#### Chapter 4

# Simultaneous Inferences and Other Topics in Regression Analysis

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### Simultaneous Inference

• Start with ordinary confidence intervals for  $\beta_0$  and  $\beta_1$ 

$$b_0 \pm t(1 - \alpha/2; n - 2)s\{b_0\}$$
  
 $b_1 \pm t(1 - \alpha/2; n - 2)s\{b_1\}$ 

- Each interval has  $100(1-\alpha)\%$  confidence level.
- What about the overall/family confidence level?
  - Level of confidence that all constructed intervals contain their true parameter values
  - Often much lower than individual  $100(1-\alpha)\%$  level

## Bonferroni Inequality

- Let  $A_1$  denote the event that the first confidence interval does not cover  $\beta_0$ , i.e.  $P(A_1) = \alpha$
- Let  $A_2$  denote the event that the second condence interval does not cover  $\beta_1$ , i.e.  $P(A_2) = \alpha$
- We want to know the probability that both estimates fall in their respective confidence intervals, i.e.  $P(\bar{A}_1 \cap \bar{A}_2)$

$$P(\overline{A}_1 \cap \overline{A}_2) = 1 - P(A_1 \cup A_2) = 1 - P(A_1) - P(A_2) + P(A_1 \cap A_2)$$

$$\Rightarrow 1 - 2\alpha \le 1 - P(A_1) - P(A_2) \le P(\overline{A}_1 \cap \overline{A}_2) \le 1 - P(A_1) = 1 - \alpha$$

$$1-2\alpha \le \text{family confidence level} \le 1-\alpha$$

#### **Bonferroni Correction for Simultaneous Inference**

- Want to have family confidence level  $100(1-\alpha)\%$  or family type I error rate  $\alpha$ .
- Consider g CIs or hypothesis tests, each using  $\alpha^* = \alpha/g$

$$\Pr\left(\bigcap_{i=1}^{g} \overline{A}_i\right) \ge 1 - g\alpha^* = 1 - \alpha$$

where  $\overline{A}_i$  is *i*-th acception region or CI

- Use confidence level  $1-\alpha/g$  for each CI
- Use significance level  $\alpha/g$  for each hypothesis test
- Increasingly conservative as *g* increases

# Joint Confidence Intervals for $eta_0$ and $eta_1$

- Want to have family confidence level  $100(1-\alpha)\%$  for joint estimation of  $\beta_0$  and  $\beta_1$
- Make each confidence interval at  $100(1-\alpha/2)\%$  confidence level.

$$(1-\alpha/2)100\%$$
 CI for  $\beta_0$ :  $b_0 \pm t(1-\alpha/4; n-2)s\{b_0\}$   
 $(1-\alpha/2)100\%$  CI for  $\beta_1$ :  $b_1 \pm t(1-\alpha/4; n-2)s\{b_1\}$ 

• E.g, if we want to find joint CIs with family confidence level 95%, we need set-up 97.5% confidence interval for each  $\beta_0$  and  $\beta_1$ 

### Simultaneous Estimation of Mean Responses

- Working-Hotelling Method:
  - Confidence Band for Entire Regression Line.
  - Can be used for any number of Confidence Intervals for means, simultaneously
- Bonferroni Method:
  - Can be used for any g Confidence Intervals for means by creating  $(1-\alpha/g)100\%$  CIs at each of g specified X levels

Working-Hotelling: 
$$\hat{Y}_h \pm Ws \left\{ \hat{Y}_h \right\}$$
  $W = \sqrt{2F(1-\alpha;2,n-2)}$   
Bonferroni:  $\hat{Y}_h \pm Bs \left\{ \hat{Y}_h \right\}$   $B = t(1-\alpha/(2g);n-2)$ 

Bonferroni: 
$$\hat{Y}_h \pm Bs \left\{ \hat{Y}_h \right\}$$
  $B = t \left( 1 - \alpha / (2g); n - 2g \right)$ 

### Regression Through the Origin

- In some applications, it is believed that the regression line goes through the origin, this implies that  $E\{Y|X\} = \beta_1 X$  (proportional relation)
- Should only be used if there is a strong theoretical reason
- Should use *t*-test for  $\beta_1$
- Degrees of freedoms for some statistics are changed!
  - Model:  $Y_i = \beta_1 X_i + \varepsilon_i$   $\varepsilon_i \sim N(0, \sigma^2)$  independent
- LSE(least square estimator) or MLE of  $oldsymbol{eta}_1$

$$Q = \sum_{i=1}^{n} (Y_i - \beta_1 X_i)^2 \Rightarrow \frac{\partial Q}{\partial \beta_1} = 0 \Rightarrow b_1 = \frac{\sum_{i=1}^{n} X_i Y_i}{\sum_{i=1}^{n} X_i^2}$$

# Regression Through the Origin

$$\sum_{i=1}^{n} X_{i} Y_{i} = n \begin{pmatrix} V \end{pmatrix} \begin{pmatrix} V \end{pmatrix}$$

$$b_{1} = \frac{\sum_{i=1}^{n} X_{i} Y_{i}}{\sum_{i=1}^{n} X_{i}^{2}} = \sum_{i=1}^{n} \left( \frac{X_{i}}{\sum_{i=1}^{n} X_{i}^{2}} \right) Y_{i} \sim N \left( \beta_{1}, \frac{\sigma^{2}}{\sum_{i=1}^{n} X_{i}^{2}} \right)$$

$$\sum_{i=1}^{n} X_{i}^{2} \qquad \stackrel{i=1}{=} \left[ \sum_{i=1}^{n} X_{i}^{2} \right] \qquad \left[ \sum_{i=1}^{n} X_{i}^{2} \right]$$

$$\hat{Y}_{i} = b_{1} X_{i} \qquad e_{i} = Y_{i} - \hat{Y}_{i} \qquad SSE = \sum_{i=1}^{n} e_{i}^{2} \sim \sigma^{2} \chi^{2} (n-1)$$

$$MSE = \frac{SSE}{n-1}, \quad s^{2} \left\{ b_{1} \right\} = \frac{MSE}{\sum_{i=1}^{n} X_{i}^{2}}, \quad \frac{b_{1} - \beta_{1}}{s \left\{ b_{1} \right\}} \sim t(n-1)$$

• 
$$H_0$$
:  $\beta_1 = 0$ 

• Test statistic 
$$t^* = \frac{b_1}{n} \stackrel{H_0}{\sim} t(n)$$

• Test statistic 
$$t^* = \frac{b_1}{s\{b_i\}} \stackrel{H_0}{\sim} t(n-1)$$

### Regression Through the Origin

• Confidence intervals(CI) or prediction interval(PI)

$$(1-\alpha)100\%$$
 CI for  $\beta_1 : b_1 \pm t(1-(\alpha/2), n-1)s\{b_1\}$ 

$$(1-\alpha)100\% \text{ CI for } \beta_1 : b_1 \pm t(1-(\alpha/2), n-1)s\{b_1\}$$

$$s^{2} \{b_{1}\} = \frac{MSE}{\sum_{i=1}^{n} X_{i}^{2}}; \quad s^{2} \{\hat{Y}_{h}\} = \frac{X_{h}^{2}}{\sum_{i=1}^{n} X_{i}^{2}} MSE; \quad s^{2} \{pred\} = MSE \left(1 + \frac{X_{h}^{2}}{\sum_{i=1}^{n} X_{i}^{2}}\right)$$

$$(1-\alpha)100\% \text{ CI for } E(Y_h) = \beta_1 X_h : \hat{Y}_h \pm t \left(1 - (\alpha/2), n-1\right) s \left\{\hat{Y}_h\right\}$$

$$(1-\alpha)100\%$$
 PI for  $Y_{h(new)}: \hat{Y}_h \pm t(1-(\alpha/2), n-1)s\{\text{pred}\}$ 

### Homework

- Page 172~(174):
  - 4.4 (b); 4.8(a),(b)
  - 4.16; 4.17(c)