

The forth homework.

1. (2 points). Let the distribution of (X, Y) is given by the following table:

X/Y	Y=0	Y=1	Y=2
X=-1	0.1	p	0.15
X=2	0.2	0.3	0.05

Find p, $\text{corr}(X, Y)$, $E(X|Y)$, $E(Y^2|X)$.

2. (4 points).

Let X, Y are i.i.d.r.v's with the following density function

$$f(x, \sigma) = \sigma^2 x \exp\left(-\frac{x^2}{2\sigma^2}\right), \quad x > 0, \sigma > 0.$$

Find the distribution of the following random variables and check their independence:
 $U = \frac{X}{2Y}$, $V = X^2 + Y^2$;

2. (6 points).

a)(2 points) Find the distribution of $\sum_{i=1}^k X_i$, if X_i are i.i.d.r.v's $\sim \text{Exp}(a)$;

b)(2 points) Find the distribution of $\frac{Y}{X}$, if X, Y are independent and $X \sim N(0, 4)$, $Y \sim N(0, 5)$;

c)(2 points) Find the distribution of $\frac{2X}{3Y}$, if X, Y are independent and $X \sim N(0, \sigma^2)$, $Y \sim \text{Exp}(a)$;

d)* (1 point) Find the distribution of $\sqrt{X^2 + Y^2}$, if X, Y are independent and $X \sim \text{Norm}(0, \sigma^2)$, $Y \sim \text{Norm}(0, 10)$.

3. (8 points) Let (X, Y) be given by the following joint density function:

$$f(x, y) = \begin{cases} c(x + y), & x \in [0, 1], y \in [0, 2]. \\ 0, & \text{else.} \end{cases}$$

a)(4 points). Find c, $F_{X,Y}(x, y)$, $\text{corr}(X, Y)$, $E(X|Y)$, $E(X^2|Y)$, $E(Y|X)$ and $E(Y^2|X)$.

b) (4 points). Find the joint distribution $U = \frac{X}{Y}$, $V = X - Y$ and check their independence.

4. (8 points)

Let (X, Y) be given by the following joint density function:

$$f(x, y) = \begin{cases} cxy, & x, y \in [0, 1], x + y > 1 \\ 0, & \text{else.} \end{cases}$$

a) (2 points) Find the joint distribution of $U = X + Y$, $V = \frac{Y}{X}$ and check their independence.

b) (4 points). Find the joint distribution of $U = 2X + Y$, $V = \frac{X^2}{2Y}$ and check their independence.

c) (2 points). Find $E(X|Y)$ and $E(Y|X)$.

5. (2 points) Let random vector (X, Y, Z) is given by some joint density $f_{x,y,z}(\cdot)$.

a) Find the expression $E(g(X)|\frac{Y}{Z} = t)$.

b)* Prove

$$\frac{D(E(Y|X))}{D(Y)} = \text{corr}^2(E(Y|X), Y).$$

1. (2 points). Let the distribution of (X, Y) is given by the following table:

X/Y	Y=0	Y=1	Y=2
X=-1	0.1	p	0.15
X=2	0.2	0.3	0.05

0.45.
0.55

Find p, $\text{corr}(X, Y)$, $E(X|Y)$, $E(Y^2|X)$.

$$p = 1 - 0.1 - 0.2 - 0.3 - 0.15 - 0.05 = 0.2$$

$$E(X) = -1 \times 0.45 + 2 \times 0.55 = 0.65 \quad E(Y) = 0 \times 0.3 + 1 \times 0.5 + 2 \times 0.2 = 0.9 \quad E(XY) = 0 \cdot 0.3 + (-1) \cdot 0.2 + 2 \cdot 0.3 + (-2) \cdot 0.15 + 4 \cdot 0.05 = 0.3.$$

$$E(X^2) = 1 \times 0.45 + 4 \times 0.55 = 2.65$$

$$E(Y^2) = 0 \times 0.3 + 1 \times 0.5 + 4 \times 0.2 = 1.3 \quad D(Y) = E(Y^2) - (E(Y))^2 = 0.49.$$

$$\text{corr}(X, Y) = \frac{E(XY) - E(X)E(Y)}{\sqrt{D(X)D(Y)}} = \frac{0.3 - 0.65 \cdot 0.9}{\sqrt{2.65 \cdot 0.49}} \approx -0.27$$

$$E(X|Y=0) = (-1) \cdot \frac{0.1}{0.3} + 2 \cdot \frac{0.2}{0.3} = 1$$

$$E(Y^2|X=-1) = 1^2 \cdot \frac{0.2}{0.45} + 2^2 \cdot \frac{0.15}{0.45} = \frac{16}{9}$$

$$E(X|Y=1) = (-1) \cdot \frac{0.2}{0.5} + 2 \cdot \frac{0.3}{0.5} = \frac{4}{5}$$

$$E(Y^2|X=2) = 1^2 \cdot \frac{0.3}{0.55} + 2^2 \cdot \frac{0.05}{0.55} = \frac{10}{11}$$

$$E(X|Y=2) = (-1) \cdot \frac{0.15}{0.2} + 2 \cdot \frac{0.05}{0.2} = -\frac{1}{4}$$

2. (4 points).

Let X, Y are i.i.d.r.v's with the following density function

$$f(x, \sigma) = \frac{1}{\sigma^2} x \exp\left(-\frac{x^2}{2\sigma^2}\right), \quad x > 0, \sigma > 0.$$

Find the distribution of the following random variables and check their independence:

$$U = \frac{X}{2Y}, \quad V = X^2 + Y^2;$$

$$\text{So: since } X, Y \text{ independent } f_{X,Y}(x, y, \sigma) = \sigma^{-4} xy e^{-\frac{x^2+y^2}{2\sigma^2}}$$

$$\left\{ \begin{array}{l} U = \frac{X}{2Y} \\ V = X^2 + Y^2 \end{array} \right. \Rightarrow \left\{ \begin{array}{l} X = 2U \sqrt{\frac{V}{4U^2+1}} \\ Y = \sqrt{\frac{V}{4U^2+1}} \end{array} \right. \quad d\left(-\frac{X}{2\sigma^2}\right) = \frac{4U^2}{4U^2+1}.$$

$$|J|^{-1} = \begin{vmatrix} \frac{1}{2Y} & -\frac{X}{2Y^2} \\ 2X & 2Y \end{vmatrix} = 1 + \frac{X^2}{Y^2} = \frac{X^2 + Y^2}{Y^2} = \frac{V}{U^2} = 4U^2 + 1.$$

$$f_V(u) = \int_{-\infty}^{\infty} f(2u \sqrt{\frac{v}{4u^2+1}}) f(\sqrt{\frac{v}{4u^2+1}}) \cdot \frac{1}{4u^2+1} dv = \int_0^{\infty} \sigma^{-4} \cdot \frac{2uv}{(4u^2+1)^2} \cdot e^{-\frac{v}{2\sigma^2}} dv.$$

$$= \frac{2u\sigma^{-4}}{(4u^2+1)^2} \left[\int_0^{\infty} ve^{-\frac{v}{2\sigma^2}} dv \right] = \frac{2u\sigma^{-4}}{(4u^2+1)^2} \cdot (2\sigma^2)^2 = \frac{8u}{(4u^2+1)^2}.$$

$$f_U(v) = \int_0^{\infty} \sigma^{-4} \cdot \frac{2uv}{(4u^2+1)^2} e^{-\frac{v}{2\sigma^2}} du = \sigma^{-4} v \cdot e^{-\frac{v}{2\sigma^2}} \int_0^{\infty} \frac{d(4u^2+1)}{4(4u^2+1)^2} = \frac{\sigma^{-4}}{4} v \cdot e^{-\frac{v}{2\sigma^2}}$$

$$f_{U,V}(u, v) = f(2u \sqrt{\frac{v}{4u^2+1}}) f(\sqrt{\frac{v}{4u^2+1}}) \cdot \frac{1}{4u^2+1} = \sigma^{-4} \cdot \frac{2uv}{(4u^2+1)^2} e^{-\frac{v}{2\sigma^2}}$$

thus we have $f_{U,V}(u, v) = f_U(u) f_V(v)$. U, V are independent.

2. (6 points).

a)(2 points) Find the distribution of $\sum_{i=1}^k X_i$, if X_i are i.i.d.r.v's $\sim Exp(a)$;

b)(2 points) Find the distribution of $\frac{Y}{X}$, if X, Y are independent and $X \sim N(0, 4)$, $Y \sim N(0, 5)$;

c)(2 points) Find the distribution of $\frac{2X}{3Y}$, if X, Y are independent and $X \sim N(0, \sigma^2)$, $Y \sim Exp(a)$;

d)* (1 point) Find the distribution of $\sqrt{X^2 + Y^2}$, if X, Y are independent and $X \sim Norm(0, \sigma^2)$, $Y \sim Norm(0, 10)$.

a). denote $X = \sum_{i=1}^n X_i$. $f_{X_i}(x_i) = a e^{-ax_i}$ $x_i \geq 0$.

$$\text{if } n=2. f_X(x) = \int_0^x a^2 e^{-at} \cdot e^{-a(x-t)} dt = a^2 e^{-ax} \cdot \int_0^x dt = a^2 x e^{-ax}$$

$$\text{if } n=3. f_X(x) = \int_0^x a^3 t^2 e^{-at} \cdot e^{-a(x-t)} dt = a^3 e^{-ax} \cdot \int_0^x t^2 dt = a^3 \frac{x^3}{3} e^{-ax}$$

repeat the process for $n=k$. $f_X(x) = a^n \cdot \frac{x^{n-1}}{(n-1)!} \cdot e^{-ax}$, $x \geq 0$.

b) let $x=u$. $y=uv$. $\Rightarrow \begin{cases} u=x \\ v=\frac{y}{x} \end{cases}$ $|J| = \begin{vmatrix} 1 & 0 \\ 0 & u \end{vmatrix} = u$.

$$f_{U,V}(u,v) = f_X(u) \cdot f_Y(uv) \cdot |J| = \frac{1}{4\pi\sqrt{5}} e^{-\left(\frac{u^2}{8} + \frac{v^2}{10}\right)} |u|$$

$$f_V(v) = \int_{-\infty}^{\infty} f_{U,V}(u,v) du = \frac{1}{2\pi\sqrt{5}} \int_0^{\infty} u \cdot e^{-u^2 \left(\frac{1}{8} + \frac{v^2}{10}\right)} du = \frac{1}{2\pi\sqrt{5}} \frac{1}{2\left(\frac{1}{8} + \frac{v^2}{10}\right)} = \frac{2\sqrt{5}}{\pi(5+4v^2)}$$

c). let $\begin{cases} u=\frac{y}{x} \\ v=\frac{2x}{3y} \end{cases} \Rightarrow \begin{cases} x=\frac{3vu}{2} \\ y=u \end{cases}$ $|J| = \begin{vmatrix} \frac{3v}{2} & \frac{3u}{2} \\ 1 & 0 \end{vmatrix} = \frac{3}{2}|u|$

$$f_{U,V}(u,v) = f_X\left(\frac{3vu}{2}\right) \cdot f_Y(u) \cdot |J| = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{9u^2v^2}{8\sigma^2}} \cdot a \cdot e^{-au} \cdot \frac{3}{2}u.$$

$$\begin{aligned} f_V(v) &= \frac{3a}{2\sqrt{2\pi}\sigma} \int_0^{+\infty} e^{-\left(\frac{9v^2}{8\sigma^2} \cdot u^2 + au\right)} \cdot u du \\ &= \frac{3a}{2\sqrt{2\pi}\sigma} \left[\int_0^{+\infty} e^{-\left(\frac{9v^2}{8\sigma^2} \cdot u^2 + au\right)} \cdot \left(\frac{9v^2}{4\sigma^2}u + a\right) \cdot \frac{4\sigma^2}{9v^2} du - \frac{4\sigma^2 a}{9v^2} \int_0^{+\infty} e^{-\left(\frac{9v^2}{8\sigma^2} \cdot u^2 + au\right)} du \right] \end{aligned}$$

$$= \frac{\sqrt{2}\sigma}{3\sqrt{\pi}v^2} - \frac{\sqrt{2}a^2\sigma}{3\sqrt{\pi}v^2} \int_0^{+\infty} e^{-\left(\frac{9v^2}{8\sigma^2} \cdot u^2 + au\right)} du \rightarrow \text{seems can't find elementary function as it's sol.}$$

d). let $\begin{cases} x=r\cos\varphi \\ y=r\sin\varphi \end{cases}$ $J = \begin{vmatrix} \cos\varphi & -r\sin\varphi \\ \sin\varphi & r\cos\varphi \end{vmatrix} = |r|$

$$f_{R,\Phi}(r, \varphi) = f_x(r\cos\varphi) \cdot f_y(r\sin\varphi) \cdot |r| = \frac{1}{2\pi\sigma\sqrt{10}} \cdot e^{-\frac{r^2\cos^2\varphi}{2\sigma^2}} \cdot e^{-\frac{r^2\sin^2\varphi}{2\sigma^2}} |r|.$$

$$f_R(|r|) = \frac{|r|}{2\pi\sigma\sqrt{10}} \int_0^{2\pi} e^{-r^2 \left(\frac{\cos^2\varphi}{2\sigma^2} + \frac{\sin^2\varphi}{2\sigma^2}\right)} d\varphi$$

$$\because \sigma^2=10 \quad f_R(r) = \frac{r}{10} \cdot e^{-\frac{r^2}{20}}, \quad r > 0.$$

$$\because \sigma^2 \neq 10. \quad f_R(r) = \frac{r}{2\pi\sqrt{10}\sigma} \int_0^{2\pi} e^{-\frac{r^2}{2} \left(\frac{\cos^2\varphi}{\sigma^2} + \frac{\sin^2\varphi}{10}\right)} d\varphi, \quad r > 0. \quad \text{the integral no elementary function expression.}$$

3. (8 points) Let (X, Y) be given by the following joint density function:

$$f(x, y) = \begin{cases} c(x+y), & x \in [0, 1], y \in [0, 2]. \\ 0, & \text{else.} \end{cases}$$

a) (4 points). Find c , $F_{X,Y}(x, y)$, $\text{corr}(X, Y)$, $E(X|Y)$, $E(X^2|Y)$, $E(Y|X)$ and $E(Y^2|X)$.

b) (4 points). Find the joint distribution $U = \frac{X}{Y}$, $V = X - Y$ and check their independence.

$$a) \int_0^1 \int_0^2 c(x+y) dy dx = \int_0^1 (2cx + 2c) dx = cx^2 + 2cx \Big|_0^1 = 3c \quad 3c = 1 \Rightarrow c = \frac{1}{3}.$$

$$F_{X,Y}(x, y) = \frac{1}{3} \int_0^x \int_0^y (u+v) du dv = \frac{1}{3} \int_0^x \left(\frac{y^2}{2} + yv \right) dv = \frac{xy}{6} (x+y), \quad x \in [0, 1], y \in [0, 2]$$

$$f_X(x) = \int_0^2 \frac{1}{3} (x+y) dy = \frac{1}{3} \cdot (xy + \frac{y^2}{2}) \Big|_0^2 = \frac{2(x+1)}{3}, \quad x \in [0, 1]$$

$$f_{Y|Y}(y) = \int_0^1 \frac{1}{3} (x+y) dx = \frac{1}{3} \cdot (xy + \frac{y^2}{2}) \Big|_0^1 = \frac{2y+1}{6}, \quad y \in [0, 2]$$

$$E(XY) = \int_0^2 \int_0^1 \frac{x+y}{3} \cdot xy dx dy = \frac{1}{3} \int_0^2 y \left(\int_0^1 (x^2 + xy) dx \right) dy = \frac{1}{3} \int_0^2 y \left(\frac{1}{3} + \frac{y}{2} \right) dy = \frac{1}{3} \cdot \left(\frac{1}{6}y^2 + \frac{1}{6}y^3 \Big|_0^2 \right) = \frac{2}{3}.$$

$$E(X) = \int_0^1 \frac{2x^2 + 2x}{3} dx = \frac{1}{3} \cdot \left(\frac{2}{3}x^3 + x^2 \right) \Big|_0^1 = \frac{5}{9}, \quad E(X^2) = \int_0^1 \frac{2x^3 + 2x^2}{3} dx = \frac{1}{3} \cdot \left(\frac{2}{4}x^4 + \frac{2}{3}x^3 \right) \Big|_0^1 = \frac{7}{18}$$

$$E(Y) = \int_0^2 \frac{2y^2 + y}{6} dy = \frac{1}{6} \cdot \left(\frac{2}{3}y^3 + \frac{y^2}{2} \right) \Big|_0^2 = \frac{11}{9}, \quad E(Y^2) = \int_0^2 \frac{2y^3 + y^2}{6} dy = \frac{1}{6} \cdot \left(\frac{1}{2}y^4 + \frac{1}{3}y^3 \right) \Big|_0^2 = \frac{16}{9}$$

$$D(X) = E(X^2) - (E(X))^2 = \frac{7}{18} - \frac{25}{81} = \frac{63-50}{162} = \frac{13}{162}$$

$$D(Y) = E(Y^2) - (E(Y))^2 = \frac{16}{9} - \frac{121}{81} = \frac{144-121}{81} = \frac{23}{81}$$

$$\text{corr}(X, Y) = \frac{\text{cov}(X, Y)}{\sqrt{D(X) \cdot D(Y)}} = \frac{\frac{2}{3} - \frac{5}{9} \cdot \frac{11}{9}}{\sqrt{\frac{13}{162} \cdot \frac{23}{81}}} = \frac{-\frac{1}{81}}{\sqrt{\frac{13 \cdot 23}{162}}} = -0.0082$$

$$E(X|Y) = g(y) = \int_{-\infty}^{+\infty} x \frac{f_{X,Y}(x,y)}{f_{Y|Y}(y)} dx = \int_0^1 x \frac{\frac{1}{3}(x+y)}{\frac{2y+1}{6}} dx = \frac{2}{2y+1} \int_0^1 x^2 + xy dx = \frac{2}{2y+1} \left(\frac{1}{3} + \frac{y}{2} \right) = \frac{2+3y}{3(2y+1)}$$

$$E(X^2|Y) = \int_{-\infty}^{+\infty} x^2 \cdot \frac{f_{X,Y}(x,y)}{f_{Y|Y}(y)} dx = \frac{2}{2y+1} \cdot \int_0^1 (x^3 + x^2 y) dx = \frac{2}{2y+1} \left(\frac{1}{4} + \frac{y}{3} \right) = \frac{3+4y}{6(2y+1)}$$

$$E(Y|X) = \int_0^2 y \cdot \frac{f_{X,Y}(x,y)}{f_X(x)} dy = \frac{1}{2(x+1)} \int_0^2 y(x+y) dy = \frac{1}{2(x+1)} \left[2x + \frac{8}{3} \right] = \frac{3x+4}{3(x+1)}$$

$$E(Y^2|X) = \int_0^2 y^2 \cdot \frac{f_{X,Y}(x,y)}{f_X(x)} dy = \frac{1}{2(x+1)} \int_0^2 y^3(x+y) dy = \frac{1}{2(x+1)} \left[\frac{8}{3}x + 4 \right] = \frac{4x+6}{3(x+1)}$$

$$b). \begin{cases} U = \frac{X}{Y} \\ V = X - Y. \end{cases} \Rightarrow \begin{cases} X = \frac{UV}{U-1} \\ Y = \frac{V}{U-1} \end{cases} \quad |J| = \begin{vmatrix} -\frac{V}{(U-1)^2} & \frac{U}{U-1} \\ -\frac{U}{(U-1)^2} & \frac{1}{U-1} \end{vmatrix} = \left| \frac{-V+UV}{(U-1)^3} \right| = \frac{|V|}{(U-1)^2}, \quad U \in \mathbb{R}^+, \quad V \in [-1, 1].$$

$$f_{U,V}(u,v) = f_{X,Y}\left(\frac{UV}{U-1}, \frac{V}{U-1}\right) \cdot |J| = \frac{V}{3} |V| \cdot \frac{U+1}{(U-1)^3}$$

$$\begin{cases} 0 \leq \frac{UV}{U-1} \leq 1 \\ 0 \leq \frac{V}{U-1} \leq 2. \end{cases} \quad \begin{cases} V \geq 1 \\ 0 \leq UV \leq U-1 \\ 0 \leq V \leq 2(U-1). \end{cases} \quad \Rightarrow 0 \leq V \leq \frac{U-1}{U} \quad \text{if } U \leq 1.$$

$$\begin{cases} 0 \geq UV \geq U-1 \\ 0 \geq V \geq 2(U-1) \end{cases} \quad \Rightarrow 0 \geq V \geq \max\left\{\frac{U-1}{U}, 2(U-1)\right\}$$

$\Rightarrow U, V$ is not independent.

$$f_{U,V}(u,v) = \begin{cases} \frac{V^2}{3} \frac{U+1}{(U-1)^3} & U \geq 1, V \in [0, \frac{U-1}{U}], \\ -\frac{V^2}{3} \frac{U+1}{(U-1)^3} & U \in (0, \frac{1}{2}], V \in [2(U-1), 0] \text{ or } U \in (\frac{1}{2}, 1), V \in [\frac{U-1}{U}, 0] \end{cases}$$

Thus we have the domain of V is dependent with U .

4. (8 points)

Let (X, Y) be given by the following joint density function:

$$f(x, y) = \begin{cases} cxy, & x, y \in [0, 1], x + y > 1 \\ 0, & \text{else.} \end{cases}$$

- a) (2 points) Find the joint distribution of $U = X + Y, V = \frac{Y}{X}$ and check their independence.
 b) (4 points). Find the joint distribution of $U = 2X + Y, V = \frac{X^2}{2Y}$ and check their independence.
 c) (2 points). Find $E(X|Y)$ and $E(Y|X)$.

$$\text{find } c: \int_0^1 \int_{1-x}^1 c(xy) dy dx = c \int_0^1 \left[x \frac{y^2}{2} \Big|_{1-x}^1 \right] dx = c \int_0^1 x^2 - \frac{x^3}{2} dx = c \cdot \left[\frac{1}{3} - \frac{1}{8} \right] = 1 \Rightarrow c = \frac{24}{5}$$

$$a). \begin{cases} U = X + Y \\ V = \frac{Y}{X} \end{cases} \Rightarrow \begin{cases} X = \frac{U}{V+1} \\ Y = \frac{UV}{V+1} \end{cases} \quad |J| = \begin{vmatrix} \frac{1}{V+1} & -\frac{U}{(V+1)^2} \\ \frac{V}{V+1} & \frac{U}{(V+1)^2} \end{vmatrix} = \frac{U}{(V+1)^2} \quad \begin{cases} V > 0 \\ V > 1. \end{cases}$$

$$f_{U,V}(u,v) = f_{X,Y}\left(\frac{u}{v+1}, \frac{uv}{v+1}\right) |J| = \frac{24u^3v}{5(v+1)^4}$$

$$\begin{cases} 0 \leq \frac{U}{V+1} \leq 1 \\ 0 \leq \frac{UV}{V+1} \leq 1 \\ V > 1 \end{cases} \Rightarrow \begin{cases} 2 \geq U > 1 \\ \frac{1}{U-1} \geq V \geq U-1 \end{cases} \quad v \text{ has domain dependent with } u. \text{ thus. } u, v. \text{ not independent}$$

$$f_{U,V}(u,v) = \frac{24u^3v}{5(v+1)^4} \cdot \begin{cases} u \in (1, 2] \\ v \in [U-1, \frac{1}{U-1}] \end{cases}$$

$$b). \begin{cases} U = 2X + Y \\ V = \frac{X^2}{2Y} \end{cases} \Rightarrow \begin{cases} X = \sqrt{4V^2 + 2VU} - 2V. \\ Y = U + 4V - 2\sqrt{4V^2 + 2VU}. \end{cases} \quad |J|^{-1} = \begin{vmatrix} 2 & 1 \\ \frac{X}{Y} & -\frac{X^2}{2Y^2} \end{vmatrix} = \frac{X^2 + XY}{Y^2}$$

$$|J| = \frac{Y^2}{X^2 + XY} = \frac{Y^2}{2VY + XY} = \frac{Y}{2V + X} = \frac{U + 4V}{\sqrt{4V^2 + 2VU}} - 2.$$

$$f_{U,V}(u,v) = f_{X,Y}(x, y) |J| = \frac{24}{5} \cdot xy \cdot \frac{y^2}{x(x+y)} = \frac{24}{5} \cdot \frac{y^3}{x+y} = \frac{24}{5} \cdot \frac{(u+4v - 2\sqrt{2v(2v+u)})^3}{(2v+u) - \sqrt{2v(2v+u)}}$$

$$X+Y > 1 \Rightarrow u + 2v - \sqrt{2v(2v+u)} > 1 \Rightarrow (u-1)^2 > (2u-4)v.$$

$$\begin{aligned} \text{if } u > 2, 0 < v < \frac{(u-1)^2}{2u-4} & \quad v \text{ has domain depend on } u. \\ u < 2, 0 > v > \frac{(u-1)^2}{2u-4} & \quad \text{thus } U, V \text{ not independent.} \end{aligned}$$

$$c). f_X(x) = \frac{24}{5} \int_{1-x}^1 xy dy = \frac{24}{5} \cdot x \cdot \frac{y^2}{2} \Big|_{1-x}^1 = \frac{24}{5} \cdot \left[\frac{x}{2} - \frac{x}{2}(1-x)^2 \right] = \frac{24}{5} (x^2 - \frac{1}{2}x^3) = \frac{24}{5} x^2 - \frac{12}{5} x^3.$$

$$f_Y(y) = f_X(y) = \frac{24}{5} x^2 - \frac{12}{5} x^3 \quad \text{by symmetry of } x, y.$$

$$E(X|Y) = \int_{1-y}^1 x \cdot \frac{f_{X,Y}(x,y)}{f_Y(y)} dx = 2 \int_{1-y}^1 x \cdot \frac{xy}{2y^2 - y^3} = \frac{2}{y(2-y)} \int_{1-y}^1 x^2 dx = \frac{2}{y(2-y)} \cdot \frac{y(3-3y+y^2)}{3} = \frac{2(y^2-3y+3)}{3(2-y)}$$

$$E(Y|X) = \int_{1-x}^1 y \cdot \frac{f_{X,Y}(x,y)}{f_X(x)} dy = \frac{2(x^2 - 3x + 3)}{3(2-x)}.$$

5. (2 points) Let random vector (X, Y, Z) is given by some joint density $f_{x,y,z}(\cdot)$.

a) Find the expression $E(g(X)|\frac{Y}{Z} = t)$.

b)* Prove

$$\frac{D(E(Y|X))}{D(Y)} = \text{corr}^2(E(Y|X), Y).$$

$$a). \text{Let } \begin{cases} W = \frac{Y}{Z} \\ Z = z \end{cases} \Rightarrow \begin{cases} Y = Wz \\ Z = z \end{cases} \Rightarrow |J| = \begin{vmatrix} z & w \\ 0 & 1 \end{vmatrix} = |z|.$$

$$f_{X,W,Z} = f_{X,Y,Z}(x, wz, z) \cdot |z|.$$

$$f_{X|W=t}(x) = \frac{f_{X,W}}{f_W} = \frac{\int f_{X,Y,Z}(x, wz, z) \cdot |z| dz}{\iint f_{X,Y,Z}(x, wz, z) \cdot |z| dz dx} \quad E(g(X) | \frac{Y}{Z} = t) = \int g(x) \cdot f_{X|W=t}(x) dx.$$

$$b) \quad \frac{D(E(Y|X))}{D(Y)} = \text{corr}^2(E(Y|X), Y).$$

$$RHS = \frac{\text{Cov}^2(E(Y|X), Y)}{D(E(Y|X)) \cdot D(Y)}.$$

it suffices to check $\text{Cov}^2(E(Y|X), Y) = D^2(E(Y|X))$. since variance > 0 . eliminate the square.

$$\text{Cov}(E(Y|X), Y) = E(E(Y|X) \cdot Y) - E(E(Y|X)) \cdot E(Y).$$

$$D(E(Y|X)) = E(E(Y|X)^2) - [E(E(Y|X))]^2$$

by the property of conditional expectation. $E(E(Y|X)) = E(Y)$. thus. $E(E(Y|X)) \cdot E(Y) = [E(E(Y|X))]^2$

$$E(E(Y|X) \cdot Y) = E \left[E(E(Y|X) \cdot Y | X) \right] \xrightarrow[\substack{E(Y|X) = \text{const.} \\ \text{when } X \text{ is fixed}}]{} E \left[E(Y|X) \cdot E(Y|X) \right] = E(E(Y|X)).$$

thus we have $\text{Cov}(E(Y|X), Y) = D(E(Y|X)) \Rightarrow LHS = RHS$.