

Let  $(X, Y)$  denote a bivariate vector with finite fourth moments. Suppose that  $Y$  may be missing, so we use  $R$  (taking values 1 and 0) to indicate whether  $Y$  is observed. We further assume that  $R$  and  $Y$  are independent given  $X$  and that  $\pi(x) = P(R = 1|X = x)$  is a known function strictly bounded away from zero. The goal is to estimate  $\mu = E[Y]$ .

The observed data from  $n$  independent subjects consist of  $(R_i Y_i, R_i, X_i)$ ,  $i = 1, \dots, n$ .

- (a) (5 points) Show that for any measurable function  $g(X)$  such that  $E[g(X)^2] < \infty$ , the following equality holds:

$$\mu = E \left[ g(X) + \frac{R}{\pi(X)}(Y - g(X)) \right].$$

- (b) (5 points) From (a), we can construct the following estimator for  $\mu$  (called difference estimator):

$$\tilde{\mu} = n^{-1} \sum_{i=1}^n \left[ g(X_i) + \frac{R_i}{\pi(X_i)}(Y_i - g(X_i)) \right].$$

Derive the asymptotic distribution for  $\tilde{\mu}$ . Show that its asymptotic variance is equal to

$$\text{Var}(Y) + E[w(X)(Y - g(X))^2],$$

where  $w(x) = \pi(x)/(1 - \pi(x))$ .

- (c) (5 points) We choose  $g(x) = \gamma x$ . Show that the optimal  $\gamma$  to minimize the variance in (b) is  $\gamma^* = \frac{E[YXw(X)]}{E[X^2w(X)]}$ . Thus, an estimator for  $\gamma^*$  using the observed data is defined as

$$\hat{\gamma} = \frac{\sum_{i=1}^n R_i Y_i X_i w(X_i) / \pi(X_i)}{\sum_{i=1}^n X_i^2 w(X_i)}.$$

Show  $\hat{\gamma} \rightarrow_{a.s.} \gamma^*$  and further derive the asymptotic distribution for  $\hat{\gamma}$ . You don't need to simplify the final expression.

- (d) (5 points) Using  $\hat{\gamma}$ , we obtain an estimator for  $\mu$  as

$$\hat{\mu} = n^{-1} \sum_{i=1}^n \left[ \hat{\gamma} X_i + \frac{R_i}{\pi(X_i)}(Y_i - \hat{\gamma} X_i) \right].$$

What is the asymptotic distribution for  $\hat{\mu}$ ? Justify your answer.

- (e) (5 points) We now consider  $g(x)$  to be any arbitrary function. Show that  $g^*(x) = E[Y|X = x]$  is the optimal function for  $g$  that minimizes the asymptotic variance in (b). Suggest an estimator using the observed data to estimate this optimal function.

## Solution

*Note: Students are expected to obtain at least 10 points out of (a)-(c), 5 points out of (d)-(e)*

(a) It follows from

$$\begin{aligned} E \left[ \frac{R}{\pi(X)} (Y - g(X)) \right] &= E \left[ E \left\{ \frac{R}{\pi(X)} (Y - g(X)) | X \right\} \right] \\ &= E \left[ E \left\{ \frac{R}{\pi(X)} | X \right\} E \{ Y - g(X) | X \} \right] = E[Y - g(X)]. \end{aligned}$$

(b) By CLT,  $\sqrt{n}(\tilde{\mu} - \mu) \rightarrow_d N(0, \tilde{\sigma}^2)$ , where

$$\begin{aligned} \tilde{\sigma}^2 &= \text{Var} \left( \frac{R}{\pi(X)} Y - \frac{R - \pi(X)}{\pi(X)} g(X) \right) \\ &= \text{Var} \left( \frac{R}{\pi(X)} Y \right) - 2 \text{Cov} \left( \frac{R}{\pi(X)} Y, \frac{R - \pi(X)}{\pi(X)} g(X) \right) + \text{Var} \left( \frac{R - \pi(X)}{\pi(X)} g(X) \right) \\ &= E[Y^2/\pi(X)] - \mu^2 - 2E \left[ \frac{R(R - \pi(X))}{\pi(X)^2} Y g(X) \right] + E \left[ \left( \frac{R - \pi(X)}{\pi(X)} g(X) \right)^2 \right] \\ &= \text{Var}(Y) + E[w(X)Y^2] - 2E[w(X)Yg(X)] + E[w(X)g(X)^2] \\ &= \text{Var}(Y) + E[w(X)(Y - g(X))^2]. \end{aligned}$$

(c) The first half is obvious and  $\hat{\gamma} \rightarrow_{a.s.} \gamma^*$  follows from SLLN. To obtain the asymptotic distribution, note

$$\sqrt{n} \left\{ \left( n^{-1} \sum_{i=1}^n \frac{R_i Y_i w(X_i)}{\pi(X_i)}, n^{-1} \sum_{i=1}^n X_i^2 w(X_i) \right) - \left( E \left[ \frac{RYw(X)}{\pi(X)} \right], E[X^2 w(X)] \right) \right\} \rightarrow_d N(0, \Sigma),$$

where

$$\Sigma = \begin{pmatrix} \text{Var}(RYw(X)/\pi(X)) & \text{Cov}(RYw(X)/\pi(X), X^2 w(X)) \\ \text{Cov}(RYw(X)/\pi(X), X^2 w(X)) & \text{Var}(X^2 w(X)) \end{pmatrix}.$$

Then we apply the delta method to obtain

$$\sqrt{n}(\hat{\gamma} - \gamma^*) \rightarrow_d N \left( 0, \frac{(1, -\gamma^*) \Sigma (1, -\gamma^*)^T}{E[X^2 w(X)]^2} \right).$$

(d) Since

$$\hat{\mu} = n^{-1} \sum_{i=1}^n \left[ \gamma^* X_i + \frac{R_i}{\pi(X_i)} (Y_i - \gamma^* X_i) \right] + (\hat{\gamma} - \gamma^*) \left[ n^{-1} \sum_{i=1}^n X_i \left\{ 1 - \frac{R_i}{\pi(X_i)} \right\} \right],$$

and note

$$\left[ n^{-1} \sum_{i=1}^n X_i \left\{ 1 - \frac{R_i}{\pi(X_i)} \right\} \right] \rightarrow_{a.s.} 0,$$

we obtain

$$\sqrt{n}(\hat{\mu} - \mu) = \sqrt{n}(n^{-1} \sum_{i=1}^n \left[ \gamma^* X_i + \frac{R_i}{\pi(X_i)} (Y_i - \gamma^* X_i) \right] - \mu) + o_p(1),$$

so  $\sqrt{n}(\hat{\mu} - \mu)$  converges in distribution to  $N(0, \tilde{\sigma}^2)$  where  $\tilde{\sigma}^2$  is given in (b) with  $g(x) = \gamma^* x$ .

(e) Since

$$E[w(X)(Y - g(X))^2] = E[w(X)Var(Y|X)] + E[w(X)(E[Y|X] - g(X))^2],$$

$\tilde{\sigma}^2$  is minimized when  $g(x) = E[Y|X = x]$ . We use the empirical data to estimate  $g$  using the kernel estimators but each subject is weighted by  $R_i/\pi(X_i)$ , i.e.,

$$\frac{\sum_{i=1}^n R_i/\pi(X_i) K_{a_n}(X_i - x) Y_i}{\sum_{i=1}^n R_i/\pi(X_i) K_{a_n}(X_i - x)}.$$