

Exercise 1: Discrete distributions

①

Let $K \in \mathbb{N}^*$, $X = \{v_1, \dots, v_K\}$ K real numbers. Let

$(\pi_b)_{b \in [1, K]}$ s.t

$$\forall b \in [1, K], \pi_b \geq 0 \text{ and } \sum_{b=1}^K \pi_b = 1$$

1) We want to generate a random variable Y on X that follow the given distribution.

$$\forall b \in [1, K], \mathbb{P}(Y = v_b) = \pi_b$$

So, if we denote Z a random variable taking value in $[1, K]$ and s.t $\mathbb{P}(Z = b) = \pi_b$, we can write:

$$Y = \sum_{b=1}^K v_b \mathbb{1}_{\{Z=b\}} \quad \text{a.s (almost surely)}$$

We want to generate the random variable Z to be able to generate Y .

We will use the inverse function method to do so.

$$\forall t \in [1, K], \quad F_Z(t) := P(Z \leq t) = P(Z \leq \lfloor t \rfloor) \\ = \sum_{b=1}^{\lfloor t \rfloor} \pi_b$$

$$\text{and } \begin{cases} \forall t > K, F_Z(t) = 1 \\ \forall t < 1, F_Z(t) = 0 \end{cases}$$

So, the inverse CDF is: $\forall y \in [0, 1]$,

$$F_Z^{-1}(y) := \min_{\substack{b \in [1, K] \\ F_Z(b) \geq y}} b$$

Finally, if we sample $U \sim \mathcal{U}([0, 1])$, then:

$F_Z^{-1}(U)$ follows the same law as Z .

$$\text{and } Y = \sum_{b=1}^K \vartheta_b \mathbb{1}_{\{Z=b\}} = \sum_{b=1}^K \vartheta_b \mathbb{1}_{\{F_Z(b-1) < U \leq F_Z(b)\}} \text{ a.s.}$$

with convention $F_Z(0) := 0$.

Exercise 2: GMM and the EM algorithm. (2)

1) Let's denote $\alpha = (\alpha_j)_{j \in [1, m]}$, $\mu = (\mu_j)_{j \in [1, m]}$
and $\Sigma = (\Sigma_j)_{j \in [1, m]}$ with $m \in \mathbb{N}^*$ fixed.

Then, $\Theta = (\alpha, \mu, \Sigma) \in \mathbb{R}^m \times \mathbb{R}^{md} \times \mathbb{R}^{md^2}$
with d the dimension of the Gaussians.

Let (x_1, \dots, x_n) the observations of n iid samples
 (X_1, \dots, X_n) of our model. We have:

$\forall \Theta$,

$$\mathcal{L}(x_1, \dots, x_n; \Theta) := \prod_{i=1}^n f_{\Theta}(x_i) = \prod_{i=1}^n \sum_{j=1}^m f_{\Theta}(x_i, z_i=j)$$

$$= \prod_{i=1}^n \sum_{j=1}^m f_{\Theta}(x_i | z_i=j) P(z_i=j)$$

$$= \prod_{i=1}^n \sum_{j=1}^m \alpha_j \frac{e^{-\frac{1}{2}(x_i - \mu_j)^T \Sigma_j^{-1} (x_i - \mu_j)}}{(2\pi)^{d/2} |\Sigma_j|^{1/2}}$$

3) Our goal is to find θ maximizing the following log-likelihood:

$$l(x_1, \dots, x_n | \theta) := \sum_{i=1}^n \log \left(\sum_{j=1}^m \alpha_j f_{\theta}(x_i | z_i = j) \right)$$

This is hard to achieve, so we use EM.

First, the complete data log-likelihood is given by:

$$\begin{aligned} l((x_i)_{i \in [1, n]}, (z_i)_{i \in [1, n]}; \theta) &= \log \left(\prod_{i=1}^n f(x_i, z_i | \theta) \right) \\ &\stackrel{(x_i, z_i) \perp}{=} \sum_{i=1}^n \log (f_{\theta}(x_i | z_i, \theta) f(z_i)) \\ &= \sum_{i=1}^n \sum_{j=1}^m z_{ij} \log (f_{\theta}(x_i, z_i = j, \theta) \alpha_j) \\ &= \sum_{i=1}^n \sum_{j=1}^m z_{ij} \log (\mathcal{N}(x_i; \mu_j, \Sigma_j) \alpha_j) \end{aligned}$$

So,

$$\mathbb{E} [l(x_i)_i, (z_i)_i; \theta | \theta^{(c)}] = \sum_{i=1}^n \sum_{j=1}^m \overbrace{\mathbb{E}[z_{ij} | (x_i)_i, \theta^{(c)}]}^{:= \tau_{ij}} \log (\mathcal{N}(x_i; \mu_j^{(c)}, \Sigma_j^{(c)}) \alpha_j^{(c)})$$

τ_{ij} : probability of observation i to be in cluster j given x_i and the parameters at step c .