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# 업데이터 통계학 스터디

Chapter 7 – Point Estimation 1

# **Relative Efficiency**

### Relative Efficiency

#### **Definition**

Given two UEs  $\hat{\theta}_1$  and  $\hat{\theta}_2$  of  $\theta$ , with variances  $V(\hat{\theta}_1)$  and  $V(\hat{\theta}_2)$ , respectively, the efficiency of  $\hat{\theta}_1$  relative to  $\hat{\theta}_2$  is defined to be the ratio.

$$eff(\hat{\theta}_1, \hat{\theta}_2) = \frac{V(\hat{\theta}_2)}{V(\hat{\theta}_1)}$$

(recall)  $Y_1, \dots, Y_n$ :  $iid\ Unif(0, \theta)$ 

$$\hat{\theta}_1 = 2\overline{Y}, \quad \hat{\theta}_2 = \frac{n+1}{n} Y_{(n)}$$

Find the efficiency of  $\hat{\theta}_1$  relative to  $\hat{\theta}_2$ .



# Consistency

# **Chebyshev Inequality**

### Consistency

<u>Idea</u>: Estimator should always get closed to the truth as number of observations increases.

#### <u>Definition</u> (Convergence in probability)

A sequence of random variables  $X_1, \dots, X_2, \dots, X_n, \dots$  converges in probability to a random variable X ( $X_n \xrightarrow{p} X$ ) if for any  $\exists_{\varepsilon} > 0$ ,

$$P(|X_n - X| > \varepsilon) \xrightarrow{n \to \infty} 0 \iff P(|X_n - X| \le \varepsilon) \xrightarrow{n \to \infty} 1$$

$$\star X_n \xrightarrow{p} c$$
 if  $P(|X_n - c| > \varepsilon) \xrightarrow{n \to \infty} 0$  or  $P(|X_n - c| \le \varepsilon) \xrightarrow{n \to \infty} 1$ 

### Consistency

#### <u>Definition</u> (Consistency)

 $\hat{\theta}_n$  based on  $X_1, \dots, X_2, \dots, X_n$  is consistent for  $\theta$  if  $\hat{\theta}_n \overset{p}{\to} \theta$  as  $n \to \infty$  for all values of  $\theta$ .

### **Tools to show Consistency**

#### Tool 1: Weak Law of Large Numbers (WLLN)

 $X_1, \dots, X_n$  are *iid* with mean  $E(X_i) = \mu < \infty$ , then

$$\bar{X} = \frac{1}{n} \sum_{i=1}^{n} X_i \xrightarrow{p} \mu$$

Proof: By Using Chebyshev Inequality.



 $Y_1, \dots, Y_n$ :  $iid\ r.\ v$ . Suppose that  $E(Y^2) < \infty$  and  $E(\log Y) < \infty$ . Proof

1. 
$$\frac{1}{n}\sum_{i=1}^{n}Y_{i}^{2} \stackrel{p}{\longrightarrow} E(Y^{2})$$

2. 
$$\frac{1}{n}\sum_{i=1}^{n} log Y_{i} \xrightarrow{p} E(log Y)$$

### **Tools to show Consistency**

#### Tool 2: Theorems on Limiting distributions

Suppose 
$$W_n \xrightarrow{p} a$$
 and  $V_n \xrightarrow{p} b$ . Then,

1. 
$$c_n W_n + d_n V_n \stackrel{p}{\longrightarrow} ca + db$$
. When  $c_n \longrightarrow c$ ,  $d_n \longrightarrow d$ 

2. 
$$W_n V_n \stackrel{p}{\longrightarrow} ab$$

3. 
$$W_n/V_n \xrightarrow{p} a/b$$
 if  $b \neq 0$ 

4. 
$$h(W_n) \xrightarrow{p} h(a)$$
 if h is continuous at a

### **Tools to show Consistency**

• If 
$$\overline{Y}_n \stackrel{p}{\longrightarrow} \mu$$
. Then,

$$\overline{Y}_n^2 \xrightarrow{p} \mu^2, \quad \sqrt{\overline{Y}_n} \xrightarrow{p} \sqrt{\mu}, \quad \log(\overline{Y}_n) \xrightarrow{p} \log(\mu) \quad (\text{if } \mu > 0)$$

- 1. Show that, Sample variance( $S_n^2$ ) is always a consistent estimator of population variance ( $\sigma^2$ ).
- 2.  $S_n^2$  is unbiased and consistent for  $\sigma^2$ . Is  $S_n$  unbiased or consistent for  $\sigma$ ?

 $Y_1, \dots, Y_n$ : iid  $Exp(\theta), \theta > 0$ .

### **Tools to show Consistency**

#### Tool 3: Weak law of variance

If  $\hat{\theta}_n$  is an UE of  $\theta$  and  $V(\hat{\theta}_n) \to 0$ ,  $\hat{\theta}_n$  is consistent for  $\theta$ .

#### Tool 3: Strong Law of variance

If  $\hat{\theta}_n$  is an estimator of  $\theta$  and  $MSE(\hat{\theta}_n) \to 0$ ,  $\hat{\theta}_n$  is consistent for  $\theta$ .

Proof: By Using Chebyshev Inequality.



 $Y_1, \dots, Y_n$ : iid  $N(\mu, \sigma^2)$ .

Are  $\overline{Y}_n \& S_n^2$  consistent estimator for  $\mu$  and  $\sigma$  respectively?

 $Y_1, \dots, Y_n$ : iid  $Unif(0, \theta)$ .

Are  $\hat{\theta}_1 = 2\overline{Y}$  and  $\hat{\theta}_2 = \frac{n+1}{n} Y_n$  consistent estimator for  $\theta$ ?

#### **Definition**

Suppose  $X_n$  is a random variable with CDF  $F_n(x)$ ,  $n=1,2,\cdots$ . Then  $X_1,X_2,\cdots$  converges in distribution to a random variable X with CDF F(x) if

$$\lim_{n\to\infty} F_n(x) = F(x)$$

#### Theorem (Central Limit Theorem)

 $Y_1, \dots, Y_n$ : random sample from a distribution with  $(\mu, \sigma^2)$ . Then,

$$Z_n = \frac{\sum_{i=1}^n Y_i - E(\sum_{i=1}^n Y_i)}{\sqrt{Var(\sum_{i=1}^n Y_i)}} = \frac{\overline{Y}_n - \mu}{\sigma/\sqrt{n}} \xrightarrow{D} N(0,1)$$

meaning that CDF of  $Z_n$  converges to the CDF of  $N(0,1) \Rightarrow$ 

$$P(Z_n \le z) \to \Phi(z) \text{ for all } z$$
 
$$P(a \le Z_n \le b) = F_{Z_n}(b) - F_{Z_n}(a) \to \Phi(b) - \Phi(a)$$

#### Theorem (Mapping Theorem)

For sequence of r.v.  $X_1, \dots, X_n$ .

If  $X_n \xrightarrow{D} X$ , then  $h(X_n) \xrightarrow{D} h(X)$  for any continuous function h.

#### Theorem (Limiting MGF Theorem)

 $X_n$  has CDF  $F_n(x)$  and MGF M(t;n) that exists for |t| < h. If there is a CDF

F(x) with MGF M(t), then  $X_n$  has a limiting distribution with CDF F(x).

 $Y_1, \dots, Y_n$ : iid Bin(n, p).  $\mu = np$  is a constant.

Find a limiting distribution of  $Y_n$ .



Theorem (Slutsky's Theorem)

$$U_n \xrightarrow{D} U$$
 and  $W_n \xrightarrow{p} 1 \Rightarrow U_n/W_n \xrightarrow{D} U$ .

Prove the following proposition.

$$U_n \xrightarrow{D} N(0,1) \Rightarrow W_n^2 \xrightarrow{D} \chi^2(1)$$

Prove the following proposition.

$$T_n \sim t(n) \Rightarrow T_n \xrightarrow{D} N(0,1)$$

Prove the following proposition. ( $\sigma^2$  is known)

$$Y_1, \dots, Y_n$$
: iid  $(\mu, \sigma^2) \Rightarrow \frac{\bar{Y}_n - \mu}{\sigma/\sqrt{n}} \xrightarrow{D} N(0,1)$ 

Prove the following proposition. ( $\sigma^2$  is unknown)

$$Y_1, \dots, Y_n$$
: iid  $(\mu, \sigma^2) \Rightarrow \frac{\bar{Y}_n - \mu}{S_n / \sqrt{n}} \xrightarrow{D} N(0,1)$