#### Nonlinear models

Bayesian statistics 8 - dynamic and nonlinear model fitting

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### Things that we learned the last time

• logistic regression with  $y_i \sim \mathcal{B}(n, p_i)$  with  $\log \operatorname{id}(p_i) = \log(\frac{p_i}{1-p_i}) = a + bx_i$  or  $\gamma(x_i - \mu_x)$  which is equivalent to  $p_i = \operatorname{logistic}(a + bx_i) = \frac{\exp(a + bx_i)}{1 + \exp(a + bx_i)} = \frac{1}{1 + \exp(-(a + bx_i))}$ 

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- "logistic ANOVA"  $p_i = \text{logistic}(\alpha + \beta_1 x_{1i} + \beta_2 x_{2i} + \beta_3 x_{3i})$ . Here the logistic function maps  $(-\infty, +\infty) \rightarrow [0, 1]$ .

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- We can also want to use the logistic function to model a known curve,  $y_i = f(x_i) + \epsilon_i$  or  $y_i = f(t_i) + \epsilon_i$

#### Organism growth basics: von Bertalanffy

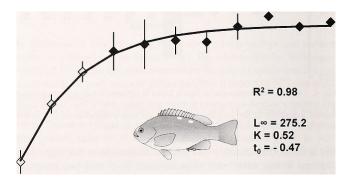


Figure 1: Von Bertalanffy growth curve fit to Girella nigricans

$$L(t) = L_{\infty}(1 - \exp(-k(t_i - t_0))) + \epsilon_i$$

#### Connection to dynamics

$$\frac{dL}{dt} = k(L_{\infty} - L)$$

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$$\frac{dL}{dt}=k(L_{\infty}-L)$$

We'll do for practical another example of organismal growth, Gompertz growth (Winsor PNAS 1932).

**Dynamics** 

$$\frac{d\ln(L)}{dt} = k(\ln(L_{\infty}) - \ln(L))$$

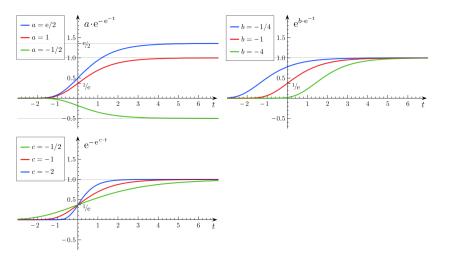
Solution

$$L(t) = ae^{-e^{b-ct}}$$

where  $a = L_{\infty}e^1$ ,  $b = kt_0$  and c = k.

Trick: note I = ln(L), solve von Bertalanffy for I, transform back.

#### The Gompertz growth curve is more logistic-like



## Mathematical cousins of von Bertalanffy: modelling saturation

Monod function (microbiology) aka Michalis-Menten (chemistry) aka Holling type II (ecology) aka . . .

$$f(x) = \frac{ax}{b+x}$$

# Another example of connection to dynamics: logistic population growth

N = population size (microbes, humans, wild boars, plants,...)

$$\frac{dN}{dt} = rN\left(1 - \frac{N}{K}\right)$$

And the solution is. . .

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And the solution is...

$$N(t) = \frac{N_0 e^{rt}}{1 + N_0 (e^{rt} - 1)/K}$$

If we use t = 1, 2, 3, ...

$$\textit{N}_1 = \frac{\textit{N}_0 e^r}{1 + \textit{N}_0 (e^r - 1) / \textit{K}}, \textit{N}_2 = \frac{\textit{N}_1 e^r}{1 + \textit{N}_1 (e^r - 1) / \textit{K}}, ...$$

aka Beverton-Holt model.

Observational noise

$$N_t = rac{N_t e^r}{1 + N_t (e^r - 1)/K} + \epsilon_t, \; \epsilon_t \sim \mathcal{N}(0, \sigma^2)$$
 i.i.d

• The effect of  $\epsilon_t$  does not accumulate.  $\epsilon_1$  does not affect  $N_8$ .

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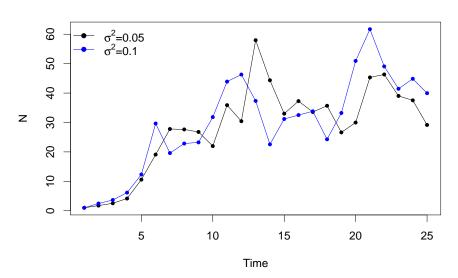
- The effect of  $\epsilon_t$  does accumulate.  $\epsilon_1$  does affect  $N_8$ .
- $\bullet$  Creates autocorrelation between  $N_t$  values
- Of course in real life you can have both (and sometimes it is hard to distinguish between the two)

### Fitting the logistic model in discrete-time with process noise I

(you can fit the observational noise model with the solution N(t) – you can't fit the process noise model solution, you have to fit the *dynamics*)

```
r=0.5
alpha=0.02
tmax=25
R=exp(r)
K=(\exp(r)-1)/alpha
N_BH=N_BH1=rep(NA,tmax)
N_BH[1] = N_BH1[1] = 1
for (t in 1:(tmax-1)){N BH[t+1]} =
  (exp(r+rnorm(1,0,sqrt(0.05))))*N_BH[t]/(1+alpha*N_BH[t])
for (t in 1:(tmax-1)){N BH1[t+1] =
  (\exp(r+r_{0.1}))*N_BH1[t]/(1+alpha*N_BH1[t])
###
par(pch=20, cex=1.5)
plot(1:tmax,N_BH,type="o",ylim=range(c(N_BH,N_BH1)),xlab="Time",ylab="N")
lines(1:tmax, N_BH1, type="o", col="blue")
legend("topleft",c(expression(paste(sigma^"2","=0.05",sep="")),
                   expression(paste(sigma^"2", "=0.1", sep=""))),
       col=c("black","blue"),lty=1,pch=16,bty="n")
```

## Fitting the logistic model in discrete-time with process noise II



#### Let's fit the model

```
logistic.data <- list(logN = log(N_BH), tmax=tmax)</pre>
cat(file="logistic.growth.txt","
model {
  r \sim dnorm(2, 0.01) ## prior on r
  alpha ~ dlnorm(1, 0.01) ## prior on alpha
  K < -(exp(r)-1)/alpha
  sigma ~ dunif(0.01,2)
  tau<-pow(sigma,-2)
  logN[1] \sim dnorm(0,1)
  N[1] < -exp(logN[1])
  #Likelihood
  for (t in 1:(tmax-1)){
  logNpred[t] <- logN[t]+ r - log(1 + alpha*N[t])</pre>
  logN[t+1] ~ dnorm(logNpred[t],tau)
  N[t+1] \leftarrow exp(logN[t+1])
```

#### Running the model I

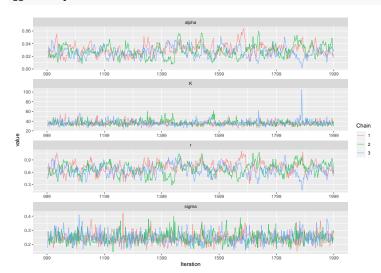
```
# Inits function
inits \leftarrow function(){list(r = rnorm(1, 0, 1),
                         alpha = rlnorm(1,0,1))
# Parameters to estimate
params <- c("r","alpha","K","sigma")</pre>
# MCMC settings
nc <- 3 ; ni <- 2000 ; nb <- 1000 ; nt <- 2
# Call JAGS, check convergence and summarize posteriors
out <- jags(logistic.data, inits, params, "logistic.growth.txt", n.thin = nt,
            n.chains = nc. n.burnin = nb. n.iter = ni)
## Compiling model graph
##
      Resolving undeclared variables
      Allocating nodes
##
## Graph information:
##
      Observed stochastic nodes: 25
     Unobserved stochastic nodes: 3
##
##
      Total graph size: 183
##
## Initializing model
```

#### Running the model II

```
print(out, dig = 3)  # Bayesian analysis
## Inference for Bugs model at "logistic.growth.txt", fit using jags,
   3 chains, each with 2000 iterations (first 1000 discarded), n.thin = 2
## n.sims = 1500 iterations saved
           mu.vect sd.vect 2.5% 25% 50% 75% 97.5% Rhat n.eff
##
## K
          36.144 5.224 28.599 32.810 35.348 38.596 48.108 1.006
                                                                  440
## alpha
          0.028 0.009 0.012 0.022 0.028 0.034 0.048 1.042 56
## r
          0.683 0.140 0.399 0.591 0.684 0.775 0.950 1.041 55
## sigma 0.245 0.040 0.179 0.218 0.242 0.268 0.338 1.000 1500
## deviance 0.842 2.927 -2.685 -1.258 0.095 2.263 8.236 1.001 1500
##
## For each parameter, n.eff is a crude measure of effective sample size,
## and Rhat is the potential scale reduction factor (at convergence, Rhat=1).
##
## DIC info (using the rule, pD = var(deviance)/2)
## pD = 4.3 and DIC = 5.1
## DIC is an estimate of expected predictive error (lower deviance is better).
```

#### Showing traceplots

```
S<-ggs(as.mcmc(out)) #R2jags
S<-filter(S,Parameter != "deviance")
ggs_traceplot(S)</pre>
```



### Showing correlations (r,K) and (r, $\alpha$ )

ggs\_pairs(S)

