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### 3. C-statistic for the time-varying AUC

For the same theta, using AUC type weight matrix seems to have a higher value.  
(AUC C > Harrell C > Uno C)

```
AUCC_Wmat <- function(Y, delta, tau){  
  
  # An index indicates whether the observation is censored  
  censor = ifelse(delta==0,1,0)  
  
  # Censoring Distribution Estimate using Kaplan-Meier Estimator  
  KM_G = survfit(formula = Surv(Y ,censor) ~ 1)  
  
  
  # Consider the effect of tau  
  ordered_time = KM_G$time[KM_G$time<tau]  
  difference = diff(ordered_time)  
  
  # Get G(Y) for each observation  
  # (Since G(Y) in KM_G is ordered we want each G(Y) to match original Y)  
  G_y = KM_G$surv[match(Y,KM_G$time)]  
  
  
  # Get the survival function for each Tk  
  S_t = c()  
  for (k in 1:length(KM_G$time)){  
    invG_y = (G_y)^(-1)  
    invG_y[is.infinite(invG_y)] = 0  
    S_tk = sum((Y>KM_G$time[k])*delta*invG_y)/length(Y)  
    S_t = c(S_t,S_tk)  
  }  
  
  # n > K  
  n <- length(Y)  
  K = length(ordered_time)  
  Wmat <- matrix(0, n, n)  
  for(i in 1:n) {
```

```

if(delta[i] != 0) {
  VALUE = matrix(0,K,n)
  for (k in 1:K){
    if (k == 1){ # In this case, k-1 = 0, t0 = 0
      if (S_t[k] ==1){
        # S(t_k) = 1, 1-S(t_k) = 0, denominator = 0
        # but we define the value = 0
        value = rep(0,n)
      }else{
        value = (Y[i]<ordered_time[k])*(ordered_time[k]<Y)*
          ordered_time[k]/(1-S_t[k])
      }

      VALUE[k,] = value
    }else{ # k > 1
      if (S_t[k] ==1 | S_t[k-1]==0){
        # S(t_k) = 1, 1-S(t_k) = 0, denominator = 0
        # but we define the value = 0
        value = rep(0,n)
      }else{
        value = (Y[i]<ordered_time[k])*(ordered_time[k]<Y)*
          difference[k-1]/(S_t[k-1]*(1-S_t[k])*KM_G$surv[k-1])
      }

      VALUE[k,] = value
    }
  }

  Wmat[i,] <- delta[i]*colSums(VALUE)/(G_y[i]*n^2)
}
}

Wmat <- Wmat/sum(Wmat)
return(Wmat)
}

```

```

theta_test = colMeans(result_test$THETA)
Wmat1 = HarrellC_Wmat(Y,delta,tau)
Wmat2 = UnoC_Wmat(Y,delta,tau)
Wmat3 = AUCC_Wmat(Y,delta,tau)
data.frame(HarrellC= HarrellC(theta_test,Wmat1),
           UnoC = HarrellC(theta_test,Wmat2),
           AUCC = HarrellC(theta_test,Wmat3))

```

```

##      HarrellC      UnoC      AUCC
## 1 0.6321131 0.6197375 0.6825279

```

## 4. Gaussian processes MCMC implementation

1. Based on what I have implemented before, I add the procedure of updating the lambda.
2. The function matrix K is used to produce the covariance matrix.

Each element (i,j) represents the kernel for input  $x_i$  and  $x_j$ .

3. My GP\_MH Sampling algorithm is even more time consuming! Needs improvement.

```
matrix_K <- function(X,lambda){
  n = dim(X)[1]
  p = dim(X)[2]
  cov_K = matrix(0,n,n)
  for (i in 1:n){
    cov_K[i,] = exp(-0.5*colSums((t(X)[,i]-t(X))^2/lambda))
  }
  return (cov_K)
}
```

```
GPMH_Sampling <- function(Y,delta,tau,
                          A,beta0,var.prop,
                          m,eta){

  accept_beta = 0
  accept_lambda = 0
  beta = beta0
  lambda = lambda0

  # What we want to record
  BETA = matrix(0,m,dim(A)[1]) # The dimension is different from LR
  LAMBDA = matrix(0,m,dim(A)[2])
  C_stat = c()
  # Dimension beta: nx1
  # Dimension lambda: px1

  for (i in 1:m){

    # Sample beta from proposal distribution
    beta.p = t(rmvnorm(1,beta,var.prop))

    # Compute C-statistics from current and last iteration
    Wmat <- HarrellC_Wmat(Y, delta, tau)
    HC.p = HarrellC(beta.p, Wmat) # Since in Gaussian Process, theta = beta
    HC = HarrellC(beta, Wmat)

    # Record C-statistics from last iteration
    C_stat = c(C_stat,HC)

    # Compute log of MH ratio

    lrMH = eta*log(HC.p) +
      # Note: I modify this part
```

```

        dmvnorm(as.numeric(beta.p),beta0,var.prop,log=T)-
        eta*log(HC) -
        dmvnorm(as.numeric(beta.p),beta0,var.prop,log=T)

    if (log(runif(1))<lrMH){
        beta = beta.p
        accept_beta = accept_beta + 1
    }
    BETA[i,] = beta

    # Sample lamda from proposal distribution
    lambda.p = exp(t(rnorm(dim(A)[2],log(lambda),rep(1,dim(A)[2]))))

    # Compute log of MH_lambda ratio
    var.prop.p = matrix_K(A,as.numeric(lambda.p))

    lrMH_lambda = dmvnorm(as.numeric(beta),beta0,var.prop.p,log=T)-
        dmvnorm(as.numeric(beta),beta0,var.prop,log=T)

    if (log(runif(1))<lrMH_lambda){
        lambda = lambda.p
        var.prop = var.prop.p
        accept_lambda = accept_lambda + 1
    }
    LAMBDA[i,] = lambda
}

return(list(BETA=BETA,
            LAMBDA = LAMBDA,
            accept_beta=accept_beta/m,
            accept_lambda=accept_lambda/m,
            C_stat = C_stat))
}

```

```

beta0 = rep(0,dim(A)[1])
lambda0 = rep(1,dim(A)[2])
var.prop = matrix_K(A,lambda0)

```

```

m = 100

```

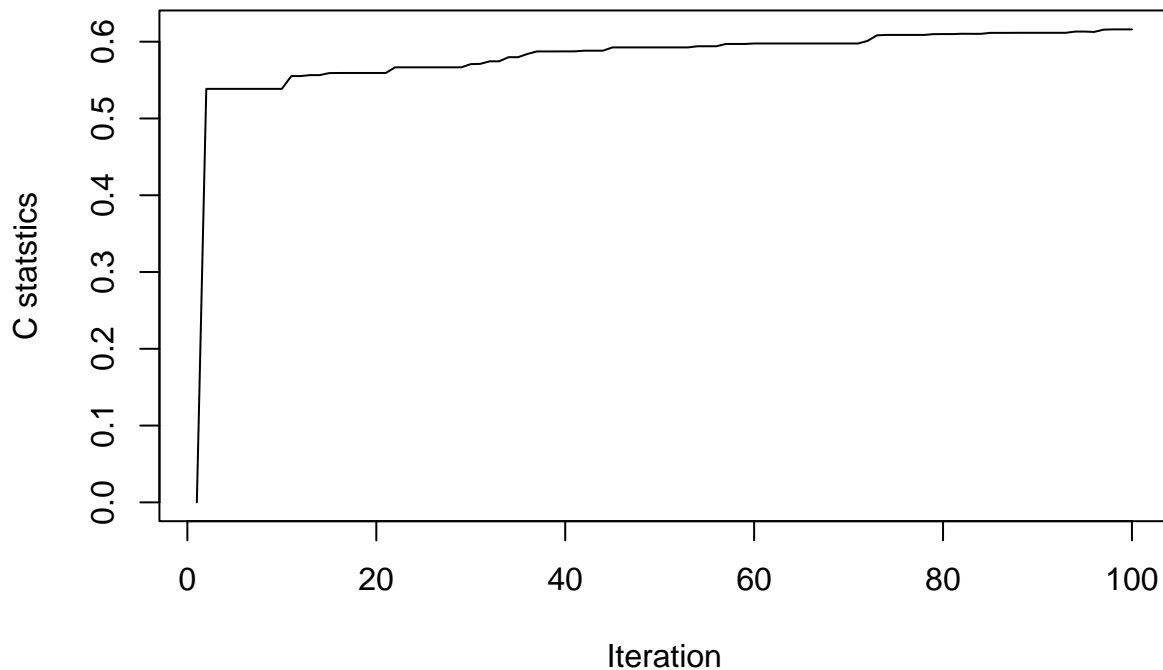
```

system.time({
    result_GP = GPMH_Sampling(Y,delta,tau,
                             A,beta0,var.prop,
                             m,eta)
})

```

```
##      user  system elapsed
## 378.13    9.06  607.60
```

```
plot(1:m,result_GP$C_stat,xlab = "Iteration",
     ylab = "C statistics",type = "l")
```



## Analysis of Why it's so time consuming

It takes a much longer time to sample multivariate normal sample for  $n = 686$

```
# Gaussian Process Case
beta0 = rep(0,dim(A)[1]) # nx1 vector
lambda0 = rep(1,dim(A)[2])
var.prop = matrix_K(A,lambda0)
system.time({ beta.p = t(rmvnorm(1,beta0,var.prop))})
```

```
##      user  system elapsed
##   3.02    0.06   3.96
```

```
# Linear Regression Case
beta0 = rep(0,dim(A)[2]) #
kappa = 10
var.prop = kappa*solve(t(A)%*%A)
system.time({ beta.p = t(rmvnorm(1,beta0,var.prop))})
```

##	user	system	elapsed
##	0	0	0