BART MCMC updates

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Overall algorithm details

The equations referred to below are displayed later in this document.

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Algorithm 1: BART Markov chain Monte Carlo
 Data: Target variables y (length n; standardised), feature matrix X (n rows and p columns
 Result: Posterior list of trees T, values of \tau, fitted values \hat{y}
 Initialisation;
 Hyper-parameter values of \alpha, \beta, \tau_{\mu}, \nu, \lambda;
 Number of trees M;
 Number of iterations N;
 Initial value \tau = 1;
 Set trees T_j; j = 1, ..., M to stumps;
 Set values of \mu to 0;
 for iterations i from 1 to N do
     for trees j from 1 to M do
         Compute partial residuals from y minus predictions of all trees except tree j;
         Grow a new tree T_i^{new} based on grow/prune/change/swap;
         Set l_{new} = \log full conditional of new tree T_j^{new} based on Equation 1 plus Equation 4;
         Set l_{old} = \log full conditional of old tree T_j based on same equations;
         Set a = \exp(l_{new} - l_{old});
         Generate u \sim U(0,1);
         if a > u then
            Set T_j = T_j^{new};
         end
         Simulate \mu values using Equation 3;
     end
     Get predictions \hat{y} from all trees;
     Update \tau using Equation 4;
 \quad \text{end} \quad
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1 Updating trees

Suppose there are n observations in a terminal node and suppose that the (partial) residuals in this terminal node are denoted R_1, \ldots, R_n . The prior distribution for these residuals is:

$$R_1, \ldots, R_n | \mu, \tau \sim N(\mu, \tau^{-1})$$

Furthermore the prior on μ is:

$$\mu \sim N(0, \tau_{\mu}^{-1})$$

Using π to denote a probability distribution, we want to find:

$$\begin{split} \pi(R_1,\dots,R_n|\tau) &= \int \pi(R_1,\dots,R_n|\mu,\tau)\pi(\mu)\partial\mu \\ &\propto \int \prod_{i=1}^n \tau^{1/2} e^{-\frac{\tau}{2}(R_i-\mu)^2} \tau_\mu^{1/2} e^{-\frac{\tau_\mu}{2}\mu^2} \partial\mu \\ &= \int \tau^{n/2} e^{-\frac{\tau}{2}\sum (R_i-\mu)^2} \tau_\mu^{1/2} e^{-\frac{\tau_\mu}{2}\mu^2} \partial\mu \\ &= \int \tau^{n/2} \tau_\mu^{1/2} e^{-\frac{1}{2}\left[\tau\left\{\sum R_i^2 + n\mu^2 - 2\mu n\bar{R}\right\} + \tau_\mu \mu^2\right]} \partial\mu \\ &= \tau^{n/2} \tau_\mu^{1/2} e^{-\frac{1}{2}\left[\tau\sum R_i^2\right]} \int e^{-\frac{1}{2}Q} \partial\mu \end{split}$$

where

$$Q = \tau n \mu^2 - 2\tau n \mu \bar{R} + \tau_{\mu} \mu^2$$

$$= (\tau_{\mu} + n\tau) \mu^2 - 2\tau n \mu \bar{R}$$

$$= (\tau_{\mu} + n\tau) \left[\mu^2 - \frac{2\tau n \mu \bar{R}}{\tau_{\mu} + n\tau} \right]$$

$$= (\tau_{\mu} + n\tau) \left[\left(\mu - \frac{2\tau n \bar{R}}{\tau_{\mu} + n\tau} \right)^2 - \left(\frac{\tau n \bar{R}}{\tau_{\mu} + n\tau} \right)^2 \right]$$

$$= (\tau_{\mu} + n\tau) \left(\mu - \frac{2\tau n \bar{R}}{\tau_{\mu} + n\tau} \right)^2 - \frac{(n\tau \bar{R})^2}{\tau_{\mu} + n\tau}$$

so therefore:

$$\int e^{-\frac{1}{2}Q} \partial \mu = \int \exp\left[-\frac{\tau_{\mu} + n\tau}{2} \left(\mu - \frac{2\tau n\bar{R}}{\tau_{\mu} + n\tau}\right)^{2} + \frac{(n\tau\bar{R})^{2}}{2(\tau_{\mu} + n\tau)}\right] \partial \mu$$

$$\propto \exp\left[\frac{1}{2} \frac{(\tau n\bar{R})^{2}}{\tau_{\mu} + n\tau}\right] (\tau_{\mu} + n\tau)^{-1/2}$$

And finally:

$$\pi(R_1, \dots, R_n | \tau) \propto (\tau_\mu + n\tau)^{-1/2} \tau^{n/2} \tau_\mu^{1/2} \exp\left[\frac{1}{2} \frac{(\tau n \bar{R})^2}{\tau_\mu + n\tau}\right] \exp\left[-\frac{\tau}{2} \sum R_i^2\right]$$
$$= \tau^{n/2} \left(\frac{\tau_\mu}{\tau_\mu + n\tau}\right)^{1/2} \exp\left[-\frac{\tau}{2} \left\{\sum R_i^2 - \frac{\tau(n\bar{R})^2}{\tau_\mu + n\tau}\right\}\right]$$

Including multiple terminal nodes

When we put back in terminal nodes we write R_{ji} where j is the terminal node and i is still the observation, so in terminal node j we have partial residuals R_{j1}, \ldots, R_{jn_j} . When we have $j = 1, \ldots, b$ terminal nodes the full conditional distribution is then:

$$\prod_{j=1}^{b} \pi(R_{j1}, \dots, R_{jn_j} | \tau) \propto \prod_{j=1}^{b} \left\{ \tau^{n_j/2} \left(\frac{\tau_{\mu}}{\tau_{\mu} + n_j \tau} \right)^{1/2} \exp \left[-\frac{\tau}{2} \left\{ \sum_{i=1}^{n_j} R_{ji}^2 - \frac{\tau(n_j \bar{R}_j)^2}{\tau_{\mu} + n_j \tau} \right\} \right] \right\}$$

which on the log scale gives:

$$\sum_{j=1}^{b} \left\{ \frac{n_j}{2} \log(\tau) + \frac{1}{2} \log \left(\frac{\tau_{\mu}}{\tau_{\mu} + n_j \tau} \right) - \frac{\tau}{2} \left[\sum_{i=1}^{n_j} R_{ji}^2 - \frac{\tau (n_j \bar{R}_j)^2}{\tau_{\mu} + n_j \tau} \right] \right\}$$

This can be simplified further to give:

$$\frac{n}{2}\log(\tau) + \frac{1}{2}\sum_{j=1}^{b}\log\left(\frac{\tau_{\mu}}{\tau_{\mu} + n_{j}\tau}\right) - \frac{\tau}{2}\sum_{j=1}^{b}\sum_{i=1}^{n_{j}}R_{ji}^{2} + \frac{\tau^{2}}{2}\sum_{j=1}^{b}\frac{S_{j}^{2}}{\tau_{\mu} + n_{j}\tau}$$
(1)

where $S_j = \sum_{i=1}^{n_j} R_{ji}$

Updating μ

The full conditional for μ_j (the terminal node parameters for node j) is similar to the above but without the integration:

$$\pi(\mu_j|\dots) \propto \prod_{i=1}^{n_j} \tau^{1/2} e^{-\frac{\tau}{2}(R_{ji} - \mu_j)^2} \tau_{\mu}^{1/2} e^{-\frac{\tau_{\mu}}{2}\mu_j^2}$$

$$\propto e^{-\frac{\tau}{2} \sum_{i=1}^{n_j} (R_{ji} - \mu_j)^2} e^{-\frac{\tau_{\mu}}{2}\mu_j^2}$$

$$\propto e^{-\frac{\tau}{2} \left[n_j \mu_j^2 - 2\mu_j \sum_{i=1}^{n_j} R_{ji}\right] - \frac{\tau_{\mu}}{2}\mu_j^2}$$

$$\propto e^{-\frac{Q}{2}}$$

Now:

$$\begin{split} Q &= n_j \tau \mu_j^2 - 2\mu_j \tau S_j + \tau_\mu \mu_j^2 \\ &= (n_j \tau + \tau_\mu) \mu_j^2 - 2\tau \mu_j S_j \\ &= (n_j \tau + \tau_\mu) \left[\mu_j^2 - \frac{2\tau \mu_j S_j}{n_j \tau + \tau_\mu} \right] \\ &\propto (n_j \tau + \tau_\mu) \left[\mu_j - \frac{\tau \mu_j S_j}{n_j \tau + \tau_\mu} \right]^2 \end{split}$$

so therefore:

$$\mu_j | \dots \sim N\left(\frac{\tau S_j}{n_j \tau + \tau_\mu}, (n_j \tau + \tau_\mu)^{-1}\right)$$
 (2)

Update for τ

I am using the shape/rate parameterisation of the gamma with prior $\tau \sim Ga(\nu/2, \nu\lambda/2)$. Letting μ_i be the prediction of the *i*th observation we get:

$$\pi(\tau|\ldots) \propto \prod_{i=1}^{n} \tau^{1/2} e^{-\frac{\tau}{2}(y_i - \mu_i)^2} \tau^{\nu/2 - 1} e^{-\tau\nu\lambda/2}$$

Letting $S = \sum_{i=1}^{n} (y_i - \mu_i)^2$ we get:

$$\pi(\tau|\ldots) \propto \tau^{n/2} e^{-\frac{\tau}{2}S} \tau^{\nu/2-1} e^{-\tau\nu\lambda/2}$$
$$= \tau^{(n+\nu)/2-1} e^{-\frac{\tau}{2}(S+\nu\lambda)}$$

so

$$\tau | \dots \sim Ga\left(\frac{n+\nu}{2} - 1, \frac{S+\nu\lambda}{2}\right)$$
 (3)

Tree prior

The tree prior used by BARTMachine says that the probability of a node being non-terminal is:

$$P(\text{node is non-terminal}) = \alpha(1+d)^{-\beta}$$

So the probability of a node being terminal is 1 minus this. A stump just has probability 1 - α . For Bart Machine $\alpha = 0.95$ and $\beta = 2$

Thus for a tree with k non-terminal nodes and b terminal nodes we have:

$$P = \prod_{i=1}^{b} \left[1 - \alpha (1 + d_i^t)^{-\beta} \right] \prod_{i=1}^{k} \left[\alpha (1 + d_i^{nt})^{-\beta} \right]$$

where d_i^t is the depth of the *i*th terminal node and d_i^{nt} is the depth of the *i*th non-terminal node. On the log scale this gives:

$$\log P = \sum_{i=1}^{b} \left[\log \left(1 - \alpha (1 + d_i^t)^{-\beta} \right) \right] + \sum_{i=1}^{b} \left[\log(\alpha) - \beta \log(1 + d_i^{nt}) \right]$$
 (4)

Mean and variance changing

An alternative, slightly more flexible, model can be created by giving each terminal node it's own precision value. Thus for tree j we have:

$$R_{1j}, \dots, R_{jn_j} | \mu_j, \tau_j \sim N(\mu_j, \tau_j^{-1})$$

The maths is much simplified by setting the prior on μ_i as:

$$\mu_j \sim N(0, (a\tau)^{-1}).$$

Chipman et al (1998) set a = 1/3 for a single tree setting, so perhaps a = 1/(3m) might be more appropriate for a multi-tree version with m trees.

We now have a prior for each terminal node:

$$\tau_i \sim Ga(\nu/2, \nu\lambda/2)$$

Chipman et al (1998) suggest making ν and λ functions of tree complexity but we don't do that here.

We don't need the subscripts j for an individual tree so the shortcut to what we require is:

$$\pi(R_1, \dots, R_n | \dots) = \int \int \pi(R_1, \dots, R_n | \mu, \tau) \pi(\mu) \pi(\tau) \partial \mu \, \partial \tau$$

$$\propto \int \int \left[\prod_{i=1}^n \tau^{1/2} \exp\left(-\frac{\tau}{2} (R_i - \mu)^2\right) \right] \tau^{1/2} \exp\left(-\frac{a\tau}{2} \mu^2\right) \tau^{\nu/2 - 1} \exp\left(-\frac{\tau \nu \lambda}{2}\right) \partial \mu \, \partial \tau$$

$$\propto \int \int \tau^{n/2} \exp\left(-\frac{\tau}{2} \sum (R_i - \mu)^2\right) \tau^{1/2} \exp\left(-\frac{a\tau}{2} \mu^2\right) \tau^{\nu/2 - 1} \exp\left(-\frac{\tau \nu \lambda}{2}\right) \partial \mu \, \partial \tau$$

Note that $\sum (R_i - \mu)^2$ can be re-written as $SS_R + n(\mu - \bar{R})^2$ where $SS_R = \sum (R_i - \bar{R})^2$.

We now get

$$\pi(R_1,\ldots,R_n|\ldots) \propto \int \int \tau^{(n+\nu-1)/2} \exp\left(-\frac{\tau}{2}\left\{SS_R + n(\mu-\bar{R})^2 + a\mu^2 + \nu\lambda\right\}\right) \partial\mu \,\partial\tau$$

Simplifying the exponent to separate out the terms with μ gives:

$$n(\mu - \bar{R})^2 + a\mu^2 = (a+n)\left[\mu + \frac{n\bar{R}}{n+a}\right]^2 + \frac{an\bar{R}^2}{n+a}$$

So we can perform the integrand with respect to μ first:

$$\pi(R_1, \dots, R_n | \dots) \propto \int \tau^{(n+\nu-2)/2} \exp\left(-\frac{\tau}{2} \left\{ SS_R + \nu\lambda + \frac{an\bar{R}^2}{n+a} \right\} \right) \int \tau^{1/2} \exp\left(-\frac{\tau(a+n)}{2} \left[\mu + \frac{n\bar{R}}{n+a}\right]^2\right) \partial\mu \, \partial\tau$$

The integrand wrt μ is proportional to $(a+n)^{-1/2}$ and also provides the full conditional for μ (putting back in the subscripts):

$$\mu_j | \dots \sim N \left(\frac{n_j \bar{R}_j}{n_j + a}, \left[\tau_j (a + n_j) \right]^{-1} \right)$$

Next we have:

$$\pi(R_1, \dots, R_n | \dots) \propto \int \frac{\tau^{(n+\nu-2)/2}}{(a+n)^{1/2}} \exp\left(-\frac{\tau}{2} \left\{ SS_R + \nu\lambda + \frac{an\bar{R}^2}{n+a} \right\} \right) \partial \tau$$

This also provides the complete conditional for (re-inserting subscripts) τ_i :

$$|\tau_j| \ldots \sim Ga\left(\frac{n_j + \nu}{2}, \frac{1}{2} \left[SS_{R_j} + \nu\lambda + \frac{an_j \bar{R}_j^2}{n_j + a} \right] \right)$$

Finally we have (again with subscripts):

$$\pi(R_{1j},\ldots,R_{jn_j}|\ldots) \propto (a+n_j)^{-1/2} \Gamma\left(\frac{n_j+\nu}{2}\right) \left[SS_{R_j}+\nu\lambda+\frac{an_j\bar{R}_j^2}{n_j+a}\right]^{-\left(\frac{n_j+\nu}{2}\right)}$$

With multiple terminal nodes for a tree we have:

$$\prod_{j=1}^{b} \pi(R_{1j}, \dots, R_{jn_j} | \dots) = \prod_{j=1}^{b} (a + n_j)^{-1/2} \Gamma\left(\frac{n_j + \nu}{2}\right) \left[SS_{R_j} + \nu\lambda - \frac{an_j \bar{R}_j^2}{n_j + a}\right]^{-(\frac{n_j + \nu}{2})}$$

$$= \left[\prod_{j=1}^{b} (a + n_j)^{-1/2}\right] \left[\prod_{j=1}^{b} \Gamma\left(\frac{n_j + \nu}{2}\right)\right] \left\{\prod_{j=1}^{b} \left[SS_{R_j} + \nu\lambda + \frac{an_j \bar{R}_j^2}{n_j + a}\right]^{-(\frac{n_j + \nu}{2})}\right\}$$

which on the log scale is:

$$\log \prod_{j=1}^{b} \pi_{j} = -\frac{1}{2} \sum_{j=1}^{b} \log(a + n_{j}) + \sum_{j=1}^{b} \log \Gamma\left(\frac{n_{j} + \nu}{2}\right) - \sum_{j=1}^{b} \left(\frac{n_{j} + \nu}{2} \log \left[SS_{R_{j}} + \nu\lambda + \frac{an_{j}\bar{R}_{j}^{2}}{n_{j} + a}\right]\right)$$