

Financial Costs of Children, Education Subsidies, and Fertility Choices in Equilibrium*

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Abstract

Macroeconomic concerns associated with low fertility have driven many countries to implement pro-natal policies. This paper examines an unconventional pro-natal policy subsidizing college students implemented recently in Japan, a leading country in this demographic issue. I first document facts suggesting that the financial costs of college education pose a significant impediment for couples in achieving their desired birth numbers in Japan. I then construct an equilibrium lifecycle model with college enrollment, fertility, and asset transfer choices. Japan's means-tested grants for college students, introduced in 2020 as a pro-natal policy, would significantly increase skilled parents' fertility by providing insurance against multiple sources of uncertainty. It would also increase the per-capita output, education mobility, and college enrollment rate, amplified through fertility margins in the long run. However, its expansion by setting a higher income threshold has limited effects on average fertility because it lowers unskilled parents' fertility by raising their expected costs of children due to the greater intergenerational mobility of education.

Keywords: Fertility, education, intergenerational linkages, pro-natal policy

JEL codes: C68, I28, J13, J24.

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1 Introduction

The total fertility rates (TFR) are below the replacement level in many developed countries. Low fertility is a root cause of demographic aging and raises macroeconomic concerns, such as labor force shortage, the sustainability of the current social security schemes, and the associated significant tax burdens on households.¹ In response, those countries have implemented various pro-natal policies that aim to remove obstacles to fertility and help people bear and raise children. A widely used instrument is unconditional cash transfers, such as baby bonuses and child benefits, and empirical studies show that they have significant effects on fertility by reducing child-rearing costs.²

In this context, the education subsidies for college students have garnered significant attention in Japan, a leading country in this demographic issue.³ Although it is unconventional as a pro-natal policy, the government introduced grants for college students in 2020 based on a pro-natal motive.⁴ One of the key underlying premises is that, as I discuss in Section 3, the financial costs arising from children’s college enrollment are a critical obstacle to fertility in Japan, a country subsidizing least in higher education among OECD countries.⁵ Another important premise is that the policy also increases the aggregate human capital or workers’ labor productivity by promoting skill acquisition, a desirable property as a countermeasure against demographic change. This is a stark difference from typical pro-natal transfers such as baby bonuses, as previous studies show that they would lower aggregate human capital because they make parents shift from the “quality” toward the “quantity” of children (e.g., Zhou, 2022; Kim et al., 2023). If this policy works as expected, it will be a powerful tool to mitigate the macroeconomic concerns of low fertility, which should be informative for countries with these concerns.

However, the consequences of the policy are not apparent for several reasons. First, the general equilibrium (GE) feedback, primarily through changes in the skill premium and tax rate, will affect children’s college enrollment decisions and parental choices on

¹Bloom et al. (2023) discuss the macroeconomic concerns of low fertility in high-income countries, including those related to the labor force shortage and social security systems.

²For example, see Milligan (2005); Cohen et al. (2013); González (2013); Malkova (2018).

³Japan’s current fertility rate has been around 1.3, far below the population replacement level. In addition, the old-age dependency ratio is more than 50% and is projected to reach 80% in 2050, which is by far the highest among the OECD countries (See, <https://data.oecd.org/>).

⁴They stated its pro-natal motive in an act enacted for the policy: “*The aim is ... to foster an environment where people can bear and raise their children with a sense of ease by alleviating the economic burden associated with higher education, thereby contributing to addressing the rapid decline in the birthrate in our country.*” (Act on Support for Higher Education Studies, enacted on May 17, 2019). See, https://www.mext.go.jp/a_menu/koutou/hutankeigen/detail/__icsFiles/afieldfile/2019/05/17/1417025_02_1.pdf (available only in Japanese).

⁵For more details, see i) in Appendix A.

the “quantity-quality.” Second, the policy will also affect the household distribution (e.g., skill, human capital, and age) in the long run, which matters to the variables of our interest via composition effects. Who adjusts fertility upon the policy change also matters to those effects, considering the intergenerational transmission of individual characteristics (e.g., human capital). Examining the consequences of the education subsidy thus requires the GE framework that captures those indirect effects of the policy or equilibrium feedback.

In this study, I first document facts suggesting that the financial costs for parents to support their children’s college enrollment are a significant impediment to fertility decisions in Japan: (1) couples are most likely to abandon having an ideal number of children because of financial costs; (2) a significant financial cost gap exists between those who have children enrolled in college and those who do not; and (3) a substantial fraction of parents desire a college education for their children.

I then construct an incomplete market GE overlapping generations (OLG) model that incorporates choices regarding college enrollment, fertility, and inter-vivo transfers (IVT). In this model, agents choose whether to attend college after graduating high school. Four factors affect their enrollment choices: (1) IVT made by their parents, (2) human capital, (3) psychic costs of education, and (4) eligibility status for government-provided student loans and grants, determined by parent’s income and student’s human capital. Agents draw (2) and (3) just before the education choices from distributions that depend on their own and parent’s characteristics (e.g., a child of parents with higher human capital tends to draw higher human capital, and a child of college graduates tends to draw a smaller psychic cost of education). At a certain point in life, they make fertility choices and choose how many children they have. They draw utility from the number of children, education spending until their children graduate high school, and IVT made for them upon their high school graduation based on altruism. They also make labor supply decisions until retirement age, and their labor productivity is subject to idiosyncratic risks.

The model is calibrated to the Japanese economy using the Japanese Panel Survey of Consumers (JPSC). I use the 60s cohort of married females and their family members, and the model replicates key moments such as the intergenerational persistence of education and the scale of parental asset transfers for college students. On average, sending a child to college requires parents to make the IVT amounting to more than 60% of the total education expenditures made until the child’s high school graduation. I then check the model’s validity and show that it replicates some important non-targeted moments, such as the benefit elasticity of fertility and fertility differentials across education levels. The benchmark model captures government-provided student loans but does not include the existing grants introduced in 2020. The existing grants are income-tested, and approximately the bottom 15% of the income distribution of households are eligible. The

payments cover approximately two-thirds of the students' average expenses.

The main findings are summarized as follows. First, introducing the existing income-tested grants into the benchmark model leads to a 3% increase in the TFR in the long run, largely driven by the fertility increase of college graduate (“skilled”) parents. Notably, the fertility increase of ex-post ineligible (relatively wealthier) parents explains the higher fertility rate. A decomposition analysis suggests that the education subsidy provides partial insurance against uncertainty regarding income and education expenditure in IVT, which makes parents comfortable having another child and increases fertility.

To understand this argument, note whether they are beneficiaries of the grants is uncertain when the parents make fertility decisions; whether their (potential) children attend college in the future is uncertain, and the eligibility status is also uncertain because it depends on their income when their children attend college. Sources of uncertainty are income risks and children’s characteristics such as human capital. The latter is translated into education expenditure in IVT because it simultaneously determines children’s enrollment and parents’ IVT choices. First, income volatility makes parents hesitate to have a child given that having a child can be considered making “consumption commitments” in that they are irreversible and force parents to make non-negligible spending over long periods (Santos and Weiss, 2016; Sommer, 2016). Further, the intergenerational persistence of education implies that college graduates expect that they will make sizable IVT for their children to support their college enrollment when they make fertility choices. However, negative income shocks may make it infeasible for them to make sufficient IVT to send their children to college, or they may substantially reduce their consumption in exchange for children’s college enrollment. Both situations are costly for parents; thus, income volatility with higher expected IVT makes college graduates tend to have fewer children. In this situation, the grants can provide partial insurance against such an unlucky state — although their children need financial support to attend college, they are poor due to the realization of negative income shocks —. As a result, the introduction of grants can increase the fertility of college graduates, including ex-post ineligible households.

Second, introducing the existing program would increase the college enrollment rate by four percentage points (p.p.) in the long run. Education mobility also increases in the sense that children of high school graduate (“unskilled”) parents are more likely to attend college by 2.5 p.p. First, the subsidy relaxes the financial constraints on students’ enrollment decisions and enables some students to attend college, which operates immediately after the policy introduction. In the long run, the skill distribution (i.e., the share of college graduates) changes among people at the age of making fertility choices. Because children of college graduates are more likely to attend college, the change in skill distribution of parent generations also contributes to the higher college enrollment rate

for their children’s generation. I also find that the fertility margins amplify the policy effects on the college enrollment rate; college graduates’ fertility increases more, and those “marginal” children born are likely to attend college due to the intergenerational persistence of education, leading to a greater share of college graduates in the long run. Higher TFR and college enrollment rates lead to a greater per-capita labor supply in efficiency units, increasing the equilibrium per-capita output by 0.7%.

I also simulate the program’s expansion by subsidizing students in broader income classes, which is actively discussed in Japan. The long-run college enrollment rate, education mobility, and per-capita output increase as the income threshold increases. However, the TFR is almost stable despite expansions because unskilled parents’ fertility decreases in expansion, which cancels out the fertility increase of skilled parents. The decomposition analysis suggests that the greater education mobility facilitated by the expansion increases the expected costs of having children for unskilled parents and lowers their fertility. In the benchmark without grants, children of unskilled parents are unlikely to attend college. Therefore, they will likely not have to make the significant IVT required to support their children’s college enrollment. If grants are introduced and their coverage expands, children of unskilled parents are more likely to attend college than in the benchmark. Given that the grants are not generous enough to cover 100% of the costs of sending their children to college, this higher probability of children going to college implies that those parents are likely to make more significant transfers upon their children’s college enrollment, making them have fewer children due to the higher expected costs.

Lastly, I examine the performance of the potential alternative programs to the existing income-tested grants: (1) grants for college students with ability tests in addition to income tests, (2) unconditional grants for college students regardless of ability and income, and (3) unconditional cash transfers to households with children until their high school graduation. I solve the stationary equilibrium with each program in an expenditure-neutral way and find that the existing scheme would lead to the highest equilibrium TFR among other scenarios. The previous decomposition analysis suggests that targeting students in unlucky households in which negative income shocks are realized would be effective, at least marginally, in increasing the TFR.

The rest of the paper is organized as follows. Section 2 reviews several strands of literature related to this study. Section 3 presents a set of facts suggesting that the financial costs of college education are a significant impediment to fertility decisions in Japan. Section 4 describes the model, which is calibrated in Section 5. Section 6 conducts numerical analysis and Section 7 concludes.

2 Related Literature

This study relates to several strands of literature. First, it is closely related to the literature on fertility choices in incomplete market models.⁶ Some studies point out that having a child can be considered making “consumption commitments”; they are irreversible choices and require at least a certain amount of resources over long periods (Santos and Weiss, 2016; Sommer, 2016). If the negative income shocks are realized when raising a child, they have to reduce their consumption substantially compared with a case of not having a child, which is costly in utility terms. Income volatility thus makes households hesitate to have children.⁷ This study builds on the lifecycle model with fertility choices and uninsurable income risks as in the literature and examines a set of policies that can potentially reduce the “risks” of having a child. I highlight the insurance effects of the income-tested grants on fertility choices, especially for college graduate parents whose children are likely to attend college; whether a child attends college or not makes a significant difference in financial burdens for parents, which is an additional “risk” of having a child. I also highlight the interaction between the subsidy and parents’ expectations for children’s education decisions. The latter is translated into the expected costs of having a child and thus affects fertility choices. Incorporating college enrollment choices and intergenerational linkages that generate the intergenerational persistence of education, which differs from the literature, enables us to obtain those new insights.

Second, this study is closely related to the literature on family-related policies using lifecycle models with fertility choices.⁸ Previous studies examine the effects of childcare subsidization (e.g., Bick, 2016), cash transfers (e.g., Kim et al., 2023; Nakakuni, 2023), both of them (e.g., Zhou, 2022), parental leave policies (e.g., Erosa et al., 2010; Kim and Yum, 2023), and tax reform (e.g., Jakobsen et al., 2022). In this context, education subsidies for college students can be classified as cash transfers targeting a specific type of household. Compared with typical pro-natal transfers such as baby bonuses, a notable feature of the college subsidy is that it is also expected to increase the average human capital by promoting skill acquisition. Previous studies show that pro-natal transfers would lower aggregate human capital because they make parents shift from the “quality” toward “quantity” of children (e.g., Zhou, 2022; Kim et al., 2023). If it works as expected, the education subsidy would be an effective tool to deal with the macroeconomic concerns

⁶For example, Schoonbroodt and Tertilt (2014); Santos and Weiss (2016); Sommer (2016).

⁷This is the case especially in the early stages of their lives, thus delaying marriage and fertility (Santos and Weiss, 2016). Delaying fertility leads to low fertility, given that the ability to reproduce declines with age (Sommer, 2016).

⁸For example, Erosa et al. (2010); Bick (2016); Yamaguchi (2019); Cavalcanti et al. (2021); Hagiwara (2021); Jakobsen et al. (2022); Zhou (2022); Nakakuni (2023).

of low fertility by increasing both the “quantity” and “quality” of the labor force in the economy, making it worth investigating. This study should provide insights into countries considering countermeasures against the demographic issue, not just Japan. It also contributes to the literature by studying the effects of education subsidies for college students on fertility, education level, and aggregate output. In doing so, I incorporate college enrollment choices into the model, which is essential for studying college subsidies.

In this literature, the closest study is [Zhou \(2022\)](#), which builds an OLG model capturing the quantity-quality trade-off and endogenous human capital accumulation for children to study the effects of family policies such as the baby bonus, childcare subsidies, as well as public education subsidies. Unlike that study, I focus on the effects of education subsidies for college students, and thus, the model is tailored to the purpose by incorporating college enrollment choices. Instead, my model exogenizes the parental time investment and the accumulation of children’s human capital before college education, which are critical in considering a broader set of family policies as [Zhou \(2022\)](#) does.

Third, this study relates to macroeconomic studies based on the quantity-quality trade-off framework⁹ and contributes to the literature by examining the effects of college subsidies building on the framework. [Kim et al. \(2023\)](#) build an OLG model with the quantity-quality trade-off to examine the role of the status externalities in education on the low fertility in South Korea. My model does not capture status externalities but is tailored to study the effects of education subsidies for college students by building a lifecycle model with college enrollment choices. From the modeling viewpoint, the closest work is [Daruich and Kozłowski \(2020\)](#). They construct a partial equilibrium lifecycle model with college enrollment, fertility, and IVT choices to investigate the roles of fertility choices and family transfers in explaining intergenerational mobility in the US. This study differs from theirs in research questions and the associated modeling strategy, examining the macroeconomic effects of education subsidies adopting the GE framework.¹⁰ Another related paper is [De la Croix and Doepke \(2004\)](#), which constructs a two-period OLG model with fertility choices to compare the macroeconomic implications of a public and private schooling regime. My study differs from theirs in constructing a lifecycle model with college enrollment choices to examine college subsidies.

⁹See, for example, [De La Croix and Doepke \(2003\)](#); [Manuelli and Seshadri \(2009\)](#); [Cordoba et al. \(2019\)](#); [Daruich and Kozłowski \(2020\)](#); [Kim et al. \(2023\)](#); [Zhou \(2022\)](#).

¹⁰Previous studies show that the general equilibrium feedback of college subsidies, primarily through the change in skill premium, is critical in assessing their macroeconomic consequences. A higher college enrollment rate facilitated by college subsidies reduces the skill premium in the long run. The change in skill premium implies the change in income distribution (e.g., [Krueger and Ludwig, 2016](#)) and also affects the incentive of college enrollment (e.g., [Abbott et al., 2019](#)). My results presented in Section 6 also highlight the roles of GE effects in explaining the long-run effects of education subsidies on fertility.

Lastly, this study contributes to the literature on macroeconomic analysis of education subsidies, especially for college students.¹¹ Given our primary focus on the performance of the education subsidy as a pro-natal policy, I incorporate fertility choices into the standard framework adopted in the literature — the incomplete market GE-OLG model with education and IVT choices —. This extended framework sheds light on the roles of fertility margins in understanding the macroeconomic consequences of the policy. I show that fertility responses amplify the policy impacts on the college enrollment rate.

3 Background: Financial Costs of Children and Fertility Choices in Japan

This section presents facts suggesting that the financial costs for parents to support their children’s college enrollment are a significant impediment to fertility decisions in Japan. This explains why education subsidies for college students are drawing attention in Japan as a pro-natal measure and motivates my quantitative analysis. First, I list the three facts below and then elaborate on each one by one:

1. Couples are most likely to abandon having an ideal number of children because of financial costs.
2. A significant financial cost gap exists between those who have children enrolled in college and those who do not.
3. A substantial fraction of parents desire a college education for their children.

Fact 1. *Couples are most likely to abandon having an ideal number of children because of financial costs.*

This fact is drawn from the National Fertility Survey (NFS), which is provided by the National Institute of Population and Social Security Research (IPSS).¹² The NFS is a cross-sectional household survey that asks respondents about their preferences or intentions regarding fertility, marriage, child-raising, and education, as well as their basic information, including their education, age, and income. It is conducted nearly every five years, and the latest survey available is in 2015, which collected 5,334 couples in which the wife is aged 18 to 49. Hereafter, I present the results focusing on married couples in

¹¹See, for example, Benabou (2002); Akyol and Athreya (2005); Kulikova (2015); Krueger and Ludwig (2016); Kotera and Seshadri (2017); Lawson (2017); Abbott et al. (2019); Lee and Seshadri (2019); Abbott (2022); Daruich (2022); Zheng and Graham (2022); Matsuda and Mazur (2022); Matsuda (2022).

¹²See, https://www.ipss.go.jp/site-ad/index_english/survey-e.asp.

which the wife is aged 25 to 39 years, leaving 2,420 couples.¹³ According to the NFS, a non-negligible gap exists between the ideal and planned numbers of children. In the 2015 survey, the ideal number of children for wives aged 25 to 39 was on average 2.38, whereas the planned number was 2.16. Fig 1 represents the distribution of the ideal and planned numbers of children, where blue (red) bars indicate the share of wives who desire (plan) to have each number of children from zero to more than five. This figure suggests that the gap originates from the downward revision of the ideal at the intensive margin. There is no significant share gap between those whose ideal number is zero and those whose planned number is zero, and a substantial fraction of wives who desire three children end up with one or two children.

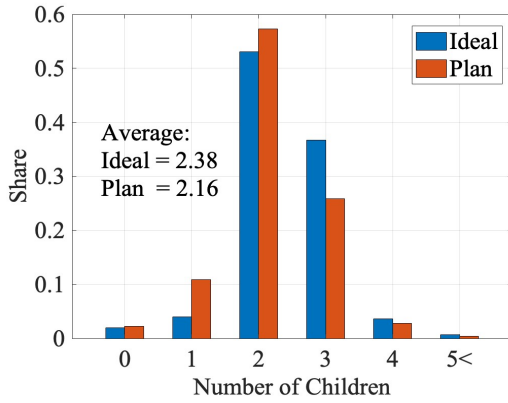


Fig 1: Distribution of ideal and planned numbers of children.

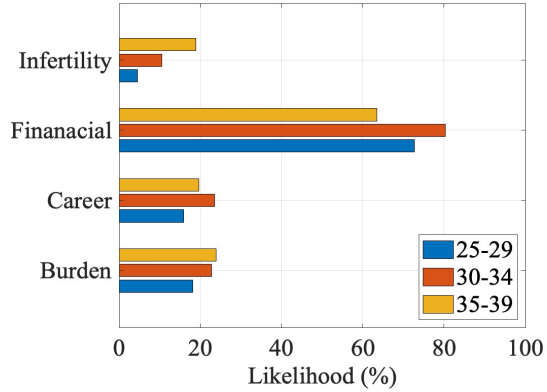


Fig 2: Reasons for the gap between the ideal and planned number of children.

Why do people abandon their ideal number of children? The NFS asks them to pick their reasons among several options, and the result indicates that they are most likely to choose the financial reason: “raising children and education are too expensive.” Fig 2 represents the likelihood¹⁴ that wives of each age group abandon their ideal number of children for a particular reason, and “Financial” represents the financial reason. Aside from this, “Career” represents “it would interfere with my job,” and “Burden” represents “I would not handle the psychological and physical burden” (arising from achieving the ideal number). On average, more than 75% of them chose the financial reason, which dominates other reasons, such as “Career” and “Burden,” chosen by only 20% of them.

¹³Its 53.1% of the sample consists of couples in which the wife is over age 40, so the sample size shrinks if we target the younger couples. I focus on wives under age 39 because here I am interested in the fertility intention of those in the stage of fertility decision. Note that the cohort fertility rate is stable after age 40 for any cohort in Japan. See, for example, p7 of <https://www.mhlw.go.jp/toukei/saikin/hw/jinkou/tokusyuu/syussyo07/d1/gaikyou.pdf> (in Japanese). Excluding those aged 18 to 24 does not affect the result significantly because they consist only of 1.5% of the observations for wives aged 18 to 49.

¹⁴I use the term “likelihood” instead of “share” because it allows respondents to choose multiple options.

To sum up, there is a non-negligible gap between the ideal and planned numbers of children for couples, and a substantial fraction of them answer the reason as “raising children and education are too expensive.” These results establish the first fact: couples are most likely to abandon having an ideal number of children because of financial costs.

Fact 2. *A significant financial cost gap exists between those who have children enrolled in college and those who do not.*

Fact 1 indicates that the financial cost of having children is a critical constraint on fertility choices in Japan. However, it is silent on how and in which cases having children is so expensive; Fact 2 addresses them from the viewpoint of education costs. To do this, I use two data sets: (1) the Survey of Children’s Learning Expenses (SCLE, 2021) and (2) the Student Life Survey (SLS, 2018), both cross-sectional household surveys and conducted by the Ministry of Education, Culture, Sports, Science and Technology (MEXT).¹⁵ The SCLE covers more than 53,000 students from preschool to high school and reports the per-student average education expenditure for each expenditure category and education stage (i.e., preschool, elementary, junior high, and high school). The expenditure category includes not only school-related ones (e.g., tuition fees and textbooks) but also extracurricular activities (e.g., cram school, music, arts, and sports). In addition to that, I use the SLS for parents’ expenditures when their children enroll in college.¹⁶

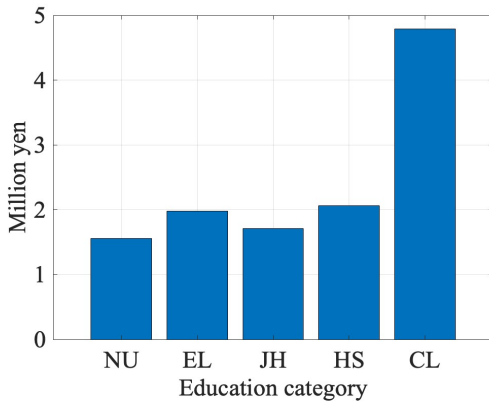


Fig 3: Average education expenditures.

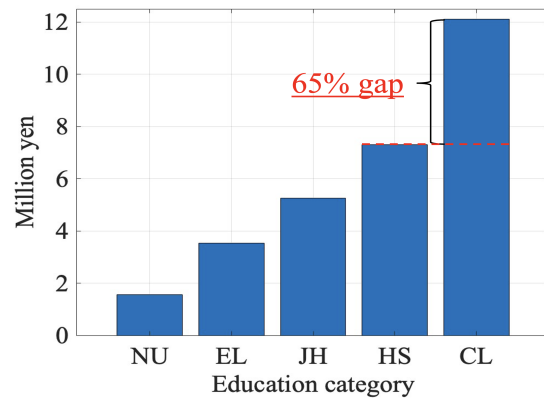


Fig 4: Cumulative education-expenditures.

Fig 3 shows the average per-child expenditure for each education category until completion. “NU,” “EL,” “JH,” “HS,” and “CL” each stand for nursery school or preschool,

¹⁵See, https://www.mext.go.jp/b_menu/toukei/chousa03/gakushuuhi/1268091.htm for the SCLE and https://www.jasso.go.jp/statistics/gakusei_chosa/_icsFiles/afieldfile/2021/03/09/data18_all.pdf (in Japanese) for the SLS.

¹⁶In 2018, the SLS correct answers from 43,394 students attending tertiary education, including college, some college, and graduate school.

elementary school, junior high school, high school, and college, only consisting of a four-year college. All expenditures are conditional on enrollment in the education stage. Fig 4 represents the cumulative education expenditures, showing that the expenditure jumps up when children attend college. Given that the high school graduation rate is approximately 100% in Japan, the figure tells us that raising one child on average takes the education costs of 7.31 million yen, which amounts to 4.5% of the individuals' average lifetime labor earnings.¹⁷ If their children attend college, they have to spend another 4.8 million yen, meaning there is more than a 60% increase in education costs if parents send their children to college. As reported in i) in Appendix A, this expensiveness of college education can, at least partly, be attributed to the fact that Japan is one of the least in subsidizing tertiary education while subsidizing a sizable portion of schooling costs up to secondary education. These observations establish Fact 2: A significant financial cost gap exists between those who have children enrolled in college and those who do not.

Fact 3. *A substantial fraction of parents desire a college education for their children.*

Facts 1 and 2 suggest that the financial costs for children enrolled in college are a critical obstacle to fertility in Japan. However, one might not still be convinced because Japan's college enrollment rate is approximately 55%, far below 100%; thus, college education costs seem relevant only to half of the population.¹⁸ The following fact answers “No, they should not.” to this argument by showing that far more fraction of parents desire a college education for their children than the current college enrollment rate.

We return to the NFS (2015), which asked respondents about the desired education level for their children. Fig 5 summarizes the results by wife's age. Here, “SC” stands for some college. The figure shows that approximately 75% of the parents desire a college education for their children, which is significantly higher than the college enrollment rate observed in any period in Japan. This observation suggests that, although college enrollment in Japan has been far below 100%, education costs for a college education can be relevant not only for a specific part of the population because many parents would like to send their children to college.

¹⁷I construct a proxy of the average individual's lifetime earnings based on the 2022 Basic Survey on Wage Structure (BSWS) by the Ministry of Health, Labour, and Welfare (MHLW). First, I compute the average monthly earnings of ordinary workers for each age unconditional on any other characteristics such as sex and education. Then, I sum up the average earnings for each age, which amounts to about 160 million yen, and regard it as a proxy of the individuals' average lifetime labor earnings.

¹⁸Dropout rates are insignificant in Japan, so the enrollment rate is almost equivalent to the graduation rate. For example, the dropout rate was 2.5% in 2021. See, https://www.mext.go.jp/content/20220603-mxt_kouhou01-000004520_03.pdf (in Japanese).

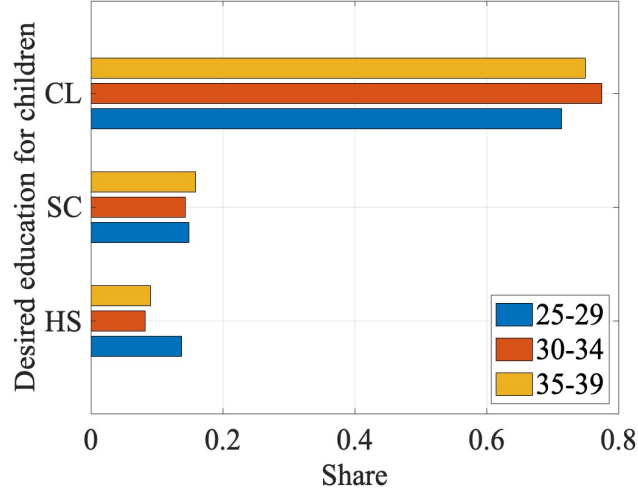


Fig 5: Wives' intention for children's education attainment.

4 Model

This section describes the incomplete market GE-OLG model incorporating choices on college enrollment, IVT, and fertility. I embed fertility choices into a framework otherwise standard in the macroeconomic literature of college subsidy (e.g., [Krueger and Ludwig, 2016](#); [Abbott et al., 2019](#); [Matsuda and Mazur, 2022](#)). Section 4.1 provides an overview of the lifecycle of this economy. Section 4.2 elaborates on preliminaries of the model and then Section 4.3 formulates households' decision problems. Section 4.4 discusses the stationary equilibrium of this economy.

4.1 Overview of the lifecycle

Fig 6 represents the households' lifecycle in this model. In this model, one period corresponds to two years. Letting j denote age, agents live with their parents until they graduate from high school at the beginning of age $j = J_E (= 18)$ and then choose whether to attend college or enter the labor market after graduating high school, represented as a node "Grad.HS" in the figure. If they do not enroll in college, they enter the labor market as a high school graduate. If they choose to attend college, it takes four years (two model-periods) to complete, and they enter the labor market as a college graduate after graduation, represented as a node "Grad.CL" in the figure.¹⁹

¹⁹This model does not consider the possibility of dropping out from college because, as I mentioned, the dropout rate is insignificant in Japan.

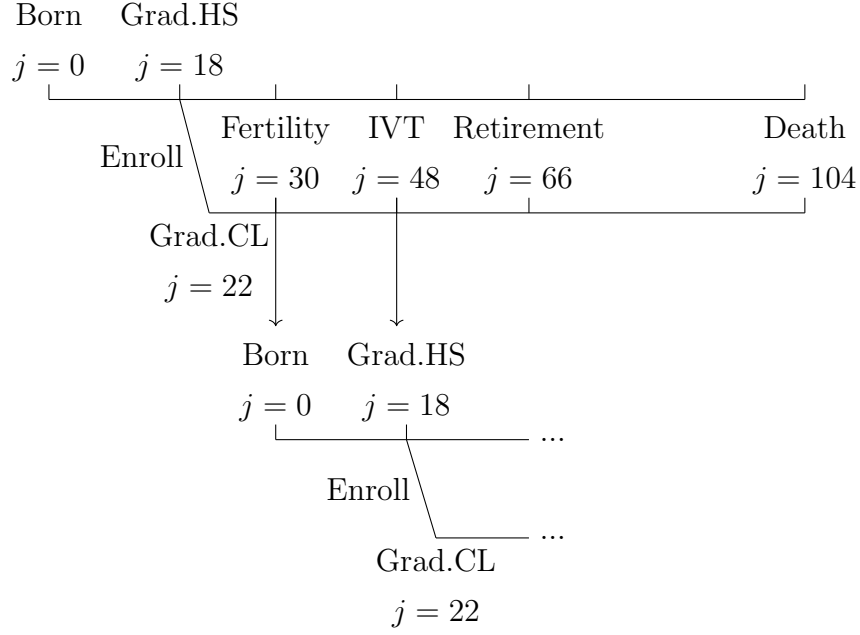


Fig 6: Model's lifecycle.

College enrollment increases their skill and potential lifetime earnings but takes some costs in terms of money and disutility. As standard in the literature, the model captures several channels of intergenerational linkage critical to the enrollment choices. First, the disutility of education (“psychic costs”) depends on their parents’ educational background. Second, parents choose the amount of IVT for their children. Third, the human capital endowment of children correlates to that of their parents, which is critical to education return. Students can fund their education costs through the IVT, labor earnings through part-time jobs, student loans with cheaper interest payments, and grants. The subsidized loans and grants can be income- or ability-tested and provided by the government.

After completing their education and entering the labor market, they choose how much to consume, save, and work, subject to idiosyncratic productivity shocks. At the beginning of age $j = J_F(= 30)$, they choose how many children they have, which is common for high school and college graduates. They draw utility from the number of children and education spending for each child, but having children requires some costs in terms of money and time. The lifecycle of a new cohort starts at this point, represented in the bottom half of the Fig 6. After their children graduate from high school, corresponding to the beginning of the age $j = J_{IVT}(= 48)$ for parents, they decide how much money to transfer to their children. This IVT decision affects the children’s college enrollment choice at the node “Grad.HS.” They are forced to retire at the beginning of age $j = J_R(= 66)$. After that period, they face mortality risks; every period, a certain fraction of them is hit by exogenous mortality shocks and exits from the economy. They can live for $J(= 104)$ years at the longest, but they exit the economy after age J .

4.2 Preliminaries

Production: A representative firm chooses labor and capital inputs in competitive factor markets to produce final goods. There are two types of labor inputs in this economy; the college graduates (skilled) and high school graduates (unskilled). Their total labor supply in efficiency units are represented as L_{CL} and L_{HS} , respectively. I allow them to be imperfect substitutes by considering the aggregate labor in efficiency units, L , is given as:

$$L = [\omega_{HS} \cdot (L_{HS})^\chi + \omega_{CL} \cdot (L_{CL})^\chi]^{1/\chi},$$

where $\omega_{HS} \equiv 1 - \omega_{CL}$ and ω_{CL} governs the relative productivity of the skilled workers.

The representative firm operates with a Cobb-Douglas production function with aggregate capital K and labor L to produce the output Y :

$$Y = ZK^\alpha L^{1-\alpha},$$

where Z represents the factor neutral productivity. Let r , w_{HS} , and w_{CL} denote the rental rate of capital and wage rates for unskilled and skilled labor. Capital depreciates at δ , and the firm has to incur the capital depreciation cost.

Demographics: As I explained above, after retirement, they face uncertainty regarding their survival in the next period. Let $\zeta_{j,j+1}$ denote the probability of surviving at age $j + 1$ conditional on surviving at age j for each $j \in \{J_R, \dots, J\}$ with $\zeta_{J,J+1} = 0$. In equilibrium, the cohort size (i.e., the number of births) grows at a constant rate g_n , which is endogenously determined in this model based on households' choices on fertility. Let μ_j denote the age distribution of the economy, determined by the cohort population growth rate and survival probabilities. The population mass of the economy is normalized to 1 (i.e., $\sum_j \mu_j = 1$) as in previous studies (e.g., [Guner et al., 2020](#)).²⁰

Intergenerational Linkages and Initial Endowments: After graduating high school (at the beginning of age 18), agents draw their human capital from a distribution $g_{h_p}^h$ varying with the parents' human capital level h_p . They also draw the psychic costs of education from a distribution g_{h,e_p}^ϕ , depending on the student's human capital and parent's education e_p . Lastly, they may receive some assets (IVT) from their parents. These three sources of intergenerational linkages shape their college enrollment choices and generate their heterogeneity.

²⁰Thus, the fertility rate matters to the age distribution, not population mass.

Preferences: Throughout their lifetime, they draw utility from consumption c and leisure l according to a utility function $u(c, l)$. They discount future utility by β . At age $j = J_F$, they choose the number of children they have, denoted by $n \in \{0, 1, \dots, N\}$. They can draw additional utility from having children in several ways. First, they draw utility from education spending q per child until the children graduate from high school, captured by a utility function $v(q)$. The utility is discounted by a function $b(n)$, increasing and concave in the number of children n . Further, based on altruistic motives, parents draw utility by transferring assets to their children after high school graduation. More specifically, parents consider children's expected lifetime utility in their education choice stage, with a discount rate $\lambda_a \cdot b(n)$, where λ_a represents the altruistic discount factor.

Costs of Children: Having children is costly in terms of money and time. First, q units of the per-child investment require $n \cdot q$ units of money, and they will make an additional expenditure $n \cdot a_{IVT}$ upon high school graduation of their children. In addition to the monetary costs, having children requires κ units of time until children graduate from high school and become independent.

Labor Earnings: Households choose hours worked to earn income. The labor earnings of a household are given by $w_e \eta_{j,z,e,h}(T - l)$; it is a product of the market wage w_e , labor efficiency or productivity $\eta_{j,z,e,h}$, and hours worked $(T - l)$. Here, T denotes the disposable time that can be devoted to work or leisure where $T = 1 - \kappa \cdot n$ if they have n children before graduating high school and $T = 1$ otherwise. Labor efficiency depends on age j , idiosyncratic productivity shock component z , education level e , and time-invariant human capital h .

Financial Markets: Financial markets are incomplete due to the lack of state-contingent claims. Households can self-insure against risks by savings, accruing interest payments at a rate of r . Households aged $j < J_R - 1$ except college students can borrow at rate $r^- = r + \iota$ where $\iota > 0$ (i.e., borrowers incur the overseeing costs ι) up to a borrowing limit \underline{A} , while we do not allow retired households' net worth to be negative. In addition, eligible students have access to student loans subsidized by the government, which entails the interest rate of $r^s = r + \iota_s$. This loan is income- and ability-tested, and eligible students can borrow up to a limit \underline{A}_s .

Government: The government raises the revenue by levying three types of taxes: consumption, labor income, and capital income taxes, where these tax rates are represented as τ_c , τ_w , and τ_a . In addition, the government collects accidental bequests and devotes

them to cover expenditures. They use this revenue to fund (1) the public pension benefits, which gives p units of money to retired households each period, (2) subsidized loans for college students, (3) grants for eligible college students, where the payment per eligible student is represented by $g(h, I)$ where $I \geq 0$ denotes the household income when the student faces education choice problem (i.e., their parent's age is J_{IVT}), (4) lump-sum transfers ψ that is introduced for replicating the progressivity of labor income tax schedule in a simple way following the literature, (5) cash benefits for households with children under 17 with per-child payment of B , and (6) other expenditures S . I do not consider grants in the benchmark to replicate the economy before grants are introduced in 2020 and set $g(h, I) = 0$ for each (h, I) . The government budget constraint is given as follows:

$$\tau_c \cdot C + \tau_w \cdot (L_{HS} + L_{CL}) + \tau_a \cdot K + Q = p \cdot \mu_{old} + (\iota - \iota_s) \cdot K_s + G + \psi + B \cdot \mu_{j \leq 17} + S, \quad (1)$$

where C , Q , μ_{old} , $\mu_{j \leq 17}$, K_s , and G represent the total consumption, total accidental bequests, population mass of retired households, that of children under age 17, the total amount of borrowing by college students, and the total grant payments.

4.3 Household Problems

This section describes problems agents face throughout their lifecycle. The first is about college enrollment choice, choosing whether to attend college or enter the labor market after graduating high school.

Education Stage: After graduating high school, they draw the time-invariant human capital h from the distribution $g_{h_p}^h$ and psychic cost ϕ from the distribution g_{h, e_p}^ϕ . They also receive IVT from their parents $a_{IVT} \geq 0$. Some of them can access subsidized loans to fund expenditures that arise during the college education stage, which is income- and ability-tested. Thus, the state variables for the students are comprised of asset a_{IVT} , the human capital h , psychic costs ϕ , and their parent's income I . They compare the expected value for entering the labor market as a high school graduate, $\mathbb{E}V^w$, with the value for enrolling in college, V_{g1} net of the psychic cost ϕ . They choose college enrollment if the latter is greater than the former; otherwise, they enter the labor market. The decision problem is formulated as follows:

$$V_{g0}(a_{IVT}, \phi, h, I) = \max_{e \in \{0,1\}} \left\{ (1 - e) \cdot \mathbb{E}_{z_0}[V^w(a_{IVT}, j = 18, z_0; e = 0, h)] + e \cdot [V_{g1}(a_{IVT}; h, I) - \phi] \right\}, \quad (2)$$

where $e \in \{0, 1\}$ indicates the education choice where $e = 1$ means college enrollment and $e = 0$ does entering the labor market as a high school graduate. V^w denotes the value

function for workers, which I formulate in the next subsection. The initial draw of z , z_0 , is uncertain and is according to the invariant distribution of z , $\bar{\pi}_z$, so the expectation operator is put next to the V^w . The value for college enrollment, V_{g1} , is defined as follows:

$$\begin{aligned} V_{g1}(a_{IVT}; h, I) &= \max_{c, l, a'} \{u(c, l) + \beta V_{g2}(a'; h, I)\}, \\ V_{g2}(a; h, I) &= \max_{c, l, a'} \{u(c, l) + \beta \mathbb{E}_{z_0}[V^w(a^s(a'), j = 22, z_0; e = 1, h)]\}. \end{aligned} \quad (3)$$

The budget constraints differ according to eligibility to the student loans. The budget constraints for eligible students are give as follows:

$$\begin{aligned} (1 + \tau_c)c + p_{CL} + a' - (1 - \tau_w)w_{HS}(1 - \bar{t} - l) \\ - \psi - g(h, I) &= \begin{cases} (1 + (1 - \tau_a)r)a & \text{if } a \geq 0, \\ (1 + r^s)a & \text{otherwise,} \end{cases} \\ a' &\geq -\underline{A}_s. \end{aligned} \quad (4)$$

The rest of the students cannot access to the student loans, which implies that their budget constraints are given as follows:

$$\begin{aligned} (1 + \tau_c)c + p_{CL} + a' &= (1 + (1 - \tau_a)r)a + (1 - \tau_w)w_{HS}(1 - \bar{t} - l) + \psi + g(h, I), \\ a' &\geq 0. \end{aligned}$$

College students draw utility from consumption and leisure and must pay tuition fees p_{CL} . Normalizing their total disposable time as 1, students must spend a \bar{t} fraction of time on study. Thus, they choose the time allocation between leisure and working over the disposable time $1 - \bar{t}$. One unit of labor supply gives college students w_{HS} units of wages.²¹ They can fund the consumption and tuition fees through (1) transfers from their parents a_{IVT} , (2) borrowing through student loans if eligible, (3) government-provided grants if eligible, and (4) labor earnings by themselves. Following the literature, I assume that fixed payments are made for 20 years (10 periods) following college graduation and transform college loans into regular bonds according to the following formula:

$$a^s(a') = a' \times \frac{r^s}{1 - (1 + r^s)^{-10}} \times \frac{1 - (1 + r^-)^{-10}}{r^-}.$$

Working Stage without Children: The remaining component in (2) and (3), V^w , represents the value function for working stage for age $j \in \{J_E, \dots, J_F - 1, J_{IVT} + 1, \dots, J_R - 1\}$. The state variables for this stage consists of asset a , age j , idiosyncratic component

²¹I assume that, while in college, there is no heterogeneity in labor efficiency and no uncertainty regarding the next period's productivity, and one unit of hours worked brings one unit of labor efficiency.

of labor productivity z , education level e , and human capital h . The uncertainty in this stage is only about the next period's productivity, which is denoted by z' following a Markov process $\pi_z(z', z)$. Households choose consumption, leisure, and savings given the state variables. The value function is formulated as follows:

$$V^w(a, j, z; e, h) = \max_{c, l, a'} \left\{ u(c, l) + \begin{cases} \beta \mathbb{E}_{z'}[V^f(a', z', e, h)] & \text{if } j = J_F - 1 \\ \beta[V^r(a', j + 1)] & \text{if } j = J_R - 1 \\ \beta \mathbb{E}_{z'}[V^w(a', j + 1, z'; e, h)] & \text{otherwise} \end{cases} \right\} \quad (5)$$

s.t.

$$\begin{aligned} (1 + \tau_c)c + a' &= (1 - \tau_w)w_e \eta_{j, z, e, h}(1 - l) + \psi + (1 + (1 - \tau_a)r)a, \\ z' &\sim \pi(z', z), \\ a' &\geq \begin{cases} 0 & \text{if } j = J_R - 1, \\ -\underline{A} & \text{otherwise.} \end{cases} \end{aligned}$$

Value functions V^f and V^r represent those for the fertility choice stage and retirement stage, respectively, which are formulated in the following subsections.

Fertility Choices and Working Stage with Children: When aged $j = J_F$, they choose whether and how many to have children. As in the previous life stages, they draw utility from consumption and leisure and decide on consumption, time allocation, and savings. In addition to that, as I mentioned, they also draw utility from the number of children n and education spending q for each child.²² Then, the value function for households aged $j = J_F$, V^f , is formulated as follows:

$$V^f(a, z, e, h) = \max_{n \in \{0, 1, \dots, N\}} \left\{ V^{wf}(a, j = J_F, z; e, h, n) \right\}$$

where, for $j = J_F, \dots, J_{IVT} - 1$,

$$V^{wf}(a, j, z; e, h, n) = \max_{c, l, q, a'} \left\{ u(c/\Lambda(n), l) + b(n) \cdot v(q) + \begin{cases} \beta \mathbb{E}_{z'}[V^{wf}(a', j + 1, z'; e, h, n)] & \text{if } j \in \{J_F, \dots, J_{IVT} - 2\} \\ \beta \mathbb{E}_{z', \phi_k, h_k}[V^{IVT}(a', z'; \phi_k, h_k, e, h, n)] & \text{if } j = J_{IVT} - 1 \end{cases} \right\}$$

s.t.

$$\begin{aligned} (1 + \tau_c)(c + nq) + a' &= Y_{wf}, \\ a' &\geq -\underline{A}, \end{aligned}$$

²²This study does not consider children's endogenous human capital accumulation through parental investments and assumes parents make educational spending on children just because it draws utility. Incorporating the endogenous human capital accumulation is left for future research.

where

$$Y_{wf} \equiv (1 - \tau_w)w_e\eta_{j,z,e,h}(1 - l - \kappa \cdot n) \\ + n \cdot B + \psi + \begin{cases} (1 + (1 - \tau_a)r)a & \text{if } a \geq 0, \\ (1 + r^-)a & \text{otherwise.} \end{cases}$$

Here, ϕ_k and h_k denote the psychic costs and human capital for their children. The household consumption is deflated by the equivalence scale $\Lambda(n)$, depending on the number of children n . The total education spending on children, nq , enters the budget constraint, and utility from children, $b(n) \cdot v(q)$, enters the objective function in the periods with children. V^{IVT} represents the value function for the stage of making IVT at $j = J_{IVT}$, which is described in the following subsection.

In addition to the next period's productivity z' , they are uncertain about their children's human capital h_k and psychic costs ϕ_k until the beginning of age $j = J_{IVT}$. Recall that the children's human capital correlates to the parent's, and the distribution of their psychic costs depends on the parent's education level and student's own human capital. Uncertainty about h_k and ϕ_k can be translated into uncertainty about expenditures on children in the form of IVT, given that those two components govern the marginal gains from IVT for parents,

$$b(n) \cdot \lambda_a \cdot \frac{\partial V_{g0}(\mathbf{x})}{\partial a_{IVT}},$$

where $\mathbf{x} = (a_{IVT}, \phi_k, h_k, I)$.²³

Inter-vivo Transfers: At age $j = J_{IVT}$, which corresponds to a period when their children graduate from high school and face the college enrollment choice, they decide on how much to transfer to their children. Some of the children's characteristics are realized at this point, such as their psychic cost ϕ_k and human capital h_k . Then, they choose the amount of per-child transfer a_{IVT} given the state variables, which is formulated as follows:

$$V^{IVT}(a, z; \phi_k, h_k, e, h, n) = \max_{c, l, a', a_{IVT}} \left\{ V^w(a - \tilde{a}_{IVT}, j = J_{IVT}, z; e, h) \right. \\ \left. + b(n) \cdot \lambda_a \cdot V_{g0}(a_{IVT}, \phi_k, h_k, I) \right\},$$

where $\tilde{a}_{IVT} = \frac{n \cdot a_{IVT}}{1 + (1 - \tau_a)r}$ and the parent's state vector pins down household income I . The budget constraint is given as follows:

$$(1 + \tau_c)c + a' + na_{IVT} = Y_{IVT},$$

$$a' \geq -\underline{A},$$

²³This representation implicitly assumes that the value function V_{g0} is differentiable for better clarity. However, in principle, this is not the case because of the discrete nature of the college enrollment choice.

where

$$Y_{IVT} \equiv (1 - \tau_w)w_e\eta_{j,z,e,h}(1 - l) + \psi + \begin{cases} (1 + (1 - \tau_a)r)a & \text{if } a \geq 0, \\ (1 + r^-)a & \text{otherwise.} \end{cases}$$

Retirement Stage: At the beginning of age J_R , they are forced to retire from the labor market. After that, they spend all their time on leisure and make consumption-saving decisions. Two points differ from previous choice problems; they receive the pension benefit p from the government and face uncertainty about the next period's survival. The value function for the retirement stage is formulated as follows:

$$\begin{aligned} V^r(a, j) &= \max_{c, a'} u(c, 1) + \beta \xi_{j,j+1} V^r(a', j + 1) \\ \text{s.t.} \\ (1 + \tau_c)c + a' &= p + (1 + (1 - \tau_a)r)a + \psi, \\ a' &\geq 0. \end{aligned}$$

4.4 Stationary Equilibrium

I solve the stationary equilibrium of this economy. In equilibrium, households make every choice to maximize their expected utility, the firm maximizes its profit, and the government budget is balanced. The stationarity implies that the distribution over state variables is invariant. Importantly, the age distribution is determined endogenously according to households' fertility choices. See [Appendix B](#) for the detailed definition of equilibrium and [Appendix C](#) for the computational algorithm for solving the equilibrium.

5 Calibration

To calibrate the model, I mainly rely on the Japanese Panel Survey of Consumers (JPSC), a panel survey of Japanese women and their household members. It started in 1993 with a representative sample of 1,500 women aged 24 – 34 and contains information about, for example, their income, educational background, marriage, fertility, and expenditures to detailed categories, including those on children's education. I focus on the cohort born in 1959-69, the oldest cohort of this survey, especially to compute the completed fertility and intergenerational mobility of education. I keep only married households as in previous works (e.g., [Daruich and Kozlowski, 2020](#)) because the model focuses on choices made within married households such as fertility and educational investments.²⁴ Unless oth-

²⁴Hence, an agent or household in this model refers to households with two individuals. Then, “children” in this model can also be interpreted as a household unit. That is, having n children can be

erwise mentioned, target moments for internally determined parameters described below are computed based on the JPSC's 1959-69 cohort data.

5.1 Targeted Moments

Preferences: Instantaneous utility for agents are given as follows:

$$u(c, l) = \frac{(c^\mu l^{1-\mu})^{1-\gamma}}{1-\gamma}.$$

μ is internally determined as 0.23 so that the households spend one-third of the total disposable time on market work. Instantaneous utility from education spending on a child is given as:

$$v(q) = \lambda_q \frac{q^{1-\gamma}}{1-\gamma},$$

where $\lambda_q = 0.62$ so that the annual educational expenditure per child amounts to 7% of average income at age 28. The utility must be always positive (or always negative) in models of altruism with endogenous fertility, and I set $\gamma = 0.5$ following the literature (e.g., [Daruich and Kozlowski, 2020](#)). The altruistic discount factor λ_a is set to 1.03 so that the annual average IVT for college students amounts to 27% of average income at age 28.²⁵ Following [Kim et al. \(2023\)](#), I let the discount function of number of children $b(n)$ be non-parametric, and assume that $b(n) = b_n$ for each $n \in \{0, 1, 2, 3, 4\}$ with $b(0) = b_0 = 0$. Those parameters are determined so that the model replicates the distribution of the completed fertility. The time for studying \bar{t} is set to 0.8 so that the income share of labor earnings for college students in the model is close to 21% (SLS, 2018). I assume that $\beta = 0.98$ as in [Zhou \(2022\)](#).

Financial Markets: I set the borrowing limits $\underline{A} = 20$ million yen and $\underline{A}_s = 2.88$ million yen.²⁶ The borrowing wedges are set $\iota = 0.055$ and $\iota_s = 0.054$ so that the model approximates the share of working households with a negative net worth (54%) and the share of students borrowing from the government-provided student loans (44%).

interpreted as reproducing $n/2$ units of households. For example, if a household has two children, it means reproducing one household unit with two individuals. If a household has one child, it means a reproduction of 0.5 units of households with two individuals.

²⁵Although some parents whose children do not enroll in college make IVT in the model, the amount is negligibly small. One reason is that the marginal gains from IVT are more significant if their children attend college, as students are financially constrained, primarily because of their limited earning ability.

²⁶The former is based on the Family Income and Expenditure Survey by the Ministry of Internal Affairs and Communications and the latter is based on the SLS (2018).

School Taste: I assume that the psychic costs ϕ are given as $\phi = \psi_{CL} \cdot \exp(-\nu \cdot h) \cdot \tilde{\phi}$. First, ψ_{CL} governs the scale of psychic costs and thus college enrollment rate. The second term $\exp(-\nu \cdot h)$ allows high ability students to have smaller psychic costs, as standard in the literature, and ν governs the education sorting by ability. Finally, $\tilde{\phi}$ is stochastic and the distribution depends on their parent’s education. Following [Daruich and Kozłowski \(2020\)](#), I assume that $\tilde{\phi}$ is distributed on an interval $[0, 1]$ and follows the following CDF:

$$G_{e_p}^{\tilde{\phi}} = \begin{cases} \tilde{\phi}^\omega & \text{if } e^p = 0 \\ 1 - (1 - \tilde{\phi})^\omega & \text{if } e^p = 1 \end{cases}$$

Here, ω governs the intergenerational transmission of school tastes. ψ_{CL} is set to 20.8 so that the model approximates the college enrollment rate of 37.7%. ν is set to 1 in the benchmark, and ω is set to 1.71 so that the intergenerational transition matrix of education in the model matches the data counterpart as closely as possible. Table 1 reports the transition matrix in the model and data. (i, j) –th entry of the matrix indicates the probability that children acquire skill j given that their parent’s skill is i in the benchmark model, and values in parentheses represent the data counterparts. The table indicates that the education level is persistent across generations; children of high school graduate parents attend college with a probability less than 0.3, whereas children of college graduate parents do with a probability approximately 0.6.

Parents/Children	HS	CL
HS	0.725 (0.798)	0.275 (0.202)
CL	0.412 (0.423)	0.588 (0.577)

Table 1: Intergenerational transition matrix of education. *Note:* (i, j) –th entry of the matrix indicates the probability that children acquire skill j given that their parent’s skill is i in the benchmark model, and values in parentheses represent the data counterparts.

Intergenerational Transmission of Human Capital: The intergenerational transmission of human capital is according to the following formula:

$$\begin{aligned} \log(h) &= \rho_h \cdot \log(h_p) + \varepsilon_h, \\ \varepsilon_h &\sim N(0, \sigma_h). \end{aligned}$$

I assume $\rho_h = 0.19$ based on [Daruich and Kozłowski \(2020\)](#) while internally determining σ_h as 0.65 so that the variance of log income at age 28 in this model is close to 0.27.

Income Process and Education Return: The efficiency labor of an agent aged j , education e , human capital h , and productivity z , $\eta_{j,z,e,h}$, is given as follows:

$$\begin{aligned}\log \eta_{j,z,e,h} &= \log[f^e(h)] + \gamma_{j,e} + z, \\ f^e(h) &= h + e \cdot (\alpha_{CL} h^{\beta_{CL}}), \\ z' &= \rho_z z + \zeta, \quad \zeta \sim N(0, \sigma_z).\end{aligned}$$

To set $\gamma_{j,e}$, I estimate the second-order polynomial of hourly wages on age using JPSC. As reported in Table 2, the income gradient on age is larger for college graduates than the rest of the workers, but the degree is modest compared with the US case (e.g., [Abbott et al., 2019](#)). I assume $\rho_z = 0.95$ and $\sigma_z = 0.02$, values in the ranges over those frequently used in the literature.²⁷ The function $f^e(h)$ indicates that the education return depends on human capital h : people with higher human capital can obtain greater returns through college education. Following [Darulich and Kozlowski \(2020\)](#), α_{CL} and β_{CL} are determined so that the model replicates the ratio of log wage between college graduates and the rest of the population at age $j = 28$ and the log wage variance for college graduates at age $j = 28$.

	HS	CL
Age	0.041	0.048
Age ² \times 10,000	−4.551	−5.364

Table 2: Parameters governing the age profile of wages. *Note:* CL indicates the college graduate households where the husband or wife is a college graduate. HS represents the rest of the population.

Production: I set the capital share $\alpha = 0.33$ and $\delta = 0.07$ following [Kitao \(2015\)](#). χ is set to 0.39 following [Matsuda and Mazur \(2022\)](#). ω_{CL} is internally determined to 0.52 so that the average wage ratio between college graduates and the rest in the model amounts to 1.36. Z is determined so that the wage rate for unskilled labor is normalized to one.

Government: Tax rates are set to $\tau_c = 0.1$, $\tau_w = 0.35$, and $\tau_a = 0.35$ in the benchmark. The lump-sum transfer is set to $\psi = 0.01$ to match the ratio between the variance of log net income and that of log gross income (0.6). The pension benefit p is set so that the government provides ¥160,000 per household per month. The cash transfer B is given as ¥10,000 per child per month, which approximates the actual payment. The other expenditure S is set so that the government budget constraint is balanced in the benchmark and fixed throughout the counterfactual experiments.

²⁷Due to data limitations, it is hard to accurately estimate the AR(1) process for z using any data source in Japan.

Miscellaneous: The survival probability $\zeta_{j,j+1}$ is set based on the Vital Statistics (2019).²⁸ Annual college tuition fees p_{CL} are set to 1.05 million yen. κ is set to 0.044 so that they spend 13.3% of their working hours on childcare.²⁹ Table 3 and Table 4 summarise the parameters externally and internally determined.

Parameter	Value	Description
\underline{A}_s	2.88 million yen	Borrowing limit for students
\underline{A}	20 million yen	Borrowing limit
p_{CL}	1.05 million yen/year	Tuition fees
κ	0.044	Time costs
$\xi_{j,j+1}$	—	survival prob.
τ_c	0.10	Consumption tax
τ_a	0.35	Capital income tax
τ_w	0.35	Labor income tax
p	¥160,000/month	Pension benefits
b	¥10,000/month	Cash transfers
α	0.33	Capital share
δ	0.07	Depreciation rate
χ	0.39	Elasticity of substitution
ρ_z	0.95	Persistence
σ_z	0.02	Transitory
ν	1.0	Education sorting by ability
γ	0.5	Curvature
β	0.98	Discount factor
ρ_h	0.19	Transmission of h

Table 3: Parameters externally determined.

²⁸See, <https://www.mhlw.go.jp/english/database/db-hw/outline/index.html>.

²⁹See, Kitao and Nakakuni (2023).

Parameter	Value	Moment	Data	Model
μ	0.23	Work hours	0.33	0.30
\bar{t}	0.8	Income share of labor earnings	0.21	0.20
ι_s	0.054	Share of students using loans	0.44	0.34
ι	0.055	Household share with negative net worth	0.54	0.45
ω_{CL}	0.52	CL–HS wage ratio	1.36	1.48
ψ	0.01	Var(log disposable income)/Var(log gross income)	0.60	0.68
λ_q	0.62	Average transfer / Average income at age 28	0.07	0.07
λ_a	1.03	Average transfer / Average income at age 28	0.27	0.27
ω	1.71	Intergenerational mobility of education	See Table 1	
σ_h	0.65	Variance of log(income) at age 28	0.27	0.24
ψ_{CL}	20.8	College enrollment rate	0.377	0.376
α_{CL}	0.1	Log wage ratio (CL–HS) at age 28	0.34	0.38
β_{CL}	0.1	Var log wage for CL at age 28	0.14	0.24
b_1	0.49	Share of one child	0.16	0.15
b_2	0.53	Share of two children	0.55	0.61
b_3	0.55	Share of three children	0.22	0.24
b_4	0.56	Share of four or more children	0.02	0.00
Z	1.99	Low skill wage	1.0	1.0

Table 4: Parameters internally determined.

5.2 Non-targeted Moments and Validation

This subsection checks the validity of the calibrated model. First, I check if the benchmark model generates a reasonable value of the benefit elasticity of fertility, which is non-targeted in calibration. Second, I check if the model performs well on other critical non-targeted moments, fertility differential across education levels and the revenue breakdown for college students.

5.2.1 The benefit elasticity of fertility

Previous works show that cash benefits such as the CB and baby bonus have a significant impact on fertility. Many of them report that the benefit elasticity of fertility, the percentage increase in fertility rate against the one percent increase in the cash transfer, is about 0.1 – 0.2. For example, [Milligan \(2005\)](#) studies a reform of Quebec’s baby bonus and shows that an extra 1,000 Canadian dollars benefit would increase fertility by 16.9%, which implies a benefit elasticity of 0.107. [Cohen et al. \(2013\)](#) uses a variation in the child

subsidy payment observed around 2003 in Israel, providing a larger subsidy for third or higher births. They show that the benefit elasticity of fertility was 0.176. [González \(2013\)](#) adopts the regression discontinuity design to study the effects of Spain’s reform in 2007, introducing a one-time payment of 2,500 euros (about 3,800 USD) for births, almost 4.5 times the monthly minimum wage for full-time workers. It finds a statistically significant impact on fertility, increasing conceptions by 5 – 6%.

To examine how this model performs in this respect, I conduct the following exercise. Let B_0 denote the per-child cash transfers for households with children under 17 in the benchmark. I solve the household problem, holding prices, tax rate, and distribution fixed, with several levels of the per-child payment $B = B_0 \cdot (1 + x)$ for some $x \in X$, where X is a set of positive real numbers. This procedure brings the implied fertility rate, and with the expansion rate x , we can compute the implied benefit elasticity of fertility for the case of the expansion rate x . I set $X = \{0.1, 0.2, \dots, 1.9, 2.0\}$, which is a reasonable range in the context of the expansion examined in the empirical studies, and compute the implied elasticity for each x . Then, I take the average of those 20 values. I find that the average elasticity is 0.138 (with a standard deviation of 0.025), which is consistent with the empirical estimates.

5.2.2 The fertility differential across education levels

More educated parents have fewer children than less educated ones. According to my sample of the JPSC, college graduate parents’ completed fertility was 1.92, which is lower than the rest’s, 2.12. This is observed in another data source that we can check the completed fertility by different education levels, the NFS. According to the NFS, college graduate wives’ completed fertility has been lower than less educated ones almost every survey year since 1977. The latest survey in 2015 reports that the completed fertility of wives with a college degree was 1.89, and that of high school graduate wives was 1.98. In this benchmark model, the completed fertility of college graduate wives is 1.79, which is lower than the high school graduate wives’, 2.28. The benchmark model captures the qualitative feature of the fertility differential across education levels, although the gap is somewhat significant compared with data counterparts. [Table 5](#) summarizes the fertility differential across education levels in the model and data.

	Model	JPSC	NFS
HS	2.28	2.12	1.98
CL	1.79	1.92	1.89

Table 5: Fertility differential across education in the benchmark model and data. *Note:* “NFS” stands for the National Fertility Survey conducted by the ISPP, and the table reports the values from the 2015 survey. It reports the completed fertility of wives with different educational backgrounds.

A factor generating the gap in this model is the differences in opportunity costs, given that having children requires a fixed fraction of time. More educated parents face the higher opportunity costs of having children as their potential earnings are higher. In this model, the intergenerational persistence of education can also generate this differential because children of educated parents are more likely to attend college, which raises financial costs in the form of asset transfers upon children’s college enrollment.

5.2.3 Composition of students’ revenue

Capturing the composition of students’ revenue – how college students finance their living expenses and tuition fees – is also important as it is critical not only to students’ education choices but also to parents’ IVT choices and, thereby, fertility choices. According to the SLS (2018), the students’ revenue consists of three parts. First, the greatest part, 61% of their revenue, is accounted for by asset transfers from their parents. Second, students’ labor earnings account for 21%, and lastly, the rest (18%) is financed by student loans. Although the revenue share of labor earnings (21%) is a targeted moment, the rest is not targeted. As Table 6 shows, the model captures the overall revenue composition as well; the IVT and loans account for 66% and 14% of their revenue, close to the data counterparts.

	IVT	Loan	Labor
Data	0.61	0.18	0.21
Model	0.66	0.14	0.20

Table 6: Composition of Students’ Revenue.

6 Numerical Analysis

This section investigates the effects of education subsidies for college students in the model with fertility choices calibrated in the previous section. Section 6.1 simulates the introduction of the existing income-tested subsidy started in 2020. Section 6.2 then

examines the mechanism through which the macroeconomic effects of the introduction are realized. Section 6.3 investigates the effects of raising the income threshold so that students in broader income classes of households are eligible. Lastly, Section 6.4 examines the performance of the potential alternative programs to the existing income-tested grants. Following the literature, I adjust the labor income tax to balance the government budget upon the introduction and expansion.

6.1 Introducing Education Subsidies

In 2020, the Japanese government introduced income-tested subsidies for college students in low-income households; until then, the government has provided only student loans. Households are eligible if their last year's annual labor income is less than a threshold value \bar{I} , and students in those households receive g amount of money each year while in college. The grant function $g(h, I)$ can be formulated as follows:

$$g(h, I) = \begin{cases} g & \text{if } I < \bar{I} \\ 0 & \text{otherwise} \end{cases} \quad (6)$$

Note that the benchmark case can be interpreted as $g(h, I) = 0$ for any (h, I) . Although the income threshold \bar{I} and payment g in the actual system vary with some characteristics of households and students, such as family structure and whether the student commutes to college from home or not, \bar{I} approximately corresponds to the 15 percentile of the household income distribution, and the payment approximately amounts to two-thirds of the average expenses of college students. I set \bar{I} and g based on this information, and income distribution and students' expenditure in the benchmark model (initial steady state). I then solve the stationary equilibrium by introducing this new grant function.

The main numerical results are as follows. First, introducing the means-tested subsidy would increase the college enrollment rate by 3.9 p.p. in the long run. Education mobility increases in the sense that children of high school graduate parents are 2.5 p.p. more likely to attend college.³⁰ Because of the higher college enrollment rate, implying a greater supply of college graduate labor, the skill premium decreases by 0.02 points. Introducing the subsidy would increase the TFR by 3% or 0.064 points. Importantly, this increase is primarily driven by a 7.4% increase in college graduate parents' fertility, whereas the fertility rate of high school graduate parents is almost stable.

Those changes in demographic structure and skill distribution affect other aggregate variables. In the long run, the higher college enrollment rate implies a higher share of

³⁰This mobility concept is a variant of an income mobility measure adopted in [Zheng and Graham \(2022\)](#).

skilled labor, and the higher TFR implies a greater share of the working-age population. As a result, per-capita labor supply in efficiency units increases by 1.3% in the long run. On the contrary, per-capita capital decreases by 1.8% for some reasons. First, introducing the subsidy reduces saving incentives for a substantial fraction of households and crowd-out IVT.³¹ In addition, the higher TFR implies a greater share of younger generations, who hold fewer assets than older ones. Despite its negative effects on capital accumulation, the positive impacts on the labor force and productivity are sufficiently greater so that the per-capita output increases by 0.7% in the long run.

6.2 Inspecting the Mechanism

Previous results imply that the subsidy introduction significantly affects the college enrollment rate and TFR, which leads to a greater output. However, the mechanism behind these changes is not so obvious because several objects aside from the grant function, such as prices, tax rate, and distribution, vary between the benchmark and new equilibrium, and each can play important roles in accounting for the overall effects on the TFR and college enrollment rate. In addition, the roles of fertility margins are also worth investigating, that is, how the results differ under exogenous fertility where the fertility choice is fixed. This is because the fertility setup would matter to the IVT and, thereby, children's education choices, given that fertility and IVT choices are joint decisions. Fertility margins can also have distributional implications through intergenerational linkages if fertility responses are heterogeneous across household characteristics. In this section, I conduct a decomposition analysis in Section 6.2.1 to discuss and understand the mechanism behind increases in the TFR and college enrollment rates upon introducing the subsidies. Next, in Section 6.2.2, I consider the roles of fertility responses in understanding the results by solving the exogenous fertility version of the model.

6.2.1 Decomposition: behavioral and composition effects

What causes the increases in the TFR and college enrollment rates? Those increases can be broken down into behavioral effects and distributional (or composition) effects, where the behavioral effects can be further broken down into the direct effects, driven only by a change in the subsidization scheme (i.e., grant function $g(h, I)$), and the indirect effects, driven by changes in factor prices (i.e., GE effects) or tax rates (i.e., Taxation effects).

³¹Note that some households, characterized on the state space, can increase savings due to the subsidy introduction. For example, households whose children do not attend college in the benchmark but attend college when the subsidy is introduced may need more assets in the new equilibrium to make some IVT upon children's college enrollment. This is because the grants are not sufficiently significant to cover 100% of the expenditures for students.

Note that the direct effects can also be interpreted as the short-run effects of the introduction, where prices, tax, and distribution are fixed. Finally, the distributional effects capture changes driven by distribution changes over state variables such as education, age, and human capital. To isolate those effects, I conduct a decomposition exercise as follows. I first solve an equilibrium by introducing the new grant function. Then, I solve household problems by replacing one of the four objectives (grant function, prices, tax rate, and state distribution) in the benchmark with that in the new equilibrium. Note that this method does not guarantee that each implied effect adds up to the overall effects because all factors except grant function are endogenous and interconnected. Table 7 summarises the decomposition results.

College Enrollment: For the long-run increase in the college enrollment rate, critical forces are direct and distributional effects, whereas GE and taxation effects do not play significant roles, at least in the case of introduction. If the grant is introduced while other variables (i.e., prices, tax rate, and distribution) are fixed as in the benchmark, the college enrollment rate increases by 2.6 p.p., more than half of the overall increase. These direct effects capture the effects of relaxing financial constraints on college enrollment decisions. Note that these direct effects do not explain the overall effects, so we need other forces accounting for the increase in college enrollment rate; that turns out to be the distributional effects.

The implied changes in the distribution in the long run, holding grant function, prices, and tax rate fixed as in the benchmark, lead to a 1.9 p.p. higher college enrollment rate. The most relevant factor is the change in the skill distribution of parents (i.e., the share of college graduates). The increase in the share of college graduates implies a higher share of those more likely to have children enrolling in college due to the intergenerational persistence of education. Then, the distributional change amplifies the short-run increase in college enrollment rate due to the direct effects.

Fertility: For the long-run increase in the TFR, vital forces are direct and GE effects. If the grant is introduced while other variables are fixed as in the benchmark, the TFR increases by 2.3% in the long run, which corresponds to three-fourths of the overall effects. As in explaining the overall effects, that 2.3% increase due to the direct effects is largely driven by the fertility increase of college graduate parents; college graduate parents increase fertility by 4.5%, whereas high school graduate parents do by 1%.

To understand these direct effects on fertility, note that parents do not know whether they are beneficiaries of the grants when they make fertility choices, eighteen years before their children graduate from high school and the eligibility is determined. They are

beneficiaries if (1) their children enroll in college and (2) their earnings, at the period when their children choose whether to attend college, are sufficiently low. Those are subject to uncertainty regarding their productivity and children’s characteristics (school tastes and human capital), which are not realized when they make fertility choices.

Now, let us look at changes in the average number of children across income distribution at the timing of IVT choices between the benchmark and new equilibrium. The average number of children for households in 1st income quintile, in which almost all children are eligible for grants if they enroll in college, remains the same. This means that the TFR increase is explained by fertility increases by ex-post ineligible households. Actually, households in 2nd and 3rd income quintiles increase fertility by 5.7%, and those in 4th and 5th income quintiles increase by 3.9%, on average. This may not be so surprising, given that I mentioned that college graduate parents increase fertility, and their income is relatively high.

Why can the subsidy introduction increase fertility, even that of ex-post ineligible households? The answer is due to an insurance effect of the income-tested subsidy. As I explained above, whether their potential children attend college is uncertain before giving birth. If their children have characteristics favoring college enrollment and thus they would like to make more IVT to support them, it significantly increases the financial costs of children. In other words, in addition to income risks throughout their working periods, they face a type of expenditure uncertainty until their children’s education choices terminate.

Consider that those “shocks” are realized simultaneously; their children have characteristics willing to attend college, and they are also willing to make a sufficient amount of IVT, but they are poor due to the realization of negative income shocks. Without the income-tested grants, those unlucky households end up with less consumption and/or IVT, which may prevent their children from enrolling in college; these come at significant utility losses. Thus, in fertility decision-making, they may choose fewer children to avoid such a situation in the future. The income-tested grants provide partial insurance against those risks and can make parents comfortable having another child. Because college graduate parents are more likely to have children with characteristics favoring college enrollment, they benefit more from this insurance, and their fertility behavior responds strongly.

Next, the decomposition result shows that the GE effect also plays a role in accounting for the TFR increase. If the prices are set to the long-run equilibrium levels while other objects are fixed as in the benchmark, the college graduates’ fertility increases by 2.5%, and the TFR then increases by 0.9%. The key is the decline of the wage rate for college graduate workers, w_{CL} . In the long-run equilibrium with the grants, w_{CL} decreases by

2.2%. First, the greater supply of college graduate workers depress w_{CL} relative to w_{HS} . In addition, greater aggregate labor supply in efficiency units and lower capital accumulation discussed in Section 6.1 imply the lower marginal productivity of labor, decreasing w_{CL} . The lower w_{CL} implies the lower opportunity costs of having children for college graduates, making some have more children.

Lastly, the distributional effect implies a 0.4% decline in the average fertility. The introduction of the grants increases the college enrollment rate, meaning that the share of college graduates increases, especially in the long run. Given that college graduates have fewer children than high school graduates as discussed in Section 5.2.2, the greater share of college graduates means a greater share of those who tend to have fewer children, which can lead to a lower average fertility.³² However, direct and GE effects are significant and lead to the higher TFR.

	Direct	GE	Tax	Dist.	All
CL share (Δ p.p.)	+2.6	0.0	0.0	+1.9	+3.9
TFR (Δ %)	+2.3	+0.9	0.0	-0.4	+3.0
HS (Δ %)	+1.0	0.0	0.0	0.0	+0.4
CL (Δ %)	+4.5	+2.5	0.0	0.0	+7.4

Table 7: Decomposition results. *Note:* Columns “Direct,” “GE,” “Tax,” and “Dist.” report the results when only grant function, prices, labor income tax rate, and distribution change, respectively. A column “All” indicates the results in the benchmark and the long-run equilibrium with the grants. Rows “CL share” represent the change in college enrollment rate in p.p. and rows “HS” and “CL” represent the percentage changes in fertility rates for high school and college graduates.

6.2.2 Roles of Fertility Responses

I next solve the exogenous fertility version of the model, a standard setup used in the literature on the education subsidy for college students. More specifically, I follow the same procedure in the previous section 6.1, except that the policy functions for fertility are fixed as in the benchmark.

Table 8 reports the main results under exogenous fertility. First, the introduction of grants results in a 2.9 p.p. increase in the college enrollment rate under exogenous fertility, which is lower than a 3.9 p.p. increase under endogenous fertility. Under both exogenous and endogenous fertility, the grants make some children who otherwise cannot enroll in college. Under endogenous fertility, in addition to that effect, the college enrollment

³²This effect is comparable with a composition effect highlighted by Zhou (2022) in understanding the effects of public education subsidies on fertility. In his model, public education subsidies increase parents with higher human capital, and they tend to have fewer children.

rate can further increase through fertility margins. As I discussed in Section 6.1, college graduate parents increase fertility than high-school graduate parents. Their children will likely attend college due to the intergenerational transmission of school tastes and human capital. And when those children become parents in the future, their children, if they have, are also likely to have similar characteristics to theirs, favoring college enrollment. Through this mechanism, the long-run share of college graduates increases, implying that the fertility margins amplify the effects on college enrollment.

From the second row onward, the changes in aggregate variables are reported. Under exogenous fertility, per-capita labor in efficiency units increases by 0.2 % compared with the benchmark mainly because of the higher share of skilled labor. However, the extent of this increase is modest compared with a 1.3% increase under endogenous fertility because the college enrollment rate is lower and working-age population share remains constant.³³ However, because of the fixed demographic structure, the negative effects on capital realized under endogenous fertility disappear. Rather, the introduction would lead to a 0.6% increase in per-capita capital under exogenous fertility. The per-capita output would increase by 0.7 %, the same degree as under endogenous fertility.

	Endogenous	Exogenous
CL share (Δ p.p.)	+3.9	+2.9
Output (Δ %)	+0.7	+0.7
Capital (Δ %)	−1.8	+0.6
Labor (Δ %)	+1.3	+0.2
Tax (Δ p.p.)	+0.04	+0.13

Table 8: Results under exogenous fertility. *Note:* Results for output, labor, and capital represent percentage changes compared with the benchmark. Those for college enrollment and labor income tax rates represent changes in percentage points compared with the benchmark.

6.3 Expansion

Finally, I consider the effects of raising the income threshold \bar{I} so that students in households of broader income classes can be eligible. Recall that I set \bar{I} so that the existing subsidy targets the students in households at the bottom 15 % of the income distribution. In this experiment, I increase \bar{I} to correspond to the 40, 50, and 60 percentile of the income distribution and solve the stationary equilibrium in each case. For students in the

³³Note that fixing policy functions for fertility does not necessarily mean that the TFR is also the same as in the benchmark because the household distribution changes, which can change the TFR even though the policy functions are fixed. However, it turns out that the TFR does not change in this case of exogenous fertility.

household at the bottom 15 %, the payment is still given by g , which amounts to two-thirds of the average expenses of students in the benchmark. For students in households higher than the 15 percentile but less than x percentile of the income distribution, where x takes either 40, 50, or 60, the payment is given by $g/2$, which amounts to one-third of the average expenses of students. In other words, the payment tapers off in income. Letting $\bar{I}_{15\%}$ and $\bar{I}_{x\%}$ denote the income level of 15 and $x \in \{40, 50, 60\}$ percentiles of the income distribution, the grant function $g(h, I; x)$ with a threshold $\bar{I}_{x\%}$ can now be formulated as:

$$g(h, I; x) = \begin{cases} g & \text{if } I < \bar{I}_{15\%} \\ g/2 & \text{if } I \in [\bar{I}_{15\%}, \bar{I}_{x\%}) \\ 0 & \text{otherwise} \end{cases} \quad (7)$$

Table 9 summarizes the results, and there are several takeaways there. First, the equilibrium college enrollment increases as the income threshold is higher. Setting the threshold at the 40, 50, and 60 percentiles of the income distribution, the enrollment rate increases by 4.7 p.p., 5.6 p.p., and 6.2 p.p., respectively, in stationary equilibrium. This is almost the case for education mobility, per-capita output, and labor income tax rates. With the income threshold at the bottom 60 % of the distribution, the per-capita output increases by 1.53%, which is more significant than 0.7% with the 15 percentile threshold. The expansion requires additional revenue, so the equilibrium tax rate should also increase by 0.3 p.p. for the case of the 60% threshold.

However, the expansion would not significantly increase the TFR. Recall that the introduction leads to a 3% increase in TFR, from the benchmark level of 2.096 to 2.160 in the long-run equilibrium with the grants. With grant functions $g(h, I; 40)$, $g(h, I; 50)$, and $g(h, I; 60)$, the equilibrium TFR would be 2.158, 2.151, and 2.157, respectively; the TFR even decreases locally by expansions. Fertility rates for each skill help us understand this situation. First, college graduate parents continue to increase fertility. With grant functions $g(h, I; 40)$, $g(h, I; 50)$, and $g(h, I; 60)$, their equilibrium fertility rate would be 1.996, 1.998, and 2.021, all higher than the fertility rate with the existing subsidy, 1.978. This result is straightforward to understand given that the insurance and GE effects contribute to the fertility increase of college graduate parents when the grants are introduced, discussed in Section 6.2. On the contrary, high school graduate parents reduce fertility instead, making the TFR remain almost constant even though the income threshold is higher. Fig 7 depicts the changes in the TFR and fertility rates for high school and college graduates, where the equilibrium TFR with the existing program (with the 15 percentile threshold) is normalized to one.

	<u>Threshold</u>			
	15%	40%	50%	60%
CL share (Δ p.p.)	+3.9	+4.7	+5.6	+6.2
HS \rightarrow CL (Δ p.p.)	+2.5	+2.6	+3.1	+3.3
Output ($\Delta\%$)	+0.70	+0.15	+1.07	+1.53
Tax (Δ p.p.)	+0.04	+0.17	+0.23	+0.30

Table 9: Main results of higher income thresholds. *Note:* Rows “CL share” and “HS \rightarrow CL” represent the changes in college enrollment rate and education mobility in the sense of probability that children of high school graduates attend college. Output changes are expressed as percentage changes. Changes in college enrollment rate, education mobility, and tax rate are represented as changes in percentage points.

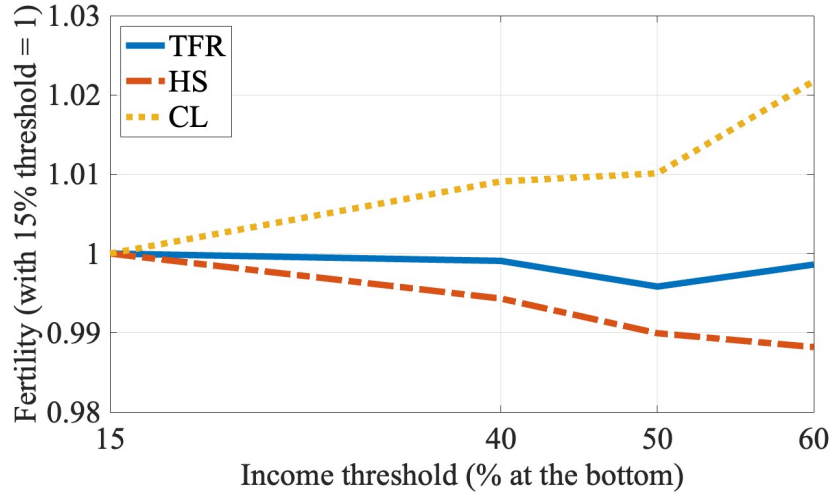


Fig 7: Changes in fertility rates with expansion. *Note:* The equilibrium TFR with the existing program (with the 15% threshold) is normalized to one. “HS” and “CL” represent fertility rates for high school and college graduates.

Why do high school graduate parents decrease fertility as the income threshold is higher? To understand this, I conduct the decomposition when the grant function is given by $g(h, I; 60)$, and the results are summarized in Table 10. Two effects are critical to explain the fertility decreases of high school graduates: the direct and GE effects.

In this case with the grant function $g(h, I; 60)$, the wage rate for high school graduates, w_{HS} , is 0.6% higher in the long run than in the benchmark, especially because of the significant supply of college graduates. This higher wage rate makes the opportunity costs of having children for high school graduates higher, which puts downward pressure on their fertility rates.

Next, the direct effects imply a lower fertility rate for high school graduates. This result is somewhat confusing given that we discuss the insurance effects of the subsidy on fertility, which is critical to account for the fertility increases of college graduates. Why

do the direct effects lead to lower fertility for high school graduates? In the benchmark without grants, children of high school graduate parents are unlikely to attend college. If the grants are introduced and their target expands, education mobility increases in the sense that children of high school graduate parents are more likely to attend college than in the benchmark, as represented in Table 9. Given that the grants are not generous enough to cover 100% of the costs to send their children to college, this higher probability of children going to college implies that those parents are more likely to have to make additional transfers upon children’s college enrollment. The expansion thus increases the expected costs of children for high school graduates, which would lower their fertility.

	Direct	GE	Tax	Dist.	All
HS ($\Delta\%$)	-0.7	-1.4	0.0	0.0	-0.8
CL ($\Delta\%$)	+9.1	+4.3	+2.4	+0.6	+13.2

Table 10: Decomposing the effects on fertility with $g(h, I, 60)$. *Note:* Rows “HS” and “CL” represent the percentage changes in fertility rates for high school and college graduates.

6.4 Reallocating Resources to Different Programs

Lastly, I examine the performance of the potential alternative programs to the existing income-tested grants. Recall that the primary goal of introducing and expanding the income-tested grants for college students is to alleviate the financial burdens of children and help people bear children as they desire, which then increases fertility rates. Thus, evaluating alternative programs in our framework is valuable to understanding effective instruments to achieve the pro-natal goal.

To this end, I consider another three scenarios, in addition to the introduction of income-tested grants to the benchmark model, examined in Section 6.1. The first scenario is to introduce grants for college students with income and ability tests so that “high-ability” students in poor households are eligible. More specifically, I keep the income threshold adopted in the existing scheme but arbitrarily set the lower bound for students’ human capital for the eligibility, \underline{h} , to its median. The payment function $g(h, I)$ for this scheme is defined as follows:

$$g(h, I) = \begin{cases} g_1 & \text{if } I \leq \bar{I} \text{ \& } h \geq \underline{h} \\ 0 & \text{otherwise} \end{cases}$$

Here, the payment g_1 is set so that the short-run expenditure upon the introduction is the same as that of the existing income-tested grants. As a result, g_1 covers approximately 100% of the average student’s expenditure in the benchmark, which is greater than the

payment in the existing scheme with only income-test because the number of eligible students is fewer.

The second scenario introduces unconditional grants for college students regardless of ability and income. The payment function is given as $g(h, I) = g_2$ for any students with (h, I) , where g_2 is set so that the short-run government expenditure is the same as the existing program. As a result, g_2 covers approximately 10% of the average students' expenditure in the benchmark, which is less significant than in the existing income-tested grants because this alternative program covers a broader range of students.

The third scenario is to increase the unconditional cash transfers to households with children B (i.e., child benefit), not reallocating money to grants for college students. As in other alternatives, the new level of the per-child payment B is set in the expenditure-neutral way, which leads to a 6.5% increase in the per-child payment.

I solve the stationary equilibrium with each scenario, and the results regarding fertility and enrollment rates are summarized in Table 11. A key highlight is that the existing scheme with income tests would lead to the highest equilibrium TFR among other scenarios. As discussed in Section 6.1, the income-tested grants would increase the TFR by 3% in the long-run equilibrium. The first, second, and third scenarios (corresponding to columns “+Ability,” “Uncond.,” and “Cash” in Table 11) increase the equilibrium TFR by 2.7%, 0.9%, and 1.0%, respectively. The previous decomposition analysis suggests that targeting students in unlucky households in which negative income shocks are realized would be effective, at least marginally, in increasing the TFR.³⁴

	Income	+Ability	Uncond.	Cash
CL share (Δ p.p.)	+3.9	+2.6	+5.1	+0.0
TFR ($\Delta\%$)	+3.0	+2.7	+0.9	+1.0
HS	+0.4	−3.5	+4.0	+2.5
CL	+7.4	+8.4	+3.9	+2.2

Table 11: Main results with several schemes with different targets. *Note:* values in each cell indicate changes from the benchmark value. Rows “HS” and “CL” indicate the percentage changes in the fertility of high school and college graduates.

³⁴A caveat is that this model does not capture the endogenous human capital accumulation before high school graduation and the dynamic complementarity of human capital. These ingredients should be critical in comparing the effects of these different programs, especially on college enrollment rates and aggregate output. The exercise in this framework is a first step, and incorporating those ingredients – fertility, college enrollment, dynamic complementarity of human capital– in one framework is left for future research.

7 Concluding Remarks

This study constructs a GE-OLG model embedding choices on college enrollment, IVT, and fertility to investigate the macroeconomic implications of education subsidies for college students. The model is calibrated to the Japanese economy using panel data. I show that the existing means-tested grants introduced in 2020 would increase the fertility of skilled parents, including ex-post ineligible ones, by providing insurance against income and expenditure risks arising from having children. They would also increase the college enrollment rate, amplified by fertility margins and distributional effects in the long run. Setting a higher income threshold would increase skilled parents' fertility further. However, the unskilled parents would decrease instead because, contrary to the skilled, the expected costs would significantly increase due to the higher probability of their children going to college, stemming from the greater education mobility thanks to the expansion.

This paper concludes by listing two points not considered in this preliminary version and to be examined further for future revision. First, while the current version focuses on the positive aspects of education subsidies, examining their welfare effects can also be relevant. In doing so, I would build on the work by [Goloso et al. \(2007\)](#) and adopt efficiency criteria that can be applied to models with endogenous fertility. Additionally, solving the transition dynamics and investigating the time horizon of the policy effects will provide another picture of the effects of education subsidies with demographic change through policy effects on fertility.

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Appendix A More on the Facts

i) Public Spending on Education in Japan

Japan is one of the least in subsidizing tertiary education, which roughly corresponds to four-year college education in Japan, given that the enrollment rates for tertiary education other than the four-year college are significantly lower than four-year college enrollment rate.³⁵

To see this, I use the cross-country data provided by the OECD³⁶ and define the *subsidization rate* for a specific education category e as follows:

$$s_e = \frac{y_e^{pub}}{y_e^{pri} + y_e^{pub}},$$

where y_e^{pri} and y_e^{pub} represent the private and public spending on education e , both represented as a share of the GDP. Public spending includes expenditures on educational institutions and educational-related subsidies for households or students. Private spending refers to expenditures financed by households and other private entities. Private spending includes expenditures on school but excludes those outside educational institutions (e.g., textbooks purchased by households, private tutoring, and student living costs).³⁷ In other words, the subsidization rate indicates what fraction of (potential) school-related costs are funded by the government.

Japan subsidizes more than 90% of the costs for primary-to-secondary education, which is higher than the OECD average. On the contrary, the subsidization rate for tertiary education was only 32%, which is the second lowest among OECD countries and less than half of the OECD average of 70%. This fact shows plenty of room for increasing subsidies for college students, which also has driven recent policy discussions on introducing and expanding education subsidies for college students.

Appendix B Equilibrium Definition

Let \mathbf{x}_j^e be an age-specific state vector for agents with education level $e \in \{HS, CL\}$ and $\mu_j^e(\mathbf{x}_j^e)$ be the measure of agents with state vector \mathbf{x}_j^e . Let $I_l(h, I)$ and $I_g(h, I)$ be indicator functions for loans and grants, respectively, returning 1 if students with (h, I) are eligible and 0 otherwise.

³⁵According to the Basic School Survey by the MEXT, the enrollment rate for some college was 4% in 2022 and is declining over the past thirty years.

³⁶I use 2018's data given that it is the latest year in which data on a significant number of countries is available.

³⁷For more detail, see the OECD data (<https://data.oecd.org/>).

Given exogenous parameters and policy rules $\{\tau_a, \tau_c, \iota, \iota_s, B, S, \psi, p, I_l(h, I), I_g(h, I)\}$, a *stationary recursive competitive equilibrium* consists of

- value functions $\{V_{g0}, V_{g1}, V_{g2}, V^w, V^f, V^{wf}, V^{IVT}, V^r\}$,
- policy functions for consumption, savings, leisure $\{c_j^e(\mathbf{x}_j^e), a_j^e(\mathbf{x}_j^e), l_j^e(\mathbf{x}_j^e)\}_{j=J_E}^J$, working hours $\{h_j^e(\mathbf{x}_j^e)\}_{j=J_E}^{J_R}$, fertility $\{n_{J_F}^e(\mathbf{x}_{J_F}^e)\}$, IVT $\{a_{J_{IVT}}^e(\mathbf{x}_{J_{IVT}}^e)\}$, and college enrollment $\{e_{J_E}(\mathbf{x}_{J_E})\}$,
- prices (r, w_{HS}, w_{CL}) ,
- labor income tax rate τ_w ,
- aggregate quantities (K, L_{HS}, L_{CL}) ,
- measures for households $\{\mu_j^e(\mathbf{x}_j^e)\}_{j=J_E}^J$,

such that:

1. The decision rules of students, workers, and retired households solve their problems, and $\{V_{g0}, V_{g1}, V_{g2}, V^w, V^f, V^{wf}, V^{IVT}, V^r\}$ are the associated value functions.
2. The representative firm maximizes its profit and optimally chooses capital and labor inputs:

$$r + \delta = \alpha \cdot Z \cdot \left(\frac{K}{L}\right)^{\alpha-1}, \quad (8)$$

$$w_{HS} = \tilde{Z} \cdot \omega_{HS} \cdot L_{HS}^{\chi-1}, \quad (9)$$

$$w_{CL} = \tilde{Z} \cdot \omega_{CL} \cdot L_{CL}^{\chi-1}, \quad (10)$$

where

$$L = [\omega_{HS} \cdot (L_{HS})^\chi + \omega_{CL} \cdot (L_{CL})^\chi]^{1/\chi},$$

$$\tilde{Z} = (1 - \alpha) \cdot Z \cdot \left(\frac{K}{L}\right)^\alpha \cdot [\omega_{HS} \cdot (L_{HS})^\chi + \omega_{CL} \cdot (L_{CL})^\chi]^{1/\chi-1}.$$

3. The labor market for each skill $e \in \{HS, CL\}$ clears:

$$L_e = \sum_{j=J_E}^{J_R} \int_{\mathbf{x}_j^e} \eta_j^e(\mathbf{x}_j^e) \cdot h_j^e(\mathbf{x}_j^e) d\mu_j^e(\mathbf{x}_j^e), \quad (11)$$

where $\eta_j^e(\mathbf{x}_j^e)$ represents the labor efficiency of agents with a state vector \mathbf{x}_j^e .

4. The capital market clears:

$$K = \sum_{j=J_E}^J \sum_e \int_{\mathbf{x}_j^e} a_j^e(\mathbf{x}_j^e) d\mu_j^e(\mathbf{x}_j^e). \quad (12)$$

5. The government budget is balanced:

$$\tau_c \cdot C + \tau_w \cdot (L_{HS} + L_{CL}) + \tau_a \cdot K + Q = p \cdot \mu_{old} + (\iota - \iota_s) \cdot K_s + G + \psi + B \cdot \mu_{j \leq 17} + S,$$

where

$$\begin{aligned} C &= \sum_{j=J_E}^J \sum_e \int_{\mathbf{x}_j^e} c_j(\mathbf{x}_j^e) d\mu_j^e(\mathbf{x}_j^e), \\ Q &= \sum_{j=J_R+1}^J \frac{1 - \zeta_{j-1,j}}{\zeta_{j-1,j}} \sum_e \int_{\mathbf{x}_j^e} a_j^e(\mathbf{x}_j^e) d\mu_j(\mathbf{x}_j^e), \\ p \cdot \mu_{old} &= \sum_{j=J_R}^J \sum_e \int_{\mathbf{x}_j^e} p d\mu_j^e(\mathbf{x}_j^e), \\ K_s &= \int_{\mathbf{x}^s} \max\{0, -a^s(\mathbf{x}^e)\} \cdot I_l(h, I) d\mathbf{x}^s \\ G &= \int_{\mathbf{x}^s} g(h, I) \cdot I_g(h, I) d\mathbf{x}^s \\ B \cdot \mu_{j \leq 17} &= \sum_{j=J_F}^{J_{IVT}-1} \sum_e \int_{\mathbf{x}_j^e} B \cdot n d\mu_j^e(\mathbf{x}_j^e), \end{aligned}$$

where \mathbf{x}^s , $\mu^s(\mathbf{x}^s)$, and $\{a^s(\mathbf{x}^s)\}$ represent a state vector for college students, measure of college students, and students' policy function for saving, respectively.

6. Distributions (measures) and households' behavior are consistent.

Appendix C Computational Algorithm: Stationary Equilibrium

For any variable or distribution x , let \tilde{x} and \hat{x} represent its guessed and model-implied values. Also, let μ_j represent the distribution over state variables for age j . Dividing the computation process into three blocks makes it easier to understand. The first block is the outer loop, searching for equilibrium prices. The other two blocks are inner loops; one is the *optimization block*, solving household problems given prices, and another is the *distribution block*, searching for the stationary distributions given prices and policy functions obtained in the optimization block. The computational algorithm proceeds as follows:

1. Guess prices $\tilde{\mathbf{p}} = (\tilde{r}, \tilde{w}_{HS}, \tilde{w}_{CL})$.

2. *Optimization block*:

- Guess the value function for agents at the beginning of age $j = 18$ (\tilde{V}_{g0}).
- Given \tilde{V}_{g0} , solve backward from the period of IVT choice to that of the education choice, which gives the model-implied value function for V_{g0} , \hat{V}_{g0} .
- Check if

$$d(\tilde{V}_{g0}, \hat{V}_{g0}) < \varepsilon, \quad (13)$$

where $d(\cdot)$ and $\varepsilon > 0$ represent an arbitrary metric function and error tolerance. If (13) is not satisfied, update \tilde{V}_{g0} and follow the same procedure until convergence. The correct V_{g0} pins down all value functions and policy functions given a set of prices $\tilde{\mathbf{p}}$.

3. *Distribution block*:

- Guess the distribution for age J_{IVT} . This $\tilde{\mu}_{J_{IVT}}$ and policy functions for IVT derive the implied distribution for agents aged 18, $\hat{\mu}_{18}$. Given $\hat{\mu}_{18}$ and policy functions, compute the implied distributions for age $j = 19, \dots, J_{IVT}$, and obtain $\hat{\mu}_{J_{IVT}}$.
- Check if

$$d(\tilde{\mu}_{J_{IVT}}, \hat{\mu}_{J_{IVT}}) < \varepsilon. \quad (14)$$

If (14) is not satisfied, update $\tilde{\mu}_{J_{IVT}}$ and follow the same procedure until convergence.

- After obtaining the correct distributions for age $j = 18, \dots, J_{IVT}$, compute the distribution for age $j_{IVT} + 1$ onward, using those distributions and policy functions.
4. After computing value functions, policy functions, and distributions, compute the implied quantities, \hat{L} and \hat{K} based on (11) and (12), which gives the implied prices $\hat{\mathbf{p}}$ based on \hat{L} , \hat{K} , (8), (9), and (10).
 5. Check if

$$d(\tilde{\mathbf{p}}, \hat{\mathbf{p}}) < \varepsilon. \tag{15}$$

If (15) is not satisfied, update $\tilde{\mathbf{p}}$, return to the optimization block, and follow the same procedure until convergence.