

AI2613: Itô Integral, Itô Formula

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1 Itô Integral

Recall that in the last lecture, we formalized a diffusion $\{X_t\}$ as

$$dX_t = \mu(t, X_t) dt + \sigma(t, X_t) dB_t, \quad (1)$$

where $\{B_t\}$ is the standard Brownian motion. With this formalization, the motion of $\{X_t\}$ in a tiny time interval $[t, t + h]$ can be viewed as a moving under the differential equation $\frac{dX_t}{dt} = \mu(t, X_t)$ with a random perturbation $\sigma(t, X_t) dB_t$. In this lecture, we provide the exact mathematical meaning of the above stochastic differential equation.

Given an ordinary differential equation $df(t) = f(t) dt$, we have that,

$$\forall T, \int_0^T df(t) = \int_0^T f(t) dt,$$

which is equivalent to

$$\forall T, f(t) = f(0) + \int_0^T f(t) dt.$$

If we apply the same process to Equation (1), we have that

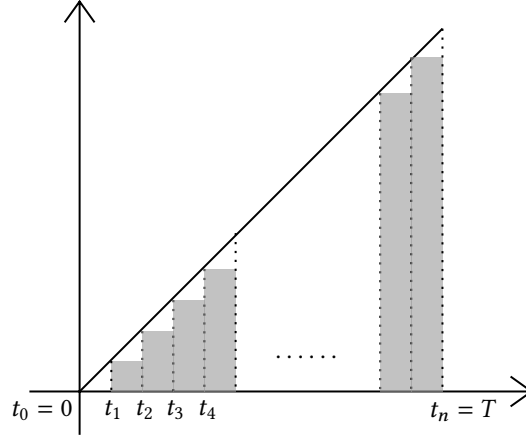
$$\forall T, X_T = X_0 + \int_0^T \mu(t, X_t) dt + \int_0^T \sigma(t, X_t) dB_t. \quad (2)$$

Let Ω be the sample space on which $\{X_t\}$ is defined. For a fixed $\omega \in \Omega$, both $\{X_t(\omega)\}$ and $\{B_t(\omega)\}$ are fixed functions with regard to t . Then Equation (2) means $\forall T, \forall \omega \in \Omega$,

$$X_t(\omega) = X_0(\omega) + \int_0^T \mu(t, X_t(\omega)) dt + \int_0^T \sigma(t, X_t(\omega)) dB_t(\omega).$$

We will omit the ω in the following if the identity holds for *almost every* ω . Note that $\int_0^T \mu(t, X_t) dt$ is the ordinary Riemann integral of $\mu(t, X_t)$. The main goal today is to rigorously define the meaning of $\int_0^T \sigma(t, X_t) dB_t$.

We first recall the Riemann integral and the Riemann-Stieltjes integral.



1.1 Riemann Integral

Consider the simple integral $\int_0^T x \, dx$. As the following figure shows, we divide $[0, T]$ into n disjoint segments $[0, t_1], (t_1, t_2], \dots, (t_{n-1}, T]$. Let $\Delta_i = t_i - t_{i-1}$ and $\Delta = \max_{i \in [n]} \Delta_i$. Then

$$\int_0^T x \, dx = \lim_{\Delta \rightarrow 0} \sum_{i=1}^n t_{i-1} (t_i - t_{i-1}) = \lim_{\Delta \rightarrow 0} -\frac{1}{2} \sum_{i=1}^n (t_i - t_{i-1})^2 + \frac{1}{2} \sum_{i=1}^n (t_i^2 - t_{i-1}^2).$$

Note that $\sum_{i=1}^n (t_i^2 - t_{i-1}^2) = T^2$ and

$$\sum_{i=1}^n (t_i - t_{i-1})^2 \leq \left(\max_{j \in [n]} \Delta_j \right) \cdot \sum_{i=1}^n (t_i - t_{i-1}) = \left(\max_{j \in [n]} \Delta_j \right) \cdot T \rightarrow 0.$$

Therefore, we have $\int_0^T x \, dx = \frac{T^2}{2}$.

1.2 Riemann-Stieltjes Integral

Then we use similar idea to calculate Riemann-Stieltjes integral¹

Let $F: [0, T] \rightarrow \mathbb{R}$ be a function with bounded derivative. Similarly, we divide $[0, T]$ into n disjoint segments $[0, t_1], (t_1, t_2], \dots, (t_{n-1}, T]$. Assume notations the same as Section 1.1. Then the Riemann-Stieltjes

¹The Riemann-Stieltjes integral has been widely used in probability theory. Let X be a random variable on sample space $[0, 1]$. Assume that the CDF of X is F and the PDF of X is f . Then the expectation of X is $E[X] = \int_0^1 X \, dF(t)$. By the definition of the Riemann-Stieltjes integral, we have

$$\int_0^1 X \, dF = \lim_{n \rightarrow \infty} X(t_i^*) (F(t_i) - F(t_{i-1})) = \lim_{n \rightarrow \infty} X(t_i^*) f(t_{i-1}) (t_i - t_{i-1}) + o(1),$$

which yields that $E[X] = \int_0^1 X_t \, dF(t) = \int_0^T X_t f(t) \, dt$.

integral of x with respect to F is defined by

$$\begin{aligned}\int_0^T x \, dF(x) &\triangleq \lim_{\Delta \rightarrow 0} \sum_{i=1}^n t_{i-1} (F(t_i) - F(t_{i-1})) \\ &= \lim_{\Delta \rightarrow 0} \sum_{i=1}^n t_{i-1} \cdot F'(t_{i-1}) (t_i - t_{i-1}) \\ &= \int_0^T x F'(x) \, dx\end{aligned}$$

We can use the same idea as Section 1.1 to calculate $\int_0^T F(x) \, dF(x)$:

$$\begin{aligned}\int_0^T F(x) \, dF(x) &= \lim_{\Delta \rightarrow 0} \sum_{i=1}^n F(t_{i-1}) (F(t_i) - F(t_{i-1})) \\ &= \frac{1}{2} (F(T)^2 - F(0)^2) - \lim_{\Delta \rightarrow 0} \frac{1}{2} \sum_{i=1}^n (F(t_i) - F(t_{i-1}))^2.\end{aligned}$$

If the derivative of F is bounded by M on $[0, T]$, we have

$$\begin{aligned}\lim_{\Delta \rightarrow 0} \sum_{i=1}^n (F(t_i) - F(t_{i-1}))^2 &\leq \lim_{\Delta \rightarrow 0} \left(\max_{j \in [n]} |F(t_j) - F(t_{j-1})| \right) \cdot \sum_{i=1}^n |F(t_i) - F(t_{i-1})| \\ &\leq \lim_{\Delta \rightarrow 0} \left(\max_{j \in [n]} |F(t_j) - F(t_{j-1})| \right) \cdot \sum_{i=1}^n M \cdot (t_i - t_{i-1}) \\ &= \lim_{\Delta \rightarrow 0} \left(\max_{j \in [n]} |F(t_j) - F(t_{j-1})| \right) \cdot MT = 0.\end{aligned}$$

1.3 Itô Integral

Consider what will happen if we substitute F with the standard Brownian motion $\{B_t\}$. We can also deduce that

$$\int_0^T B_t \, dB_t = \frac{1}{2} B_T^2 - \frac{1}{2} \lim_{\Delta \rightarrow 0} \sum_{i=1}^n (B_{t_i} - B_{t_{i-1}})^2.$$

However, as $B_t(\omega)$ could be non-differentiable with regard to t , we can not calculate this integral as we do in the Riemann-Stieltjes integral. The term $\lim_{\Delta \rightarrow 0} \sum_{i=1}^n (B_{t_i} - B_{t_{i-1}})^2$ might not vanish! In this following, we aim at understanding what it is.

Let $Q_n = \sum_{i=1}^n (B_{t_i} - B_{t_{i-1}})^2$. It is a random variable, and thus we first compute its expectation and variance. Recall that $\Delta_i = t_i - t_{i-1}$. We know $B_{t_i} - B_{t_{i-1}} \sim \mathcal{N}(0, \Delta_i)$. Then

$$\mathbb{E}[Q_n] = \sum_{i=1}^n \mathbb{E}[(B_{t_i} - B_{t_{i-1}})^2] = \sum_{i=1}^n \Delta_i = T,$$

and

$$\begin{aligned}\text{Var}[Q_n] &= \sum_{i=1}^n \text{Var}\left[(B_{t_i} - B_{t_{i-1}})^2\right] = \sum_{i=1}^n \mathbb{E}\left[(B_{t_i} - B_{t_{i-1}})^4\right] - \sum_{i=1}^n \mathbb{E}\left[(B_{t_i} - B_{t_{i-1}})^2\right]^2 \\ &= 2 \sum_{i=1}^n \Delta_i^2 \leq 2 \left(\max_{i \in [n]} \Delta_i\right) \cdot \sum_{i=1}^n \Delta_i.\end{aligned}$$

The variance tends to 0 as $\Delta \rightarrow 0$ (or equivalently as $n \rightarrow \infty$ if we choose each Δ_i to have equal length). This means that Q_n converges to T in the following mean square sense.

Definition 1 (Mean Square Convergence). Let Z_1, Z_2, \dots and Z be random variables that $\mathbb{E}[Z^2] < \infty$ and $\mathbb{E}[Z_n^2] < \infty$ for $n \geq 1$. We say Z_n converges to Z in mean square, or Z is the mean square limit of $\{Z_n\}$, if $\lim_{n \rightarrow \infty} \mathbb{E}[(Z_n - Z)^2] = 0$. This is denoted as $Z_n \xrightarrow{L_2} Z$.

Then we can define the Itô integral.

Definition 2. Assume that $\{X_t\}$ is a “nice enough” stochastic process². Then we define the integral $\int_0^T X_t dB_t$ as the mean square limit of

$$\sum_{i=1}^n X_{t_{i-1}} (B_{t_i} - B_{t_{i-1}}).$$

This is called the Itô integral of $\{X_t\}$ with respect to $\{B_t\}$.

With Definition 2, we can verify that

$$\int_0^T B_t dB_t = \frac{1}{2} B_T^2 - \frac{1}{2} T.$$

More generally, we may define $\int_0^T X_t dB_t$ as the mean square limit of

$$\sum_{i=1}^n X_{t_i^*} (B_{t_i} - B_{t_{i-1}}),$$

where $t_i^* = \alpha t_{i-1} + (1 - \alpha)t_i$ with $\alpha \in [0, 1]$. The Itô integral defined in Definition 2 corresponds to the case that $\alpha = 1$. By choosing $\alpha = \frac{1}{2}$, we have the definition of Stratonovich integral and it holds that $\int_0^T B_t dB_t = \frac{1}{2} B_T^2$ with Stratonovich integral.

2 Itô Formula

Recall that in the example in Section 1.3, we have

$$Q_n = \sum_{i=1}^n (B_{t_i} - B_{t_{i-1}})^2, \quad \mathbb{E}[Q_n] = T \quad \text{and} \quad \text{Var}[Q_n] \rightarrow 0.$$

²This means that the process has some nice properties, such as that X_t is independent with $\{B_u\}_{u>t}$ for all t . Some other technical requirements can be found in any standard textbook on stochastic differential equations (e.g., [Kle12])

In other words,

$$\lim_{n \rightarrow \infty} Q_n = \int_0^T (\mathrm{d}B_t)^2 = T = \int_0^T \mathrm{d}t.$$

This suggests that $(\mathrm{d}B_t)^2 \approx \mathrm{d}t$ holds. By definition of $\{B_t\}$, we know that $\mathrm{d}B_t = B_{t+\mathrm{d}t} - B_t \sim \mathcal{N}(0, \mathrm{d}t)$. Let X and Y be two random variables that $X \sim \mathcal{N}(0, \mathrm{d}t)$ and $Y = X^2$. Then the formula $(\mathrm{d}B_t)^2 \approx \mathrm{d}t$ tells us that Y is well concentrated on $\mathrm{d}t$.

With this observation, we then (heuristically) deduce the chain rule under the definition of Itô integral.

2.1 Classical Chain Rule of Differentiation

For any differentiable function f , we have $\mathrm{d}f(t) = f'(t) \mathrm{d}t$. This can be verified via Taylor's expansion:

$$\mathrm{d}f(t) = f(t + \mathrm{d}t) - f(t) = f'(t) \mathrm{d}t + \frac{1}{2} f''(t) (\mathrm{d}t)^2 + o((\mathrm{d}t)^2) \xrightarrow{\mathrm{d}t \rightarrow 0} f'(t) \mathrm{d}t.$$

For two differentiable functions g and f , the chain rule of differentiation is

$$\frac{\mathrm{d}f(g(t))}{\mathrm{d}t} = f'(g(t)) \cdot g'(t).$$

We can also derive this using the Taylor expansion:

$$\begin{aligned} \mathrm{d}f(g(t)) &= f(g(t + \mathrm{d}t)) - f(g(t)) \\ &= f(g(t) + \mathrm{d}g(t)) - f(g(t)) \\ &= f'(g(t)) \mathrm{d}g(t) + \frac{1}{2} f''(g(t)) (\mathrm{d}g(t))^2 + o((\mathrm{d}g(t))^2). \end{aligned}$$

Then it follows that

$$\frac{\mathrm{d}f(g(t))}{\mathrm{d}t} = f'(g(t)) \cdot g'(t) + o(\mathrm{d}t) \xrightarrow{\mathrm{d}t \rightarrow 0} f'(g(t)) \cdot g'(t).$$

2.2 The Chain Rule with Itô Integral

Consider a diffusion $\{X_t\}$ such that

$$\mathrm{d}X_t = \mu_t \mathrm{d}t + \sigma_t \mathrm{d}B_t,$$

where μ_t and σ_t are the abbreviations of $\mu(t, X_t)$ and $\sigma(t, X_t)$ respectively. By the Taylor's expansion,

$$\begin{aligned} \mathrm{d}f(X_t) &= f(X_t + \mathrm{d}X_t) - f(X_t) \\ &= f'(X_t) \mathrm{d}X_t + \frac{1}{2} f''(X_t) (\mathrm{d}X_t)^2 + o((\mathrm{d}X_t)^2) \\ &= f'(X_t) (\mu_t \mathrm{d}t + \sigma_t \mathrm{d}B_t) + \frac{1}{2} f''(X_t) (\mu_t \mathrm{d}t + \sigma_t \mathrm{d}B_t)^2 + o((\mathrm{d}X_t)^2) \\ &\xrightarrow{\mathrm{d}t \rightarrow 0} \left(f'(X_t) \mu_t + \frac{1}{2} f''(X_t) \sigma_t^2 \right) \mathrm{d}t + f'(X_t) \sigma_t \mathrm{d}B_t \end{aligned}$$

This chain rule with Itô integral is called Itô formula. Then we see some examples of Itô formula.

Example 1. We can use Itô formula to calculate B_T^2 . Note that

$$d(B_t)^2 = 2B_t dB_t + (dB_t)^2 = 2B_t dB_t + dt.$$

Therefore, $B_T^2 = 2 \int_0^T B_t dB_t + T$. This provides a quick justification that $\int_0^T B_t dB_t = \frac{1}{2}B_T^2 - \frac{1}{2}T$.

Example 2 (Geometric Brownian Motion). The geometric Brownian motion is $Y_t = e^{B_t}$ where $\{B_t\}$ is the standard Brownian motion. Then it follows from the Itô formula that

$$dY_t = e^{B_t} \left(dB_t + \frac{1}{2} (dB_t)^2 \right) = Y_t dB_t + \frac{1}{2} Y_t dt.$$

Example 3 (Ornstein-Uhlenbeck Process). Let $\{X_t\}$ be a Ornstein-Uhlenbeck process that $dX_t = -X_t dt + 2dB_t$ with $X_0 = 0$. We can calculate X_t according to this equation. Let $f(t, X_t) = e^t \cdot X_t$. Using Taylor expansion, we have

$$\begin{aligned} df(t, X_t) &= f(t + dt, X_{t+dt}) - f(t, X_t) \\ &= X_t e^t dt + e^t dX_t \\ &= X_t e^t dt - e^t \cdot X_t dt + 2e^t dB_t = 2e^t dB_t. \end{aligned}$$

Then we have $e^T X_T = \int_0^T 2e^t dB_t$ and $X_t = e^{-t} \int_0^T 2e^t dB_t$. Since $\int_0^T 2e^t dB_t$ can be viewed as the limit of the weighted sum of many Gaussian random variables, one can verify that $\int_0^T 2e^t dB_t \sim \mathcal{N}\left(0, 4 \int_0^T e^{2t} dt\right)$.

References

[Kle12] Fima C Klebaner. *Introduction to stochastic calculus with applications*. World Scientific Publishing Company, 2012. 4