

LECTURE 6

§7. HOTELLING DISTRIBUTION.

Hotelling's T -squared distribution (T^2) is a multivariate distribution proportional to the F -distribution and arises importantly as the distribution of a set of statistics which are natural generalizations of the statistics underlying Student's t -distribution. Hotelling's T -squared statistic (T^2) is a generalization of Student's t -statistic that is used in multivariate hypothesis testing.

Multivariate method that is the multivariate counterpart of Student's t and which also forms the basis for certain multivariate control charts is based on Hotelling's T^2 distribution, which was introduced by Hotelling (1947). Recall that

$$t = \frac{\bar{x} - \mu}{S/\sqrt{n}}$$

has a t -distribution, provided that η is normally distributed. If we wanted to test the hypothesis that $\mu = \mu_0$, we would then have

$$t = \frac{\bar{x} - \mu}{S/\sqrt{n}}$$

so that

$$t^2 = \frac{(\bar{x} - \mu)^2}{S^2/n} = n(\bar{x} - \mu)(S^2)^{-1}(\bar{x} - \mu).$$

When T^2 is generalized to m variables it becomes

$$T^2 = n(\bar{\mathbf{x}} - \mu_0)S^{-1}(\bar{\mathbf{x}} - \mu_0),$$

with

$$\bar{\mathbf{x}} = \begin{bmatrix} \bar{x}_1 \\ \bar{x}_2 \\ \cdot \\ \cdot \\ \cdot \\ \bar{x}_m \end{bmatrix} \quad \bar{\mu} = \begin{bmatrix} \mu_1^0 \\ \mu_2^0 \\ \cdot \\ \cdot \\ \cdot \\ \mu_m^0 \end{bmatrix}.$$

S^{-1} is the inverse of the sample variance-covariance matrix, S , and n is the sample size upon which each \bar{x}_i , $i = 1, 2, \dots, m$, is based. (The diagonal elements of S are the variances and the off-diagonal elements are the covariances for the m variables).

The continuous probability distribution, concentrated on the positive semi-axis $(0, \infty)$, with density

$$f_k(x) = \frac{\Gamma\left(\frac{n+1}{2}\right) x^{k/2-1} \left(1 + \frac{x}{n}\right)^{-(n+1)/2}}{\Gamma\left(\frac{n-k+1}{2}\right) \Gamma(k/2) \cdot n^{k/2}}. \quad (7.1)$$

depending on two integer parameters (the number of degrees of freedom) n and k . If a k -dimensional random vector Y has the normal distribution with null vector of means and covariance matrix Σ and if

$$S = \frac{1}{n} \sum_{i=1}^n Z_i' Z_i$$

where the random vectors Z_i are independent, distributed as Y and also independent of Y , then the random variable $T^2 = Y' S^{-1} Y$ has the Hotelling -distribution with n degrees of freedom (Y is a column vector and Y' means transposition). If $k = 1$, then

$$T^2 = \frac{Y^2}{\frac{1}{n} \chi_n^2} = t_n^2,$$

where the random variable t_n has the Student distribution with n degrees of freedom. If in the definition of the random variable T^2 it is assumed that Y has the normal distribution with parameters (ν, Σ) and Z_i has the normal distribution with parameters $(0, \Sigma)$, then the corresponding distribution is called a non-central Hotelling T^2 -distribution with degrees of freedom and non-centrality parameter ν .

Hotelling's T^2 -distribution is used in mathematical statistics in the same situation as Student's t -distribution, but then in the multivariate case. If the results of observations X_1, X_2, \dots, X_n are independent normally-distributed random vectors with mean vector ν and non-degenerate covariance matrix Σ , then the statistic

$$T^2 = n(\bar{X} - \mu)' S^{-1} (\bar{X} - \mu),$$

where

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

and

$$S = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})(X_i - \bar{X})'$$

has the Hotelling T^2 -distribution with $n - 1$ degrees of freedom. This fact forms the basis of the Hotelling test. For numerical calculations one uses tables of the beta-distribution or of the Fisher F -distribution, because the random variable

$$\frac{n - k + 1}{nk} T^2$$

has the F -distribution with k and $n - k + 1$ degrees of freedom.

The Hotelling -distribution was proposed by H. Hotelling [1] for testing equality of means of two normal populations.

Let us verify, that the function in (7.1) is density function. Recall definition of the Beta-function:

$$B(m, n) = \int_0^1 x^{m-1} (1 - x)^{n-1} dx.$$

It is well-known that the values of this function can be calculated using Γ -function:

$$B(m, n) = \frac{\Gamma(m) \Gamma(n)}{\Gamma(n + m)}.$$

Therefore, using Beta-function, (7.1) we can rewrite in the following form:

$$\frac{1}{B\left(\frac{k}{2}, \frac{n - k + 1}{2}\right)} \frac{x^{k/2-1} \left(1 + \frac{x}{n}\right)^{-(n+1)/2}}{n^{k/2}}.$$

If we make the change of variables

$$x = \frac{y}{1 + y} \quad \text{and} \quad dx = \frac{1}{(1 + y)^2} dy$$

we get

$$B(m, n) = \int_0^1 x^{m-1} (1 - x)^{n-1} dx = \int_0^{+\infty} y^{m-1} (1 + y)^{-(m+n)} dy.$$

Therefore,

$$\int_0^{+\infty} f_k(x) dx = \frac{1}{B\left(\frac{k}{2}, \frac{n - k + 1}{2}\right)} \int_0^{+\infty} \frac{x^{k/2-1} \left(1 + \frac{x}{n}\right)^{-(n+1)/2}}{n^{k/2}} dx =$$

If we make change of variable $\frac{x}{n} = z$, we obtain

$$= \frac{1}{B\left(\frac{k}{2}, \frac{n - k + 1}{2}\right)} \int_0^{+\infty} z^{k/2-1} (1 + z)^{-(n+1)/2} dz = \frac{1}{B\left(\frac{k}{2}, \frac{n - k + 1}{2}\right)} B\left(\frac{k}{2}, \frac{n - k + 1}{2}\right) = 1.$$