

## 4. Linear Model Selection and Regularization

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[jeshan49.github.io/eemp2/](https://jeshan49.github.io/eemp2/)

- Lecture<sup>1</sup>:
  - Subset Selection
  - Shrinkage/Regularization
  - Dimension Reduction
- Tutorial:
  - Reproducing results from the lecture using:
    - Forward/Backward Subset Selection
    - Ridge and Lasso Regression
    - Principal Component and Partial Least Squares Regression

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<sup>1</sup>Some of the figures in this presentation are taken from “An Introduction to Statistical Learning, with applications in R” (Springer, 2013) with permission from the authors: G. James, D. Witten, T. Hastie and R. Tibshirani

- In the regression setting, the standard linear model

$$Y = \beta_0 + \beta_1 X_1 + \dots \beta_p X_p + \varepsilon \quad (1)$$

is commonly used to describe the relationship between the response and the input

- This linear model has an obvious advantage compared to non-linear methods in terms of *model interpretability*
- In addition, it is surprisingly competitive in relation to non-linear methods in many settings in terms of *prediction accuracy*
- Today, we will discuss alternative fitting strategies to least squares that may improve the fit

# Why use alternative fitting strategies over least squares?

Let us first consider *prediction accuracy*:

- Recall the bias-variance trade-off: In general, too simple (complex) models will have high (low) bias and low (high) variance
- Suppose now that the true relationship between the input and output is approx. linear
- Then, our linear model in (1) will have low bias. In addition, if  $N \gg p$ , it will have low variance as well
- However, if  $N > p$ , there is a lot of variability in the fit, and hence high variance. In addition, if  $N < p$ , this variance is infinite
- By constraining or shrinking the coefficients, we can reduce the variance substantially at the cost of a small increase in bias

Let us now consider *model interpretability*:

- Often some or many of the  $p$  predictors are not associated with the response
- Including these predictors leads to unnecessary in the resulting model
- This is because the least square fit is extremely unlikely to yield exact zero coefficients
- By setting some of the coefficients to zero, we obtain a more easily interpretable model
- We will thus consider methods that automatically perform feature selection

# Three alternative methods

Today we'll discuss three alternative methods:

**1** Subset selection:

- Identify a subset of the  $p$  predictors that we believe are related to  $Y$ . Then we fit using least squares on this subset

**2** Shrinkage:

- Fit on all  $p$  predictors using least squares subject to a constraint on the size of the coefficients
- This shrinkage/regularization reduces the variance

**3** Dimension reduction:

- Projecting the  $p$  predictors into a  $M$ -dimensional subspace, where  $M < p$
- We then fit the model with the  $M$  predictors using least squares

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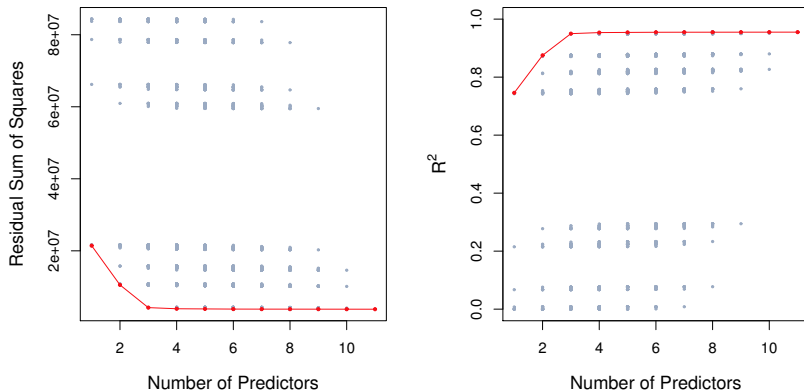
**Algorithm 1** Best subset selection

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- 1: Let  $\mathcal{M}_0$  denote the null model, which contains no predictors.
  - 2: **for**  $k = 1$  to  $p$  **do**
  - 3:   (a) fit all  $\binom{p}{k}$  models that contain exactly  $k$  predictors.
  - 4:   (b) Pick the best among these  $\binom{p}{k}$  models, and call it  $\mathcal{M}_k$ .  
Here the best is defined as having the smallest RSS or highest  $R^2$ .
  - 5: **end for**
  - 6: Select a single best model from  $\mathcal{M}_0, \dots, \mathcal{M}_p$  using CV,  $C_p$ ,  $AIC$ ,  $BIC$  or Adj.  $R^2$ .
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# Example



**Figure:** For each possible model containing a subset of ten predictors in the Credit data set, the RSS and  $R^2$  are displayed. The red frontier tracks the best model for a given number of predictors. (See ISLR p. 206)

# Some notes on best subset selection

- The same idea of best subset selection can be applied to a wide array of models, e.g. logistic regression
- While the method is simple, it suffers from computational limitations:
  - If  $p = 10$  we must fit  $2^{10} = 1,024$  models
  - If  $p = 20$ , we must fit  $2^{20} = 1,048,576$  models
- Thus, best subset selection becomes unfeasible for  $p > 40$
- In addition, the method may suffer from overfitting and high variance of coefficient estimates for large  $p$
- We will now consider more computationally efficient compromises, with a smaller search space - *Stepwise selection*

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**Algorithm 2** Forward stepwise selection

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- 1: Let  $\mathcal{M}_0$  denote the null model, which contains no predictors.
  - 2: **for**  $k = 0$  to  $p - 1$  **do**
  - 3:   (a) Consider all  $p - k$  models that augment the predictors in  $\mathcal{M}_k$  with one additional predictor.
  - 4:   (b) Pick the best among these  $p - k$  models, and call it  $\mathcal{M}_{k+1}$ .  
Here the best is defined as having the smallest RSS or highest  $R^2$ .
  - 5: **end for**
  - 6: Select a single best model from  $\mathcal{M}_0, \dots, \mathcal{M}_p$  using CV,  $C_p$ ,  $AIC$ ,  $BIC$  or Adj.  $R^2$ .
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# Some notes on forward stepwise selection

- Forward stepwise selection has a clear computational advantage over best subset selection:
  - For  $p = 20$ , the latter fits  $2^{20} = 1,048,576$  models whereas the former only fits  $1 + \sum_{k=0}^{20-1} (20 - k) = 1 + 20(20 + 1)/2 = 211$  models
- Furthermore, it may perform better due to its smaller search space
- However, it may fail to selection the best model
  - Suppose that  $p = 3$  and the best model is a two-variable model with  $X_2, X_3$
  - If the best one-variable model is with  $X_1$ , then forward stepwise selection will fail in finding the best model

# Backward stepwise selection

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**Algorithm 3** Backward stepwise selection

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- 1: Let  $\mathcal{M}_p$  denote the full model, which contains all  $p$  predictors.
  - 2: **for**  $k = p$  to 1 **do**
  - 3:   (a) Consider all  $k$  models contain all but one of the predictors in  $\mathcal{M}_k$ , for a total of  $k - 1$  predictors.
  - 4:   (b) Pick the best among these  $k$  models, and call it  $\mathcal{M}_{k-1}$ . Here the best is defined as having the smallest RSS or highest  $R^2$ .
  - 5: **end for**
  - 6: Select a single best model from  $\mathcal{M}_0, \dots, \mathcal{M}_p$  using CV,  $C_p$ ,  $AIC$ ,  $BIC$  or Adj.  $R^2$ .
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# Some notes on backward stepwise selection

- Like forward stepwise selection, backward stepwise selection only fits  $1 + p(p + 1)/2$  models
  - Thus, it can be used with large  $p$
- However, once again, it is not guaranteed to select the best model containing a subset of  $p$  predictors
- Furthermore, backward stepwise selection can only be used in settings in which  $N > p$
- In contrast, we can always use forward selection up to a particular number of predictors

# Choosing the optimal model

- Each preceding method yield an optimal model for  $1, \dots, p$
- Thus, at the end, we need to select one best model from these candidates
- We do not want to use  $RSS$  nor  $R^2$  as they are directly related to the training error
  - as we know, the training error can be a poor estimate of the test error
- Thus, we consider two alternative approaches:
  - 1 Indirectly estimate test error by making an adjustment to the training error
  - 2 Directly estimate the test error by using the validation set or the cross-validation approach

## $C_p$ , $AIC$ , $BIC$ , and Adj. $R^2$

- Let  $d$  be the # of predictors and  $\hat{\sigma}^2 = RSS/(N - p - 1)$  an estimate of  $Var[\varepsilon]$  from the full model

Mallow's  $C_p$ :

$$C_p = \frac{1}{N}(RSS + 2d\hat{\sigma}^2) \quad (2)$$

Akaike information criterion:

$$AIC = \frac{1}{N\hat{\sigma}^2}(RSS + 2d\hat{\sigma}^2) = \frac{1}{\hat{\sigma}^2} * C_p \quad (3)$$

Bayesian information criterion:

$$BIC = \frac{1}{N\hat{\sigma}^2}(RSS + \log(N)d\hat{\sigma}^2) \quad (4)$$

Adjusted  $R^2$ :

$$\text{Adj. } R^2 = 1 - \frac{RSS/(N - d - 1)}{TSS/(N - 1)} \quad (5)$$



# Example (Best subset selection)

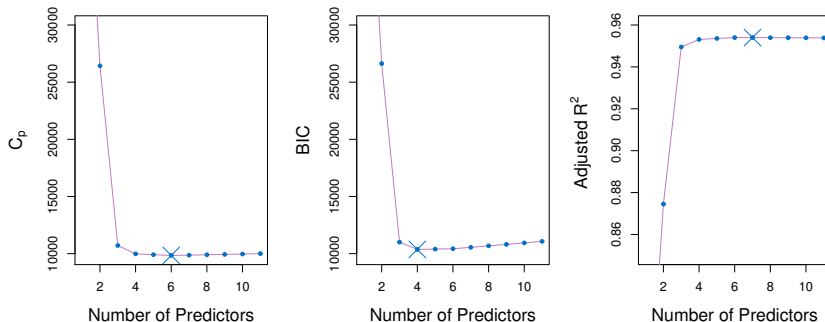
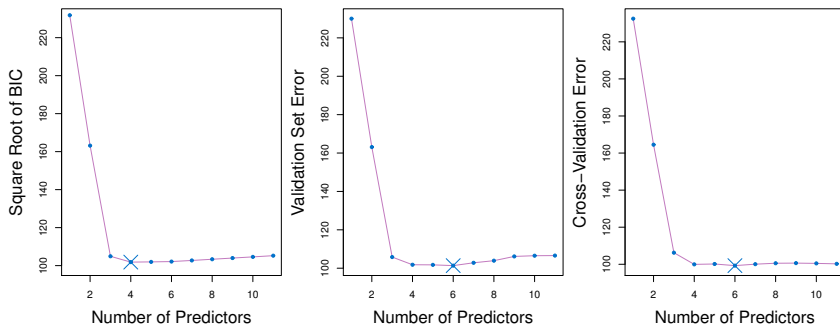


Figure:  $C_p$ ,  $BIC$ , and Adj.  $R^2$  for the best models of each size of the Credit data set. (See ISLR p. 211)

- $C_p$ : *income, limit, rating, cards, age, student*
- $BIC$ : *income, limit, cards, student*
- Adj.  $R^2$ : *income, limit, rating, cards, age, student, gender*

# Validation and Cross-Validation

- As we already know, we can also use the validation set approach or k-fold CV for the task of model selection



**Figure:** Credit data set. Left: Square root of BIC, Center: Validation set errors, Right: 10-fold CV errors (See ISLR p. 214)

# Shrinkage methods - Ridge Regression

- As an alternative to subset selection, we can fit a model on all  $p$  predictors, whilst shrinking the coefficients toward zero
- We will see that this can reduce the variance of our model

Ridge regression solves:

$$\min_{\beta} \sum_{i=1}^N \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j^2 x_{ij} \right)^2 + \lambda \sum_{j=1}^p \beta_j^2 = RSS + \lambda \sum_{j=1}^p \beta_j^2 \quad (6)$$

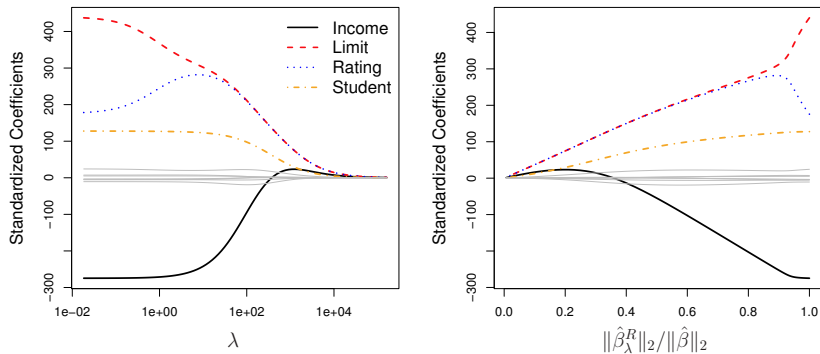
where  $\lambda \geq 0$  is a tuning/hyper parameter; it controls the magnitude of regularization vs. fit

- We find an estimate  $\hat{\beta}_{\lambda}^R$  for many  $\lambda$ , and then choose the optimal  $\lambda$  by CV - This is not computational expensive

# Notes on Ridge regression

- Shrinkage is applied to  $\beta_1, \dots, \beta_p$  but not to  $\beta_0$ 
  - This is because we want to shrink the estimated association of each variable with the response
- Standard least square coefficient estimates are scale invariant
  - $X_j \hat{\beta}_j$  will remain the same
- Ridge coefficient estimates can change substantially
  - $X_j \hat{\beta}_{j,\lambda}^R$  may not only depend on  $\lambda$  and its predictor's scale, but also on other predictors' scale
- Thus, it is best to apply ridge regression after standardizing the predictors

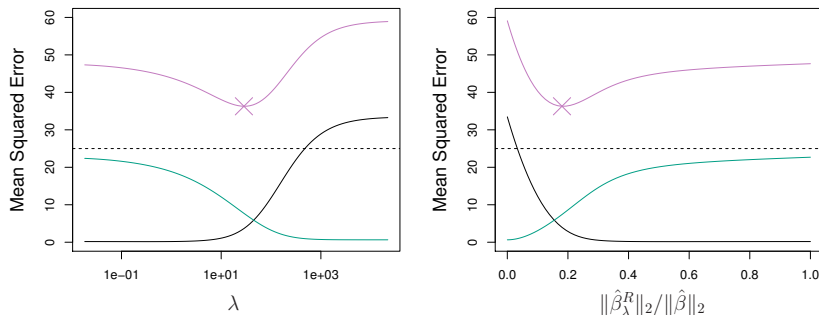
# Example



**Figure:** The standardized ridge regression coefficients are displayed for the Credit data set, as a function of  $\lambda$  and  $\|\hat{\beta}_{\lambda}^R\|_2 / \|\hat{\beta}\|_2$  (See ISLR p. 216)

# Why does Ridge regression improve over Least squares?

- As  $\lambda$  increases, the flexibility of the fit decreases, leading to decreased variance but increased bias



**Figure:** Squared bias (black), variance (green), and test MSE (purple) for the ridge regression predictions on a simulated data set, as a function of  $\lambda$  and  $\|\hat{\beta}_\lambda^R\|_2 / \|\hat{\beta}\|_2$ . The dashed line indicates minimum MSE (See ISLR p. 218)

# Shrinkage methods - Lasso

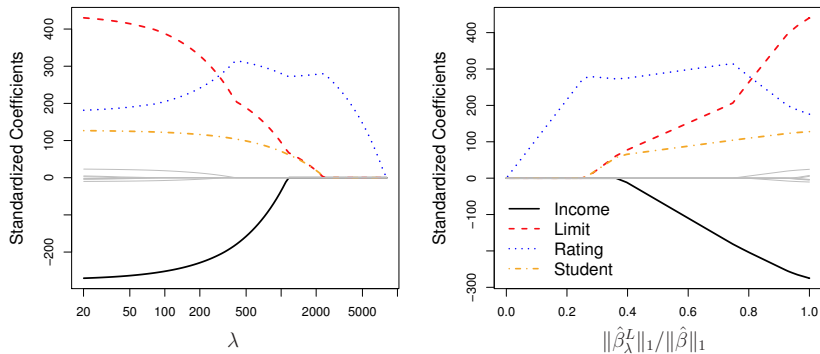
- Ridge regression has one obvious limitation: None of the coefficient estimates will be exactly zero, unless  $\lambda = \infty$

Lasso solves:

$$\min_{\beta} \sum_{i=1}^N \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j^2 x_{ij} \right)^2 + \lambda \sum_{j=1}^p |\beta_j| = RSS + \lambda \sum_{j=1}^p |\beta_j| \quad (7)$$

- The penalty term of the lasso has the effect of setting coefficient estimates exactly zero for finite  $\lambda$
- We say that the lasso yields sparse models – models that involve only a subset of the variables
- We find an estimate  $\hat{\beta}_{\lambda}^L$  for many  $\lambda$ , and then choose the optimal  $\lambda$  by CV

# Example



**Figure:** The standardized lasso coefficients are displayed for the Credit data set, as a function of  $\lambda$  and  $\|\hat{\beta}_\lambda^L\|_1 / \|\hat{\beta}\|_1$  (See ISLR p. 220)



# Another representation for Ridge regression and Lasso

- One can show that (i) *best subset selection*, (ii) *ridge regression*, and (iii) *the lasso* can be formulated as follows

Best subset selection:

$$\min_{\beta} \left\{ \sum_{i=1}^N \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j^2 x_{ij} \right)^2 \right\} \text{ s.t. } \sum_{j=1}^p \mathbb{I}(\beta_j \neq 0) \leq s \quad (8)$$

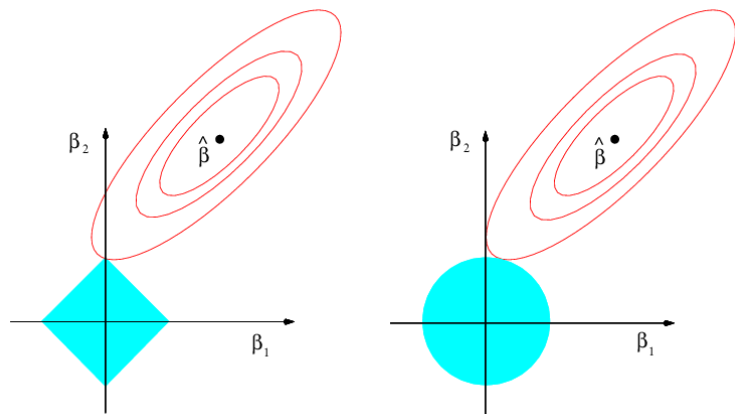
Ridge regression:

$$\min_{\beta} \left\{ \sum_{i=1}^N \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j^2 x_{ij} \right)^2 \right\} \text{ s.t. } \sum_{j=1}^p \beta_j^2 \leq s \quad (9)$$

Lasso:

$$\min_{\beta} \left\{ \sum_{i=1}^N \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j^2 x_{ij} \right)^2 \right\} \text{ s.t. } \sum_{j=1}^p |\beta_j| \leq s \quad (10)$$

# The variable selection property of Lasso



**Figure:** Contours of the error and constraint functions for the lasso (left) and ridge regression (right). The solid blue areas are the constraint regions,  $|\beta_1| + |\beta_2| \leq s$  and  $\beta_1^2 + \beta_2^2 \leq s$  (See ISLR p. 222)

# Comparing Ridge regression and Lasso

- Lasso has an advantage over ridge regression in terms of model interpretability
- But what about prediction accuracy?
- Lasso implicitly assumes that some of the predictors are unrelated to the response
- If this assumption holds, then the lasso can perform better. If not, then ridge regression will in general perform better
- It is possible to combine both approaches in one method
  - ElasticNet regression: a convex combination of ridge regression and lasso

# Dimension reduction methods

- Methods that transform the  $p$  predictors onto  $M$  dimensions and then fit a least square model on the transformed variables
- Let  $Z_1, \dots, Z_M$  represent  $M < p$  linear combinations of the  $p$  predictors

$$Z_m = \sum_{j=1}^p \phi_{jm} X_j \quad \forall m \quad (11)$$

for some constants  $\phi_{1m}, \dots, \phi_{pm}$

- We then fit the linear regression model by least squares:

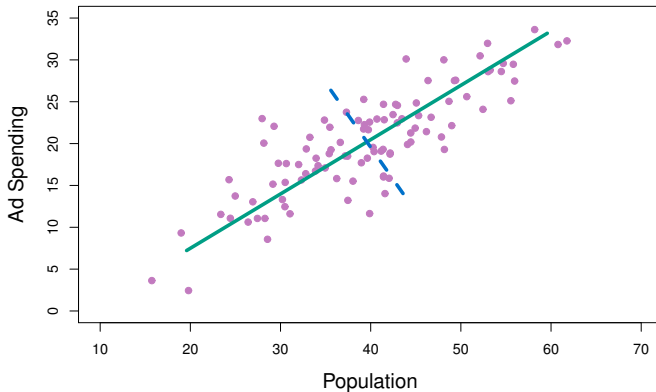
$$y_i = \theta_0 + \sum_{m=1}^M \theta_m z_{im} + \varepsilon_i, i = 1, \dots, n \quad (12)$$

- With this approach, we are reducing the dimension of the problem from  $p + 1$  to  $M + 1$

# Principal Component Regression (PCR)

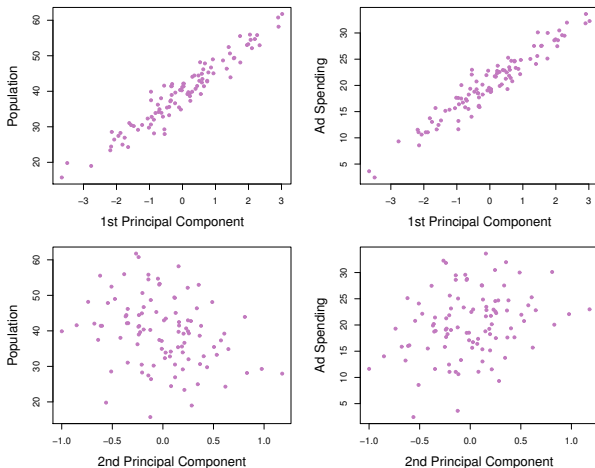
- To perform PCR, we apply principal component analysis (PCA)
  - PCA is a technique for reducing the dimension of our data
- Our goal: Ending up with  $M < p$  principal components which summarizes the majority of the variation in our data
- **Assumption:** The direction in which  $X_1, \dots, X_p$  show the most variation are the directions that are associated with the response
- 1st PC ( $Z_1$ ): The linear combination of predictors with the largest variance
  - Or: the line that is as close as possible to the data
- $i + 1$ th PC ( $Z_{i+1}$ ): The linear combination of predictors with the largest variance subject to being uncorrelated with the  $i$ th PC
  - Or: The line that is as close as possible to the data subject to being orthogonal to the  $i$ th PC

# Example



**Figure:** The population size and ad spending for 100 different cities are shown as purple circles. First PC (green), second PC (blue, dashed) (See ISLR p. 230)

# Example (cont'd)



**Figure:** Plots of the first (top) and second (bottom) PC scores vs. population (left) and ad spending (right) (See ISLR pp. 233-234)

# Partial Least Squares (PLS)

- PCR identifies  $Z_1, \dots, Z_M$  in an unsupervised way – i.e. without considering the response
- Thus, it may be that the directions/components are not the best predictors of the response
- Unlike PCR, PLS defines  $Z_1, \dots, Z_M$  in a supervised way:
- PLS computes  $Z_1$  by defining each  $\phi_{j1}$  from (11) to be the coefficient from the simple linear regression of  $Y$  onto  $X_j$ 
  - The coefficient is proportional to the correlation
  - Thus, PLS places most weight on variables that are strongly correlated with the response
- Subsequent directions are found by taking residuals and repeating the process



- With both PCR and PLS, we standardize the predictors before applying the methods
- With both PCR and PLS, we generally locate the optimal number of directions by cross-validation
- In general, the supervised dimension reduction by PLS can reduce bias
- However, this can come at the cost of an increase in variance
- PCR, PLS and ridge regression performs in practice similar in terms of prediction accuracy

James, G., Witten, D., Hastie, T., & Tibshirani, R. (2013). An introduction to statistical learning (Vol. 112). **Chapter 6**