

# Solutions to Stochastic Differential Equations by Øksendal

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## 2. Some Mathematical Preliminaries

### Exercise 2.1

Suppose that  $X : \Omega \rightarrow \mathbb{R}$  is a function which takes only countably many values  $a_1, a_2, \dots \in \mathbb{R}$ .

(a) Show that  $X$  is a random variable if and only if

$$X^{-1}(a_i) \in \mathcal{F} \text{ for all } i \in \mathbb{N}.$$

(b) Suppose that  $X$  is a random variable. Show that

$$E[|X|] = \sum_{i=1}^{\infty} |a_i| P(X = a_i).$$

(c) If  $X$  is a random variable and  $E[|X|] < \infty$ , show that

$$E[X] = \sum_{i=1}^{\infty} a_i P(X = a_i).$$

(d) If  $X$  is a random variable and  $f : \mathbb{R} \rightarrow \mathbb{R}$  is measurable and bounded, show that

$$E[f(X)] = \sum_{i=1}^{\infty} f(a_i) P(X = a_i).$$

*Solution.*

For (a), suppose first that  $X$  is a random variable. Since  $\{a_i\}$  are Borel sets,  $X^{-1}(a_i) \in \mathcal{F}$  for all  $i \in \mathbb{N}$ . Conversely, assume that  $X^{-1}(a_i) \in \mathcal{F}$  for all  $a_i$ . Since the range of  $X$  is  $\{a_i\}_{i \in \mathbb{N}}$ , for any Borel set  $B \subset \mathbb{R}$ ,  $X^{-1}(B) = \bigcup_{a_i \in B} X^{-1}(a_i) \in \mathcal{F}$ , by the definition of  $\sigma$ -algebra. Thus,  $X$  is a random variable.

For (b), since  $X$  takes only countably many values, so does  $|X|$  with  $\{|a_i|\}_{i \in \mathbb{N}}$ . By the definition of expectation, we have

$$E[|X|] = \sum_{i=1}^{\infty} |a_i| P(X = a_i)$$

in the extended sense.

For (c), since  $E[|X|] < \infty$  and  $X$  is a random variable, the series converges absolutely and is well-defined. Hence

$$E[X] = \sum_{i=1}^{\infty} a_i P(X = a_i).$$

For (d), since  $f$  is measurable,  $f^{-1}(B)$  is Borel and  $X^{-1}f^{-1}(B)$  is measurable.  $f(X)$  takes

only countably many values,  $f(a_1), f(a_2), \dots$ . The definition of expectation gives us

$$E[f(X)] = \sum_{i=1}^{\infty} f(a_i)P(f(X) = f(a_i)) = \sum_{i=1}^{\infty} f(a_i)P(X = a_i).$$

■

### Exercise 2.2

$X : \Omega \rightarrow \mathbb{R}$  is a random variable. The distribution function  $F$  of  $X$  is defined as

$$F(x) = P(X \leq x).$$

(a) Prove that  $F$  has the following properties:

(i)  $0 \leq F \leq 1$ ,  $\lim_{x \rightarrow -\infty} F(x) = 0$  and  $\lim_{x \rightarrow \infty} F(x) = 1$ .

(ii)  $F$  is non-decreasing.

(iii)  $F$  is right-continuous.

(b)  $g : \mathbb{R} \rightarrow \mathbb{R}$  is measurable such that  $E[|g(X)|] < \infty$ . Show that

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)dF(x).$$

(c) Let  $p(x) \geq 0$  be measurable on  $\mathbb{R}$  be the density of  $X$ , i.e.,

$$F(x) = \int_{-\infty}^x p(t)dt.$$

Find density of  $B_t^2$ .

*Solution.*

For (a), since  $P$  is a probability measure,  $0 \leq P(S) \leq 1$  for any  $S \in \mathcal{F}$ . In particular,  $0 \leq P(X \leq x) \leq 1$  for all  $x \in \mathbb{R}$ . Also, we can take  $x_n \searrow -\infty$  and  $|X \leq x_n| \searrow \emptyset$  as  $n \rightarrow \infty$ . Hence

$$\lim_{x \rightarrow -\infty} F(x) = \lim_{n \rightarrow \infty} P(X \leq x_n) = P(\emptyset) = 0.$$

Similarly, we can take  $x_n \nearrow \infty$  and  $|X \leq x_n| \nearrow \Omega$  as  $n \rightarrow \infty$ . Hence

$$\lim_{x \rightarrow \infty} F(x) = \lim_{n \rightarrow \infty} P(X \leq x_n) = P(\Omega) = 1.$$

(i) is proved. For (ii),  $F$  is non-decreasing because if  $x_1 < x_2$ , then

$$F(x_1) = P(X \leq x_1) \leq P(X \leq x_2) = F(x_2).$$

For (iii), let  $h > 0$ .

$$F(x+h) - F(x) = P(X \leq x+h) - P(X \leq x) = P(x < X \leq x+h).$$

For any  $y > x$ , there exists  $h > 0$  such that  $y > x + h$ . Thus  $(x, x + h] \searrow \emptyset$  as  $h \rightarrow 0$ . Hence

$$F(x + h) - F(x) = P(x < X \leq x + h) \rightarrow P(\emptyset) = 0$$

as  $h \rightarrow 0$ . Therefore,  $F$  is right-continuous.

For (b), by definition of expectation, the left-hand side is

$$E[g(X)] = \int_{\mathbb{R}} g(x) d\mu_X(x),$$

where  $\mu_X(B) = P(X^{-1}(B))$  for any Borel set  $B \subset \mathbb{R}$ .

For (c),

$$F(x) = P(B_t^2 \leq x) = P(B_t \leq \sqrt{x}) = \int_{-\infty}^{\sqrt{x}} \frac{1}{\sqrt{2\pi t}} \exp\left(-\frac{u^2}{2t}\right) du.$$

Hence,

$$p(u) = \frac{d}{dx} F(x) = \frac{1}{\sqrt{2\pi t}} \exp\left(-\frac{x}{2t}\right) \frac{1}{2\sqrt{x}}.$$

■

### Exercise 2.3

Let  $\{\mathcal{F}_i\}_{i \in I}$  be a collection of  $\sigma$ -algebras on  $\Omega$ . Prove that

$$\mathcal{F} = \bigcap_{i \in I} \mathcal{F}_i$$

is again a  $\sigma$ -algebra.

*Solution.*

First, since  $\mathcal{F}_i$  are  $\sigma$ -algebras, they contain  $\emptyset$  and hence  $\emptyset \in \mathcal{F}$ . For any  $A \in \mathcal{F}$ ,  $A \in \mathcal{F}_i$  for all  $i \in I$  and hence  $A^c \in \mathcal{F}_i$  for all  $i \in I$ . Thus  $A^c \in \mathcal{F}$ . Finally, for any countable collection  $\{A_n\}_{n \in \mathbb{N}} \subset \mathcal{F}$ , we have  $A_n \in \mathcal{F}_i$  for all  $i \in I$  and all  $n \in \mathbb{N}$ . Then  $\bigcup_n A_n \in \mathcal{F}_i$  for all  $i \in I$ . Hence  $\bigcup_n A_n \in \mathcal{F}$ . Therefore,  $\mathcal{F}$  is a  $\sigma$ -algebra. ■

### Exercise 2.4

- (a) Let  $X : \Omega \rightarrow \mathbb{R}$  be a random variable such that  $E[|X|^p] < \infty$  for some  $p \in (0, \infty)$ . Prove the Chebyshev's inequality:

$$P(|X| \geq \lambda) \leq \frac{1}{\lambda^p} E[|X|^p]$$

for any  $\lambda > 0$ .

- (b) Suppose there exists  $k > 0$  such that  $M = E[\exp(k|X|)] < \infty$ . Prove that  $P(|X| \geq \lambda) \leq Me^{-k\lambda}$  for any  $\lambda > 0$ .

*Solution.*

For (a), directly estimate that

$$P(|X| \geq \lambda) = \int_{\Omega} \chi_{\{|X|^p \geq \lambda^p\}} dP \leq \int_{\Omega} \frac{|X|^p}{\lambda^p} dP = \frac{1}{\lambda^p} E[|X|^p].$$

(b) is similar:

$$P(|X| \geq \lambda) = \int_{\Omega} \chi_{\{\exp(k|X|) \geq \exp(k\lambda)\}} dP \leq \int_{\Omega} \exp(k|X|) \exp(-k\lambda) dP = M \exp(-k\lambda).$$

■

### Exercise 2.5

Let  $X, Y : \Omega \rightarrow \mathbb{R}$  be two independent random variables and assume for simplicity that  $X, Y$  are bounded. Prove that

$$E[XY] = E[X] E[Y].$$

*Solution.*

For any  $\epsilon > 0$ , by definition of the expectation, we can find simple functions  $s$  and  $t$  on  $\Omega$  such that

$$\int |s - X| dP < \epsilon, \quad \int |t - Y| dP < \epsilon, \quad \Rightarrow \quad \left| E[X] - \int s dP \right| < \epsilon, \quad \left| E[Y] - \int t dP \right| < \epsilon,$$

where  $s$  and  $t$  can be written as

$$s = \sum_{i=1}^n s_i \chi_{X^{-1}[s_i, s_{i+1})} \quad \text{and} \quad t = \sum_{j=1}^m t_j \chi_{Y^{-1}[t_j, t_{j+1})},$$

with  $s_i$  and  $t_j$  being arranged in ascending order. Thus,

$$\begin{aligned} \int s t dP &= \sum_{i=1}^n \sum_{j=1}^m s_i t_j P(\{X \in [s_i, s_{i+1})\} \cap \{Y \in [t_j, t_{j+1})\}) \\ &= \sum_{i=1}^n \sum_{j=1}^m s_i t_j P(X \in [s_i, s_{i+1})) P(Y \in [t_j, t_{j+1})) \\ &= \left( \sum_{i=1}^n s_i P(X \in [s_i, s_{i+1})) \right) \left( \sum_{j=1}^m t_j P(Y \in [t_j, t_{j+1})) \right) = \left( \int s dP \right) \left( \int t dP \right). \end{aligned}$$

Also,

$$\left| E[XY] - \int s t dP \right| \leq \left| \int |X - s| |t| dP \right| + \left| \int |Y - t| |X| dP \right|.$$

$X$  and  $Y$  are bounded, say by  $M$  and  $N$  respectively. Then  $t$  is also bounded by  $N$  from our construction. Thus

$$\left| E[XY] - \int s t dP \right| \leq M\epsilon + N\epsilon.$$

Combine the results above, we arrive at

$$\begin{aligned}
|E[XY] - E[X]E[Y]| &\leq \left| E[XY] - \int s dP \right| + \left| E[X]E[Y] - \int s dP \int t dP \right| \\
&\leq (M + N)\epsilon + \left| E[X] - \int s dP \right| \left| \int t dP \right| + \left| E[Y] - \int t dP \right| |E[X]| \\
&\leq (M + N)\epsilon + \epsilon N + \epsilon M.
\end{aligned}$$

Since  $\epsilon$  is arbitrary, we conclude that  $E[XY] = E[X]E[Y]$ . ■

### Exercise 2.6

Let  $(\Omega, \mathcal{F}, P)$  be a probability space and  $A_1, \dots \in \mathcal{F}$  be sets such that

$$\sum_{i=1}^{\infty} P(A_i) < \infty.$$

Prove the Borel-Cantelli lemma:

$$P\left(\bigcap_{m=1}^{\infty} \bigcup_{i=m}^{\infty} A_i\right) = 0.$$

*Solution.*

Set  $B_m = \bigcup_{i=m}^{\infty} A_i$  be measurable. Then

$$P(B_m) \leq \sum_{i=m}^{\infty} P(A_i) \rightarrow 0$$

as  $m \rightarrow \infty$  by the assumption. Thus

$$P\left(\bigcap_{m=1}^{\infty} B_m\right) \leq \lim_{n \rightarrow \infty} P\left(\bigcap_{m=1}^n B_m\right) \leq \lim_{n \rightarrow \infty} P(B_n) = 0.$$

■

### Exercise 2.7

(a) Suppose  $G_1, \dots, G_n$  are disjoint sets in  $\mathcal{F}$  such that  $\bigcup_{i=1}^n G_i = \Omega$ . Prove that the family

$$\mathcal{G} = \{G \mid G \text{ is a union of some } G_i\} \cup \{\emptyset\}$$

is a  $\sigma$ -algebra.

(b) Prove that every finite  $\sigma$ -algebra is of type  $\mathcal{G}$  as in (a).

(c) Let  $\mathcal{F}$  be a finite  $\sigma$ -algebra on  $\Omega$  and  $X : \Omega \rightarrow \mathbb{R}$  be  $\mathcal{F}$ -measurable. Prove that  $X$  is simple.

*Solution.*

For (a), first,  $\emptyset \in \mathcal{G}$  by definition. Let  $G \in \mathcal{G}$ . Then  $G = \cup_{i \in I} G_i$  for some  $I \subset \{1, \dots, n\}$ , with the convention that  $\cup_{i \in \emptyset} G_i = \emptyset$ . Then  $G^c = \cup_{i \notin I} G_i \in \mathcal{G}$ . Lastly, for countably many  $G_i \in \mathcal{G}$ , since  $\mathcal{G}$  is finite, there are in fact finitely many distinct  $G_i$  and the union must lie in  $\mathcal{G}$  by the definition. Hence  $\mathcal{G}$  is a  $\sigma$ -algebra.

For (b), let  $\mathcal{F}$  be a finite  $\sigma$ -algebra. Consider the collection

$$\mathcal{S} = \{S \in \mathcal{F} \mid S \cap F = \emptyset \text{ or } S \text{ for all } F \in \mathcal{F}\}.$$

Since  $\mathcal{F}$  is finite,  $\mathcal{S}$  is also finite. We first check that every distinct sets in  $\mathcal{S}$  are disjoint. Suppose not. There are  $S_1, S_2 \in \mathcal{S}$  such that  $S_1 \cap S_2$  is non-empty. Then  $S_1 \cap S_2 = S_1 = S_2$ , contradicting the assumption that  $S_1$  and  $S_2$  are distinct. Thus every distinct sets in  $\mathcal{S}$  are disjoint. Next, we check that  $\cup_{S \in \mathcal{S}} S = \Omega$ . If not, let  $A = \Omega \setminus \cup_{S \in \mathcal{S}} S$  be non-empty and  $A \cap F$  is a non-empty proper subset of  $A$  for some  $F \in \mathcal{F}$ . But then  $A \cap F$  or  $A \cap F^c$  must satisfy the condition that there is some  $F' \in \mathcal{F}$  such that  $A \cap F \cap F'$  or  $A \cap F^c \cap F'$  is non-empty, proper subset of  $A \cap F$  or  $A \cap F^c$  respectively. Note that  $F' \neq F$  and the process continues. In the end, we can find a infinite sequence of distinct sets lying in  $\mathcal{F}$ , contradicting the finiteness of  $\mathcal{F}$ . Thus  $\cup_{S \in \mathcal{S}} S = \Omega$ . Finally, by (a),

$$\mathcal{G} = \{G \mid G \text{ is a union of some } S \in \mathcal{S}\} \cup \{\emptyset\}$$

is a  $\sigma$ -algebra. It remains to show that  $\mathcal{G} = \mathcal{F}$ . Clearly,  $\mathcal{G} \subset \mathcal{F}$  since  $\mathcal{S} \subset \mathcal{F}$ . For any  $F \in \mathcal{F}$ , we can write  $F = \cup_{i=1}^n S_i$  for some  $S_i \in \mathcal{S}$ . Thus  $F \in \mathcal{G}$ . We end up with  $\mathcal{G} = \mathcal{F}$ .

For (c), suppose that  $X$  can take infinitely many values  $\{a_i\}_{i \in I}$ . Since  $X$  is  $\mathcal{F}$ -measurable,  $X^{-1}(\{a_i\}) \in \mathcal{F}$  for all  $i \in I$ . In particular,  $X^{-1}(\{a_i\})$  and  $X^{-1}(\{a_j\})$  are disjoint for all  $i \neq j$ . This implies that  $\mathcal{F}$  contains infinitely many disjoint sets, contradicting the finiteness of  $\mathcal{F}$ . Thus  $X$  can only take finitely many values and is simple. ■

### Exercise 2.8

Let  $B_t$  be Brownian motion on  $\mathbb{R}$ ,  $B_0 = 0$ . Put  $E = E^0$ .

(a) Prove that

$$E \left[ e^{iuB_t} \right] = e^{-\frac{u^2 t}{2}} \text{ for all } u \in \mathbb{R}.$$

(b) Use the power series expansion of the exponential function to show that

$$E \left[ B_t^{2k} \right] = \frac{(2k)!}{2^k k!} t^k \text{ for all } k \in \mathbb{N}.$$

(c) Prove that

$$E \left[ f(B_t) \right] = \int_{\mathbb{R}} f(x) \frac{1}{\sqrt{2\pi t}} e^{-\frac{x^2}{2t}} dx$$

for all measurable functions  $f$  on  $\mathbb{R}$  such that the integral is finite. Deduce (b) by setting  $f(x) = x^{2k}$ .

(d) Now suppose that  $B_t$  is a  $n$ -dimensional Brownian motion. Prove that

$$E^x [|B_t - B_s|^4] = n(n+2) |t-s|^2$$

for all  $n \in \mathbb{N}$  and  $0 \leq s, t \leq T$ .

*Solution.*

For (a), directly compute the expectation:

$$E [e^{iuB_t}] = \int_{-\infty}^{\infty} e^{iux} \frac{1}{\sqrt{2\pi t}} e^{-\frac{x^2}{2t}} dx = e^{-\frac{u^2 t}{2}} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi t}} e^{-\frac{(x-iut)^2}{2t}} dx = e^{-\frac{u^2 t}{2}}.$$

For (b), note that the power series expansion of the exponential function gives

$$e^{iuB_t} = \sum_{k=0}^{\infty} \frac{(iuB_t)^k}{k!} \Rightarrow E [e^{iuB_t}] = \sum_{k=0}^{\infty} \frac{(iu)^k}{k!} E [B_t^k].$$

The right-hand side is

$$e^{-\frac{u^2 t}{2}} = \sum_{k=0}^{\infty} \frac{1}{k!} \left( -\frac{u^2 t}{2} \right)^k = \sum_{k=0}^{\infty} \frac{(-1)^k u^{2k}}{2^k k!} t^k.$$

For those two expressions to be equal, as a function of  $u$ ,  $E [B_t^k] = 0$  for odd  $k$  and we may rewrite the first expression as

$$E [e^{iuB_t}] = \sum_{k=0}^{\infty} \frac{(-1)^k u^{2k}}{(2k)!} E [B_t^{2k}].$$

By comparing the coefficients,

$$E [B_t^{2k}] = \frac{(-1)^k t^k}{2^k k!} \cdot \frac{(2k)!}{(-1)^k} = \frac{(2k)!}{2^k k!} t^k.$$

For (c), it is clear that  $B_t$  has the density  $p(x) = \frac{1}{\sqrt{2\pi t}} e^{-\frac{x^2}{2t}}$ . It follows that

$$E [f(B_t)] = \int f(x) \frac{1}{\sqrt{2\pi t}} e^{-\frac{x^2}{2t}} dx.$$

By setting  $f : x \mapsto x^{2k}$ , we have

$$\begin{aligned} E [B_t^{2k}] &= \int x^{2k} \frac{1}{\sqrt{2\pi t}} e^{-\frac{x^2}{2t}} dx \\ &= x^{2k+1} \frac{1}{\sqrt{2\pi t}} e^{-\frac{x^2}{2t}} \Big|_{-\infty}^{\infty} - \int 2kx^{2k} \frac{1}{\sqrt{2\pi t}} e^{-\frac{x^2}{2t}} - x^{2k} \frac{1}{\sqrt{2\pi t}} e^{-\frac{x^2}{2t}} \frac{x^2}{t} dx \\ &= \frac{1}{t} E [B_t^{2(k+1)}] - 2k E [B_t^{2k}] \Rightarrow E [B_t^{2(k+1)}] = (2k+1)t E [B_t^{2k}]. \end{aligned}$$



And also when  $k = 1$ ,  $E[B_t^2] = t$ . Suppose  $E[B_t^{2k}] = \frac{(2k)!}{2^k k!} t^k$ . Then

$$E[B_t^{2(k+1)}] = (2k+1)t \frac{(2k)!}{2^k k!} t^k = \frac{(2(k+1))!}{2^{k+1} (k+1)!} t^{k+1}.$$

The conclusion follows by induction.

For (d), if  $n = 1$ ,  $B_t - B_s \sim N(0, |t - s|)$  and  $E^x[|B_t - B_s|^4] = 3|t - s|^2$ . Now suppose that for  $n$ -dimensional  $B_t$ ,  $E^x[|B_t - B_s|^4] = n(n+2)|t - s|^2$ . Then for  $n+1$ -dimensional  $B_t$ ,

$$\begin{aligned} E^x[|B_t - B_s|^4] &= \int_{\mathbb{R}^n} \int_{\mathbb{R}} (|y|^2 + z^2)^2 \frac{1}{\sqrt{(2\pi)^{n+1} |t - s|^{n+1}}} \exp\left(-\frac{|y|^2 + z^2}{2|t - s|}\right) dy dz \\ &= \int_{\mathbb{R}^n} |y|^4 \frac{1}{\sqrt{(2\pi)^n |t - s|^n}} \exp\left(-\frac{|y|^2}{2|t - s|}\right) dy \\ &\quad + 2|t - s| \int_{\mathbb{R}^n} |y|^2 \frac{1}{\sqrt{(2\pi)^n |t - s|^n}} \exp\left(-\frac{|y|^2}{2|t - s|}\right) dy + 3|t - s|^2 \\ &= n(n+2)|t - s|^2 + 2n|t - s|^2 + 3|t - s|^2 = (n+1)(n+3)|t - s|^2. \end{aligned}$$

Thus by induction, the conclusion holds for all  $n \in \mathbb{N}$ . It follows from the Kolmogorov's continuity theorem that we can always set  $B_t$  to be a continuous process. ■

### Exercise 2.9

Let  $(\Omega, \mathcal{F}, P) = ([0, \infty), \mathcal{B}, \mu)$  be a probability space where  $\mu$  is a probability measure such that there is no mass at single points. Define

$$X_t(\omega) = \begin{cases} 1 & \text{if } t = \omega, \\ 0 & \text{if } t \neq \omega. \end{cases} \quad \text{and} \quad Y_t(\omega) = 0 \text{ for all } (t, \omega) \in [0, \infty) \times [0, \infty).$$

Prove that  $\{X_t\}$  and  $\{Y_t\}$  have the same distributions and  $X_t$  is a version of  $Y_t$ . And yet  $t \mapsto Y_t(\omega)$  is continuous for all  $\omega$ , while  $t \mapsto X_t(\omega)$  is discontinuous for all  $\omega$ .

*Solution.*

First, given any  $t$ ,  $X_t$  and  $Y_t$  are both random variables. For  $t_1, \dots, t_k \in [0, \infty)$ , consider the sets  $F_1, \dots, F_k \in \mathcal{B}$ . Since  $Y_t$  is constant,

$$P(Y_{t_1} \in F_1, \dots, Y_{t_k} \in F_k) = \mathbb{1}_{\{0 \in \cap_{i=1}^k F_i\}}.$$

Also, since  $\mu$  has no mass at single points,

$$P(X_{t_1} \in F_1, \dots, X_{t_k} \in F_k) = \mathbb{1}_{\{0 \in \cap_{i=1}^k F_i\}}.$$

Thus,  $X_t$  and  $Y_t$  have the same distributions. Furthermore, for any  $t \in [0, \infty)$ ,  $\{X_t = Y_t\} = \Omega \setminus \{t\}$ , which has zero measure and hence  $X_t$  is a version of  $Y_t$ . Now since  $Y_t$  is constant,  $t \mapsto Y_t(\omega)$  is continuous for all  $\omega$ . On the other hand, for any  $\omega \in [0, \infty)$ ,  $X(t, \omega)$  is discontinuous

at  $t = \omega$ , proving that  $t \mapsto X_t(\omega)$  is discontinuous for all  $\omega$ . ■

### Exercise 2.10

Prove that the Brownian motion  $B_t$  has stationary increments, i.e., given  $h > 0$ , the process  $\{B_{t+h} - B_t\}$  has the same distributions for all  $t$ .

*Solution.*

For any  $h > 0$ ,  $B_{t+h} - B_t \sim N(0, h)$ . The stationarity follows immediately. ■

### Exercise 2.11

If  $B_t = (B_t^{(1)}, \dots, B_t^{(n)})$  is an  $n$ -dimensional Brownian motion, then the component processes  $B_t^{(i)}$ ,  $1 \leq i \leq n$ , are independent Brownian motions.

*Solution.*

Since the  $B_t$  is continuous almost surely, the component processes  $B_t^{(i)}$  are continuous almost surely as well. Now we may regard the component processes as projections of  $B_t$  and hence they are normally distributed.  $E[B_t^{(i)}] = 0$  since  $E[B_t] = 0$ . Also,  $\text{Cov}(B_t^{(i)}, B_t^{(j)}) = t\delta_{ij}$  as  $\text{Var}(B_t) = tI$ . Since for  $i \neq j$  the covariance is zero, the component processes are independent. ■

### Exercise 2.12

Let  $B_t$  be a Brownian motion and fix  $t_0 \geq 0$ . Prove that the process  $\tilde{B}_t = B_{t+t_0} - B_{t_0}$  is a Brownian motion.

*Solution.*

First, it is clear that  $\tilde{B}_0 = B_{t_0} - B_{t_0} = 0$ . Since  $B_t$  is almost surely continuous,  $B_{t+t_0} - B_{t_0}$  is also almost surely continuous. Also,  $B_{t+t_0} - B_{t_0} \sim N(0, t)$ . For  $s < t < u$ ,

$$\text{Cov}(\tilde{B}_u - \tilde{B}_t, \tilde{B}_t - \tilde{B}_s) = \text{Cov}(B_{u+t_0} - B_{t+t_0}, B_{t+t_0} - B_{s+t_0}) = 0.$$

Thus  $\tilde{B}_t$  has independent increments. This shows that  $\tilde{B}_t$  is a Brownian motion. ■

### Exercise 2.13

Let  $B_t$  be 2-dimensional Brownian motion and put

$$D_\rho = \{x \in \mathbb{R}^2 \mid |x| < \rho\}$$

for  $\rho > 0$ . Compute  $P^0(B_t \in D_\rho)$ .

*Solution.*

Since  $B_t \sim N(0, tI)$ ,

$$P^0(B_t \in D_\rho) = \int_{D_\rho} \frac{1}{2\pi t} \exp\left(-\frac{|x|^2}{2t}\right) dx = \int_0^{2\pi} \int_0^\rho \frac{1}{2\pi t} \exp\left(-\frac{r^2}{2t}\right) r dr d\theta = 1 - \exp\left(-\frac{\rho^2}{2t}\right).$$

■

**Exercise 2.14**

Let  $B_t$  be  $n$ -dimensional Brownian motion and  $K \subset \mathbb{R}^n$  be a measure zero set under the Lebesgue measure. Prove that the expected total length of time that  $B_t$  spends in  $K$  is zero.

*Solution.*

Given  $\omega \in \Omega$ , the process  $B_t(\omega)$  spends

$$\int_0^\infty \mathbb{1}_{\{B_t(\omega) \in K\}} dt$$

amount of time in  $K$ . The expected total length of time that  $B_t$  spends in  $K$  is

$$\int_\Omega \int_0^\infty \mathbb{1}_{\{B_t(\omega) \in K\}} dt dP(\omega) = \int_0^\infty \int_\Omega \mathbb{1}_{\{B_t(\omega) \in K\}} dP(\omega) dt = 0$$

by the Fubini-Tonelli theorem, since  $K$  has measure zero under the Lebesgue measure and  $B_t^{-1}(K)$  is measure zero under the probability measure  $P$  for all  $t \geq 0$ . ■

**Exercise 2.15**

Let  $B_t$  be an  $n$ -dimensional Brownian motion starting at 0 and let  $U \in \mathbb{R}^{n \times n}$  be a orthogonal matrix, i.e.,  $U^T U = I$ . Prove that  $\tilde{B}_t = UB_t$  is also a Brownian motion.

*Solution.*

Since  $B_t$  is a Brownian motion and  $U$  is a linear transformation on a finite-dimensional space,  $\tilde{B}_t$  must be continuous almost surely and  $\tilde{B}_0 = UB_0 = 0$ . Also, since  $B_t \sim N(0, tI)$ , we have  $\tilde{B}_t \sim N(0, tU^T U) = N(0, tI)$ . For  $s < t < u$ ,

$$\text{Cov}(\tilde{B}_u - \tilde{B}_t, \tilde{B}_t - \tilde{B}_s) = \text{Cov}(UB_u - UB_t, UB_t - UB_s) = U \text{Cov}(B_u - B_t, B_t - B_s) U^T = 0.$$

Thus  $\tilde{B}_t$  has independent increments. Therefore,  $\tilde{B}_t$  is a Brownian motion. ■

**Exercise 2.16**

Let  $B_t$  be a Brownian motion on  $\mathbb{R}$  and  $c > 0$ . Prove that the process  $X_t = \frac{1}{c} B_{c^2 t}$  is a Brownian motion.

*Solution.*

First,  $X_0 = \frac{1}{c} B_0 = 0$ . Suppose  $\omega \in \Omega$  is such that  $B_t(\omega)$  is continuous. For  $\epsilon > 0$ , there exists  $\delta > 0$  such that  $|B_t(\omega) - B_s(\omega)| < c\epsilon$  as long as  $|t - s| < \delta$ . Then we may set  $\delta' = \delta/c^2$  and see that

$$|X_t(\omega) - X_s(\omega)| = \frac{1}{c} |B_{c^2 t}(\omega) - B_{c^2 s}(\omega)| < \epsilon$$

whenever  $|t - s| < \delta' \Leftrightarrow |c^2 t - c^2 s| < \delta$ . Thus  $X_t$  is continuous almost surely. Also, since  $B_{c^2 t} \sim N(0, c^2 tI)$ ,  $X_t \sim N(0, tI)$ . For  $s < t < u$ ,

$$\text{Cov}(X_u - X_t, X_t - X_s) = \text{Cov}\left(\frac{1}{c} B_{c^2 u} - \frac{1}{c} B_{c^2 t}, \frac{1}{c} B_{c^2 t} - \frac{1}{c} B_{c^2 s}\right) = \frac{1}{c^2} \text{Cov}(B_{c^2 u} - B_{c^2 t}, B_{c^2 t} - B_{c^2 s}) = 0.$$

We conclude that  $X_t$  is a Brownian motion. ■

### Exercise 2.17

Let  $B_t$  be a Brownian motion on  $\mathbb{R}$ . Show that the quadratic variation process is  $\langle B, B \rangle_t^2(\omega) = t$  almost surely by the following steps:

- (a) Define  $\Delta B_k = B_{t_{k+1}} - B_{t_k}$  and put  $Y(t, \omega) = \sum_{t_k \leq t} (\Delta B_k)^2$  for a partition  $\{t_k\}_{k=1}^n$  of  $[0, t]$ . Show that

$$E[(Y_t - t)^2] = 2 \sum_{t_k \leq t} (\Delta t_k)^2$$

and deduce that  $Y(t, \cdot) \rightarrow t$  in  $L^2$  as  $\Delta t_k \rightarrow 0$ .

- (b) Use (a) to show that almost every paths of  $B_t$  does not have a bounded total variation on  $[0, t]$ .

*Solution.*

For (a), note that  $\Delta B_k$  are independent and  $\Delta B_k \sim N(0, \Delta t_k)$ .

$$\begin{aligned} E[(Y_t - t)^2] &= E\left[\left(\sum_{t_k \leq t} (\Delta B_k)^2 - \Delta t_k\right)^2\right] = E\left[\sum_{t_k \leq t} ((\Delta B_k)^2 - \Delta t_k)^2\right] \\ &= \sum_{t_k \leq t} E[(\Delta B_k)^4] - 2E[(\Delta B_k)^2] \Delta t_k + (\Delta t_k)^2 \\ &= \sum_{t_k \leq t} 3(\Delta t_k)^2 - 2(\Delta t_k)^2 + (\Delta t_k)^2 \\ &= 2 \sum_{t_k \leq t} (\Delta t_k)^2 \leq 2 \max_k \Delta t_k \sum_{t_k \leq t} \Delta t_k = 2t \max_k \Delta t_k \rightarrow 0 \end{aligned}$$

as  $\max_k \Delta t_k \rightarrow 0$ , the right-hand side converges to 0 and hence  $Y(t, \cdot) \rightarrow t$  in  $L^2(P)$ . Hence  $\langle B, B \rangle_t^2(\omega) = t$  almost surely.

For (b), let  $\mathcal{P}$  be a partition of  $[0, t]$ . With respect to  $\mathcal{P}$ , define the total variation of  $B_t$  as

$$Z_t^{\mathcal{P}}(\omega) = \sum_{t_k \leq t} |B_{t_{k+1}}(\omega) - B_{t_k}(\omega)| \leq N(\mathcal{P}) \sum_{t_k \leq t} |B_{t_{k+1}}(\omega) - B_{t_k}(\omega)|^2 \rightarrow \infty$$

as  $\|\mathcal{P}\| \rightarrow 0$  by the Cauchy inequality and (a) that  $Y(t, \cdot) \rightarrow t$  in  $L^2(P)$ . ■

### Exercise 2.18

Let  $\Omega = \{1, 2, 3, 4, 5\}$  and  $\mathcal{U} = \{\{1, 2, 3\}, \{3, 4, 5\}\}$  be a collection of subsets of  $\Omega$ .

- (a) Find  $\sigma(\mathcal{U})$ .

- (b) Define  $X : \Omega \rightarrow \mathbb{R}$  by

$$X(1) = X(2) = 0, \quad X(3) = 10, \quad X(4) = X(5) = 1.$$

Is  $X$   $\sigma(\mathcal{U})$ -measurable?

(c) Define  $Y : \Omega \rightarrow \mathbb{R}$  by

$$Y(1) = 0, \quad Y(2) = Y(3) = Y(4) = Y(5) = 1.$$

Find  $\sigma(Y)$ .

*Solution.*

For (a),

$$\sigma(\mathcal{U}) = \{\emptyset, \{1, 2\}, \{3\}, \{4, 5\}, \{1, 2, 3\}, \{3, 4, 5\}, \{1, 2, 4, 5\}, \Omega\}.$$

For (b), it is not hard to verify that  $X$  is  $\sigma(\mathcal{U})$ -measurable.

For (c),

$$\sigma(Y) = \{\emptyset, \{1\}, \{2, 3, 4, 5\}, \Omega\}.$$

■

### Exercise 2.19

*Prove that every convergent sequence is a Cauchy sequence.*

*Solution.*

Let  $\{x_n\}$  be a convergent sequence with limit  $x$ . For any  $\epsilon > 0$ , we may find  $N \in \mathbb{N}$  such that for all  $n \geq N$ ,  $d(x_n, x) < \epsilon/2$ . Then for all  $m, n \geq N$ ,

$$d(x_n, x_m) \leq d(x_n, x) + d(x_m, x) < \epsilon/2 + \epsilon/2 = \epsilon.$$

Hence  $\{x_n\}$  is a Cauchy sequence.

■

### Exercise 2.20

*Let  $B_t$  be a 1-dimensional Brownian motion,  $\sigma \in \mathbb{R}$  and  $0 \leq s < t$ . Prove that*

$$E [\exp(\sigma(B_s - B_t))] = \exp\left(\frac{\sigma^2(s - t)}{2}\right).$$

*Solution.*

Since  $B_s - B_t \sim N(0, t - s)$ ,

$$\begin{aligned} E [\exp(\sigma(B_s - B_t))] &= \int_{-\infty}^{\infty} e^{\sigma x} \frac{1}{\sqrt{2\pi(t-s)}} e^{-\frac{x^2}{2(t-s)}} dx \\ &= e^{\frac{\sigma^2(t-s)}{2}} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi(t-s)}} e^{-\frac{(x - \sigma(t-s))^2}{2(t-s)}} dx = e^{\frac{\sigma^2(t-s)}{2}}. \end{aligned}$$

■

### 3. Itô Integrals

#### Exercise 3.1

Prove that

$$\int_0^t s dB_s = tB_t - \int_0^t B_s ds.$$

*Solution.*

Let  $\mathcal{P}_k = \{0 = t_0 < t_1 < \dots < t_k = t\}$  be a partition of the interval  $[0, t]$ . Denote  $B_{t_i}$  as  $B_i$ . We have that

$$\sum_i t_i \Delta B_i = t_{k-1} B_k - \sum_i \Delta t_i B_{i+1}.$$

Now take  $\|\mathcal{P}_k\| \rightarrow 0$ . Then

$$\int_0^t s dB_s = tB_t - \int_0^t B_s ds.$$

■

#### Exercise 3.2

Prove that

$$\int_0^t B_s^2 dB_s = \frac{1}{3} B_t^3 - \int_0^t B_s ds.$$

*Solution.*

Let  $\mathcal{P}_k = \{0 = t_0 < \dots < t_k = t\}$  be a partition. Denote  $B_{t_i}$  as  $B_i$  and  $\Delta B_i = B_{t_{i+1}} - B_{t_i}$ . Now

$$\begin{aligned} B_t^3 &= \sum_i B_{i+1}^3 - B_i^3 = \sum_i (B_i + \Delta B_i)^3 - B_i^3 \\ &= \sum_i B_i^3 + 3B_i^2 \Delta B_i + 3B_i (\Delta B_i)^2 + (\Delta B_i)^3 - B_i^3 \\ &= \sum_i 3B_i^2 \Delta B_i + 3B_i (\Delta B_i)^2 + (\Delta B_i)^3. \end{aligned}$$

Now take  $\|\mathcal{P}_k\| \rightarrow 0$ . Then

$$\sum_i 3B_i^2 (\Delta B_i) \rightarrow 3 \int_0^t B_s^2 dB_s$$

by the definition. For the second term, let  $X_t = 3B_t$ . Notice that by (d) of [exercise 2.8](#),

$$\begin{aligned} E \left[ \left( \sum_i X_i (\Delta B_i)^2 - \sum_i X_i \Delta t_i \right)^2 \right] &= E \left[ \sum_i X_i^2 \left( (\Delta B_i)^2 - \Delta t_i \right)^2 + 2 \sum_{i>j} X_i X_j \left( (\Delta B_i)^2 - \Delta t_i \right) \left( (\Delta B_j)^2 - \Delta t_j \right) \right] \\ &= \sum_i E \left[ X_i^2 \left( (\Delta B_i)^2 - \Delta t_i \right)^2 \right] = \sum_i E \left[ X_i^2 \right] (3(\Delta t_i)^2 - 2(\Delta t_i)^2 + (\Delta t_i)^2) \\ &= \sum_i E \left[ X_i^2 \right] (\Delta t_i)^2 \rightarrow 0 \end{aligned}$$

as  $\|\mathcal{P}_k\| \rightarrow 0$  since  $E[X_i^2]$  is bounded. But  $\sum_i X_i \Delta t_i \rightarrow \int_0^t X_s ds$  by definition, so

$$\sum_i 3B_i(\Delta B_i)^2 \rightarrow 3 \int_0^t B_s ds.$$

Finally, we note that

$$\begin{aligned} E \left[ \left( \sum_i (\Delta B_i)^3 \right)^2 \right] &= E \left[ \sum_i (\Delta B_i)^6 + 2 \sum_{i>j} (\Delta B_i)^3 (\Delta B_j)^3 \right] \\ &= \sum_i E [(\Delta B_i)^6] = 15 \sum_i (\Delta t_i)^3 \rightarrow 0 \end{aligned}$$

as  $\|\mathcal{P}_k\| \rightarrow 0$ . Hence we see that

$$B_t^3 = 3 \int_0^t B_s^2 dB_s + 3 \int_0^t B_s ds \quad \Rightarrow \quad \int_0^t B_s^2 dB_s = \frac{1}{3} B_t^3 - \int_0^t B_s ds.$$

■

### Exercise 3.3

Let  $X_t : \Omega \rightarrow \mathbb{R}^n$  is a stochastic process and  $\mathcal{H}_t$  is the filtration of  $X_t$ .

- (a) Show that if  $X_t$  is a martingale with respect to some filtration  $\mathcal{N}_t$ , then it is also a martingale with respect to  $\mathcal{H}_t$ .
- (b) Show that if  $X_t$  is a martingale with respect to  $\mathcal{H}_t$ , then

$$E[X_t] = E[X_0]$$

for all  $t \geq 0$ .

- (c) Give an example of stochastic process  $X_t$  such that  $E[X_t] = E[X_0]$  for all  $t \geq 0$  but  $X_t$  is not a martingale with respect to  $\mathcal{H}_t$ .

*Solution.*

For (a), for  $X_t$  being a martingale with respect to  $\mathcal{N}_t$ , we have  $\mathcal{H}_t \subset \mathcal{N}_t$  for all  $t \geq 0$ . Also,  $X_t$  is integrable. Finally,

$$E[X_s | \mathcal{H}_t] = E[E[X_s | \mathcal{N}_t] | \mathcal{H}_t] = E[X_t | \mathcal{H}_t] = X_t$$

for all  $s \geq t$  by the tower property. Hence  $X_t$  is a martingale with respect to  $\mathcal{H}_t$  as well.

For (b), note that  $E[X_t | \mathcal{H}_0] = X_0$ . Hence  $E[X_t] = E[E[X_t | \mathcal{H}_0]] = E[X_0]$ .

For (c), consider the process defined by

$$X_0 = \begin{cases} -1 & \text{with prob.} = 0.5 \\ 1 & \text{with prob.} = 0.5 \end{cases}, \quad X_t = (t+1) \operatorname{sgn}(X_0) \text{ for } t > 0.$$

Then  $E[X_t] = t \cdot 0.5 + (-t) \cdot 0.5 = 0$  for all  $t \geq 0$ , but  $E[X_t | X_0] = E[(t+1) \operatorname{sgn}(X_0) | X_0] = (t+1) \operatorname{sgn}(X_0) \neq X_0$  for  $t > 0$ . Hence  $X_t$  is not a martingale with respect to  $\mathcal{H}_t$ . ■

### Exercise 3.4

Check whether the following processes are martingale with respect to  $\{\mathcal{F}_t\}$ .

(a)  $X_t = B_t + 4t$ .

(b)  $X_t = B_t^2$ .

(c)  $X_t = t^2 B_t - 2 \int_0^t s B_s ds$ .

(d)  $X_t = B_1(t)B_2(t)$ , where  $(B_1(t), B_2(t))$  is a 2-dimensional Brownian motion.

*Solution.*

For (a),

$$E[X_s | \mathcal{F}_t] = E[B_s | \mathcal{F}_t] + 4s = B_t + 4s \neq X_t.$$

Hence  $X_t$  is not a martingale.

For (b),

$$\begin{aligned} E[X_s | \mathcal{F}_t] &= E[B_s^2 | \mathcal{F}_t] = E[(B_t + (B_s - B_t))^2 | \mathcal{F}_t] \\ &= E[B_t^2 + 2B_t(B_s - B_t) + (B_s - B_t)^2 | \mathcal{F}_t] \\ &= B_t^2 + (s - t) \neq X_t. \end{aligned}$$

Hence  $X_t$  is not a martingale.

For (c),

$$\begin{aligned} E[X_s | \mathcal{F}_t] &= E\left[s^2 B_s - 2 \int_0^s u B_u du \mid \mathcal{F}_t\right] = s^2 B_t - 2 \int_0^t u B_u du - 2 \int_t^s u E[B_u | \mathcal{F}_t] du \\ &= s^2 B_t - 2 \int_0^t u B_u du - 2 B_t \int_t^s u du = t^2 B_t - 2 \int_0^t u B_u du = X_t. \end{aligned}$$

Hence  $X_t$  is a martingale.

For (d),

$$E[B_1(s)B_2(s) | \mathcal{F}_t] = E[B_1(s) | \mathcal{F}_t] E[B_2(s) | \mathcal{F}_t] = B_1(t)B_2(t).$$

Hence  $X_t$  is a martingale. ■

### Exercise 3.5

Prove that  $M_t = B_t^2 - t$  is an  $\mathcal{F}_t$ -martingale.

*Solution.*

Compute that

$$E[M_s | \mathcal{F}_t] = E[B_s^2 | \mathcal{F}_t] - s = E[(B_t + (B_s - B_t))^2 | \mathcal{F}_t] - s = B_t^2 + (s - t) - s = B_t^2 - t = M_t.$$



Hence  $M_t$  is a martingale. ■

### Exercise 3.6

Prove that  $N_t = B_t^3 - 3tB_t$  is a martingale.

*Solution.*

Compute that

$$\begin{aligned}
 E[N_s | \mathcal{F}_t] &= B_t^3 - 3tB_t + E[B_s^3 - B_t^3 | \mathcal{F}_t] - 3E[sB_s - tB_t | \mathcal{F}_t] \\
 &= N_t + E[(B_s - B_t)^3 + 3B_s^2B_t - 3B_sB_t^2 | \mathcal{F}_t] - 3B_t(s - t) \\
 &= N_t + 3E[B_s^2B_t - B_sB_t^2 | \mathcal{F}_t] - 3B_t(s - t) \\
 &= N_t + 3B_tE[(B_s - B_t)^2 - 2B_t(B_s - B_t) + B_t^2 | \mathcal{F}_t] - 3B_t^3 - 3B_t(s - t) \\
 &= N_t + 3B_t(s - t) + 3B_t^3 - 3B_t^3 - 3B_t(s - t) = N_t.
 \end{aligned}$$

Hence  $N_t$  is a martingale. ■

### Exercise 3.7

A famous result from Itô gives the following formula

$$n! \int_{0 \leq u_1 \leq u_2 \leq \dots \leq u_n \leq t} dB_{u_1} dB_{u_2} \dots dB_{u_n} = t^{n/2} h_n\left(\frac{B_t}{\sqrt{t}}\right),$$

where  $h_n$  is the Hermite polynomial of degree  $n$ , defined by

$$h_n(x) = (-1)^n e^{x^2/2} \frac{d^n}{dx^n} \left[ e^{-x^2/2} \right].$$

- (a) Verify that for each  $n \in \mathbb{N}$ , the integrand satisfies the requirement of definition 3.1.4; in other words, the Itô integral is well-defined.
- (b) Verify the formula for  $n = 1, 2, 3$ .
- (c) Use (b) to give another proof of [exercise 3.6](#).

*Solution.*

For (a), we show this by induction. If  $n = 1$ , the integral becomes  $\int_0^t dB_{u_1}$  and the integrand is constant. Since  $(t, \omega) \rightarrow 1$  is clearly  $\mathcal{B} \times \mathcal{F}$ -measurable,  $\omega \rightarrow 1$  is also  $\mathcal{F}_t$ -measurable for each  $t \geq 0$  and thus  $\mathcal{F}_t$ -adapted. Finally,  $E\left[\int_0^t 1^2 dt\right] = t < \infty$ , so the integrand is in  $L^2$  and the Itô integral is well-defined. Suppose the claim holds for  $n$ , for  $n + 1$ , the integrand is

$$f(u_1, \omega) = \int_{u_1}^t \int_{u_2}^t \dots \int_{u_n}^t dB_{u_{n+1}} \dots dB_{u_3} dB_{u_2}.$$

■