

# May 15 Lec—Exponential, Chi-Square, Midterm Review

A fumbled-up quiz question:

Choose one ball from 3 reds and 6 blues. If you get a red, flip an unfair coin with  $P(H|R) = p$ ,  $P(T|R) = 1 - p$ . If you get a blue, flip another unfair coin with  $P(T|B) = q$ ,  $P(H|B) = 1 - q$ . If  $H$  is +2 and  $T$  is -1, find the expected value.

$$P(H) = P(R)P(H|R) + P(B)P(H|B) = \frac{1}{3} \cdot p + \frac{2}{3}(1 - q)$$

$$P(T) = P(R)P(T|R) + P(B)P(T|B) = \frac{1}{3}(1 - p) + \frac{2}{3}q$$

$$\mathbb{E}(\text{score}) = 2 \left( \frac{1}{3}p + \frac{2}{3}(1 - q) \right) - \left( \frac{1}{3}(1 - p) + \frac{2}{3}q \right) = p - 2q + 1$$

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## Exponential distributions

Recall that

$$\int_0^{\infty} x^a e^{-bx} dx = \frac{\Gamma(a+1)}{b^{a+1}}$$

for  $a > -1, b > 0$ .

### Definition: Exponential distribution

Exponential distributions are gamma distributions with  $\alpha = 1$ . We write  $X \sim \text{Exponential}(\beta)$ , and the pdf of  $X$  is

$$f_X(x) = \begin{cases} \frac{1}{\beta} e^{-x/\beta} & x \in (0, \infty) \\ 0 & \text{otherwise} \end{cases}$$

The cdf of  $X$  is

$$F_X(t) = \int_{-\infty}^x f_X(t) dt = \begin{cases} 1 - e^{-x/\beta} & x \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

### Remark

Just writing

$$f(x) = \frac{1}{\beta} e^{-x/\beta}$$

without specifying the domain does *not* make  $f(x)$  a pdf.

Exponential distributions are used to model lifetimes.

### Definition: Survival function

$$\begin{aligned} S(x) &= P(X > x) = 1 - F_X(x) \\ &= \begin{cases} 1 & x \leq 0 \\ e^{-x/\beta} & x \geq 0 \end{cases} \end{aligned}$$

Because  $S(a+b) = S(a)S(b)$  for  $a, b > 0$  due to the properties of the exponential distribution, we have

$$\begin{aligned} P(X > a+b) &= P(X > a)P(X > b) \\ \frac{P(X > a+b)}{P(X > a)} &= P(X > b) \end{aligned}$$

Since  $P(X > a+b) = P([X > a+b] \cap [X > a])$ ,

$$\frac{P([X > a+b] \cap [X > a])}{P(X > a)} = P(X > b)$$

The right hand side is  $P(X > a+b|X > a)$ , so we have

$$P(X > a+b|X > a) = P(X > b)$$

which is the famous **memoryless property**.

### Remark

Exercise for those that love analysis: It can be prove that any continuous function that satisfies  $F(a+b) = F(a)F(b)$  is exponential. So this is the only distribution with this property.

## Properties of the exponential distribution

1.  $\mathbb{E}(X) = \beta$
2.  $V(X) = \beta^2$

### Proof

Obviously elementary from the fact that  $X$  also has a gamma pdf.

### Example

Let  $X \sim E(\beta)$  with  $\mathbb{E}(X) = \beta$ . Consider

$$Y = \begin{cases} 10 & 0 < x < 2 \\ 15 & 2 < x < 4 \\ 50 & x > 4 \end{cases}$$

Find  $\mathbb{E}(Y)$ .

We see that  $\beta = 4$   $p_Y(10) = F_X(2) = 1 - e^{-1/2}$ ,

$$p_Y(15) = F_X(4) - F_X(2) = e^{-1/2} - e^{-1}$$

$$p_Y(50) = 1 - F_X(4) = e^{-1}.$$

So  $\mathbb{E}(Y) =$

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## Chi-square distribution

Let  $Z \sim N(0,1)$ . We find the distribution of  $X = Z^2$ .

We first find the cdf of  $X$ .

$$F_X(x) = P(X = Z^2 \leq x) = 0, \quad x < 0$$

If  $x > 0$ ,

$$\begin{aligned} F_X(x) &= P(Z^2 \leq x) \\ &= P(-\sqrt{x} \leq Z \leq \sqrt{x}) \\ &= \Phi_Z(\sqrt{x}) - \Phi_Z(-\sqrt{x}) \end{aligned}$$

By symmetry,

$$= 2\Phi_Z(\sqrt{x}) - 1$$

So we have

$$F_X(x) = \begin{cases} 0 & x < 0 \\ 2\Phi_Z(\sqrt{x}) - 1 & x \geq 0 \end{cases}$$

The pdf of  $X$  is

$$f_X(x) = \begin{cases} 0 & x < 0 \\ \frac{1}{\sqrt{x}} \Phi'_Z(\sqrt{x}) & x > 0 \end{cases}$$

Recall that  $\Phi_Z(t) = f_Z(t) = \frac{1}{\sqrt{2\pi}} e^{-t^2/2}$

So

$$f_X(x) = \begin{cases} 0 & x < 0 \\ \frac{1}{\sqrt{2\pi}} x^{-1/2} e^{-x/2} & x > 0 \end{cases}$$

We therefore observe that  $X \sim G\left(\alpha = \frac{1}{2}, \beta = 2\right)$

(Aside: From here we learn that  $\frac{1}{\Gamma(\frac{1}{2}) \cdot (\frac{1}{2})^{1/2}} = \frac{1}{\sqrt{2\pi}}$  so  $\Gamma\left(\frac{1}{2}\right) = \sqrt{\pi}$ .)

### Definition: Chi-square distribution

$X$  is said to have a chi-square distribution with  $p$  degrees of freedom ( $X \sim \chi^2(p)$ ,  $p$  a positive integer) if  $X \sim G\left(\alpha = \frac{p}{2}, \beta = 2\right)$ , i.e. the pdf of  $X$  is

$$f_X(x) = \begin{cases} \frac{1}{\Gamma(\frac{p}{2}) 2^{p/2}} x^{p/2-1} e^{-x/2} & x > 0 \\ 0 & \text{otherwise} \end{cases}$$

### Expected value and variance

If  $X \sim \chi^2(p)$ , then

1.  $\mathbb{E}(X) = 2 \cdot \frac{p}{2} = p$
2.  $V(X) = 2^2 \cdot \frac{p}{2} = 2p$

### Example

Find  $\mathbb{E}(Z^{2n})$ .

We have  $\mathbb{E}(Z^{2n}) = \mathbb{E}((Z^2)^n) = \mathbb{E}(X^n)$  where  $X = Z^2 \sim G(\frac{1}{2}, 2)$  and

$$\begin{aligned}\mathbb{E}(X^n) &= \int_0^\infty x^n f_X(x) dx \\&= \int_0^\infty x^n \cdot \frac{1}{\sqrt{2\pi}} x^{-1/2} e^{-x/2} dx \\&= \frac{1}{\sqrt{2\pi}} \int_0^\infty x^{n-1/2} e^{-x/2} dx \\&= \frac{1}{\sqrt{2\pi}} \cdot \frac{\Gamma(n + \frac{1}{2})}{(\frac{1}{2})^{n+1/2}} \\&= \frac{2^n}{\sqrt{\pi}} \Gamma\left(n + \frac{1}{2}\right)\end{aligned}$$

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## Midterm Review

1. Consider the jumbled up REDDDIITT. What is the probability that in any shuffling of the letters the two T's are together?

There are  $\frac{9!}{3!2!2!}$  unique placements up to the same letters.

There are 8 possible positions of the two T's. For the other 7 letters, there are  $\frac{7!}{3!2!}$  ways of numbering them up to the same letters. So the probability of this event happening is

$$\frac{8 \cdot \frac{7!}{3!2!}}{\frac{9!}{3!2!2!}}$$

What if we specify that the two T's are together and only two of the three D's are together? (this would be very complicated)

Side note: The most probable/most likely outcome is the outcome with the highest probability, not the expected value.

2. Consider  $X$  with mgf  $\frac{e^{-2t}}{8}(1 + e^{2t})^3$ . What is the distribution of  $X$ ?

$$\frac{e^{-2t}}{8}(1 + e^{2t})^3 = e^{-2t} \left( \frac{1}{2} + \frac{1}{2}e^{2t} \right)^3$$

$$X = -2 + 2Y \text{ where } Y \sim B\left(3, \frac{1}{2}\right)$$

3. Let  $X \sim B\left(4, \frac{1}{5}\right)$ ,  $X$  = number of heads in 4 tosses. If head, triple your current fortune (10). If tail, half your current fortune. Let  $Y$  be your fortune after the 4 tosses. Find the expected fortune.

$$Y = 10 \cdot 3^X \cdot \left(\frac{1}{2}\right)^{4-X} = \frac{10}{16} \cdot 6^X$$

$$\mathbb{E}(Y) = \frac{10}{16} \mathbb{E}(6^X) = \frac{10}{16} \mathbb{E}(e^{\ln(6)X}) = \frac{10}{16} m_X(\ln(6))$$

$$= \frac{10}{16} \left( \frac{4}{5} + \frac{1}{5}e^{\ln(6)} \right)^4 = 10$$

4. Consider the probability function  $p_Y(n) = \left(\frac{1}{2}\right)^{n+1}$ . It isn't a geometric distribution. But consider  $X = Y + 1$ . We see that  $X \in \{1, 2, 3, \dots\}$  and  $P(X = k) = P(Y = k - 1)$ .