

## 1 Detecting defects

a)

We define E the event of find at least one defective cookie. The complimentary probability of event E,  $Pr(\overline{E})$  is not to find defective cookies. Since we pick each cookie independently and uniformly at random, not finding defective cookies after k tries is equal to:

$$Pr(\overline{E}) = (1 - p)^k \implies Pr(E) = 1 - (1 - p)^k \quad (1)$$

Given  $p \geq a$  and  $k \geq \frac{\ln(100)}{a}$  it is true that:

$$\begin{aligned} k &\geq \frac{\ln(100)}{p} \implies \\ k \cdot p &\geq \ln(100) \implies \\ -k \cdot p &\leq \ln(0.01) \implies \\ e^{-k \cdot p} &\leq 0.01 \xrightarrow{(1-x)^k \leq e^{-kx}} \\ (1 - p)^k &\leq 0.01 \implies \\ 1 - (1 - p)^k &\geq 0.99 \xrightarrow{(1)} \\ Pr(E) &= 0.99 \end{aligned}$$

b)

We define W the event all the unreliable workers have been detected and fired and  $W_i$  the event that the worker i has been detected as reliable. Therefore :

$$Pr(W) = 1 - Pr\left[\bigcup_{i \in n} W_i\right]$$

Assuming that the  $Pr(W) \geq 0.99$  it implies that:

$$Pr\left[\bigcup_{i \in n} W_i\right] \leq 0.01$$

Based on union bound inequality, we know that:

$$Pr\left[\bigcup_{i \in n} W_i\right] \leq n \cdot Pr(W_i)$$

With regards to subquestion (a) for a reliable worker  $W_i$  it is true that  $Pr(W_i) \leq (1 - a)^k$ . So:

$$\begin{aligned} n \cdot (1 - a)^k &\leq 0.01 \xrightarrow{(1-x)^k \leq e^{-kx}} \\ n \cdot e^{-a \cdot k} &\leq 0.01 \implies \\ e^{-a \cdot k} &\leq \frac{0.01}{n} \implies \end{aligned}$$

$$\begin{aligned}
-a \cdot k &\leq \ln\left(\frac{0.01}{n}\right) \implies \\
-a \cdot k &\leq \ln\left(\frac{1}{100 \cdot n}\right) \implies \\
k &\geq \frac{\ln(100 \cdot n)}{a} \implies \\
k_{min} &= \frac{\ln(100 \cdot n)}{a}
\end{aligned}$$

## 2 Jensen's Inequality

a)

Given that  $f$  is concave function we can assume that  $f$  is a twice differentiable function and  $f''(x) \leq 0$ . We define  $\mu = E[X]$  where  $X$  is a random value. In order to proof the Jensen's inequality change for concave functions we will use the Taylor expansion of  $f$  function based on the fact that  $f$  is twice differentiable. So:

$$\begin{aligned}
f(x) &= f(\mu) + f'(\mu) \cdot (x - \mu) + \frac{f''(c) \cdot (x - \mu)^2}{2} \xrightarrow{f''(x) \leq 0} \\
&\leq f(\mu) + f'(\mu) \cdot (x - \mu)
\end{aligned}$$

Using linearity of expectations to the previous equation and based on the identity of  $E[cX] = c[X]$  where  $c$  is a constant the previous inequality turn into:

$$\begin{aligned}
E[f(x)] &\leq E[f(\mu) + f'(\mu) \cdot (x - \mu)] \implies \\
&\leq E[f(\mu)] + f'(\mu) \cdot E[(x - \mu)] \implies \\
&\leq f(\mu) + f'(\mu) \cdot (E[x] - E[\mu]) \implies \\
&\leq f(E[X])
\end{aligned}$$

So Jensen's Inequality for a concave  $f$  function turns into  $E[f(x)] \leq f(E[X])$

b)

We define arithmetic mean of  $n$  numbers as  $A(x_1, \dots, x_n) = \frac{x_1 + \dots + x_n}{n} = \frac{1}{n} \cdot \sum_{i=1}^n x_i$

Also, we define geometric mean of  $n$  numbers as  $G(x_1, \dots, x_n) = (x_1 \cdot \dots \cdot x_n)^{\frac{1}{n}} = \prod_{i=1}^n (x_i)^{\frac{1}{n}}$

We want to prove that  $\frac{1}{n} \cdot \sum_{i=1}^n x_i \geq \prod_{i=1}^n (x_i)^{\frac{1}{n}}$ .

Based on subquestion (a) we define  $f(x)$  as a concave function, so it is true that:

$$\sum_{i=1}^n p_i \cdot f(x_i) \leq f\left(\sum_{i=1}^n p_i \cdot x_i\right) \tag{2}$$

where  $\sum_{i=1}^n p_i = 1$ . We choose  $f(x) = \ln(x)$  since  $\ln(x)$  is a concave function since it is a twice differentiable function with  $f''(x) = -\frac{1}{x^2} \leq 0$ . So equation (1) will turn into:

$$\begin{aligned}
 \sum_{i=1}^n p_i \cdot \ln(x_i) &\leq \ln\left(\sum_{i=1}^n p_i \cdot x_i\right) \implies \\
 \sum_{i=1}^n \ln(x_i^{p_i}) &\leq \ln\left(\sum_{i=1}^n p_i \cdot x_i\right) \implies \\
 \ln\left(\prod_{i=1}^n x_i^{p_i}\right) &\leq \ln\left(\sum_{i=1}^n p_i \cdot x_i\right) \xrightarrow{\ln(x) \nearrow} \\
 \prod_{i=1}^n (x_i)^{p_i} &\leq \sum_{i=1}^n p_i \cdot x_i \xrightarrow{p_i = \frac{1}{n}, \forall i \in [1, \dots, n]} \\
 \prod_{i=1}^n (x_i)^{\frac{1}{n}} &\leq \sum_{i=1}^n \frac{1}{n} \cdot x_i \implies \\
 \prod_{i=1}^n (x_i)^{\frac{1}{n}} &\leq \frac{1}{n} \cdot \sum_{i=1}^n x_i
 \end{aligned}$$

c)

We choose  $f(x) = \sin(x)$  which is a twice differentiable function and  $f''(x) = -\sin(x) \leq 0$ , so  $f$  is a concave function. It is true that:  $\sin(60) = \sin\left(\frac{A+B+C}{3}\right) = \frac{\sqrt{3}}{2}$ . So based on the fact that  $f$  is concave we know that:

$$\begin{aligned}
 E[f(x)] &\leq f(E[X]) \implies \\
 E[\sin(x)] &\leq \sin(E[X]) \xrightarrow{x = \frac{A+B+C}{3}} \\
 E\left[\sin\left(\frac{A+B+C}{3}\right)\right] &\leq \frac{\sqrt{3}}{2} \implies \\
 \frac{\sin(A) + \sin(B) + \sin(C)}{3} &\leq \frac{\sqrt{3}}{2} \implies \\
 \sin(A) + \sin(B) + \sin(C) &\leq 3 \cdot \frac{\sqrt{3}}{2}
 \end{aligned}$$

### 3 Random children

a)

We define  $X$  the random variable of the event that a girl was born after a number of succeeding boy's births. Assuming that  $(1-p)$  is the probability that a boy was born and  $p$  the probability that a girl was born then the probability of the birth of a girl after  $k$  boy's birth is the following:

$$Pr(X = k) = (1-p)^{k-1} \cdot p$$

The expected value  $E[X]$  is given as:

$$\begin{aligned}
 E[X] &= \sum_{k=1}^{\infty} k \cdot Pr(X = k) \implies \\
 &= \sum_{k=1}^{\infty} k \cdot (1-p)^{k-1} \cdot p \implies \\
 &= \frac{p}{1-p} \cdot \sum_{k=1}^{\infty} k \cdot (1-p)^k \xrightarrow{c=1-p} \\
 &= \frac{p}{1-p} \cdot \sum_{k=1}^{\infty} k \cdot c^k \implies \\
 &= \frac{p}{1-p} \cdot c \cdot \sum_{k=1}^{\infty} k \cdot c^{k-1} \implies \\
 &= \frac{p}{1-p} \cdot c \cdot \sum_{k=1}^{\infty} \frac{d}{dc}(c^k) \implies \\
 &= \frac{p}{1-p} \cdot c \cdot \frac{d}{dc} \sum_{k=1}^{\infty} c^k \implies \\
 &= \frac{p}{1-p} \cdot c \cdot \frac{d}{dc} \left( \frac{c}{1-c} \right) \implies \\
 &= \frac{p}{1-p} \cdot c \cdot \frac{1}{(1-c)^2} \implies \\
 &= \frac{1}{p}
 \end{aligned}$$

So the total number of births since the first arrival of a daughter if  $p = \frac{1}{2}$  are equal to  $E[X] = \frac{1}{p} = 2$ . So the expected value of a girl is 1 and the expected value of a boy is also 1.

b)

Following the same formula as in subquestion (a), the total number of births since the first arrival of a daughter is equal to:  $E[X] = \frac{1}{p} = \frac{10}{4} = 2.5$ . So the expected value of a girl is 1 and the expected value of a boy is 1.5.

c)

Assuming that after  $k$  births the number of girls are going to be born are either 1 or 0, since either a girl will be born and stop the process or  $k$  boys will be born and stop the process, the distribution of having a girl is a Bernoulli. So if we define  $X$  the number of girls who are probable to be born then:

$$\begin{aligned}
 E[X] &= 1 \cdot Pr(X = 1) + 0 \cdot Pr(X = 0) \implies \\
 E[X] &= Pr(X = 1) = \sum_{i=1}^k (1-p)^{i-1} \cdot p = 1 - (1-p)^k = 1 - \left(\frac{1}{2}\right)^k
 \end{aligned}$$

Based on the assumption that we define Y the number of boys who are probable to be born then the expected value is:

$$\begin{aligned}
 E[Y] &= \sum_{i=1}^k i \cdot \Pr(Y = i) \implies \\
 E[Y] &= \sum_{i=1}^{k-1} [i \cdot (1-p)^i \cdot p] + k \cdot (1-p)^k \implies \\
 E[Y] &= p \cdot (1-p) \cdot \sum_{i=1}^{k-1} [i \cdot (1-p)^{i-1}] + k \cdot (1-p)^k \implies \\
 E[Y] &= p \cdot (1-p) \cdot \sum_{i=1}^{k-1} \frac{d}{di} \cdot (1-p)^i + k \cdot (1-p)^k \implies \\
 E[Y] &= p \cdot (1-p) \cdot \frac{d}{di} \cdot \sum_{i=1}^{k-1} (1-p)^i + k \cdot (1-p)^k \implies \\
 E[Y] &= 1 - \left(\frac{1}{2}\right)^k
 \end{aligned}$$

d)

Using the same formulas of subquestion (c) the expected value of girls  $E[X]$  is:

$$E[X] = \Pr(X = 1) = \sum_{i=1}^k (1-p)^{i-1} \cdot p = 1 - (1-p)^k = 1 - (0.6)^k$$

While the expected value of boys  $E[Y]$  is:

$$\begin{aligned}
 E[Y] &= \sum_{i=1}^{k-1} [i \cdot (1-p)^i \cdot p] + k \cdot (1-p)^k \implies \\
 E[Y] &= \sum_{i=1}^{k-1} [i \cdot (0.6)^i \cdot 0.4] + k \cdot 0.6^k \implies \\
 E[Y] &= \frac{3}{2} \cdot 5^{-k} \cdot (5^k - 3^k)
 \end{aligned}$$

## 4 Random counter

a)

We define n the number of the counter and k the number of bits we need to represent them. Assuming that the counter works deterministically then:

$$k = \lfloor \log_2(n) \rfloor + 1$$

For example if the number of counter is  $n = 5$  then  $k = \lfloor \log_2(5) \rfloor + 1 = \lfloor 2.32 \rfloor + 1 = 3$ . We will need 3 bits.

b)

With regards to total expectation we need to compute  $E[2^{X_i}]$ . We are going to condition the expected value based on the previous value of  $X$ ,  $X_{i-1}$ . So:

$$E[2^{X_i}] = E[E[2^{X_i} | X_{i-1} = x_{i-1}]] \quad (3)$$

For the inner expected value of the previous equation  $E[2^{X_i} | X_{i-1} = x_{i-1}]$ , we are going to compute it thinking intuitively. In the previous step the counter had a value  $X_{i-1} = x_{i-1}$ . The counter either changed to the next number of the counter  $x_{i-1} + 1$  with a probability of  $p = 2^{-x_{i-1}}$  or stayed in the same value  $x_{i-1}$  with a probability of  $1 - p = 1 - 2^{-x_{i-1}}$ . So, the inner expected value of  $E[2^{X_i} | X_{i-1} = x_{i-1}]$  will be:

$$\begin{aligned} E[2^{X_i} | X_{i-1} = x_{i-1}] &= 2^{x_{i-1}+1} \cdot 2^{-(x_{i-1})} + 2^{x_{i-1}} \cdot (1 - 2^{-x_{i-1}}) \implies \\ E[2^{X_i} | X_{i-1} = x_{i-1}] &= 1 + 2^{x_{i-1}} \end{aligned}$$

Using equation (3):

$$\begin{aligned} E[2^{X_i}] &= E[E[2^{X_i} | X_{i-1} = x_{i-1}]] \implies \\ E[2^{X_i}] &= E[1 + 2^{x_{i-1}}] = 1 + E[2^{x_{i-1}}] \end{aligned}$$

Working iteratively in the previous equation over the  $E[2^{x_{i-1}}]$  terms, it is true that:

$$E[2^{X_i}] = E[1 + 2^{x_{i-1}}] = 1 + E[2^{x_{i-1}}] = 1 + 1 + E[2^{x_{i-2}}] = \dots = i + 1$$

c)

Assuming that:

$$\begin{aligned} 2_i^X &\leq 10 \cdot E[2_i^X] \xrightarrow{E[2_i^X]=m+1, i=m} \\ 2^{X_i} &\leq 10 \cdot (m + 1) \implies \\ X_i &\leq \log_2(10 \cdot (m + 1)) \implies \\ \log_2(X_i) &\leq \log_2(\log_2(10 \cdot (m + 1))) \end{aligned}$$

So, with asymptotic notation the number of bits we will need is  $O(\log_2(\log_2(10 \cdot (m + 1))))$

d)

Following the same idea as in subquestion (b), with regards to total expectation we need to compute  $E[2^{2 \cdot X_i}]$ . We are going to condition the expected value based on the previous value of  $X$ ,  $X_{i-1}$ . So:

$$E[2^{2 \cdot X_i}] = E[E[2^{2 \cdot X_i} | X_{i-1} = x_{i-1}]] \quad (4)$$

The inner expected value of  $E[2^{2 \cdot X_i} | X_{i-1} = x_{i-1}]$  will be:

$$\begin{aligned} E[2^{2 \cdot X_i} | X_{i-1} = x_{i-1}] &= 2^{2 \cdot x_{i-1} + 1} \cdot 2^{-(x_{i-1})} + 2^{2 \cdot x_{i-1}} \cdot (1 - 2^{-x_{i-1}}) \implies \\ E[2^{2 \cdot X_i} | X_{i-1} = x_{i-1}] &= 3 \cdot 2^{x_{i-1}} + 2^{2 \cdot (x_{i-1})} \end{aligned}$$

Using equation (4):

$$\begin{aligned} E[2^{2 \cdot X_i}] &= E[E[2^{2 \cdot X_i} | X_{i-1} = x_{i-1}]] \implies \\ E[2^{2 \cdot X_i}] &= E[3 \cdot 2^{x_{i-1}} + 2^{2 \cdot (x_{i-1}-1)}] = 3 \cdot E[2^{x_{i-1}}] + E[2^{2 \cdot x_{i-1}}] = 3i + E[2^{2 \cdot x_{i-1}}] \end{aligned}$$

Working iteratively in the previous equation over the  $E[2^{2 \cdot x_{i-1}}]$  terms, it is true that:

$$\begin{aligned} E[2^{2 \cdot X_i}] &= E[3 \cdot 2^{x_{i-1}} + 2^{2 \cdot (x_{i-1}-1)}] = 3i + E[2^{2 \cdot x_{i-1}}] = 3i + 3(i-1) + E[2^{2 \cdot x_{i-2}}] \implies \\ E[2^{2 \cdot X_i}] &= 3i + 3(i-1) + 3(i-2) + \dots + 3(i-i) \implies \\ E[2^{2 \cdot X_i}] &= 3 \cdot (i(i+1) - \frac{(i+1)i}{2}) = 3 \cdot \frac{i(i+1)}{2} \end{aligned}$$

As long as the variance we know that  $Var(2^X) = E[2^{2 \cdot X}] - (E[2^X])^2 = 3 \cdot \frac{i(i+1)}{2} - (i+1)^2$