

Exercise 1: Discrete distributions.

①

Let $K \in \mathbb{N}^*$, $X = \{v_0, \dots, v_K\}$ K real numbers. Let

$$(\pi_b)_{b \in [1, K]} \text{ s.t.}$$

$$\forall b \in [1, K], \pi_b \geq 0 \text{ and } \sum_{b=1}^K \pi_b = 1$$

1) We want to generate a random variable Y on X that follow the given distribution.

$$\forall b \in [1, K], P(Y = v_b) = \pi_b$$

So, if we denote Z a random variable taking value in $[1, K]$ and s.t $P(Z = b) = \pi_b$, we can write:

$$Y = \sum_{b=1}^K v_b \mathbf{1}_{\{Z=b\}} \quad \text{a.s (almost surely)}$$

We want to generate the random variable Z to be able to generate Y .

We will use the inverse function method to do so.

$$\forall t \in [1, K], \quad F_Z(t) := P(Z \leq t) = P(Z \leq \lfloor t \rfloor) \\ = \sum_{k=1}^{\lfloor t \rfloor} \pi_k$$

and $\begin{cases} \forall t > K, F_Z(t) = 1 \\ \forall t < 1, F_Z(t) = 0 \end{cases}$

So, the inverse CDF is: $\forall y \in [0, 1]$,

$$F_Z^{-1}(y) := \min_{\substack{k \in [1, K] \\ F_Z(k) \geq y}} k$$

Finally, if we sample $U \sim U[0, 1]$, then:

$F_Z^{-1}(U)$ follows the same law as Z .

$$\text{and } Y = \sum_{k=1}^K \vartheta_k \mathbb{1}_{\{Z=k\}} = \sum_{k=1}^K U_k \mathbb{1}_{\{F_Z(k-1) < U \leq F_Z(k)\}}$$

with convention $F_Z(0) = 0$.

Exercise 2: GMM and the EM algorithm. (2)

1) Let's denote $\alpha = (\alpha_j)_{j \in [1, m]}$, $\nu = (\nu_j)_{j \in [1, m]}$ and $\Sigma = (\Sigma_j)_{j \in [1, m]}$ with $m \in \mathbb{N}^*$ fixed.

Then, $\Theta = (\alpha, \nu, \Sigma) \in \mathbb{R}^m \times \mathbb{R}^{md} \times \mathbb{R}^{md^2}$

with d the dimension of the Gaussians.

Let (x_1, \dots, x_n) the observations of n iid samples (X_1, \dots, X_n) of our model. We have:

For,

$$\mathcal{L}(x_1, \dots, x_n; \Theta) := \prod_{i=1}^n f_\Theta(x_i) = \prod_{i=1}^n \sum_{j=1}^m f_\Theta(x_i | z_i=j)$$

$$= \prod_{i=1}^n \sum_{j=1}^m f_\Theta(x_i | z_i=j) \pi_j$$

$= \prod_{i=1}^n \sum_{j=1}^m \alpha_j \frac{e^{-\frac{1}{2}(\mathbf{x}_i - \boldsymbol{\nu}_j)^T \boldsymbol{\Sigma}_j^{-1} (\mathbf{x}_i - \boldsymbol{\nu}_j)}}{(2\pi)^{\frac{d}{2}} |\boldsymbol{\Sigma}_j|^{1/2}}$

3) Our goal is to find θ maximizing the following log-likelihood:

$$l_{(x_1, \dots, x_n)}(\theta) := \sum_{i=1}^n \log \left(\sum_{j=1}^m \alpha_j f_\theta(x_i | z_i=j) \right)$$

This is hard to achieve, so we use EM.

First, the complete data log-likelihood is given by:

$$\begin{aligned} l((x_i)_{i \in [1, n]}, (z_i)_{i \in [1, n]}; \theta) &= \log \left(\prod_{i=1}^n f(x_i, z_i; \theta) \right) \\ &\stackrel{(x_i, z_i) \perp\!\!\!\perp}{=} \sum_{i=1}^n \log \left(f_\theta(x_i | z_i, \theta) f(z_i) \right) \\ &= \sum_{i=1}^n \sum_{j=1}^m z_{ij} \log \left(f_\theta(x_i, z_i=j; \theta_j) \alpha_j \right) \\ &= \sum_{i=1}^n \sum_{j=1}^m z_{ij} \log \left(\mathcal{N}(x_i; \mu_j, \Sigma_j) \alpha_j \right) \end{aligned}$$

So,

$$\mathbb{E} [l((x_i)_i, (z_i)_i; \theta) | \theta^{(c)}] = \sum_{i=1}^n \sum_{j=1}^m \mathbb{E}[z_{ij} | (x_i)_i, \theta^{(c)}] \log \mathcal{N}(x_i; \mu_j^{(c)}, \Sigma_j^{(c)})$$

z_{ij} : probability of observation i to be in cluster j given x_i and the parameters at step c .