

Structural Econometrics and Inverse Reinforcement Learning: Inferring preferences and beliefs from human behavior*

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Abstract

This article, for *Oxford Research Encyclopedia of Economics and Finance*, summarizes the literature on structural estimation of dynamic discrete choice models and points out close parallels to the rapidly growing literature on inverse reinforcement learning. Both literatures are focused on inferring underlying preferences (and sometimes beliefs) of human decision makers from observations of their states and actions.

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1 Introduction

This article compares and contrasts methods developed in two different literatures for inferring preferences (and sometimes also beliefs) of human decision makers from observations on their behavior over time: one developed by economists that can be called “structural econometrics” (SE) and another developed by computer scientists in machine learning (ML) and artificial intelligence (AI) that is broadly under the general rubric called “reinforcement learning” (RL). Both SE and RL are huge literatures and we will not attempt to survey them in any breadth. Instead, we will focus on a particular subarea of SE known as *dynamic discrete choice* (DDC) which is most clearly parallel to a subareas of RL called “inverse reinforcement learning” (IRL).¹ Both literatures rely on the presumption that the agents whose preferences we are interested in recovering are *rational* i.e. they take a sequence of actions over time to maximize an internally perceived reward (or “utility”) given their subjective beliefs of how their current actions cause or at least influence potentially uncertain future outcomes that affect their rewards. The goals of DDC and IRL are identical: to invert or infer their underlying objective or reward function (and sometimes also their beliefs) from observations of the sequence of states and actions of a sample of decision makers (DMs) that could be individuals, firms, or other intelligent agents (e.g. animals). Beyond the intellectual interest in whether this type of inversion and inference is possible what are the practical motivations for DDC and IRL modeling?

In economics, structural models are used to test and evaluate economic theories, and to make *counterfactual predictions* of the effect of changes in economic policies or other changes in the environment on behavior and welfare (i.e. who is hurt and who is helped by the policy change?). Structural models are distinguished from reduced-form models that provide statistical summaries of agents’ behavior, but do not attempt to infer their preferences and beliefs. Reduced-form models can make counterfactual predictions of a policy change when there has been sufficient historical variation in policy, so it can be included as an explanatory variable in an econometric model that predicts economic behavior. However if policies have not changed in the past reduced form models are unable to predict their impact on behavior whereas structural models can produce accurate “out-of-sample” predictions of policy changes on behavior and welfare in the absence of past variation in the policy.

For example, for decades Denmark has imposed one of the highest tax rates on new cars of any country in the world. [Gillingham et al. \(2022\)](#) used a structural model of equilibrium ownership and trading of cars in Denmark to predict the impact of reducing this “registration tax”. The model predicts that by reducing the registration tax and increasing the gas tax, it would be possible of Denmark to 1) improve the welfare of its citizens, 2) increase total tax revenue from cars (registration plus gas tax), and 3) reduce total CO₂ pollution. It can be substantially faster and cheaper to use structural models to predict and evaluate the effects of various changes in economic

¹To space we will rely heavily on abbreviations throughout this article, but refer readers to table 1 for a summary of the various abbreviations and acronyms used throughout this article.

policies than to actually carry them out and learning over time by trial and error. Auto producers learned this lesson decades ago when they developed finite element simulations of car crashes that were sufficiently realistic that they were adopted as a faster and cheaper way to evaluate new car designs. Structural models can be used to “crash test” policy changes in a similar way.

The practical motivation for inferring the DM’s reward function in IRL stems from the desire to produce intelligent agents that can perform well in situations with complex, poorly-defined, or hard to specify goals. They do this by inferring rewards that motivate humans to perform these same tasks, using limited human *demonstration data*. ML and RL have had astounding success in training machines to understand, predict, and generate text, artwork, video, spoken language, and have made impressive strides in robotics. This success depends not only on advances in computer power and algorithms such as deep neural networks (DNN) that have the flexibility to make good predictions from a variety of unstructured inputs, but also on the existence of large amounts of training data, *and* a well defined objective function that the algorithm is trained to optimize.

As [Russell and Norvig \(2022\)](#) note, most of the existing ML methods use *supervised learning* where the algorithm “learns by passively observing example input/output pairs provided by a ‘teacher’.” (p. 840). Here the reward function is known: to predict the choices of the teacher as well as possible according to some metric. However the labelled data provided by a teacher is often quite limited, and this is where RL has proven so successful. When a reward function is known RL algorithms can produce behavior that approximately maximizes the reward function without an explicit teacher. A good example is chess, where [Russell and Norvig \(2022\)](#) note “there are relatively few examples (about 10^8) compared to the space of all possible chess positions (about 10^{40}).” When the reward function is known (e.g. in chess, 1 for winning and 0 for losing), RL has been able to learn superhuman performance via *self-play*, i.e. repeatedly playing the algorithm against computerized opponents in “offline training” resulting in performance that exceeds the best human and computerized chess players in “online” real-time play, see e.g. [Silver et al. \(2018\)](#).

There are many challenging tasks in AI and robotics where it is unclear what reward function should be to train intelligent behavior. [Christiano et al. \(2017\)](#) provide a concrete example: “suppose that we wanted to use reinforcement learning to train a robot to clean a table or scramble an egg. It’s not clear how to construct a suitable reward function, which will need to be a function of the robot’s sensors.” However the option of “manually programming their behavior has become increasingly challenging and expensive.” [Osa et al. \(2018\)](#). The IRL literature circumvents this by inferring a reward function from limited observations of human choices for a task we would like to automate. If it is possible to estimate a reward function that “rationalizes” human choices, RL can use it to train the algorithm to optimize it much more rapidly and cheaply under a wider range of scenarios that were encountered in the demonstration data set. The result is intelligent behavior consistent with human preferences, including situations that are outside the limited set in the data it was trained on.²

²There is also a literature on “inverse optimal control” (IOC) originated by [Kalman \(1964\)](#) that is also related to structural estimation and IRL. One of the early practical successes in this literature, [Mombaur et al. \(2010\)](#), was

For example [Barnes et al. \(2024\)](#) used data on trips from Google Maps to “provide routes that best reflect travelers’ latent preferences” which can only be inferred from their “physical behavior, which implicitly trade-off factors including traffic conditions, distance, hills, safety, scenery, road conditions, etc.” They showed how to massively scale IRL to estimate a “route preference function” with 360 million parameters using a training dataset of 110 million trips sampled from Google Maps. Using this reward function, they used RL to train Google Maps to provide route recommendations more consistent with human preferences resulting “in a policy that achieves a 16-24% improvement in route quality at a global scale, and to the best of our knowledge, represents the largest published study or IRL algorithms in a real-world setting to date.”³

The remainder of this article is organized as follows. Section 2 provides background on dynamic programming (DP) and Markovian decision processes (MDP) which is the common ancestor of both the SE and RL literatures. We introduce the key concepts of value and policy function and the Bellman equation which characterizes the solution for an optimal dynamic policy, and introduce RL as a class of iterative stochastic algorithms for solving the MDP problem. Section 3 describes how the DP framework is adapted to inference for the reward function in the SE literature, focusing on the *dynamic discrete choice* (DDC), and deriving the equations for inference of preferences and beliefs by the method of maximum likelihood. We also discuss the *identification problem* that puts inherent limits on our ability to uniquely recover the true preferences and beliefs of a decision maker in the absence of strong additional assumptions on their functional form. Section 4 summarizes the literature on IRL and its relation to the DDC and section 5 discusses several unsolved problems.

2 Background on DP and MDPs

The SE and RL literatures both build on a body of work that provides a recursive characterization of optimal sequential decision making under uncertainty, commonly referred to as *dynamic programming*, (DP) the name coined by one of the prominent early contributors to this theory, [Bellman \(1957\)](#). We will briefly review a special class of DP problems known as stationary, infinite horizon Markovian Decision Problems (MDP) introduced by Bellman and [Howard \(1960\)](#) that is the predominant underpinning of SE and RL and we refer the reader to [Bertsekas \(2017\)](#) for a first rate treatment of DP and MDPs.⁴

in robotics where IOC was used to “generate natural overall locomotion trajectories from an initial rest position an orientation to a given target rest position and orientation.” [Abbeel \(2016\)](#) extend the IOC approach to use DNNs to provide more flexible approximations to the underlying reward or cost function that “can learn complex, nonlinear cost representations ... and can be applied to high-dimensional systems with unknown dynamics” resulting in a method that “outperforms prior IOC algorithms on a set of simulated benchmarks, and achieves good results on several real-world tasks.” See [Azar et al. \(2020\)](#) for further connections between IOC and IRL.

³While 110 million observations may not seem “limited”, it is a tiny fraction of all trips navigated using Google Maps. [Jeon et al. \(2021\)](#) cites additional examples where “defining a reward function beforehand is particularly challenging and IRL is simply more pragmatic.” These examples include robotic manipulation, autonomous driving, and clinical motion analysis. See [Zhao and Liang \(2023\)](#) who used Adversarial IRL (AIRL) to estimate context-dependent route preferences using data on trips by taxi drivers in Shanghai.

⁴Note that DP and MDPs involve the optimization of discounted utility which was introduced by [Samuelson \(1937\)](#) combined with the expected utility concept introduced by [von Neumann and Morgenstern \(1944\)](#) for problems

Table 1. List of abbreviations and acronyms

| Abbreviation | Full Name |
|--------------|---|
| AI | Artificial Intelligence |
| AIRL | Adversarial Inverse Reinforcement Learning |
| AGI | Artificial General Intelligence |
| AL | Apprenticeship Learning |
| BC | Behavioral Cloning |
| BIRL | Bayesian Inverse Reinforcement Learning |
| CCP | Conditional Choice Probability |
| CCS | Conditional Choice Simulator |
| DDC | Dynamic Discrete Choice |
| DM | Decision Maker (Agent) |
| DNN | Deep Neural Network |
| DP | Dynamic Programming |
| DSE | Dynamic Structural Estimation |
| DU | Discounted Utility |
| EU | Expected Utility |
| IOC | Inverse Optimal Control |
| IL | Imitation Learning |
| IRL | Inverse Reinforcement Learning |
| MCMC | Markov Chain Monte Carlo |
| MDP | Markovian Decision Process |
| ML | Machine Learning |
| MPE | Markov Perfect Equilibrium |
| MPEC | Mathematical Programming with Equilibrium Constraints |
| NFXP | Nested Fixed Point |
| PI | Policy Iteration |
| RFE | Reduced Form Estimation |
| RL | Reinforcement Learning |
| RTDP | Real Time Dynamic Programming |
| SA | Successive Approximations |
| TD | Temporal Difference |

MDPs are defined by 1) a state space S , 2) an action space A , and 3) the objects $\{r, p, \beta\}$, where $r(s, a)$ is the *reward or utility function* providing the payoff to a decision maker (DM), $p(s'|s, a)$ is a *transition density* for next period's state when the state is s and action is a and $\beta \in (0, 1)$ is a discount factor. To simplify exposition we assume that both S and A are finite sets, with $|S|$

involving uncertainty. Though Samuelson “had concerns about its descriptive realism, and it was never empirically validated as the appropriate model for intertemporal choice” (Frederick et al. (2002)), and laboratory tests including the famous *Allais Paradox* (see Machina (1987)) constitute evidence against the hypothesis of expected utility maximization, the MDP framework remains highly relevant for modern work in economics and AI and other areas because MDP model are sufficiently flexible to provide good approximations to a huge range of sequential decision problems. Though there is research that relaxes the discounted utility and expected utility assumptions, there are unsolved challenges in applying recursive methods to characterize and numerically approximate optimal strategies in for these “non-separable” problems.

and $|A|$ elements, respectively.⁵ The DM is presumed to choose a function $\pi(a|s)$ referred to as a *strategy*, *decision rule*, or *policy* that specifies the probability of taking action a in state s . For any $s \in S$ the support of π is restricted to a subset $A(s) \subset A$ of choices that are feasible for the DM to take in state s . Note that strategies can be either *mixed* (i.e. π is a probability) or *pure* (π is a function). The DM’s objective is to choose a strategy π to maximize the expected discounted stream of rewards under the policy π ,

$$V(s) = \max_{\pi} V_{\pi}(s), \quad \text{where } V_{\pi}(s) = E_{\pi} \left\{ \sum_{t=0}^{\infty} \beta^t r(s_t, a_t) | s_0 = s \right\}, \quad (1)$$

where E_{π} denotes the expectation operator over the Markov process $\{s_t, a_t\}$ implied by the transition probability p and decision rule π . [Bellman \(1957\)](#), [Howard \(1960\)](#) and others proved that an optimal policy can be defined recursively in terms of what is now known as *Bellman’s equation*

$$V(s) = \max_{a \in A(s)} \left[r(s, a) + \beta \sum_{s' \in S} V(s') p(s'|s, a) \right] \equiv \Gamma(V)(s), \quad (2)$$

so V is the fixed point $V = \Gamma(V)$ where the operator $\Gamma(V)$, known as the “Bellman operator”, is the function of s implicitly defined by the right hand side of (2). The Bellman operator can also be shown to be a *contraction mapping* (so for any vectors $V, W \in R^{|S|}$ we have $\|\Gamma(V) - \Gamma(W)\| \leq \beta \|V - W\|$ where β is less than 1 and $\|\cdot\|$ is the Euclidean norm). This implies a unique solution V to the Bellman equation. The optimal policy can be recovered from V as the choice $a^* \in A(s)$ that attains the maximum in Bellman’s equation

$$\pi(a|s) = \underset{a \in A(s)}{\operatorname{argmax}} \left[r(s, a) + \beta \sum_{s' \in S} V(s') p(s'|s, a) \right]. \quad (3)$$

Note that equation (3) implies that the optimal policy is generically a pure strategy, i.e. except for rare case of ties, for each state s there is a unique $a^* \in A(s)$ that maximizes the right hand side of (3), so an optimal policy chooses this action with probability 1.

An MDP can be considered as a “game against nature” where “nature’s move” is encoded in the transition probability $p(s'|s, a)$ and the optimal strategy π can be considered as a “best reply” to nature. This framework can be extended to dynamic games where there are multiple players in addition to nature. In this setting the reward and transition probability functions for each player can depend on the actions of their opponents. [Maskin and Triole \(2001\)](#) introduced the concept of *Markov Perfect equilibrium* (MPE) which is a generalization of the concept of Nash equilibrium in static games where each player’s strategy is a best reply to all of their opponents and nature. A MPE is defined by the solution to a *system of Bellman equations*, one for each player (except

⁵The general theory carries over essentially unchanged if S or A have an infinite number or even a continuum of elements provided certain boundedness and continuity conditions hold. However we will see that very different methods are used to solve MDPs with large state spaces.

nature). Unlike the single Bellman equation in a single agent game against nature, this system of Bellman equations can have multiple solutions, corresponding to multiple MPE. Further, MPE strategies can be mixed whereas the optimal policy in a single agent MDP is generically pure. There are no known algorithms that are guaranteed to find even a single MPE, since unlike the single agent case, the system of Bellman equations is no longer a contraction mapping.⁶ We now summarize several standard algorithms for solving the Bellman equation in the single agent setting.

2.1 Successive Approximations

As is well known, any contraction mapping on a Euclidean space has a unique fixed point, and the contraction property guarantees that the method of *successive approximations* i.e. the sequence $\{V_t\}$ given by $V_t = \Gamma(V_{t-1})$, for $t = 1, 2, \dots$ starting from any initial starting guess V_0 will converge to the fixed point V . Successive approximations converge to V at a geometric rate in β and can be very slow when β is close to 1.

2.2 Policy Iteration

In many cases, a much faster method for solving the MDP is the *policy iteration* (PI) algorithm due to Howard (1960). This algorithm cycles between *policy valuation* and *policy improvement* steps. At iteration t of the PI algorithm there is a candidate policy $\pi_t(a|s)$. The policy valuation step calculates the value implied by this policy as $V_{\pi_t} \in R^{|S|}$ as the solution to the linear system

$$V_{\pi_t} = r_{\pi_t} + \beta E_{\pi_t} V_{\pi_t}, \quad (4)$$

where $r_{\pi_t} \in R^{|S|}$ is the vector whose s^{th} element is $r(s, \pi_t(s))$ and E_{π_t} is the $|S| \times |S|$ matrix whose with element in row s and column s' given by $E_{\pi_t}(s, s') = p(s'|s, \pi_t(a|s))$. It is easy to show that for any feasible policy π that E_π is a Markov transition probability matrix with matrix norm $\|E_\pi\| = 1$, which implies that equation (4) has a unique solution given by $V_{\pi_t} = [I - \beta E_{\pi_t}]^{-1} r_{\pi_t}$.⁷ Given the value V_{π_t} the policy improvement step calculates a new policy π_{t+1} given by the right hand side of equation (3) except with V_{π_t} substituted for V . If $\pi_{t+1} = \pi_t$ then policy iteration has converged and it is not hard to show that $V_{\pi_t} = V$, the unique solution to Bellman's equation (2). Puterman and Brumelle (1979) showed that PI is mathematically equivalent to using Newton's method to find a zero of the function $F(V) = V - \Gamma(V)$, and as well known Newton's method is quadratically convergent. In fact, for discrete MDPs, PI converges in a finite number of iterations and typically the number of iterations is quite small, regardless of how close β is to 1. The main downside of PI is that for problems where the state space is large, the work required to solve the linear system (4) for V_{π_t} at each policy valuation step involves $O(|S|^3)$ operations. Thus, in the

⁶Iskhakov et al. (2015) introduced the *recursive lexicographical search* (RLS) that can provably find all MPE in the subclass of dynamic directional games.

⁷Alternatively we can show that the policy-specific Bellman operator Γ_{π_t} in equation (4) is a contraction.

absence of sparsity or other “special structure” for the matrices E_{π_t} , alternative algorithms must be relied on for problems where the number of states $|S|$ exceeds several hundred thousand.

2.3 Avoiding the Curse of Dimensionality with function approximation

The computational burden of solving an MDP when the state space is large was recognized early on: Bellman referred to the problem as the *curse of dimensionality*. For example in problems where s is a continuous vector in R^d , if we approximate Bellman’s equation using interpolation with a finite grid N points in each dimension, then $|S| = N^d$ and the total work to solve the MDP is $O(N^{3d})$, which increases exponentially fast in d . Even when the state space S is naturally discrete and finite, $|S|$ can be very large: for example in classic board games such as chess, the number of possible states (board positions) is estimated to be approximately 10^{44} . So it is clearly infeasible to use PI or even SA to try to solve MDPs.

Bellman and Dreyfus (1959) were among the first to recognize that one way to deal with the curse of dimensionality is to approximate the value function parametrically. For example we could represent it using a linear combination of fixed basis functions or approximate it with a neural network. Let ϕ denote the vector of parameters defining a parametric approximation to V with V_ϕ denoting the value function evaluated at parameter ϕ . Then an alternative way to approximate the solution to Bellman’s equation is to convert it into a nonlinear least squares problem where $\hat{\phi}$ is chosen to minimize the squared “Bellman residuals” over a grid of points (s_1, \dots, s_J) in S

$$\hat{\phi} = \underset{\phi}{\operatorname{argmin}} \sum_{j=1}^J [V_\phi(s_j) - \Gamma(V_\phi)(s_j)]^2. \quad (5)$$

The implied value function $V_{\hat{\phi}}$ is an approximate solution to Bellman’s equation, (2). We can also consider a modification of PI that we call *parametric policy iteration* (PPI), Benitez-Silva et al. (2000), where we use a parametric approximation, $V_{\hat{\phi}_t}$ and find $\hat{\phi}_t$ to minimize a least squares criterion similar to (5) to approximate the solution $V_{\pi_t} = r_{\pi_t} + \beta E_{\pi_t} V_{\pi_t}$ to each policy valuation step t , (4). When V_ϕ is approximated as a linear combination of “basis functions”, $\hat{\phi}_t$ can be computed by linear regression rather than a nonlinear regression as in (5) when we replace the nonlinear Bellman operator Γ with the linear operator E_{π_t} . Most of the modern methods for solving high dimensional MDPs involve the use of function approximation, either for the value function or the policy function, or sometimes both simultaneously.

2.4 Reinforcement Learning

This is a class of stochastic iterative algorithms for solving for the value function V and the optimal policy π . A prominent example is *Q-learning* introduced by Watkins (1989) as a “general formal model of an animal’s behavioural choices in its environment” and how it improves over time in response to feedback. In this sense Q-learning embodies a type of “learning” though for

our purposes it can be viewed as a stochastic algorithm for solving MDPs since “Q” is just the choice-specific value function in an alternative way of expressing the Bellman equation

$$Q(s, a) = r(s, a) + \beta \sum_{s'} V(s') p(s'|s, a), \quad (6)$$

$$V(s) = \max_{a \in A(s)} Q(s, a). \quad (7)$$

Note that by substituting equation (7) into equation (6) we can represent Q as a fixed point $Q = \Lambda(Q)$ to a mapping Λ that can also be shown to be a contraction mapping. $Q(s, a)$ represents the value to taking action a in state s and the optimal policy is to choose the action with the highest $Q(s, a)$, so $V(s)$ is the largest Q value in state s . Watkins proposed a stochastic iterative algorithm for updating Q given by

$$Q_{t+1}(s_t, a_t) = Q_t(s_t, a_t) + \alpha_t \left[r(s_t, a_t) + \beta \max_{a \in A(s_t)} [Q_t(\tilde{s}_{t+1}, a)] - Q_t(s_t, a_t) \right], \quad (8)$$

where (s_t, a_t) represent the state and action of the DM at iteration t , and \tilde{s}_{t+1} is a random draw from the transition probability $p(s'|s_t, a_t)$. The iteration (8) is a type of *stochastic approximation* [Robbins and Munro \(1951\)](#). The step size sequence $\{\alpha_t\}$ must converge to 0 but at a sufficiently slow rate to guarantee the probability 1 convergence of the stochastic sequence $\{Q_t\}$ to $Q = \Lambda(Q)$ the unique Q function implied by the modified version of Bellman’s in equations (6) and (7) as shown by [Watkins and Dayan \(1992\)](#) and [Tsitsiklis \(1994\)](#).⁸

Equation (8) does not specify how a_t is chosen: a typical choice is the *greedy policy* $a_t = \operatorname{argmax}_{a \in A(s_t)} Q_t(s_t, a)$. However a key condition for the convergence of Q-learning is that every state-action pair (s, a) must be sampled an infinite number of times. This implies the usual exploration/exploitation tradeoff since suboptimal actions must be taken infinitely often for order for Q learning to converge, but this need not happen under a greedy policy. One solution to this problem is to use a ε -greedy policy that takes the optimal action a_t with probability $1 - \varepsilon$, and with probability ε a randomly selected suboptimal action (i.e. an action a with $Q_t(s_t, a) < Q_t(s_t, a_t)$). However unless ε converges to 0 with t , Q_t will converge to a Q that does not satisfy the Bellman equation (7) since an ε -greedy policy is suboptimal.

[Haarnoja et al. \(2017\)](#) introduced *soft-Q learning* building on the work of [Ziebart et al. \(2008\)](#) on *maximum entropy reinforcement learning* that provides an alternative way to generate continuous exploration, ensuring that all actions are explored infinitely often. These result in policies that converge to mixed strategies that are self-consistent in the sense that the corresponding Q satisfy a modified version of the Bellman equation (2) they refer to as the *soft Bellman equation*. The term “maximum entropy” refers to the inclusion of an “entropy regularization” term in the agent’s objective function to result in optimal policies that are generically mixed. Specifically, suppose instead of the original objective in equation (1), we assume the agent’s objective includes an entropy

⁸The conditions on the step size are 1) $\sum_t \alpha_t = \infty$ and 2) $\sum_t \alpha_t^2 < \infty$. Note $\alpha_t = 1/t$ satisfies these conditions.

term, $H(\pi(s)) = -\sum_{a \in A(s)} \pi(a|s) \log(\pi(a|s))$, so the agent maximizes

$$V(s) = \max_{\pi} V_{\pi}(s), \quad \text{where } V_{\pi}(s) = E_{\pi} \left\{ \sum_{t=0}^{\infty} \beta^t [r(s_t, a_t) + \sigma H(\pi(s_t))] | s_0 = s \right\}, \quad (9)$$

where σ is the weight on entropy. Note that when $\sigma = 0$ we obtain the original MDP objective (1) and the optimal π is generically pure and has 0 entropy, but as $\sigma \rightarrow \infty$ the agent’s optimal policy is simply to choose each feasible action with equal probability, resulting in a π with maximum entropy. The solution V to the modified DP problem (9) turns out to satisfy a “soft Bellman equation” which turns out to be identical to a version of the “smoothed Bellman equation” derived by Rust (1988), see equations (16) and (15) of section 3.

Q-learning is described as *model-free* since, unlike SA or PI, it does not require explicit calculation of expectations using the transition probability $p(s'|s, a)$: instead all that is needed is be able to *simulate new states using* p . In Watkins’ example of how an animal learns to adapt and thrive in its environment, it would be the environment that “draws” the successor state s_{t+1} needed to implement the update in Q_t in equation (8). When trained in this manner Q-learning is indeed “model-free” and so are other RL algorithms that are trained “online”.⁹

In many situations online Q-learning and RTDP produce policies that are too noisy and converge too slowly to be useful in practical real time applications, especially when the number of states is large. As Jiang and Xie (2024) note, their “decisions can lead to undesirable outcomes, especially at the early stages of learning when the algorithm has little knowledge of the environment. This is not a problem when the environment is a simulator, but can lead to serious consequences when people are part of the environment.” As a result RL training is increasingly done *offline* with sufficient computer power to rapidly conduct many thousands or millions of training iterations to produce a much better solution than could be achieved from a limited amount on real-time, online experience. However when models are trained offline using computer simulations, they are no longer “model-free” since the computer needs the transition probability $p(s'|s, a)$ to make simulated draws of the successor states, s_{t+1} , needed by the Q-learning algorithm. The ability to simulate from the true $p(s'|s, a)$ (or a good approximation to it) and for sufficient number of iterations, is the key to producing good solutions using RL. In section 4 we return to this issue, showing that when we have sufficient data on state/action transitions in data on *past behavior* IRL can use this information to learn agents’ rewards in an offline training in a way that is model-free.

Tabular (i.e. finite state/action) Q-learning, (8), is far too slow for MDPs with large state or action spaces. One scalable solution to this problem is to use a Deep Q-Network (DQN), which is a deep convolutional neural network with parameters ϕ that parameterizes the Q-function as $Q_{\phi}(s, a)$ similar to the way V is approximated by V_{ϕ} in (5). This transforms the problem from iteratively updating Q-values, as in (8) to using stochastic gradient descent to find ϕ that minimizes the mean

⁹See Barto et al. (1995) for other examples of online, model-free algorithms for solving DP problems they refer to as *real time dynamic programming* (RTDP).

squared Bellman error $\rho(\phi)$ corresponding to (7) given by

$$\rho(\phi) = \sum_{i=1}^N \left[r(s_i, a_i) + \beta \sum_{s'} \max_{a \in A(s_i)} Q_\phi(s', a) p(s'|s_i, a_i) - Q_\phi(s_i, a_i) \right]^2, \quad (10)$$

over a grid of N points $\{(s_i, a_i)\}$. However the problem with directly minimizing $\rho(\phi)$ over ϕ in (10) is that forming a grid is challenging and the transition probability p may be unknown. Mnih et al. (2015) introduced a stochastic iterative procedure that converges to a $\hat{\phi}$ that attempt to approximately minimize the Bellman error $\rho(\phi)$ to find optimal strategies for the game of chess, and later for other board games see Silver et al. (2018), using *self-play*. This is an iterative process that involves simulating a trial policy π to generate a simulated data set \mathcal{D} of (s, a, s') sequences, and then using stochastic gradient descent to minimize the an approximation version of the squared Bellman error (??) using the mean squared *temporal difference error*

$$\rho(\phi) = \mathbb{E}_{(s,a,s') \sim \mathcal{D}} \left[\left(r(s, a) + \beta \max_{a'} Q_{\hat{\phi}}(s', a') - Q_\phi(s, a) \right)^2 \right], \quad (11)$$

where $\mathbb{E}_{(s,a,s') \sim \mathcal{D}}$ denotes the sample average over the simulated data set \mathcal{D} . Note that in (11) $\hat{\phi}$ is a lagged value of the network parameters from previous training iterations. Use of $\hat{\phi}$ instead of ϕ in (11) is motivated by the fact that the combination of off-policy learning, bootstrapping from subsequent estimates, and the use of a flexible nonlinear function approximator (the ‘deadly triad’) can lead to instability and non-convergence as noted by (Sutton and Barto, 2020). Two innovations in the DQN avoided these problems. The first was *experience replay*, whereby transitions are stored in a large “replay buffer” \mathcal{D} . The network is then trained on mini-batches of transitions sampled uniformly randomly from \mathcal{D} , breaking the strong temporal correlations in the sequence of observations (e.g. persistent streaks in certain parts of the state space), thereby stabilizing the updates by conforming more closely to the i.i.d. assumptions underlying most stochastic gradient methods. The second innovation is the use of a separate *target network* with parameters $\hat{\phi}$, $Q_{\hat{\phi}}$, is a lagged version of the current or “online” network Q_ϕ . Keeping the target network’s parameters $\hat{\phi}$ fixed for several iterations of the overall stochastic descent algorithm, keeps the optimization target stable, preventing the online network from attempting to converge to a non-stationary objective. Together, these innovations mitigated the deadly triad and enabled Q-learning to scale to very high-dimensional problems that were previously intractable.¹⁰

¹⁰Mnih et al. (2015) demonstrated that a single deep reinforcement learning agent could learn to play a diverse suite of Atari 2600 games directly from raw pixel inputs and game scores. It reused the same neural network architecture and hyperparameters across all tasks. Without any game-specific engineering, the DQN agent achieved human-level performance in several games. Despite this achievement, it should be highlighted that they did not demonstrate that their DQN converged to the optimal value function as the convergence of the approach is still an open question. Proofs of convergence for stochastic RL algorithms involving function approximation do exist: Tsitsiklis and Roy (1997) proved the probability 1 convergence for Temporal Difference (TD) methods in the case where V_ϕ is parameterized as a linear combination of basis functions.

Later studies confirmed that minimizing Bellman error in deep RL does not ensure recovery of the optimal value function. Empirical results show that agents can achieve low Bellman error on replay data while learning suboptimal policies due to overfitting and distribution shift (Fu et al., 2019). Theoretical work provides explicit examples where many Q-functions satisfy the Bellman equation on sampled data but only one is optimal, and Bellman error can be minimized by poor value functions in the finite data regime (Fujimoto et al., 2022). Despite these gaps, deep RL results in policies that exhibit human level and often superhuman levels of performance in a variety of high dimensional problems where standard methods such as value or policy iteration are infeasible.

3 Structural Estimation of DDCs

The development of new recursive methods for solving DPs and MDPs in the 1950s and 60s sparked a revolution in economics because these tools enabled economists to model sequential decision making under uncertainty, dynamic strategic interactions, and dynamic equilibria in markets which are key features of nearly all economic problems. Starting in the 1960s these tools facilitated the development of new dynamic economic theories that showed how expectations of the future affect the current choices of rational forward looking agents, resulting in new dynamic economic theories embodying *rational expectations* Muth (1961). Since this new framework had the flexibility to predict a wide range of behavior, the natural next step was to develop empirical methods for estimating and testing how well these new theories can explain data. This led to a new area of econometrics that can be called *dynamic structural estimation* (DSE) which differed from existing econometric methods that could be classified as *reduced form estimation* (RFE). The term “structural” refers to the attempt to infer the underlying preferences and beliefs of economic agents rather than simply summarizing their behavior, the main goal of reduced form econometrics. In terms of the MDP framework, we can consider an agent’s decision rule $\pi(a|s)$ to be the “reduced form model” because it captures their behavior, whereas structural econometrics tries to infer the underlying *structure* i.e. the preferences and beliefs $\{\beta, r, p\}$ as well as the implied optimal decision rule π .¹¹

Robert E. Lucas (1976) provide a practical rationale for why we do structural econometrics. He argued that reduced-form econometric models produce misleading predictions of the effect of economic policy changes, whereas structural models that estimate the underlying preferences of consumers and technologies of firms in production of goods, are more likely to accurately predict the effects of policy changes. Policy evaluation requires counterfactual predictions of how economic agents and the economy will react to changes in economic policy and the economic environment.

¹¹The earliest development of structural econometrics dates back to work at the Cowles Foundation in the 1940s where econometric methods were developed for inference on *linear simultaneous equation systems* which are static models of market equilibrium. In this work the structure was the coefficients of linear market supply and demand curves, whereas the reduced form were linear regressions that predict how equilibrium prices and quantities (p, q) shift with various variables x that shift the supply and demand curves. This was the earliest work showing how it could be possible to infer the structure (i.e. the coefficients of the supply and demand curves) using only data (p, q, x) on the intersections of the supply and demand curves in different market states x .

To explain Lucas’s argument in the MDP framework, let ρ denote a set of environment or policy variables that affect the preferences (or profits) and beliefs of economic agents. We can write the reward function as $r(s, a, \rho)$ and beliefs as $p(s'|s, a, \rho)$ to emphasize the dependence of these quantities on ρ . We can solve the MDP model for different values of ρ to make counterfactual predictions of how behavior $\pi(a|s, \rho)$ and welfare $V(s, \rho)$ change in response to changes in ρ . As noted, when there is insufficient variation in ρ (or we want to consider changes in ρ beyond the range of historical experience), reduced form methods cannot provide good estimates of how $\pi(a|s, \rho)$ as ρ changes, nor can they predict how it affects welfare. However it is sometimes possible to know how ρ affects $r(s, a, \rho)$ and $p(s'|s, a, \rho)$ even when there has been little or no variation in ρ in the past. The introduction discussed the example of new car taxation in Denmark discussed in the introduction where ρ is the tax rate on new cars. Under the assumption that $r(s, a, \rho)$ additively separable in (s, a) and ρ Gillingham et al. (2022) used structural estimation to recover the component of $r(s, a, \rho)$ that depends only on (s, a) , and with this, they could calculate $r(s, a, \rho)$ for any other possible tax rate ρ , due to their assumption about its functional form. In contrast, we rarely have a good idea of the functional form of $\pi(a|s, \rho)$ but structural estimation of $r(s, a, \rho)$ and $p(s'|s, a, \rho)$ enables us to calculate $\pi(a|s, \rho)$ for any policy ρ by solving the MDP. These solutions constitute counterfactual predictions of how car holding and trading behavior and welfare changes in response to changes in the new car tax rate ρ .¹²

The earliest work on estimation of dynamic structural models includes Sargent (1978), who estimated the dynamic linear quadratic model¹³ of firms’ demand for workers by maximum likelihood. In the remainder of this section we provide a selective summary of the literature on structural estimation of DDCs that has the closest connection to the literature on IRL. Key references include Rust (1987), Rust (1988), Rust (1994), Aguirregabiria and Mira (2002), and the survey by Aguirregabiria and Mira (2010).

The goal is to infer the structural objects $\{\beta, r, p\}$ from data on the trajectories of choices and states on a random sample of individuals using a modified version of the finite state and action MDP discussed in section 2. The modification is motivated by the *statistical degeneracy* of the optimal decision rule $\pi(a|s)$ in equation (3), i.e. it is generically a pure strategy (i.e. deterministic function of s). For most choices we actually observe, different individuals will choose different actions in the same observed state s . In order to explain this heterogeneity in behavior, we consider a modified MDP with an augmented state variable $x = (s, \epsilon)$ where the individual observes both s and ϵ (and thus uses a pure strategy) but we only observe s but not ϵ , which is a vector of idiosyncratic preference shocks that we do not observe. As a result, from our standpoint individuals appear to be using a mixed strategy. If there is sufficient variation in the unobserved ϵ component, we can derive a *conditional choice probability* (CCP) $\pi(a|s)$ from the overall (pure) strategy $\pi(a|x)$ that

¹²Since changes in tax rates ρ also affect prices of new and used cars, a counterfactual prediction also has to account for how ρ affects the entire equilibrium in the Danish car market, so both new and used car prices are implicit functions of ρ . These equilibrium changes are also accounted for in the structural model of Gillingham et al. (2022).

¹³That is, the law of motion for s_t is a linear autoregression and r is a quadratic function of (s_t, a_t) .

admits positive probabilities for all feasible choices a .

Building on the contributions of [McFadden \(1973\)](#) for static discrete choice models, we can derive an especially simple expression for the CCP $\pi(a|s)$ given by the *multinomial logit model* or “softmax” function in the RL literature. Suppose the choice sets depend only on s and are finite, and assume that for each s , ε is a vector with the same length as $|A(s)|$ the number of actions, so for $\varepsilon(a)$ is the component of ε associated with action $a \in A(s)$ so the reward for action a in state a $r(x, a)$ can be written as $r(s, a) + \varepsilon(a)$. We call this the Additive Separability (AS) assumption. We also assume that the transition density $p(x'|x, a)$ for (x, ε) given a factors as $f(\varepsilon'|s')p(s'|s, d)$, which we call the Conditional Independence (CI) assumption. Finally, we assume ε has a multivariate Gumbel or Type 1 extreme value distribution given by

$$F(\varepsilon|s) = \prod_{a \in A(s)} \exp\{-\exp\{-(\varepsilon(a) - \mu)/\sigma\}\}, \quad (12)$$

where we set $\mu = -\sigma\gamma$ where γ is Euler’s constant, which implies that $E\{\varepsilon(a)|s\} = 0$ for $a \in A(s)$. We call this the Extreme Value (EV) assumption. Let $V(x) = V(s, \varepsilon)$ be the optimal value for this MDP, the solution to

$$V(s, \varepsilon) = \max_{\pi} V_{\pi}(s, \varepsilon), \quad \text{where } V_{\pi}(s, \varepsilon) = E_{\pi} \left\{ \sum_{t=0}^{\infty} \beta^t [r(s_t, a_t) + \varepsilon_t(a_t)] \mid s_0 = s, \varepsilon_0 = \varepsilon \right\}, \quad (13)$$

Under assumptions AS, CI and EV, we can write the Bellman equation (2) for $V(s, \varepsilon)$ and show it has the representation

$$V(s, \varepsilon) = \max_{a \in A(s)} [Q(s, a) + \varepsilon(a)], \quad (14)$$

where $Q(s, a)$ is given by

$$Q(s, a) = r(s, a) + \beta \sum_{s'} \sigma \log \left(\sum_{a' \in A(s')} \exp\{Q(s', a')/\sigma\} \right) p(s'|s, a). \quad (15)$$

The derivation of (14) and (15) makes use of an important property of the extreme value distribution for ε , namely that the expectation of the expectation of $V(s, \varepsilon)$ with respect to ε which has an analytical expression, i.e. what is often called the log-sum or “smoothed max function” $V(s)$ given by

$$\begin{aligned} V(s) &\equiv E \{V(\tilde{s}, \tilde{\varepsilon}) \mid \tilde{s} = s\} = \int_{\varepsilon} \max_{a \in A(s)} [Q(s, a) + \varepsilon(a)] F(d\varepsilon|s) \\ &= \sigma \log \left(\sum_{a \in A(s)} \exp\{Q(s, a)/\sigma\} \right). \end{aligned} \quad (16)$$

Using equations (15) and (16) we can derive a “smoothed Bellman equation” for V

$$V(s) = \sigma \log \left(\sum_{a \in A(s)} \exp \left\{ [r(s, a) + \beta \sum_{s'} V(s') p(s'|s, a)] / \sigma \right\} \right), \quad (17)$$

Equation (15) defines a “smoothed Q ” as the fixed point of the smoothed contraction mapping $Q_\sigma = \Lambda_\sigma(Q_\sigma)$, and (17) defines V as a fixed point of the smoothed Bellman operator (also a contraction mapping) $V_\sigma = \Gamma_\sigma(V_\sigma)$. These operators are equivalent to the corresponding operators for MDP without the state variable ε in equations (6) and (7) and it is not hard to show that Q_σ and V_σ converge to Q and V as $\sigma \downarrow 0$.

The representation of $V(x, \varepsilon)$ in equation (14) means that the decision rule $\pi(a|x, \varepsilon)$ is isomorphic to a static discrete choice model where $Q(s, a) + \varepsilon(a)$ is the utility/payoff from action a . McFadden (1973) showed that the CCP $\pi(a|s)$ is the conditional probability that the DM chooses action $a \in A(s)$ given the observed state s and has the classic multinomial logit formula given by

$$\begin{aligned} \pi(a|s) &= \int I \left\{ Q(s, a) + \varepsilon(a) \geq \max_{a' \in A(s)} [Q(s, a') + \varepsilon(a')] \right\} F(d\varepsilon|s) \\ &= \frac{\exp\{Q(s, a)/\sigma\}}{\sum_{a' \in A(s)} \exp\{Q(s, a')/\sigma\}}. \end{aligned} \quad (18)$$

It is easy to show that as $\sigma \downarrow 0$ $\pi(a|s)$ converges to a degenerate probability that equals 1 if $Q(s, a)$ is the largest Q value and 0 otherwise, i.e. to the optimal pure strategy π in equation (3) for MDPs without the ε variable. It follows that the choice-specific values $Q(s, a)$ in (15) and CCPs $\pi(a|s)$ in (18) are exactly the same as the “soft Q ” values and optimal policy derived in the RL literature under the assumption that agents care about entropy of the policy π , see the objective (9) in section 2.4.

Equation (18) can be considered as a version of the *policy improvement* step for MDPs in equation (3). Is there an analogous equation for the policy valuation step (4) for regular MDPs where the value for any policy π is the solution to a linear system? Aguirregabiria and Mira (2002) showed that the value V_π corresponding to any CCP π is given by

$$V_\pi(s) = \sum_{a \in A(s)} \pi(a|s) [r(s, a) + E\{\varepsilon(a)|s, a\} + \beta \sum_{s'} V_\pi(s') p(s'|s, a)], \quad (19)$$

where $E\{\varepsilon(a)|s, a\}$ is the conditional expectation of $\varepsilon(a)$ given that action a is chosen in state s under decision rule π . Under the extreme value assumption for ε , we have $E\{\varepsilon(a)|s, a\} = -\sigma \log(\pi(a|s))$. Using this, it is easy to see that the right side hand of (19) equals the smoothed max function defining the smoothed Bellman operator in equation (17). Thus, we can solve the smoothed Bellman equation $V_\sigma = \Gamma_\sigma(V_\sigma)$ using policy iteration just as in the case of MDPs, where we cycle between policy valuation steps (19) and policy improvement steps (18) using the Q_π

implied by V_π in equation (15). In fact, just as in the case of MDPs, policy iteration is equivalent to Newton's method for solving the system $V_\sigma - \Gamma_\sigma(V_\sigma) = 0$.

Equation (18) implies that $\sigma \log(\pi(a|s)) = Q(s, a) - V(s) \equiv \mathcal{A}(s, a)$, known as the *advantage function* in RL. Using $\mathcal{A}(s, a)$ we can rewrite the nonlinear fixed point equation (15) for Q as a linear system depending on $r(s, a)$ and $\log(\pi(a|s))$,

$$Q(s, a) = r(s, a) + \beta \sum_{s'} \sum_{a' \in A(s')} [-\sigma \log(\pi(a'|s')) + Q(s', a')] \pi(a'|s') p(s'|s, a), \quad (20)$$

Equation (20) implies that we can decompose Q as $Q(s, a) = Q_r(s, a) + Q_\varepsilon(s, a)$ where $Q_r(s, a)$ is the component of discounted value from observed rewards r and $Q_\varepsilon(s, a)$ is the component from unobserved rewards ε . Further substituting expression (19) in place of (17) in the equation for Q in (15), we can show that Q_r and Q_ε are each the solution to the following linear systems

$$Q_r(s, a) = r(s, a) + \beta \sum_{s'} \sum_{a' \in A(s')} Q_r(s', a') \pi(a'|s') p(s'|s, a) \quad (21)$$

$$Q_\varepsilon(s, a) = \beta \sum_{s'} \sum_{a' \in A(s')} [-\sigma \log(\pi(a'|s')) + Q_\varepsilon(s', a')] \pi(a'|s') p(s'|s, a). \quad (22)$$

This decomposition can be exploited to reduce the computational burden of estimating DDC models as we will show below.

3.1 Estimation Methods for DDCs

The most common estimation method for DDCs is some form of maximum likelihood. Suppose we observe a random sample of *trajectories* i.e. data on the consecutive states and actions taken by N individuals $\{(s_{it}, a_{it}) | t = 0, \dots, T_i, i = 1, \dots, N\}$. Suppose that we smoothly parametrize (r, p) using a vector unknown parameters θ , which we denote by r_θ and p_θ .¹⁴ Using the implicit function theorem, we can show that V_θ and Q_θ , the contraction fixed points given in equations (16) and (15), are continuously differentiable function of θ . Using these fixed points and the logit formula for $\pi(a|s)$ in equation (18) we can write a full likelihood function for the data as a function of the unknown parameters θ given by

$$L(\theta) = \prod_{i=1}^N \prod_{t=1}^{T_i} \pi_\theta(a_{it}|s_{it}) p_\theta(s_{it}|s_{it-1}, a_{it-1}). \quad (23)$$

The smoothness of V_θ and Q_θ in θ implies that the CCP π_θ and the likelihood $L(\theta)$ are both smooth functions of θ so the standard asymptotic efficiency properties of maximum likelihood apply.

Rust (1988) introduced the Nested Fixed Point algorithm (NFXP) to compute the maximum likelihood estimator (MLE) of θ . It consists of a standard outer hill climbing algorithm to maximize

¹⁴The discount factor β is often treated as known, though in some cases it is also estimated and thus part of θ .

$L(\theta)$ over θ , but at each evaluation of θ it calculates the contraction fixed point $Q_\theta = \Lambda_\theta(Q_\theta)$ that defines the choice-specific values Q_θ where Λ_θ is the operator defined on the right hand side of equation (15). Notice that $L(\theta)$ depends on θ via Q_θ . Su and Judd (2012) proposed an alternative strategy for computing the MLE based on the MPEC algorithm that maximizes a likelihood $L(\theta, Q)$ expressed as a function of (θ, Q) , subject to the fixed point constraint $Q_\theta = \Lambda_\theta(Q_\theta)$.¹⁵

Estimation of DDC models is often done in a two step fashion, where a parametric model of the transition probability $p(s'|s, d)$ is estimated in the first stage using a partial likelihood that does not involve the CCP π in (23), and in the second stage the parameters of r are estimated using another partial likelihood involving only π , treating the estimated parameters of p as if they were known. A final Newton step iteration for maximization of the full likelihood function (23) can be done using the estimated parameters of p and r from steps 1 and 2 as starting values to obtain fully efficient parameter estimates with the correct asymptotic covariance matrix, i.e. it accounts for “first stage” estimation effort in the parameters of p .

The computational burden of NFXP stems from the need to compute the inner fixed point $Q_\theta = \Lambda_\theta(Q_\theta)$ for each trial value of θ in the outer hill-climbing algorithm that maximizes $L(\theta)$. Newton’s method (e.g. policy iteration) is typically used to solve for Q_θ (or alternatively for V_θ in equation (19)). Even though Newton’s method converges rapidly, it requires $O(|S|^3)$ operations per Newton/policy iteration step, making NFXP burdensome for large state spaces S . Therefore alternative estimation methods have been proposed that reduce the total number of policy iteration steps required to estimate θ .

Hotz and Miller (1993) proposed the “CCP estimator” for r that uses a non-parametric estimate of $\pi(a|s)$ to avoid repeated solution of Q over the search for θ that NFXP requires. Consider the case where r_θ is linear in parameters, i.e. we can write $r_\theta(s, a) = r(s, a) * \theta \equiv \sum_{k=1}^K r_k(s, a) \theta_k$ for known functions $\vec{r} = (r_1, \dots, r_K)$ called *features* in the IRL literature. Then (21) implies that $Q_r(s, a) = R(s, a) * \theta$ where R is the solution to

$$R(s, a) = \vec{r}(s, a) + \beta \sum_{s'} \sum_{a' \in A(s')} R(s', a') \pi(a'|s') p(s'|s, a) = \vec{r}(s, a) + \beta ER(s, a), \quad (24)$$

where $ER(s, a)$ is a shorthand for the conditional expectation of $R(s, a)$ the middle of expression in (24). It follows that we only need to solve (24) once using a non-parametric estimate of $\hat{\pi}$ along with Q_ε in (22) to serve as a “correction term” to obtain the following form for the CCP

$$\pi_\theta(a|s) = \frac{\exp\{\hat{R}(s, a) * \theta + \hat{Q}_\varepsilon(s, a)\}}{\sum_{a' \in A(s)} \exp\{\hat{R}(s, a') * \theta + \hat{Q}_\varepsilon(s, a')\}}, \quad (25)$$

where $\hat{R}(s, a)$ is the solution to (24) and \hat{Q}_ε is the solution to (22) using non-parametric first stage

¹⁵Su and Judd (2012) claim that the MPEC algorithm is substantially faster than NFXP, but Iskhakov et al. (2016) showed that this was due to using SA to compute the fixed point $Q_\theta = \Lambda_\theta(Q_\theta)$ which is slow when β is close to 1. When Newton’s method is used to compute this fixed point, the cpu times for MPEC and NFXP are roughly comparable with NFXP being faster on test problems with large sample sizes.

estimators $\hat{\pi}(a|s)$ and $\hat{p}(s'|s, a)$. The CCP estimator $\hat{\theta}$ maximizes the partial likelihood $L_{\pi}(\theta)$ given by

$$L_{\pi}(\theta) = \prod_{i=1}^N \prod_{t=1}^{T_i} \pi_{\theta}(a_{it}|s_{it}), \quad (26)$$

where π is given in (25).¹⁶ Thus, when r is a linear combination of known features, we only need to compute \hat{R} and \hat{Q}_{ε} once at the start of the estimation, requiring $K + 1$ solutions of systems with $|S|$ equations and unknowns, where K is the number of features. This can be further reduced to a single solution if we assume that $r(s, a_s)$ is known for some “normalizing alternative” $a_s \in A(s)$ as we discuss further below in the section on identification of DDC models.

A cost of the CCP estimator is that it is less efficient than using NFXP to estimate the same partial likelihood criterion (26) due to the noise in the first stage non-parametric estimates $\hat{\pi}$ that are used to construct the \hat{Q}_{ε} function in (25). Aguirregabiria and Mira (2002) (AM) addressed this problem by proposing an iterative estimator called *nested pseudo likelihood* (NPL) that uses an initial non-parametric estimate $\hat{\pi}$ to compute the value function $V_{\hat{\pi}}$ implied by this initial estimate by solving for it in the policy valuation step in equation (19), which we denote as $\hat{V}_{\theta, \hat{\pi}}$ since it depends both on $\hat{\pi}$ and the parameters θ entering the reward function. They propose to estimate $\hat{\theta}$ using a pseudo-likelihood similar to (26) except that the CCP $\pi_{\theta}(a|s)$ is evaluated using (18) using estimated Q functions given by

$$\hat{Q}_{\theta}(s, a) = r_{\theta}(s, a) + \beta \sum_{s'} \hat{V}_{\theta, \hat{\pi}}(s'|s, a) p(s'|s, a), \quad (27)$$

Thus, the likelihood $L_{\pi}(\theta)$ as depends implicitly on the first stage non-parametric CCP estimate $\hat{\pi}$, which can be regarded as a “nuisance parameter”. In principle, estimation noise in $\hat{\pi}$ will contaminate and affect the asymptotic covariance matrix for the parameters of interest θ . However AM established a key *zero Jacobian property* that at the fixed point $\partial V(s)/\partial \pi = 0$ which implies that the first stage nuisance parameter $\hat{\pi}$ is “Neyman orthogonal” with respect to the parameters of interest θ in the sense of Chernozhukov et al. (2022). It follows that the NPL estimator of θ is unaffected by noise in the first stage $\hat{\pi}$ and is the same as if it was treated as known. AM showed that NPL can be iterated: using $\hat{\theta}$, they can compute an updated estimate of the CCPs ($\hat{\pi} = \pi_{\hat{\theta}}$) and repeat the policy valuation step to get a new set of Q functions (27), and use these in the likelihood (26) to get an updated estimate $\hat{\theta}$. They refer to NPL as “swapping” the policy iteration and maximization steps with respect to what NFXP does (i.e. it has maximization as the “outside” loop and policy iteration as the “inside” one), and doing so reduces the total number of policy iteration steps required to estimate θ . They showed that if this iterative version of NPL converges, it produces the same estimate $\hat{\theta}$ that NFXP produces from maximization of the partial likelihood (26). Thus the NPL estimator is as efficient as NFXP but at lower computational cost,

¹⁶See Chernozhukov et al. (2022) for a modified version of the CCP estimator that is “locally robust” i.e. it corrects for sampling error in the first stage estimates of the functions $\hat{H}(s, a)$ and $\hat{Q}_{\varepsilon}(s, a)$.

and it is more efficient than the CCP estimator but at a higher computational cost.¹⁷

Bayesian approaches have been proposed for inference in DDC models using Markov Chain Monte Carlo methods that simultaneously simulate the likelihood as well as stochastic simulations to compute estimates of value functions that are similar to “Monte Carlo tree search” used by RL algorithms such as the Q-learning algorithm discussed in section 2.4, see [Imai et al. \(2009\)](#) and [Norets \(2009\)](#). There are also non-Bayesian simulation estimators such as [Bajari et al. \(2007\)](#) that estimate θ as the value that minimizes the degree of suboptimality in observed actions relative to other actions not taken using simulation estimates of Q function evaluated at the actions the DM actually an other actions not chosen.

3.2 Function Approximation Methods for Large State Spaces

When the state space S is very large (i.e. the state space is continuous or the number of states $|S|$ is in the hundreds of thousands or higher), the all of the approaches discussed above that require one or more solutions of large linear systems with $|S|$ equations and unknowns become computationally infeasible. An alternative approach is to use parametric methods to approximate the value function V as $V_\phi(s) = f(s, \phi)$ where $f(s, \phi)$ is a member of a flexible family such as neural networks where ϕ denote the weights (parameters) defining the network, and then recasting the fixed point problems (15) and (16) as potentially more tractable nonlinear least squares problems. For example we can approximate V as $V_{\hat{\phi}}$ where $\hat{\phi}$ are the network weights that solve the nonlinear least squares problem of minimizing the squared Bellman residuals in equation (5).

$$\rho(\theta, \phi) = \sum_{j=1}^N [V_\phi(s_j) - \Gamma_{\theta, \sigma}(V_\phi)(s_j)]^2. \quad (28)$$

Note that the smoothed Bellman operator Γ_σ defined on the right hand side of (16) depends on r and p which in turn are function of the model parameters θ , so we write it as $\Gamma_{\theta, \sigma}$. It follows that the minimizing parameters $\hat{\phi}$ are implicit functions of θ .

Let $\rho(\theta, \phi)$ denote the squared Bellman residuals defined over a grid of N points in the state space Then a “nested least squares” estimator similar to NFXP can be used estimate θ in large state spaces, where we maximize $L(\theta)$, but in the inner fixed point problem has been transformed into the problem of minimizing $\rho(\theta, \phi)$ over the parameters ϕ . $\rho(\theta, \phi)$ for each trial θ . The CCP estimator can also be implemented in a similar fashion where we parametrize R and Q_ε by parameters ϕ_R and ϕ_ε and solve two nonlinear least squares problems to approximate the fixed points in (24) and (22), respectively. Note, that these would only need to solved *once* at the start of estimation, and the resulting solutions $R_{\hat{\phi}_R}$ and $Q_{\hat{\phi}_\varepsilon}$ can then used to evaluate $\pi_\theta(a|s)$ in (25).

[Luo and Sang \(2024\)](#) introduced a penalized likelihood estimator where $\hat{\theta}$ and $\hat{\phi}$ are simultane-

¹⁷See also [Dearing and Blevins \(2025\)](#) who also proposed a similar iterative procedure called *efficient psuedo likelihood* (EPL) that is also an efficient sequential estimator that can estimate parameters of dynamic games.

ously chosen to maximize

$$(\hat{\theta}, \hat{\phi}) = \underset{\theta, \phi}{\operatorname{argmax}} \log(L(\theta)) - \lambda \rho(\theta, \phi), \quad (29)$$

where $\lambda \geq 0$ is a penalization parameter. Note that the likelihood depends on both θ and ϕ since both the CCP π and Q are defined in terms of V_ϕ which depends on the network weights, ϕ , in addition to the structural parameters of interest, θ . The consistency of the penalized likelihood estimator requires the use of a *sieve*, i.e the dimension of ϕ must grow to infinity at certain rate with the sample size to guarantee that the error in approximating the true fixed point by V_ϕ converges to zero. As a result they refer to the estimator (29) as SEES, for Sieve-based Efficient Estimator for Structural models. See [Peter Arcidiacono and James \(2013\)](#) and [Kristensen et al. \(2021\)](#) for other examples of sieve-based estimators for DDC models.

A drawback of the SEES estimator is that the ϕ parameters determining the value function V_ϕ become nuisance parameters, and for fixed θ the ϕ that maximize (29) will be an implicit function of θ . As a result, computation of the asymptotic covariance matrix for θ requires inverting an analog of an information matrix for (θ, ϕ) jointly which can be challenging to do when ϕ is high dimensional. [Nguyen \(2025\)](#) introduced an estimator called *neural network efficient estimator* (NNES) that maximizes a penalized criterion similar to the SEES estimator (29) but in a sequential fashion similar to the NPL estimator of AM, benefitting from the zero Jacobian property of the NPL estimator and ensuring Neyman orthogonality between the parameters of interest θ and the nuisance parameters ϕ . NNES is initialized from an initial non-parametric estimator of the CCPs, $\hat{\pi}$, but instead of solving for the implied value function V exactly as a system of linear equations (19), it uses a neural network parameterization of the value function V_ϕ and finds $\hat{\phi}$ that minimizes the sum of squared Bellman residuals $\rho(\theta, \phi)$ in (28). The resulting value function, which also depends on θ (so we write $V_{\theta, \hat{\phi}}$) is then used to form the \hat{Q}_θ values per (27) which are in turn used to calculate the CCPs $\pi_\theta(a|s)$ and the pseudo likelihood $L_\pi(\theta)$ following the same approach as the NPL estimator. In essence, NNES is a convex combination of the NPL and SEES estimators: 1) like NPL it is iterative, starting from an initial nonparametric estimator $\hat{\pi}$, 2) unlike NPL but similar to SEES, it does not solve for $V_{\hat{\pi}}$ in equation (19) exactly via standard linear algebra, but rather minimizes the square Bellman residuals $\rho(\theta, \phi)$. The main benefit of the NNES estimator over SEES is that the zero Jacobian/Neyman orthogonality property implies that to get the covariance matrix for θ we only have to invert a matrix of the same dimension as θ rather than take the (θ, θ) block of the inverse of a matrix with dimension equal to the sum of the number of parameters in θ and ϕ . For flexible and deep neural networks, the dimension of ϕ can be in the thousands, so this can be a significant advantage,

3.3 The Identification Problem

At its core, structural estimation of DDC models can be regarded as a process of inverting observed behavior to uncover preferences. The observed behavior from the DDC model is captured by the

CCP $\pi(a|s)$, and it is referred to as a “reduced form” model of the agent’s behavior since with enough data on (a, s) pairs, we can estimate π nonparametrically without any attempt to uncover the agent’s preferences or beliefs by solving the MDP problem to derive π as outlined above. In econometrics this is referred to as reduced-form estimation (RFE), whereas in the ML literature it is referred to as “behavioral cloning” (BC) “imitation learning” (IM), or “apprenticeship learning” (AL). The analysis of identification treats π as known because with infinite data we could learn it perfectly. The preceding section showed how to derive π from the structure $\{\beta, r, p\}$, a process that we can summarize by the mapping $\pi = f(\beta, r, p)$. The identification problem concerns the invertibility of this mapping: given π is there a unique underlying structure $\{\beta, r, p\}$ that implies it? Unfortunately, without further restrictions the answer is negative, as shown by [Hotz and Miller \(1993\)](#), [Rust \(1994\)](#) and [Magnac and Thesmar \(2002\)](#). To understand why, consider the first step of the inversion process, uncovering Q from π . Using the logit form of π in (18) we have

$$\log(\pi(a|s)/\pi(a'|s)) = Q(s, a) - Q(s, a'), \quad s \in S, a, a' \in A. \quad (30)$$

We already see that information on π is insufficient to recover all the Q values: only the *differences* in Q are identified but not all $|S||A|$ possible values of $Q(s, a)$. Indeed, the structure $\{\beta, r, p\}$ contains $1 + |S|(|A| - 1) + |S|^2|A|$ parameters which is more than then $|S|(|A| - 1)$ parameters of π , so we will generally have “more equations than unknowns” so the underlying structure is not uniquely identified.

Thus, it is clear that in order to identification requires additional restrictions on $\{\beta, r, p\}$. For example we might impose the assumption that r and p have a parametric functional forms that depend on a relatively small number of parameters θ . For example [Anderson et al. \(2025\)](#) are able to identify the subjective beliefs of professional tennis servers about how their serve strategies affects their probability of winning a tennis game. These beliefs take the form of a transition $p(s'|s, a)$ where s denotes the point state of the game and a denotes serve direction. Identification is possible because it is reasonable to assume r is known: i.e. the server earns a reward of 1 if they win the game and 0 otherwise, and due to reasonable parametric restrictions on p .

In many applications researchers are willing to assume agents have rational expectations which implies that their subjective beliefs about state transitions $p(s'|s, a)$ correspond to the actual probability governing these transitions and so p can be estimated non-parametrically given sufficient data, and thus can be treated as known but with restrictions on r this assumption is testable.¹⁸ Many studies typically assume that the discount factor β is also known, however under the assumption of rational expectations and restrictions on rewards, β can be identified, see [Abbring and Daljord \(2020\)](#).

Assuming $\{\beta, p\}$ are known, identification of r depends on the assumption that either it has a known parametric functional form r_θ where the dimension of θ is less than $|S|(|A| - 1)$, or that

¹⁸[Anderson et al. \(2025\)](#) strongly reject the hypothesis the many professional tennis servers have rational beliefs about their own ability and those of their opponent, as embodied by $p(s'|s, a)$.

for each $s \in S$ there is a normalizing or anchor action a_s such that $r(s, a_s)$ is known. The latter assumption amounts to $|S|$ restrictions that imply there are only $|S|(|A| - 1)$ unknown rewards to be recovered, the same number of free values of π . Hotz and Miller (1993) and Kang et al. (2025) show that in this “exactly identified” case, we can calculate $Q(s, a_s)$ from knowledge of $r(s, a_s)$ and $\pi(a_s|s)$ using equation (20) and then (30) allows us to recover all remaining Q values. Using the Q values we can back out the remaining unknown rewards r from the smooth Bellman equation (15).

Note that the assumption that we know $r(s, a_s)$ for some anchor action a_s in each state $s \in S$ is far from innocuous. If this assumption is wrong, even though the DDC model has enough parameters to perfectly fit π , the estimated rewards will not equal the true rewards. In this case, counterfactual behavior predicted by the model with an incorrectly identified reward function can differ from what would actually happen under the true reward. Another way to say this is that in the absence of any prior restrictions on rewards, we can only identify a set of “observationally equivalent” rewards of the form $r(s, a) + h(s) - \beta \sum_{s'} h(s') p(s'|s, a)$ for any function $h : S \rightarrow R$ that is called the partially identified set. This constitutes an inherent limit on what can be learned about agents’ rewards and emphasizes the importance of having accurate prior knowledge of at least some aspects of the reward function to obtain credible counterfactual predictions and welfare impacts of changes in policy and the environment. Kalouptsi et al. (2021) provides examples of certain counterfactuals that are identified even when r is only partially identified. They conclude that their “results call for caution while leaving room for optimism: although counterfactual behavior and welfare can be sensitive to identifying restrictions imposed on the model, there exists important classes of counterfactuals that are robust to such restrictions.” (p. 385).

4 Inverse Reinforcement Learning

The IRL literature was initiated by Ng and Russell (2000) who identified two motivations for inferring the underlying rewards that are presumed to generate observed behavior: 1) “the potential use of reinforcement learning and related methods as potential models for human and animal learning” and 2) to provide a reward function that RL can use to construct an “intelligent agent that can behave successfully in a particular domain.” Though they note that imitation and apprenticeship learning (IL and AL) can be used to learn a policy π , “the reward function often provides a much more parsimonious description of behaviour” so that IRL can constitute an effective form of AL. IRL has become an important method in many areas of robotics including development of autonomous drones and vehicles (AV) where there is insufficient data on decisions to human drivers under multitudes of different road conditions to develop effective AVs. Wang et al. (2025) note that IRL has “found great success in predicting trajectories through inferring cost or reward functions from expert demonstrations, and then using these functions to guide the behavior of self-driving vehicles in unseen driving environments.”

This section provides a selective review of the IRL literature focusing on methods that overlap most closely with the structural DDC literature reviewed in section 3. We refer readers to surveys

by Arora and Doshi (2021), Gleave and Toyer (2022), and Adams et al. (2022) for a broader and more in depth view of IRL. The latter survey provides further examples of how IRL uses “an expert to demonstrate the successful task and then a highly parameterized reward function could be transferred to the robot for optimization” and how IRL can generate more effective policies than AL, particularly for new, counterfactual environments where “IRL has the added advantage that the reward function can be transferred to new environments with different dynamics.”.

There are three broad classes of IRL algorithms: 1) max-margin, 2) Bayesian methods, and 3) maximum entropy methods. Ng and Russell (2000) introduced three different types of max-margin estimators, which search for a reward function for which payoffs under the policy of the observed expert are higher than any other policy (i.e. behavior), an approach that shares similarities with the “moment inequality” simulation estimator proposed by Bajari et al. (2007) discussed in section 3. They recognized the identification problem discussed in the previous section, namely there may be many reward functions that satisfy the max-margin inequalities (payoff inequalities) including the trivial solution $r = 0$.

Bayesian IRL (BIRL) was introduced by Ramachandran and Amir (2007). They assumed π is a mixed strategy with the softmax form (18) that depends on Q functions which are implicit functions or r per the Bellman equations (6) and (7) and interpret the σ parameter in the softmax function for π as controlling for the “degree of confidence we have in the [DM’s] ability to choose actions with high value.” This allowed them to formulate a likelihood for the sequence of observed actions by the DM that depends on r . Using Bayes’ Rule, they computed a posterior distribution for r , treating it as the unknown parameter with a prior distribution. Since it is infeasible to directly compute the posterior, BIRL uses MCMC to generate draws from the posterior distribution, with policy iteration used to solve for Q for each draw of r . This makes BIRL method computationally intensive for large state spaces similar to NFXP. Overall BIRL is similar to the Bayesian MCMC methods discussed in section 3, except that those methods use stochastic updates of Q similar to those produced by Q-learning to gradually compute Q over the course of the MCMC simulations rather than use policy iteration to find Q for each MCMC draw of r .

A seminal contribution to the IRL literature is Ziebart et al. (2008), who introduced *maximum entropy IRL* as a means of justifying the use of the softmax/logit probability for $\pi(a|s)$ as the probability with maximum entropy subject to certain constraints that make expectations of certain functions known as *features* equal the sample averages of these features.¹⁹ They proposed to estimate the parameters of a reward function that is a linear combination of a set of known features using a partial likelihood for observed (a, s) pairs that is identical to the partial likelihood (26) used in the DDC literature.²⁰ They applied this method to model the route choices of 25 Yellow cab drivers in Pittsburgh with a reward function that involves four features: road type, speed, lanes and transitions. They showed that the estimated model could correctly predict nearly 80% of the cab

¹⁹Ziebart et al. (2010) extended these ideas to MDPs and called it *Maximum causal entropy RL*.

²⁰The full likelihood (23) is not frequently used as the estimation criterion in IRL because of the model-free or non-parametric approach that bypasses the need to explicitly specify or estimate the transition probability $p(s'|s, a)$.

paths in a holdout or evaluation sample and that their new method was a significant improvement over other IRL methods such as max-margin.

These ideas evolved into the already cited work of Haarnoja et al. (2017) who characterized π as the policy that maximizes the expected discounted entropy-regularized payoff given in equation (9). Various criteria are used to estimate r such as “occupancy matching” (Ho and Ermon (2016), Yue et al. (2023)). For our purposes we focus on IRL where the likelihood function is the estimation criterion, such as Ermon et al. (2015), Zeng et al. (2023), Zeng et al. (2024), Barzegary and Yoganarasimhan (2022) and Kang et al. (2025).²¹ This results in a nearly complete equivalence between the DDC and IRL literatures, with identical softmax/logit formulas for the policy $\pi(a|s)$ (18) that depend on the choice-specific or “soft Q ” values that satisfy the “soft Bellman equation” in equation (15). Overall, DDC and maximum entropy IRL result in mathematically identical optimization problems for recovering r from data on states and actions of a sample of DMs. The main difference is in the algorithms used to solve the DM’s DP problem.

In DDC the Q functions are computed by non-stochastic algorithms such as SA or PI, or in large state spaces by parametrizing Q or the value function V and finding parameters ϕ that minimize the squared “Bellman residuals” as in the SEES estimator, (29). All of these methods can be described as *model-based* because they require knowledge of the transition probability $p(s'|s, a)$, or at least p must be estimable from data. Adams et al. (2022) notes that most IRL methods are model-based. However, we will focus on model-free IRL that does not require specification of a function form for p or even require that it be explicitly estimated non-parametrically. Instead, new offline IRLs inspired model-free methods such as Q-learning are able to approximate the Q functions using only realizations from p without knowing p itself. But how is it possible to generate realizations s' from $p(s'|s, a)$ without knowing p itself?

Model-free IRL solves this problem in a clever way: it uses already realized transitions in demonstration datasets to provide the “simulator” needed by RL algorithms to train the Q and values functions. The insight is that the data on observed states and decisions provides the “environment” and the state transitions needed by RL to find the value and Q functions needed to recover r , reminiscent of the idea of *experience replay* in the deep Q learning literature, Mnih et al. (2015).²² To get a better understanding of these methods, we will provide a more in-depth discussion of two recent contributions the IRL literature that can recover r from data in a model-free manner. One can be regarded as an IRL version of the CCP estimator and the other is similar to the SEES estimator (29). Both of these methods use parametric function approximation methods (expressing

²¹Zeng et al. (2023) and Zeng et al. (2024) maximize a discounted log-likelihood function.

²²Also related is the distinction between online and offline RL. IRL generally uses offline RL since the estimation of rewards is based on already collected datasets rather than being done in real time. See the survey of offline RL by Levine et al. (2020) who note “Offline reinforcement learning algorithms hold tremendous promise for making it possible to turn large datasets into powerful decision making engines.”. Yue et al. (2023) note that offline IRL also focuses on “learning from a previously collected dataset without online interaction with the environment” and “holds tremendous promise for safety-sensitive applications where manually identifying an appropriate reward is difficult but historical datasets of human demonstrations are readily available (e.g., in healthcare, autonomous driving, robotics, etc.).”

V and Q as outputs of neural networks or linear combinations of a set of basis functions) that enables them to handle problems with large state spaces.

Adusumilli and Eckhardt (2025) (AE) introduce a model-free two step CCP estimator similar the one described in the previous section but using RL-inspired methods to avoid estimating or making parametric assumptions about $p(s'|s, a)$. They do assume that the reward function is a linear combination of known “features”, i.e. $r_\theta(s, a) = \sum_{k=1}^K \theta_k r_k(s, a)$. They use an method from RL, *temporal difference* (TD) to approximate the functions R and Q_ϵ (defined in equations (22) and (24), respectively) that enter as covariates in the CCP π in equation (25), using transitions in the data to estimate these functions without having to explicitly estimate $p(s'|s, a)$. To accomodate empirical applications with large state spaces they propose parameterizing R and Q_ϵ as linear combinations of a set of known basis functions that depend on (s, a) . Let $\nu(s, a) = \{\nu_1(s, a), \dots, \nu_J(s, a)\}$ denote J basis functions that are used to approximate R nad Q_ϵ by a regression that uses the basis functions evaluated at values of (s, a) in the data set as regressors. We will describe AE’s approach to approximating R , since the same procedure is used to approximate Q_ϵ and is omitted for brevity. Let R_ϕ denote an approximation to R based on a $J \times 1$ parameter vector ϕ of the form $R_\phi(s, a) = \nu(s, a)\phi = \sum_{j=1}^J \theta_j \nu_j(s, a)$. AE propose to to choose $\hat{\phi}$ to minimize the mean-squared TD error given by

$$\hat{\phi} = \underset{\phi}{\operatorname{argmin}} E \{ [\bar{r}(s, a) + \beta R_\phi(s', a') - R_\phi(s, a)]^2 \}, \quad (31)$$

where E denotes the expectation operator over (s, a, s', a') . Using results from Tsitsiklis and Roy (1997) on the convergence of TD learning with linear function approximation, AE show that $\hat{\phi}$ can be computed using the sample expectation operator E_N in (31) using the formula

$$\hat{\phi} = [E_N \{ \nu(s, a) [\nu(s, a) - \beta \nu(s', a')]^T \}]^{-1} E_N \{ \nu(s, a) \bar{r}(s, a) \}, \quad (32)$$

and the implied estimate $\hat{R}(s, a) = \phi(s, a)\hat{\phi}$ of the features or “covariates” for the θ in (25). We also have to approximate the second “correction term” $\hat{Q}_\epsilon(s, a)$, which can be estimated via regression similar to (32) but it requires a first stage non-parametric estimate of $\pi(a|s)$ to construct estimates of the expected choice-specific unobserved shocks $E\{\varepsilon(a)|s, a\} = -\log(\pi(a|s))$.²³ Using the estimated $\hat{R}(s, a)$ and $\hat{Q}_\epsilon(s, a)$ values, they can compute a second state pseudo-maximum likelihood (PMLE) estimate of θ by maximizing the partial likelihood (26) using the formula for $\pi_\theta(a|s)$ in (25). Consistency of this first stage estimator requires specification of a sieve with J basis functions $\{\nu_1, \dots, \nu_J\}$ that increases with sample size N at an appropriate rate and with sufficient linear independence to allow any function R to be perfectly approximated by its projection $\Pi_\nu R$ in the limit as $J \rightarrow \infty$. The PMLE is affected by “estimation noise” in the \hat{R} and \hat{Q}_ϵ functions that can introduce bias and noise. They propose an additional step that uses the PMLE to create a modified score (i.e. gradient) of the partial log-likelihood (26) to create a robust (Neyman orthogonal) version of the score function which they used to their preferred estimator $\hat{\theta}$ as the

²³AE normalize σ to 1 since it cannot be separately identified from the θ parameters.

value that sets the modified score to zero following the approach of [Chernozhukov et al. \(2022\)](#).

AE proposed a different estimator they call AVI (for Approximate Value Iteration) that allows for nonlinear parametrizations of the R and Q_ε functions which are computed using an iterative procedure for updating the parameter $\hat{\phi}_k$ is the estimated parameters determining $\hat{R}_k(s, a) = R_{\hat{\phi}_k}(s, a)$ as follows

$$\hat{\phi}_{k+1} = \underset{\phi}{\operatorname{argmin}} E_N \left\{ \left[\bar{r}(s, a) + \beta R_{\hat{\phi}_k}(s', a') - R_\phi(s, a) \right]^2 \right\}. \quad (33)$$

These iterations continue until the sequence $\{\hat{\phi}_k\}$ converges and the resulting estimator of R is $\hat{R} = R_{\hat{\phi}_k}$ where $\hat{\phi}_k$ is the last value of the iterative procedure (33), and similarly for \hat{Q}_ε .²⁴ AE conclude that “A range of Monte Carlo simulations using a dynamic firm entry problem, a dynamic firm entry game and two versions of the famous Rust (1987) engine replacement problem show that the proposed algorithms work well in practice.”²⁵

[Kang et al. \(2025\)](#) developed an estimator they call GLADIUS (Gradient-based Learning with Ascent-Descent for Inverse Utility learning from Samples) that is 1) model-free, 2) does not require a first stage non-parametric estimation of $\pi(a|s)$, and 3) does not require a parametric specification for rewards. Instead, GLADIUS is an iterative algorithm that uses gradient ascent/descent to maximize a penalized likelihood function similar to the SEES estimator (29). However instead of parametrizing r as a function of “structural parameters” θ and using using auxilliary parameters ϕ to approximate the Q functions, GLADIUS relies on flexible parametrization of the Q function as $Q_{\theta_1}(s, a)$ and the conditional expectation of the value function $E\{V(s')|s, a\}$ as $EV_{\theta_2}(s, a)$ for parameters θ_1 and θ_2 of flexible approximations to Q and EV such as DNNs. Once GLADIUS converges, returning parameters $\hat{\theta}_1$ and $\hat{\theta}_2$ rewards are estimated as the function

$$r(s, a) = Q_{\hat{\theta}_1}(s, a) - \beta EV_{\hat{\theta}_2}(s, a), \quad (34)$$

i.e. they are “backed out” from the fixed point equation for smoothed Q function, (15).

Both GLADIUS and SEES use an objective function that maximizes a penalized likelihood where the penalty term is a “mean squared Bellman error” for the Q functions. That is, if we express Q as a fixed point $Q = \Lambda_\sigma(Q)$, Q minimizes the quantity $\rho_{BE}(Q) = E\{[Q(s, a) - \Lambda_\sigma(Q)(s, a)]^2\}$ over a joint distribution of (s, a) pairs. SEES is a model-based approach that uses an assumed parametric functional form for rewards, $r_\theta(s, a)$ and knowledge (or estimated version thereof) of

²⁴It is possible to extend their approach to reward functions $r_\theta(s, a)$ where θ does not enter as a linear combination of a set of known features $\bar{r}(s, a)$. However in that case, we can no longer express $Q_r(s, a) = R(s, a) * \theta$, and in general θ will enter Q_r in a non-linear fashion. The TD estimator can still be adapted to work in this situation except that now there must be a nested TD subroutine that recomputes \hat{Q}_r as a function of θ each time θ is updated in an outer optimization algorithm that searches over θ to maximize the partial likelihood function (26).

²⁵See [Khawaja and Srivastava \(2025\)](#) for an alternative approach to AE’s TD methods that they call RLTD-CCS that combines the TD algorithm from RL with the Conditional Choice Simulator (CCS) estimator of [Hotz et al. \(1994\)](#) that relies on forward simulations of paths to produce a Monte Carlo estimate the $Q_\theta(s, a)$ and hence is a model-based approach to estimation.

the transition probability' $p(s'|s, a)$ to compute the Bellman operator $\Lambda_{\theta, p, \sigma}$, where we subscript Λ_σ to emphasize its dependence on θ as well as p . Thus, SEES calculates the mean squared Bellman error $\rho_{BE}(Q)$ numerically using deterministic methods such as numerical quadrature to compute the conditional expectations of the Q functions. such as numerical quadrature to compute the conditional expectations of the Q functions.

GLADIUS is a model-free approach inspired by Q-learning that uses transitions (s', s, a) in the data to construct an estimate of the mean squared Bellman error $\rho_{BE}(Q)$. That is, GLADIUS seeks to approximate the quantity

$$\Lambda_\sigma(Q)(s, a) = r(s, a) + \beta \sum_{s'} V(s') p(s'|a, d) \quad (35)$$

by averaging realizations of the form

$$\tilde{\Lambda}_\sigma(Q)(s', s, a) = r(s, a) + \beta V(s'), \quad (36)$$

using observed (s, a, s') transitions in the sample. It follows that $E\{\tilde{\Lambda}_\sigma(s', s, a)|s, a\} = \Lambda_\sigma(Q)(s, a)$ so we can use transitions (s', s, a) in the data to calculate $\Lambda_\sigma(Q)(s', s, a)$ values, and then use these as dependent variables in a regression to estimate its conditional expectation $\Lambda_\sigma(s, a)$. This suggests a strategy of estimating Q by searching for θ_1 to minimize the mean squared temporal difference error $\rho_{TD}(Q)$ given by

$$\rho_{TD}(Q) = E\{[\Lambda_\sigma(Q)(s', s, a) - Q(s, a)]^2\}, \quad (37)$$

where, using the terminology of the RL literature, $\tilde{\Lambda}_\sigma(Q)(s', s, a) - Q(s, a)$ is the TD error.

However [Kang et al. \(2025\)](#) note an important complication: the true fixed point Q will minimize the mean squared Bellman error $\rho_{BE}(Q)$ but it will not necessarily minimize the mean squared TD error. To see, this, the note the following identity

$$\rho_{TD}(Q) = \rho_{BE}(Q) + \beta^2 E\{[V(s') - EV(s, a)]^2\}, \quad (38)$$

where the second term reflects expected ‘‘heteroscedasticity’’ in the TD error, i.e. it is the average of the conditional variances given by $H(V)(s, a) = E\{[V(s') - EV(s, a)]^2|s, a\}$. Since V is a function of Q from (16), we write V as $V(Q)$ and the heteroscedasticity term as $H(Q)$. It follows from (38) that the Q that minimizes $\rho_{TD}(Q)$ will not necessarily be the same as the Q that minimizes $\rho_{BE}(Q)$. To deal with this issue, they express the mean squared Bellman error as $\rho_{BE}(Q) = \rho_{TD}(Q) - H(Q)$. Note that $H(Q)$ is also an unknown quantity they note that it can be approximated as the solution to

$$H(Q)(s, a) = \min_{v \in R} \sum_{s'} [V(Q)(s') - v]^2 p(s'|s, a). \quad (39)$$

Of course the v that minimizes (39) is $EV(s, a)$. [Kang et al. \(2025\)](#) parametrize Q and EV by

DNNs as Q_{θ_1} and EV_{θ_2} , respectively, to obtain an estimation criterion that is a penalized likelihood function similar to SEES, except it is a max-min problem given by

$$(\hat{\theta}_1, \hat{\theta}_2) = \underset{\theta_1}{\operatorname{argmax}} \underset{\theta_2}{\operatorname{argmin}} \log(L(\theta_1)) - \lambda \rho(\theta_1, \theta_2), \quad (40)$$

where λ is a penalization parameter, and $\rho(\theta_1, \theta_2)$ is the estimate of the mean squared Bellman error given by

$$\rho(\theta_1, \theta_2) = \frac{1}{N} \sum_i^N [\Lambda_\sigma(Q_{\theta_1}(s'_i, s_i, a_i) - Q_{\theta_1}(s_i, a_i))^2 - \beta^2 [V(Q_{\theta_1})(s'_i) - EV_{\theta_2}(s_i, a_i)]^2] \quad (41)$$

Computation of the penalty term can be further reduced to limiting the summation to a subset of states s_i where the corresponding action a_i equals the *normalizing action*. The reason is that Kang et al. (2025) impose the identifying assumption that for each s there is an action a_s called the normalizing action for which $r(s, a_s)$ is known. From our discussion of identification in section 3.2, this restriction on rewards implies that we can determine all Q values by enforcing the fixed point constraint $Q = \Lambda_\sigma(Q)$ only for the subset of pair (s, a_s) which enables us to determine the corresponding Q at the normalizing action, $Q(s, a_s)$ for all $s \in S$. Then given $Q(s, a_s)$ the remaining $Q(s, a)$ values are determined from $\pi(a|s)$ using the Hotz-Miller inversion formula (30).

Kang et al. (2025) establish an important result on the global convergence of GLADIUS. First, under a “realizability assumption” that there exists θ_1^* such that $Q^* = Q_{\theta_1^*}^*$, where Q^* is the true Q function (as well as some other regularity conditions), they show that if $\hat{Q}_{T,N}$ is the estimate of Q^* after T iterations of the GLADIUS algorithm using a data set with N observations on agent’s states and transitions (s', s, a) , then

$$E_N \left\{ [Q^*(s, a) - \hat{Q}_{T,N}(s, a)]^2 \right\} \leq O(1/T) + O(1/N). \quad (42)$$

They note that “To the best of our knowledge, no prior work has proposed an algorithm that guarantees global optimum convergence of the minimization problem that involves the [mean squared Bellman residual term]”. They compare GLADIUS to several other DDC estimation methods using the Rust (1987) bus engine problem and a high dimensional version of this problem where they added between 2 and 100 additional irrelevant state variables (i.e. variables that have no effect on cost of maintaining the bus or its mileage transitions, and thus on the Q values and decision rule π) that take on 20 possible values each, thus massively increasing the size of the state space. They conclude that “We find that, on average, our method performs quite well in recovering rewards in both low and high-dimensional settings. Further, it has better/on-par performance compared to other benchmark algorithms in this area (including algorithms that assume the parametric form of the reward function and knowledge of state transition probabilities) and is able to recover rewards

even in settings where other algorithms are not viable.”²⁶

We have provided a very selective survey of the IRL literature, going into depth on two recent contributions to this literature that are particularly closely related to and we think particularly promising new methods for structural estimation of DDC models. We have emphasized the distinction between model-based and model-free estimation methods, but we wish to make clear that not all IRL methods are model-free. For example [Zeng et al. \(2024\)](#) use non-parametric methods to estimate p to provide a “generative world model” in a model-based version of IRL and “Through extensive experiments, we demonstrate that our algorithm outperforms existing benchmarks for offline IRL and Imitation Learning, especially on high-dimensional robotics control tasks.” Thus, it is not obvious that model-free approaches are necessarily superior to model-based approaches, though we agree that good models of p should be used, since “inaccurate models of the world obtained from finite data with limited coverage could compound inaccuracy in estimated rewards.” It is also clear that if we are interested in doing a counterfactual that involves changes p model-based methods will probably be necessary to estimate/construct this new p and then train the model to calculate the new optimal decision rule $\pi(a|s)$ corresponding to counterfactual p .

We also do not want to leave an impression that the two studies we have focused on are the first to use function approximation methods such as DNNs to estimate models with large scale state spaces in the IRL literature. As the survey by [Gleave and Toyer \(2022\)](#) notes: “algorithms based on MCE IRL have scaled to high-dimensional environments. Maximum Entropy Deep IRL [Wulfmeier et al. \(2016\)](#) was one of the first extensions, and is able to learn rewards in gridworlds from pixel observations. More recently, Guided Cost Learning [Abbeel \(2016\)](#) and Adversarial IRL [Fu et al. \(2018\)](#) have scaled to MuJoCo continuous control tasks. Given its popularity and accomplishments we focus on MCE IRL in the remainder of this document; we refer the reader to [Jeon et al. \(2020\)](#) for a broader overview of reward learning.” For example, in the introduction we noted two large scale studies using IRL: the study by [Barnes et al. \(2024\)](#) that uses a version of IRL called Receding Inverse Horizon Planning to infer preferences over route choice using data from Google Maps, and the study by [Zhao and Liang \(2023\)](#) who used AIRL and showed that it outperforms competing methods in an application to inferring context-dependent preferences for route choice using data on taxi drivers in Shanghai. However all this work indicates the promise of IRL methods for estimation of large scale DDC problems.

5 Conclusion

This selective survey has drawn connections between the literature on structural estimation of dynamic discrete choice (DDC) models and the literature on inverse reinforcement learning (IRL), and their common ancestor: Markovian Decision Processes (MDP). MDPs provide a recipe for

²⁶[Kang et al. \(2025\)](#) also note GLADIUS corrects a problem in IRL methods that use “occupancy matching” as the estimation criterion such as [Ho and Ermon \(2016\)](#), [Garg et al. \(2022\)](#), and [Yue et al. \(2023\)](#). They show that occupancy matching is not guaranteed to consistency estimate the true Q function, and “This implies that r cannot be inferred from Q using the Bellman equation after deriving Q using occupancy matching.”

how to model the behavior of an intelligent agent who makes sequential decisions to maximize the expected, discounted value of a stream of rewards. Though MDPs can be viewed as specialized in some respects and some of their assumptions (expected utility, discounted utility) are inconsistent with empirical evidence on human decision making, it is a powerful way to approximate intelligent behavior that has opened up a huge range of empirical applications.

Despite the common ancestry, there are clear differences in the practical reasons for doing DDC and IRL modeling. In economics, DDC and other types of structural models (i.e. dynamic, stochastic general equilibrium models in macroeconomics) are used to provide counterfactual simulations to assess the impact on economics outcomes and welfare. Structural models overcome a limitation of reduced form models, which treat policy variables as a covariate and rely on historical variation in policies to make counterfactual predictions. Structural models are able to produce counterfactual predictions of the impact of novel policies that have never been enacted or even considered previously. This is why structural models are increasingly used for policymaking because it is cheaper and faster to “crash test” new policy ideas via counterfactual simulations in an “offline” environment rather than trying out new policies “online” and learning by trial and error.

In the AI and ML literatures, IRL has proven useful as a way to supply a reward function used by Reinforcement Learning (RL) algorithms for offline training of intelligent algorithms in robotics and AI. There is a wide range of complex, hard to formalize tasks where the appropriate reward function is unclear *a priori*. IRL is able to infer this reward function from limited human “demonstration data.” For example, [Christiano et al. \(2017\)](#) notes that “For sophisticated reinforcement learning (RL) systems to interact usefully with real-world environments, we need to communicate complex goals to these systems. In this work, we explore goals defined in terms of (non-expert) human preferences between pairs of trajectory segments. We show that this approach can effectively solve complex RL tasks without access to the reward function, including Atari games and simulated robot locomotion, while providing feedback on less than 1% of our agent’s interactions with the environment. This reduces the cost of human oversight far enough that it can be practically applied to state-of-the-art RL systems. To demonstrate the flexibility of our approach, we show that we can successfully train complex novel behaviors with about an hour of human time. These behaviors and environments are considerably more complex than any which have been previously learned from human feedback.”

Another important difference in the IRL and DDC literatures is that the latter is more focused on statistical inference of the estimated reward function and deriving approximate distributions that reflect how estimates are affected by sampling errors. As a result, many DDC applications involve estimation of smaller, more parsimonious models with the goal of using them to gain new insights into human decision making and test behavioral hypotheses. The IRL literature is more focused on uncovering a reward function that can be used to train large scale applications using RL. IRL has been used to uncover reward functions for concrete problems such as route choice that can require estimation of thousands and sometimes even millions of parameters. The large number

of parameters and complexity of such reward functions makes it challenging to evaluate them from the perspective of statistical inference and insight into human behavior. However the “proof is in the pudding” — the highly sophisticated and intelligent behavior in robotic and AI applications that have been learned after extensive RL training using the reward functions provided by IRL (which are in turn learned from more limited observations of human behavior).

The DDC and IRL literatures have both greatly benefitted from the use of Deep Neural Networks (DNNs) to approximate value functions in problems with large state spaces. Though it is still not formally established, the use of DNNs and other function approximation methods do seem to break Bellman’s “curse of dimensionality” (at least in a practical sense) allowing us to formulate and estimate increasingly large and more realistic models.

A challenging area for further work is to relax and test the strong rationality assumptions underlying DDC and IRL models. There are two key assumptions: 1) rational expectations (i.e. the subjective beliefs about the decision-dependent law of motion for state variables coincides with the true one), and 2) the ability to perfectly optimize given one’s beliefs. If these assumptions do not hold, rewards inferred from human behavior will be distorted by our attempt to “rationalize” behavior that may not be fully rational. The notion that human behavior is more likely to be only “boundedly rational” goes back to the work of [Simon \(1957\)](#), and it is exemplified by the fact that chess algorithms trained using RL such as Alpha-Zero consistently beat the best human Grand Masters. RL was able to do this without any need for human data and IRL because 1) the reward function is known *a priori* (i.e. 1 if win, 0 if lose), and 2) the algorithm as trained via self-play.

Self-play is an example of the “model-free” approach used in IRL and RL that avoids taking a stand on the “law of motion” of the state variables. Though estimation using fewer assumptions is generally an attribute, it does rely on the implicit assumption that agents we are modeling have rational expectations. To the extent that humans are boundedly rational, we would like to understand whether suboptimal behavior is due to distorted or irrational beliefs or the inability to perfectly optimize. In simpler environments that are less strategically complex than chess, humans may be more likely to perfectly optimize but if their beliefs are not rational, the resulting behavior will still be suboptimal from the standpoint of a rational observer.

[Anderson et al. \(2025\)](#) showed that elite tennis professionals use suboptimal serve strategies, and demonstrated how their behavior can be explained by distorted subjective beliefs about their own strengths and weaknesses as servers as well as their opponents. This opens a path for DDC and IRL models to improve human performance using data and model-free approaches that enables them to correct distorted subjective beliefs. However there are many high stake strategic situations (e.g. a wartime scenario) where the relevant data may be too scarce, leaving any decision maker no choice but to rely on an implicit “mental model” of the situation. In such situations neither DDC nor IRL methods may prove useful because of their dependence on good models of beliefs. More work needs to be done on how it might possible to construct relatively accurate mental models (i.e. beliefs about how different actions affect outcomes), in unique situations where relevant data

is extremely limited.

We conclude that the limitations confronting DDC and IRL are not as much computational but rather inherent limitations on our ability to infer underlying rewards and beliefs from data on states and actions as we noted in our discussion of the identification problem in section 2.4. The solution to the identification problem requires us to impose strong prior restrictions (assumptions) about rewards and beliefs. If these assumptions are wrong, counterfactual predictions from these models may not be accurate even though the models provide a good fit to the demonstration data used to estimate them. However practical experience suggests that even if our assumptions may not be precisely correct and human decision makers may not be not fully rational, the assumptions used to identify DDC and IRL models appear to be sufficiently good approximations to produce realistic counterfactual simulations and leverage limited human demonstration data to enable us to train artificially intelligent agents to perform well in a variety of tasks in a variety of unfamiliar situations that previously were tasks only human beings could do well.

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