

Exercise 3 IML

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(1)(2) (3)

$$\underset{\mathcal{O}(w,b)}{\operatorname{arg\,min}} \|w\|^2 \quad \text{s.t.} \quad \forall i: y_i(\langle w, x_i \rangle + b) \geq 1 \quad \text{: Hard Svm}$$

$$\begin{array}{c} \text{מ}' \quad \text{P} \quad \text{Q}, \text{R} \quad \text{A} \quad \text{B} \quad \text{C} \quad \text{D} \\ \downarrow \quad \downarrow \quad \downarrow \quad \downarrow \quad \downarrow \quad \downarrow \quad \downarrow \\ \text{f}'' \quad \text{f}''' \quad \text{f}'''' \quad \text{f}''''' \quad \text{f}'''''' \quad \text{f}''''''' \end{array}$$

$$\underset{\mathcal{A} \leq \mathcal{B}}{\operatorname{arg\,min}} \frac{1}{2} r^T Q r \quad \text{s.t.} \quad A r \leq b$$

$$A r \leq b \quad \text{N.W.} \quad \text{f}''(x) \quad \text{f}''''(x) \quad \text{f}'''''(x) \quad \text{f}''''''(x)$$

$$y_i(\langle w, x_i \rangle + b) \geq 1 \iff -y_i(\langle w, x_i \rangle + b) \leq -1 \iff$$

$$\iff -y_i(x_i^T w + b) \leq -1 \iff -y_i x_i^T w - y_i b \leq -1$$

$$Y = \begin{bmatrix} w \\ b \end{bmatrix} \quad . \quad A = \begin{bmatrix} y_1 x_1^T & \dots & y_n x_n^T & y_1 \\ \vdots & & & \vdots \\ y_1 x_n^T & \dots & y_n x_n^T & y_n \end{bmatrix} \quad \text{R.H.S.}$$

$$-y_i x_i^T w - y_i b = A e_i \leq -1$$

$$\|w\|^2 = w^T I w \quad \text{C} \quad \text{Q.E.D.} \quad \text{arg\,min}_{\mathcal{O}(w,b)} \|w\|^2 \quad \text{R.H.S.}$$

$$Q = \begin{bmatrix} I & 0 \\ 0 & 0 \end{bmatrix}$$

$$I_{d \times d} \quad \text{def}$$

$$\begin{aligned} \operatorname{arg\,min}_w w^T I w &= \operatorname{arg\,min}_w \underbrace{\left[\begin{matrix} w \\ b \end{matrix} \right]^T}_{= \operatorname{arg\,min}_w w^T Q w + b^T v} \underbrace{\left[\begin{matrix} w \\ b \end{matrix} \right]}_{= Q \cdot w + v} \\ &= \operatorname{arg\,min}_v v^T Q \cdot v + b^T v \end{aligned}$$

Soft SVM 2

$$\arg \min \frac{1}{2} \|w\|^2 + \frac{1}{m} \sum_i \xi_i \quad \text{s.t. } y_i \langle w, x_i \rangle \geq b - \xi_i, \quad \xi_i \geq 0$$

: sum of loss function Soft SVM model 13

$$\arg \min \frac{1}{2} \|w\|^2 + \frac{1}{m} \sum_i \text{hinge}(y_i \langle w, x_i \rangle)$$

$$f(x) = \sum_i w_i x_i + b \quad \text{hinge} \quad \text{error}$$

$$y_i \langle w, x_i \rangle \geq b - \xi_i, \quad \xi_i \geq 0$$

$$y_i \langle w, x_i \rangle \geq 1 \quad \text{for } w \in \mathbb{R}^d \quad \text{sign} \quad \text{if}$$

$$\text{if } y_i \langle w, x_i \rangle < 1 \quad \text{then } \xi_i = 1 - y_i \langle w, x_i \rangle \quad \text{otherwise } \xi_i = 0$$

①

$$\text{hinge} \rightarrow 0 \quad \text{if } y_i \langle w, x_i \rangle < 1 \quad \text{else}$$

$$1 \rightarrow 2 \quad 1 - y_i \langle w, x_i \rangle \quad \text{if } y_i \langle w, x_i \rangle < 1 \quad \text{else}$$

$$\text{hinge}(y_i \langle w, x_i \rangle) \rightarrow 0 \quad \text{if } y_i \langle w, x_i \rangle < 1 \quad \text{else}$$

$$\text{if } y_i \langle w, x_i \rangle < 1 \quad \xi_i = \text{hinge}(y_i \langle w, x_i \rangle) \quad \text{else } \xi_i = 0$$

estimators: likelihood l-wogn i 1034 P. 203 1) k . k . 3

: \propto Likelihood \rightarrow \propto \propto $f(\mathbf{y})$

$$L(\theta | X, y) = \sum_{i=1}^m f_{x,y|\theta}(x_i, y_i)$$

$$\text{Joint Posterior} \propto \prod_{i=1}^n N\left(x_i | \mu_{y_i}, \sigma_{y_i}^2\right) \cdot M_{\text{ukt}}(y_i, \pi)$$

$$Q(\Theta | X_{1:y}) = \log \left(\prod_{i=1}^m M(x_i | y_i, \sigma^2_{y_i}) \cdot M_{\text{wt}}(y_i, \pi) \right)$$

$$= \sum_{i=1}^m \log\left(N(x_i | y_i, \sigma^2_{y_i})\right) + \log\left(\text{Mult}(y_i, \pi)\right)$$

$$= \sum_{i=1}^m \log \left(\frac{1}{\sqrt{(2\pi)^d |\sigma^2|}} \exp \left(-\frac{1}{2} (x_i - y_i)^T (\theta^*)^{-1} (x_i - y_i) \right) \right) + \log(\eta_{y_i})$$

$$= \sum_{i=1}^m \log(\pi_{y_i}) - \frac{1}{2} \log(2\pi) - \frac{1}{2} \log(\sigma^2) - \frac{1}{2} (X_i - \mu)^T (\sigma^2)^{-1} (X_i - \mu)$$

$$= \sum_k \left[n_k \log(\sigma_k) - \frac{1}{2} \sum_i |y_i - \hat{y}_i|^2 + \frac{1}{2} \log(2\pi) + \frac{m}{2} \log(\sigma^2) \right]$$

132), $\eta_{\text{no}} \approx 1$ \rightarrow $\eta_{\text{no}} \approx 1$, $\eta_{\text{no}} \approx 1$

• $\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$

$$\sum_{k=1}^K \pi_k = 1 \quad | \quad \mathcal{T} \in [0, 1]^K \quad \text{and} \quad \pi_k \geq 0$$

Ein optimales Modell muss die Wahrscheinlichkeit der Beobachtung maximieren. Das ist dann der Fall, wenn $\sum_{k=1}^K \pi_k = 1$.

$$L = l(\theta|x,y) - \lambda g(\pi)$$

$$\frac{\partial L}{\partial \pi_k} = \frac{\partial}{\partial \pi_k} \cdot l(\theta|x,y) - \lambda \frac{\partial^2}{\partial \pi_k^2} g(\pi) = \frac{n_k}{\pi_k} - 1 = 0$$

$$\boxed{\pi_k = \frac{n_k}{\lambda}}$$

Die Parameter π_k sind frei, λ ist ein festes Maß

$$\sum_{i=1}^m \frac{n_i}{\lambda} = 1 \iff m = \lambda$$

$$\hat{\pi}_k^{\text{MLE}} = \frac{n_k}{m} \quad \text{ist} \quad \text{MLE}$$

$$\hat{y}_k^{\text{MLE}} = \frac{1}{n_k} \cdot \sum_{i=1}^m I_{y_i=k} x_i \quad \hat{\sigma}^2 = \frac{1}{m} \sum_{i=1}^m (x_i - \hat{y}_i^{\text{MLE}})^2 (x_i - \hat{y}_i^{\text{MLE}})^T$$

Wahrscheinlichkeit einer Beobachtung ist das Produkt der Wahrscheinlichkeiten der einzelnen Beobachtungen.

Die Wahrscheinlichkeit einer Beobachtung ist das Produkt der Wahrscheinlichkeiten der einzelnen Beobachtungen.

$$\prod_{i=1}^m \prod_{j=1}^{n_k} N(x_i | \mu_{y_j}, \sigma_{y_j}^2) \cdot \text{Mult}(y_i, \pi)$$

Log-Likelihood \rightarrow Punkt

$$= \sum_k \left[n_k \log(\pi_k) - \frac{1}{2} \sum_{i:y_i=k} (x_i - \bar{x}_i)^T (\bar{x}_i) \right] - \frac{m}{2} \log(2\pi) \cdot \frac{m}{2} \log |\Omega|$$

(Maximum Likelihood Estimation 11/2023)

$$\hat{\pi}_k = \frac{n_k}{m}$$

$$\hat{\mu}_{kj}^{MLE} = \frac{1}{n_k} \sum_{i:y_i=k} x_{ij}$$

$$(\hat{\Omega}^2)_{kj}^{MLE} = \frac{1}{n_k} \sum_{i:y_i=k} (x_{ij} - \bar{x}_{kj})^2$$

לעומת象 likelihood, מוגדר פונקציית האנווקה כ**המינימום של סכום שטחים**

$$\prod_{i=1}^m \text{Pois}(x_i | \mu_{y_i}, \lambda_{y_i}) \cdot \text{Mult}(y_i, \pi)$$

(jc. בז' שאלת ר' בונס פ' 6.1) :JJ. Log-likelihood

$$L(\theta | X, y) = \log \left(\prod_{i=1}^m \text{Pois}(x_i | \mu_{y_i}, \lambda_{y_i}) \cdot \text{Mult}(y_i, \pi) \right)$$

$$= \sum_{i=1}^m \log(\pi_{y_i}) + \log \left(\frac{\lambda_{y_i}^{x_i} \cdot e^{-\lambda_{y_i}}}{x_i!} \right)$$

$$= \sum_{i=1}^m \log(\pi_{y_i}) + x_i \log(\lambda_{y_i}) - \lambda_{y_i} - \log(x_i!)$$

$$\frac{\partial L}{\partial \lambda} = \frac{1}{\lambda} \cdot \sum_{i:y_i=k} x_i - n_k = 0$$

$$\lambda_k^{MLE} = \frac{1}{n_k} \sum_{i=1}^m \sum_{y_i=k} x_i$$

$$\hat{\mu}_k^{\text{MLE}} = \frac{n_k}{m}$$

2.3. *likelyhood*, *prob.*, *probabil.* *prob.*

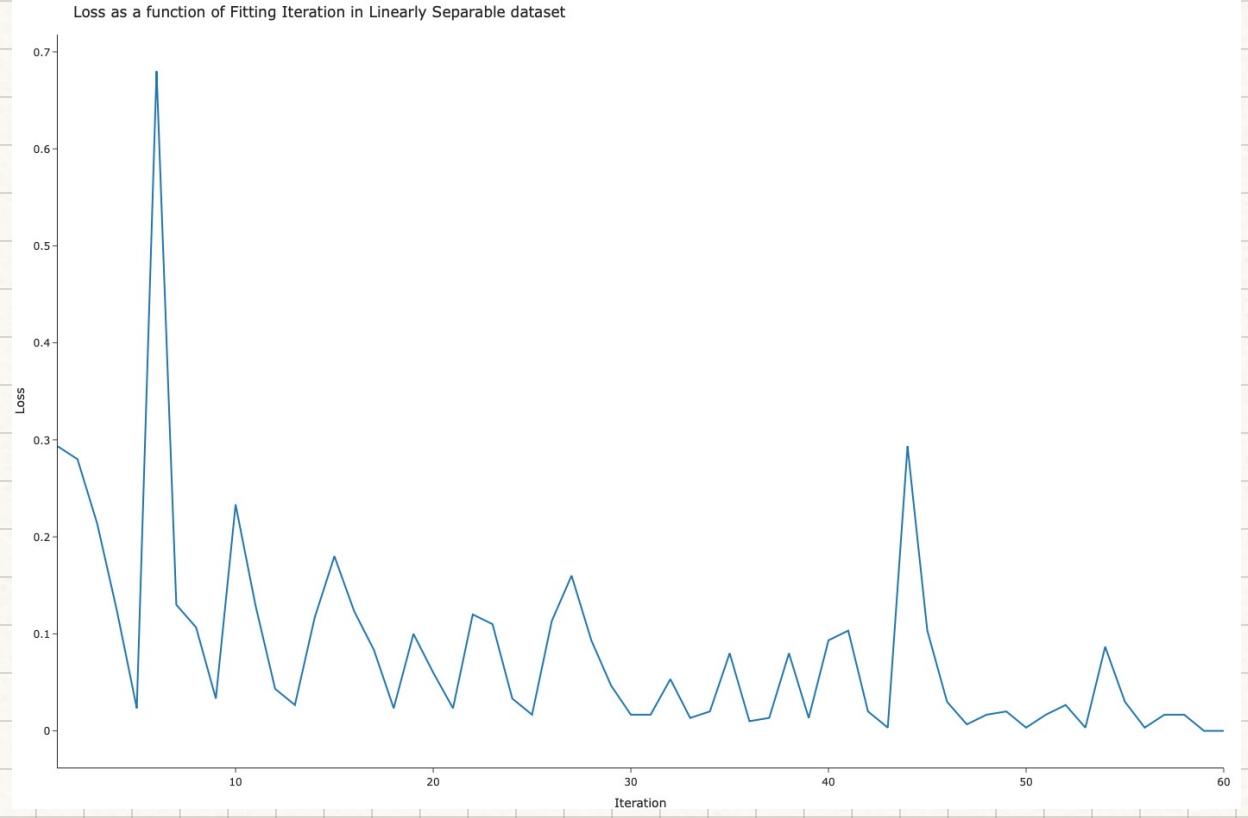
$$\prod_{j=1}^m \text{Pois}(X_i | \mu_{y_i}, \lambda_{y_i}) \cdot \text{Mult}(y_i, \pi)$$

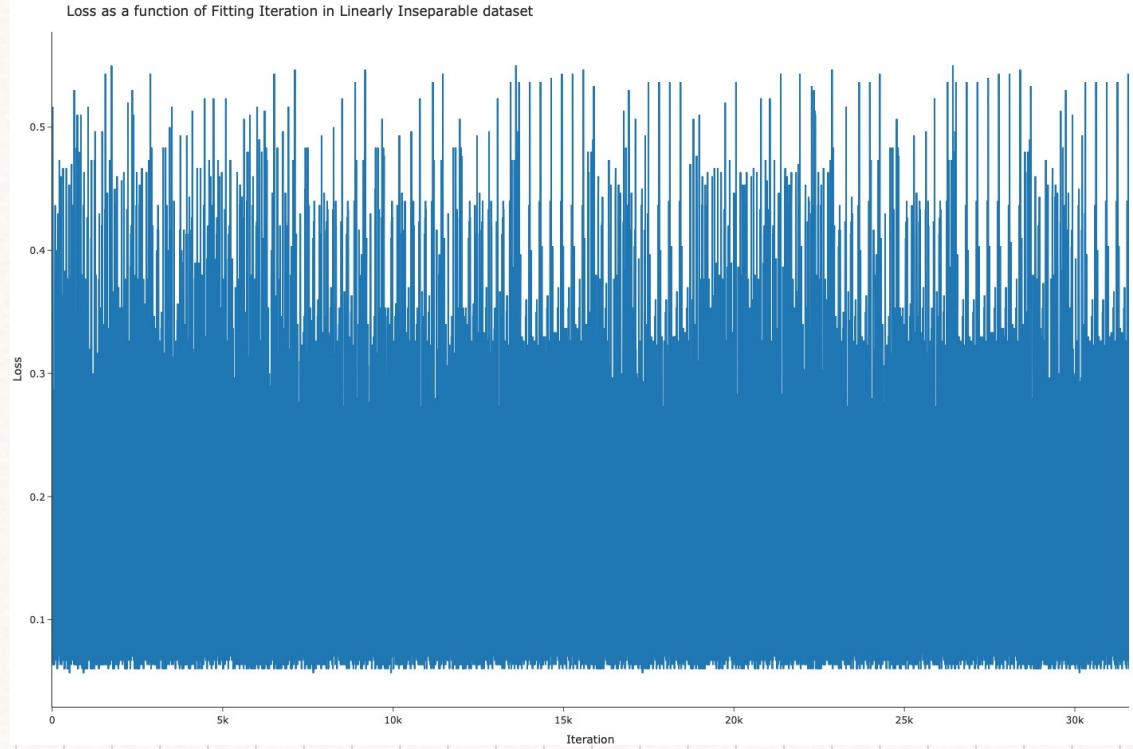
Now I like liposuction -- , you know plastic surgery

$$\log(p_{ij}) = \sum_{k=1}^m x_{ik} \log(\lambda_k) - \sum_{j=1}^n x_{kj} \log(\lambda_j)$$

$$\lambda_{ij}^{\text{MLE}} = \frac{1}{n_k} \sum_{i=1}^{n_k} \mathbb{1}_{y_i=k} x_{ij}$$

$$\text{vif} - \hat{H}_L = \frac{n_L}{m}$$

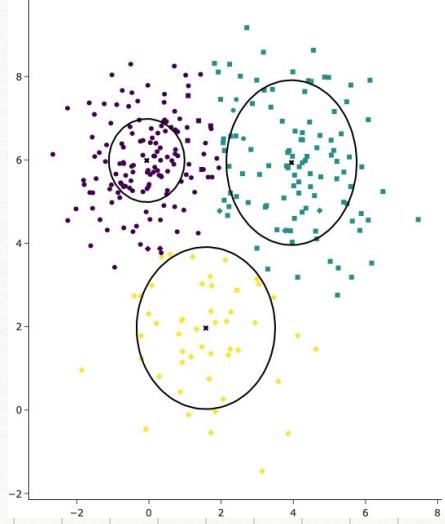




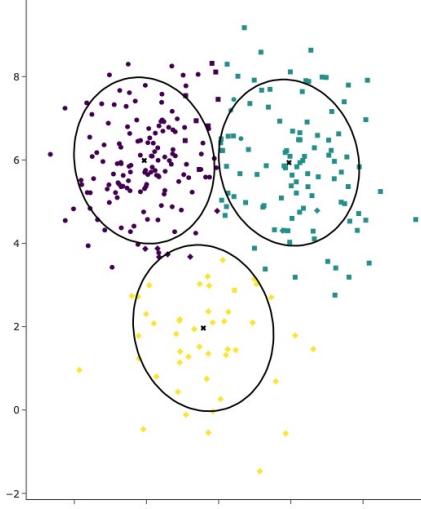
2

תפקידו של ה-Perceptron הוא למדוד אם ה- $\sum_i w_i x_i \geq b$.
 מושג זה נקרא פונקציית ה- sign .

Gaussian Dataset: gaussian1.npy
Gaussian Naive Bayes with 95.33% Accuracy

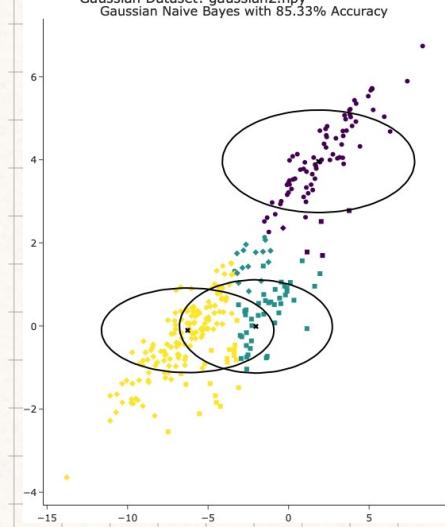


LDA with 93.33% Accuracy

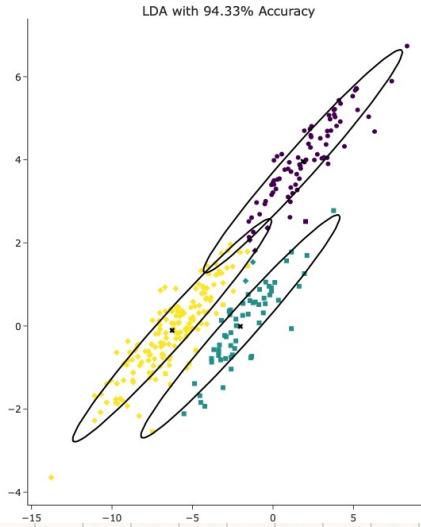


3

Gaussian Dataset: gaussian2.npy
Gaussian Naive Bayes with 85.33% Accuracy



LDA with 94.33% Accuracy



4

Gaussian Naive Bayes

Naive Bayes classifier
Decision Boundary
Gaussian Naive Bayes
LDA

Naive Bayes classifier
Decision Boundary
Gaussian Naive Bayes
LDA