

1.

Let

$$y = \Sigma^{-\frac{1}{2}}(x - \mu),$$

where $\Sigma^{\frac{1}{2}}$ is the unique symmetric positive definite square root of Σ satisfying

$$\Sigma^{\frac{1}{2}}\Sigma^{\frac{1}{2}} = \Sigma.$$

Then $x = \mu + \Sigma^{\frac{1}{2}}y$, and the Jacobian determinant of this transformation is

$$\left| \frac{\partial x}{\partial y} \right| = \left| \Sigma^{\frac{1}{2}} \right| = |\Sigma|^{\frac{1}{2}}.$$

Substituting $x = \mu + \Sigma^{\frac{1}{2}}y$ gives:

$$\begin{aligned} \int_{\mathbb{R}^k} f(x) dx &= \frac{1}{(2\pi)^{\frac{k}{2}} |\Sigma|^{\frac{1}{2}}} \int_{\mathbb{R}^k} e^{\left[-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)\right]} dx \\ &= \frac{1}{(2\pi)^{\frac{k}{2}} |\Sigma|^{\frac{1}{2}}} \int_{\mathbb{R}^k} e^{\left[-\frac{1}{2}y^T y\right]} |\Sigma|^{\frac{1}{2}} dy \\ &= \frac{1}{(2\pi)^{\frac{k}{2}}} \int_{\mathbb{R}^k} e^{-\frac{1}{2}\|y\|^2} dy. \end{aligned}$$

The integrand factorizes across coordinates:

$$\int_{\mathbb{R}^k} e^{-\frac{1}{2}\|y\|^2} dy = \prod_{i=1}^k \int_{-\infty}^{\infty} e^{-\frac{1}{2}y_i^2} dy_i = (2\pi)^{\frac{k}{2}}.$$

Therefore:

$$\int_{\mathbb{R}^k} f(x) dx = \frac{(2\pi)^{\frac{k}{2}}}{(2\pi)^{\frac{k}{2}}} = 1.$$

2.

(a)

Write the trace explicitly:

$$\text{trace}(AB) = \sum_{i=1}^n \sum_{j=1}^n A_{ij} B_{ji}.$$

Then, for each entry A_{pq} ,

$$\frac{\partial}{\partial A_{pq}} \text{trace}(AB) = B_{qp}.$$

Hence, the matrix of all partial derivatives satisfies:

$$\frac{\partial}{\partial A} \text{trace}(AB) = B^T.$$

(b)

Using the cyclic invariance of the trace ($\text{trace}(XYZ) = \text{trace}(ZXY)$):

$$\text{trace}(xx^T A) = \text{trace}(Axx^T) = \sum_{i,j} A_{ij} x_j x_i = x^T A x.$$

(c)

Let $x_1, \dots, x_n \in \mathbb{R}^k$ be i.i.d. with density

$$f(x; \mu, \Sigma) = \frac{1}{(2\pi)^{\frac{k}{2}} |\Sigma|^{\frac{1}{2}}} e^{\left[-\frac{1}{2}(x-\mu)^T \Sigma^{-1} (x-\mu)\right]},$$

where $\mu \in \mathbb{R}^k$ and $\Sigma \in \mathbb{R}^{k \times k}$ is symmetric positive definite ($\Sigma > 0$). The log-likelihood is

$$l(\mu, \Sigma) = \sum_{i=1}^n \log f(x_i; \mu, \Sigma) = -\frac{nk}{2} \log |\Sigma| - \frac{1}{2} \sum_{i=1}^n (x_i - \mu)^T \Sigma^{-1} (x_i - \mu).$$

For convenience define $S_i \equiv (x_i - \mu)(x_i - \mu)^T$ and $S \equiv \sum_{i=1}^n S_i$. We will use

trace identities:

$$(x - \mu)^T \Sigma^{-1} (x - \mu) = \text{trace}(\Sigma^{-1} (x_i - \mu)(x_i - \mu)^T) = \text{trace}(\Sigma^{-1} S_i).$$

Write the μ -dependent part:

$$l(\mu, \Sigma) = \text{const w.r.t. } \mu - \frac{1}{2} \sum_{i=1}^n (x_i - \mu)^T \Sigma^{-1} (x_i - \mu).$$

Differentiate (vector derivative). Use the identity for symmetric A :

$$\frac{\partial}{\partial \mu} (x_i - \mu)^T A (x_i - \mu) = -2A(x_i - \mu).$$

Since Σ^{-1} is symmetric,

$$\frac{\partial l}{\partial \mu} = -\frac{1}{2} \sum_{i=1}^n (-2\Sigma^{-1} (x_i - \mu)) = \Sigma^{-1} \sum_{i=1}^n (x_i - \mu).$$

Set $\frac{\partial l}{\partial \mu} = 0$. Because Σ^{-1} is invertible,

$$\sum_{i=1}^n (x_i - \mu) = 0 \quad \Rightarrow \quad \hat{\mu} = \bar{x} = \frac{1}{n} \sum_{i=1}^n x_i.$$

Check: second derivative (Hessian) with respect to μ is

$$\frac{\partial^2 l}{\partial \mu \partial \mu^T} = -n\Sigma^{-1},$$

which is negative definite because $\Sigma^{-1} \succ 0$. Thus the critical point $\hat{\mu}$ is a (strict)

local maximum in μ , and because the objective is concave in μ it is the global

maximizer for fixed Σ .

We now maximize with respect to $\Sigma \succ 0$. Use the log-likelihood written in

trace form:

$$l(\mu, \Sigma) = -\frac{n}{2} \log |\Sigma| - \frac{1}{2} \sum_{i=1}^n \text{trace}(\Sigma^{-1} S_i) + \text{const.}$$

(“const” depends on data but not on Σ .)

We use standard matrix derivative facts (for Σ symmetric and $\Sigma \succ 0$):

$$\frac{\partial}{\partial \Sigma} \log |\Sigma| = (\Sigma^{-1})^T = \Sigma^{-1}, \quad \frac{\partial}{\partial \Sigma} \text{trace}(\Sigma^{-1} S) = -(\Sigma^{-1} S \Sigma^{-1})^T = -\Sigma^{-1} S \Sigma^{-1}.$$

(Transposes drop because Σ^{-1} and S are symmetric when appropriate; the

formulas above are general.)

Differentiate:

$$\frac{\partial l}{\partial \Sigma} = -\frac{n}{2}\Sigma^{-1} + \frac{1}{2}\sum_{i=1}^n \Sigma^{-1}S_i\Sigma^{-1}.$$

Set equal to zero:

$$-\frac{n}{2}\Sigma^{-1} + \frac{1}{2}\sum_{i=1}^n \Sigma^{-1}S_i\Sigma^{-1} = 0.$$

Multiply on the left and right by Σ (which is invertible):

$$-nI + \sum_{i=1}^n S_i\Sigma^{-1} = 0.$$

Right-multiply by Σ or rearrange to

$$\sum_{i=1}^n S_i = n\Sigma.$$

Therefore the stationary point for Σ is

$$\hat{\Sigma} = \frac{1}{n}\sum_{i=1}^n (x_i - \hat{u})(x_i - \hat{u})^T.$$

(We substituted $u = \hat{u}$ because \hat{u} was found earlier.)

We must argue the stationary point is actually the global maximum on

domain $\Sigma \succ 0$.

Two standard justifications:

1. Concavity in μ and Σ^{-1} . The log-likelihood $l(\mu, \Sigma)$ is a concave function of μ (quadratic, negative definite Hessian) and is concave in the precision matrix $\Lambda = \Sigma^{-1}$ (linear in $\log \det$ term becomes $\log \det(\Lambda^{-1}) = -\log \det \Lambda$ plus linear term $\text{trace}(\Lambda S)$; more directly, l is concave in Λ). Therefore the stationary point in the domain $\Sigma \succ 0$ is the unique global maximum.

2. Second-variation in Σ . One can compute the second variation of l in directions H with small symmetric H and show it is negative definite at the solution; equivalently, the Hessian operator of l on the space of symmetric matrices is negative definite at the stationary point. This confirms a (strict) local maximum; combined with concavity argument above yields global maximality.

Hence, $\hat{\mu}, \hat{\Sigma}$ above are the MLEs.

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- The MLE for the covariance uses factor $1/n$. The unbiased sample covariance uses $1/(n-1)$; $\frac{1}{n-1} \sum (x_i - \bar{x})(x_i - \bar{x})^T$, but that is **not** the MLE.
 - Domain: the MLE $\hat{\Sigma}$ is positive semidefinite; with nondegenerate data it is positive definite. If data are degenerate (e.g., points lie in a lower-dimensional

affine subspace), $\hat{\Sigma}$ may be singular and the MLE over $\Sigma \succ 0$ does not exist (or is on the boundary).

- As a final boxed summary:

$$\hat{\mu} = \bar{x} = \frac{1}{n} \sum_{i=1}^n x_i, \quad \hat{\Sigma} = \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})(x_i - \bar{x})^T$$