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업데이터 통계학 스터디

Chapter 3 – Continuous Random Variables 1

A continuous random variable differs from a discrete random variable in that it takes on an uncountably infinite number of possible outcomes.

Example

Choose a number from an interval [0,1] at random.

- 1. What is the probability that you choose 0.5? P(x=0.5)=0
- 2. What is the probability that you choose 0.382915? $\rho(x=0.382915)=0$
- 3. What is the probability that the number is less than 0.5? $\frac{1}{2}$
- 4. Let X be the random number chosen from [0,1]. Guess what is the CDF of X? $F_{*}(x) = \begin{cases} \gamma & (\infty,0) \\ \gamma & (0,1) \end{cases}$
- 5. Draw the CDF of X.

The PMF of discrete random variable X is P(X = x).

However, P(X = x) = 0 for continuous random variables. i.e., the probability that X takes on any value x equals 0.

This time: The probability distribution of continuous random variable X is defined via $P(a < X \le b)$ and it means the probability that X falls in some interval (a, b].

Definition

A random variable X is called a continuous random variable if its CDF $F_X(x)$ is a continuous function for all $x \in \mathbb{R}$.

<u>NOTE</u>: In discrete case, the discontinuity points have positive probability. However, the probability mass is zero for each point if the random variable is of continuous type. Let us define probability function for continuous random variable that works as probability mass function of discrete random variable.

We want to make the function $f_X(x)$ such that the probability of an interval is the area under the curve. In other words, for any interval (a, b], the pdf of X, $f_X(x)$ satisfies

$$P(a < X \le b) = \int_{a}^{b} f_X(x) \, dx$$

For each $x \in S_X$, consider the probability of a short interval

$$P(x < X \le x + \Delta)$$

The probability can be computed by

$$\Delta(width) * f_X(x)(height)$$

approximately. Thus, the probability density function is defined as

$$f_X(x) = \lim_{\Delta \to 0} \frac{P(x < X \le x + \Delta)}{\Delta}$$

$$f_X(x) = \lim_{\Delta \to 0} \frac{P(x < X \le x + \Delta)}{\Delta}$$
$$= \lim_{\Delta \to 0} \frac{F_X(x + \Delta) - F_X(x)}{\Delta}$$
$$= F'_X(x)$$

$$\therefore f_X(x) = F'_X(x)$$

Definition.

Let X be a continuous random variable. Then the probability density function (PDF) of X is a function $f_X(x)$ defined on real numbers such that

- 1. $f_X(x)$ is non-negative for any $x \in \mathbb{R}$.
- 2. The area under the curve $f_X(x)$ is 1, that is:

$$\int_{\mathbb{R}} f_X(x) dx = \int_{-\infty}^{\infty} f_X(x) dx = 1$$

We call $S_X = \{x \in \mathbb{R} : f_X(x) > 0\}$ the support of $f_X(x)$. Which means $f_X(x) = 0$ for any x outside S_X . Thus, we can write

$$\int_{\mathbb{R}} f_X(x) dx = \int_{-\infty}^{\infty} f_X(x) dx = \int_{S_X} f_X(x) dx = 1$$

3. For any $A \subseteq S_X$,

$$P(X \in A) = \int_A f_X(x) \, dx$$

Let X be a continuous random variable whose PDF is:

$$f_X(x) = 3x^2$$
, $(0 < x < 1)$

- 1. Verify that $f_X(x)$ is a valid PDF. $\begin{cases} 0 & \text{Non-negative} & \text{(i)} \\ \text{(i)} & \text{(i)} &$
- 2. What is the probability that X falls between 1/2 and 1? That is, what is P(1/2 < X < 1)? $\int_{\frac{1}{2}}^{1} 3x^{2} dx = \frac{1}{8}$
- 3. What is P(X = 1/2)?

NOTE: $P(a \le X \le b) = P(a < X \le b) = P(a \le X < b) = P(a < X < b)$.

Let X be a continuous random variable whose PDF is:

$$f_X(x) = \frac{x^3}{4}, \qquad (0 < x < c)$$

What value of the constant c makes $f_X(x)$ a valid PDF?

$$\int_{0}^{C} \frac{x^{2}}{4} dx = \left[\frac{x^{4}}{16} \right]_{0}^{C} = \frac{c^{4}}{16} = 1, \quad \therefore C = 2$$

Cumulative Distribution Function

CDF of a random variable X always exists.; $P(X \le x)$.

Let X is a continuous random variable (c.r.v.). Then,

$$F_X(x) = P(X \le x) = P(-\infty < X < x) = \int_{-\infty}^x f_X(t) dt$$

For $x \in \mathbb{R}$.

NOTE: For c.r.v., $F_X(x)$ is a non-decreasing continuous function.

Let X be a continuous random variable whose PDF is:

$$f_X(x) = 3x^2$$
, $(0 < x < 1)$

What is the cumulative distribution function $F_X(x)$?

$$F_{x}(x) = \begin{cases} \chi^{3} & (0,1) = \chi^{3} \cdot \overline{I}_{(0,1)}(\lambda) + \overline{I}_{(1,\infty)}(\lambda) \\ 0 & (-\infty,0) \end{cases}$$

$$I_{n} dicator \quad \text{function} \quad \text{where } \lambda = \frac{1}{2} \text{ with } \lambda$$

Let X be a continuous random variable whose PDF is:

$$f_X(x) = \frac{x^3}{4}, \qquad (0 < x < 2)$$

What is the cumulative distribution function $F_X(x)$?

$$F_{X}(\gamma) = \begin{cases} \frac{1}{16} \chi^{4} & (0, 2) = \frac{1}{16} \chi^{4}. \int_{(0, 2)} (\chi) + \int_{(2, \infty)} (\chi) \\ 0 & (-\infty, 0) \end{cases}$$

$$= \frac{1}{16} \chi^{4}. \int_{(0, 2)} (\chi) + \int_{(2, \infty)} (\chi)$$

Mean (Expectation value)

Definition

Let X be a continuous random variable with its PDF $f_X(x)$. The expected value or mean of a continuous random variable X is:

$$\mu_X = E(X) = \int_{-\infty}^{\infty} x f_X(x) dx$$

Let u(x) is a real valued function. Then the expected value of u(X) is:

$$E[u(X)] = \int_{-\infty}^{\infty} u(x) f_X(x) dx$$

Variance / Standard Deviation

Definition

The variance of a continuous random variable X is:

$$\sigma_X^2 = Var(X) = E[(X - \mu_X)^2] = \int_{-\infty}^{\infty} (x - \mu_X)^2 f_X(x) dx$$

The **standard deviation** of a continuous random variable *X* is:

$$\sigma_X = \sqrt{Var(X)}$$

Variance / Standard Deviation

We can use the shortcut formula for variance.

$$E[(X - \mu_X)^2] = E(X^2) - {\mu_X}^2$$

And,

$$Var(aX + b) = a^2 Var(X)$$

Linearity of Expectation Value

Theorem

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Suppose that X is a random variable with finite expectation. Then the following statements are true.

- 1. If c is a constant in the interval of X, then E(c) = c.
- 2. If c is a constant and u is a function, then

$$E[cu(X)] = cE[u(X)].$$

$$= 3E(X) + 6E(X) + 1$$

3. For any positive integer k,

$$E\left[\sum_{i=1}^k c_i u_i(X)\right] = \sum_{i=1}^k c_i E[u_i(X)]$$

Suppose X is a continuous random variable with the following probability density function:

$$f_X(x) = 3x^2 \quad (0 < x < 1)$$

- 1. What is the mean of X? $\xi(x) = \int_{0}^{1} x \cdot 3x^{2} dx = \left[\frac{3}{4}x^{4}\right]_{0}^{1} = \frac{3}{4}$
- 2. What is the variance of *X*?

$$E(x^2) = \int_0^1 \chi^2 \cdot 3\chi^2 \, d\chi = \left[\frac{3}{5} \chi^5 \right]_0^1 = \frac{3}{5}$$

:
$$Vor(X) = \frac{3}{5} - \frac{9}{16} = \frac{3}{20}$$

Let Y be a continuous random variable with PDF

$$g_Y(y) = 2y \ (0 < y < 1)$$

- 1. What is the mean of Y? $E(Y) = \int_{0}^{1} y \cdot y dy = \left[\frac{2}{3}y^{3}\right]_{0}^{1} = \frac{2}{3}$
- 2. What is the variance of *Y*?

$$E(Y^{2}) = \int_{0}^{1} y^{2} \cdot 2y \, dy = \left[\frac{2}{4}y^{4}\right]_{0}^{1} = \frac{1}{2}$$

$$V_{or}(Y) = \frac{1}{2} - \frac{4}{9} = \frac{1}{18}$$

Moment Generating Function (MGF)

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$$E((x-\mu)^k)$$
: k -th central moment 34 34

Definition

The moment generating function of a continuous random variable X, if it exists, is:

$$M_X(t) = E(e^{tX}) = \int_{-\infty}^{\infty} e^{tx} f_X(x) dx, \qquad t \in (-h, h), \qquad \exists_h > 0$$

Differentiating the moment generating function provides us with a way of finding the mean and the variance:

$$\mu_X = E(X) = M_X'(0), \qquad \sigma_X^2 = Var(X) = M_X''(0) - \{M_X'(0)^2\}$$

* Gomma (
$$\alpha_{i,\delta}$$
)
$$\sim M_{x(t)} = \frac{1}{(1-\mu t)^{d}}$$

Suppose *X* is a continuous random variable with the following PDF:

$$f_X(x) = xe^{-x} \quad (0 \le x < \infty)$$

~ Gramma (2,1) \to Mx(t)= $\frac{1}{(1-t)^2}$, $M_X = 2$

Use the MGF, find the mean and variance of X.

$$M_{x}(t) = E(e^{tX}) = \int_{0}^{\infty} e^{tx} \cdot \chi e^{-x} d\chi$$

$$= \int_{0}^{\infty} \chi \cdot e^{x(t-1)} d\chi \leftarrow t - 1 < 0$$

$$= \left[\frac{1}{t-1} \chi e^{x(t-1)} - \frac{1}{(t-1)^{2}} e^{x(t-1)} \right]_{0}^{\infty}$$

$$= \frac{1}{(1-t)^{2}}$$

$$M_{x}'(t) = \frac{2}{(1-t)}$$
 $M_{x} = M_{x}'(0) = 2$

* exponential distribution

 $X \sim \text{Exp}(\lambda)$, $f_{x}(x) = \lambda e^{-\lambda x}$. $M_{x}(t) = \frac{\lambda}{\lambda - t}$ $E(x) = \frac{1}{\lambda}$ $V_{\text{or}}(x) = \frac{1}{\lambda^{2}}$

Suppose *X* has the PDF:

$$f_X(x) = e^{-x-1} \quad (-1 < x < \infty)$$

1.
$$P(X>1) = \int_{1}^{\infty} e^{-x^{-1}} dx = \left[-e^{-x_{-1}} \right]_{1}^{\infty} = 1$$
2. $M_{X}(t) = E(e^{tX}) = \int_{-1}^{\infty} e^{tx} e^{-x^{-1}} dx = \int_{1}^{\infty} e^{(t-1)x-1} dx = \left[\frac{1}{t-1} e^{(t-1)x-1} \right]_{-1}^{\infty} = \frac{e^{-t}}{1-t}$
3. $\mu_{X} = M_{Y}(0) = \frac{-e^{-t}(1-t)^{3}}{(1-t)^{3}} \Big|_{t=0} = \frac{te^{-t}}{(1-t)^{3}} \Big|_{t=0} = 0$, $E(x^{2}) = M_{X}(0) = \frac{(-te^{-t}(1-t)^{2} + te^{-t} - 2(1-t))}{(1-t)^{4}} \Big|_{t=0} = 1$
4. $\sigma_{X} = E(x^{3}) - \{E(x)\}_{1}^{2} = 1$

5.
$$F_X(x) = \int_{-1}^{x} \{x(t) dt = \int_{-1}^{x} e^{t-1} dt = [-e^{t-1}]_{-1}^{x} = -e^{x-1} + [-e^{t$$

6.
$$q_1, q_2, q_3$$