

# May 21 Lec—Multivariate Distributions

## Intro, jp(d)fs, Marginal Distributions

### Examples related to beta distributions

1. Let  $X$  be a continuous random variable such that

$$f(x) = \begin{cases} cx^2(1-x) & 0 < x < 1 \\ 0 & \text{otherwise} \end{cases}$$

Find  $C$  and  $\mathbb{E}(\min(X, \frac{1}{2}))$ .

We note that  $X \sim B(\alpha = 3, \beta = 2)$ . So  $C = \frac{\Gamma(3+2)}{\Gamma(3)\Gamma(2)} = \frac{4!}{2!1!} = 12$ .

$$\begin{aligned} \mathbb{E}\left(\min\left(X, \frac{1}{2}\right)\right) &= \int_0^1 12x^2(1-x) \cdot \min\left(x, \frac{1}{2}\right) dx \\ &= \int_0^{1/2} 12x^2(1-x) \cdot x dx + \int_{1/2}^1 12x^2(1-x) \cdot \frac{1}{2} dx \\ &\quad \dots \end{aligned}$$

2.  $m_X(t) = \frac{e^{3t}}{1-2t}$  for  $t < \frac{1}{2}$

Compute  $\mathbb{E}(X)$ ,  $V(X)$ , and  $P(X > 5)$ .

Let  $Y \sim G(\alpha = 1, \beta = 2)$  (i.e. exponential,  $\beta = 2$ ; or chi-square, degree of freedom  $p = 2$ ). We have  $X = Y + 3$ .

$$\mathbb{E}(X) = \mathbb{E}(Y) + 3 = 5$$

$$V(X) = V(Y) = 4$$

$$P(X > 5) = P(Y > 2) = e^{-\frac{2}{2}} = e^{-1} \quad (\text{using the survival function of an exponential distribution})$$

The quiz tonight will cover Chapter 4.

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## Multivariate distributions

We'll mostly talk about bivariate distributions.

This chapter will use Calc 3 knowledge, but the professor thinks that you

should still understand even if you haven't taken it before.

Motivation for some discussion about double integrals:

Given a uniform probability distribution with nontrivial support on  $(0,2)$ , you know that to find the constant  $\lambda$  of the pdf we would need to do

$$\int_0^2 \lambda dx = 1 \implies \lambda = \frac{1}{2}$$

Given an isosceles right triangle with leg lengths 1, suppose it has uniform density  $C$  and total mass 1. Then  $\iint_{\text{triangle}} C dA = 1 \implies C = 2$ .

### Definition: Joint probability function

Let  $X$  and  $Y$  be two *discrete* random variables. The joint probability function (jpf) of  $X$  and  $Y$  is the function

$$\begin{aligned} p(x, y) &= P(X = x, Y = y) \\ &= P(\{X = x\} \cap \{Y = y\}) \end{aligned}$$

### Properties of the jpf

1.  $0 \leq p(x, y) \leq 1$
2. The set  $\{(x, y) \mid p(x, y) > 0\}$ , called the **support** of  $(X, Y)$ , is discrete.
3.  $\sum_{x, y} p(x, y) = 1$

### Example

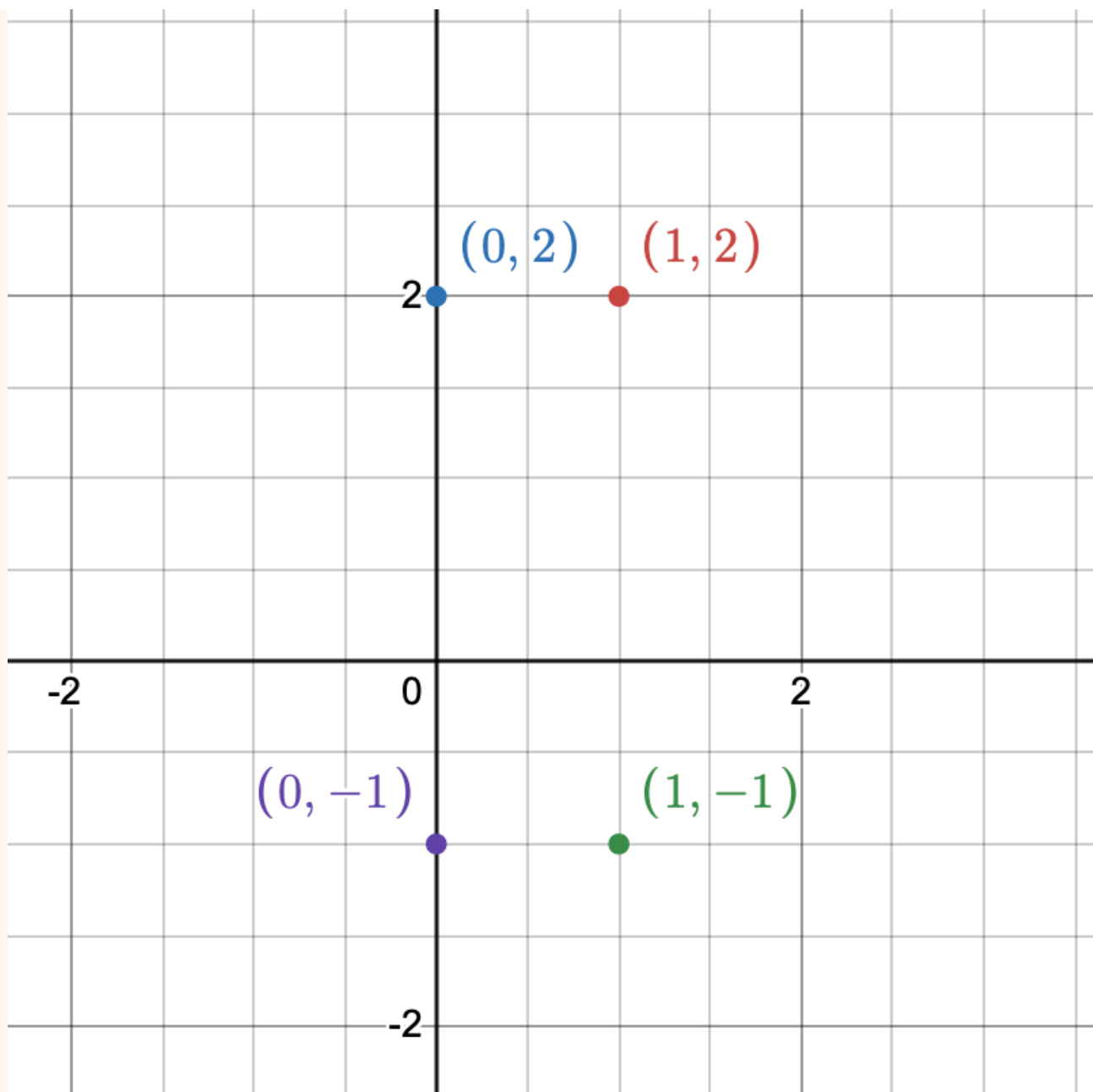
Urn A contains 3 Red and 7 Blue. Urn B contains 6 Red and 4 Blue.

An (unfair) coin is such that  $P(H) = \frac{1}{3}$ .

Flip the coin. If the result is  $H$ , set  $X = 1$ , and draw a ball from Urn A. Otherwise, set  $X = 0$  and draw a ball from Urn B.

If the ball is Red, set  $Y = 2$ . Otherwise, set  $Y = -1$ .

Picture of the support:



We calculate  $P(X = 0, Y = -1)$

$$= P(Y = -1|X = 0) \cdot P(X = 0)$$

$$= \frac{4}{10} \cdot \frac{2}{3} = \frac{8}{30}$$

Similarly,  $P(X = 0, Y = 2)$

$$= P(Y = 2|X = 0) \cdot P(X = 0)$$

$$= \frac{6}{10} \cdot \frac{2}{3} = \frac{12}{30}$$

We continue to calculate the other ones:

$$\begin{aligned}P(X = 1, Y = -1) &= P(Y = -1|X = 1) \cdot P(X = 1) \\&= \frac{7}{10} \cdot \frac{1}{3} = \frac{7}{30}\end{aligned}$$

$$\begin{aligned}P(X = 1, Y = 2) &= P(Y = 2|X = 1) \cdot P(X = 1) \\&= \frac{3}{10} \cdot \frac{1}{3} = \frac{3}{30}\end{aligned}$$

We can summarize the above in a table:

$y\text{-value} \setminus x\text{-value}$	0	1
-1	$\frac{8}{30}$	$\frac{7}{30}$
-2	$\frac{12}{30}$	$\frac{3}{30}$

The total weight over all the cells is 1, as expected.

Now let's see the analog of all this stuff in terms of continuous random variables.

### Definition: Jointly continuous random variable

Two continuous random variables  $X$  and  $Y$  are jointly described by the **joint probability density function** (jpdf) which is a function  $f(x, y)$  such that

1.  $f(x, y) \geq 0$
2.  $\iint_{\mathbb{R}^2} f(x, y) dA = 1$  (i.e. the volume under the surface formed by  $f(x, y)$  over the entire  $xy$ -plane must total to 1)

### Example

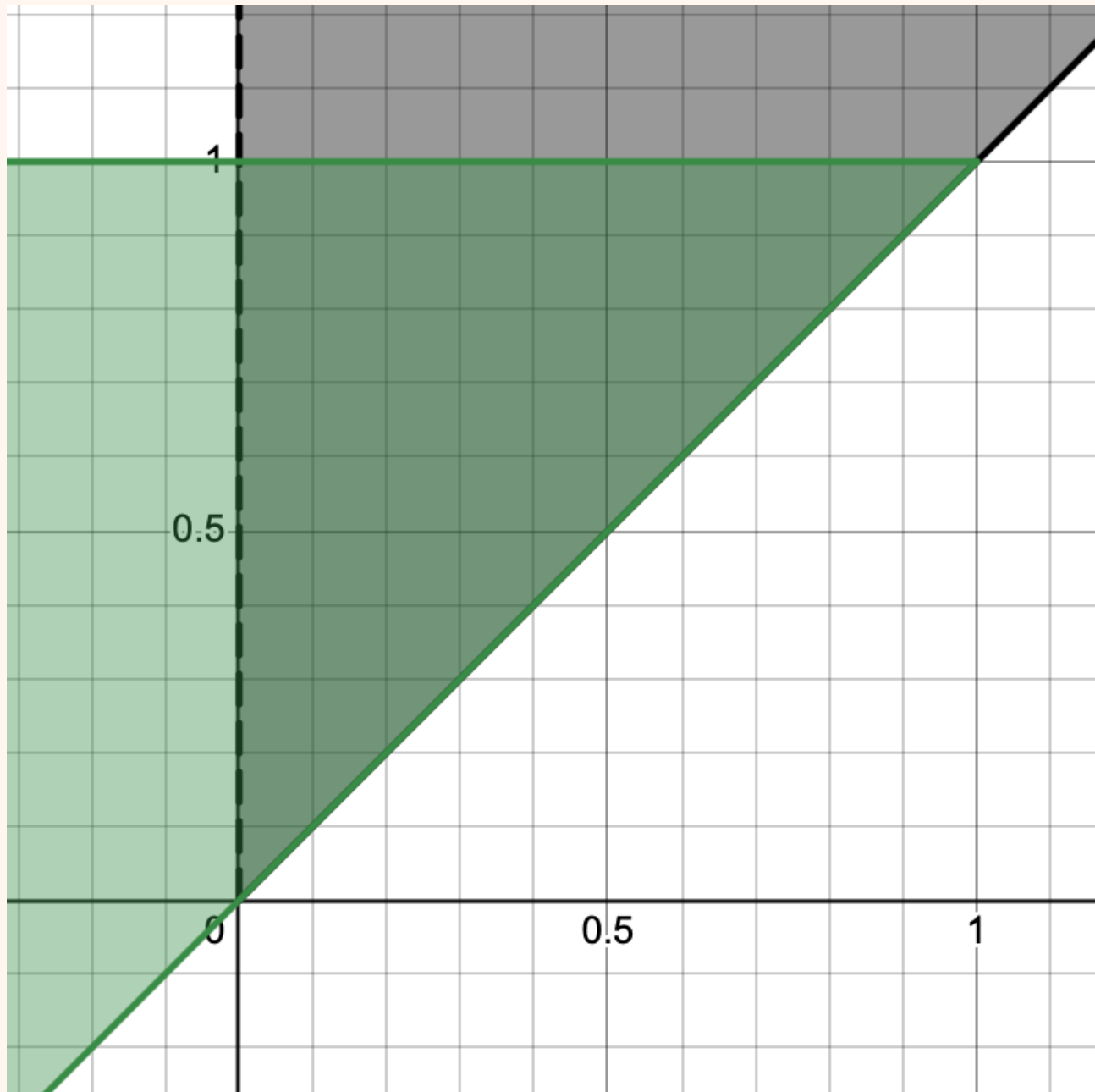
$$f(x, y) = \begin{cases} Cx & 0 < x \leq y \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

Find  $C$  such that  $f$  is a jpdf ( $C > 0$ ).

Before you do any work with double integrals, *draw the domain (support)!!!!!!!*

As always,  $x$  is the kernel.

The domain is the overlap of the black area ( $0 < x \leq y$ ) and the green area ( $x \leq y \leq 1$ ).



You can also sketch the area by thinking, for any fixed  $y$  in  $(0,1]$ , what are all the permitted values of  $x$ . (or for any fixed  $x$  in  $(0,1]$ , what are all the permitted values of  $y$ )

We now need to compute  $\iint_D Cx \, dA$ . To do this, we need to parametrize  $D$ , by making one variable in terms of the other.

The first way is to parametrize  $x$  in terms of  $y$ . That means, if we let  $y$  be the free variable (not depending on anything), what are all the possible values of  $x$  given fixed  $y$ ?

We have that  $y$  can range from 0 to 1, and for any  $y$ ,  $x$  can vary from 0 to  $y$ .

We then conclude that with this parametrization, we need to "sum" with respect to  $x$  first because it depends on  $y$ , and then we "sum" with respect to  $y$ . This means, we make  $x$  the inner integral and  $y$  the outer one.

$$\begin{aligned}\iint_D Cx \, dA &= \int_0^1 \left( \int_0^y Cx \, dx \right) dy \\ &= \int_0^1 C \cdot \frac{1}{2} y^2 \, dy = C \cdot \frac{1}{6}\end{aligned}$$

We can also parametrize  $D$  in another way, by letting  $x$  be the free variable and let  $y$  depend on  $x$ .  $x$  can range from 0 to 1, and for any fixed  $x$ ,  $y$  can range from  $x$  to 1.

$$\begin{aligned}\iint_D Cx \, dA &= \int_0^1 \left( \int_x^1 Cx \, dy \right) dx \\ &= C \int_0^1 \left( x \int_x^1 dy \right) dx \\ &= C \cdot \int_0^1 x(1-x) \, dx = C \cdot \frac{\Gamma(2)\Gamma(2)}{\Gamma(4)} = C \cdot \frac{1}{6}\end{aligned}$$

Note that the two integrals are the same (Fubini's Theorem). Regardless of the parametrization we chose, we deduce  $C = 6$ .

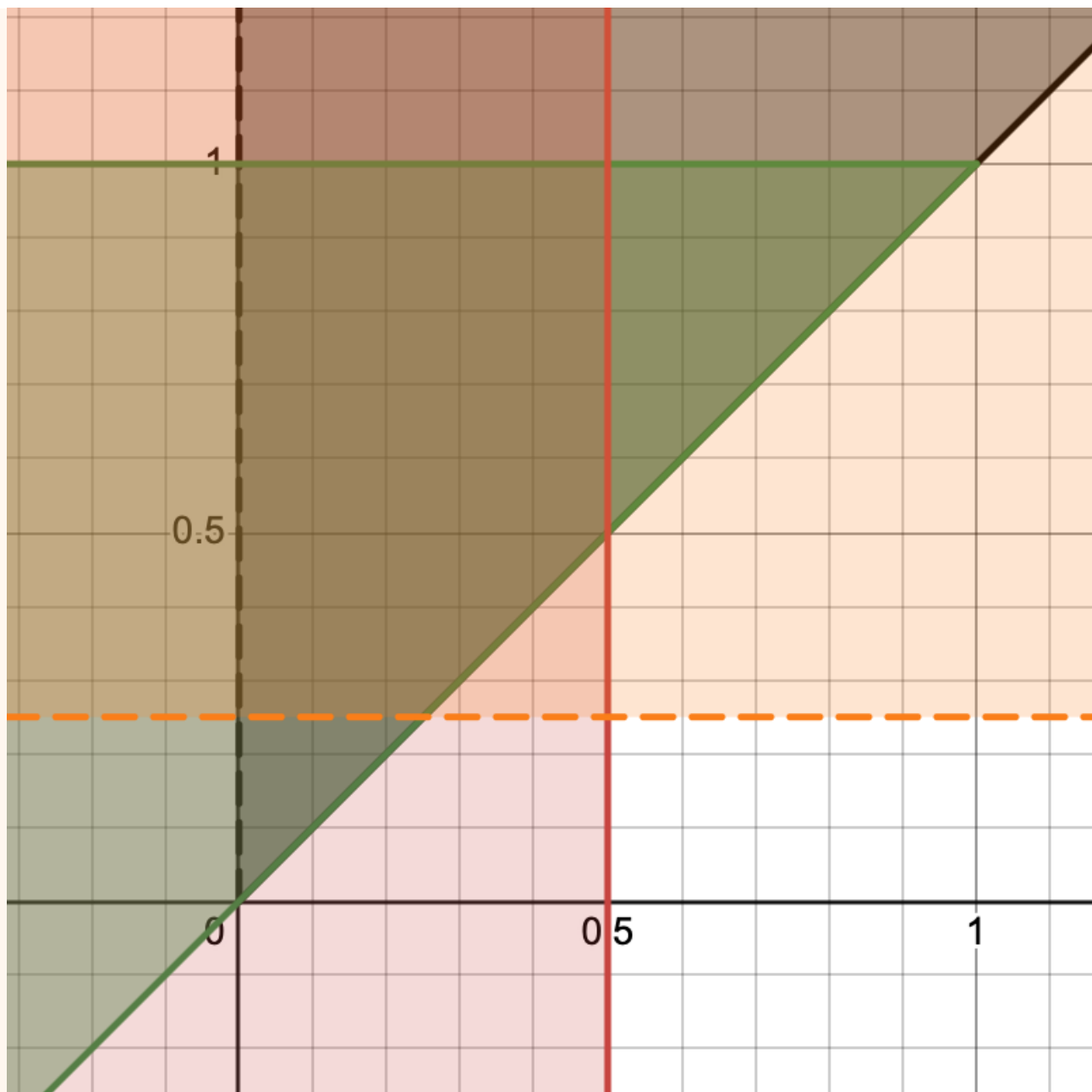
### Another continuous example

From this point forward, we will use this particular jpdf a lot.  
For

$$f(x, y) = \begin{cases} 6x & 0 < x < y < 1 \\ 0 & \text{otherwise} \end{cases}$$

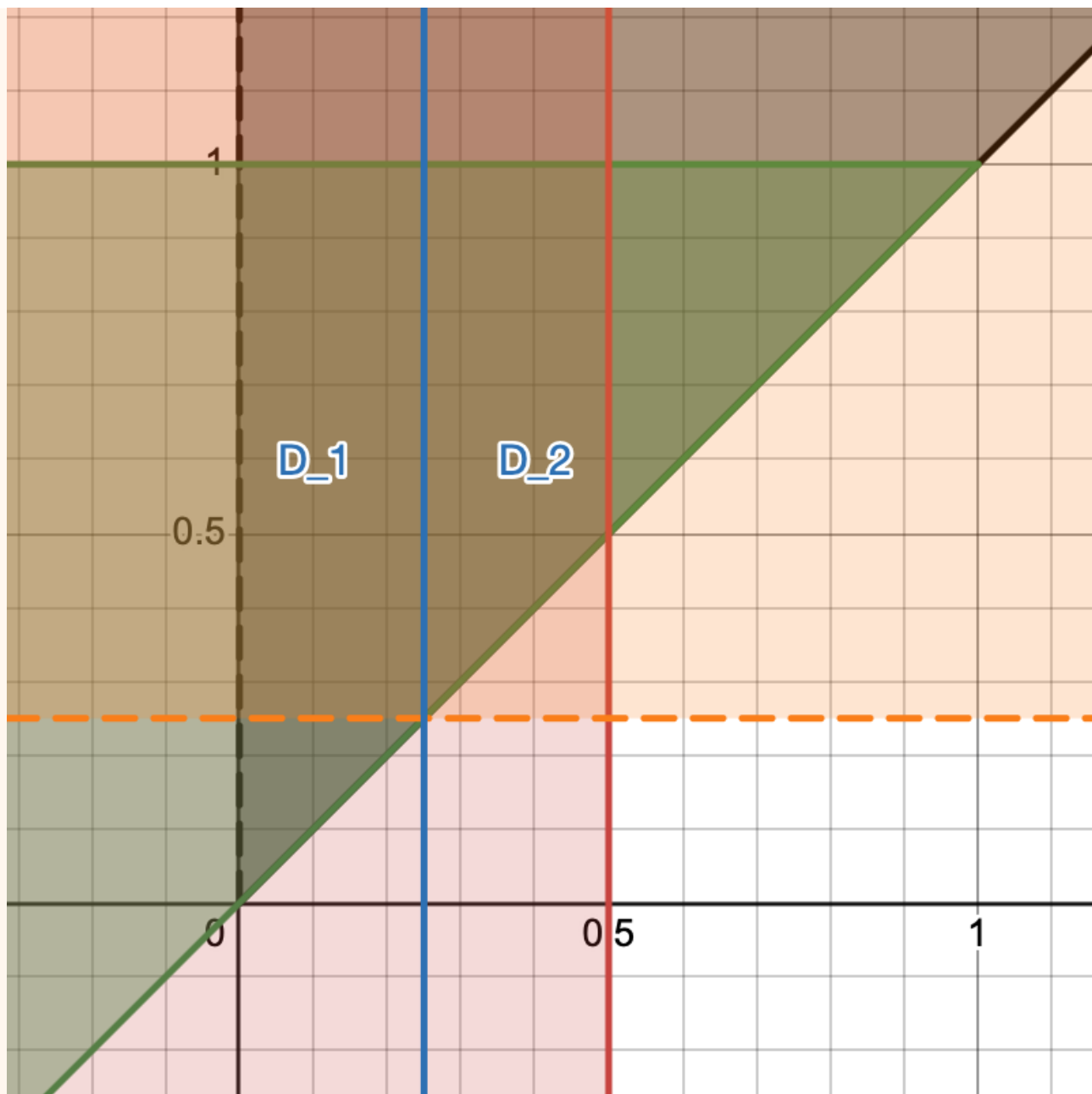
find  $P(X \leq \frac{1}{2}, Y > \frac{1}{4})$ .

This area is represented below as the area where all four colors (green, black, light orange, light red) overlap, sorry for the messy picture : (



Note that the function is 0 on the non-green, non-black areas of the graph.

We can split this area into two areas  $D_1$ ,  $D_2$  to parametrize it better:



For  $D_1$ ,  $x$  ranges from 0 to  $\frac{1}{4}$  and for any fixed  $x$ ,  $y$  ranges from  $\frac{1}{4}$  to 1.

For  $D_2$ ,  $x$  ranges from  $\frac{1}{4}$  to  $\frac{1}{2}$  and for any fixed  $x$ ,  $y$  ranges from  $x$  to 1.

So

$$\begin{aligned}
 P\left(X \leq \frac{1}{2}, y > \frac{1}{4}\right) &= \iint_{D_1} 6x \, dA + \iint_{D_2} 6x \, dA \\
 &= \int_0^{\frac{1}{4}} \int_{\frac{1}{4}}^1 6x \, dy \, dx + \int_{\frac{1}{4}}^{\frac{1}{2}} \int_x^1 6x \, dy \, dx
 \end{aligned}$$



For the remainder of today, we will talk about different  $X$  and  $Y$  distributions and their resulting joint distribution.

### Definition: Marginal distribution

1. Let  $p(x, y)$  be the jpf of two discrete random variables  $X, Y$ . The marginal probability function of  $X$  is

$$p_X(x) = P(X = x) = \sum_y p(x, y)$$

Similarly, the marginal probability function of  $Y$  is

$$p_Y(y) = P(Y = y) = \sum_x p(x, y)$$

2. Let  $f(x, y)$  be the jpdf of two continuous random variables  $X, Y$ . The marginal pdf of  $X$  is

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

and the marginal pdf of  $Y$  is

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx$$

### Example of discrete marginal pfs (see above)

$y\text{-value} \setminus x\text{-value}$	-1	0	1	$p_Y(y)$
0	$\frac{1}{4}$	$\frac{1}{6}$	$\frac{1}{12}$	$\frac{1}{2}$
1	$\frac{1}{4}$	0	$\frac{1}{4}$	$\frac{1}{2}$
$p_X(x)$	$\frac{1}{2}$	$\frac{1}{6}$	$\frac{1}{3}$	

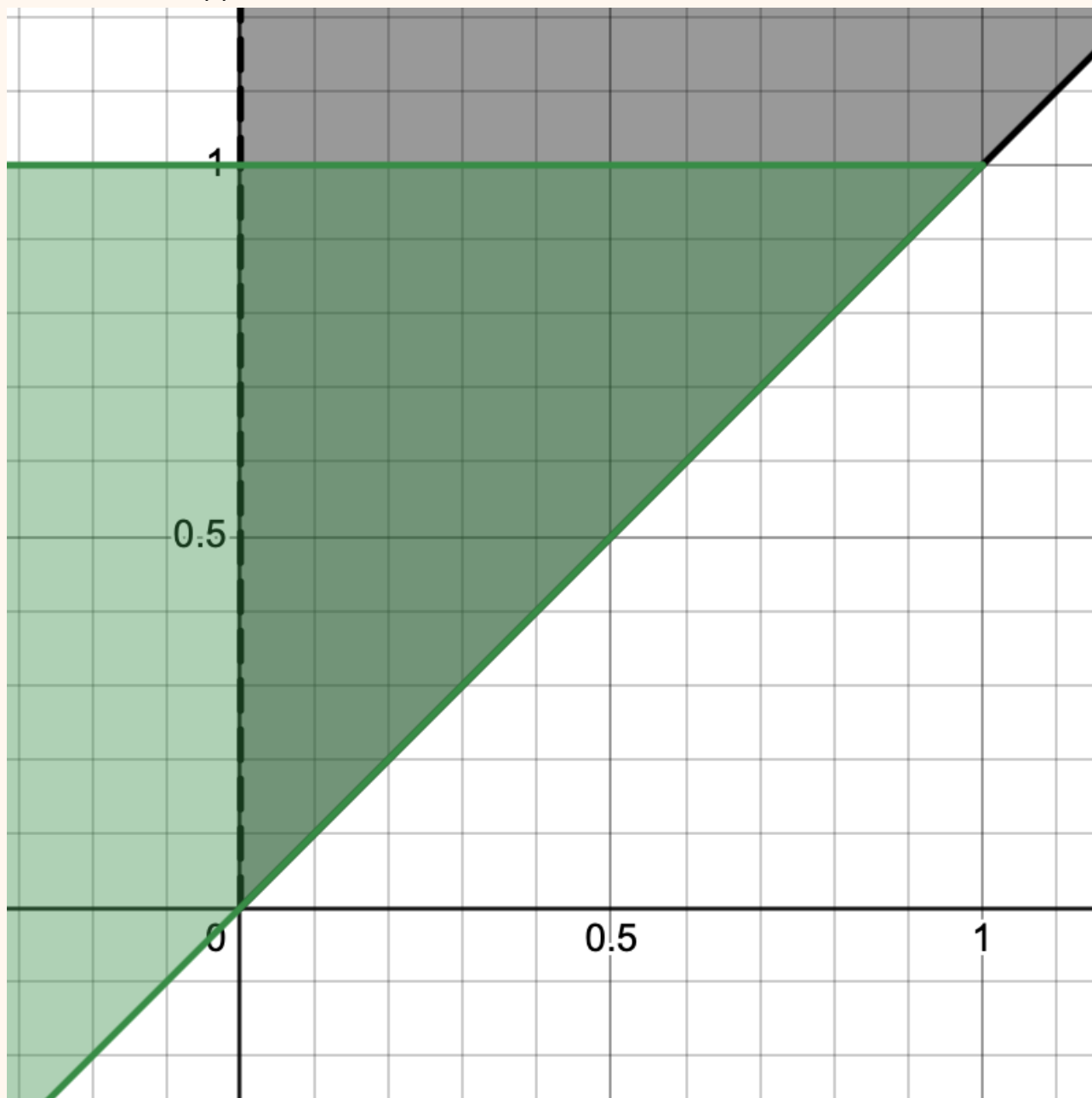
Challenge yourself with making up a randomized experiment that has this distribution. (It might help to note that  $Y$  is a Bernoulli variable,  $p = \frac{1}{2}$ .)

### A continuous example

Take again

$$f(x, y) = \begin{cases} 6x & 0 < x < y < 1 \\ 0 & \text{otherwise} \end{cases}$$

Picture of support:



We have

$$\begin{aligned} f_X(x) &= \int_{-\infty}^{\infty} f(x, y) dy \\ &= \begin{cases} \int_x^1 6x dy & x \in (0, 1) \\ 0 & \text{otherwise} \end{cases} \end{aligned}$$

$$= \begin{cases} 6x(1-x) & x \in (0, 1) \\ 0 & \text{otherwise} \end{cases}$$

so  $X \sim \text{Beta}(\alpha = 2, \beta = 2)$ .

Similarly, we have

$$\begin{aligned} f_Y(y) &= \int_{-\infty}^{\infty} f(x, y) dx \\ &= \begin{cases} \int_0^y 6x dx & y \in (0, 1) \\ 0 & \text{otherwise} \end{cases} \\ &= \begin{cases} 3y^2 & y \in (0, 1) \\ 0 & \text{otherwise} \end{cases} \end{aligned}$$

so  $Y \sim \text{Beta}(\alpha = 3, \beta = 1)$ .

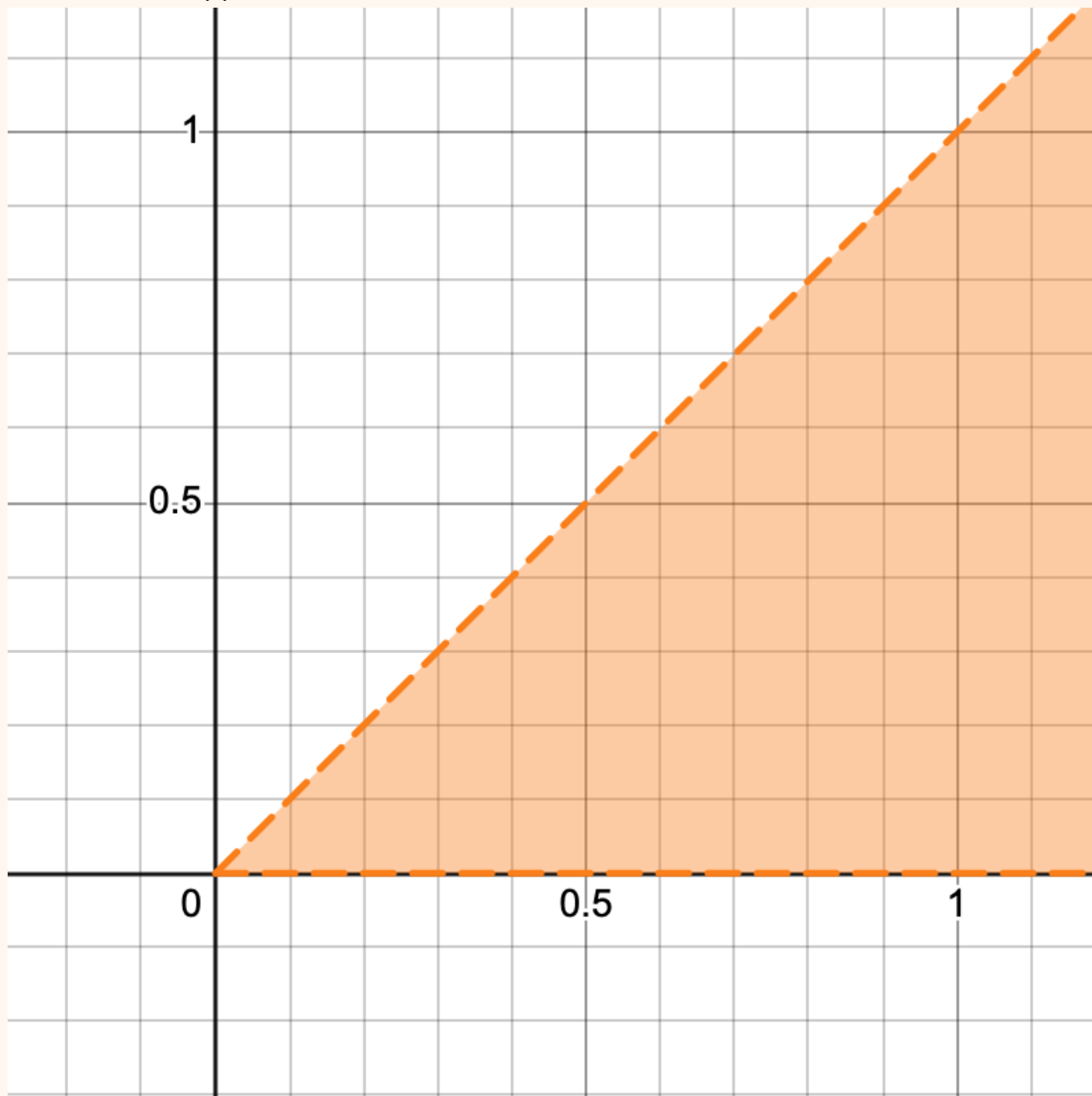
### Yet another example

Let

$$f(x, y) = \begin{cases} Ce^{-x} & 0 < y < x \\ 0 & \text{otherwise} \end{cases}$$

Find  $C$  such that  $f$  is a jpdf, and find  $f_X$ ,  $f_Y$ .

Picture of support:



Parametrizing, we have that  $x \in (0, \infty)$  and  $y \in (0, x)$ .

We need to find  $C$  such that  $C \cdot \iint_D e^{-x} dA = 1$ .

$$\begin{aligned} \iint_D e^{-x} dx &= \int_0^\infty \int_0^x e^{-x} dy dx \\ &= \int_0^\infty x e^{-x} dx = \Gamma(2) = 1 \end{aligned}$$

So  $C = 1$ , and

$$f(x, y) = \begin{cases} e^{-x} & 0 < y < x \\ 0 & \text{otherwise} \end{cases}$$

$$f_X(x) = \begin{cases} \int_0^x e^{-x} dy & x \in (0, \infty) \\ 0 & \text{otherwise} \end{cases} = \begin{cases} xe^{-x} & x \in (0, \infty) \\ 0 & \text{otherwise} \end{cases}$$

$$f_Y(y) = \begin{cases} \int_y^\infty e^{-x} dx & y \in (0, \infty) \\ 0 & \text{otherwise} \end{cases} = \begin{cases} e^{-y} & y \in (0, \infty) \\ 0 & \text{otherwise} \end{cases}$$

(so  $Y \sim \text{Exponential}(\text{mean} = 1)$ .)

### Definition: Conditional distribution

- Let  $p$  be the jpdf of a pair  $(X, Y)$  of discrete random variables. The conditional pf of  $Y$  given  $X = x$  is defined as

$$p(y|X = x) = P(Y = y|X = x) = \frac{p(x, y)}{p_X(x)}$$

for  $p_X(x) \neq 0$ .

Similarly, the conditional pf of  $X$  given  $Y = y$  is defined as

$$p(x|Y = y) = \frac{p(x, y)}{p_Y(y)}$$

- Let  $f$  be the jpdf of a pair  $(X, Y)$  of continuous random variables. The conditional pdf of  $Y$  given  $X = x$  is

$$f(y|X = x) = \frac{f(x, y)}{f_X(x)}$$

and the conditional pdf of  $X$  given  $Y = y$  is

$$f(x|Y = y) = \frac{f(x, y)}{f_Y(y)}$$

### Examples

$y\text{-value} \backslash x\text{-value}$	-1	0	1	$p_Y(y)$
0	$\frac{1}{4}$	$\frac{1}{6}$	$\frac{1}{12}$	$\frac{1}{2}$
1	$\frac{1}{4}$	0	$\frac{1}{4}$	$\frac{1}{2}$
$p_X(x)$	$\frac{1}{2}$	$\frac{1}{6}$	$\frac{1}{3}$	

Refer to the table above.

Let us find  $p(x|Y = y)$  for  $y = 0, 1$ .

$$p(-1|Y = 0) = \frac{\frac{1}{4}}{\frac{1}{2}} = \frac{1}{2}$$

$$p(0|Y = 0) = \frac{\frac{1}{6}}{\frac{1}{2}} = \frac{1}{3}$$

$$p(1|Y = 0) = \frac{\frac{1}{12}}{\frac{1}{2}} = \frac{1}{6}$$

Similarly, for  $y = 1$ :

$$p(-1|Y = 1) = \frac{\frac{1}{4}}{\frac{1}{2}} = \frac{1}{2}$$

$$p(0|Y = 1) = \frac{0}{\frac{1}{2}} = 0$$

$$p(1|Y = 1) = \frac{\frac{1}{4}}{\frac{1}{2}} = \frac{1}{2}$$