

STA2001 Probability and Statistics (I)

Lecture 11

Tianshi Chen

The Chinese University of Hong Kong, Shenzhen

Review of the last lecture

Key concepts and/or techniques:

- ▶ Bivariate RV: (X, Y) or X and Y with range

$$\bar{S} \subseteq \bar{S}_X \times \bar{S}_Y \subseteq \mathbb{R}^2$$

- ▶ Joint pmf $f(x, y) : \bar{S} \rightarrow (0, 1]$

$$\bar{S} \subseteq \bar{S}_X \times \bar{S}_Y$$

Joint pmf $f(x, y) : \bar{S} \rightarrow (0, 1]$

- ▶ How to derive Marginal pmf from the joint pmf

$$f_X(x) = P_X(X=x, Y \in \bar{S}_Y(x)) \rightarrow \text{对 } y \text{ 求和}$$

$$f_X(x) = P_X(X=x) \triangleq P(\{X=x, Y \in \bar{S}_Y(x)\}) = \sum_{y \in \bar{S}_Y(x)} f(x, y)$$

三项分布

- ▶ Trinomial distribution: $(X, Y) \sim \text{Trinomial}(n, p_X, p_Y)$

$$f(x, y) = \frac{n!}{x!y!(n-x-y)!} p_X^x p_Y^y (1-p_X-p_Y)^{n-x-y}, (x, y) \in \bar{S},$$

$$\bar{S} = \{(x, y) | x + y \leq n, x = 0, 1, \dots, n, y = 0, 1, \dots, n\}$$

- ▶ X and Y are independent if $f(x, y) = f_X(x)f_Y(y)$

Review of the last lecture

Definition[Joint pmf]

The function $f(x, y) : \bar{S} \rightarrow (0, 1]$ is called the joint probability mass function (joint pmf) of X and Y or (X, Y) , if

1. $f(x, y) > 0$ for $(x, y) \in \bar{S}$,

2. $\sum_{(x,y) \in \bar{S}} f(x, y) = 1$,

3. For $A \subseteq \bar{S}$,

$$P[(X, Y) \in A] \triangleq P(\{(X, Y) \in A\}) = \sum_{(x,y) \in A} f(x, y)$$

which defines the probability function for a set A . In particular, taking $A = \{(x, y)\}$ yields the probability of $X = x$ and $Y = y$, i.e.,

$$P(X = x, Y = y) = f(x, y)$$

Review of the last lecture

Definition[Marginal pmf]

Let (X, Y) be a bivariate RV, or X and Y be two RVs, and have the joint pmf $f(x, y) : \bar{S} \rightarrow (0, 1]$. Sometimes, we are interested in the pmf of X or Y alone, which is called the marginal pmf of X or Y and described by

For $x \in \bar{S}_X$,

$$\begin{aligned} f_X(x) &= P_X(X = x) \triangleq P(\{X = x, Y \in \bar{S}_Y(x)\}) \\ &= \sum_{y \in \bar{S}_Y(x)} f(x, y) \end{aligned}$$

where

$$\bar{S}_Y(x) = \{y | (x, y) \in \bar{S}\} \text{ for the given } x \in \bar{S}_X.$$

Review of the last lecture

It is crucial to understand the following definitions

$$\overline{S}, \overline{S_X}, \overline{S_Y}, \overline{S_X}(y), \overline{S_Y}(x)$$

$$\overline{S} = \{\text{all possible values of } (X, Y)\}$$

$$\overline{S_X} = \{\text{all possible values of } X\} = \{x \mid (x, y) \in \overline{S}\}$$

$$\overline{S_Y} = \{\text{all possible values of } Y\} = \{y \mid (x, y) \in \overline{S}\}$$

$$\overline{S_X}(y) = \{x \mid (x, y) \in \overline{S}\} \text{ for a given } y \in \overline{S_Y}$$

$$\overline{S_Y}(x) = \{y \mid (x, y) \in \overline{S}\} \text{ for a given } x \in \overline{S_X}$$

Review of the last lecture

Definition

The random variables X and Y are said to be independent if for every $x \in \overline{S}_X$ and $y \in \overline{S}_Y$

indc.

$$f(x, y) = f_X(x)f_Y(y)$$

or equivalently,

$$P(X = x, Y = y) = P_X(X = x)P_Y(Y = y).$$

X and Y are said to be dependent if otherwise.

When X and Y are independent,

not rectangular not inde

$$\overline{S} = \overline{S}_X \times \overline{S}_Y, \quad \overline{S} \text{ is said to be rectangular}$$

which is a necessary condition for independence of X and Y .

Section 4.2 The correlation coefficient

相关系数

Motivation

Study the relation between two RVs (random phenomena)

Covariance of X and Y

Definition

Let X and Y be RVs with joint pmf $f(x, y) : \bar{S} \rightarrow (0, 1]$
Take

$$g(X, Y) = (X - E(X))(Y - E(Y))$$

$$\text{Cov}(X, Y) \triangleq E[(X - E(X))(Y - E(Y))]$$

$$= \sum_{(x,y) \in \bar{S}} (x - E(X))(y - E(Y))f(x, y)$$

Moreover, we have $\text{Cov}(X, Y) = E[(X - E(X))(Y - E(Y))]$
 $= E[XY] - E[X]E[Y]$

$$\text{Cov}(X, Y) \triangleq E(XY) - E(X)E(Y)$$

Covariance of X and Y 协方差

► $\text{Cov}(X, Y) = E(XY) - E(X)E(Y)$ Cov(X,Y)=0
不相关

When $\text{Cov}(X, Y) = 0$, X and Y are uncorrelated.

When $Cov(X, Y) > 0$, X and Y are positively correlated. $Cov(X, Y) > 0$, 正相关

When $\text{Cov}(X, Y) < 0$, X and Y are negatively correlated. $\text{Cov}(X, Y) < 0$, 负相关

- Interpretation: Roughly speaking, a positive or negative covariance indicate that the values of $X - E(X)$ and $Y - E(Y)$ obtained in a single experiment “tend” to have the same or the opposite sign respectively.

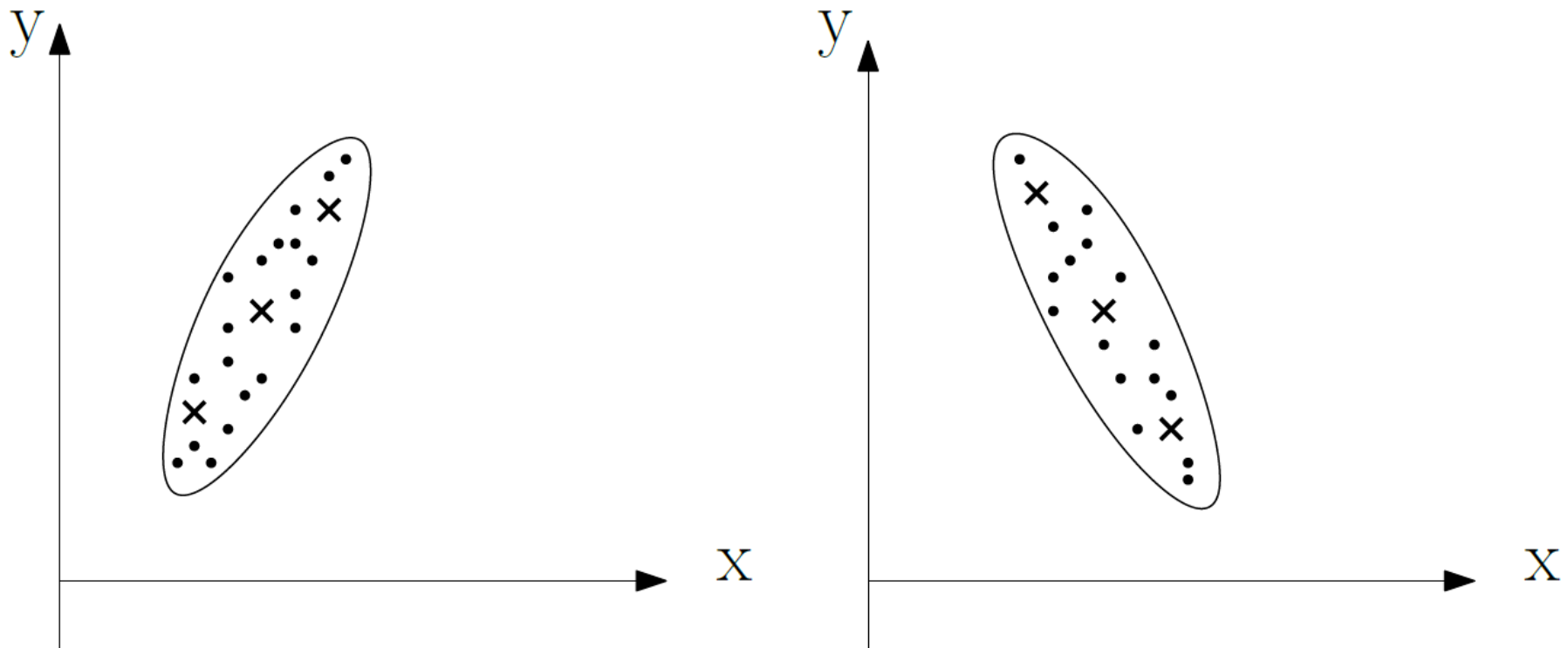
$$\text{Cov}(X, Y) = \sum_{(x,y) \in \bar{S}} (x - E(X))(y - E(Y))f(x, y)$$

$$x - u_x \text{ and } y - u_y$$

have same/opposite sign

Example 1 [Positively Correlated and Negatively Correlated RVs]

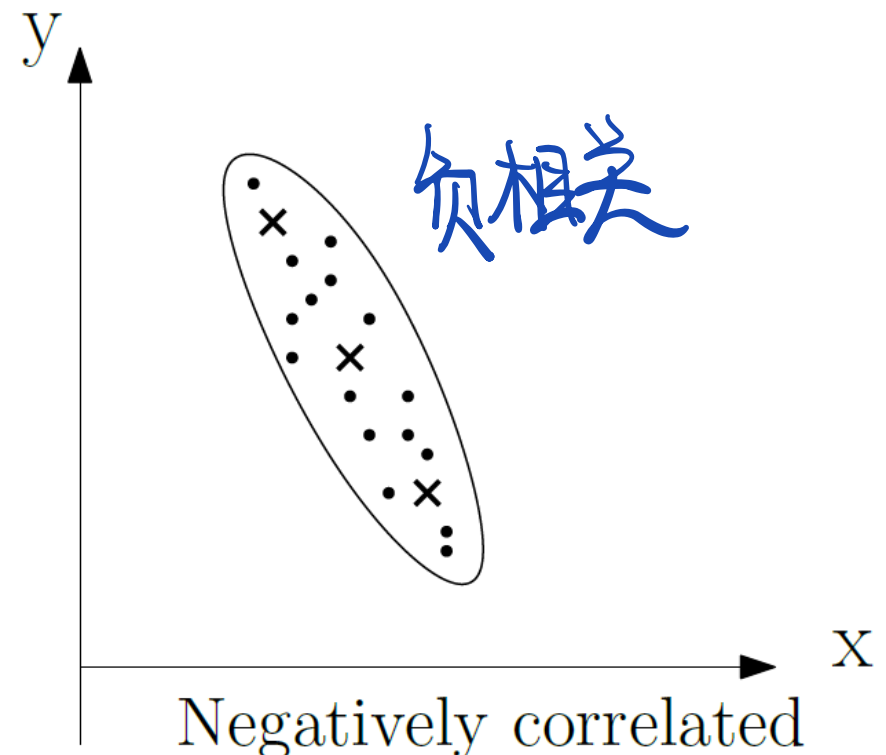
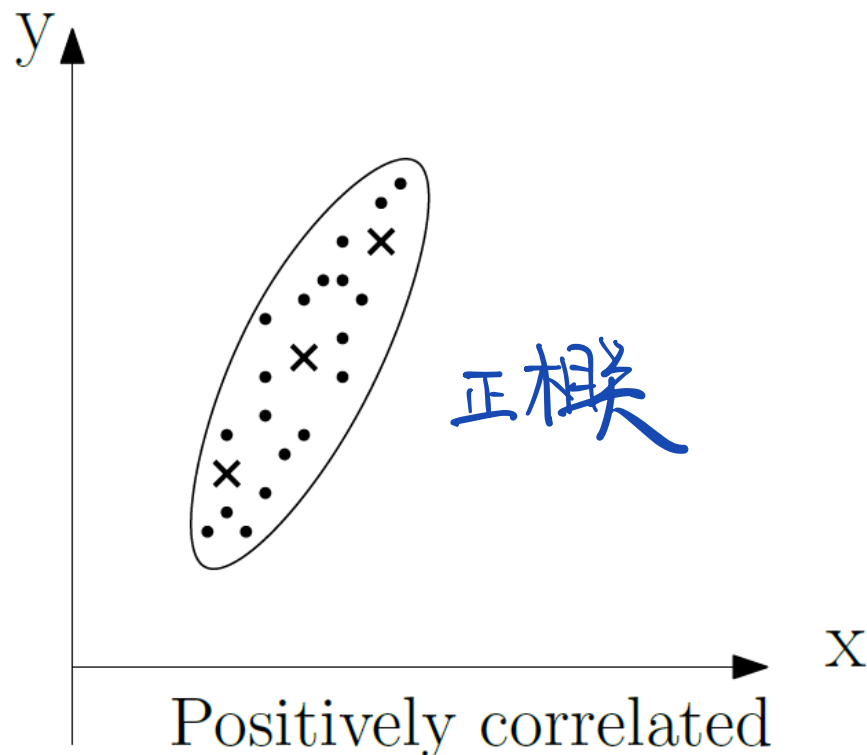
Assume that X and Y are uniformly distributed over the ellipses.



Question: which figure shows that X and Y are positively correlated?

Example 1 [Positively Correlated and Negatively Correlated RVs]

Assume that X and Y are uniformly distributed over the ellipses.



Independence \Rightarrow Uncorrelation

独立

- If X and Y are independent, we have

$$f(x, y) = f_X(x)f_Y(y) \Rightarrow \bar{S} = \bar{S}_X \times \bar{S}_Y$$

$$E(XY) = \sum_{(x,y) \in \bar{S}} xyf(x, y) = \sum_{x \in \bar{S}_X} \sum_{y \in \bar{S}_Y} xyf_X(x)f_Y(y)$$

$$= \sum_{x \in \bar{S}_X} xf_X(x) \left[\sum_{y \in \bar{S}_Y} yf_Y(y) \right] = E(X)E(Y)$$

inde $E_{XY} = Z_X \cdot Z_Y$

Independence \Rightarrow Uncorrelation

Therefore

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = 0$$

Independence of two RVs \Rightarrow uncorrelation of two RVs.

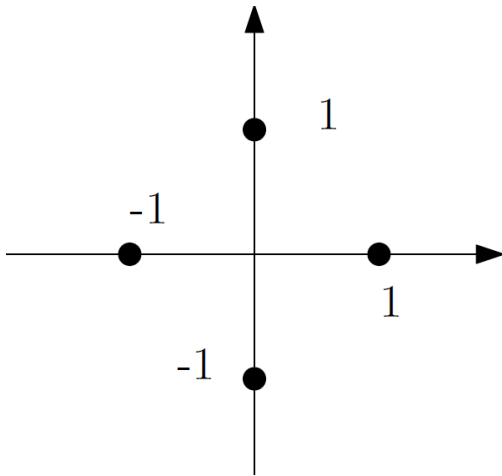
独立 \Rightarrow 无相关性
 \Leftarrow

However, the converse is not true, i.e, there exist X and Y which are uncorrelated but not independent.

Example 2 [Uncorrelation \nRightarrow Independence]

Question

Let (X, Y) be a bivariate RV that takes values $(1, 0)$, $(0, 1)$, $(-1, 0)$, $(0, -1)$, each with probability $\frac{1}{4}$, as shown in the figure below



Question 1

What are the marginal pmf of X and Y ?

Question 2

What is $\text{Cov}(X, Y)$?

Question 3

Are X and Y independent?

Example 2 [Uncorrelation \nRightarrow Independence]

To find marginal pmf of X and Y , $\overline{S_X} = \overline{S_Y} = \{-1, 0, 1\}$

$$f_X(x) = \begin{cases} \frac{1}{4}, & x = 1 \\ \frac{1}{2}, & x = 0 \\ \frac{1}{4}, & x = -1 \end{cases} \quad f_Y(y) = \begin{cases} \frac{1}{4}, & y = 1 \\ \frac{1}{2}, & y = 0 \\ \frac{1}{4}, & y = -1 \end{cases}$$

$$\text{Cov}(X, Y) = E(\overset{0}{XY}) - E(\overset{0}{X})E(\overset{0}{Y}) = 0 - 0 \times 0 = 0$$

which shows that X and Y are uncorrelated. 无关系

$$f_X(0)f_Y(1) = \frac{1}{2} \times \frac{1}{4} = \frac{1}{8} \neq f(0, 1) = \frac{1}{4}$$

which shows that X and Y are NOT independent.

找一个反例

Correlation Coefficient

Definition

The correlation coefficient of X and Y that have nonzero variance is defined as

$$\rho(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X)}\sqrt{\text{Var}(Y)}}$$

Handwritten notes in blue ink:

- 相关系数 (Correlation Coefficient) written above the fraction.
- $\frac{\text{Cov}(X, Y)}{\sigma_X \cdot \sigma_Y}$ written above the fraction.
- Arrows pointing from σ_X and σ_Y to $\sqrt{\text{Var}(X)}$ and $\sqrt{\text{Var}(Y)}$ respectively.

- Interpretation : $\rho > 0$ (or $\rho < 0$) indicate the values of $X - E[X]$ and $Y - E[Y]$ “tend” to have the same (or negative, respectively) sign.

Properties of the Correlation Coefficient

- ▶ It is a normalized version of $\text{Cov}(X, Y)$ and in fact $-1 \leq \rho(X, Y) \leq 1$.

$$\rho \in [-1, 1]$$

- ▶ $\rho = 1$ (resp. $\rho = -1$) if and only if there exists a positive (resp. negative) constant c such that

线性关系

$$Y - E(Y) = c(X - E(X)),$$

and the size of $|\rho|$ provides a normalized measure of the extent to which this is true.

$$\rho = 1 \quad Y - E(Y) = c(X - E(X))$$

Proof for Properties of Correlation Coefficient (1/3)

To prove $|\rho(X, Y)| \leq 1$ is equivalent to prove that

$$\text{Cov}(X, Y)^2 \leq \text{Var}(X)\text{Var}(Y)$$

Proof for Properties of Correlation Coefficient (1/3)

To prove $|\rho(X, Y)| \leq 1$ is equivalent to prove that

$$\text{Cov}(X, Y)^2 \leq \text{Var}(X)\text{Var}(Y)$$

To this goal, we consider

$$E((V + tW)^2) \geq 0,$$

where $t \in \mathbb{R}$, $V = X - E(X)$, $W = Y - E(Y)$. Then we have

$$\begin{aligned} E((V + tW)^2) &= E(V^2 + 2tVW + t^2W^2) \\ &= E(V^2) + 2tE(VW) + t^2E(W^2) \geq 0 \end{aligned}$$

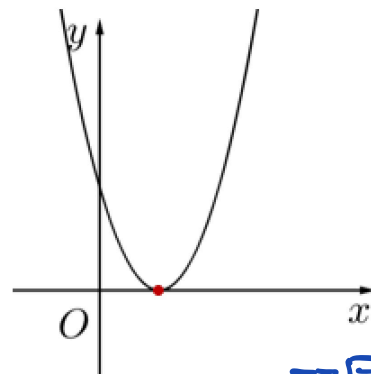
Proof for Properties of Correlation Coefficient (2/3)

Noting that

$$E(V^2) = \text{Var}(X), E(W^2) = \text{Var}(Y), E(VW) = \text{Cov}(X, Y)$$

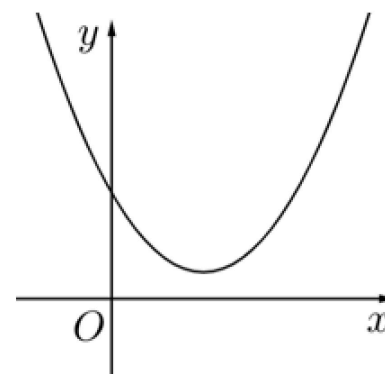
yields that for $t \in \mathbb{R}$,

$$E((V + tW)^2) = \text{Var}(X) + 2\text{Cov}(X, Y)t + \text{Var}(Y)t^2 \geq 0$$



One Solution

双同解



No Real Solution

无实解

Since the above equation is true for any $t \in \mathbb{R}$, it must hold that

$$4\text{Cov}(X, Y)^2 - 4\text{Var}(X)\text{Var}(Y) \leq 0, \text{ i.e., } \text{Cov}(X, Y)^2 \leq \text{Var}(X)\text{Var}(Y)$$

which implies that $|\rho(X, Y)| \leq 1$.

Proof for Properties of Correlation Coefficient (3/3)

When $\text{Cov}(X, Y)^2 - \text{Var}(X)\text{Var}(Y) = 0$, $|\rho(X, Y)| = 1$ implying

$$\begin{aligned} E((V + tW)^2) &= \text{Var}(X) + 2\text{Cov}(X, Y)t + \text{Var}(Y)t^2 = 0 \\ &= \text{Var}(X) \pm 2\sqrt{\text{Var}(X)}\sqrt{\text{Var}(Y)}t + \text{Var}(Y)t^2 = 0 \\ &= (\sqrt{\text{Var}(X)} \pm \sqrt{\text{Var}(Y)}t)^2 = 0 \end{aligned}$$

$$t^* = \mp \frac{\sqrt{\text{Var}(X)}}{\sqrt{\text{Var}(Y)}}.$$

Inserting the above t^* back in $E((V + tW)^2)$ yields

$$E((V + t^*W)^2) = 0 \implies V = -t^*W$$

$$X - E(X) = \pm \frac{\sqrt{\text{Var}(X)}}{\sqrt{\text{Var}(Y)}}(Y - E(Y)),$$

where $\rho = 1$ (resp. $\rho = -1$) corresponds to $+$ (resp. $-$).

Example

Question

Consider n independent tosses of a coin with probability of a head equal to p . Let X and Y be the number of heads and of tails in the n tosses, respectively. Calculate the correlation coefficient of X and Y .

Example

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Consider n independent tosses of a coin with probability of a head equal to p . Let X and Y be the number of heads and of tails in the n tosses, respectively. Calculate the correlation coefficient of X and Y .

$$X + Y = n \Rightarrow E[X] + E[Y] = n \Rightarrow X - E(X) = -[Y - E(Y)]$$

Example

Question

Consider n independent tosses of a coin with probability of a head equal to p . Let X and Y be the number of heads and of tails in the n tosses, respectively. Calculate the correlation coefficient of X and Y .

$$\textcircled{1} X + Y = n$$

$$\textcircled{2} EX + EY = n$$

$$\textcircled{1} - \textcircled{2} \rightarrow$$

$$X + Y = n \Rightarrow E[X] + E[Y] = n \Rightarrow X - E(X) = -[Y - E(Y)]$$

$$\text{Cov}(X, Y) = E[(X - E(X))(Y - E(Y))]$$

$$X - EX = -[Y - EY]$$

$$= -E[(Y - E(Y))^2] = -\text{Var}(Y)$$

$$\text{Var}(X) = E[(X - E(X))^2] = E[(Y - E(Y))^2] = \text{Var}(Y)$$

$$\rho(X, Y) = \frac{-\text{Var}(Y)}{\sqrt{\text{Var}(X)}\sqrt{\text{Var}(Y)}} = \frac{-\text{Var}(Y)}{\sqrt{\text{Var}(Y)}\sqrt{\text{Var}(Y)}} = -1$$

Section 4.3 Conditional distribution

Conditional Distribution 条件分布

Motivation: the conditional probability distribution is a probability distribution that describes the distribution of probability of events of a RV given the occurrence of a particular event

Assume that X and Y have a joint pmf $f(x, y) : \bar{S} \rightarrow (0, 1]$.

The marginal pmf of X and Y are

$$f_X(x) : \bar{S}_X \rightarrow (0, 1] \quad f_Y(y) : \bar{S}_Y \rightarrow (0, 1]$$

$$\bar{S}_X = \{\text{all possible values of } X \text{ in } \bar{S}\}$$

$$\bar{S}_X(y) = \{x | (x, y) \in \bar{S}\} \text{ for } y \in \bar{S}_Y$$

$$\bar{S}_Y = \{\text{all possible values of } Y \text{ in } \bar{S}\}$$

$$\bar{S}_Y(x) = \{y | (x, y) \in \bar{S}\} \text{ for } x \in \bar{S}_X$$

Conditional Distribution

By definition,

$$f(x, y) = P(X = x, Y = y)$$

$$\triangleq P(\{X = x, Y = y\}), (x, y) \in \bar{S}$$

$$f_X(x) = P_X(X = x)$$

$$\triangleq P(\{X = x, Y \in \bar{S}_Y(x)\}) = \sum_{y \in \bar{S}_Y(x)} f(x, y)$$

$$f_Y(y) = P_Y(Y = y)$$

$$\triangleq P(\{X \in \bar{S}_X(y), Y = y\}) = \sum_{x \in \bar{S}_X(y)} f(x, y)$$

Conditional Distribution

Let

$$A = \{X = x, Y \in \overline{S_Y}(x)\}$$

$$B = \{X \in \overline{S_X}(y), Y = y\}$$

Then for $(x, y) \in \overline{S}$,

$$A \cap B = \{X = x, Y = y\}$$

and recall the conditional probability of event A given event B is

$$P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{f(x, y)}{f_Y(y)}$$

under the assumption $P(B) > 0$, i.e., $f_Y(y) > 0$.

Conditional pmf

Definition

Conditional pmf of X given $Y = y$ is defined by

$$g(x|y) = \frac{f(x, y)}{f_Y(y)}, \quad x \in \overline{S}_X(y)$$

provided that $f_Y(y) > 0$.

$$g(x|y) = \frac{f(x, y)}{f_Y(y)}$$

Similarly, the conditional pmf of Y given that $X = x$ is defined by

$$h(y|x) = \frac{f(x, y)}{f_X(x)}, \quad y \in \overline{S}_Y(x)$$

provided that $f_X(x) > 0$.

$$h(y|x) = \frac{f(x, y)}{f_X(x)}$$

- ▶ What is the interpretation of the conditional pmf $g(x|y)$?
- ▶ What if X and Y are independent?

Some Remarks

性质 Conditional pmf is a well-defined pmf:

1. $h(y|x) > 0$

2. $\sum_{y \in \overline{S}_Y(x)} h(y|x) = 1$

$$\sum_{y \in \overline{S}_Y(x)} h(y|x) = \sum_{y \in \overline{S}_Y(x)} \frac{f(x, y)}{f_X(x)} = \frac{\sum_{y \in \overline{S}_Y(x)} f(x, y)}{f_X(x)} = \frac{f_X(x)}{f_X(x)} = 1$$

3. for $A \subseteq \overline{S}_Y(x)$

$$P(Y \in A | X = x) = \frac{P(X = x, Y \in A)}{P(X = x)}$$

$$= \frac{\sum_{y \in A} f(x, y)}{f_X(x)} = \sum_{y \in A} h(y|x)$$

Therefore, $h(y|x)$ (resp. $g(x|y)$) determines the distribution of probability of events of Y (resp. X) given $X = x$ (resp. $Y = y$).

Some Remarks

inde.

If X and Y are independent, then $f(x, y) = f_X(x)f_Y(y)$ and thus

$$g(x|y) = f_X(x), \text{ and } h(y|x) = f_Y(y),$$

which implies

- ▶ the occurrence of the event $Y = y$ does not change the probability of the occurrence of events of X
- ▶ the occurrence of the event $X = x$ does not change the probability of the occurrence of events of Y

Now, the implication of independent RVs becomes clear.

Example 1

Question

Let X and Y have the joint pmf

$$f(x, y) = \frac{x + y}{21}, \quad x = 1, 2, 3; \quad y = 1, 2.$$

We have showed

$$f_X(x) = \frac{2x + 3}{21}, \quad x = 1, 2, 3$$

$$f_Y(y) = \frac{y + 2}{7}, \quad y = 1, 2.$$

Q1: What is the conditional pmf of X given $Y = y$?

Q2: What is the conditional pmf of Y given $X = x$?

Q3: What is $P(1 \leq X \leq 2 | Y = 1)$?

Example 1

Q1: The conditional pmf of X given $Y = y$ is

$$g(x|y) = \frac{f(x,y)}{f_Y(y)} = \frac{x+y}{21} \bigg/ \left(\frac{y+2}{7}\right) = \frac{x+y}{3(y+2)},$$
$$x = 1, 2, 3; \quad y = 1, 2.$$

Q2: The conditional pmf of Y given $X = x$ is

$$h(y|x) = \frac{f(x,y)}{f_X(x)} = \frac{x+y}{21} \bigg/ \left(\frac{2x+3}{21}\right) = \frac{x+y}{2x+3}$$
$$x = 1, 2, 3; \quad y = 1, 2.$$

Q3:

$$P(1 \leq X \leq 2 | Y = 1) = \sum_{x=1}^2 g(x|1) = \sum_{x=1}^2 \frac{x+1}{3(1+2)} = \frac{5}{9}$$

Conditional Mathematical Expectation

- ▶ Let $g(Y)$ be a function of Y .

Then the conditional expectation of $g(Y)$ given $X = x$

$$E(g(Y)|X = x) = \sum_{y \in \overline{S_Y}(x)} g(y)h(y|x)$$

- ▶ When $g(Y) = Y$,

$$E[g(Y)|X=x] = \sum_{y \in \overline{S_Y}(x)} g(y)h(y|x)$$

$$E(Y|X = x) = \sum_{y \in \overline{S_Y}(x)} yh(y|x) \rightarrow \text{conditional mean}$$

$$\begin{aligned} E[Y|X=x] \\ = \sum_{y \in \overline{S_Y}(x)} y h(y|x) \rightarrow \text{conditional mean} \end{aligned}$$

Conditional Mathematical Expectation

$$g(Y) = \{Y - E[Y|X=x]\}^2$$

► When $g(Y) = [Y - E(Y|X = x)]^2$

$$\text{Var}(Y|X = x) \triangleq E\{[Y - E(Y|X = x)]^2 | X = x\}$$

$\text{Var}[Y|X=x]$
 $= \sum_{y \in \bar{S}_Y(x)} \{Y - E[Y|X=x]\}^2 \cdot h(y|x)$

$$= \sum_{y \in \bar{S}_Y(x)} [y - E(Y|X = x)]^2 h(y|x)$$

$$= E(Y^2 | X = x) - [E(Y | X = x)]^2$$

→ conditional variance

$$E[Y^2 | X=x] - E^2[Y | X=x]$$

Example 1, continued

Question

Let X and Y have the joint pmf

$$f(x, y) = \frac{x + y}{21}, \quad x = 1, 2, 3; \quad y = 1, 2.$$

We have showed

$$f_X(x) = \frac{2x + 3}{21}, \quad x = 1, 2, 3$$
$$f_Y(y) = \frac{y + 2}{7}, \quad y = 1, 2.$$

Q1: What is the expectation of Y given $X = 3$?

Q2: What is the variance of Y given $X = 3$?

Example 1, continued

$$Q1 : E(Y|X = 3) = \sum_{y \in \overline{S_Y}(3)} yh(y|3) = \sum_{y=1}^2 y\left(\frac{3+y}{9}\right) = \frac{14}{9}$$

$$Q2 : Var(Y|X = 3) = \sum_{y \in \overline{S_Y}(3)} [y - E(Y|X = 3)]^2 h(y|3)$$

$$= \sum_{y=1}^2 \left(y - \frac{14}{9}\right)^2 \frac{3+y}{9} = \frac{20}{81}$$