

Ex Suppose a radio tube is tested.

Let probability of positive test =  $3/4$

Now test =  $1/4$

If a large supply TN a positive tube appears.

$X$  = No. of tests required to terminate the test

$$S = \{+, -+, --+, ---+, \dots\}$$

$$\therefore P(n) = P(X=n) = \left(\frac{1}{4}\right)^{n-1} \left(\frac{3}{4}\right)$$

Check  $\sum p(n) = \frac{3}{4} \left(1 + \frac{1}{4} + \frac{1}{16} + \dots\right)$   
 $= \frac{3}{4} \cdot \frac{1}{1 - 1/4} = 1.$

Expected value of a M

$$E(X) = \sum_{i=1}^{\infty} x_i p(x_i) \quad | \quad E(X) = \int_{-\infty}^{\infty} x f(x) dx$$

Variance of a R.V.

$$V(x) = E[(X - E(x))^2]$$

$\sqrt{V(x)}$  is called standard dev of  $x$ .

## Binomial distribution

$$D = 0.2$$

$$N = 0.8.$$

choose 3 items.

$$\{ S = \{ DDD, DDN, DND, DNN, NDD, NDN, NND, NNN \} \}$$

$$\text{or } S_1 \times S_2 \times S_3$$

$$S_1 = \{ D, N \}.$$

How many defective items were found?

$DDD$	$\rightarrow (0.2)^3$	$P[X=3]$
$DDN$	$\rightarrow (0.2)^2(0.8)$	$P[X=2]$
$DND$	$\rightarrow (0.2)^2(0.8)$	$P[X=2]$
$DNN$	$\rightarrow (0.2)(0.8)^2$	$P[X=1]$
$NDD$	$\rightarrow (0.8)(0.2)^2$	$P[X=1]$
$NDN$	$\rightarrow (0.8)^2(0.2)$	$P[X=1]$
$NND$	$\rightarrow (0.8)^2(0.2)$	$P[X=1]$
$NNN$	$\rightarrow (0.8)^3$	$P[X=0]$

$$\therefore P(0) = (0.8)^3 = 0.512$$

$$P(1) = (0.2)(0.8)^2 = 0.128 \times 3$$

$$P(2) = (0.2)^2(0.8) = 0.032 \times 3$$

$$P(3) = (0.2)^3 = 0.008$$

$$\sum P(x) = 1.$$

$$\text{Also } (0.8 + 0.2)^3 = 1$$

Binomial r.v.

$$P(X) = np$$

Let  $\mathcal{E}$  be an exp. Let  $A$  be an event of

A.S. Let  $P(A) = p$  and  $P(\bar{A}) = 1-p$ . Consider  $n$  independent repetitions of  $\mathcal{E}$  st.  $P(A) = p$  for all repetitions.

Let the r.v.  $X$  be :

$X = \text{No. of times event } A \text{ occurs}$ .

Then  $X$  is Bin r.v. and individual exp. is called Bernoulli trial.

Then  $P(X=k) = P(k) = {}^n C_k p^k (1-p)^{n-k}$

$$k = 0, 1, 2, \dots, n$$

The separately shots to be called Bin. shot.

Ex. A radio tube is inserted in a set. Prob. of a tube functioning  $> 500$  hrs. = 0.2. 20 tubes are tested. What is the probability that exactly 'k' of these will function for more than 500 hrs.

Then  $P(X=k) = {}^{20} C_k (0.2)^k (0.8)^{n-k}$

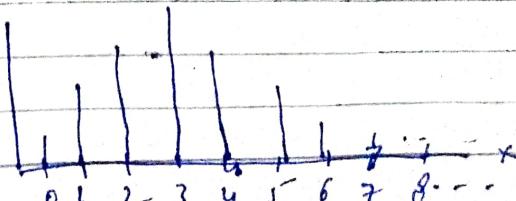
Work out :  $P(X=0) = 0.012$

$$P(X=1) = 0.058$$

:

$$P(X=10) = 0.002.$$

etc



## Continuous RV

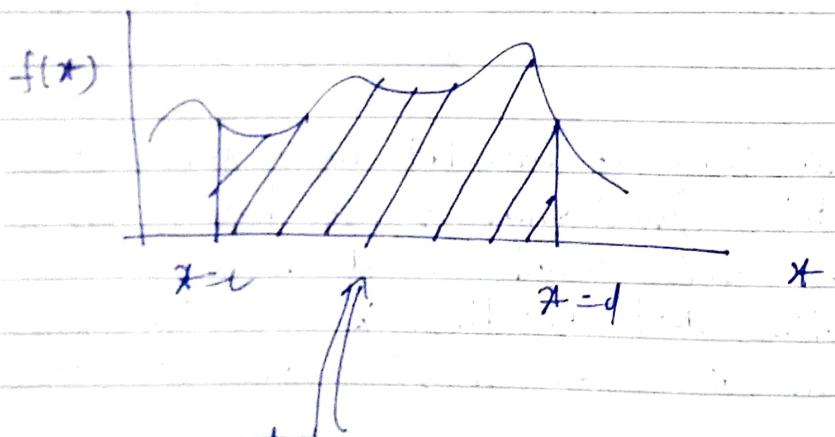
$X$  is said to be a continuous RV if a function 'f', called the probability density  $f(x)$  or  $f(x) dx$  of  $X$ , satisfying the following conditions:

(i)  $f(x) \geq 0 \quad \forall x$

(ii)  $\int_{-\infty}^{\infty} f(x) dx = 1$

(iii) If  $a, b \in (-\infty, \infty)$  &  $a < b$

$$P(a \leq X \leq b) = \int_a^b f(x) dx.$$



$$\int_c^d f(x) dx = \text{Area under pdf}$$

Note  $\int_c^c f(x) dx = 0 \quad \forall c$

Ex: Let p.d.f of  $X$  if  $f(x)$  is given as:

$$f(x) = 2x, \quad 0 < x < 1$$

= 0 Elsewhere.

Clearly  $f(x) \geq 0 \wedge \int_0^1 2x dx$

$$= \left[ 2 \cdot \frac{x^2}{2} \right]_0^1$$

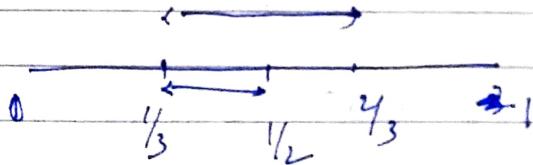
$$= [x^2]_0^1 = 1.$$

$\therefore f(x)$  is a p.d.f.

Find  $P(X \leq \frac{1}{2})$ .  $= \int_0^{\frac{1}{2}} 2x dx = [x^2]_0^{\frac{1}{2}}$   
 $= \frac{1}{4}$ .

Ex Find  $P(X \leq \frac{1}{2} | \frac{1}{3} \leq X \leq \frac{2}{3})$

=



$$\frac{\int_{\frac{1}{3}}^{\frac{1}{2}} 2x dx}{\int_{\frac{1}{3}}^{\frac{2}{3}} 2x dx} = \frac{\frac{5}{36}}{\frac{1}{3}} = \boxed{\frac{15}{12}}$$

## Cumulative Distribution Function (CDF)

Let  $X$  be a r.v. CDF of  $X$ ,  $F(x)$  is defined as

$$F(x) = P(X \leq x)$$

for Disc. R.V

$$F(x) = \sum_j p(x_j) ; x_j \leq x$$

for Con. R.V

$$F(x) = \int_{-\infty}^x f(x) dx.$$

Ex:  $f(x) = 2x$ ;  $0 < x < 1$ .

$$F(x) = \int_0^x 2x dx = x^2 \quad 0 < x \leq 1 \\ = 1 \quad x > 1$$

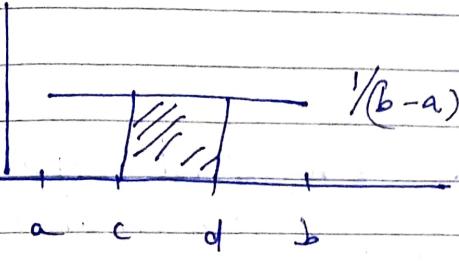
Note If follows that

$$\boxed{f(x) = \frac{d}{dx} F(x)}$$

## Uniform Distribution

$$f(x) = \frac{1}{b-a} \quad a \leq x \leq b$$

0 otherwise



$$\begin{aligned} E(X) &= \int_{-\infty}^{\infty} x f(x) dx \\ &= \frac{1}{b-a} \int_{c}^{d} x dx \\ &= \frac{1}{b-a} \frac{(b-a)(b+a)}{2} \\ &= (a+b)/2 \end{aligned}$$

$$\begin{aligned} P(a \leq x \leq d) &= \int_c^d \frac{1}{b-a} dx \\ &= \frac{1}{b-a} [x]_c^d \\ &= \frac{d-c}{b-a}. \end{aligned}$$

If  $d=b$  &  $a=c$  then  $P(a \leq x \leq b) = 1$

CDF of U.D.

$$\begin{aligned} F(x) &= \int_{-\infty}^x f(x) dx \\ &= \int_a^x \frac{1}{b-a} dx = \frac{1}{b-a} (x-a) \\ &= \frac{x-a}{b-a}; \quad a \leq x \leq b. \end{aligned}$$

## Normal or Gaussian Distribution

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2} \left(\frac{x-\mu}{\sigma}\right)^2}$$

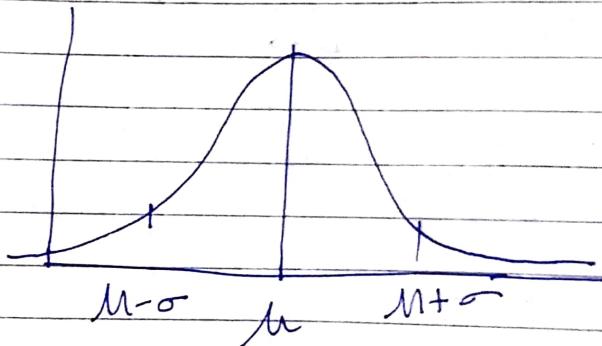
$\therefore -\infty < x < \infty$

$$= N(\mu, \sigma^2)$$

$\mu$  = mean of  $X$

$\sigma$  = SD of  $X$

It serves as an excellent approximation to a large class of distributions which have great practical importance.



$$\text{At } x = \mu$$

Case  $\sigma = 1$

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(x-\mu)^2}$$

At  $\mu = 0$

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} = N(0, 1)$$

## Points of inflection

When the derivative changes sign

$$\cancel{x+\sigma} \cdot x = \mu + \sigma$$

i. Large  $\sigma \Rightarrow$  flat graph

## Expected value

$$E(x) = \frac{1}{\sqrt{2\pi}\sigma} \int_{-\infty}^{\infty} x e^{-\frac{1}{2} \left(\frac{x-\mu}{\sigma}\right)^2} dx -$$

$$\text{Let } z = \frac{x-\mu}{\sigma} \Rightarrow dx = \sigma dz.$$

$$\therefore E(x) = \frac{1}{\sqrt{2\pi}\sigma} \int_{-\infty}^{\infty} (\sigma z + \mu) e^{-\frac{1}{2} z^2} \sigma dz$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} (\sigma z + \mu) e^{-\frac{z^2}{2}} dz$$

$$= \frac{1}{\sqrt{2\pi}} \sigma \int_0^{\infty} z e^{-\frac{z^2}{2}} dz + \underbrace{\frac{\mu}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{z^2}{2}} dz}_{= 1}$$

$$\therefore \boxed{E(x) = \mu}$$

CDF of Normal Dist (standardized)

$$\phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-t^2/2} dt.$$

Also  $P(a \leq x \leq b) = \phi(b) - \phi(a)$

$\phi(x)$  is tabulated and hence can be used to compute  $P(a \leq x \leq b)$ .

If  $x$  has distribution

$$N(\mu, \sigma^2)$$

$$Y = ax + b \text{ then } Y \text{ is dist}$$

$$N(a\mu + b, a^2\sigma^2)$$

$$\bar{y} \quad Y = \frac{x - \mu}{\sigma} = \left( \frac{1}{\sigma} \right) x - \left( \frac{\mu}{\sigma} \right) \text{ then}$$

$$Y = N\left(\frac{1}{\sigma} \cdot \mu - \frac{\mu}{\sigma}, \left(\frac{1}{\sigma}\right)^2 \sigma^2\right)$$

$$= N\left(\frac{\mu}{\sigma} - \frac{\mu}{\sigma}, 1\right)$$

$$= N(0, 1)$$

$$\begin{aligned} \therefore P(a \leq x \leq b) &= P\left(\frac{a - \mu}{\sigma} \leq \bar{y} \leq \frac{b - \mu}{\sigma}\right) \\ &= \phi\left(\frac{b - \mu}{\sigma}\right) - \phi\left(\frac{a - \mu}{\sigma}\right) \end{aligned}$$

## Exponential Distribution

$$f(x) = \begin{cases} \alpha e^{-\alpha x}, & x > 0, \alpha > 0 \\ 0, & \text{otherwise.} \end{cases}$$

CDF

$$F(x) = P(X \leq x) = \int_0^x \alpha e^{-\alpha t} dt$$

$$= \begin{cases} 1 - e^{-\alpha x}, & x \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

$$\begin{aligned} \therefore P(X > x) &= 1 - F(x) \\ &= 1 - (1 - e^{-\alpha x}) \\ &= e^{-\alpha x}. \end{aligned}$$

Expected Value

$$E(x) = \int_0^\infty x \cdot \alpha e^{-\alpha x} dx$$

$$\text{Let } \alpha e^{-\alpha x} dx = dv \quad \text{and} \quad t = u.$$

Solving  $\boxed{E(x) = \frac{1}{\alpha}}$

$\therefore$  The reciprocal of ' $\alpha$ ' is the mean of  $x$ .

Ex. Find  $P(X > \frac{1}{\alpha})$ .

$$\Rightarrow e^{-\alpha X} \rightarrow \frac{1}{\alpha}$$

We know  $P(X > z) = e^{-\alpha z}$

$$\therefore P(X > \frac{1}{\alpha}) = e^{-\alpha \cdot \frac{1}{\alpha}} = \frac{1}{e} \\ = \frac{1}{e} < \frac{1}{2}.$$

## The Gamma Distribution

$$f(x) = \begin{cases} \frac{x^{\alpha}}{\Gamma(\alpha)} e^{-x}, & x > 0 \\ 0, & \text{otherwise} \end{cases}$$

There are 2 parameters

- (i)  $\sigma^2$
- (ii)  $\alpha$

$P(x)$  is a Gamma function in  $x$ .  
defined as.

$$\Gamma(p) = \int_0^\infty x^{p-1} e^{-x} dt, \quad p > 0$$

Properties by parts

$$\Gamma(p) = (p-1) \Gamma(p-1)$$

$$\therefore \Gamma(n) = (n-1) \Gamma(n-1) = (n-1)(n-2) \dots \Gamma(1)$$

$$\therefore \Gamma(n) = (n-1)!$$

$$\boxed{\Gamma(n) = \int_0^\infty e^{-t} t^{n-1} dt}$$

Some Properties :

At  $x=1$

$$1 - e^{-\alpha x}$$

$$f(x) = \frac{\alpha}{\Gamma(1)} (\alpha x)^{1-1} e^{-\alpha x}$$

$$= \alpha e^{-\alpha x}$$

= Exponential distribution.

## Probability of S.A

Let  $\Omega$  be a sample space &  $A \subseteq \Omega$  be an event. If  $p(A) = P(\{x \in \Omega : x \in A\})$  is called the probability of  $A$ . The numbers  $p(A)$  are called by

(i)  $p(A) \geq 0 \forall A$

(ii)  $\sum_{A \in \Omega} p(A) = 1$

Let

$P$  is called the "Probability Function".

The collection of ordered pairs  $(x, p(x))$  is called the "Probability Distribution of  $X$ ".



Probability distribution can be plotted as graphs

