

A Bayesian Approach to Choosing the Best Physiochemical Properties in a Strain of Red Wine

by

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Introduction: Description of the problem.

The wine industry has flourished over the years, giving the opportunity to many renowned vineyards to become rich and famous. One such brand of wine is the “*Vinho Verde*”, that is aptly named after the demarcated region in northwestern Portugal in which it is grown, bottled and sold. Under this name that carries a long and successful history are lots of varieties of the wine and as it is natural, when there are a lot of varieties, there are also a lot of opinions. Wine connoisseurs say that some varieties of “*Vinho Verde*” are more successful than others. And even though, taste is a very subjective sense, the aim of this report is to try and identify the factors that may have contributed to the success of the red “*Vinho Verde*” and then utilize this model in order to create the best strain of red “*Vinho Verde*” as it is statistically possible.

The Red Wine Data

In order to do our analysis, we were provided with a set of data in an Excel file. However, as we are informed^[1], this selection of data was quite limited, due to privacy and logistic issues. Therefore, in total, our data is comprised of 12 variables (*columns*) and 1599 samples (*lines*). The first 11 columns include the physiochemical data, or in other words, our inputs. These data include objective tests performed on the wines, in order to determine their most common physio-chemical characteristics. The last column presents the sensory output. Here the wine experts were asked to taste the wine and afterwards, grade their samples on a scale from 0 to 10. Then, a median of at least 3 evaluations was taken and put in as a value in the last column of our Excel file. This last column consists of some subjective values, as the human taste is not easily quantifiable. Nevertheless, the results were a median of (at least three) experts, and therefore we shall trust their score and judgement.

Fortunately (as we can easily conclude in *R*)^[2] there were no missing data, something that will make this task a lot easier.

Therefore, for the sake of summarizing our data, we will create the following Table (*Table 1*), thus providing a quick statistical overview of the data given.

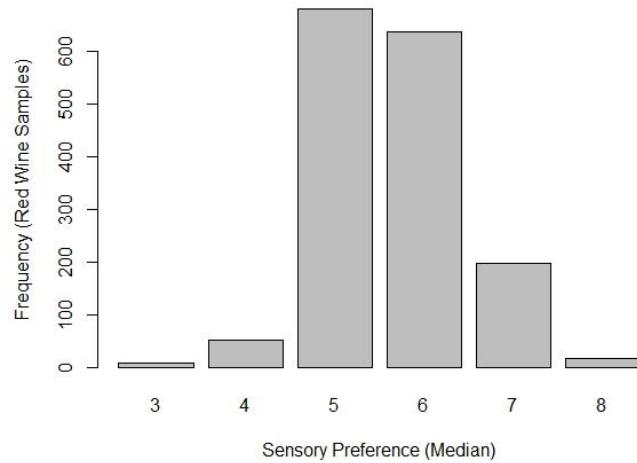
1. Information taken from: P. Cortez, A. Cerdeira, F. Almeida, T. Matos and J. Reis. Modeling wine preferences by data mining from physicochemical properties. In Decision Support Systems, Elsevier, 47(4): 547-553, 2009, pp. 6 - 7.

2. Read Appendix 2. (8. *The R Code*)

Table 1. The Physiochemical Data Statistics

Attribute (units) ^[3]	Min	Max	Mean
fixed acidity (g(tartaric acid)/dm³)	4.6	15.9	8.319637
volatile acidity (g(acetic acid)/dm³)	0.12	1.58	0.5278205
citric acid (g/dm³)	0	1	0.2709756
residual sugar (g/dm³)	0.9	15.5	2.538806
chlorides (g(sodium chloride)/dm³)	0.012	0.611	0.08746654
free sulfur dioxide (g/dm³)	0.001	0.072	0.01587492
total sulfur dioxide (g/dm³)	0.006	0.289	0.04646779
density (g/dm³)	0.99007	1.00369	0.9967467
pH	2.74	4.01	3.311113
sulphates (g(potassium sulphate)/dm³)	0.33	2	0.6581488
alcohol (% volume)	8.4	14.9	10.42298

Furthermore, we shall create a histogram of the sensory data (*Figure 1*). Here we can see that most of the median grades fall in the middle of the plot, while extreme values are not as common. It should be noted that there are only whole numbers, without any decimals.

Figure 1. Sensory Data Plot

3. Unit values were taken from: P. Cortez, A. Cerdeira, F. Almeida, T. Matos and J. Reis. Modeling wine preferences by data mining from physicochemical properties. In Decision Support Systems, Elsevier, 47(4): 547-553, 2009, pp. 23, Table 1.

More on Physiochemical Data in Appendix 1, pp. 9-11.

Model specification

For our analysis, we shall use the *multinomial logistic regression*^[4]. Since we are not given any specific information about our coefficients and how they might affect the scores of our judges, we will use low information priors on the betas (β). Therefore, our model will have the following general form:

$$y_i \sim \text{Multinomial}(p_{i,K})$$

for $i = 1599$, our number of observations and $K = 1:6$, our different categories scores (3-8). In this model, our reference category will be 6, which was altered in R to be the most frequent category (the one that had the most observations). Thus, the link function for each level will be:

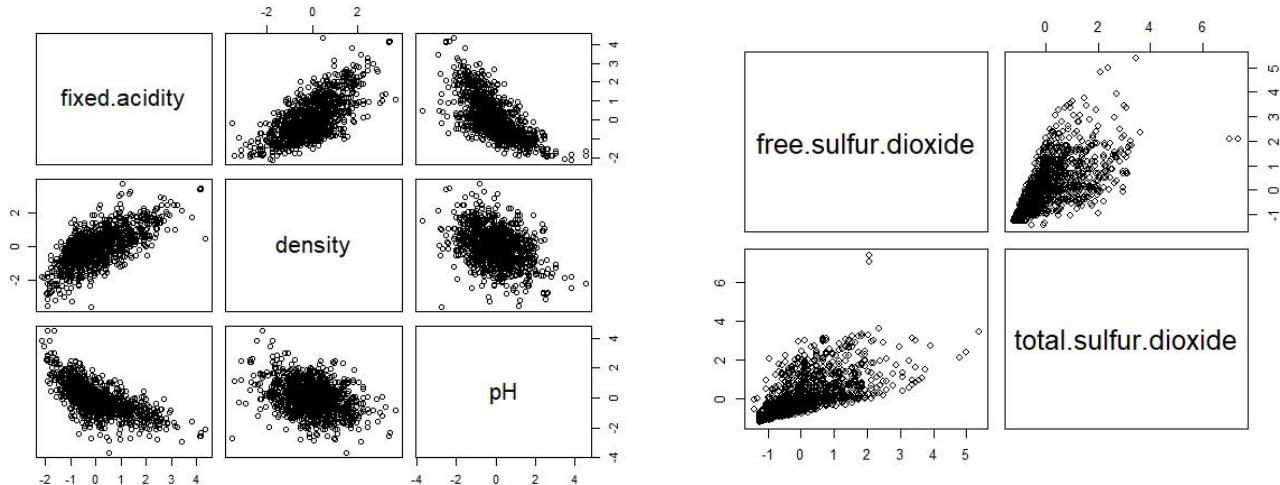
$$\text{Level1: } \ln\left(\frac{p_{i1}}{p_{i6}}\right) = \beta_{10} + \beta_{11} \cdot X_{i1} + \beta_{12} \cdot X_{i2} + \cdots + \beta_{111} \cdot X_{i11}$$

$$\text{Level2: } \ln\left(\frac{p_{i2}}{p_{i6}}\right) = \beta_{20} + \beta_{21} \cdot X_{i1} + \beta_{22} \cdot X_{i2} + \cdots + \beta_{211} \cdot X_{i11}$$

⋮

$$\text{Level5: } \ln\left(\frac{p_{i5}}{p_{i6}}\right) = \beta_{50} + \beta_{51} \cdot X_{i1} + \beta_{52} \cdot X_{i2} + \cdots + \beta_{511} \cdot X_{i11}$$

Initial analysis shows that there might be some form of collinearity between *fixed acidity*, *density* and *pH*, as well as *free* and *total sulfur dioxide*, just as the graphs below indicate:



We are not certain which levels they might affect, but we expect the majority of those trouble-making variables to not coexist, or to do so up to some degree, by the time we select our variables. It should also be noted that to help us with our convergence, we have centered our

explanatory variables, by subtracting the mean and dividing the standard deviation of the observations. Therefore, our final results will be bereft of units of measurement.

Variable Selection^[5]

Variable selection was done using the *BIC prior*, the *Empirical Bayes* method, the *g-prior* and the *hyper-g* prior on the coefficients of our model, by using both the uniform and beta-binomial priors. Most of these methods were run in both R (by using BAS package) and OpenBUGS software and they did, or at least they should produce similar results that differ only because of the MCMC component in OpenBUGS. Since beta-binomial is more flexible than the uniform on our model priors, because in a sense it lets us choose how parsimonious we want our final model to be, this analysis is based on beta-binomial priors for our model.

After running the models in R, we got the following DIC values:

MODEL	DIC VALUE
<i>Original Model</i>	3036.0
<i>Reduced Model (Variable Selection through BIC prior on betas with beta-binomial on the model)</i>	3063.0
<i>Reduced Model (Variable Selection through g-prior (of Liang et al.) prior on betas with beta-binomial on the model)</i>	3058.0
<i>Reduced Model (Variable Selection through hyper-g prior on betas with beta-binomial on the model)</i>	3017.0

The problem with the reduced models which used *BIC prior* and *g-prior* on the betas is that they were too strict with the selection of the statistically significant coefficients up to the point where their final models did not represent the data as well as they should and thus, had much higher DIC values than the original model.

Therefore, since it is the best option of the three, we ended up choosing *hyper-g* priors (as they were modified by Liang et al.) on our model's parameters, with *beta-binomial* priors on the model, in order to make it a bit more parsimonious, if possible. In other words, for the variable selection we had:

$$\beta_{jk} \sim N(0, G \cdot (X^T \cdot X)^{-1} \cdot \sigma^2)$$

for $j = 1, 2, 3, \dots, 11$, our number of estimates and $k = 1, 2, \dots, 5$, our number of levels.

which had the following hyper-prior specification:

```
a ← 3
bw ← a/2 - 1
w ~ dbeta( 1, bw )
G ← w/(1 - w)
```

After running our model in BAS package in R, so as to not do any unnecessary mistakes, according to our results, this was the final form of our model's link function, after disposing the superfluous variables:

$$\ln\left(\frac{p_{i1}}{p_{i6}}\right) = -8.8 + 1.7 \cdot Z_{i1} + 1.5 \cdot Z_{i2} + 0.6 \cdot Z_{i5} + 1.6 \cdot Z_{i6} - 3.7 \cdot Z_{i7} + 1.6 \cdot Z_{i9} + \\ - 1.7 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i2}}{p_{i6}}\right) = -3 + 0.5 \cdot Z_{i2} + 0.2 \cdot Z_{i4} - 0.7 \cdot Z_{i7} + 0.4 \cdot Z_{i9}$$

$$\ln\left(\frac{p_{i3}}{p_{i6}}\right) = 0.1 - 0.4 \cdot Z_{i2} - 0.2 \cdot Z_{i5} + 0.3 \cdot Z_{i6} - 0.6 \cdot Z_{i7} + 0.4 \cdot Z_{i10} + 0.8 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i4}}{p_{i6}}\right) = -2.1 + 0.5 \cdot Z_{i1} - 0.8 \cdot Z_{i2} + 0.3 \cdot Z_{i4} - 0.5 \cdot Z_{i5} + 0.3 \cdot Z_{i6} - \\ - 0.9 \cdot Z_{i7} - 0.4 \cdot Z_{i8} + 0.9 \cdot Z_{i10} + 1.4 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i5}}{p_{i6}}\right) = -6.5 - 0.7 \cdot Z_{i1} + 0.7 \cdot Z_{i3} - 0.1 \cdot Z_{i4} - 2 \cdot Z_{i5} + 0.4 \cdot Z_{i6} - \\ - 1.7 \cdot Z_{i7} - 1.1 \cdot Z_{i9} + 1.2 \cdot Z_{i10} + 2.4 \cdot Z_{i11}$$

And now, we are ready to produce some analysis and inference on the exponentiated values of these parameters.

Posterior analysis and interpretation of the results

To summarize our exponentiated parameters, we are giving the following table^[6]:

B01	B02	B03	B04	B05	B11	B14	B15	B21	B22
0.0	0.0	1.2	0.1	0.0	6.9	1.6	0.5	4.8	1.7
B23	B24	B35	B42	B44	B45	B51	B53	B54	B55
0.7	0.5	2.2	1.3	1.4	1.0	2.0	0.8	0.6	0.2
B61	B63	B64	B65	B71	B72	B73	B74	B75	B84
7.5	1.3	1.4	1.7	0.1	0.5	0.5	0.4	0.2	0.7
B91	B92	B95	B103	B104	B105	B111	B113	B114	B115
6.3	1.5	0.4	1.5	2.4	3.4	0.2	2.2	4.1	11.1

So let us see what they mean (well, at least for level 1):

6. The result of this table came from the final model, which can also be found in Appendix 1, pp.122-124.

For level 1:

$B01 = \exp\{\beta_{01}\} \cong 0$: If we do not include any of our physiochemical characteristics in this level at all, then there is almost (**it is not exactly zero**) no chance that our wine will get a score of 3, but rather that of 5. It is kind of nonsensical as a result, but maybe the other factors that are not included in this level are enough to make this distinction between these two scores.

$B11 = \exp\{\beta_{11}\} = 6.9$: Is saying to us that should we keep all the other variables of this level fixed and simultaneously increase the value of fixed acidity by 1 (no units of measurement, remember?), then our odds of receiving a score of 3, rather than 5 will increase by a huge 590%.

$B21 = \exp\{\beta_{21}\} = 4.8$: Tells us that should we increase the value of volatile acidity by 1 unit (no units of measurement), all while keeping the other variables constant, then the odds of being in category 1 that gets a median score of 3, will increase drastically by 380%, instead of being in category 6, which gets a median score of 5.

$B51 = \exp\{\beta_{51}\} = 2$: Tells us that by keeping all other variables of this level fixed, but also increasing the value of chlorides by 1 (no units of measurement), our odds of receiving a score of 3 instead of 5 will increase by 100%.

$B61 = \exp\{\beta_{61}\} = 7.5$: Is an indication that should we keep all other variables of this level constant in their values, while increasing the value of free sulfur dioxide by 1 (no units of measurement), our odds of receiving a score of 3 instead of 5 will increase by an enormous 650%.

$B71 = \exp\{\beta_{71}\} = 0.1$: This value points out that if we keep all the other variables of this level fixed and simultaneously increase the value of total sulfur dioxide by 1 (no units of measurement), then the odds of getting a score of 3 rather than 5 will decrease by 90%.

$B91 = \exp\{\beta_{91}\} = 6.3$: Tells us that if we keep all the other variables of this level fixed, all the while increasing the value of pH by 1 (no units of measurement), then the odds of getting a score of 3 rather than 5 will increase by an astonishing 530%.

$B111 = \exp\{\beta_{111}\} = 0.2$: Is saying to us that should we keep all the other variables of this level fixed, but also increase the value alcohol by 1 (no units of measurement), then the odds of getting a score of 3 rather than 5 will diminish by 80%.

And this analysis follows in the same manner for all other levels^[7]. So, what can we surmise from these results?

1. It appears that alcohol as well as sulphates appealed to the taste of our judges and in fact those were the two factors that could, in moderation, make or break our wine's score.
2. There was no good reason to keep a high level of volatile acid in our wine, as it usually made our wine's grade drop to a lower level. In fact, it is not even included in level 5.
3. The same thing can be said about chlorides, as higher quantities of this physiochemical characteristic only make our wine's score drop to a lower value.

7. For more analytical results, please check: Appendix 1, pp.124-129.

4. This can be also extended to pH. It is preferable to keep low values of this physiochemical characteristic in our wines.
5. Free sulfur dioxide adds to the taste, but should be kept in moderation in our wines as it can potentially change our final score to the lowest.
6. Density appears only in level 4 and ruins the chances for our wine to get a higher score, so it should be kept at low values in our wine.
7. Total sulfur dioxide does not contribute to any grade whatsoever, apart from maybe that of 5. Thus, it should be kept at low levels in our wines.

And apart from those variables, the others of our model either contributed or took away from the scores of each level.

Another good question would be, what if we wanted to produce only wines that scored 6 and above? What good physiochemical characteristics should we pay special attention to?

In order to answer that question, we ran a simple binomial model^[8], which divided our data into two categories (5 and below, versus 6 and above). Here we used a stricter approach and considered a BIC prior on the betas, with a beta-binomial on the model. The results showed us that free sulfur dioxide, sulphates and alcohol worked in favour of our score and thus should be kept in high values in our model, while volatile acidity, chlorides and total sulfur dioxide should receive values that are as low as possible.

This result along with the one above should work together, when we “fine tune” the physiochemical characteristics in order to produce the best tasting red variant of “*Vinho Verde*”, so as to know what variables should be taken with greater care.

Final comments and conclusion

All in all, this was our suggestion based on the analysis done. Making a fine bottle of wine is an art in itself, put too much alcohol in and you will have poison, put too little and you will have grape juice. And even though those physiochemical characteristics are very important to have in small or large quantities in our wines, they are but a small fraction of what makes a wine unique and successful. The climate, the type of grapes, the soil, the way it is bottled, all these and more play an important role in the final product and its score.

Therefore, this recipe is not going to guarantee us a good quality of wine, because it is so much more than that. It does however provide us with a good indication that we are on the right (or the wrong) track to making a tasteful red “*Vinho Verde*” that people can enjoy. And at the end of the day, that is what really matters.

8. More details in: Appendix 1, pp.129-137.

APPENDIX 1

(A more Analytical Solution with Tables & Diagrams)

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The Data:

We are given a set of data that are related to the red variant of the Portuguese "Vinho Verde" wine. In total, the researchers took 1599 samples and collected their findings in an Excel file.

The file itself consists of 12 columns and 1599 rows. The first 11 columns are the most common physiochemical characteristics of each sample and we are informed that their collection was done in a laboratory that specializes in that procedure and thus can be trusted. These shall be our explanatory variables, and so by definition, they will try to give an explanation, on why the score (our response variable) was given that particular value.

The last column (column 12) contains the response variable of the data set, for each of the 1599 samples that were taken. This value is a median of at least three experts that have tasted a variety of that wine and given a certain score.

The Explanatory Variables:

For our convenience, the table below gives us an overview of all the 11 explanatory variables. It should be noted that the units of the variable *free sulfur dioxide*, as well as the units of *total sulfur dioxide* were transformed from mg/dm^3 to g/dm^3 .

Table 1. Physiochemical Characteristics

Χαρακτηριστικό (μονάδες μέτρησης)	Min	Max	Μέσος
fixed acidity (<i>g(tartaric acid)/dm³</i>)	4.6	15.9	8.319637
volatile acidity (<i>g(acetic acid)/dm³</i>)	0.12	1.58	0.5278205
citric acid (<i>g/dm³</i>)	0	1	0.2709756
residual sugar (<i>g/dm³</i>)	0.9	15.5	2.538806
chlorides (<i>g(sodium chloride)/dm³</i>)	0.012	0.611	0.08746654
free sulfur dioxide (<i>g/dm³</i>)	0.001	0.072	0.01587492
total sulfur dioxide (<i>g/dm³</i>)	0.006	0.289	0.04646779
density (<i>g/dm³</i>)	0.99007	1.00369	0.9967467
pH	2.74	4.01	3.311113
sulphates (<i>g(potassium sulphate)/dm³</i>)	0.33	2	0.6581488
alcohol (% volume)	8.4	14.9	10.42298

A bit more on those parameters:

- **Fixed Acidity (*g(tartaric acid)/dm³*):** Fixed acids are fruit acids (nonvolatile) that are organic to grapes. This value represents the most acids, which are contained in a bottle of wine that are constant, meaning that when a person opens the bottle, these acids do

not evaporate as easily as others. Here, it is a continuous variable, which is measured in grams of tartaric acid, over cubic decimeter. It takes values from 4.6, to 15.9 and has a mean of 8.319637.

- **Volatile Acidity ($g(acetic\ acid)/dm^3$):** Volatile acidity is an unstable acid formed by dissolving carbon dioxide in water. It is the basis of carbonated beverages and is related to the carbonate group of compounds. This value represents the amount of acetic acid in a bottle of wine. Should this value be high enough, it gives the wine the unpleasant taste of vinegar. In our data, it is a continuous variable, which is measured in grams of acetic acid, per cubic decimeter. It takes values from 0.12, to 1.58 and has a mean value of 0.5278205.
- **Citric Acid (g/dm^3):** In small quantities, it adds a sense of “freshness” to the wines. It is used in many foods, confections, and soft drinks to improve their stability in natural metals. Here, it is a continuous variable, which is measured in grams of citric acid, over cubic decimeter. It takes values from 0, to 1 and has a mean value of 0.2709756.
- **Residual Sugar (g/dm^3):** It is the amount of residual sugar that remains after the procedure of fermentation has stopped. Usually, it takes values between $1\ g/dm^3$ and $45\ g/dm^3$, with the latter value representing the sweet wines. In our data, it is a continuous variable, which is measured in grams of residual sugar, per cubic decimeter. It takes values from 0.9, to 15.5 and has a mean value of 2.538806. Thus, we can safely say that these samples did not contain any sweet wines.
- **Chlorides ($g(sodium\ chloride)/dm^3$):** Is the amount of salt, which is contained in a wine. Here, it is a continuous variable, which is measured in grams of sodium chloride, per cubic decimeter. It takes values from 0.012, to 0.611 and has a mean value of 0.08746654.
- **Free Sulfur Dioxide (g/dm^3):** Is a chemical component that is used to shield the wine from undesirable microbes and unwanted oxidation, which spoil the good taste of the wine. Here we have altered its unit of measurement from mg/dm^3 to g/dm^3 . Thus, in our data, it is a continuous variable, which is measured in grams of free sulfur dioxide, per cubic decimeter. It takes values from 0.001, to 0.072 and has a mean value of 0.01587492.
- **Total Sulfur Dioxide (g/dm^3):** This physicochemical parameter contains the previous variable (meaning, free sulfur dioxide), as well as the bound form of sulfur dioxide in its value. Sulfur dioxide is mostly difficult to detect through the common senses, however in large quantities (greater than $0.05\ g/dm^3$), it becomes perceptible through its odor and its taste. Once again, we have changed the unit of measurement from mg/dm^3 to g/dm^3 . Thus, in our data, it is a continuous variable, which is measured in grams of total sulfur dioxide, per cubic decimeter. It takes values from 0.006, to 0.289 and has a mean

value of 0.04646779. Therefore, we can see that there were samples where it was perceptible and perhaps distasteful.

- **Density (g/dm^3):** It measures the density of the wine as a liquid. Dry wines have less density than sweet ones. Density is defined in a qualitative manner as a measure of the relative ‘heaviness’ of an object with constant volume. Here, it is a continuous variable, which is measured in grams over cubic decimeter. It takes values from 0.99007, to 1.00369 and has a mean value of 0.9967467. Hence, here there are only dry wines.
- **pH:** This parameter refers to the degree of acidity in our wine. At first it took values on a scale from 0 to 14, but today this scale has been expanded. For the red wine a $pH < 3.4$ is recommended. In our data, it is a continuous variable, which is measured in pH . It takes values from 2.74, to 4.01 and has a mean value of 3.311113.
- **Sulphates ($g(\text{potassium sulphate})/dm^3$):** It is an additive mainly used in foods. It helps the wine to preserve its “freshness”, while also improving the taste. Just like the sulfur dioxide, this also helps shielding the wine from harmful microbes and oxidization. They are also known to cause the annoying headache, that people commonly know as hangover. In these data, it is a continuous variable, which is measured in grams of potassium sulphate per cubic decimeter. It takes values from 0.33, to 2 and has a mean value of 0.6581488. Therefore, if we consider that the maximum limit in the USA is $0.35\ g/dm^3$, then we can certainly say that although the wines had a good taste, they must have left the wine experts with one of the greatest hangovers in the history of hangovers.
- **Alcohol (% volume):** This parameter measures the percentage of alcohol in % of volume. This is what separates the wines from grape juice. In our data, it takes values from 8.4% to 14.9%, while having a mean value of 10.42298%.

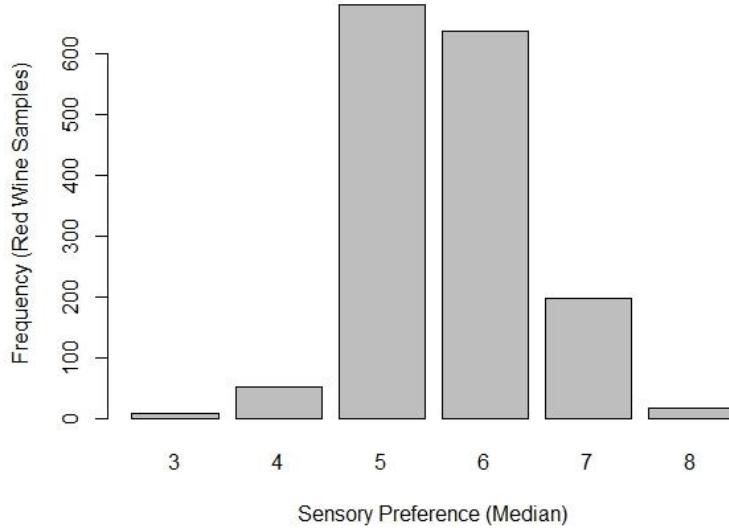
Response Variable:

The wine connoisseurs did not know which wine brand of “Vinho Verde” they were sampling each time. When they tried one wine and graded it according to their taste, this did not give them any information about the previous or the next sample. Thus, we can safely assume that those trials were independent to each other.

All in all, the sampling procedure would unfold in the following manner:

At least three experts would taste a certain type of wine, which had the same physiochemical characteristics and then evaluated it, based on their preference and knowledge in the field of wines. Then the researchers would collect those scores, for those specific physiochemical parameters and take a median value of at least three of them. That is how each line of column 12 was created. A histogram of our response variable would be:

Diagram 1. Sensory preference



As we can see in this histogram, the categories are positive integers. There are no intermediate values like 4.5, or 6.856 for instance. Most of the median values are either scored as 5 or 6, while extreme values of 3 and 8 do not have a lot of observations. Thus, the majority of those medians refer to medium tasting wine. Furthermore, even though the scores should range from 1 to 10, we only have observations for values 3 to 8 and so in total we have 6 observed categories, instead of 11.

The Model (Multinomial Logistic Regression):

After almost two months of suffering and agony, due to the striking the inexperience of the author, the final model to which we concluded was the Multinomial Logistic Regression. There were of course other models that were considered, but then rejected before we ended up with this one.

The most obvious of all the models was at first the Ordinal Logistic Regression (Proportional-Odds Logit Model), because our response variable is ordered, meaning that a grade of 8 denotes a better tasting wine, than if it received a grade of 7, or 6, or God forbid 2. However, this model was rejected, because the proportional-odds (parallel lines) assumption was not in effect. So, for example if we ran that in SPSS, we would get:

Test of Parallel Lines ^a				
Model	-2 Log Likelihood	Chi-Square	df	Sig.
Null Hypothesis	3074.767			
General	2919.061 ^b	155.707 ^c	44	.000

The null hypothesis states that the location parameters (slope coefficients) are the same across response categories.

a. Link function: Logit.

b. The log-likelihood value cannot be further increased after maximum number of step-halving.

c. The Chi-Square statistic is computed based on the log-likelihood value of the last iteration of the general model. Validity of the test is uncertain.

Obviously this is a violation, which means that the coefficients of the model are not the same for each level, and so instead of having something more humane, like:

$$\ln\left(\frac{\pi_1}{\pi_2 + \pi_3 + \dots + \pi_K}\right) = a_1 + \beta_1 \cdot X_1 + \beta_2 \cdot X_2 + \dots + \beta_p \cdot X_p$$

$$\ln\left(\frac{\pi_1 + \pi_2}{\pi_3 + \pi_4 + \pi_5}\right) = a_2 + \beta_1 \cdot X_1 + \beta_2 \cdot X_2 + \dots + \beta_p \cdot X_p$$

⋮

$$\ln\left(\frac{\pi_1 + \pi_2 + \dots + \pi_4}{\pi_5}\right) = a_5 + \beta_1 \cdot X_1 + \beta_2 \cdot X_2 + \dots + \beta_p \cdot X_p$$

Which has only 16 total variables ($p = 11$ here), instead we have:

$$\text{Level 1} = a_1 + \beta_{11} \cdot X_1 + \beta_{12} \cdot X_2 + \dots + \beta_{1p} \cdot X_p$$

$$\text{Level 2} = a_2 + \beta_{21} \cdot X_1 + \beta_{22} \cdot X_2 + \dots + \beta_{2p} \cdot X_p$$

⋮

$$\text{Level 5} = a_5 + \beta_{51} \cdot X_1 + \beta_{52} \cdot X_2 + \dots + \beta_{5p} \cdot X_p$$

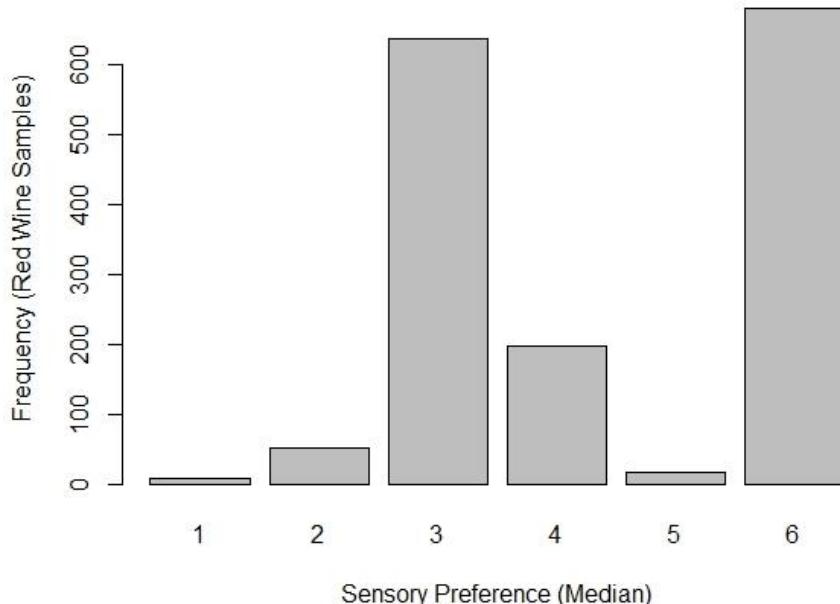
That are a total of 60 variables to estimate. But thankfully there are ways to do that as well. And just like good wine, these ways take time.

So, the proportional-odds model does not work, what can one do then? Well, the first thing that should come to mind, would be a Multinomial Logistic Regression, which is a more general case of the model that did not fit our data.

This model does not suppose an order between the levels, nor does it impose a common value for our covariates. Therefore, in our wine data set, we will ignore that the value 2 is greater than 1, or that 5 is better than 4, 3 and 2, for instance, and instead treat them as independent categories.

The idea of how this model works is it separates the multinomial model of K categories, into a series of (binary) logistic $K - 1$ models, for each pair of response categories. In order to compare the log-odds model, we need a common baseline category which will be our point of reference.

Usually, the most “meaningful” category is chosen to be the reference category, or the most frequent category, or either the first or the last category of our response variable. Almost any category, as long as it is not an outlier or something rare, that has very few observations. In our data, the most frequent category is the third one and so it will be chosen to be the reference category. What we shall do now, is switch their order and place the reference category in the last position, which will help us later in the programming part. Therefore, if we see a graph after we organize our data, we will have:



Notice that our categories are individual observations and are not grouped, so the response is a variable, which denotes the level (so 1, 2, 3, etc.) at which we currently find ourselves for each of the 1599 observations. Therefore, we have:

$$y_i \sim \text{Categorical}(p_{i,K})$$

or equally

$$y_i \sim Multinomial(p_{i,K}, N_i = 1)$$

In total there are 11 explanatory variables and six categories (3-8), so if we make the sixth category our reference group (because it has the largest frequency in the group), then the baseline-category probability will be:

$$p_{i6} = \frac{1}{1 + \sum_{k=1}^5 \exp\{\beta_{0k} + \beta_{1k} \cdot X_{i1} + \dots + \beta_{11k} \cdot X_{i11}\}}$$

While the other probabilities will be:

$$p_{ij} = \frac{\exp\{\beta_{j0} + \beta_{j1} \cdot X_{i1} + \dots + \beta_{j11} \cdot X_{i11}\}}{1 + \sum_{k=1}^5 \exp\{\beta_{k0} + \beta_{k1} \cdot X_{i1} + \dots + \beta_{k11} \cdot X_{i11}\}}, \quad \text{for } j = 1, 2, \dots, 5$$

Or alternatively, for each level we will have:

$$\ln\left(\frac{p_{i1}}{p_{i6}}\right) = \beta_{10} + \beta_{11} \cdot X_{i1} + \beta_{12} \cdot X_{i2} + \dots + \beta_{111} \cdot X_{i11}$$

$$\ln\left(\frac{p_{i2}}{p_{i6}}\right) = \beta_{20} + \beta_{21} \cdot X_{i1} + \beta_{22} \cdot X_{i2} + \dots + \beta_{211} \cdot X_{i11}$$

⋮

$$\ln\left(\frac{p_{i5}}{p_{i6}}\right) = \beta_{50} + \beta_{51} \cdot X_{i1} + \beta_{52} \cdot X_{i2} + \dots + \beta_{511} \cdot X_{i11}$$

As for the last probability:

$$p_{i6} = 1 - \sum_{j=1}^5 p_{i5}$$

Since all these probabilities will eventually sum to 1.

Running the original model:

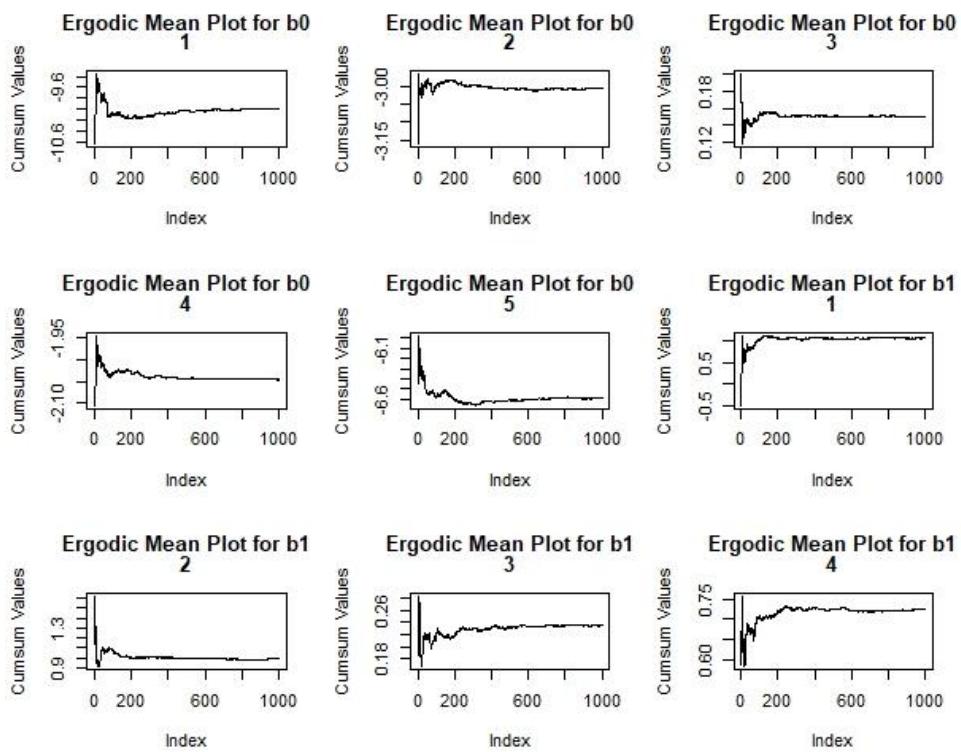
Before we start doing variable selection, we should do a trial run and see whether the model we chose converges for the parameters that we wish to estimate. Here of course we want to

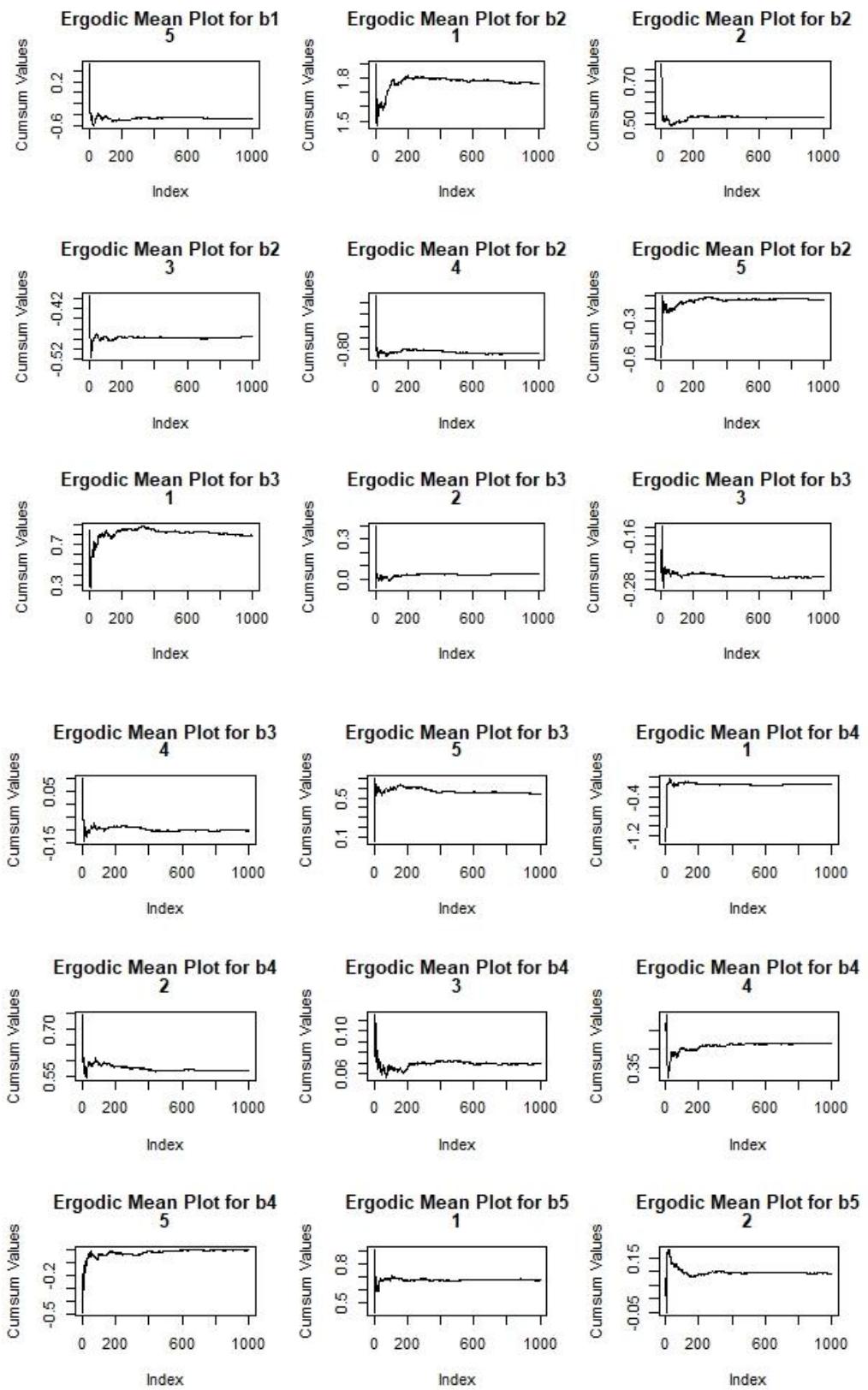
estimate the coefficients of the physiochemical data, but we are given no information as to a set of values that they should or should not take. Therefore, we are going to be vague in our priors chosen. By setting the mean of all β s equal to zero (because it is the center of the real numbers set \mathbb{R} that goes from $-\infty$ to ∞) and the precision to a low value (here 0.0001 or 1.0E-4), we state that we have low information priors on those coefficients and so we shall let the data determine their real values. Also, before running our program, we shall center the values of the physicochemical data, to assist us in both the speed of the convergence and autocorrelation problems.

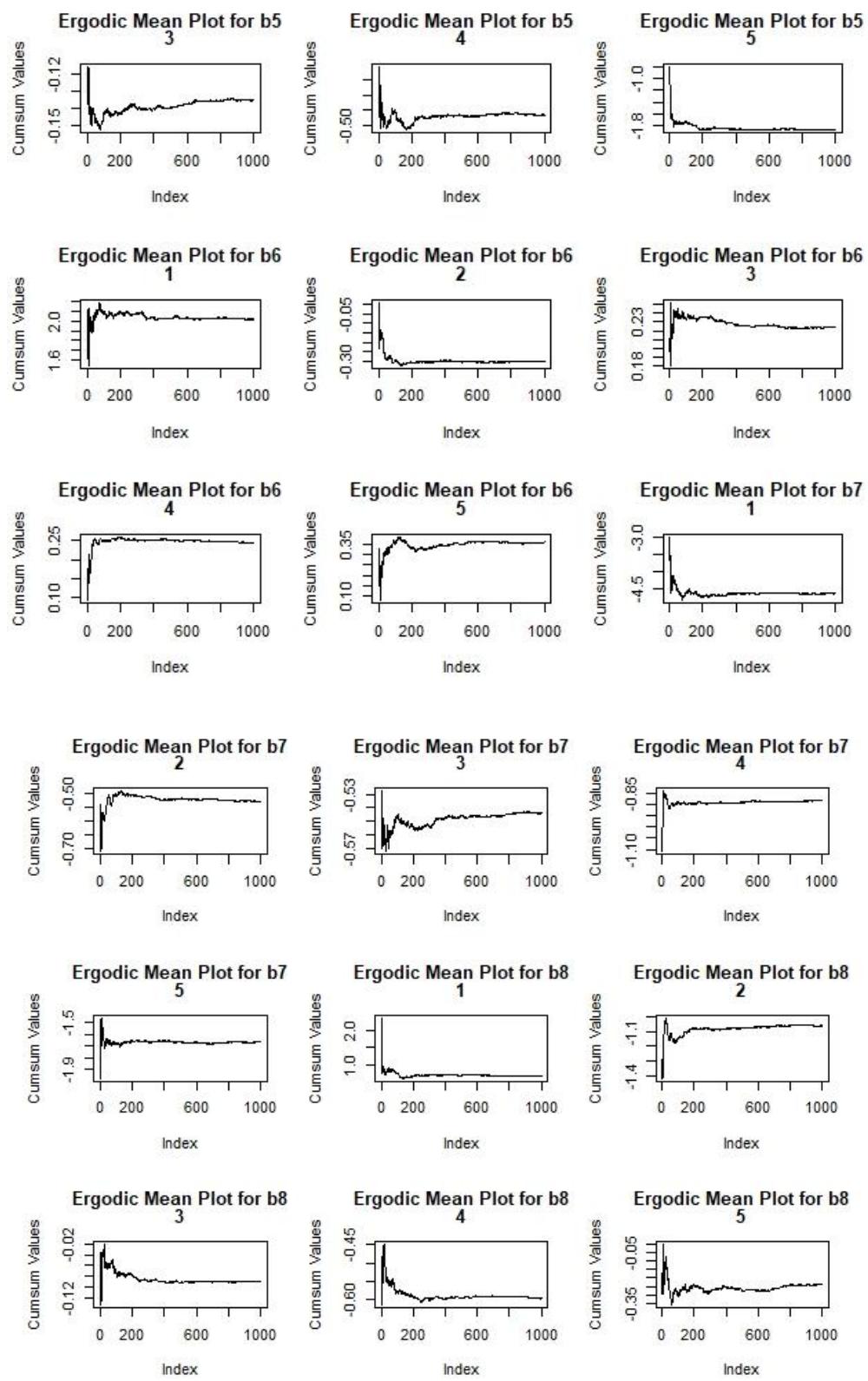
Having done some initial runs, the plots showed some significant autocorrelation problems up until 30 lags. Therefore, we will also apply a thin of 30, to rectify this problem. This procedure, even though we have only 1 chain is quite slow, and therefore we shall do the following: We will discard the first 1000 iterations as a burn-in period and then run our program for a total of 30000 iterations, which according to the thinning value that we have set, should return 1000 samples to us.

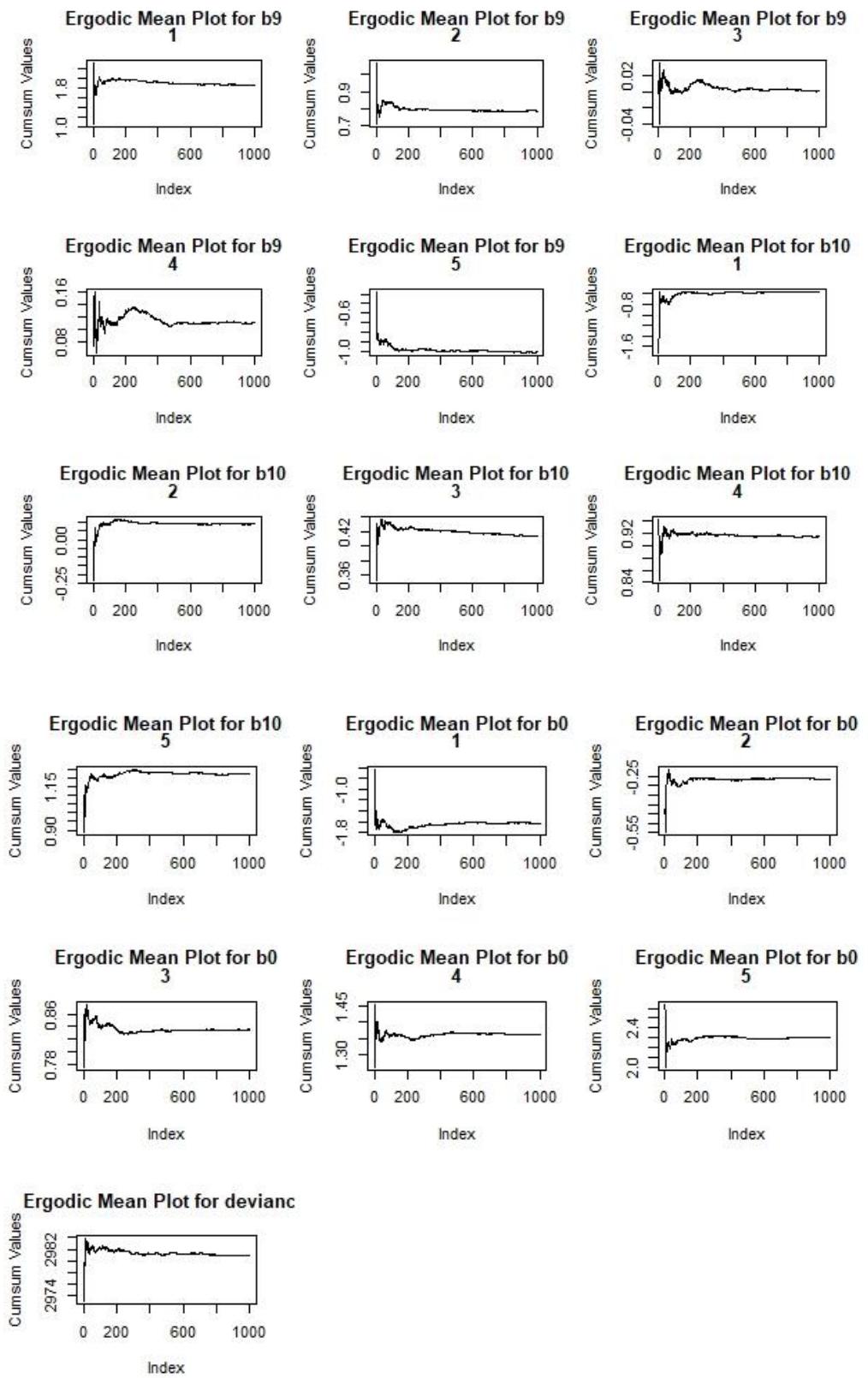
Now having done all that, we run our program based on the *multinomial.txt* model and get the following results.

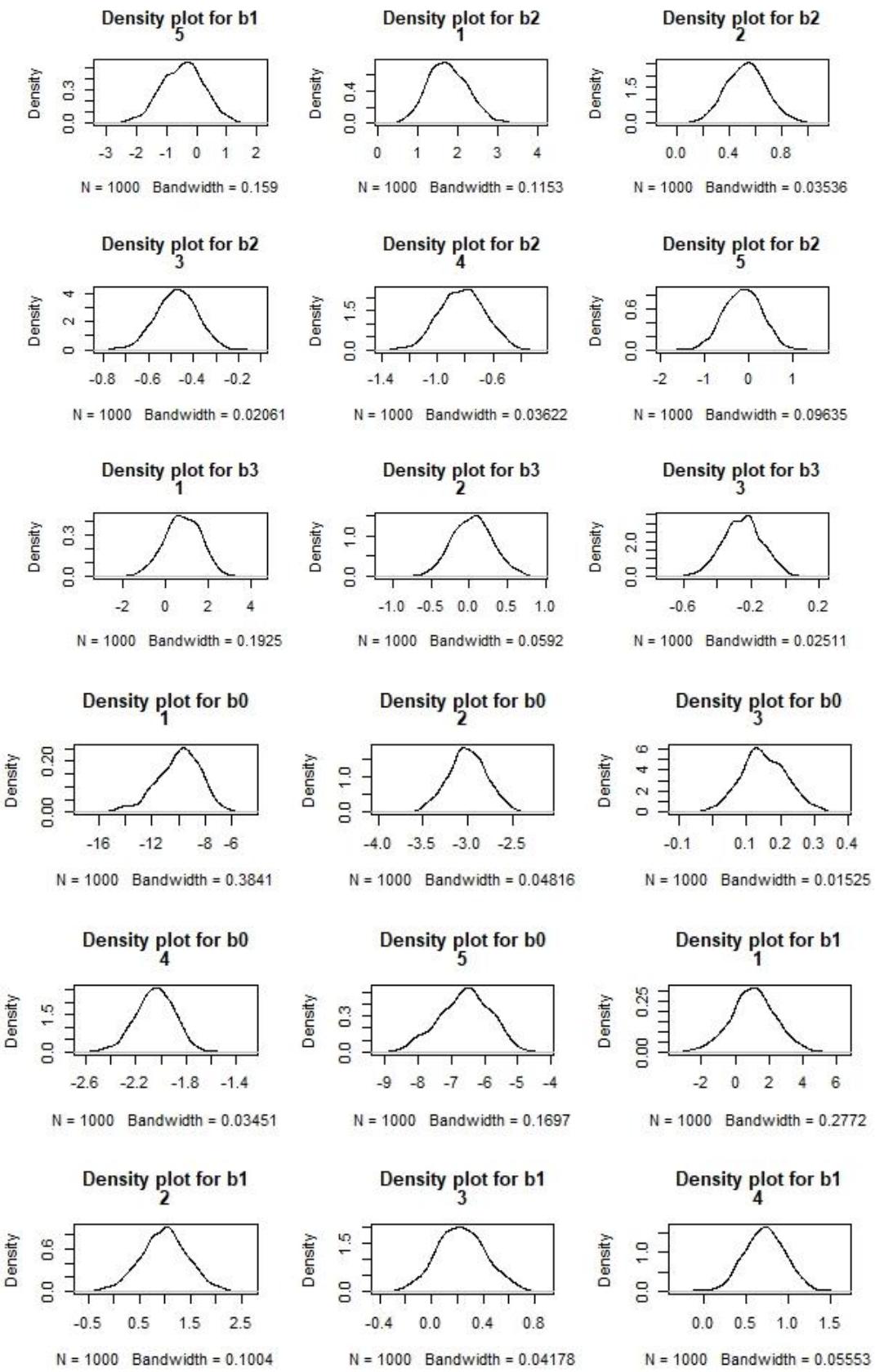
Convergence plots on the multinomial logistic regression (original model):

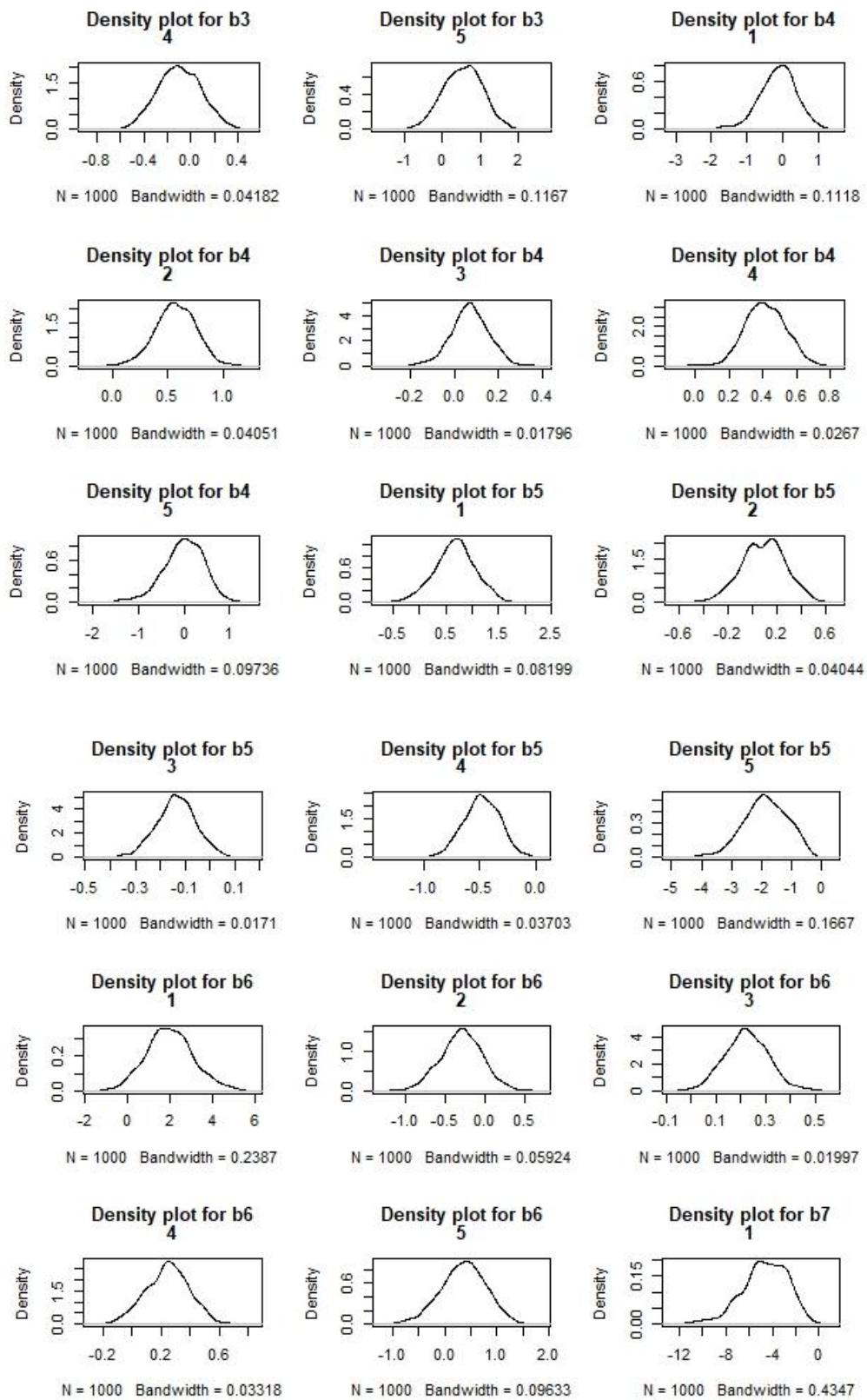


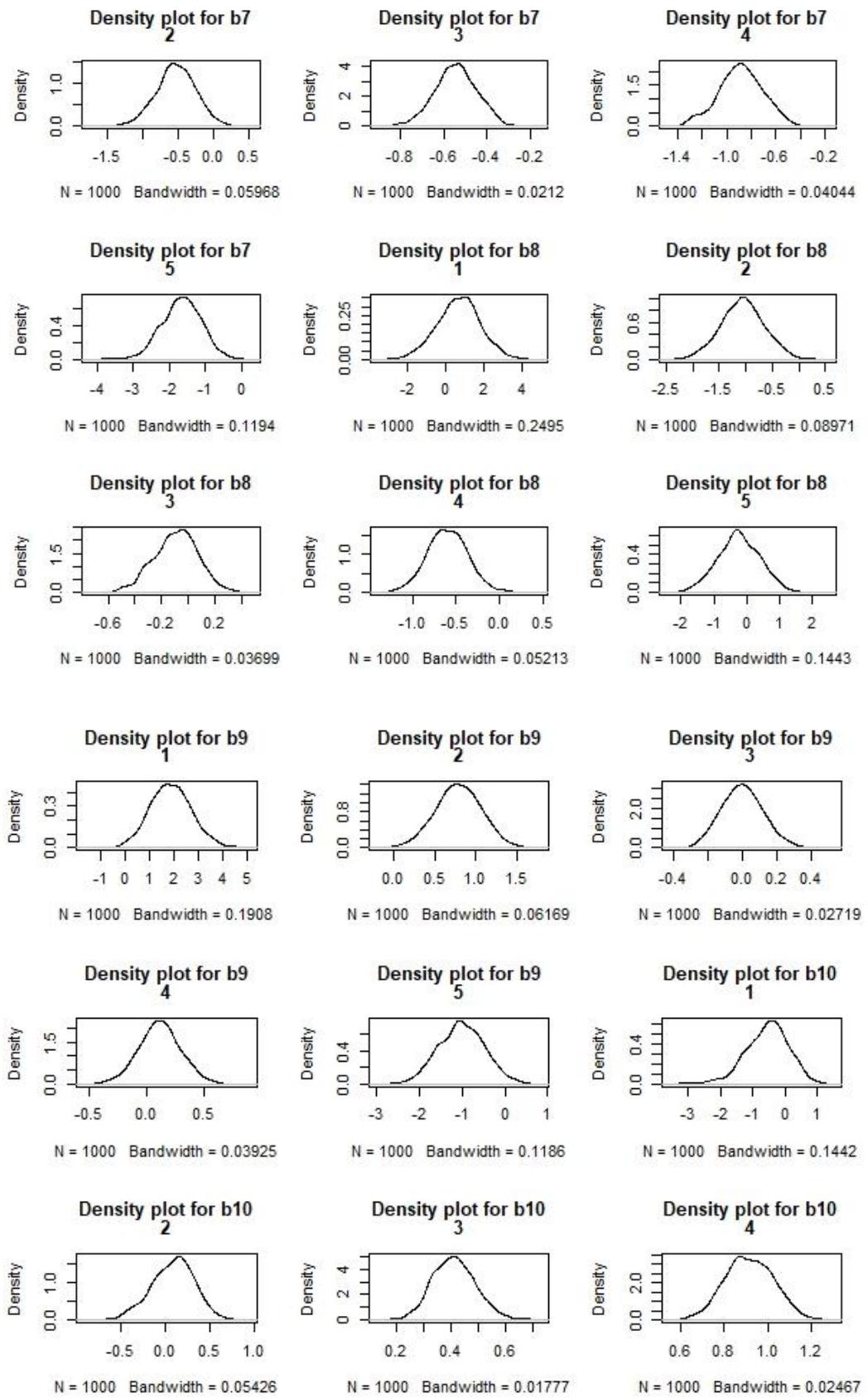


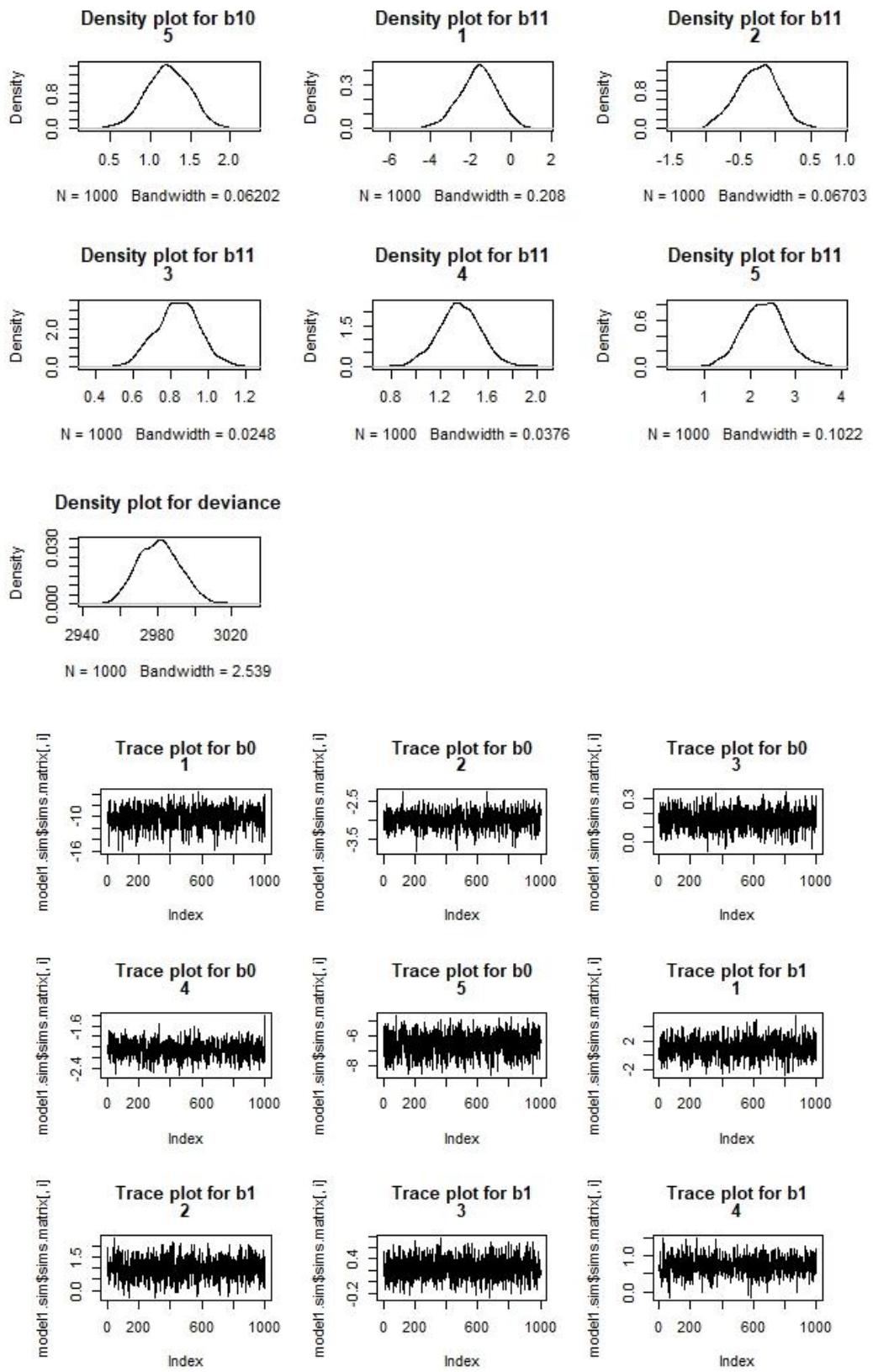


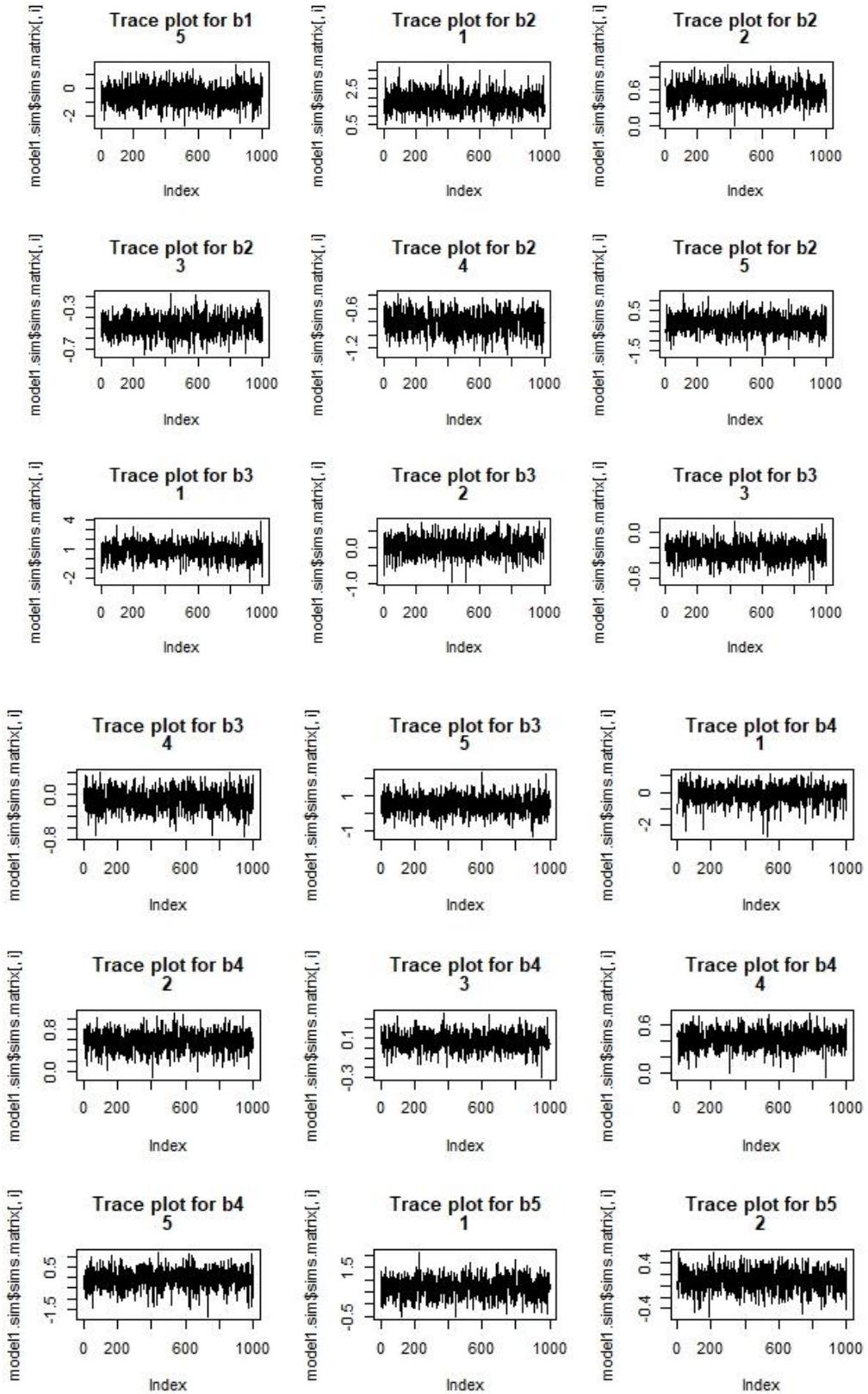


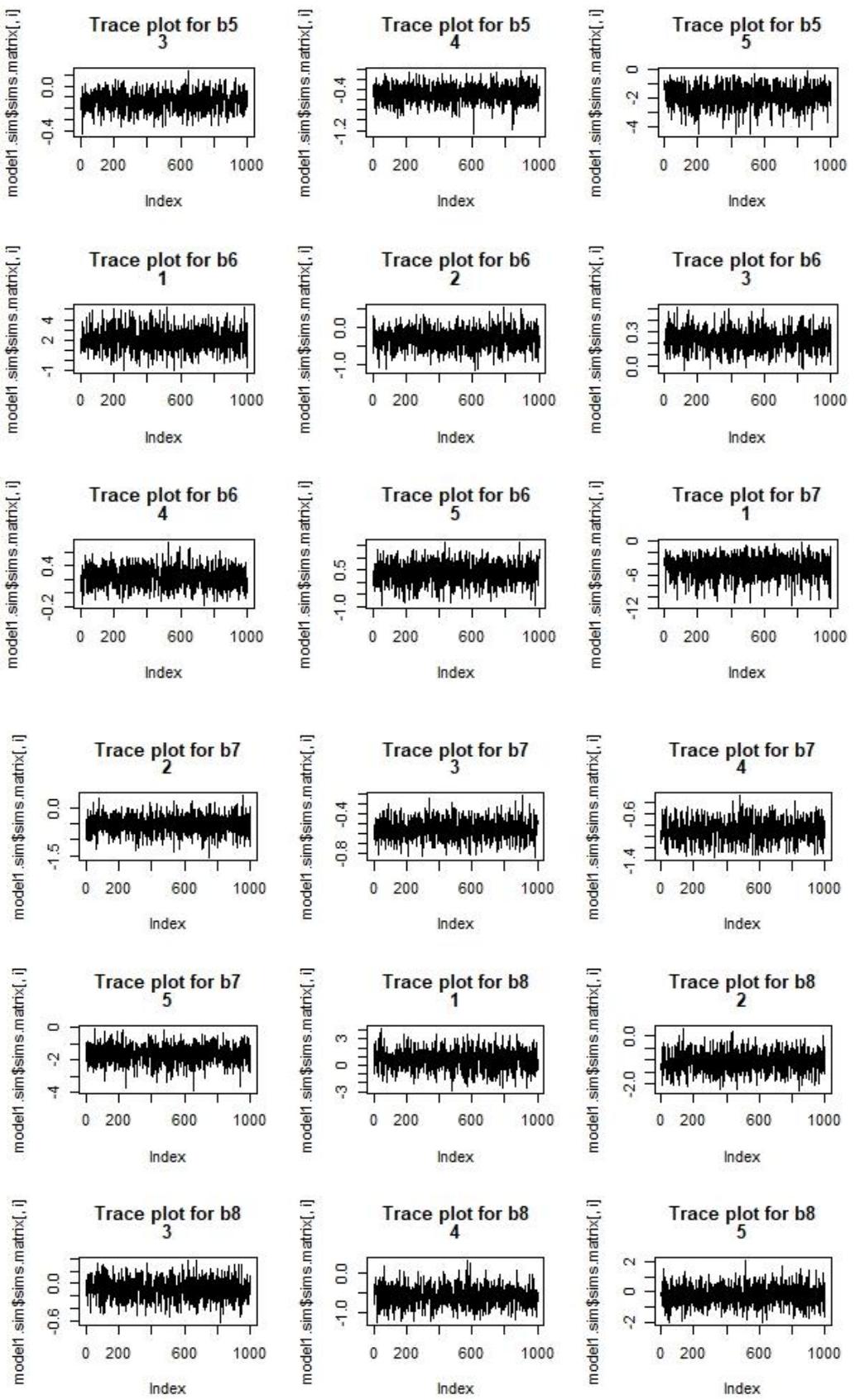


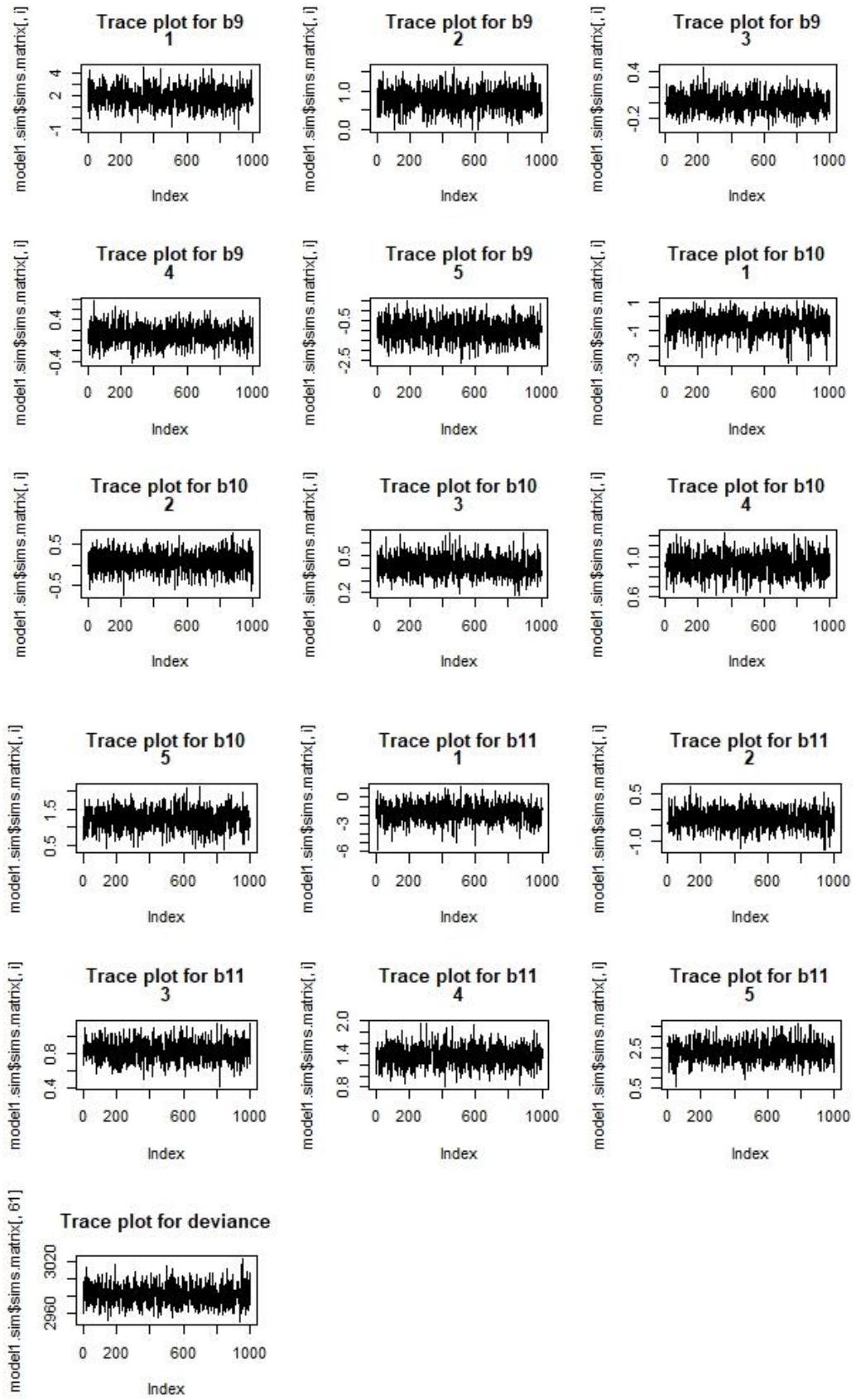


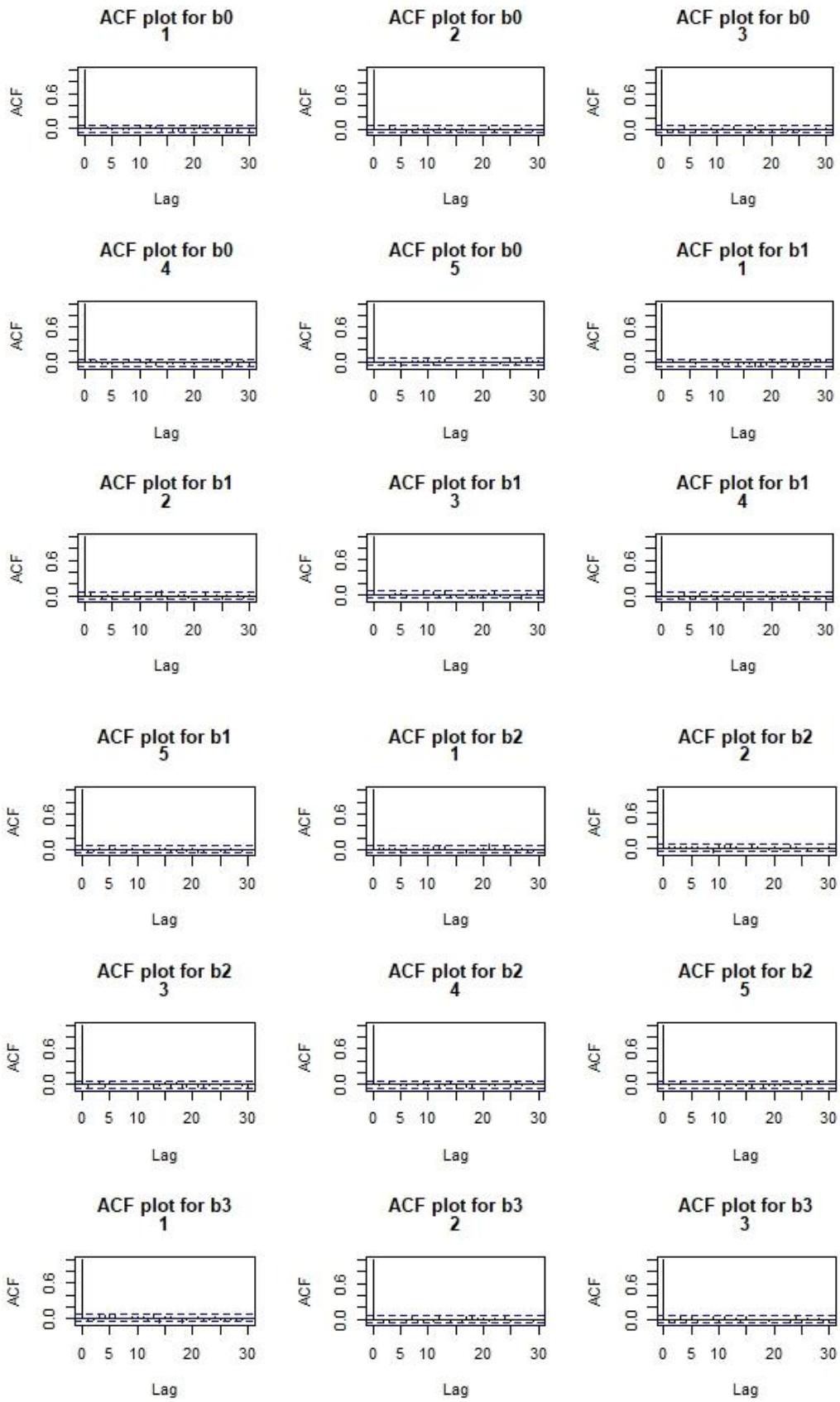


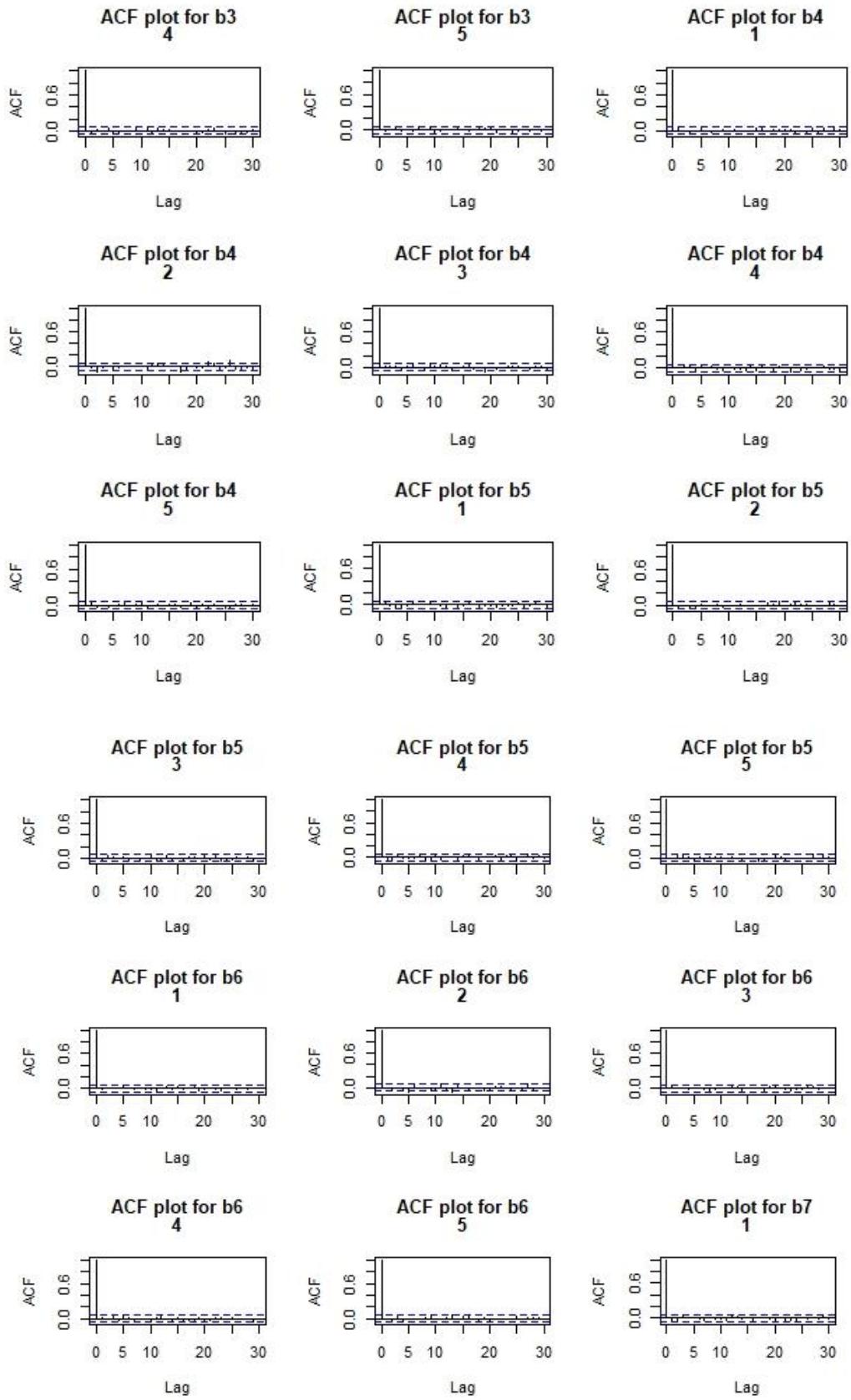


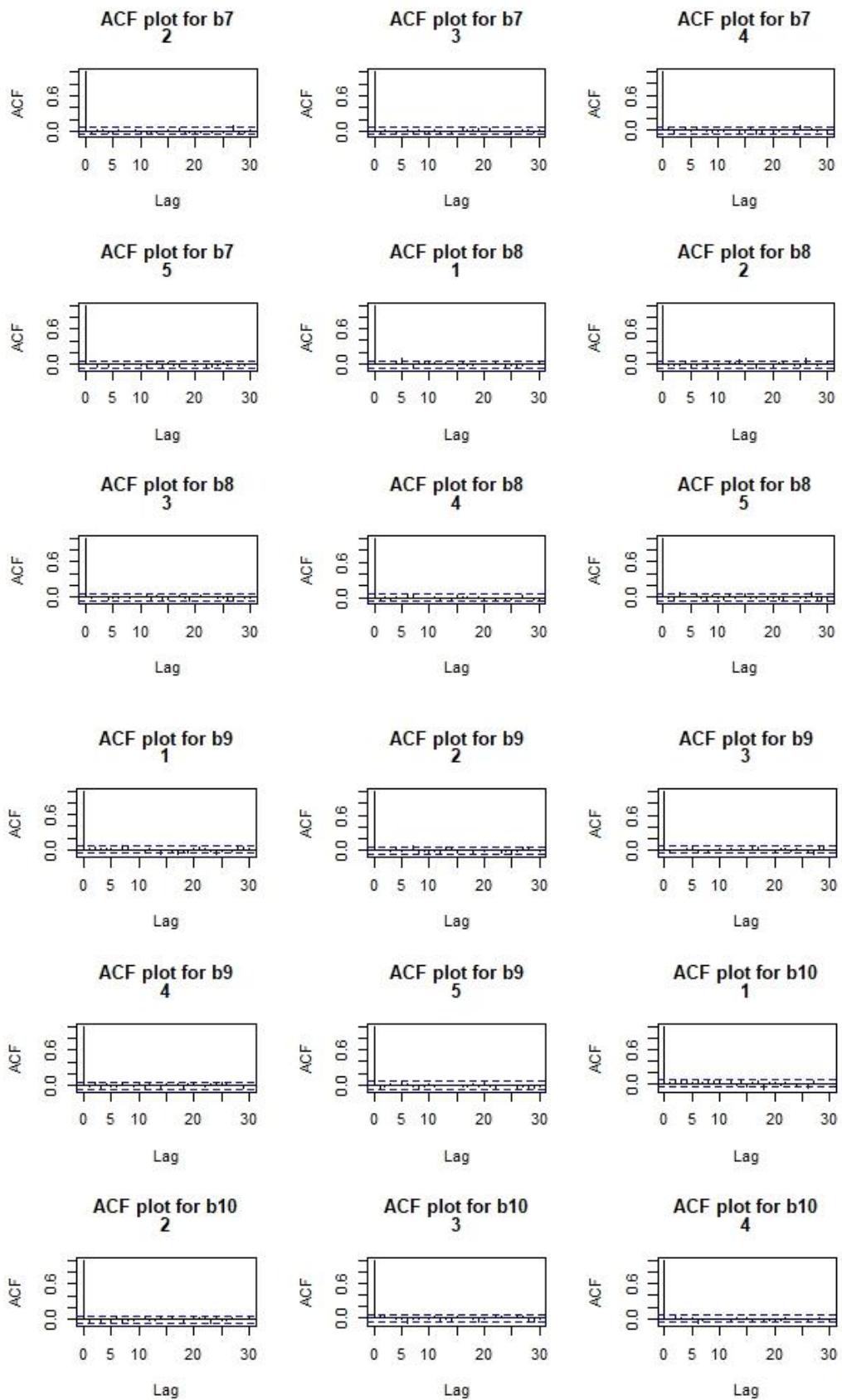


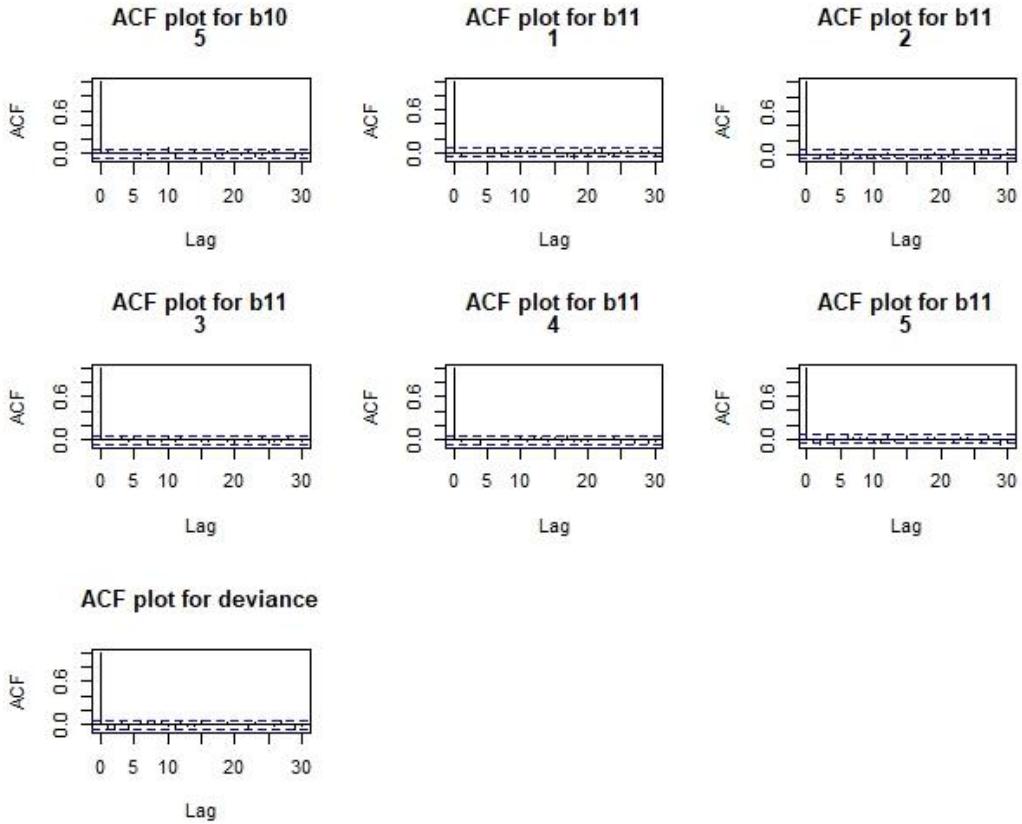












All those graphs indicate convergence for our model. While some parameters behave better than others, their erratic behaviour can be fixed by adding more iterations, which should smooth out our final results.

Results of the initial run:

After the pilot run, which is described above, we get the following results:

```
Inference for Bugs model at "C:\Users\30697\Desktop\multinomial.txt",
fit using OpenBUGS,
1 chains, each with 31000 iterations (first 1000 discarded), n.thin =
30
n.sims = 1000 iterations saved
      mean    sd   2.5%   25%   50%   75%  97.5%
b0[1] -10.0  1.7 -14.0 -11.1 -9.8 -8.8 -7.1
b0[2]  -3.0  0.2 -3.4 -3.1 -3.0 -2.9 -2.6
b0[3]   0.2  0.1  0.0  0.1  0.1  0.2  0.3
b0[4]  -2.0  0.2 -2.4 -2.1 -2.0 -1.9 -1.8
b0[5]  -6.6  0.8 -8.1 -7.1 -6.5 -6.0 -5.2
b1[1]   1.1  1.3 -1.5  0.3  1.1  1.9  3.6
b1[2]   1.0  0.5  0.1  0.7  1.0  1.3  1.9
b1[3]   0.2  0.2 -0.1  0.1  0.2  0.4  0.6
```

b1[4]	0.7	0.2	0.2	0.6	0.7	0.9	1.2
b1[5]	-0.5	0.7	-1.9	-1.0	-0.4	0.0	0.9
b2[1]	1.8	0.5	0.8	1.4	1.7	2.1	2.8
b2[2]	0.5	0.2	0.2	0.4	0.5	0.6	0.8
b2[3]	-0.5	0.1	-0.7	-0.5	-0.5	-0.4	-0.3
b2[4]	-0.8	0.2	-1.1	-0.9	-0.8	-0.7	-0.5
b2[5]	-0.1	0.4	-1.0	-0.4	-0.1	0.2	0.7
b3[1]	0.8	0.9	-1.0	0.2	0.8	1.4	2.3
b3[2]	0.0	0.3	-0.5	-0.1	0.0	0.2	0.6
b3[3]	-0.3	0.1	-0.5	-0.3	-0.3	-0.2	0.0
b3[4]	-0.1	0.2	-0.4	-0.2	-0.1	0.0	0.3
b3[5]	0.5	0.5	-0.5	0.2	0.6	0.9	1.5
b4[1]	-0.1	0.5	-1.3	-0.5	-0.1	0.2	0.8
b4[2]	0.6	0.2	0.2	0.5	0.6	0.7	0.9
b4[3]	0.1	0.1	-0.1	0.0	0.1	0.1	0.2
b4[4]	0.4	0.1	0.2	0.3	0.4	0.5	0.6
b4[5]	0.0	0.4	-1.0	-0.3	0.0	0.3	0.8
b5[1]	0.7	0.4	-0.1	0.4	0.7	0.9	1.4
b5[2]	0.1	0.2	-0.3	0.0	0.1	0.2	0.4
b5[3]	-0.1	0.1	-0.3	-0.2	-0.1	-0.1	0.0
b5[4]	-0.5	0.2	-0.8	-0.6	-0.5	-0.4	-0.2
b5[5]	-1.9	0.7	-3.4	-2.3	-1.9	-1.3	-0.6
b6[1]	2.0	1.1	0.0	1.3	2.0	2.7	4.3
b6[2]	-0.3	0.3	-0.8	-0.5	-0.3	-0.1	0.2
b6[3]	0.2	0.1	0.1	0.2	0.2	0.3	0.4
b6[4]	0.2	0.1	0.0	0.1	0.2	0.3	0.5
b6[5]	0.4	0.4	-0.6	0.1	0.4	0.7	1.2
b7[1]	-4.6	1.9	-8.9	-5.8	-4.5	-3.1	-1.4
b7[2]	-0.5	0.3	-1.1	-0.7	-0.5	-0.3	0.0
b7[3]	-0.5	0.1	-0.7	-0.6	-0.5	-0.5	-0.4
b7[4]	-0.9	0.2	-1.3	-1.0	-0.9	-0.8	-0.5
b7[5]	-1.7	0.5	-2.7	-2.0	-1.6	-1.3	-0.6
b8[1]	0.7	1.1	-1.6	-0.1	0.7	1.4	2.8
b8[2]	-1.1	0.4	-1.9	-1.3	-1.1	-0.8	-0.3
b8[3]	-0.1	0.2	-0.4	-0.2	-0.1	0.0	0.2
b8[4]	-0.6	0.2	-1.0	-0.8	-0.6	-0.4	-0.1
b8[5]	-0.2	0.6	-1.5	-0.7	-0.3	0.2	1.0
b9[1]	1.9	0.8	0.2	1.3	1.8	2.4	3.6
b9[2]	0.8	0.3	0.2	0.6	0.8	1.0	1.3
b9[3]	0.0	0.1	-0.2	-0.1	0.0	0.1	0.3
b9[4]	0.1	0.2	-0.2	0.0	0.1	0.2	0.5
b9[5]	-1.0	0.5	-2.0	-1.4	-1.0	-0.7	0.0
b10[1]	-0.6	0.7	-2.1	-1.0	-0.5	-0.1	0.6
b10[2]	0.1	0.2	-0.4	-0.1	0.1	0.3	0.5
b10[3]	0.4	0.1	0.3	0.4	0.4	0.5	0.6
b10[4]	0.9	0.1	0.7	0.8	0.9	1.0	1.1
b10[5]	1.2	0.3	0.7	1.0	1.2	1.4	1.7
b11[1]	-1.6	1.0	-3.8	-2.2	-1.6	-1.0	0.2

```

b11[2]      -0.3  0.3   -0.9  -0.5  -0.3  -0.1  0.3
b11[3]       0.8  0.1    0.6   0.8   0.8   0.9  1.1
b11[4]       1.4  0.2    1.0   1.3   1.4   1.5  1.7
b11[5]       2.3  0.5    1.4   2.0   2.3   2.6  3.2
deviance 2981.0 11.2 2960.5 2972.6 2980.7 2988.6 3003.6

DIC info (using the rule, pD = Dbar-Dhat)
pD = 55.2 and DIC = 3036.0
DIC is an estimate of expected predictive error (lower deviance is
better).

```

We can discern that with the exception of the 7th estimate for the β of total sulfur dioxide, which affected all the levels negatively, the other estimates either had a positive or a negative effect on each level. It could mean that if the levels that had a positive estimated value, then they had the correct dosage of a physiochemical characteristic, which worked well when it was combined with all the other components, while if they had a negative value, then that particular physiochemical deviant was not playing well with the others.

According to our results, our full model is:

$$\ln\left(\frac{p_{i1}}{p_{i6}}\right) = -10 + 1.1 \cdot Z_{i1} + 1.8 \cdot Z_{i2} + \dots - 1.6 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i2}}{p_{i6}}\right) = -3 + 1 \cdot Z_{i1} + 0.5 \cdot Z_{i2} + \dots - 0.3 \cdot Z_{i11}$$

⋮

$$\ln\left(\frac{p_{i5}}{p_{i6}}\right) = -6.6 - 0.5 \cdot Z_{i1} - 0.1 \cdot Z_{i2} + \dots + 2.3 \cdot Z_{i11}$$

where Z_{ip} are our centered values of X_{ip} , meaning that:

$$Z_{ip} = \frac{X_{ip} - \bar{X}_p}{S_p}$$

where \bar{X}_p is the mean value of the physiochemical characteristic p (the whole column of 1599 observations), while S_p is the standard deviation of the physiochemical characteristic p.

Variable Selection:

Now we are getting to the good part. The aim of this study is after all, to take the model above and select only the most significant physiochemical characteristics of those that we have data on. We will do it, by using the following methods:

Variable Selection through BIC on prior coefficients (using BAS package in R):

Just as it says, this method uses the Bayes Information Criterion (BIC) to pick and choose the most important variables in our model. This procedure will be done for centered values. We can run our code by using a bunch of prior distributions on the model. In our problem we will do it by running the following two:

1. Using a Uniform Distribution on the priors:

Which is as if we are specifying that we want to choose different models with the same probability which will be imposed on their dimensions.

Since there is no way to implement the BAS package for the multinomial logistic regression as it is, we separate our problem into 5 binomial problems (the number of levels that we have) in our code. And if we run the code for it, then we shall get the following results:

For level 1 (category 1 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 234 models

	post mean	post SD	post p(B != 0)
Intercept	-5.528401	0.764418	1.000000
fixed.acidity	0.022590	0.156528	0.055957
volatile.acidity	1.068056	0.289777	0.995752
citric.acid	0.042818	0.213934	0.069580
residual.sugar	0.006080	0.078205	0.039355
chlorides	0.055310	0.156334	0.140918
free.sulfur.dioxide	0.057447	0.363146	0.092334
total.sulfur.dioxide	-0.821427	0.943397	0.595020
density	0.132811	0.341706	0.168555
pH	0.025142	0.156214	0.059082
sulphates	0.002636	0.102711	0.041406
alcohol	-0.102082	0.400745	0.096729

> **summary(res1.1)**

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.00000000	1.0000000	1.000000	1.0000000
fixed.acidity	0.05595703	0.0000000	0.00000	0.0000000
volatile.acidity	0.99575195	1.0000000	1.00000	1.0000000
citric.acid	0.06958008	0.0000000	0.00000	0.0000000
residual.sugar	0.03935547	0.0000000	0.00000	0.0000000
chlorides	0.14091797	0.0000000	0.00000	0.0000000
free.sulfur.dioxide	0.09233398	0.0000000	0.00000	0.0000000
total.sulfur.dioxide	0.59501953	1.0000000	0.00000	1.0000000
density	0.16855469	0.0000000	0.00000	1.0000000
pH	0.05908203	0.0000000	0.00000	0.0000000

sulphates	0.04140625	0.0000000	0.00000	0.0000000
alcohol	0.09672852	0.0000000	0.00000	0.0000000
BF	NA	0.9975965	1.00000	0.2595098
PostProbs	NA	0.2343000	0.23200	0.0568000
R2	NA	0.2611000	0.19870	0.2979000
dim	NA	3.0000000	2.00000	4.0000000
logmarg	NA	-45.1674217	-45.16502	-46.5139760
		model 4	model 5	
Intercept	1.0000000	1.0000000		
fixed.acidity	0.0000000	0.0000000		
volatile.acidity	1.0000000	1.0000000		
citric.acid	0.0000000	0.0000000		
residual.sugar	0.0000000	0.0000000		
chlorides	1.0000000	0.0000000		
free.sulfur.dioxide	0.0000000	0.0000000		
total.sulfur.dioxide	1.0000000	1.0000000		
density	0.0000000	0.0000000		
pH	0.0000000	0.0000000		
sulphates	0.0000000	0.0000000		
alcohol	0.0000000	1.0000000		
BF	0.1765086	0.1445988		
PostProbs	0.0442000	0.0291000		
R2	0.2905000	0.2867000		
dim	4.0000000	4.0000000		
logmarg	-46.8994009	-47.0988074		

For level 2 (category 2 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 263 models

	post mean	post SD	post p(B != 0)
Intercept	-2.889649	0.203378	1.000000
fixed.acidity	-0.006037	0.064136	0.051025
volatile.acidity	0.575179	0.150088	0.992920
citric.acid	0.005414	0.058890	0.048926
residual.sugar	0.044970	0.116062	0.164014
chlorides	0.002428	0.030505	0.039404
free.sulfur.dioxide	-0.030289	0.132021	0.089160
total.sulfur.dioxide	-0.560638	0.243236	0.920166
density	-0.025939	0.106533	0.087939
pH	0.062857	0.145760	0.199561
sulphates	0.003964	0.039327	0.044189
alcohol	0.077925	0.177308	0.200439

> summary(res1.2)

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.00000000	1.0000	1.0000000	1.0000000

	0.05102539	0.0000	0.0000000	0.0000000
fixed.acidity	0.05102539	0.0000	0.0000000	0.0000000
volatile.acidity	0.99291992	1.0000	1.0000000	1.0000000
citric.acid	0.04892578	0.0000	0.0000000	0.0000000
residual.sugar	0.16401367	0.0000	0.0000000	0.0000000
chlorides	0.03940430	0.0000	0.0000000	0.0000000
free.sulfur.dioxide	0.08916016	0.0000	0.0000000	0.0000000
total.sulfur.dioxide	0.92016602	1.0000	1.0000000	1.0000000
density	0.08793945	0.0000	0.0000000	0.0000000
pH	0.19956055	0.0000	0.0000000	1.0000000
sulphates	0.04418945	0.0000	0.0000000	0.0000000
alcohol	0.20043945	0.0000	1.0000000	0.0000000
BF	NA	1.0000	0.3167589	0.2574623
PostProbs	NA	0.3694	0.1068000	0.0936000
R2	NA	0.0961	0.1074000	0.1063000
dim	NA	3.0000	4.0000000	4.0000000
logmarg	NA	-178.6481	-179.7977127	-180.0049804
	model 4	model 5		
Intercept	1.0000000	1.0000000		
fixed.acidity	0.0000000	0.0000000		
volatile.acidity	1.0000000	1.0000000		
citric.acid	0.0000000	0.0000000		
residual.sugar	1.0000000	0.0000000		
chlorides	0.0000000	0.0000000		
free.sulfur.dioxide	0.0000000	0.0000000		
total.sulfur.dioxide	1.0000000	1.0000000		
density	0.0000000	1.0000000		
pH	0.0000000	0.0000000		
sulphates	0.0000000	0.0000000		
alcohol	0.0000000	0.0000000		
BF	0.1789942	0.0885787		
PostProbs	0.0591000	0.0282000		
R2	0.1044000	0.1007000		
dim	4.0000000	4.0000000		
logmarg	-180.3685001	-181.0719622		

For level 3 (category 3 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 75 models

	post mean	post SD	post p(B != 0)
Intercept	0.1442064	0.0659399	1.0000000
fixed.acidity	0.0061386	0.0374363	0.0517578
volatile.acidity	-0.4072339	0.0867200	1.0000000
citric.acid	-0.0337026	0.0871079	0.1713867
residual.sugar	0.0016025	0.0149511	0.0333496

chlorides	-0.0413485	0.0774618	0.2664551
free.sulfur.dioxide	0.2243361	0.1368383	0.8024414
total.sulfur.dioxide	-0.5984030	0.1281603	0.9992188
density	0.0014770	0.0186418	0.0305176
pH	-0.0002957	0.0143161	0.0306152
sulphates	0.3505989	0.0760334	0.9984375
alcohol	0.8283591	0.0850498	0.9998047
> summary(res1.3)			
	P(B != 0 Y)	model 1	model 2
Intercept	1.00000000	1.0000	1.00000000
fixed.acidity	0.05175781	0.0000	0.00000000
volatile.acidity	1.00000000	1.0000	1.00000000
citric.acid	0.17138672	0.0000	0.00000000
residual.sugar	0.03334961	0.0000	0.00000000
chlorides	0.26645508	0.0000	1.00000000
free.sulfur.dioxide	0.80244141	1.0000	1.00000000
total.sulfur.dioxide	0.99921875	1.0000	1.00000000
density	0.03051758	0.0000	0.00000000
pH	0.03061523	0.0000	0.00000000
sulphates	0.99843750	1.0000	1.00000000
alcohol	0.99980469	1.0000	1.00000000
BF	NA	1.0000	0.4045382
PostProbs		NA	0.4576
R2		NA	0.1784
dim		NA	6.0000
logmarg		NA	-768.5041 -769.4090773 -770.237319
		model 4	model 5
Intercept	1.0000000	1.0000000	
fixed.acidity	0.0000000	0.0000000	
volatile.acidity	1.0000000	1.0000000	
citric.acid	1.0000000	1.0000000	
residual.sugar	0.0000000	0.0000000	
chlorides	0.0000000	0.0000000	
free.sulfur.dioxide	1.0000000	0.0000000	
total.sulfur.dioxide	1.0000000	1.0000000	
density	0.0000000	0.0000000	
pH	0.0000000	0.0000000	
sulphates	1.0000000	1.0000000	
alcohol	1.0000000	1.0000000	
BF	0.1416244	0.1080856	
PostProbs	0.0528000	0.0435000	
R2	0.1802000	0.1760000	
dim	7.0000000	6.0000000	
logmarg	-770.4586449	-770.7289000	

For level 4 (category 4 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 110 models

	post mean	post SD	post p(B != 0)
Intercept	-1.937507	0.156390	1.000000
fixed.acidity	0.061361	0.136611	0.228418
volatile.acidity	-0.981452	0.164900	0.998682
citric.acid	-0.005383	0.079777	0.056348
residual.sugar	0.016653	0.066310	0.085840
chlorides	-0.122066	0.191919	0.362207
free.sulfur.dioxide	0.116267	0.209230	0.286084
total.sulfur.dioxide	-0.589230	0.222370	0.991211
density	-0.008337	0.085674	0.058789
pH	-0.028556	0.092393	0.126807
sulphates	0.644060	0.132453	0.997705
alcohol	1.750716	0.170207	0.999121

> **summary(res1.4)**

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.00000000	1.0000	1.00000000	1.00000000
fixed.acidity	0.22841797	0.0000	0.00000000	0.00000000
volatile.acidity	0.99868164	1.0000	1.00000000	1.00000000
citric.acid	0.05634766	0.0000	0.00000000	0.00000000
residual.sugar	0.08583984	0.0000	0.00000000	0.00000000
chlorides	0.36220703	0.0000	1.00000000	0.00000000
free.sulfur.dioxide	0.28608398	0.0000	0.00000000	1.00000000
total.sulfur.dioxide	0.99121094	1.0000	1.00000000	1.00000000
density	0.05878906	0.0000	0.00000000	0.00000000
pH	0.12680664	0.0000	0.00000000	0.00000000
sulphates	0.99770508	1.0000	1.00000000	1.00000000
alcohol	0.99912109	1.0000	1.00000000	1.00000000
BF	NA	1.0000	0.5284735	0.3574823
PostProbs	NA	0.2460	0.1499000	0.0959000
R2	NA	0.5192	0.5250000	0.5242000
dim	NA	5.0000	6.0000000	6.0000000
logmarg	NA	-239.7406	-240.3783653	-240.7692722
		model 4	model 5	
Intercept	1.0000000	1.0000000		
fixed.acidity	1.0000000	1.0000000		
volatile.acidity	1.0000000	1.0000000		
citric.acid	0.0000000	0.0000000		
residual.sugar	0.0000000	0.0000000		
chlorides	0.0000000	1.0000000		
free.sulfur.dioxide	0.0000000	0.0000000		
total.sulfur.dioxide	1.0000000	1.0000000		
density	0.0000000	0.0000000		
pH	0.0000000	0.0000000		
sulphates	1.0000000	1.0000000		

alcohol	1.0000000	1.0000000
BF	0.2508113	0.1633278
PostProbs	0.0718000	0.0455000
R2	0.5235000	0.5298000
dim	6.0000000	7.0000000
logmarg	-241.1236571	-241.5525986

For level 5 (category 5 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 401 models

	post mean	post SD	post p(B != 0)
Intercept	-6.90153	1.48388	1.00000
fixed.acidity	-0.07906	0.37641	0.10449
volatile.acidity	-0.41466	0.65142	0.35020
citric.acid	1.00256	0.99446	0.60117
residual.sugar	-1.17122	1.11341	0.61919
chlorides	-1.19416	1.49193	0.52090
free.sulfur.dioxide	0.02615	0.22948	0.07310
total.sulfur.dioxide	-1.58583	0.86360	0.88843
density	-0.32705	0.74951	0.22988
pH	-0.23728	0.51044	0.23784
sulphates	1.26924	0.54658	0.92886
alcohol	2.48258	0.87481	0.99849

> **summary(res1.5)**

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.0000000	1.00000	1.000000	1.0000000
fixed.acidity	0.1044922	0.00000	0.000000	0.0000000
volatile.acidity	0.3501953	0.00000	0.000000	1.0000000
citric.acid	0.6011719	1.00000	1.000000	0.0000000
residual.sugar	0.6191895	1.00000	0.000000	1.0000000
chlorides	0.5208984	1.00000	1.000000	0.0000000
free.sulfur.dioxide	0.0730957	0.00000	0.000000	0.0000000
total.sulfur.dioxide	0.8884277	1.00000	1.000000	1.0000000
density	0.2298828	0.00000	1.000000	0.0000000
pH	0.2378418	0.00000	0.000000	0.0000000
sulphates	0.9288574	1.00000	1.000000	1.0000000
alcohol	0.9984863	1.00000	1.000000	1.0000000
BF	NA	1.00000	0.755376	0.8293052
PostProbs	NA	0.12020	0.098900	0.0944000
R2	NA	0.66940	0.666000	0.6280000
dim	NA	7.00000	7.000000	6.0000000
logmarg	NA	-47.30155	-47.582095	-47.4887219
	model 4	model 5		
Intercept	1.0000000	1.0000000		

fixed.acidity	0.0000000	0.0000000
volatile.acidity	0.0000000	1.0000000
citric.acid	1.0000000	0.0000000
residual.sugar	1.0000000	1.0000000
chlorides	0.0000000	0.0000000
free.sulfur.dioxide	0.0000000	0.0000000
total.sulfur.dioxide	1.0000000	1.0000000
density	0.0000000	0.0000000
pH	0.0000000	1.0000000
sulphates	1.0000000	1.0000000
alcohol	1.0000000	1.0000000
BF	0.8464082	0.3194614
PostProbs	0.0886000	0.0386000
R2	0.6282000	0.6557000
dim	6.0000000	7.0000000
logmarg	-47.4683084	-48.4426737

The summaries above show us the top 5 models for each of the categories, while the coefficients are for the MAP, (or in other words HMP) models. We will choose the best model based on the Bayes Factor (BF). Therefore, if it is 1, then this is the best model. Thus, for instance, for the first table, we saw that model 2 was the best, with a very close second alternative from model 1. According to the results above, we can see that our model is of the following form:

$$\ln\left(\frac{p_{i1}}{p_{i6}}\right) = -5.53 + 1.07 \cdot Z_{i2}$$

$$\ln\left(\frac{p_{i2}}{p_{i6}}\right) = -2.89 + 0.58 \cdot Z_{i2} - 0.56 \cdot Z_{i7}$$

$$\ln\left(\frac{p_{i3}}{p_{i6}}\right) = 0.14 - 0.41 \cdot Z_{i2} + 0.22 \cdot Z_{i6} - 0.6 \cdot Z_{i7} + 0.35 \cdot Z_{i10} + 0.83 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i4}}{p_{i6}}\right) = -1.94 - 0.98 \cdot Z_{i2} - 0.6 \cdot Z_{i7} + 0.64 \cdot Z_{i10} + 1.75 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i5}}{p_{i6}}\right) = -6.9 + 1 \cdot Z_{i3} - 1.17 \cdot Z_{i4} - 1.19 \cdot Z_{i5} - 1.6 \cdot Z_{i7} + 1.3 \cdot Z_{i10} + 2.48 \cdot Z_{i11}$$

We can see that the physiochemical characteristics of *Fixed Acidity*, *pH* and *Density* were not considered useful at all, while others were used to a degree, depending on the level on which they were.

2. Using a Beta-Binomial Distribution on the priors:

Same as before, only this time instead of imposing a *Uniform* distribution on the priors, we shall try the *Beta-Binomial*. Here we will choose the hyperparameter a to be smaller than the hyperparameter b , thus stating that we would prefer a more parsimonious model in the end. Based on our code, we get the following results:

For level 1 (category 1 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 67 models

	post mean	post SD	post p(B != 0)
Intercept	-5.259e+00	5.402e-01	1.000e+00
fixed.acidity	5.495e-03	7.699e-02	1.099e-02
volatile.acidity	1.060e+00	2.660e-01	9.886e-01
citric.acid	1.628e-03	5.875e-02	7.715e-03
residual.sugar	9.681e-04	2.973e-02	7.373e-03
chlorides	5.733e-03	5.071e-02	1.675e-02
free.sulfur.dioxide	-3.446e-03	9.055e-02	1.255e-02
total.sulfur.dioxide	-1.524e-01	4.575e-01	1.291e-01
density	1.011e-02	9.218e-02	1.636e-02
pH	4.531e-03	6.775e-02	9.375e-03
sulphates	7.002e-05	2.980e-02	4.297e-03
alcohol	-7.579e-03	1.075e-01	1.113e-02

> **summary(res1.1)**

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.000000000	1.00000	1.0000000	1.0000000
fixed.acidity	0.010986328	0.00000	0.0000000	0.0000000
volatile.acidity	0.988574219	1.00000	1.0000000	1.0000000
citric.acid	0.007714844	0.00000	0.0000000	0.0000000
residual.sugar	0.007373047	0.00000	0.0000000	0.0000000
chlorides	0.016748047	0.00000	0.0000000	0.0000000
free.sulfur.dioxide	0.012548828	0.00000	0.0000000	0.0000000
total.sulfur.dioxide	0.129101562	0.00000	1.0000000	0.0000000
density	0.016357422	0.00000	0.0000000	1.0000000
pH	0.009375000	0.00000	0.0000000	0.0000000
sulphates	0.004296875	0.00000	0.0000000	0.0000000
alcohol	0.011132813	0.00000	0.0000000	0.0000000
BF	NA	1.00000	0.9975965	0.1151104
PostProbs	NA	0.79260	0.1122000	0.0114000
R2	NA	0.19870	0.2611000	0.2198000
dim	NA	2.00000	3.0000000	3.0000000
logmarg	NA	-45.16502	-45.1674217	-47.3268794
	model 4		model 5	

Intercept	1.0000000	1.00000000
fixed.acidity	0.0000000	0.00000000
volatile.acidity	1.0000000	1.00000000
citric.acid	0.0000000	0.00000000
residual.sugar	0.0000000	0.00000000
chlorides	1.0000000	0.00000000
free.sulfur.dioxide	0.0000000	1.00000000
total.sulfur.dioxide	0.0000000	0.00000000
density	0.0000000	0.00000000
pH	0.0000000	0.00000000
sulphates	0.0000000	0.00000000
alcohol	0.0000000	0.00000000
BF	0.1116504	0.08472875
PostProbs	0.0103000	0.00840000
R2	0.2193000	0.21400000
dim	3.0000000	3.00000000
logmarg	-47.3573984	-47.63331571

For level 2 (category 2 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 92 models

	post mean	post SD	post p(B != 0)
Intercept	-2.8982978	0.1996788	1.0000000
fixed.acidity	-0.0024105	0.0314922	0.0149414
volatile.acidity	0.5786542	0.1617787	0.9791992
citric.acid	0.0002903	0.0243806	0.0083984
residual.sugar	0.0109902	0.0586060	0.0430176
chlorides	0.0004225	0.0148975	0.0106445
free.sulfur.dioxide	-0.0228008	0.1140097	0.0520508
total.sulfur.dioxide	-0.5490599	0.2592015	0.8873535
density	-0.0060057	0.0498721	0.0220703
pH	0.0224283	0.0935222	0.0670410
sulphates	0.0007647	0.0180731	0.0102051
alcohol	0.0350480	0.1284128	0.0828613

> **summary(res1.2)**

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.000000000	1.0000	1.0000000	1.0000000
fixed.acidity	0.014941406	0.0000	0.0000000	0.0000000
volatile.acidity	0.979199219	1.0000	1.0000000	1.0000000
citric.acid	0.008398437	0.0000	0.0000000	0.0000000
residual.sugar	0.043017578	0.0000	0.0000000	0.0000000
chlorides	0.010644531	0.0000	0.0000000	0.0000000
free.sulfur.dioxide	0.052050781	0.0000	0.0000000	0.0000000
total.sulfur.dioxide	0.887353516	1.0000	1.0000000	1.0000000

density	0.022070313	0.0000	0.0000000	0.0000000
pH	0.067041016	0.0000	0.0000000	1.0000000
sulphates	0.010205078	0.0000	0.0000000	0.0000000
alcohol	0.082861328	0.0000	1.0000000	0.0000000
BF	NA	1.0000	0.3167589	0.2574623
PostProbs	NA	0.6815	0.0532000	0.0363000
R2	NA	0.0961	0.1074000	0.1063000
dim	NA	3.0000	4.0000000	4.0000000
logmarg	NA	-178.6481	-179.7977127	-180.0049804
		model 4	model 5	
Intercept	1.000000e+00	1.0000000		
fixed.acidity	0.000000e+00	0.0000000		
volatile.acidity	1.000000e+00	1.0000000		
citric.acid	0.000000e+00	0.0000000		
residual.sugar	0.000000e+00	1.0000000		
chlorides	0.000000e+00	0.0000000		
free.sulfur.dioxide	0.000000e+00	0.0000000		
total.sulfur.dioxide	0.000000e+00	1.0000000		
density	0.000000e+00	0.0000000		
pH	0.000000e+00	0.0000000		
sulphates	0.000000e+00	0.0000000		
alcohol	0.000000e+00	0.0000000		
BF	9.554042e-03	0.1789942		
PostProbs	3.430000e-02	0.0297000		
R2	5.430000e-02	0.1044000		
dim	2.000000e+00	4.0000000		
logmarg	-1.832989e+02	-180.3685001		

For level 3 (category 3 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 56 models

	post mean	post SD	post p(B != 0)
Intercept	1.463e-01	6.624e-02	1.000e+00
fixed.acidity	1.712e-03	2.020e-02	1.626e-02
volatile.acidity	-4.043e-01	8.166e-02	9.999e-01
citric.acid	-1.981e-02	6.667e-02	1.019e-01
residual.sugar	8.029e-04	1.073e-02	1.812e-02
chlorides	-2.484e-02	6.359e-02	1.583e-01
free.sulfur.dioxide	1.822e-01	1.523e-01	6.465e-01
total.sulfur.dioxide	-5.678e-01	1.360e-01	9.998e-01
density	7.787e-04	1.646e-02	1.660e-02
pH	-3.233e-05	1.012e-02	1.792e-02
sulphates	3.393e-01	7.310e-02	9.986e-01
alcohol	8.348e-01	8.997e-02	9.984e-01

> summary(res1.3)

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.00000000	1.0000	1.000000	1.0000000
fixed.acidity	0.01625977	0.0000	0.000000	0.0000000
volatile.acidity	0.99990234	1.0000	1.000000	1.0000000
citric.acid	0.10190430	0.0000	0.000000	0.0000000
residual.sugar	0.01811523	0.0000	0.000000	0.0000000
chlorides	0.15825195	0.0000	0.000000	1.0000000
free.sulfur.dioxide	0.64648438	1.0000	0.000000	1.0000000
total.sulfur.dioxide	0.99975586	1.0000	1.000000	1.0000000
density	0.01660156	0.0000	0.000000	0.0000000
pH	0.01791992	0.0000	0.000000	0.0000000
sulphates	0.99858398	1.0000	1.000000	1.0000000
alcohol	0.99838867	1.0000	1.000000	1.0000000
BF	NA	1.0000	0.176709	0.4045382
PostProbs	NA	0.4693	0.239100	0.1029000
R2	NA	0.1784	0.172600	0.1814000
dim	NA	6.0000	5.000000	7.0000000
logmarg	NA	-768.5041	-770.237319	-769.4090773
		model 4	model 5	
Intercept	1.0000000	1.00000000		
fixed.acidity	0.0000000	0.00000000		
volatile.acidity	1.0000000	1.00000000		
citric.acid	1.0000000	0.00000000		
residual.sugar	0.0000000	0.00000000		
chlorides	0.0000000	1.00000000		
free.sulfur.dioxide	0.0000000	0.00000000		
total.sulfur.dioxide	1.0000000	1.00000000		
density	0.0000000	0.00000000		
pH	0.0000000	0.00000000		
sulphates	1.0000000	1.00000000		
alcohol	1.0000000	1.00000000		
BF	0.1080856	0.08753025		
PostProbs	0.0477000	0.03740000		
R2	0.1760000	0.17580000		
dim	6.0000000	6.00000000		
logmarg	-770.7289000	-770.93983893		

For level 4 (category 4 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 80 models

	post mean	post SD	post p(B != 0)
Intercept	-1.928005	0.154828	1.000000
fixed.acidity	0.027436	0.090846	0.110107

volatile.acidity	-1.002758	0.155906	0.999756	
citric.acid	-0.001380	0.039418	0.020313	
residual.sugar	0.007085	0.043262	0.038379	
chlorides	-0.063223	0.149847	0.189697	
free.sulfur.dioxide	0.052857	0.151766	0.131348	
total.sulfur.dioxide	-0.538386	0.188332	0.988525	
density	-0.001178	0.040672	0.020801	
pH	-0.011699	0.059048	0.055273	
sulphates	0.626892	0.123760	0.998682	
alcohol	1.760819	0.156445	0.999951	
> summary(res1.4)				
	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.00000000	1.0000	1.00000000	1.00000000
fixed.acidity	0.11010742	0.0000	0.00000000	0.00000000
volatile.acidity	0.99975586	1.0000	1.00000000	1.00000000
citric.acid	0.02031250	0.0000	0.00000000	0.00000000
residual.sugar	0.03837891	0.0000	0.00000000	0.00000000
chlorides	0.18969727	0.0000	1.00000000	0.00000000
free.sulfur.dioxide	0.13134766	0.0000	0.00000000	1.00000000
total.sulfur.dioxide	0.98852539	1.0000	1.00000000	1.00000000
density	0.02080078	0.0000	0.00000000	0.00000000
pH	0.05527344	0.0000	0.00000000	0.00000000
sulphates	0.99868164	1.0000	1.00000000	1.00000000
alcohol	0.99995117	1.0000	1.00000000	1.00000000
BF	NA	1.0000	0.5284735	0.3574823
PostProbs	NA	0.5808	0.1132000	0.0698000
R2	NA	0.5192	0.5250000	0.5242000
dim	NA	5.0000	6.0000000	6.0000000
logmarg	NA	-239.7406	-240.3783653	-240.7692722
		model 4	model 5	
Intercept	1.000000	1.0000000		
fixed.acidity	1.000000	0.0000000		
volatile.acidity	1.000000	1.0000000		
citric.acid	0.000000	0.0000000		
residual.sugar	0.000000	0.0000000		
chlorides	0.000000	0.0000000		
free.sulfur.dioxide	0.000000	0.0000000		
total.sulfur.dioxide	1.000000	1.0000000		
density	0.000000	0.0000000		
pH	0.000000	1.0000000		
sulphates	1.000000	1.0000000		
alcohol	1.000000	1.0000000		
BF	0.2508113	0.1176858		
PostProbs	0.0495000	0.0237000		
R2	0.5235000	0.5218000		
dim	6.0000000	6.0000000		
logmarg	-241.1236571	-241.8803398		

For level 5 (category 5 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 288 models

	post mean	post SD	post p(B != 0)
Intercept	-5.69923	1.40947	1.00000
fixed.acidity	-0.02627	0.21167	0.03550
volatile.acidity	-0.60438	0.71479	0.46704
citric.acid	0.49775	0.78559	0.33643
residual.sugar	-0.66771	0.99523	0.36714
chlorides	-0.54429	1.20641	0.21753
free.sulfur.dioxide	-0.00152	0.12386	0.02661
total.sulfur.dioxide	-0.88077	0.96336	0.52803
density	-0.12867	0.45461	0.09658
pH	-0.15009	0.40770	0.15039
sulphates	0.70311	0.68441	0.57793
alcohol	2.18824	0.71022	0.99844

> **summary(res1.5)**

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.00000000	1.00000000	1.00000	1.00000000
fixed.acidity	0.03549805	0.00000000	0.00000	0.00000000
volatile.acidity	0.46704102	1.00000000	0.00000	1.00000000
citric.acid	0.33642578	0.00000000	1.00000	0.00000000
residual.sugar	0.36713867	0.00000000	1.00000	1.00000000
chlorides	0.21752930	0.00000000	0.00000	0.00000000
free.sulfur.dioxide	0.02661133	0.00000000	0.00000	0.00000000
total.sulfur.dioxide	0.52802734	0.00000000	1.00000	1.00000000
density	0.09658203	0.00000000	0.00000	0.00000000
pH	0.15039062	0.00000000	0.00000	0.00000000
sulphates	0.57792969	0.00000000	1.00000	1.00000000
alcohol	0.99843750	1.00000000	1.00000	1.00000000
BF	NA	0.04817807	1.00000	0.9797934
PostProbs	NA	0.17370000	0.07660	0.0747000
R2	NA	0.47450000	0.62820	0.6280000
dim	NA	3.00000000	6.00000	6.00000000
logmarg	NA	-50.50115974	-47.46831	-47.4887219
		model 4	model 5	
Intercept	1.00000000	1.0000000		
fixed.acidity	0.00000000	0.0000000		
volatile.acidity	0.00000000	1.0000000		
citric.acid	0.00000000	0.0000000		
residual.sugar	0.00000000	0.0000000		
chlorides	0.00000000	0.0000000		
free.sulfur.dioxide	0.00000000	0.0000000		
total.sulfur.dioxide	1.00000000	1.0000000		
density	0.00000000	0.0000000		

pH	0.00000000	0.0000000
sulphates	1.00000000	1.0000000
alcohol	1.00000000	1.0000000
BF	0.07164376	0.2830567
PostProbs	0.05470000	0.0478000
R2	0.51840000	0.5740000
dim	4.00000000	5.0000000
logmarg	-50.10435766	-48.7304164

Just like before, we see that in the end our model will be looking like this:

$$\ln\left(\frac{p_{i1}}{p_{i6}}\right) = -5.3 + 1.06 \cdot Z_{i2}$$

$$\ln\left(\frac{p_{i2}}{p_{i6}}\right) = -2.89 + 0.57 \cdot Z_{i2} - 0.55 \cdot Z_{i7}$$

$$\ln\left(\frac{p_{i3}}{p_{i6}}\right) = 0.15 - 0.4 \cdot Z_{i2} + 0.18 \cdot Z_{i6} - 0.57 \cdot Z_{i7} + 0.34 \cdot Z_{i10} + 0.83 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i4}}{p_{i6}}\right) = -1.93 - 1 \cdot Z_{i2} - 0.54 \cdot Z_{i7} + 0.63 \cdot Z_{i10} + 1.76 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i5}}{p_{i6}}\right) = -5.7 + 0.5 \cdot Z_{i3} - 0.67 \cdot Z_{i4} - 0.88 \cdot Z_{i7} + 0.7 \cdot Z_{i10} + 2.19 \cdot Z_{i11}$$

Here we did not use the physiochemical characteristics of *Fixed Acidity*, *chlorides*, *pH* and *Density* at all. This selection was a bit more parsimonious than the last one.

Variable Selection through Empirical Bayes (using OpenBUGS):

The outlawed approach that will get us in trouble if we talk to any old school Bayesians. This method is assuming that:

$$\beta_{jk} \sim N\left(\tilde{\beta}_{jk}, n \cdot \tilde{S}_{\beta_{jk}}^2\right)$$

for $j = 1, 2, 3, \dots, 11$, our number of estimates and $k = 1, 2, \dots, 5$, our number of levels.

where, $\tilde{\beta}_{jk}$ is the posterior mean and $\tilde{S}_{\beta_{jk}}^2$ is the posterior variance of the full model. In our code, we have set our constants to always be present in the final result. Due to time constraints, we ran each model for 11000 iterations with a thinning step of 30.

1. Using a Uniform Distribution on the priors:

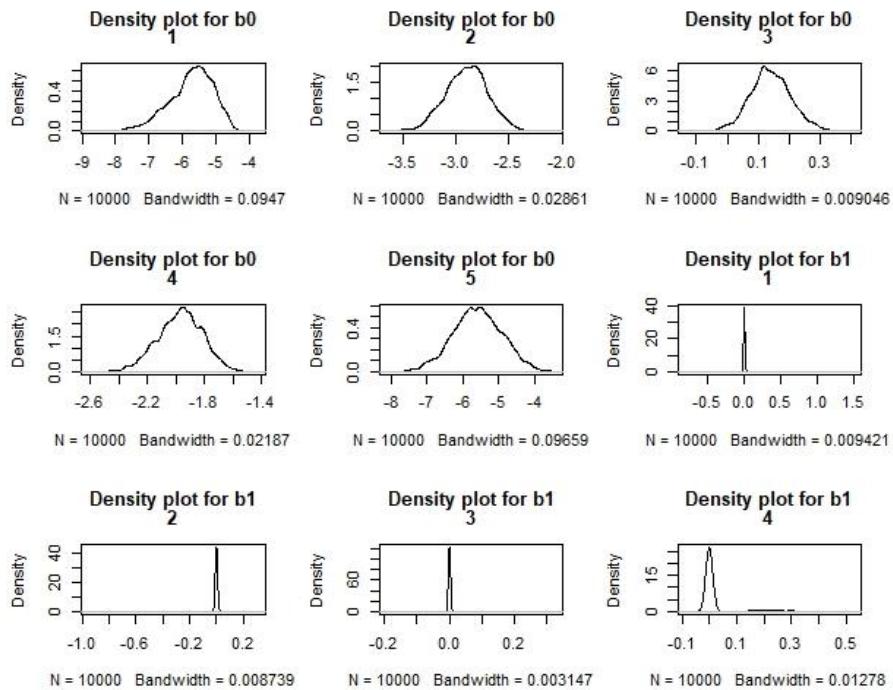
Here the results are:

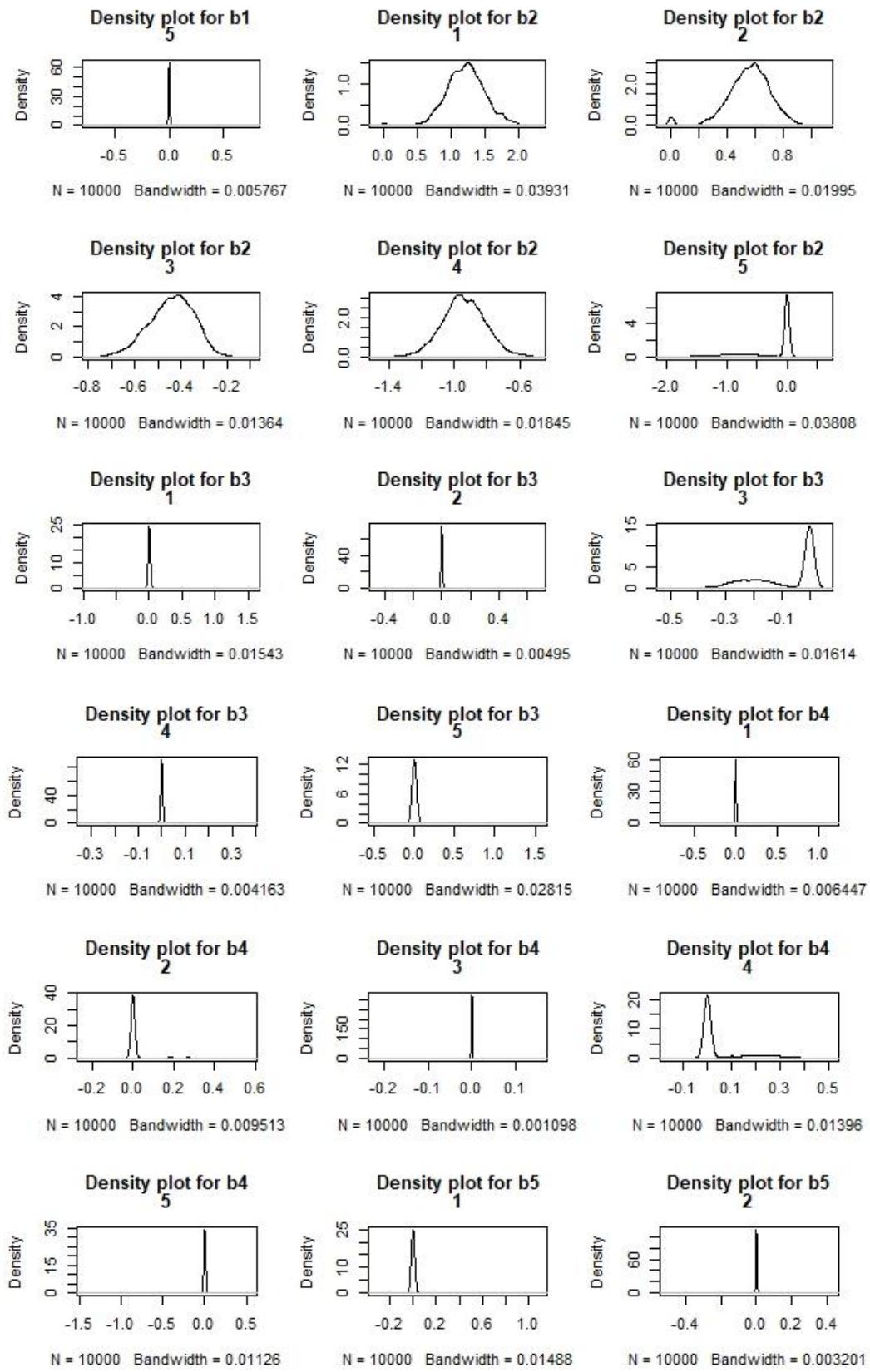
```
Inference for Bugs model at "C:\Users\30697\Desktop\empirical.txt",
fit using OpenBUGS,
 1 chains, each with 11000 iterations (first 1000 discarded), n.thin =
30
  n.sims = 333 iterations saved

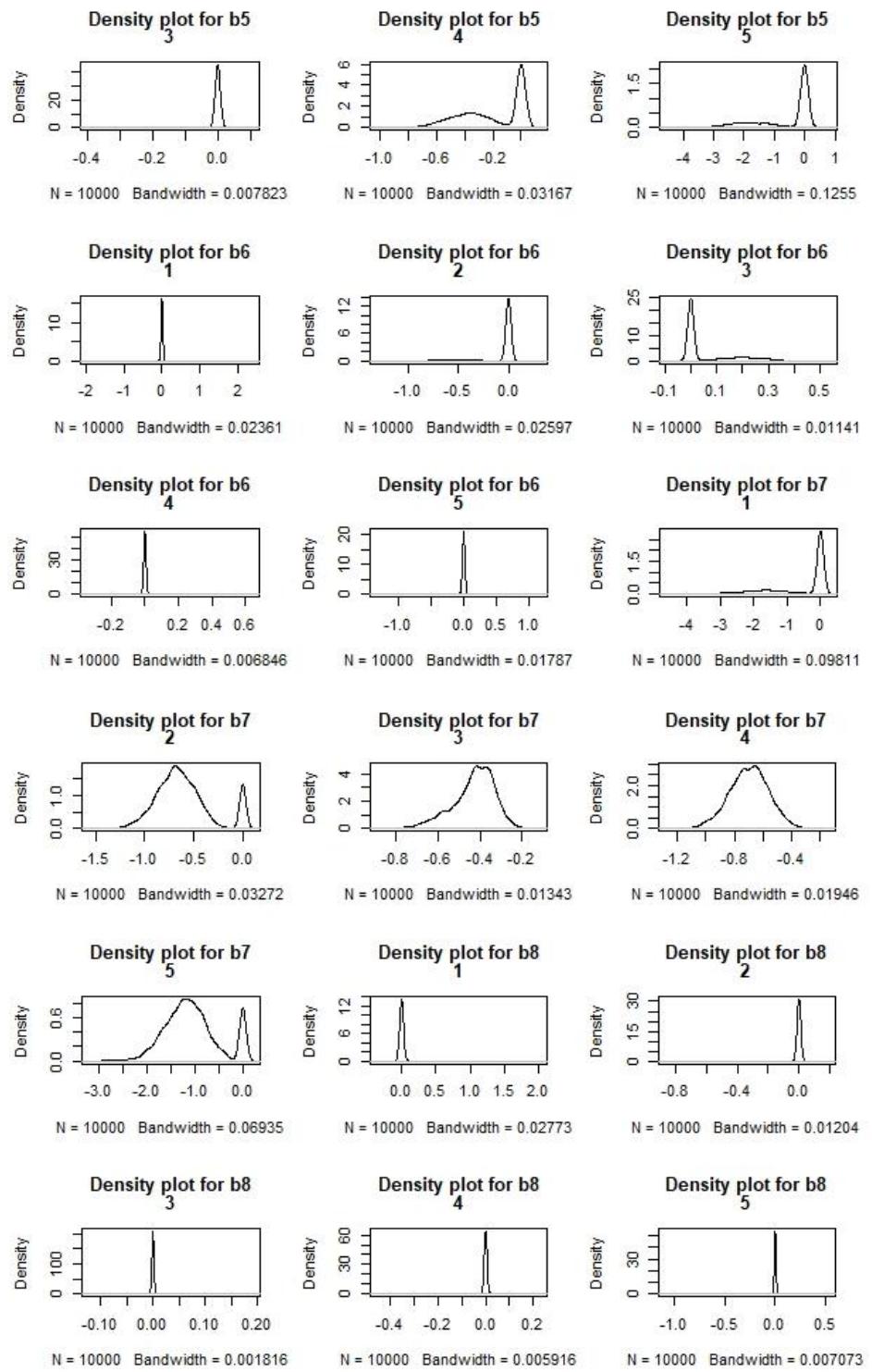
      mean    sd   2.5%   25%   50%   75%  97.5%
gb0[1] -5.7  0.6  -6.9  -6.1  -5.6  -5.3  -4.7
gb0[2] -2.9  0.2  -3.3  -3.0  -2.9  -2.8  -2.5
gb0[3]  0.1  0.1   0.0   0.1   0.1   0.2   0.3
gb0[4] -2.0  0.2  -2.3  -2.1  -2.0  -1.9  -1.7
gb0[5] -5.6  0.7  -6.9  -6.0  -5.6  -5.2  -4.2
gb1[1]  0.0  0.0   0.0   0.0   0.0   0.0   0.0
gb1[2]  0.0  0.1   0.0   0.0   0.0   0.0   0.0
gb1[3]  0.0  0.0   0.0   0.0   0.0   0.0   0.0
gb1[4]  0.0  0.1   0.0   0.0   0.0   0.0   0.3
gb1[5]  0.0  0.0   0.0   0.0   0.0   0.0   0.0
gb2[1]  1.2  0.3   0.7   1.0   1.2   1.4   1.8
gb2[2]  0.5  0.2   0.0   0.5   0.6   0.6   0.8
gb2[3] -0.4  0.1  -0.7  -0.5  -0.4  -0.4  -0.3
gb2[4] -0.9  0.1  -1.2  -1.0  -1.0  -0.9  -0.7
gb2[5] -0.2  0.4  -1.4   0.0   0.0   0.0   0.0
gb3[1]  0.0  0.2   0.0   0.0   0.0   0.0   0.1
gb3[2]  0.0  0.0   0.0   0.0   0.0   0.0   0.0
gb3[3] -0.1  0.1  -0.3  -0.2   0.0   0.0   0.0
gb3[4]  0.0  0.0   0.0   0.0   0.0   0.0   0.0
gb3[5]  0.1  0.2   0.0   0.0   0.0   0.0   0.8
gb4[1]  0.0  0.0   0.0   0.0   0.0   0.0   0.0
gb4[2]  0.0  0.1   0.0   0.0   0.0   0.0   0.3
gb4[3]  0.0  0.0   0.0   0.0   0.0   0.0   0.0
gb4[4]  0.0  0.1   0.0   0.0   0.0   0.0   0.3
gb4[5]  0.0  0.1   0.0   0.0   0.0   0.0   0.0
gb5[1]  0.0  0.1   0.0   0.0   0.0   0.0   0.3
gb5[2]  0.0  0.0   0.0   0.0   0.0   0.0   0.0
gb5[3]  0.0  0.0  -0.2   0.0   0.0   0.0   0.0
gb5[4] -0.2  0.2  -0.6  -0.4  -0.1   0.0   0.0
gb5[5] -0.6  0.9  -2.8  -1.2   0.0   0.0   0.0
gb6[1]  0.0  0.2   0.0   0.0   0.0   0.0   0.0
gb6[2] -0.1  0.2  -0.7   0.0   0.0   0.0   0.0
gb6[3]  0.1  0.1   0.0   0.0   0.0   0.1   0.3
gb6[4]  0.0  0.1   0.0   0.0   0.0   0.0   0.2
gb6[5]  0.0  0.1  -0.2   0.0   0.0   0.0   0.0
gb7[1] -0.5  0.8  -2.6  -0.8   0.0   0.0   0.0
gb7[2] -0.6  0.3  -1.1  -0.8  -0.7  -0.5   0.0
gb7[3] -0.4  0.1  -0.7  -0.5  -0.4  -0.4  -0.3
```

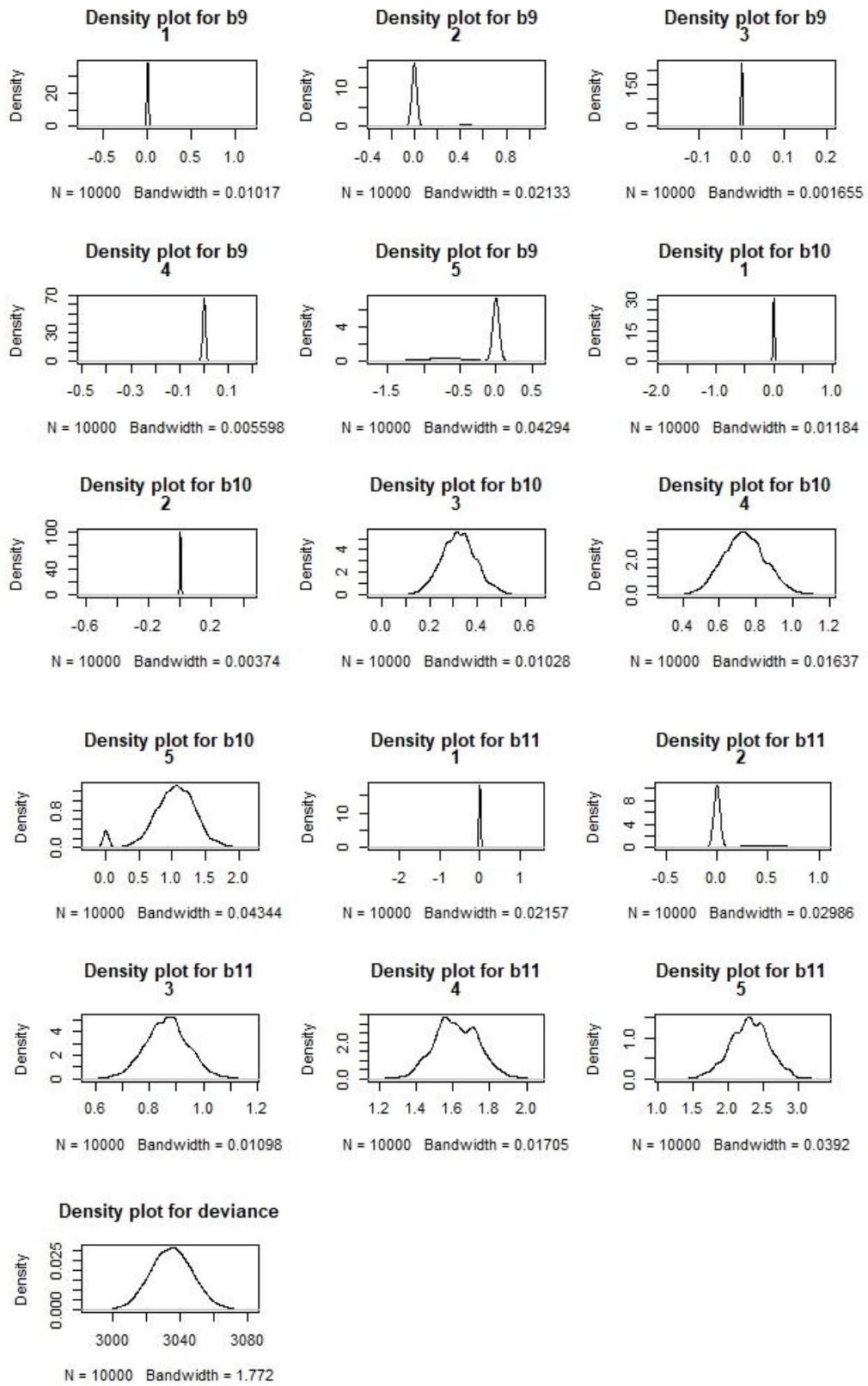
gb7[4]	-0.7	0.1	-1.0	-0.8	-0.7	-0.6	-0.4
gb7[5]	-1.1	0.6	-2.1	-1.5	-1.1	-0.7	0.0
gb8[1]	0.0	0.2	0.0	0.0	0.0	0.0	0.8
gb8[2]	0.0	0.1	-0.3	0.0	0.0	0.0	0.0
gb8[3]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb8[4]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb8[5]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb9[1]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb9[2]	0.1	0.2	0.0	0.0	0.0	0.0	0.5
gb9[3]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb9[4]	0.0	0.0	-0.1	0.0	0.0	0.0	0.0
gb9[5]	-0.1	0.3	-1.0	0.0	0.0	0.0	0.0
gb10[1]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb10[2]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb10[3]	0.3	0.1	0.2	0.3	0.3	0.4	0.5
gb10[4]	0.7	0.1	0.5	0.7	0.7	0.8	1.0
gb10[5]	1.0	0.3	0.0	0.8	1.0	1.2	1.6
gb11[1]	0.0	0.1	0.0	0.0	0.0	0.0	0.0
gb11[2]	0.1	0.2	0.0	0.0	0.0	0.0	0.7
gb11[3]	0.9	0.1	0.7	0.8	0.9	0.9	1.0
gb11[4]	1.6	0.1	1.4	1.5	1.6	1.7	1.8
gb11[5]	2.3	0.3	1.8	2.1	2.3	2.5	2.8
deviance	3036.4	13.0	3011.6	3028.3	3036.6	3044.3	3060.2

DIC info (using the rule, $pD = \text{var}(\text{deviance}) / 2$)
 $pD = 84.3$ and DIC = 3120.7
DIC is an estimate of expected predictive error (lower deviance is better).









Then in this case, the model will be looking like this:

$$\ln\left(\frac{p_{i1}}{p_{i6}}\right) = -5.7 + 1.2 \cdot Z_{i2} - 0.5 \cdot Z_{i7}$$

$$\ln\left(\frac{p_{i2}}{p_{i6}}\right) = -2.9 + 0.5 \cdot Z_{i2} - 0.1 \cdot Z_{i6} - 0.6 \cdot Z_{i7} + 0.1 \cdot Z_{i9} + 0.1 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i3}}{p_{i6}}\right) = 0.1 - 0.4 \cdot Z_{i2} - 0.1 \cdot Z_{i3} + 0.1 \cdot Z_{i6} - 0.4 \cdot Z_{i7} + 0.3 \cdot Z_{i10} + 0.9 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i4}}{p_{i6}}\right) = -2 - 0.9 \cdot Z_{i2} - 0.2 \cdot Z_{i5} - 0.7 \cdot Z_{i7} + 0.7 \cdot Z_{i10} + 1.6 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i5}}{p_{i6}}\right) = -5.60 - 0.2 \cdot Z_{i2} + 0.1 \cdot Z_{i3} - 0.6 \cdot Z_{i5} - 1.1 \cdot Z_{i7} + 1 \cdot Z_{i10} + 2.3 \cdot Z_{i11}$$

Surely, had we taken more iterations, the results would look better. Here we did not use the physiochemical characteristics of *Fixed Acidity*, *Residual Sugar* and *Density* at all, while others were used up to a certain degree at each level.

2. Using a Beta-Binomial Distribution on the priors:

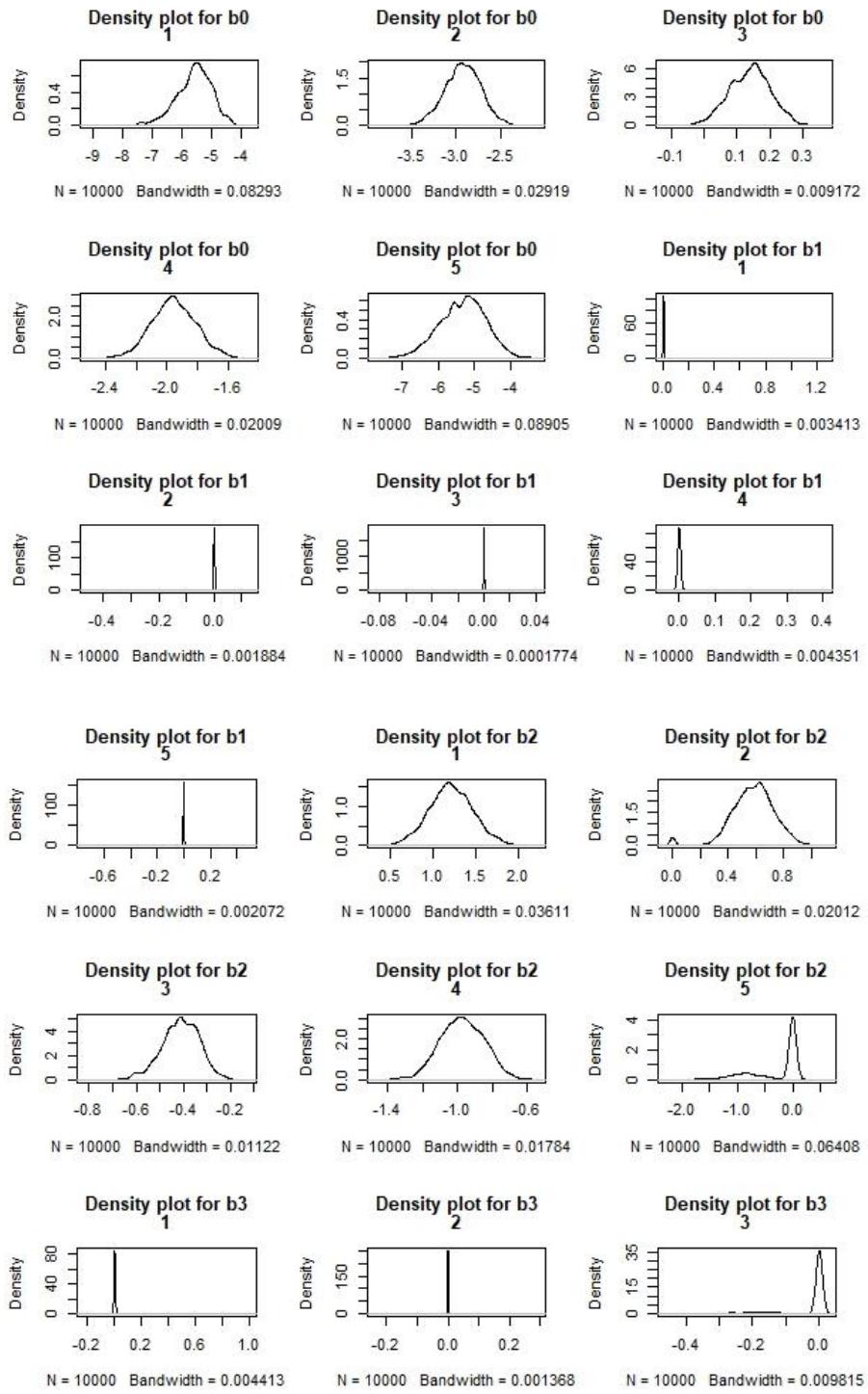
```
Inference for Bugs model at "C:\Users\30697\Desktop\empirical.txt",
fit using OpenBUGS,
 1 chains, each with 11000 iterations (first 1000 discarded), n.thin =
30
  n.sims = 333 iterations saved
      mean   sd   2.5%   25%   50%   75% 97.5%
gb0[1]  -5.6  0.6  -7.0  -5.9  -5.5  -5.2  -4.5
gb0[2]  -2.9  0.2  -3.3  -3.1  -2.9  -2.8  -2.5
gb0[3]   0.1  0.1   0.0   0.1   0.1   0.2   0.3
gb0[4]  -2.0  0.1  -2.2  -2.0  -2.0  -1.9  -1.7
gb0[5]  -5.3  0.6  -6.6  -5.8  -5.3  -4.9  -4.2
gb1[1]   0.0  0.0   0.0   0.0   0.0   0.0   0.0
gb1[2]   0.0  0.0   0.0   0.0   0.0   0.0   0.0
gb1[3]   0.0  0.0   0.0   0.0   0.0   0.0   0.0
gb1[4]   0.0  0.0   0.0   0.0   0.0   0.0   0.0
gb1[5]   0.0  0.0   0.0   0.0   0.0   0.0   0.0
gb2[1]   1.2  0.3   0.7   1.0   1.2   1.4   1.7
gb2[2]   0.6  0.2   0.3   0.5   0.6   0.7   0.9
gb2[3]  -0.4  0.1  -0.6  -0.5  -0.4  -0.4  -0.3
gb2[4]  -1.0  0.1  -1.2  -1.1  -1.0  -0.9  -0.7
gb2[5]  -0.3  0.4  -1.3  -0.6   0.0   0.0   0.0
gb3[1]   0.0  0.0   0.0   0.0   0.0   0.0   0.0
gb3[2]   0.0  0.0   0.0   0.0   0.0   0.0   0.0
```

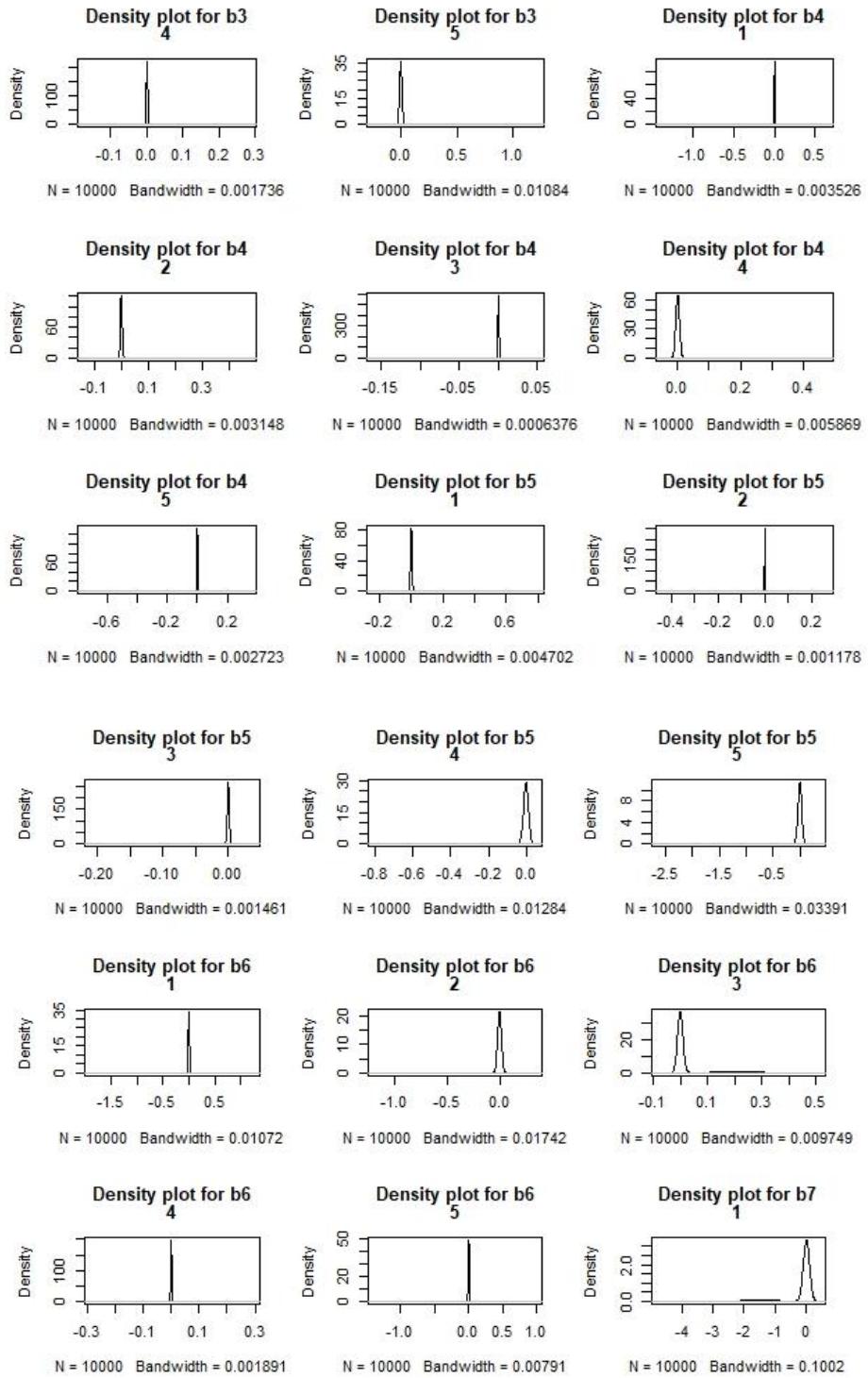
gb3[3]	0.0	0.1	-0.3	0.0	0.0	0.0	0.0
gb3[4]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb3[5]	0.0	0.1	0.0	0.0	0.0	0.0	0.0
gb4[1]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb4[2]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb4[3]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb4[4]	0.0	0.0	0.0	0.0	0.0	0.0	0.2
gb4[5]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb5[1]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb5[2]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb5[3]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb5[4]	0.0	0.1	-0.4	0.0	0.0	0.0	0.0
gb5[5]	0.0	0.2	-0.5	0.0	0.0	0.0	0.0
gb6[1]	0.0	0.1	0.0	0.0	0.0	0.0	0.0
gb6[2]	0.0	0.1	-0.5	0.0	0.0	0.0	0.0
gb6[3]	0.0	0.1	0.0	0.0	0.0	0.0	0.3
gb6[4]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb6[5]	0.0	0.1	0.0	0.0	0.0	0.0	0.0
gb7[1]	-0.3	0.7	-2.6	0.0	0.0	0.0	0.0
gb7[2]	-0.6	0.3	-1.1	-0.8	-0.7	-0.5	0.0
gb7[3]	-0.4	0.1	-0.6	-0.4	-0.4	-0.3	-0.3
gb7[4]	-0.7	0.1	-0.9	-0.8	-0.7	-0.6	-0.4
gb7[5]	-0.9	0.7	-2.0	-1.4	-1.0	0.0	0.0
gb8[1]	0.0	0.1	0.0	0.0	0.0	0.0	0.0
gb8[2]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb8[3]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb8[4]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb8[5]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb9[1]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb9[2]	0.0	0.1	0.0	0.0	0.0	0.0	0.1
gb9[3]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb9[4]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb9[5]	0.0	0.1	0.0	0.0	0.0	0.0	0.0
gb10[1]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb10[2]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb10[3]	0.3	0.1	0.1	0.2	0.3	0.3	0.4
gb10[4]	0.7	0.1	0.5	0.6	0.7	0.7	0.9
gb10[5]	0.8	0.4	0.0	0.7	0.9	1.2	1.5
gb11[1]	0.0	0.1	0.0	0.0	0.0	0.0	0.0
gb11[2]	0.1	0.2	0.0	0.0	0.0	0.0	0.6
gb11[3]	0.9	0.1	0.7	0.8	0.9	0.9	1.0
gb11[4]	1.6	0.1	1.4	1.6	1.6	1.7	1.9
gb11[5]	2.3	0.3	1.8	2.2	2.3	2.5	2.8
deviance	3051.3	11.7	3030.4	3043.2	3050.2	3058.4	3076.7

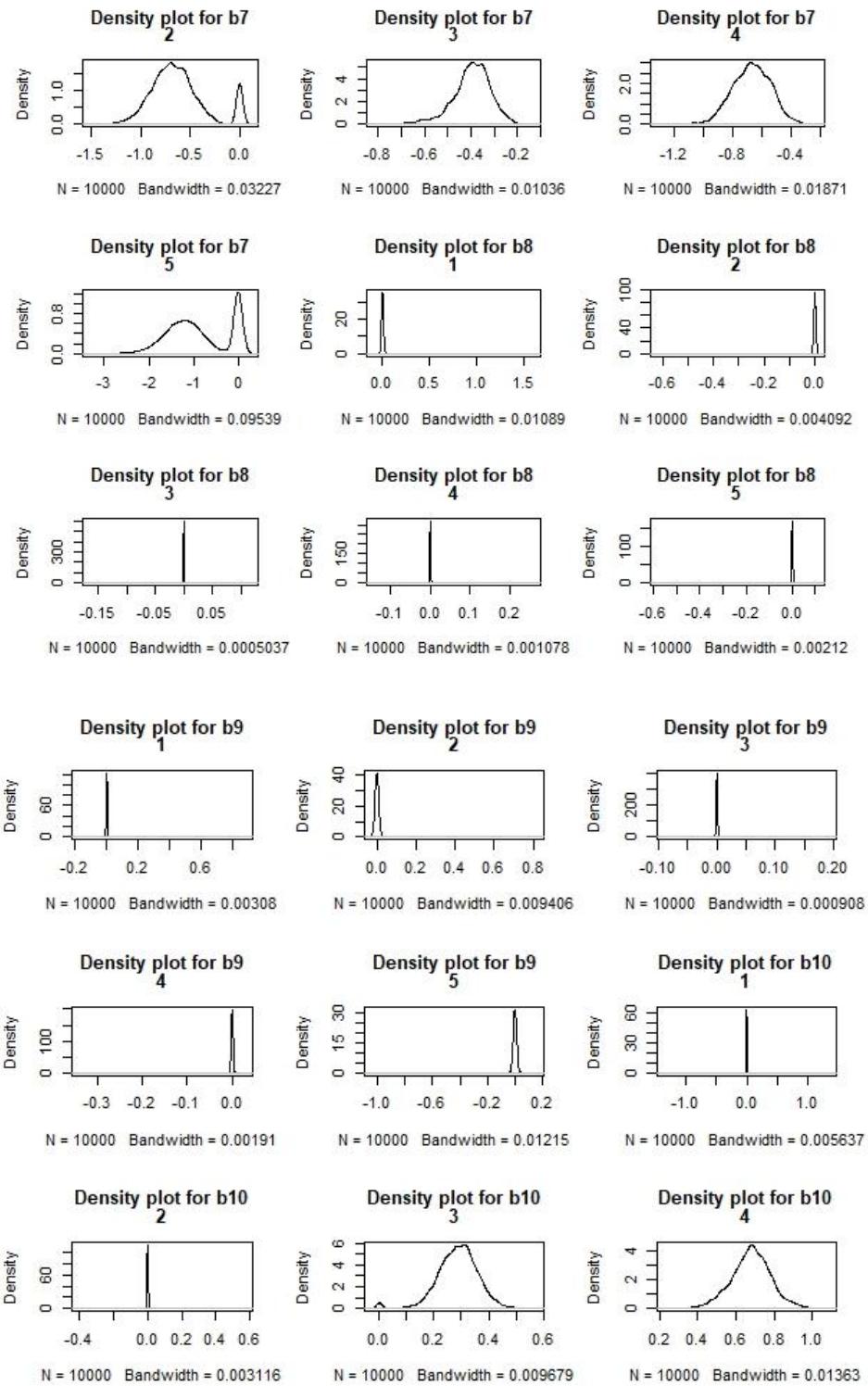
DIC info (using the rule, pD = var(deviance)/2)

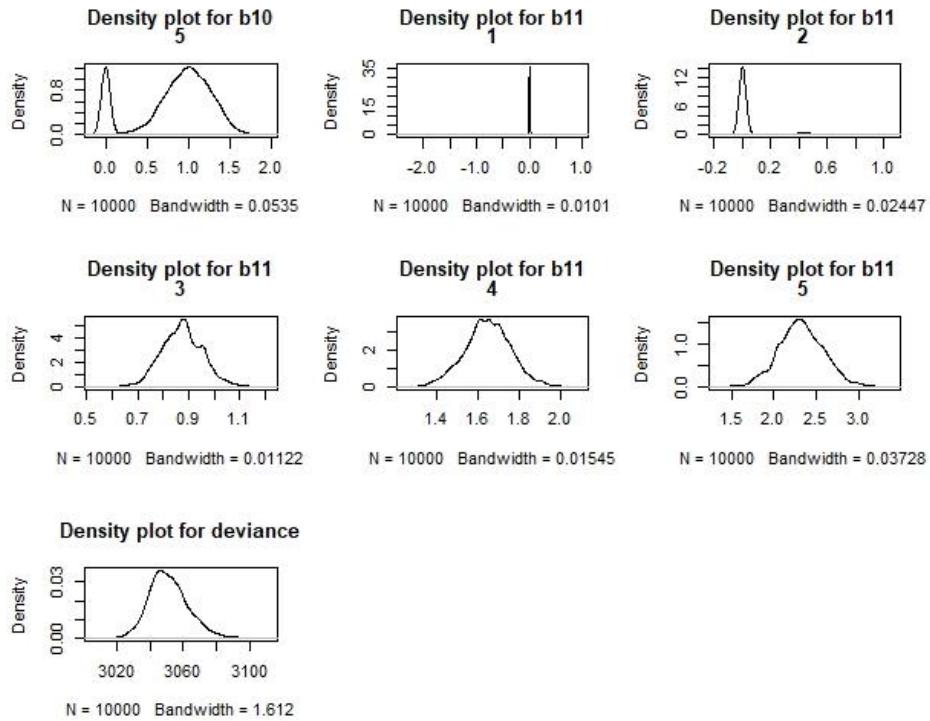
pD = 68.2 and DIC = 3119.5

DIC is an estimate of expected predictive error (lower deviance is better).









And so, this model will be:

$$\begin{aligned}
 \ln\left(\frac{p_{i1}}{p_{i6}}\right) &= -5.6 + 1.2 \cdot Z_{i2} - 0.3 \cdot Z_{i7} \\
 \ln\left(\frac{p_{i2}}{p_{i6}}\right) &= -2.9 + 0.6 \cdot Z_{i2} - 0.6 \cdot Z_{i7} + 0.1 \cdot Z_{i11} \\
 \ln\left(\frac{p_{i3}}{p_{i6}}\right) &= 0.1 - 0.4 \cdot Z_{i2} - 0.4 \cdot Z_{i7} + 0.3 \cdot Z_{i10} + 0.9 \cdot Z_{i11} \\
 \ln\left(\frac{p_{i4}}{p_{i6}}\right) &= -2 - 1 \cdot Z_{i2} - 0.7 \cdot Z_{i7} + 0.7 \cdot Z_{i10} + 1.6 \cdot Z_{i11} \\
 \ln\left(\frac{p_{i5}}{p_{i6}}\right) &= -5.3 - 0.3 \cdot Z_{i2} - 0.9 \cdot Z_{i7} + 0.8 \cdot Z_{i10} + 2.3 \cdot Z_{i11}
 \end{aligned}$$

Which is a more parsimonious model than the last one, as it did not contain the physiochemical characteristics of *Fixed Acidity*, *Citric Acid*, *Residual Sugar*, *Chlorides*, *Free Sulphur Dioxide*, *Density* and *pH*.

Variable Selection through g-prior on prior coefficients (using BAS package in R):

Another alternative for the choice of a prior distribution of β s when no prior information is available, is the Zellner's g-prior. Here, we will have:

$$\beta_{jk} \sim N(0, n \cdot (X^T \cdot X)^{-1} \cdot \sigma^2)$$

The algorithm uses the modified version of Liang et al. of this prior, with improper priors for the intercept of the model, as well as the precision (*tau*). Just like before, the process of variable selection will be done in the following two ways:

1. Using a Uniform Distribution on the priors:

The results here for each group, will be:

For level 1 (category 1 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 237 models

	post mean	post SD	post p(B != 0)
Intercept	-5.523561	0.753000	1.000000
fixed.acidity	0.021996	0.157864	0.058594
volatile.acidity	1.066413	0.289899	0.994531
citric.acid	0.042204	0.208483	0.073779
residual.sugar	0.006579	0.083114	0.043604
chlorides	0.051107	0.149910	0.131201
free.sulfur.dioxide	0.056498	0.366369	0.096045
total.sulfur.dioxide	-0.813376	0.943329	0.589404
density	0.137584	0.343356	0.179639
pH	0.019555	0.139275	0.051758
sulphates	0.002821	0.087830	0.034473
alcohol	-0.098715	0.394548	0.095264

> **summary(res2.1)**

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.000000000	1.0000000	1.00000	1.0000000
fixed.acidity	0.05859375	0.0000000	0.00000	0.0000000
volatile.acidity	0.99453125	1.0000000	1.00000	1.0000000
citric.acid	0.07377930	0.0000000	0.00000	0.0000000
residual.sugar	0.04360352	0.0000000	0.00000	0.0000000
chlorides	0.13120117	0.0000000	0.00000	0.0000000
free.sulfur.dioxide	0.09604492	0.0000000	0.00000	0.0000000
total.sulfur.dioxide	0.58940430	0.0000000	1.00000	1.0000000
density	0.17963867	0.0000000	0.00000	1.0000000
pH	0.05175781	0.0000000	0.00000	0.0000000
sulphates	0.03447266	0.0000000	0.00000	0.0000000

alcohol	0.09526367	0.0000000	0.00000	0.0000000
BF	NA	0.9791067	1.00000	0.2620167
PostProbs	NA	0.2254000	0.22420	0.0613000
R2	NA	0.1987000	0.26110	0.2979000
dim	NA	2.0000000	3.00000	4.0000000
logmarg	NA	-45.3435427	-45.32243	-46.6617752
		model 4	model 5	
Intercept	1.0000000	1.0000000		
fixed.acidity	0.0000000	0.0000000		
volatile.acidity	1.0000000	1.0000000		
citric.acid	0.0000000	0.0000000		
residual.sugar	0.0000000	0.0000000		
chlorides	1.0000000	0.0000000		
free.sulfur.dioxide	0.0000000	0.0000000		
total.sulfur.dioxide	1.0000000	0.0000000		
density	0.0000000	1.0000000		
pH	0.0000000	0.0000000		
sulphates	0.0000000	0.0000000		
alcohol	0.0000000	0.0000000		
BF	0.1783827	0.1125876		
PostProbs	0.0410000	0.0308000		
R2	0.2905000	0.2198000		
dim	4.0000000	3.0000000		
logmarg	-47.0462518	-47.5064518		

For level 2 (category 2 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 275 models				
	post mean	post SD	post p(B != 0)	
Intercept	-2.887121	0.204377	1.000000	
fixed.acidity	-0.005761	0.068728	0.057129	
volatile.acidity	0.571058	0.154100	0.988916	
citric.acid	0.004545	0.057720	0.050342	
residual.sugar	0.050481	0.121995	0.183984	
chlorides	0.002467	0.030742	0.040381	
free.sulfur.dioxide	-0.035058	0.140407	0.100586	
total.sulfur.dioxide	-0.552940	0.249282	0.909424	
density	-0.032040	0.117474	0.107178	
pH	0.064811	0.147766	0.206152	
sulphates	0.004552	0.041501	0.048584	
alcohol	0.082297	0.180961	0.212842	

> summary(res2.2)				
	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.000000000	1.00000	1.0000000	1.0000000
fixed.acidity	0.05712891	0.00000	0.0000000	0.0000000

volatile.acidity	0.98891602	1.0000	1.000000	1.0000000
citric.acid	0.05034180	0.0000	0.000000	0.0000000
residual.sugar	0.18398438	0.0000	0.000000	0.0000000
chlorides	0.04038086	0.0000	0.000000	0.0000000
free.sulfur.dioxide	0.10058594	0.0000	0.000000	0.0000000
total.sulfur.dioxide	0.90942383	1.0000	1.000000	1.0000000
density	0.10717773	0.0000	0.000000	0.0000000
pH	0.20615234	0.0000	0.000000	1.0000000
sulphates	0.04858398	0.0000	0.000000	0.0000000
alcohol	0.21284180	0.0000	1.000000	0.0000000
BF	NA	1.0000	0.317415	0.2577731
PostProbs	NA	0.3190	0.101500	0.0961000
R2	NA	0.0961	0.107400	0.1063000
dim	NA	3.0000	4.000000	4.0000000
logmarg	NA	-179.6678	-180.815324	-181.0234544
		model 4	model 5	
Intercept	1.0000000	1.00000000		
fixed.acidity	0.0000000	0.00000000		
volatile.acidity	1.0000000	1.00000000		
citric.acid	0.0000000	0.00000000		
residual.sugar	1.0000000	0.00000000		
chlorides	0.0000000	0.00000000		
free.sulfur.dioxide	0.0000000	0.00000000		
total.sulfur.dioxide	1.0000000	1.00000000		
density	0.0000000	1.00000000		
pH	0.0000000	0.00000000		
sulphates	0.0000000	0.00000000		
alcohol	0.0000000	0.00000000		
BF	0.1791273	0.08857708		
PostProbs	0.0679000	0.03490000		
R2	0.1044000	0.10070000		
dim	4.0000000	4.00000000		
logmarg	-181.3874374	-182.09166084		

For level 3 (category 3 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

	Based on the top 66 models	post mean	post SD	post p(B != 0)
Intercept	0.1443183	0.0659369	1.0000000	
fixed.acidity	0.0056524	0.0359002	0.0476562	
volatile.acidity	-0.4062907	0.0867942	0.9993652	
citric.acid	-0.0330804	0.0850380	0.1722168	
residual.sugar	0.0011866	0.0130161	0.0263672	
chlorides	-0.0427213	0.0784243	0.2743652	
free.sulfur.dioxide	0.2228540	0.1371573	0.7988770	

total.sulfur.dioxide	-0.5970115	0.1282187	0.9991699
density	0.0010487	0.0177876	0.0345215
pH	-0.0003257	0.0125768	0.0264648
sulphates	0.3512195	0.0755211	0.9992676
alcohol	0.8273962	0.0845784	0.9999023
> summary(res2.3)			
	P(B != 0 Y)	model 1	model 2
Intercept	1.000000000	1.0000	1.0000000
fixed.acidity	0.04765625	0.0000	0.0000000
volatile.acidity	0.99936523	1.0000	1.0000000
citric.acid	0.17221680	0.0000	0.0000000
residual.sugar	0.02636719	0.0000	0.0000000
chlorides	0.27436523	0.0000	1.0000000
free.sulfur.dioxide	0.79887695	1.0000	1.0000000
total.sulfur.dioxide	0.99916992	1.0000	1.0000000
density	0.03452148	0.0000	0.0000000
pH	0.02646484	0.0000	0.0000000
sulphates	0.99926758	1.0000	1.0000000
alcohol	0.99990234	1.0000	1.0000000
BF	NA	1.0000	0.4048037
PostProbs	NA	0.4276	0.1916000
R2	NA	0.1784	0.1814000
dim	NA	6.0000	7.0000000
logmarg		NA -770.4470	-771.3513829 -772.1823228
		model 4	model 5
Intercept	1.0000000	1.0000000	
fixed.acidity	0.0000000	0.0000000	
volatile.acidity	1.0000000	1.0000000	
citric.acid	1.0000000	1.0000000	
residual.sugar	0.0000000	0.0000000	
chlorides	0.0000000	0.0000000	
free.sulfur.dioxide	1.0000000	0.0000000	
total.sulfur.dioxide	1.0000000	1.0000000	
density	0.0000000	0.0000000	
pH	0.0000000	0.0000000	
sulphates	1.0000000	1.0000000	
alcohol	1.0000000	1.0000000	
BF	0.1416914	0.1079878	
PostProbs	0.0723000	0.0433000	
R2	0.1802000	0.1760000	
dim	7.0000000	6.0000000	
logmarg	-772.4011336	-772.6727669	

For level 4 (category 4 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

	Based on the top 106 models	post mean	post SD	post p(B != 0)
Intercept	-1.938019	0.156460	1.000000	
fixed.acidity	0.056255	0.132291	0.206592	
volatile.acidity	-0.980949	0.163226	0.999170	
citric.acid	-0.006592	0.067517	0.050146	
residual.sugar	0.018000	0.068111	0.094287	
chlorides	-0.126321	0.194024	0.373975	
free.sulfur.dioxide	0.117227	0.209602	0.288721	
total.sulfur.dioxide	-0.593707	0.219810	0.995605	
density	-0.007231	0.076584	0.063037	
pH	-0.031068	0.094747	0.134766	
sulphates	0.646540	0.129095	0.999805	
alcohol	1.748001	0.164713	0.999658	

> **summary(res2.4)**

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.000000000	1.0000	1.00000000	1.00000000
fixed.acidity	0.20659180	0.0000	0.00000000	0.00000000
volatile.acidity	0.99916992	1.0000	1.00000000	1.00000000
citric.acid	0.05014648	0.0000	0.00000000	0.00000000
residual.sugar	0.09428711	0.0000	0.00000000	0.00000000
chlorides	0.37397461	0.0000	1.00000000	0.00000000
free.sulfur.dioxide	0.28872070	0.0000	0.00000000	1.00000000
total.sulfur.dioxide	0.99560547	1.0000	1.00000000	1.00000000
density	0.06303711	0.0000	0.00000000	0.00000000
pH	0.13476562	0.0000	0.00000000	0.00000000
sulphates	0.99980469	1.0000	1.00000000	1.00000000
alcohol	0.99965820	1.0000	1.00000000	1.00000000
BF	NA	1.0000	0.5316337	0.3590766
PostProbs	NA	0.2611	0.1467000	0.0971000
R2	NA	0.5192	0.5250000	0.5242000
dim	NA	5.0000	6.0000000	6.0000000
logmarg	NA	-241.0493	-241.6810716	-242.0734905
		model 4	model 5	
Intercept	1.0000000	1.0000000		
fixed.acidity	1.0000000	0.0000000		
volatile.acidity	1.0000000	1.0000000		
citric.acid	0.0000000	0.0000000		
residual.sugar	0.0000000	0.0000000		
chlorides	0.0000000	1.0000000		
free.sulfur.dioxide	0.0000000	1.0000000		
total.sulfur.dioxide	1.0000000	1.0000000		
density	0.0000000	0.0000000		
pH	0.0000000	0.0000000		
sulphates	1.0000000	1.0000000		
alcohol	1.0000000	1.0000000		
BF	0.2516325	0.1408933		
PostProbs	0.0623000	0.0441000		

R2	0.5235000	0.5294000
dim	6.0000000	7.0000000
logmarg	-242.4290568	-243.0090236

For level 5 (category 5 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 405 models

	post mean	post SD	post p(B != 0)
Intercept	-6.89717	1.48251	1.00000
fixed.acidity	-0.05734	0.32288	0.07422
volatile.acidity	-0.36479	0.62581	0.30776
citric.acid	1.04250	0.97959	0.63057
residual.sugar	-1.17978	1.11055	0.62251
chlorides	-1.22246	1.51072	0.52515
free.sulfur.dioxide	0.02373	0.21598	0.06611
total.sulfur.dioxide	-1.56124	0.87070	0.87734
density	-0.34645	0.75974	0.23804
pH	-0.22178	0.49156	0.22397
sulphates	1.25975	0.55290	0.92310
alcohol	2.46388	0.87950	0.99824

> **summary(res2.5)**

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.00000000	1.00000	1.0000000	1.0000000
fixed.acidity	0.07421875	0.00000	0.0000000	0.0000000
volatile.acidity	0.30776367	0.00000	0.0000000	0.0000000
citric.acid	0.63056641	1.00000	1.0000000	1.0000000
residual.sugar	0.62250977	1.00000	0.0000000	1.0000000
chlorides	0.52514648	1.00000	1.0000000	0.0000000
free.sulfur.dioxide	0.06611328	0.00000	0.0000000	0.0000000
total.sulfur.dioxide	0.87734375	1.00000	1.0000000	1.0000000
density	0.23803711	0.00000	1.0000000	0.0000000
pH	0.22397461	0.00000	0.0000000	0.0000000
sulphates	0.92309570	1.00000	1.0000000	1.0000000
alcohol	0.99824219	1.00000	1.0000000	1.0000000
BF	NA	1.00000	0.7499369	0.8118072
PostProbs	NA	0.13470	0.1025000	0.0986000
R2	NA	0.66940	0.6660000	0.6282000
dim	NA	7.00000	7.0000000	6.0000000
logmarg	NA	-47.45245	-47.7402123	-47.6609386
		model 4	model 5	
Intercept	1.0000000	1.0000000		
fixed.acidity	0.0000000	0.0000000		
volatile.acidity	1.0000000	1.0000000		
citric.acid	0.0000000	0.0000000		
residual.sugar	1.0000000	1.0000000		

chlorides	0.0000000	0.0000000
free.sulfur.dioxide	0.0000000	0.0000000
total.sulfur.dioxide	1.0000000	1.0000000
density	0.0000000	0.0000000
pH	0.0000000	1.0000000
sulphates	1.0000000	1.0000000
alcohol	1.0000000	1.0000000
BF	0.7928453	0.3153408
PostProbs	0.0940000	0.0375000
R2	0.6280000	0.6557000
dim	6.0000000	7.0000000
logmarg	-47.6845734	-48.6065476

Thus, if we select the best model of them all, which has BF equal to 1, then we would get the following model:

$$\ln\left(\frac{p_{i1}}{p_{i6}}\right) = -5.5 + 1.07 \cdot Z_{i2} - 0.81 \cdot Z_{i7}$$

$$\ln\left(\frac{p_{i2}}{p_{i6}}\right) = -2.89 + 0.57 \cdot Z_{i2} - 0.55 \cdot Z_{i7}$$

$$\ln\left(\frac{p_{i3}}{p_{i6}}\right) = 0.14 - 0.41 \cdot Z_{i2} + 0.22 \cdot Z_{i6} - 0.59 \cdot Z_{i7} + 0.35 \cdot Z_{i10} + 0.83 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i4}}{p_{i6}}\right) = -1.94 - 0.98 \cdot Z_{i2} - 0.59 \cdot Z_{i7} + 0.65 \cdot Z_{i10} + 1.75 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i5}}{p_{i6}}\right) = -6.9 + 1.04 \cdot Z_{i3} - 1.18 \cdot Z_{i4} - 1.22 \cdot Z_{i5} - 0.88 \cdot Z_{i7} + 1.26 \cdot Z_{i10} + 2.46 \cdot Z_{i11}$$

This method did not use the physiochemical characteristics of *Fixed Acidity*, *Density* and *pH*.

2. Using a Beta-Binomial Distribution on the priors:

Here we will have the following results for each group:

For level 1 (category 1 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

	Based on the top 90 models	post mean	post SD	post p(B != 0)
Intercept	-5.2720092	0.5508172	1.0000000	
fixed.acidity	0.0039901	0.0587205	0.0102051	

volatile.acidity	1.0594639	0.2660126	0.9891602
citric.acid	0.0042257	0.0656519	0.0103027
residual.sugar	0.0013579	0.0346564	0.0093262
chlorides	0.0080483	0.0579613	0.0245605
free.sulfur.dioxide	-0.0023368	0.1130868	0.0171387
total.sulfur.dioxide	-0.1798177	0.5021136	0.1481934
density	0.0161466	0.1173092	0.0254395
pH	0.0048286	0.0631537	0.0126465
sulphates	0.0006452	0.0339597	0.0064941
alcohol	-0.0123864	0.1397073	0.0154785

> summary(res2.1)

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.000000000	1.0000000	1.00000	1.0000000
fixed.acidity	0.010205078	0.0000000	0.00000	0.0000000
volatile.acidity	0.989160156	1.0000000	1.00000	1.0000000
citric.acid	0.010302734	0.0000000	0.00000	0.0000000
residual.sugar	0.009326172	0.0000000	0.00000	0.0000000
chlorides	0.024560547	0.0000000	0.00000	1.0000000
free.sulfur.dioxide	0.017138672	0.0000000	0.00000	0.0000000
total.sulfur.dioxide	0.148193359	0.0000000	1.00000	0.0000000
density	0.025439453	0.0000000	0.00000	0.0000000
pH	0.012646484	0.0000000	0.00000	0.0000000
sulphates	0.006494141	0.0000000	0.00000	0.0000000
alcohol	0.015478516	0.0000000	0.00000	0.0000000
BF	NA	0.9791067	1.00000	0.1098552
PostProbs	NA	0.7650000	0.11090	0.0152000
R2	NA	0.1987000	0.26110	0.2193000
dim	NA	2.0000000	3.00000	3.0000000
logmarg	NA	-45.3435427	-45.32243	-47.5310203
		model 4	model 5	
Intercept	1.0000000	1.0000000		
fixed.acidity	0.0000000	0.0000000		
volatile.acidity	1.0000000	1.0000000		
citric.acid	0.0000000	0.0000000		
residual.sugar	0.0000000	0.0000000		
chlorides	0.0000000	0.0000000		
free.sulfur.dioxide	0.0000000	1.0000000		
total.sulfur.dioxide	0.0000000	0.0000000		
density	1.0000000	0.0000000		
pH	0.0000000	0.0000000		
sulphates	0.0000000	0.0000000		
alcohol	0.0000000	0.0000000		
BF	0.1125876	0.08361466		
PostProbs	0.0146000	0.00960000		
R2	0.2198000	0.21400000		
dim	3.0000000	3.0000000		
logmarg	-47.5064518	-47.80396441		

For level 2 (category 2 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 91 models

	post mean	post SD	post p(B != 0)
Intercept	-2.8981454	0.1999721	1.0000000
fixed.acidity	-0.0022860	0.0319280	0.0129883
volatile.acidity	0.5756445	0.1649886	0.9758301
citric.acid	0.0003056	0.0221035	0.0083496
residual.sugar	0.0119482	0.0617194	0.0453125
chlorides	0.0003418	0.0128964	0.0082031
free.sulfur.dioxide	-0.0211946	0.1094167	0.0468750
total.sulfur.dioxide	-0.5478193	0.2597361	0.8856934
density	-0.0070800	0.0555156	0.0235840
pH	0.0247246	0.0973603	0.0745605
sulphates	0.0007053	0.0170610	0.0093262
alcohol	0.0324694	0.1238535	0.0768066

> **summary(res2.2)**

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.000000000	1.0000	1.000000	1.000000e+00
fixed.acidity	0.012988281	0.0000	0.000000	0.000000e+00
volatile.acidity	0.975830078	1.0000	1.000000	1.000000e+00
citric.acid	0.008349609	0.0000	0.000000	0.000000e+00
residual.sugar	0.045312500	0.0000	0.000000	0.000000e+00
chlorides	0.008203125	0.0000	0.000000	0.000000e+00
free.sulfur.dioxide	0.046875000	0.0000	0.000000	0.000000e+00
total.sulfur.dioxide	0.885693359	1.0000	1.000000	0.000000e+00
density	0.023583984	0.0000	0.000000	0.000000e+00
pH	0.074560547	0.0000	0.000000	0.000000e+00
sulphates	0.009326172	0.0000	0.000000	0.000000e+00
alcohol	0.076806641	0.0000	1.000000	0.000000e+00
BF	NA	1.0000	0.317415	9.500522e-03
PostProbs	NA	0.6874	0.046600	4.180000e-02
R2	NA	0.0961	0.107400	5.430000e-02
dim	NA	3.0000	4.000000	2.000000e+00
logmarg	NA	-179.6678	-180.815324	-1.843242e+02
		model 4	model 5	
Intercept	1.0000000	1.0000000		
fixed.acidity	0.0000000	0.0000000		
volatile.acidity	1.0000000	1.0000000		
citric.acid	0.0000000	0.0000000		
residual.sugar	0.0000000	1.0000000		
chlorides	0.0000000	0.0000000		
free.sulfur.dioxide	0.0000000	0.0000000		
total.sulfur.dioxide	1.0000000	1.0000000		
density	0.0000000	0.0000000		

pH	1.0000000	0.0000000
sulphates	0.0000000	0.0000000
alcohol	0.0000000	0.0000000
BF	0.2577731	0.1791273
PostProbs	0.0416000	0.0270000
R2	0.1063000	0.1044000
dim	4.0000000	4.0000000
logmarg	-181.0234544	-181.3874374

For level 3 (category 3 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 58 models

	post mean	post SD	post p(B != 0)
Intercept	1.466e-01	6.601e-02	1.000e+00
fixed.acidity	1.770e-03	2.070e-02	2.153e-02
volatile.acidity	-4.055e-01	8.240e-02	9.996e-01
citric.acid	-2.212e-02	6.921e-02	1.188e-01
residual.sugar	4.647e-04	8.390e-03	9.521e-03
chlorides	-2.279e-02	6.124e-02	1.459e-01
free.sulfur.dioxide	1.772e-01	1.528e-01	6.308e-01
total.sulfur.dioxide	-5.631e-01	1.368e-01	9.990e-01
density	5.051e-04	1.500e-02	1.621e-02
pH	6.238e-05	9.232e-03	1.377e-02
sulphates	3.388e-01	7.159e-02	1.000e+00
alcohol	8.361e-01	8.533e-02	9.995e-01

> summary(res2.3)

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.000000000	1.0000	1.0000000	1.0000000
fixed.acidity	0.021533203	0.0000	0.0000000	0.0000000
volatile.acidity	0.999609375	1.0000	1.0000000	1.0000000
citric.acid	0.118798828	0.0000	0.0000000	0.0000000
residual.sugar	0.009521484	0.0000	0.0000000	0.0000000
chlorides	0.145898438	0.0000	0.0000000	1.0000000
free.sulfur.dioxide	0.630810547	1.0000	0.0000000	1.0000000
total.sulfur.dioxide	0.999023438	1.0000	1.0000000	1.0000000
density	0.016210938	0.0000	0.0000000	0.0000000
pH	0.013769531	0.0000	0.0000000	0.0000000
sulphates	1.000000000	1.0000	1.0000000	1.0000000
alcohol	0.999462891	1.0000	1.0000000	1.0000000
BF	NA	1.0000	0.1763485	0.4048037
PostProbs	NA	0.4766	0.2392000	0.0753000
R2	NA	0.1784	0.1726000	0.1814000
dim	NA	6.0000	5.0000000	7.0000000
logmarg	NA	-770.4470	-772.1823228	-771.3513829
	model 4		model 5	

Intercept	1.0000000	1.0000000
fixed.acidity	0.0000000	0.0000000
volatile.acidity	1.0000000	1.0000000
citric.acid	1.0000000	0.0000000
residual.sugar	0.0000000	0.0000000
chlorides	0.0000000	1.0000000
free.sulfur.dioxide	0.0000000	0.0000000
total.sulfur.dioxide	1.0000000	1.0000000
density	0.0000000	0.0000000
pH	0.0000000	0.0000000
sulphates	1.0000000	1.0000000
alcohol	1.0000000	1.0000000
BF	0.1079878	0.08741685
PostProbs	0.0520000	0.04740000
R2	0.1760000	0.17580000
dim	6.0000000	6.0000000
logmarg	-772.6727669	-772.88409712

For level 4 (category 4 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 75 models

	post mean	post SD	post p(B != 0)
Intercept	-1.928137	0.154825	1.000000
fixed.acidity	0.033836	0.103107	0.128760
volatile.acidity	-1.001509	0.156026	1.000000
citric.acid	-0.003752	0.049809	0.026221
residual.sugar	0.007770	0.045452	0.041504
chlorides	-0.054911	0.141177	0.164941
free.sulfur.dioxide	0.059618	0.159883	0.148486
total.sulfur.dioxide	-0.539286	0.194703	0.985303
density	-0.003803	0.051549	0.023779
pH	-0.012482	0.060765	0.057129
sulphates	0.623905	0.121529	0.999463
alcohol	1.760783	0.158483	0.999756

> summary(res2.4)

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.000000000	1.0000	1.000000000	1.000000000
fixed.acidity	0.12875977	0.0000	0.000000000	0.000000000
volatile.acidity	1.000000000	1.0000	1.000000000	1.000000000
citric.acid	0.02622070	0.0000	0.000000000	0.000000000
residual.sugar	0.04150391	0.0000	0.000000000	0.000000000
chlorides	0.16494141	0.0000	1.000000000	0.000000000
free.sulfur.dioxide	0.14848633	0.0000	0.000000000	1.000000000
total.sulfur.dioxide	0.98530273	1.0000	1.000000000	1.000000000
density	0.02377930	0.0000	0.000000000	0.000000000

pH	0.05712891	0.0000	0.0000000	0.0000000
sulphates	0.99946289	1.0000	1.0000000	1.0000000
alcohol	0.99975586	1.0000	1.0000000	1.0000000
BF	NA	1.0000	0.5316337	0.3590766
PostProbs	NA	0.5545	0.0997000	0.0838000
R2	NA	0.5192	0.5250000	0.5242000
dim	NA	5.0000	6.0000000	6.0000000
logmarg	NA	-241.0493	-241.6810716	-242.0734905
		model 4	model 5	
Intercept	1.0000000	1.0000000		
fixed.acidity	1.0000000	0.0000000		
volatile.acidity	1.0000000	1.0000000		
citric.acid	0.0000000	0.0000000		
residual.sugar	0.0000000	0.0000000		
chlorides	0.0000000	0.0000000		
free.sulfur.dioxide	0.0000000	0.0000000		
total.sulfur.dioxide	1.0000000	1.0000000		
density	0.0000000	0.0000000		
pH	0.0000000	1.0000000		
sulphates	1.0000000	1.0000000		
alcohol	1.0000000	1.0000000		
BF	0.2516325	0.1179672		
PostProbs	0.0627000	0.0241000		
R2	0.5235000	0.5218000		
dim	6.0000000	6.0000000		
logmarg	-242.4290568	-243.1866198		

For level 5 (category 5 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 303 models

	post mean	post SD	post p(B != 0)
Intercept	-5.7972473	1.4221738	1.0000000
fixed.acidity	-0.0429014	0.2731041	0.0518066
volatile.acidity	-0.5973103	0.7091462	0.4698730
citric.acid	0.5579142	0.8388904	0.3642090
residual.sugar	-0.6606628	0.9853124	0.3676270
chlorides	-0.6449986	1.2808107	0.2627441
free.sulfur.dioxide	-0.0001341	0.1298969	0.0284668
total.sulfur.dioxide	-0.9263280	0.9597479	0.5592773
density	-0.1727900	0.5335677	0.1228027
pH	-0.1626270	0.4346457	0.1562988
sulphates	0.7545693	0.6846952	0.6170898
alcohol	2.1752184	0.7135057	0.9988281

> summary(res2.5)

P(B != 0 Y)	model 1	model 2	model 3
---------------	---------	---------	---------

Intercept	1.00000000	1.00000000	1.00000000	1.00000000
fixed.acidity	0.05180664	0.00000000	0.00000000	0.00000000
volatile.acidity	0.46987305	1.00000000	1.00000000	0.00000000
citric.acid	0.36420898	0.00000000	0.00000000	0.00000000
residual.sugar	0.36762695	0.00000000	1.00000000	0.00000000
chlorides	0.26274414	0.00000000	0.00000000	0.00000000
free.sulfur.dioxide	0.02846680	0.00000000	0.00000000	0.00000000
total.sulfur.dioxide	0.55927734	0.00000000	1.00000000	1.00000000
density	0.12280273	0.00000000	0.00000000	0.00000000
pH	0.15629883	0.00000000	0.00000000	0.00000000
sulphates	0.61708984	0.00000000	1.00000000	1.00000000
alcohol	0.99882812	1.00000000	1.00000000	1.00000000
BF	NA	0.04219674	0.9766424	0.06508855
PostProbs	NA	0.15410000	0.0703000	0.06050000
R2	NA	0.47450000	0.6280000	0.51840000
dim	NA	3.00000000	6.0000000	4.00000000
logmarg	NA	-50.82635091	-47.6845734	-50.39294531
	model 4	model 5		
Intercept	1.00000	1.00000000		
fixed.acidity	0.00000	0.00000000		
volatile.acidity	0.00000	1.00000000		
citric.acid	1.00000	0.00000000		
residual.sugar	1.00000	0.00000000		
chlorides	0.00000	0.00000000		
free.sulfur.dioxide	0.00000	0.00000000		
total.sulfur.dioxide	1.00000	1.00000000		
density	0.00000	0.00000000		
pH	0.00000	0.00000000		
sulphates	1.00000	1.00000000		
alcohol	1.00000	1.00000000		
BF	1.00000	0.2707585		
PostProbs	0.05350	0.0501000		
R2	0.62820	0.5740000		
dim	6.00000	5.0000000		
logmarg	-47.66094	-48.9674665		

And so, the model this time will be:

$$\ln\left(\frac{p_{i1}}{p_{i6}}\right) = -5.27 + 1.06 \cdot Z_{i2} - 0.17 \cdot Z_{i7}$$

$$\ln\left(\frac{p_{i2}}{p_{i6}}\right) = -2.9 + 0.58 \cdot Z_{i2} - 0.55 \cdot Z_{i7}$$

$$\ln\left(\frac{p_{i3}}{p_{i6}}\right) = 0.15 - 0.41 \cdot Z_{i2} + 1.78 \cdot Z_{i6} - 0.56 \cdot Z_{i7} + 0.34 \cdot Z_{i10} + 0.84 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i4}}{p_{i6}}\right) = -1.93 - 1 \cdot Z_{i2} - 0.54 \cdot Z_{i7} + 0.62 \cdot Z_{i10} + 1.76 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i5}}{p_{i6}}\right) = -5.8 + 0.55 \cdot Z_{i3} - 0.66 \cdot Z_{i4} - 0.93 \cdot Z_{i7} + 0.75 \cdot Z_{i10} + 2.18 \cdot Z_{i11}$$

Looking at the model above, we see that we did not use the physiochemical characteristics of *Fixed Acidity*, *chlorides*, *pH* and *Density* at all.

Variable Selection through (Liang et al.) g-prior (using OpenBUGS):

This will be the same method as before, only this time it will be implemented in OpenBUGS. Once more we will use the version of Liang et al. Here:

$$G \leftarrow (n = 1599)$$

This method should produce similar results to the previous one, the only minor differences, should be because it uses MCMC. We will repeat the procedure by:

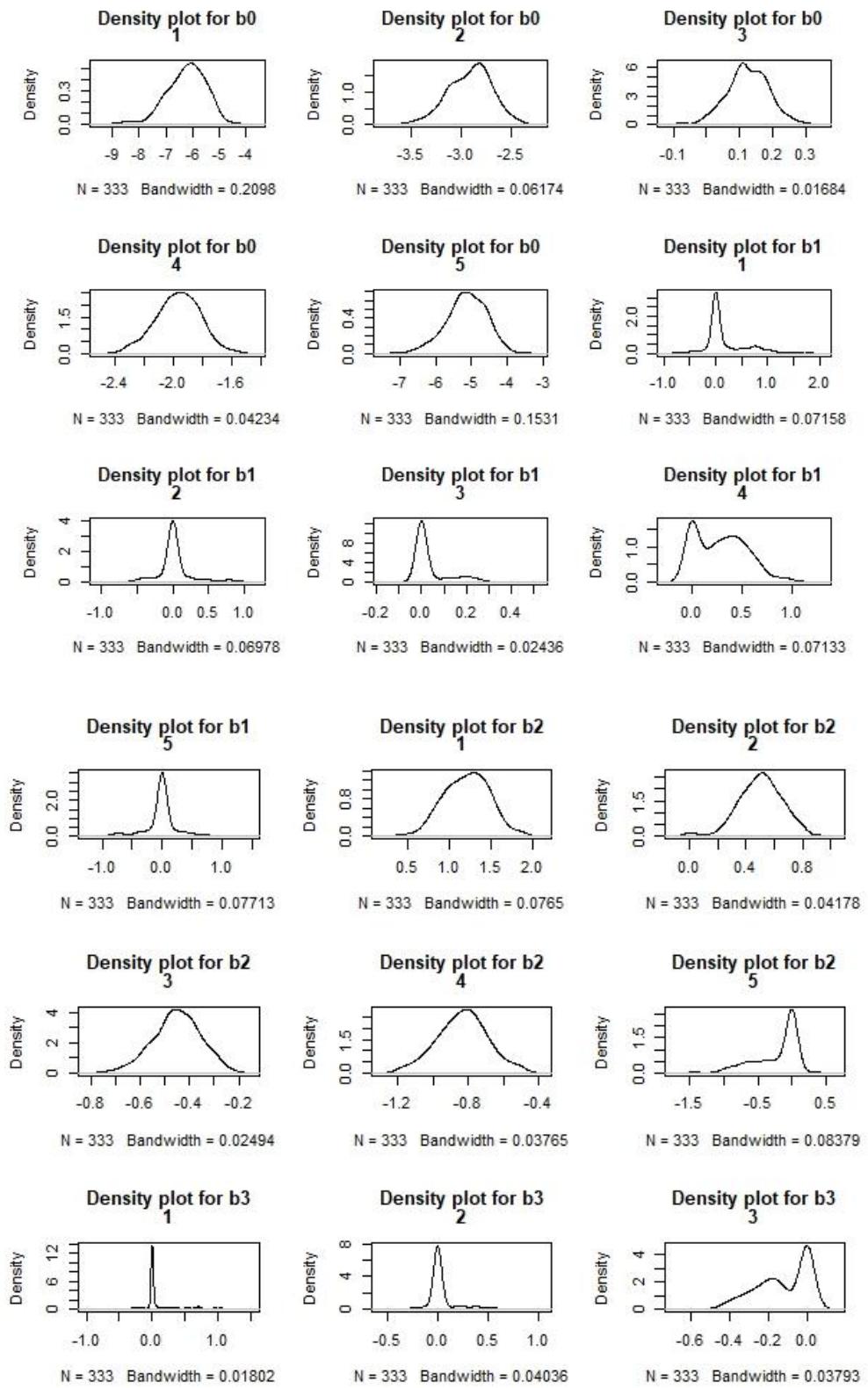
1. Using a Uniform Distribution on the priors:

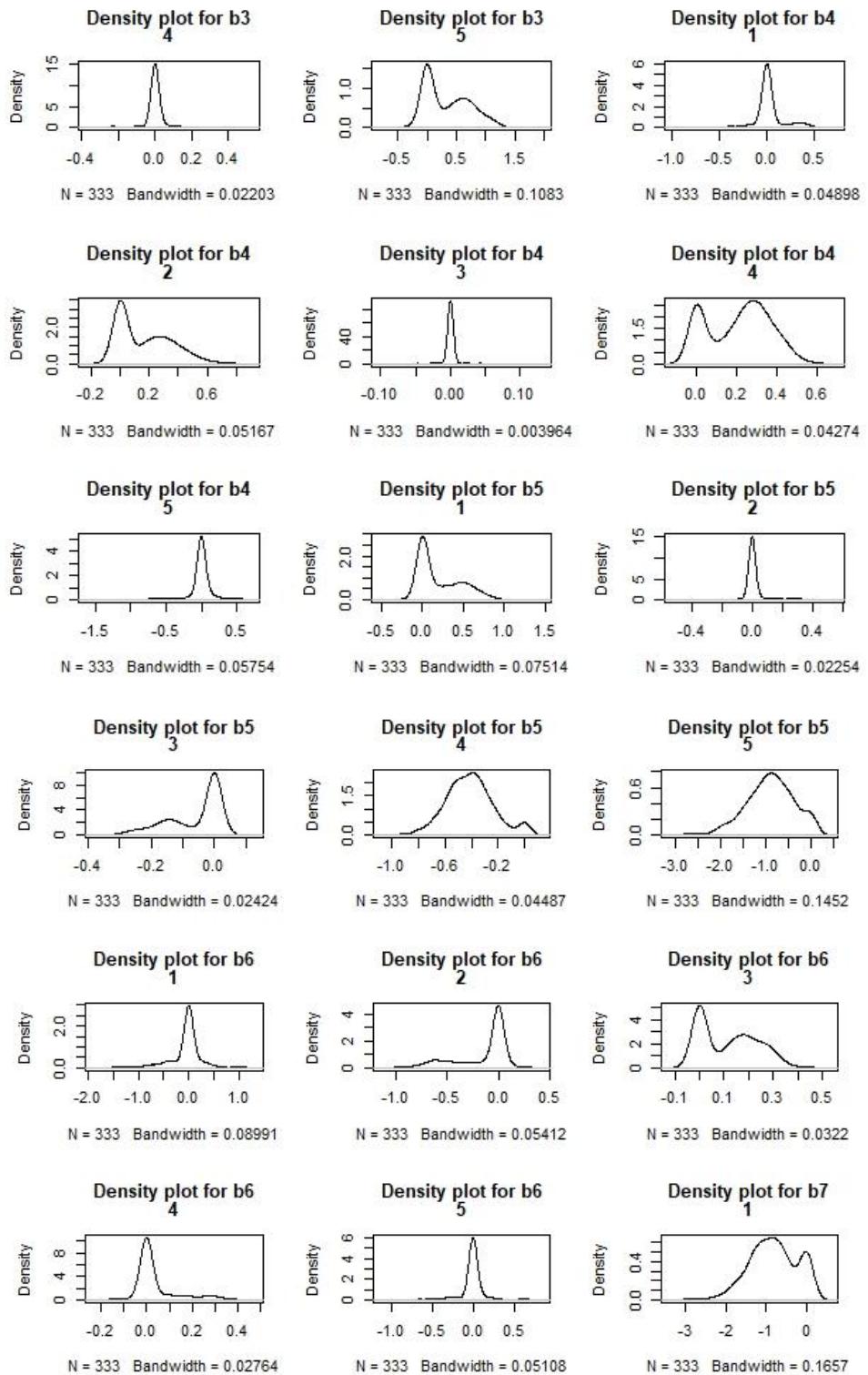
```
Inference for Bugs model at "C:\Users\30697\Desktop\gprior.txt", fit
using OpenBUGS,
 1 chains, each with 11000 iterations (first 1000 discarded), n.thin =
30
  n.sims = 333 iterations saved

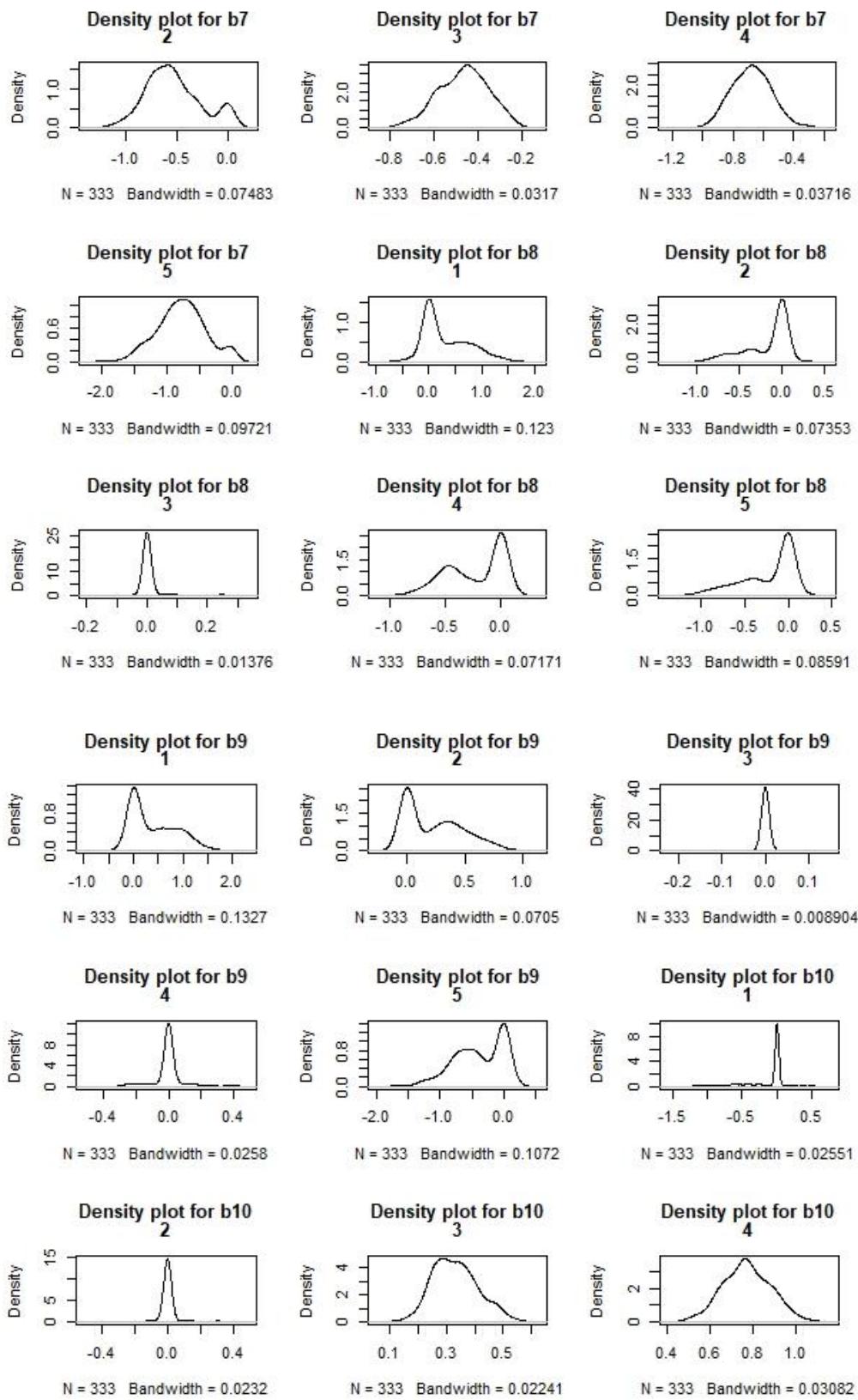
      mean    sd   2.5%   25%   50%   75%  97.5%
alpha[1]  -6.2  0.8  -8.1  -6.7  -6.2  -5.7  -5.1
alpha[2]  -2.9  0.2  -3.3  -3.1  -2.9  -2.8  -2.5
alpha[3]   0.1  0.1   0.0   0.1   0.1   0.2   0.3
alpha[4]  -2.0  0.2  -2.3  -2.1  -2.0  -1.9  -1.7
alpha[5]  -5.1  0.6  -6.3  -5.5  -5.1  -4.7  -4.1
gb1[1]   0.2  0.4  -0.4   0.0   0.0   0.3   1.3
gb1[2]   0.0  0.2  -0.5   0.0   0.0   0.0   0.8
gb1[3]   0.0  0.1   0.0   0.0   0.0   0.0   0.3
gb1[4]   0.3  0.3   0.0   0.0   0.3   0.5   0.9
gb1[5]   0.0  0.3  -0.7   0.0   0.0   0.0   0.6
gb2[1]   1.2  0.3   0.7   1.0   1.2   1.4   1.7
gb2[2]   0.5  0.2   0.2   0.4   0.5   0.6   0.8
gb2[3]  -0.5  0.1  -0.6  -0.5  -0.4  -0.4  -0.3
gb2[4]  -0.8  0.1  -1.1  -0.9  -0.8  -0.7  -0.5
gb2[5]  -0.2  0.3  -1.0  -0.4   0.0   0.0   0.0
gb3[1]   0.1  0.3  -0.4   0.0   0.0   0.1   1.0
gb3[2]   0.0  0.1  -0.2   0.0   0.0   0.0   0.5
gb3[3]  -0.1  0.1  -0.4  -0.2  -0.1   0.0   0.0
gb3[4]   0.0  0.1  -0.2   0.0   0.0   0.0   0.2
gb3[5]   0.3  0.4  -0.1   0.0   0.2   0.7   1.1
```

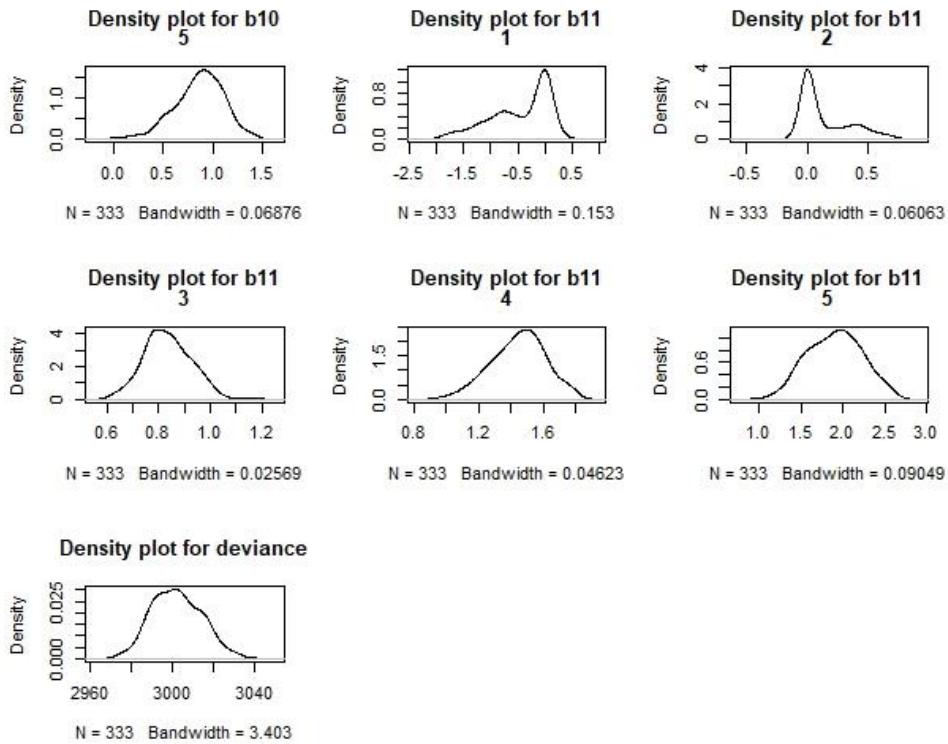
gb4[1]	0.0	0.2	-0.4	0.0	0.0	0.0	0.4
gb4[2]	0.2	0.2	0.0	0.0	0.1	0.3	0.6
gb4[3]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb4[4]	0.2	0.2	0.0	0.0	0.2	0.3	0.5
gb4[5]	0.0	0.2	-0.6	0.0	0.0	0.0	0.3
gb5[1]	0.2	0.3	0.0	0.0	0.0	0.4	0.8
gb5[2]	0.0	0.1	-0.1	0.0	0.0	0.0	0.3
gb5[3]	-0.1	0.1	-0.3	-0.1	0.0	0.0	0.0
gb5[4]	-0.4	0.2	-0.8	-0.5	-0.4	-0.3	0.0
gb5[5]	-0.9	0.5	-2.0	-1.2	-0.9	-0.5	0.0
gb6[1]	-0.1	0.3	-0.9	0.0	0.0	0.0	0.5
gb6[2]	-0.1	0.3	-0.8	-0.3	0.0	0.0	0.1
gb6[3]	0.1	0.1	0.0	0.0	0.1	0.2	0.3
gb6[4]	0.0	0.1	-0.1	0.0	0.0	0.0	0.3
gb6[5]	0.0	0.2	-0.5	0.0	0.0	0.0	0.3
gb7[1]	-0.8	0.6	-2.0	-1.2	-0.8	-0.4	0.0
gb7[2]	-0.5	0.3	-1.0	-0.7	-0.6	-0.4	0.0
gb7[3]	-0.5	0.1	-0.7	-0.5	-0.5	-0.4	-0.3
gb7[4]	-0.7	0.1	-0.9	-0.8	-0.7	-0.6	-0.4
gb7[5]	-0.8	0.4	-1.5	-1.0	-0.8	-0.5	0.0
gb8[1]	0.3	0.4	-0.3	0.0	0.0	0.6	1.3
gb8[2]	-0.2	0.3	-0.8	-0.3	0.0	0.0	0.1
gb8[3]	0.0	0.0	0.0	0.0	0.0	0.0	0.2
gb8[4]	-0.2	0.3	-0.7	-0.5	-0.2	0.0	0.0
gb8[5]	-0.2	0.3	-0.9	-0.4	0.0	0.0	0.0
gb9[1]	0.4	0.5	-0.1	0.0	0.2	0.8	1.4
gb9[2]	0.2	0.3	0.0	0.0	0.2	0.4	0.8
gb9[3]	0.0	0.0	-0.1	0.0	0.0	0.0	0.0
gb9[4]	0.0	0.1	-0.2	0.0	0.0	0.0	0.2
gb9[5]	-0.4	0.4	-1.2	-0.7	-0.4	0.0	0.0
gb10[1]	-0.1	0.3	-1.1	-0.1	0.0	0.0	0.3
gb10[2]	0.0	0.1	-0.1	0.0	0.0	0.0	0.3
gb10[3]	0.3	0.1	0.2	0.3	0.3	0.4	0.5
gb10[4]	0.8	0.1	0.5	0.7	0.8	0.8	1.0
gb10[5]	0.9	0.3	0.3	0.7	0.9	1.0	1.3
gb11[1]	-0.5	0.5	-1.7	-0.8	-0.3	0.0	0.1
gb11[2]	0.1	0.2	-0.1	0.0	0.0	0.3	0.7
gb11[3]	0.8	0.1	0.7	0.8	0.8	0.9	1.0
gb11[4]	1.4	0.2	1.1	1.3	1.5	1.6	1.8
gb11[5]	1.9	0.3	1.3	1.7	1.9	2.1	2.5
deviance	3002.2	12.1	2980.8	2993.1	3001.7	3011.3	3026.2

DIC info (using the rule, $pD = \text{var}(\text{deviance})/2$)
 $pD = 73.0$ and DIC = 3075.2
DIC is an estimate of expected predictive error (lower deviance is better).









Therefore, the model that we get will be:

$$\ln\left(\frac{p_{i1}}{p_{i6}}\right) = -6.2 + 1.2 \cdot Z_{i2} + 0.2 \cdot Z_{i5} - 0.8 \cdot Z_{i7} + 0.3 \cdot Z_{i8} + 0.4 \cdot Z_{i9} - 0.5 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i2}}{p_{i6}}\right) = -2.9 + 0.5 \cdot Z_{i2} + 0.4 \cdot Z_{i4} - 0.5 \cdot Z_{i7} - 0.2 \cdot Z_{i8} + 0.2 \cdot Z_{i9} + 0.1 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i3}}{p_{i6}}\right) = 0.1 - 0.5 \cdot Z_{i2} - 0.1 \cdot Z_{i3} - 0.9 \cdot Z_{i5} + 0.1 \cdot Z_{i6} - 0.5 \cdot Z_{i7} + 0.3 \cdot Z_{i10} + 0.8 \cdot Z_{i11}$$

$$\begin{aligned} \ln\left(\frac{p_{i4}}{p_{i6}}\right) = & -2 + 0.3 \cdot Z_{i1} - 0.8 \cdot Z_{i2} + 0.2 \cdot Z_{i4} - 0.4 \cdot Z_{i5} - 0.7 \cdot Z_{i7} - 0.2 \cdot Z_{i8} + \\ & + 0.8 \cdot Z_{i10} + 1.4 \cdot Z_{i11} \end{aligned}$$

$$\begin{aligned} \ln\left(\frac{p_{i5}}{p_{i6}}\right) = & -5.1 - 0.2 \cdot Z_{i2} + 0.3 \cdot Z_{i3} - 0.9 \cdot Z_{i5} - 0.8 \cdot Z_{i7} - 0.2 \cdot Z_{i8} - 0.4 \cdot Z_{i9} + \\ & + 0.9 \cdot Z_{i10} + 1.9 \cdot Z_{i11} \end{aligned}$$

This is not an ideal result, as it does not resemble the result of the BAS package and some variables that were previously excluded were included in this method. This could be due to the MCMC that is used, or an error in the code itself. Perhaps more iterations could smooth out the results and make this model look somewhat more like the one in BAS above.

2. Using a Beta-Binomial Distribution on the priors:

```
Inference for Bugs model at "C:\Users\30697\Desktop\gprior.txt", fit  
using OpenBUGS,  
 1 chains, each with 11000 iterations (first 1000 discarded), n.thin =  
30  
n.sims = 333 iterations saved  
      mean    sd   2.5%   25%   50%   75% 97.5%  
alpha[1] -6.4  0.8  -7.9  -6.9  -6.3  -5.8 -5.0  
alpha[2] -2.9  0.2  -3.3  -3.1  -2.9  -2.8 -2.5  
alpha[3]  0.1  0.1   0.0   0.1   0.1   0.2  0.3  
alpha[4] -2.0  0.2  -2.3  -2.1  -2.0  -1.9 -1.7  
alpha[5] -5.1  0.6  -6.3  -5.5  -5.1  -4.7 -4.1  
gb1[1]   0.2  0.4  -0.5   0.0   0.0   0.3  1.2  
gb1[2]   0.1  0.2  -0.3   0.0   0.0   0.1  0.8  
gb1[3]   0.0  0.1   0.0   0.0   0.0   0.0  0.3  
gb1[4]   0.3  0.3   0.0   0.0   0.4   0.6  0.8  
gb1[5]   0.0  0.3  -0.8   0.0   0.0   0.0  0.7  
gb2[1]   1.2  0.3   0.7   1.0   1.2   1.4  1.8  
gb2[2]   0.5  0.2   0.2   0.4   0.5   0.6  0.8  
gb2[3]  -0.4  0.1  -0.7  -0.5  -0.4  -0.4 -0.3  
gb2[4]  -0.8  0.1  -1.1  -0.9  -0.8  -0.7 -0.6  
gb2[5]  -0.2  0.3  -1.0  -0.5   0.0   0.0  0.1  
gb3[1]   0.1  0.3  -0.4   0.0   0.0   0.1  1.0  
gb3[2]   0.0  0.1  -0.2   0.0   0.0   0.0  0.4  
gb3[3]  -0.1  0.1  -0.4  -0.2  -0.1   0.0  0.0  
gb3[4]   0.0  0.1  -0.3   0.0   0.0   0.0  0.3  
gb3[5]   0.4  0.4   0.0   0.0   0.3   0.7  1.2  
gb4[1]   0.0  0.2  -0.6   0.0   0.0   0.0  0.5  
gb4[2]   0.2  0.2   0.0   0.0   0.2   0.4  0.6  
gb4[3]   0.0  0.0   0.0   0.0   0.0   0.0  0.1  
gb4[4]   0.2  0.2   0.0   0.0   0.3   0.3  0.5  
gb4[5]   0.0  0.2  -0.5   0.0   0.0   0.0  0.4  
gb5[1]   0.3  0.3   0.0   0.0   0.2   0.5  0.9  
gb5[2]   0.0  0.1  -0.1   0.0   0.0   0.0  0.3  
gb5[3]  -0.1  0.1  -0.3  -0.1   0.0   0.0  0.0  
gb5[4]  -0.4  0.2  -0.7  -0.5  -0.4  -0.3  0.0  
gb5[5]  -0.9  0.5  -1.9  -1.2  -0.8  -0.5  0.0  
gb6[1]   0.0  0.3  -0.8   0.0   0.0   0.0  0.7  
gb6[2]  -0.1  0.2  -0.7  -0.2   0.0   0.0  0.1  
gb6[3]   0.1  0.1   0.0   0.0   0.1   0.2  0.3  
gb6[4]   0.0  0.1  -0.1   0.0   0.0   0.0  0.3  
gb6[5]  -0.1  0.2  -0.7   0.0   0.0   0.0  0.4  
gb7[1]  -0.9  0.5  -2.0  -1.2  -0.9  -0.6  0.0  
gb7[2]  -0.6  0.3  -1.1  -0.8  -0.6  -0.4  0.0  
gb7[3]  -0.5  0.1  -0.7  -0.5  -0.5  -0.4  -0.2  
gb7[4]  -0.7  0.2  -1.0  -0.8  -0.7  -0.6  -0.4  
gb7[5]  -0.8  0.4  -1.5  -1.0  -0.8  -0.5  0.0
```

```

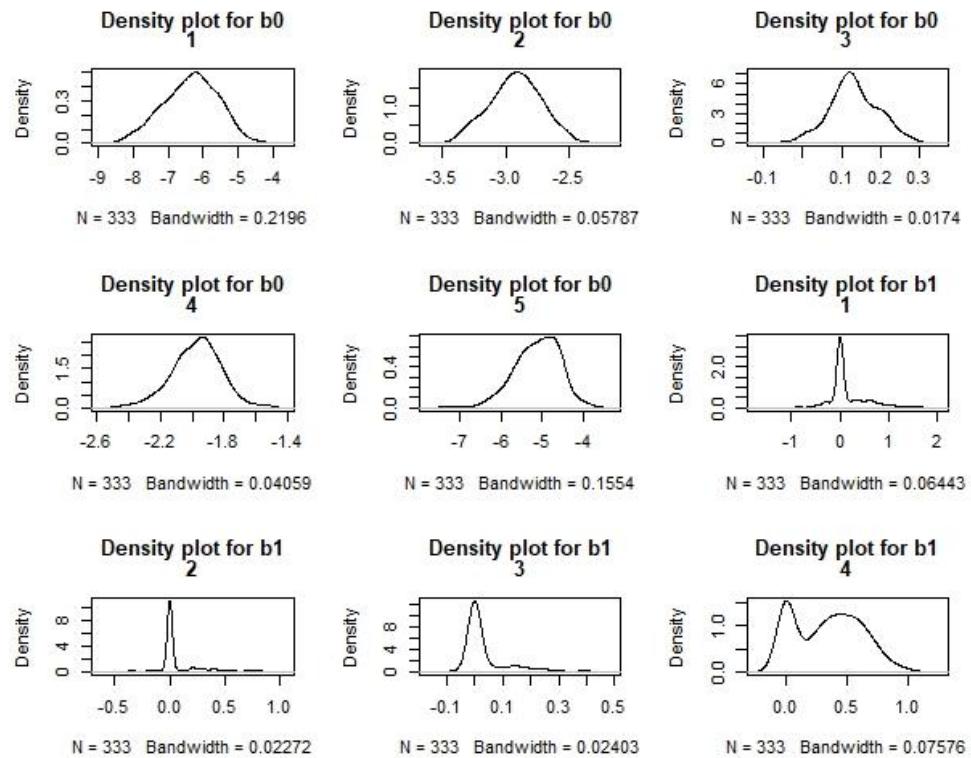
gb8[1]      0.4  0.4  -0.2   0.0   0.2   0.7  1.3
gb8[2]     -0.2  0.3  -0.8  -0.4  -0.1   0.0  0.0
gb8[3]      0.0  0.0   0.0   0.0   0.0   0.0  0.2
gb8[4]     -0.3  0.3  -0.8  -0.5  -0.3   0.0  0.0
gb8[5]     -0.3  0.4  -1.2  -0.6  -0.1   0.0  0.0
gb9[1]      0.4  0.5  -0.1   0.0   0.3   0.7  1.4
gb9[2]      0.3  0.3   0.0   0.0   0.3   0.5  0.7
gb9[3]      0.0  0.0  -0.1   0.0   0.0   0.0  0.0
gb9[4]      0.0  0.1  -0.2   0.0   0.0   0.0  0.3
gb9[5]     -0.4  0.4  -1.1  -0.6  -0.3   0.0  0.0
gb10[1]    -0.2  0.4  -1.2  -0.4   0.0   0.0  0.4
gb10[2]     0.0  0.1  -0.2   0.0   0.0   0.0  0.3
gb10[3]     0.3  0.1   0.2   0.3   0.3   0.4  0.5
gb10[4]     0.8  0.1   0.5   0.7   0.8   0.9  1.0
gb10[5]     0.9  0.3   0.4   0.7   0.9   1.1  1.3
gb11[1]    -0.5  0.5  -1.7  -0.9  -0.4   0.0  0.0
gb11[2]     0.1  0.2  -0.1   0.0   0.0   0.1  0.6
gb11[3]     0.8  0.1   0.7   0.8   0.8   0.9  1.0
gb11[4]     1.4  0.2   1.1   1.3   1.4   1.5  1.7
gb11[5]     1.8  0.3   1.2   1.6   1.8   2.0  2.4
deviance 3001.1 14.0 2973.4 2991.6 3001.6 3009.8 3031.2

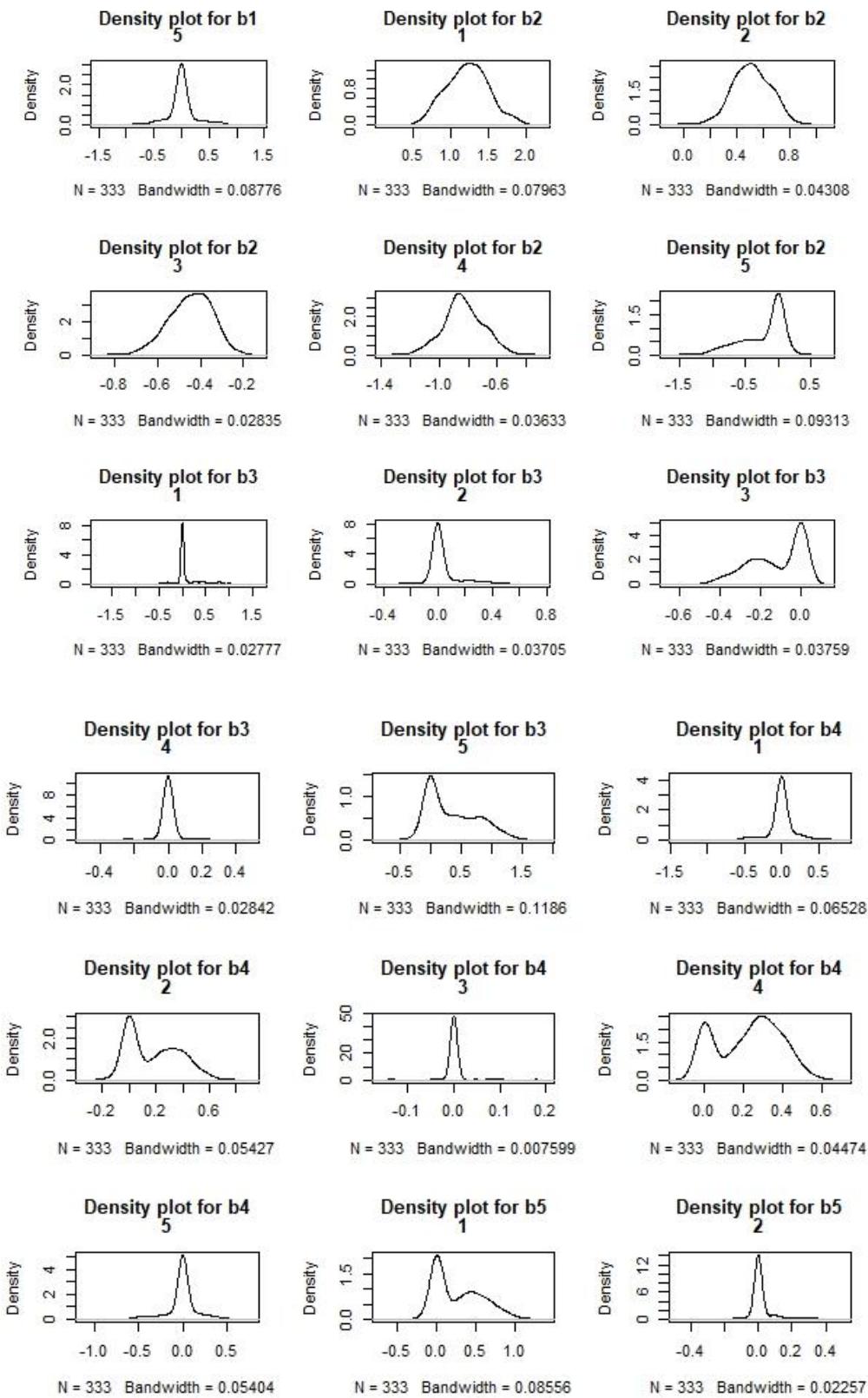
```

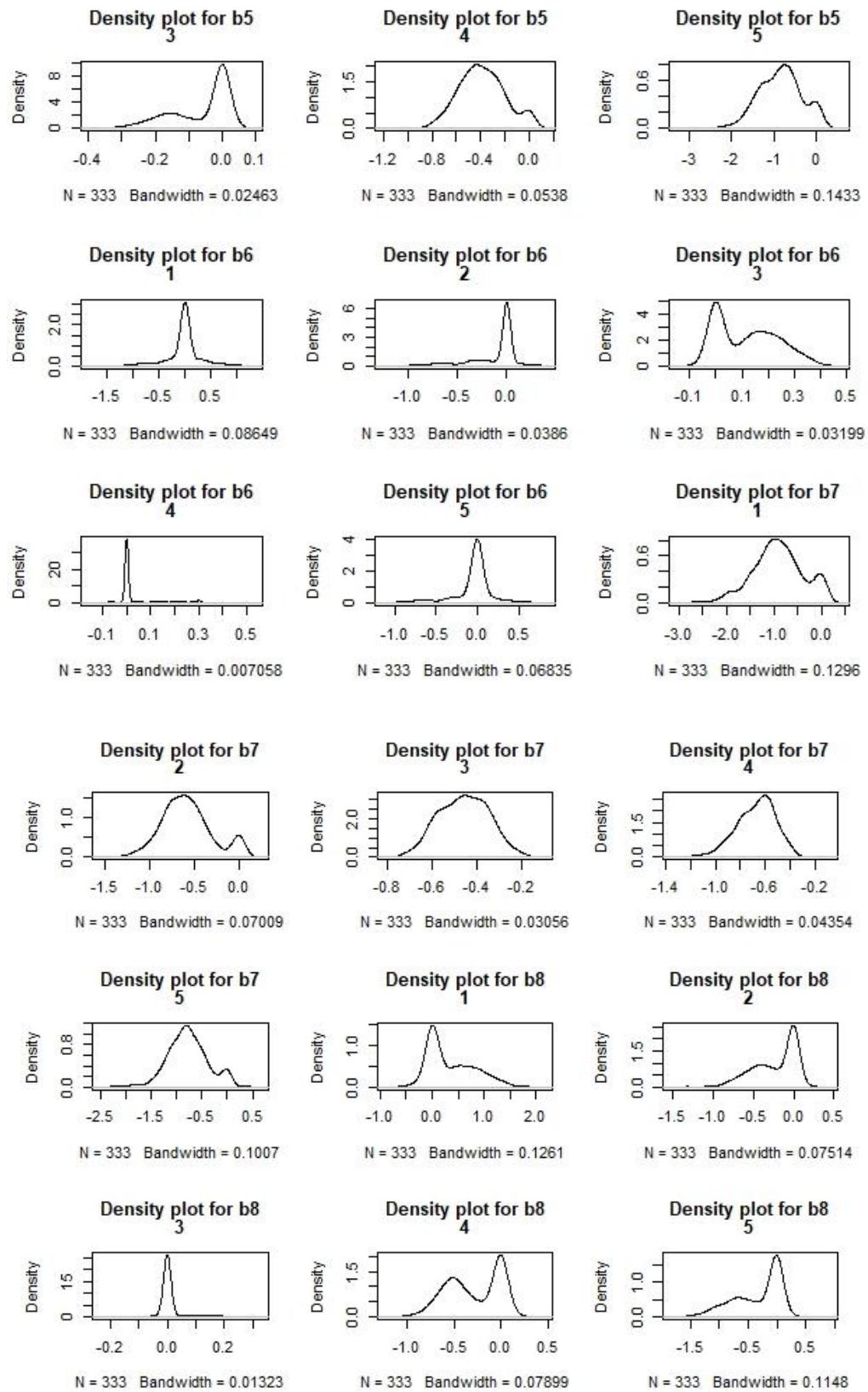
DIC info (using the rule, $pD = \text{var}(\text{deviance})/2$)

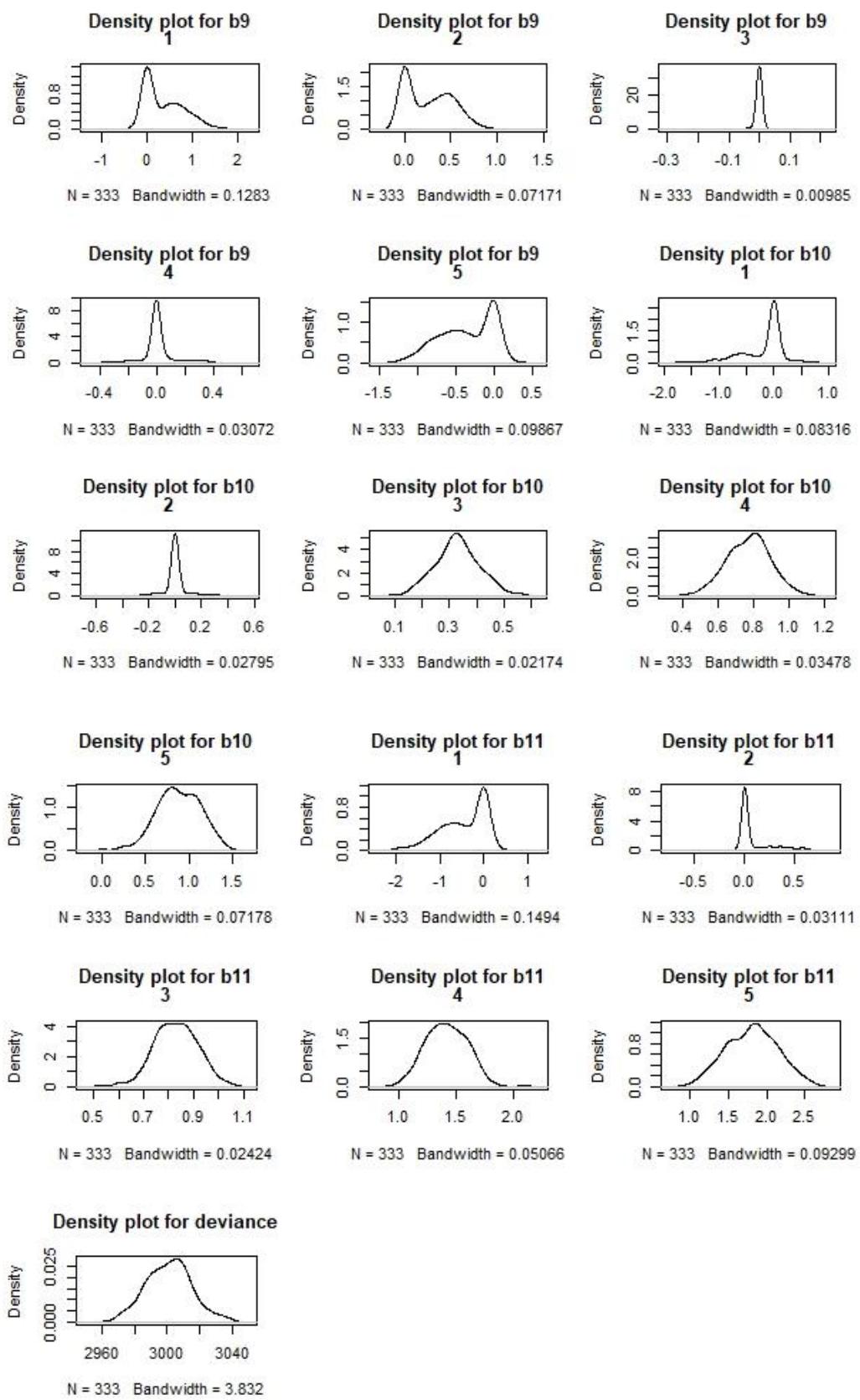
$pD = 97.7$ and DIC = 3098.8

DIC is an estimate of expected predictive error (lower deviance is better).









And so, this model is of the following form:

$$\ln\left(\frac{p_{i1}}{p_{i6}}\right) = -6.4 + 0.2 \cdot Z_{i1} + 1.2 \cdot Z_{i2} + 0.3 \cdot Z_{i5} - 0.9 \cdot Z_{i7} + 0.4 \cdot Z_{i8} + 0.4 \cdot Z_{i9} - \\ - 0.2 \cdot Z_{i10} - 0.5 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i2}}{p_{i6}}\right) = -2.9 + 0.1 \cdot Z_{i1} + 0.5 \cdot Z_{i2} + 0.2 \cdot Z_{i4} - 0.6 \cdot Z_{i7} - 0.2 \cdot Z_{i8} + 0.3 \cdot Z_{i9} + 0.1 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i3}}{p_{i6}}\right) = 0.1 - 0.4 \cdot Z_{i2} - 0.1 \cdot Z_{i3} - 0.1 \cdot Z_{i5} + 0.1 \cdot Z_{i6} - 0.5 \cdot Z_{i7} + 0.3 \cdot Z_{i10} + 0.8 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i4}}{p_{i6}}\right) = -2 + 0.3 \cdot Z_{i1} - 0.8 \cdot Z_{i2} + 0.2 \cdot Z_{i4} - 0.4 \cdot Z_{i5} - 0.7 \cdot Z_{i7} - 0.3 \cdot Z_{i8} + \\ + 0.8 \cdot Z_{i10} + 1.4 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i5}}{p_{i6}}\right) = -5.1 - 0.2 \cdot Z_{i2} + 0.4 \cdot Z_{i3} - 0.9 \cdot Z_{i5} - 0.1 \cdot Z_{i6} - 0.8 \cdot Z_{i7} - 0.3 \cdot Z_{i8} - 0.4 \cdot Z_{i9} + \\ + 0.9 \cdot Z_{i10} + 1.8 \cdot Z_{i11}$$

These are pretty similar results to the Uniform prior we calculated above.

Variable Selection through hyper-g on prior coefficients (using BAS package in R):

Similar method to the one above, meaning that once more:

$$\beta_{jk} \sim N(0, G \cdot (X^T \cdot X)^{-1} \cdot \sigma^2)$$

Only now, we have a different hyperprior, which is:

$$a \leftarrow 3 \\ bw \leftarrow a/2 - 1 \\ w \sim dbeta(1, bw) \\ G \leftarrow w/(1 - w)$$

And so, we repeat the procedure once again, by:

1. Using a Uniform Distribution on the priors:

Here we shall have:

For level 1 (category 1 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 597 models

	post mean	post SD	post p(B != 0)
Intercept	-7.52315	1.49508	1.00000
fixed.acidity	0.25773	0.64414	0.54600
volatile.acidity	0.84906	0.33209	0.99683
citric.acid	0.23838	0.46934	0.49570
residual.sugar	-0.01754	0.20251	0.42070
chlorides	0.28271	0.24538	0.78384
free.sulfur.dioxide	0.72020	0.71562	0.80117
total.sulfur.dioxide	-1.82046	1.24840	0.99258
density	0.37208	0.54611	0.59717
pH	0.44995	0.52270	0.69106
sulphates	-0.02880	0.27377	0.36753
alcohol	-0.74725	0.72534	0.81191

> summary(res3.1)

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.0000000	1.00000	1.0000000	1.0000000
fixed.acidity	0.5459961	1.00000	1.0000000	0.0000000
volatile.acidity	0.9968262	1.00000	1.0000000	1.0000000
citric.acid	0.4957031	0.00000	0.0000000	1.0000000
residual.sugar	0.4207031	0.00000	0.0000000	1.0000000
chlorides	0.7838379	1.00000	1.0000000	1.0000000
free.sulfur.dioxide	0.8011719	1.00000	1.0000000	1.0000000
total.sulfur.dioxide	0.9925781	1.00000	1.0000000	1.0000000
density	0.5971680	0.00000	1.0000000	1.0000000
pH	0.6910645	1.00000	1.0000000	1.0000000
sulphates	0.3675293	0.00000	0.0000000	0.0000000
alcohol	0.8119141	1.00000	1.0000000	1.0000000
BF	NA	1.00000	0.7122447	0.5627455
PostProbs	NA	0.03990	0.0251000	0.0218000
R2	NA	0.42060	0.4245000	0.4289000
dim	NA	8.00000	9.0000000	10.0000000
logmarg	NA	-38.72036	-39.0596942	-39.2952883
		model 4	model 5	
Intercept	1.0000000	1.0000000		
fixed.acidity	1.0000000	1.0000000		
volatile.acidity	1.0000000	1.0000000		
citric.acid	0.0000000	1.0000000		
residual.sugar	0.0000000	0.0000000		
chlorides	1.0000000	1.0000000		
free.sulfur.dioxide	1.0000000	1.0000000		
total.sulfur.dioxide	1.0000000	1.0000000		
density	0.0000000	0.0000000		
pH	1.0000000	1.0000000		
sulphates	1.0000000	0.0000000		

alcohol	1.0000000	1.0000000
BF	0.5837728	0.6661712
PostProbs	0.0195000	0.0184000
R2	0.4211000	0.4239000
dim	9.0000000	9.0000000
logmarg	-39.2586040	-39.1265691

For level 2 (category 2 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 839 models

	post mean	post SD	post p(B != 0)
Intercept	-2.86369	0.20785	1.00000
fixed.acidity	0.10712	0.29831	0.42842
volatile.acidity	0.46271	0.14223	0.99438
citric.acid	0.05396	0.15641	0.35137
residual.sugar	0.19549	0.17493	0.69678
chlorides	0.02573	0.08736	0.32349
free.sulfur.dioxide	-0.06036	0.17042	0.36519
total.sulfur.dioxide	-0.44209	0.23341	0.91304
density	-0.21645	0.29171	0.56333
pH	0.22298	0.23757	0.65645
sulphates	0.04183	0.11490	0.35405
alcohol	0.06646	0.19198	0.43115

> **summary(res3.2)**

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.0000000	1.0000	1.0000000	1.0000000
fixed.acidity	0.4284180	1.0000	1.0000000	1.0000000
volatile.acidity	0.9943848	1.0000	1.0000000	1.0000000
citric.acid	0.3513672	0.0000	0.0000000	0.0000000
residual.sugar	0.6967773	1.0000	1.0000000	1.0000000
chlorides	0.3234863	0.0000	0.0000000	1.0000000
free.sulfur.dioxide	0.3651855	0.0000	0.0000000	0.0000000
total.sulfur.dioxide	0.9130371	1.0000	1.0000000	1.0000000
density	0.5633301	1.0000	1.0000000	1.0000000
pH	0.6564453	1.0000	1.0000000	1.0000000
sulphates	0.3540527	0.0000	1.0000000	0.0000000
alcohol	0.4311523	0.0000	0.0000000	0.0000000
BF	NA	1.0000	0.7531036	0.7124517
PostProbs	NA	0.0199	0.0142000	0.0132000
R2	NA	0.1303	0.1344000	0.1342000
dim	NA	7.0000	8.0000000	8.0000000
logmarg	NA	-176.7948	-177.0783963	-177.1338871
	model 4	model 5		
Intercept	1.0000000	1.0000000		

fixed.acidity	0.0000000	0.0000000
volatile.acidity	1.0000000	1.0000000
citric.acid	0.0000000	0.0000000
residual.sugar	1.0000000	1.0000000
chlorides	0.0000000	0.0000000
free.sulfur.dioxide	0.0000000	0.0000000
total.sulfur.dioxide	1.0000000	1.0000000
density	1.0000000	1.0000000
pH	1.0000000	0.0000000
sulphates	0.0000000	1.0000000
alcohol	0.0000000	0.0000000
BF	0.6853099	0.3956922
PostProbs	0.0115000	0.0103000
R2	0.1218000	0.1184000
dim	6.0000000	6.0000000
logmarg	-177.1727279	-177.7219624

For level 3 (category 3 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 125 models

	post mean	post SD	post p(B != 0)
Intercept	0.141387	0.066142	1.000000
fixed.acidity	0.039904	0.087145	0.292480
volatile.acidity	-0.421703	0.098488	0.999707
citric.acid	-0.107425	0.133124	0.531250
residual.sugar	0.008734	0.033869	0.184473
chlorides	-0.081154	0.088043	0.574951
free.sulfur.dioxide	0.239161	0.108734	0.929688
total.sulfur.dioxide	-0.590159	0.112643	0.999951
density	0.006282	0.045699	0.161572
pH	-0.005276	0.036958	0.161084
sulphates	0.367034	0.076644	0.998877
alcohol	0.797093	0.088799	0.999561

> summary(res3.3)

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.0000000	1.0000	1.0000000	1.0000000
fixed.acidity	0.2924805	0.0000	1.0000000	0.0000000
volatile.acidity	0.9997070	1.0000	1.0000000	1.0000000
citric.acid	0.5312500	0.0000	1.0000000	0.0000000
residual.sugar	0.1844727	0.0000	0.0000000	0.0000000
chlorides	0.5749512	1.0000	0.0000000	0.0000000
free.sulfur.dioxide	0.9296875	1.0000	1.0000000	1.0000000
total.sulfur.dioxide	0.9999512	1.0000	1.0000000	1.0000000
density	0.1615723	0.0000	0.0000000	0.0000000

pH	0.1610840	0.0000	0.0000000	0.0000000
sulphates	0.9988770	1.0000	1.0000000	1.0000000
alcohol	0.9995605	1.0000	1.0000000	1.0000000
BF	NA	1.0000	0.4438559	0.4659427
PostProbs	NA	0.1681	0.0775000	0.0694000
R2	NA	0.1814	0.1825000	0.1784000
dim	NA	7.0000	8.0000000	6.0000000
logmarg	NA	-765.9311	-766.7433254	-766.6947627
		model 4	model 5	
Intercept	1.0000000	1.0000000		
fixed.acidity	0.0000000	0.0000000		
volatile.acidity	1.0000000	1.0000000		
citric.acid	1.0000000	1.0000000		
residual.sugar	0.0000000	0.0000000		
chlorides	0.0000000	1.0000000		
free.sulfur.dioxide	1.0000000	1.0000000		
total.sulfur.dioxide	1.0000000	1.0000000		
density	0.0000000	0.0000000		
pH	0.0000000	0.0000000		
sulphates	1.0000000	1.0000000		
alcohol	1.0000000	1.0000000		
BF	0.3621207	0.3908064		
PostProbs	0.0669000	0.0653000		
R2	0.1802000	0.1823000		
dim	7.0000000	8.0000000		
logmarg	-766.9468479	-766.8706131		

For level 4 (category 4 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 139 models

	post mean	post SD	post p(B != 0)
Intercept	-1.96680	0.16004	1.00000
fixed.acidity	0.29379	0.29372	0.67393
volatile.acidity	-0.86458	0.18231	0.99985
citric.acid	-0.04244	0.14466	0.25996
residual.sugar	0.15445	0.18189	0.54561
chlorides	-0.22809	0.20644	0.70747
free.sulfur.dioxide	0.32244	0.23688	0.77344
total.sulfur.dioxide	-0.72767	0.23563	0.99893
density	-0.21311	0.31194	0.46133
pH	-0.04897	0.14938	0.32085
sulphates	0.67686	0.13398	0.99961
alcohol	1.57815	0.21830	0.99917

> summary(res3.4)

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.0000000	1.0000	1.0000000	1.0000000
fixed.acidity	0.6739258	1.0000	1.0000000	0.0000000
volatile.acidity	0.9998535	1.0000	1.0000000	1.0000000
citric.acid	0.2599609	0.0000	0.0000000	0.0000000
residual.sugar	0.5456055	1.0000	1.0000000	0.0000000
chlorides	0.7074707	1.0000	0.0000000	1.0000000
free.sulfur.dioxide	0.7734375	1.0000	1.0000000	1.0000000
total.sulfur.dioxide	0.9989258	1.0000	1.0000000	1.0000000
density	0.4613281	1.0000	1.0000000	0.0000000
pH	0.3208496	0.0000	0.0000000	1.0000000
sulphates	0.9996094	1.0000	1.0000000	1.0000000
alcohol	0.9991699	1.0000	1.0000000	1.0000000
BF	NA	1.0000	0.3934193	0.4334659
PostProbs	NA	0.1100	0.0526000	0.0499000
R2	NA	0.5436	0.5381000	0.5350000
dim	NA	10.0000	9.0000000	8.0000000
logmarg	NA	-236.4561	-237.3889340	-237.2919970
		model 4	model 5	
Intercept	1.0000000	1.0000000		
fixed.acidity	1.0000000	1.0000000		
volatile.acidity	1.0000000	1.0000000		
citric.acid	0.0000000	1.0000000		
residual.sugar	0.0000000	1.0000000		
chlorides	1.0000000	1.0000000		
free.sulfur.dioxide	1.0000000	1.0000000		
total.sulfur.dioxide	1.0000000	1.0000000		
density	0.0000000	1.0000000		
pH	0.0000000	0.0000000		
sulphates	1.0000000	1.0000000		
alcohol	1.0000000	1.0000000		
BF	0.4033167	0.2719545		
PostProbs	0.0476000	0.0335000		
R2	0.5348000	0.5440000		
dim	8.0000000	11.0000000		
logmarg	-237.3640880	-237.7581753		

For level 5 (category 5 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 151 models

	post mean	post SD	post p(B != 0)
Intercept	-8.18184	1.73365	1.00000
fixed.acidity	-0.56748	0.64713	0.73643
volatile.acidity	-0.03611	0.38643	0.45522

citric.acid	1.05922	0.70389	0.95942
residual.sugar	-0.83269	0.66020	0.93423
chlorides	-1.03723	0.96602	0.97109
free.sulfur.dioxide	0.17914	0.38303	0.54082
total.sulfur.dioxide	-1.24010	0.66914	0.99634
density	-0.08360	0.59972	0.50195
pH	-0.57133	0.54758	0.82681
sulphates	0.81304	0.35972	0.99888
alcohol	1.63723	0.68058	1.00000

> summary(res3.5)

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.0000000	1.00000	1.0000000	1.0000000
fixed.acidity	0.7364258	1.00000	1.0000000	1.0000000
volatile.acidity	0.4552246	0.00000	0.0000000	1.0000000
citric.acid	0.9594238	1.00000	1.0000000	1.0000000
residual.sugar	0.9342285	1.00000	1.0000000	1.0000000
chlorides	0.9710937	1.00000	1.0000000	1.0000000
free.sulfur.dioxide	0.5408203	1.00000	0.0000000	1.0000000
total.sulfur.dioxide	0.9963379	1.00000	1.0000000	1.0000000
density	0.5019531	0.00000	0.0000000	0.0000000
pH	0.8268066	1.00000	1.0000000	1.0000000
sulphates	0.9988770	1.00000	1.0000000	1.0000000
alcohol	1.0000000	1.00000	1.0000000	1.0000000
BF	NA	1.00000	0.8888327	0.7741582
PostProbs	NA	0.10440	0.1008000	0.0809000
R2	NA	0.71730	0.7116000	0.7180000
dim	NA	10.00000	9.0000000	11.0000000
logmarg	NA	-32.84751	-32.9653535	-33.1034863
		model 4	model 5	
Intercept	1.0000000	1.000000		
fixed.acidity	1.0000000	1.000000		
volatile.acidity	0.0000000	1.000000		
citric.acid	1.0000000	1.000000		
residual.sugar	1.0000000	1.000000		
chlorides	1.0000000	1.000000		
free.sulfur.dioxide	1.0000000	0.000000		
total.sulfur.dioxide	1.0000000	1.000000		
density	1.0000000	0.000000		
pH	1.0000000	1.000000		
sulphates	1.0000000	1.000000		
alcohol	1.0000000	1.000000		
BF	0.6939117	0.565809		
PostProbs	0.0793000	0.069800		
R2	0.7178000	0.711700		
dim	11.0000000	10.000000		
logmarg	-33.2129179	-33.417006		

And so, the model will be:

$$\ln\left(\frac{p_{i1}}{p_{i6}}\right) = -7.52 + 0.25 \cdot Z_{i1} + 0.85 \cdot Z_{i2} + 0.28 \cdot Z_{i5} + 0.72 \cdot Z_{i6} - 1.8 \cdot Z_{i7} + 0.3 \cdot Z_{i8} + \\ + 0.44 \cdot Z_{i9} - 0.75 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i2}}{p_{i6}}\right) = -2.86 + 0.1 \cdot Z_{i1} + 0.46 \cdot Z_{i2} + 0.2 \cdot Z_{i4} - 0.44 \cdot Z_{i7} - 0.22 \cdot Z_{i8} + 0.22 \cdot Z_{i9}$$

$$\ln\left(\frac{p_{i3}}{p_{i6}}\right) = 0.14 - 0.42 \cdot Z_{i2} - 0.08 \cdot Z_{i5} + 0.24 \cdot Z_{i6} - 0.6 \cdot Z_{i7} + 0.37 \cdot Z_{i10} + 0.8 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i4}}{p_{i6}}\right) = -1.97 + 0.29 \cdot Z_{i1} - 0.86 \cdot Z_{i2} + 0.15 \cdot Z_{i4} - 0.22 \cdot Z_{i5} + 0.32 \cdot Z_{i6} - \\ - 0.73 \cdot Z_{i7} - 0.21 \cdot Z_{i8} + 0.68 \cdot Z_{i10} + 1.58 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i5}}{p_{i6}}\right) = -8.18 - 0.57 \cdot Z_{i1} + 1.06 \cdot Z_{i3} - 0.83 \cdot Z_{i4} - 1.04 \cdot Z_{i5} + 0.18 \cdot Z_{i6} - \\ - 1.24 \cdot Z_{i7} - 0.57 \cdot Z_{i9} + 0.81 \cdot Z_{i10} + 1.63 \cdot Z_{i11}$$

2. Using a Beta-Binomial Distribution on the priors:

Our results will be:

For level 1 (category 1 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

	Based on the top 714 models	post mean	post SD	post p(B != 0)
Intercept	-6.8430230	1.5042012	1.0000000	
fixed.acidity	0.1617056	0.5145016	0.3564453	
volatile.acidity	0.8787058	0.3153759	0.9950684	
citric.acid	0.1836641	0.4138383	0.3675293	
residual.sugar	-0.0004211	0.1671873	0.2753418	
chlorides	0.2022830	0.2381807	0.5709473	
free.sulfur.dioxide	0.4963420	0.6921159	0.5653809	
total.sulfur.dioxide	-1.5038692	1.2120729	0.9023438	
density	0.3061322	0.4860572	0.4829102	
pH	0.2768490	0.4610304	0.4469727	
sulphates	-0.0142461	0.2225155	0.2402344	
alcohol	-0.5392789	0.6986689	0.5900391	
> summary(res3.1)				
	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.0000000	1.000000000	1.000000000	1.00000
fixed.acidity	0.3564453	0.000000000	0.000000000	1.00000
volatile.acidity	0.9950684	1.000000000	1.000000000	1.00000

citric.acid	0.3675293	0.000000000	0.000000000	0.00000
residual.sugar	0.2753418	0.000000000	0.000000000	0.00000
chlorides	0.5709473	0.000000000	0.000000000	1.00000
free.sulfur.dioxide	0.5653809	0.000000000	0.000000000	1.00000
total.sulfur.dioxide	0.9023438	0.000000000	1.000000000	1.00000
density	0.4829102	0.000000000	0.000000000	0.00000
pH	0.4469727	0.000000000	0.000000000	1.00000
sulphates	0.2402344	0.000000000	0.000000000	0.00000
alcohol	0.5900391	0.000000000	0.000000000	1.00000
BF	NA	0.001661367	0.01156175	1.00000
PostProbs	NA	0.036800000	0.03670000	0.02450
R2	NA	0.198700000	0.26110000	0.42060
dim	NA	2.000000000	3.000000000	8.00000
logmarg	NA	-45.120475228	-43.18041339	-38.72036
		model 4	model 5	
Intercept	1.000000000	1.0000000		
fixed.acidity	0.000000000	0.0000000		
volatile.acidity	1.000000000	1.0000000		
citric.acid	0.000000000	0.0000000		
residual.sugar	0.000000000	0.0000000		
chlorides	0.000000000	1.0000000		
free.sulfur.dioxide	0.000000000	0.0000000		
total.sulfur.dioxide	1.000000000	1.0000000		
density	1.000000000	0.0000000		
pH	0.000000000	0.0000000		
sulphates	0.000000000	0.0000000		
alcohol	0.000000000	0.0000000		
BF	0.02169067	0.0180953		
PostProbs	0.01430000	0.0134000		
R2	0.29790000	0.2905000		
dim	4.000000000	4.0000000		
logmarg	-42.55123353	-42.7324634		

For level 2 (category 2 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

	post mean	post SD	post p(B != 0)
Intercept	-2.87874	0.20794	1.00000
fixed.acidity	0.03418	0.19330	0.19663
volatile.acidity	0.48114	0.15086	0.98267
citric.acid	0.02062	0.10388	0.15469
residual.sugar	0.10450	0.15533	0.39712
chlorides	0.01143	0.05975	0.15415
free.sulfur.dioxide	-0.04813	0.15391	0.21890
total.sulfur.dioxide	-0.46098	0.23781	0.89351
density	-0.09611	0.21153	0.27876
pH	0.13334	0.19911	0.42666

sulphates	0.01528	0.07241	0.14238
alcohol	0.07382	0.17420	0.29150
> summary(res3.2)			
	P(B != 0 Y)	model 1	model 2
Intercept	1.0000000	1.0000000	1.0000000
fixed.acidity	0.1966309	0.0000000	0.0000000
volatile.acidity	0.9826660	1.0000000	1.0000000
citric.acid	0.1546875	0.0000000	0.0000000
residual.sugar	0.3971191	0.0000000	0.0000000
chlorides	0.1541504	0.0000000	0.0000000
free.sulfur.dioxide	0.2188965	0.0000000	0.0000000
total.sulfur.dioxide	0.8935059	1.0000000	1.0000000
density	0.2787598	0.0000000	0.0000000
pH	0.4266602	0.0000000	0.0000000
sulphates	0.1423828	0.0000000	0.0000000
alcohol	0.2915039	0.0000000	1.0000000
BF	NA	0.3970406	0.7446754
PostProbs	NA	0.1388000	0.0623000
R2	NA	0.0961000	0.1074000
dim	NA	3.0000000	4.0000000
logmarg	NA	-178.0274924	-177.3985826
	model 3	model 4	model 5
Intercept	1.0000000	1.0000000	1.0000
fixed.acidity	0.0000000	0.0000000	0.0000
volatile.acidity	1.0000000	1.0000000	1.0000
citric.acid	0.0000000	0.0000000	0.0000
residual.sugar	0.0000000	1.0000000	1.0000
chlorides	0.0000000	0.0000000	0.0000
free.sulfur.dioxide	0.0000000	0.0000000	0.0000
total.sulfur.dioxide	1.0000000	1.0000000	1.0000
density	0.0000000	0.0000000	0.0000
pH	1.0000000	0.0000000	1.0000
sulphates	0.0000000	0.0000000	0.0000
alcohol	0.0000000	0.0000000	0.0000
BF	0.6489227	0.4632912	1.0000
PostProbs	0.0552000	0.0343000	0.0272
R2	0.1063000	0.1044000	0.1156
dim	4.0000000	4.0000000	5.0000
logmarg	-177.5362175	-177.8731751	-177.1038

For level 3 (category 3 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 118 models

post mean	post SD	post p(B != 0)
-----------	---------	----------------

Intercept	0.142465	0.066061	1.000000
fixed.acidity	0.027626	0.075117	0.203418
volatile.acidity	-0.413797	0.095587	0.999902
citric.acid	-0.082520	0.124862	0.406543
residual.sugar	0.006284	0.029072	0.133008
chlorides	-0.067377	0.086408	0.468799
free.sulfur.dioxide	0.228102	0.121249	0.870215
total.sulfur.dioxide	-0.583919	0.121366	0.998096
density	0.004795	0.039211	0.117285
pH	-0.003445	0.031899	0.126123
sulphates	0.358153	0.077227	0.998193
alcohol	0.800701	0.087438	0.999707

> summary(res3.3)

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.0000000	1.0000	1.0000000	1.0000000
fixed.acidity	0.2034180	0.0000	0.0000000	0.0000000
volatile.acidity	0.9999023	1.0000	1.0000000	1.0000000
citric.acid	0.4065430	0.0000	0.0000000	1.0000000
residual.sugar	0.1330078	0.0000	0.0000000	0.0000000
chlorides	0.4687988	1.0000	0.0000000	0.0000000
free.sulfur.dioxide	0.8702148	1.0000	1.0000000	1.0000000
total.sulfur.dioxide	0.9980957	1.0000	1.0000000	1.0000000
density	0.1172852	0.0000	0.0000000	0.0000000
pH	0.1261230	0.0000	0.0000000	0.0000000
sulphates	0.9981934	1.0000	1.0000000	1.0000000
alcohol	0.9997070	1.0000	1.0000000	1.0000000
BF	NA	1.0000	0.4659427	0.3621207
PostProbs	NA	0.1854	0.1850000	0.0697000
R2	NA	0.1814	0.1784000	0.1802000
dim	NA	7.0000	6.0000000	7.0000000
logmarg	NA	-765.9311	-766.6947627	-766.9468479
		model 4	model 5	
Intercept	1.0000000	1.0000000		
fixed.acidity	1.0000000	0.0000000		
volatile.acidity	1.0000000	1.0000000		
citric.acid	1.0000000	1.0000000		
residual.sugar	0.0000000	0.0000000		
chlorides	0.0000000	1.0000000		
free.sulfur.dioxide	1.0000000	1.0000000		
total.sulfur.dioxide	1.0000000	1.0000000		
density	0.0000000	0.0000000		
pH	0.0000000	0.0000000		
sulphates	1.0000000	1.0000000		
alcohol	1.0000000	1.0000000		
BF	0.4438559	0.3908064		
PostProbs	0.0528000	0.0443000		
R2	0.1825000	0.1823000		
dim	8.0000000	8.0000000		

logmarg -766.7433254 -766.8706131

For level 4 (category 4 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 138 models

	post mean	post SD	post p(B != 0)
Intercept	-1.95849	0.15929	1.00000
fixed.acidity	0.21308	0.26791	0.53838
volatile.acidity	-0.89002	0.17810	0.99976
citric.acid	-0.02680	0.11912	0.17441
residual.sugar	0.10976	0.16537	0.40825
chlorides	-0.19564	0.20780	0.60122
free.sulfur.dioxide	0.25818	0.24706	0.62329
total.sulfur.dioxide	-0.68087	0.23985	0.99795
density	-0.13858	0.27207	0.32080
pH	-0.04653	0.13498	0.26587
sulphates	0.65850	0.13567	0.99844
alcohol	1.62077	0.20550	0.99980

> summary(res3.4)

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.0000000	1.0000	1.00000000	1.0000000
fixed.acidity	0.5383789	1.0000	0.00000000	0.0000000
volatile.acidity	0.9997559	1.0000	1.00000000	1.0000000
citric.acid	0.1744141	0.0000	0.00000000	0.0000000
residual.sugar	0.4082520	1.0000	0.00000000	0.0000000
chlorides	0.6012207	1.0000	0.00000000	1.0000000
free.sulfur.dioxide	0.6232910	1.0000	0.00000000	0.0000000
total.sulfur.dioxide	0.9979492	1.0000	1.00000000	1.0000000
density	0.3208008	1.0000	0.00000000	0.0000000
pH	0.2658691	0.0000	0.00000000	0.0000000
sulphates	0.9984375	1.0000	1.00000000	1.0000000
alcohol	0.9998047	1.0000	1.00000000	1.0000000
BF	NA	1.0000	0.05858977	0.1362797
PostProbs	NA	0.0771	0.05530000	0.0479000
R2	NA	0.5436	0.51920000	0.5250000
dim	NA	10.0000	5.00000000	6.0000000
logmarg	NA	-236.4561	-239.29324998	-238.4491009
		model 4	model 5	
Intercept	1.0000000	1.0000000		
fixed.acidity	1.0000000	0.0000000		
volatile.acidity	1.0000000	1.0000000		
citric.acid	0.0000000	0.0000000		
residual.sugar	0.0000000	0.0000000		
chlorides	1.0000000	1.0000000		
free.sulfur.dioxide	1.0000000	1.0000000		
total.sulfur.dioxide	1.0000000	1.0000000		
density	0.0000000	0.0000000		
pH	0.0000000	1.0000000		

sulphates	1.0000000	1.0000000
alcohol	1.0000000	1.0000000
BF	0.4033167	0.4334659
PostProbs	0.0451000	0.0421000
R2	0.5348000	0.5350000
dim	8.0000000	8.0000000
logmarg	-237.3640880	-237.2919970

For level 5 (category 5 with reference level 6):

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 169 models

	post mean	post SD	post p(B != 0)
Intercept	-8.14219	1.73035	1.00000
fixed.acidity	-0.53744	0.64231	0.69438
volatile.acidity	-0.04779	0.38350	0.42656
citric.acid	1.05143	0.70281	0.94722
residual.sugar	-0.84141	0.66956	0.92119
chlorides	-1.04793	0.97491	0.96411
free.sulfur.dioxide	0.16790	0.37331	0.50361
total.sulfur.dioxide	-1.23927	0.66381	0.99512
density	-0.09844	0.60212	0.48501
pH	-0.55161	0.54722	0.79897
sulphates	0.82627	0.36340	0.99741
alcohol	1.64879	0.68175	0.99995

> **summary(res3.5)**

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.0000000	1.0000000	1.00000	1.0000000
fixed.acidity	0.6943848	1.0000000	1.00000	1.0000000
volatile.acidity	0.4265625	0.0000000	0.00000	0.0000000
citric.acid	0.9472168	1.0000000	1.00000	1.0000000
residual.sugar	0.9211914	1.0000000	1.00000	1.0000000
chlorides	0.9641113	1.0000000	1.00000	1.0000000
free.sulfur.dioxide	0.5036133	0.0000000	1.00000	1.0000000
total.sulfur.dioxide	0.9951172	1.0000000	1.00000	1.0000000
density	0.4850098	0.0000000	0.00000	1.0000000
pH	0.7989746	1.0000000	1.00000	1.0000000
sulphates	0.9974121	1.0000000	1.00000	1.0000000
alcohol	0.9999512	1.0000000	1.00000	1.0000000
BF	NA	0.8888327	1.00000	0.6939117
PostProbs	NA	0.1108000	0.09930	0.0729000
R2	NA	0.7116000	0.71730	0.7178000
dim	NA	9.0000000	10.00000	11.0000000
logmarg	NA	-32.9653535	-32.84751	-33.2129179
	model 4	model 5		
Intercept	1.0000000	1.0000000		

fixed.acidity	1.0000000	1.0000000
volatile.acidity	1.0000000	0.0000000
citric.acid	1.0000000	1.0000000
residual.sugar	1.0000000	1.0000000
chlorides	1.0000000	1.0000000
free.sulfur.dioxide	1.0000000	0.0000000
total.sulfur.dioxide	1.0000000	1.0000000
density	0.0000000	1.0000000
pH	1.0000000	1.0000000
sulphates	1.0000000	1.0000000
alcohol	1.0000000	1.0000000
BF	0.7741582	0.6133596
PostProbs	0.0729000	0.0645000
R2	0.7180000	0.7119000
dim	11.0000000	10.0000000
logmarg	-33.1034863	-33.3363111

And therefore, the model will have the following form:

$$\ln\left(\frac{p_{i1}}{p_{i6}}\right) = -6.84 + 0.16 \cdot Z_{i1} + 0.88 \cdot Z_{i2} + 0.2 \cdot Z_{i5} + 0.5 \cdot Z_{i6} - 1.5 \cdot Z_{i7} + 0.28 \cdot Z_{i9} + \\ -0.54 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i2}}{p_{i6}}\right) = -2.88 + 0.98 \cdot Z_{i2} + 0.4 \cdot Z_{i4} - 0.89 \cdot Z_{i7} + 0.43 \cdot Z_{i9}$$

$$\ln\left(\frac{p_{i3}}{p_{i6}}\right) = 0.14 - 0.41 \cdot Z_{i2} - 0.07 \cdot Z_{i5} + 0.23 \cdot Z_{i6} - 0.58 \cdot Z_{i7} + 0.36 \cdot Z_{i10} + 0.8 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i4}}{p_{i6}}\right) = -1.96 + 0.21 \cdot Z_{i1} - 0.89 \cdot Z_{i2} + 0.11 \cdot Z_{i4} - 0.2 \cdot Z_{i5} + 0.26 \cdot Z_{i6} - \\ - 0.68 \cdot Z_{i7} - 0.14 \cdot Z_{i8} + 0.66 \cdot Z_{i10} + 1.62 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i5}}{p_{i6}}\right) = -8.14 - 0.54 \cdot Z_{i1} + 1.05 \cdot Z_{i3} - 0.84 \cdot Z_{i4} - 1.04 \cdot Z_{i5} + 0.17 \cdot Z_{i6} - \\ - 1.24 \cdot Z_{i7} - 0.55 \cdot Z_{i9} + 0.83 \cdot Z_{i10} + 1.65 \cdot Z_{i11}$$

Which is at least more parsimonious than the last one.

Variable Selection through (Liang et al.) hyper-g (using OpenBUGS):

Just like before, only this time in OpenBUGS. So, as always, we will select our most important variables by:

1. Using a Uniform Distribution on the priors:

```
Inference for Bugs model at "C:\Users\30697\Desktop\hyperg.txt", fit  
using OpenBUGS,  
 1 chains, each with 11000 iterations (first 1000 discarded), n.thin =  
30  
n.sims = 333 iterations saved  
      mean    sd   2.5%   25%   50%   75% 97.5%  
alpha[1] -6.4  0.8  -8.2  -7.0  -6.3  -5.8 -4.9  
alpha[2] -2.9  0.2  -3.3  -3.0  -2.9  -2.8 -2.5  
alpha[3]  0.1  0.1   0.0   0.1   0.1   0.2  0.2  
alpha[4] -2.0  0.1  -2.2  -2.1  -2.0  -1.8 -1.7  
alpha[5] -5.2  0.6  -6.3  -5.6  -5.2  -4.8 -4.2  
gb1[1]   0.2  0.5  -0.6   0.0   0.0   0.4  1.4  
gb1[2]   0.0  0.3  -0.5   0.0   0.0   0.0  0.9  
gb1[3]   0.0  0.1   0.0   0.0   0.0   0.0  0.3  
gb1[4]   0.3  0.2   0.0   0.0   0.3   0.5  0.8  
gb1[5]   0.0  0.3  -0.7   0.0   0.0   0.0  0.7  
gb2[1]   1.2  0.3   0.7   1.0   1.2   1.4  1.8  
gb2[2]   0.5  0.1   0.2   0.4   0.5   0.6  0.8  
gb2[3]  -0.5  0.1  -0.6  -0.5  -0.5  -0.4 -0.3  
gb2[4]  -0.8  0.1  -1.1  -0.9  -0.8  -0.7 -0.6  
gb2[5]  -0.2  0.3  -0.9  -0.4   0.0   0.0  0.1  
gb3[1]   0.1  0.3  -0.4   0.0   0.0   0.0  1.0  
gb3[2]   0.0  0.1  -0.2   0.0   0.0   0.0  0.5  
gb3[3]  -0.1  0.1  -0.4  -0.2  -0.1   0.0  0.0  
gb3[4]   0.0  0.1  -0.3   0.0   0.0   0.0  0.3  
gb3[5]   0.4  0.4  -0.1   0.0   0.2   0.7  1.2  
gb4[1]   0.0  0.2  -0.5   0.0   0.0   0.0  0.4  
gb4[2]   0.2  0.2   0.0   0.0   0.0   0.3  0.5  
gb4[3]   0.0  0.0   0.0   0.0   0.0   0.0  0.0  
gb4[4]   0.2  0.2   0.0   0.0   0.2   0.3  0.5  
gb4[5]   0.0  0.2  -0.6   0.0   0.0   0.0  0.3  
gb5[1]   0.3  0.3   0.0   0.0   0.2   0.5  0.8  
gb5[2]   0.0  0.1  -0.1   0.0   0.0   0.0  0.2  
gb5[3]  -0.1  0.1  -0.2  -0.1   0.0   0.0  0.0  
gb5[4]  -0.4  0.2  -0.7  -0.5  -0.4  -0.3  0.0  
gb5[5]  -0.9  0.6  -2.1  -1.3  -0.9  -0.5  0.0  
gb6[1]  -0.1  0.3  -0.9   0.0   0.0   0.0  0.8  
gb6[2]  -0.1  0.2  -0.7  -0.2   0.0   0.0  0.0  
gb6[3]   0.1  0.1   0.0   0.0   0.1   0.2  0.4  
gb6[4]   0.0  0.1   0.0   0.0   0.0   0.0  0.4  
gb6[5]   0.0  0.2  -0.6   0.0   0.0   0.0  0.4  
gb7[1]  -0.9  0.6  -2.1  -1.3  -0.9  -0.4  0.0  
gb7[2]  -0.5  0.3  -1.0  -0.7  -0.6  -0.4  0.0  
gb7[3]  -0.4  0.1  -0.7  -0.5  -0.4  -0.4  -0.3  
gb7[4]  -0.7  0.2  -1.0  -0.8  -0.7  -0.6  -0.4  
gb7[5]  -0.8  0.4  -1.4  -1.0  -0.8  -0.5  0.0
```

```

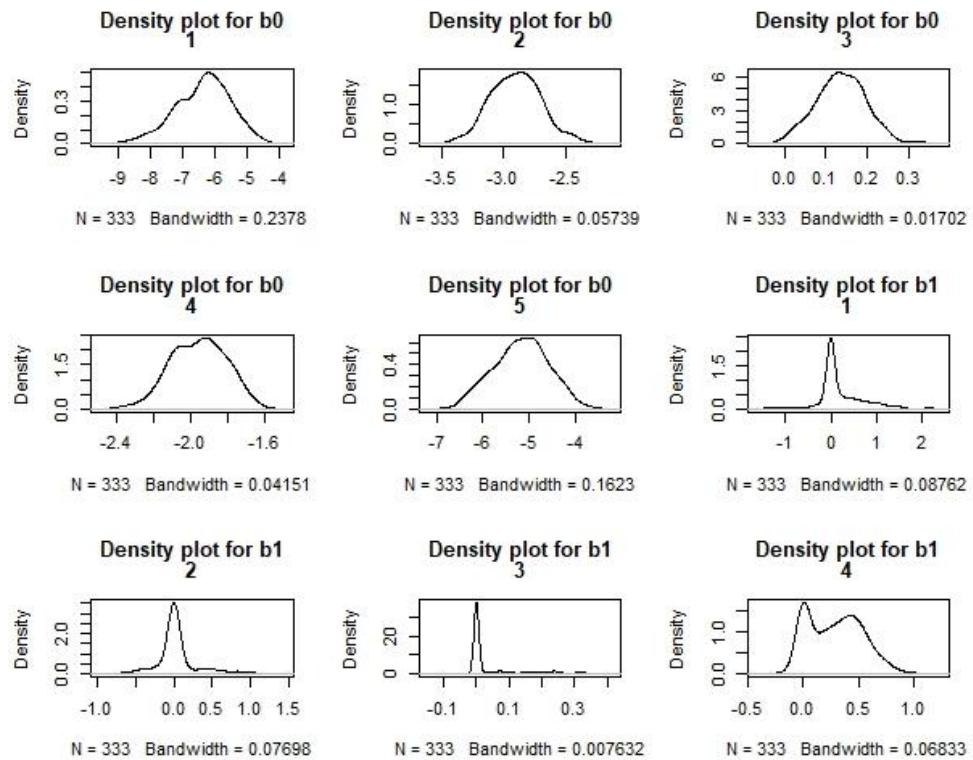
gb8[1]      0.4  0.5  -0.1   0.0   0.2   0.7  1.4
gb8[2]     -0.2  0.2  -0.8  -0.3   0.0   0.0  0.0
gb8[3]      0.0  0.0   0.0   0.0   0.0   0.0  0.2
gb8[4]     -0.2  0.2  -0.7  -0.4  -0.2   0.0  0.0
gb8[5]     -0.3  0.4  -1.2  -0.5   0.0   0.0  0.1
gb9[1]      0.4  0.5  -0.1   0.0   0.2   0.8  1.6
gb9[2]      0.2  0.3   0.0   0.0   0.2   0.4  0.8
gb9[3]      0.0  0.0  -0.1   0.0   0.0   0.0  0.0
gb9[4]      0.0  0.1  -0.2   0.0   0.0   0.0  0.2
gb9[5]     -0.4  0.4  -1.1  -0.7  -0.3   0.0  0.0
gb10[1]    -0.1  0.4  -1.2  -0.2   0.0   0.0  0.3
gb10[2]     0.0  0.1  -0.2   0.0   0.0   0.0  0.2
gb10[3]     0.3  0.1   0.2   0.3   0.3   0.4  0.5
gb10[4]     0.8  0.1   0.6   0.7   0.8   0.9  1.0
gb10[5]     0.8  0.3   0.3   0.7   0.8   1.0  1.4
gb11[1]    -0.5  0.6  -1.9  -0.9  -0.4   0.0  0.1
gb11[2]     0.1  0.2   0.0   0.0   0.0   0.3  0.7
gb11[3]     0.8  0.1   0.7   0.8   0.8   0.9  1.0
gb11[4]     1.5  0.2   1.1   1.4   1.5   1.6  1.8
gb11[5]     1.9  0.3   1.2   1.7   1.9   2.1  2.5
deviance 3002.1 12.5 2979.2 2993.4 3001.8 3010.6 3026.4

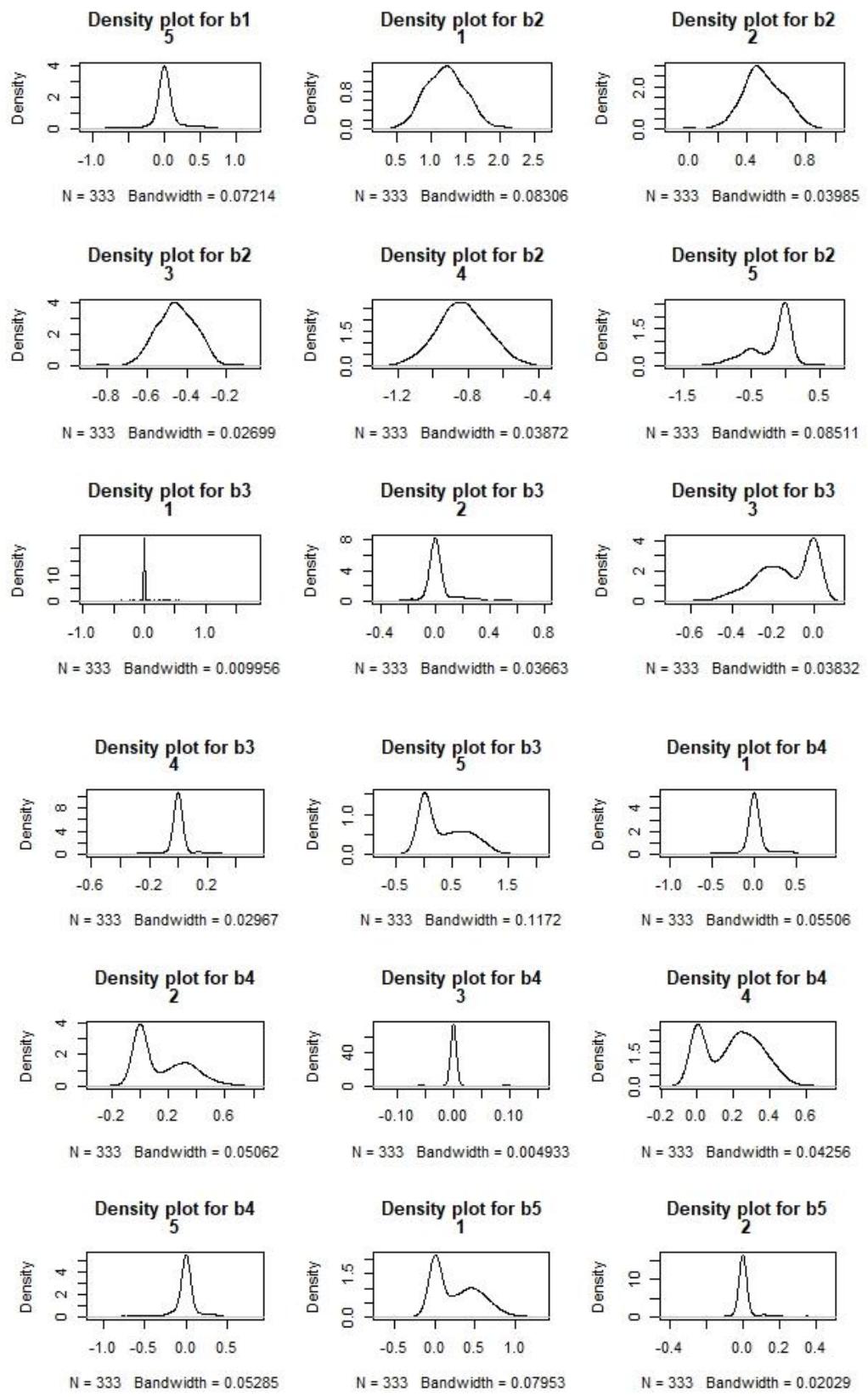
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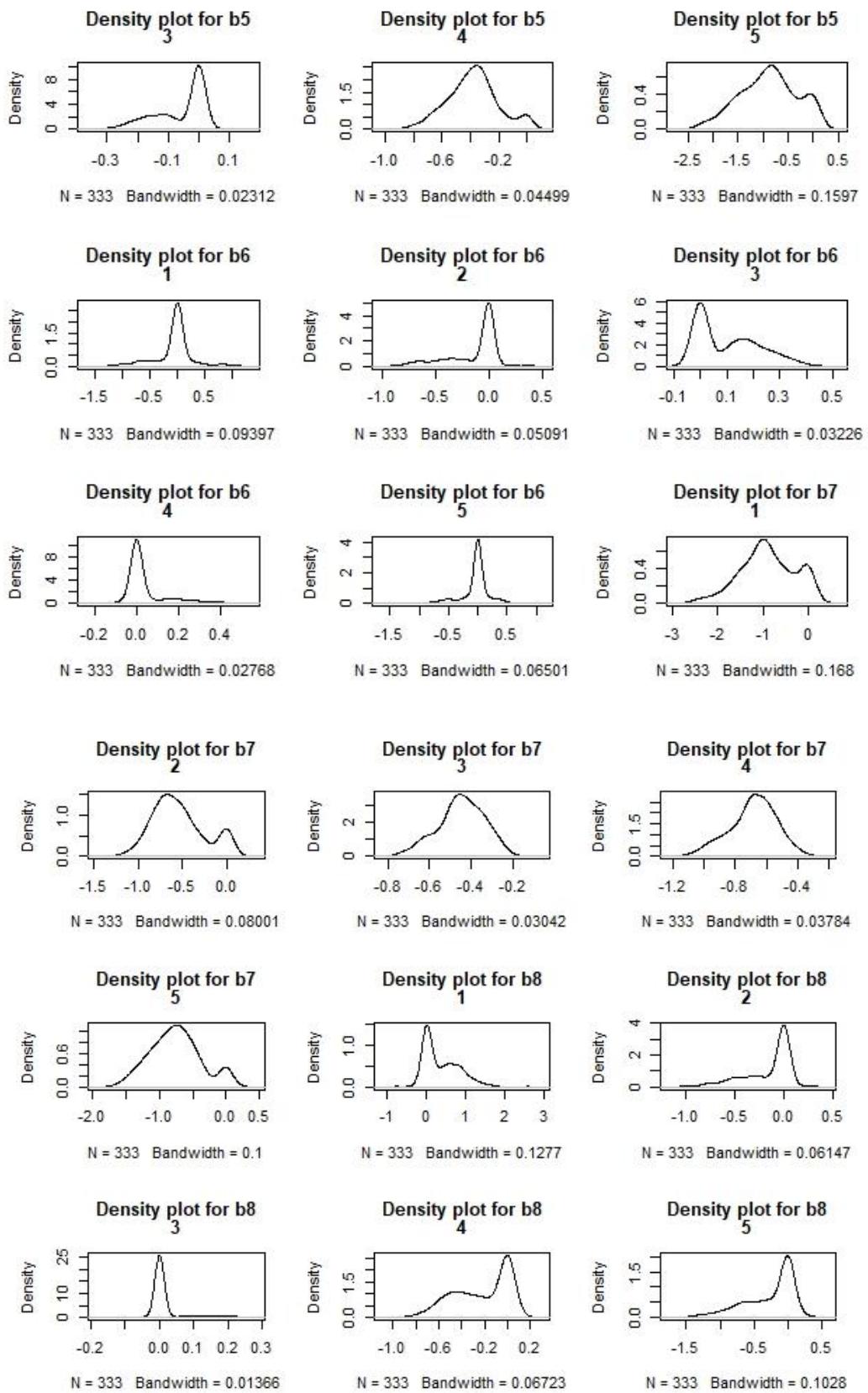
DIC info (using the rule, $pD = \text{var}(\text{deviance})/2$)

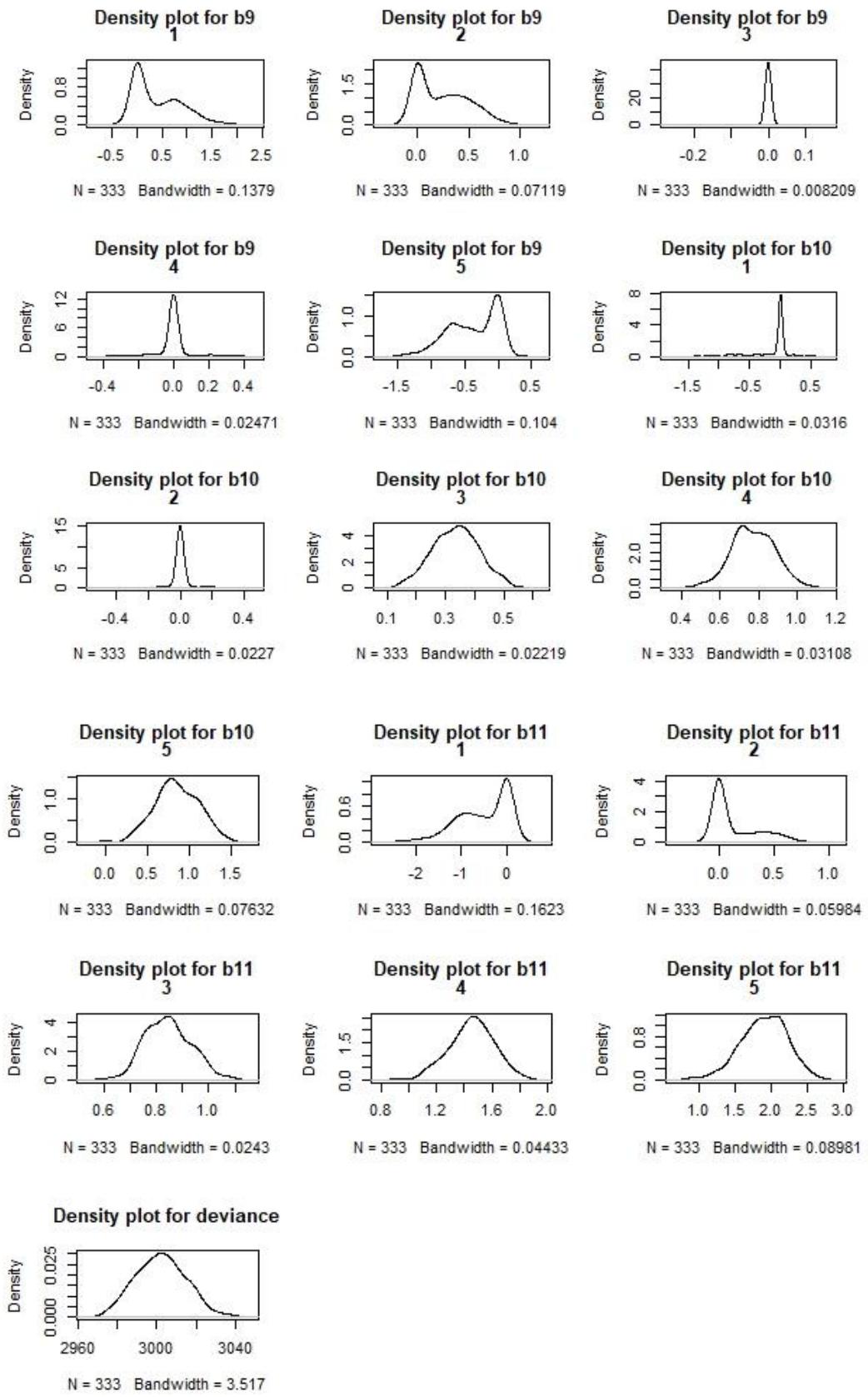
$pD = 77.9$ and DIC = 3080.0

DIC is an estimate of expected predictive error (lower deviance is better).









And so, this model will look like this:

$$\ln\left(\frac{p_{i1}}{p_{i6}}\right) = -6.4 + 0.2 \cdot Z_{i1} + 1.2 \cdot Z_{i2} + 0.1 \cdot Z_{i3} + 0.3 \cdot Z_{i5} - 0.1 \cdot Z_{i6} - 0.9 \cdot Z_{i7} + \\ + 0.4 \cdot Z_{i8} + 0.4 \cdot Z_{i9} - 0.1 \cdot Z_{i10} - 1.9 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i2}}{p_{i6}}\right) = -2.9 + 0.5 \cdot Z_{i2} + 0.2 \cdot Z_{i4} - 0.1 \cdot Z_{i6} - 0.5 \cdot Z_{i7} - 0.2 \cdot Z_{i8} + 0.2 \cdot Z_{i9} - 0.1 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i3}}{p_{i6}}\right) = 0.1 - 0.5 \cdot Z_{i2} - 0.1 \cdot Z_{i3} - 0.1 \cdot Z_{i5} + 0.1 \cdot Z_{i6} - 0.4 \cdot Z_{i7} + 0.3 \cdot Z_{i10} + 0.8 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i4}}{p_{i6}}\right) = -2 + 0.3 \cdot Z_{i1} - 0.8 \cdot Z_{i2} + 0.2 \cdot Z_{i4} - 0.4 \cdot Z_{i5} - 0.7 \cdot Z_{i7} - 0.2 \cdot Z_{i8} + \\ + 0.8 \cdot Z_{i10} + 1.5 \cdot Z_{i11}$$

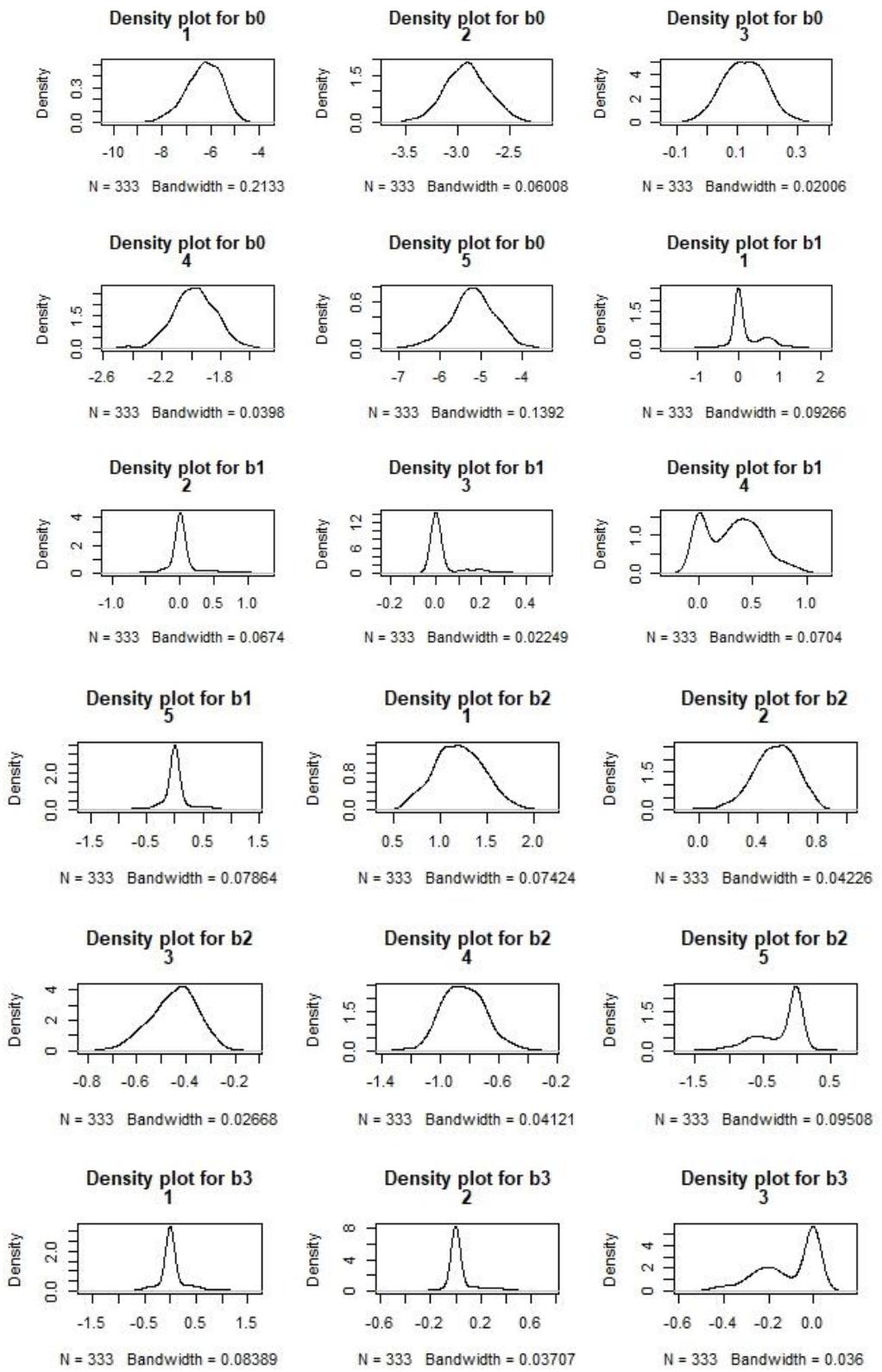
$$\ln\left(\frac{p_{i5}}{p_{i6}}\right) = -5.2 - 0.2 \cdot Z_{i2} + 0.4 \cdot Z_{i3} - 0.9 \cdot Z_{i5} - 0.8 \cdot Z_{i7} - 0.3 \cdot Z_{i8} - 0.4 \cdot Z_{i9} + \\ + 0.8 \cdot Z_{i10} + 1.9 \cdot Z_{i11}$$

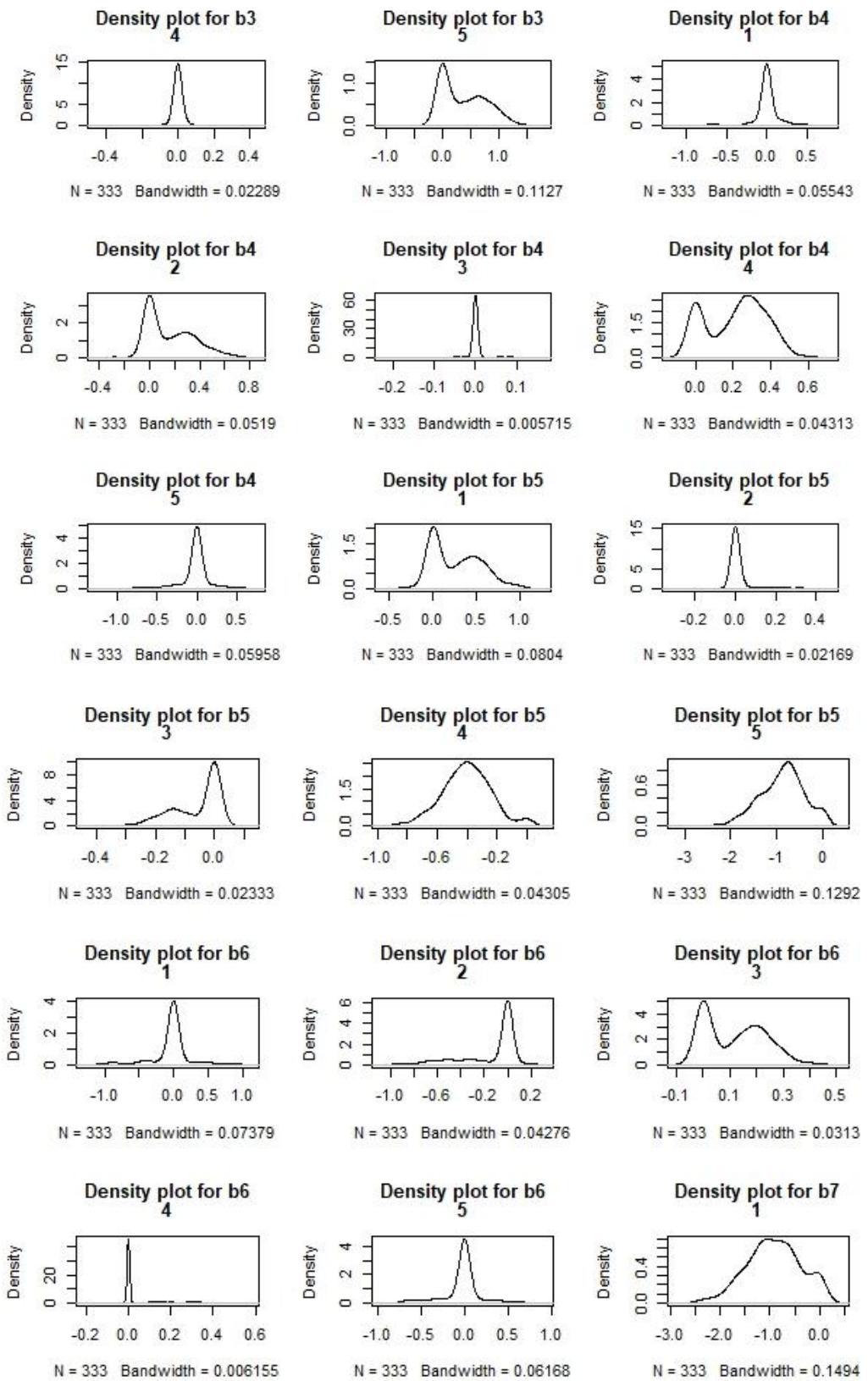
2. Using a Beta-Binomial Distribution on the priors:

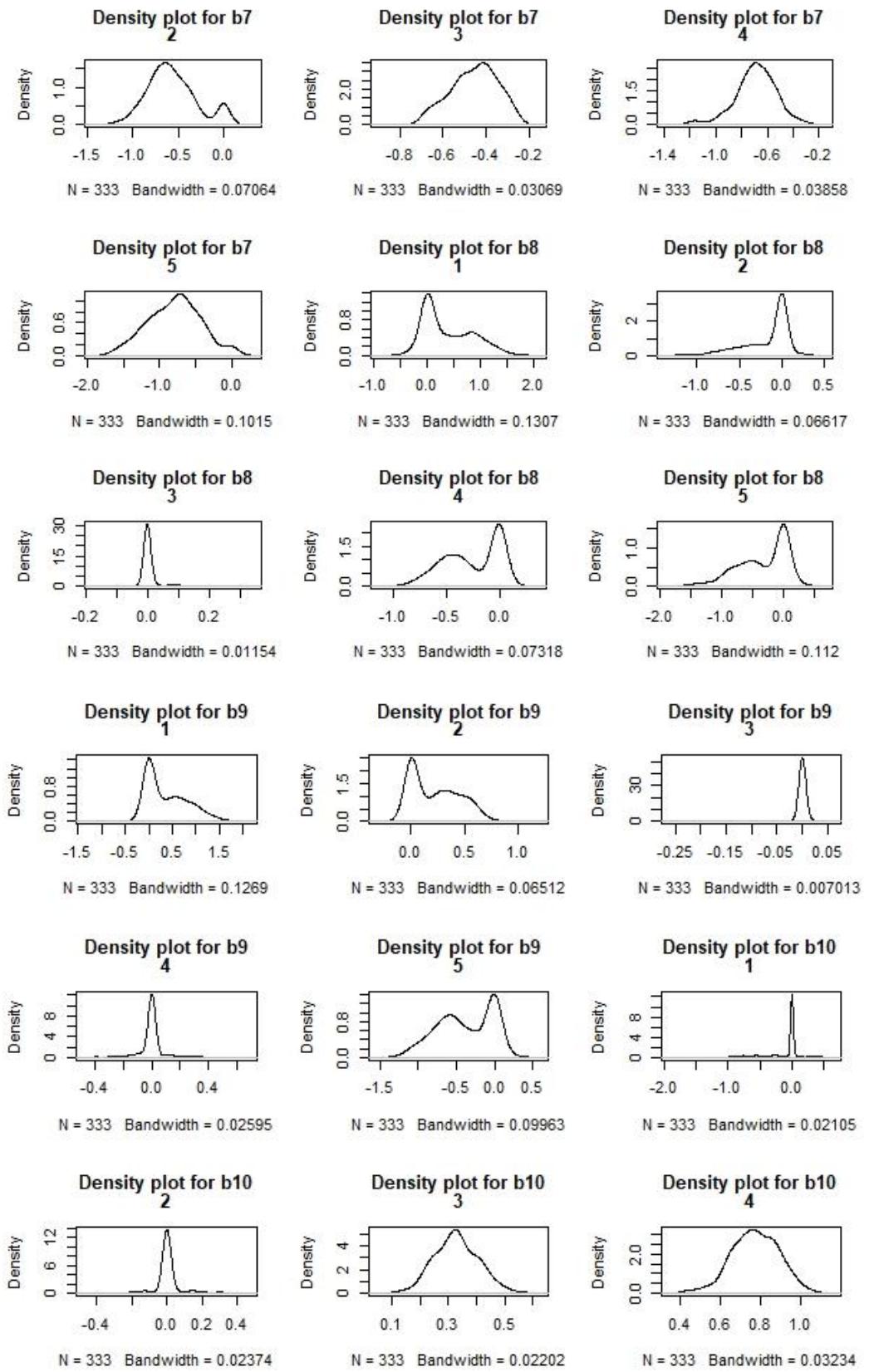
```
Inference for Bugs model at "C:\Users\30697\Desktop\hyperg.txt", fit
using OpenBUGS,
 1 chains, each with 11000 iterations (first 1000 discarded), n.thin =
30
n.sims = 333 iterations saved
      mean    sd   2.5%   25%   50%   75% 97.5%
alpha[1] -6.3  0.8  -7.9  -6.8  -6.2  -5.7 -5.0
alpha[2] -2.9  0.2  -3.3  -3.1  -2.9  -2.8 -2.5
alpha[3]  0.1  0.1   0.0   0.1   0.1   0.2  0.3
alpha[4] -2.0  0.1  -2.3  -2.1  -2.0  -1.9 -1.7
alpha[5] -5.2  0.5  -6.3  -5.5  -5.2  -4.9 -4.2
gb1[1]   0.2  0.4  -0.4   0.0   0.0   0.4  1.2
gb1[2]   0.0  0.2  -0.3   0.0   0.0   0.0  0.8
gb1[3]   0.0  0.1   0.0   0.0   0.0   0.0  0.3
gb1[4]   0.3  0.2   0.0   0.0   0.3   0.5  0.8
gb1[5]   0.0  0.3  -0.6   0.0   0.0   0.0  0.7
gb2[1]   1.2  0.3   0.7   1.0   1.2   1.4  1.7
gb2[2]   0.5  0.2   0.2   0.4   0.5   0.6  0.8
gb2[3]  -0.4  0.1  -0.6  -0.5  -0.4  -0.4 -0.3
gb2[4]  -0.8  0.1  -1.1  -0.9  -0.8  -0.7 -0.5
gb2[5]  -0.2  0.3  -1.1  -0.5   0.0   0.0  0.1
gb3[1]   0.0  0.3  -0.5   0.0   0.0   0.0  0.9
gb3[2]   0.0  0.1  -0.1   0.0   0.0   0.0  0.4
gb3[3]  -0.1  0.1  -0.4  -0.2   0.0   0.0  0.0
```

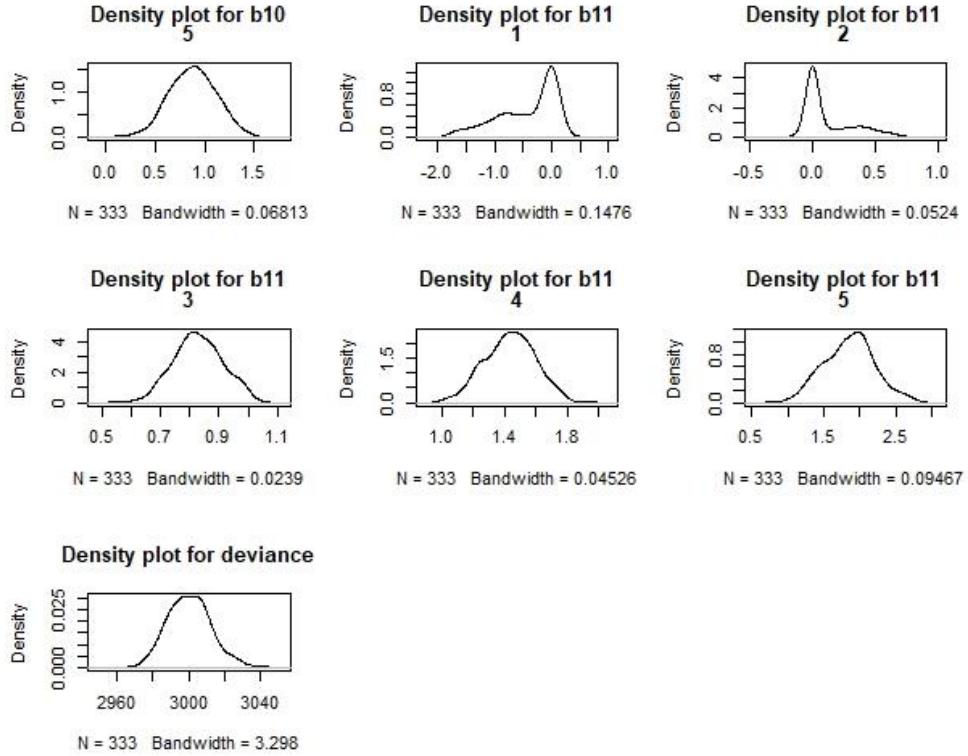
gb3[4]	0.0	0.1	-0.2	0.0	0.0	0.0	0.3
gb3[5]	0.4	0.4	0.0	0.0	0.3	0.7	1.2
gb4[1]	0.0	0.2	-0.7	0.0	0.0	0.0	0.4
gb4[2]	0.2	0.2	0.0	0.0	0.1	0.3	0.6
gb4[3]	0.0	0.0	0.0	0.0	0.0	0.0	0.0
gb4[4]	0.2	0.2	0.0	0.0	0.2	0.3	0.5
gb4[5]	0.0	0.2	-0.7	0.0	0.0	0.0	0.4
gb5[1]	0.3	0.3	0.0	0.0	0.2	0.5	0.9
gb5[2]	0.0	0.1	-0.1	0.0	0.0	0.0	0.3
gb5[3]	-0.1	0.1	-0.2	-0.1	0.0	0.0	0.0
gb5[4]	-0.4	0.2	-0.7	-0.5	-0.4	-0.3	0.0
gb5[5]	-0.9	0.5	-1.9	-1.2	-0.8	-0.6	0.0
gb6[1]	0.0	0.3	-0.8	0.0	0.0	0.0	0.5
gb6[2]	-0.1	0.2	-0.7	-0.2	0.0	0.0	0.1
gb6[3]	0.1	0.1	0.0	0.0	0.1	0.2	0.3
gb6[4]	0.1	0.1	0.0	0.0	0.0	0.0	0.4
gb6[5]	0.0	0.2	-0.6	0.0	0.0	0.0	0.5
gb7[1]	-0.9	0.5	-2.0	-1.3	-0.9	-0.6	0.0
gb7[2]	-0.6	0.3	-1.0	-0.7	-0.6	-0.4	0.0
gb7[3]	-0.5	0.1	-0.7	-0.5	-0.4	-0.4	-0.3
gb7[4]	-0.7	0.2	-1.1	-0.8	-0.7	-0.6	-0.4
gb7[5]	-0.8	0.4	-1.5	-1.0	-0.8	-0.5	0.0
gb8[1]	0.4	0.5	-0.2	0.0	0.2	0.8	1.3
gb8[2]	-0.2	0.3	-0.8	-0.3	0.0	0.0	0.1
gb8[3]	0.0	0.0	0.0	0.0	0.0	0.0	0.1
gb8[4]	-0.3	0.3	-0.8	-0.5	-0.2	0.0	0.0
gb8[5]	-0.3	0.4	-1.2	-0.6	-0.2	0.0	0.2
gb9[1]	0.4	0.5	-0.1	0.0	0.2	0.7	1.4
gb9[2]	0.2	0.2	0.0	0.0	0.2	0.4	0.7
gb9[3]	0.0	0.0	-0.1	0.0	0.0	0.0	0.0
gb9[4]	0.0	0.1	-0.2	0.0	0.0	0.0	0.2
gb9[5]	-0.4	0.4	-1.1	-0.7	-0.4	0.0	0.0
gb10[1]	-0.1	0.3	-0.9	-0.1	0.0	0.0	0.3
gb10[2]	0.0	0.1	-0.2	0.0	0.0	0.0	0.2
gb10[3]	0.3	0.1	0.2	0.3	0.3	0.4	0.5
gb10[4]	0.8	0.1	0.5	0.7	0.8	0.9	1.0
gb10[5]	0.9	0.2	0.4	0.7	0.9	1.0	1.3
gb11[1]	-0.4	0.5	-1.6	-0.8	-0.2	0.0	0.1
gb11[2]	0.1	0.2	-0.1	0.0	0.0	0.2	0.6
gb11[3]	0.8	0.1	0.7	0.8	0.8	0.9	1.0
gb11[4]	1.4	0.2	1.1	1.3	1.4	1.5	1.7
gb11[5]	1.8	0.4	1.2	1.6	1.9	2.1	2.6
deviance	3000.3	12.3	2978.0	2991.9	2999.6	3007.6	3025.5

DIC info (using the rule, pD = var(deviance)/2)
pD = 76.2 and DIC = 3076.6
DIC is an estimate of expected predictive error (lower deviance is better).









And so, this model will be something like this:

$$\ln\left(\frac{p_{i1}}{p_{i6}}\right) = -6.3 + 0.2 \cdot Z_{i1} + 1.2 \cdot Z_{i2} + 0.3 \cdot Z_{i5} - 0.9 \cdot Z_{i7} + 0.4 \cdot Z_{i8} + 0.4 \cdot Z_{i9} - 0.1 \cdot Z_{i10} - 0.4 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i2}}{p_{i6}}\right) = -2.9 + 0.5 \cdot Z_{i2} + 0.2 \cdot Z_{i4} - 0.1 \cdot Z_{i6} - 0.6 \cdot Z_{i7} - 0.2 \cdot Z_{i8} + 0.2 \cdot Z_{i9} + 0.1 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i3}}{p_{i6}}\right) = 0.1 - 0.4 \cdot Z_{i2} - 0.1 \cdot Z_{i3} - 0.1 \cdot Z_{i5} + 0.1 \cdot Z_{i6} - 0.5 \cdot Z_{i7} + 0.3 \cdot Z_{i10} + 0.8 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i4}}{p_{i6}}\right) = -2 + 0.3 \cdot Z_{i1} - 0.8 \cdot Z_{i2} + 0.2 \cdot Z_{i4} - 0.4 \cdot Z_{i5} - 0.7 \cdot Z_{i7} - 0.3 \cdot Z_{i8} + 0.8 \cdot Z_{i10} + 1.4 \cdot Z_{i11}$$

$$\ln\left(\frac{p_{i5}}{p_{i6}}\right) = -5.2 - 0.2 \cdot Z_{i2} + 0.4 \cdot Z_{i3} - 0.9 \cdot Z_{i5} - 0.8 \cdot Z_{i7} - 0.3 \cdot Z_{i8} - 0.4 \cdot Z_{i9} + 0.9 \cdot Z_{i10} + 1.8 \cdot Z_{i11}$$

Conclusion (of Variable Selection):

We saw different methods and their results, as far as variable selection is concerned. The idea is simple, and their implementation differs in the form of the prior distribution that we choose each time for the parameters β .

What we would like to do, is to pick only the most necessary variables for our final model. While other methods (such as g-prior or BIC prior) are worth considering, they get rid of too many variables to the point of making our new model redundant when compared to the original. The table below is a synopsis of their Deviance Information Criterion values:

MODEL	DIC VALUE
<i>Original Model</i>	3036.0
<i>Reduced Model (Variable Selection through BIC prior on betas with beta-binomial on the model)</i>	3063.0
<i>Reduced Model (Variable Selection through g-prior (of Liang et al.) prior on betas with beta-binomial on the model)</i>	3058.0

And thus, we conclude that none of these models are better than the original. Therefore, we shall choose the *hyper-g prior* on our β s, which utilizes the *beta-binomial* for the model so as to at least have a modicum of parsimony in the final result. We will be using the following variables on each level:

1. *Level 1 (Category 1 with reference level of 6):* used the intercept (β_0), fixed acidity (X_1), volatile acidity (X_2), chlorides (X_5), free sulfur dioxide (X_6), total sulfur dioxide (X_7), pH (X_9) and alcohol (X_{11}).
2. *Level 2 (Category 2 with reference level of 6):* used the intercept (β_0), volatile acidity (X_2), residual sugar (X_4), total sulfur dioxide (X_7) and pH (X_9).
3. *Level 3 (Category 3 with reference level of 6):* used the intercept (β_0), volatile acidity (X_2), chlorides (X_5), free sulfur dioxide (X_6), total sulfur dioxide (X_7), sulphates (X_{10}) and alcohol (X_{11}).
4. *Level 4 (Category 4 with reference level of 6):* used the intercept (β_0), fixed acidity (X_1), volatile acidity (X_2), residual sugar (X_4), chlorides (X_5), free sulfur dioxide (X_6), total sulfur dioxide (X_7), density (X_8), sulphates (X_{10}) and alcohol (X_{11}).
5. *Level 5 (Category 5 with reference level of 6):* used the intercept (β_0), fixed acidity (X_1), citric acid (X_3), residual sugar (X_4), chlorides (X_5), free sulfur dioxide (X_6), total sulfur dioxide (X_7), pH (X_9), sulphates (X_{10}) and alcohol (X_{11}).

It looks like all the physiochemical characteristics are used in this model, although not on all levels. Time to run our model one more time, without all the superfluous variables that were included in the original.

Posterior analysis and interpretation of the results:

Should we run this much simpler model, which contains only the statistically significant variables for each level, we would get the following convergence diagnostics and graphs:

```
GEWEKE CONVERGENCE DIAGNOSTIC (Z-score)
=====
Iterations used = 1001:50951
Thinning interval = 50
Sample size per chain = 1000

$chain1

Fraction in 1st window = 0.1
Fraction in 2nd window = 0.5

      b01      B01      b02      B02      b03      B03      b04      B04
-0.20476 -0.26875  1.91301  1.59428 -0.15938 -0.18603  0.56523  0.50987
      b05      B05      b103     B103     b104     B104     b105     B105
  0.06036  0.27307  1.65632  1.61372  0.77819  0.76217  0.85338  1.05632
      b11      B11      b111     B111     b113     B113     b114     B114
  0.39281  0.03356  0.16221  0.27720 -1.30230 -1.20170 -0.65869 -0.61384
      b115     B115     b14      B14      b15      B15      b21      B21
-0.95878 -0.71558  0.77899  0.66690  0.60275  0.20317  0.21143 -0.00659
      b22      B22      b23      B23      b24      B24      b35      B35
-1.80271 -2.03212  0.45908  0.46769  0.77868  0.61238 -0.21261 -0.33850
      b42      B42      b44      B44      b45      B45      b51      B51
  0.25064 -0.00200  1.30299  1.12780 -0.05491  0.03058 -0.21426 -0.37251
      b53      B53      b54      B54      b55      B55      b61      B61
-0.42969 -0.50951 -0.06167 -0.09627 -0.12096  0.05724  0.32938  0.82647
      b63      B63      b64      B64      b65      B65      b71      B71
  1.61524  1.66812  0.61970  0.70173  0.67484  1.20563 -0.25496  0.14864
      b72      B72      b73      B73      b74      B74      b75      B75
  0.48214  0.30838 -0.08973 -0.00715 -0.89042 -0.85021 -0.26975  0.15530
      b84      B84      b91      B91      b92      B92      b95      B95
-0.20844 -0.18607 -1.15582 -1.55224  0.64378  0.44617  0.33552  0.51150
deviance
-1.10274
```

HEIDELBERGER AND WELCH STATIONARITY AND INTERVAL HALFWIDTH TESTS
=====

```
Iterations used = 1001:50951
```

```

Thinning interval = 50
Sample size per chain = 1000

Precision of halfwidth test = 0.1

$chain1

      Stationarity start      p-value
      test          iteration
b01    passed        1      0.7111
B01    passed        1      0.7551
b02    passed        1      0.2048
B02    passed        1      0.2323
b03    passed        1      0.8287
B03    passed        1      0.8587
b04    passed        1      0.3420
B04    passed        1      0.4746
b05    passed        1      0.9057
B05    passed        1      0.7010
b103   passed        1      0.5418
B103   passed        1      0.5634
b104   passed        1      0.7611
B104   passed        1      0.7623
b105   passed        1      0.5258
B105   passed        1      0.4630
b11    passed        1      0.9614
B11    passed        1      0.9472
b111   passed        1      0.4913
B111   passed        1      0.3787
b113   passed        1      0.7410
B113   passed        1      0.7564
b114   passed        1      0.9030
B114   passed        1      0.9311
b115   passed        1      0.9153
B115   passed        1      0.9494
b14    passed        1      0.7338
B14    passed        1      0.6890
b15    passed        1      0.9916
B15    passed        1      0.9115
b21    passed        1      0.3615
B21    passed        1      0.3369
b22    passed        1      0.1210
B22    passed        1      0.0957
b23    passed        1      0.9838
B23    passed        1      0.9843
b24    passed        1      0.8962
B24    passed        1      0.9292
b35    passed        1      0.8978
B35    passed        1      0.9563
b42    passed        1      0.5880
B42    passed        1      0.5401
b44    passed        1      0.6320
B44    passed        1      0.6705

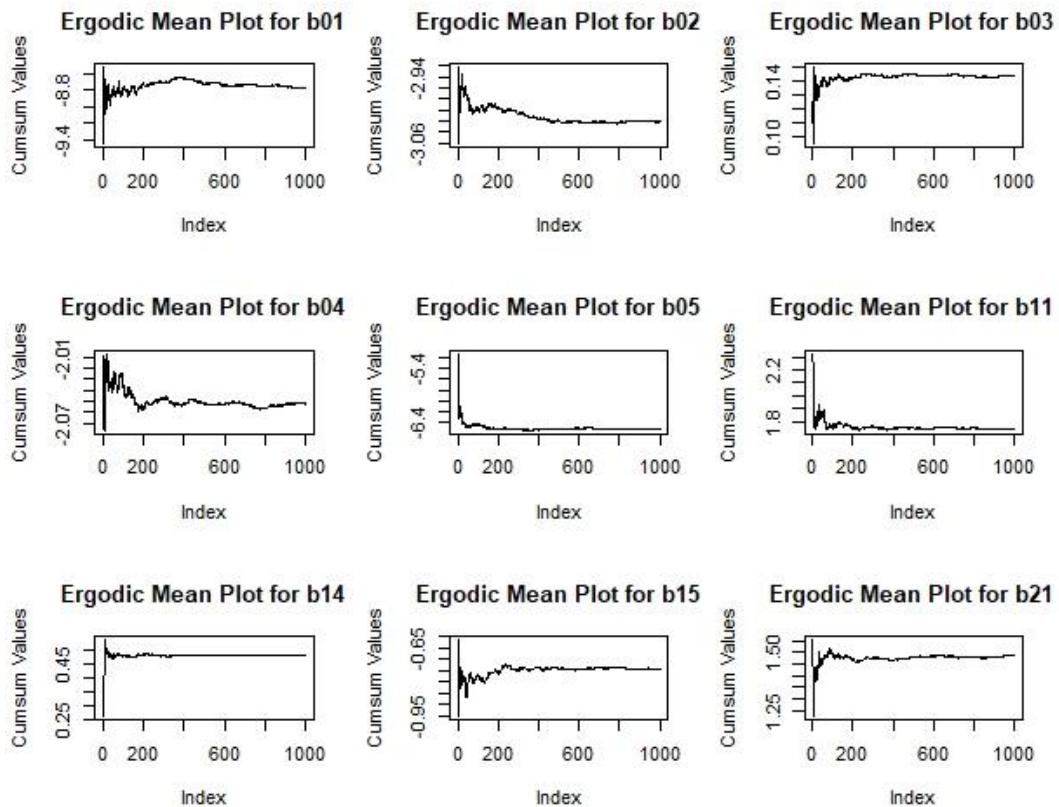
```

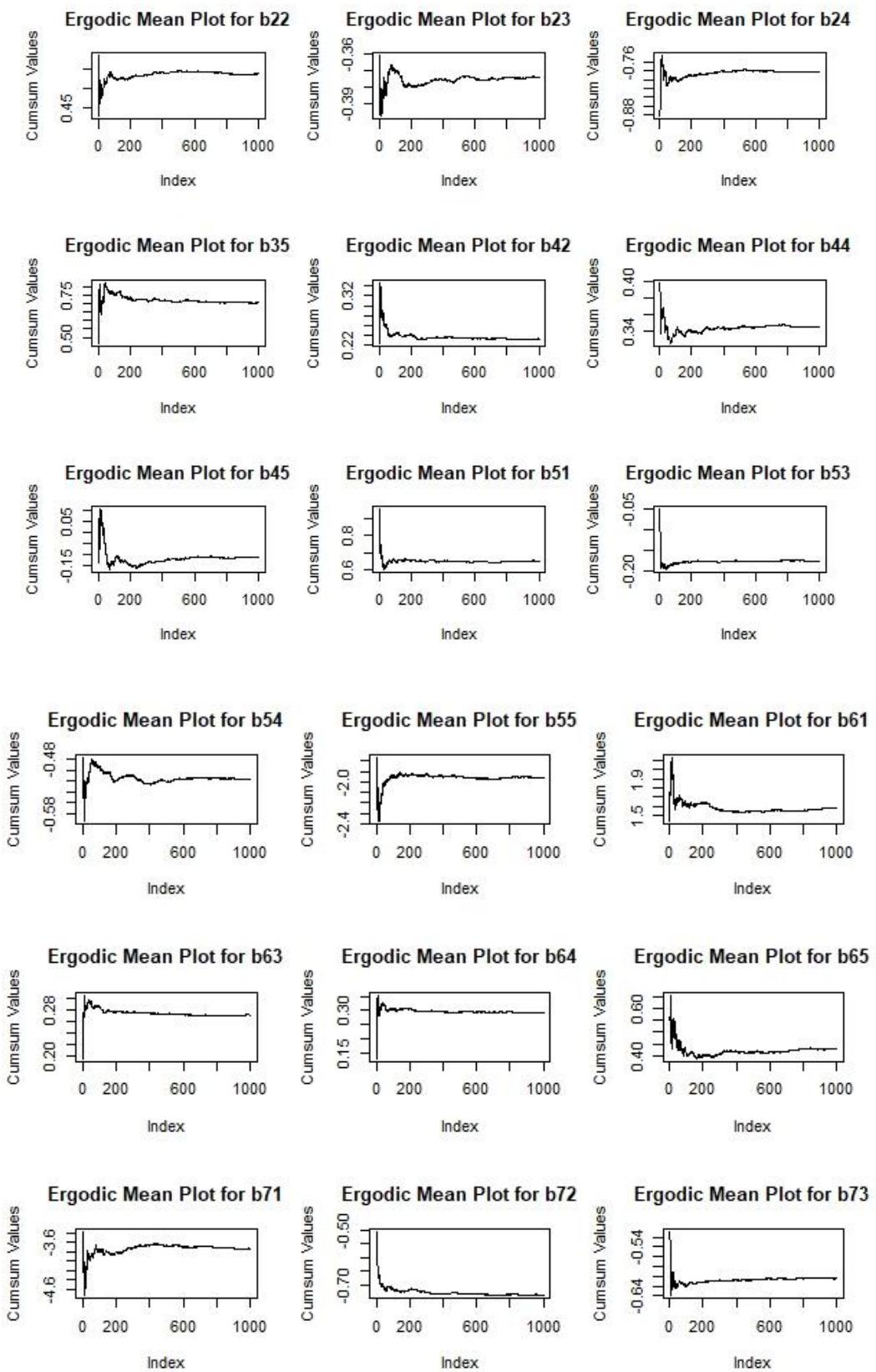
b45	passed	1	0.1784
B45	passed	1	0.1814
b51	passed	1	0.1671
B51	passed	1	0.2293
b53	passed	1	0.9044
B53	passed	1	0.8696
b54	passed	1	0.7560
B54	passed	1	0.7141
b55	passed	1	0.6099
B55	passed	1	0.6156
b61	passed	1	0.4318
B61	passed	1	0.0534
b63	passed	1	0.3285
B63	passed	1	0.3292
b64	passed	1	0.5903
B64	passed	1	0.5913
b65	passed	1	0.8239
B65	passed	1	0.8248
b71	passed	1	0.2836
B71	passed	1	0.8461
b72	passed	1	0.1399
B72	passed	1	0.3212
b73	passed	1	0.2287
B73	passed	1	0.1963
b74	passed	1	0.4157
B74	passed	1	0.4185
b75	passed	1	0.7521
B75	passed	1	0.8142
b84	passed	1	0.9514
B84	passed	1	0.9771
b91	passed	1	0.3171
B91	passed	1	0.2509
b92	passed	1	0.7585
B92	passed	1	0.7838
b95	passed	1	0.8984
B95	passed	1	0.9168
deviance	passed	201	0.2551

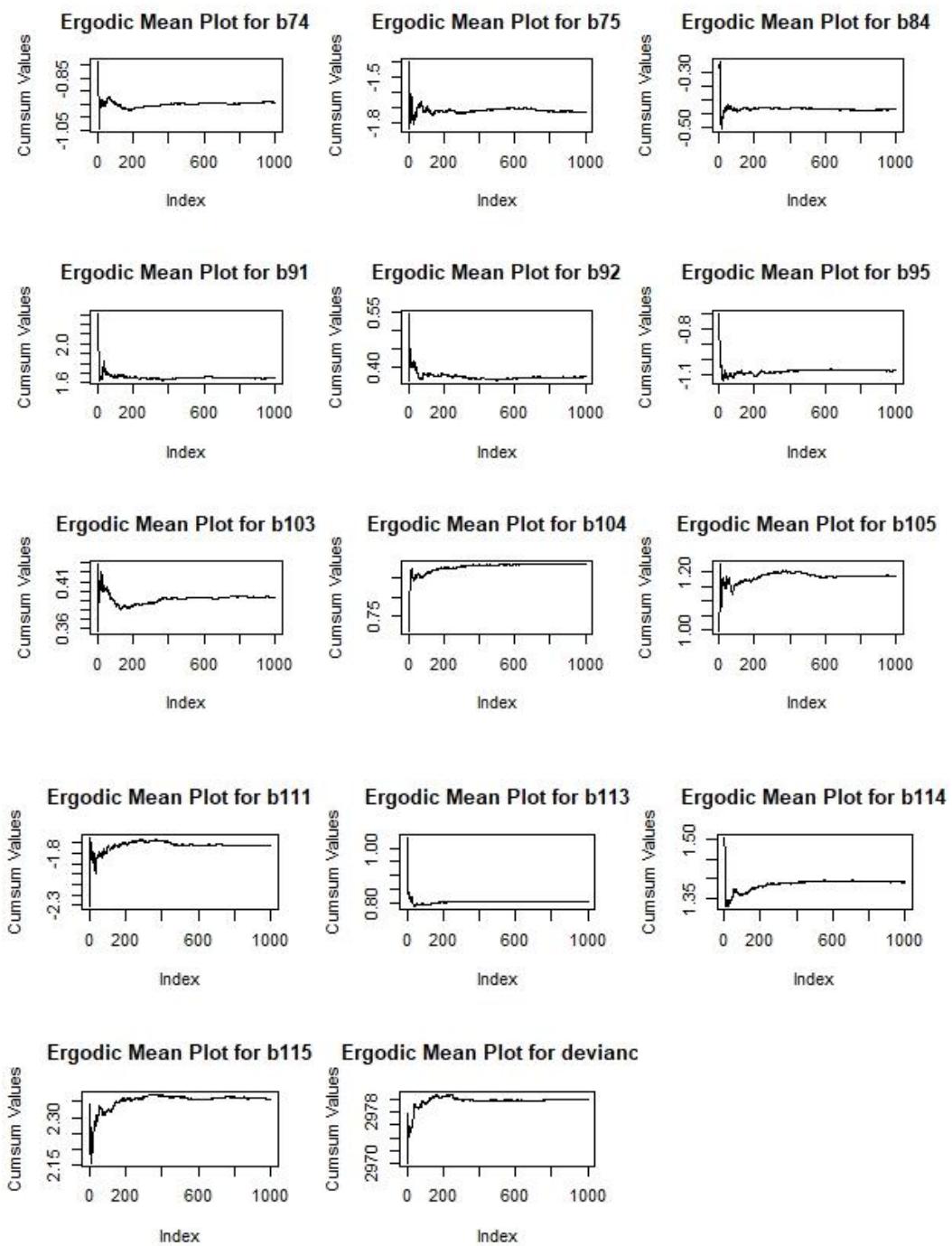
		Halfwidth	Mean	Halfwidth
	test			
b01	passed	-8.77e+00	1.88e-01	
B01	failed	3.33e-04	4.79e-05	
b02	passed	-3.02e+00	1.25e-02	
B02	passed	4.97e-02	6.18e-04	
b03	passed	1.43e-01	3.92e-03	
B03	passed	1.16e+00	4.55e-03	
b04	passed	-2.05e+00	1.01e-02	
B04	passed	1.30e-01	1.32e-03	
b05	passed	-6.50e+00	6.65e-02	
B05	passed	1.94e-03	1.13e-04	
b103	passed	3.94e-01	4.76e-03	
B103	passed	1.49e+00	7.15e-03	
b104	passed	8.86e-01	7.02e-03	

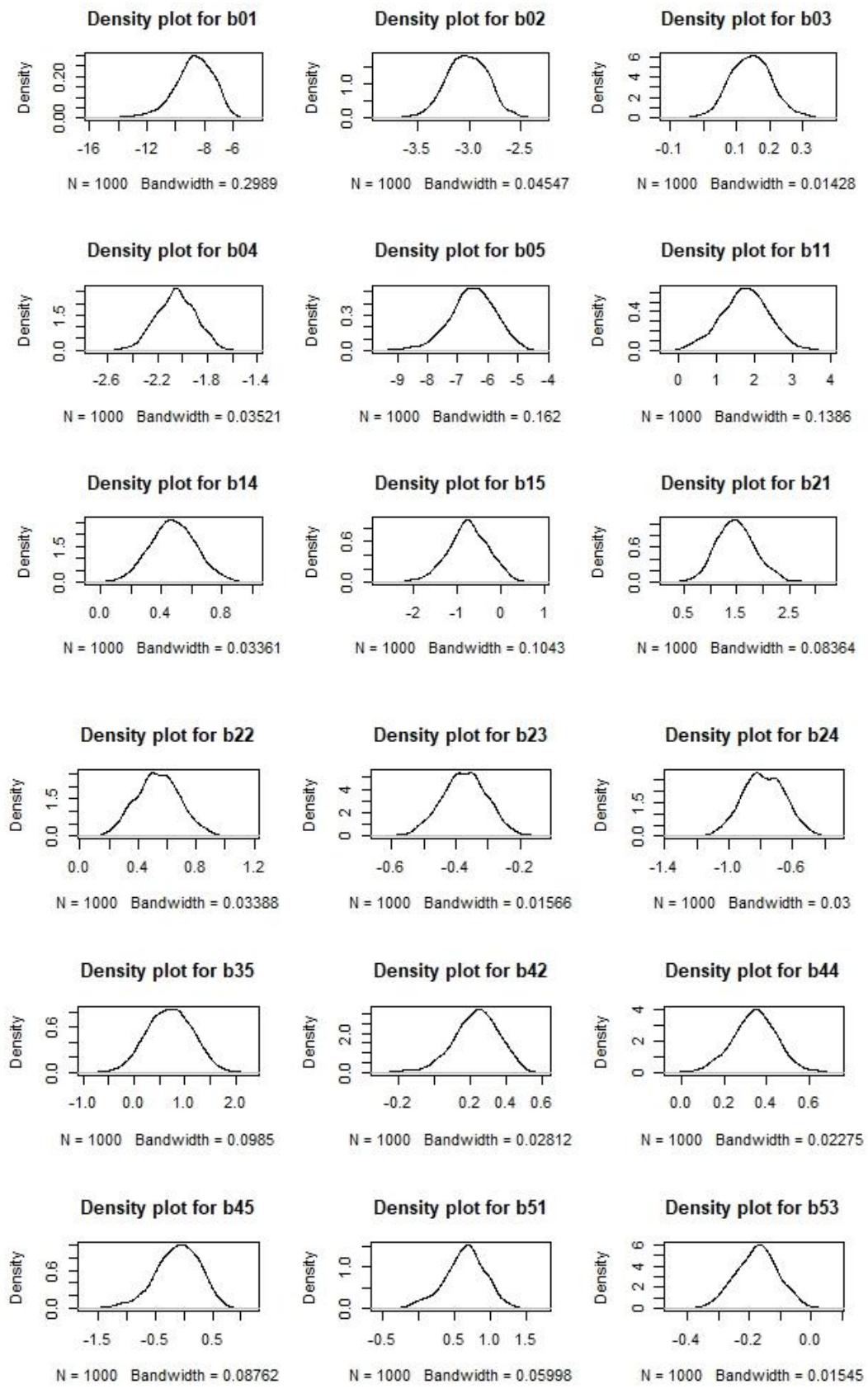
B104	passed	2.44e+00	1.71e-02
b105	passed	1.19e+00	1.91e-02
B105	passed	3.39e+00	6.90e-02
b11	passed	1.75e+00	4.21e-02
B11	passed	6.91e+00	3.02e-01
b111	passed	-1.72e+00	7.76e-02
B111	passed	2.34e-01	1.49e-02
b113	passed	8.04e-01	5.02e-03
B113	passed	2.24e+00	1.13e-02
b114	passed	1.39e+00	9.49e-03
B114	passed	4.06e+00	3.90e-02
b115	passed	2.36e+00	2.67e-02
B115	passed	1.11e+01	3.09e-01
b14	passed	4.78e-01	1.09e-02
B14	passed	1.63e+00	1.78e-02
b15	passed	-7.43e-01	3.46e-02
B15	passed	5.28e-01	1.67e-02
b21	passed	1.49e+00	3.48e-02
B21	passed	4.76e+00	1.83e-01
b22	passed	5.39e-01	1.27e-02
B22	passed	1.73e+00	2.22e-02
b23	passed	-3.74e-01	4.38e-03
B23	passed	6.90e-01	3.02e-03
b24	passed	-7.83e-01	8.23e-03
B24	passed	4.61e-01	3.78e-03
b35	passed	7.04e-01	2.70e-02
B35	passed	2.22e+00	6.17e-02
b42	passed	2.33e-01	7.79e-03
B42	passed	1.27e+00	9.68e-03
b44	passed	3.45e-01	7.33e-03
B44	passed	1.42e+00	1.04e-02
b45	failed	-1.12e-01	2.40e-02
B45	passed	9.60e-01	2.19e-02
b51	passed	6.48e-01	1.90e-02
B51	passed	1.99e+00	3.68e-02
b53	passed	-1.74e-01	4.24e-03
B53	passed	8.42e-01	3.56e-03
b54	passed	-5.18e-01	1.09e-02
B54	passed	6.04e-01	6.46e-03
b55	passed	-1.96e+00	4.62e-02
B55	passed	1.80e-01	7.64e-03
b61	passed	1.58e+00	8.66e-02
B61	failed	7.48e+00	1.43e+00
b63	passed	2.71e-01	5.28e-03
B63	passed	1.32e+00	6.98e-03
b64	passed	2.91e-01	8.85e-03
B64	passed	1.35e+00	1.21e-02
b65	passed	4.29e-01	2.62e-02
B65	passed	1.67e+00	4.43e-02
b71	passed	-3.74e+00	1.78e-01
B71	failed	5.82e-02	7.55e-03
b72	passed	-7.38e-01	1.17e-02
B72	passed	4.87e-01	5.64e-03

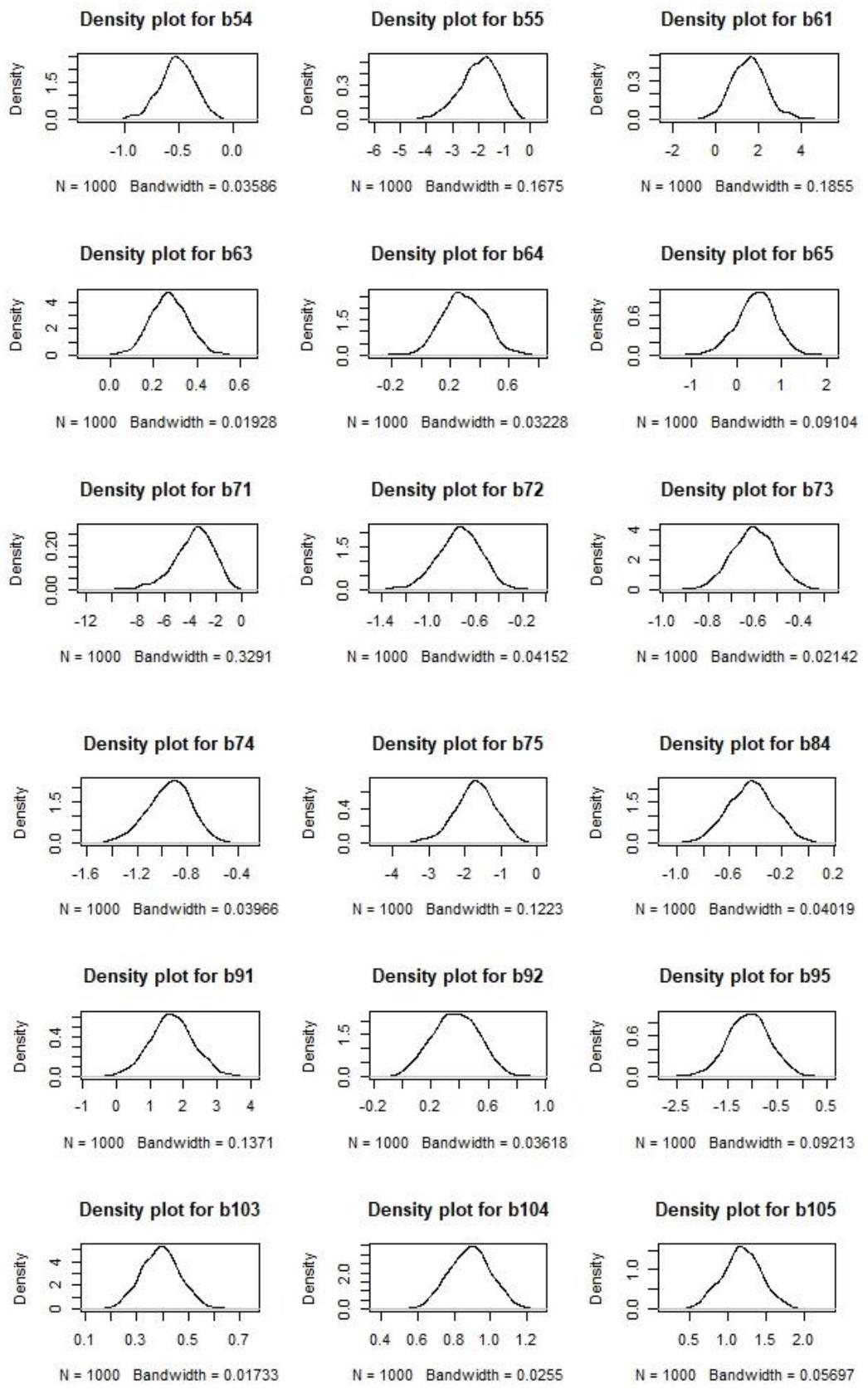
b73	passed	-6.05e-01	5.61e-03
B73	passed	5.49e-01	3.22e-03
b74	passed	-9.44e-01	1.09e-02
B74	passed	3.95e-01	4.25e-03
b75	passed	-1.73e+00	3.96e-02
B75	passed	2.08e-01	8.39e-03
b84	passed	-4.36e-01	1.34e-02
B84	passed	6.57e-01	8.86e-03
b91	passed	1.65e+00	4.78e-02
B91	passed	6.35e+00	3.41e-01
b92	passed	3.72e-01	9.92e-03
B92	passed	1.47e+00	1.46e-02
b95	passed	-1.07e+00	2.55e-02
B95	passed	3.73e-01	9.76e-03
deviance	passed	2.98e+03	7.54e-01

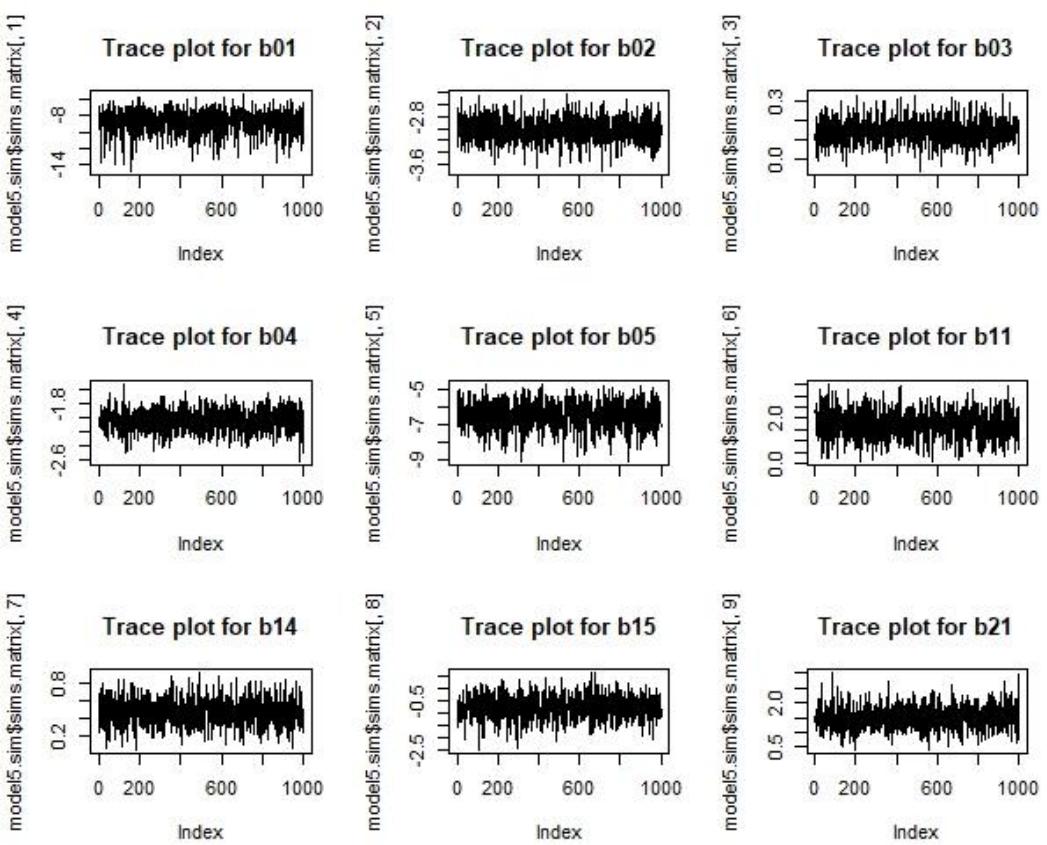
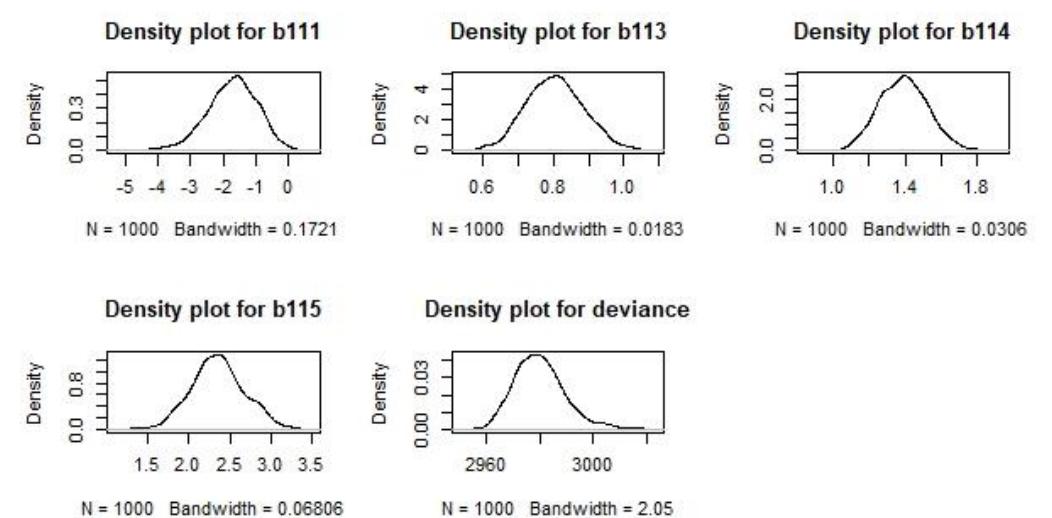


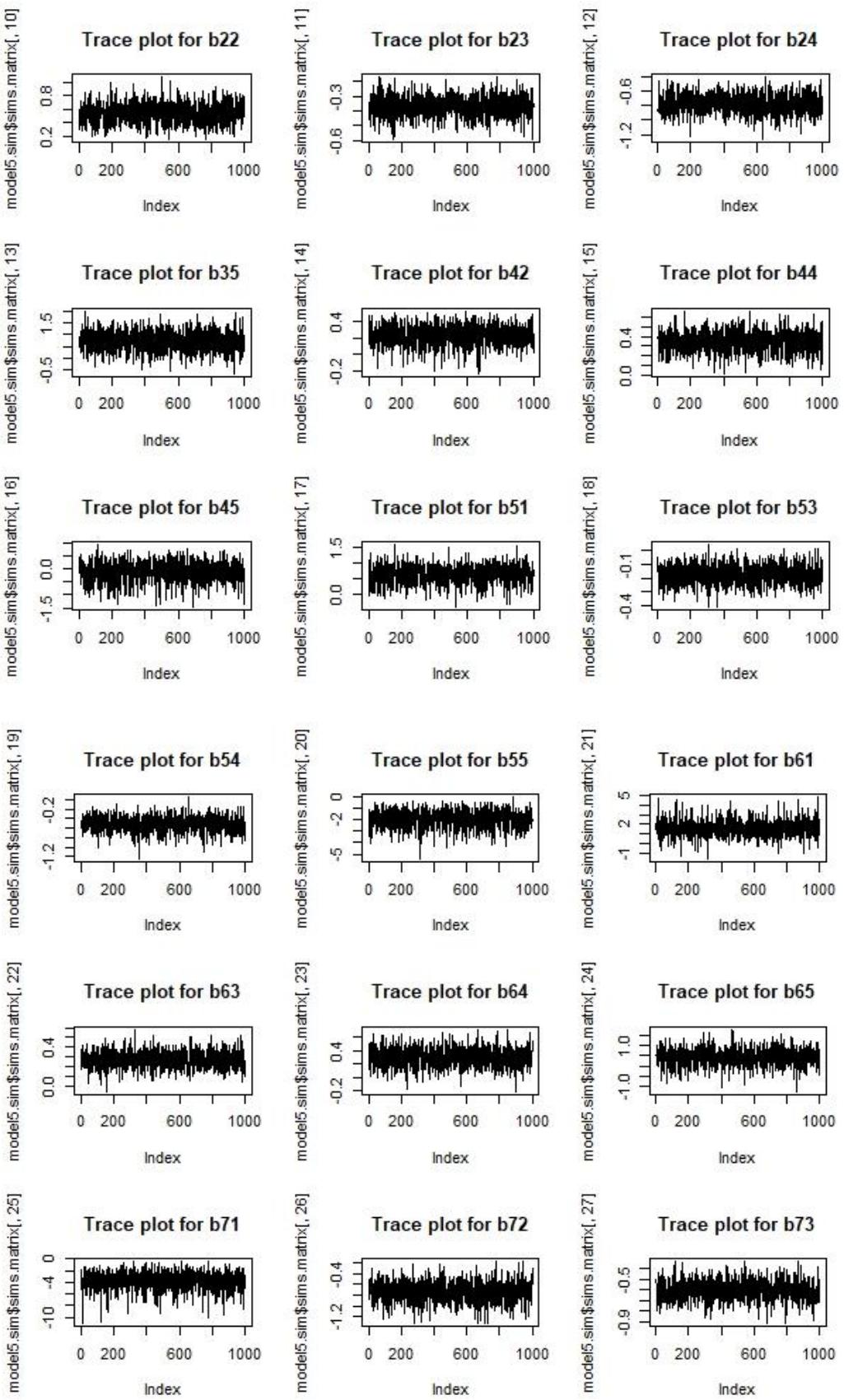


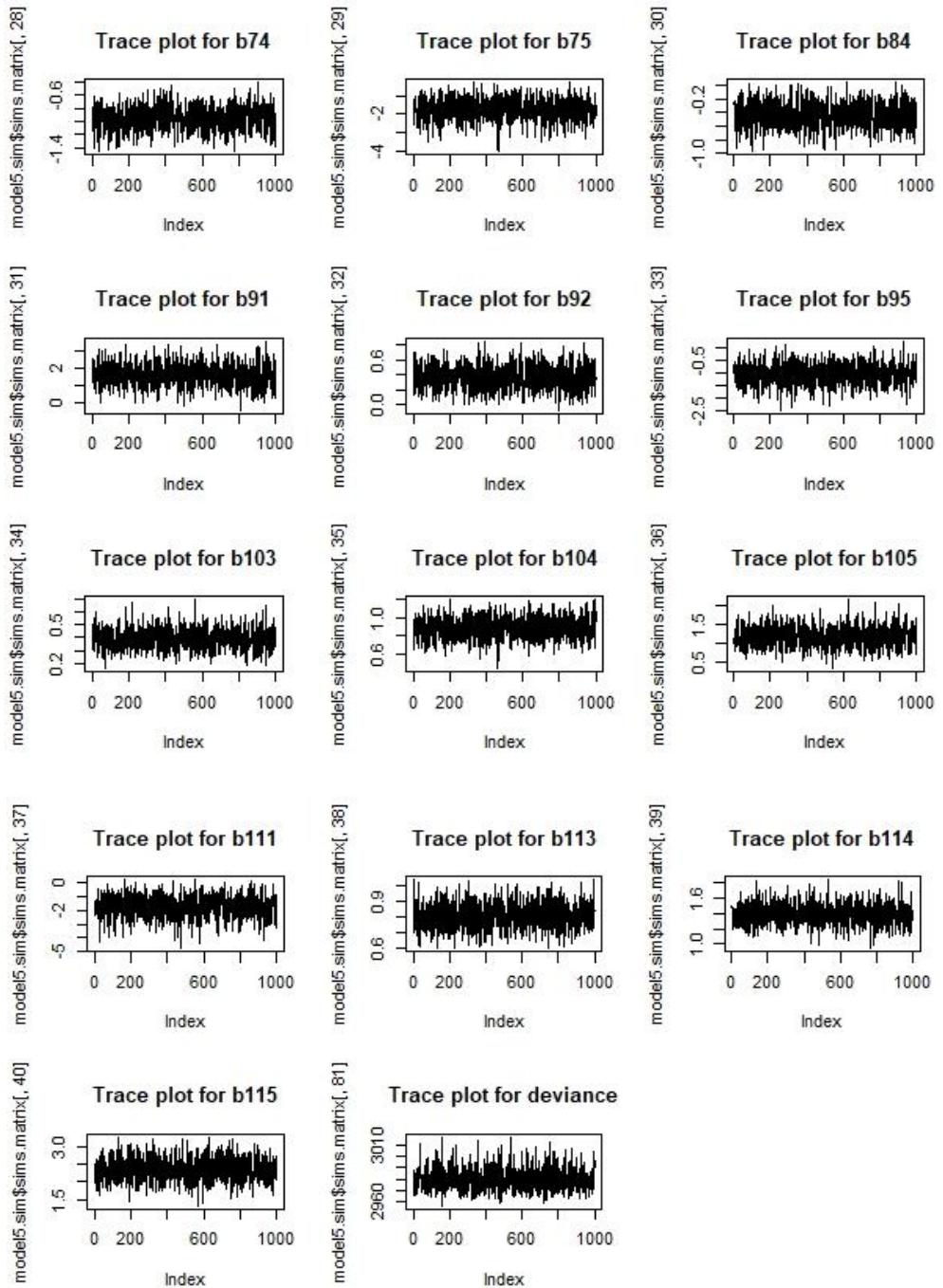


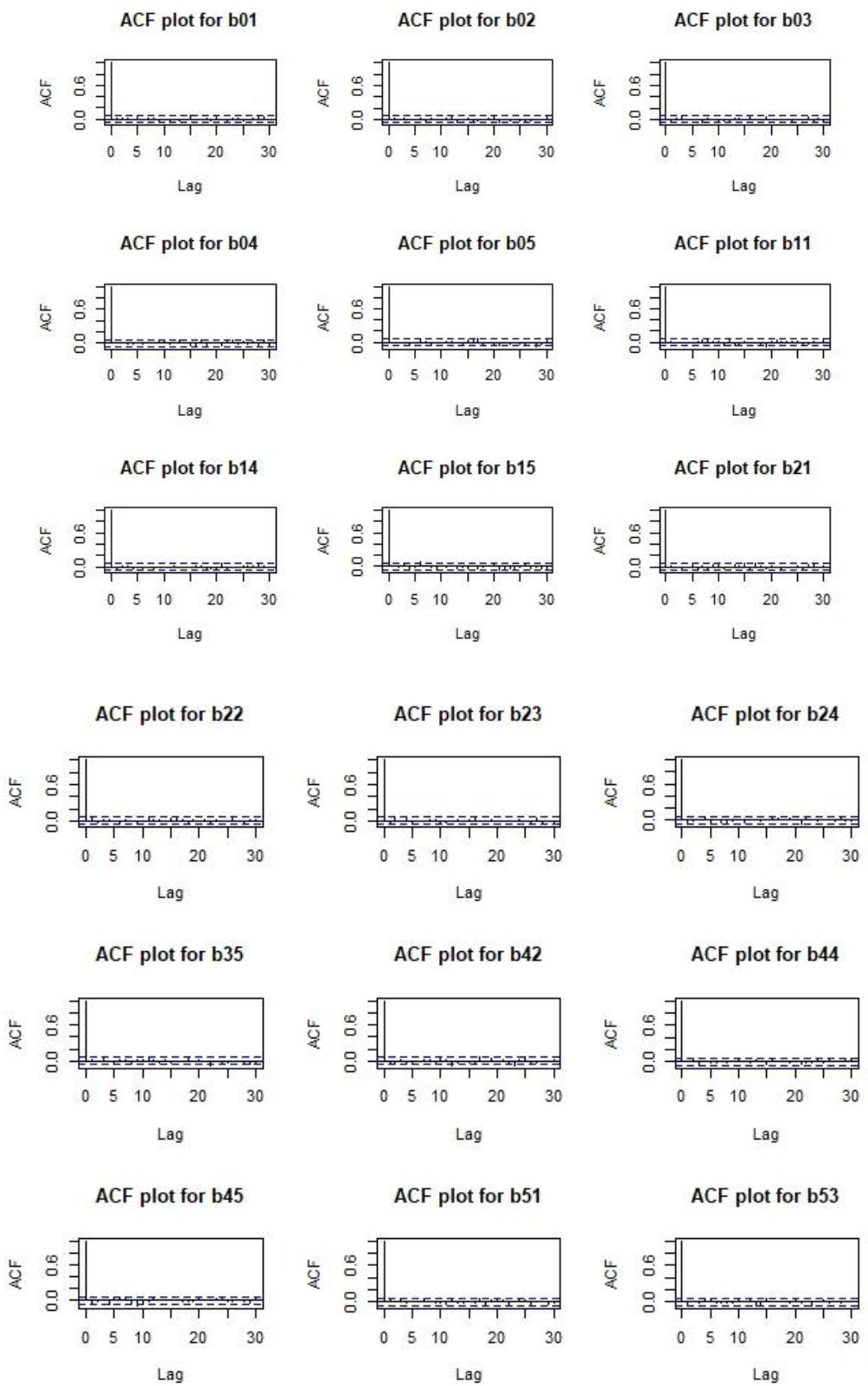


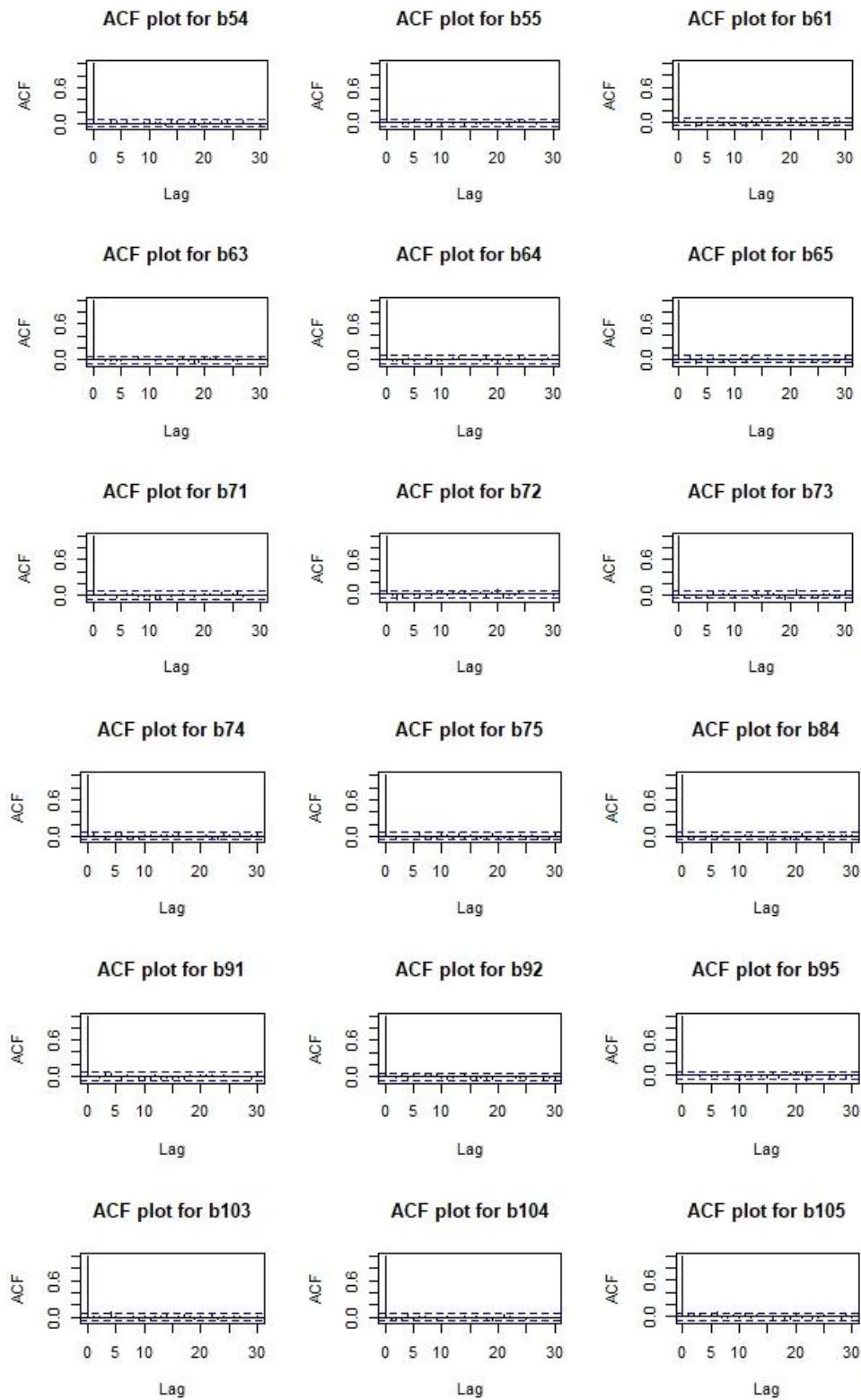


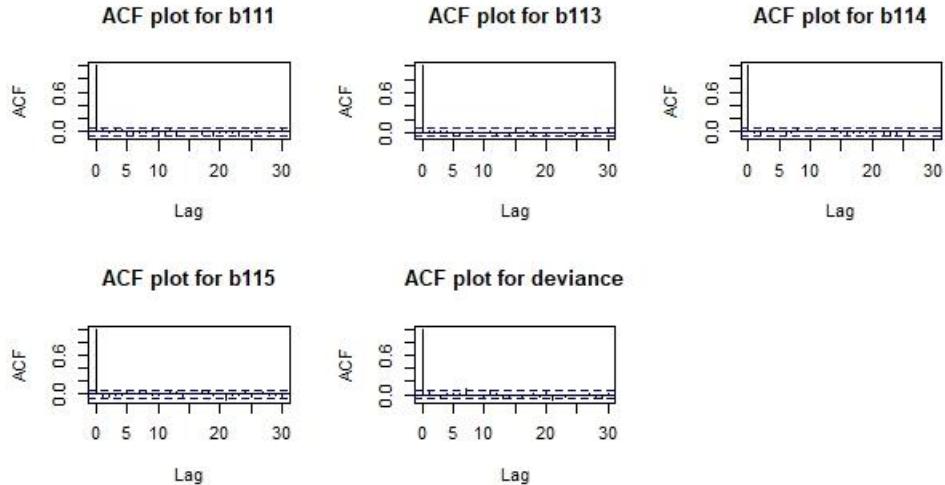












Even though admittedly the Geweke diagnostic could be a little better (maybe by taking a longer burn-in period) and the Heidelberger-Welch test could be improved by taking more iterations, all in all, the graphs and tests indicate a successful convergence for our model.

The posterior results of our model are:

```
Inference for Bugs model at "C:\Users\30697\Desktop\multinomial2.txt", fit
using OpenBUGS,
 1 chains, each with 51000 iterations (first 1000 discarded), n.thin = 50
 n.sims = 1000 iterations saved

      mean    sd   2.5%   25%   50%   75%  97.5%
b01    -8.8  1.4  -12.1  -9.5  -8.6  -7.8  -6.5
b02    -3.0  0.2  -3.4  -3.2  -3.0  -2.9  -2.6
b03     0.1  0.1   0.0   0.1   0.1   0.2   0.3
b04    -2.1  0.2  -2.4  -2.2  -2.1  -1.9  -1.8
b05    -6.5  0.7  -8.1  -7.0  -6.5  -6.0  -5.1
b11     1.7  0.6   0.5   1.3   1.8   2.2   2.9
b14     0.5  0.1   0.2   0.4   0.5   0.6   0.8
b15    -0.7  0.5  -1.7  -1.0  -0.7  -0.4   0.1
b21     1.5  0.4   0.8   1.2   1.5   1.7   2.3
b22     0.5  0.2   0.2   0.4   0.5   0.6   0.8
b23    -0.4  0.1  -0.5  -0.4  -0.4  -0.3  -0.2
b24    -0.8  0.1  -1.0  -0.9  -0.8  -0.7  -0.5
b35     0.7  0.4  -0.1   0.4   0.7   1.0   1.5
b42     0.2  0.1   0.0   0.2   0.2   0.3   0.5
b44     0.3  0.1   0.1   0.3   0.3   0.4   0.5
b45    -0.1  0.4  -1.0  -0.4  -0.1   0.2   0.6
```

b51	0.6	0.3	0.0	0.5	0.7	0.8	1.2
b53	-0.2	0.1	-0.3	-0.2	-0.2	-0.1	0.0
b54	-0.5	0.2	-0.9	-0.6	-0.5	-0.4	-0.2
b55	-2.0	0.7	-3.6	-2.4	-1.9	-1.4	-0.7
b61	1.6	0.9	-0.1	1.0	1.6	2.1	3.5
b63	0.3	0.1	0.1	0.2	0.3	0.3	0.4
b64	0.3	0.1	0.0	0.2	0.3	0.4	0.6
b65	0.4	0.4	-0.4	0.2	0.5	0.7	1.2
b71	-3.7	1.6	-7.5	-4.6	-3.5	-2.7	-1.3
b72	-0.7	0.2	-1.1	-0.9	-0.7	-0.6	-0.4
b73	-0.6	0.1	-0.8	-0.7	-0.6	-0.5	-0.4
b74	-0.9	0.2	-1.3	-1.1	-0.9	-0.8	-0.6
b75	-1.7	0.6	-3.0	-2.1	-1.7	-1.3	-0.7
b84	-0.4	0.2	-0.8	-0.6	-0.4	-0.3	-0.1
b91	1.6	0.6	0.4	1.2	1.6	2.0	2.9
b92	0.4	0.2	0.1	0.3	0.4	0.5	0.7
b95	-1.1	0.4	-1.9	-1.3	-1.1	-0.8	-0.3
b103	0.4	0.1	0.2	0.3	0.4	0.4	0.5
b104	0.9	0.1	0.7	0.8	0.9	1.0	1.1
b105	1.2	0.3	0.7	1.0	1.2	1.4	1.7
b111	-1.7	0.8	-3.4	-2.2	-1.7	-1.2	-0.3
b113	0.8	0.1	0.7	0.7	0.8	0.9	1.0
b114	1.4	0.1	1.1	1.3	1.4	1.5	1.7
b115	2.4	0.3	1.8	2.1	2.4	2.6	3.0
B01	0.0	0.0	0.0	0.0	0.0	0.0	0.0
B02	0.0	0.0	0.0	0.0	0.0	0.1	0.1
B03	1.2	0.1	1.0	1.1	1.2	1.2	1.3
B04	0.1	0.0	0.1	0.1	0.1	0.1	0.2
B05	0.0	0.0	0.0	0.0	0.0	0.0	0.0
B11	6.9	4.5	1.6	3.8	5.8	8.7	18.4
B14	1.6	0.2	1.2	1.5	1.6	1.8	2.2
B15	0.5	0.2	0.2	0.4	0.5	0.7	1.1
B21	4.8	2.0	2.2	3.4	4.4	5.6	9.7
B22	1.7	0.3	1.3	1.5	1.7	1.9	2.3
B23	0.7	0.0	0.6	0.7	0.7	0.7	0.8
B24	0.5	0.1	0.4	0.4	0.5	0.5	0.6
B35	2.2	1.0	0.9	1.5	2.0	2.8	4.6
B42	1.3	0.2	1.0	1.2	1.3	1.4	1.6
B44	1.4	0.1	1.2	1.3	1.4	1.5	1.7
B45	1.0	0.4	0.4	0.7	0.9	1.2	1.7
B51	2.0	0.6	1.0	1.6	2.0	2.3	3.3
B53	0.8	0.1	0.7	0.8	0.8	0.9	1.0
B54	0.6	0.1	0.4	0.5	0.6	0.7	0.8
B55	0.2	0.1	0.0	0.1	0.2	0.2	0.5
B61	7.5	10.6	0.9	2.7	4.8	8.1	32.9
B63	1.3	0.1	1.1	1.2	1.3	1.4	1.5
B64	1.4	0.2	1.0	1.2	1.3	1.5	1.8
B65	1.7	0.7	0.7	1.2	1.6	2.0	3.3
B71	0.1	0.1	0.0	0.0	0.0	0.1	0.3
B72	0.5	0.1	0.3	0.4	0.5	0.5	0.7
B73	0.5	0.1	0.5	0.5	0.5	0.6	0.7
B74	0.4	0.1	0.3	0.3	0.4	0.4	0.5
B75	0.2	0.1	0.1	0.1	0.2	0.3	0.5

```

B84      0.7  0.1   0.5   0.6   0.6   0.7   0.9
B91      6.3  4.4   1.5   3.4   5.2   7.8   17.6
B92      1.5  0.2   1.1   1.3   1.4   1.6   2.0
B95      0.4  0.2   0.2   0.3   0.3   0.4   0.8
B103     1.5  0.1   1.3   1.4   1.5   1.6   1.7
B104     2.4  0.3   2.0   2.2   2.4   2.6   3.0
B105     3.4  0.9   2.0   2.8   3.3   3.9   5.5
B111     0.2  0.2   0.0   0.1   0.2   0.3   0.7
B113     2.2  0.2   1.9   2.1   2.2   2.4   2.6
B114     4.1  0.6   3.1   3.6   4.0   4.4   5.3
B115    11.1 3.7   5.8   8.6   10.5  12.8  19.6
deviance 2979.9 9.4 2964.1 2973.1 2979.1 2985.3 3002.2

DIC info (using the rule, pD = Dbar-Dhat)
pD = 37.4 and DIC = 3017.0
DIC is an estimate of expected predictive error (lower deviance is better).

```

Immediately we see that the DIC value (3017.0) is far smaller than that of the full model (3036.0), and so we conclude that this is indeed a better model than the original. Therefore, according to the table above, the final model will look like this:

$$\begin{aligned}
\ln\left(\frac{p_{i1}}{p_{i6}}\right) &= -8.8 + 1.7 \cdot Z_{i1} + 1.5 \cdot Z_{i2} + 0.6 \cdot Z_{i5} + 1.6 \cdot Z_{i6} - 3.7 \cdot Z_{i7} + 1.6 \cdot Z_{i9} + \\
&\quad - 1.7 \cdot Z_{i11} \\
\ln\left(\frac{p_{i2}}{p_{i6}}\right) &= -3 + 0.5 \cdot Z_{i2} + 0.2 \cdot Z_{i4} - 0.7 \cdot Z_{i7} + 0.4 \cdot Z_{i9} \\
\ln\left(\frac{p_{i3}}{p_{i6}}\right) &= 0.1 - 0.4 \cdot Z_{i2} - 0.2 \cdot Z_{i5} + 0.3 \cdot Z_{i6} - 0.6 \cdot Z_{i7} + 0.4 \cdot Z_{i10} + 0.8 \cdot Z_{i11} \\
\ln\left(\frac{p_{i4}}{p_{i6}}\right) &= -2.1 + 0.5 \cdot Z_{i1} - 0.8 \cdot Z_{i2} + 0.3 \cdot Z_{i4} - 0.5 \cdot Z_{i5} + 0.3 \cdot Z_{i6} - \\
&\quad - 0.9 \cdot Z_{i7} - 0.4 \cdot Z_{i8} + 0.9 \cdot Z_{i10} + 1.4 \cdot Z_{i11} \\
\ln\left(\frac{p_{i5}}{p_{i6}}\right) &= -6.5 - 0.7 \cdot Z_{i1} + 0.7 \cdot Z_{i3} - 0.1 \cdot Z_{i4} - 2 \cdot Z_{i5} + 0.4 \cdot Z_{i6} - \\
&\quad - 1.7 \cdot Z_{i7} - 1.1 \cdot Z_{i9} + 1.2 \cdot Z_{i10} + 2.4 \cdot Z_{i11}
\end{aligned}$$

More iterations would improve this model significantly (by reducing the MCMC error), but unfortunately, we have run out of time. However, levels with more observations (like for instance level 3: Category 3 with reference level 6) are generally better off as far as results are concerned.

Now to interpret them, but first, we have to remember that in order to center our values, we subtracted the mean and divided by the standard deviation of each variable. Therefore, we have no units of measurement for our variables.

Using OpenBUGS, we have calculated the exponentiated values of our parameters, as they will be used to interpret our results in the following manner:

For level 1:

$B01 = \exp\{\beta_{01}\} \cong 0$: If we do not include any of our physiochemical characteristics in this level at all, then there is almost (it is not exactly zero) no chance that our wine will take a score of 3, but rather that of 5. It is kind of nonsensical as a result, but maybe the other factors that are not included in this level are enough to make this distinction between these two scores.

$B11 = \exp\{\beta_{11}\} = 6.9$: Is saying to us that should we keep all the other variables of this level fixed and simultaneously increase the value of fixed acidity by 1 (no units of measurement, remember?), then our odds of receiving a score of 3, rather than 5 will increase by a huge 590%.

$B21 = \exp\{\beta_{21}\} = 4.8$: Tells us that should we increase the value of volatile acidity by 1 unit (no units of measurement), all while keeping the other variables constant, then the odds of being in category 1 that gets a median score of 3, will increase drastically by 380%, instead of being in category 6, which gets a median score of 5.

$B51 = \exp\{\beta_{51}\} = 2$: Tells us that by keeping all other variables of this level fixed, but also increasing the value of chlorides by 1 (no units of measurement), our odds of receiving a score of 3 instead of 5 will increase by 100%.

$B61 = \exp\{\beta_{61}\} = 7.5$: Is an indication that should we keep all other variables of this level constant in their values, while increasing the value of free sulfur dioxide by 1 (no units of measurement), our odds of receiving a score of 3 instead of 5 will increase by an enormous 650%.

$B71 = \exp\{\beta_{71}\} = 0.1$: This value points out that if we keep all the other variables of this level fixed and simultaneously increase the value of total sulfur dioxide by 1 (no units of measurement), then the odds of getting a score of 3 rather than 5 will decrease by 90%.

$B91 = \exp\{\beta_{91}\} = 6.3$: Tells us that if we keep all the other variables of this level fixed, all the while increasing the value of pH by 1 (no units of measurement), then the odds of getting a score of 3 rather than 5 will increase by an astonishing 530%.

$B111 = \exp\{\beta_{111}\} = 0.2$: Is saying to us that should we keep all the other variables of this level fixed, but also increase the value alcohol by 1 (no units of measurement), then the odds of getting a score of 3 rather than 5 will diminish by 80%.

For level 2:

$B02 = \exp\{\beta_{02}\} \cong 0$: Should all the physiochemical characteristics be gone from this level, then there is almost (it is not exactly zero) no chance at all of receiving a score of 4, but instead, we will receive that of 5.

$B22 = \exp\{\beta_{22}\} = 1.7$: Keeping all other variables constant, this value is an indication that should we increase the volatile acidity by 1 (again, no units of measurement), then the odds of receiving a score of 4, rather than 5 will increase by 70%.

$B42 = \exp\{\beta_{42}\} = 1.3$: Is telling us that if we keep all other variables constant in their values and increase the value of residual sugar by 1 (no units of measurement), then the odds of getting a score of 4, and not 5 will increase by 30%.

$B72 = \exp\{\beta_{72}\} = 0.5$: Once more, should other variables remain unchanged and we increase the total sulfur dioxide by 1 (no units of measurement), then our odds of receiving a score of 4 instead of 5 will decrease by 50%.

$B92 = \exp\{\beta_{92}\} = 1.5$: Says that should all other variables remain unchanged in their values and we increase the value of pH by 1 (no units of measurement), then then our odds of receiving a score of 4 rather than 5 will increase by 50%.

For level 3:

$B03 = \exp\{\beta_{03}\} = 1.2$: If all the other variables of this level are zero, then the probability of receiving a score of 6, rather than 5 is increased by 20%. This is what this value means.

$B23 = \exp\{\beta_{23}\} = 0.7$: Should all other variables remain fixed, then if we increase volatile acidity by 1 (no units of measurement), then the odds of scoring 6, rather than 5 with the judges will decrease by 30%.

$B53 = \exp\{\beta_{53}\} = 0.8$: Is saying that should all other variables remain fixed and we increase chlorides by 1 (no units of measurement), then the odds of receiving a score of 6, rather than 5 will diminish by 20%.

$B63 = \exp\{\beta_{63}\} = 1.3$: If we keep all the other variables of this level fixed at their values and increase the free sulfur dioxide by 1 (no units of measurement), then the odds of getting a score of 6 instead of 5 will increase by 30%.

$B73 = \exp\{\beta_{73}\} = 0.5$: Keeping all the other variables at fixed values, should we increase the total sulfur dioxide by 1 (no units of measurement yet again), then the odds of receiving a score of 6 and not 5 will decrease by 50%.

$B103 = \exp\{\beta_{103}\} = 1.5$: If all other variables remain unchanged and we increase the sulphates by 1 (no units of measurement), then the odds of receiving a score of 6, rather than 5 will increase 50%.

$B113 = \exp\{\beta_{113}\} = 2.2$: Is telling us that should we keep all other variables fixed and increase the alcohol by 1 (no units of measurement), then the odds of getting a score of 6 rather than 5 will increase by a whopping 120%.

For level 4:

$B_{04} = \exp\{\beta_{04}\} = 0.1$: Should we equate all the other variables of this level to zero, then the probability of receiving a score of 7 rather than 5 will decrease by 90%.

$B_{14} = \exp\{\beta_{14}\} = 1.6$: This value tells us that if we keep all other physiochemical characteristics of this level at fixed values, all the while increasing the fixed acidity by 1 (no units of measurement), then our odds of scoring a 7 instead of 5 with the experts will increase by 60%.

$B_{24} = \exp\{\beta_{24}\} = 0.5$: If we keep all other variables constant in their values and increase volatile acidity by 1 (there are no units of measurement), then our odds of receiving a score of 7 instead of 5 will decrease by 50%.

$B_{44} = \exp\{\beta_{44}\} = 1.4$: This value tells us that should we keep all other variables of this level fixed and increase the value of residual sugar by 1 (no units of measurement), then our odds of getting a score of 7 and not 5 will increase by 40%.

$B_{54} = \exp\{\beta_{54}\} = 0.6$: This indicates that if we keep all other variables of this level fixed and increase the value of chlorides by 1 (no units of measurement), then our odds of getting a score of 7 instead of 5 will decrease by 40%.

$B_{64} = \exp\{\beta_{64}\} = 1.4$: This says to us that if we keep all other variables of this level fixed, while increasing the value of free sulfur dioxide by 1 (no units of measurement), then our odds of getting a score of 7 rather than 5 will increase by 40%.

$B_{74} = \exp\{\beta_{74}\} = 0.4$: Keeping all other variables fixed in their values and increasing total sulfur dioxide by 1 (no units of measurement), will decrease our odds of receiving a grade of 7 instead of 5 by 60%.

$B_{84} = \exp\{\beta_{84}\} = 0.7$: If we keep all other physiochemical characteristics fixed and increase the value of density by 1 (no units of measurement), then our odds of getting a score of 7 rather than 5 will decrease by 30%.

$B_{104} = \exp\{\beta_{104}\} = 2.4$: If all other variables remain unchanged and we increase the sulphates by 1 (once again, no units of measurement), then the odds of receiving a score of 7, instead of a 5 will increase 140%.

$B_{114} = \exp\{\beta_{114}\} = 4.1$: Should we keep all other variables fixed and increase the alcohol by 1 unit (no units of measurement), then the odds of receiving a score of 7 rather than 5 will increase by a whole 310%.

For level 5:

$B_{05} = \exp\{\beta_{05}\} \cong 0$: Should we equate all the other variables of this level to zero, then the probability of receiving a score of 8, rather than 5 is nearly impossible (it is not exactly zero).

$B_{15} = \exp\{\beta_{15}\} = 0.5$: This result is telling us that should we increase fixed acidity by 1 (no units of measurement), while keeping all other variables of this level fixed, then our odds of getting a score of 8 rather than 5 will decrease by 50%

$B_{35} = \exp\{\beta_{35}\} = 2.2$: Should all other variables remain fixed, then if we increase citric acid by 1 (again, no units of measurement), then the odds of receiving a score of 8, rather than 5 will increase by 120%.

$B_{45} = \exp\{\beta_{45}\} = 1$: If we keep all the other variables of this level fixed at their values and increase the residual sugar by 1 (and again, no units of measurement), then the odds of receiving a score of 8 and 5 are exactly the same. Well not exactly, plus this is bound to change if we take more iterations.

$B_{55} = \exp\{\beta_{55}\} = 0.2$: This value tells us that if we keep all the other variables of this level fixed, while increasing the value of chlorides by 1 (no units of measurement), then our odds of getting a score of 8 rather than 5 will decrease by 80%.

$B_{65} = \exp\{\beta_{65}\} = 1.7$: Having all the other variables at fixed values, should we increase the value of free sulfur dioxide by 1 (no units of measurement yet again), then the odds of receiving a score of 8 and not 5 will increase by 70%.

$B_{75} = \exp\{\beta_{75}\} = 0.2$: Keeping all the other variables at fixed values, should we increase the total sulfur dioxide by 1 (no units of measurement yet again), then the odds of receiving a score of 8 and not 5 will decrease by 80%.

$B_{95} = \exp\{\beta_{95}\} = 0.4$: Is telling us that should we keep all the other physiochemical characteristics of this level constant in their values, while increasing pH by 1 (no units of measurement), then our odds of scoring an 8 instead of 5 with the judges will decrease by 60%.

$B_{105} = \exp\{\beta_{105}\} = 3.4$: If all other variables remain unchanged and we increase the value of sulphates by 1 (no units of measurement), then the odds of receiving a score of 8, rather than 5 will increase drastically by 240%.

$B_{115} = \exp\{\beta_{115}\} = 11.1$: Should we keep all other variables fixed and increase the value of alcohol by 1 (no units of measurement), then the odds of receiving a score of 8 rather than 5 will increase by an insane amount of 1010%.

So, what can we surmise from these results?

- a. It appears that alcohol as well as sulphates appealed to the taste of our judges and in fact those were the two factors that could, in moderation, make or break our wine's score.
- b. There was no good reason to keep a high level of volatile acid in our wine, as it usually made our wine's grade drop to a lower level. In fact, it is not even included in level 5.
- c. The same thing can be said about chlorides, as they only make our wine's score drop to a lower value.

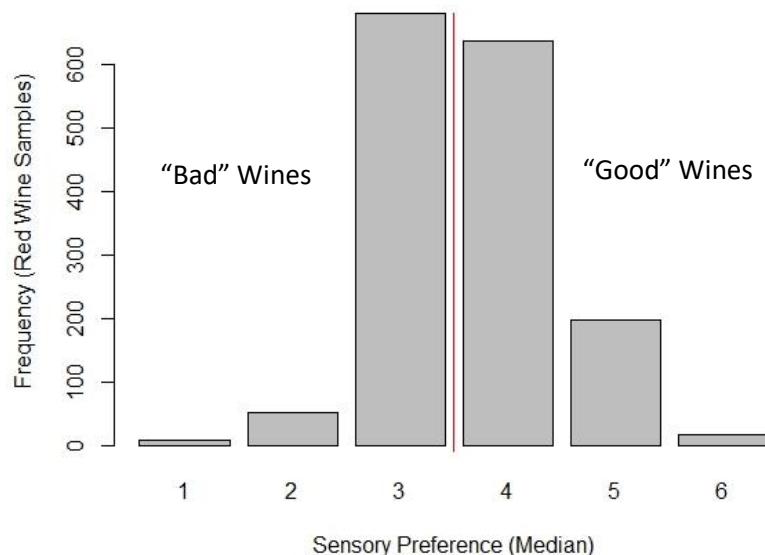
- d. This can be also extended to pH. It is preferable to keep low values of this physiochemical characteristic in our wines.
- e. Free sulfur dioxide is always tasteful, but should be kept in moderation in our wines as it can suddenly change our final score to the lowest.
- f. Density appears only in level 4 and ruins the chances for our wine to get a higher score, so it should be kept at low values in our wine.
- g. Total sulfur dioxide does not contribute to any grade whatsoever, apart from maybe that of 5. Thus, it should be kept at low levels in our wines.

And apart from those variables, the others of our model either contributed or took away from the scores of each level.

What if we wanted to produce only wines of scores 6-8?

Taste is a very fickle way of measurement. A wine that was an 8 today could be a 7 tomorrow and vice versa, except of course if these experts are infallible. As rational human beings, what we would like, should we be in the shoes of these winemakers, is to at least produce wines that would get high scores. So, we would love to see our wine get scores of 6, 7 or even 8, and not 5 or lower. What would we advise these people to look out for as far as physiochemical characteristics are concerned? The answer will be given by a simple binomial logistic regression. We shall split our data into two groups, below 5.5 and above 5.5 and try to determine the most important variables and how they will affect our final scores.

Graphically, this will look like so:



Therefore, if we run our variable selection package (BAS), for a BIC prior on betas and beta-binomial prior on the model, we will end up with the following results:

Marginal Posterior Summaries of Coefficients:

Using BMA

Based on the top 66 models

	post mean	post SD	post p(B != 0)
Intercept	-8.580e+00	6.374e+00	1.000e+00
fixed.acidity	9.771e-03	3.783e-02	8.374e-02
volatile.acidity	-3.016e+00	4.562e-01	9.985e-01
citric.acid	-1.291e-01	4.542e-01	1.003e-01
residual.sugar	4.875e-04	6.607e-03	1.421e-02
chlorides	-2.830e+00	2.447e+00	6.329e-01
free.sulfur.dioxide	1.224e+01	1.326e+01	5.115e-01
total.sulfur.dioxide	-1.471e+01	3.647e+00	9.994e-01
density	2.779e-01	6.254e+00	1.387e-02
pH	-2.608e-02	1.648e-01	3.696e-02
sulphates	2.486e+00	5.058e-01	1.000e+00
alcohol	8.894e-01	7.855e-02	9.997e-01

> summary(res)

	P(B != 0 Y)	model 1	model 2	model 3
Intercept	1.000000000	1.0000	1.00000000	1.00000000
fixed.acidity	0.08374023	0.0000	0.00000000	0.00000000
volatile.acidity	0.99848633	1.0000	1.00000000	1.00000000
citric.acid	0.10034180	0.0000	0.00000000	0.00000000
residual.sugar	0.01420898	0.0000	0.00000000	0.00000000
chlorides	0.63291016	1.0000	1.00000000	0.00000000
free.sulfur.dioxide	0.51147461	1.0000	0.00000000	1.00000000
total.sulfur.dioxide	0.99941406	1.0000	1.00000000	1.00000000
density	0.01386719	0.0000	0.00000000	0.00000000
pH	0.03696289	0.0000	0.00000000	0.00000000
sulphates	0.99995117	1.0000	1.00000000	1.00000000
alcohol	0.99970703	1.0000	1.00000000	1.00000000
BF	NA	1.0000	0.4427971	0.2644016
PostProbs	NA	0.2743	0.2653000	0.1504000
R2	NA	0.2464	0.2423000	0.2418000
dim	NA	7.0000	6.0000000	6.0000000
logmarg	NA	-856.4953	-857.3098941	-857.8255366
		model 4	model 5	
Intercept	1.000000000	1.0000000		
fixed.acidity	0.000000000	1.0000000		
volatile.acidity	1.000000000	1.0000000		
citric.acid	0.000000000	1.0000000		
residual.sugar	0.000000000	0.0000000		
chlorides	0.000000000	0.0000000		
free.sulfur.dioxide	0.000000000	0.0000000		

total.sulfur.dioxide	1.00000000	1.0000000
density	0.00000000	0.0000000
pH	0.00000000	0.0000000
sulphates	1.00000000	1.0000000
alcohol	1.00000000	1.0000000
BF	0.09341534	0.0559124
PostProbs	0.13200000	0.0320000
R2	0.23760000	0.2438000
dim	5.00000000	7.0000000
logmarg	-858.86595013	-859.3792196

The variables that are unimportant in this model, are *Fixed Acidity*, *Citric Acid*, *Residual Sugar*, *Density* and *pH*. Thus, rerunning our model (*binomial2.txt*), without those superfluous variables, will give us the following results:

```
GEWEKE CONVERGENCE DIAGNOSTIC (Z-score)
=====
Iterations used = 1001:11000
Thinning interval = 1
Sample size per chain = 10000

$chain1

Fraction in 1st window = 0.1
Fraction in 2nd window = 0.5

      b[1]      B[1]      b[2]      B[2]      b[3]      B[3]      b[4]      B[4]
-0.0163 -0.0777 -1.3401 -1.3972 -0.8335 -0.8397  0.3659  0.3377
      b[5]      B[5]      b[6]      B[6]      b[7]      B[7] deviance
-0.5331 -0.6139 -0.0422 -0.0269  0.6717  0.6789 -1.4709
```

```
HEIDELBERGER AND WELCH STATIONARITY AND INTERVAL HALFWIDTH TESTS
=====
```

```
Iterations used = 1001:11000
Thinning interval = 1
Sample size per chain = 10000
```

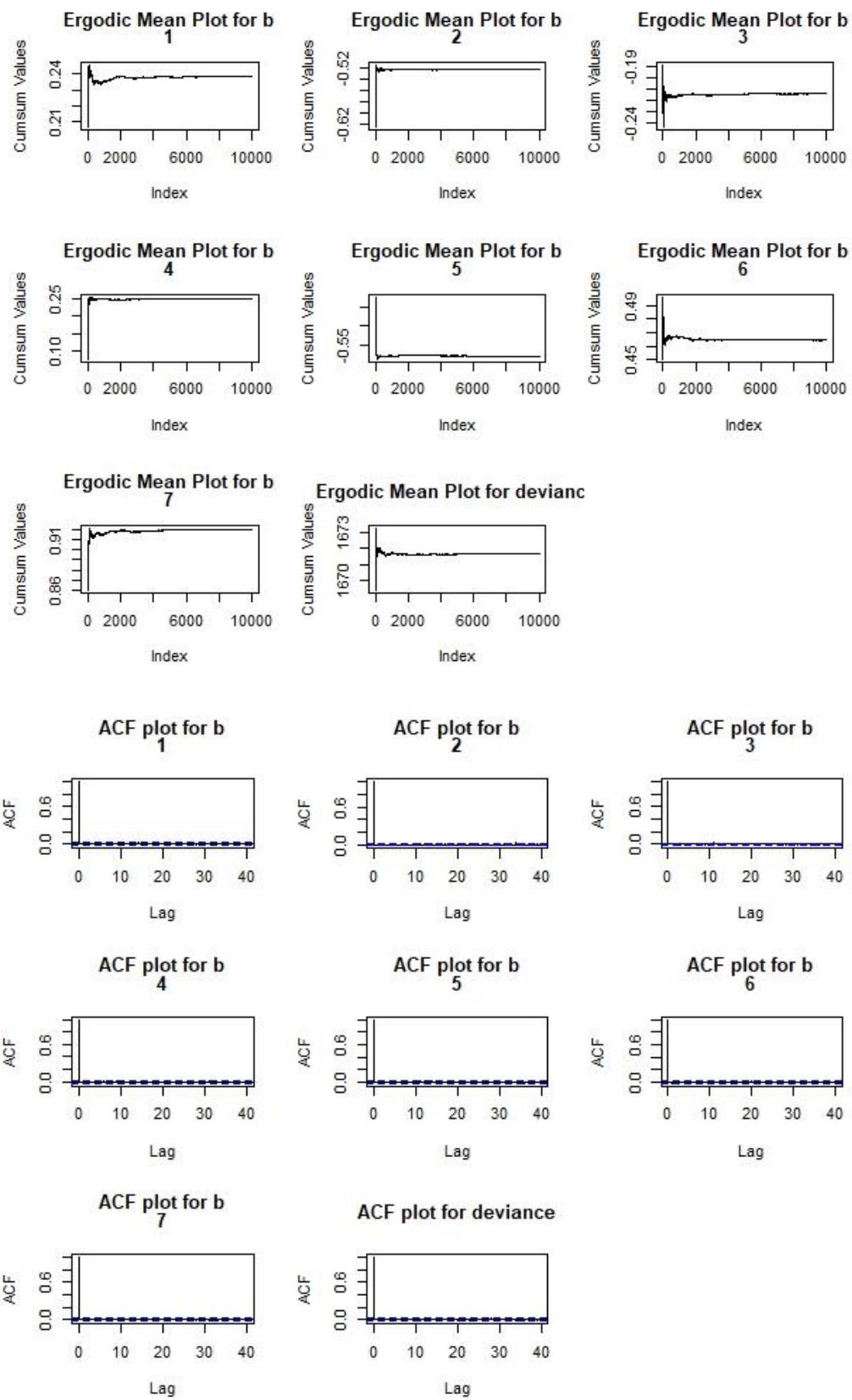
```
Precision of halfwidth test = 0.1
```

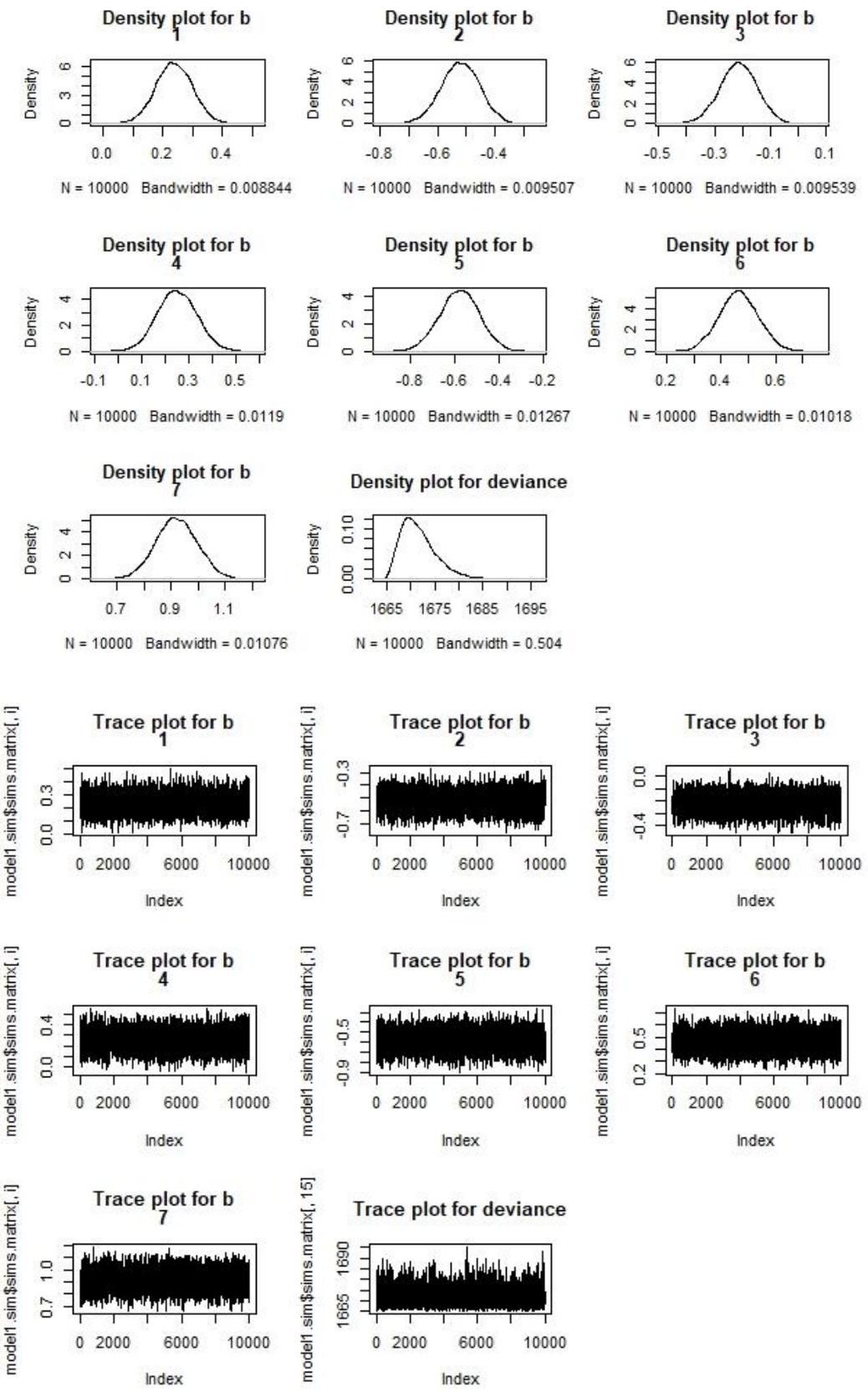
```
$chain1

      Stationarity start      p-value
      test          iteration
b[1]    passed        1      0.5710
B[1]    passed        1      0.6295
b[2]    passed        1      0.0818
```

B[2]	passed	1	0.0814
b[3]	passed	1	0.8495
B[3]	passed	1	0.8663
b[4]	passed	1	0.6204
B[4]	passed	1	0.5536
b[5]	passed	1	0.1061
B[5]	passed	1	0.1318
b[6]	passed	1	0.3401
B[6]	passed	1	0.3277
b[7]	passed	1	0.2239
B[7]	passed	1	0.2310
deviance	passed	2001	0.1183
		Halfwidth	Mean
		test	Halfwidth
b[1]	passed	0.239	0.00149
B[1]	passed	1.272	0.00191
b[2]	passed	-0.522	0.00155
B[2]	passed	0.595	0.00092
b[3]	passed	-0.214	0.00169
B[3]	passed	0.809	0.00135
b[4]	passed	0.252	0.00205
B[4]	passed	1.291	0.00266
b[5]	passed	-0.580	0.00223
B[5]	passed	0.562	0.00125
b[6]	passed	0.465	0.00189
B[6]	passed	1.596	0.00292
b[7]	passed	0.920	0.00185
B[7]	passed	2.516	0.00473
deviance	passed	1671.706	0.10683

And as for plots, we get:





All of these indicate convergence. The results are:

```
Inference for Bugs model at "C:\Users\30697\Desktop\binomial2.txt",
fit using OpenBUGS,
 1 chains, each with 11000 iterations (first 1000 discarded)
n.sims = 10000 iterations saved

      mean   sd   2.5%   25%   50%   75%  97.5%
b[1]    0.2  0.1   0.1   0.2   0.2   0.3   0.4
b[2]   -0.5  0.1  -0.7  -0.6  -0.5  -0.5  -0.4
b[3]   -0.2  0.1  -0.3  -0.3  -0.2  -0.2  -0.1
b[4]    0.3  0.1   0.1   0.2   0.3   0.3   0.4
b[5]   -0.6  0.1  -0.8  -0.6  -0.6  -0.5  -0.4
b[6]    0.5  0.1   0.3   0.4   0.5   0.5   0.6
b[7]    0.9  0.1   0.8   0.9   0.9   1.0   1.1
B[1]   1.3  0.1   1.1   1.2   1.3   1.3   1.4
B[2]   0.6  0.0   0.5   0.6   0.6   0.6   0.7
B[3]   0.8  0.1   0.7   0.8   0.8   0.8   0.9
B[4]   1.3  0.1   1.1   1.2   1.3   1.4   1.5
B[5]   0.6  0.0   0.5   0.5   0.6   0.6   0.7
B[6]   1.6  0.1   1.4   1.5   1.6   1.7   1.8
B[7]   2.5  0.2   2.2   2.4   2.5   2.6   2.9
deviance 1671.7 3.7 1666.4 1668.9 1671.0 1673.7 1680.5

DIC info (using the rule, pD = Dbar-Dhat)
pD = 6.9 and DIC = 1679.0
DIC is an estimate of expected predictive error (lower deviance is better).
```

And so, our model will be:

$$\text{logit}(p_i) = 0.2 - 0.5 \cdot Z_{i2} - 0.2 \cdot Z_{i5} + 0.3 \cdot Z_{i6} - 0.6 \cdot Z_{i7} + 0.5 \cdot Z_{i10} + 0.9 \cdot Z_{i11}$$

The interpretation of those results is based on the exponentiated values of the parameters. Therefore:

$B[1] = 1.3$: If we equate all physiochemical parameters of the model above to zero, then the probability of producing a wine that will be in the upper echelon, which gets a score of 6-8, will be increased by 30%. So, if we do nothing, then our wine will supposedly get categorized as a good wine. This makes little to no sense, as a wine with alcohol of 0 is in reality, grape juice.

$B[2] = 0.6$: Keeping all other variables fixed, while increasing the volatile acidity by 1 (no units of measurement, because we've centered our values), then the odds of characterizing our wine as a “good” wine (i.e. the wine that gets a score of 6-8), will diminish by 40%.

$B[3] = 0.8$: Having kept all the other variables constant and at the same time increasing chlorides by 1 (no units of measurement), our odds of producing a “good” wine have been decreased by 20%.

$B[4] = 1.3$: Should we keep all other variables fixed and increase only free sulfur dioxide by 1 (no units of measurement), then our odds of scoring 6-8 for our wine will increase by 30%.

$B[5] = 0.6$: If we keep all other physiochemical parameters fixed at their values and increase only total sulfur dioxide by 1 (no units of measurement), then our odds for our wine to be categorized as a “good” wine will diminish by 40%.

$B[6] = 1.6$: Keeping all other parameters constant, while increasing sulphates by 1 (no unit of measurement), will increase our odds of having our wine score 6-8 points by 60%.

$B[7] = 2.5$: Should we keep all other parameters of our model fixed, while increasing alcohol by 1 (no unit of measurement), then the odds of producing a good wine will increase by a whole 150%.

Therefore, if we want to “fine tune” our wine and get those good ratings that we need, then we would have to increase, up to a logical extent (we do not want to produce poison) the physiochemical parameters of free sulfur dioxide, sulphates and alcohol, all while keeping a low value for volatile acidity, chlorides and total sulfur dioxide. These are the most important variables for the best wines in that region.

That is how we will theoretically produce a high scoring wine in the Vinho Verde region of Portugal.

Conclusion and final comments:

This was our suggestion, if we can get these physiochemical properties in a bottle of wine, we are almost certain that this bottle could achieve high scores among the judges. However, we have to remember that good wine does not depend exclusively on those physiochemical characteristics, but it needs so much more to succeed. The climate, the type of grapes, the soil, the way it is bottled, all these and more play an important role in the final product and its score.

Thus, these variables that we suggest should be a roadmap and not the destination, meaning that should a bottle of Vinho Verde contain these characteristics, then we have good indications that we are producing a good red wine that people can enjoy. And that is what this paper was all about.

APPENDIX 2

(Models and Algorithms)

All the models are advised to be put in a different .txt and named as the parentheses indicate.

1. The Original Model: Multinomial Logistic Regression (multinomial.txt)

```

model {
  for (i in 1:1599){                                     ## loop over observations
    y[i,1] ~ dcat(p[i, 1:6])                            ## multinomial outcome
    for (j in 1:5){                                      ## loop over categories
      p[i,j] <- expeta[i,j]/(1 + expeta[i,1] +
                                expeta[i,2] + expeta[i,3] +
                                expeta[i,4] + expeta[i,5])

      expeta[i,j] <- exp(eta[i,j])
      eta[i,j] <- b0[j] +
        b1[j]*x[i,1] +
        b2[j]*x[i,2] +
        b3[j]*x[i,3] +
        b4[j]*x[i,4] +
        b5[j]*x[i,5] +
        b6[j]*x[i,6] +
        b7[j]*x[i,7] +
        b8[j]*x[i,8] +
        b9[j]*x[i,9] +
        b10[j]*x[i,10] +
        b11[j]*x[i,11]

    }
    expeta[i,6] <- 1.0                                    ## baseline (reference) category
    p[i,6] <- 1/(1 + expeta[i,1] +
                  expeta[i,2] + expeta[i,3] +
                  expeta[i,4] + expeta[i,5])
  }
## -----
##     Priors
## -----
for (k in 1:5){                                         ## low-information priors
  b0[k] ~ dnorm(0.0, 1.0E-4)
  b1[k] ~ dnorm(0.0, 1.0E-4)
  b2[k] ~ dnorm(0.0, 1.0E-4)
  b3[k] ~ dnorm(0.0, 1.0E-4)
  b4[k] ~ dnorm(0.0, 1.0E-4)
  b5[k] ~ dnorm(0.0, 1.0E-4)
  b6[k] ~ dnorm(0.0, 1.0E-4)
  b7[k] ~ dnorm(0.0, 1.0E-4)
  b8[k] ~ dnorm(0.0, 1.0E-4)
}

```

```
b9[k] ~ dnorm(0.0, 1.0E-4)
b10[k] ~ dnorm(0.0, 1.0E-4)
b11[k] ~ dnorm(0.0, 1.0E-4)
}
#####
## end of model #####

```

2. Empirical Variable Selection through the Multinomial Logistic Regression (empirical.txt)

```
model {
  for (i in 1:1599){                                     ## loop over observations
    y[i,1] ~ dcat(p[i, 1:6])                           ## multinomial outcome

    for (j in 1:5){                                     ## loop over categories
      p[i,j] <- expeta[i,j]/(1 + expeta[i,1] +
                                expeta[i,2] + expeta[i,3] +
                                expeta[i,4] + expeta[i,5])

      expeta[i,j] <- exp(eta[i,j])
      eta[i,j] <- gb0[j] +
        gb1[j]*x[i,1] +
        gb2[j]*x[i,2] +
        gb3[j]*x[i,3] +
        gb4[j]*x[i,4] +
        gb5[j]*x[i,5] +
        gb6[j]*x[i,6] +
        gb7[j]*x[i,7] +
        gb8[j]*x[i,8] +
        gb9[j]*x[i,9] +
        gb10[j]*x[i,10] +
        gb11[j]*x[i,11]

    }
    expeta[i,6] <- 1.0                                 ## baseline (reference) category
    p[i,6] <- 1/(1 + expeta[i,1] +
                  expeta[i,2] + expeta[i,3] +
                  expeta[i,4] + expeta[i,5])
  }
}

## -----
##     Priors
## -----
for (k in 1:5){
  b0[k] ~ dnorm(prop.mean.beta0[k], taub0[k])
  taub0[k] <- (gamma0[k]/1599 + (1-gamma0[k]))/pow(prop.sd.beta0[k], 2)
  gamma0[k] <- 1.0                                    ## constants are always in the model
  gb0[k] <- gamma0[k]*b0[k]

  b1[k] ~ dnorm(prop.mean.beta1[k], taub1[k])
  taub1[k] <- (gamma1[k]/1599 + (1-gamma1[k]))/pow(prop.sd.beta1[k], 2)
  gamma1[k] ~ dbern(0.5)                             ## Uniform
#  gamma1[k] ~ dbern(p1)                            ## Beta-binomial
  gb1[k] <- gamma1[k]*b1[k]

  b2[k] ~ dnorm(prop.mean.beta2[k], taub2[k])
  taub2[k] <- (gamma2[k]/1599 + (1-gamma2[k]))/pow(prop.sd.beta2[k], 2)
  gamma2[k] ~ dbern(0.5)                             ## Uniform
#  gamma2[k] ~ dbern(p2)                            ## Beta-binomial
  gb2[k] <- gamma2[k]*b2[k]

  b3[k] ~ dnorm(prop.mean.beta3[k], taub3[k])
}
```

```

taub3[k] <- (gamma3[k]/1599 + (1-gamma3[k]))/pow(prop.sd.beta3[k], 2)
gamma3[k] ~ dbern(0.5) ## Uniform
# gamma3[k] ~ dbern(p3) ## Beta-binomial
gb3[k] <- gamma3[k]*b3[k]

b4[k] ~ dnorm(prop.mean.beta4[k], taub4[k])
taub4[k] <- (gamma4[k]/1599 + (1-gamma4[k]))/pow(prop.sd.beta4[k], 2)
gamma4[k] ~ dbern(0.5) ## Uniform
# gamma4[k] ~ dbern(p4) ## Beta-binomial
gb4[k] <- gamma4[k]*b4[k]

b5[k] ~ dnorm(prop.mean.beta5[k], taub5[k])
taub5[k] <- (gamma5[k]/1599 + (1-gamma5[k]))/pow(prop.sd.beta5[k], 2)
gamma5[k] ~ dbern(0.5) ## Uniform
# gamma5[k] ~ dbern(p5) ## Beta-binomial
gb5[k] <- gamma5[k]*b5[k]

b6[k] ~ dnorm(prop.mean.beta6[k], taub6[k])
taub6[k] <- (gamma6[k]/1599 + (1-gamma6[k]))/pow(prop.sd.beta6[k], 2)
gamma6[k] ~ dbern(0.5) ## Uniform
# gamma6[k] ~ dbern(p6) ## Beta-binomial
gb6[k] <- gamma6[k]*b6[k]

b7[k] ~ dnorm(prop.mean.beta7[k], taub7[k])
taub7[k] <- (gamma7[k]/1599 + (1-gamma7[k]))/pow(prop.sd.beta7[k], 2)
gamma7[k] ~ dbern(0.5) ## Uniform
# gamma7[k] ~ dbern(p7) ## Beta-binomial
gb7[k] <- gamma7[k]*b7[k]

b8[k] ~ dnorm(prop.mean.beta8[k], taub8[k])
taub8[k] <- (gamma8[k]/1599 + (1-gamma8[k]))/pow(prop.sd.beta8[k], 2)
gamma8[k] ~ dbern(0.5) ## Uniform
# gamma8[k] ~ dbern(p8) ## Beta-binomial
gb8[k] <- gamma8[k]*b8[k]

b9[k] ~ dnorm(prop.mean.beta9[k], taub9[k])
taub9[k] <- (gamma9[k]/1599 + (1-gamma9[k]))/pow(prop.sd.beta9[k], 2)
gamma9[k] ~ dbern(0.5) ## Uniform
# gamma9[k] ~ dbern(p9) ## Beta-binomial
gb9[k] <- gamma9[k]*b9[k]

b10[k] ~ dnorm(prop.mean.beta10[k], taub10[k])
taub10[k] <- (gamma10[k]/1599 + (1-gamma10[k]))/pow(prop.sd.beta10[k], 2)
gamma10[k] ~ dbern(0.5) ## Uniform
# gamma10[k] ~ dbern(p10) ## Beta-binomial
gb10[k] <- gamma10[k]*b10[k]

b11[k] ~ dnorm(prop.mean.beta11[k], taub11[k])
taub11[k] <- (gamma11[k]/1599 + (1-gamma11[k]))/pow(prop.sd.beta11[k], 2)
gamma11[k] ~ dbern(0.5) ## Uniform
# gamma11[k] ~ dbern(p11) ## Beta-binomial
gb11[k] <- gamma11[k]*b11[k]
}

```

```
# For the Beta-Binomial
# p1 ~ dbeta(2,10)
# p2 ~ dbeta(2,10)
# p3 ~ dbeta(2,10)
# p4 ~ dbeta(2,10)
# p5 ~ dbeta(2,10)
# p6 ~ dbeta(2,10)
# p7 ~ dbeta(2,10)
# p8 ~ dbeta(2,10)
# p9 ~ dbeta(2,10)
# p10 ~ dbeta(2,10)
# p11 ~ dbeta(2,10)
}
##### end of model #####
```

3. Variable Selection using g-prior through the Multinomial Logistic Regression (gprior.txt)

```
model {
  for (i in 1:1599){                                     ## loop over observations
    y[i,1] ~ dcat(p[i, 1:6])                           ## multinomial outcome

    for (j in 1:5){                                     ## loop over categories
      p[i,j] <- expeta[i,j]/(1 + expeta[i,1] +
                                expeta[i,2] + expeta[i,3] +
                                expeta[i,4] + expeta[i,5])

      expeta[i,j] <- exp(eta[i,j])
      eta[i,j] <- alpha[j] +
        gb1[j]*x[i,1] +
        gb2[j]*x[i,2] +
        gb3[j]*x[i,3] +
        gb4[j]*x[i,4] +
        gb5[j]*x[i,5] +
        gb6[j]*x[i,6] +
        gb7[j]*x[i,7] +
        gb8[j]*x[i,8] +
        gb9[j]*x[i,9] +
        gb10[j]*x[i,10] +
        gb11[j]*x[i,11]

    }
    expeta[i,6] <- 1.0                                 ## baseline (reference) category
    p[i,6] <- 1/(1 + expeta[i,1] +
                  expeta[i,2] + expeta[i,3] +
                  expeta[i,4] + expeta[i,5])
  }
}

## -----
##     Priors
## -----
# Liang g-prior
logtau ~ dflat()
tau <- exp(logtau)
#####
for (j in 1:5) {
  alpha[j] ~ dflat()
}
#####
B1[1:5] ~ dnorm(mean.beta1[1:5], T1[1:5, 1:5])
for (j in 1:5) {
  mean.beta1[j] <- (1-gamma1[j])*prop.mean.beta1[j]
  for (k in 1:5) {
    T1[j,k] <- gamma1[j]*gamma1[k]*tau*XTX[j,k]/G + (1-
      gamma1[j]*gamma1[k])*equals(j,k)*pow(prop.sd.beta1[k],-2)
  }
}
for (j in 1:5) {
  gamma1[j] ~ dbern(0.5) # Uniform
#  gamma1[j] ~ dbern(p1) # Beta-Binomial
```

```

gb1[j] <- gamma1[j]*B1[j]
}
#####
B2[1:5] ~ dmnorm(mean.beta2[1:5], T2[1:5, 1:5])
for (j in 1:5) {
  mean.beta2[j] <- (1-gamma2[j])*prop.mean.beta2[j]
  for (k in 1:5) {
    T2[j,k] <- gamma2[j]*gamma2[k]*tau*XTX[j,k]/G + (1-
gamma2[j]*gamma2[k])*equals(j,k)*pow(prop.sd.beta2[k],-2)
  }
}
for (j in 1:5) {
  gamma2[j] ~ dbern(0.5) # Uniform
#  gamma2[j] ~ dbern(p1) # Beta-Binomial
  gb2[j] <- gamma2[j]*B2[j]
}
#####
B3[1:5] ~ dmnorm(mean.beta3[1:5], T3[1:5, 1:5])
for (j in 1:5) {
  mean.beta3[j] <- (1-gamma3[j])*prop.mean.beta3[j]
  for (k in 1:5) {
    T3[j,k] <- gamma3[j]*gamma3[k]*tau*XTX[j,k]/G + (1-
gamma3[j]*gamma3[k])*equals(j,k)*pow(prop.sd.beta3[k],-2)
  }
}
for (j in 1:5) {
  gamma3[j] ~ dbern(0.5) # Uniform
#  gamma3[j] ~ dbern(p1) # Beta-Binomial
  gb3[j] <- gamma3[j]*B3[j]
}
#####
B4[1:5] ~ dmnorm(mean.beta4[1:5], T4[1:5, 1:5])
for (j in 1:5) {
  mean.beta4[j] <- (1-gamma4[j])*prop.mean.beta4[j]
  for (k in 1:5) {
    T4[j,k] <- gamma4[j]*gamma4[k]*tau*XTX[j,k]/G + (1-
gamma4[j]*gamma4[k])*equals(j,k)*pow(prop.sd.beta4[k],-2)
  }
}
for (j in 1:5) {
  gamma4[j] ~ dbern(0.5) # Uniform
#  gamma4[j] ~ dbern(p1) # Beta-Binomial
  gb4[j] <- gamma4[j]*B4[j]
}
#####
B5[1:5] ~ dmnorm(mean.beta5[1:5], T5[1:5, 1:5])
for (j in 1:5) {
  mean.beta5[j] <- (1-gamma5[j])*prop.mean.beta5[j]
  for (k in 1:5) {
    T5[j,k] <- gamma5[j]*gamma5[k]*tau*XTX[j,k]/G + (1-
gamma5[j]*gamma5[k])*equals(j,k)*pow(prop.sd.beta5[k],-2)
  }
}

```

```

for (j in 1:5) {
  gamma5[j] ~ dbern(0.5) # Uniform
#  gamma5[j] ~ dbern(p1) # Beta-Binomial
  gb5[j] <- gamma5[j]*B5[j]
}
#####
B6[1:5] ~ dmnorm(mean.beta6[1:5], T6[1:5, 1:5])
for (j in 1:5) {
  mean.beta6[j] <- (1-gamma6[j])*prop.mean.beta6[j]
  for (k in 1:5) {
    T6[j,k] <- gamma6[j]*gamma6[k]*tau*XTX[j,k]/G + (1-
gamma6[j]*gamma6[k])*equals(j,k)*pow(prop.sd.beta6[k],-2)
  }
}
for (j in 1:5) {
  gamma6[j] ~ dbern(0.5) # Uniform
#  gamma6[j] ~ dbern(p1) # Beta-Binomial
  gb6[j] <- gamma6[j]*B6[j]
}
#####
B7[1:5] ~ dmnorm(mean.beta7[1:5], T7[1:5, 1:5])
for (j in 1:5) {
  mean.beta7[j] <- (1-gamma7[j])*prop.mean.beta7[j]
  for (k in 1:5) {
    T7[j,k] <- gamma7[j]*gamma7[k]*tau*XTX[j,k]/G + (1-
gamma7[j]*gamma7[k])*equals(j,k)*pow(prop.sd.beta7[k],-2)
  }
}
for (j in 1:5) {
  gamma7[j] ~ dbern(0.5) # Uniform
#  gamma7[j] ~ dbern(p1) # Beta-Binomial
  gb7[j] <- gamma7[j]*B7[j]
}
#####
B8[1:5] ~ dmnorm(mean.beta8[1:5], T8[1:5, 1:5])
for (j in 1:5) {
  mean.beta8[j] <- (1-gamma8[j])*prop.mean.beta8[j]
  for (k in 1:5) {
    T8[j,k] <- gamma8[j]*gamma8[k]*tau*XTX[j,k]/G + (1-
gamma8[j]*gamma8[k])*equals(j,k)*pow(prop.sd.beta8[k],-2)
  }
}
for (j in 1:5) {
  gamma8[j] ~ dbern(0.5) # Uniform
#  gamma8[j] ~ dbern(p1) # Beta-Binomial
  gb8[j] <- gamma8[j]*B8[j]
}
#####
B9[1:5] ~ dmnorm(mean.beta9[1:5], T9[1:5, 1:5])
for (j in 1:5) {
  mean.beta9[j] <- (1-gamma9[j])*prop.mean.beta9[j]
  for (k in 1:5) {

```

```

T9[j,k] <- gamma9[j]*gamma9[k]*tau*XTX[j,k]/G + (1-
gamma9[j]*gamma9[k])*equals(j,k)*pow(prop.sd.beta9[k],-2)
}
}
for (j in 1:5) {
  gamma9[j] ~ dbern(0.5) # Uniform
#  gamma9[j] ~ dbern(p1) # Beta-Binomial
  gb9[j] <- gamma9[j]*B9[j]
}
#####
B10[1:5] ~ dmnorm(mean.beta10[1:5], T10[1:5, 1:5])
for (j in 1:5) {
  mean.beta10[j] <- (1-gamma10[j])*prop.mean.beta10[j]
  for (k in 1:5) {
    T10[j,k] <- gamma10[j]*gamma10[k]*tau*XTX[j,k]/G + (1-
gamma10[j]*gamma10[k])*equals(j,k)*pow(prop.sd.beta10[k],-2)
  }
}
for (j in 1:5) {
  gamma10[j] ~ dbern(0.5) # Uniform
#  gamma10[j] ~ dbern(p1) # Beta-Binomial
  gb10[j] <- gamma10[j]*B10[j]
}
#####
B11[1:5] ~ dmnorm(mean.beta11[1:5], T11[1:5, 1:5])
for (j in 1:5) {
  mean.beta11[j] <- (1-gamma11[j])*prop.mean.beta11[j]
  for (k in 1:5) {
    T11[j,k] <- gamma11[j]*gamma11[k]*tau*XTX[j,k]/G + (1-
gamma11[j]*gamma11[k])*equals(j,k)*pow(prop.sd.beta11[k],-2)
  }
}
for (j in 1:5) {
  gamma11[j] ~ dbern(0.5) # Uniform
#  gamma11[j] ~ dbern(p1) # Beta-Binomial
  gb11[j] <- gamma11[j]*B11[j]
}
#####
# For the Beta-Binomial
#  p1 ~ dbeta(2,10)
#  p2 ~ dbeta(2,10)
#  p3 ~ dbeta(2,10)
#  p4 ~ dbeta(2,10)
#  p5 ~ dbeta(2,10)
#  p6 ~ dbeta(2,10)
#  p7 ~ dbeta(2,10)
#  p8 ~ dbeta(2,10)
#  p9 ~ dbeta(2,10)
#  p10 ~ dbeta(2,10)
#  p11 ~ dbeta(2,10)

# hyper-prior
G <- 1599

```

```
}
```

```
#####
##### end of model #####
#####
```

4. Variable Selection using hyper-g through the Multinomial Logistic Regression (hyperg.txt)

```
model {
  for (i in 1:1599){                                     ## loop over observations
    y[i,1] ~ dcat(p[i, 1:6])                           ## multinomial outcome

    for (j in 1:5){                                     ## loop over categories
      p[i,j] <- expeta[i,j]/(1 + expeta[i,1] +
                                expeta[i,2] + expeta[i,3] +
                                expeta[i,4] + expeta[i,5])

      expeta[i,j] <- exp(eta[i,j])
      eta[i,j] <- alpha[j] +
        gb1[j]*x[i,1] +
        gb2[j]*x[i,2] +
        gb3[j]*x[i,3] +
        gb4[j]*x[i,4] +
        gb5[j]*x[i,5] +
        gb6[j]*x[i,6] +
        gb7[j]*x[i,7] +
        gb8[j]*x[i,8] +
        gb9[j]*x[i,9] +
        gb10[j]*x[i,10] +
        gb11[j]*x[i,11]

    }
    expeta[i,6] <- 1.0                                 ## baseline (reference) category
    p[i,6] <- 1/(1 + expeta[i,1] +
                  expeta[i,2] + expeta[i,3] +
                  expeta[i,4] + expeta[i,5])
  }
}

## -----
##     Priors
## -----
# Liang hyper-g
logtau ~ dflat()
tau <- exp(logtau)
#####
for (j in 1:5) {
  alpha[j] ~ dflat()
}
#####
B1[1:5] ~ dnorm(mean.beta1[1:5], T1[1:5, 1:5])
for (j in 1:5) {
  mean.beta1[j] <- (1-gamma1[j])*prop.mean.beta1[j]
  for (k in 1:5) {
    T1[j,k] <- gamma1[j]*gamma1[k]*tau*XTX[j,k]/G + (1-
      gamma1[j]*gamma1[k])*equals(j,k)*pow(prop.sd.beta1[k],-2)
  }
}
for (j in 1:5) {
  gamma1[j] ~ dbern(0.5) # Uniform
#  gamma1[j] ~ dbern(p1) # Beta-Binomial
```

```

gb1[j] <- gamma1[j]*B1[j]
}
#####
B2[1:5] ~ dmnorm(mean.beta2[1:5], T2[1:5, 1:5])
for (j in 1:5) {
  mean.beta2[j] <- (1-gamma2[j])*prop.mean.beta2[j]
  for (k in 1:5) {
    T2[j,k] <- gamma2[j]*gamma2[k]*tau*XTX[j,k]/G + (1-
gamma2[j]*gamma2[k])*equals(j,k)*pow(prop.sd.beta2[k],-2)
  }
}
for (j in 1:5) {
  gamma2[j] ~ dbern(0.5) # Uniform
#  gamma2[j] ~ dbern(p1) # Beta-Binomial
  gb2[j] <- gamma2[j]*B2[j]
}
#####
B3[1:5] ~ dmnorm(mean.beta3[1:5], T3[1:5, 1:5])
for (j in 1:5) {
  mean.beta3[j] <- (1-gamma3[j])*prop.mean.beta3[j]
  for (k in 1:5) {
    T3[j,k] <- gamma3[j]*gamma3[k]*tau*XTX[j,k]/G + (1-
gamma3[j]*gamma3[k])*equals(j,k)*pow(prop.sd.beta3[k],-2)
  }
}
for (j in 1:5) {
  gamma3[j] ~ dbern(0.5) # Uniform
#  gamma3[j] ~ dbern(p1) # Beta-Binomial
  gb3[j] <- gamma3[j]*B3[j]
}
#####
B4[1:5] ~ dmnorm(mean.beta4[1:5], T4[1:5, 1:5])
for (j in 1:5) {
  mean.beta4[j] <- (1-gamma4[j])*prop.mean.beta4[j]
  for (k in 1:5) {
    T4[j,k] <- gamma4[j]*gamma4[k]*tau*XTX[j,k]/G + (1-
gamma4[j]*gamma4[k])*equals(j,k)*pow(prop.sd.beta4[k],-2)
  }
}
for (j in 1:5) {
  gamma4[j] ~ dbern(0.5) # Uniform
#  gamma4[j] ~ dbern(p1) # Beta-Binomial
  gb4[j] <- gamma4[j]*B4[j]
}
#####
B5[1:5] ~ dmnorm(mean.beta5[1:5], T5[1:5, 1:5])
for (j in 1:5) {
  mean.beta5[j] <- (1-gamma5[j])*prop.mean.beta5[j]
  for (k in 1:5) {
    T5[j,k] <- gamma5[j]*gamma5[k]*tau*XTX[j,k]/G + (1-
gamma5[j]*gamma5[k])*equals(j,k)*pow(prop.sd.beta5[k],-2)
  }
}

```

```

for (j in 1:5) {
  gamma5[j] ~ dbern(0.5) # Uniform
#  gamma5[j] ~ dbern(p1) # Beta-Binomial
  gb5[j] <- gamma5[j]*B5[j]
}
#####
B6[1:5] ~ dmnorm(mean.beta6[1:5], T6[1:5, 1:5])
for (j in 1:5) {
  mean.beta6[j] <- (1-gamma6[j])*prop.mean.beta6[j]
  for (k in 1:5) {
    T6[j,k] <- gamma6[j]*gamma6[k]*tau*XTX[j,k]/G + (1-
gamma6[j]*gamma6[k])*equals(j,k)*pow(prop.sd.beta6[k],-2)
  }
}
for (j in 1:5) {
  gamma6[j] ~ dbern(0.5) # Uniform
#  gamma6[j] ~ dbern(p1) # Beta-Binomial
  gb6[j] <- gamma6[j]*B6[j]
}
#####
B7[1:5] ~ dmnorm(mean.beta7[1:5], T7[1:5, 1:5])
for (j in 1:5) {
  mean.beta7[j] <- (1-gamma7[j])*prop.mean.beta7[j]
  for (k in 1:5) {
    T7[j,k] <- gamma7[j]*gamma7[k]*tau*XTX[j,k]/G + (1-
gamma7[j]*gamma7[k])*equals(j,k)*pow(prop.sd.beta7[k],-2)
  }
}
for (j in 1:5) {
  gamma7[j] ~ dbern(0.5) # Uniform
#  gamma7[j] ~ dbern(p1) # Beta-Binomial
  gb7[j] <- gamma7[j]*B7[j]
}
#####
B8[1:5] ~ dmnorm(mean.beta8[1:5], T8[1:5, 1:5])
for (j in 1:5) {
  mean.beta8[j] <- (1-gamma8[j])*prop.mean.beta8[j]
  for (k in 1:5) {
    T8[j,k] <- gamma8[j]*gamma8[k]*tau*XTX[j,k]/G + (1-
gamma8[j]*gamma8[k])*equals(j,k)*pow(prop.sd.beta8[k],-2)
  }
}
for (j in 1:5) {
  gamma8[j] ~ dbern(0.5) # Uniform
#  gamma8[j] ~ dbern(p1) # Beta-Binomial
  gb8[j] <- gamma8[j]*B8[j]
}
#####
B9[1:5] ~ dmnorm(mean.beta9[1:5], T9[1:5, 1:5])
for (j in 1:5) {
  mean.beta9[j] <- (1-gamma9[j])*prop.mean.beta9[j]
  for (k in 1:5) {

```

```

T9[j,k] <- gamma9[j]*gamma9[k]*tau*XTX[j,k]/G + (1-
gamma9[j]*gamma9[k])*equals(j,k)*pow(prop.sd.beta9[k],-2)
}
}
for (j in 1:5) {
  gamma9[j] ~ dbern(0.5) # Uniform
#  gamma9[j] ~ dbern(p1) # Beta-Binomial
  gb9[j] <- gamma9[j]*B9[j]
}
#####
B10[1:5] ~ dmnorm(mean.beta10[1:5], T10[1:5, 1:5])
for (j in 1:5) {
  mean.beta10[j] <- (1-gamma10[j])*prop.mean.beta10[j]
  for (k in 1:5) {
    T10[j,k] <- gamma10[j]*gamma10[k]*tau*XTX[j,k]/G + (1-
gamma10[j]*gamma10[k])*equals(j,k)*pow(prop.sd.beta10[k],-2)
  }
}
for (j in 1:5) {
  gamma10[j] ~ dbern(0.5) # Uniform
#  gamma10[j] ~ dbern(p1) # Beta-Binomial
  gb10[j] <- gamma10[j]*B10[j]
}
#####
B11[1:5] ~ dmnorm(mean.beta11[1:5], T11[1:5, 1:5])
for (j in 1:5) {
  mean.beta11[j] <- (1-gamma11[j])*prop.mean.beta11[j]
  for (k in 1:5) {
    T11[j,k] <- gamma11[j]*gamma11[k]*tau*XTX[j,k]/G + (1-
gamma11[j]*gamma11[k])*equals(j,k)*pow(prop.sd.beta11[k],-2)
  }
}
for (j in 1:5) {
  gamma11[j] ~ dbern(0.5) # Uniform
#  gamma11[j] ~ dbern(p1) # Beta-Binomial
  gb11[j] <- gamma11[j]*B11[j]
}
#####
# For the Beta-Binomial
#  p1 ~ dbeta(2,10)
#  p2 ~ dbeta(2,10)
#  p3 ~ dbeta(2,10)
#  p4 ~ dbeta(2,10)
#  p5 ~ dbeta(2,10)
#  p6 ~ dbeta(2,10)
#  p7 ~ dbeta(2,10)
#  p8 ~ dbeta(2,10)
#  p9 ~ dbeta(2,10)
#  p10 ~ dbeta(2,10)
#  p11 ~ dbeta(2,10)
#####
# hyper-prior
a <- 3

```

```
bw <- a/2-1
w ~ dbeta( 1, bw )
G <- w/(1-w)
}
##### end of model #####
```

5. The Reduced Model: Multinomial Logistic Regression Without the Unnecessary Variables
hyper-g, beta-binomial case (multinomial2.txt)

```

model {
  for (i in 1:1599){                                     ## loop over observations
    y[i,1] ~ dcat(p[i, 1:6])                           ## multinomial outcome

    for (j in 1:5){                                     ## loop over categories
      p[i,j] <- expeta[i,j]/(1 + expeta[i,1] +
                                expeta[i,2] + expeta[i,3] +
                                expeta[i,4] + expeta[i,5])

      expeta[i,j] <- exp(eta[i,j])
    }

    eta[i,1] <- b01 + b11*x[i,1] + b21*x[i,2] + b51*x[i,5] + b61*x[i,6] +
                b71*x[i,7] + b91*x[i,9] + b111*x[i,11]
    eta[i,2] <- b02 + b22*x[i,2] + b42*x[i,4] + b72*x[i,7] + b92*x[i,9]
    eta[i,3] <- b03 + b23*x[i,2] + b53*x[i,5] + b63*x[i,6] + b73*x[i,7] +
                b103*x[i,10] + b113*x[i,11]
    eta[i,4] <- b04 + b14*x[i,1] + b24*x[i,2] + b44*x[i,4] + b54*x[i,5] +
                b64*x[i,6] + b74*x[i,7] + b84*x[i,8] + b104*x[i,10] +
                b114*x[i,11]
    eta[i,5] <- b05 + b15*x[i,1] + b35*x[i,3] + b45*x[i,4] + b55*x[i,5] +
                b65*x[i,6] + b75*x[i,7] + b95*x[i,9] + b105*x[i,10] +
                b115*x[i,11]

    expeta[i,6] <- 1.0                                    ## baseline (reference) category
    p[i,6] <- 1/(1 + expeta[i,1] +
                  expeta[i,2] + expeta[i,3] +
                  expeta[i,4] + expeta[i,5])
  }
}

## -----
##     Priors
## -----

b01 ~ dnorm(0.0, 1.0E-4)                               ## low-information priors
b02 ~ dnorm(0.0, 1.0E-4)
b03 ~ dnorm(0.0, 1.0E-4)
b04 ~ dnorm(0.0, 1.0E-4)
b05 ~ dnorm(0.0, 1.0E-4)
b11 ~ dnorm(0.0, 1.0E-4)
b14 ~ dnorm(0.0, 1.0E-4)
b15 ~ dnorm(0.0, 1.0E-4)
b21 ~ dnorm(0.0, 1.0E-4)
b22 ~ dnorm(0.0, 1.0E-4)
b23 ~ dnorm(0.0, 1.0E-4)
b24 ~ dnorm(0.0, 1.0E-4)
b35 ~ dnorm(0.0, 1.0E-4)
b42 ~ dnorm(0.0, 1.0E-4)
b44 ~ dnorm(0.0, 1.0E-4)
b45 ~ dnorm(0.0, 1.0E-4)
b51 ~ dnorm(0.0, 1.0E-4)

```

```

b53 ~ dnorm(0.0, 1.0E-4)
b54 ~ dnorm(0.0, 1.0E-4)
b55 ~ dnorm(0.0, 1.0E-4)
b61 ~ dnorm(0.0, 1.0E-4)
b63 ~ dnorm(0.0, 1.0E-4)
b64 ~ dnorm(0.0, 1.0E-4)
b65 ~ dnorm(0.0, 1.0E-4)
b71 ~ dnorm(0.0, 1.0E-4)
b72 ~ dnorm(0.0, 1.0E-4)
b73 ~ dnorm(0.0, 1.0E-4)
b74 ~ dnorm(0.0, 1.0E-4)
b75 ~ dnorm(0.0, 1.0E-4)
b84 ~ dnorm(0.0, 1.0E-4)
b91 ~ dnorm(0.0, 1.0E-4)
b92 ~ dnorm(0.0, 1.0E-4)
b95 ~ dnorm(0.0, 1.0E-4)
b103 ~ dnorm(0.0, 1.0E-4)
b104 ~ dnorm(0.0, 1.0E-4)
b105 ~ dnorm(0.0, 1.0E-4)
b111 ~ dnorm(0.0, 1.0E-4)
b113 ~ dnorm(0.0, 1.0E-4)
b114 ~ dnorm(0.0, 1.0E-4)
b115 ~ dnorm(0.0, 1.0E-4)

```

```

## -----
##      Exponentiated Values
## -----

```

```

B01 <- exp(b01)
B02 <- exp(b02)
B03 <- exp(b03)
B04 <- exp(b04)
B05 <- exp(b05)
B11 <- exp(b11)
B14 <- exp(b14)
B15 <- exp(b15)
B21 <- exp(b21)
B22 <- exp(b22)
B23 <- exp(b23)
B24 <- exp(b24)
B35 <- exp(b35)
B42 <- exp(b42)
B44 <- exp(b44)
B45 <- exp(b45)
B51 <- exp(b51)
B53 <- exp(b53)
B54 <- exp(b54)
B55 <- exp(b55)
B61 <- exp(b61)
B63 <- exp(b63)
B64 <- exp(b64)
B65 <- exp(b65)
B71 <- exp(b71)
B72 <- exp(b72)

```

```
B73 <- exp(b73)
B74 <- exp(b74)
B75 <- exp(b75)
B84 <- exp(b84)
B91 <- exp(b91)
B92 <- exp(b92)
B95 <- exp(b95)
B103 <- exp(b103)
B104 <- exp(b104)
B105 <- exp(b105)
B111 <- exp(b111)
B113 <- exp(b113)
B114 <- exp(b114)
B115 <- exp(b115)
}
#####
##### end of model #####
####
```

6. The Binomial Logistic Regression (binomial.txt)

```

model {
  for (i in 1:1599){
    y[i,1] ~ dbern(p[i])
    logit(p[i]) <- ystar[i]
    ystar[i] <- b[1] +
      b[2]*x[i,1] +
      b[3]*x[i,2] +
      b[4]*x[i,3] +
      b[5]*x[i,4] +
      b[6]*x[i,5] +
      b[7]*x[i,6] +
      b[8]*x[i,7] +
      b[9]*x[i,8] +
      b[10]*x[i,9] +
      b[11]*x[i,10] +
      b[12]*x[i,11]
  }
## -----
##     Priors
## -----
for (k in 1:12){
  b[k] ~ dnorm(0.0, 1.0E-4)      ## low information priors
}
#####
##### end of model #####

```

7. The Binomial Logistic Regression Without the Unnecessary Variables (binomial2.txt)

```
model {
  for (i in 1:1599){                      ## loop over observations
    y[i,1] ~ dbern(p[i])                  ## binary outcome
    logit(p[i]) <- ystar[i]               ## logit link
    ystar[i] <- b[1] +                    ## regression structure for covariates
          b[2]*x[i,2] +
          b[3]*x[i,5] +
          b[4]*x[i,6] +
          b[5]*x[i,7] +
          b[6]*x[i,10] +
          b[7]*x[i,11]
  }
## -----
##     Priors
## -----
for (k in 1:7){
  b[k] ~ dnorm(0.0, 1.0E-4)           ## low information priors
}

## -----
##   Exponentiated Values
## -----
for (k in 1:7){
  B[k] <- exp(b[k])
}
}

##### end of model #####
```

8. The R Code:

```
#####
# Initial Statistics & Plots for our Analysis.
#####
## Load the data into R.
wines <- read.csv("C:\\\\Users\\\\30697\\\\Desktop\\\\06_winequality-red.csv",
header = T, sep=";", stringsAsFactors = FALSE)

## Check for missing values (no missing values).
any(is.na(wines))

## Change the values from mg/dm^3 to g/dm^3 for free and total sulfur
## dioxide.
wines$free.sulfur.dioxide <- wines$free.sulfur.dioxide*0.001
wines$total.sulfur.dioxide <- wines$total.sulfur.dioxide*0.001

## Get some statistics on our parameters.
lapply(wines, mean)
lapply(wines, min)
lapply(wines, max)

## Plot of the sensory data.
plot(factor(wines$quality), xlab = "Sensory Preference (Median)", ylab =
"Frequency (Red Wine Samples)")

## Quality with 6 Categories (3-8).
table(wines$quality)

## Shift the response variable from 3 until 8, to 1 until 6, placing
## the third category (our reference category) to the last position.
wines[wines[,12]==3,12] <- 1
wines[wines[,12]==4,12] <- 2
wines[wines[,12]==6,12] <- 3
wines[wines[,12]==7,12] <- 4
wines[wines[,12]==5,12] <- 6 # our reference level
wines[wines[,12]==8,12] <- 5

# Used for BAS later.
wines.bas <- wines

## Check if it worked.
table(wines$quality)
plot(factor(wines$quality), xlab = "Sensory Preference (Median)", ylab =
"Frequency (Red Wine Samples)")

## Turn our data frame into a matrix and then into a list.
y <- as.matrix(wines$quality)
```

```

wines <- scale(wines[,-12], center=TRUE, scale=TRUE) # centering our
values (reduces run time by ~ 5-10 minutes)
wines1 <- as.matrix(wines)
lwinres <- list(y=y,x=wines1[,1:11])

#####
# Run our Model in R.
#####
## Load libraries.
library(R2WinBUGS)
library(BRugs)

## Set the directory.
openbugs.dir <- "D:\\OpenBUGS323"

## A list of initials (one chain).
inits1 <- list(
  list(b0=rep(0, times=5), b1=rep(0, times=5), b2=rep(0, times=5),
       b3=rep(0, times=5), b4=rep(0, times=5), b5=rep(0, times=5),
       b6=rep(0, times=5), b7=rep(0, times=5), b8=rep(0, times=5),
       b9=rep(0, times=5), b10=rep(0, times=5),
       b11=rep(0, times=5)) )

## The parameters of interest.
parameter.names <- c( 'b0', 'b1', 'b2', 'b3', 'b4', 'b5', 'b6', 'b7',
'b8', 'b9', 'b10', 'b11')

## The model (in OpenBUGS). (Run time for 1 chain ~ 85 minutes)
model1.sim <- bugs( lwinres, inits1, model.file =
"C:\\\\Users\\\\30697\\\\Desktop\\\\multinomial.txt", parameters =
parameter.names,
  n.chains = 1, n.iter = 31000, n.burnin=1000, n.thin=30,
bugs.directory = openbugs.dir, debug=F, program="OpenBUGS")

#####
# Check for convergence using plots and CODA
#####
library(coda)
temp<-as.mcmc.list(model1.sim)
summary(temp)
summary(temp)$statistics[, "Naive SE"]
summary(temp)$statistics[, "Time-series SE"]
batchSE(temp)
effectiveSize(temp)

```

```

## Ineffective size, some are < 200, but it's okay.

## 1000 samples (iterations) & 1 chain allow for the following test:
geweke.diag(temp)
heidel.diag(temp)

## Imperfect convergence.
pairs(wines.bas)
pairs(wines.bas[c(1,8,9)])
pairs(wines.bas[c(6,7)])


#-----
## Convergence plots for the OpenBUGS program.
#-----
## Ergodic Plots for betas.
#-----
# dev.off() # delete all previous plots
par(mfrow=c(3,3)) # 9 plots per picture

for (i in 1:5){
plot(cumsum(modell.sim$sims.matrix[,i])/1:length(modell.sim$sims.matri
x[,i]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b0",i)))
}

for (i in 6:10){
plot(cumsum(modell.sim$sims.matrix[,i])/1:length(modell.sim$sims.matri
x[,i]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b1",i-5)))
}

for (i in 11:15){
plot(cumsum(modell.sim$sims.matrix[,i])/1:length(modell.sim$sims.matri
x[,i]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b2",i-10)))
}

for (i in 16:20){
plot(cumsum(modell.sim$sims.matrix[,i])/1:length(modell.sim$sims.matri
x[,i]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b3",i-15)))
}

```

```

for (i in 21:25){
plot(cumsum(model1.sim$sims.matrix[,i])/1:length(model1.sim$sims.matri
x[,i]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b4",i-20)))
}

for (i in 26:30){
plot(cumsum(model1.sim$sims.matrix[,i])/1:length(model1.sim$sims.matri
x[,i]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b5",i-25)))
}

for (i in 31:35){
plot(cumsum(model1.sim$sims.matrix[,i])/1:length(model1.sim$sims.matri
x[,i]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b6",i-30)))
}

for (i in 36:40){
plot(cumsum(model1.sim$sims.matrix[,i])/1:length(model1.sim$sims.matri
x[,i]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b7",i-35)))
}

for (i in 41:45){
plot(cumsum(model1.sim$sims.matrix[,i])/1:length(model1.sim$sims.matri
x[,i]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b8",i-40)))
}

for (i in 46:50){
plot(cumsum(model1.sim$sims.matrix[,i])/1:length(model1.sim$sims.matri
x[,i]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b9",i-45)))
}

for (i in 51:55){
plot(cumsum(model1.sim$sims.matrix[,i])/1:length(model1.sim$sims.matri
x[,i]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b10",i-50)))
}

```

```

for (i in 56:60){
plot(cumsum(modell.sim$sims.matrix[,i])/1:length(modell.sim$sims.matrix[,i]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b0",i-55)))
}

plot(cumsum(modell.sim$sims.matrix[,61])/1:length(modell.sim$sims.matrix[,61]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for deviance")))

#-----
## Density Plots for betas.
#-----
# dev.off() # delete all previous plots
par(mfrow=c(3,3)) # 9 plots per picture

for (i in 1:5){
  plot(density(modell.sim$sims.matrix[,i]), main = paste(c("Density
plot for b0",i)))
}

for (i in 6:10){
  plot(density(modell.sim$sims.matrix[,i]), main = paste(c("Density
plot for b1",i-5)))
}

for (i in 11:15){
  plot(density(modell.sim$sims.matrix[,i]), main = paste(c("Density
plot for b2",i-10)))
}

for (i in 16:20){
  plot(density(modell.sim$sims.matrix[,i]), main = paste(c("Density
plot for b3",i-15)))
}

for (i in 21:25){
  plot(density(modell.sim$sims.matrix[,i]), main = paste(c("Density
plot for b4",i-20)))
}

for (i in 26:30){
  plot(density(modell.sim$sims.matrix[,i]), main = paste(c("Density
plot for b5",i-25)))
}

```

```

for (i in 31:35){
  plot(density(modell.sim$sims.matrix[,i]), main = paste(c("Density
plot for b6",i-30)))
}

for (i in 36:40){
  plot(density(modell.sim$sims.matrix[,i]), main = paste(c("Density
plot for b7",i-35)))
}

for (i in 41:45){
  plot(density(modell.sim$sims.matrix[,i]), main = paste(c("Density
plot for b8",i-40)))
}

for (i in 46:50){
  plot(density(modell.sim$sims.matrix[,i]), main = paste(c("Density
plot for b9",i-45)))
}

for (i in 51:55){
  plot(density(modell.sim$sims.matrix[,i]), main = paste(c("Density
plot for b10",i-50)))
}

for (i in 56:60){
  plot(density(modell.sim$sims.matrix[,i]), main = paste(c("Density
plot for b11",i-55)))
}

plot(density(modell.sim$sims.matrix[,61]), main = paste(c("Density
plot for deviance")))

#-----
## Trace Plots for betas.
#-----
# dev.off() # delete all previous plots
par(mfrow=c(3,3)) # 9 plots per picture

for (i in 1:5){
  plot(modell.sim$sims.matrix[,i], type = "l", main = paste(c("Trace
plot for b0",i)))
}

for (i in 6:10){
  plot(modell.sim$sims.matrix[,i], type = "l", main = paste(c("Trace
plot for b1",i-5)))
}

```

```
}

for (i in 11:15){
  plot(model1.sim$sims.matrix[,i], type = "l", main = paste(c("Trace
plot for b2",i-10)))
}

for (i in 16:20){
  plot(model1.sim$sims.matrix[,i], type = "l", main = paste(c("Trace
plot for b3",i-15)))
}

for (i in 21:25){
  plot(model1.sim$sims.matrix[,i], type = "l", main = paste(c("Trace
plot for b4",i-20)))
}

for (i in 26:30){
  plot(model1.sim$sims.matrix[,i], type = "l", main = paste(c("Trace
plot for b5",i-25)))
}

for (i in 31:35){
  plot(model1.sim$sims.matrix[,i], type = "l", main = paste(c("Trace
plot for b6",i-30)))
}

for (i in 36:40){
  plot(model1.sim$sims.matrix[,i], type = "l", main = paste(c("Trace
plot for b7",i-35)))
}

for (i in 41:45){
  plot(model1.sim$sims.matrix[,i], type = "l", main = paste(c("Trace
plot for b8",i-40)))
}

for (i in 46:50){
  plot(model1.sim$sims.matrix[,i], type = "l", main = paste(c("Trace
plot for b9",i-45)))
}

for (i in 51:55){
  plot(model1.sim$sims.matrix[,i], type = "l", main = paste(c("Trace
plot for b10",i-50)))
}

for (i in 56:60){
```

```

plot(model1.sim$sims.matrix[,i], type = "l", main = paste(c("Trace
plot for b11",i-55)))
}

plot(model1.sim$sims.matrix[,61], type = "l", main = paste(c("Trace
plot for deviance")))

#-----
## Autocorrelation Plots for betas.
#-----

# dev.off() # delete all previous plots
par(mfrow=c(3,3)) # 9 plots per picture

for (i in 1:5){
  acf(model1.sim$sims.matrix[,i], main = paste(c("ACF plot for
b0",i)))
}

for (i in 6:10){
  acf(model1.sim$sims.matrix[,i], main = paste(c("ACF plot for b1",i-
5)))
}

for (i in 11:15){
  acf(model1.sim$sims.matrix[,i], main = paste(c("ACF plot for b2",i-
10)))
}

for (i in 16:20){
  acf(model1.sim$sims.matrix[,i], main = paste(c("ACF plot for b3",i-
15)))
}

for (i in 21:25){
  acf(model1.sim$sims.matrix[,i], main = paste(c("ACF plot for b4",i-
20)))
}

for (i in 26:30){
  acf(model1.sim$sims.matrix[,i], main = paste(c("ACF plot for b5",i-
25)))
}

for (i in 31:35){
  acf(model1.sim$sims.matrix[,i], main = paste(c("ACF plot for b6",i-
30)))
}

```

```

for (i in 36:40){
  acf(model1.sim$sims.matrix[,i], main = paste(c("ACF plot for b7",i-
35)))
}

for (i in 41:45){
  acf(model1.sim$sims.matrix[,i], main = paste(c("ACF plot for b8",i-
40)))
}

for (i in 46:50){
  acf(model1.sim$sims.matrix[,i], main = paste(c("ACF plot for b9",i-
45)))
}

for (i in 51:55){
  acf(model1.sim$sims.matrix[,i], main = paste(c("ACF plot for
b10",i-50)))
}

for (i in 56:60){
  acf(model1.sim$sims.matrix[,i], main = paste(c("ACF plot for
b11",i-55)))
}

acf(model1.sim$sims.matrix[,i], main = paste(c("ACF plot for
deviance")))

#####
# Variable Selection using BAS.
=====
# install.packages("BAS")
library(BAS)

=====
# Preparing our data before running BAS
=====
# We will separate the multinomial problem
# into 5 binomial logistic regresions, in
# order to use the package BAS.
# Our reference category will be 6.
-----

## Seperate our categories and make the response a binomial.
## E.g: Category 6 = 0 (baseline), Category 1 = 1

```

```

## For category 1, with reference level 6.
category1.6 <- wines.bas[wines.bas[,12]==1|wines.bas[,12]==6,]
category1.6[category1.6[,12]==6,12] <- 0

## For category 2, with reference level 6.
category2.6 <- wines.bas[wines.bas[,12]==2|wines.bas[,12]==6,]
category2.6[category2.6[,12]==6,12] <- 0
category2.6[category2.6[,12]==2,12] <- 1

## For category 3, with reference level 6.
category3.6 <- wines.bas[wines.bas[,12]==3|wines.bas[,12]==6,]
category3.6[category3.6[,12]==6,12] <- 0
category3.6[category3.6[,12]==3,12] <- 1

## For category 4, with reference level 6.
category4.6 <- wines.bas[wines.bas[,12]==4|wines.bas[,12]==6,]
category4.6[category4.6[,12]==6,12] <- 0
category4.6[category4.6[,12]==4,12] <- 1

## For category 5, with reference level 6.
category5.6 <- wines.bas[wines.bas[,12]==5|wines.bas[,12]==6,]
category5.6[category5.6[,12]==6,12] <- 0
category5.6[category5.6[,12]==5,12] <- 1

## Check that our datasets have the proper values
# table(wines.bas$quality)
# table(category1.6$quality)
# table(category2.6$quality)
# table(category3.6$quality)
# table(category4.6$quality)
# table(category5.6$quality)

#=====
# Variable Selection through BIC, using BAS
#=====

## For category 1, with reference level 6.
##-----
n <- nrow(category1.6)

## Uniform Prior
#res1.1 <- bas.glm(quality~, family='binomial', data=category1.6,
betaprior=bic.prior(n), modelprior = uniform())

## Beta-Binomial Prior
res1.1 <- bas.glm(quality~, family='binomial', data=category1.6,
betaprior=bic.prior(n), modelprior = beta.binomial(2,10))

```

```

coef(res1.1)
summary(res1.1)

##-----
## For category 2, with reference level 6.
##-----
n <- nrow(category2.6)

## Uniform Prior
#res1.2 <- bas.glm(quality~, family='binomial', data=category2.6,
betaprior=bic.prior(n), modelprior = uniform())

## Beta-Binomial Prior
res1.2 <- bas.glm(quality~, family='binomial', data=category2.6,
betaprior=bic.prior(n), modelprior = beta.binomial(2,10))

coef(res1.2)
summary(res1.2)

##-----
## For category 3, with reference level 6.
##-----
n <- nrow(category3.6)

## Uniform Prior
#res1.3 <- bas.glm(quality~, family='binomial', data=category3.6,
betaprior=bic.prior(n), modelprior = uniform())

## Beta-Binomial Prior
res1.3 <- bas.glm(quality~, family='binomial', data=category3.6,
betaprior=bic.prior(n), modelprior = beta.binomial(2,10))

coef(res1.3)
summary(res1.3)

##-----
## For category 4, with reference level 6.
##-----
n <- nrow(category4.6)

## Uniform Prior
#res1.4 <- bas.glm(quality~, family='binomial', data=category4.6,
betaprior=bic.prior(n), modelprior = uniform())

## Beta-Binomial Prior
res1.4 <- bas.glm(quality~, family='binomial', data=category4.6,
betaprior=bic.prior(n), modelprior = beta.binomial(2,10))

```

```

coef(res1.4)
summary(res1.4)

##-----
## For category 5, with reference level 6.
##-----
n <- nrow(category5.6)

## Uniform Prior
#res1.5 <- bas.glm(quality~, family='binomial', data=category5.6,
betaprior=bic.prior(n), modelprior = uniform())

## Beta-Binomial Prior
res1.5 <- bas.glm(quality~, family='binomial', data=category5.6,
betaprior=bic.prior(n), modelprior = beta.binomial(2,10))

coef(res1.5)
summary(res1.5)

#=====
# Variable Selection through g-prior, using BAS
#=====

## For category 1, with reference level 6.
##-----
n <- nrow(category1.6)

## Uniform Prior
#res2.1 <- bas.glm(quality~, family='binomial', data=category1.6,
betaprior=g.prior(n), modelprior = uniform())

## Beta-Binomial Prior
res2.1 <- bas.glm(quality~, family='binomial', data=category1.6,
betaprior=g.prior(n), modelprior = beta.binomial(2,10))

coef(res2.1)
summary(res2.1)

##-----
## For category 2, with reference level 6.
##-----
n <- nrow(category2.6)

## Uniform Prior
#res2.2 <- bas.glm(quality~, family='binomial', data=category2.6,
betaprior=g.prior(n), modelprior = uniform())

## Beta-Binomial Prior

```

```

res2.2 <- bas.glm(quality~, family='binomial', data=category2.6,
betaprior=g.prior(n), modelprior = beta.binomial(2,10))

coef(res2.2)
summary(res2.2)

##-----
## For category 3, with reference level 6.
##-----
n <- nrow(category3.6)

## Uniform Prior
#res2.3 <- bas.glm(quality~, family='binomial', data=category3.6,
betaprior=g.prior(n), modelprior = uniform())

## Beta-Binomial Prior
res2.3 <- bas.glm(quality~, family='binomial', data=category3.6,
betaprior=g.prior(n), modelprior = beta.binomial(2,10))

coef(res2.3)
summary(res2.3)

##-----
## For category 4, with reference level 6.
##-----
n <- nrow(category4.6)

## Uniform Prior
#res2.4 <- bas.glm(quality~, family='binomial', data=category4.6,
betaprior=g.prior(n), modelprior = uniform())

## Beta-Binomial Prior
res2.4 <- bas.glm(quality~, family='binomial', data=category4.6,
betaprior=g.prior(n), modelprior = beta.binomial(2,10))

coef(res2.4)
summary(res2.4)

##-----
## For category 5, with reference level 6.
##-----
n <- nrow(category5.6)

## Uniform Prior
#res2.5 <- bas.glm(quality~, family='binomial', data=category5.6,
betaprior=g.prior(n), modelprior = uniform())

## Beta-Binomial Prior

```

```

res2.5 <- bas.glm(quality~, family='binomial', data=category5.6,
betaprior=g.prior(n), modelprior = beta.binomial(2,10))

coef(res2.5)
summary(res2.5)

#=====
# Variable Selection through hyper-g, using BAS
#=====

## Here alpha = 3.

##-----
## For category 1, with reference level 6.
##-----

## Uniform Prior
#res3.1 <- bas.glm(quality~, family='binomial', data=category1.6,
betaprior=hyper.g(3), modelprior = uniform())

## Beta-Binomial
res3.1 <- bas.glm(quality~, family='binomial', data=category1.6,
betaprior=hyper.g(3), modelprior = beta.binomial(2,10))

coef(res3.1)
summary(res3.1)

##-----
## For category 2, with reference level 6.
##-----

## Uniform Prior
#res3.2 <- bas.glm(quality~, family='binomial', data=category2.6,
betaprior=hyper.g(3), modelprior = uniform())

## Beta-Binomial
res3.2 <- bas.glm(quality~, family='binomial', data=category2.6,
betaprior=hyper.g(3), modelprior = beta.binomial(2,10))

coef(res3.2)
summary(res3.2)

##-----
## For category 3, with reference level 6.
##-----

## Uniform Prior
#res3.3 <- bas.glm(quality~, family='binomial', data=category3.6,
betaprior=hyper.g(3), modelprior = uniform())

```

```

## Beta-Binomial
res3.3 <- bas.glm(quality~, family='binomial', data=category3.6,
betaprior=hyper.g(3), modelprior = beta.binomial(2,10))

coef(res3.3)
summary(res3.3)

#-----#
## For category 4, with reference level 6.
#-----#
## Uniform Prior
#res3.4 <- bas.glm(quality~, family='binomial', data=category4.6,
betaprior=hyper.g(3), modelprior = uniform())

## Beta-Binomial
res3.4 <- bas.glm(quality~, family='binomial', data=category4.6,
betaprior=hyper.g(3), modelprior = beta.binomial(2,10))

coef(res3.4)
summary(res3.4)

#-----#
## For category 5, with reference level 6.
#-----#
## Uniform Prior
#res3.5 <- bas.glm(quality~, family='binomial', data=category5.6,
betaprior=hyper.g(3), modelprior = uniform())

## Beta-Binomial
res3.5 <- bas.glm(quality~, family='binomial', data=category5.6,
betaprior=hyper.g(3), modelprior = beta.binomial(2,10))

coef(res3.5)
summary(res3.5)

#####
# Variable Selection using OpenBUGS.
=====
# Preparing our data to run the Empirical Bayes model in OpenBUGS.
=====
## Load libraries.
library(R2WinBUGS)
library(BRugs)

```

```

## Set the directory.
openbugs.dir <- "D:\\OpenBUGS323"

## Posterior Mean Estimates
prop.mean.beta0 <- c(-10, -3, 0.2, -2, -6.6)
prop.mean.beta1 <- c(1.1, 1, 0.2, 0.7, -0.5)
prop.mean.beta2 <- c(1.8, 0.5, -0.5, -0.8, -0.1)
prop.mean.beta3 <- c(0.8, 0, -0.3, -0.1, 0.5)
prop.mean.beta4 <- c(-0.1, 0.6, 0.1, 0.4, 0)
prop.mean.beta5 <- c(0.7, 0.1, -0.1, -0.5, -1.9)
prop.mean.beta6 <- c(2, -0.3, 0.2, 0.2, 0.4)
prop.mean.beta7 <- c(-4.6, -0.5, -0.5, -0.9, -1.7)
prop.mean.beta8 <- c(0.7, -1.1, -0.1, -0.6, -0.2)
prop.mean.beta9 <- c(1.9, 0.8, 0, 0.1, -1)
prop.mean.beta10 <- c(-0.6, 0.1, 0.4, 0.9, 1.2)
prop.mean.beta11 <- c(-1.6, -0.3, 0.8, 1.4, 2.3)

## Posterior Standard Deviation Estimates
prop.sd.beta0 <- c(1.7, 0.2, 0.1, 0.2, 0.8)
prop.sd.beta1 <- c(1.3, 0.5, 0.2, 0.2, 0.7)
prop.sd.beta2 <- c(0.5, 0.2, 0.1, 0.2, 0.4)
prop.sd.beta3 <- c(0.9, 0.3, 0.1, 0.2, 0.5)
prop.sd.beta4 <- c(0.5, 0.2, 0.1, 0.1, 0.4)
prop.sd.beta5 <- c(0.4, 0.2, 0.1, 0.2, 0.7)
prop.sd.beta6 <- c(1.1, 0.3, 0.1, 0.1, 0.4)
prop.sd.beta7 <- c(1.9, 0.3, 0.1, 0.2, 0.5)
prop.sd.beta8 <- c(1.1, 0.4, 0.2, 0.2, 0.6)
prop.sd.beta9 <- c(0.8, 0.3, 0.1, 0.2, 0.5)
prop.sd.beta10 <- c(0.7, 0.2, 0.1, 0.1, 0.3)
prop.sd.beta11 <- c(1, 0.3, 0.1, 0.2, 0.5)

## A new list for our data.
lwines1 <- list(y=y, x=wines1[,1:11], prop.mean.beta0=prop.mean.beta0,
                 prop.mean.beta1=prop.mean.beta1,
                 prop.mean.beta2=prop.mean.beta2,
                 prop.mean.beta3=prop.mean.beta3,
                 prop.mean.beta4=prop.mean.beta4,
                 prop.mean.beta5=prop.mean.beta5,
                 prop.mean.beta6=prop.mean.beta6,
                 prop.mean.beta7=prop.mean.beta7,
                 prop.mean.beta8=prop.mean.beta8,
                 prop.mean.beta9=prop.mean.beta9,
                 prop.mean.beta10=prop.mean.beta10,
                 prop.mean.beta11=prop.mean.beta11,
                 prop.sd.beta0=prop.sd.beta0, prop.sd.beta1=prop.sd.beta1,
                 prop.sd.beta2=prop.sd.beta2, prop.sd.beta3=prop.sd.beta3,
                 prop.sd.beta4=prop.sd.beta4, prop.sd.beta5=prop.sd.beta5,

```

```

prop.sd.beta6=prop.sd.beta6, prop.sd.beta7=prop.sd.beta7,
prop.sd.beta8=prop.sd.beta8, prop.sd.beta9=prop.sd.beta9,
prop.sd.beta10=prop.sd.beta10,
prop.sd.beta11=prop.sd.beta11)

## A list of initials (one chain).
inits2 <- list(
  list(b0=rep(0, times=5), b1=rep(0, times=5), b2=rep(0, times=5),
       b3=rep(0, times=5), b4=rep(0, times=5),
       b5=rep(0, times=5), b6=rep(0, times=5),
       b7=rep(0, times=5), b8=rep(0, times=5),
       b9=rep(0, times=5), b10=rep(0, times=5),
       b11=rep(0, times=5),
       gamma1=rep(0, times=5),
       gamma2=rep(0, times=5), gamma3=rep(0, times=5),
       gamma4=rep(0, times=5), gamma5=rep(0, times=5),
       gamma6=rep(0, times=5), gamma7=rep(0, times=5),
       gamma8=rep(0, times=5), gamma9=rep(0, times=5),
       gamma10=rep(0, times=5), gamma11=rep(0, times=5) #,
#       p1=0.5, p2=0.5, p3=0.5, p4=0.5, p5=0.5, p6=0.5,
#       p7=0.5, p8=0.5, p9=0.5, p10=0.5, p11=0.5
    ) )

## The parameters of interest.
parameter.names1 <- c( 'gb0', 'gb1', 'gb2', 'gb3', 'gb4', 'gb5',
'gb6', 'gb7', 'gb8', 'gb9', 'gb10', 'gb11')

## The model for the empirical Bayes method (in OpenBUGS). (Run time ~
## 3:30 hours)
model2.sim <- bugs( lwinres1, inits2, model.file =
"C:\\\\Users\\\\30697\\\\Desktop\\\\empirical.txt", parameters =
parameter.names1,
  n.chains = 1, n.iter = 11000, n.burnin=1000, n.thin=30,
bugs.directory = openbugs.dir, debug=F, program="OpenBUGS")

#####
## Density Plots for betas.
#####

# dev.off() # delete all previous plots
par(mfrow=c(3,3)) # 9 plots per picture

for (i in 1:5){
  plot(density(model2.sim$sims.matrix[,i]), main = paste(c("Density
plot for b0",i)))
}

for (i in 6:10){

```

```

    plot(density(model2.sim$sims.matrix[,i]), main = paste(c("Density
plot for b1",i-5)))
}

for (i in 11:15){
    plot(density(model2.sim$sims.matrix[,i]), main = paste(c("Density
plot for b2",i-10)))
}

for (i in 16:20){
    plot(density(model2.sim$sims.matrix[,i]), main = paste(c("Density
plot for b3",i-15)))
}

for (i in 21:25){
    plot(density(model2.sim$sims.matrix[,i]), main = paste(c("Density
plot for b4",i-20)))
}

for (i in 26:30){
    plot(density(model2.sim$sims.matrix[,i]), main = paste(c("Density
plot for b5",i-25)))
}

for (i in 31:35){
    plot(density(model2.sim$sims.matrix[,i]), main = paste(c("Density
plot for b6",i-30)))
}

for (i in 36:40){
    plot(density(model2.sim$sims.matrix[,i]), main = paste(c("Density
plot for b7",i-35)))
}

for (i in 41:45){
    plot(density(model2.sim$sims.matrix[,i]), main = paste(c("Density
plot for b8",i-40)))
}

for (i in 46:50){
    plot(density(model2.sim$sims.matrix[,i]), main = paste(c("Density
plot for b9",i-45)))
}

for (i in 51:55){
    plot(density(model2.sim$sims.matrix[,i]), main = paste(c("Density
plot for b10",i-50)))
}

```

```

for (i in 56:60){
  plot(density(model2.sim$sims.matrix[,i]), main = paste(c("Density
plot for b11",i-55)))
}

plot(density(model2.sim$sims.matrix[,61]), main = paste(c("Density
plot for deviance")))

#=====
# Preparing our data to run the g-prior (of Liang et al.) model in
# OpenBUGS.
#=====

## Load libraries.
library(R2WinBUGS)
library(BRugs)

## Set the directory.
openbugs.dir <- "D:\\OpenBUGS323"

## Posterior Mean Estimates
prop.mean.beta1 <- c(1.1, 1, 0.2, 0.7, -0.5)
prop.mean.beta2 <- c(1.8, 0.5, -0.5, -0.8, -0.1)
prop.mean.beta3 <- c(0.8, 0, -0.3, -0.1, 0.5)
prop.mean.beta4 <- c(-0.1, 0.6, 0.1, 0.4, 0)
prop.mean.beta5 <- c(0.7, 0.1, -0.1, -0.5, -1.9)
prop.mean.beta6 <- c(2, -0.3, 0.2, 0.2, 0.4)
prop.mean.beta7 <- c(-4.6, -0.5, -0.5, -0.9, -1.7)
prop.mean.beta8 <- c(0.7, -1.1, -0.1, -0.6, -0.2)
prop.mean.beta9 <- c(1.9, 0.8, 0, 0.1, -1)
prop.mean.beta10 <- c(-0.6, 0.1, 0.4, 0.9, 1.2)
prop.mean.beta11 <- c(-1.6, -0.3, 0.8, 1.4, 2.3)

## Posterior Standard Deviation Estimates
prop.sd.beta1 <- c(1.3, 0.5, 0.2, 0.2, 0.7)
prop.sd.beta2 <- c(0.5, 0.2, 0.1, 0.2, 0.4)
prop.sd.beta3 <- c(0.9, 0.3, 0.1, 0.2, 0.5)
prop.sd.beta4 <- c(0.5, 0.2, 0.1, 0.1, 0.4)
prop.sd.beta5 <- c(0.4, 0.2, 0.1, 0.2, 0.7)
prop.sd.beta6 <- c(1.1, 0.3, 0.1, 0.1, 0.4)
prop.sd.beta7 <- c(1.9, 0.3, 0.1, 0.2, 0.5)
prop.sd.beta8 <- c(1.1, 0.4, 0.2, 0.2, 0.6)
prop.sd.beta9 <- c(0.8, 0.3, 0.1, 0.2, 0.5)
prop.sd.beta10 <- c(0.7, 0.2, 0.1, 0.1, 0.3)
prop.sd.beta11 <- c(1, 0.3, 0.1, 0.2, 0.5)

## The wines data are already centered, so

```

```

## here we only calculate the XTX matrix
XTX <- t(wines1) %*% wines1

## A new list for our data.
lwinres2 <- list(y=y, XTX=XTX, x=wines1[,1:11],
                  prop.mean.beta1=prop.mean.beta1,
                  prop.mean.beta2=prop.mean.beta2,
                  prop.mean.beta3=prop.mean.beta3,
                  prop.mean.beta4=prop.mean.beta4,
                  prop.mean.beta5=prop.mean.beta5,
                  prop.mean.beta6=prop.mean.beta6,
                  prop.mean.beta7=prop.mean.beta7,
                  prop.mean.beta8=prop.mean.beta8,
                  prop.mean.beta9=prop.mean.beta9,
                  prop.mean.beta10=prop.mean.beta10,
                  prop.mean.beta11=prop.mean.beta11,
                  prop.sd.beta1=prop.sd.beta1,
                  prop.sd.beta2=prop.sd.beta2, prop.sd.beta3=prop.sd.beta3,
                  prop.sd.beta4=prop.sd.beta4, prop.sd.beta5=prop.sd.beta5,
                  prop.sd.beta6=prop.sd.beta6, prop.sd.beta7=prop.sd.beta7,
                  prop.sd.beta8=prop.sd.beta8, prop.sd.beta9=prop.sd.beta9,
                  prop.sd.beta10=prop.sd.beta10,
                  prop.sd.beta11=prop.sd.beta11)

## A list of initials (one chain).
inits3 <- list(
  list(alpha=rep(0.6, times=5), B1=rep(0, times=5),
       B2=rep(0, times=5), B3=rep(0, times=5), B4=rep(0, times=5),
       B5=rep(0, times=5), B6=rep(0, times=5), B7=rep(0, times=5),
       B8=rep(0, times=5), B9=rep(0, times=5),
       B10=rep(0, times=5), B11=rep(0, times=5), logtau=0,
       gamma1=rep(0, times=5),
       gamma2=rep(0, times=5), gamma3=rep(0, times=5),
       gamma4=rep(0, times=5), gamma5=rep(0, times=5),
       gamma6=rep(0, times=5), gamma7=rep(0, times=5),
       gamma8=rep(0, times=5), gamma9=rep(0, times=5),
       gamma10=rep(0, times=5), gamma11=rep(0, times=5) #,
       p1=0.5, p2=0.5, p3=0.5, p4=0.5, p5=0.5, p6=0.5,
       p7=0.5, p8=0.5, p9=0.5, p10=0.5, p11=0.5
      ) )

## The parameters of interest.
parameter.names2 <- c('alpha', 'gb1', 'gb2', 'gb3', 'gb4', 'gb5',
                      'gb6', 'gb7', 'gb8', 'gb9', 'gb10', 'gb11')

## The model for the gprior method (in OpenBUGS). (WARNING!!! Run time
~ 13 hours)

```

```

model3.sim <- bugs( lwines2, inits3, model.file =
"C:\\\\Users\\\\30697\\\\Desktop\\\\gprior.txt", parameters =
parameter.names2,
      n.chains = 1, n.iter = 11000, n.burnin=1000, n.thin=30,
bugs.directory = openbugs.dir, debug=F, program="OpenBUGS")

#####
## Density Plots for betas.
#####

# dev.off() # delete all previous plots
par(mfrow=c(3,3)) # 9 plots per picture

for (i in 1:5){
  plot(density(model3.sim$sims.matrix[,i]), main = paste(c("Density
plot for b0",i)))
}

for (i in 6:10){
  plot(density(model3.sim$sims.matrix[,i]), main = paste(c("Density
plot for b1",i-5)))
}

for (i in 11:15){
  plot(density(model3.sim$sims.matrix[,i]), main = paste(c("Density
plot for b2",i-10)))
}

for (i in 16:20){
  plot(density(model3.sim$sims.matrix[,i]), main = paste(c("Density
plot for b3",i-15)))
}

for (i in 21:25){
  plot(density(model3.sim$sims.matrix[,i]), main = paste(c("Density
plot for b4",i-20)))
}

for (i in 26:30){
  plot(density(model3.sim$sims.matrix[,i]), main = paste(c("Density
plot for b5",i-25)))
}

for (i in 31:35){
  plot(density(model3.sim$sims.matrix[,i]), main = paste(c("Density
plot for b6",i-30)))
}

for (i in 36:40){

```

```

    plot(density(model3.sim$sims.matrix[,i]), main = paste(c("Density
plot for b7",i-35)))
}

for (i in 41:45){
    plot(density(model3.sim$sims.matrix[,i]), main = paste(c("Density
plot for b8",i-40)))
}

for (i in 46:50){
    plot(density(model3.sim$sims.matrix[,i]), main = paste(c("Density
plot for b9",i-45)))
}

for (i in 51:55){
    plot(density(model3.sim$sims.matrix[,i]), main = paste(c("Density
plot for b10",i-50)))
}

for (i in 56:60){
    plot(density(model3.sim$sims.matrix[,i]), main = paste(c("Density
plot for b11",i-55)))
}

plot(density(model3.sim$sims.matrix[,61]), main = paste(c("Density
plot for deviance")))

#=====
# Preparing our data to run the hyper-g (of Liang et al.) model in
# OpenBUGS.
#=====

## Load libraries.
library(R2WinBUGS)
library(BRugs)

## Set the directory.
openbugs.dir <- "D:\\OpenBUGS323"

## Posterior Mean Estimates
prop.mean.beta1 <- c(1.1, 1, 0.2, 0.7, -0.5)
prop.mean.beta2 <- c(1.8, 0.5, -0.5, -0.8, -0.1)
prop.mean.beta3 <- c(0.8, 0, -0.3, -0.1, 0.5)

```

```

prop.mean.beta4 <- c(-0.1, 0.6, 0.1, 0.4, 0)
prop.mean.beta5 <- c(0.7, 0.1, -0.1, -0.5, -1.9)
prop.mean.beta6 <- c(2, -0.3, 0.2, 0.2, 0.4)
prop.mean.beta7 <- c(-4.6, -0.5, -0.5, -0.9, -1.7)
prop.mean.beta8 <- c(0.7, -1.1, -0.1, -0.6, -0.2)
prop.mean.beta9 <- c(1.9, 0.8, 0, 0.1, -1)
prop.mean.beta10 <- c(-0.6, 0.1, 0.4, 0.9, 1.2)
prop.mean.beta11 <- c(-1.6, -0.3, 0.8, 1.4, 2.3)

## Posterior Standard Deviation Estimates
prop.sd.betal <- c(1.3, 0.5, 0.2, 0.2, 0.7)
prop.sd.beta2 <- c(0.5, 0.2, 0.1, 0.2, 0.4)
prop.sd.beta3 <- c(0.9, 0.3, 0.1, 0.2, 0.5)
prop.sd.beta4 <- c(0.5, 0.2, 0.1, 0.1, 0.4)
prop.sd.beta5 <- c(0.4, 0.2, 0.1, 0.2, 0.7)
prop.sd.beta6 <- c(1.1, 0.3, 0.1, 0.1, 0.4)
prop.sd.beta7 <- c(1.9, 0.3, 0.1, 0.2, 0.5)
prop.sd.beta8 <- c(1.1, 0.4, 0.2, 0.2, 0.6)
prop.sd.beta9 <- c(0.8, 0.3, 0.1, 0.2, 0.5)
prop.sd.beta10 <- c(0.7, 0.2, 0.1, 0.1, 0.3)
prop.sd.beta11 <- c(1, 0.3, 0.1, 0.2, 0.5)

## The wines data are already centered, so
## here we only calculate the XTX matrix
XTX <- t(wines1) %*% wines1

## A new list for our data.
lwines2 <- list(y=y, XTX=XTX, x=wines1[,1:11],
                 prop.mean.betal=prop.mean.betal,
                 prop.mean.beta2=prop.mean.beta2,
                 prop.mean.beta3=prop.mean.beta3,
                 prop.mean.beta4=prop.mean.beta4,
                 prop.mean.beta5=prop.mean.beta5,
                 prop.mean.beta6=prop.mean.beta6,
                 prop.mean.beta7=prop.mean.beta7,
                 prop.mean.beta8=prop.mean.beta8,
                 prop.mean.beta9=prop.mean.beta9,
                 prop.mean.beta10=prop.mean.beta10,
                 prop.mean.beta11=prop.mean.beta11,
                 prop.sd.betal=prop.sd.betal,
                 prop.sd.beta2=prop.sd.beta2, prop.sd.beta3=prop.sd.beta3,
                 prop.sd.beta4=prop.sd.beta4, prop.sd.beta5=prop.sd.beta5,
                 prop.sd.beta6=prop.sd.beta6, prop.sd.beta7=prop.sd.beta7,
                 prop.sd.beta8=prop.sd.beta8, prop.sd.beta9=prop.sd.beta9,
                 prop.sd.beta10=prop.sd.beta10,
                 prop.sd.beta11=prop.sd.beta11)

## A list of initials (one chain).

```

```

inits4 <- list(
  list(alpha=rep(0.6, times=5), B1=rep(0, times=5),
       B2=rep(0, times=5), B3=rep(0, times=5),
       B4=rep(0, times=5), B5=rep(0, times=5),
       B6=rep(0, times=5), B7=rep(0, times=5),
       B8=rep(0, times=5), B9=rep(0, times=5),
       B10=rep(0, times=5), B11=rep(0, times=5),
       logtau=0, w=0.7,
       gamma1=rep(0, times=5),
       gamma2=rep(0, times=5), gamma3=rep(0, times=5),
       gamma4=rep(0, times=5), gamma5=rep(0, times=5),
       gamma6=rep(0, times=5), gamma7=rep(0, times=5),
       gamma8=rep(0, times=5), gamma9=rep(0, times=5),
       gamma10=rep(0, times=5), gamma11=rep(0, times=5) #,
#       p1=0.5, p2=0.5, p3=0.5, p4=0.5, p5=0.5, p6=0.5,
#       p7=0.5, p8=0.5, p9=0.5, p10=0.5, p11=0.5
    ) )

## The parameters of interest.
parameter.names2 <- c( 'alpha', 'gb1', 'gb2', 'gb3', 'gb4', 'gb5',
'gb6', 'gb7', 'gb8', 'gb9', 'gb10', 'gb11')

## The model for the gprior method (in OpenBUGS). (WARNING!!! Run time
~ 13 hours)
model4.sim <- bugs( lwinex2, inits4, model.file =
"C:\\\\Users\\\\30697\\\\Desktop\\\\hyperg.txt", parameters =
parameter.names2,
  n.chains = 1, n.iter = 11000, n.burnin=1000, n.thin=30,
bugs.directory = openbugs.dir, debug=F, program="OpenBUGS")

#-----
## Density Plots for betas.
#-----

# dev.off() # delete all previous plots
par(mfrow=c(3,3)) # 9 plots per picture

for (i in 1:5){
  plot(density(model4.sim$sims.matrix[,i]), main = paste(c("Density
plot for b0",i)))
}

for (i in 6:10){
  plot(density(model4.sim$sims.matrix[,i]), main = paste(c("Density
plot for b1",i-5)))
}

for (i in 11:15){

```

```

    plot(density(model4.sim$sims.matrix[,i]), main = paste(c("Density
plot for b2",i-10)))
}

for (i in 16:20){
    plot(density(model4.sim$sims.matrix[,i]), main = paste(c("Density
plot for b3",i-15)))
}

for (i in 21:25){
    plot(density(model4.sim$sims.matrix[,i]), main = paste(c("Density
plot for b4",i-20)))
}

for (i in 26:30){
    plot(density(model4.sim$sims.matrix[,i]), main = paste(c("Density
plot for b5",i-25)))
}

for (i in 31:35){
    plot(density(model4.sim$sims.matrix[,i]), main = paste(c("Density
plot for b6",i-30)))
}

for (i in 36:40){
    plot(density(model4.sim$sims.matrix[,i]), main = paste(c("Density
plot for b7",i-35)))
}

for (i in 41:45){
    plot(density(model4.sim$sims.matrix[,i]), main = paste(c("Density
plot for b8",i-40)))
}

for (i in 46:50){
    plot(density(model4.sim$sims.matrix[,i]), main = paste(c("Density
plot for b9",i-45)))
}

for (i in 51:55){
    plot(density(model4.sim$sims.matrix[,i]), main = paste(c("Density
plot for b10",i-50)))
}

for (i in 56:60){
    plot(density(model4.sim$sims.matrix[,i]), main = paste(c("Density
plot for b11",i-55)))
}

```

```

plot(density(model4.sim$sims.matrix[,61]), main = paste(c("Density
plot for deviance")))

#####
# Run our simpler (reduced) Model in R.
#####
#####

## Load libraries.
library(R2WinBUGS)
library(BRugs)

## Set the directory.
openbugs.dir <- "D:\\OpenBUGS323"

## A list of initials (one chain).
inits5 <- list(
  list(b01=0, b02=0, b03=0, b04=0, b05=0, b11=0, b14=0, b15=0,
       b21=0, b22=0, b23=0, b24=0, b35=0, b42=0, b44=0, b45=0,
       b51=0, b53=0, b54=0, b55=0, b61=0, b63=0, b64=0, b65=0,
       b71=0, b72=0, b73=0, b74=0, b75=0, b84=0, b91=0, b92=0,
       b95=0, b103=0, b104=0, b105=0, b111=0, b113=0, b114=0,
       b115=0) )

## The parameters of interest.
parameter.names3 <- c( 'b01', 'b02', 'b03', 'b04', 'b05', 'b11',
                       'b14', 'b15', 'b21', 'b22', 'b23', 'b24',
                       'b35', 'b42', 'b44', 'b45', 'b51', 'b53',
                       'b54', 'b55', 'b61', 'b63', 'b64', 'b65',
                       'b71', 'b72', 'b73', 'b74', 'b75', 'b84',
                       'b91', 'b92', 'b95', 'b103', 'b104',
                       'b105', 'b111', 'b113', 'b114', 'b115',
                       'B01', 'B02', 'B03', 'B04', 'B05', 'B11',
                       'B14', 'B15', 'B21', 'B22', 'B23', 'B24',
                       'B35', 'B42', 'B44', 'B45', 'B51', 'B53',
                       'B54', 'B55', 'B61', 'B63', 'B64', 'B65',
                       'B71', 'B72', 'B73', 'B74', 'B75', 'B84',
                       'B91', 'B92', 'B95', 'B103', 'B104',
                       'B105', 'B111', 'B113', 'B114', 'B115')

## The model (in OpenBUGS). (Run time for 1 chain ~ 50 minutes)
model5.sim <- bugs( lwinis, inits5, model.file =
"C:\\\\Users\\\\30697\\\\Desktop\\\\multinomial2.txt", parameters =
parameter.names3,

```

```

n.chains = 1, n.iter = 51000, n.burnin=1000, n.thin=50,
bugs.directory = openbugs.dir, debug=F, program="OpenBUGS")

#####
# Check for convergence using plots and CODA
#####
=====

library(coda)
temp<-as.mcmc.list(model5.sim)
summary(temp)
summary(temp)$statistics[, "Naive SE"]
summary(temp)$statistics[, "Time-series SE"]
batchSE(temp)
effectiveSize(temp)

## 1000 samples (iterations) allow for the following tests:
geweke.diag(temp)
heidel.diag(temp)

#####
## Ergodic Plots for betas.
#####
par(mfrow=c(3,3)) # 9 plots per picture

plot(cumsum(model5.sim$sims.matrix[,1])/1:length(model5.sim$sims.matrix[,1]),
      type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b01")))
plot(cumsum(model5.sim$sims.matrix[,2])/1:length(model5.sim$sims.matrix[,2]),
      type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b02")))
plot(cumsum(model5.sim$sims.matrix[,3])/1:length(model5.sim$sims.matrix[,3]),
      type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b03")))
plot(cumsum(model5.sim$sims.matrix[,4])/1:length(model5.sim$sims.matrix[,4]),
      type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b04")))
plot(cumsum(model5.sim$sims.matrix[,5])/1:length(model5.sim$sims.matrix[,5]),
      type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b05")))
plot(cumsum(model5.sim$sims.matrix[,6])/1:length(model5.sim$sims.matrix[,6]),
      type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b06")))

```

```

type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b11")))
plot(cumsum(model5.sim$sims.matrix[,7])/1:length(model5.sim$sims.matri
x[,7]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b14")))
plot(cumsum(model5.sim$sims.matrix[,8])/1:length(model5.sim$sims.matri
x[,8]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b15")))
plot(cumsum(model5.sim$sims.matrix[,9])/1:length(model5.sim$sims.matri
x[,9]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b21")))
plot(cumsum(model5.sim$sims.matrix[,10])/1:length(model5.sim$sims.matri
x[,10]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b22")))
plot(cumsum(model5.sim$sims.matrix[,11])/1:length(model5.sim$sims.matri
x[,11]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b23")))
plot(cumsum(model5.sim$sims.matrix[,12])/1:length(model5.sim$sims.matri
x[,12]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b24")))
plot(cumsum(model5.sim$sims.matrix[,13])/1:length(model5.sim$sims.matri
x[,13]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b35")))
plot(cumsum(model5.sim$sims.matrix[,14])/1:length(model5.sim$sims.matri
x[,14]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b42")))
plot(cumsum(model5.sim$sims.matrix[,15])/1:length(model5.sim$sims.matri
x[,15]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b44")))
plot(cumsum(model5.sim$sims.matrix[,16])/1:length(model5.sim$sims.matri
x[,16]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b45")))
plot(cumsum(model5.sim$sims.matrix[,17])/1:length(model5.sim$sims.matri
x[,17]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b51")))
plot(cumsum(model5.sim$sims.matrix[,18])/1:length(model5.sim$sims.matri
x[,18]),

```

```

type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b53")))
plot(cumsum(model5.sim$sims.matrix[,19])/1:length(model5.sim$sims.matr
ix[,19]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b54")))
plot(cumsum(model5.sim$sims.matrix[,20])/1:length(model5.sim$sims.matr
ix[,20]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b55")))
plot(cumsum(model5.sim$sims.matrix[,21])/1:length(model5.sim$sims.matr
ix[,21]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b61")))
plot(cumsum(model5.sim$sims.matrix[,22])/1:length(model5.sim$sims.matr
ix[,22]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b63")))
plot(cumsum(model5.sim$sims.matrix[,23])/1:length(model5.sim$sims.matr
ix[,23]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b64")))
plot(cumsum(model5.sim$sims.matrix[,24])/1:length(model5.sim$sims.matr
ix[,24]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b65")))
plot(cumsum(model5.sim$sims.matrix[,25])/1:length(model5.sim$sims.matr
ix[,25]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b71")))
plot(cumsum(model5.sim$sims.matrix[,26])/1:length(model5.sim$sims.matr
ix[,26]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b72")))
plot(cumsum(model5.sim$sims.matrix[,27])/1:length(model5.sim$sims.matr
ix[,27]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b73")))
plot(cumsum(model5.sim$sims.matrix[,28])/1:length(model5.sim$sims.matr
ix[,28]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b74")))
plot(cumsum(model5.sim$sims.matrix[,29])/1:length(model5.sim$sims.matr
ix[,29]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b75")))
plot(cumsum(model5.sim$sims.matrix[,30])/1:length(model5.sim$sims.matr
ix[,30]),

```

```

type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b84")))
plot(cumsum(model5.sim$sims.matrix[,31])/1:length(model5.sim$sims.matr
ix[,31]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b91")))
plot(cumsum(model5.sim$sims.matrix[,32])/1:length(model5.sim$sims.matr
ix[,32]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b92")))
plot(cumsum(model5.sim$sims.matrix[,33])/1:length(model5.sim$sims.matr
ix[,33]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b95")))
plot(cumsum(model5.sim$sims.matrix[,34])/1:length(model5.sim$sims.matr
ix[,34]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b103")))
plot(cumsum(model5.sim$sims.matrix[,35])/1:length(model5.sim$sims.matr
ix[,35]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b104")))
plot(cumsum(model5.sim$sims.matrix[,36])/1:length(model5.sim$sims.matr
ix[,36]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b105")))
plot(cumsum(model5.sim$sims.matrix[,37])/1:length(model5.sim$sims.matr
ix[,37]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b111")))
plot(cumsum(model5.sim$sims.matrix[,38])/1:length(model5.sim$sims.matr
ix[,38]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b113")))
plot(cumsum(model5.sim$sims.matrix[,39])/1:length(model5.sim$sims.matr
ix[,39]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b114")))
plot(cumsum(model5.sim$sims.matrix[,40])/1:length(model5.sim$sims.matr
ix[,40]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for b115")))
plot(cumsum(model5.sim$sims.matrix[,81])/1:length(model5.sim$sims.matr
ix[,81]),
type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean Plot
for deviance")))

```

```

##-----#
## Density Plots for betas.
##-----#
# dev.off() # delete all previous plots
par(mfrow=c(3,3)) # 9 plots per picture

plot(density(model5.sim$sims.matrix[,1]), main = paste(c("Density plot
for b01")))
plot(density(model5.sim$sims.matrix[,2]), main = paste(c("Density plot
for b02")))
plot(density(model5.sim$sims.matrix[,3]), main = paste(c("Density plot
for b03")))
plot(density(model5.sim$sims.matrix[,4]), main = paste(c("Density plot
for b04")))
plot(density(model5.sim$sims.matrix[,5]), main = paste(c("Density plot
for b05")))
plot(density(model5.sim$sims.matrix[,6]), main = paste(c("Density plot
for b11")))
plot(density(model5.sim$sims.matrix[,7]), main = paste(c("Density plot
for b14")))
plot(density(model5.sim$sims.matrix[,8]), main = paste(c("Density plot
for b15")))
plot(density(model5.sim$sims.matrix[,9]), main = paste(c("Density plot
for b21")))
plot(density(model5.sim$sims.matrix[,10]), main = paste(c("Density
plot for b22")))
plot(density(model5.sim$sims.matrix[,11]), main = paste(c("Density
plot for b23")))
plot(density(model5.sim$sims.matrix[,12]), main = paste(c("Density
plot for b24")))
plot(density(model5.sim$sims.matrix[,13]), main = paste(c("Density
plot for b35")))
plot(density(model5.sim$sims.matrix[,14]), main = paste(c("Density
plot for b42")))
plot(density(model5.sim$sims.matrix[,15]), main = paste(c("Density
plot for b44")))
plot(density(model5.sim$sims.matrix[,16]), main = paste(c("Density
plot for b45")))
plot(density(model5.sim$sims.matrix[,17]), main = paste(c("Density
plot for b51")))
plot(density(model5.sim$sims.matrix[,18]), main = paste(c("Density
plot for b53")))
plot(density(model5.sim$sims.matrix[,19]), main = paste(c("Density
plot for b54")))
plot(density(model5.sim$sims.matrix[,20]), main = paste(c("Density
plot for b55")))
plot(density(model5.sim$sims.matrix[,21]), main = paste(c("Density
plot for b61")))

```

```

plot(density(model5.sim$sims.matrix[,22]), main = paste(c("Density
plot for b63")))
plot(density(model5.sim$sims.matrix[,23]), main = paste(c("Density
plot for b64")))
plot(density(model5.sim$sims.matrix[,24]), main = paste(c("Density
plot for b65")))
plot(density(model5.sim$sims.matrix[,25]), main = paste(c("Density
plot for b71")))
plot(density(model5.sim$sims.matrix[,26]), main = paste(c("Density
plot for b72")))
plot(density(model5.sim$sims.matrix[,27]), main = paste(c("Density
plot for b73")))
plot(density(model5.sim$sims.matrix[,28]), main = paste(c("Density
plot for b74")))
plot(density(model5.sim$sims.matrix[,29]), main = paste(c("Density
plot for b75")))
plot(density(model5.sim$sims.matrix[,30]), main = paste(c("Density
plot for b84")))
plot(density(model5.sim$sims.matrix[,31]), main = paste(c("Density
plot for b91")))
plot(density(model5.sim$sims.matrix[,32]), main = paste(c("Density
plot for b92")))
plot(density(model5.sim$sims.matrix[,33]), main = paste(c("Density
plot for b95")))
plot(density(model5.sim$sims.matrix[,34]), main = paste(c("Density
plot for b103")))
plot(density(model5.sim$sims.matrix[,35]), main = paste(c("Density
plot for b104")))
plot(density(model5.sim$sims.matrix[,36]), main = paste(c("Density
plot for b105")))
plot(density(model5.sim$sims.matrix[,37]), main = paste(c("Density
plot for b111")))
plot(density(model5.sim$sims.matrix[,38]), main = paste(c("Density
plot for b113")))
plot(density(model5.sim$sims.matrix[,39]), main = paste(c("Density
plot for b114")))
plot(density(model5.sim$sims.matrix[,40]), main = paste(c("Density
plot for b115")))
plot(density(model5.sim$sims.matrix[,81]), main = paste(c("Density
plot for deviance")))

```

```

##-----
## Trace Plots for betas.
##-----
# dev.off() # delete all previous plots
par(mfrow=c(3,3)) # 9 plots per picture

```

```
plot(model5.sim$sims.matrix[,1], type = "l", main = paste(c("Trace
plot for b01")))
plot(model5.sim$sims.matrix[,2], type = "l", main = paste(c("Trace
plot for b02")))
plot(model5.sim$sims.matrix[,3], type = "l", main = paste(c("Trace
plot for b03")))
plot(model5.sim$sims.matrix[,4], type = "l", main = paste(c("Trace
plot for b04")))
plot(model5.sim$sims.matrix[,5], type = "l", main = paste(c("Trace
plot for b05")))
plot(model5.sim$sims.matrix[,6], type = "l", main = paste(c("Trace
plot for b11")))
plot(model5.sim$sims.matrix[,7], type = "l", main = paste(c("Trace
plot for b14")))
plot(model5.sim$sims.matrix[,8], type = "l", main = paste(c("Trace
plot for b15")))
plot(model5.sim$sims.matrix[,9], type = "l", main = paste(c("Trace
plot for b21")))
plot(model5.sim$sims.matrix[,10], type = "l", main = paste(c("Trace
plot for b22")))
plot(model5.sim$sims.matrix[,11], type = "l", main = paste(c("Trace
plot for b23")))
plot(model5.sim$sims.matrix[,12], type = "l", main = paste(c("Trace
plot for b24")))
plot(model5.sim$sims.matrix[,13], type = "l", main = paste(c("Trace
plot for b35")))
plot(model5.sim$sims.matrix[,14], type = "l", main = paste(c("Trace
plot for b42")))
plot(model5.sim$sims.matrix[,15], type = "l", main = paste(c("Trace
plot for b44")))
plot(model5.sim$sims.matrix[,16], type = "l", main = paste(c("Trace
plot for b45")))
plot(model5.sim$sims.matrix[,17], type = "l", main = paste(c("Trace
plot for b51")))
plot(model5.sim$sims.matrix[,18], type = "l", main = paste(c("Trace
plot for b53")))
plot(model5.sim$sims.matrix[,19], type = "l", main = paste(c("Trace
plot for b54")))
plot(model5.sim$sims.matrix[,20], type = "l", main = paste(c("Trace
plot for b55")))
plot(model5.sim$sims.matrix[,21], type = "l", main = paste(c("Trace
plot for b61")))
plot(model5.sim$sims.matrix[,22], type = "l", main = paste(c("Trace
plot for b63")))
plot(model5.sim$sims.matrix[,23], type = "l", main = paste(c("Trace
plot for b64")))
plot(model5.sim$sims.matrix[,24], type = "l", main = paste(c("Trace
plot for b65")))
```

```

plot(model5.sim$sims.matrix[,25], type = "l", main = paste(c("Trace
plot for b71")))
plot(model5.sim$sims.matrix[,26], type = "l", main = paste(c("Trace
plot for b72")))
plot(model5.sim$sims.matrix[,27], type = "l", main = paste(c("Trace
plot for b73")))
plot(model5.sim$sims.matrix[,28], type = "l", main = paste(c("Trace
plot for b74")))
plot(model5.sim$sims.matrix[,29], type = "l", main = paste(c("Trace
plot for b75")))
plot(model5.sim$sims.matrix[,30], type = "l", main = paste(c("Trace
plot for b84")))
plot(model5.sim$sims.matrix[,31], type = "l", main = paste(c("Trace
plot for b91")))
plot(model5.sim$sims.matrix[,32], type = "l", main = paste(c("Trace
plot for b92")))
plot(model5.sim$sims.matrix[,33], type = "l", main = paste(c("Trace
plot for b95")))
plot(model5.sim$sims.matrix[,34], type = "l", main = paste(c("Trace
plot for b103")))
plot(model5.sim$sims.matrix[,35], type = "l", main = paste(c("Trace
plot for b104")))
plot(model5.sim$sims.matrix[,36], type = "l", main = paste(c("Trace
plot for b105")))
plot(model5.sim$sims.matrix[,37], type = "l", main = paste(c("Trace
plot for b111")))
plot(model5.sim$sims.matrix[,38], type = "l", main = paste(c("Trace
plot for b113")))
plot(model5.sim$sims.matrix[,39], type = "l", main = paste(c("Trace
plot for b114")))
plot(model5.sim$sims.matrix[,40], type = "l", main = paste(c("Trace
plot for b115")))
plot(model5.sim$sims.matrix[,81], type = "l", main = paste(c("Trace
plot for deviance")))

##-----
## Autocorrelation Plots for betas.
##-----
# dev.off() # delete all previous plots
par(mfrow=c(3,3)) # 9 plots per picture

acf(model5.sim$sims.matrix[,1], main = paste(c("ACF plot for b01")))
acf(model5.sim$sims.matrix[,2], main = paste(c("ACF plot for b02")))
acf(model5.sim$sims.matrix[,3], main = paste(c("ACF plot for b03")))
acf(model5.sim$sims.matrix[,4], main = paste(c("ACF plot for b04")))
acf(model5.sim$sims.matrix[,5], main = paste(c("ACF plot for b05")))
acf(model5.sim$sims.matrix[,6], main = paste(c("ACF plot for b11")))

```

```

acf(model5.sim$sims.matrix[,7], main = paste(c("ACF plot for b14")))
acf(model5.sim$sims.matrix[,8], main = paste(c("ACF plot for b15")))
acf(model5.sim$sims.matrix[,9], main = paste(c("ACF plot for b21")))
acf(model5.sim$sims.matrix[,10], main = paste(c("ACF plot for b22")))
acf(model5.sim$sims.matrix[,11], main = paste(c("ACF plot for b23")))
acf(model5.sim$sims.matrix[,12], main = paste(c("ACF plot for b24")))
acf(model5.sim$sims.matrix[,13], main = paste(c("ACF plot for b35")))
acf(model5.sim$sims.matrix[,14], main = paste(c("ACF plot for b42")))
acf(model5.sim$sims.matrix[,15], main = paste(c("ACF plot for b44")))
acf(model5.sim$sims.matrix[,16], main = paste(c("ACF plot for b45")))
acf(model5.sim$sims.matrix[,17], main = paste(c("ACF plot for b51")))
acf(model5.sim$sims.matrix[,18], main = paste(c("ACF plot for b53")))
acf(model5.sim$sims.matrix[,19], main = paste(c("ACF plot for b54")))
acf(model5.sim$sims.matrix[,20], main = paste(c("ACF plot for b55")))
acf(model5.sim$sims.matrix[,21], main = paste(c("ACF plot for b61")))
acf(model5.sim$sims.matrix[,22], main = paste(c("ACF plot for b63")))
acf(model5.sim$sims.matrix[,23], main = paste(c("ACF plot for b64")))
acf(model5.sim$sims.matrix[,24], main = paste(c("ACF plot for b65")))
acf(model5.sim$sims.matrix[,25], main = paste(c("ACF plot for b71")))
acf(model5.sim$sims.matrix[,26], main = paste(c("ACF plot for b72")))
acf(model5.sim$sims.matrix[,27], main = paste(c("ACF plot for b73")))
acf(model5.sim$sims.matrix[,28], main = paste(c("ACF plot for b74")))
acf(model5.sim$sims.matrix[,29], main = paste(c("ACF plot for b75")))
acf(model5.sim$sims.matrix[,30], main = paste(c("ACF plot for b84")))
acf(model5.sim$sims.matrix[,31], main = paste(c("ACF plot for b91")))
acf(model5.sim$sims.matrix[,32], main = paste(c("ACF plot for b92")))
acf(model5.sim$sims.matrix[,33], main = paste(c("ACF plot for b95")))
acf(model5.sim$sims.matrix[,34], main = paste(c("ACF plot for b103")))
acf(model5.sim$sims.matrix[,35], main = paste(c("ACF plot for b104")))
acf(model5.sim$sims.matrix[,36], main = paste(c("ACF plot for b105")))
acf(model5.sim$sims.matrix[,37], main = paste(c("ACF plot for b111")))
acf(model5.sim$sims.matrix[,38], main = paste(c("ACF plot for b113")))
acf(model5.sim$sims.matrix[,39], main = paste(c("ACF plot for b114")))
acf(model5.sim$sims.matrix[,40], main = paste(c("ACF plot for b115")))
acf(model5.sim$sims.matrix[,81], main = paste(c("ACF plot for
deviance")))

```

```

#####
# Initial Statistics & Plots for our Analysis for the Binomial Case
#####
## Load the data into R.
wines <- read.csv("C:\\\\Users\\\\30697\\\\Desktop\\\\06_winequality-red.csv",
header = T, sep=";", stringsAsFactors = FALSE)

```

```

## Change the values from mg/dm^3 to g/dm^3 for free and total sulfur
## dioxide.
wines$free.sulfur.dioxide <- wines$free.sulfur.dioxide*0.001
wines$total.sulfur.dioxide <- wines$total.sulfur.dioxide*0.001

## Plot of the sensory data.
plot(factor(wines$quality), xlab = "Sensory Preference (Median)", ylab
= "Frequency (Red Wine Samples)")

## Quality with 6 Categories (3-8).
table(wines$quality)

## Shift the response variable from 3 until 8, to 1 until 6.
wines[wines[,12]==3,12] <- 1
wines[wines[,12]==4,12] <- 2
wines[wines[,12]==5,12] <- 3
wines[wines[,12]==6,12] <- 4
wines[wines[,12]==7,12] <- 5
wines[wines[,12]==8,12] <- 6

## Check if it worked.
table(wines$quality)
plot(factor(wines$quality), xlab = "Sensory Preference (Median)", ylab
= "Frequency (Red Wine Samples)")

## Split the data response into binary
wines$quality <- ifelse(wines$quality<=3, 0, 1)
y <- wines$quality
table(wines$quality)
# 10+53+681=744 zeros
# 638+199+18=855 ones

#####
## Variable Selection using BAS
#####
library(BAS)
wines.bas2 <- wines

n <- 1599
res <- bas.glm(quality~., family='binomial', data=wines.bas2,
betaprior=g.prior(n), modelprior = beta.binomial(2,10))

coef(res)
summary(res)

## Turn our data frame into a matrix and then into a list.

```

```

wines <- scale(wines[,-12], center=TRUE, scale=TRUE) # if we want to
center our values
wines1 <- as.matrix(wines)
y <- as.matrix(y)
lwin <- list(y=y,x=wines1[,1:11])

#####
# Run our Binomial Model in R.
#####
## Load libraries.
library(R2WinBUGS)
library(BRugs)

## Set the directory.
openbugs.dir <- "D:\\OpenBUGS323"

## A list of initials (one chain).
inits6 <- list(
  list(b=rep(0, times=7)) )

## The parameters of interest.
parameter.names6 <- c('b', 'B')

## The model (in OpenBUGS).
model6.sim <- bugs( lwin, inits6, model.file =
"C:\\Users\\30697\\Desktop\\binomial2.txt", parameters =
parameter.names6,
n.chains = 1, n.iter = 11000, n.burnin=1000, n.thin=1, bugs.directory
= openbugs.dir, debug=F, program="OpenBUGS")

#####
# Check for convergence using plots and CODA
#####
library(coda)
temp<-as.mcmc.list(model5.sim)
summary(temp)
summary(temp)$statistics[, "Naive SE"]
summary(temp)$statistics[, "Time-series SE"]
batchSE(temp)
effectiveSize(temp)

## Automated Convergence Tests.
codamenu()
2
temp
2

```

```

1
4
3
2
4
2
8
4
Y

## Ergodic Plots for bs.
for (i in 1:7){

plot(cumsum(model1.sim$sims.matrix[,i])/1:length(model1.sim$sims.matri
x[,i]), type="l", ylab = "Cumsum Values", main = paste(c("Ergodic Mean
Plot for b",i)))
}

plot(cumsum(model1.sim$sims.matrix[,15])/1:length(model1.sim$sims.matr
ix[,15]), type="l", ylab = "Cumsum Values", main = paste(c("Ergodic
Mean Plot for deviance")))

## Autocorrelation Plots for bs.
for (i in 1:7){
  acf(model1.sim$sims.matrix[,i], main = paste(c("ACF plot for
b",i)))
}
acf(model1.sim$sims.matrix[,15], main = paste(c("ACF plot for
deviance")))

## Density Plots for bs.
for (i in 1:7){
  plot(density(model1.sim$sims.matrix[,i]), main = paste(c("Density
plot for b",i)))
}
plot(density(model1.sim$sims.matrix[,15]), main = paste(c("Density
plot for deviance")))

## Trace Plots for b0s.
for (i in 1:7){
  plot(model1.sim$sims.matrix[,i], type = "l", main = paste(c("Trace
plot for b",i)))
}
plot(model1.sim$sims.matrix[,15], type = "l", main = paste(c("Trace
plot for deviance")))

```

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