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1 About

STATY is an educational project designed and developed by Danijela Markovic and Oskar Kärcher with the aim of improving data literacy among students of natural and social sciences.

STATY lets you focus on the analysis and interpretation of the results from many different statistical and machine learning methods using top-notch libraries without any need to write a single line of code:

- Prepare your data for analysis with data cleaning, data imputation and variable transformation
- > Examine different visualizations of your variables, even of geospatial data, and find patterns in your data
- Explore web scraping, text data analysis and further data analysis tools like Distribution fitting, Correlation analysis or ANOVA to dive deeper into the characteristics of your included variables
- Use the modelling method that suits your data best: Time series models like MA, ARMA or ARIMA, panel data models with fixed and random effects, or further regression models including multiple linear regression, logistic regression, generalized additive models, random forest, boosted regression trees and artificial neural networks

Start by simply uploading your data and letting **STATY** do the work for you!

2 Access

You can access **STATY** right here!

https://quant-works.de/staty/

STATY will open in a tab of your browser.

3 Features

STATY has many different methods for specific types of data. You can select the respective feature for analysing your data without much effort! In every feature available for data analysis all output tables (as xlsx-file) and figures are available for download.

3.1 General

Before diving into the different methodological applications, here are some basic information.

3.1.1 Notation

Throughout the documentation the following notation will be used:

n = sample size

k = number of (explanatory) variables

 $X = matrix \ of \ (explanatory) \ variables$

 $x_k = (explanatory) variable k$

 $x_i = i^{th}$ observation of the matrix X

 $x_{ik} = i^{th}$ observation of (explanatory) variable k

Y = vector for the response variable

 $y_i = i^{th}$ observation of the response variable Y

 $\beta = coefficient vector$

 $\beta_k = coefficient of explanatory variable k$

 $\hat{y}_i = i^{th}$ model prediction for the i^{th} observation of the response variable

 $\varepsilon = error$

 $\varepsilon_i = i^{th} \; model \; error \; of \; the \; i^{th} \; observation \; of \; the \; response \; variable$

N = number of groups in panel/grouped data

T = number of time periods in panel data

3.1.2 Data upload

You upload you own data in the sidebar. Here you have different upload settings, so that STATY identifies your data correctly (Figure 1). Please check your data prior to uploading, so that you can select the decimal separator, column separator, thousands separator and the encoding accordingly.

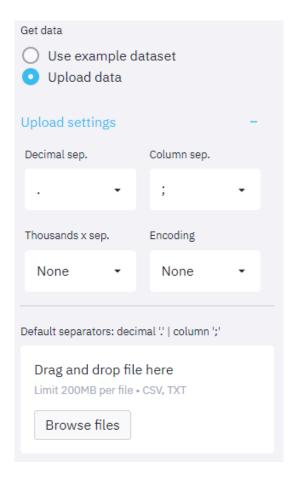


Figure 1 Data upload settings.

3.1.3 Data summary

After uploading your data, **STATY** offers you many different parameters for summarizing your data before further data processing and analyses. Besides information about missing values, duplicates, number of unique values and the data types, you have the option to inspect main summary statistics:

$$Mean_k = \frac{1}{n} \sum_i x_{ik} = \bar{x}_k$$

 $Mode_k = value \ of \ variable \ k \ that \ occurs \ most \ of ten$

 $Median_k = 50\%$ quantile

Standard deviation_k =
$$\sqrt{\frac{1}{n-1} \sum_{i} (x_{ik} - \bar{x}_k)^2} = s_k$$

$$Variance_k = \frac{1}{n-1} \sum_{i} (x_{ik} - \bar{x}_k)^2 = s_k^2$$

 $Skewness_k = see \underline{here}$

 $Kurtosis_k = see \underline{here}$

$$Min_k = \min_i x_{ik}$$

$$p\% - quantiles_k: \frac{\textit{number of values of } x_k \leq p - quantile}{n} \geq p \ \& \ \frac{\textit{number of values of } x_k \geq p - quantile}{n} \geq 1 - p$$

$$Max_k = \max_i x_{ik}$$

3.1.4 Data processing

Prepare your data for the final data analysis with several different data processing options.

Data cleaning

Duplicate rows, rows with missing values, as well as manually selected rows and columns can be deleted. Rows and specific columns that should be kept can also be selected.

Data filtering

You can filter data by the categories of a variable.

Data imputation

Missing values for numeric variables can be replaced by either the mean, mode, median or some random value of the respective variable. For other variable types, missing values can only be replaced by the mode or a random value. The imputation can be also conducted for grouped data in some cases.

Data transformation

Variables can be transformed with log, sqrt, by squaring, centering, standardizing, and normalizing.

The log and sqrt transformation are only calculated for positive values. If the minimum of the selected variable is <0, then $\min(x_i) + 1$ is added to every realization of the selected variable.

If a variable should be centred, the mean of the variable is subtracted.

Please note that standardization and normalization are only carried out, if the standard deviation is $\neq 0$ and $(\max - \min) \neq 0$, respectively.

Additionally, variables (without NAs) can be transformed into integer-type (numeric) categories, which are either based on the sorted values of the selected variable or on manually assignable categories for up to maximal 5 categories, multiplied and also divided.

Variables selected for transformation will be not replaced in the data set but added as new columns. Transformation options will be shown based on the uploaded data.

3.2 Uni- and bivariate data

For data with one or two variables, or for analyses of maximum two variables, specific tools can be used to explore the relationships between variables, the distribution type of included variables and differences between groups of a discrete variable. You can use frequency analysis, ANOVA, distribution fitting, correlation analysis, linear regression techniques and contingency tables to find out more about your data.

3.2.1 Univariate frequency analysis

With the univariate frequency analysis, you can determine the absolute, relative, and cumulative frequencies of each group or class of a discrete or continuous variable. The respective frequencies can be displayed as bar chart or table. Depending on the variable type, you have additional settings for continuous and discrete variables. For discrete variables, you can additionally specify the order of the x-labels (Figure 2).

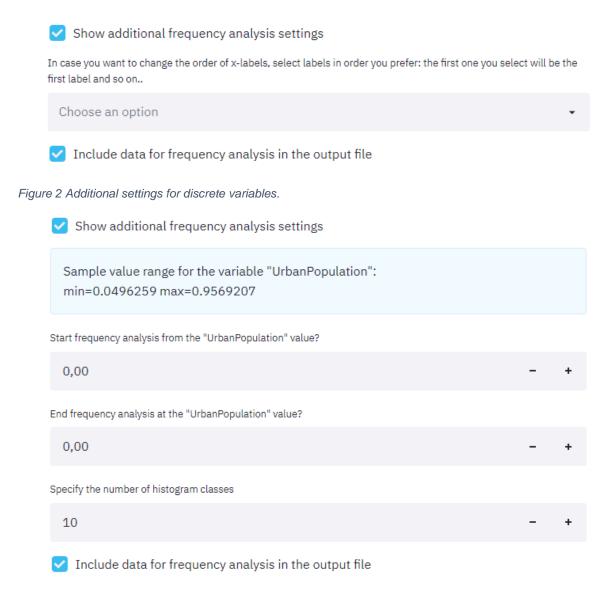


Figure 3 Additional settings for continuous variables.

For continuous variables, you can select the value range used for the frequency analysis and the number of classes (Figure 3). Furthermore, it can be selected whether the data used for the analysis should be included in the output file available for download or not for both cases.

The relative and cumulative frequency for each group x_j of variable x is determined as follows:

$$h_j$$
 (relative frequency) = $\frac{n_j}{n}$,

where n_j is the absolute frequency of group/class x_j

$$F(x_j) = \sum_{i \le j} h_i$$
, where F is the cumulative distribution function (cdf).

3.2.2 ANOVA

With Analysis Of Variance (ANOVA) you can compare the means of several groups to test whether the observations of the groups were drawn from populations with the same mean (null hypothesis) or not (alternative hypothesis). The ANOVA table consists of the sum of squares between, sum of squares within and sum of squares total, which are used to calculate the final F-statistic:

$$SS_{between} = \sum_{i=1}^{N} n_i (x_i - \bar{x})^2, \qquad df = N-1$$
 $SS_{within} = \sum_{i=1}^{N} (n_i - 1)s_i^2, \qquad df = n-N$
 $SS_{total} = SS_{between} + SS_{total}$
 $MS_{between} = \frac{SS_{between}}{N-1}$
 $MS_{within} = \frac{SS_{within}}{n-N}$
 F $statistic = \frac{MS_{between}}{MS_{within}}$

For the ANOVA test result to be reliable, the following assumptions must be fulfilled:

- (1) Independent samples of the response variable
- (2) Homogeneity of variance across groups
- (3) Normal distribution across groups

In order to perform one-way ANOVA, you have to select the target variable that you want to analyse as well as the variable that should be used as a classifier variable (Figure 4).



Figure 4 ANOVA settings.

Furthermore, it can be selected whether the data used for the analysis should be included in the output file available for download or not. Assumptions can be checked by using the provided boxplots and residual distribution plots based on the regression problem $y \sim x_{categories}$. Here, the standardized residuals are calculated as $\frac{residuals-residuals_{mean}}{residuals_{standard\ deviation}}$.

3.2.3 Hypothesis testing

With hypothesis testing by using the z-test or one sample location t-test you can test whether the deviation of the mean value based on your sample from a specified value or population mean is statistically significant or not. By using the two sample location t-test you can compare two mean values by grouping your data and test for a statistically significant deviation. Thus, you can identify whether the unknown population means for these two independent groups are equal or not.

Based on the test that you want to conduct, you must provide different information. For the z-test you must enter the population mean value and the population variance (Figure 5).

| Select the test | | | | | |
|--------------------------------|---|---|-------------------------------|---|---|
| z-test | | | | | • |
| Select the test variable | | | | | |
| UrbanPopulation | | | | | • |
| Enter sollwert/population mean | | | Enter the population variance | | |
| 0,40000000 | - | + | 0,00000000 | - | + |

Figure 5 Settings for the z-test.

$$z \, statistic = \frac{\bar{x} - \mu_0}{\sqrt{\frac{\sigma^2}{n}}}$$

where

n = sample size

 $\mu_0 = population mean$

 $\sigma^2 = population variance$

p-value = according to z statistic

The test statistic follows the normal distribution.

For the one sample location t-test only the population mean must be provided (Figure 6).

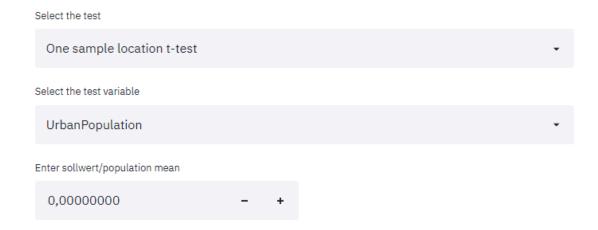


Figure 6 Settings for one sample location t-test.

$$t \ statistic = \frac{\bar{x} - \mu_0}{\sqrt{\frac{s^2}{n}}}$$

where

n = sample size

 $\mu_0 = population mean$

 $s^2 = sample \ variance$

DOF (degrees of freedom) = n - 1

p-value = according to t statistic

The test statistic follows a t-distribution.

For the two sample location t-test you have to provide information about the variance assumption (either unequal or equal variance for the two samples) and the variable that should be used to split the data into two samples. If unequal variance was selected, Welch-test is performed. In case of more than two realisations of the variable used for grouping, you can reclassify your data, whereas reclassification options depend on the variable type (Figure 7).

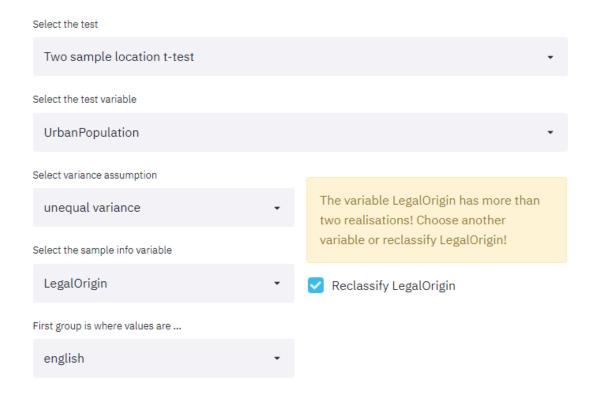


Figure 7 Settings for two sample location t-test.

Equal variance

$$t \ statistic = \sqrt{\frac{n_1 * n_2}{n_1 + n_2}} \ \frac{\bar{x}_1 - \bar{x}_2}{S}$$

where

 n_1 = sample size of first sample

 n_2 = sample size of second sample

 $\bar{x}_1 = mean \ value \ of \ first \ sample$

 \bar{x}_2 = mean value of second sample

$$s = \sqrt{\frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{n_1 + n_2 - 2}}$$

with

 $s_1^2 = variance of first sample$

 $s_2^2 = variance of second sample$

DOF (degrees of freedom) = $n_1 + n_2 - 2$

p-value = according to t statistic and test problem

<u>Unequal variance (Welch test)</u>

$$t \; statistic = \frac{\bar{x}_1 - \bar{x}_2}{s}$$

where

 $n_1 = sample \ size \ of \ first \ sample$

 n_2 = sample size of second sample

 $\bar{x}_1 = mean \ value \ of \ first \ sample$

 \bar{x}_2 = mean value of second sample

$$S = \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}$$

with

 $s_1^2 = variance of first sample$

 $s_2^2 = variance of second sample$

$$DOF (degrees of freedom) = \frac{\left(\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}\right)^2}{\frac{\left(\frac{s_1^2}{n_1}\right)^2}{n_1 - 1} + \frac{\left(\frac{s_2^2}{n_2}\right)^2}{n_2 - 1}}$$

p-value = according to t statistic and test problem

The test statistic follows a t-distribution.

3.2.4 Distribution fitting

For the variables in your data set you can determine the distribution that fits your data best. The best fit is based on the Chi-Square Goodness of Fit Test and the selected distributions (Figure 8). You can find a list of the continuous distributions included here. You can additionally specify the range and the number of bins for the histogram, which will be used to determine the best fit (Figure 8).

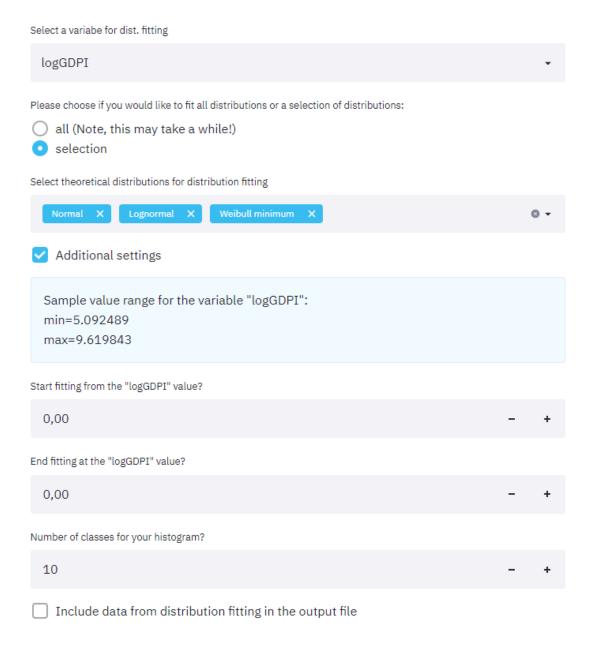


Figure 8 Settings for distribution fitting.

The output of the distribution fitting is given in tabular form and includes the test results and the relative and expected relative frequencies for each class of the best fit distribution.

SSD (sum of squared differences) =
$$\sum_{i=1}^{\# classes} (n_i - n_{expected,i})^2$$

where n_i and $n_{expected,i}$ are the absolute frequency and the expected absolute frequency of the respective distribution of group i.

$$Chi - squared \ (test \ statistic) = \sum\nolimits_{i=1}^{\# \ classes} \frac{\left(n_i - n_{expected,i}\right)^2}{n_{expected,i}}$$

DOF (degrees of freedom) =
$$m_1 - m_2 - 1$$

where m_1 corresponds to the number of classes with non-zero observations, m_2 is the number of parameters needed for the respective distribution.

p-value = corresponding to the test statistic

The test statistic follows a χ^2 distribution.

Moreover, it can be selected whether the data used for the analysis should be included in the output file available for download or not (Figure 8).

3.2.5 Correlation analysis

The correlation coefficient for two variables can be determined by using the method of Pearson, Kendall, or Spearman.

Pearson:

$$r = \frac{Cov(x,y)}{std(x)*std(y)}$$

Kendall:

$$r = \frac{n_c - n_d}{\binom{n}{2}}$$

where n_c is the number of concordant pairs and n_d the number of discordant pairs.

Spearman:

$$r = \frac{Cov(rank(x), rank(y))}{std(rank(x))*std(rank(y))}$$

While Pearson is a parametric method, Kendall and Spearman are non-parametric methods based on the order and ranks of observation pairs.

To perform the correlation analysis, you can select the variables for which the coefficients should be calculated (Figure 9). Please note non-numerical variables will be automatically excluded.

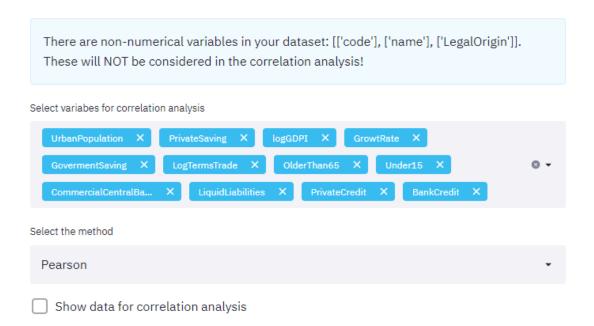


Figure 9 Correlation analysis settings.

3.2.6 Regression techniques

Having two numeric variables, you can use linear regression models to further explore the relationship between the two selected variables. You can choose from different types of linear regressions: Simple Linear Regression, Linear-Log Regression, Log-Linear Regression, Log-Log Regression and Polynomial Regression.

All regression techniques predict the selected response variable by establishing a linear relationship between the explanatory and the response variable based on the chosen technique (Figure 10). The optional detailed output includes ANOVA/ F-test and coefficient estimates. It can also be selected whether the data used for the analysis should be included in the output file available for download or not

| Please choose if you would like to apply all regression techniques or a selection of techniques: | |
|--|-----|
| all | |
| selection | |
| | |
| Select regression techniques | |
| Simple Linear Regres X | O - |
| | |
| Show detailed output per technique? | |
| | |
| ☐ Include data in the output file | |

Figure 10 Settings for Regression techniques.

Simple Linear Regression:

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$
 for $i = 1 \dots n$

Linear-Log Regression:

$$y_i = \beta_0 + \beta_1 \log(x_i) + \varepsilon_i$$
 for $i = 1 ... n$

Log-Linear Regression:

$$\log (y_i) = \beta_0 + \beta_1 x_i + \varepsilon_i \text{ for } i = 1 \dots n$$

Log-Log Regression:

$$\log (y_i) = \beta_0 + \beta_1 \log (x_i) + \varepsilon_i$$
 for $i = 1 \dots n$

Polynomial Regression:

$$y_i = \beta_0 + \beta_1 x_i + \sum_{j=2}^m x_i^j + \varepsilon_i$$
 for $i = 1 \dots n$

where m determines the maximum power of the transformed explanatory variable (Figure 11);

Select regression techniques



Figure 11 Settings for the polynomial regression.

and where

n = number of observations

y = response variable

x = explanatory variables

 $\beta_1 = coefficient of explanatory variable x$

 $\beta_0 = intercept$

 $\varepsilon = error$.

The coefficients are estimated with the Ordinary Least Squares (OLS) estimator. For the OLS estimator to be the Best Linear Unbiased Estimator (BLUE), the following assumptions must be fulfilled:

- (1) $E(\varepsilon_i) = 0 \ \forall i \ (mean \ error \ of \ 0)$
- (2) $Var(\varepsilon_i) = \sigma^2 \forall i \ (homoskedasticity)$
- (3) $Cov(\varepsilon_i, \varepsilon_j) = 0 \ \forall i \neq j \ (uncorrelated \ errors).$

When these assumptions are met, the OLS estimator is BLUE according to the Gauss-Markov theorem.

For the definition and description of the output, please see 3.3.1 Multiple Linear Regression where the same estimator is used.

3.2.7 Contingency tables and association measures

In case of a bivariate or a two-dimensional frequency distribution of two discrete variables, you can perform a contingency analysis to identify the strength of the association between the two variables. In case you want to include a continuous variable, you have the option to either reclassify your data or use your data as is (Figure 12).

You can try some of these options:

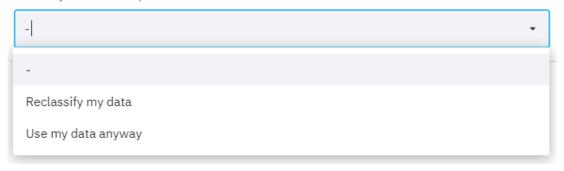


Figure 12 Options for continuous variables in contingency tables and association measures.

The reclassification allows you to specify the range of the data that should be used and the number of created classes (Figure 13).

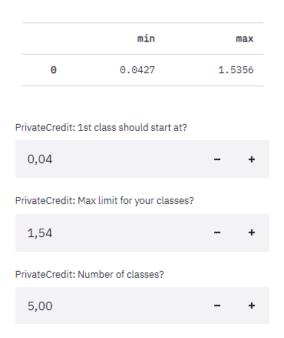


Figure 13 Classification options for continuous variables.

The strength of association is measured with the corrected Pearson contingency coefficient, which is based on the chi-squared value. You can also optionally select to display the marginal frequencies for the contingency analysis, show the used data for the analysis and whether the data should be included in the output file.

$$n_{expected,i} = \frac{marginal\ frequency\ of\ x\ for\ combination\ i*marginal\ frequency\ of\ y\ for\ combination\ i}{n}$$

$$\chi^2 = \sum\nolimits_{i=1}^{\#\,combinations} \frac{\left(n_i - n_{expected,i}\right)^2}{n_{expected,i}}$$

 $p-value = corresponding to \chi^2 - statistic$

$$K = \sqrt{\frac{\chi^2}{\chi^2 + n}}$$

$$K_{cor} = \frac{K}{K_{max}}$$

 $K_{max} = \sqrt{\frac{\min(k,l)-1}{\min(k,l)}}$, where k = #rows and l = #columns of the contingency table

3.3 Multivariate data

Having multivariate data, you can use machine learning models to predict a continuous or binary response variable. With the help of the models, you are also able to identify variable importance and influence directions of the explanatory variables.

To adjust the models to your specific data set, you can select several settings for the included modelling techniques. Moreover, you have the option to tune and validate your models by running hyperparameter-tuning and a model validation, respectively.

3.3.1 Multiple Linear Regression

The modelling technique Multiple Linear Regression can be optionally selected.

Definition

Multiple Linear Regression (MLR) is a regression technique that uses several explanatory variables to predict a response variable by establishing linear relationships between the explanatory variables and the response variable:

$$y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \dots + \beta_k x_{ik} + \varepsilon_i$$
 for $i = 1 \dots n$, or in matrix form $Y = X\beta + \varepsilon$, where

n = number of observations

k = number of explanatory variables

y = response variable

 $x_i = explanatory variables, j = 1 ... k$

 $\beta_i = coefficient of explanatory variable x_i$

 $\beta_0 = intercept$

 $\varepsilon = error$.

The coefficients are estimated with the Ordinary Least Squares (OLS) estimator. For the OLS estimator of the MLR to be the Best Linear Unbiased Estimator (BLUE), the following assumptions must be fulfilled:

(1)
$$E(\varepsilon_i) = 0 \ \forall i \ (mean \ error \ of \ 0)$$

(2)
$$Var(\varepsilon_i) = \sigma^2 \forall i \ (homoskedasticity)$$

(3)
$$Cov(\varepsilon_i, \varepsilon_j) = 0 \ \forall i \neq j \ (uncorrelated \ errors).$$

When these assumptions are met, the OLS estimator is BLUE according to the Gauss-Markov theorem.

Settings

For MLR, it can be selected whether an intercept should be included in the model or not and which covariance type should be used (Figure 14). If the homoskedasticity assumption is violated (see Output for heteroskedasticity tests), estimations for the standard error of coefficients and consequently for the statistical tests are unreliable. Besides a non-robust covariance matrix Σ , you can select a robust covariance matrix Σ (HC = heteroskedasticity-consistent) to derive more reliable conclusions from statistical tests.

Multiple Linear Regression settings Adjust settings for Multiple Linear Regression Include intercept Yes Covariance type non-robust HC0 HC1 HC2 HC3

Figure 14 Settings for Multiple Linear Regression including intercept and covariance estimator choices.

Non-robust

$$\Sigma = \sigma^2 (X^T X)^{-1}$$

HC0

$$\Sigma = (X^{T}X)^{-1}X^{T} diag((y_{i} - \hat{y}_{i})^{2}) X(X^{T}X)^{-1}$$

<u>HC1</u>

$$\Sigma = \frac{n}{n-k-1} (X^T X)^{-1} X^T \operatorname{diag}((y_i - \hat{y}_i)^2) X (X^T X)^{-1} (\operatorname{degree\ of\ freedom\ adjustment})$$

HC2

$$\Sigma = (X^T X)^{-1} X^T \operatorname{diag}\left(\frac{(y_i - \hat{y}_i)^2}{1 - h_{ii}}\right) X (X^T X)^{-1} \left(\operatorname{leverage adjustment}\right)$$

HC3

$$\Sigma = (X^T X)^{-1} X^T \operatorname{diag}\left(\frac{(y_i - \hat{y}_i)^2}{(1 - h_{ii})^2}\right) X(X^T X)^{-1} \text{ (squared leverage adjustment)}$$

where

 $\hat{y}_i = prediction of the model for y_i$

 $diag((y_i - \hat{y}_i)^2) = diagonal\ matrix\ with\ squared\ residuals\ on\ the\ diagonal$ $h_{ii} = x_i(X^TX)^{-1}x_i^T$, which are the diagonal values (leverage values) of the hat matrix $X(X^TX)^{-1}X^T$.

Output

Regression statistics

$$R^2 = 1 - \frac{\sum_i (y_i - \hat{y}_i)^2}{\sum_i (y_i - \bar{y})^2} = \frac{\sum_i (\hat{y}_i - \bar{y})^2}{\sum_i (y_i - \bar{y})^2}$$

$$Adj. R^2 = 1 - \frac{(1-R^2)(n-1)}{n-k-1}$$

Multiple correlation coefficient = $\sqrt{R^2}$

Residual standard error =
$$\sqrt{\frac{\sum_{i}(y_{i}-\hat{y}_{i})^{2}}{n-k-1}}$$

$$Log \; likelihood = \max_{\beta} \log \left(\prod_{i} \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left(\frac{-(y_i - x_i \beta)^2}{2\sigma^2} \right) \right)$$

AIC (Akaike Information Criterion) = $2(k + 1) - 2 \log likelihood$ BIC (Bayesian Information Criterion) = $(k + 1) \ln(n) - 2 \log likelihood$

ANOVA

Explained Sum of Squares (ESS) =
$$\sum_{i} (\hat{y}_i - \bar{y})^2$$

Residual Sum of Squares (SSR) =
$$\sum_{i} (y_i - \hat{y}_i)^2$$

Total Sum of Squares (SST) =
$$\sum_{i} (y_i - \bar{y})^2$$

Mean Sum of Squares (MS) =
$$\frac{Sum \ of \ Squares}{DF}$$

$$F \ statistic \ (non-robust) = \frac{MS_{explained}}{MS_{residual}}$$

 $F \ statistic \ (robust) = \ Wald \ statistic$

The test statistic follows a F-distribution.

Coefficients

 $Coefficients = \beta$ estimates of the OLS estimator

Standard error = diagonal of the covariance matrix Σ

$$t-statistic = \frac{coefficient}{standard\ error}$$
 $p-value = corresponding\ to\ t-statistic$
 $lower\ 95\%\ confidence\ pprox\ coefficient\ -1.96*standard\ error$
 $upper\ 95\%\ confidence\ pprox\ coefficien\ +1.96*standard\ error$

Heteroskedasticity test

Breusch-Pagan test: No interactions are included. Estimation of γ for (e.g., with two explanatory variables)

$$(y - \hat{y})^2 = \gamma_0 + \gamma_1 x_1 + \gamma_2 x_2 + \varepsilon$$

White test (without interactions): No interactions, but squared explanatory variables are added. Estimation of γ for (e.g., with two explanatory variables)

$$(y - \hat{y})^2 = \gamma_0 + \gamma_1 x_1 + \gamma_2 x_2 + \gamma_3 x_1^2 + \gamma_4 x_2^2 + \varepsilon$$

White test (with interactions): Squared explanatory variables and interactions of the explanatory variables are added. Estimation of γ for (e.g., with two explanatory variables)

$$(y - \hat{y})^2 = \gamma_0 + \gamma_1 x_1 + \gamma_2 x_2 + \gamma_3 x_1^2 + \gamma_4 x_2^2 + \gamma_5 x_1 x_2 + \varepsilon$$

For the tests, the following test statistic is used:

 $Test\ statistic = nR^2$.

The test statistic follows a χ^2 distribution.

Variable importance (via permutation)

For the determination of the permutation based variable importance the data set is shuffled 10 times for each explanatory variable. R² is used as scoring metric to determine the importance:

$$importance = score_{full \; model} - \frac{_1}{_{10}} \; \textstyle \sum_{i=1}^{10} score_i.$$

Please note that the importance might be misleading for strongly correlated explanatory variables.

Analytical plots

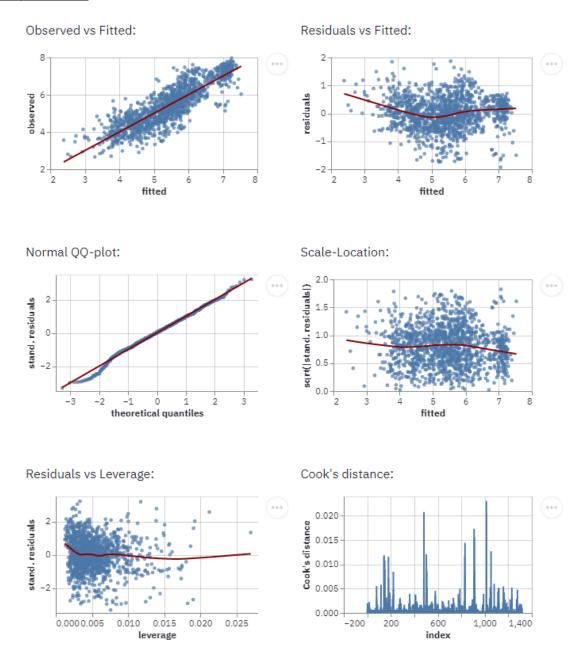


Figure 15 Graphical output of Multiple Linear Regression.

The leverage in Figure 15 corresponds to the diagonal values of the hat matrix and Cook's distance is based on a re-calculated regression where observation i was removed:

$$Cook's\ distance_i = \frac{\sum_{j=1, j\neq i}^n (\hat{y}_j - \hat{y}_{j,without\ obs.i})^2}{(k+1)\sigma^2}.$$

3.3.2 Logistic Regression

The modelling technique Logistic Regression can be optionally selected if the response variable is binary.

Definition

Logistic Regression (LR) is a regression technique that models the probability of occurrence for a response variable with two realizations. Modelling the probability is approached by at first formulating the regression problem using logit:

$$logit(y_{0/1}) = log(\frac{P(y=1)}{1 - P(y=1)}) = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_k x_k = X\beta$$

where

k = number of explanatory variables

y = binary response variable

 $x_i = explanatory\ variables, j = 1...k$

 $\beta_i = coefficient \ of \ explanatory \ variable \ x_i$

 $\beta_0 = intercept.$

After transforming the above equation, the probability for P(y = 1) can be estimated

$$P(y=1) = \frac{1}{1 + exp(-(\beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_k x_k))} = \frac{1}{1 + exp(-X\beta)}.$$

The coefficients are determined with the Maximum Likelihood Estimator (MLE).

Settings

For LR, it can be selected whether an intercept should be included in the model or not and which covariance type should be used (Figure 16). Currently, only "HC0" is available.

Logistic Regression settings

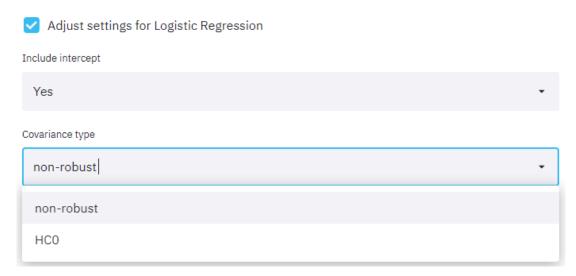


Figure 16 Settings for Logistic Regression including intercept and covariance estimator choices.

Output

Regression statistics

AUC ROC = area under the receiver operating characteristic curve

Pseudo
$$R^2 = 1 - \frac{LLF}{LL-Null}$$

$$Log \ likelihood \ (LLF) = \max_{\beta} \sum\nolimits_{i} y_{i} \log \left(\frac{1}{1 + \exp\left(-(x_{i}\beta)\right)}\right) + (1 - y_{i}) \log \left(1 - \frac{1}{1 + \exp\left(-(x_{i}\beta)\right)}\right)$$

 $LL-Null = Log\ likelihood\ of\ the\ constant-only\ model$

 $Residual\ deviance = -2\ LLF$

 $Null\ deviance = -2\ LL - Null$

 $LLR (Log \ likelihood \ ratio) = -2 (LL - Null - LLF)$

 $LLR \ p-value = \chi^2 \ probability \ corresponding \ to \ LLR$

AIC (Akaike Information Criterion) = 2(k + 1) - 2 Log likelihood

BIC (Bayesian Information Criterion) = $(k + 1) \ln(n) - 2 \log likelihood$

Coefficients

Coefficients = β estimates of the MLE

Standard error = diagonal of the covariance matrix Σ

$$t - statistic = \frac{coefficient}{standard\ error}$$

p-value = corresponding to t - statistic

lower 95% confidence \approx coefficient -1.96 * standard error

upper 95% *confidence* ≈ *coefficien* + 1.96 * *standard error*

Variable importance (via permutation)

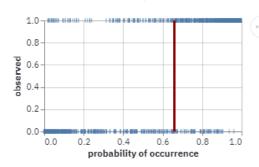
For the determination of the permutation based variable importance the data set is shuffled 10 times for each explanatory variable. AUC ROC is used as scoring metric to determine the importance:

 $importance = score_{full\ model} - \frac{1}{10} \sum_{i=1}^{10} score_i.$

Please note that the importance might be misleading for strongly correlated explanatory variables.

Analytical plots

Observed vs. Probability of Occurrence:



ROC curve:

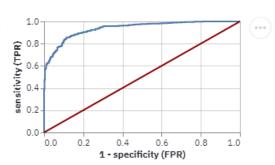


Figure 17 Graphical output of Logistic Regression.

In Figure 17 the threshold corresponds to the Youden Index:

$$threshold = \max_{t \in thresholds} \left(sensitivity_t + specificity_t - 1 \right).$$

Partial probability plots

Partial probability plots are based on univariate LR modelling with each selected explanatory variable.

3.3.3 Generalized Additive Models

The modelling technique Generalized Additive Models can be optionally selected.

Definition

Generalized Additive Models (GAMs) are flexible and additive modelling techniques for identifying non-linear relationships between the response and explanatory variables by using smoothing functions, e.g., based on splines. GAMs have the following general form:

$$g(E(Y|X)) = \beta_0 + f_1(X_1) + f_2(X_2) + \dots + f_k(X_k)$$

where

k = number of explanatory variables

Y = response variable

 $X_i = explanatory variables, j = 1 ... k$

 $f_i = smooth\ feature\ function\ for\ X_i$

g = link function

 $\beta_0 = intercept$

E(Y|X) = expected value of Y.

Y follows an exponential family distribution, here either normal or binomial for continuous and binomial response variables, respectively. The link function is defined accordingly, such that for normal distributions the identity function and for binomial distributions the logit function is used:

$$E(Y|X) = \beta_0 + f_1(X_1) + f_2(X_2) + \dots + f_k(X_k)$$

or

$$logit(E(Y_{0/1}|X)) = log(\frac{P(Y=1|X)}{P(Y=0|X)}) = \beta_0 + f_1(X_1) + f_2(X_2) + \dots + f_k(X_k).$$

The smooth feature functions f_j are based on penalized B-splines, for which several parameters can be set. The smoothing parameter estimation problem is solved by using the generalized cross-validation (GCV) and un-biased risk estimator (UBRE) score for normal and binomial distributions, respectively.

Settings

For each feature function, the spline settings can be specified, and it can be decided whether an intercept should be included or not (Figure 18). The number of splines or basis functions used for the estimation of the feature function, the spline order, which determines the degree of the piecewise polynomial functions (degree = spline order - 1), and the lambda value

controlling for the penalty and thus overfitting can be set manually. For the splines, a second derivative smoothing is used. Please note that lambda controls for the bias-variance trade-off and higher values of lambda will lead to smoother functions.

Generalized Additive Models settings Adjust settings for Generalized Additive Models Include intercept Yes Search for lambda No Number of splines (Social support) Spline order (Social support) Lambda (Social support) 20 0,600 Spline order (Generosity) Number of splines (Generosity) Lambda (Generosity) 20 3 0,600

Figure 18 Settings for Generalized Additive Models including intercept and spline parameter choices.

Additionally, a search for the best lambda value can be conducted, which either aims at minimizing GCV or UBRE depending on the response variable.

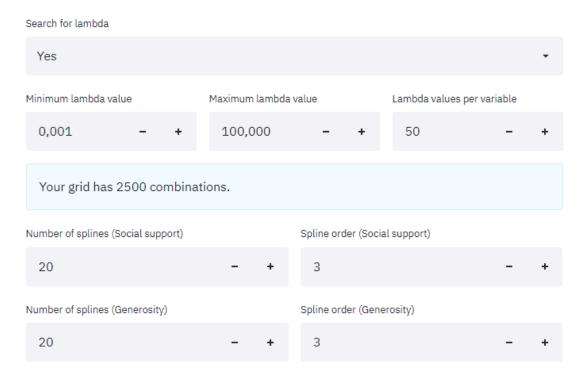


Figure 19 Lambda search settings.

For the lambda search, the interval boundaries for possible lambda values and the number of values taken from this interval for each selected explanatory variable can be specified (Figure 19). The values taken from the interval are evenly spaced. Besides the settings for the lambda search, further spline parameters can still be specified for each explanatory variable.

Output (continuous response variable)

Regression statistics

Log likelihood

AIC (Akaike Information Criterion) = 2(Effective DF + 1) - 2 Log likelihood

$$AICc \ (second \ order \ AIC) = AIC + \frac{2 \ (Effective \ DF+1)^2 + 2 \ (Effective \ DF+1)}{n - (Effective \ DF+1) - 1}$$

GCV (Generalized Cross – Validation score) = minimized GCV score of fitted GAM

Scale = residual standard error squared

 $Pseudo R^2 = deviance explained$

Feature significance

Coefficients = *shown if intercept included*

 $Lambda = final\ lambda\ value$

rank = number of splines

edof = effective degrees of freedom after regularization

p-value = significance of corresponding feature

Variable importance (via permutation)

For the determination of the permutation based variable importance the data set is shuffled 10 times for each explanatory variable. R² is used as scoring metric to determine the importance:

$$importance = score_{full\ model} - \frac{1}{10} \sum_{i=1}^{10} score_i.$$

Please note that the importance might be misleading for strongly correlated explanatory variables.

Partial dependence plots (one-way)

The partial dependence plots display the contribution of each feature to the overall prediction while keeping the other predictors constant and the 95% confidence intervals.

Output (binary response variable)

Regression statistics

 $UBRE\ (Un-Biased\ Risk\ Estimator)=minimized\ UBRE\ score\ of\ fitted\ GAM$ Scale=1

 $AUC\ ROC = area under the receiver operating characteristic curve$

Variable importance (via permutation)

For the determination of the permutation based variable importance the data set is shuffled 10 times for each explanatory variable. AUC ROC is used as scoring metric to determine the importance:

$$importance = score_{full\ model} - \frac{1}{10} \sum_{i=1}^{10} score_i.$$

Please note that the importance might be misleading for strongly correlated explanatory variables.

Analytical plots

See Output for Logistic Regression.

Partial dependence plots (one-way)

See Output (continuous response variable) for Generalized Additive Models.

3.3.4 Random Forest

The modelling technique Random Forest can be optionally selected.

Definition

Random Forest (RF) is a tree-based approach that combines several decision trees. The combination of the single trees creates the forest, where each single tree is built based on a bootstrap sample, i.e., a sample consisting of randomly drawn observations with replacement. The single decision trees either use a (randomly selected) subset of the included explanatory variables or all of them for splitting at each node. Thus, two sources of randomness can be included into this modelling technique, which aims at decreasing the variance of the model. The quality of the split is measured by the mean squared error and Gini impurity for continuous and binary response variables, respectively. The prediction of the random forest model is determined by taking the average of the predictions of all trees. Consequently, for binary response variables the probability predictions are averaged.

Hyperparameters

For RF models, parameters controlling the size of the forest and the characteristic of each single tree can be specified. Here, the number of trees, maximum tree depth, maximum number of features and sample rate can be tuned (Figure 20).

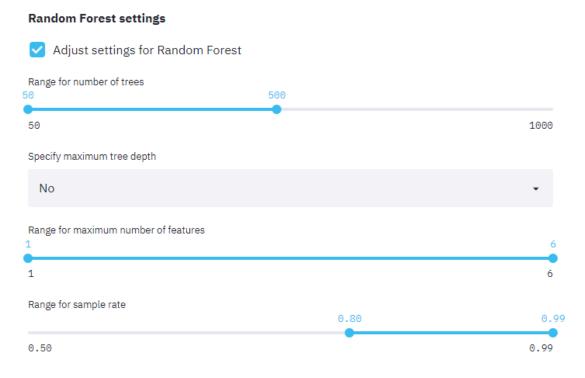


Figure 20 Range options for hyperparameters of Random Forest.

If the maximum tree depth is not specified, no restriction on the maximum tree depth is used and if only one explanatory variable is included, the range for maximum number of features, i.e. the number of features to consider at a split, cannot be specified.

For the search methods grid-search and random grid-search, a grid based on the ranges for each hyperparameter must be created. The grid is formed by using the minimum and maximum of the specified range (Figure 20) and a pre-defined step size (Table 1).

| Hyperparameter | Default value | Range | Default range | Step size |
|----------------------------|---------------|--------------|---------------|-----------|
| Number of trees | 100 | 50 – 1000 | 50 – 500 | 5 |
| Maximum tree depth | None | 1 – 50 | None/ 2 – 10 | 1 |
| Maximum number of features | #features | 1 – #featues | 1 – #featues | 1 |
| Sample rate | 0.99 | 0.5 – 0.99 | 0.8 - 0.99 | 0.05 |

Table 1 Hyperparameter specifications for Random Forest.

Please note that for the search methods Bayes optimization and sequential model-based optimization the definition of step sizes is not necessary as the next hyperparameter combination is automatically proposed within the specified ranges (see 3.3.7 Hyperparameter-tuning).

Output (continuous response variable)

Regression information

OOB (out of bag) score = estimation of the generalization score (R^2)

Regression statistics

Mean squared error (MSE) =
$$\frac{1}{n} \sum_{i} (y_i - \hat{y}_i)^2$$

Root mean squared error (RMSE) =
$$\sqrt{\frac{1}{n} \sum_{i} (y_i - \hat{y}_i)^2}$$

Mean absolute error (MAE) =
$$\frac{1}{n} \sum_{i} |y_i - \hat{y}_i|$$

Residual standard error =
$$\sqrt{\frac{\sum_{i}(y_{i} - \hat{y}_{i})^{2}}{n-k-1}}$$

Variable importance (via permutation)

For the determination of the permutation based variable importance the data set is shuffled 10 times for each explanatory variable. R² is used as scoring metric to determine the importance:

$$importance = score_{full\ model} - \frac{1}{10} \sum_{i=1}^{10} score_i.$$

Please note that the importance might be misleading for strongly correlated explanatory variables.

Variable importance (impurity-based)

The impurity-based variable importance is a regression tree specific feature. The importance

corresponds to the (normalized) total reduction of the used splitting criterion. The sum of the importance of each single explanatory variable is one. Please note that impurity-based variable importance can be misleading for high cardinality features, i.e., features that have many unique values. As alternative the permutation-based variable importance can be considered.

Partial dependence plots (one-way)

The partial dependence plots can be interpreted as the expected change in the response variable along the gradient of the considered explanatory variable. Here, the 'brute' method over the whole range of the respective gradient with a grid resolution of 100 was used.

Output (binary response variable)

Regression information

OOB (out of bag) score = estimation of the generalization <math>score (mean accuracy)

Regression statistics

AUC ROC = area under the receiver operating characteristic curve

$$AP = \sum_{m} (R_m - R_{m-1})P_m$$
 , $R_m = recall \ and \ P_m = precision \ at \ m^{th} \ threshold$

 $AUC\ PRC = area\ under\ the\ precision - recall\ curve$

$$LOG - LOSS = \frac{\text{Log likelihood}}{n}$$

Variable importance (via permutation)

For the determination of the permutation based variable importance the data set is shuffled 10 times for each explanatory variable. AUC ROC is used as scoring metric to determine the importance:

$$importance = score_{full\ model} - \frac{1}{10} \sum_{i=1}^{10} score_i.$$

Please note that the importance might be misleading for strongly correlated explanatory variables.

Variable importance (impurity-based)

See Output (continuous response variable) for Random Forest.

Analytical plots

See Output for Logistic Regression.

Partial dependence plots (one-way)

See Output (continuous response variable) for Random Forest.

3.3.5 Boosted Regression Trees

The modelling technique (gradient) Boosted Regression Trees can be optionally selected.

Definition

Boosted Regression Trees (BRT) is a tree-based machine learning technique. For evaluating the learning of the BRT model, a loss function is defined. Here, the least squares and deviance are implemented as loss functions for continuous and binary response variables, respectively. During learning, a certain number of trees are sequentially fitted on the negative gradient of the given loss function, with each tree being typically built on a bootstrap data set (Figure 21).

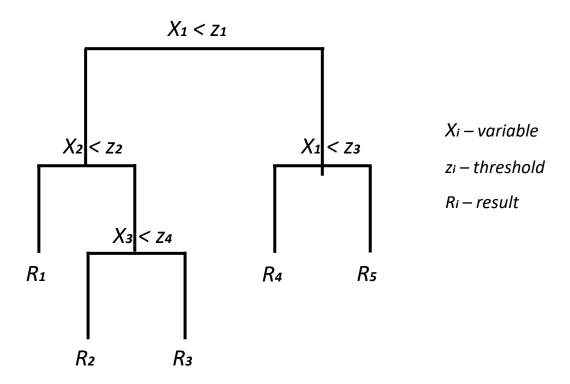


Figure 21 Example of a decision tree $f^t(x)$ based on a bootstrap data set.

The structure of the trees can be defined by the number of nodes or the maximum tree depth and the influence of each tree in the final model can be controlled through a learning rate:

$$f(x) = \sum_{t=1}^{N_t} \lambda f^t(x)$$
, $\lambda = learning \ rate$, $N_t = number \ of \ trees$.

Hyperparameters

In BRT models, the parameters number of trees, learning rate, maximum tree depth and the sample rate for building the single trees can be tuned (Figure 22). Please note that a small learning rate leads to a slow learner, which tends to perform well, and that larger values for the maximum tree depth can imply variable interactions.

Boosted Regression Trees settings



Figure 22 Range options for hyperparameters of Boosted Regression Trees.

For the search methods grid-search and random grid-search, a grid based on the ranges for each hyperparameter must be created. The grid is formed by using the minimum and maximum of the specified range (Figure 22) and a pre-defined step size (Table 2).

| Hyperparameter | Default value | Range | Default range | Step size |
|--------------------|---------------|-------------|---------------|-----------|
| Number of trees | 100 | 50 – 1000 | 50 – 500 | 5 |
| Learning rate | 0.1 | 0.001 – 0.1 | 0.001 - 0.02 | 0.001 |
| Maximum tree depth | 3 | 1 – 30 | 2 – 10 | 1 |
| Sample rate | 1 | 0.5 – 1.0 | 0.8 – 1.0 | 0.05 |

Table 2 Hyperparameter specifications for Boosted Regression Trees.

Please note that for the search methods Bayes optimization and sequential model-based optimization the definition of step sizes is not necessary as the next hyperparameter combination is automatically proposed within the specified ranges (see 3.3.7 Hyperparameter tuning).

Output (continuous response variable)

Regression statistics

Mean squared error (MSE) =
$$\frac{1}{n} \sum_{i} (y_i - \hat{y}_i)^2$$

Root mean squared error (RMSE) =
$$\sqrt{\frac{1}{n} \sum_{i} (y_i - \hat{y}_i)^2}$$

Mean absolute error (MAE) =
$$\frac{1}{n} \sum_{i} |y_i - \hat{y}_i|$$

Residual standard error =
$$\sqrt{\frac{\sum_{i}(y_{i}-\hat{y}_{i})^{2}}{n-k-1}}$$

Variable importance (via permutation)

For the determination of the permutation based variable importance the data set is shuffled 10 times for each explanatory variable. R² is used as scoring metric to determine the importance:

$$importance = score_{full\ model} - \frac{1}{10} \sum_{i=1}^{10} score_i.$$

Please note that the importance might be misleading for strongly correlated explanatory variables.

Variable importance (impurity-based)

The impurity-based variable importance is a regression tree specific feature. The importance corresponds to the (normalized) total reduction of the used splitting criterion. The sum of the importance of each single explanatory variable is one. Please note that impurity-based variable importance can be misleading for high cardinality features, i.e., features that have many unique values. As alternative the permutation-based variable importance can be considered.

Partial dependence plots (one-way)

The partial dependence plots can be interpreted as the expected change in the response variable along the gradient of the considered explanatory variable. Here, the 'brute' method over the whole range of the respective gradient with a grid resolution of 100 was used.

Output (binary response variable)

Regression statistics

AUC ROC = area under the receiver operating characteristic curve

$$AP = \sum_{m} (R_m - R_{m-1}) P_m$$
 , $R_m = recall \ and \ P_m = precision \ at \ m^{th} \ threshold$

 $AUC\ PRC = area\ under\ the\ precision - recall\ curve$

$$LOG - LOSS = \frac{\text{Log likelihood}}{n}$$

Training score

$$Deviance = 2 LOG - LOSS$$

Variable importance (via permutation)

For the determination of the permutation based variable importance the data set is shuffled 10 times for each explanatory variable. AUC ROC is used as scoring metric to determine the

importance:

 $importance = score_{full\ model} - \frac{1}{10} \sum_{i=1}^{10} score_i.$

Please note that the importance might be misleading for strongly correlated explanatory variables.

Variable importance (impurity-based)

See Output (continuous response variable) for Boosted Regression Trees.

Analytical plots

See Output for Logistic Regression.

Partial dependence plots (one-way)

See Output (continuous response variable) for Boosted Regression Trees.

3.3.6 Artificial Neural Networks

The modelling technique Artificial Neural Networks can be optionally selected.

Definition

Artificial Neural Networks (ANNs) is a machine learning technique that resembles the biological neural system, i.e., ANNs include neurons with a specified number of layers that are linked by so-called activation functions (Figure 23).

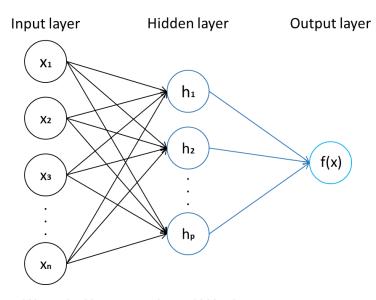


Figure 23 Artificial Neural Network with p neurons in one hidden layer.

The single observations x_i are linearly weighted, adjusted with a bias value and afterwards transformed by the activation function $f_{activation}$. For example, for the first neuron in the hidden layer we get

$$h_1 = \, f_{activation}(w_1x_1 + w_2x_2 + \cdots + w_nx_n + bias), \label{eq:h1}$$

so that h_1 is now the input for the next layer. When all neurons of the last hidden layer have been calculated, the values of these neurons are transformed one last time by the output function f_{output} to get the wanted output/ prediction. For continuous and binary response variables the identity and logistic function are used as output functions, respectively.

The learning of the ANN consists of updating the weights, i.e., the weight optimization, by using a specific solver algorithm, e.g., gradient-based optimization, and several different hyperparameters. For evaluating the learning of the ANN model, the least squares and log-loss are implemented as loss functions for continuous and binary response variables, respectively.

Please note that Artificial Neural Networks are sensitive to different scales of the data. Therefore, the data is re-scaled by standardization, such that the mean is 0 and the variance

is 1. In case of testing during hyperparameter-tuning or model validation, the same scaling is applied to the test data set.

Hyperparameters

In ANN models, the parameters weight optimization solver, maximum number of iterations, activation function, hidden layer sizes, learning rate and L² regularization can be tuned (Figure 24).

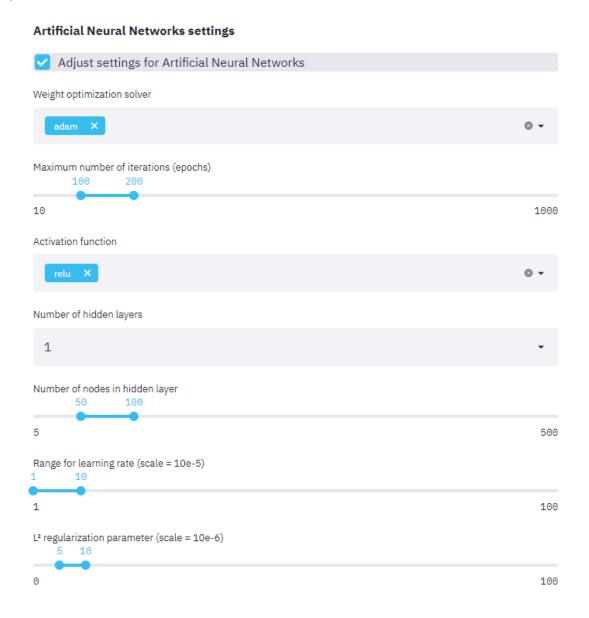


Figure 24 Range options for hyperparameters of Artificial Neural Networks.

Please note that, for the solver 'adam', the maximum number of iterations determines the number of epochs, i.e., it defines how often the data will be seen by the model. Also, the learning rate, which is used to control the weight updates throughout weight optimization, is only used for the solver 'adam'. L² regularization is used to control for the weight size and thus penalizing complex models and controlling overfitting.

For the search methods grid-search and random grid-search, a grid based on the ranges for each hyperparameter must be created. The grid is formed by using the minimum and maximum of the specified range (Figure 24) and a pre-defined step size (Table 3). For the number of hidden layers and the corresponding nodes all combinations are considered.

| Hyperparameter | Default value | Range | Default range | Step size |
|-------------------------------|---------------|-----------------------------------|------------------|-----------|
| Solver | adam | adam | adam | - |
| Maximum iterations | 200 | 10 – 1000 | 100 – 200 | 5 |
| Activation function | relu | relu, identity, logistic, tanh | relu | - |
| Hidden layers | 1 | 1 – 3 | 1 | - |
| Nodes per hidden layer | 100 | 5 – 500 | 50 – 100 | - |
| Learning rate | 0.001 | 0.0001 - 0.01 | 0.0001 - 0.002 | 0.0005 |
| L ² regularization | 0.0001 | 0 – 0.001 | 0.00001 - 0.0002 | 0.00005 |

Table 3 Hyperparameter specifications for Artificial Neural Networks.

Please note that for the search methods Bayes optimization and sequential model-based optimization the definition of step sizes is not necessary as the next hyperparameter combination is automatically proposed within the specified ranges (see 3.3.7 Hyperparameter-tuning).

Output (continuous response variable)

Regression statistics

Mean squared error (MSE) =
$$\frac{1}{n} \sum_{i} (y_i - \hat{y}_i)^2$$

Root mean squared error (RMSE) =
$$\sqrt{\frac{1}{n} \sum_{i} (y_i - \hat{y}_i)^2}$$

Mean absolute error (MAE) =
$$\frac{1}{n} \sum_{i} |y_i - \hat{y}_i|$$

Residual standard error =
$$\sqrt{\frac{\sum_{i}(y_i - \hat{y}_i)^2}{n-k-1}}$$

Best loss = loss from loss cruve

Loss curve

Loss value for each iteration.

Variable importance (via permutation)

For the determination of the permutation based variable importance the data set is shuffled

10 times for each explanatory variable. R² is used as scoring metric to determine the importance:

$$importance = score_{full\ model} - \frac{1}{10} \sum_{i=1}^{10} score_i.$$

Please note that the importance might be misleading for strongly correlated explanatory variables.

Partial dependence plots (one-way)

The partial dependence plots can be interpreted as the expected change in the response variable along the gradient of the considered explanatory variable. Here, the 'brute' method over the whole range of the respective gradient with a grid resolution of 100 was used.

Output (binary response variable)

Regression statistics

AUC ROC = area under the receiver operating characteristic curve

$$AP = \sum_{m} (R_m - R_{m-1}) P_m$$
 , $R_m = recall \ and \ P_m = precision \ at \ m^{th} \ threshold$

 $AUC \ PRC = area \ under \ the \ precision - recall \ curve$

$$LOG - LOSS = \frac{\text{Log likelihood}}{n}$$

Best loss = loss from loss cruve

Loss curve

Loss value for each iteration. Please note that only for the solver 'adam' a loss curve is displayed.

Variable importance (via permutation)

For the determination of the permutation based variable importance the data set is shuffled 10 times for each explanatory variable. AUC ROC is used as scoring metric to determine the importance:

$$importance = score_{full\ model} - \frac{1}{10} \sum_{i=1}^{10} score_i.$$

Please note that the importance might be misleading for strongly correlated explanatory variables.

Analytical plots

See Output for Logistic Regression.

Partial dependence plots (one-way)

See Output (continuous response variable) for Artificial Neural Networks.

3.3.7 Hyperparameter-tuning

Hyperparameter-tuning is an optional setting. It can be conducted for Random Forest, Boosted Regression Trees and Artificial Neural Networks.

Definition

The parameters that influence the learning of a machine learning method are called hyperparameters. For these hyperparameters, different values are worth considering to get the best model possible. However, finding the specific values is a great challenge. Here, hyperparameter-tuning tries to give a solution to this problem.

The aim of hyperparameter-tuning is to find the best possible set of hyperparameters for a modelling technique across a broad range of possible combinations of values called the parameter space/ grid. The best parameter set from the space/ grid is the one that yields the highest R² or AUC for continuous or binary variables, respectively, or the best value for the corresponding objective function. Under hyperparameter-tuning settings you can adjust the ranges for single hyperparameters and thus construct the parameter space/ grid from which the best parameter set will be selected.

Search methods

There are several strategies for finding the best hyperparameter combination across the parameter space/ grid. You can select from the classical grid-search, random grid-search, Bayesian optimization and sequential model-based optimization using decision trees.

Grid-search

Full search across all hyperparameter combinations defined by the parameter grid.

Random grid-search

Search across randomly selected hyperparameter combinations. The selected number of iterations determines how many combinations are considered.

Bayes optimization

Bayesian hyperparameter search using Gaussian process regression to approximate the objective function in order to propose the next hyperparameter combination. The selected number of iterations determines how many combinations are considered.

Sequential model-based optimization

Sequential hyperparameter search using decision trees to approximate the objective function in order to propose the next hyperparameter combination. The selected number of iterations determines how many combinations are considered.

Tuning process

For the hyperparameter-tuning the underlying data set is randomly split into training and test data. The training and test data consist of 80% and 20% of the full data set, respectively. Only the training data is used for tuning the model and finding the best possible hyperparameter combination. For each proposed hyperparameter combination, cross-validation is performed to determine the performance. The number of folds for the cross-validation can be set under hyperparameter-tuning settings (Figure 25). If the response variable is binary, a stratified cross-validation is conducted. By selecting the number of iterations, you can limit the number of hyperparameter combinations that are considered during the tuning for random grid-search, Bayesian optimization and sequential model-based optimization (Figure 25). Please note that high numbers for iterations will result in searches that take a lot of time. After determining the best model either by using mean % VE (percentage of explained variance, continuous) or AUC (binary) from the cross-validation runs, the model is tested on the unseen test data.

Hyperparameter-tuning Yes WARNING: Hyperparameter-tuning can take a lot of time! Hyperparameter-search method random grid-search Select number for n-fold cross-validation 5 2 10 Select number of iterations for search 20 1000

Figure 25 Hyperparameter-tuning settings.

Output

The results of the hyperparameter-tuning are shown in tabular form and consist of the final hyperparameter set for each modelling technique and the tuning performance, including the utilized scoring metric (R² or AUC), the number of iterations/ considered models, the mean performance across the cross-validation runs and the corresponding standard deviation, and the performance on the test data. Please note that the performance score on the test data is

usually smaller than the mean performance from the cross-validation, since the model makes predictions for unseen data.

3.3.8 Model comparison

Different performance metrics for model comparisons are provided for continuous and binary response variables.

Continuous response variable

% VE (proportion of variance explained) =
$$1 - \frac{\sum_{i}(y_i - \hat{y}_i)^2}{\sum_{i}(y_i - \bar{y})^2}$$

Mean squared error (MSE) =
$$\frac{1}{n} \sum_{i} (y_i - \hat{y}_i)^2$$

Root mean squared error (RMSE) =
$$\sqrt{\frac{1}{n} \sum_{i} (y_i - \hat{y}_i)^2}$$

Mean absolute error (MAE) =
$$\frac{1}{n} \sum_{i} |y_i - \hat{y}_i|$$

$$MaxErr (maximum \ residual \ error) = \max_{i} (|y_i - \hat{y}_i|)$$

EVRS (explained variance regression score) =
$$1 - \frac{Var\{y - \hat{y}\}}{Var\{y\}}$$

SSR (sum of squared residuals) =
$$\sum_{i} (y_i - \hat{y}_i)^2$$

Binary response variable

Threshold-independent metrics

AUC ROC = area under the receiver operating characteristic curve

$$AP = \sum_{m} (R_m - R_{m-1}) P_m$$
 , $R_m = recall \ and \ P_m = precision \ at \ m^{th} \ threshold$

 $AUC\ PRC = area\ under\ the\ precision - recall\ curve$

$$LOG - LOSS = \frac{\text{Log likelihood}}{n}$$

Threshold

$$threshold \ (Yoden \ Index) = \max_{t \in thresholds} (sensitivity_t + specificity_t - 1) \, .$$

Threshold-dependent metrics

$$\mathit{TPR}$$
 (true positive rate, sensitivity, recall) = $\frac{\mathit{TP}}{\mathit{TP}+\mathit{FN}}$

$$FNR$$
 (false negative rate) = $\frac{FN}{TP+FN}$

TNR (true negative rate, specificity) =
$$\frac{TN}{TN+FP}$$

$$FPR (false positive rate) = \frac{FP}{TN+FP}$$

$$TSS$$
 (true skill statistic) = $sensitivity + specificity - 1$

$$PREC (precision) = \frac{TP}{TP+FP}$$

$$F1 (F1 \, score) = \frac{2*precision*recall}{precision+recall}$$

$$KAPPA (Cohen's Kappa) = \frac{\frac{TP+TN}{n} - \left(\frac{TP+FP}{n} * \frac{TP+FN}{n} + \frac{FN+TN}{n} * \frac{FP+TN}{n}\right)}{1 - \left(\frac{TP+FP}{n} * \frac{TP+FN}{n} + \frac{FN+TN}{n} * \frac{FP+TN}{n}\right)}$$

$$ACC (accuracy) = \frac{TP + FN}{n}$$

$$BAL\ ACC\ (balanced\ accuracy)=\frac{TPR+FNR}{2}$$

3.3.9 Model predictions

Model predictions are always provided for the original data (default or uploaded in the sidebar menu) and optionally also for newly uploaded data (Figure 26), i.e., if you want to make predictions based on the calibrated model using the original data. In case of variable transformations in 3.1.4 Data processing, STATY automatically checks whether the transformed variables are included in the new data set and conducts the transformation accordingly. Rows with missing values in your new data are also automatically deleted.

Model predictions Use model prediction for new data Yes Drag and drop file here Limit 200MB per file • CSV, TXT WHR_2021.csv 129.0KB X Loading data... done! All variables are available for predictions! WARNING: Your new data set includes NAs. Rows with NAs are automatically deleted!

Figure 26 Model prediction settings for new data.

3.3.10 Model validation

Model validation is an optional setting. If model validation is used, it will be performed for all selected modelling techniques.

Definition

The aim of model validation is to evaluate a trained model with a test data set and thus test the generalization ability. Prior to the evaluation the model is calibrated with a separate training data set and with final hyperparameters, if hyperparameter-tuning was conducted, or default hyperparameters. Afterwards, the prediction performance of the model for the test data set is determined.

Validation process

For randomly splitting the data into a training and test data set, you can select the fraction for the training data size (Figure 27). The respective modelling technique will be trained with the training data set and afterwards evaluated on the test data set. This process will be repeated according to the number for validation runs (Figure 27). To assure comparability among the selected modelling techniques, the training and test data will be the same throughout the single validation runs.

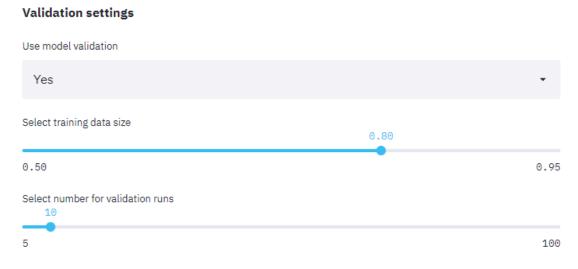


Figure 27 Validation settings.

Output

The results of the model validation for the selected modelling techniques are shown in tabular and graphical form containing the main performance measures (mean and standard deviation), boxplots of the residual distribution and the % of the explained variance, variable importance (mean and standard deviation) and main quantiles of the residuals across all validation runs. Please note that values of the performance measures are usually below those obtained for the full model.

3.4 Panel data

STATY provides different linear models tailored to panel data like the entity fixed effects, time fixed effects, two-ways fixed effects, random entity effects and pooled model.

You have the option to adjust the covariance matrix for each model and to carry out a model validation. Before starting, you have to specify the corresponding variables for entity and time.

3.4.1 Entity Fixed Effects

The modelling technique Entity Fixed Effects can be optionally selected.

Definition

Entity Fixed Effects (EFE) is a linear regression technique that uses several explanatory variables and entity-specific effects to predict a continuous response variable:

$$y_{it} = \beta_1 x_{it1} + \dots + \beta_k x_{itk} + \alpha_i + \varepsilon_{it}$$

where

i = entity index

t = time index

k = number of explanatory variables

y = response variable

 $x_i = explanatory \ variables, j = 1 \dots k$

 $\beta_i = coefficient \ of \ explanatory \ variable \ x_i$

 α_i = entity effect capturing unobserved heterogeneity

 $\varepsilon = idiosyncratic error.$

The coefficients are estimated with the within estimator. The estimator applies OLS to the following transformed regression problem:

$$y_{it} - \overline{y}_i = (x_{it} - \overline{x}_i)^T \beta + (\varepsilon_{it} - \overline{\varepsilon}_i).$$

For the estimator of EFE to be consistent and efficient, given the observed data and entity effects the following assumptions must be fulfilled:

(1)
$$E(\varepsilon_{it}|x_{it},\alpha_i) = 0$$
 (strict exogeneity)

(2)
$$E(\varepsilon_{it}^2|x_{it}) = \sigma_{\varepsilon}^2 \rightarrow (independent \ and \ identically \ distributed$$

(3)
$$E(\varepsilon_{it}\varepsilon_{js}|x_{it}) = 0. \rightarrow error over i and t, \varepsilon_{it} \sim iid(0, \sigma_{\varepsilon}^2))$$

Moreover, for the EFE model with the within estimator it is allowed that the entity effects are correlated with the observations: $E(\alpha_i|x_{it}) \neq 0$, as the within estimator eliminates α_i . Another

possibility to obtain entity effects is by using Least Squares Dummy Variable (LSDV) with dummy variables for the entities.

Covariance type

The within estimator includes the OLS estimator, for which certain assumptions must hold for getting reliable standard errors and statistical test results. Therefore, you can select either a non-robust (homoscedastic) covariance matrix Σ or robust (heteroskedastic, clustered) covariance matrix Σ (Figure 28).

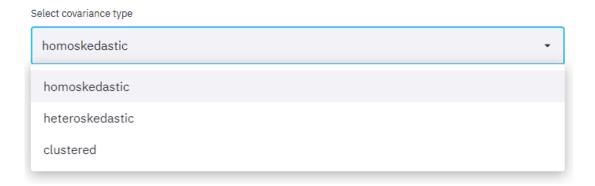


Figure 28 Covariance types for panel data models.

If 'clustered' is selected, either one- or two-way clustering is possible (Figure 29). If the same variable is used for clustering as in the effects model, the degrees of freedom used in estimating the effects are not counted.

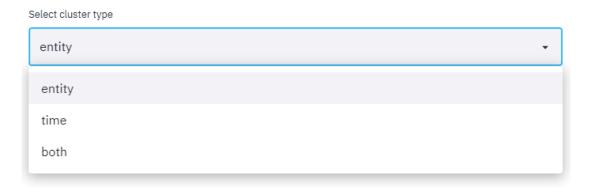


Figure 29 Options for the clustered covariance estimator.

Homoskedastic

$$\begin{split} \Sigma &= s^2 \Sigma_{XX}^{-1}, \\ where \ \Sigma_{XX} &= \sum\nolimits_{i=1}^N \sum\nolimits_{t=1}^{T_i} x_{it}^T x_{it} \\ and \ N &= number \ of \ entities \\ and \ s^2 &= (n-k) \sum\nolimits_{i=1}^N \sum\nolimits_{t=1}^{T_i} \hat{\varepsilon}_{it}^2 \end{split}$$

Heteroskedastic

$$\Sigma = \frac{n}{(n-k)} \Sigma_{XX}^{-1} \hat{S} \Sigma_{XX}^{-1},$$

where
$$\hat{S} = \sum_{i=1}^{N} \sum_{t=1}^{T_i} \hat{\varepsilon}_{it}^2 x_{it}^T x_{it}$$

Clustered (one-way)

$$\Sigma = \frac{n}{(n-k)} \Sigma_{XX}^{-1} \hat{S}_G \Sigma_{XX}^{-1},$$

where
$$\hat{S}_G = \frac{G}{G-1} * \frac{n-1}{n} \sum_{g=1}^{G} \xi_g^T \xi_g$$
, $(G = number\ of\ groups)$

where
$$\xi_g = \sum_{it \in G_g} \hat{\varepsilon}_{it}^2 x_{it}^T x_{it}$$

where it ϵ G_g indicates that the observation belongs to group g

Clustered (two-way)

$$\hat{S}_G = \hat{S}_{G_1} + \hat{S}_{G_2} - \hat{S}_{G_1 \cap G_2}$$

Output

Regression statistics

$$R^{2} = 1 - \frac{\sum_{i}(y_{i} - \hat{y}_{i})^{2}}{\sum_{i}(y_{i} - \bar{y})^{2}} = 1 - \frac{SSR}{SST}$$

$$R^2$$
 (between) = $Corr^2(\hat{y}_i, \bar{y}_i)$, $\hat{y}_i = \bar{x}_i^T \hat{\beta}$

$$R^2$$
 (within) = $Corr^2(\hat{y}_{it} - \hat{y}_i, y_{it} - \bar{y}_i)$, $\hat{y}_{it} = \bar{x}_{it}^T \hat{\beta}$

$$R^2$$
 (overall) = $Corr^2(\hat{y}_{it}, y_{it})$

Log likelihood

$$SST = \sum_{i=1}^{N} \sum_{t=1}^{T} (y_{it} - \bar{y}_i)^2$$

$$SST (overall) = \sum_{i=1}^{N} \sum_{t=1}^{T} (y_{it} - \bar{y})^2$$

Overall performance (with effects)

see 3.3.8 Model comparison

Coefficients

Coefficients = β estimates of the within estimator

Standard error = diagonal of the covariance matrix Σ

$$t - statistic = \frac{coefficient}{standard\ error}$$

p-value = corresponding to t-statistic

lower 95% confidence \approx coefficient -1.96 * standard error

upper 95% *confidence* ≈ *coefficien* + 1.96 * *standard error*

Entity effects/ Time effects/ Combined effects

intercept for each entity, time period and their combinations

F-tests and Hausman-test

F-test (non-robust):

 $F=rac{(SSR_R-SSR_u)/k}{SSR_u/df_u}$ where SSR_R is the restricted sum of squares from the model where the coefficients on all exogenous variables is zero, excluding a constant if one was included; SSR_u is the unrestricted residual sum of squares; k is the number of non-constant regressors in the model; df_u is the residual degree of freedom in the unrestricted model

F-test (robust):

 $W=\hat{eta}^T \Sigma^{-1} \hat{eta}$ where W is the Wald statistic using the estimated parameter covariance and thus inheriting the robustness from the selected covariance estimator; \hat{eta} as coefficient vector does not include the model constant; Σ is the estimated covariance matrix; W is additionally divided by the number of restrictions and consequently following the F-distribution

F-test (poolability):

 $F = \frac{(SSR_{pool} - SSR_{effect})/(df_{pool} - df_{effect})}{SSR_{effect}/df_{effect}} \text{ where } SSR_{pool} \text{ is the residual sum of squares from a no-effect (pooled) model; } SSR_{effect} \text{ is the residual sum of squares from a model with effects; } df_{pool} \text{ is the residual degree of freedom in the pooled regression; } df_{effect} \text{ is the residual degree of freedom from the model with effects; the test statistic follows an } F_{k,dfu} \text{ distribution where } k = df_{pool} - df_{effect}; \text{ testing whether a pooled model is more appropriate (null hypothesis); please note that the underlying test is only correct under the assumption of homoskedasticity}$

Hausman-test:

 $Wu-Hausman\ statistic = (\hat{\beta}_{FE}-\hat{\beta}_{RE})^T \left(Var(\hat{\beta}_{FE})-Var(\hat{\beta}_{RE})\right)^{-1} (\hat{\beta}_{FE}-\hat{\beta}_{RE})$ where $\hat{\beta}_{FE}$ are the coefficients of the fixed effects model; $\hat{\beta}_{RE}$ are the coefficients of the random effects model; testing whether both models are consistent (null hypothesis, RE model is preferred because of higher efficiency); the test statistic follows a $\chi^2(k)$ distribution; please note that the Hausman-test is only conducted for EFE and a homoskedastic covariance type.

3.4.2 Time Fixed Effects

The modelling technique Time Fixed Effects can be optionally selected.

Definition

Time Fixed Effects (TFE) is a linear regression technique that uses several explanatory variables and time-specific effects to predict a continuous response variable:

$$y_{it} = \beta_1 x_{it1} + \dots + \beta_k x_{itk} + \gamma_t + \varepsilon_{it}$$

where

i = entity index

t = time index

k = number of explanatory variables

y = response variable

 $x_i = explanatory variables, j = 1 ... k$

 $\beta_i = coefficient of explanatory variable x_i$

 $\gamma_t = time\ effect$

 $\varepsilon = idiosyncratic error.$

TFE uses the same assumptions and estimator as EFE. Please see 3.4.1 Entity Fixed Effects for details.

Covariance type

See 3.4.1 Entity Fixed Effects for details.

Output

Please see 3.4.1 Entity Fixed Effects for details.

$$\frac{\text{Regression statistics}}{SST = \sum_{i=1}^{N} \sum_{t=1}^{T} (y_{it} - \overline{y_t})^2}$$

3.4.3 Two-ways Fixed Effects

The modelling technique Two-ways Fixed Effects can be optionally selected.

Definition

Two-ways Fixed Effects (TWFE) is a linear regression technique that uses several explanatory variables and entity- as well as time-specific effects to predict a continuous response variable:

$$y_{it} = \beta_1 x_{it1} + \dots + \beta_k x_{itk} + \alpha_i + \gamma_t + \varepsilon_{it}$$

where

i = entity index

t = time index

k = number of explanatory variables

y = response variable

 $x_i = explanatory variables, j = 1 ... k$

 $\beta_i = coefficient \ of \ explanatory \ variable \ x_i$

 $\alpha_i = entity \ effect$

 $\gamma_t = time\ effect$

 $\varepsilon = idiosyncratic error.$

TWFE uses the same assumptions and estimator as EFE and TFE. Please see 3.4.1 Entity Fixed Effects for details.

Covariance type

See 3.4.1 Entity Fixed Effects for details.

Output

Please see 3.4.1 Entity Fixed Effects for details.

$$\frac{\text{Regression statistics}}{SST = \sum_{i=1}^{N} \sum_{t=1}^{T} (y_{it} - \overline{y}_i - \overline{y}_t)^2}$$

3.4.4 Random Effects

The modelling technique Random Effects can be optionally selected.

Definition

Random Effects (RE) is a linear regression technique that uses several explanatory variables and random entity-specific effects to predict a continuous response variable:

$$y_{it} = \beta_1 x_{it1} + \dots + \beta_k x_{itk} + \alpha_i + \varepsilon_{it}$$

where

i = entity index

t = time index

k = number of explanatory variables

y = response variable

 $x_i = explanatory\ variables, j = 1...k$

 $\beta_i = coefficient \ of \ explanatory \ variable \ x_i$

 α_i = entity effect

 $\varepsilon = idiosyncratic error.$

The coefficients are estimated with the quasi-demeaned estimator. The estimator applies OLS to the following equation:

$$y_{it} - \theta_i \overline{y}_i = (x_{it} - \theta_i \overline{x}_i)^T \beta + (1 - \theta_i) \alpha_i + (\varepsilon_{it} - \theta_i \overline{\varepsilon}_i)$$

where θ_i is the parameter for the demeaning, which is estimated by running a specific prior regression (using the between estimator). Note, that $\theta_i=0$ leads to the pooled model and $\theta_i=1$ to the within estimation.

Similar assumptions as for the Entity Fixed Effects (EFE) model must be fulfilled:

(1)
$$E(\varepsilon_{it}|x_{it},\alpha_i) = 0$$
 (strict exogeneity)

(2)
$$E(\varepsilon_{it}^2|x_{it}) = \sigma_{\varepsilon}^2 \rightarrow (independent \ and \ identically \ distributed)$$

(3)
$$E(\varepsilon_{it}\varepsilon_{js}|x_{it}) = 0. \rightarrow error over i and t, \varepsilon_{it} \sim iid(0, \sigma_{\varepsilon}^2)$$

In contrast to the EFE model with the within estimator where it is allowed that the entity effects are correlated with the observations: $E(\alpha_i|x_{it}) \neq 0$, this correlation is not allowed in the RE estimator: $E(\alpha_i|x_{it}) = 0$. The RE model assumes an independent and identically distributed effect: $\alpha_i \sim iid(0, \sigma_\alpha^2)$).

Covariance type

Please see 3.4.1 Entity Fixed Effects for details.

Output

Please also see 3.4.1 Entity Fixed Effects for details.

$$\frac{\text{Regression statistics}}{SST = \sum_{i=1}^{N} \sum_{t=1}^{T} (\tilde{y}_{it} - \bar{\tilde{y}})^2, \qquad \tilde{y}_{it} = (y_{it} - \theta_i \bar{y}_i)}$$

Variance decomposition

$$\hat{\sigma}_{\varepsilon}^{2}(idiosyncratic\ error) = \frac{\sum_{i=1}^{N} \sum_{t=1}^{T} \hat{\varepsilon}_{it}^{2}}{\sum_{i=1}^{N} T_{i} - N - k + 1}$$

 $\hat{\sigma}_{\alpha}^2(individual\ error) = \max \bigl\{0, \frac{SSR_b}{N-k} - \frac{\hat{\sigma}_{\mathcal{E}}^2}{\bar{T}}\bigr\} \ \text{where} \ SSR_b \ \text{is the residual sum of squares of the between estimator and} \ \bar{T} = \frac{n}{\sum_{l=1}^n r_l^{-1}}$

$$\theta_i = 1 - \sqrt{\frac{\hat{\sigma}_{\varepsilon}^2}{T_i \hat{\sigma}_{\alpha}^2 + \hat{\sigma}_{\varepsilon}^2}}$$

3.4.5 Pooled

See 3.3.1 Multiple Linear Regression.

3.4.6 Model predictions

See 3.3.9 Model predictions.

Please note that when models with effects are used and predictions should be made for new data, effects are only included in the predictions if the effect of the corresponding entity or time period could be identified by the calibrated model with the original data. Consequently, if no effects are available, the prediction is solely based on the selected explanatory variables.

3.4.7 Model validation

See 3.3.10 Model validation.

3.5 Time series data

Follows soon!

3.5.1 Diagnosis plots and tests

3.5.2 Time series detrending and seasonal adjustment

3.5.3 Moving Average (MA)

3.5.4 Autoregressive model (AR)

3.5.5 Autoregressive Moving Average (ARMA)

| 3.5.6 Non-seasonal | Autoregressive I | Integrated M | oving Average | e (non-seasonal | ARIMA) |
|--------------------|------------------|--------------|---------------|-----------------|--------|
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3.5.7 Seasonal Autoregressive Integrated Moving Average (seasonal ARIMA)

3.6 Web scraping and text data

With web scraping and text data you can get automatic summaries of web pages, have access to stock prices and can create word clouds for any text you want (Figure 30).

What analysis would you like to perform?

-
Web-Page summary
Stock data analysis
WordCloud

Figure 30 Analysis options for web scraping and text data.

3.6.1 Web-Page summary

The web page summary provides a small web page preview of the inserted link and an automatic web page summary consisting of five sentences (Figure 31). The summary is created by using natural language processing and a neural network language model.



Figure 31 Web-Page summary input area.

3.6.2 Stock data analysis

Stock data analysis uses stock data from yahoo finance. You can insert any stock ticker symbol to get price information for specific time periods and select further stocks for comparison (Figure 32).

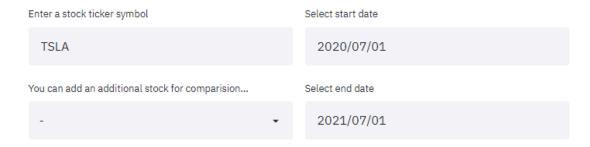


Figure 32 Input for stock data analysis.

3.6.3 WordCloud

Based on provided links to web pages or text inputs you can create word clouds with a selectable main colour (Figure 33). Words that you would like to exclude can be selected prior to creating the word cloud. If a link to a web page is provided, a web page preview is shown additionally.

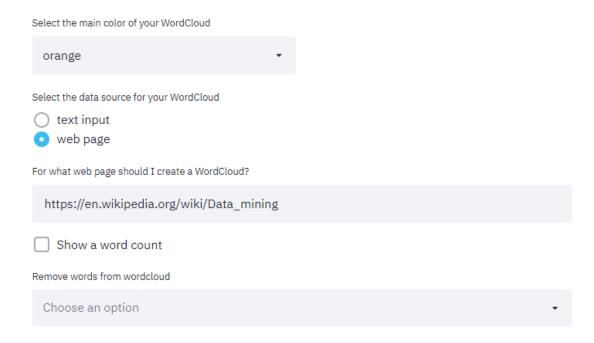


Figure 33 Options for the creation of a word cloud.

3.7 Geospatial data

Geospatial data can be graphically displayed with STATY. For the correct identification, a country name or country code has to be provided. If a time variable is included, the temporal development of the selected variable for each geographical region can be animated (Figure 34).

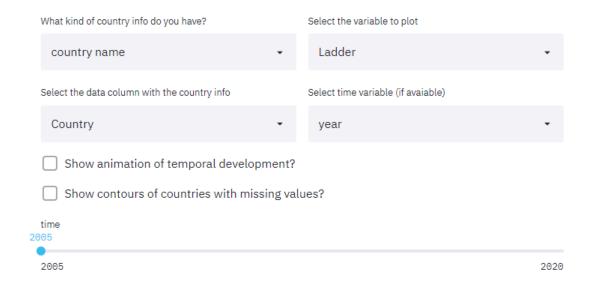


Figure 34 Settings for geospatial data visualization.

Additionally, histograms and boxplots for the variable of interest are created.

4 Default data

4.1 Uni- and bivariate data

Data source

The data come from the World Bank's study on financial intermediation and growth.

Citation

Levine, Ross; Loayza, Norman; Beck, Thorsten.1999. Financial intermediation and growth: Causality and causes (English). Policy, Research working paper; no. WPS 2059 Washington, D.C.: World Bank Group.

Variables in the dataset

UrbanPopulation Share of urban population

Private saving rate as the ratio of gross private savings and

PrivateSaving

GPDI

logGDPI Log of real per capita GPDI

GrowtRate Growth rate of real GPDI
GovernmentSaving Government saving as share of real GDP

LogTermsTrade Log of terms of trade

OlderThan65 Share of population over 65 in total population
Under15 Share of population under 15 in total population

CommercialCentralBank Commercial-Central Bank, period average

LiquidLiabilities Liquid Liabilities, period average
PrivateCredit Private Credit, period average
BankCredit Bank credit, period average

LegalOrigin English, French, German or Scandinavian

4.2 Multivariate data

Data source

The data come from the Gallup World Poll surveys from 2018 to 2020. For more details see the World Happiness Report 2021.

Citation

Helliwell, John F., Richard Layard, Jeffrey Sachs, and Jan-Emmanuel De Neve, eds. 2021. World Happiness Report 2021. New York: Sustainable Development Solutions Network.

Variables in the dataset

Country country name

Year year ranging from 2005 to 2020

Ladder happiness score or subjective well-being with the best possible

life being a 10, and the worst possible life being a 0

Log GDP per capita in purchasing power parity at constant 2017 international dollar

prices

Social support the national average of the binary responses (either 0 or 1) to the

question regarding relatives or friends to count on

Healthy life based on the data extracted from the World Health Organization's

expectancy at birth Global Health Observatory data repository

choices

Generosity residual of regressing national average of response to the

question regarding money donations in the past month on GDP

per capita

Perceptions of the national average of the survey responses to the

corruption corresponding question

Positive affect the average of three positive affect measures (happiness, laugh

and enjoyment)

Negative affect the average of three negative affect measures (worry, sadness

(worry, sadness and a

anger)

and anger)

4.3 Panel data

Data source

This is the original 11-firm data set from Grunfeld's Ph.D. thesis (*Grunfeld, 1958, The Determinants of Corporate Investment, Department of Economics, University of Chicago*). For more details see online complements for the article The Grunfeld Data at 50.

Citation

Kleiber C, Zeileis A (2010). "The Grunfeld Data at 50," German Economic Review, 11(4), 404-417. doi:10.1111/j.1468-0475.2010.00513.x

Variables in the dataset

| invest | Gross investment, defined as additions to plant and equipment |
|---------|---|
| | plus maintenance and repairs in millions of dollars deflated by the |
| | implicit price deflator of producers' durable equipment (base |
| | 1947). |
| value | Market value of the firm, defined as the price of common shares |
| | at December 31 (or, for WH, IBM and CH, the average price of |
| | December 31 and January 31 of the following year) times the |
| | number of common shares outstanding plus price of preferred |
| | shares at December 31 (or average price of December 31 and |
| | January 31 of the following year) times number of preferred |
| | shares plus total book value of debt at December 31 in millions of |
| | dollars deflated by the implicit GNP price deflator (base 1947). |
| capital | Stock of plant and equipment, defined as the accumulated sum of |
| | net additions to plant and equipment deflated by the implicit price |
| | deflator for producers' durable equipment (base 1947) minus |
| | depreciation allowance deflated by depreciation expense deflator |
| | (10 years moving average of wholesale price index of metals and |
| | metal products, base1947). |
| firm | General Motors (GM), US Steel (US), General Electric (GE), |
| | Chrysler (CH), Atlantic Refining (AR), IBM, Union Oil (UO), |
| | Westinghouse (WH), Goodyear (GY), Diamond Match (DM), |
| | American Steel (AS). |
| year | Year ranging from 1935 to 1954. |

4.4 Time series data

Data source

The data come from Box & Jenkins (1970), but we use the version that is integrated in the R package 'astsa' which is a companion to the book 'Time Series Analysis and Its Applications' by Shumway & Stoffer's (2017).

Citation

Box, G.E.P. and G.M. Jenkins (1970). Time Series Analysis, Forecasting, and Control.

Oakland, CA: Holden-Day

Shumway, R.H, and D.S. Stoffer (2017). Time Series Analysis and Its Applications: With R

Examples. New York: Springer

Variables in the dataset

Air passengers The monthly totals of international airline passengers

Date Month ranging from January 1949 to December 1960

4.5 Geospatial data

See 4.2 Multivariate data.

5 Contact

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6 Disclaimer

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