

## STAT 305 Assignment 5

Name: Kevin Nguyen  
 Student Number: 97703045

### Question 1

1. We have the log-likelihood

$$\begin{aligned}
 LL(\theta) &= \log \left[ n! \prod_{r,c} \frac{\theta_{r,c}^{n_{r,c}}}{n_{r,c}!} \right] \\
 &= \log(n!) + \log \left[ \prod_{r,c} \frac{\theta_{r,c}^{n_{r,c}}}{n_{r,c}!} \right] \\
 &= \log(n!) + \sum_{r,c} \log \left[ \frac{\theta_{r,c}^{n_{r,c}}}{n_{r,c}!} \right] \\
 &= \log(n!) + \left[ \sum_{r,c} n_{r,c} \log(\theta_{r,c}) - \log(n_{r,c}!) \right]
 \end{aligned}$$

2. Under the null,  $\theta_{r,c} = \phi_r \rho_c$

i ) finding the MLE,  $\hat{\phi}_r$

$$\begin{aligned}
 \mathcal{L} &= \log(n!) + \left[ \sum_{r,c} n_{r,c} \log(\phi_r \rho_c) - \log(n_{r,c}!) \right] \\
 \text{constraint : } \sum_r \phi_r &= 1 \\
 \frac{\partial}{\partial \phi_r} &= \frac{\partial}{\partial \phi_r} \left[ \mathcal{L} - \lambda \left( \sum_r \phi_r - 1 \right) \right] \\
 &= \left[ \sum_{r,c} \frac{\partial}{\partial \phi_r} (n_{r,c} \log(\phi_r \rho_c)) \right] - \lambda \\
 &= \sum_{r,c} \frac{n_{r,c}}{\phi_r} - \lambda = \frac{1}{\phi_r} \sum_{r,c} n_{r,c} - \lambda \\
 &= \frac{n_{r.}}{\phi_r} - \lambda = 0 \\
 \phi_r &= \frac{n_{r.}}{\lambda} \propto n_{r.} \\
 \hat{\phi}_r &= \frac{n_{r.}}{\sum_r n_{r.}} = \frac{n_{r.}}{n} \quad \square
 \end{aligned}$$

to make sure  $\phi_r$  sums to 1

ii ) finding the MLE,  $\hat{\rho}_c$

$$\mathcal{L} = \log(n!) + \left[ \sum_{r,c} n_{r,c} \log(\phi_r \rho_c) - \log(n_{r,c}!) \right]$$

constraint :  $\sum_c \rho_c = 1$

$$\begin{aligned} \frac{\partial}{\partial \rho_c} &= \frac{\partial}{\partial \rho_c} \left[ \mathcal{L} - \lambda \left( \sum_r \phi_r - 1 \right) \right] \\ &= \left[ \sum_{r,c} \frac{\partial}{\partial \rho_c} (n_{r,c} \log(\phi_r \rho_c)) \right] - \lambda \\ &= \sum_{r,c} \frac{n_{r,c}}{\rho_c} - \lambda = \frac{1}{\rho_c} \sum_{r,c} n_{r,c} - \lambda \\ &= \frac{n_{..c}}{\rho_c} - \lambda = 0 \\ \rho_c &= \frac{n_{..c}}{\lambda} \propto n_{..c} \\ \hat{\rho}_c &= \frac{n_{..c}}{\sum_c n_{..c}} = \frac{n_{..c}}{n} \quad \square \end{aligned}$$

to make sure  $\rho_c$  sums to 1

3. Under the null,  $\theta_{r,c}$  depends on  $\phi_r$  and  $\rho_c$ , which have the constraint of  $\sum_r \phi_r + \sum_c \rho_c = 1$ . Thus, since there are 3 categories in  $r$ , there are 2 free variables for  $\phi_r$  (ex. if you know  $\phi_{\text{red}}$  and  $\phi_{\text{green}}$ ,  $\phi_{\text{blue}}$  is 1 minus those 2 numbers). Similarly, since there are 2 categories in  $c$ , there are 1 free variables for  $\rho_c$ . Thus we have 3 free variable.
4. Under the alternative, the parameters are  $\theta_{r,c}$  for  $r = 1, 2, 3$  and  $c = 1, 2$ , it's obvious to see there's 6 combination of  $(r, c)$ . We have the constraint that  $\sum \theta_{r,c} = 1$  so similar to the question above, we only need the first 5 quantity to figure out the last value. Thus there are 5 free parameters.
5. We can obtain the Wilks statistic via the following R code based on the following formula.

$$W = 2 \sum_{r,c} Y_{r,c} \ln \left( \frac{Y_{r,c}}{E_{r,c}} \right)$$

First, we need to compute all the expected value

$$\begin{aligned} e_{11} &= \phi_1 \cdot \rho_1 \cdot n = \frac{n_r}{n} \cdot \frac{n_{..c}}{n} \cdot n = \frac{50}{150} \cdot \frac{25}{150} \cdot 150 = \frac{25}{3} \\ e_{12} &= \phi_1 \cdot \rho_2 \cdot n = \frac{50}{150} \cdot \frac{125}{150} \cdot 150 = \frac{125}{3} \end{aligned}$$

$\dots$

Since  $\phi_1 = \phi_2 = \phi_3 = 1/3$ , expected value depends on either  $\rho_c$  and thus every cell in the Monday column has the same expected value and every cell in the Friday column.

In R,

```
observed <- c(3, 47, 5, 45, 17, 33)
# order is {(1,1), (1,2), (2,1), (2,2), (3,1), (3,2)}

a <- 1/18 * 150
# expected value for every cell in the Monday column

b <- 5/18 *150
# expected value for every cell in the Friday column

expected <- c(a, b, a, b, a, b)
wilks <- 2 * sum(observed * log(observed / expected))
print(wilks) # yields 15.85977
```

So we get a test statistic of 15.85977. From the results of the previous questions, the difference in free parameter between the null and alternative hypothesis is  $5 - 3 = 2$ . Thus  $W \sim \chi^2_2$ . The critical value of a 5% confidence level can be obtained via **qchisq(0.95, 2)**. This gives a critical value of 5.991465. As you can see, our test statistics exceed the critical value, so we reject the null hypothesis. Alternatively, if we look at the p-value, we get 0.0007196556 (via R by **2\*pchisq(15.85977, 2, lower.tail = FALSE)**). This means that we would reject for any confidence level above 0.007%.