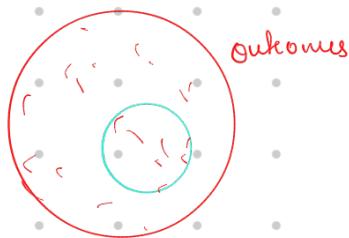


4th Sept

★ Recap

$$\Pr[X=a] \equiv \Pr[\text{outcomes } \omega : X(\omega)=a]$$



$$E[X] = \sum_a a \Pr[X=a]$$

↳ To simplify these eqⁿ we break them into simpler random variable.

i.e. $X = X_1 + X_2 + \dots + X_n$

$E[X] = E[X_1] + E[X_2] + \dots + E[X_n]$

↳ Linearity of Expectation.

* Proof of Linearity of Expectation

↳ prove for smaller expression

i.e. if $X = Y+Z$

$$E[X] = E[Y] + E[Z]$$

we can use this to prove for multiple variables.

i.e. $X = Y+Z+W$

$$X = Y+X'$$

$$= E[X] = E[Y] + E[X']$$

$$= E[X] = E[Y] + E[Z] + E[W]$$

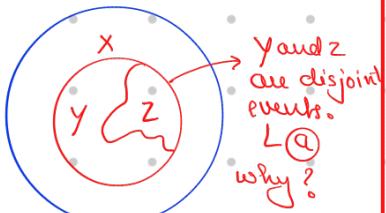
Proof

$$\text{LHS} = E[X]$$

$$\sum_a a \Pr[X=a] \Rightarrow \sum_a [\Pr\{Y+Z=a\}]$$

↳ for $Y+Z$ to be equal to a only if $y=b$, then $Z=a-b$

$$[Y+Z=a] = \bigcup_b [Y=b, Z=a-b].$$



by ④ we know that probability of union of disjoint events is sum of all events

i.e. $\sum_a a \left(\sum_{b,c} \Pr[Y=b, Z=c] \right)$

$$= \sum_a a \left(\sum_{b,c} \Pr[Y=b, Z=c] \right) \because a \geq b+c$$

our eqⁿ is of form

$$\sum_a \sum_{b,c} f(b,c)$$

which is equivalent to

$\sum_{b,c} f(b,c) \because$ we are basically finding and summing all possible pairs of b, c which equal to some a .

$$= \sum_{b,c} [\Pr[Y=b, Z=c] \cdot f(b,c)]$$

$$= \sum_{b,c} b \underbrace{\Pr[Y=b, Z=c]}_{\textcircled{b}} + \sum_{b,c} c \underbrace{\Pr[Y=b, Z=c]}_{\textcircled{c}}$$

we can say that

$$[Y=b] = \bigcup_c [Y=b, Z=c].$$

$$\Pr[Y=b] = \sum_c \Pr[Y=b, Z=c]. \rightarrow \textcircled{d}$$

↳ Saying as we toss a coin and ask what is the probability of first coin toss is heads.

↳ $\Pr[\text{first heads}] = \sum \Pr[\text{first head, other heads}]$

$$\Pr[Y=b] = \Pr[Y=b, Z=H] + \Pr[Y=b, Z=T]$$

by using \textcircled{b}

$$\sum_b b \underbrace{\Pr[Y=b, Z=c]}_{\textcircled{b}}$$

$$\sum_b b \cdot \Pr[Y=b] \rightarrow \text{by } \textcircled{d}$$

Similarly for \textcircled{c} we get $\sum_c c \cdot \Pr[Z=c]$

hence

$$E[X] = \sum_b b \cdot \Pr[Y=b] + \sum_c c \cdot \Pr[Z=c]$$

$$E[X] = E[Y] + E[Z] \text{ hence proved.}$$

Prove: $\mathbb{E}X = \mathbb{E}Y \cdot \mathbb{E}Z$

$$\hookrightarrow \sum_a \sum_{\substack{b,c \\ b+c=a}} \Pr[Y=b, Z=c] \quad \text{using the similar steps as before just we get } b+c=a$$

but now we cannot decompose this but suppose we could, then.

$$\Rightarrow \Pr[Y=b] \cdot \Pr[Z=c] \quad \text{(a)}$$

$$\Rightarrow \sum_{b,c} \Pr[Y=b] \Pr[Z=c] \quad (\text{b}, \text{c})$$

$$\Rightarrow \sum_{b,c} b \Pr[Y=b] \cdot c \Pr[Z=c] \quad \downarrow \text{how is it possible}$$

$$\Rightarrow \sum_b b \Pr[Y=b] \cdot \sum_c \Pr[Z=c] \cdot c \quad \begin{aligned} &\text{assume } \sum_{i=1}^n y_i \\ &= (x_1, y_1) + \\ &(x_2, y_2) + \\ &(x_3, y_3) + \\ &(x_4, y_4) \\ &\Rightarrow (x_1+x_2)(y_1+y_2) \end{aligned}$$

$$\Rightarrow \mathbb{E}Y \cdot \mathbb{E}Z.$$

for this to happen step (a) must be possible but in general it is not possible.

Example

$$X = \begin{cases} 1 & \text{if H} \\ 0 & \text{if T} \end{cases}$$

$Y = X, Z = X$ is $X = YZ$?

↳ yes for this binary ex.

by $\mathbb{E}X = \mathbb{E}Y$.

$$\text{so, } \mathbb{E}X = \mathbb{E}Y \cdot \mathbb{E}Z$$

$$= \frac{1}{2} = \frac{1}{2} \cdot \frac{1}{2}$$

$\Rightarrow Y = YZ$ which is not possible

hence this only works if Y and Z are independent i.e.

if Y and Z are independent.

$$\Pr[Y=b, Z=c] = \Pr[Y=b] \cdot \Pr[Z=c].$$

for sum, this is true regardless.

Example

(1) Toss 2 coins

$$X = \begin{cases} 1 & \text{if 1st coin is H} \\ 0 & \text{otherwise} \end{cases} \quad Y = \begin{cases} 1 & \text{if 2nd coin is H} \\ 0 & \text{otherwise} \end{cases}$$

$$\Pr[X=1, Y=1] = \frac{1}{4} = \Pr[X=H] \cdot \Pr[Y=H].$$

$\therefore X$ and Y are independent.

(2) when throwing a dice

$$X = \begin{cases} 1 & \text{if number is even} \\ 0 & \text{otherwise} \end{cases}$$

$$Y = \begin{cases} 1 & \text{number is greater than 3} \\ 0 & \text{otherwise} \end{cases}$$

X and Y are not independent.

i.e.

$$\Pr[X=1, Y=1] \neq \Pr[X=\text{even}] \cdot \Pr[Y=\text{gt 3}].$$

(3) Toss n coins

$X = \# \text{ heads}$

$$X = X_1 + X_2 + \dots + X_n$$

$$X_i = \begin{cases} 1 & \text{if } i^{\text{th}} \text{ coin toss is H} \\ 0 & \text{otherwise} \end{cases}$$

$$\mathbb{E}X = \mathbb{E}X_1 + \mathbb{E}X_2 + \dots + \mathbb{E}X_n$$

$$\mathbb{E}X = \frac{1}{2} + \frac{1}{2} + \dots + \frac{1}{2} = n \cdot \frac{1}{2}.$$

(4) n letters in n envelopes

$X = \# \text{ number of letter that go in correct envelope}$

$$X = X_1 + X_2 + \dots + X_n$$

$$X_i = \begin{cases} 1 & \text{if } i^{\text{th}} \text{ letter is in correct envelope} \\ 0 & \text{otherwise.} \end{cases}$$

$$\Pr[X_i=1] = \frac{1}{n} \quad \begin{aligned} &\text{equally likely to go in any} \\ &\text{envelope} \end{aligned}$$

$$\text{i.e. } \mathbb{E}X_i = \frac{1}{n}.$$

$$\mathbb{E}X = \mathbb{E}X_1 + \mathbb{E}X_2 + \dots + \mathbb{E}X_n$$

$$= n \cdot \frac{1}{n}$$

$$\boxed{\mathbb{E}X = 1}$$

⑤ we are given an array of length n with distinct elements and we want to find the max

i.e. $\text{Max} = A[0]$
for $i \geq 1$ to $n-1$
if $A[i] > \text{max}$
 $\text{max} = A[i]$ ⑥

6 In worst case we will execute ⑥ $n-1$ times.

6 Start with a random permutation of $1, \dots, n$

$X = \# \text{ of times } ⑥ \text{ is executed.}$

what's $\mathbb{E}X$.

$X = X_1 + X_2 + \dots + X_{n-1}$ not independent because
if max is at j the every $i \leq j$ there is a selection.
 $X_i = \begin{cases} 1 & \text{if } ⑥ \text{ is executed in } i^{\text{th}} \text{ iteration} \\ 0 & \text{otherwise.} \end{cases}$

$$\mathbb{E}X = \mathbb{E}X_1 + \mathbb{E}X_2 + \dots + \mathbb{E}X_{n-1}$$

$$\mathbb{E}X_i = \Pr[X_i = 1]. 1$$

↳ we will execute ⑥ at i^{th} iteration iff

$$A[i] > A[0] \text{ to } A[i-1]$$

$$\text{i.e. } \Pr[X_i = 1] = \frac{1}{i}$$

$$\mathbb{E}X = 1 + \frac{1}{2} + \frac{1}{3} + \dots + \frac{1}{n-1} \approx \log n.$$

$$\boxed{\mathbb{E}X \approx \log n}$$

⑥ Random Walk.

$X = \text{Position after } n \text{ steps.}$

$$X = Y_1 + Y_2 + \dots + Y_n. \quad \text{Independent.}$$

$Y_i = \begin{cases} 1 & \text{if move right on } i^{\text{th}} \text{ step.} \\ -1 & \text{if move left on } i^{\text{th}} \text{ step.} \end{cases}$

$$\mathbb{E}X_n = \mathbb{E}Y_1 + \mathbb{E}Y_2 + \dots + \mathbb{E}Y_n.$$

$$\mathbb{E}Y_i = 0 \quad \because 1 = Y_2 + (-1) \cdot Y_2 \quad \text{— ⑦}$$

$$\mathbb{E}X = 0.$$

now, if I want to calculate the distance
we will need to ignore the sign
i.e. $\mathbb{E}(X_n^2)$

$$\text{Let } X_n = Y_1 + Y_2 + \dots + Y_n.$$

$$X_n^2 = (Y_1 + Y_2 + \dots + Y_n)^2$$

$$= X_n^2 = Y_1^2 + Y_2^2 + \dots + Y_n^2 + 2 \sum_{i \neq j} Y_i Y_j$$

$$\mathbb{E}(X_n^2) = \mathbb{E}(Y_1^2) + \mathbb{E}(Y_2^2) + \dots + 2 \sum_{i \neq j} \mathbb{E}(Y_i Y_j)$$

$$\mathbb{E}(Y_i Y_j) = 1.$$

$\mathbb{E}(Y_i Y_j) = \mathbb{E}Y_i \cdot \mathbb{E}Y_j \quad \because \text{they are independent}$
and $\mathbb{E}Y_i, \mathbb{E}Y_j = 0$
by ⑦

i.e.

$$\mathbb{E}(X_n^2) = 1 + 1 + \dots + n \text{ times} + 0$$

$$\boxed{\mathbb{E}X_n^2 = n}$$

⑦ Randomize Quicksort.

① pick a random element

② divide the array in 3 parts. 

③ Sort both pieces.

Every pair is going to be compared at most once
i.e. running time is at most n^2 .

↓ remaining.