

What to do today?

Part I. Introduction and Preparation

- I.1. General Introduction
- I.2. Review on Matrix Algebra (Chp 2.1-4, Supplement 2A)
- I.3. Introduction to R (More at the 1st tutorial)
- I.4. Multivariate Random Variables and Distributions
(Chp 1, 2.5-6, 3)**

Part II. Inference under Multivariate Normal Distribution (Chp 4-7)

II.1 Multivariate Normal Distribution (Chp 4)

- II.1.1 Multivariate Normal Distribution $MN_p(\mu, \Sigma)$ (Chp 4.1-2)**

Part III. Commonly-Used Multivariate Analysis Methods (Chp 8-11)

Part IV. Other Topics (Chp 12)

Part I.4.1 Random Vectors and Matrices (Chp 2.5): Review

- ▶ random vector (multivariate random variable).

$$\mathbf{X} = \begin{pmatrix} X_1 \\ X_2 \\ \vdots \\ X_p \end{pmatrix} = (X_1, X_2, \dots, X_p)'$$

is a p-dim random vector if X_1, \dots, X_p are r.v.s.

- ▶ distribution. The cdf of \mathbf{X} is the joint cumulative distribution function (joint cdf) of X_1, \dots, X_p : for $\mathbf{x} = (x_1, x_2, \dots, x_p)'$,

$$F(\mathbf{x}) = P(\mathbf{X} \leq \mathbf{x}) = P(X_1 \leq x_1, \dots, X_p \leq x_p).$$

- ▶ Suppose the study has n subjects with $\mathbf{X}_1, \dots, \mathbf{X}_n$ their observations on the p -dim r.v. $\mathbf{X} = (X_1, \dots, X_p)'$. A $n \times p$ random matrix:

$$\begin{pmatrix} X_{11} & X_{12} & \dots & X_{1p} \\ X_{21} & X_{22} & \dots & X_{2p} \\ \vdots & \vdots & \dots & \vdots \\ X_{n1} & X_{n2} & \dots & X_{np} \end{pmatrix} = \begin{pmatrix} \mathbf{X}'_1 \\ \mathbf{X}'_2 \\ \vdots \\ \mathbf{X}'_n \end{pmatrix}$$

Part I.4.2 Mean Vectors and Covariance Matrices: Review

- ▶ **expectation (population mean)** of p -dim random vector $\mathbf{X} = (X_1, \dots, X_p)'$:

$$E(\mathbf{X}) = (E(X_1), E(X_2), \dots, E(X_p))'.$$

denoted by $\boldsymbol{\mu} = (\mu_1, \dots, \mu_p)'$.

- ▶ **(population) variance matrix.** p -dim r.v. \mathbf{X} 's variance:

$$V(\mathbf{X}) = E[\mathbf{X} - E(\mathbf{X})][\mathbf{X} - E(\mathbf{X})]' = E(\mathbf{X}\mathbf{X}') - (E\mathbf{X})(E\mathbf{X})',$$

denoted by $\boldsymbol{\Sigma} = (\sigma_{ij})$ with $\sigma_{ij} = \sigma_{ji} = \text{Cov}(X_i, X_j)$.

- ▶ **population correlation matrix.** A standardized variance-covariance matrix: $\boldsymbol{\rho} = (\rho_{ij})$ with $\rho_{ij} = \text{cor}(X_i, X_j) = \sigma_{ij}/\sqrt{\sigma_{ii}}/\sqrt{\sigma_{jj}}$ and thus $\rho_{ii} = 1$.

With $\mathbf{V} = \text{diag}(\sigma_{11}, \dots, \sigma_{pp})$, $\boldsymbol{\rho} = \mathbf{V}^{-1/2} \boldsymbol{\Sigma} \mathbf{V}^{-1/2}$, $\boldsymbol{\Sigma} = \mathbf{V}^{1/2} \boldsymbol{\rho} \mathbf{V}^{1/2}$

- The **covariance** of two random vectors \mathbf{X} and \mathbf{Y} is

$$\text{Cov}(\mathbf{X}, \mathbf{Y}) = E[\mathbf{X} - E(\mathbf{X})][\mathbf{Y} - E(\mathbf{Y})]'$$
.

- **linear combinations of r.v.s.**

- Suppose $Y = c_1X_1 + c_2X_2 + \dots + c_pX_p = \mathbf{c}'\mathbf{X}$.

- $E(Y) = c_1E(X_1) + c_2E(X_2) + \dots + c_pE(X_p) = \mathbf{c}'\mathbf{X}$.

- $V(Y) = \mathbf{c}'V(\mathbf{X})\mathbf{c} = \mathbf{c}'\Sigma\mathbf{c}$.

If X_1, \dots, X_p are indpt,

$$V(Y) = \mathbf{c}'\text{diag}(\sigma_1^2, \dots, \sigma_p^2)\mathbf{c} = \sum_{j=1}^p c_j^2 \sigma_j^2.$$

- Suppose $Z_j = c_{j1}X_1 + c_{j2}X_2 + \dots + c_{jp}X_p$ for $j = 1, \dots, q$, and
 $\mathbf{Z} = (Z_1, \dots, Z_q)' = \mathbf{C}_{q \times p}\mathbf{X}_{p \times 1}$.

- $E(\mathbf{Z}) = \mathbf{C}E(\mathbf{X}) = \mathbf{C}\boldsymbol{\mu}$.

- $V(\mathbf{Z}) = \mathbf{C}V(\mathbf{X})\mathbf{C}' = \mathbf{C}\Sigma\mathbf{C}'$.

- Suppose $\mathbf{U} = \mathbf{A}\mathbf{X}$ and $\mathbf{W} = \mathbf{B}\mathbf{Y}$.

- $\text{Cov}(\mathbf{U}, \mathbf{W}) = \mathbf{A}\text{Cov}(\mathbf{X}, \mathbf{Y})\mathbf{B}'$

Part I.4.3 Descriptive Multivariate Analysis: Review

Summary Statistics: Suppose a study has n iid observations $\mathbf{X}_1, \dots, \mathbf{X}_n$ on a p -dim r.v. $\mathbf{X} = (X_1, \dots, X_p)'$ with $E(\mathbf{X}) = \mu$ and $V(\mathbf{X}) = \Sigma$.

- ▶ $E(\mathbf{X}_i) = \mu$ and $V(\mathbf{X}_i) = \Sigma$ for $i = 1, \dots, n$.
- ▶ **sample mean vector** $\bar{\mathbf{x}} = (\bar{x}_1, \dots, \bar{x}_p)'$ with $\bar{x}_j = \sum_{i=1}^n x_{ij}/n$.
- ▶ **sample variance matrix**

$$\mathbf{S}_n = \begin{pmatrix} s_{11} & s_{12} & \dots & s_{1p} \\ s_{21} & s_{22} & \dots & s_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ s_{p1} & s_{p2} & \dots & s_{pp} \end{pmatrix}$$

with $s_{jk} = \sum_{i=1}^n (x_{ij} - \bar{x}_j)(x_{ik} - \bar{x}_k)/n$.

- ▶ **sample correlation matrix**

$$\mathbf{R} = \begin{pmatrix} 1 & r_{12} & \dots & r_{1p} \\ r_{21} & 1 & \dots & r_{2p} \\ \vdots & \dots & \vdots & \\ r_{p1} & r_{p2} & \dots & 1 \end{pmatrix}$$

with $r_{jk} = \frac{s_{jk}}{\sqrt{s_{jj}}\sqrt{s_{kk}}}$.

Part I.4.4 More on Descriptive Multivariate Analysis (Chp 3.3)

Consider $\mathbf{X}_1, \dots, \mathbf{X}_n$ be a random sample from the population with population mean μ and variance Σ .

- ▶ That is, $\mathbf{X}_1, \dots, \mathbf{X}_n$ are iid observations on \mathbf{X} with $E(\mathbf{X}) = \mu$ and $V(\mathbf{X}) = \Sigma$.
- ▶ The sample mean $\bar{\mathbf{X}} = (\mathbf{X}_1 + \dots + \mathbf{X}_n)/n$ is an *unbiased* estimator of μ : $E(\bar{\mathbf{X}}) = \mu$.
- ▶ The sample variance matrix $\mathbf{S}_n = \left(\sum_{i=1}^n \mathbf{X}_i \mathbf{X}_i' - n\bar{\mathbf{X}}\bar{\mathbf{X}}' \right)/n$ is an *biased* estimator of Σ : $E(\mathbf{S}_n) = \frac{n-1}{n}\Sigma$.
- ▶ **(unbiased) sample variance-covariance matrix:**

$$\mathbf{S} = \frac{n}{n-1} \mathbf{S}_n = \frac{1}{n-1} \sum_{i=1}^n (\mathbf{X}_i - \bar{\mathbf{X}})(\mathbf{X}_i - \bar{\mathbf{X}})'$$

Part II.1.1 Multivariate Normal Distribution $MN_p(\mu, \Sigma)$ (Chp 4.1-2)

The most important distribution in all of Statistics is the normal (Gaussian) distribution.

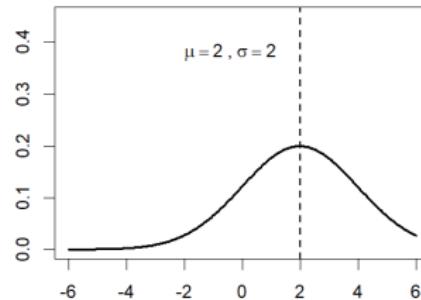
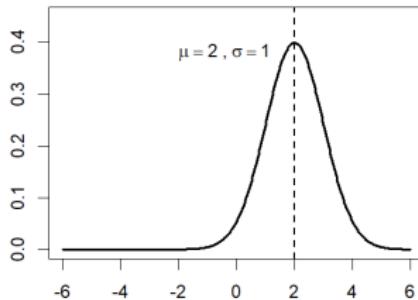
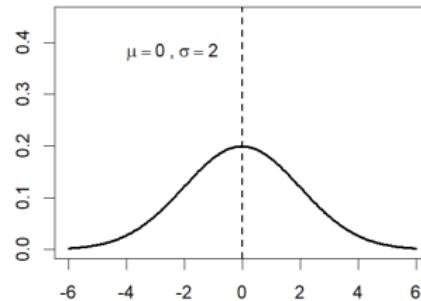
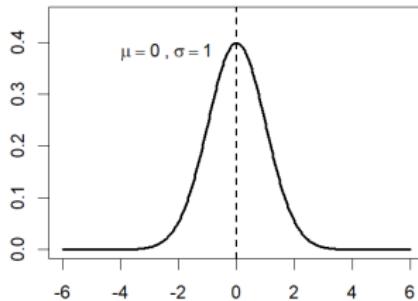
Review on univariate normal distribution $N(\mu, \sigma^2)$:

Definition. A r.v. X has a *normal* distribution if its pdf

$$f(x; \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left\{ -\frac{1}{2} \left(\frac{x-\mu}{\sigma} \right)^2 \right\}, \quad -\infty < x < \infty,$$

where $\sigma > 0$. Denote it by $X \sim N(\mu, \sigma^2)$.

- ▶ If $X \sim N(\mu, \sigma^2)$, $E(X) = \mu$ and $V(X) = \sigma^2$.
- ▶ $N(\mu, \sigma^2)$: a family of distributions.
 - ▶ e.g. $N(0, 1)$, the standard normal distribution.
 $F(x)$ of $N(0, 1)$ is often denoted by $\Phi(x)$ and the rv by Z .



More about the normal distributions ...

- ▶ The pdf is symmetric about μ .
- ▶ As μ changes, the mode of the pdf curve shifts accordingly.
- ▶ As σ increases, the spread of the pdf curve increases.

Part II.1.1 Multivariate Normal Distribution

$MN_p(\mu, \Sigma)$ (Chp 4.1-2)

What is a multivariate normal distribution?

Definition. A r.v. X has a *normal* distribution if its pdf

$$f(x; \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{1}{2}(x-\mu)\sigma^{-2}(x-\mu)\right\}, \quad -\infty < x < \infty,$$

where $\sigma > 0$. Denote it by $X \sim N(\mu, \sigma^2)$.

Part II.1.1 Multivariate Normal Distribution

$MN_p(\mu, \Sigma)$ (Chp 4.1-2)

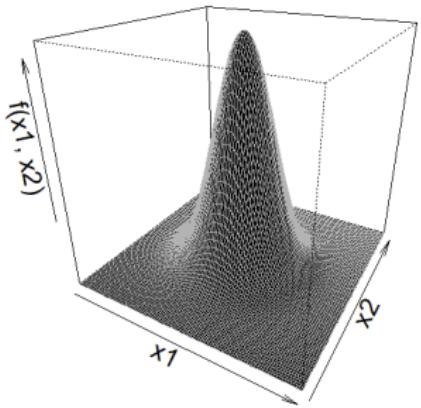
II.1.1A. Multivariate normal distribution

Definition. A p-dim r.v. \mathbf{X} has a *normal* distribution if its pdf

$$f(\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{|2\pi\boldsymbol{\Sigma}|}} \exp \left\{ -\frac{1}{2} (\mathbf{x}-\boldsymbol{\mu})' \boldsymbol{\Sigma}^{-1} (\mathbf{x}-\boldsymbol{\mu}) \right\}, \quad -\infty < \mathbf{x} < \infty,$$

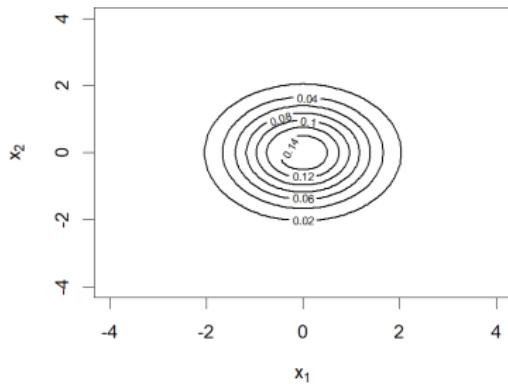
where $\boldsymbol{\Sigma}$ is positive definite. Denote it by $\mathbf{X} \sim MN_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$.

- ▶ If $\mathbf{X} \sim MN_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, $E(\mathbf{X}) = \boldsymbol{\mu}$ and $V(\mathbf{X}) = \boldsymbol{\Sigma}$.
- ▶ $MN_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$: a family of (multivariate) distributions.
 - ▶ e.g. $MN_p(\mathbf{0}, \mathbf{I})$, the standard (multivariate) normal distribution.

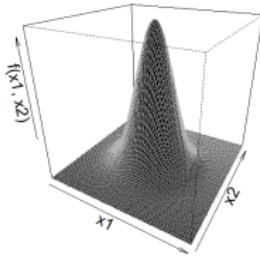


(a) $\text{BN}(\mu, \Sigma)$

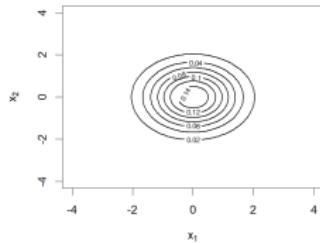
$\mu = (0, 0)'$ and $\Sigma = \text{diag}(1, 1)$



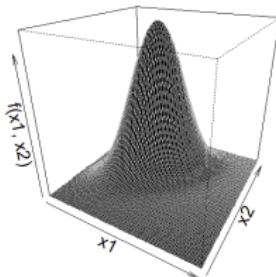
(b) Contour



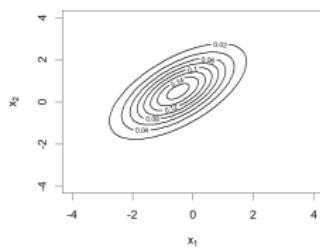
(a) $\text{BN}(\mu, \Sigma)$



(b) Contour



(c) $\text{BN}(\mu, \Sigma)$

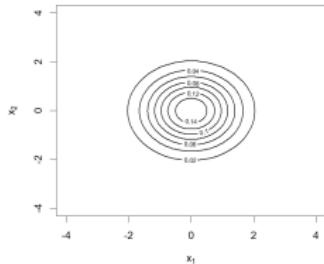


(d) Contour

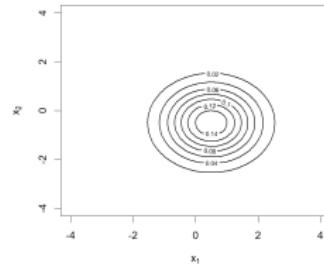
- ▶ (1) $\mu = (0, 0)'$ and $\Sigma = \text{diag}(1, 1)$; (2) $\mu = (0.5, -0.5)'$ and
 $\Sigma = \begin{pmatrix} 1.2 & 0.75 \\ 0.75 & 1.2 \end{pmatrix}$

More about the normal distributions ...

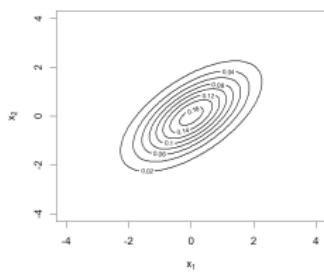
- ▶ The pdf is symmetric about μ .
- ▶ As μ changes, the mode of the pdf curve shifts accordingly.
- ▶ As Σ changes, the spread and/or shape of the pdf surface change accordingly.



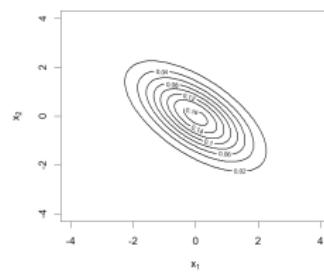
Contour of (1)



Contour of (2)



Contour of (3)



Contour of (4)

- ▶ (1) $\mu = (0, 0)'$ and $\Sigma = \text{diag}(1, 1)$; (2) $\mu = (0.5, -0.5)'$ and $\Sigma = \text{diag}(1, 1)$;
- ▶ (3) $\mu = (0, 0)'$ and $\Sigma = \begin{pmatrix} 1.2 & 0.75 \\ 0.75 & 1.2 \end{pmatrix}$; (4) $\mu = (0, 0)'$ and $\Sigma = \begin{pmatrix} 1.2 & -0.75 \\ -0.75 & 1.2 \end{pmatrix}$

II.1.1B. Shape of Multivariate Normal Density

Suppose $\mathbf{X} \sim BN(\boldsymbol{\mu}, \boldsymbol{\Sigma})$: its pdf

$$f(\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{|2\pi\boldsymbol{\Sigma}|}} \exp \left\{ -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})' \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right\}, \quad -\infty < \mathbf{x} < \infty,$$

where $\boldsymbol{\mu} = (\mu_1, \mu_2)'$ and $\boldsymbol{\Sigma} = (\sigma_{ij}) = \begin{pmatrix} \sigma_1^2 & \rho\sigma_1\sigma_2 \\ \rho\sigma_1\sigma_2 & \sigma_2^2 \end{pmatrix}$.

- ▶ If $\rho = 0$, $f(\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \phi(x_1; \mu_1, \sigma_1^2)\phi(x_2; \mu_2, \sigma_2^2)$: $\phi(x; \mu, \sigma^2)$ is the pdf of $X \sim N(\mu, \sigma^2)$.
- ▶ The density $f(\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma})$ is a constant for all $\mathbf{x} = (x_1, x_2)'$ satisfy $(\mathbf{x} - \boldsymbol{\mu})' \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) = c^2$. (It defines an ellipse centered at $\boldsymbol{\mu} = (\mu_1, \mu_2)'$.)

Normal density contour:

$$\{\mathbf{x} : (\mathbf{x} - \boldsymbol{\mu})' \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) = c^2\}$$

- ▶ It defines an ellipse centered at $\boldsymbol{\mu} = (\mu_1, \mu_2)'$.
- ▶ The ellipse's axes are $\pm c \sqrt{\lambda_j} \mathbf{e}_j$ for $j = 1, 2$, where λ_j, \mathbf{e}_j are *eigenvalue* and *eigenvector* pairs of $\boldsymbol{\Sigma}$.

II.1.1C. Important Properties of Multivariate Normal Distribution

- If $\mathbf{X} \sim MN_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, $\mathbf{Y} = \mathbf{A}'\mathbf{X} + \mathbf{d} \sim MN_p(\mathbf{A}'\boldsymbol{\mu} + \mathbf{d}, \mathbf{A}'\boldsymbol{\Sigma}\mathbf{A})$.

e.g. Using $\mathbf{d} = -\boldsymbol{\mu}$ and $\mathbf{A} = \boldsymbol{\Sigma}^{-1/2}$,

$$\mathbf{Y} = \mathbf{A}'\mathbf{X} + \mathbf{d} \sim MN_p(\mathbf{0}, \mathbf{I}).$$

- $\mathbf{X} \sim MN_p(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \iff \mathbf{a}'\mathbf{X} \sim N(\mathbf{a}'\boldsymbol{\mu}, \mathbf{a}'\boldsymbol{\Sigma}\mathbf{a})$ for any $\mathbf{a} \in \mathcal{R}^p$.

More generally, $\mathbf{X} \sim MN_p(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \iff$

$$\mathbf{A}'\mathbf{X} \sim MN_q(\mathbf{A}'\boldsymbol{\mu}, \mathbf{A}'\boldsymbol{\Sigma}\mathbf{A})$$
 for any $\mathbf{A} \in \mathcal{R}^{p \times q}$.

The normality is preserved under any linear transformation.

- If $\mathbf{X} \sim MN_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ and

$$\mathbf{X} = \begin{pmatrix} \mathbf{X}_1 \\ \mathbf{X}_2 \end{pmatrix}, \quad \boldsymbol{\mu} = \begin{pmatrix} \boldsymbol{\mu}_1 \\ \boldsymbol{\mu}_2 \end{pmatrix}, \quad \boldsymbol{\Sigma} = \begin{pmatrix} \boldsymbol{\Sigma}_{11} & \boldsymbol{\Sigma}_{12} \\ \boldsymbol{\Sigma}_{21} & \boldsymbol{\Sigma}_{22} \end{pmatrix},$$

$\mathbf{X}_1 \sim MN_{p_1}(\boldsymbol{\mu}_1, \boldsymbol{\Sigma}_{11})$ and $\mathbf{X}_2 \sim MN_{p_2}(\boldsymbol{\mu}_2, \boldsymbol{\Sigma}_{22})$.

- If $\mathbf{X}_1 \sim MN_{p_1}(\boldsymbol{\mu}_1, \boldsymbol{\Sigma}_{11})$ and $\mathbf{X}_2 \sim MN_{p_2}(\boldsymbol{\mu}_2, \boldsymbol{\Sigma}_{22})$, then \mathbf{X}_1 and \mathbf{X}_2 are independent \iff

$$\mathbf{X} = \begin{pmatrix} \mathbf{X}_1 \\ \mathbf{X}_2 \end{pmatrix} \sim MN_{p_1+p_2}\left(\begin{pmatrix} \boldsymbol{\mu}_1 \\ \boldsymbol{\mu}_2 \end{pmatrix}, \begin{pmatrix} \boldsymbol{\Sigma}_{11} & \mathbf{0} \\ \mathbf{0} & \boldsymbol{\Sigma}_{22} \end{pmatrix}\right).$$

- ▶ If $\mathbf{X} = \begin{pmatrix} \mathbf{x}_1 \\ \mathbf{x}_2 \end{pmatrix} \sim MN_{p_1+p_2}\left(\begin{pmatrix} \boldsymbol{\mu}_1 \\ \boldsymbol{\mu}_2 \end{pmatrix}, \begin{pmatrix} \boldsymbol{\Sigma}_{11} & \boldsymbol{\Sigma}_{12} \\ \boldsymbol{\Sigma}_{21} & \boldsymbol{\Sigma}_{22} \end{pmatrix}\right)$ with $\boldsymbol{\Sigma}_{22}$ invertible (i.e. $|\boldsymbol{\Sigma}_{22}| > 0$), the conditional distribution $\mathbf{x}_1 | \mathbf{x}_2 = \mathbf{x}_2$ is
 $MN_{p_1}(\boldsymbol{\mu}_1 + \boldsymbol{\Sigma}_{12}\boldsymbol{\Sigma}_{22}^{-1}(\mathbf{x}_2 - \boldsymbol{\mu}_2), \boldsymbol{\Sigma}_{11} - \boldsymbol{\Sigma}_{12}\boldsymbol{\Sigma}_{22}^{-1}\boldsymbol{\Sigma}_{21}).$
- ▶ If $\mathbf{X} \sim MN_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ with $\boldsymbol{\Sigma}$ invertible,
 $(\mathbf{X} - \boldsymbol{\mu})' \boldsymbol{\Sigma}^{-1} (\mathbf{X} - \boldsymbol{\mu}) \sim \chi_p^2$, the Chi-square distribution with degree of freedom p .

What will we study next?

- ▶ *Part I. Introduction and Preparation*
- ▶ **Part II. Inference under Multivariate Normal Distribution (Chp 4-7)**
 - ▶ **II.1 Multivariate Normal Distribution (Chp 4)**
 - ▶ **II.2 Inferences on Mean Vector (Chp 5)**
 - ▶ **II.3 Comparisons of Several Mean Vector (Chp 6)**
 - ▶ **II.4 Multivariate Linear Regression (Chp 7)**
- ▶ *Part III. Commonly-Used Multivariate Analysis Methods (Chp 8-11)*
- ▶ *Part IV. Other Topics (Chp 12)*