
Bayesian and Empirical Bayesian Forests

Matt Taddy

University of Chicago Booth School of Business

TADDY@CHICAGOBOOTH.EDU

Chun-Sheng Chen

eBay

CHUNSCHEN@EBAY.COM

Jun Yun

eBay

JUNYU@EBAY.COM

Mitch Wyle

eBay

MWYLE@EBAY.COM

Abstract

We derive ensembles of decision trees through a nonparametric Bayesian model, allowing us to view random forests as samples from a posterior distribution. This insight provides large gains in interpretability, and motivates a class of Bayesian forest (BF) algorithms that yield small but reliable performance gains. Based on the BF framework, we are able to show that high-level tree hierarchy is stable in large samples. This leads to an empirical Bayesian forest (EBF) algorithm for building approximate BFs on massive distributed datasets and we show that EBFs outperform sub-sampling based alternatives by a large margin.

1. Introduction

Decision trees are a fundamental machine learning tool. They partition the feature (input) space into regions of response homogeneity, such that the response (output) value associated with any point in a given partition can be predicted from the average for that of its neighbors. The classification and regression tree (CART) algorithm of (Breiman et al., 1984) is a common recipe for building trees; it grows greedily through a series of partitions on features, each of which maximizes reduction in some measure of impurity at the current tree leaves (terminal nodes; i.e., the implied input space partitioning). The development of random forests (RF) by (Breiman, 2001), which predict through the average of many CART trees fit to bootstrap data resamples,

is an archetype for the successful strategy of tree ensemble learning. For prediction problems with training sets that are large relative to the number of inputs, properly trained ensembles of trees can predict out-of-the-box as well as any carefully tuned, application-specific alternative.

This article makes three contributions to understanding and application of decision tree ensembles (or, *forests*).

Bayesian forest: A nonparametric Bayesian (npB) point-of-view allows interpretation of forests as a sample from a posterior over trees. Imagine CART applied to a data generating process (DGP) with finite support: the tree greedily splits support to minimize impurity of the partitioned response distributions (terminating at some minimum-leaf-probability threshold). We present a nonparametric Bayesian model for DGPs based on multinomial draws from (large) finite support, and derive the Bayesian forest (BF) algorithm for sampling from the distribution of CART trees implied by the posterior over DGPs. Random forests are an approximation to this BF exact posterior sampler, and we show in examples that BFs provide a small but reliable gain in predictive performance over RFs.

Posterior tree variability: Based upon this npB framework, we derive results on the stability of CART over different DGP realizations. We find that, conditional on the data allocated to a given node on the sample CART tree, the probability that the next split for a posterior DGP realization matches the observed full-sample CART split is

$$p(\text{split matches sample CART}) \gtrsim 1 - \frac{p}{\sqrt{n}} e^{-n}, \quad (1)$$

where p is the number of possible split locations and n the number of observations on the node. Even if p grows with n , the result indicates that partitioning can be stable conditional on the data being split. This conditioning is key:

CART’s well known instability is due to its recursive nature, such that a single split different from sample CART at some node removes any expectation of similarity below that node. However, for large samples, (1) implies that we will see little variation at the top hierarchy of trees in a forest. We illustrate such stability in our examples.

Empirical Bayesian forests: the npB forest interpretation and tree-stability results lead us to propose empirical Bayesian forests (EBF) as an algorithm for building approximate BFs on massive distributed datasets. Traditional empirical Bayesian analysis fixes parameters in high levels of a hierarchical model at their marginal posterior mode, and quantifies uncertainty for the rest of the model conditional upon these fixed estimates. EBFs work the same way: we fit a single shallow CART *trunk* to the sample data, and then sample a BF ensemble of *branches* at each terminal node of this trunk. The initial CART trunk thus maps observations to their branch, and each branch BF is fit in parallel without any communication with the other branches. With little posterior variability about the trunk structure, an EBF sample should look similar to the (much more costly, or even infeasible) full BF sample. In a number of experiments, we compare EBFs to the common distributed-computing strategy of fitting forests to data subsamples and find that the EBFs lead to a large improvement in predictive performance. This type of strategy is key to efficient machine learning with Big Data: focus the ‘Big’ on the pieces of models that are most difficult to learn.

Bayesian forests are introduced in Section 2 along with a survey of Bayesian tree models, Section 3 investigates tree stability in theory and practice, and Section 4 presents the empirical Bayesian forest framework. Throughout, we use publicly available data on home prices in California to illustrate our ideas. We also provide a variety of other data analyses to benchmark performance, and close with description of how EBF algorithms are being built and perform in large-scale machine learning at eBay.com.

2. Bayesian forests

Informally, write dgp to represent the stochastic process defined over a set of possible DGPs. A Bayesian analogue to classical ‘distribution-free’ nonparametric statistical analysis (e.g., [Hollander & Wolfe, 1999](#)) has two components:

1. set a nonparametric statistic $\mathcal{T}(dgp)$ that is of interest in your application regardless of the true DGP,
2. and build a flexible model for the DGP, so that the posterior distribution on $\mathcal{T}(dgp)$ can be derived from posterior distribution on possible DGPs.

In the context of this article, $\mathcal{T}(dgp)$ refers to a CART tree. Indeed, trees are useful precisely because they are good

predictors regardless of the underlying data distribution – they do not rely upon distributional assumptions to share information across training observations. Our DGP model, detailed below, leads to a posterior for dgp that is represented through random weighting of observed support. A Bayesian forest contains CART fits corresponding to each draw of support weights, and the BF ensemble prediction is an approximate posterior mean.

2.1. Nonparametric model for the DGP

We employ a Dirichlet-multinomial sampling model in nonparametric Bayesian analysis. The approach dates back to [Ferguson \(1973\)](#). [Chamberlain & Imbens \(2003\)](#) provide an overview in the context of econometric problems. [Rubin \(1981\)](#) proposed the Bayesian bootstrap as an algorithm for sampling from versions of the implied posterior, and it has since become closely associated with this model.

Use $\mathbf{z}_i = \{\mathbf{x}_i, y_i\}$ to denote the features and response for observation i . We suppose that data are drawn *independently* from a finite L possible values,

$$dgp = p(\mathbf{z}) = \sum_{l=1}^L \omega_l \mathbb{1}_{[\mathbf{z}=\zeta_l]} \quad (2)$$

where $\omega_l \geq 0 \forall l$ and $\sum_l \omega_l = 1$. Thus the generating process for observation i draws l_i from a multinomial with probability ω_{l_i} , and this indexes one of the L support points. Since L can be arbitrarily large, this so-far implies no restrictive assumptions beyond that of independence.

The conjugate prior for ω is a Dirichlet distribution, written $\text{Dir}(\omega; \nu) \propto \prod_{l=1}^L \omega_l^{\nu_l-1}$. We will parametrize the prior with a single concentration parameter $\nu = a > 0$, such that $\mathbb{E}[\omega_l] = a/La = 1/L$ and $\text{var}(\omega_l) = (L-1)/[L^2(La+1)]$. Suppose you have the observed sample $\mathbf{Z} = [\mathbf{z}_1 \cdots \mathbf{z}_n]'$. For convenience, we allow $\zeta_l = \zeta_k$ for $l \neq k$ in the case of repeated values. Write $l_1 \dots l_n = 1 \dots n$ so that $\mathbf{z}_i = \zeta_{l_i}$ and $\mathbf{Z} = [\zeta_{l_1} \cdots \zeta_{l_n}]'$. Then the posterior distribution for ω has $\omega_i = a + 1$ for $i \leq n$ and $\omega_l = a$ for $l > n$, so that

$$p(\omega) \propto \prod_{i=1}^n \omega_{l_i}^a \prod_{l=n+1}^L \omega_l^{a-1}. \quad (3)$$

This, in turn, defines our posterior for the data generating process through our sampling model in (2).

There are many possible strategies for specification of a and ζ_l for $l > n$.¹ The non-informative prior that arises as $a \rightarrow 0$ is a convenient default: in this limit, $\omega_l = 0$ with probability one for $l > n$.² We apply this limiting

¹The unobserved ζ_l act as data we imagine we might have seen, to smooth the posterior away from the data we have actually observed. See [Poirier \(2011\)](#) for discussion.

²For $l > n$ the posterior has $\mathbb{E}[\omega_l] = 0$ with variance $\text{var}(\omega_l) = \lim_{a \rightarrow 0} a[n+a(L-1)]/[(n+La)^2(n+La+1)] = 0$.

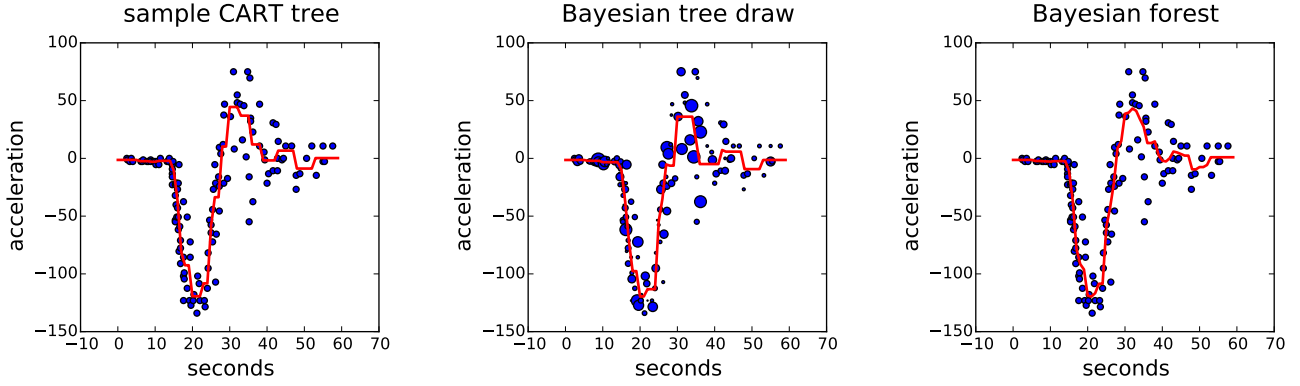


Figure 1. Motorcycle data illustration. The lines show each of CART fit to the data sample, CART fit to a sample with weights drawn IID from an exponential (point size proportional to weight), and the BF predictor which averages over 100 weighted CART fits.

prior throughout, such that our posterior for the data generating process is a multinomial draw from the observed data points, with a uniform $\text{Dir}(\mathbf{1})$ distribution on the $\omega = [\omega_1 \dots \omega_n]'$ sampling probabilities. We will also find it convenient to parametrize un-normalized ω via IID exponential random variables: $\theta = [\theta_1 \dots \theta_n]$, where $\theta_i \stackrel{\text{iid}}{\sim} \text{Exp}(1)$ in the posterior and $\omega_i = \theta_i / |\theta|$ with $|\theta| = \sum_i \theta_i$.

2.2. CART as posterior functional

Conditional upon θ , the population tree $\mathcal{T}(\text{dgp})$ is defined through a weighted-sample CART fit. In particular, given data $\mathbf{Z}^\eta = \{\mathbf{X}^\eta, \mathbf{y}^\eta\}$ in node η , sort through all dimensions of all observations in \mathbf{Z}^η to find the split that minimizes the average of some ω -weighted impurity metric across the two new child nodes. For example, in the case of regression trees, the impurity to minimize is weighted-squared error

$$\mathcal{I}(\mathbf{y}^\eta) = \sum_{i \in \eta} \theta_i (y_i - \mu_\eta)^2 \quad (4)$$

with $\mu_\eta = \sum_i \theta_i y_i / |\theta^\eta|$. For our classification trees, we minimize Gini impurity. The split candidates are restricted to satisfy a minimum leaf probability, such that every node in the tree must have $|\theta^\eta|$ greater than some threshold.³ This procedure is repeated on every currently-terminal tree node until it is no longer possible to split while satisfying the minimum probability threshold. To simplify notation, we refer to the resulting CART tree as $\mathcal{T}(\theta)$.

2.3. Posterior sampling

Following (Rubin, 1981) we can sample from the posterior on $\mathcal{T}(\theta)$ via the Bayesian bootstrap in Algorithm 1.

We’ve implemented BF through simple adjustment of the ensemble module of python’s `scikit-learn` (Pe-

³In practice this can be replaced with a threshold on the minimum number of observations at each leaf.

Algorithm 1 Bayesian Forest

```

for  $b = 1$  to  $B$  do
  draw  $\theta^b \stackrel{\text{iid}}{\sim} \text{Exp}(\mathbf{1})$ 
  run weighted-sample CART to get  $\mathcal{T}_b = \mathcal{T}(\theta^b)$ 
end for
    
```

dregosa et al., 2011).⁴ As a quick illustration, Figure 1 shows three fits for the conditional mean velocity of a motorcycle helmet after impact in a crash: sample CART $\mathcal{T}(\mathbf{1})$, a single draw of $\mathcal{T}(\theta)$, and the BF average prediction (data are from the MASS R package, Venables & Ripley, 2002).

Note that the Bayesian forest differs from Breiman’s random forest only in that the weights are drawn from an exponential (or Dirichlet, when normalized) distribution rather than a Poisson (or multinomial) distribution. To the extent that RF sampling provides a coarse approximation to the BF samples, the former is a convenient approximation. Moreover, we will find little difference in predictive performance between BFs and RFs, so that one should feel free to use readily available RF software while still relying on the ideas of this paper for intuition and interpretation.

2.4. Bayesian tree-as-parameter models

Other Bayesian analyses of DTs treat the tree as a parameter which governs the DGP, rather than a functional thereof, and thus place some set of restrictions on the distributional relationship between inputs and outputs.

The original Bayesian tree model is the Bayesian CART (BCART) of Chipman et al. (1998). BCART defines a likelihood where response values for observations allocated (via \mathbf{x}) to the same leaf node are IID from some parametric family (e.g., for regression trees, observations in

⁴Replace the variable `sample_counts` in `forest.py` to be drawn exponential rather than binomial when bootstrapping.

each leaf are IID Gaussian with shared variance and mean). The BCART authors also propose linear regression leaves, while the Treed Gaussian Processes of Gramacy & Lee (2008) use Gaussian process regression at each leaf. The models are fit via Markov chain Monte Carlo (above examples) or sequential Monte Carlo (Taddy et al., 2011) algorithms that draw from the posterior by proposing small changes to the tree (e.g., grow or prune).⁵

Monte Carlo tree sampling use the same incremental moves that employed in CART. Unfortunately, this means that they tend to get stuck in locally-optimal regions of tree-space. Bayesian Additive Regression Trees (BART; Chipman et al., 2010) replace a single tree with the sum of many small trees. An input vector is allocated to a leaf in each tree, and the corresponding response distribution has mean equal to the average of each leaf node value and variance shared across observations. Original BART has Gaussian errors, and extensions include Gaussian mixtures. Since BART only samples short trees, it is fast and mixes well.

Easy sampling comes at a potentially steep price: the assumption of homoskedastic additive errors.⁶ Despite this restriction, empirical studies have repeatedly shown BART to outperform alternative flexible prediction rules. Many response variables (especially after, say, log transformation) have the property that they are well fit by flexible regression with homoskedastic errors. Whenever the model assumptions in BART are close enough to true, it should outperform methods which do not make those assumptions.

In contrast, the npB interpretation of forests (BF or RF) makes it clear that they are suited to applications where the response distribution defies parametric representation, such that CART fit is the most useful DGP summary available. We often encounter this situation in application. For example, internet transaction data combines discrete point masses with an incredibly fat right tail (e.g., see Taddy et al., 2014). In academia it is common to transform such data before analysis, but businesses wish to target the response on the scale measured (e.g., clicks or dollars) and need to build a predictor that does well on that scale.

2.5. Friedman example

A common simulation experiment in evaluating flexible predictors is based around the Friedman (1991) function,

$$y = f(\mathbf{x}) + \varepsilon \quad (5)$$

$$= 10\sin(\pi x_1 x_2) + 20(x_3 - 0.5)^2 + 10x_4 + 5x_5 + \varepsilon.$$

where $\varepsilon \sim N(0, 1)$ and $x_j \sim U(0, 1)$.

⁵The Bayesian bootstrap is also a potential sampling tool in this tree-as-parameter setting. See Clyde & Lee (2001) for details on the technique and its relation to model averaging.

⁶For classification, this is manifest through a probit link.

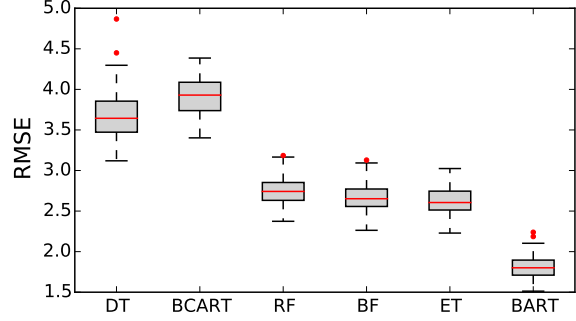


Figure 2. Friedman experiment predictive RMSE over 100 runs.

For our experiment, we follow previous authors by including as features for training the spurious $x_6 \dots x_p$. Each regression models are fit to 100 random draws from (5) and tested at 1000 new \mathbf{x} locations. Root mean square error (RMSE) is calculated between predicted and true $f(\mathbf{x})$.

Results over 100 repeats are shown in Figure 5.⁷ As forecast, the only model which assumes the true homoskedastic error structure, BART, well outperforms all others. The two forests, BF and RF, are both a large improvement over DT. The BF average RMSE is only about 1% better than the RF's, since Bayesian and classical bootstrapping differ little in practice. BCART does very poorly: worse than a single DT. We hypothesize that this is due to the notoriously poor mixing of the BCART MCMC, such that this fit is neither finding a posterior mean (as it is intended to) or optimizing to a local posterior mode (as DT does).

We also include the extremely randomized trees (ET) of Geurts et al. (2006), which are similar to RFs except that (a) instead of optimizing greedy splits, a few candidate splits are chosen randomly and the best is used and (b) the full unweighted data sample is used to fit each tree. ET slightly outperforms both BF and RF; in our experience this happens in small datasets where the restriction of population support to observed support (as assumed in our npB analysis) is invalid and the forest posteriors are over-fit.

2.6. California housing data

As more realistic example, we consider the California housing data of Pace & Barry (1997) consisting of median home price along with eight features (location, income, housing stock) for 20640 census blocks. Since prices tend

⁷In this and the next example, CART-based algorithms had minimum-leaf-samples set at 3 and the ensembles contain 100 trees. BART and BCART run at their `bayestree` and `tgpr` R package defaults, except that BART draws only 200 trees after a burn-in of 100 MCMC iterations. This is done to get comparable compute times; for these settings BART requires around 75 seconds per fold with the California housing data, compared to BF's 10 seconds when run in serial, and 3 seconds running on 4 cores.

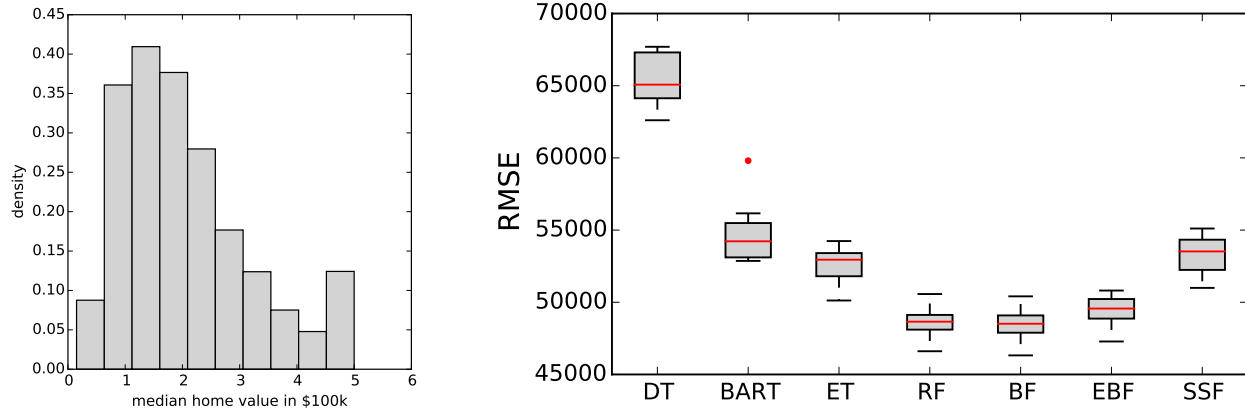


Figure 3. California housing data. The response sample is on the left, and the right panel shows predictive RMSE across 10-fold CV.

to vary with covariates on a multiplicative scale, standard analyses take log median price as the response of interest. Instead, we will model the conditional expectation of dollar median price. This is relevant to applications where prediction of raw transaction data is of primary interest. The marginal response distribution for median home price is shown in the left panel of Figure 3.⁸

Figure 3 shows results from a 10-fold cross-validation (CV) experiment, with details in Table 1. Except for DT and BCART, which still perform worse than all others⁹, results are reversed from those for the small-sample and homoskedastic Friedman data. Both RF and BF do much better than BART and its restrictive error model. BF again offers a small gain over RF. Since the larger sample size makes observed support a better approximation to population support, the forests outperform ET. The EBF (empirical Bayesian forest) and SSF (sub-sample forest) predictors are based on distributed computing algorithms that we introduce later in Section 3. At this point we note only that the massively scalable EBF is amongst the top performers; the next section helps explain why.

	BF	RF	EBF	ET	SSF	BART	DT
RMSE	48.2	48.5	49.4	52.5	53.1	54.8	65.6
%WTB	0.0	0.5	2.4	8.7	10.0	13.4	35.9

Table 1. Average RMSE in \$10k and % worse-than-best for the California housing data 10-fold CV experiment.

3. Understanding the posterior over trees

Theory on decision trees is sparse. The original CART book of Breiman et al. (1984) provides consistency results; they show that any partitioner that is able to learn enough

to split the DGP support into very low probability leaves will be able to reproduce the conditional response distribution of that DGP. However, this result says nothing about the structure of the underlying trees, nor does it say anything about the ability of a tree to predict when there is not enough data to finely partition the DGP. Others have focused on the frequentest properties of individual split decisions. In the CHAID work of Kass (1980), splits are based upon χ^2 tests at each leaf node. Loh (2002) and Hothorn et al. (2006) are generalizations, both of which combat multiple testing and other biases inherent in tree-building through a sequence of hypothesis tests. However, such contributions provide little intuition in the setting where we are not working from a no-split null hypothesis distribution.

Despite this lack of theory, it is generally thought that there is large uncertainty (sampling or posterior) about the structure of decision trees. For example, Geurts & Wehenkel (2000) present extensive simulation of tree uncertainty and find that the location and order of splits is often no-better than random (this motivates work by the same authors on extremely randomized trees). The intuition behind such randomness is clear: the probability of a tree having any given branch structure is the product of conditional probabilities for each successive split. After a enough steps, any specific tree approaches probability of zero.

However, it is possible that elements of the tree structure are stable. For example, in the context of boosting, Appel et al. (2013) argue that the conditionally optimal split locations for internal nodes can be learned from subsets of the full data allocated to each node. They use this to propose a faster boosting algorithm. In this section we make a related claim: in large samples, there is little posterior variation for the top of the tree. We make the point first in theory, then through empirical demonstration.

⁸Values appear to have been capped at \$500k

⁹We’ve left off BCART’s average RMSE of \$82k, 70% WTB.

3.1. Probability of the sample CART tree

We focus on regression trees for this derivation, wherein node impurity is measured as the sum of squared errors. Consider a simplified setup with each $x_j \in \{0, 1\}$ a binary random variable (possibly created through discretization of a continuous input). We'll investigate here the probability that the impurity minimizing split on a given node is the same for a given realization of posterior DGP weights as it is for the unweighted data sample.

Suppose $\{\mathbf{z}_1 \dots \mathbf{z}_n\}$ are the data to be partitioned at some tree node, with $\mathbf{z}_i = [y_i, x_{i1}, \dots, x_{ip}]'$. Say that $\mathcal{F}_j = \{i : x_{ij} = 0\}$ and $\mathcal{T}_j = \{i : x_{ij} = 1\}$ are the partitions implied by splitting on a given x_j . The resulting impurity is

$$\sigma_j^2(\boldsymbol{\theta}) = \frac{1}{n} \sum_i \theta_i [y_i - \mu_j(x_{ij})]^2, \quad (6)$$

$$\mu_j(0) = \sum_{i \in \mathcal{F}_j} y_i \theta_i / |\boldsymbol{\theta}_{\mathcal{F}_j}|, \mu_j(1) = \sum_{i \in \mathcal{T}_j} y_i \theta_i / |\boldsymbol{\theta}_{\mathcal{T}_j}|.$$

We could use the Bayesian bootstrap to simulate the posterior for σ_j implied by the exponential posterior on $\boldsymbol{\theta}$, but an analytic expression is not available. Instead, we follow the approach used in Lancaster (2003), Poirier (2011), and Taddy et al. (2014): derive a first-order Taylor approximation to the function σ_j and describe the posterior for that related functional.

In particular, the $1 \times n$ gradient of σ_j with respect to $\boldsymbol{\theta}$ is

$$\nabla \sigma_j^2 = \nabla \frac{1}{n} \left[\sum_i \theta_i y_i^2 - \frac{1}{|\boldsymbol{\theta}_{\mathcal{F}_j}|} (\mathbf{y}'_{\mathcal{F}_j} \boldsymbol{\theta}_{\mathcal{F}_j})^2 - \frac{1}{|\boldsymbol{\theta}_{\mathcal{T}_j}|} (\mathbf{y}'_{\mathcal{T}_j} \boldsymbol{\theta}_{\mathcal{T}_j})^2 \right] \quad (7)$$

which has elements $\nabla_i \sigma_j^2 = (y_i - \mu(x_{ij}))^2 / n$

The Taylor approximation is then

$$\begin{aligned} \sigma_j^2 &\approx \tilde{\sigma}_j^2 = \sigma_j^2(\mathbf{1}) + \nabla \sigma_j^2|_{\boldsymbol{\theta}=\mathbf{1}} (\boldsymbol{\theta} - \mathbf{1}) \\ &= \frac{1}{n} \sum_i \theta_i (y_i - \bar{y}_j(x_{ij}))^2 \end{aligned} \quad (8)$$

with $\bar{y}_j(0) = \frac{1}{n_{\mathcal{F}_j}} \sum_{i \in \mathcal{F}_j} y_i$ and $\bar{y}_j(t) = \frac{1}{n_{\mathcal{T}_j}} \sum_{i \in \mathcal{T}_j} y_i$ the observed response averages in each partition.

Suppose that $\sigma_1(\mathbf{1}) < \sigma_j(\mathbf{1}) \forall j$, so that variable '1' is that selected for splitting based on the unweighted data sample. Then we can quantify variability about this selection by looking at differences in approximate impurity,

$$\begin{aligned} \Delta_j(\boldsymbol{\theta}) &= \tilde{\sigma}_1^2 - \tilde{\sigma}_j^2 \\ &= \frac{1}{n} \sum_i \theta_i [\bar{y}_1^2(x_{i1}) - \bar{y}_j^2(x_{ij}) - \\ &\quad 2y_i (\bar{y}_1(x_{i1}) - \bar{y}_j(x_{ij}))] \\ &\equiv \frac{1}{n} \sum_i \theta_i d_{ji}. \end{aligned} \quad (9)$$

Say $\mathbf{d}_j = [d_{j1} \dots d_{jn}]'$ is the vector of squared error differences. Then the total difference has mean $\mathbb{E} \Delta_j = \bar{d}_j$ and variance $\text{var} \Delta_j = \mathbf{d}'_j \mathbf{d}_j / n^2$. Since Δ_j is the mean of independent Exponential random variables with known means and variances, the central limit theorem applies and it converges in distribution to a Gaussian:

$$\sqrt{n}(\Delta_j(\boldsymbol{\theta}) - \bar{d}_j) \rightsquigarrow N(0, \mathbf{d}'_j \mathbf{d}_j / n). \quad (10)$$

The weighted-sample impurity-minimizing split matches that for the unweighted-sample if and only if all Δ_j are negative, which occurs with probability (note $\bar{d}_j < 0$)

$$\begin{aligned} p(\Delta_j < 0 : j = 2 \dots p) &\geq 1 - \sum_{j=2}^p p(\Delta_j > 0) \\ &\rightsquigarrow 1 - \sum_{j=2}^p \Phi \left(-\frac{\sqrt{n} |\bar{d}_j|}{\sqrt{\mathbf{d}'_j \mathbf{d}_j / n}} \right) \\ &\geq 1 - \frac{1}{\sqrt{2\pi}} \sum_{j=2}^p \frac{\exp(-nz_j^2/2)}{z_j \sqrt{n}} \end{aligned} \quad (11)$$

where $z_j = |\bar{d}_j| (\mathbf{d}'_j \mathbf{d}_j / n)^{-\frac{1}{2}}$ is sample mean increase in impurity over the sample standard deviation of impurity. This ratio is bounded in probability, so that that the exponential bound goes to zero very quickly with n . In particular, ignoring variation in z_j across variables, we arrive at the approximate lower bound on the probability of the sample split

$$p(\text{posterior split matches sample split}) \gtrsim 1 - \frac{p}{\sqrt{n}} e^{-n}.$$

Even allowing for $p \approx n \times d$, with d some underlying continuous variable dimension and p the input dimension discretization on these variables, the probability of the observed split goes to one at order $\mathcal{O}(n^{-1})$ if $d < e^n / n^{3/2}$.

Given this, why is there any uncertainty at all about trees? The answer is recursion: each split is conditionally stable given the sample at the current node, but the probability of any specific sequence of splits is roughly (we don't actually have independence) the product of individual node split probabilities. This will get small as we move deeper down the tree, and given one split different from the sample CART the rest of the tree will grow arbitrarily far from this modal structure. In addition, the sample size going into our probability bounds is shrinking exponentially with each partition, whereas the dimension of eligible split variables is reduced only by one at each level.

Regardless of overall tree variability, we can take from this section an expectation that for large samples the high-level tree structure varies little across posterior DGP realizations. The next section shows this to be the case in practice.

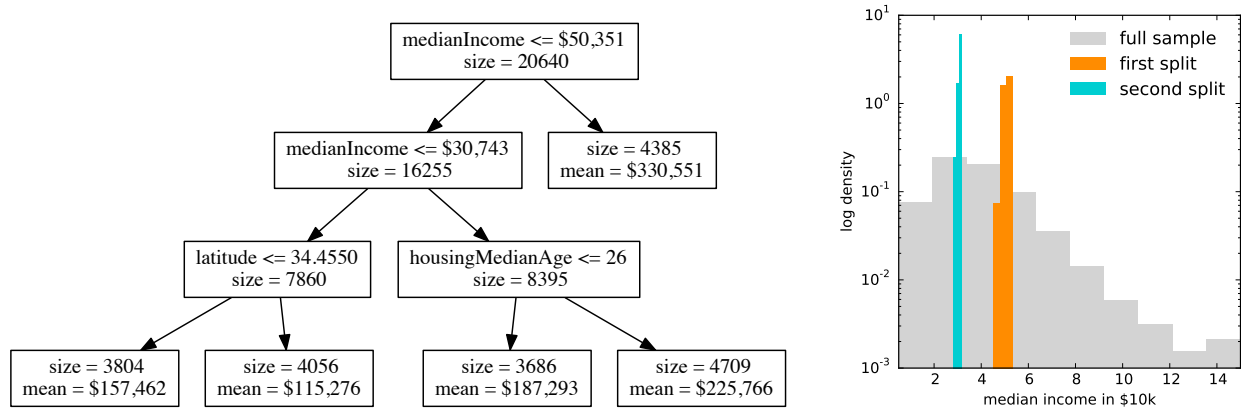


Figure 4. California housing data. The left shows the CART fit on unweighted data with a minimum of 3500 samples-per-leaf. The right panel shows the distribution of first and second split locations – always on median income – across 100 draws from the BF posterior.

3.2. Trunk stability in California

We’ll illustrate the stability of high-level tree structure on our California housing data. Consider a ‘trunk’ DT fit to the unweighted data with no less than 3500 census blocks (out of 20640 total) in each leaf partition. This leads to the tree on the left of Figure 4 with five terminal nodes. It splits on the obvious variable of median income, and also manages to divide California into its north and south regions (34.455 degrees north is just north of Santa Barbara).

To investigate tree uncertainty, we can apply the BF algorithm and repeatedly fit similar CART trees to randomly-weighted data. Running a 100 tree BF, we find that the sample tree occurs 62% of the time. The second most common tree, occurring 28% of the time, differs only in that it splits on median income again instead of on housing median age. Thus 90% of the posterior weight is on trees that split on income twice, and then latitude. Moreover, a striking 100% of trees have first two splits on median income. From this, we can produce the plot in the right panel of Figure 4 showing the locations of these first two splits: each split-location posterior is tightly concentrated around the corresponding sample CART split.

4. Empirical Bayesian forests

Empirical Bayes (EB) is an established framework with a successful track record in fast approximate Bayesian inference; see, e.g., Efron (2010) for an overview. In parametric models, EB can be interpreted as fixing at their marginal posterior maximum the parameters at higher levels of a hierarchical model. For example, in the simple setting of many group means (the average student test score for each of many schools) shrunk to a global center (the outcome for an imaginary ‘average school’), an EB procedure will shrink each group mean toward the overall sample average (each school towards all-student average). Kass & Steffey

(1989) investigate such procedures and show that, under fairly general conditions, the EB conditional posterior for each group mean quickly approaches the fully Bayesian unconditional posterior as the sample size grows. *This occurs because there is little uncertainty about the global mean.*

CART trees are not a parametric model, but they are hierarchical and admit an interpretation similar to those studied in Kass & Steffey (1989). The data which reaches any given interior node is a function of the partitioning implied by nodes shallower in the tree. Moreover, due to the greedy algorithm through which CART grows, a given shallow trunk is unaffected by changes to the tree structure below. Finally, Section 3 demonstrated that, like high levels in a parametric hierarchical model, there is relatively little uncertainty about high levels of the tree.

An empirical Bayesian forest (EBF) takes advantage of this structure by fixing the highest levels in the hierarchy – the earliest CART splits – at a single estimate of high posterior probability. We fit a single shallow CART *trunk* to the unweighted sample data, and sample a BF ensemble of *branches* at each terminal node of this trunk. The initial CART trunk maps observations to their branch, and each branch BF deals with a dataset of manageable size and can be fit in parallel without any communication with the other branches. That is, EBF replaces the BF posterior sample of trees with a conditional posterior sample that takes the pre-fit trunk as fixed. Since the trunk has relatively low variance for large samples (precisely the setting where such distribution is desirable), the EBF should provide predictions similar to that of the full BF at a fraction of the cost.¹⁰

In contrast, consider the ‘sub-sample forest’ (SSF) algorithm, which replaces the full data with random sub-samples of manageable size. Independent forests are fit

¹⁰This is especially helpful, because the earliest splits in the tree – requiring impurity search over

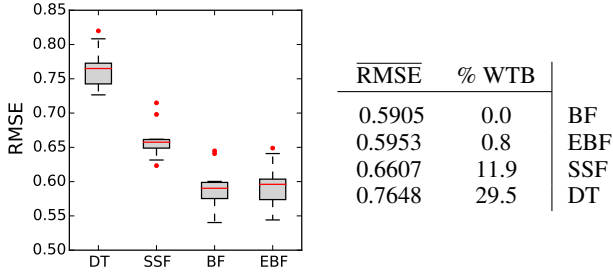


Figure 5. Wine Data: 10-fold OOS prediction experiment Root Mean Square Error and % Worse than Best for the mean RMSE.

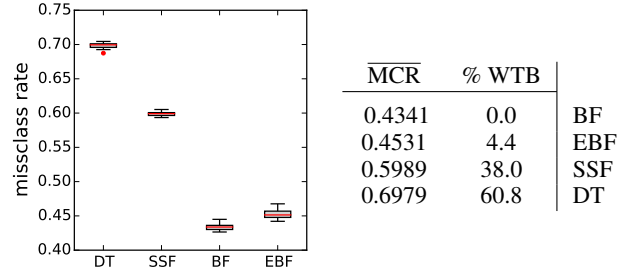


Figure 6. Beer Data: 10-fold OOS prediction experiment Miss-Classification Rate and % Worse than Best for the average MCR.

to each sub-sample and predictions are averaged across all forests. SSF is a commonly applied strategy (e.g., see mention, but not recommendation, of it in Panda et al., 2009), but it implies using only partial data for learning deep tree structure. Although the tree trunks are stable, the full tree is highly uncertain and learning such structure is precisely where you want to use a full Big Data sample.

4.1. Out-of-sample experiments

In the California housing experiment of Figure 3 and Table 4, EBF¹¹ predictions are only 2% worse than those from the full BF. In contrast, SSF predictions are 10% worse.

We consider two additional prediction problems. The first example, taken from (Cortez et al., 1998), involves prediction of an ‘expert’ quality ranking on the scale of 0-10 for wine based upon 11 continuous attributes (physiochemical properties of the wine) plus wine color (red or white). There are 4898 observations. Results are in Figure 5: EBF is only 1% worse than the full BF, while SSF is 12% worse.

The second example is from the Nielson Consumer Panel data, available for academic research through the Kilts Center at Chicago Booth, and our sample contains 73,128 purchases of light beer in a number of US markets during 2004. The response of interest is brand choice, of a possible five major light beer labels: Budweiser, Miller, Coors, Natural, and Busch. Each purchase is associated with a household that is codified through 16 standard demographic categorical variables (maximum age, total income, main occupation, etc). Applying classification forests and DTs (based on Gini impurity) leads to the results in Figure 6: EBF is only 4.4% worse than the BF, while SSF is 38% worse.

4.2. Choosing the trunk depth

From the distributed computing perspective which motivates this work, you should fix the trunk only as deep as you must. In our full scale applications (next section), this

¹¹EBFs use five node trunks in this Section. The SSFs are fit on data split into five equally sized subsets.

means setting branch size so that the working memory required to fit a forest at each branch is slightly less than the memory available on each machine.

This leaves open questions, e.g., on the trade-off between number of trees in the forest and depth of the trunk. We don’t yet have rigorous answers, but the npB framework can help in further investigation. In the interim, Table 2 shows the effect on OOS error from doubling and halving the minimum leaf size for the EBFs in our three examples.

	CA housing			Wine			Beer		
MLS 10 ³	6	3	1.5	2	1	0.5	20	10	5
% WTb	1.6	2.4	4.3	0.3	0.8	2.2	1.0	4.4	7.6

Table 2. % Worse than Best on OOS error for different minimum leaf sizes (MSL) specified in thousands of observations.

4.3. Scaling for Big Data

To close, we note that this work is motivated by the need for reliable forest fitting algorithms that can be deployed on millions or hundreds of millions of observations, as encountered when analyzing internet commerce data. A number of additional engineering details are required for such deployment, but the basic approach is straightforward: an initial trunk is fit (possibly to a data subset)¹², and this trunk acts as a sorting function to map observations to separate locations corresponding to each branch. Forests are then fit at each location (machine) for each branch.

Preliminary work at eBay.com applies EBFs for prediction of ‘Bad Buyer Experiences’ (e.g. complaints, returns, or shipping problems) on the site. Training on a relatively small sample of 12 million transactions, the EBF algorithm using 32 branch chunks is able to provide a 1.3% drop in misclassification over the SSF alternatives. This amounts to more than 20,000 extra detected BBE occurrences over the short sample window, potentially giving eBay the opportunity to try and stop them before they occur.

¹²Note that the original CART trunk can itself be fit in distribution, e.g. using the MLLib library for Apache Spark.

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