

习题一 事件与概率

一. 填空题

1. **答** $A = B$.
2. **答** $C_5^1 \cdot C_5^1 = 25$.
3. **答** $\{0, 1, 2, 3, 4, 5, 6, 7, 8, 9, 10\}$.
4. **答** $A_1 A_1 A_1 + A_1 A_1 A_2 + A_1 A_2 A_2 + A_1 A_1 A_3$.
5. **答** 不大于二次 (≤ 2).

二. 计算题

1. **答** (1) $A_1 A_2 A_3$; (2) $A_1 A_2 A_3$; (3) $A_1 \cup A_2 \cup A_3$.
2. **解** 每一个部件用 3 只强度较弱的钉的概率为 $\frac{1}{C_{50}^3}$, 则发生一个部件较弱的概率为 $\frac{10}{C_{50}^3} = \frac{1}{1960}$.

3. **解一** A 表事件“取到二只球中至少有一只是白球.
基本事件总数

$$n = C_{10}^2 = 45.$$

A 所包含基本事件数

$$C_4^1 \times C_6^1 + C_4^2 = 30, \quad P(A) = \frac{30}{45} = \frac{2}{3}.$$

$$\text{解二} \quad P(A) = \frac{C_4^1 C_6^1 + C_4^2 C_6^0}{C_{10}^2} = \frac{24 + 6}{45} = \frac{2}{3}.$$

$$\text{解三} \quad P(A) = 1 - P(\bar{A}) = 1 - \frac{C_6^2 \cdot C_4^0}{C_{10}^2} = 1 - \frac{15}{45} = \frac{2}{3}.$$

4. **解** 这是一古典概型概率问题, 设 A 表示“3 卷一套的放在一起”, B 表示“4 卷一套的放在一起”, C 表示“两套各自放在一起”, D 表示“两套按卷次顺序排好”.

3 卷一套的放在一起, 可把 3 卷看成一个整体, 总共有 8 个位置, 不同的放法共有 $8!$ 种, 3 卷一套之间可以任意排, 共有 $3!$ 种放法, 所以

$$P(A) = \frac{8! \times 3!}{10!} = \frac{1}{15}.$$

5. **解** A 表事件“每一个班级各分配到一名优秀生”.
基本事件总数

$$r = C_{15}^5 \cdot C_{10}^5 \cdot C_5^5 = \frac{15!}{5! \cdot 5! \cdot 5!}.$$

A 所包含的基本件数

$$r = C_3^1 \cdot C_{12}^4 \cdot C_2^1 \cdot C_8^4 \cdot C_1^1 \cdot C_4^4 = \frac{3! \times 12!}{4! \cdot 4! \cdot 4!},$$

$$P(A) = \frac{r}{n} = \frac{25}{91}.$$

习题二 概率性质 条件概率

一. 填空题

1. 答 0.6.

2. 答 0.8.

3. 答 0.5.

若 A 与 B 相互独立, 则 $P(AB) = P(A)P(B)$, 由概率的加法公式

$$P(A \cup B) = P(A) + P(B) - P(AB) = P(A) + P(B) - P(A)P(B),$$

$$\text{则得 } P(B) = \frac{P(A \cup B) - P(A)}{1 - P(A)} = \frac{0.7 - 0.4}{1 - 0.4} = \frac{0.3}{0.6} = 0.5.$$

4. 答 $\frac{17}{30}$.

二. 计算题

1. 解 A : “抽到的一人为男人”, B : “抽到的一人为色盲者”.

$$\text{则 } P(A) = \frac{3}{5}, \quad P(B|A) = \frac{5}{100} = \frac{1}{20},$$

$$P(\bar{A}) = \frac{2}{5}, \quad P(B|\bar{A}) = \frac{25}{10000} = \frac{1}{400}.$$

于是

$$P(B) = P(A)P(B|A) + P(\bar{A})P(B|\bar{A}) = \frac{3}{5} \times \frac{1}{20} + \frac{2}{5} \times \frac{1}{400} = \frac{31}{1000}.$$

2. 解 (1) 因为 $AB \subset A$, $AB \subset B$, 所以 $P(AB) \leq P(A)$, $P(AB) \leq P(B)$,
 $A \subset B$ 时 $P(AB)$ 最大, 其值为 $P(A) = 0.6$;

(2) 由 $P(A \cup B) = P(A) + P(B) - P(AB)$ 知, $P(A \cup B)$ 最大时, $P(AB)$ 最小, 当 $A \cup B = \Omega$ 时, $P(A \cup B)$ 最大, 此时 $P(AB)$ 最小为 0.3.

3. 解 设 $A = \{\text{父亲得病}\}$, $B = \{\text{母亲得病}\}$, $C = \{\text{孩子得病}\}$.

已知 $P(C) = 0.6$, $P(B|C) = 0.5$, $P(A|BC) = 0.4$,

则所求为 $P(BC\bar{A})$.

$$\begin{aligned} P(BC\bar{A}) &= P(C)P(B|C)P(\bar{A}|BC) \\ &= 0.6 \times 0.5 \times (1 - 0.4) \\ &= 0.18. \end{aligned}$$

4. **证** 由概率性质 $AB \subset C$, 有 $P(AB) \leq P(C)$.

$$P(A \cup B) = P(A) + P(B) - P(AB)$$

即
$$P(AB) = P(A) + P(B) - P(A \cup B) \leq P(C)$$

$$1 \geq P(A \cup B) \geq P(A) + P(B) - P(C)$$

故
$$P(A) + P(B) - P(C) \leq 1.$$

5. **证**
$$P(B|A) = \frac{P(BA)}{P(A)} = \frac{P(A) + P(B) - P(A+B)}{P(A)}$$

因为 $P(A+B) \leq 1$, 所以

$$P(A+B) \geq \frac{p_1 + p_2 - 1}{p_1} = 1 - \frac{1 - p_2}{p_1}.$$

习题三 全概率公式 贝叶斯公式

一. 计算题

1. **解** 令事件 $A_1 = \{\text{警报系统 } A \text{ 有效}\}$, 事件 $A_2 = \{\text{警报系统 } B \text{ 有效}\}$,

由 $P(A_1) = 0.93$, $P(A_2) = 0.94$, $P(A_2 | \bar{A}_1) = 0.85$, 得

$$P(\bar{A}_1 A_2) = P(\bar{A}_1) P(A_2 | \bar{A}_1) = 0.07 \times 0.85 = 0.0595,$$

$$P(A_1 A_2) = P(A_2) - P(\bar{A}_1 A_2) = 0.94 - 0.0595 = 0.8805,$$

$$\begin{aligned} \text{于是 } P(A_1 \cup A_2) &= P(A_1) + P(A_2) - P(A_1 A_2) \\ &= 0.93 + 0.94 - 0.8805 = 0.9895. \end{aligned}$$

或用下面方法计算

$$\begin{aligned} P(A_1 \cup A_2) &= 1 - P(\overline{A_1 \cup A_2}) = 1 - P(\bar{A}_1 \bar{A}_2) = 1 - P(\bar{A}_1) P(\bar{A}_2 | A_1) \\ &= 1 - P(\bar{A}_1) [1 - P(\bar{A}_2 | A_1)] = 1 - 0.07 \times (1 - 0.85) \\ &= 1 - 0.07 \times 0.15 = 0.9895. \end{aligned}$$

2. **解** 设 $A_1 = \text{"从甲袋中取出的为白球"}$, $A_2 = \text{"从甲袋中取出的为红球"}$, $B = \text{"从乙袋中取出的为白球"}$, 则

$$P(A_1) = \frac{n}{m+n}, \quad P(A_2) = \frac{m}{n+m}$$

$$P(B|A_1) = \frac{N+1}{M+N+1}, \quad P(B|A_2) = \frac{N}{M+N+1}.$$

则由全概率公式, 有

$$P(B) = P(A_1) \cdot P(B|A_1) + P(A_2) \cdot P(B|A_2) = \frac{n(N+1) + m \cdot N}{(m+n) \cdot (M+N+1)}.$$

3. **解** 设 $A = \text{"抽得次品"}$, B_1, B_2, B_3 分别表示抽得甲、乙、丙机床的产品. 由全概率公式

$$\begin{aligned} P(A) &= P(B_1) P(A|B_1) + P(B_2) P(A|B_2) + P(B_3) P(A|B_3) \\ &= 0.2 \times 0.2 + 0.3 \times 0.3 + 0.5 \times 0.1 = 0.18. \end{aligned}$$

由贝叶斯公式, 所求概率为

$$P(B_2|A) = \frac{P(B_2) P(A|B_2)}{P(A)} = \frac{0.3 \times 0.3}{0.18} = 0.5.$$

4. **解** 令事件 $A = \{\text{原发信息}0\}$, $\bar{A} = \{\text{原发信息}1\}$. $B = \{\text{收到信息为}0\}$, 首先求 $P(B)$.

已知 $P(A) = \frac{7}{10}$, $P(\bar{A}) = \frac{3}{10}$, $P(B|A) = 0.98$, $P(B|\bar{A}) = 0.01$, 利用全概率公式, 得

$$P(B) = P(A)P(B|A) + P(\bar{A})P(B|\bar{A}) = \frac{7}{10} \times 0.98 + \frac{3}{10} \times 0.01 = 0.689.$$

再利用贝叶斯公式, 得

$$P(A|B) = \frac{P(AB)}{P(B)} = \frac{P(A)P(B|A)}{P(B)} = \frac{\frac{7}{10} \times 0.98}{0.689} = 0.9956.$$

5. **解** 设 A_i 为“第 i 项考核合格” ($i=1, 2, 3, 4$). 已知

$$P(A_1) = 0.6, P(A_2) = 0.8, P(A_3) = 0.91, P(A_4) = 0.95,$$

因为 4 项考核相互独立, 且被淘汰的对立事件为 4 过, 所以:

- (1) 这项招工的淘汰率为

$$\begin{aligned} P(\overline{A_1 A_2 A_3 A_4}) &= 1 - P(A_1)P(A_2)P(A_3)P(A_4) \\ &= 1 - 0.6 \times 0.8 \times 0.91 \times 0.95 \\ &= 0.585. \end{aligned}$$

- (2) 通过第一、三项被淘汰的概率为

$$\begin{aligned} P(A_1 A_3 \overline{A_2 A_4}) &= P(A_1)P(A_3)P(\overline{A_2 A_4}) \\ &= P(A_1)P(A_3)[1 - P(A_2 A_4)] \\ &= P(A_1)P(A_3)[1 - P(A_2)P(A_4)] \\ &= 0.69 \times 0.91 \times [1 - 0.8 \times 0.95] = 0.131. \end{aligned}$$

- (3) 若考核按顺序进行, 应聘者一项考核不合格就被淘汰, 不再参加后面项目考核, 这种情况下的淘汰率为

$$\begin{aligned} &P\{\bar{A}_1 \cup (A_1 \bar{A}_2) \cup (A_1 A_2 \bar{A}_3) \cup (A_1 A_2 A_3 \bar{A}_4)\} \\ &= P(\bar{A}_1) + P(A_1 \bar{A}_2) + P(A_1 A_2 \bar{A}_3) + P(A_1 A_2 A_3 \bar{A}_4) \\ &= P(\bar{A}_1) + P(A_1)P(\bar{A}_2) + P(A_1)P(A_2)P(\bar{A}_3) \\ &\quad + P(A_1)P(A_2)P(A_3)P(\bar{A}_4) \\ &= (1 - 0.6) + 0.6 \times (1 - 0.8) + 0.6 \times 0.8 \times (1 - 0.91) \\ &\quad + 0.6 \times 0.8 \times 0.91 \times (1 - 0.95) \\ &= 0.585. \end{aligned}$$

习题四 离散型随机变量

一. 填空题

1. 答 0.2.

2. 答 $\frac{27}{8}e^{-3}$ 或 $3.375e^{-3}$.

二. 计算题

1. 解

ξ	0	1	2	3
P	$\frac{3}{4}$	$\frac{9}{44}$	$\frac{9}{220}$	$\frac{1}{220}$

2. 解 ξ 的取值范围为 $\xi=1, \xi=2, \xi=3, \xi=4$,

$$P\{\xi=1\} = \frac{c_3^3 \times c_6^1}{c_9^4} = \frac{1}{21}, \quad P\{\xi=2\} = \frac{c_3^2 \times c_6^2}{c_9^4} = \frac{5}{14},$$

$$P\{\xi=3\} = \frac{c_3^1 \times c_6^3}{c_9^4} = \frac{3 \times 20}{126} = \frac{10}{21}, \quad P\{\xi=4\} = \frac{c_3^0 \times c_6^4}{c_9^4} = \frac{5}{42}.$$

即

ξ	1	2	3	4
p	$\frac{1}{21}$	$\frac{5}{14}$	$\frac{10}{21}$	$\frac{5}{42}$

3. 解 (1) 此题为贝努利概型. 每次试验成功的概率 $p=0.3$, 设 5 次独立试验中成功的次数为 $X, X \sim B(5, 0.3)$. 要求 $P\{X \geq 3\}$, 有

$$\begin{aligned} P\{X \geq 3\} &= C_5^3 \times 0.3^3 \times 0.7^2 + C_5^4 \times 0.3^4 \times 0.7 + 0.3^5 \\ &\approx 0.1323 + 0.02835 + 0.00243 \\ &\approx 0.163. \end{aligned}$$

(2) 此题也为贝努利概型.

$$p=0.3, n=7, X \sim B(7, 0.3).$$

$$\begin{aligned} P\{X \geq 3\} &= 1 - P\{X \leq 2\} \\ &= 1 - [0.7^7 + C_7^1 \times 0.3 \times 0.7^6 + C_7^2 \times 0.3^2 \times 0.7^5] \\ &\approx 0.353. \end{aligned}$$

$$\begin{aligned}
 4. \quad \text{解} \quad \sum_{k=1}^{\infty} \frac{1}{k(k+1)} &= \sum_{k=1}^{\infty} \left(\frac{1}{k} - \frac{1}{k+1} \right) = \lim_{n \rightarrow \infty} \sum_{k=1}^n \left(\frac{1}{k} - \frac{1}{k+1} \right) \\
 &= \lim_{n \rightarrow \infty} \left(1 - \frac{1}{n+1} \right) = 1.
 \end{aligned}$$

又由分布律的性质有

$$1 = \sum_{k=1}^{\infty} P\{X=k\} = \sum_{k=1}^{\infty} \frac{A}{k(k+1)} = A \sum_{k=1}^{\infty} \frac{1}{k(k+1)} = A.$$

所以 $A=1$.

$$\begin{aligned}
 5. \quad \text{解} \quad P\{\xi=n\} &= \begin{cases} 0.3 \times (0.42)^{k-1}, & n=2k-1 \\ 0.28 \times (0.42)^{k-1}, & n=2k \end{cases} \\
 & n=1, 2, \dots
 \end{aligned}$$

习题五 分布函数 连续型随机变量

一. 填空题

1. **答** 由公式 $P\{X=x_0\}=F(x_0)-F(x_0-0)$, 算出

$$P\{X=-1\}=0.4-0=0.4,$$

$$P\{X=1\}=0.8-0.4=0.4,$$

$$P\{X=3\}=1-0.8=0.2.$$

所以 X 的概率分布为

X	-1	1	3
p	0.4	0.4	0.2

2. **答** $\begin{cases} x, & 0 \leq x \leq \sqrt{2} \\ 0, & \end{cases}$ (没有等号也对).

二. 计算题

1. **解** 因 $F(+\infty)=1$, 可知 $A=1$.

因为 $F(x)$ 是连续函数, 所以有

$$\lim_{x \rightarrow (-\pi/2)^+} F(x) = C = 0, \quad \lim_{x \rightarrow 0^-} F(x) = B = 1.$$

2. **解** $P\{0 < \xi \leq x\} = cx,$

$$F(x) = \begin{cases} 0, & x < 0 \\ \frac{x}{a}, & 0 \leq x < a \\ 1, & a \leq x \end{cases}.$$

3. **解** 设 $F(x)$ 表示 X 的分布函数, 当 $x < 0$ 时, 则

$$F(x) = P(X \leq x) = \int_{-\infty}^x f(u) du = \int_{-\infty}^x \frac{1}{2} e^u du = \frac{1}{2} e^x;$$

当 $x \geq 0$ 时, 则

$$\begin{aligned} F(x) &= P(X \leq x) = \int_{-\infty}^x f(u) du = \int_{-\infty}^0 \frac{1}{2} e^u du + \int_0^x \frac{1}{2} e^{-u} du \\ &= \frac{1}{2} + \left(\frac{1}{2} - \frac{1}{2} e^{-x} \right) = 1 - \frac{1}{2} e^{-x}, \end{aligned}$$

综合表示为

$$F(x) = \begin{cases} \frac{1}{2} e^x, & x < 0 \\ 1 - \frac{1}{2} e^{-x}, & x \geq 0 \end{cases}.$$

4. **解** (1) $P\{\xi \geq 2\} = 1 - P\{\xi < 2\} = 1 - F(2) = e^{-4}$.

(2) $P\{-3 \leq \xi < 4\} = F(4) - F(-3) = 1 - e^{-8} - 0 = 1 - e^{-8}$.

(3) 由 $P\{\xi \geq a\} = 1 - F(a) = P\{\xi < a\} = F(a)$,

知 $2(1 - e^{-2a}) = 1$, 解得 $a = \frac{\ln 2}{2}$.

5. **解** 首先求一只电子管工作 1000 小时以上的概率.

$$p = \int_{1000}^{+\infty} \frac{1}{1000} e^{-\frac{x}{1000}} dx = e^{-1} \approx 0.3679.$$

只有当 5 只电子管皆工作在 1000 小时以上, 仪器才能工作 1000 小时以上. 又“每只电子管工作 1000 小时以上”是相互独立的, 所以所求概率为

$$p^5 \approx 0.00673.$$

此概率很小.

习题六

随机变量的函数分布

一. 填空题

1.

答
$$\psi(y) = \begin{cases} \frac{1}{2}, & 1 \leq y \leq 3 \\ 0, & \end{cases}$$

2.

答案
$$\begin{array}{c|ccc} Y & 0 & \sqrt{2}/2 & 1 \\ \hline p & 9/16 & 5/16 & 1/8 \end{array}$$

提示
$$P\left\{Y = \frac{\sqrt{2}}{2}\right\} = P\left\{X = \frac{\pi}{4}\right\} + P\left\{X = -\frac{\pi}{4}\right\} = \frac{5}{16}, \text{ 等.}$$

二. 计算题

1. **解**

η	1	2	5	10
p	$\frac{1}{8}$	$\frac{1}{3}$	$\frac{5}{12}$	$\frac{1}{8}$

2. 解 $y = 2x^2 + 1, x^2 = \frac{y-1}{2}.$

见图 2-1, 当 $y \geq 1$ 时, Y 的分布函数

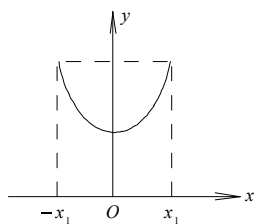


图 2-1

$$F_Y(y) = P\{Y \leq y\} = P\left\{X^2 \leq \frac{y-1}{2}\right\} = P\{-x_1 \leq X \leq x_1\}.$$

其中 $x_1 = \sqrt{\frac{y-1}{2}}$

从而

$$\begin{aligned} F_Y(y) &= \int_{-x_1}^{x_1} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx = \int_{-\sqrt{\frac{y-1}{2}}}^{\sqrt{\frac{y-1}{2}}} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \\ &= 2 \int_0^{\sqrt{\frac{y-1}{2}}} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx. \end{aligned}$$

对变上限求导, 得

$$f_Y(y) = \frac{1}{2\sqrt{\pi(y-1)}} e^{-\frac{(y-1)}{4}}.$$

当 $y < 1$ 时, $F_Y(y) = P\{Y \leq y\} = 0$, 从而 $f_Y(y) = 0$.

故
$$f_Y(y) = \begin{cases} \frac{1}{2\sqrt{\pi(y-1)}} e^{-\frac{(y-1)}{4}}, & y \geq 1 \\ 0, & y < 1 \end{cases}.$$

3. 解 K 服从 $U(0, 5)$ 分布, 其概率密度为

$$f(x) = \begin{cases} 1/5, & 0 < x < 5 \\ 0, & \text{其它} \end{cases}$$

方程 $4x^2 + 4Kx + K + 2 = 0$ 有实根的条件是:

$$\Delta = (4K)^2 - 4 \times 4(k+2) = 16(K-2)(K+1) \geq 0$$

解得 $K \geq 2$ 或 $K \leq -1$.

即 $\{\text{方程有实根}\} = \{K \geq 2\} \cup \{K \leq -1\}$.

故 $P\{\text{方程有实根}\} = P\{K \geq 2\} + P\{K \leq -1\}$

$$= \int_2^5 1/5 dx + 0 = 3/5.$$

4. 证 $P\{|X| < a\} = P\{-a < X < a\} = \int_{-a}^a f(x) dx = 2 \int_0^a f(x) dx$
 $= 2[F(a) - F(0)] = 2F(a) - 1.$

5. 解 (1) $Y = 2X + 3$, 于是

$$y = 2x + 3, \quad x = \frac{y-3}{2}, \quad x' = \frac{1}{2},$$

故
$$f_Y(y) = \begin{cases} \frac{1}{2} \left(\frac{y-3}{2}\right)^3 e^{-\left(\frac{y-3}{2}\right)^2}, & y \geq 3; \\ 0, & y < 3 \end{cases}$$

(2) $Y = X^2$, 由 $y = x^2$, 得出

$$x_1 = \sqrt{y}, \quad x_1' = \frac{1}{2\sqrt{y}}; \quad x_2 = -\sqrt{y} < 0, \quad x_2' = -\frac{1}{2\sqrt{y}}$$

故
$$\begin{aligned} f_Y(y) &= f(\sqrt{y})(\sqrt{y})' + f(-\sqrt{y})|(-\sqrt{y})'| \\ &= \frac{1}{2\sqrt{y}} (\sqrt{y})^3 e^{-(\sqrt{y})^2} + 0 \cdot \frac{1}{2\sqrt{y}} \\ &= \begin{cases} \frac{1}{2} y e^{-y}, & y > 0; \\ 0, & y \leq 0 \end{cases} \end{aligned}$$

(3) $Y = \ln X$, 于是

$$y = \ln x, \quad x = e^y, \quad x' = e^y,$$

故
$$f_Y(y) = f(e^y) e^y = 2 e^{4y} e^{-e^{2y}} \quad (-\infty < y < +\infty).$$

一. 填空题

1. **答案** $\frac{1}{36}$.

提示 $\sum_{i=1}^3 \sum_{j=1}^3 cij = 1$.

2.

答

$$f(x, y) = \begin{cases} \frac{1}{\pi}, & (x, y) \in D \\ 0, & (x, y) \notin D \end{cases},$$

设二维随机变量 (X, Y) 服从平面区域 D 上的均匀分布, 且区域 D 的面积为 A , 则 (X, Y) 的联合密度函数为

$$f(x, y) = \begin{cases} \frac{1}{A}, & (x, y) \in D \\ 0, & (x, y) \notin D \end{cases},$$

而区域 $D = \{(x, y) : x^2 + y^2 \leq 1\}$ 的面积为 π , 因此 (X, Y) 的联合密度函数为

$$f(x, y) = \begin{cases} \frac{1}{\pi}, & (x, y) \in D \\ 0, & (x, y) \notin D \end{cases}.$$

二. 计算题

1. **解** $F(x, y)$ 不可能是某二维随机变量的联合分布函数. 因

$$\begin{aligned} P\{0 < \xi \leq 2, 0 < \eta \leq 1\} &= F(2, 1) - F(0, 1) - F(2, 0) + F(0, 0) \\ &= 1 - 1 - 1 + 0 = -1 < 0. \end{aligned}$$

故 $F(x, y)$ 不可能是某二维随机变量的联合分布函数.

2. **解** (1)

$\xi \backslash \eta$	0	1
-1	1/2	0
0	1/3	1/6

$$(2) F(x, y) = \begin{cases} 0, & x < -1 \text{ 或 } y < 0 \\ 1/2, & -1 \leq x < 0, y \geq 0 \\ 5/6, & x \geq 0, 0 \leq y < 1 \\ 1, & x \geq 0, y \geq 1 \end{cases}.$$

$$\begin{aligned}
3. \quad & \text{解} \quad P\{\max\{X, Y\} \geq 0\} \\
& = P\{X, Y \text{ 至少一个大于等于 } 0\} \\
& = P\{X \geq 0\} + P\{Y \geq 0\} - P\{X \geq 0, Y \geq 0\} \\
& = \frac{4}{7} + \frac{4}{7} - \frac{3}{7} = \frac{5}{7}.
\end{aligned}$$

$$\begin{aligned}
4. \quad & \text{解} \quad (Y, Z) \text{ 可能取值为 } (0, 0), (0, 1), (1, 0), (1, 1), \text{ 且} \\
& P\{Y=0, Z=0\} = P\{X \leq 1, X \leq 2\} = P\{X \leq 1\} = F(1) = 1 - e^{-1} \\
& P\{Y=0, Z=1\} = P\{X \leq 1, X > 2\} = 0 \\
& P\{Y=1, Z=0\} = P\{X > 1, X \leq 2\} = P\{1 < X \leq 2\} \\
& \quad = F(2) - F(1) = e^{-1} - e^{-2} \\
& P\{Y=1, Z=1\} = P\{X > 1, X > 2\} = P\{X > 2\} \\
& \quad = \int_2^{+\infty} e^{-x} dx = e^{-2}
\end{aligned}$$

故

$Z \backslash Y$	0	1
0	$1 - e^{-1}$	$e^{-1} - e^{-2}$
1	0	e^{-2}

$$\begin{aligned}
5. \quad & \text{解} \quad F(x, y) = \begin{cases} \int_0^x \int_0^y 12e^{-(3x+4y)}, & x \geq 0, y \geq 0 \\ 0, & \text{其它} \end{cases} \\
& \quad = (1 - e^{-3x})(1 - e^{-4y})
\end{aligned}$$

$$\begin{aligned}
P\{0 \leq \xi < 1, 0 \leq \eta < 2\} & = F(1, 2) - F(1, 0) - F(0, 2) + F(0, 0) \\
& = (1 - e^{-3})(1 - e^{-8}).
\end{aligned}$$

一. 填空题

1. 解 ① ~ ⑦ 分别填: $\frac{5}{12}, \frac{1}{4}, \frac{1}{3}, \frac{1}{6}, \frac{1}{4}, \frac{5}{12}, \frac{1}{6}$.

2. 答 $\frac{1}{2\pi} e^{-\frac{1}{2}(x^2+y^2)}$.

二. 计算题

1. 解 当 $x > 0$ 时

$$f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \int_0^{+\infty} e^{-y} dy = e^{-x}.$$

当 $x \leq 0$ 时, $f(x, y) = 0$, 故 $f_X(x) = 0$. 得

$$f_X(x) = \begin{cases} e^{-x}, & x > 0 \\ 0, & \text{其它} \end{cases}.$$

同理 $f_Y(y) = \begin{cases} \int_0^x e^{-y} dy, & y > 0 \\ 0, & \text{其它} \end{cases} = \begin{cases} ye^{-y}, & y > 0 \\ 0, & \text{其它} \end{cases}.$

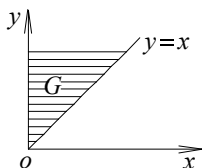


图 3-9

2. 解 (1) 本题是已知了 X_1 与 X_2 的边缘分布律, 条件 $P\{X_1 X_2 = 0\} = 1$, 求出联合分布. 列表如下:

$X_2 \backslash X_1$	-1	0	1	$P\{X_2 = j\}$
0	1/4	0	1/4	1/2
1	0	1/2	0	1/2
$P\{X_1 = i\}$	1/4	1/2	1/4	1

由已知 $P\{X_1 X_2 = 0\} = 1$, 即等价于 $P\{X_1 X_2 \neq 0\} = 0$, 可知

$$P\{X_1 = 1, X_2 = 1\} = 0, P\{X_1 = -1, X_2 = 1\} = 0.$$

再由 $p_{\cdot 1} = p_{-11} + p_{11} + p_{01}$, 得 $p_{01} = \frac{1}{2}$,

$$p_{-10} = p_{\cdot 1} - p_{-11} = \frac{1}{4}, p_{10} = p_{\cdot 1} - p_{11} = \frac{1}{4},$$

从而得 $p_{00} = 0$.

(2) 由于 $p_{-10} = \frac{1}{4} \neq p_{\cdot 1} \cdot p_{\cdot 0} = \frac{1}{4} \cdot \frac{1}{2} = \frac{1}{8}$,

所以知 X_1 与 X_2 不独立.

3. 解 (1) 根据 $\int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(x, y) dx dy = 1$

$$\begin{aligned} 1 &= \int_0^{+\infty} \int_0^{+\infty} k e^{-(5x+6y)} dx dy \\ &= k \left(\int_0^{+\infty} e^{-5x} dx \right) \cdot \left(\int_0^{+\infty} e^{-6y} dy \right) \\ &= k \times \frac{1}{5} \times \frac{1}{6} = \frac{1}{30} k \end{aligned}$$

故 $k=30$.

(2) 当 $x > 0$ 时

$$f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \int_0^{+\infty} 30 e^{-5x-6y} dy$$

$$\text{即} \quad f_X(x) = \begin{cases} 5 e^{-5x}, & x > 0 \\ 0, & x \leq 0 \end{cases},$$

同理当 $y > 0$ 时

$$f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \int_0^{+\infty} 30 e^{-5x-6y} dx$$

$$\text{即} \quad f_Y(y) = \begin{cases} 5 e^{-6y}, & y > 0 \\ 0, & y \leq 0 \end{cases},$$

因为 $f(x, y) = f_X(x) f_Y(y)$, 故 X, Y 相互独立.

4. 解 先写出 (X, Y, Z) 的联合分布列, 再写出 (X, Z) 的联合分布列 (用求边缘分布列的方法).

当 $z=1$ 时			当 $z=0$ 时		
$\begin{array}{c cc} Y \backslash X & 0 & 1 \\ \hline 0 & (1-p)^2 & 0 \\ 1 & 0 & p^2 \end{array}$			$\begin{array}{c cc} Y \backslash X & 0 & 1 \\ \hline 0 & 0 & p(1-p) \\ 1 & p(1-p) & 0 \end{array}$		

所以, (X, Z) 的联合分布列为

$\begin{array}{c cc} X \backslash Z & 0 & 1 \\ \hline 0 & p(1-p) & (1-p)^2 \\ 1 & p(1-p) & p^2 \end{array}$		
---	--	--

对 $X=i, Z=j, i, j=0, 1$, 若 X, Z 独立, 应有 $p_{ij} = p_i \cdot p_j$. 由

$$P\{X=1, Z=1\} = p^2 = P\{X=1\} P\{Z=1\} = p[p^2 + (1-p)^2]$$

解得 $p=1/2$, 经验证, 对 (X, Z) 的一切 (i, j) , 均满足

$$p_{ij} = p_i \cdot p_j,$$

所以, 当 $p=1/2$ 时, X 与 Z 相互独立.

习题九 条件分布 二维随机变量的函数分布

一. 填空题

1.

答

Z	1	4
P	0.16	0.84

由 X 的分布律可得 X^2 的分布律

X^2	1	4
P	0.4	0.6

而 Y^2 的分布律与 X^2 的分布律相同, 故随机变量 Z 的取值仅为 1 与 4. 于是

$$P(Z=1)=P(X^2=1, Y^2=1)=P(X^2=1)P(Y^2=1)=0.16,$$

$$P(Z=4)=1-P(Z=1)=0.84.$$

2. 答

ξ	0	1
$P\{\xi \eta \neq 1\}$	3/7	4/7

二. 计算题

1. 解

ξ	0	1
$\eta=1$	0.1	0.3
$\eta=2$	0.2	0.1
$\eta=3$	0.1	0.2

2. 解

ζ	-2	-1	$-\frac{1}{2}$	$-\frac{1}{4}$	0
P	0	4/15	1/5	0	8/15

评分参考: 每列 2 分.

3.

解

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)}.$$

注意在 y 处 x 值位于 $|x| \leq \sqrt{R^2 - y^2}$ 这个范围内, $f(x, y)$ 才有非零值, 故在此范围内, 有

$$f_{X|Y}(x|y) = \frac{\frac{1}{\pi R^2}}{\frac{2}{\pi R^2} \cdot \sqrt{R^2 - y^2}} = \frac{1}{2\sqrt{R^2 - y^2}},$$

即 $Y = y$ 时 X 的条件分布密度为

$$f_{X|Y}(x|y) = \begin{cases} \frac{1}{2\sqrt{R^2 - y^2}}, & |x| \leq \sqrt{R^2 - y^2} \\ 0, & \text{其它} \end{cases}.$$

同法可得 $X = x$ 时 Y 的条件分布密度为

$$f_{X|Y}(y|x) = \begin{cases} \frac{1}{2\sqrt{R^2 - x^2}}, & |y| \leq \sqrt{R^2 - x^2} \\ 0, & \text{其它} \end{cases}.$$

由于条件分布密度与边缘分布密度不相等, 所以 X 与 Y 不独立.

4.

解

$$F(z) = P\{\xi + \eta < z\} = \iint_{x+y < z} \varphi(x, y) dx dy,$$

当 $z < 1$ 时, $F(z) = 0$;

当 $1 \leq z < 2$ 时, $F(z) = \int_1^z dy \int_0^{z-y} 6 \times dx = (z-1)^3$;

当 $z \geq 2$ 时, $F(z) = 1$;

$$F(z) = \begin{cases} 0, & z < 1 \\ (z-1)^3, & 1 \leq z < 2 \\ 1, & z \geq 2 \end{cases}$$

$$\varphi(z) = \begin{cases} 3(z-1)^2, & 1 \leq z \leq 2 \\ 0, & \text{其它} \end{cases}.$$

习题十 随机变量的数学期望

一. 填空题

1. 答 2.4.

2. 答 2.

二. 计算题

1. 解 $E(X) = \int_{-\infty}^{+\infty} xf(x) dx$

$$\begin{aligned} &= \int_{-\infty}^0 0 dx + \int_0^{1500} \frac{x^2}{(1500)^2} dx + \int_{1500}^{3000} x \frac{3000-x}{(1500)^2} dx + \int_{3000}^{+\infty} 0 dx \\ &= \int_0^{1500} \frac{x^2}{(1500)^2} dx + \int_{1500}^{3000} \frac{(3000-x)x}{(1500)^2} dx = 1500. \end{aligned}$$

2. 解 $E(\xi) = \sum_{k=1}^3 x_k p_k = (-2) \times 0.4 + 0 \times 0.3 + 2 \times 0.3 = -0.2$

$$\begin{aligned} E(\xi - E(\xi))^3 &= \sum_{k=1}^3 (x_k - E\xi)^3 p_k \\ &= [-2 - (-0.2)]^3 \times 0.4 + (0 + 0.2)^3 \times 0.3 \\ &\quad + (2 + 0.2)^3 \times 0.3 \\ &= 0.864. \end{aligned}$$

3. 解 $E(V) = E(IR) = E(I) \cdot E(R)$

$$\begin{aligned} &= \int_{-\infty}^{+\infty} x \varphi_I(x) dx \int_{-\infty}^{+\infty} y \varphi_R(y) dy \\ &= \int_0^1 x \cdot 3x^2 dx \cdot \int_0^2 y \cdot \frac{1}{2} y dy = \frac{3}{4} x^4 \Big|_0^1 \cdot \frac{1}{6} y^3 \Big|_0^2 \\ &= \frac{3}{4} \times \frac{8}{6} = 1. \end{aligned}$$

4.

$$\text{解 } P\{Y = -1\} = P\{X < 10\} = P\left\{\frac{X - \mu}{1} < \frac{10 - \mu}{1}\right\} = \Phi(10 - \mu),$$

$$P\{Y = -5\} = P\{X > 12\} = P\left\{\frac{X - \mu}{1} > \frac{12 - \mu}{1}\right\} = 1 - \Phi(12 - \mu),$$

$$\begin{aligned} P\{Y = 20\} &= P\{10 \leq X \leq 12\} = P\left\{\frac{10 - \mu}{1} \leq \frac{X - \mu}{1} \leq \frac{12 - \mu}{1}\right\} \\ &= \Phi(12 - \mu) - \Phi(10 - \mu). \end{aligned}$$

所以, 随机变量 Y 的分布律为

Y	-5	-1	20
P	$1 - \Phi(12 - \mu)$	$\Phi(10 - \mu)$	$\Phi(12 - \mu) - \Phi(10 - \mu)$

由此得

$$\begin{aligned} E(Y) &= (-5) \times [1 - \Phi(12 - \mu)] + (-1) \times \Phi(10 - \mu) \\ &\quad + 20 \times [\Phi(12 - \mu) - \Phi(10 - \mu)] \\ &= 25\Phi(12 - \mu) - 21\Phi(10 - \mu) - 5. \end{aligned}$$

令 $g(\mu) = 25\Phi(12 - \mu) - 21\Phi(10 - \mu) - 5$, 则

$$\begin{aligned} g'(\mu) &= -25\Phi(12 - \mu) + 21\Phi(10 - \mu) \\ &= -25 \times \frac{1}{\sqrt{2\pi}} e^{-\frac{(12-\mu)^2}{2}} + 21 \times \frac{1}{\sqrt{2\pi}} e^{-\frac{(10-\mu)^2}{2}} \\ &= (-25 \times e^{2\mu-22} + 21) \times \frac{1}{\sqrt{2\pi}} e^{-\frac{(10-\mu)^2}{2}}. \end{aligned}$$

令 $g'(\mu) = 0$, 得 $g(\mu)$ 的唯一驻点

$$\mu_0 = 11 + \frac{1}{2} \ln \frac{21}{25} \approx 10.9128.$$

而且可以判定 μ_0 点是函数 $g(\mu)$ 的极大值点, 从而也是 $g(\mu)$ 的最大值点. 因此, 当 $\mu_0 = 11 + \frac{1}{2} \ln \frac{21}{25} \approx 10.9128$ 时, 零件的利润为最大.

一. 填空题

1. 答 $\frac{1}{6}$.

2. 答 由题设知 X 的概率密度函数为

$$f(x) = \begin{cases} \frac{1}{3}, & x \in [-1, 2] \\ 0, & \text{其它} \end{cases},$$

由连续型随机变量的数学期望公式有

$$\begin{aligned} E(Y) &= \int_{-\infty}^{+\infty} Y(x) f(x) dx = \frac{1}{3} \int_{-1}^2 Y(x) dx \\ &= \frac{1}{3} \left[\int_{-1}^0 Y(x) dx + \int_0^2 Y(x) dx \right] = \frac{1}{3} (-1 + 2) = \frac{1}{3}. \end{aligned}$$

$$\begin{aligned} E(Y^2) &= \int_{-\infty}^{+\infty} Y^2(x) f(x) dx = \frac{1}{3} \int_{-1}^2 Y^2(x) dx \\ &= \frac{1}{3} \left[\int_{-1}^0 Y^2(x) dx + \int_0^2 Y^2(x) dx \right] = \frac{1}{3} (1 + 2) = 1. \end{aligned}$$

于是 $D(Y) = E(Y^2) - E^2(Y) = 1 - \frac{1}{9} = \frac{8}{9}.$

二. 计算题

1. **解** 设 ξ 表示 k 次所抽球的号码和, ξ_i 表示第 i 次所抽球的号码,

$$\text{则 } \xi = \sum_{i=1}^k \xi_i$$

$$p(\xi_i = j) = \frac{1}{n} \quad (j = 1, 2, \cdots, n)$$

$$E(\xi_i) = \sum_{j=1}^n j \cdot \frac{1}{n} = \frac{1}{n} \sum_{j=1}^n j = \frac{1}{n} \frac{n(n+1)}{2} = \frac{n+1}{2}$$

$$\begin{aligned} E(\xi_i^2) &= \sum_{j=1}^n j^2 \cdot \frac{1}{n} = \frac{1}{n} \sum_{j=1}^n j^2 = \frac{1}{n} \cdot \frac{n(n+1)(2n+1)}{6} \\ &= \frac{(n+1)(2n+1)}{6} \end{aligned}$$

$$\begin{aligned} D(\xi_i) &= E(\xi_i^2) - (E(\xi_i))^2 \\ &= \frac{(n+1)(2n+1)}{6} - \frac{(n+1)^2}{4} = \frac{n^2 - 1}{12} \end{aligned}$$

$$E(\xi) = E\left(\sum_{i=1}^k \xi_i\right) = \sum_{i=1}^k E(\xi_i) = \frac{k(n+1)}{2}$$

因 $\xi_1, \xi_2, \cdots, \xi_k$ 相互独立, 故

$$D(\xi) = D\left(\sum_{i=1}^k \xi_i\right) = \sum_{i=1}^k D(\xi_i) = \frac{k(n^2 - 1)}{12}.$$

2. **解** 随机变量 X 的所有可能取值为 3, 4, 5, 取各个值的概率为

$$P(X=3) = \frac{1}{C_5^3} = \frac{1}{10}, \quad P(X=4) = \frac{C_3^2}{C_5^3} = \frac{3}{10}, \quad P(X=5) = \frac{C_4^2}{C_5^3} = \frac{6}{10},$$

则 X 的概率分布为

X	3	4	5
p_k	$\frac{1}{10}$	$\frac{3}{10}$	$\frac{6}{10}$

$$\text{而} \quad E(X) = 3 \times \frac{1}{10} + 4 \times \frac{3}{10} + 5 \times \frac{6}{10} = \frac{45}{10} = 4.5,$$

$$E(X^2) = 9 \times \frac{1}{10} + 16 \times \frac{3}{10} + 25 \times \frac{6}{10} = \frac{207}{10} = 20.7,$$

$$\text{所以} \quad D(X) = E(X^2) - [E(X)]^2 = 20.7 - (4.5)^2 = 0.45.$$

$$\begin{aligned}
 3. \quad \text{证} \quad E[(\xi - c)^2] &= E(\xi^2 - 2\xi c + c^2) = E(\xi^2) - 2cE(\xi) + c^2 \\
 &= E(\xi^2) + [c - E(\xi)]^2 - [E(\xi)]^2 \\
 &= [c - E(\xi)]^2 + D(\xi) > D(\xi) \quad (c \neq E(\xi)).
 \end{aligned}$$

4. 解 AB 的弦长为 $\eta = 2R |\sin \xi|$, ξ 的概率密度为

$$\varphi(\theta) = \begin{cases} \frac{1}{\pi}, & -\frac{\pi}{2} < \theta < \frac{\pi}{2} \\ 0, & \text{其它} \end{cases}$$

$$E(\eta) = \int_{-\pi/2}^{\pi/2} 2R |\sin \theta| \cdot \frac{1}{\pi} d\theta = \frac{4R}{\pi}$$

$$E(\eta^2) = \int_{-\pi/2}^{\pi/2} 4R^2 \sin^2 \theta \cdot \frac{1}{\pi} d\theta = 2R^2$$

$$D(\eta) = E(\eta^2) - E(\eta)^2 = 2R^2 \left(1 - \frac{8}{\pi^2} \right).$$

习题十二

协方差 相关系数

一. 填空题

1. **答** 由于

$$\begin{aligned}\text{Cov}(2X, 3Y) &= E\{[2X - E(2X)][3Y - E(3Y)]\} \\ &= 6E\{[X - E(X)][Y - E(Y)]\} = 6\text{Cov}(X, Y)\end{aligned}$$

又 $\rho_{XY} = \frac{\text{Cov}(X, Y)}{\sqrt{D(X)}\sqrt{D(Y)}}$, 故

$$\text{Cov}(2X, 3Y) = 6\rho_{XY} \sqrt{D(X)} \sqrt{D(Y)}$$

$$\begin{aligned}\text{得 } D(2X - 3Y) &= D(2X) + D(3Y) - 2\text{Cov}(2X, 3Y) \\ &= 4D(X) + 9D(Y) - 12\rho_{XY} \sqrt{D(X)} \sqrt{D(Y)} \\ &= 4 \times 4 + 9 \times 9 - 12 \times 0.5 \times 2 \times 3 \\ &= 97 - 36 = 61.\end{aligned}$$

2. **答** 0.1.

二. 计算题

1. **解** 由 $P(XY=1)=\frac{1}{5}$, 得 $P(X=1, Y=1)=\frac{1}{5}$, 利用 X, Y 的边缘概率分布, 有

$$P(X=0, Y=1)=P(Y=1)-P(X=1, Y=1)=\frac{1}{4}-\frac{1}{5}=\frac{1}{20},$$

$$P(X=0, Y=2)=P(X=0)-P(X=0, Y=1)=\frac{1}{3}-\frac{1}{20}=\frac{17}{60},$$

$$P(X=1, Y=2)=P(X=1)-P(X=1, Y=1)=\frac{2}{3}-\frac{1}{5}=\frac{7}{15}.$$

则得 (X, Y) 的联合概率分布为

$\begin{array}{c} X \\ Y \end{array}$	0	1	$P(Y=y_j)$
1	$\frac{1}{20}$	$\frac{1}{5}$	$\frac{1}{4}$
2	$\frac{17}{60}$	$\frac{7}{15}$	$\frac{3}{4}$
$P(X=x_i)$	$\frac{1}{3}$	$\frac{2}{3}$	1

XY 的概率分布为

XY	0	1	2
p_k	$\frac{1}{3}$	$\frac{1}{5}$	$\frac{7}{15}$

X 和 Y 的协方差为

$$\begin{aligned}
 \text{cov}(X, Y) &= E(XY) - E(X)E(Y) \\
 &= 1 \times \frac{1}{5} + 2 \times \frac{7}{15} - \left(1 \times \frac{2}{3}\right) \left(1 \times \frac{1}{4} + 2 \times \frac{3}{4}\right) \\
 &= \frac{1}{5} + \frac{14}{15} - \frac{2}{3} \times \frac{7}{4} \\
 &= \frac{17}{15} - \frac{14}{12} = -\frac{1}{30}.
 \end{aligned}$$

$$2. \quad \text{解} \quad E(\eta_1) = aE(\xi_1) + bE(\xi_2) = 0, \quad E(\eta_2) = aE(\xi_1) - bE(\xi_2) = 0,$$

$$D(\eta_1) = D(\eta_2) = a^2 D(\xi_1) + b^2 D(\xi_2) = a^2 + b^2.$$

$$\begin{aligned} (1) \quad E(\eta_1^2 + \eta_2^2) &= E(\eta_1^2) + E(\eta_2^2) \\ &= D(\eta_1) + E^2(\eta_1) + D(\eta_2) + E^2(\eta_2) \\ &= 2(a^2 + b^2), \end{aligned}$$

$$E(\eta_1^2 - \eta_2^2) = E(\eta_1^2) - E(\eta_2^2) = 0.$$

$$\begin{aligned} (2) \quad \rho &= \frac{\text{cov}(2\eta_1, 3\eta_2)}{\sqrt{D(2\eta_1)} \sqrt{D(3\eta_2)}} = \frac{6\text{cov}(a\xi_1 + b\xi_2, a\xi_1 - b\xi_2)}{\sqrt{2^2 D(\eta_1)} \sqrt{3^2 D(\eta_2)}} \\ &= \frac{a^2 D(\xi_1) - b^2 D(\xi_2)}{a^2 + b^2} = \frac{a^2 - b^2}{a^2 + b^2}. \end{aligned}$$

$$3. \quad \text{解} \quad p(AB) = p(A) \cdot p\left(\frac{B}{A}\right) = \frac{1}{2} \times \frac{1}{4} = \frac{1}{8}, \quad p(\bar{A}) = \frac{3}{4}$$

$$p(AB) = p(B) \text{ 及 } p(B) = \frac{1}{2}, \quad p(\bar{B}) = \frac{1}{2}$$

$$p(\bar{A}\bar{B}) = 1 - p(A \cup B) = 1 - \frac{1}{4} - \frac{1}{2} + \frac{1}{8} = \frac{3}{8}$$

$$p(\bar{A}B) = p(B) \cdot p\left(\frac{\bar{A}}{B}\right) = \frac{1}{2} \times \frac{3}{4} = \frac{3}{8}$$

$$p(A\bar{B}) = p(A) \cdot p\left(\frac{\bar{B}}{A}\right) = \frac{1}{2} \times \frac{1}{4} = \frac{1}{8}$$

故 (ξ, η) 的联合分布律为

η	$\xi = 0$	$\xi = 1$	p_{oj}
0	3/8	1/8	4/8
1	3/8	1/8	4/8
p_{io}	6/8	2/8	

$$E(\xi) = \frac{2}{8} = \frac{1}{4}, \quad E(\eta) = \frac{4}{8} = \frac{1}{2}, \quad E(\xi\eta) = \frac{1}{8}$$

$$\text{cov}(\xi, \eta) = E(\xi\eta) - E(\xi) \cdot E(\eta) = \frac{1}{8} - \frac{1}{4} \times \frac{1}{2} = 0.$$

4. 证 $E(X)=E(\cos Z)$

$$= \cos\left(-\frac{\pi}{2}\right) \times 0.3 + \cos\frac{\pi}{2} \times 0.3 + \cos 0 \times 0.4 = 0.4.$$

$$E(Y)=E(\sin Z)=\sin\left(-\frac{\pi}{2}\right) \times 0.3 + \sin 0 \times 0.4 + \sin\frac{\pi}{2} \times 0.3 = 0.$$

$$D(X)=E(X^2)-[E(X)]^2=0.24, \quad D(Y)=0.6,$$

$$E(XY)=E[\cos Z \sin Z]=\frac{1}{2}E(\sin 2Z)=0.$$

$$\text{cov}(X, Y)=E(XY)-E(X) \cdot E(Y)=0.$$

所以 $\rho_{XY}=0$, 即 X, Y 不相关.

$$P\{X=1\}=P\{\cos Z=1\}=P\{Z=0\}=0.4,$$

$$P\{Y=1\}=P\{\sin Z=1\}=P\left\{Z=\frac{\pi}{2}\right\}=0.6,$$

$$P\{X=1\} \cdot P\{Y=1\}=0.24,$$

$$P\{X=1, Y=1\}=P\{\cos Z=1, \sin Z=1\}=0,$$

$$P\{X=1, Y=1\} \neq P\{X=1\} \cdot P\{Y=1\}.$$

故 X, Y 不相互独立.

一. 填空题

1. 答 0.5.

2. 答 $\frac{1}{12}$.

二. 计算题

1. 解 设各零件的重量为 $X_i (i=1, 2, \dots, 5000)$, 已知

$$\mu = E(X_i) = 0.5 \text{ kg}, \quad \sqrt{D(X_i)} = \sigma = 0.1 \text{ kg},$$

总重量 $Z = \sum_{i=1}^{5000} X_i$, 故所求概率为

$$\begin{aligned} P\{Z > 2510\} &= P\left\{\frac{Z - 5000 \times 0.5}{0.1\sqrt{5000}} > \frac{2510 - 5000 \times 0.5}{0.1\sqrt{5000}}\right\} \\ &\approx 1 - \Phi\left(\frac{10}{0.1\sqrt{5000}}\right) \\ &= 1 - \Phi(1.414) = 1 - 0.9214 = 0.0787. \end{aligned}$$

2. 解 (1) 因为 X 的分布律为

$$P\{X=x\} = p^x (1-p)^{1-x}, \quad x=0, 1,$$

所以 X_1, X_2, \dots, X_5 的联合概率密度为

$$\prod_{i=1}^5 P\{X=x_i\} = \prod_{i=1}^5 p^{x_i} (1-p)^{1-x_i} = p^{\sum_{i=1}^5 x_i} (1-p)^{5-\sum_{i=1}^5 x_i}.$$

(2) $X_1 + X_2$, $\max_{1 \leq i \leq 5} X_i$, $(X_5 - X_1)^2$ 都是统计量, $X_5 + 2p$ 不是统计量 (因 p 是未知参数).

3. **解** $X_1 - 2X_2 \sim N(0, 5)$, $3X_3 - 4X_4 \sim N(0, 25)$.

故 $\frac{1}{\sqrt{5}}(X_1 - 2X_2) \sim N(0, 1)$,

$$\frac{1}{\sqrt{5}}(3X_1 - 4X_2) \sim N(0, 1),$$

从而 $\frac{1}{5}(X_1 - 2X_2)^2 + \frac{1}{25}(3X_1 - 4X_2)^2 \sim \chi^2(2)$

应取 $a = \frac{1}{5}$, $b = \frac{1}{25}$,

此时 X 服从自由度为 2 的 χ^2 分布.

4. **解** 由于 $\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{X_i - \mu}{\sigma} \sim N(0, 1)$

$$\sum_{i=1}^n \left(\frac{Y_i - \bar{Y}}{\sigma} \right)^2 \sim \chi^2(n-1)$$

依题意 $\frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{X_i - \mu}{\sigma}$ 与 $\sum_{i=1}^n \left(\frac{Y_i - \bar{Y}}{\sigma} \right)^2$ 相互独立, 所以

$$\frac{\frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{X_i - \mu}{\sigma}}{\sqrt{\frac{\sum_{i=1}^n \left(\frac{Y_i - \bar{Y}}{\sigma} \right)^2}{n}}} = \frac{\sum_{i=1}^n \frac{X_i - \mu}{\sigma}}{\sqrt{\sum_{i=1}^n (Y_i - \bar{Y})^2}}$$

服从 $t(n-1)$ 分布.

习题十四 点估计

一. 计算题

1. **解** (1) 因为 $E(X) = \frac{a+8}{2}$, 得 $a = 2E(X) - 8$, 所以 $a = 2(\bar{X}) - 8$;

(2) 因为 $E(X) = \frac{3+b}{2}$, 得 $b = 2E(X) - 3$, 所以 $b = 2(\bar{X}) - 3$.

2. **解** $\because E(X) = \int_{-\infty}^{+\infty} x\varphi(x)dx = 0$,

\therefore 不能利用 $E(X)$ 构造 θ 的估计.

$$E(X^2) = \int_{-\infty}^{+\infty} x^2\varphi(x)dx = 2\theta^2$$

$$\therefore \theta^2 = \frac{1}{2}E(X^2)$$

$$\theta = \sqrt{\frac{1}{2}E(X^2)} = \sqrt{\frac{1}{2n} \sum_{i=1}^n x_i^2}.$$

3.

解
$$L(\sigma) = \frac{1}{(2\sigma)^n} \exp\left(-\frac{\sum_{i=1}^n |x_i|}{\sigma}\right)$$

$$\ln L(\sigma) = -n \ln(2\sigma) - \frac{\sum_{i=1}^n |x_i|}{\sigma} = -n \ln 2 - n \ln \sigma - \frac{\sum_{i=1}^n |x_i|}{\sigma}$$

$$\frac{\partial \ln L(\sigma)}{\partial \sigma} = -\frac{n}{\sigma} + \frac{\sum_{i=1}^n |x_i|}{\sigma^2} = 0$$

则
$$\hat{\sigma} = \frac{1}{n} \sum_{i=1}^n |x_i|.$$

4. **解** 令

$$E(a\theta_1 + 2a\theta_2) = \theta, \quad aE(\theta_1) + 2aE(\theta_2) = \theta$$

由题设 $E(\theta_1) = \theta, E(\theta_2) = \theta$, 所以有 $a\theta + 2a\theta = \theta$, 解得 $a = \frac{1}{3}$.

即应选择 $a = \frac{1}{3}$.

5. 解 $\because E(a_1) = \frac{a}{5} + \frac{3a}{10} + \frac{a}{2} = a, D(a_1) = \frac{1}{25} + \frac{9}{100} + \frac{1}{4} = 0.38;$
 $E(a_2) = \frac{a}{3} + \frac{a}{4} + \frac{5a}{12} = a, D(a_2) = \frac{1}{9} + \frac{1}{16} + \frac{1}{144} = 0.347;$
 $E(a_3) = \frac{a}{3} + \frac{a}{6} + \frac{a}{2} = a, D(a_3) = \frac{1}{9} + \frac{1}{36} + \frac{1}{4} = 0.389,$
 $\therefore a_1, a_2, a_3$ 均为 a 的无偏估计量, a_2 的方差最小.

6. 解 (1) $E(X) = \int_0^{+\infty} k \left(\frac{x}{b}\right)^k e^{-\left(\frac{x}{b}\right)^k} dx$

令 $\left(\frac{x}{b}\right)^k = u$, 则 $\frac{k}{b} \left(\frac{x}{b}\right)^{k-1} dx = du$,

于是 $E(X) = \int_0^{+\infty} b u^{\frac{1}{k}} e^{-u} du = b \Gamma\left(\frac{1}{k} + 1\right).$

其中 k 为已知.

以 \bar{x} 代 $E(X)$, \hat{b} 代 b , 得 b 的矩估计

$$\hat{b} = \frac{\bar{x}}{\Gamma\left(\frac{1}{k} + 1\right)}.$$

(2) 似然函数

$$L(b) = \left(\frac{k}{b}\right)^n \left(\frac{x_1, \dots, x_n}{b^n}\right)^{k-1} e^{-\sum_{i=1}^n \left(\frac{x_i}{b}\right)^k},$$

$$\begin{aligned} \ln L &= n \ln k - n \ln b + (k-1) \sum_{i=1}^n \ln x_i - (k-1)n \ln b - \sum_{i=1}^n \frac{x_i^k}{b^k} \frac{\partial \ln L}{\partial b} \\ &= -\frac{n}{b} - \frac{(k-1)n}{b} + \frac{1}{b^{k+1}} \sum_{i=1}^n x_i^k = 0. \end{aligned}$$

解和 b 的极大似然估计 $\hat{b} = \left(\frac{1}{n} \sum_{i=1}^n x_i^k\right)^{1/k}.$

习题十五 区间估计

一. 填空题

1. **答** $P\{\hat{\theta}_1 \leq \theta \leq \hat{\theta}_2\} = 0.95.$

2. **答** $(5.204, 36.667).$

$$n=10, S^2=11, \alpha=0.05, \chi_{0.025}^2(9)=19.023, \chi_{0.975}^2(9)=2.700,$$

因此
$$\left(\frac{(n-1)S^2}{\chi_{\frac{\alpha}{2}}^2(n-1)}, \frac{(n-1)S^2}{\chi_{1-\frac{\alpha}{2}}^2(n-1)} \right) = (5.204, 36.667).$$

3. **答** $\frac{4\sigma^2}{L^2} Z_{\frac{\alpha}{2}}^2.$

由 $2\frac{\sigma}{\sqrt{n}} Z_{\frac{\alpha}{2}} \leq L$, 得 $n \geq \left(\frac{2\sigma}{L} Z_{\frac{\alpha}{2}} \right)^2.$

二. 计算题

1. **解** $\bar{x}=1509.5, S=32.226, n=16, t_{0.025}(15)=2.1315.$

μ 的置信设为 0.95 的置信区间为

$$\begin{aligned} 1509.5 \pm \frac{32.226}{\sqrt{16}} \times 2.1315 &= 1509.5 \pm 17.172 \\ &= [1492.33, 1526.67]. \end{aligned}$$

2. **解** 如果总体期望 μ 未知时, 则 σ^2 的置信区间公式为

$$I = \left[\frac{1}{\lambda_2} \sum_{i=1}^{12} (x_i - \bar{x})^2, \frac{1}{\lambda_1} \sum_{i=1}^{12} (x_i - \bar{x})^2 \right]$$

由题知
$$\sum_{i=1}^{12} (x_i - \bar{x})^2 = \sum_{i=1}^{12} x_i^2 - 12\bar{x}^2 = 4.05$$

$$\lambda_1 = \chi_{0.95}^2(11) = 4.58, \lambda_2 = \chi_{0.05}^2(11) = 19.68$$

故所求的置信区间为 $[0.21, 0.88].$

3. **解** $\mu_1 - \mu_2 = \mu_d$ 因这是成对数据样本, 距离差 $d \sim N(\mu_d, \sigma_d^2)$, μ_d, σ_d^2 未知, 求 μ_d 的 99% 的置信区间.

由所给数据, 算得距离差样本均值为

$$\bar{d} = 1112.5, s_d = 1454.49.$$

选 t 分布随机变量

$$T = \frac{\bar{d} - \mu_d}{S_d / \sqrt{n}} \sim t(n-1),$$

得 $\mu_d = \mu_1 - \mu_2$ 的 $1-\alpha$ 的置信区间为

$$\left(\bar{d} \pm t_{\alpha/2}(n-1) \cdot \frac{S_d}{\sqrt{n}} \right).$$

由 $1-\alpha=0.99$, $\alpha=0.01$, $\frac{\alpha}{2}=0.005$, $n=8$, 查 t 分布表得

$$t_{0.005}(7) = 3.499,$$

于是 μ_d 的置信区间为

$$(1112.5 \pm 3.499 \times 1454.49 / \sqrt{8}), \text{ 即 } (-687, 2912).$$

4. **解** 由所给数据算出

$$\bar{x}_1 = 98.40, \bar{x}_2 = 110.71, s_1^2 = 8.73^2, s_2^2 = 32.19^2, n_1 = 5, n_2 = 7.$$

因为是求方差比的区间估计, 故选用 F 分布变量, 即

$$F = \frac{S_1^2}{S_2^2} \bigg/ \frac{\sigma_1^2}{\sigma_2^2} \sim F(n_1-1, n_2-1).$$

对于置信度 $1-\alpha$, 取双侧概率相等的置信区间为

$$\left(\frac{S_1^2}{S_2^2} \cdot \frac{1}{F_{\alpha/2}(n_1-1, n_2-1)}, \frac{S_1^2}{S_2^2} \cdot F_{\alpha/2}(n_1-1, n_2-1) \right).$$

本题所给 $1-\alpha=0.90$, $\alpha=0.10$, $\frac{\alpha}{2}=0.05$, $n_1=5$, $n_2=7$.

查 F 分布表得

$$F_{0.05}(6, 4) = 6.16, F_{0.05}(4, 6) = 4.53.$$

$$\frac{s_1^2}{s_2^2} = \frac{8.73^2}{32.19^2} = 0.0736.$$

于是 $\frac{\sigma_1^2}{\sigma_2^2}$ 的 0.90 的置信区间为

$$\left(0.0736 \times \frac{1}{4.53}, 0.0736 \times 6.16 \right), \text{ 即 } (0.016, 0.453).$$

习题十六 假设检验

一. 填空题

1. **答** $T = \frac{\bar{X}}{Q} \sqrt{n(n-1)}.$

2. **解** 填 $\frac{s_1^2}{\lambda s_2^2};$

$$\{F < F_{1-\alpha/2}(n_1-1, n_2-1) \text{ 或 } F > F_{\alpha/2}(n_1-1, n_2-1)\}.$$

$$F = \frac{s_1^2 \sigma_1^2}{s_2^2 \sigma_2^2} \sim F(n_1-1, n_2-1)$$

当 H_0 成立时 $\left(\frac{\sigma_1^2}{\sigma_2^2} = \lambda\right),$

$$F = \frac{s_1^2}{\lambda s_2^2} \sim F(n_1-1, n_2-1).$$

从而拒绝域 W 为上式所示.

3. **答** $-0.53.$

二. 计算题

1. **解** (1) $\bar{x} = 499, s = 16.031, n = 9.$

$$t = \frac{(\bar{x} - \mu_0)}{s} \sqrt{n} = \frac{999 - 500}{16.031} \sqrt{9} = 0.1871,$$

$$\alpha = 0.05, \quad t_{0.025}(8) = 2.306.$$

因 $|t| < t_{0.025}(8)$, 故接受 H_0 , 认为该天每袋平均质量可视为 500g.

$$(2) \quad \chi^2 = \frac{8 \times s^2}{\sigma_0^2} = \frac{8 \times 16.031^2}{10^2} = 21.258,$$

$$\alpha = 0.05, \quad \chi_{0.05}^2(8) = 15.507,$$

因 $\chi^2 > \chi_{0.05}^2(8) = 15.507$, 故拒绝 H_0 , 认为该天标准差超过 10g.

2. 解 (1) $H_0: \mu = \mu_0 = 15.25$, 这是个双侧检验问题; σ 已知, 故可

选统计量 $U = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}}$, 代入观察值得

$$U = \frac{15.06 - 15.25}{\sqrt{0.05}/\sqrt{6}} = -2.08.$$

相应的拒绝域为 $W_1: \{U/|U| \geq u_{1-\frac{\alpha}{2}} = U_{0.975} = 1.96\}$.

而由于 $U = -2.08$ 落入拒绝域 W_1 中, 故在显著水平 $\alpha = 0.05$ 下, 拒绝 H_0 .

(2) 该问题 $H_0: \mu \leq \mu_0 = 15.25$, 这是个单侧检验问题; 所选统计量 U 以及 U 值与上题相同; 但拒绝域为:

$$W_1 = (U_{1-\alpha}, +\infty) = (U_{0.95}, +\infty) = (1.65, +\infty).$$

注意到 U 值不落入 W_1 , 故不拒绝 H_0 , 即接受 $H_0: \mu \leq 15.25$.

3. 解 作假设 $H_0: \sigma_1 = \sigma_2$.

$$\frac{1}{50-1} \sum_{i=1}^{50} (x_i - \bar{x})^2 = \frac{50 \times s_1^2}{49} = \frac{50 \times 0.0139}{49} = 0.0142 \text{ (大)}$$

$$\frac{1}{50-1} \sum_{i=1}^{52} (y_i - \bar{y})^2 = \frac{52 \times s_2^2}{51} = \frac{52 \times 0.0053}{51} = 0.0054 \text{ (小)}$$

代入统计量得 $F = \frac{0.0142}{0.0054} = 2.63$, 查 F 表得

$$F_{\frac{\alpha}{2}}(50-1, 52-1) = F_{0.05}(49, 51) = 1.59,$$

$$F_{\frac{\alpha}{2}}(50-1, 52-1) = F_{0.01}(49, 51) = 1.93,$$

故 $F_{\frac{\alpha}{2}} = F_{0.025} = \frac{1}{2}(F_{0.05} + F_{0.01}) = \frac{1}{2}(1.59 + 1.93) = 1.76$

由 $F = 2.63 > 1.76 = F_{\frac{\alpha}{2}}$, 故假设 $H_0: \sigma_1 = \sigma_2$ 被否定, 即甲、乙两

段岩心磁化率测定数据的标准差在 $\alpha = 5\%$ 时有显著差异.

4. **解** 这问题即是在 $\alpha = 0.05$ 下检验两正态总体均值有无显著差, 应该先检验方差齐性 $H_0: \sigma_1^2 = \sigma_2^2; H_1: \sigma_1^2 \neq \sigma_2^2$,

$$n_1 = 10, \mathbf{x}_1 = 9.4, (n_1 - 1)s_1^2 = 40.4, s_1^2 = 4.4889,$$

$$n_2 = 8, \mathbf{x}_2 = 8.125, (n_2 - 1)s_2^2 = 36.88, s_2^2 = 5.2686.$$

$$F = \frac{s_2^2}{s_1^2} = 1.1737 < 4.20 = F_{0.975}(7, 9) = F_{1-\frac{\alpha}{2}}(n_2 - 1, n_1 - 1),$$

故接受 H_0 , 认为两正态总体具有方差齐性. 再检验

$$H'_0: \mu_1 - \mu_2 = 0; H'_1: \mu_1 - \mu_2 \neq 0.$$

$$\begin{aligned} t &= \frac{\mathbf{x}_1 - \mathbf{x}_2}{\sqrt{\frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{n_1 + n_2 - 2}} \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \\ &= \frac{9.4 - 8.125}{\sqrt{\frac{40.4 + 36.88}{16}} \sqrt{\frac{1}{10} + \frac{1}{8}}} = 1.247, \end{aligned}$$

由于 $|t| = 1.247 < 2.12 = t_{0.975}(16) = t_{1-\frac{\alpha}{2}}(n_1 + n_2 - 2)$,

故接受 H'_0 , 即认为两个班组生产的导线的平均电阻无显著差异.