

Chapter 9 Sols

① Formula 9.4 states:

$$S_{ij} = \prod_{k=i}^j (1-h_k) \quad \text{Let } i=1$$

and we obtain $S_j = \prod_{k=1}^j (1-h_k)$
Prob. of living past age j .

The relationship between this expression and 9.1 is given by (9.2). More explicitly, the probability of living past age j can be written as the probability of living exactly j years, plus the probability of living exactly $j+1$ years and so forth. In math and using slightly different notation from 9.2, we have:

$$S_j = f_j + f_{j+1} + f_{j+2} + \dots$$

$$\Rightarrow S_j = \sum_{k \geq j} f_k$$

② We start from 9.4 with $i=1$
 $S_i = \prod_{k=1}^j (1-h_k)$, and assume no censoring, thus our data is only

a record of n survival times t_1, t_2, t_3, \dots .
Then, we estimate h_k by $\hat{h}_k = y_k / n_k$ to get

$$\begin{aligned}\hat{S}_j &= \prod_{k=1}^j (1 - \hat{h}_k) = \prod_{k=1}^j \left(1 - \frac{y_k}{n_k}\right) \\ &= \prod_{k=1}^j \left(\frac{n_k - y_k}{n_k}\right)\end{aligned}$$

Then, if we let $t_{(1)} < t_{(2)} < \dots < t_{(n)}$ be the ordered survival times and use a time unit small enough s.t. there are no ties, then $n_k = n - k + 1$ $\forall k$

The subject whose death will happen in the current time period.

and $y_k = 1$ (since there are no ties). Thus,

$$\hat{S}_j = \prod_{k=1}^j \left(\frac{n - k + 1 - 1}{n - k + 1}\right) = \prod_{k=1}^j \left(\frac{n - k}{n - k + 1}\right)$$

which exactly matches 9.17 if all $d(k)$'s are 1, matching our non-censored data assumption.

③ See Jupyter Notebook.

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⑤ Because the numbers on the margins are not constant from month to month, so we need to control for the changes in n_A and n_B from period to period. We

could have used a Binomial model if n_A and n_B were proportional (say with $p = \frac{n_A}{n_A + n_B}$) for all months.

⑥ - ⑧ we skipped the sections of these 3 questions.