Lecture 1 - Linear Regression

Model Selection

When 2 models both fit the training data equally well, which one should we choose?

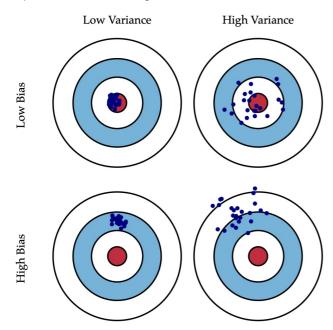
- · use validation data and choose the better one
 - 。 一般来说 train : valid : test = 8:1:1
 - 。 一般over-fit时error on validation data会上升
- choose a simpler one (if without validation data)
 - 。 Occam's Razor(奥卡姆剃刀准则): 如无必要, 勿增实体
 - Inductive bias

Bias-Variance Decomposition

- D is a training set and $D=\{(x_1,y_1),\ldots,(x_n,y_n)\}$ $(D\sim P(D))$
 - 。 也就是说用由分布P(D)采样出来的D训练一个模型f(x;D)
- 对于一个测试数据(x,y), 平方误差的期望为 $E_D[f(x;D)-y]^2$
 - \circ which means average error of f(x;D) over infinitely many $D\sim P(D)$
 - \circ we define $ar{f}(x) := E_D[f(x;D) y]$
 - o now we have:

$$egin{aligned} E_D[f(x;D)-y]^2 &= E_D[f(x;D)-ar{f}(x)+ar{f}(x)-y]^2 \ &= E_D[f(x;D)-ar{f}(x)]^2 + 2E_D[f(x;D)-ar{f}(x)][ar{f}(x)-y] + E_D[ar{f}(x)-y]^2 \ &= E_D[f(x;D)-ar{f}(x)]^2 + (ar{f}(x)-y)^2 \ &= ext{Variance} + ext{Bias}^2 \ &(E_D[f(x;D)-ar{f}(x)]=0) \end{aligned}$$

- \circ Variance: error caused by over-fitting particular D can be decreased by increasing the size n of D 简单来说即是测试的预测与平均预测之差的平方
- Bias: error caused by the model's ability to fit the data, or the model is not complex enough
 即平均预测与标签也就是实际值之差
- Trade-off
 - o simple model: high bias, low variance
 - o complex model: low bias, high variance



- · One more thing
 - $\circ \ ar{f}(x) = E_D(f(x;D))$ is an average model over infinitely sample small D
 - $\hat{f}(x) = f(x; D, |D| \to \infty)$ is a model trained over an infinitely large D
 - 。 虽然上述推导使用了第一个函数,但在实际过程中数据自然越多越好,那么也就是第二个模型往往表现更好
 - o example:

$$P(D) = \begin{cases} 1/3 & \{(-1,1),(0,0)\} \\ 1/3 & \{(1,1),(0,0)\} \\ 1/3 & \{(-1,1),(1,1)\} \end{cases}$$

- $\begin{tabular}{l} $\bar{f}(x)=1/3$, $variance>0$, $bias^2=1$ \\ $\bar{f}(x)=2/3$, $variance=0$, $bias^2=2/3$ \\ \end{tabular}$

Linear Regression

Basic Model

- D is a training set and $D=\{(x_1,y_1),\ldots,(x_n,y_n)\}.$ $x_i\in\mathbb{R}^d,\;y_i\in\mathbb{R},\;i=1,2,\ldots,n$ We are going to train a model $y = w^T x + b, \ w \in \mathbb{R}^d, \ b \in \mathbb{R}$
- How to determine w, b
 - ERM(Empirical Risk Minimization)
- · Loss function
 - \circ squared loss $(f(x_i)-y_i)^2$
- · Objective:

$$egin{aligned} & \min_{w,b} \sum_{i=1}^n (f(x_i) - y_i)^2 \ & = \min_{w,b} \sum_{i=1}^n (w^T x_i + b - y_i)^2 := L(w,b) \end{aligned}$$

Gradient:

$$egin{aligned} rac{\partial L}{\partial w} &= 2\sum_{i=1}^n (w^Tx_i + b - y_i)x_i \ rac{\partial L}{\partial b} &= 2\sum_{i=1}^n (w^Tx_i + b - y_i) \end{aligned}$$

注意这里是一个向量对标量求梯度,结果还是一个向量

Gradient Descent

$$w \leftarrow w - \alpha \frac{\partial L}{\partial w}$$
$$b \leftarrow b - \alpha \frac{\partial L}{\partial b}$$

lpha is the learning rate and it's a hyperparameter

• 常用公式(已知 $w \in \mathbb{R}^d$)

$$egin{aligned} rac{\partial w^T w}{\partial w} &= 2w \ & rac{\partial w^T x}{\partial x} &= w \ & rac{\partial x^T A x}{\partial x} &= (A + A^T) x \end{aligned}$$

Least Squares

- · Least squares actually has closed-form solution! (which also regards as Normal Equation)
- we need some vectorized variables:

$$egin{aligned} x &= egin{bmatrix} x_1^T, 1 \ x_2^T, 1 \ \dots \ x_n^T, 1 \end{bmatrix} \in \mathbb{R}^{n imes (d+1)} \ \hat{w} &= egin{bmatrix} w \ b \end{bmatrix} \in \mathbb{R}^{d+1} \ y &= egin{bmatrix} y_1 \ \dots \ y_n \end{bmatrix} \in \mathbb{R}^n \end{aligned}$$

· the loss function now looks like this:

$$L(\hat{w}) = (y - x\hat{w})^T (y - x\hat{w})$$

because we know that:

$$x\hat{w} = egin{bmatrix} f(x_1) \ \dots \ f(x_n) \end{bmatrix}$$

• to compute \hat{w} we want the gradient to be 0

$$\begin{split} \frac{\partial L(\hat{w})}{\partial \hat{w}} &= \frac{\partial [(y - x\hat{w})^T (y - x\hat{w})]}{\partial \hat{w}} \\ &= \frac{\partial (y^T y - y^T x \hat{w} - \hat{w}^T x^T y + \hat{w}^T x^T x \hat{w})}{\partial \hat{w}} \\ &= \frac{\partial (y^T y - 2y^T x \hat{w} + \hat{w}^T x^T x \hat{w})}{\partial \hat{w}} \\ &= -2x^T y + 2x^T x \hat{w} = 0 \end{split}$$

- now we focus on x^Tx :
 - 。我们知道 x^Tx 是一个实对称矩阵,那么它是半正定的,也就是说该矩阵的特征值都是大于等于0的
 - 。 如果 x^Tx 不满秩,那么显然这个方程里 \hat{w} 是可能有多于一个解的,造成这种结果的可能原因是:
 - d+1 > n:
 - $ullet rank(x^Tx) = rank(x) \leq min\{d+1,n\}$
 - x has repeated columns or rows
 - \circ 如果 x^Tx 满秩,则特征值都大于0,我们可以对其进行特征值分解:

$$x^T x = U \Lambda U^T$$

$$= U \begin{bmatrix} \lambda_1 & & & \\ & \ddots & & \\ & & \lambda_{d+1} \end{bmatrix} U^T$$

$$(x^T x)^{-1} = U \Lambda^{-1} U^T$$

$$= U \begin{bmatrix} \lambda_1^{-1} & & & \\ & \ddots & & \\ & & \lambda_{d+1}^{-1} \end{bmatrix} U^T$$

- 那么这个方程有**唯一解** $\hat{w} = (x^T x)^{-1} x^T y$
- 但依然有一个问题,就是 λ_{d+1} 可能很接近0,也就是 x^Tx 很接近一个非奇异矩阵,这就导致 λ_{d+1}^{-1} 会相当大,因此会导致计算上的问题(numerical issues and instability)

L2 Regularization

• Linear Regression with L2 Regularization (Ridge Regression):

$$minL(\hat{w}) + \lambda ||\hat{w}||_2^2$$

- $\circ \lambda$ is a positive **constant hyperparameter** which can control the strength of regularization (weight decay)
- 。 简单来说就是避免某个w特别突出,使所有的w更平均一些
- · loss function with L2 Regularization

$$J(\hat{w}) = (y - x\hat{w})^T (y - x\hat{w}) + \lambda w^T w \ (\lambda > 0)$$

· compute the gradient:

$$egin{aligned} rac{\partial J(\hat{w})}{\partial \hat{w}} &= -2x^Ty + 2x^Tx\hat{w} + 2\lambda\hat{w} \ &= -2x^Ty + 2(x^Tx + \lambda I)\hat{w} = 0 \end{aligned}$$

• 这是我们再对 $x^Tx + \lambda I$ 进行特征值分解,就可以得到:

$$x^T x + \lambda I = U(\Lambda + \lambda I)U^T$$

= $U\begin{bmatrix} \lambda_1 + \lambda & & & \\ & \dots & & \\ & & \lambda_{d+1} + \lambda \end{bmatrix}U^T$

so $x^Tx + \lambda I$ is always **non-singular** and \hat{w} is a unique solution!

$$\hat{w} = (x^T x + \lambda I)^{-1} x^T y$$

L1 Regularization

• Linear Regression with L1 Regularization (Lasso Regression):

$$egin{aligned} min\,L(\hat{w}) + \lambda ||\hat{w}||_1 \ = &min\,L(\hat{w}) + \lambda \sum_{i=1}^{d+1} |\hat{w}_i| \end{aligned}$$

- \circ induce sparsity of \hat{w} (many dimensions of \hat{w} will be 0)
- · Lasso stands for Least Absolute Shrinkage and Selection Operator
 - feature selection: 当参数很多时,用L1正则化可以筛选出真正有用的参数

Geometric view of Linear Regression

- Ideally, we want $X\hat{W}=y$
- 但是显然我们没办法找到一个model可以fit任何数据
- Consider the column space of X as col(X), spanned by columns of X
 - $\circ X\hat{W}$ lies in col(X), but y may not
 - \circ let $X\hat{W}=\hat{y}$, this \hat{W} is the least square solution to $X\hat{W}=y$
 - \circ Because \hat{y} minimizes $||y-X\hat{W}||^2$, which is the object of Linear Regression, $\hat{y}(X\hat{W})$ should satisfy $y-\hat{y}\perp col(X)\leftrightarrow X^T(y-\hat{y})=0 \leftrightarrow X^TX\hat{W}=X^Ty$