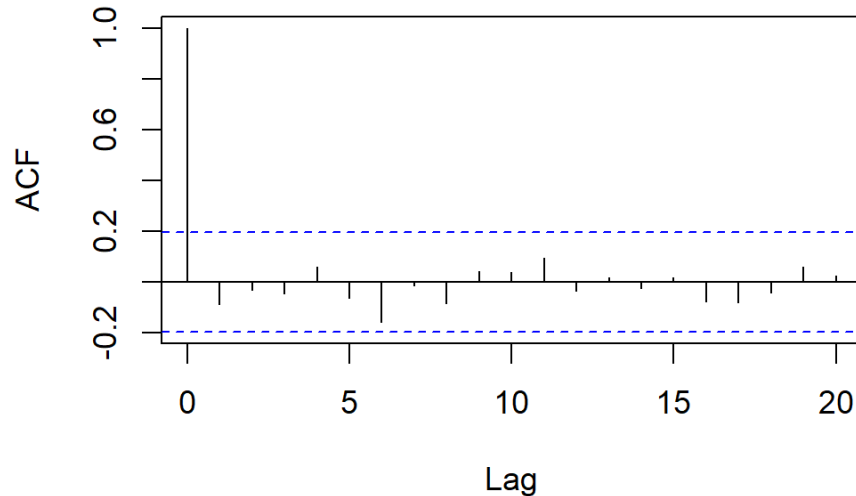
오차의 정규성을 포기한 회귀 모델이 많은데,  
등분산은 유지하려 하는 경향은 있다!

# Residual Plot

## ACF (autocorrelation function) Plot

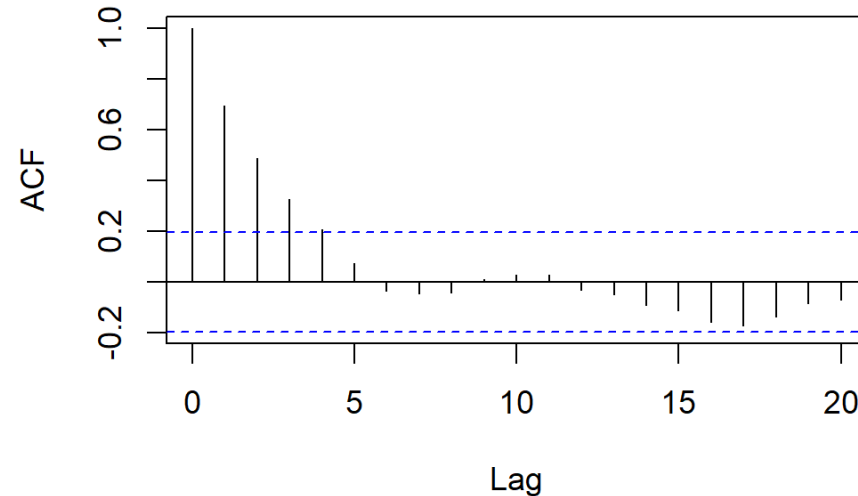
오차의 독립성 체크: ACF 값이 전체적으로 낮으면 만족. 첫번째 값은 높음

ACF of Residuals (No Autocorrelation)



독립 가정 만족

ACF of Residuals (Autocorrelation)



독립 가정 불만족

# Goodness of Fit Test

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Coefficient of Determination ( $R^2$ )

$$R^2 = \frac{SSR}{SST} = \frac{SST - SSE}{SST} = 1 - \frac{SSE}{SST} \leq 1$$

- 전체 변동 (SST) 중 모델이 설명하는 변동 (SSR)의 비율이 얼마큼 되는가.
- 1에 가까워질수록 모델의 설명력이 좋다고 할 수 있다.



# 수고하셨습니다!

해당 세션자료는 KUBIG Github에서 보실 수 있습니다!  
다음은 이번 주차 과제 설명이 있습니다!