

1. A special case of a normal family is one in which the mean and the variance are related, the  $N(\theta, a\theta)$  family. If we are interested in testing this relationship, regardless of the value of  $\theta$ , we are again faced with a nuisance parameter problem.
  - (a) Find the LRT of  $H_0 : a = 1$  versus  $H_1 : a \neq 1$  based on a sample  $X_1, \dots, X_n$  from a  $N(\theta, a\theta)$  family, where  $\theta$  is unknown.
  - (b) A similar question can be asked about a related family, the  $N(\theta, a\theta^2)$  family. Thus if  $X_1, \dots, X_n$  are iid  $N(\theta, a\theta^2)$ , where  $\theta$  is unknown, find the LRT of  $H_0 : a = 1$  versus  $H_1 : a \neq 1$ .
2. Show that for a random sample  $X_1, \dots, X_n$  from a  $N(0, \sigma^2)$  population, the most powerful test of  $H_0 : \sigma = \sigma_0$  versus  $H_1 : \sigma = \sigma_1$ , where  $\sigma_0 < \sigma_1$ , is given by

$$\phi(\sum X_i^2) = \begin{cases} 1 & \text{if } \sum X_i^2 > c \\ 0 & \text{if } \sum X_i^2 \leq c \end{cases}$$

For a given value of  $\alpha$ , the size of the Type I Error, show how the value of  $c$  is explicitly determined.

3. Let  $X_1, X_2, \dots, X_n$  be a random sample from  $\text{Uniform}(\theta, \theta + 1)$ . To test  $H_0 : \theta = 0$  versus  $H_1 : \theta > 0$ , reject  $H_0$  if  $X_{(n)} \geq 1$  or  $X_{(1)} > k$  where  $k$  is a constant.
  - (a) Determine  $k$  such that the test has size  $\alpha$ .
  - (b) Find an expression for the power function of the test in part (a).
  - (c) Prove that the test is UMP size  $\alpha$  test.
4. Let  $\{f_\theta(x)\}$  be a family of density functions with parameter  $\theta \in \mathbb{R}$ . Assume that  $f_\theta(x) > 0$  for all  $\theta$  and  $x$  and that  $\frac{\partial^2}{\partial \theta \partial x} \log f_\theta(x)$  exists. Show that this family has MLR in  $x$  is equivalent to one of the following conditions:
  - (a)  $\frac{\partial^2}{\partial \theta \partial x} \log f_\theta(x) \geq 0$  for all  $x$  and  $\theta$ ;
  - (b)  $f_\theta(x) \frac{\partial^2}{\partial \theta \partial x} f_\theta(x) \geq \frac{\partial}{\partial \theta} f_\theta(x) \frac{\partial}{\partial x} f_\theta(x)$  for all  $x$  and  $\theta$ .
5. Consider tests for  $H_0 : \theta = 0$  versus  $H_1 : \theta \neq 0$  based on a single observation  $X$  from  $N(\theta, 1)$ . Consider using  $Y = |X|$  in a test.
  - (a) Find the probability density  $f(y|\theta)$  of  $Y$  and show that it depends only on  $|\theta|$ .
  - (b) Show that the probability density  $f(y|\theta)$  have MLR.
  - (c) Find the UMP size  $\alpha$  test based on  $Y$ .
  - (d) The UMP test in part (c) is not most powerful compared with tests based on  $X$ . Find a level  $\alpha$  test based on  $x$  that have better power at  $\theta = -1$ . What is the difference of power at  $\theta = -1$  for  $\alpha = 0.05$ ?
6. Let  $f(x|\theta)$  be the Cauchy scale pdf

$$f(x|\theta) = \frac{\theta}{\pi} \frac{1}{\theta^2 + x^2}, \theta > 0.$$

- (a) Show that this family does not have an MLR.

- (b) If  $X$  is one observation from  $f(x|\theta)$ , show that  $|X|$  is sufficient for  $\theta$ , and that the distribution of  $|X|$  does have an MLR.
7. Consider the hypothesis test  $H_0 : \mu = 0$  versus  $H_1 : \mu = \mu_1$  for a normal population  $N(\mu, 1)$  with unknown mean  $\mu$ .
- (a) Given a sample size  $n$ , find test such that the Type-I and Type-II error probability are equal  $\alpha = 1 - \beta$ .
  - (b) For the test in part(a), if we want to control  $\alpha = 1 - \beta = \gamma$ . How much sample do we need?
  - (c) Find the Wald'd sequential test such that the Type-I and Type-II error probability are equal  $\alpha = 1 - \beta$ . Using approximated formula to compute the expected number of sample when the test is terminated.
  - (d) Compare the two sample sizes.