

# The Decision Theory Framework

## Statistical Theory

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# Statistics as a Random Game

# Statistics as a Random Game?

**Nature** and a **statistician** decide to play a game. What's in the box?

- A *family of distributions*  $\mathcal{F}$ , usually assumed to admit densities (frequencies). This is the variant of the game we decide to play.
- A *parameter space*  $\Theta \subseteq \mathbb{R}^p$  which parametrizes the family  $\mathcal{F} = \{F_\theta\}_{\theta \in \Theta}$ . This represents the space of possible plays/moves available to Nature.
- A *data space*  $\mathcal{X}$ , on which the parametric family is supported. This represents the space of possible outcomes following a play by Nature.
- An *action space*  $\mathcal{A}$ , which represents the space of possible *actions* or *decisions* or plays/moves available to the statistician.
- A *loss function*  $\mathcal{L} : \Theta \times \mathcal{A} \rightarrow \mathbb{R}^+$ . This represents how much the statistician has to pay nature when losing.
- A *set  $\mathcal{D}$  of decision rules*. Any  $\delta \in \mathcal{D}$  is a (measurable) function  $\delta : \mathcal{X} \rightarrow \mathcal{A}$ . These represent the possible strategies available to the statistician.

# Statistics as a Random Game?

How the game is played:

- First we agree on the rules:
  - ❶ Fix a parametric family  $\{F_\theta\}_{\theta \in \Theta}$
  - ❷ Fix an action space  $\mathcal{A}$
  - ❸ Fix a loss function  $\mathcal{L}$
- Then we play:
  - ❶ Nature selects (plays)  $\theta_0 \in \Theta$ .
  - ❷ The statistician observes  $\mathbf{X} \sim F_{\theta_0}$
  - ❸ The statistician plays  $\alpha \in \mathcal{A}$  in response.
  - ❹ The statistician has to pay nature  $\mathcal{L}(\theta_0, \alpha)$ .

Framework proposed by A. Wald in 1939. Encompasses three basic statistical problems:

- Point estimation
- Hypothesis testing
- Interval estimation

# Point Estimation as a Game

In the problem of point estimation we have:

- 1 Fixed parametric family  $\{F_\theta\}_{\theta \in \Theta}$
- 2 Fixed an action space  $\mathcal{A} = \Theta$
- 3 Fixed loss function  $\mathcal{L}(\theta, \alpha)$  (e.g.  $\|\theta - \alpha\|^2$ )

The game now evolves simply as:

- 1 Nature picks  $\theta_0 \in \Theta$
- 2 The statistician observes  $\mathbf{X} \sim F_{\theta_0}$
- 3 The statistician plays  $\delta(\mathbf{X}) \in \mathcal{A} = \Theta$
- 4 The statistician loses  $\mathcal{L}(\theta_0, \delta(\mathbf{X}))$

Notice that in this setup  $\delta$  is an *estimator* (it is a statistic  $\mathcal{X} \rightarrow \Theta$ ).

The statistician always loses.

→ Is there a good strategy  $\delta \in \mathcal{D}$  for the statistician to restrict his losses?

→ Is there an optimal strategy?

# Risk (Expected Loss)

# Risk of a Decision Rule

Statistician would like to pick strategy  $\delta$  so as to minimize his losses. But losses are random, as they depend on  $\mathbf{X}$ .

## Definition (Risk)

Given a parameter  $\theta \in \Theta$ , the *risk* of a decision rule  $\delta : \mathcal{X} \rightarrow \mathcal{A}$  is the expected loss incurred when employing  $\delta$ :  $R(\theta, \delta) = \mathbb{E}_{\theta} [\mathcal{L}(\theta, \delta(\mathbf{X}))]$ .

## Key notion of decision theory

*decision rules should be compared by comparing their risk functions*

## Example (Mean Squared Error)

In point estimation, the mean squared error

$$MSE(\delta(\mathbf{X})) = \mathbb{E}_{\theta} [\|\theta - \delta(\mathbf{X})\|^2]$$

is the risk corresponding to a squared error loss function.



# Coin Tossing Revisited

Consider the “coin tossing game” with quadratic loss:

- Nature picks  $\theta \in [0, 1]$
- We observe  $n$  variables  $X_i \stackrel{iid}{\sim} \text{Bernoulli}(\theta)$ .
- Action space is  $\mathcal{A} = [0, 1]$
- Loss function is  $\mathcal{L}(\theta, \alpha) = (\theta - \alpha)^2$ .

Consider 3 different decision procedures  $\{\delta_j\}_{j=1}^3$ :

- 1  $\delta_1(\mathbf{X}) = \frac{1}{n} \sum_{i=1}^n X_i$
- 2  $\delta_2(\mathbf{X}) = X_1$
- 3  $\delta_3(\mathbf{X}) = \frac{1}{2}$

Let us compare these using their associated risks as benchmarks.

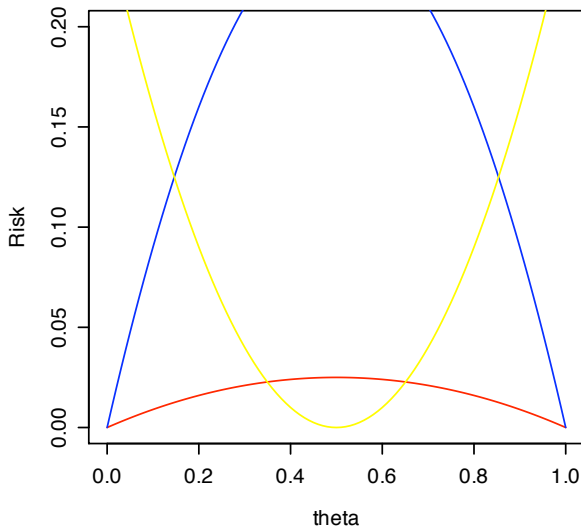
# Coin Tossing Revisited

Risks associated with different decision rules:

$$R_j(\theta) = R(\theta, \delta_j(\mathbf{X})) = \mathbb{E}_\theta[(\theta - \delta_j(\mathbf{X}))^2]$$

- $R_1(\theta) = \frac{1}{n}\theta(1 - \theta)$
- $R_2(\theta) = \theta(1 - \theta)$
- $R_3(\theta) = \left(\theta - \frac{1}{2}\right)^2$

# Coin Tossing Revisited – Every dog has its day



$$R_1(\theta), R_2(\theta), R_3(\theta)$$

# Admissibility and Inadmissibility

# Risk of a Decision Rule

## Definition (Inadmissible Decision Rule)

Let  $\delta$  be a decision rule for the experiment  $(\{F_\theta\}_{\theta \in \Theta}, \mathcal{L})$ . If there exists a decision rule  $\delta^*$  that strictly dominates  $\delta$ , i.e.

$$R(\theta, \delta^*) \leq R(\theta, \delta), \quad \forall \theta \in \Theta \quad \& \quad \exists \theta' \in \Theta : R(\theta', \delta^*) < R(\theta', \delta),$$

then  $\delta$  is called an *inadmissible decision rule*.

$R_2(\theta) > R_1(\theta)$  so  $R_2(\theta)$  is inadmissible.

- An inadmissible decision rule is a “silly” strategy since we can find a strategy that always does at least as well and sometimes better.
- However “silly” is with respect to  $\mathcal{L}$  and  $\Theta$ . (it may be that our choice of  $\mathcal{L}$  is “silly” !!!)
- If we change the rules of the game (i.e. different loss or different parameter space) then domination may break down.

# Risk of a Decision Rule

## Example (Exponential Distribution)

Let  $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Exponential}(\lambda)$ ,  $n \geq 2$ . The MLE of  $\lambda$  is

$$\hat{\lambda} = \frac{1}{\bar{X}}$$

with  $\bar{X}$  the empirical mean. Observe that

$$\mathbb{E}_\lambda[\hat{\lambda}] = \frac{n\lambda}{n-1}.$$

It follows that  $\tilde{\lambda} = (n-1)\hat{\lambda}/n$  is an unbiased estimator of  $\lambda$ . Observe now that

$$MSE_\lambda(\tilde{\lambda}) < MSE_\lambda(\hat{\lambda})$$

since  $\tilde{\lambda}$  is unbiased and  $\text{Var}_\lambda(\tilde{\lambda}) < \text{Var}_\lambda(\hat{\lambda})$ . Hence the MLE is an inadmissible rule for quadratic loss.

# Risk of a Decision Rule

Notice that the parameter space in this example is  $(0, \infty)$ . In such cases, quadratic loss tends to penalize over-estimation more heavily than under-estimation (the maximum possible under-estimation is bounded!).

Different loss function might change the result!

# Risk of a Decision Rule

## Example

If we consider another loss:

$$\mathcal{L}(a, b) = a/b - 1 - \log(a/b)$$

where, for each fixed  $a$ ,  $\lim_{b \rightarrow 0} \mathcal{L}(a, b) = \lim_{b \rightarrow \infty} \mathcal{L}(a, b) = \infty$ . Now, for  $n > 1$ ,

$$\begin{aligned} R(\lambda, \tilde{\lambda}) &= \mathbb{E}_{\lambda} \left[ \frac{n\lambda\bar{X}}{n-1} - 1 - \log \left( \frac{n\lambda\bar{X}}{n-1} \right) \right] \\ &= \underbrace{\mathbb{E}_{\lambda} [\lambda\bar{X} - 1 - \log(\lambda\bar{X})]}_{R(\lambda, \hat{\lambda})} + \underbrace{\frac{\mathbb{E}_{\lambda}(\lambda\bar{X})}{n-1} - \log \left( \frac{n}{n-1} \right)}_{g(n)} \end{aligned}$$

where we wrote  $\bar{X} = \frac{n-1}{n}\bar{X} + \frac{1}{n}\bar{X}$ .



## Example (Exponential Distribution)

Note that  $\mathbb{E}_\lambda[\bar{X}] = \lambda^{-1}$ , so

$$g(n) = \frac{1}{n-1} - \log\left(\frac{n}{n-1}\right).$$

We claim that  $g(n) > 0$  for  $n \geq 2$ . Using  $\log x = \int_1^x t^{-1} dt$ , this follows if

$$\begin{aligned} \frac{1}{x} &> \log(x+1) - \log x, \quad x > 1 \\ \iff \frac{1}{x} &> \int_x^{x+1} t^{-1} dt, \quad x > 1 \end{aligned}$$

which holds by a rectangle area bound on the integral, as follows:

$$\frac{1}{x} = [(x+1) - x] \frac{1}{x} = \int_x^{x+1} \frac{1}{x} dt > \int_x^{x+1} \frac{1}{t} dt, \quad \text{when } x > 1$$

Consequently,  $R(\lambda, \tilde{\lambda}) > R(\lambda, \hat{\lambda})$  and  $\hat{\lambda}$  dominates  $\tilde{\lambda}$ .

# Criteria for Choosing Decision Rules

## Definition (Admissible Decision Rule)

A decision rule  $\delta$  is *admissible* for the experiment  $(\{F_\theta\}_{\theta \in \Theta}, \mathcal{L})$  if it is not strictly dominated by any other decision rule.

- In non-trivial problems, it may not be easy at all to decide whether a given decision rule is admissible.
- Stein's paradox ("one of the most striking post-war results in mathematical statistics" -Brad Efron)

Admissibility is a minimal requirement - what about the opposite end (optimality) ?

- In almost any non-trivial experiment, there will be no decision rule that makes risk uniformly smallest over  $\theta$
- Narrow down class of possible decision rules by unbiasedness/symmetry/... considerations, and try to find *uniformly dominating* rules of all other rules (next week!).

# Minimax Rules

# Minimax Decision Rules

- Another approach to good procedures is to use global rather than local criteria (with respect to  $\theta$ ).

Rather than look at **risk at every  $\theta$**   $\leftrightarrow$  Concentrate on **maximum risk**

## Definition (Minimax Decision Rule)

Let  $\mathcal{D}$  be a class of decision rules for an experiment  $(\{F_\theta\}_{\theta \in \Theta}, \mathcal{L})$ . If  $\delta \in \mathcal{D}$  is such that

$$\sup_{\theta \in \Theta} R(\theta, \delta) \leq \sup_{\theta \in \Theta} R(\theta, \delta'), \quad \forall \delta' \in \mathcal{D},$$

then  $\delta$  is called a minimax decision rule.

- A minimax rule  $\delta$  satisfies  $\sup_{\theta \in \Theta} R(\theta, \delta) = \inf_{\kappa \in \mathcal{D}} \sup_{\theta \in \Theta} R(\theta, \kappa)$ .
- In the minimax setup, a rule is *preferable* to another if it has smaller maximum risk.

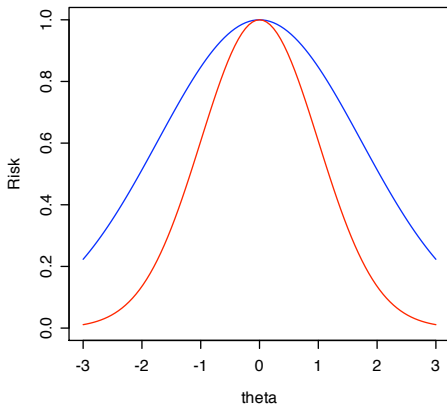
# Minimax Decision Rules

A few comments on minimaxity:

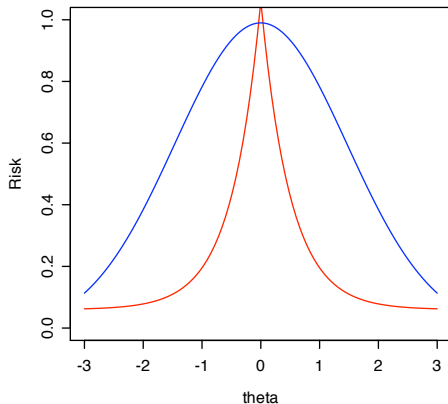
- Motivated as follows: we do not know anything about  $\theta$  so let us insure ourselves against the worst thing that can happen.
- Makes sense if you are in a zero-sum game: if your opponent chooses  $\theta$  to maximize  $\mathcal{L}$  then one should look for minimax rules. But is nature really an opponent?
- If there is no reason to believe that nature is trying to “do her worst”, then the minimax principle is overly conservative: it places emphasis on the “bad  $\theta$ ”.
- Minimax rules may not be unique, and may not even be admissible. A minimax rule may very well dominate another minimax rule.
- A unique minimax rule is (obviously) admissible.
- Minimavity can lead to counterintuitive results. A rule may dominate another rule, except for a small region in  $\Theta$ , where the other rule achieves a smaller supremum risk.

# Minimax Decision Rules

Inadmissible minimax rule



Counterintuitive minimax rule



# Bayes Rules

# Bayes Decision Rules

- Suppose we have some prior belief about the value of  $\theta$ . How can this be factored in our risk-based considerations?

Rather than look at **risk at every  $\theta$**   $\leftrightarrow$  Concentrate on **average risk**

## Definition (Bayes Risk)

Let  $\pi(\theta)$  be a probability density (frequency) on  $\Theta$  and let  $\delta$  be a decision rule for the experiment  $(\{F_\theta\}_{\theta \in \Theta}, \mathcal{L})$ . The  $\pi$ -Bayes risk of  $\delta$  is defined as

$$r(\pi, \delta) = \int_{\Theta} R(\theta, \delta) \pi(\theta) d\theta = \int_{\Theta} \int_{\mathcal{X}} \mathcal{L}(\theta, \delta(\mathbf{x})) F_\theta[d\mathbf{x}] \pi(\theta) d\theta$$

The prior  $\pi(\theta)$  places different emphasis for different values of  $\theta$  based on our prior belief/knowledge.



# Bayes Decision Rules

- Bayes principle: a decision rule is *preferable* to another if it has smaller Bayes risk (depends on the prior  $\pi(\theta)$ !).

## Definition (Bayes Decision Rule)

Let  $\mathcal{D}$  be a class of decision rules for an experiment  $(\{F_\theta\}_{\theta \in \Theta}, \mathcal{L})$  and let  $\pi(\cdot)$  be a probability density (frequency) on  $\Theta$ . If  $\delta \in \mathcal{D}$  is such that

$$r(\pi, \delta) \leq r(\pi, \delta') \quad \forall \delta' \in \mathcal{D},$$

then  $\delta$  is called a *Bayes decision rule* with respect to  $\pi$ .

- The minimax principle aims to minimize the **maximum risk**.
- The Bayes principle aims to minimize the **average risk**
- Sometime no Bayes rule exists because the infimum may not be attained for any  $\delta \in \mathcal{D}$ . However in such cases  $\forall \epsilon > 0 \exists \delta_\epsilon \in \mathcal{D}$ :  
 $r(\pi, \delta_\epsilon) < \inf_{\delta \in \mathcal{D}} r(\pi, \delta) + \epsilon$ .

# Admissibility of Bayes Rules

Rule of thumb: Bayes rules are nearly always admissible.

## Theorem (Discrete Case Admissibility)

*Assume that  $\Theta = \{\theta_1, \dots, \theta_t\}$  is a finite space and that the prior  $\pi(\theta_i) > 0$ ,  $i = 1, \dots, t$ . Then a Bayes rule with respect to  $\pi$  is admissible.*

## Proof.

Let  $\delta$  be a Bayes rule, and suppose that  $\kappa$  strictly dominates  $\delta$ . Then

$$\begin{aligned} R(\theta_j, \kappa) &\leq R(\theta_j, \delta), \quad \forall j \\ R(\theta_j, \kappa)\pi(\theta_j) &\leq R(\theta_j, \delta)\pi(\theta_j), \quad \forall \theta \in \Theta \\ \sum_j R(\theta_j, \kappa)\pi(\theta_j) &< \sum_j R(\theta, \delta)\pi(\theta_j) \end{aligned}$$

which is a contradiction (strict inequality follows by strict domination and the fact that  $\pi(\theta_j)$  is always positive). □

# Admissibility of Bayes Rules

## Theorem (Uniqueness and Admissibility)

*If a Bayes rule is unique, it is admissible.*

## Proof.

Suppose that  $\delta$  is a unique Bayes rule and assume that  $\kappa$  strictly dominates it. Then,

$$\int_{\Theta} R(\theta, \kappa) \pi(\theta) d\theta \leq \int_{\Theta} R(\theta, \delta) \pi(\theta) d\theta.$$

as a result of strict domination and by  $\pi(\theta)$  being non-negative. This implies that  $\kappa$  either improves upon  $\delta$ , or  $\kappa$  is a Bayes rule. Either possibility contradicts our assumption. □

# Admissibility of Bayes Rules

## Theorem (Continuous Case Admissibility)

*Let  $\Theta \subset \mathbb{R}^d$ . Assume that the risk functions  $R(\theta, \delta)$  are continuous in  $\theta$  for all decision rules  $\delta \in \mathcal{D}$ . Suppose that  $\pi$  places positive mass on any open subset of  $\Theta$ . Then a Bayes rule with respect to  $\pi$  is admissible.*

## Proof.

Let  $\kappa$  be a decision rule that strictly dominates  $\delta$ . Let  $\Theta_0$  be the set on which  $R(\theta, \kappa) < R(\theta, \delta)$ . Given a  $\theta_0 \in \Theta_0$ , we have  $R(\theta_0, \kappa) < R(\theta_0, \delta)$ . By continuity, there must exist an  $\epsilon > 0$  such that  $R(\theta, \kappa) < R(\theta, \delta)$  for all  $\theta$  satisfying  $\|\theta - \theta_0\| < \epsilon$ . It follows that  $\Theta_0$  is open and hence, by our assumption,  $\pi[\Theta_0] > 0$ . Therefore, it must be that

$$\int_{\Theta_0} R(\theta, \kappa) \pi(\theta) d\theta < \int_{\Theta_0} R(\theta, \delta) \pi(\theta) d\theta$$

# Admissibility of Bayes Rules

Observe now that

$$\begin{aligned}r(\pi, \kappa) &= \int_{\Theta} R(\theta, \kappa) \pi(\theta) d\theta \\&= \int_{\Theta_0} R(\theta, \kappa) \pi(\theta) d\theta + \int_{\Theta_0^c} R(\theta, \kappa) \pi(\theta) d\theta \\&< \int_{\Theta_0} R(\theta, \delta) \pi(\theta) d\theta + \int_{\Theta_0^c} R(\theta, \delta) \pi(\theta) d\theta \\&= r(\pi, \delta),\end{aligned}$$

since  $\int_{\Theta_0^c} R(\theta, \kappa) \pi(\theta) d\theta \leq \int_{\Theta_0^c} R(\theta, \delta) \pi(\theta) d\theta$ , while we have strict inequality on  $\Theta_0$ , contradicting our assumption that  $\delta$  is a Bayes rule.  $\square$

- The continuity assumption and the assumption on  $\pi$  ensure that  $\Theta_0$  is not an isolated set, and has positive measure, so that it “contributes” to the integral.

# Randomised Rules

# Randomised Decision Rules

Given

- decision rules  $\delta_1, \dots, \delta_k$
- probabilities  $\pi_i \geq 0$ ,  $\sum_{i=1}^k p_i = 1$

we may define a new decision rule

$$\delta_* = \sum_{i=1}^k p_i \delta_i$$

called a *randomised decision rule*. Interpretation:

Given data  $\mathbf{X}$ , choose a rule  $\delta_i$  with probability  $p_i$  independently of  $\mathbf{X}$ . If  $\delta_j$  is the outcome ( $1 \leq j \leq k$ ), then take action  $\delta_j(\mathbf{X})$ .

→ Risk of  $\delta_*$  is average risk:  $R(\theta, \delta_*) = \sum_{i=1}^k p_i R(\theta, \delta_i)$

- Appears artificial but often minimax rules are randomised
- Examples of randomised rules with  $\sup_{\theta} R(\theta, \delta_*) < \sup_{\theta} R(\theta, \delta_i) \forall i$

# Summary

Decision theory gives us a tool to compare different estimators / statistical procedures inside parametric models:

In order to use Decision Theory, we have to choose an appropriate loss function.

Comparing risk function is hard because there is no canonical ordering on positive functions ! We saw three possibilities:

- Admissibility : corresponding to a partial order.
- Minimax : ordering risk functions according to their maximum.
- Bayes' rules : corresponding to a weighting of the different  $\theta$ .

Amazingly, Bayes' rules and admissible rules have a very close relationship.

We presented randomized decisions which might appear silly but are useful for minimaxity.