

Croon's Measurement Error Correction

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Abstract

I summarize the procedure in Croon (2002) to correct for measurement error and I accompany it with an implementation in Stata.

In the following lines we explain very briefly how to correct for measurement error estimates from standard linear regressions estimated through OLS.

Let y_i be a (demeaned) outcome of interest, X_i be a (demeaned) vector of characteristics or controls, and θ_i be a (demeaned) vector of factors. X_i is of dimension $1 \times K$ and θ_i is of dimension $1 \times K$. The objective is to estimate the coefficients in the following model

$$y_i = \mathbf{X}_i\beta + \theta_i\gamma + \varepsilon_i \quad (1)$$

where β (vector) and γ (vector) are estimands and $\mathbb{E}[\varepsilon_i|\mathbf{X}_i, \theta_i] = 0$.

Plausibly, θ_i is measured with error, since it is a factor score based on multiple measures arbitrarily chosen by the researcher.

Croon (2002) proposes a correction. Let the relationship between the “true factor scores” and the “estimated” factor scores be given by

$$\hat{\theta}_i = \theta_i + V_i \quad (2)$$

where $\mathbb{E}[V_i] = 0$, $\mathbb{E}[V_i|\mathbf{X}_i, \theta_i] = 0$. Let the covariance matrix of \mathbf{R}_i and \mathbf{S}_i be given by

$$\Sigma = \begin{bmatrix} \Sigma_{\mathbf{R}_i, \mathbf{R}_i} & \Sigma_{\mathbf{R}_i, \mathbf{S}_i} \\ \Sigma_{\mathbf{S}_i, \mathbf{R}_i} & \Sigma_{\mathbf{S}_i, \mathbf{S}_i} \end{bmatrix}. \quad (3)$$

If we assume (4) is the data generating process, the OLS estimator is inconsistent. To see this substitute (2) into (1) and obtain the following:

$$y_i = \beta_0 + \mathbf{X}_i\beta + \theta_i\gamma + \varepsilon_i - \gamma V_i \quad (4)$$

so that

$$\text{plim} \begin{bmatrix} \hat{\beta}^{OLS} \\ \hat{\gamma}^{OLS} \end{bmatrix} = A \begin{bmatrix} \beta \\ \gamma \end{bmatrix} \quad (5)$$

where $A = \begin{bmatrix} \Sigma_{\hat{\theta}_i, \hat{\theta}_i} & \Sigma_{\theta_i, \mathbf{X}_i} \\ \Sigma_{\mathbf{X}_i, \theta_i} & \Sigma_{\mathbf{X}_i, \mathbf{X}_i} \end{bmatrix}^{-1} \begin{bmatrix} \Sigma_{\theta_i, \theta_i} & \Sigma_{\theta_i, \mathbf{X}_i} \\ \Sigma_{\mathbf{X}_i, \theta_i} & \Sigma_{\mathbf{X}_i, \mathbf{X}_i} \end{bmatrix}$. Importantly

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$$\Sigma_{\hat{\theta}_i, \hat{\theta}_i} = \Sigma_{\theta_i, \theta_i} + \Sigma_{V_i, V_i}. \quad (6)$$

Now note that $\Sigma_{\theta_i, \mathbf{X}_i} = \Sigma_{\hat{\theta}_i, \mathbf{X}_i}$, so that we can estimate a . To get it straight, θ_i is the vector of factor scores in the data. Thus, it's easy to obtain all the components in A except for the component in (6). They come from computing the empirical covariance matrices between the factor scores, θ_i , and the rest of the data in the estimation (controls), \mathbf{X}_i . To obtain $\Sigma_{\hat{\theta}_i, \hat{\theta}_i}$ calculate the variance of the factor scores, which depends on the factor extraction method. Once all the components of A have been estimated, premultiplying the OLS estimates by its inverse yields consistent estimates:

$$A^{-1} \text{plim} \begin{bmatrix} \hat{\beta}^{OLS} \\ \hat{\gamma}^{OLS} \end{bmatrix} = \begin{bmatrix} \beta \\ \gamma \end{bmatrix}. \quad (7)$$

Put differently, premultiplying the OLS estimates by A^{-1} is Croon's correction.

References

Croon, M. (2002). Using Predicted Latent Scores in General Latent Structure Models. *Latent Variable and Latent Structure Models*, 195–223.