

YSU ASDS, Statistics, Fall 2019

Lecture 21

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Contents

- ▶ Maximum Likelihood Estimation
- ▶ Some Topics we'll cover soon
- ▶ Confidence Intervals (CI)

Last Lecture ReCap

- ▶ What is the problem that MLE is solving?

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- ▶ What are the remarkable properties of the MLE?

Properties of the MLE, Cont'd

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$$\widehat{g(\theta)}^{MLE} = g\left(\hat{\theta}^{MLE}\right).$$

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Example Find the MLE for σ in $\mathcal{N}(\mu, \sigma^2)$ Model.

Solution: OTB

Some topics we will talk about soon

- ▶ Multivariate Normal and MLE for MVNormal
- ▶ Kullback-Leibler Divergence and its relation to MLE
- ▶ MLE for the Mixture Model, EM Algorithm
- ▶ Bayesian Estimation: MAP and Bayes Estimator

Confidence Intervals

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Prelude No. 2

But the good news is that even when we cannot exactly find the True value of our Parameter using $\hat{\theta}$, if $\hat{\theta}$ possesses some good properties, we believe that the Estimate obtained is a good approximation/Estimate for θ^* .

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And you were asking nice questions about how small is the *error* or how much sure are we in our Estimate (for the Unknown Parameter), and how large n needs to be to have a good estimate.

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Here we want to develop the theory of Confidence Intervals, which will contain answers to these questions.

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CI Problem Setting

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- ▶ which has the possible smallest length.

Let us state this in Mathematical terms. We will consider here only 1D case, i.e., we will assume $\theta \in \Theta \subset \mathbb{R}$.

Random Intervals and CI

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Example: Assume $X \sim \text{Pois}(2.3)$. Then

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Example: Let X_1, X_2, \dots, X_n are IID r.v.s. Then

$$(\bar{X} - 0.1, \bar{X} + 0.1)$$

is a Random Interval.

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The usual values of the confidence level are 90%, 95%, 99%, so the usual values of α are 0.1, 0.05 and 0.01.

Definition: Assume $0 < \alpha < 1$, and let $L = L(x_1, \dots, x_n, \alpha)$, $U = U(x_1, \dots, x_n, \alpha)$ be two functions with $L(x_1, \dots, x_n, \alpha) \leq U(x_1, \dots, x_n, \alpha)$ for all $(x_1, \dots, x_n, \alpha)$.

Definition: Assume $0 < \alpha < 1$, and let $L = L(x_1, \dots, x_n, \alpha)$, $U = U(x_1, \dots, x_n, \alpha)$ be two functions with $L(x_1, \dots, x_n, \alpha) \leq U(x_1, \dots, x_n, \alpha)$ for all $(x_1, \dots, x_n, \alpha)$. The random interval

$$(L, U) = (L(X_1, \dots, X_n, \alpha), U(X_1, \dots, X_n, \alpha))$$

is called a **confidence interval (or confidence interval estimator) for θ of confidence level $1 - \alpha$** , if for any $\theta \in \Theta$,

$$\mathbb{P}(L < \theta < U) \geq 1 - \alpha.$$

In the case we have a realization/observation of X_1, \dots, X_n , say, x_1, \dots, x_n , then the interval

$$\left(L(x_1, \dots, x_n, \alpha), U(x_1, \dots, x_n, \alpha) \right)$$

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Going back to our CI, CI of the confidence level $1 - \alpha$ is a Random Interval that contains θ in more than $(1 - \alpha) \cdot 100\%$ of cases.

CI, Interpretation

Note: It is important to understand, that in the CI definition

$$\mathbb{P}(L < \theta < U) \geq 1 - \alpha$$

θ is not our r.v., θ is our unknown constant Parameter, so we do not read this as “with high Probability, θ is in (L, U) ”.

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So, if we will have/generate different observations, we will have different Intervals² (L, U) , and we want to have that most of the time that interval contains our unknown Parameter value.

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CI, R Simulation

Example: Consider an example: our Model is $Exp(\lambda)$, and we have an observation from it. Let us take a Random Sample for the general case: X_1, X_2, \dots, X_n from $Exp(\lambda)$.

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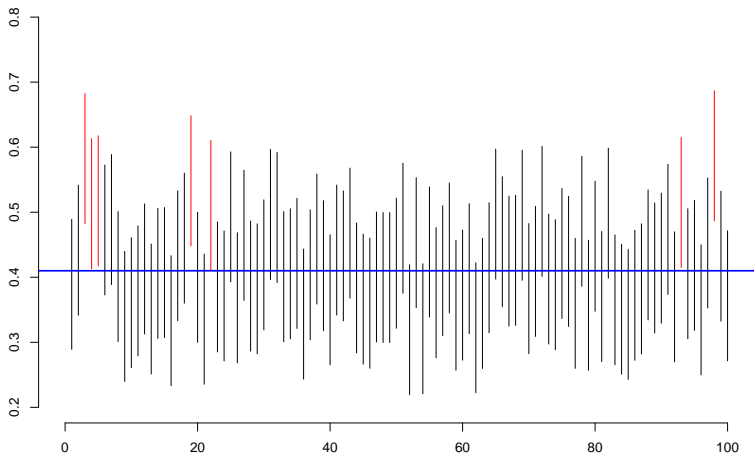
Now, let us take as CI

$$\left(\frac{1}{\bar{X}} - 0.1, \frac{1}{\bar{X}} + 0.1 \right)$$

and do some simulations:

CI, R Simulation

Exponential Model, CI, $(1/\text{mean} - 0.1, 1/\text{mean} + 0.1)$



CI, R Simulation, Code

#CI Idea, Exponential Model

```
lambda <- 0.41
```

```
conf.level <- 0.95; a = 1 - conf.level
```

```
sample.size <- 50; no.of.intervals <- 100
```

```
epsilon <- 0.1
```

```
plot.new()
```

```
plot.window(xlim = c(0,no.of.intervals), ylim = c(0.2,0.8))
```

```
axis(1); axis(2)
```

```
title("Exponential Model, CI, (1/mean - 0.1, 1/mean + 0.1)")
```

```
for(i in 1:no.of.intervals){
```

```
  x <- rexp(sample.size, rate = lambda)
```

```
  lo <- 1/mean(x) - epsilon; up <- 1/mean(x) + epsilon
```

```
  if(lo > lambda || up < lambda){
```

```
    segments(c(i), c(lo), c(i), c(up), col = "red")
```

```
  }
```

```
  else{
```

```
    segments(c(i), c(lo), c(i), c(up))
```

```
  }
```

```
}
```

```
abline(h = lambda, lwd = 2, col = "blue")
```

Methods to obtain Confidence Intervals

We will consider several methods to construct CIs:

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And then we will talk about Asymptotic CIs.

Prob Refresher, Chebyshev Inequality

Recall the Cheby Inequality: If X is a r.v. with finite Mean $\mathbb{E}(X)$ and Variance $Var(X)$, then for any $\varepsilon > 0$,

$$\mathbb{P}\left(|X - \mathbb{E}(X)| \geq \varepsilon\right) \leq \frac{Var(X)}{\varepsilon^2},$$

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or, which is the same,

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CI for the Mean, Variance is given, Cheby Method

Example: Assume X_1, X_2, \dots, X_n are Independent r.v. with the same Mean $\mathbb{E}(X_k) = \mu$ and the same Variance $\text{Var}(X_k) = \sigma^2$.

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so if we will plug \bar{X} in the Cheby Inequality, we will obtain

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for any $\varepsilon > 0$.

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for any $\varepsilon > 0$. Now, take $\frac{\sigma^2}{n \cdot \varepsilon^2} = \alpha$. Here, σ , n and α are known, so this equality will give us the value for ε :

$$\varepsilon = \frac{\sigma}{\sqrt{n \cdot \alpha}}.$$

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Plugging this value into the above Cheby Inequality, we will get

$$\mathbb{P}\left(|\bar{X} - \mu| < \frac{\sigma}{\sqrt{n \cdot \alpha}}\right) \geq 1 - \alpha,$$

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This means, that the interval

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Some Notes

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The CI length obtained above is

$$\frac{2\sigma}{\sqrt{n \cdot \alpha}}.$$

Note: If we increase n , the CI gets narrower. This is intuitive: if we collect more data, we can estimate the parameter more precisely, we can enclose it in a smaller length interval.

Some Notes

Two notes about the obtained CI - in fact, these notes will work also for other cases too:

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Note: If we increase the Confidence Level, i.e., if we decrease α , then the length of CI increases. This is intuitive too: if we want to be more sure where our unknown Parameter is lying, we will get a larger interval.

CI for the Proportion, Cheby Method

Example: Now, let us construct a CI of CLevel $1 - \alpha$ for p in the *Bernoulli*(p) Model.

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Here, on the RHS, we have the unknown parameter value p , which is not desirable. To get rid of that, we use the estimate

$$p(1-p) \leq \frac{1}{4}, \text{ so } \mathbb{P}(|\bar{X} - p| < \varepsilon) \geq 1 - \frac{1}{4n \cdot \varepsilon^2}.$$

CI for the Proportion, Cheby Method, Cont'd

On the RHS, for the CI, we want to have $1 - \alpha$. So, as in the previous Example, we choose

$$\frac{1}{4n \cdot \varepsilon^2} = \alpha,$$

CI for the Proportion, Cheby Method, Cont'd

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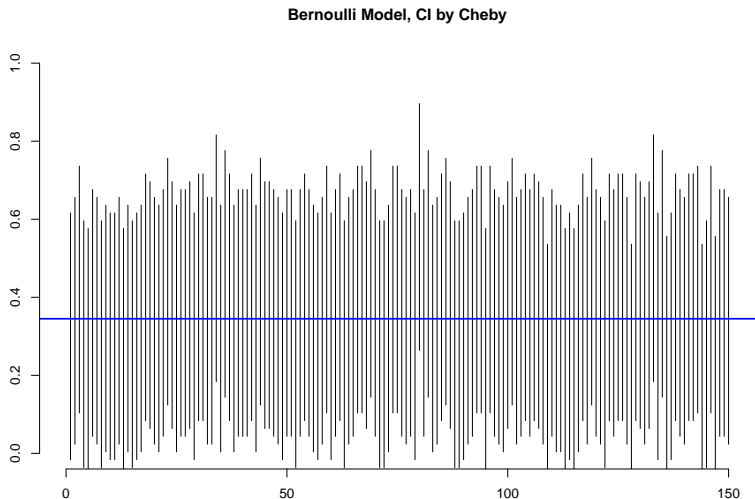
$$\mathbb{P}\left(\bar{X} - \frac{1}{2\sqrt{n \cdot \alpha}} < p < \bar{X} + \frac{1}{2\sqrt{n \cdot \alpha}}\right) \geq 1 - \alpha.$$

This means that the interval

$$\left(\bar{X} - \frac{1}{2\sqrt{n \cdot \alpha}}, \bar{X} + \frac{1}{2\sqrt{n \cdot \alpha}}\right)$$

is a CI for p of level $1 - \alpha$.

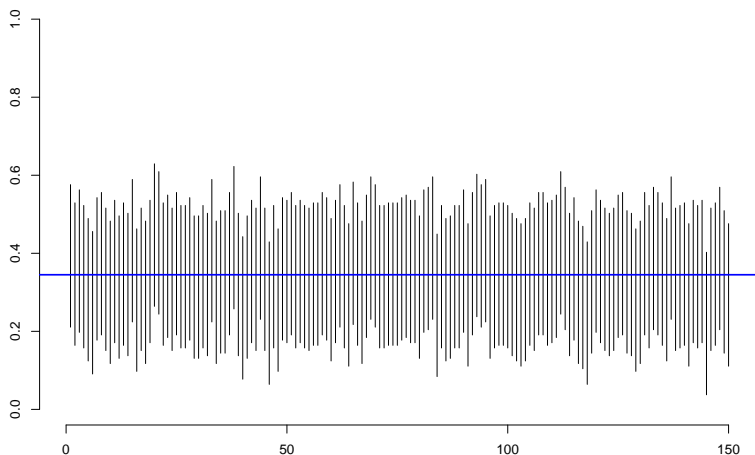
CI for Bernoulli, R Simulation



Sample Size = 50, $CL = 95\%$

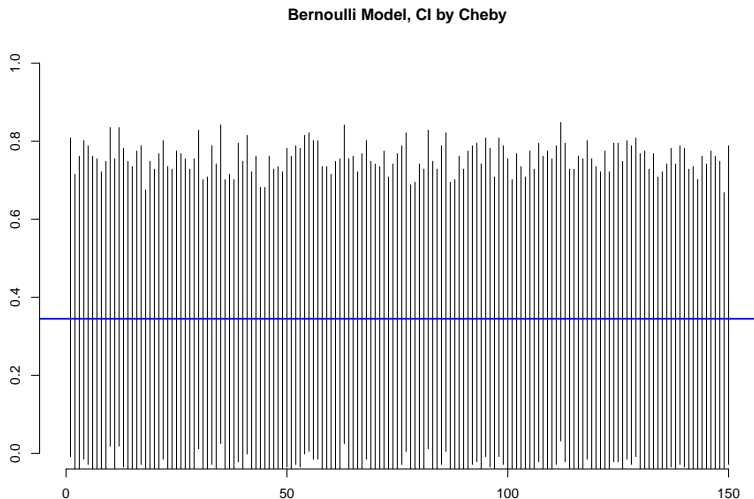
CI for Bernoulli, R Simulation

Bernoulli Model, CI by Cheby



Sample Size = 150, $CL = 95\%$

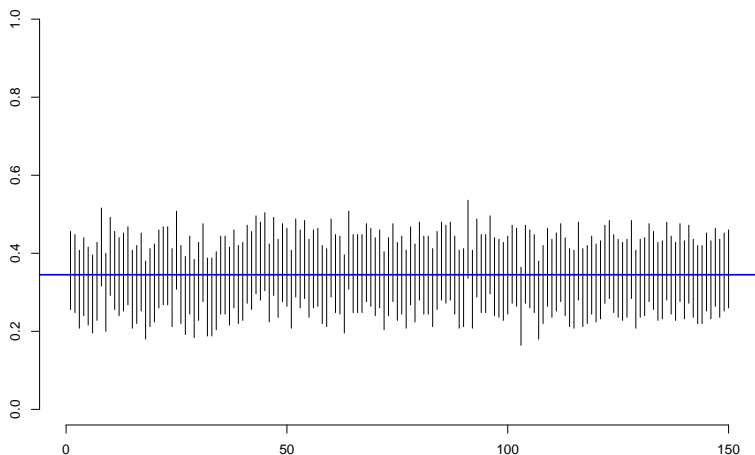
CI for Bernoulli, R Simulation



Sample Size = 150, $CL = 99\%$

CI for Bernoulli, R Simulation

Bernoulli Model, CI by Cheby



Sample Size = 250, $CL = 90\%$

CI, R Simulation, Code

```
#CI Idea, Bernoulli Model
```

```
p <- 0.345
```

```
conf.level <- 0.9; a = 1 - conf.level
```

```
sample.size <- 250; no.of.intervals <- 150
```

```
ME <- 1/(2*sqrt(sample.size*a)) #Margin of Error
```

```
plot.new()
```

```
plot.window(xlim = c(0,no.of.intervals), ylim = c(0,1))
```

```
axis(1); axis(2)
```

```
title("Bernoulli Model, CI by Cheby")
```

```
for(i in 1:no.of.intervals){
```

```
  x <- rbinom(sample.size, size = 1, prob = p)
```

```
  lo <- mean(x) - ME
```

```
  up <- mean(x) + ME
```

```
  if(lo > p || up < p){
```

```
    segments(c(i), c(lo), c(i), c(up), col = "red")
```

```
  }
```

```
  else{
```

```
    segments(c(i), c(lo), c(i), c(up))
```

```
  }
```

```
}
```

```
abline(h = p, lwd = 2, col = "blue")
```


CI for the Proportion, Cheby Method

Recall that if we have a Random Sample

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then the interval

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Note: Here

$$\frac{1}{2\sqrt{n \cdot \alpha}}$$

is called the **Margin of Error** (for the Interval Estimate of p).

Examples

Example: Assume we are interested in the proportion of smokers in AUA. We ask 120 persons at AUA and learn that 55 of them are smokers. Construct a CI for the proportion of smokers in AUA of 95% confidence level.

Solution: OTB

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Example: Continuing the above Example: now assume we want to find that Proportion within the Error Margin 0.1, with the CL 95%. At least, how many persons at AUA we need to ask?

Solution: OTB