

Homework 3 : Part A

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Exercise 1

1. Let $\gamma \in \mathbb{N}^*$ and $X = \{1, \dots, 6\}$. We consider the Hastings-Metropolis algorithm with target π associated to the proposal distribution:

$$q(x, \cdot) = \mathcal{U}(\{X_n - \gamma, X_n - \gamma + 1, \dots, X_n - 1, X_n + 1, \dots, X_n + \gamma\})$$

Algorithm 1 Metropolis Hastings algorithm

Given one initial point in $x_0 \in X$

for $n=1:N$

-Generate $Y_n \sim q(x_n, \cdot)$

- $X_{n+1} = Y_n$ with probability $\alpha(X_n, Y_n)$

and $X_{n+1} = X_n$ with probability $1 - \alpha(X_n, Y_n)$.

end

Since $\forall x \in X, \forall y \in X, q(x, y) = \frac{1}{\gamma}$, the acceptance ratio is equal to

$$\alpha(x, y) = \begin{cases} 1 \wedge \frac{\pi(y)}{\pi(x)} & \text{if } y \in X \\ 0 & \text{otherwise} \end{cases}$$

To avoid technical discussions, we will consider the extended distribution π on \mathbb{N} , with $\pi(k) = 0$ if $k \notin X$ so that :

$$\alpha(x, y) = 1 \wedge \frac{\pi(y)}{\pi(x)}$$

And the transition kernel is equal, for $y \neq x$ to:

$$\begin{aligned} p(x, y) &= q(x, y)\alpha(x, y) \\ &= \frac{1}{\gamma} \left(1 \wedge \frac{\pi(y)}{\pi(x)}\right) \end{aligned}$$

and :

$$p(x, x) = 1 - \sum_{y \neq x} q(x, y)\alpha(x, y)$$

First :

$$\begin{aligned}
 p(x, y \neq x) &= \frac{1}{\gamma} (1 \wedge \frac{\pi(y)}{\pi(x)}) \\
 &\geq \frac{1}{\gamma} (1 \wedge \frac{\pi(y)}{\max_x \pi(x)}) \\
 &= \frac{\pi(y)}{\gamma \max_x \pi(x)} \\
 &\geq \frac{\pi(y)}{\gamma}
 \end{aligned}$$

The third equality is due to $\frac{\pi(y)}{\max_x \pi(x)} \leq 1$. The fourth inequality is due to $\max_x \pi(x) \leq 1$.

Second,

$$\begin{aligned}
 p(x, x) &= 1 - \sum_{y \neq x} q(x, y) \alpha(x, y) \\
 &= q(x, x) + \sum_{y \neq x} q(x, y) - \sum_{y \neq x} q(x, y) \alpha(x, y) \\
 &= q(x, x) + \sum_{x \neq y} q(x, y) (1 - \alpha(x, y)) \\
 &\geq q(x, x) \\
 &= \frac{1}{\gamma} \\
 &\geq \frac{\pi(x)}{\gamma}
 \end{aligned}$$

The second inequality is obtained by upper bounding the sum terms by 1, and using $|X| = 6$.

Let $\epsilon = \frac{1}{\gamma}$, then

$$\forall x, \forall y \quad p(x, y) \geq \epsilon \pi(y)$$

which can be written in terms of subsets B of X :

$$p(x, B) \geq \epsilon \pi(B)$$

2. Given the inequality written above, the algorithm satisfies the Doeblin condition. The Doeblin Lemma gives : $\Delta(P) \leq (1 - \epsilon)$

Given one initial distribution ξ , we have in the general case:

$$\|\xi P^n - \pi\|_{TV} \leq (\Delta(P))^n \|\xi - \pi\|_{TV}$$

which yields here :

$$\|\xi P^n - \pi\|_{TV} \leq (1 - \frac{1}{\gamma})^n \|\xi - \pi\|_{TV}$$

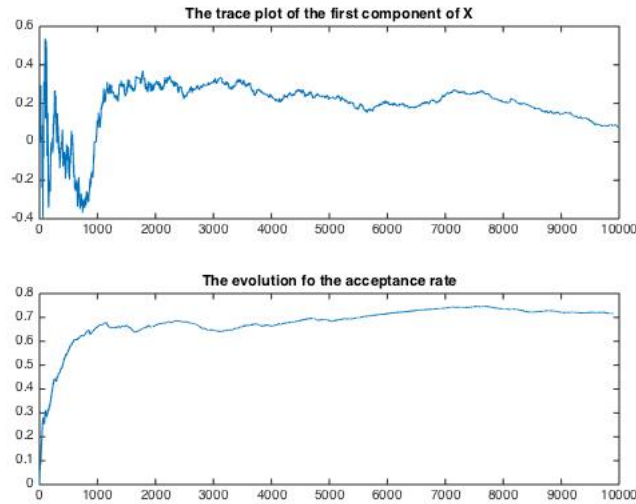


Figure 1: The naive symmetric random walk Hastings-Metropolis chain

Where ξP^n is actually the distribution of X_n . We have a geometric ergodicity

3.

Exercise 3

1.

2.

Exercise 3

We choose for the numerical simulation the parameters:

$$N : = 1000$$

$$c : = 2$$

$$n_0 : = 100$$

1. In figure 1, we see how the sequence X_1^n varies too much and takes a long time to converge to the mean which is 0 in this case. The second plot shows that the acceptance ratio also converges, slowly though, up to around 70%. So there is room to do better.

2. In figure 2, we see that the sequence X_n converges abruptly to the mean after n_0 . The second plot shows that the acceptance ratio also converges rapidly up to around 100%.

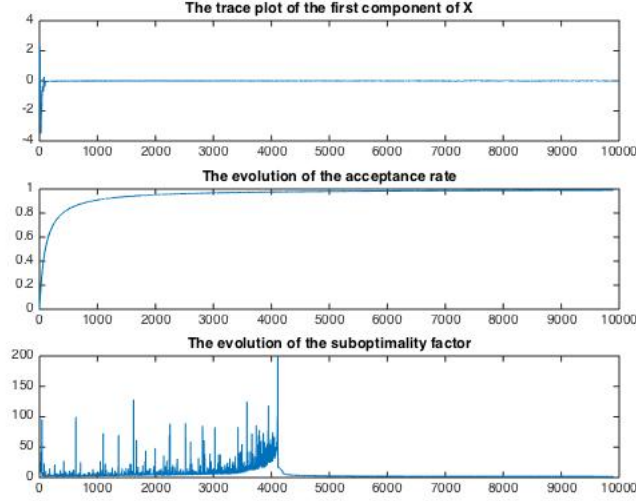


Figure 2: The adaptive symmetric random walk Hastings-Metropolis chain

Further, we can see that the suboptimality factor depend on n in 2 ways. First, it varies a lot and stays strictly over 1 (the optimal case). In the second phase, it starts decreasing very slowly to 1.

3. The suboptimality factor of the first algorithm is constant independent from c :

$$n \mapsto d \frac{\text{trace}(\Gamma_{\pi}^2)}{\text{trace}(\Gamma_{\pi})^2} > 1$$

The adaptive algorithm yields a sequence that adapts and thus convergences, slowly, to the optimal case. In conclusion, the adaptive random walk is much better than the naive one.