Charitable Giving, Tax Reform, and Self-selection of Tax Report: Evidence from South Korea

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Introduction

- ▶ 世界の多くの国で寄付に対する税制上の優遇がなされている。
- ▶ 税制優遇が厚生に与える影響について、寄付の価格弾力性が厚生評価のための重要な指標であると考えられている。(e.g. Saez, 2004)
 - ▶ 直感的に言えば寄付の価格弾力性が-1を下回ると、\$1の税制優遇が \$1以上の寄付を生み出すこととなる。
 - ▶ 「寄付の税制優遇で寄付の価格が下がり、寄付が促進されるだろう」
- ▶ 既存研究の多くは課税申告データを用いて寄付の価格弾力性を推定してきた。(e.g. Almunia et al., 2020, Auten et al., 2002, and so on)

Introduction

- ▶ しかし、課税申告データは申告された寄付のみしか記録していない。
 - この論文の扱う第1の問題: 実際の寄付量と申告された寄付量は異なるものであり、推定バイアスが起きる (Fack and Landais (2016);
 Gillitzer and Skov (2018))。
 - ▶ この問題に対応するため、課税申告データではなく、韓国のパネルデータを使って分析を行った。
- ▶ 納税者は寄付を申告するかどうか自ら選択でき、申告する際には寄付証明などの提出などのコストがかかる。
 - ▶ この論文の扱う第2の問題:寄付の申告コストが無視されると推定バイアスが起きる。
 - ← 寄付の申告をしなければ税制優遇は受けられない
 - ► この論文では操作変数法 (IV) と Control Function (CF) アプローチを つかってこの問題に対処。
- ▶ 差の差法 (DID) をメインの識別戦略として用いて、寄付の価格弾力性の大きさについて韓国のデータを用いて調べた。

Introduction

分析の結果

- 1. ベースラインの結果では、寄付の価格弾力性が Intensive Margins で-1.4 以下、Extensive Margins で-1.7 以下となった。
 - ▶ 寄付の申告コストを考慮しない結果では Intensive Margins で-0.9 となった。
- 2. 寄付の申告者に絞った分析結果では寄付の価格弾力性は Intensive Margins で-1.2~-1.6 となった。

既存研究の多くでは Intensive Margins の寄付の価格弾力性が約-1 であり、これよりも概ね弾力的な推定結果が得られた。

- (Not in the paper) The estimated declaration cost of giving is KRW
 (\$).
- 4. (Not in the paper) Given our estimates, increasing the subsidy on charitable giving will be desirable in Korea.

In Korea, income tax payers can receive tax relief for their charitable giving.

- For the application of tax relief, tax payers have to submit a certificate for charitable giving.
- Wage earners pay their income tax by withholding tax and declare their charitable giving via their company.
 - Wage earners can submit the certificate at any time.
- Non wage earners, such as the self-employed, pay their income tax by tax-return and declare their charitable giving via the National Tax Service.
 - Non wage earners have to retain the certificate until they submit tax return.

Our major price variation comes from the 2014 tax reform.

- ▶ Before 2014, tax deduction (所得控除) was used for tax relief on charitable giving.
 - ▶ I.e. the giving price depended on income level.
- ► After 2014, tax credit (税額控除) started to be used for tax relief on charitable giving.
 - ► The tax credit rate was determined as 15%
 - ► Giving price is 0.85, irrespective of income level.

Model

- ▶ Consider private consumption (x_i) and charitable giving (g_i) .
- ► The budget constraint is

$$x_i + g_i = y_i - R_i K - R_i T(y_i, g_i) - (1 - R_i) T(y_i)$$

where y_i is pre-tax total income, R_i is a dummy of declaration of tax relief and $T(y_i)$ and $T(y_i, g_i)$ are respectively the amount of tax when i does not declare tax relief and when i declares tax relief.

Tax payers declare their charitable giving if its benefit exceeds its cost.

$$R_{i} = \begin{cases} 1 \text{ if } T(y_{i}, g_{i}) - T(y_{i}) > K \\ 0 \text{ if } T(y_{i}, g_{i}) - T(y_{i}) \leq K. \end{cases}$$
 (1)

Tax deduction system (until 2013)

$$T(y_i,g_i)=T(y_i-g_i)$$

- ▶ In 2012 and 2013, the marginal tax rate was the same, though it was different from ones before 2011.
- ▶ The logged relative giving price is $R_i \ln(1 T'(y_i g_i))$.

Tax credit system (from 2014)

$$T(y_i,g_i)=T(y_i)-R_img_i$$

- ightharpoonup m is tax credit rate and is m = 0.15.
- ▶ The logged relative giving price is $R_i \ln(1 0.15) = R_i \ln 0.85$.

Note: The logged relative giving price for the non-declared is $\ln 1 = 0$.

Source of endogeneity

- 1. Usage of tax return data only captures declared charitable giving.
 - If the charitable giving is not declared, tax relief has a little effect.
 - Some papers use survey data to deal with this problem (e.g. Rehavi and Shack, 2013).
 - Following them, we use survey panel data of Korea.
- 2. If the declaration cost is ignored, the estimation should be biased.
 - ► The giving price depends not only on marginal tax rate and tax credit rate, but also on the declaration behavior.
 - As far as we know, only Almunia et al. (2020) deal with this problem, though they used tax return data.
 - We use the different declaration cost btw wage earners and the others as an instrumental variable (IV).

We use the Korean annual financial panel survey, called the National Survey of Tax and Benefit (hereafter, NaSTab).

- ► The subjects of this survey are general households and household members living in 15 cities and provinces nationwide.
- ▶ This survey is based on a face-to-face interview.
- Data is constructed as the subjects represent the population of Korean society.
- ▶ We exclude the subject of the sample, whose age is under 23, since they are not likely to have income or assets.
- ▶ We use data from 2013 to 2017.

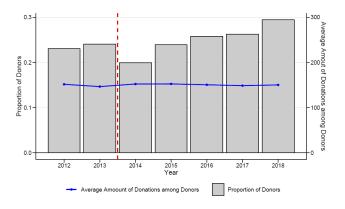


Figure: Proportion of Donors and Average Donations among Donors

- ▶ About $20\sim30\%$ of people make a donation.
- ▶ The average amount of donations among donors is about 1.5 million KRW.

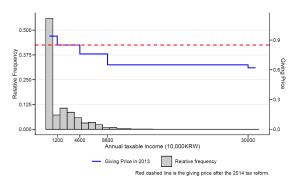


Figure: Income Distribution and Relative Giving Price in 2013

In 2014, relative giving price

- decreases for people whose income is less than 12 million KRW.
- is the same for people whose income is between 12 million KRW and 46 million KRW.
- increase for people whose income is more than 46 million KRW.



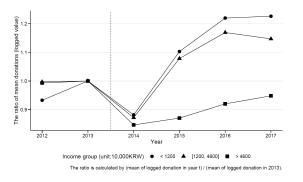


Figure: Average Logged Giving in Three Income Groups

Compared to 2012 and 2013, the amount of charitable giving after 2014

- increases for people whose income is less than 12 million KRW.
- relatively increases for people whose income is between 12 million KRW and 46 million KRW.
- decreases for people whose income is more than 46 million KRW.



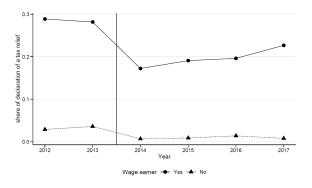


Figure: Share of Declaration of Tax Relief

- Wage earners are more likely to declare their charitable giving to receive tax relief.
 - \rightarrow This reflects the difference of the declaration cost.

	N	Mean	Std.Dev.	Min	Median	Max
Charitable Donations						
Annual chariatable giving (unit: 10,000KRW)	40064	36.64	153.72	0.00	0.00	10000.00
Dummary of donation > 0	40064	0.24	0.43	0.00	0.00	1.00
Income, giving price, and tax report						
Annual taxable labor income (unit: 10,000KRW)	40054	1674.04	2733.18	0.00	0.00	91772.00
First giving relative price	40063	0.86	0.04	0.62	0.85	0.94
Dummy of declaration of a tax relief	40064	0.11	0.31	0.00	0.00	1.00
Individual Characteristics						
Age	40064	54.20	16.31	24.00	52.00	104.00
Wage earner dummy	29753	0.54	0.50	0.00	1.00	1.00

Table: Summary Statistics

Statistical Model

- ▶ Main identification source is tax reform in 2014. (DID-like strategy)
 - ▶ Below 12 million KRW: Giving price decreases.
 - ▶ Btw 12 and 46 million KRW: Giving price is the same.
 - Above 46 million KRW: Giving price increases.
- ► To capture the difference of declaration cost, we use a dummy to show whether a subject is a wage earner or not as IV.

Statistical Model

▶ We estimate the following two-way fixed effect model:

$$\ln g_{it} = \varepsilon_p R_{it} \ln p_{it}(y_{it}, g_{it}) + \varepsilon_y \ln y_{it} + \mathbf{X}_{it} \beta + \mu_i + \iota_t + u_{it}$$
 (5)

where μ_i , ι_t and u_{it} are an individual fixed effect, a year fixed effect, and a error term, respectively.

- X_{it} is a vector of covariates including square of age, industry dummy, and area dummy.
- ► In the literature, estimations in terms of **intensive** and **extensive** margins are common.
 - ▶ **Intensive margins**: estimate (5) only for $g_{it} > 0$.
 - **Extensive margins**: estimate (5) but dependent variable is $1[g_{it} > 0]$.
 - \rightarrow Following the literature, we estimate both of them.

Statistical Model: Endogeneity of p_{it}

▶ The giving price (compared to private good) is

$$p_{it}(y_{it}, g_{it}) = \begin{cases} 1 - T'_t(y_{it} - g_{it}) & \text{if } t < 2014 \\ 0.85 & \text{if } t \ge 2014 \end{cases}$$
 (6')

➤ Since the giving price is endogenous to the amount of giving before 2014, we use "the first-price of giving", which is defined as

$$p_{it}^f(y_{it}) = p_{it}(y_{it}, 0)$$

instead of $p_{it}(y_{it}, 0)$ in the estimation.

 $ightharpoonup p_{it}(y_{it},0)$ is called as "the last-price of giving".

Statistical Model: Endogeneity of R_{it}

- Since donors can choose whether they declare charitable giving or not, declaration, R_{it} , is endogenous.
- ► To overcome this, we estimate

$$\ln g_{it} = \varepsilon_p R_{it} \ln p_{it}^f(y_{it}) + \varepsilon_y \ln y_{it} + \mathbf{X}_{it} \beta + \mu_i + \iota_t + u_{it}$$
 (7)

where $R_{it} \ln p_{it}^f(y_{it})$ is instrumented by $WageEarner_{it} \times \ln p_{it}^f(y_{it})$.

- This estimation is based on 2SLS.
- WageEarner_{it} is a dummy to show whether a subject is a wage earner or not.
 - WageEarner_{it} should not correlate to u_{it} when we control incomes and industry dummies.

Statistical Model: Endogeneity of R_{it}

- ▶ In alternative models, we use propensity score to declare $P(Z_{it})$ as an instrument.
 - ► This estimation method is called "control function (CF)" approach.
 - ▶ The propensity score is estimated by a probit model

$$R_{it} = 1[\delta_0 + Z_{it}\delta_1 + u_{it1} > 0],$$
 (8)

where $Z_{it} \equiv \{WageEarner_{it}, \ln p_{it}^f(y_{it}), \ln y_{it}, \mathbf{X}_{it}\}.$

- ► The propensity score is $P(Z_{it}) = Φ(\hat{\delta}_0 + Z_{it}\hat{\delta}_1)$.
- We consider two cases:
 - ① pooled probit model (δ_0, δ_1) is constant.
 - ② separated probit model (δ_0, δ_1) can vary by year.
- ▶ In addition, we also estimate the following by OLS:

$$\ln g_{it} = \varepsilon_p P(Z_{it}) \ln p_{it}^f(y_{it}) + \varepsilon_y \ln y_{it} + \boldsymbol{X}_{it} \boldsymbol{\beta} + \mu_i + \iota_t + u_{it}$$
 (9)

Results: Intensive Margins

	FE-2SLS			OLS		
	(1)	(2)	(3)	(4)	(5)	
Applying tax relief x log(first price)	-1.429***	-1.506***	-1.598***			
	(0.398)	(0.354)	(0.361)			
PS of applying tax relief x log(first price)				-1.584***	-1.564***	
				(0.371)	(0.353)	
log(income)	1.162	1.102	1.030	1.037	1.013	
	(1.112)	(1.084)	(1.085)	(1.110)	(1.116)	
Num.Obs.	7080	7080	7080	7080	7080	
R2	0.820	0.820	0.820	0.820	0.820	
R2 Adj.	0.693	0.693	0.693	0.693	0.694	
FE: area	X	X	X	X	X	
FE: industry	X	X	X	X	X	
FE: panelid	X	X	X	X	X	
FE: year	X	X	X	X	X	
Square of age	X	X	X	X	X	
Instrument	Wage earner x Price	PS x Price	PS x Price			
Method of PS	-	Pool	Separate	Pool	Separate	

^{*} p < 0.1, ** p < 0.05, *** p < 0.01

Table: First-Price Elasticities (Intensive Margins)

The estimated giving price elasticity in terms of intensive margins is about -1.5.

Results: Extensive Margins

	FE-2SLS			OLS		
	(1)	(2)	(3)	(4)	(5)	
Applying tax relief x log(first price)	-0.445***	-0.509***	-0.710***			
	(0.172)	(0.124)	(0.113)			
PS of applying tax relief x log(first price)				-0.416***	-0.546***	
				(0.111)	(0.097)	
log(income)	2.105***	2.074***	1.975***	1.955***	1.832***	
	(0.280)	(0.263)	(0.257)	(0.281)	(0.279)	
Implied price elasticity	-1.863***	-2.129***	-2.975***	-1.743***	-2.286***	
	(0.721)	(0.518)	(0.475)	(0.465)	(0.407)	
Num.Obs.	26922	26922	26922	26922	26922	
R2	0.679	0.681	0.687	0.663	0.663	
R2 Adj.	0.569	0.572	0.580	0.547	0.547	
FE: area	X	X	X	X	X	
FE: industry	X	X	X	X	X	
FE: panelid	X	X	X	X	X	
FE: year	X	X	X	X	X	
Square of age	X	X	X	X	X	
Instrument	Wage earner x Price	PS x Price	PS x Price			
Method of PS		Pool	Separate	Pool	Separate	

^{*} p < 0.1, ** p < 0.05, *** p < 0.01

Table: First-Price Elasticities (Extensive Margins)

The estimated giving price elasticity in terms of extensive margins is about -1.7 \sim -2.9.



Results

- ► The results show that the giving price elasticity in Korea is more elastic than many papers in the literature show.
 - ▶ In the literature, the giving price elasticity in terms of intensive margins is typically -1.
- ► The elasticity in terms of extensive margins captures the behavior of whether people donate or not.
 - The elastic extensive margins elasticity shows that reducing giving price induce people to donate.
 - ▶ IV を使わない場合の推定結果が得られるならば、ここで IV を使わない推定結果と比べたときの含意を書く。

Results

We try several robustness checks.

- Estimation excluding 2013 and 2014 data: to eliminate announcement effect.
- 2. Estimation using the last-price
- Estimation using the subsample of those who have applied for tax relief
 - We use Semykina and Wooldridge (2010)'s way of the correction of sample selection bias.
 - ► Estimated elasticity (intensive margins) is around -1.2~-1.6.
 - Subsample analysis enables us to use several methods to deal with issues related to tax deduction system.
 - e.g. fluctuation of income level (Randolph, 1995, and so on.)
 - \rightarrow usage of k-th difference model / lead and lag.

Most of result shows that the giving price elasticity is less than

- ▶ -1.4 in terms of intensive margins and
- ▶ -1.7 in terms of extensive margins.

Following Almunia et al. (2020), we can derive the welfare implication by specifying the following utility maximization.

$$\max_{x_i, g_i, R_i} U(x_i, g_i, G) = x_i - R_i K + \theta u(g_i) + V(G)$$
s.t. $x_i + g_i = R_i (y - T(y, g_i)) + (1 - R_i)(y - T(y)),$
and $G = g_i + G_{-i},$

where $\theta \in [\underline{\theta}, \overline{\theta}]$ is the parameter to show the preference for donation, which follows the density $f(\cdot)$.

- G is the total donation (including the governmental provision).
- $ightharpoonup G_{-i}$ is the total donation except *i*.
- ightharpoonup For simplicity, assume tax schedule is now linear and tax rate is au.

▶ Then, the optimization will be

$$\max_{g_i,R_i}(1-\tau)y-R_i\rho g_i-(1-R_i)g_i-R_iK+\theta u(g_i)+V(G).$$

- ▶ Denote $g(p; \theta)$ and $g(1; \theta)$ as the donation when $R_i = 1$ and 0.
- As a result of the optimization, the indirect utility of those who declare giving and do not will respectively be

$$\nu(p;\theta) = \theta_i u(g(p;\theta)) - pg(p;\theta)$$

$$\nu(1;\theta) = \theta_i u(g(1;\theta)) - g(1;\theta).$$

- **Depends** on θ and giving price p, three types of individuals exist.
 - 1. Non-donor: $\theta \leq \theta_0$
 - 2. Donor but non-declarer: $\theta \in (\theta_0, \theta(p)]$
 - ightarrow Denote their total donation as $g^0(p) \equiv \int_{ heta_0}^{ heta(p)} g(1; heta) f(heta) d heta.$
 - 3. Donor and declarer: $\theta > \theta(p)$
 - ightarrow Denote their total donation as $g^1(p) \equiv \int_{\theta(q)}^{\bar{\theta}} g(p;\theta) f(\theta) d\theta$.
- Social welfare can be written as

$$egin{aligned} W &= V(G) + \int_{ heta(
ho)}^{ar{ heta}} (
u(
ho; heta) - K) f(heta) d heta \ &+ \int_{ heta_0}^{ heta(
ho)}
u(1; heta) f(heta) d heta + \lambda [ty - (1-
ho)g^1(
ho) - G_g] \end{aligned}$$

where λ is the marginal cost of public finance.

Using $\lambda = V'$ (Saez, 2004), the effect of changing giving price p on the welfare W can be shown as

$$\frac{dW}{dp} = \lambda(g_p^0 + g_p^1) + (\lambda - 1)g^1 - \lambda(1 - p)g_p^1.$$

 $ightharpoonup \frac{dW}{dp} < 0$ is equivalent to

$$\epsilon \equiv -rac{pg_p^1}{g^1} > rac{\lambda-1}{\lambda} + rac{g_p^0}{g^1}.$$

- From the subsample analysis, ϵ is 1.304 \sim 1.603.
- Assuming $\lambda \in [1,2]$, $\frac{\lambda-1}{\lambda} \in [0,\frac{1}{2}]$.
- From the data, $\frac{g_p^0}{g^1}$ is 0.00003.
- Our result suggests that more generous tax relief will increase the welfare in Korea.