MG-RAST Nested Data

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Have some cheesecake.

1 DP Data Mining – Nice tool to have

For $i = 1, \dots, n$,

$$Y_{i} \equiv \{X_{i1}, X_{i2}, \cdots, X_{iK}\}$$

$$Y_{i} \sim \text{Multinomial}(N_{i}, p_{i})$$

$$X_{ik} \in \mathbb{N}^{+} \cup \{0\}$$

$$N_{i} = \sum_{k=1}^{K} X_{ik}$$

Model

$$p_i \sim \operatorname{Dirichlet}(\theta_i)$$
 $\theta_i \sim G$
 $G \sim \operatorname{DP}(\alpha G_0)$
 $\alpha \sim \operatorname{Gamma}(a, b)$
 $G_0 \equiv \operatorname{Normal}_+(\mu, \Sigma)$

Or if that's not quite appropriate for modeling the p_i , model, for example,

$$p_i \sim \operatorname{Dirichlet}(\theta + O_i)$$
 $O_i \sim G$
 $G \sim \operatorname{DP}(\alpha G_0)$
 $\alpha \sim \operatorname{Gamma}(a, b)$
 $G_0 \equiv \pi \delta_0 + (1 - \pi) \operatorname{Normal}(\mu, \Sigma)$

Choices for G_0 are important. There's lots of data to do this in an empirical fashion. Inference on θ locates interesting pies Y_i (clusters, perhaps). This could be applied across any single hierarchical data level. And then I don't have to look through this data for interesting things myself.

2 Three level nested data – DP hierarchical modeling

For $i = 1, \dots, n$,

$$Y_{i}^{(1)} \equiv \{X_{i1}^{(1)}, X_{i2}^{(1)}, \cdots, X_{iQ^{(1)}}^{(1)}\}$$

$$Y_{i}^{(1)} \sim \text{Multinomial}(N_{i}, p_{i}^{(1)})$$

$$X_{ij}^{(1)} \in \mathbb{N}^{+} \cup \{0\}$$

$$N_{i} = \sum_{j=1}^{Q^{(1)}} X_{ij}^{(1)}$$

$$Y_{ij}^{(2)} \equiv \{X_{ij1}^{(2)}, X_{ij2}^{(2)}, \cdots, X_{ijQ^{(2)}}^{(2)}\}$$

$$Y_{ij}^{(2)} \sim \text{Multinomial}(X_{ij}^{(1)}, p_{ij}^{(2)})$$

$$X_{ijk}^{(2)} \in \mathbb{N}^+ \cup \{0\}$$

$$X_{ij}^{(1)} = \sum_{k=1}^{Q^{(2)}} X_{ijk}^{(2)}$$

$$Y_{ijk}^{(3)} \equiv \{X_{ijk1}^{(3)}, X_{ijk2}^{(3)}, \cdots, X_{ijkQ^{(3)}}^{(3)}\}$$

$$Y_{ijk}^{(3)} \sim \text{Multinomial}(X_{ijk}^{(2)}, p_{ijk}^{(3)})$$

$$X_{ijkl}^{(3)} \in \mathbb{N}^+ \cup \{0\}$$

$$X_{ijk}^{(2)} = \sum_{l=1}^{Q^{(3)}} X_{ijkl}^{(3)}$$

This is a three level nested data structure:

- Individual i has a 1^{st} level multinomial random variable, $Y_i^{(1)}$.
- The j^{th} category in $Y_i^{(1)}$ (with $X_{ij}^{(1)}$ counts) may itself be subdivided into a multinomial random variable, $Y_{ij}^{(2)}$.

- The k^{th} category in $Y_{ij}^{(2)}$ (with $X_{ijk}^{(2)}$ counts) may itself be subdivided into a multinomial random variable, $Y_{ijk}^{(3)}$.
- The l^{th} category in $Y_{ijk}^{(3)}$ will have $X_{ijkl}^{(2)}$ counts.

Consider just the top two levels (ignore the bottom third level), and model

- The DPs cluster the *i*'s element wise on the A adjustments.
- Each of the elements in the A adjustment become a new multinomial random variable. The clusters get the same $\theta^{(2)}$.
- The extension from the second to the third level would be analogous to that of the extension from the first to the second.
- I view this as some kind of bubbling process. When there's a cluster in a top level category, those in the cluster grow a separate parameter in the next level.
- Writing this out is all well and good... fitting this thing is going to be something else.

• I also am not sure if a difference between a category at one level means that there should be a difference in the next level breakdown of that category... but that's what we're doing here.