

# Progress Report

Objective of Score matching on non-negative data

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- 1. Review of Score Matching
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  - (1) Weighted Score Matching
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## Score Function

$$\nabla_x \log p(x) = \left( \frac{\partial}{\partial x_1} \log p(x), \dots, \frac{\partial}{\partial x_m} \log p(x) \right)^\top \quad (x \in \mathbb{R}^m)$$

**Key property:** The score function is *invariant under normalization*.

Let  $p(x) = \frac{\tilde{p}(x)}{Z}$ , where  $Z = \int \tilde{p}(x) dx$  is the partition function. Then:

$$\nabla_x \log p(x) = \nabla_x \log \tilde{p}(x) \quad \text{since } \log p(x) = \log \tilde{p}(x) - \log Z$$

The normalization constant  $Z$  disappears under differentiation.

# Objective of Score Matching

## Score Matching Loss

$$J(p) = \int_{\mathbb{R}^m} \|\nabla_x \log p(x) - \nabla_x \log q(x)\|^2 q(x) dx$$

This loss measures the discrepancy between the score functions of the model  $p$  and the data distribution  $q$ , using an expectation under  $q(x)$ .

$$\hat{\theta} = \operatorname{argmin}_{\theta} J(p_{\theta}, q) = \operatorname{argmin}_{\theta} \int_{\mathbb{R}^m} \|\nabla_x \log p_{\theta}(x) - \nabla_x \log q(x)\|^2 q(x) dx$$

We estimate the model parameter  $\theta$  by minimizing the score-matching objective.

# Approach to Reformulating the Objective

We begin by expressing the score matching loss as:

$$J = \int_{\mathbb{R}^m} F(x) q(x) dx \iff J = \mathbb{E}_{x \sim q}[F(x)]$$

This suggests we can approximate  $J$  using the empirical average:

$$\mathbb{E}_{x \sim q}[F(x)] \approx \frac{1}{n} \sum_{i=1}^n F(x_i)$$

However, in its current form,  $F(x)$  still depends on  $q(x)$ , which is unknown.

**Goal:** Reformulate the objective so that

$$J = \int_{\mathbb{R}^m} F'(x) q(x) dx$$

where  $F'(x)$  no longer involves  $q(x)$ . This enables unbiased empirical estimation.

# Fisher Divergence as a Decomposed Divergence

- Expanding the squared norm in  $J(p)$  gives:

$$\int \|\nabla_x \log p(x)\|^2 q(x) dx - 2 \int \nabla_x \log p(x)^\top \nabla_x \log q(x) q(x) dx + \int \|\nabla_x \log q(x)\|^2 q(x) dx$$

- This can be written in the form:

$$J(p) = g(q) + d(p, q)$$

where

$$g(q) = \int \|\nabla_x \log q(x)\|^2 q(x) dx,$$

$$d(p, q) = \int \|\nabla_x \log p(x)\|^2 q(x) dx - 2 \int \nabla_x \log p(x)^\top \nabla_x \log q(x) q(x) dx$$

# Structure of $J(p) = d(p, q)$

Since  $g(q)$  does not depend on the model  $p$ , minimizing  $J(p)$  is equivalent to minimizing  $d(p, q)$ . **Thus, we redefine  $J(p) := d(p, q)$  for the remainder of this presentation.:**

$$J(p) = \int \|\nabla_x \log p(x)\|^2 q(x) dx - 2 \int \nabla_x \log p(x)^\top \nabla_x \log q(x) q(x) dx$$

This objective consists of two terms:

- The first term depends only on the model  $p$ , via its score function.
- The second term couples the model score  $\nabla_x \log p(x)$  with the data score  $\nabla_x \log q(x)$ , and is problematic because  $\nabla_x \log q(x)$  is unknown.

**Goal:** Eliminate the dependence on the unknown  $\nabla_x \log q(x)$  using integration by parts.

# Step 1: Substituting the Score Function

We begin with the problematic term:

$$-2 \int \nabla_x \log p(x)^\top \nabla_x \log q(x) q(x) dx.$$

Using the identity:

$$\nabla_x \log p(x) = \frac{\nabla_x q(x)}{q(x)} \quad \Rightarrow \quad \nabla_x \log q(x) \cdot q(x) = \nabla_x q(x),$$

we rewrite:

$$-2 \int \nabla_x \log p(x)^\top \nabla_x \log q(x) q(x) dx = -2 \int \nabla_x \log p(x)^\top \nabla_x q(x) dx.$$

**This step assumes:**

- $q(x) > 0$  almost everywhere,
- $q \in C^1(\mathbb{R}^d)$ ,
- The integral is well-defined and finite.



## Step 2: Integration by Parts

We now handle the term:

$$-2 \int \nabla_x \log p(x)^\top \nabla_x q(x) dx.$$

Component-wise:

$$= -2 \sum_{i=1}^m \int \frac{\partial}{\partial x_i} \log p(x) \cdot \frac{\partial q(x)}{\partial x_i} dx.$$

By integration by parts:

$$= 2 \sum_{i=1}^m \int \frac{\partial^2}{\partial x_i^2} \log p(x) \cdot q(x) dx = 2 \int \Delta_x \log p(x) \cdot p(x) dx.$$

**Assumption:** boundary term vanishes,

$$\lim_{\|x\| \rightarrow \infty} \frac{\partial}{\partial x_i} \log p(x) \cdot q(x) = 0.$$

# Final Form of $J(p)$

Combining the two terms, we have:

$$J(p) = \int \|\nabla_x \log p(x)\|^2 q(x) dx + 2 \int \Delta_x \log p(x) q(x) dx.$$

## Key features:

- The expression depends only on the model  $p$ ,
- The expectation is taken under  $q(x)$ , which can be approximated from data.

This forms the basis of the score matching objective.

# Empirical Score Matching Objective

From the previous analysis, we obtained the following quantity to minimize:

$$J(p) = \int (\|\nabla_x \log p(x)\|^2 + 2\Delta_x \log p(x)) q(x) dx.$$

This is an expectation over the data distribution  $p(x)$ , which is unknown.

However, given i.i.d. samples  $x^{(1)}, \dots, x^{(n)} \sim p(x)$ , we approximate it as:

$$\hat{J}(p) = \frac{1}{n} \sum_{i=1}^n \left( \|\nabla_x \log p(x^{(i)})\|^2 + 2\Delta_x \log p(x^{(i)}) \right).$$

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# Score Matching for Non-negative Data

## Weighted Score Matching

Let  $h_1, \dots, h_m : \mathbb{R}_+ \rightarrow \mathbb{R}_+$  be a.s. positive functions that are absolutely continuous on every bounded subinterval of  $\mathbb{R}_+$ , and set  $h(x) = [h_1(x_1), \dots, h_m(x_m)]^\top$ , which is absolutely continuous on  $\mathbb{R}_+^m$ .

Then, the weighted score matching objective is defined as

$$J_h(p) = \int_{\mathbb{R}_+^m} \left\| h(x)^{1/2} \odot \nabla_x \log p(x) - h(x)^{1/2} \odot \nabla_x \log q(x) \right\|^2 q(x) dx,$$

where  $h(x)^{1/2} = [h_1(x_1)^{1/2}, \dots, h_m(x_m)^{1/2}]^\top$ , and  $\odot$  is denote the element-wise product

$$(\mathbf{y} \odot \mathbf{z} = [y_1 \cdot z_1, \dots, y_m \cdot z_m] \quad \mathbf{y}, \mathbf{z} \in \mathbb{R}^m)$$

# Rewriting Weighted Score Matching Objective

We expand the weighted score matching loss:

$$J_h(p) = \int_{\mathbb{R}_+^m} \left\| h(x)^{1/2} \odot (\nabla_x \log p(x) - \nabla_x \log q(x)) \right\|^2 q(x) dx - C(q) =$$

$$\int_{\mathbb{R}_+^m} \left\| h(x)^{1/2} \odot \nabla_x \log p(x) \right\|^2 q(x) dx - 2 \int_{\mathbb{R}_+^m} (h(x) \odot \nabla_x \log p(x))^\top (h(x) \odot \nabla_x \log q(x)) q(x) dx -$$

where  $C(q) = \int \left\| h(x)^{1/2} \odot \nabla_x \log q(x) \right\|^2 q(x) dx$  is constant in  $p$ .

**Goal:** Eliminate dependence on  $\nabla_x \log q(x)$  via integration by parts.

# Integration by Parts with Weight Function

We now handle the cross term:

$$\begin{aligned} & \int_{\mathbb{R}_+^m} \left( h(x)^{1/2} \odot \nabla_x \log p(x) \right)^\top \left( h(x)^{1/2} \odot \nabla_x \log q(x) \right) q(x) dx \\ &= \int_{\mathbb{R}_+^m} (h(x) \odot \nabla_x \log p(x))^\top \nabla_x \log q(x) \cdot q(x) dx = \int (h(x) \odot \nabla_x \log p(x))^\top \nabla_x q(x) dx \end{aligned}$$

Then, by component-wise integration by parts:

$$= - \int \operatorname{div}(h(x) \odot \nabla_x \log p(x)) \cdot q(x) dx + \text{boundary term}$$

**Assumption:** boundary term vanishes or is handled separately.

# Final Objective: Weighted Score Matching

Combining all terms, we get:

$$J_h(p) = \int_{\mathbb{R}_+^m} \left( \left\| h(x)^{1/2} \odot \nabla_x \log p(x) \right\|^2 + 2 \cdot \text{div}(h(x) \odot \nabla_x \log p(x)) \right) q(x) dx$$

**Empirical Approximation:** Given samples  $x^{(1)}, \dots, x^{(n)} \sim q(x)$ , we approximate  $J_h(p)$  by:

$$\hat{J}_h(p) = \frac{1}{n} \sum_{i=1}^n \left( \left\| h(x^{(i)})^{1/2} \odot \nabla_x \log p(x^{(i)}) \right\|^2 + 2 \cdot \text{div}(h(x^{(i)}) \odot \nabla_x \log p(x^{(i)})) \right)$$



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# Proposed Method: Score Matching on Domain $M$

We propose a modified score matching objective over a domain  $M \subset \mathbb{R}^m$  with boundary  $\partial M$ .

## Proposed objective

$$J_{\text{prop}}(p) = \int_M \|\nabla \log p(x)\|^2 q(x) dx + 2 \int_M \Delta \log p(x) \cdot q(x) dx + 2B(p, q)$$

where  $B(p, q)$  denotes the **boundary term** arising from applying Stokes' theorem to the cross term.

**Key difference from conventional score matching:** The boundary term  $B(p, q)$  is preserved rather than discarded.

# Boundary Term in the Proposed Objective

The boundary term  $B(p, q)$  has the following explicit form:

$$B(p, q) = - \int_{\partial M} (\nabla \log p(x) \cdot \mathbf{n}(x)) q(x) dS(x)$$

## Explanation:

- $\mathbf{n}(x)$ : outward unit normal vector on the boundary  $\partial M$ ,
- $dS(x)$ : surface measure on  $\partial M$ ,

**Interpretation:** It penalizes mismatch between the model flow and the data density near the boundary.

## Case $M = \mathbb{R}_+^m$

For simplicity, let us assume:

$$M = \mathbb{R}_+^m = \{(x_1, \dots, x_m) \mid x_1, \dots, x_m \geq 0\}$$

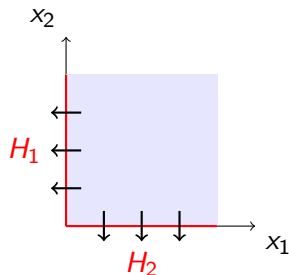
$$H_i = \{(x_1, \dots, x_m) \in \mathbb{R}_+^m \mid x_i = 0\}$$

Then, the each component is

$$\partial M = \sum_{i=1}^m H_i$$

$$\mathbf{n}(x) = -e_i \quad (x \in H_i)$$

$$dS(x) = dx_{-i} = dx_1 \dots dx_{i-1} dx_{i+1} \dots dx_m \quad (x \in H_i)$$



# Transforming $B(p, q)$ and Preparing for Empirical Estimation

We start from the boundary term:

$$B(p, q) = - \int_{\partial M} (\nabla \log p(x) \cdot \mathbf{n}(x)) q(x) dS(x)$$

In the case  $M = \mathbb{R}_+^m$ , this becomes:

$$= \sum_{i=1}^m \int_{H_i} \frac{\partial}{\partial x_i} \log p(x) \cdot q(x_i = 0, x_{-i}) dx_{-i}$$

To make empirical estimation feasible, we rewrite the joint density using the identity:

$$q(x_i = 0, x_{-i}) = q(x_i = 0) \cdot q(x_{-i} \mid x_i = 0)$$

Hence:

$$= \sum_{i=1}^m q(x_i = 0) \int_{H_i} \frac{\partial}{\partial x_i} \log p(x) \cdot q(x_{-i} \mid x_i = 0) dx_{-i}$$

# Empirical Approximation of the Boundary Term

We consider the following form of the boundary term:

$$B(p, q) = \sum_{i=1}^m q(x_i = 0) \cdot \mathbb{E}_{x_{-i} \sim q(x_{-i} | x_i = 0)} \left[ \frac{\partial}{\partial x_i} \log p(x_i = 0, x_{-i}) \right]$$

Since exact sampling from  $x_i = 0$  is infeasible, we approximate using small  $\epsilon > 0$ :

**Density at the boundary:**

$$q(x_i = 0) = \lim_{\epsilon \rightarrow 0} \frac{\mathbb{P}(0 \leq x_i < \epsilon)}{|\{0 \leq x_i < \epsilon\}|} \approx \frac{n_i^\epsilon}{n\epsilon} \quad \left( n_i^\epsilon = \sum_{j=1}^n \mathbf{1}_{\{x|0 \leq x_i < \epsilon\}}(x^{(j)}) \right)$$

**Conditional expectation:**

$$\mathbb{E}_{x_{-i} \sim q(x_{-i} | x_i = 0)} \left[ \frac{\partial}{\partial x_i} \log p(x) \right] \approx \frac{1}{n_i^\epsilon} \sum_{j=1}^n \frac{\partial}{\partial x_i} \log p(x^{(j)}) \mathbf{1}_{\{x|0 \leq x_i < \epsilon\}}(x^{(j)})$$

# Final Empirical Objective with Boundary Term

**Final approximation of  $B(p, q)$ :**

$$\hat{B}(p, q) = \frac{1}{n\epsilon} \sum_{i=1}^m \sum_{j=1}^n \frac{\partial}{\partial x_i} \log p(x^{(j)}) \cdot \mathbf{1}_{\{0 \leq x_i^{(j)} < \epsilon\}}(x^{(j)})$$

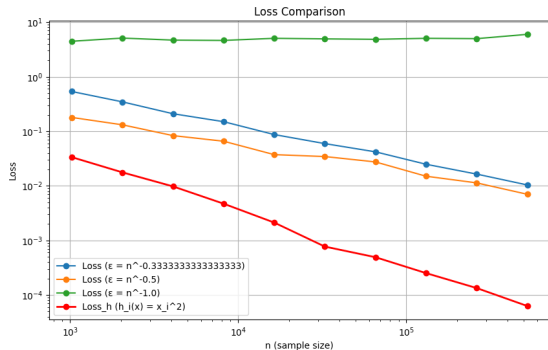
**Thus, the full empirical objective becomes:**

$$\hat{J}(p) = \frac{1}{n} \sum_{j=1}^n \sum_{i=1}^m \left( \frac{\partial}{\partial x_i} \log p(x^{(j)}) \right)^2 + 2 \cdot \frac{\partial^2}{\partial x_i^2} \log p(x^{(j)}) + \frac{2}{\epsilon} \cdot \frac{\partial}{\partial x_i} \log p(x^{(j)}) \cdot \mathbf{1}_{\{0 \leq x_i^{(j)} < \epsilon\}}(x^{(j)})$$

**where**  $x^{(j)} \sim q(x)$ , and  $\epsilon > 0$  controls boundary proximity.

# Experimental Comparison: Weighted vs. Proposed Score Matching

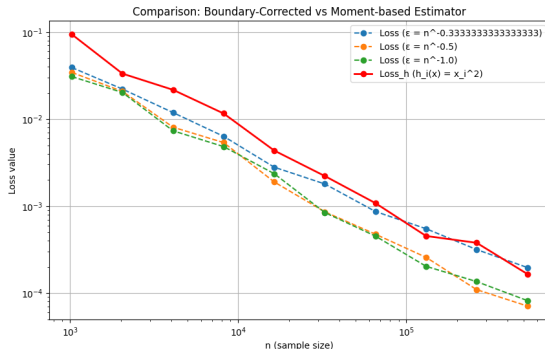
## Exponential Distribution



### Estimated Parameters:

- Weighted:  $\hat{\theta}_w = 2\bar{x}/\bar{x}^2$
- Proposed:  $\hat{\theta}_p = |H_i^\epsilon|/n\epsilon$

## Truncated Normal Distribution



### Estimated Parameters:

- Weighted:  $\hat{\theta}_{iw} = \bar{x}_i^4 / 3\bar{x}_i^2$
- Proposed:  $\hat{\theta}_p = n\epsilon \bar{x}_i^2 / (\sum x_i^\epsilon + n\epsilon)$



 Yu, S., Drton, M., & Shojaie, A. (2021).  
Generalized Score Matching for General Domains.

*Information and Inference: A Journal of the IMA*, 11(2), 739–780.

<https://doi.org/10.1093/imaiai/iaaa041>