

EECS 553 HW8

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November 2024

1 Problem 1

(a): To avoid confusion, I will still use the original notation. We first notice that :

$$P(y = k|\mathbf{x}, \pi') = \frac{P(\mathbf{x}|y = k, \pi')P(y = k|\pi')}{P(\mathbf{x}|\pi')} \quad (1)$$

$$= \frac{P(\mathbf{x}|y = k, \pi)\pi'_k}{P(\mathbf{x}|\pi')} \quad (2)$$

because we know that $\phi_k(\mathbf{x})$ is unaffected by the shift. Now we also know that

$$P(\mathbf{x}|y = k, \pi) = \frac{P(y = k|\mathbf{x}, \pi)P(\mathbf{x}|\pi)}{\pi_k}$$

Now if we substitute this term into (1), we will observe that

$$P(y = k|\mathbf{x}, \pi') = P(y = k|\mathbf{x}, \pi) \cdot \frac{\pi'_k}{\pi_k} \cdot \frac{P(\mathbf{x}|\pi)}{P(\mathbf{x}|\pi')}$$

Observe that

$$P(\mathbf{x}|\pi') = \sum_{j=0}^1 P(\mathbf{x}|y = j, \pi')\pi'_j \quad (3)$$

$$= \sum_{j=0}^1 P(\mathbf{x}|y = j, \pi)\pi'_j \quad (\text{because } \phi_k(\mathbf{x}) \text{ is unaffected by shift}) \quad (4)$$

$$= P(\mathbf{x}|\pi) \sum_{j=0}^1 \frac{P(y = j|\mathbf{x}, \pi)\pi'_j}{\pi_j} \quad (5)$$

Therefore, we finally conclude that

$$P(y = k|\mathbf{x}, \pi') = \frac{\frac{\pi'_k}{\pi_k} P(y = k|\mathbf{x}, \pi)}{\frac{\pi'_0}{\pi_0} P(y = 0|\mathbf{x}, \pi) + \frac{\pi'_1}{\pi_1} P(y = 1|\mathbf{x}, \pi)}$$

which is exactly the same expression if we use the shorthand notation. This completes the proof.

(b) & (c): For the incomplete data likelihood function, we have that

$$l(\pi') = \log p(\mathbb{X}|\pi') \quad (6)$$

$$= \log \prod_{i=1}^m [\phi_1(\mathbf{x}_i)\pi' + \phi_0(\mathbf{x}_i)(1 - \pi')] \quad (7)$$

$$= \sum_{i=1}^m \log \left(\phi_1(\mathbf{x}_i)\pi' + \phi_0(\mathbf{x}_i)(1 - \pi') \right) \quad (8)$$

Now for the complete data log likelihood function, we have that

$$\log P(\mathbb{X}, \mathbb{Z} | \pi') = \log P(\mathbb{Z} | \mathbb{X}, \pi') + \log P(\mathbb{X} | \pi')$$

Now we need to find $\log P(\mathbb{Z} | \mathbb{X}, \pi')$, that is :

$$\log P(\mathbb{Z} | \mathbb{X}, \pi') = \sum_{i=1}^m y_i \log P(y_i = 1 | \mathbb{X}, \pi') + (1 - y_i) \log P(y_i = 0 | \mathbb{X}, \pi')$$

Thus, combine the $\log P(\mathbb{Z} | \mathbb{X}, \pi')$ and $\log P(\mathbb{X} | \pi')$ will give us the full log-likelihood.

E-Step

Now to compute $Q(\pi'; \pi'^{(t)})$, we observe that we need to compute that

$$E(y_i | \mathbb{X}, \pi'^{(t)}) = P^{(t)}(y = 1 | \mathbb{X}, \pi'^{(t)})$$

that denote the randomness of $Q(\pi'; \pi'^{(t)})$. We also notice that by Bayes rule, we have that

$$P(y_i = 1 | \mathbb{X}, \pi') = \frac{\phi_1(\mathbf{x}_i) \pi'}{(1 - \pi') \phi_0(\mathbf{x}_i) + \pi' \phi_1(\mathbf{x}_i)}$$

Thus, plug the above term into $Q(\pi'; \pi'^{(t)})$ and after some simplification, we find that

$$Q(\pi'; \pi'^{(t)}) = \sum_{i=1}^m P^{(t)}(y_i = 1 | \mathbb{X}, \pi'^{(t)}) \cdot [\log \phi_1(\mathbf{x}_i) + \log \pi' - \log((1 - \pi') \phi_0(\mathbf{x}_i) + \pi' \phi_1(\mathbf{x}_i))] \quad (9)$$

$$+ (1 - P^{(t)}(y_i = 1 | \mathbb{X}, \pi'^{(t)})) \cdot [\log \phi_0(\mathbf{x}_i) + \log(1 - \pi') - \log((1 - \pi') \phi_0(\mathbf{x}_i) + \pi' \phi_1(\mathbf{x}_i))] \quad (10)$$

$$+ \sum_{i=1}^m \log(\phi_1(\mathbf{x}_i) \pi' + \phi_0(\mathbf{x}_i) (1 - \pi')) \quad (11)$$

M-Step

Now take derivative respect to π' , we will finally simplify to the below expression that:

$$\frac{\partial Q}{\partial \pi'} = \frac{1}{\pi' (1 - \pi')} \sum_{i=1}^m P^{(t)}(y_i = 1 | \mathbf{x}_i, \pi'^{(t)}) - \frac{m}{1 - \pi'} = 0$$

Thus, we can then conclude that the final M-step update reduces to:

$$\pi'^{(t+1)} = \frac{1}{m} \sum_{i=1}^m P^{(t)}(y_i = 1 | \mathbf{x}_i, \pi'^{(t)})$$

where $P^{(t)}(y_i = 1 | \mathbf{x}_i, \pi'^{(t)})$ is obtained after we replacing π'_k with $\pi'^{(t)}$ and \mathbf{x} with \mathbf{x}_i . This completes part(b) and part(c).

(d) Please check the py and ipynb file I uploaded on canvas to see the code, or you can see it in the end of the file. The unadjust LR has accuracy of 0.83, the EM-adjusted LR has accuracy of 0.9, the accuracy of Clairvoyant adjusted LR is 0.9.

2 Problem 2

Given that $\epsilon_i \sim \text{Laplacian}(\beta)$, we conclude that $y_i \sim \text{Laplacian}(w^T x_i, \beta)$ that is

$$p(y_i) = \frac{\beta}{2} \exp(-\beta |y_i - w^T x_i|)$$

Thus, the likelihood can be written as

$$L(\mathbf{y} | w) = \prod_{i=1}^n \frac{\beta}{2} \exp(-\beta |y_i - w^T x_i|) \quad (12)$$

$$= \left(\frac{\beta}{2}\right)^n \exp\left(-\beta \sum_{i=1}^n |y_i - w^T x_i|\right) \quad (13)$$

so the log-likelihood could be written as

$$l(w|\mathbf{y}) = -\beta \sum_{i=1}^n |y_i - w^T x_i| + n \log \left(\frac{\beta}{2} \right)$$

Thus, to maximize the likelihood and find w to achieve this, it is same to solve the minimization problem that

$$\min_w \frac{1}{n} \sum_{i=1}^n |y_i - w^T x_i| - \frac{1}{\beta} \log \left(\frac{\beta}{2} \right)$$

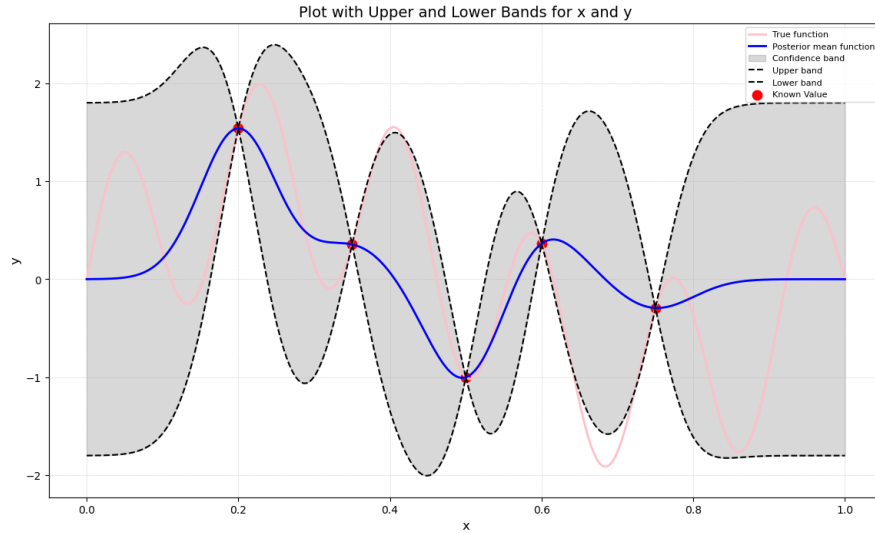
and notice the last term is unrelated to w , therefore, it is same to solve problem of empirical risk minimization that

$$\min_w \frac{1}{n} \sum_{i=1}^n L(y_i, f(x_i))$$

that the loss function is $L(y, t) = |y - t|$ and $f(x_i) = w^T x_i$. This completes the proof.

3 Problem 3

(a): please check the below picture.



(b): Please check the code I uploaded in canvas, or at the end of the page. The next value of x that going to evaluate is 0.2474.

(c): The next point that PI method suggest to evaluate is 0.2004.

(d): The next point that EI method suggest to evaluate is 0.234. Notice here I did not use the exact formula for calculating expectation, instead, I sample points from the normal distribution and then average them to get the EI for each point, which make sense by law of large number.(Monte-Carlo)

(e): All three method will approximately converge to 0.2276 with function value of 1.99. The detail of algorithm, number of iteration, epsilon can be seen in the below code I attached.