solutions01

September 16, 2016

0.1 Linear regression

0.1.1 Part A

```
\begin{split} f(\beta) &= \sum_{i=1}^{N} \{ \frac{\omega_i}{2} (y_i - x_i \beta)^2 \} \\ &= (y - X\beta)^T \omega (y - X\beta) \\ &= y^T \omega y - 2y^T \omega X\beta + \beta^T X^T \omega X\beta \\ \text{Therefore,} \\ \frac{\partial f\beta}{\partial \beta} &= 0 - 2y^T \omega X + 2\beta^T X^T \omega X = 0 \\ X^T \omega y &= (X^T \omega X)\beta \end{split}
```

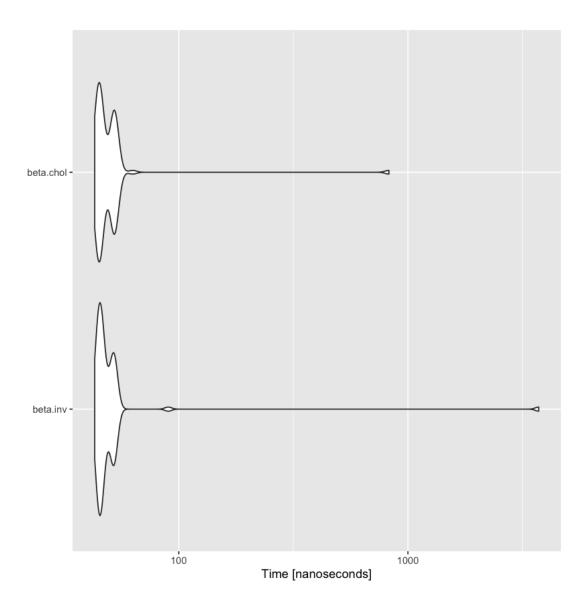
0.1.2 Part B

pseudo code for my method (Cholesky Factorization) $X^T W X = L L^T$

0.1.3 Part C

```
In [23]: # generate simulated data
P = 1000 # variable size
N = 2000 # sample size
set.seed(1)
beta = rep(1,P)
X = matrix(rnorm(N*P), nrow=N)
y = X%*%beta + rnorm(N,sd=0.05)
W = diag(rep(1, N))
```

In [26]: # experiment on Cholesky method
 beta.chol = Cholesky.method(X, y, W)

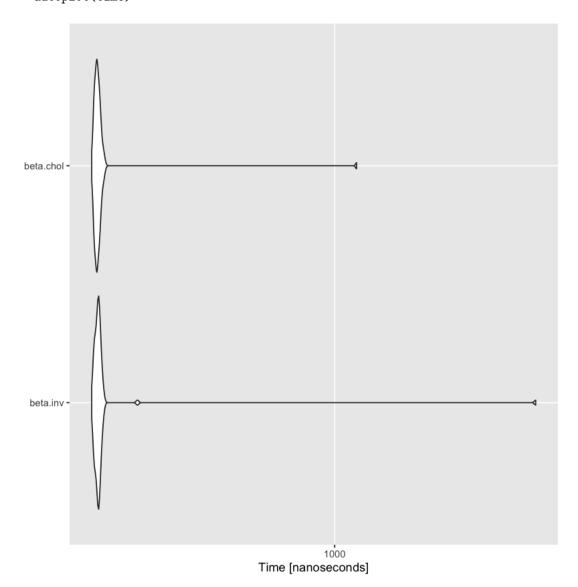


0.1.4 Part D

```
In [28]: # generate sparse data
    P = 1000  # variable size
    N = 2000  # sample size
    sparse = 0.1  # dense of sparse matrix
    set.seed(1)
    beta = rep(1,P)
    X = matrix(rnorm(N*P), nrow=N)
    mask = matrix(rbinom(N*P, 1, sparse), nrow=N)
    X = mask * X
    y = X%*%beta + rnorm(N,sd=0.05)
    W = diag(rep(1, N))
In [29]: # experiment on inversion method for sparse data beta.inv = inversion.method(X, y, W)
```

In [30]: # experiment on Cholesky method for sparse data beta.chol = Cholesky.method(X, y, W)

In [31]: # timing time = microbenchmark(beta.inv, beta.chol, times=100) autoplot(time)



Generalized linear regression 0.2

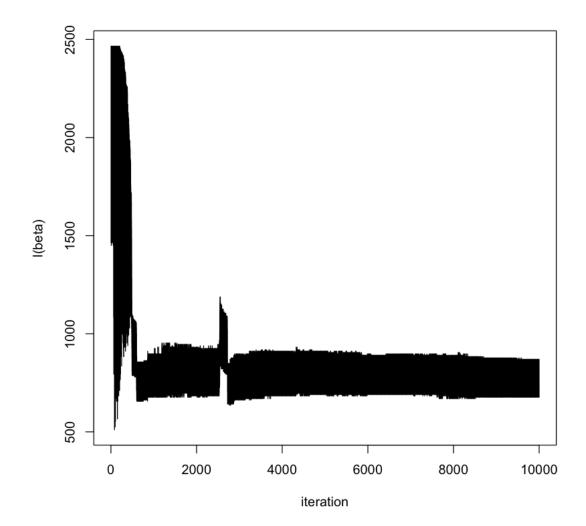
0.2.1 Part A

Negative log likelihood:
$$\begin{split} l(\beta) &= -log\{\prod_{i=1}^N p(y_i|\beta)\} \\ &= -log\{\prod_{i=1}^N C*\omega_i^{y_i}*(1-\omega_i)^{m_i-y_i}\} \end{split}$$

$$\begin{split} &= -\sum_{i=1}^{N} \{logC + y_i * log\omega_i + (m_i - y_i) * log(1 - \omega_i)\} \\ &= -\sum_{i=1}^{N} \{y_i * log\frac{\omega_i}{1 - \omega_i} + m_i * log(1 - \omega_i) + C\} \\ &= -\sum_{i=1}^{N} \{y_i * log\frac{\omega_i}{1 - \omega_i} + m_i * log(1 - \omega_i) + C\} \\ &= -\sum_{i=1}^{N} \{y_i * loge^{x\beta} - m_i * log(1 + e^{x\beta})\} \\ &= -\sum_{i=1}^{N} \{y_i * loge^{x\beta} - m_i * log(1 + e^{x\beta})\} \\ &\text{Gradient of the expression:} \\ &\nabla l(\beta) = \frac{\partial l(\beta)}{\partial \beta} \\ &= -\sum_{i=1}^{N} \{\frac{y_i}{\omega_i * (1 - \omega_i)} - \frac{m_i}{1 - \omega_i}\} \frac{\partial \omega_i}{\partial \beta} \\ &= -\sum_{i=1}^{N} \{\frac{y_i}{\omega_i * (1 + e^{x_i * \beta})^2} \\ &\text{Therefore, } \nabla l(\beta) = -\sum_{i=1}^{N} (y_i - m_i * \omega_i) * x_i = -X^T * (y - m * w) \end{split}$$

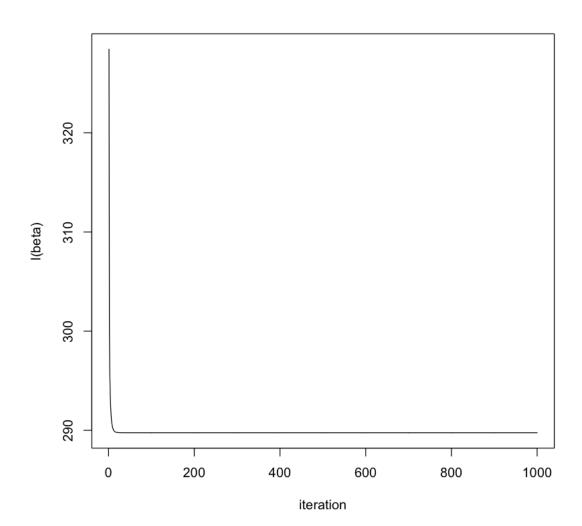
0.2.2 Part B

```
In [21]: # import functions for gradient descent
         source('~/Box Sync/PhDCourses/SDS385Statistical_models_for_big_data/SDS385/solutions/exercises
In [22]: setwd('~/Box Sync/PhDCourses/SDS385Statistical_models_for_big_data/SDS385/data')
         # import wdbs.csv data
         data <- read.csv('wdbc.csv', header=F)</pre>
         y.BM <- data[ ,2]
         y <- rep(0, length(y.BM))
         y[y.BM == "B"] <- 1
         y <- as.matrix(y)</pre>
         x.variables <- as.matrix(data[ , 3:12])</pre>
         x <- cbind(x.variables, rep(1,dim(x.variables)[1]))</pre>
         colnames(x) <- NULL
In [24]: # gradient descent
         betas = matrix(1, 11, 1)
         GD.results <- gradient.descent(x, y, betas, step.size=0.01, max.iter=10000)
In [26]: # plot l(beta)
         1.beta.GD = GD.results[[2]]
         plot(1:10000, l.beta.GD, type='l', xlab='iteration', ylab='l(beta)')
```



The log likelihood value never converage even with 10k iterations when using wdbc.csv data. So next I tried gradient descent algorithm on simulated data –

```
In [18]: # generate simulated data
    set.seed(666)
    x1 = rnorm(1000)  # some continuous variables
    x2 = rnorm(1000)
    x = cbind(as.matrix(x1), as.matrix(x2))
    x = cbind(x, rep(1,1000))  # add one column for intercept
    z = 1 + 2*x[ ,1] + 3*x[ ,2]  # linear combination with a bias
    prob = 1/(1+exp(-z))  # pass through an inv-logit function
    y = as.matrix(rbinom(1000,1,prob))
In [20]: # gradient descent on simulated data.
    betas = matrix(1, 3, 1)
    GD.results <- gradient.descent(x, y, betas, step.size=0.01, max.iter=1000)</pre>
```



0.2.3 Part C

The second-order Taylot approximation:
$$q(\beta;\beta_0) = l(\beta_0) + (\beta - \beta_0)l'(\beta) + \frac{1}{2}(\beta - \beta_0)^2 l''(\beta)$$
 Derive $l''(\beta)$ from $l'(\beta)$
$$l''(\beta) = -\sum_{i=1}^{N} \{x_i^T(y_i - m_i * \omega_i)\}$$

$$l'''(\beta) = -\sum_{i=1}^{N} \{x_i^T(0 - m_i) * \frac{\partial \omega_i}{\partial \beta}\}$$

$$= \sum \{x_i^T m_i \frac{e^{x_i \beta}}{1 + e^{x_i \beta}} x_i \frac{1}{1 + e^{x_i \beta}}\}$$

$$= \sum \{x_i^T m_i \omega_i (1 - \omega_i) x_i\}$$

```
Therefore, l''(\beta) = X^T S X, where S is a diagonal matrix with elements m_i \omega_i (1 - \omega_i) Thus, q(\beta; \beta_0) = l(\beta_0) - (\beta - \beta_0)^T X^T (y - m_i \omega_i) + \frac{1}{2} (\beta - \beta_0)^T X^T S X (\beta - \beta_0) q(\beta; \beta_0) = C - (\beta^T X^T - \beta_0^T X^T) (y - m_i \omega_i) + \frac{1}{2} (\beta^T X^T - \beta_0^T X^T) S (X \beta - X \beta_0) = \frac{1}{2} (X \beta - X \beta_0 - S^{-1} (y - m \omega))^T S (X \beta - X \beta_0 - S^{-1} (y - m \omega)) + C Compare the above equation with the original format -z = X \beta + S^{-1} (y - m \omega) W = S
```

0.2.4 Part D

