

23.11.2021

Seminar W9 - 831 X_1, X_2, \dots, X_n i.i.d., $X_i \sim X$ We want to find estimators for the parameters of the law of X

method of moments: $\bar{v}_k = \frac{1}{n} \sum_{i=1}^n X_i^k$

$\bar{v}_k = \begin{cases} \sum_{i=1}^n x_i^k \cdot f_X(x) \\ \int_{\mathbb{R}} x^k \cdot f_X(x) dx, \text{ if } X \text{ continuous} \end{cases}$

method of maximal likelihood.

→ minimising the likelihood function

Exercise 1. Using the method of moments, find estimators for the parameter(s) of the distribution of the characteristic X , if:

- (a) $X \sim \text{Geo}(p)$, $p \in (0, 1)$;
- (b) $X \sim \text{Exp}(\lambda)$, $\lambda > 0$;
- (c) $X \sim \text{Bino}(n, p)$, $n \in \mathbb{N}^*$, $p \in (0, 1)$;
- (d) $X \sim \text{Unif}[a, b]$, $a < b$;
- (e) $X \sim \text{NBin}(n, p)$, $n \in \mathbb{N}^*$, $p \in (0, 1)$;
- (f) $X \sim \text{Gamma}(a, b)$, $a, b > 0$;
- (g) $X \sim \text{Pareto}(\alpha, \beta)$, $\alpha > 2$, $\beta > 0$;

• If $X \sim \text{Gamma}(a, b)$, $a, b > 0$, then:

$$f_X(x) = \frac{1}{\Gamma(a)b^a} \cdot x^{a-1} e^{-\frac{x}{b}} \cdot 1_{(0, \infty)}(x)$$

where:

$$\Gamma(a) = \int_0^\infty x^{a-1} e^{-x} dx, \quad a > 0$$

$$E(X) = ab, \quad V(X) = ab^2$$

Sol. (f) $X \sim \text{Gamma}(a, b)$

$$\begin{cases} \bar{v}_1 = v_1 \\ \bar{v}_2 = v_2 \end{cases}$$

$$\begin{aligned} v_1 &= E(X) = ab \\ v_2 &= ab^2 \end{aligned}$$

$$V(X) = E(X^2) - E(X)^2 = v_2 - v_1^2$$

$$V(X) = ab^2 \Rightarrow \bar{V}_2 = V(X) + \bar{V}_1^2 = ab^2 + (ab)^2$$

$$\Rightarrow \begin{cases} \bar{V}_1 = ab \\ \bar{V}_2 = ab^2 + (ab)^2 \end{cases}$$

We solve the system, considering a and b as our variables

$$\Rightarrow \begin{cases} a = \frac{\bar{V}_1}{b} \\ \bar{V}_2 = ab^2 + (ab)^2 \end{cases} \Rightarrow \begin{cases} a = \frac{\bar{V}_1}{b} \\ \bar{V}_2 = \frac{\bar{V}_1}{b} \cdot b^2 + \bar{V}_1^2 \end{cases} \Rightarrow$$

$$\Rightarrow \begin{cases} a = \frac{\bar{V}_1}{b} \\ \bar{V}_2 = b\bar{V}_1 + \bar{V}_1^2 \end{cases} \Rightarrow \begin{cases} b = \frac{\bar{V}_2 - \bar{V}_1^2}{\bar{V}_1} \\ a = \frac{\bar{V}_1}{b} \end{cases} \Rightarrow \begin{cases} b = \frac{\bar{V}_2 - \bar{V}_1^2}{\bar{V}_1} \\ a = \frac{\bar{V}_1^2}{\bar{V}_2 - \bar{V}_1^2} \end{cases}$$

Therefore the estimators for a and b are:

$$\bar{a} = \frac{\bar{V}_1^2}{\bar{V}_2 - \bar{V}_1^2} \quad \bar{b} = \frac{\bar{V}_2 - \bar{V}_1^2}{\bar{V}_1}$$

(a) $X \sim \text{Geo}(p)$

$$\bar{V}_1 = \bar{V}_1 \quad \bar{V}_1 = E(X) = \frac{1-p}{p}$$

$$\bar{V}_1 = \frac{1-p}{p} \Rightarrow p\bar{V}_1 = 1-p \Rightarrow p\bar{V}_1 + p = 1 \Rightarrow p(\bar{V}_1 + 1) = 1 \Rightarrow$$

$$\Rightarrow p = \frac{1}{\bar{V}_1 + 1} \Rightarrow \bar{p} = \frac{1}{\bar{V}_1 + 1}$$

(b) $X \sim \text{Exp}(\lambda)$

$$E(X) = \frac{1}{\lambda} \quad \left(V(X) = \frac{1}{\lambda^2} \quad \text{for which it may be known} \right)$$

$$\bar{v}_1 = E(X) = \frac{1}{\lambda} \quad \text{"new"}$$

$$\Rightarrow \bar{\lambda} = \frac{1}{\bar{v}_1} = \frac{n}{x_1 + x_2 + \dots + x_n} \quad \text{"old"}$$

$$(g) \quad X \sim \text{Pareto}(\alpha, \beta)$$

$$v_2 = E(X^2) = V(X) + (E(X))^2$$

$$\begin{cases} \bar{v}_1 = v_1 = E(X) = \frac{\alpha \beta}{\alpha - 1} \\ \bar{v}_2 = v_2 = V(X) + E(X)^2 = \frac{\alpha \beta^2}{(\alpha - 1)^2 \cdot (\alpha - 2)} + \frac{\alpha^2 \beta^2}{(\alpha - 1)^2} = \\ = \frac{\alpha \beta^2}{(\alpha - 1)^2} \left(\frac{1}{\alpha - 2} + \alpha \right) \end{cases}$$

$$\begin{cases} \beta = \frac{\bar{v}_1 \cdot (\alpha - 1)}{\alpha} \\ \bar{v}_2 = \frac{\alpha}{(\alpha - 1)^2} \cdot \frac{\alpha^2 - 2\alpha + 1}{\alpha - 2} \cdot \frac{\bar{v}_1^2 \cdot (\alpha - 1)^2}{\alpha^2} \end{cases}$$

$$\Rightarrow \begin{cases} \beta = \frac{\bar{v}_1 \cdot (\alpha - 1)}{\alpha} \\ \bar{v}_2 = \frac{\alpha \cdot (\alpha - 1)^2 \cdot \bar{v}_1^2}{\alpha^2 \cdot (\alpha - 2)} \end{cases}$$

$$\Rightarrow \begin{cases} \beta = \frac{\bar{v}_1 \cdot (\alpha - 1)}{\alpha} \\ \bar{v}_2 = \frac{(\alpha - 1)^2}{\alpha(\alpha - 2)} \cdot \frac{\bar{v}_1^2}{\alpha} \Rightarrow \frac{\alpha^2 - 2\alpha + 1}{\alpha^2 - 2\alpha} = \frac{\bar{v}_2}{\bar{v}_1^2} \Rightarrow \end{cases}$$

$$\Rightarrow \frac{\alpha^2 - 2\alpha + 1}{(\alpha^2 - 2\alpha + 1) - (\alpha^2 - 2\alpha)} = \frac{\bar{v}_2}{\bar{v}_1^2}$$

$$\Rightarrow (\alpha - 1)^2 = \frac{\bar{y}_2}{\bar{y}_2 - \bar{y}_1^2} \Rightarrow \alpha = 1 + \sqrt{\frac{\bar{y}_2}{\bar{y}_2 - \bar{y}_1^2}}$$

$$\Rightarrow \beta = \bar{y}_1 \cdot \frac{\alpha - 1}{\alpha} = \bar{y}_1 \cdot \frac{\sqrt{\frac{\bar{y}_2}{\bar{y}_2 - \bar{y}_1^2}}}{1 + \sqrt{\frac{\bar{y}_2}{\bar{y}_2 - \bar{y}_1^2}}} =$$

$$= \frac{\bar{y}_1 \sqrt{\bar{y}_2}}{\sqrt{\bar{y}_2 - \bar{y}_1^2} + \sqrt{\bar{y}_2}}$$

(c) $X \sim \text{Bin}(n, p)$

$$E(X) = np \quad V(X) = np(1-p)$$

$$\begin{cases} \bar{y}_1 = y_1 \\ \bar{y}_2 = y_2 \end{cases} \quad y_2 = V(X) + E(X)^2$$

$$y_1 = np \quad y_2 = np(1-p) + (np)^2$$

$$\Rightarrow \begin{cases} \bar{y}_1 = np \\ \bar{y}_2 = np(1-p) + (np)^2 \end{cases} \Leftrightarrow \begin{cases} n = \frac{\bar{y}_1}{p} \\ \bar{y}_2 = \frac{\bar{y}_1}{p} \cdot p \cdot (1-p) + \frac{\bar{y}_1^2}{p^2} p^2 \end{cases}$$

$$\Leftrightarrow \begin{cases} n = \frac{\bar{y}_1}{p} \\ \bar{y}_2 = \bar{y}_1(1-p) + \bar{y}_1^2 \end{cases} \Leftrightarrow \begin{cases} n = \frac{\bar{y}_1}{p} \\ 1-p = \frac{\bar{y}_2 - \bar{y}_1^2}{\bar{y}_1} \end{cases} \Leftrightarrow$$

$$\Leftrightarrow \begin{cases} n = \frac{\bar{y}_1}{p} \\ p = 1 - \frac{\bar{y}_2 - \bar{y}_1^2}{\bar{y}_1} \end{cases} \Rightarrow \bar{p} = \frac{\bar{y}_1 - \bar{y}_2 + \bar{y}_1^2}{\bar{y}_1}$$

$$\bar{n} = \frac{\bar{y}_1^2}{\bar{y}_1 - \bar{y}_2 + \bar{y}_1^2}$$

Method of maximum likelihood

x_1, \dots, x_n i.i.d.

$$\theta = (\theta_1, \dots, \theta_k)$$

$$L(x_1, x_2, \dots, x_n; \theta) = \prod_{i=1}^n f_{x_i}$$

θ maximal likelihood estimator if it achieves the maximum for the likelihood function L

In order to do this, it suffices to find the value of θ for which $\ln L$ achieves its maximum.

We will find this value of θ by imposing the condition-

$$\frac{\partial \ln L}{\partial \theta_i} = 0$$

Exercise 2. Using the maximum likelihood method, estimate the parameters of the distribution of the characteristic X , if:

- (a) $X \sim \text{Unif}[a, b]$, $a < b$, for the sample data: $x_1, \dots, x_n \in [a, b]$.
- (b) $X \sim \text{Bino}(10, p)$, $p \in (0, 1)$, for the sample data: $x_1, \dots, x_n \in \{0, 1, \dots, 10\}$ not all zero and not all ten.
- (c) $X \sim \text{Exp}(\lambda)$, $\lambda > 0$, for the sample data: $x_1, \dots, x_n > 0$.
- (d) $X \sim \text{Geo}(p)$, $p \in (0, 1)$, for the sample data: $x_1, \dots, x_n \in \mathbb{N}$ not all zero.

- If $X \sim \text{Bino}(n, p)$, $n \in \mathbb{N}^*$, $p \in (0, 1)$, then:

$$X \left(\binom{n}{k} p^k (1-p)^{n-k} \right)_{k \in \overline{0, n}}$$

$$E(X) = np, V(X) = np(1-p)$$

$$(b) \quad X \sim \text{Bino}(10, p)$$

$$\begin{aligned} L(X_1, \dots, X_m; p) &= \prod_{i=1}^m f_{X_i} = \prod_{i=1}^m \underbrace{\binom{10}{X_i} p^{X_i} (1-p)^{10-X_i}}_{f_{X_i}} = \\ &= \underbrace{\left(\prod_{i=1}^m \binom{10}{X_i} \right)}_{=: K} \cdot p^{X_1 + \dots + X_m} \cdot (1-p)^{10m - (X_1 + \dots + X_m)} \end{aligned}$$

$$\begin{aligned} \ln L(X_1, \dots, X_m; p) &= \ln K + (X_1 + \dots + X_m) \cdot \ln p + \\ &+ (10m - (X_1 + \dots + X_m)) \cdot \ln(1-p) \end{aligned}$$

$$\frac{\partial \ln L}{\partial p} = \frac{X_1 + \dots + X_m}{p} - \frac{10m - (X_1 + \dots + X_m)}{1-p}$$

Using maximal likelihood we get:

$$\frac{X_1 + \dots + X_m}{p} = \frac{10m - (X_1 + \dots + X_m)}{1-p} \quad | \quad /m$$

$$\Rightarrow \frac{\bar{X}_1}{p} = \frac{10 - \bar{X}_1}{1-p} \Rightarrow \frac{\bar{X}_1}{10 - \bar{X}_1} = \frac{p}{1-p} \Rightarrow$$

$$\Rightarrow \frac{\bar{X}_1}{10 - \bar{X}_1 + \bar{X}_1} = \frac{p}{1-p+p} \Rightarrow p = \frac{\bar{X}_1}{10} \Rightarrow \bar{p} = \frac{1}{10} \cdot \bar{X}_1$$

$$(c) \quad X \sim \text{Exp}(\lambda)$$

$$(d) \quad X \sim \text{Geo}(p)$$

• If $X \sim \text{Exp}(\lambda)$, $\lambda > 0$ then:

$$f_X(x) = \lambda e^{-\lambda x} \cdot 1_{[0, \infty)}(x)$$

$$E(X) = \frac{1}{\lambda}, \quad V(X) = \frac{1}{\lambda^2}$$

$$(c) \quad X \sim \text{Exp}(\lambda)$$

$$\begin{aligned} L(x_1, \dots, x_n; \lambda) &= \prod_{i=1}^n f_{X_i} = \prod_{i=1}^n \left(\lambda \cdot e^{-\lambda x_i} \cdot 1_{[0, \infty)}(x_i) \right) = \\ &= \lambda^n \cdot \underbrace{\left(\prod_{i=1}^n e^{-\lambda x_i} \right)}_{= e^{-\lambda(x_1 + \dots + x_n)}} \cdot \underbrace{\prod_{i=1}^n 1_{[0, \infty)}(x_i)}_{=: K} = \end{aligned}$$

$$= K \lambda^n \cdot e^{-\lambda(x_1 + \dots + x_n)}$$

$$\Rightarrow \ln L(x_1, \dots, x_n; \lambda) = \ln K + n \ln(\lambda) - \lambda(x_1 + \dots + x_n)$$

$$\Rightarrow \frac{\partial \ln L}{\partial \lambda} = -(x_1 + \dots + x_n) + \frac{n}{\lambda}$$

$$\Rightarrow \frac{n}{\lambda} = x_1 + \dots + x_n \Rightarrow \lambda = \frac{n}{x_1 + \dots + x_n}$$

$$\Rightarrow \bar{\lambda} = \frac{n}{x_1 + \dots + x_n} = \frac{1}{\bar{x}} = \frac{1}{\bar{y}_1}$$

$$(d) \quad X \sim \text{Geo}(p)$$

• If $X \sim \text{Geo}(p)$, $p \in (0, 1)$, then

$$X \left(\binom{k}{p(1-p)^k} \right)_{k \in \mathbb{N}}$$

$$E(X) = \frac{1-p}{p}, \quad V(X) = \frac{1-p}{p^2}$$

$$L(x_1, \dots, x_n; p) = \prod_{i=1}^n f_{X_i} = \prod_{i=1}^n p(1-p)^{x_i} = p^n \cdot \prod_{i=1}^n (1-p)^{x_i} =$$

$$= p^n \cdot (1-p)^{x_1 + \dots + x_n}$$

$$\ln L = n \ln(p) + \ln(1-p) \cdot (x_1 + \dots + x_n)$$

$$\frac{\partial \ln L}{\partial p} = \frac{n}{p} - \frac{1}{1-p} \cdot (x_1 + \dots + x_n)$$

$$\Rightarrow \frac{n}{p} - \frac{1}{1-p} (x_1 + \dots + x_n) = 0$$

$$\Rightarrow \frac{n}{p} = \frac{1}{1-p} (x_1 + \dots + x_n)$$

$$\Rightarrow n(1-p) = p(x_1 + \dots + x_n)$$

$$\Rightarrow p(n + x_1 + \dots + x_n) = n$$

$$\Rightarrow p = \frac{n}{n + x_1 + \dots + x_n}$$

$$\Rightarrow \bar{p} = \frac{n}{n + x_1 + \dots + x_n} = \frac{1}{1 + \bar{x}} = \frac{1}{1 + \bar{y}_n}$$

$$(a) \quad X \sim \text{Unif}([a, b]) \Rightarrow f_X = \frac{1}{b-a} \cdot 1_{[a, b]}$$

$$L(x_1, \dots, x_n; a, b) = \prod_{i=1}^n \frac{1}{b-a} = \frac{1}{(b-a)^n}$$

The maximum of L is achieved when $b-a$ achieves its minimum

$$x_1, \dots, x_n \in [a, b] \Rightarrow \begin{aligned} \bar{a} &= \min \{x_1, \dots, x_n\} \\ \bar{b} &= \max \{x_1, \dots, x_n\} \end{aligned}$$