

NNTI: Exercise Sheet 3

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Problem 3.1 - Linear Regression

(a)

(b)

(c)

The Mean Squared Error (MSE) equation is given by:

$$\text{MSE}_{\text{train}} = \frac{1}{m} \sum_{i=1}^m (\hat{y}_{\text{train}}^{(i)} - y_{\text{train}}^{(i)})^2$$

for a linear regression model $Y = wX + b$, where $\hat{y}_{\text{train}}^{(i)}$ is the predicted value, $y_{\text{train}}^{(i)}$ is the actual value, and m is the number of training examples.

To minimize this equation, we need to find the value of w where the gradient of MSE with respect to w is equal to zero. For that, we calculate the derivative of MSE with respect to w and set it to zero:

$$\frac{\partial \text{MSE}_{\text{train}}}{\partial w} = 0 \rightarrow \frac{\partial}{\partial w} \left[\frac{1}{m} \sum_{i=1}^m (\hat{y}_{\text{train}}^{(i)} - y_{\text{train}}^{(i)})^2 \right] = 0 \quad (1)$$

From the linear regression model equation, we know that $\hat{y}_{\text{train}}^{(i)} = wX^{(i)} + b$. We can plug this into (2):

$$\frac{\partial}{\partial w} \left[\frac{1}{m} \sum_{i=1}^m (wX^{(i)} + b - y_{\text{train}}^{(i)})^2 \right] = \frac{2}{m} \sum_{i=1}^m (\hat{y}_{\text{train}}^{(i)} - y_{\text{train}}^{(i)})X^{(i)} \quad (2)$$

3. **Setting the Derivative to Zero:**

$$\frac{2}{m} \sum_{i=1}^m (\hat{y}_{\text{train}}^{(i)} - y_{\text{train}}^{(i)})X^{(i)} = 0$$

4. **Simplify the Equation:**

$$\sum_{i=1}^m (\hat{y}_{\text{train}}^{(i)} - y_{\text{train}}^{(i)})X^{(i)} = 0$$

5. **Expand the Summation:**

$$\sum_{i=1}^m \hat{y}_{\text{train}}^{(i)}X^{(i)} - \sum_{i=1}^m y_{\text{train}}^{(i)}X^{(i)} = 0$$

6. **Move the Second Term to the Other Side:**

$$\sum_{i=1}^m \hat{y}_{\text{train}}^{(i)}X^{(i)} = \sum_{i=1}^m y_{\text{train}}^{(i)}X^{(i)}$$

7. **Substitute $\hat{y}_{train}^{(i)} = wX^{(i)} + b$:**

$$\sum_{i=1}^m (wX^{(i)} + b)X^{(i)} = \sum_{i=1}^m y_{train}^{(i)} X^{(i)}$$

8. **Expand the First Term:**

$$\sum_{i=1}^m w(X^{(i)})^2 + bX^{(i)} = \sum_{i=1}^m y_{train}^{(i)} X^{(i)}$$

9. **Move the Second Term to the Other Side:**

$$\sum_{i=1}^m w(X^{(i)})^2 = \sum_{i=1}^m y_{train}^{(i)} X^{(i)} - bX^{(i)}$$

10. **Divide by the Sum of Squares of $X^{(i)}$:**

$$w = \frac{\sum_{i=1}^m y_{train}^{(i)} X^{(i)} - bX^{(i)}}{\sum_{i=1}^m (X^{(i)})^2}$$

This is the expression for w that minimizes the MSE.

Problem 3.2 - PCA as Autoencoder

Problem 3.3 - PCA

See attached .ipynb solution in .zip file.

Problem 3.4

Give an appropriate positive constant c such that $f(n) \leq c \cdot g(n)$ for all $n > 1$.

1. $f(n) = n^2 + n + 1$, $g(n) = 2n^3$
2. $f(n) = n\sqrt{n} + n^2$, $g(n) = n^2$
3. $f(n) = n^2 - n + 1$, $g(n) = n^2/2$

Solution

We solve each solution algebraically to determine a possible constant c .

Part One

$$\begin{aligned} n^2 + n + 1 &= \\ &\leq n^2 + n^2 + n^2 \\ &= 3n^2 \\ &\leq c \cdot 2n^3 \end{aligned}$$

Thus a valid c could be when $c = 2$.

Part Two

$$\begin{aligned}n^2 + n\sqrt{n} &= \\&= n^2 + n^{3/2} \\&\leq n^2 + n^{4/2} \\&= n^2 + n^2 \\&= 2n^2 \\&\leq c \cdot n^2\end{aligned}$$

Thus a valid c is $c = 2$.

Part Three

$$\begin{aligned}n^2 - n + 1 &= \\&\leq n^2 \\&\leq c \cdot n^2/2\end{aligned}$$

Thus a valid c is $c = 2$.

Problem 3.5

Let $\Sigma = \{0, 1\}$. Construct a DFA A that recognizes the language that consists of all binary numbers that can be divided by 5.

Let the state q_k indicate the remainder of k divided by 5. For example, the remainder of 2 would correlate to state q_2 because $7 \bmod 5 = 2$.

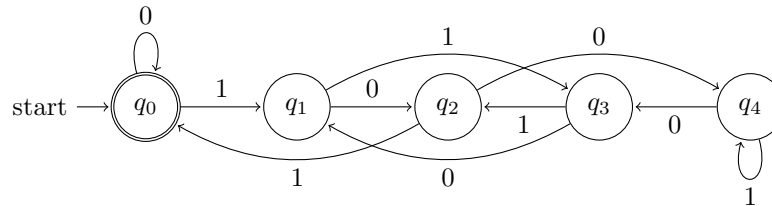


Figure 1: DFA, A , this is really beautiful, ya know?

Justification

Take a given binary number, x . Since there are only two inputs to our state machine, x can either become $x0$ or $x1$. When a 0 comes into the state machine, it is the same as taking the binary number and multiplying it by two. When a 1 comes into the machine, it is the same as multiplying by two and adding one.

Using this knowledge, we can construct a transition table that tell us where to go:

	$x \bmod 5 = 0$	$x \bmod 5 = 1$	$x \bmod 5 = 2$	$x \bmod 5 = 3$	$x \bmod 5 = 4$
$x0$	0	2	4	1	3
$x1$	1	3	0	2	4

Therefore on state q_0 or ($x \bmod 5 = 0$), a transition line should go to state q_0 for the input 0 and a line should go to state q_1 for input 1. Continuing this gives us the Figure 1.

Problem 3.6

Write part of **Quick-Sort**($list, start, end$)

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1: function QUICK-SORT( $list, start, end$ )
2:   if  $start \geq end$  then
3:     return
4:   end if
5:    $mid \leftarrow$  PARTITION( $list, start, end$ )
6:   QUICK-SORT( $list, start, mid - 1$ )
7:   QUICK-SORT( $list, mid + 1, end$ )
8: end function

```

Algorithm 1: Start of QuickSort

Problem 3.7

Suppose we would like to fit a straight line through the origin, i.e., $Y_i = \beta_1 x_i + e_i$ with $i = 1, \dots, n$, $E[e_i] = 0$, and $\text{Var}[e_i] = \sigma_e^2$ and $\text{Cov}[e_i, e_j] = 0, \forall i \neq j$.

Part A

Find the least squares estimator for $\hat{\beta}_1$ for the slope β_1 .

Solution

To find the least squares estimator, we should minimize our Residual Sum of Squares, RSS:

$$\begin{aligned} RSS &= \sum_{i=1}^n (Y_i - \hat{Y}_i)^2 \\ &= \sum_{i=1}^n (Y_i - \hat{\beta}_1 x_i)^2 \end{aligned}$$

By taking the partial derivative in respect to $\hat{\beta}_1$, we get:

$$\frac{\partial}{\partial \hat{\beta}_1} (RSS) = -2 \sum_{i=1}^n x_i (Y_i - \hat{\beta}_1 x_i) = 0$$

This gives us:

$$\begin{aligned} \sum_{i=1}^n x_i (Y_i - \hat{\beta}_1 x_i) &= \sum_{i=1}^n x_i Y_i - \sum_{i=1}^n \hat{\beta}_1 x_i^2 \\ &= \sum_{i=1}^n x_i Y_i - \hat{\beta}_1 \sum_{i=1}^n x_i^2 \end{aligned}$$

Solving for $\hat{\beta}_1$ gives the final estimator for β_1 :

$$\hat{\beta}_1 = \frac{\sum x_i Y_i}{\sum x_i^2}$$

Part B

Calculate the bias and the variance for the estimated slope $\hat{\beta}_1$.

Solution

For the bias, we need to calculate the expected value $E[\hat{\beta}_1]$:

$$\begin{aligned} E[\hat{\beta}_1] &= E\left[\frac{\sum x_i Y_i}{\sum x_i^2}\right] \\ &= \frac{\sum x_i E[Y_i]}{\sum x_i^2} \\ &= \frac{\sum x_i (\beta_1 x_i)}{\sum x_i^2} \\ &= \frac{\sum x_i^2 \beta_1}{\sum x_i^2} \\ &= \beta_1 \frac{\sum x_i^2 \beta_1}{\sum x_i^2} \\ &= \beta_1 \end{aligned}$$

Thus since our estimator's expected value is β_1 , we can conclude that the bias of our estimator is 0.

For the variance:

$$\begin{aligned} \text{Var}[\hat{\beta}_1] &= \text{Var}\left[\frac{\sum x_i Y_i}{\sum x_i^2}\right] \\ &= \frac{\sum x_i^2}{\sum x_i^2 \sum x_i^2} \text{Var}[Y_i] \\ &= \frac{\sum x_i^2}{\sum x_i^2 \sum x_i^2} \text{Var}[Y_i] \\ &= \frac{1}{\sum x_i^2} \text{Var}[Y_i] \\ &= \frac{1}{\sum x_i^2} \sigma^2 \\ &= \frac{\sigma^2}{\sum x_i^2} \end{aligned}$$

Problem 3.8

Prove a polynomial of degree k , $a_k n^k + a_{k-1} n^{k-1} + \dots + a_1 n^1 + a_0 n^0$ is a member of $\Theta(n^k)$ where $a_k \dots a_0$ are nonnegative constants.

Proof. To prove that $a_k n^k + a_{k-1} n^{k-1} + \dots + a_1 n^1 + a_0 n^0$, we must show the following:

$$\exists c_1 \exists c_2 \forall n \geq n_0, c_1 \cdot g(n) \leq f(n) \leq c_2 \cdot g(n)$$

For the first inequality, it is easy to see that it holds because no matter what the constants are, $n^k \leq a_k n^k + a_{k-1} n^{k-1} + \dots + a_1 n^1 + a_0 n^0$ even if $c_1 = 1$ and $n_0 = 1$. This is because $n^k \leq c_1 \cdot a_k n^k$ for any nonnegative constant, c_1 and a_k .

Taking the second inequality, we prove it in the following way. By summation, $\sum_{i=0}^k a_i$ will give us a new constant, A . By taking this value of A , we can then do the following:

$$\begin{aligned} a_k n^k + a_{k-1} n^{k-1} + \dots + a_1 n^1 + a_0 n^0 &= \\ &\leq (a_k + a_{k-1} \dots a_1 + a_0) \cdot n^k \\ &= A \cdot n^k \\ &\leq c_2 \cdot n^k \end{aligned}$$

where $n_0 = 1$ and $c_2 = A$. c_2 is just a constant. Thus the proof is complete. \square

Problem 3.18

Evaluate $\sum_{k=1}^5 k^2$ and $\sum_{k=1}^5 (k-1)^2$.

Problem 3.19

Find the derivative of $f(x) = x^4 + 3x^2 - 2$

Problem 3.6

Evaluate the integrals $\int_0^1 (1-x^2)dx$ and $\int_1^\infty \frac{1}{x^2}dx$.