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Chapter

BIVARIATE DISTRIBUTION

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3.1. JOINT PROBABILITY

Two random variable X and Y are said to be jointly distributed if they are defined on same probability space. The joint probability function is denoted by $P_{XY}(x, y)$ or $f_{XY}(x, y)$.

3.2. JOINT PROBABILITY MASS FUNCTION

Let X and Y be random variables on a sample space S with respective image sets $X(S) = \{x_1, x_2, \dots, x_n\}$ and $Y(S) = \{y_1, y_2, \dots, y_m\}$. The function p on $X(S) \times Y(S)$ defined as $p_{ij} = P(X = x_i \cap Y = y_j) = p(x_i, y_j)$ is called *joint probability function* of X and Y where $X(S) \times Y(S) = \{x_1, x_2, \dots, x_n\} \times \{y_1, y_2, \dots, y_m\}$.

The tabular representation is given below:

$x \rightarrow$	$y \downarrow$	y_1	y_2	y_3	...	y_j	...	y_m	Total
x_1		p_{11}	p_{12}	p_{13}	...	p_{1j}	...	p_{1m}	p_1
x_2		p_{21}	p_{22}	p_{23}	...	p_{2j}	...	p_{2m}	p_2
x_3		p_{31}	p_{32}	p_{33}	...	p_{3j}	...	p_{3m}	p_3
\vdots		\vdots	\vdots	\vdots		\vdots		\vdots	\vdots
x_i		p_{i1}	p_{i2}	p_{i3}	...	p_{ij}	...	p_{im}	p_i
\vdots		\vdots	\vdots	\vdots		\vdots		\vdots	\vdots
x_n		p_{n1}	p_{n2}	p_{n3}	...	p_{nj}	...	p_{mn}	p_n
Total		p_1	p_2	p_3	...	p_j	...	p_m	

Also

$$\sum_{i=1}^n \sum_{j=1}^m p(x_i, y_j) = 1$$

$$\text{BIVARIAITE DISTRIBUTION}$$

$$2. (a) F(-\infty, y) = \lim_{x \rightarrow -\infty} F(x, y) = 0 \quad \forall y \in R$$

$$(b) F(x, -\infty) = \lim_{y \rightarrow -\infty} F(x, y) = 0 \quad \forall x \in R$$

150 MARGINAL AND CONDITIONAL PROBABILITY FUNCTION

Suppose the joint distribution of two random variable X and Y is given then the probability distribution of X is determined as follows

$$f_X(x) = p_X(x_i) = P(X = x_i) = p_{i1} + p_{i2} + \dots + p_{ij} + \dots + p_{im}$$

$$= \sum_{j=1}^m p_{ij} = p_i$$

and is known as *marginal probability function of X* .

Similarly

$$f_Y(y) = p_Y(y_j) = P(Y = y_j) = \sum_{i=1}^n p_{ij} = p_j$$

is called *marginal probability function of Y* .

$$f_{XY}(x_i y_j) = P(X = x_i \cap Y = y_j) = \frac{P(X = x_i \cap Y = y_j)}{P(Y = y_j)} = \frac{p_{ij}}{p_j} = \frac{p_{ij}}{p_j}$$

Also

This called *conditional probability function of X when $Y = y_j$* is given

$$\text{Similarly } f_{Y|X}(y|x) = P(Y = y_j | X = x_i) = \frac{P(x_i, y_j)}{p(x_i)} = \frac{p_{ij}}{p_i}$$

is conditional probability function of Y when $X = x_i$ is given

Note: Two random variables X and Y are said to be *independent if*

$$P(X = x_i, Y = y_j) = P(X = x_i) P(Y = y_j)$$

otherwise they are *dependent*.

151 JOINT PROBABILITY DISTRIBUTION FUNCTION

Let (X, Y) be a two dimensional random variable then their joint distribution function is denoted by $F_{X,Y}(x, y)$ and is defined as

$$F_{X,Y}(x, y) = P(X \leq x, Y \leq y) \quad \forall x, y \in R$$

where $\int_{-\infty}^x \int_{-\infty}^y f_{X,Y}(x, y) dx dy = 1$, in case of continuous random variable

$$\text{or } \sum_x \sum_y f_{X,Y}(x, y) = 1, \text{ in case of discrete random variable}$$

where $\sigma_X > 0$ and $\sigma_Y > 0$

Note: $-1 \leq \rho(X, Y) \leq 1$

Conditional Expectation

If (X, Y) are joint discrete random variable then conditional expectation of $g(X, Y)$ given $X = x$ is defined as

1. If $x_1 < x_2$ and $y_1 < y_2$ then

$$P[x_1 < X \leq x_2, y_1 < Y \leq y_2] = F(x_2, y_2) - F(x_2, y_1) - F(x_1, y_2) + F(x_1, y_1) \geq 0 \quad [\text{Rectangle Rule}]$$

152 PROPERTIES OF JOINT DISTRIBUTION FUNCTION

A joint distribution function of two variables has following properties:

1. If $x_1 < x_2$ and $y_1 < y_2$ then

$$\text{In particular } E[Y|X = x] = \sum_j g(x_i, y_j) f_{Y|X}(y_j | X)$$

SOLVED EXAMPLES

bivariate probability distribution of X and Y find:

(iii) $P(Y=4)$

- Example 3.1. For the following bivariate probability distribution of X and Y find:
- (i) $P(X \leq 2, Y=3)$
 - (ii) $P(X < 2, Y \leq 3)$
 - (iv) $P(Y \leq 5)$

$X \backslash Y$	1	2	3	4	5	6
0	0	0	$\frac{1}{32}$	$\frac{2}{32}$	$\frac{2}{32}$	$\frac{3}{32}$
1	$\frac{1}{16}$	$\frac{1}{16}$	$\frac{1}{8}$	$\frac{1}{8}$	$\frac{1}{8}$	$\frac{1}{8}$
2	$\frac{1}{32}$	$\frac{1}{32}$	$\frac{1}{64}$	$\frac{1}{64}$	0	$\frac{2}{64}$

Solution. The marginal distributions are given below:

$X \backslash Y$	1	2	3	4	5	6	$p_{X(x)}$
0	0	0	$\frac{1}{32}$	$\frac{2}{32}$	$\frac{2}{32}$	$\frac{3}{32}$	$\frac{8}{32}$
1	$\frac{1}{16}$	$\frac{1}{16}$	$\frac{1}{8}$	$\frac{1}{8}$	$\frac{1}{8}$	$\frac{1}{8}$	$\frac{10}{32}$
2	$\frac{1}{32}$	$\frac{1}{32}$	$\frac{1}{64}$	$\frac{1}{64}$	0	$\frac{2}{64}$	$\frac{8}{64}$

(i) $P(X \leq 2, Y=3) = P(X=0, Y=3) + P(X=1, Y=3) + P(X=2, Y=3)$

$$= \frac{1}{32} + \frac{1}{16} + \frac{1}{8} = \frac{11}{64}$$

(ii) $P(X \leq 1) = P(X=0) + P(X=1) = \frac{8}{32} + \frac{10}{32} = \frac{7}{8}$

(iii) $P(Y=y) = \frac{13}{64}$

(iv) $P(Y \leq 5) = P(Y=1) + P(Y=2) + P(Y=3) + P(Y=4) + P(Y=5)$

$$\begin{aligned} &= \frac{3}{32} + \frac{3}{32} + \frac{11}{64} + \frac{13}{64} + \frac{6}{32} \\ &= \frac{6+6+11+13+12}{64} = \frac{48}{64} \end{aligned}$$

- Example 3.2. Let X and Y have joint p.d.f.

$X \backslash Y$	-1	0	1
0	b	$2b$	b
1	$3b$	$2b$	b

Find marginal distribution of X and Y . Also find conditional distribution of X given $Y=1$.

Solution.

$Y \backslash X$	-1	0	1	$p_{Y(y)}$
0	b	$2b$	b	$4b$
1	$3b$	$2b$	b	$6b$
$p_{Y(y)}$	$6b$	$5b$	$4b$	$15b$

Marginal distribution of X

$$P(X=-1) = 6b, \quad P(X=0) = 5b, \quad P(X=1) = 4b$$

Marginal distribution of Y

$$P(Y=0) = 4b, \quad P(Y=1) = 6b, \quad P(Y=2) = 5b$$

Conditional distribution of X when $Y=1$

$$P(X=x|Y=1) = \frac{P(X=x \cap Y=1)}{P(Y=1)}$$

$$= \begin{cases} \frac{P(X=-1 \cap Y=1)}{P(Y=1)} \\ \frac{P(X=0 \cap Y=1)}{P(Y=1)} \\ \frac{P(X=1 \cap Y=1)}{P(Y=1)} \end{cases}$$

$$= \begin{cases} \frac{3b}{6b} = \frac{1}{2} & \text{when } X=-1/Y=1 \\ \frac{2b}{6b} = \frac{1}{3} & \text{when } X=0/Y=1 \\ \frac{b}{6b} = \frac{1}{6} & \text{when } X=1/Y=1 \end{cases}$$

$$\begin{aligned} &P(X \leq 2, Y \leq 3) = P(X=0, Y=1) + P(X=0, Y=2) + P(X=1, Y=1) \\ &\quad + P(X=1, Y=2) + P(X=2, Y=1) \\ &= 0 + 0 + \frac{1}{32} + \frac{1}{16} + \frac{1}{16} + \frac{1}{8} = \frac{1+2+2+4}{32} = \frac{9}{32} \end{aligned}$$

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Example 3.3. The joint probability distribution of X and Y is given in the following table.

	Y	1	3	9
X	2	$\frac{1}{8}$	$\frac{1}{24}$	$\frac{1}{12}$
	4	$\frac{1}{4}$	$\frac{1}{4}$	0
	6	$\frac{1}{8}$	$\frac{1}{24}$	$\frac{1}{12}$

(a) Find the marginal probability distribution of Y .

(b) Find the conditional distribution of Y .

(c) Find covariance of X and Y .

(d) Are X and Y independent?

Solution.

$X \backslash Y$	1	3	9	$f_X(x)$
2	$\frac{1}{8}$	$\frac{1}{24}$	$\frac{1}{12}$	$\frac{6}{24}$
4	$\frac{1}{4}$	$\frac{1}{4}$	0	$\frac{2}{4}$
6	$\frac{1}{8}$	$\frac{1}{24}$	$\frac{1}{12}$	$\frac{6}{24}$
$f_Y(y)$	$\frac{4}{8}$	$\frac{8}{24}$	$\frac{2}{12}$	1

(a) Marginal probability distribution of Y

$$P(Y=1) = \frac{4}{8} = \frac{1}{2}, \quad P(Y=2) = \frac{8}{24} = \frac{1}{3}, \quad P(Y=3) = \frac{2}{12} = \frac{1}{6}$$

(b) Conditional probability of Y when $X=4$

$$P(Y=y|X=4) = \frac{P(Y=y \cap X=4)}{P(X=4)}$$

$$P(Y=1|X=4) = \frac{P(Y=1 \cap X=4)}{P(X=4)} = \frac{1}{4} = \frac{1}{2}$$

- Solution.** (i) It is given that X denotes the number of the downturned face of the first tetrahedron and Y the larger of the numbers on the two downturned faces. Then (X, Y) takes the values: $(1, 1), (1, 2), (1, 3), (1, 4), (2, 2), (2, 3), (2, 4), (3, 3), (3, 4), (4, 4)$

$$P(Y=3|X=4) = \frac{P(Y=3 \cap X=4)}{P(X=4)} = \frac{\frac{1}{2}}{\frac{1}{2}} = 1$$

$$P(Y=9|X=4) = \frac{P(Y=9 \cap X=4)}{P(X=4)} = \frac{0}{\frac{1}{2}} = 0$$

$$(c) \text{ To find } \text{Cov}(X, Y) \text{ we will first find } E(XY), E(X) \text{ and } E(Y)$$

$$E(X) = \sum x f_X(x) = 2 \times \frac{6}{24} + 4 \times \frac{2}{4} + 6 \times \frac{6}{24} = 4$$

$$E(Y) = \sum y f_Y(y) = 1 \times \frac{4}{8} + 3 \times \frac{8}{24} + 9 \times \frac{2}{12} = 3$$

$$\begin{aligned} \text{Cov}(X, Y) &= E(XY) - E(X)E(Y) = 12 - 12 = 0 \\ f_{X,Y}(4, 3) &= \frac{1}{4}, \quad f_X(4) = \frac{2}{4}, \quad f_Y(3) = \frac{8}{24} \\ f_X(4)f_Y(3) &= \frac{2}{4} \times \frac{8}{24} = \frac{1}{6} \\ f_{X,Y}(4, 3) &\neq f_X(4)f_Y(3) \end{aligned}$$

$\therefore X$ and Y are not independent.

Example 3.4. Two tetrahedra with sides numbered 1 to 4 are tossed. Let X denote the number on the downturned face of the first tetrahedron and Y the larger of the downturned numbers. Find

- (i) the joint density function of X and Y ,
(ii) the marginal density function of X and Y ,

- (iii) $P[X \leq 2, Y \leq 3]$,

- (iv) the conditional distribution of Y given $X=2$ and $X=3$,

- (v) $E[Y/X=2]$ and $E[Y/X=3]$,

- (vi) $p(X, Y)$

Solution. (i) It is given that X denotes the number of the downturned face of the first tetrahedron

$$P(Y=1|X=4) = \frac{P(Y=1 \cap X=4)}{P(X=4)} = \frac{1}{4} = \frac{1}{2}$$

The joint discrete density function of X and Y is given below:

	(1, 1)	(1, 2)	(1, 3)	(1, 4)	(2, 2)	(2, 3)	(2, 4)	(3, 3)	(3, 4)	(4, 4)
(x, y)	$\frac{1}{16}$	$\frac{1}{16}$	$\frac{1}{16}$	$\frac{1}{16}$	$\frac{2}{16}$	$\frac{1}{16}$	$\frac{3}{16}$	$\frac{1}{16}$	$\frac{4}{16}$	$\frac{4}{16}$
$f_{X,Y}(x, y)$	$\frac{1}{16}$									

We can also write the above data as follows:

$X \backslash Y$	1	2	3	4	$f_Y(x)$
1	$\frac{1}{16}$	0	0	0	$\frac{1}{16}$
2	$\frac{1}{16}$	$\frac{2}{16}$	0	0	$\frac{3}{16}$
3	$\frac{1}{16}$	$\frac{1}{16}$	$\frac{3}{16}$	0	$\frac{5}{16}$
4	$\frac{1}{16}$	$\frac{1}{16}$	$\frac{1}{16}$	$\frac{4}{16}$	$\frac{7}{16}$
$f_X(x)$	$\frac{4}{16}$	$\frac{4}{16}$	$\frac{4}{16}$	$\frac{4}{16}$	1

(i) Marginal distribution of X . From the above table, we get

$$f_X(1) = \frac{4}{16}, \quad f_X(2) = \frac{4}{16}, \quad f_X(3) = \frac{4}{16}, \quad f_X(4) = \frac{4}{16}$$

Marginal distribution of Y . From the above table, we get

$$f_Y(1) = \frac{1}{16}, \quad f_Y(2) = \frac{3}{16}, \quad f_Y(3) = \frac{5}{16}, \quad f_Y(4) = \frac{7}{16}$$

(ii) $P[X \leq 2, Y \leq 3] = P[X = 1, Y = 1] + P[X = 1, Y = 2] + P[X = 1, Y = 3] + P[X = 2, Y = 2]$

$$= \frac{1}{16} + \frac{1}{16} + \frac{1}{16} + \frac{1}{16} = \frac{6}{16} = \frac{3}{8}$$

(iv) Conditional distribution of Y given $X = 3$. We have

$$P[Y = y \cap X = 2] = P[X = 2]P[Y = y | X = 2]$$

$$\therefore P[Y = y | X = 2] = \frac{P[Y = y \cap X = 2]}{P[X = 2]}$$

Now

$$P[Y = 2 | X = 2] = \frac{P[Y = 2 \cap X = 2]}{P[X = 2]} = \frac{2/16}{4/16} = \frac{1}{2},$$

$$P[Y = 3 | X = 2] = \frac{P[Y = 3 \cap X = 2]}{P[X = 2]} = \frac{1/16}{4/16} = \frac{1}{4},$$

$$P[Y = 4 | X = 2] = \frac{P[Y = 4 \cap X = 2]}{P[X = 2]} = \frac{1/16}{4/16} = \frac{1}{4}$$

(v) By definition

$$E[Y | X = 2] = \sum_{y_j} y_j P[Y = y_j | X = 2]$$

$$= 2P[Y = 2 | X = 2] + 3P[Y = 3 | X = 2] + 4P[Y = 4 | X = 2]$$

[Use part (iv)]

$$E[Y | X = 3] = \sum_{y_j} y_j P[Y = y_j | X = 3]$$

= $3P[Y = 3 | X = 3] + 4P[Y = 4 | X = 3]$

[Use part (iv)]

(vi) Now we shall find $\rho(X, Y)$ i.e., the correlation coefficient between X and Y . From the table, we have

$$E[X] = \sum x f_X(x) = 1 \times \frac{4}{16} + 2 \times \frac{4}{16} + 3 \times \frac{4}{16} + 4 \times \frac{4}{16} = \frac{5}{2}$$

$$E[X^2] = \sum x^2 f_X(x) = 1^2 \times \frac{4}{16} + 2^2 \times \frac{4}{16} + 3^2 \times \frac{4}{16} + 4^2 \times \frac{4}{16} = \frac{15}{2}$$

$$E[Y] = \sum y f_Y(y) = 1 \times \frac{1}{16} + 2 \times \frac{3}{16} + 3 \times \frac{5}{16} + 4 \times \frac{7}{16} = \frac{25}{8}$$

$$E[Y^2] = \sum y^2 f_Y(y) = 1^2 \times \frac{1}{16} + 2^2 \times \frac{3}{16} + 3^2 \times \frac{5}{16} + 4^2 \times \frac{7}{16} = \frac{85}{8}$$

$$E[XY] = \sum xy f_{X,Y}(x, y) = \left(1 \times \frac{1}{16} + 2 \times \frac{1}{16} + 3 \times \frac{1}{16} + 4 \times \frac{1}{16} \right) + \left(4 \times \frac{2}{16} + 6 \times \frac{1}{16} + 8 \times \frac{1}{16} \right) + \left(9 \times \frac{3}{16} + 12 \times \frac{1}{16} \right) + 16 \times \frac{4}{16} = \frac{135}{16}$$

$$= \frac{1}{16} + \frac{1}{16} + \frac{1}{16} + \frac{1}{16} + \frac{2}{16} + \frac{1}{16} + \frac{3}{16} + \frac{1}{16} + \frac{4}{16} + \frac{8}{16} + \frac{1}{16} + \frac{3}{16} + \frac{1}{16} + \frac{4}{16} + \frac{16}{16} = \frac{135}{16}$$

$$\text{Var}[X] = E[X^2] - (E[X])^2 = \frac{15}{2} - \frac{25}{4} = \frac{5}{4}$$

$$\text{Var}[Y] = E[Y^2] - (E[Y])^2 = \frac{85}{8} - \left(\frac{25}{8} \right)^2 = \frac{55}{64}$$

$$\text{Cov}(X, Y) = E[XY] - E[X]E[Y] = \frac{135}{16} - \frac{5}{2} \times \frac{25}{8} = \frac{5}{8}$$

Hence,

$$\rho(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}[X]\text{Var}[Y]}} = \frac{5/8}{\sqrt{5/4 \times 55/64}} = \frac{2}{\sqrt{11}}$$

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Example 3.5. Three fair coins are tossed. Let X denote the number of heads on the first two coins, and let Y denote the numbers of tails on the last two coins.

(a) Find the joint distribution of Y given that $X = 1$.

(b) Find the conditional distribution of Y .

(c) Find $\text{Cov}(X, Y)$.

(d) Find $\rho(X, Y)$.

Solution. Eight (2^3) possible outcomes of the experiment are:

	BHH	HHT	HTH	HTT	THH	TTT	TTH	THT
$X:$	2	2	1	1	1	0	0	1
$Y:$	0	1	1	2	0	2	1	1

(a) The joint distribution of X and Y is given below:

$Y \backslash X$	0	1	2	$f_Y(y)$
0	0	$\frac{1}{8}$	$\frac{1}{8}$	$\frac{2}{8} = \frac{1}{4}$
1	$\frac{1}{8}$	$\frac{2}{8}$	$\frac{1}{8}$	$\frac{4}{8} = \frac{1}{2}$
2	$\frac{1}{8}$	$\frac{1}{8}$	0	$\frac{2}{8} = \frac{1}{4}$
$f_X(x)$	$\frac{2}{8} = \frac{1}{4}$	$\frac{4}{8} = \frac{1}{2}$	$\frac{2}{8} = \frac{1}{4}$	1

(b) To find the conditional distribution of Y given that $X = 1$

i.e., to find $f_{Y|X}(y|X) = \frac{f_{YX}(y|X)}{f_X(1)} = 2f(1, y)$, since $f_X(1) = \frac{1}{2}$

$Y \backslash X$	0	1	2	$f_Y(y)$
0	0	$\frac{1}{6}$	0	$\frac{2}{6} = \frac{1}{3}$
1	$\frac{2}{6}$	0	$\frac{2}{6} = \frac{1}{3}$	$\frac{4}{6} = \frac{2}{3}$
2	0	$\frac{2}{6} = \frac{1}{3}$	0	$\frac{2}{6} = \frac{1}{3}$
$f_X(x)$	$\frac{4}{6} = \frac{2}{3}$	$\frac{2}{6} = \frac{1}{3}$	$\frac{2}{6} = \frac{1}{3}$	1

(a) The joint discrete density function of X and Y is given below:

$Y \backslash X$	0	1	2	$f_{Y X}(y X=1)$
0	0	$\frac{1}{6}$	0	$\frac{2}{6} = \frac{1}{3}$
1	$\frac{2}{6}$	0	$\frac{2}{6} = \frac{1}{3}$	$\frac{4}{6} = \frac{2}{3}$
2	0	$\frac{2}{6} = \frac{1}{3}$	0	$\frac{2}{6} = \frac{1}{3}$

(b) The conditional distribution of Y given $X = 1$ is given by

$$f_{Y|X}(y|X=1) = \frac{f_{YX}(y|X=1)}{f_X(1)} = \frac{f(1, y)}{4/6} = \frac{3}{2} f(1, y)$$

$$\begin{aligned} f_{Y|X}(X=1) &= \begin{cases} 2 \times \frac{1}{8} = \frac{1}{4} & \text{at } y=0, \\ 2 \times \frac{2}{8} = \frac{1}{4} & \text{at } y=1, \\ 2 \times \frac{1}{8} = \frac{1}{4} & \text{at } y=2, \end{cases} \\ &= \begin{cases} \frac{3}{2} \times \frac{2}{6} = \frac{1}{2}, & y=2 \\ \frac{3}{2} \times \frac{2}{6} = \frac{1}{2}, & y=3. \end{cases} \end{aligned}$$

(c) $E[X] = 0 \times \frac{1}{4} + 1 \times \frac{1}{2} + 2 \times \frac{1}{4} = 1$. Similarly, $E[Y] = 1$.

$$E[XY] = 1 \times \frac{2}{8} + 2 \times \frac{1}{8} + 2 \times \frac{1}{8} + 4 \times 0 = \frac{6}{8} = \frac{3}{4}$$

Hence,

$$\text{cov}(X, Y) = E[XY] - E[X]E[Y] = \frac{3}{4} - 1 = -\frac{1}{4}$$

$$\text{E}[X^2] = 0^2 \times \frac{1}{4} + 1^2 \times \frac{1}{2} + 2^2 \times \frac{1}{4} = \frac{3}{2}$$

$$\text{Var}[X] = \text{E}[X^2] - [E[X]]^2 = \frac{3}{2} - 1 = \frac{1}{2}. \text{ Also } \text{var}[Y] = \frac{1}{2}$$

$$\rho(X, Y) = \frac{\text{cov}(X, Y)}{\sigma_X \sigma_Y} = \frac{-1/4}{1/2} = -0.5$$

Hence,

Example 3.6. Consider a sample of size 2 drawn without replacement from an urn containing three balls, numbered 1, 2 and 3. Let X be the smaller of the two numbers drawn and Y the larger.

(a) Find the joint discrete density function of X and Y .

(b) Find the conditional distribution of Y given $X = 1$.

(c) Find $\rho(X, Y)$.

Solution. (a) (X, Y) takes the values:

(1, 2) (1, 3) (2, 1) (2, 3) (3, 1) (3, 2)

The distributions of X and Y are given below:

X	1	1	2	2	1	2
Y	2	3	2	3	3	3

The distributions of X and Y are given below:

$Y \backslash X$	2	2	2	$f_Y(y)$
2	0	$\frac{2}{6}$	0	$\frac{2}{6} = \frac{1}{3}$
3	$\frac{2}{6}$	0	$\frac{2}{6} = \frac{1}{3}$	$\frac{4}{6} = \frac{2}{3}$
$f_X(x)$	$\frac{4}{6} = \frac{2}{3}$	$\frac{2}{6} = \frac{1}{3}$	$\frac{2}{6} = \frac{1}{3}$	1

(a) The joint discrete density function of X and Y is given below:

$Y \backslash X$	2	2	2	$f_{Y X}(y X=1)$
2	0	$\frac{2}{6}$	0	$\frac{2}{6} = \frac{1}{3}$
3	$\frac{2}{6}$	0	$\frac{2}{6} = \frac{1}{3}$	$\frac{4}{6} = \frac{2}{3}$
$f_X(x)$	$\frac{4}{6} = \frac{2}{3}$	$\frac{2}{6} = \frac{1}{3}$	$\frac{2}{6} = \frac{1}{3}$	1

(b) The conditional distribution of Y given $X = 1$ is given by

$$f_{Y|X}(y|X=1) = \frac{f_{YX}(y|X=1)}{f_X(1)} = \frac{f(1, y)}{4/6} = \frac{3}{2} f(1, y)$$

$$\begin{aligned} f_{Y|X}(X=1) &= \begin{cases} 2 \times \frac{1}{8} = \frac{1}{4} & \text{at } y=2, \\ 2 \times \frac{2}{8} = \frac{1}{4} & \text{at } y=3. \end{cases} \\ &= \begin{cases} \frac{3}{2} \times \frac{2}{6} = \frac{1}{2}, & y=2 \\ \frac{3}{2} \times \frac{2}{6} = \frac{1}{2}, & y=3. \end{cases} \end{aligned}$$

$$\overline{\overline{160}} \quad E[X] = 1 \times \frac{4}{6} + 2 \times \frac{2}{6} = \frac{4}{3}, E[X^2] = 1^2 \times \frac{4}{6} + 2^2 \times \frac{2}{6} = 2,$$

$$(c) \quad E[Y] = 2 \times \frac{2}{6} + 3 \times \frac{4}{6} = \frac{8}{3}, E[Y^2] = 2^2 \times \frac{2}{6} + 3^2 \times \frac{4}{6} = \frac{22}{3}$$

$$E[XY] = 2 \times \frac{2}{6} + 3 \times \frac{2}{6} + 4 \times 0 + 6 \times \frac{2}{6} = \frac{11}{3}$$

$$\therefore \text{Cov}(X, Y) = E[XY] - E[X]E[Y] = \frac{11}{3} - \frac{4}{3} \times \frac{8}{3} = \frac{1}{9}$$

$$\text{var}[X] = E[X^2] - \{E[X]\}^2 = 2 - \left(\frac{4}{3}\right)^2 = \frac{2}{9}$$

Now

$$\text{var}[Y] = E[Y^2] - \{E[Y]\}^2 = \frac{22}{3} - \left(\frac{8}{3}\right)^2 = \frac{2}{9}$$

$$\text{Hence, } \rho(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{\text{var}(X)\text{var}(Y)}} = \frac{1/9}{\sqrt{2/9}\sqrt{2/9}} = \frac{1}{9} \times \frac{9}{2} = \frac{1}{2} = 0.5$$

Example 3.7. X and Y are two random variable having joint density function $= \frac{1}{27}(2x+y)$,

where x and y can assume only integer values 0, 1 and 2. Find conditional distribution of Y for $X=x$.

Solution. $f(x, y) = \frac{1}{27}(2x+y)$ $x=0, 1, 2; y=0, 1, 2$

The joint probability distribution is given below

$X \backslash Y$	0	1	2	$f_{X,Y}(x, y)$
0	0	$\frac{1}{27}$	$\frac{2}{27}$	$\frac{3}{27}$
	1	$\frac{2}{27}$	$\frac{3}{27}$	$\frac{4}{27}$
	2	$\frac{4}{27}$	$\frac{5}{27}$	$\frac{6}{27}$
$f_Y(y)$	$\frac{6}{27}$	$\frac{9}{27}$	$\frac{12}{27}$	1

The marginal probability distribution of X is

$$f_X(x) = \sum_y f(x, y)$$

$$f_X(0) = \frac{3}{27}, f_X(1) = \frac{9}{27}, f_X(2) = \frac{15}{27}$$

The condition distribution of Y for $X=x$ is given by
 $f_{Y|X}(y|x) = \frac{f(x, y)}{f_X(x)}$

$X \backslash Y$	0	1	2
0	0	$\frac{1}{3}$	$\frac{2}{3}$
1	$\frac{2}{9}$	$\frac{3}{9}$	$\frac{4}{9}$
2	$\frac{4}{15}$	$\frac{5}{15}$	$\frac{6}{15}$

EXERCISE 3.1

1. Two tetrahedral with sides numbered 1 to 4 are tossed. Let X denote the smaller of down turned face of two tetrahedra and Y the larger of two.

- (a) Find joint density function of X and Y

- (b) Find $P(X \geq 2, Y \geq 2)$

- (c) Find mean and variance of X and Y.

- (d) Find the conditional distribution of Y given X for each possible value of X.

- (e) Find correlation coefficient of X and Y.

[Ans. (b) $\frac{9}{16}$ (c) $\bar{X} = \frac{15}{8}$, $\bar{Y} = \frac{25}{8}$, variance $X = \frac{55}{64}$ variance $Y = \frac{55}{64}$ (e) 0.45]

2. Three coins are tossed. Let X denote the number of heads on first two and Y denote the number of heads on last two.

- (a) Find joint distribution of X and Y.

- (b) Find $E[Y|X=1]$ [Ans. (a)

[Ans. (a)

$X \backslash Y$	0	1	2
0	$\frac{1}{8}$	$\frac{1}{8}$	0
1	$\frac{1}{8}$	$\frac{2}{8}$	$\frac{1}{8}$
2	0	$\frac{1}{8}$	$\frac{1}{8}$

- (b) 1 (c) 0.51

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3. Consider a sample of size 2 drawn without replacement from an urn containing three balls, numbered 1, 2 and 3. Let X be the number on the first ball drawn and Y the larger of the two numbers drawn.

(a) Find joint discrete density function of X and Y .(b) Find $P[X = 1 | Y = 3]$ (c) Find $\text{cov } p[X, Y]$ (d) Find $p[X, Y]$

[Ans. (a)

$Y \backslash X$	1	2	3
2	$\frac{1}{6}$	$\frac{1}{6}$	0
1	$\frac{1}{6}$	$\frac{1}{6}$	$\frac{2}{6}$
3	$\frac{1}{6}$	$\frac{1}{6}$	$\frac{2}{6}$

$$(b) \frac{1}{4} \quad (c) \frac{1}{6} \quad (d) \frac{\sqrt{3}}{4}$$

4. An urn contain four balls: Two of the balls are numbered with 1 and other two are numbered with 2. Two balls are drawn from the urn without replacement. Let X denote the smaller of the numbers on drawn balls and Y the larger:

- (a) Find joint density of X and Y
 (b) Find marginal distribution of Y
 (c) Find $\text{cov}(X, Y)$

[Ans. (a)

$Y \backslash X$	1	2
1	$\frac{1}{6}$	0
2	$\frac{4}{6}$	$\frac{1}{6}$

$$(b) \frac{5}{6} \quad (c) \frac{1}{5}$$

5. The joint density of X and Y is

$$f(x, y) = \frac{x+2y}{27}$$

- where x and y can assume only integer values 0, 1, 2. Find conditional distribution of Y for

$$X = x.$$

$$\boxed{\text{Ans. } f_{Y/X}(y/x) = \frac{x+2y}{3x+6}, x, y = 1, 2, 3}$$

- b. Two discrete random variable X and Y have
- $$P[X = 0, Y = 0] = \frac{2}{9}, \quad P[X = 0, Y = 1] = \frac{1}{9}$$
- $$P[X = 1, Y = 0] = \frac{1}{9}, \quad P[X = 1, Y = 1] = \frac{5}{9}$$

Examine whether X and Y are independent.

[Ans. No]

36. JOINT CONTINUOUS DENSITY FUNCTION

A two dimensional random variable (X, Y) is said to be continuous iff there exist a function

$$f_{X,Y}(x, y) \geq 0 \text{ such that}$$

$$F_{X,Y}(x, y) = \int_{-\infty}^x \int_{-\infty}^y f(u, v) du dv$$

The function $f_{X,Y}(x, y)$ or $f(x, y)$ is called joint probability density function which satisfy following properties

$$(i) f(x, y) \geq 0 \text{ for all } (x, y) \in \mathbb{R}^2$$

$$(ii) \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = 1$$

Note: $F_{X,Y}(x, y) = P(X \leq x, Y \leq y)$

37. MARGINAL AND CONDITIONAL PROBABILITY DENSITY FUNCTION

If X and Y are jointly continuous random variable with joint probability density function $f_{X,Y}(x, y)$ then $f_X(x)$ and $f_Y(y)$ are called marginal probability density function and are defined as

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx$$

The conditional probability function of Y given $X = x$ is defined as

$$f_{Y|X}(y/x) = \frac{f_{X,Y}(x, y)}{f_X(x)} \text{ if } f_X(x) > 0$$

$$f_{X|Y}(x/y) = \frac{f_{X,Y}(x, y)}{f_Y(y)} \text{ if } f_Y(y) > 0$$

Similarly

$$(b) \frac{5}{6} \quad (c) \frac{1}{5}$$

38. CONDITIONAL CUMULATIVE DISTRIBUTION

If X and Y are joint continuous random variables with joint probability density function $f_{X,Y}(x, y)$ then conditional cumulative distribution of Y , given $X = x$ is defined as

$$F_{Y|X}(y/x) = \int_{-\infty}^y f_{Y|X}(z/x) dz$$

Similarly

$$F_{X|Y}(x/y) = \int_{-\infty}^x f_{X|Y}(z/y) dz$$

for all y such that $f_Y(y) > 0$.

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Note: If X and Y are independent then
 $f_{X,Y}(x,y) = f_X(x)f_Y(y)$

$$\begin{aligned} f_{X,Y}(x,y) &= f_X(x) \\ f_{Y|X}(y|x) &= f_Y(y) \end{aligned}$$

Expectation

$f(X, Y)$ be a two dimensional continuous random variable with joint probability density function
 $f(X, Y)$. The expectation of $g(X, Y)$ denoted as $E[g(X, Y)]$

$$E[g(X, Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f(x, y) dx dy$$

In particular,

$$E(X) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x f(x, y) dx dy$$

$$E(Y) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} y f(x, y) dx dy$$

$$\begin{aligned} &= k \int_1^2 x \left[\int_x^2 y dy \right] dx = k \int_1^2 x \cdot \frac{1}{2} (4-x^2) dx \\ &= \frac{k}{2} \int_1^2 (4x - x^3) dx = \frac{k}{2} \left[2x^2 - \frac{x^4}{4} \right]_1^2 \\ &= \frac{k}{2} \left[2(4-1) - \frac{1}{4}(16-1) \right] = \frac{k}{2} \left(6 - \frac{15}{4} \right) = \frac{9k}{8} \end{aligned}$$

$$E(XY) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy f(x, y) dx dy$$

Joint Moment Generating Function

The joint moment generating function of a two-dimensional random variable (X, Y) , denoted by
 $m_{X,Y}(t_1, t_2)$, is defined as

$$m_{X,Y}(t_1, t_2) = E[e^{t_1 X + t_2 Y}],$$

where t_1 and t_2 are real parameters.

Remark 1. We observe that

$$M_X(t_1) = M_{X,Y}(t_1, 0) = \lim_{t_2 \rightarrow 0} M_{X,Y}(t_1, t_2)$$

$$M_Y(t_2) = M_{X,Y}(0, t_2) = \lim_{t_1 \rightarrow 0} M_{X,Y}(t_1, t_2)$$

Recall that $M_X(t_1), M_Y(t_2)$ are moment generating functions of X and Y , respectively.

Remark 2. If X and Y are independent random variables, then

$$M_{X,Y}(t_1, t_2) = M_X(t_1) \cdot M_Y(t_2).$$

Proof. We have

$$M_{X,Y}(t_1, t_2) = E[e^{t_1 X + t_2 Y}] = E[e^{t_1 X} \cdot e^{t_2 Y}].$$

Taking $g_1(x) = e^{t_1 x}$ and $g_2(y) = e^{t_2 y}$ and using the independence of X and Y , it follows that

$$\begin{aligned} M_{X,Y}(t_1, t_2) &= E[e^{t_1 X}] \cdot E[e^{t_2 Y}] \\ &= M_X(t_1) \cdot M_Y(t_2). \end{aligned}$$

SOLVED EXAMPLES

Example 3.8. Find k so that $f(x, y) = kxy, 1 \leq x \leq y \leq 2$ will be a joint probability density function. We have

$$\begin{aligned} 1 &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = \int_1^2 \int_x^2 kxy dx dy \\ &= k \int_1^2 x \left[\int_x^2 y dy \right] dx = k \int_1^2 x \cdot \frac{1}{2} (4-x^2) dx \\ &= \frac{k}{2} \int_1^2 (4x - x^3) dx = \frac{k}{2} \left[2x^2 - \frac{x^4}{4} \right]_1^2 \\ &= \frac{k}{2} \left[2(4-1) - \frac{1}{4}(16-1) \right] = \frac{k}{2} \left(6 - \frac{15}{4} \right) = \frac{9k}{8} \end{aligned}$$

Hence, $1 = \frac{9k}{8} \Rightarrow k = \frac{8}{9}$.

Example 3.9. Find K so that $f(x, y) = K(x+y), 0 < x < 1$ and $0 < y < 1$, is a joint probability density function.

Solution. We have

$$\begin{aligned} 1 &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = \int_0^1 \int_0^1 K(x+y) dx dy \\ &= K \int_0^1 \left[\left(xy + \frac{1}{2} y^2 \right) \right]_0^1 dx = K \int_0^1 \left(x + \frac{1}{2} \right) dx \\ &= K \left(\frac{1}{2} + \frac{1}{2} \right) = K. \text{ Hence } K = 1. \end{aligned}$$

Example 3.10. Let the joint p.d.f of X and Y be

$$\begin{aligned} f(x, y) &= (x+y), 0 \leq x \leq 1, 0 \leq y \leq 1 \\ &= 0, \text{ otherwise.} \end{aligned}$$

Find $P[0 < X < \frac{1}{2}, 0 < Y < \frac{1}{4}]$, $E[X], E[Y], E[XY], E[X+Y], \rho[X, Y]$.

Solution. We have

$$\begin{aligned} P[0 < X < \frac{1}{2}, 0 < Y < \frac{1}{4}] &= \int_0^{1/2} \left[\int_0^{1/4} (x+y) dy \right] dx = \int_0^{1/2} \left[xy + \frac{1}{2} y^2 \right]_0^{1/4} dx \\ &= \int_0^{1/2} \left(\frac{1}{4}x + \frac{1}{32} \right) dx = \frac{1}{4} \times \frac{1}{8} + \frac{1}{64} = \frac{3}{64}. \end{aligned}$$

$$E[Y^2] = \frac{5}{12}.$$

$$E[X] = \int_0^1 \int_0^1 xy(x+y) dx dy = \int_0^1 \int_0^1 x(x+y) dx dy$$

$$\text{var}[X] = E[X^2] - (E[X])^2 = \frac{5}{12} - \left(\frac{7}{12}\right)^2 = \frac{11}{144}.$$

Now

$$\text{var}[Y] = \frac{11}{144}$$

Similarly,

$$E[Y] = \int_0^1 \int_0^1 yf(x,y) dx dy = \frac{7}{12}.$$

Similarly,

$$E[XY] = \int_0^1 \int_0^1 xyf(x,y) dx dy \int_0^1 \int_0^1 xy(x+y) dx dy$$

Now

$$= \int_0^1 \left[x^2 \left(\frac{1}{2}y^2 \right) + x \left(\frac{1}{3}y^3 \right) \right] dx = \int_0^1 \left(\frac{1}{2}x^2 + \frac{1}{3}x \right) dx$$

$$= \frac{1}{6} + \frac{1}{6} = \frac{1}{3}.$$

$$E[X+Y] = \int_0^1 \int_0^1 (x+y)f(x,y) dx dy = \int_0^1 \int_0^1 (x+y)^2 dx dy$$

$$= \int_0^1 \int_0^1 (x^2 + 2xy + y^2) dx dy$$

$$= \int_0^1 \left[x^2y + \left(\frac{1}{3}y^3 \right) \right] dx \\ = \int_0^1 \left(x^2 + x + \frac{1}{3} \right) dx = \frac{1}{3} + \frac{1}{2} + \frac{1}{3} = \frac{7}{6}.$$

$$\text{We have } \text{cov}[X, Y] = E[XY] - E[X]E[Y] = \frac{1}{3} - \left(\frac{7}{12}\right)^2 = -\frac{1}{144}.$$

Now

$$E[X^2] = \int_0^1 \int_0^1 x^2 f(x,y) dx dy = \int_0^1 \int_0^1 x^2 (x+y) dx dy$$

$$= \int_0^1 \left[x^3 y + x^2 \left(\frac{1}{2}y^2 \right) \right] dx$$

$$= \int_0^1 \left(x^3 + \frac{1}{2}x^2 \right) dx = \frac{1}{4} + \frac{1}{6} = \frac{5}{12}$$

- (i) Find the marginal density functions of X and Y.
(ii) Find the conditional density function of Y given X = x and that of X given Y = y
(iii) Are X and Y independent?

Solution. It may be observed that

$$f(x,y) \geq 0 \text{ and } \int_0^1 \int_0^x 2 dy dx = \int_0^1 2x dx = 1$$

(i) The marginal density function of X is

$$f_X(x) = \int_{-\infty}^{\infty} f(x,y) dy = \int_0^x 2 dy = 2x, 0 < x < 1$$

$$= 0, \text{ elsewhere.}$$

The marginal density function of Y is

$$f_Y(y) = \int_{-\infty}^{\infty} f(x,y) dx = \int_y^1 2 dx = 2(1-y), 0 < y < 1$$

$$= 0, \text{ elsewhere.}$$

(ii) The conditional density function of Y given X = x is

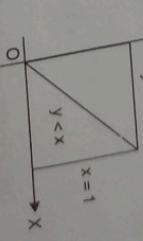
$$f_{Y|X}(y|x) = \frac{f(x,y)}{f_X(x)} = \frac{2}{2x} = \frac{1}{x}, 0 < x < 1.$$

The conditional density function of X given Y = y is

$$f_{X|Y}(x|y) = \frac{f(x,y)}{f_Y(y)} = \frac{2}{2(1-y)} = \frac{1}{1-y}, 0 < y < 1.$$

$$(iii) f_X(x)f_Y(y) = 2x \times 2(1-y) = 4x(1-y) \neq f(x,y).$$

Hence X and Y are not independent.



Example 3.12. If X and Y are two random variables having joint density function,

$$\begin{aligned} f(x, y) &= \frac{1}{8}(6-x-y), & 0 < x < 2, & 2 < y < 4 \\ &= 0, \text{ otherwise.} \end{aligned}$$

- (i) $P[X < 1 \cap Y < 3]$, (ii) $P[X + Y < 3]$, (iii) $P[X < 1 \mid Y < 3]$.
Find also the marginal and conditional distributions.

Solution. We have

$$\begin{aligned} (i) P[X < 1 \cap Y < 3] &= \int_0^1 \int_0^3 f(x, y) dx dy \\ &= \int_0^1 \int_0^3 \frac{1}{8}(6-x-y) dx dy \\ &= \frac{1}{8} \int_0^1 \left\{ (6-x) - \frac{1}{2}(9-4) \right\} dx = \frac{1}{8} \int_0^1 \left(\frac{7}{2} - x \right) dx \\ &= \frac{1}{8} \left(\frac{7}{2} - \frac{1}{2} \right) = \frac{3}{8} \end{aligned}$$

$$(ii) P[X + Y < 3] = \int_0^1 \int_{2-x}^{3-x} \frac{1}{8}(6-x-y) dx dy$$

$$\begin{aligned} &= \frac{1}{8} \int_0^1 \left[(6-x)(1-x) - \frac{1}{2}((3-x)^2 - 4) \right] dx \\ &= \frac{1}{16} \int_0^1 (x^2 - 8x + 7) dx = \frac{1}{16} \left(\frac{1}{3} - 4 + 7 \right) = \frac{5}{24} \end{aligned}$$

- (iii) Firstly, we calculate $P[Y < 3]$. We have

$$\begin{aligned} P[Y < 3] &= \int_0^2 \int_0^3 \frac{1}{8}(6-x-y) dx dy \\ &= \frac{1}{8} \int_0^2 \left\{ (6-x)(1-x) - \frac{1}{2}((3-x)^2 - 4) \right\} dx \\ &= \frac{1}{16} \int_0^2 (x^2 - 8x + 7) dx = \frac{1}{16} \left(\frac{1}{3} - 4 + 7 \right) = \frac{5}{24}. \end{aligned}$$

Now $P[X < 1 \mid Y < 3] = \frac{P[X < 1 \cap Y < 3]}{P[Y < 3]} = \frac{3/8}{5/8} = \frac{3}{5}$.

- (iv) The marginal distribution of X is given by

$$f_X(x) = \int_0^4 f(x, y) dy = \frac{1}{8} \int_0^4 (6-x-y) dy$$

Now

$$E[X] = \int_0^4 x f_X(x) dx = \int_0^4 x \left(x + \frac{1}{2} \right) dx = \frac{1}{3} + \frac{1}{4} = \frac{7}{12}.$$

$$\begin{aligned} \text{The marginal distributions of } Y \text{ is given by} \\ &= \int_0^2 \frac{1}{8}(6-x-y) dx = \frac{1}{8} \left[6x - 2 - \frac{1}{2}x^2 - 2y \right] \\ &= \begin{cases} \frac{1}{4}(5-y), & 2 < y < 4 \\ 0, & \text{otherwise.} \end{cases} \end{aligned}$$

The conditional distributions of X and Y are

$$\begin{aligned} f_{X|Y}(x \mid y) &= \frac{f(x, y)}{f_Y(y)} = \frac{6-x-y}{2(5-y)}, & 0 < x < 2, & 2 < y < 4 \\ f_{Y|X}(y \mid x) &= \frac{f(x, y)}{f_X(y)} = \frac{6-x-y}{2(3-x)}, & 0 < x < 2, & 2 < y < 4. \end{aligned}$$

Example 3.13. The probability density function of a continuous bivariate distribution is given by

$$\begin{aligned} f(x, y) &= x + y, \text{ where } 0 \leq x \leq 1, 0 \leq y \leq 1 \\ &= 0, \text{ otherwise.} \end{aligned}$$

Find the marginal distributions and the correlation coefficient of x and y .

Solution. The marginal distribution of X is

$$f_X(x) = \int_0^{\infty} f(x, y) dy = \int_0^1 (x+y) dy$$

$$= \left[xy + \frac{1}{2}y^2 \right]_0^1 = x + \frac{1}{2}.$$

$$f_X(x) = x + \frac{1}{2}, \quad 0 \leq x \leq 1$$

$$= 0, \text{ otherwise}$$

$$f_Y(y) = y + \frac{1}{2}, \quad 0 \leq y \leq 1$$

$$= 0, \text{ otherwise.}$$

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$$\text{By symmetry, } E[X] = E[Y] = \frac{7}{12}.$$

$$\begin{aligned} E[XY] &= \int_0^1 \int_0^1 f_{XY}(x+y) dx dy = \int_0^1 \left(x^2 \cdot \frac{1}{2} + x \cdot \frac{1}{3} \right) dx \\ &= \frac{1}{6} + \frac{1}{6} = \frac{1}{3}. \end{aligned}$$

$$\begin{aligned} \text{cov}(X, Y) &= E[XY] - E[X]E[Y] = \frac{1}{3} - \frac{7}{12} \times \frac{7}{12} = -\frac{1}{144}. \\ \therefore E[X^2] &= \int_0^1 x^2 f_X(x) dx = \int_0^1 x^2 \left(x + \frac{1}{2} \right) dx = \frac{1}{4} + \frac{1}{6} = \frac{5}{12} \end{aligned}$$

Now

$$\begin{aligned} \text{var}[X] &= E[X^2] - \{E[X]\}^2 = \frac{5}{12} - \left(\frac{7}{12}\right)^2 = \frac{11}{144}. \\ \therefore \text{var}[Y] &= E[X^2] - \{E[Y]\}^2 = \frac{11}{144}. \end{aligned}$$

$$\begin{aligned} \text{By symmetry, } \text{var}[Y] &= \frac{11}{144}. \\ \text{The correlation coefficient of } X \text{ and } Y \text{ is given by} \end{aligned}$$

$$\rho(X, Y) = \frac{\text{cov}(X, Y)}{\sigma_X \sigma_Y} = \frac{-1/144}{11/144} = -\frac{1}{11}.$$

Example 3.14. Two random variables X and Y have the joint density

$$f(x, y) = 2 - x - y, 0 < x < 1, 0 < y < 1$$

Find the marginal and conditional density functions of X and Y , and show that $\rho[X, Y] = -1/11$; find $E[X|y]$.

Solution. The marginal density function of X is given by

$$f_X(x) = \int_x^\infty f(x, y) dy = \int_x^1 (2 - x - y) dy = 2 - x - \frac{1}{2} = \frac{3}{2} - x.$$

Thus

$$f_X(x) = \begin{cases} \frac{3}{2} - x, & 0 < x < 1 \\ 0, & \text{otherwise.} \end{cases}$$

By symmetry,

$$\begin{aligned} f_Y(y) &= \frac{3}{2} - y, 0 < y < 1 \\ &= 0, \text{ otherwise.} \end{aligned}$$

The conditional density functions of X and Y are

$$f_{XY}(x|y) = \frac{f(x, y)}{f_Y(y)} = \frac{2 - x - y}{\frac{3}{2} - y}, 0 < x, y < 1;$$

$$f_{YX}(y|x) = \frac{f(x, y)}{f_X(x)} = \frac{2 - x - y}{\frac{3}{2} - x}, 0 < x, y < 1.$$

Now

$$\begin{aligned} E[X|y] &= \int_0^1 x \frac{f(x, y)}{f_Y(y)} dx = \int_0^1 \frac{x(2-x-y)}{\frac{3}{2}-y} dx \\ &= \frac{1}{\left(\frac{2}{3}-y\right)} \left[\frac{1}{2} - \frac{1}{3} - \frac{1}{2}y \right] = \frac{4-3y}{9-6y}, 0 < y < 1. \end{aligned}$$

Now we shall determine the correlation coefficient of X and Y

$$E[X] = \int_0^1 x f_X(x) dx = \int_0^1 x \left(\frac{3}{2} - x \right) dx = \frac{3}{2} \cdot \frac{1}{2} - \frac{1}{3} = \frac{5}{12}.$$

$$\begin{aligned} \text{By symmetry, } E[Y] &= E[X] = \frac{5}{12}. \\ \text{Now} \quad E[XY] &= \int_0^1 \int_0^1 xy(2-x-y) dx dy \\ &= \int_0^1 \left(2x \cdot \frac{1}{2} - x^2 \cdot \frac{1}{2} - x \cdot \frac{1}{3} \right) dx = \frac{1}{2} - \frac{1}{6} - \frac{1}{6} = \frac{1}{6}. \end{aligned}$$

$$\begin{aligned} \text{cov}(X, Y) &= E[XY] - E[X]E[Y] = \frac{1}{6} - \frac{5}{12} \cdot \frac{5}{12} = -\frac{1}{144}. \\ \therefore E[X^2] &= \int_0^1 x^2 f_X(x) dx = \int_0^1 x^2 \left(\frac{3}{2} - x \right) dx \\ &= \frac{3}{2} \cdot \frac{1}{2} - \frac{1}{4} = \frac{1}{4}. \end{aligned}$$

$$\text{var}[X] = E[X^2] - \{E[X]\}^2 = \frac{1}{4} - \left(\frac{5}{12}\right)^2 = \frac{11}{144}.$$

tion.

$$\begin{aligned} \text{var}[Y] &= E[Y^2] - \{E[Y]\}^2 = \frac{1}{4} - \left(\frac{5}{12}\right)^2 = \frac{11}{144}. \end{aligned}$$

in

$$\begin{aligned} \rho(X, Y) &= \frac{\text{cov}(X, Y)}{\sigma_X \sigma_Y} = \frac{-1/144}{11/144} = -\frac{1}{11}. \end{aligned}$$

$$\text{Hence, } \rho(X, Y) = -\frac{1}{11}.$$

Hence,

EXERCISE 3.2

1. If $f(x, y) = e^{-(x+y)}$, $x \geq 0, y \geq 0$
 $= 0$ elsewhere

is joint probability density function of random variable X and Y , find:

- (a) $P[X < 1]$
(b) $P[X > 4]$
(c) $P[X + Y < 1]$
[Ans. (a) $1 - e^{-1}$, (b) $\frac{1}{2}$, (c) $1 - 2e^{-1}$]

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2. The joint distribution of X and Y is

$$f(x, y) = \begin{cases} K(x^2 + y^2), & 0 \leq x \leq y \leq 1 \\ 0 & \text{elsewhere} \end{cases}$$

$$\left[\text{Ans. } K = \frac{1}{3} \right]$$

3. Determine K and find marginal densities of X and Y .
 Given below:

$$f(x, y) = \begin{cases} \frac{8}{9} xy & 1 \leq x \leq y < 2 \\ 0 & \text{elsewhere} \end{cases}$$

- (a) Find marginal density function of X and Y :
 (b) Find conditional density function of Y given $X = x$ and conditional density function of X given $Y = y$.

[Ans.]

$$f_X(x) = \frac{4}{9} x(4 - x^2), \quad 1 \leq x \leq 2$$

$$f_Y(y) = \frac{4}{9} y(y^2 - 1), \quad 1 \leq y \leq 2$$

$$f_{Y/X}(y|x) = \frac{2y}{4 - x^2}, \quad x \leq y \leq 2$$

$$f_{X/Y}(x|y) = \frac{2y}{y^2 - 1}, \quad 1 \leq x \leq y.$$

4. If X and Y are two random variables having joint density function

$$f(x, y) = \begin{cases} \frac{1}{8}(6 - x - y), & 0 < x < 2, 2 < y < 2 \\ 0 & \text{otherwise} \end{cases}$$

Find: (a) $P(X < 1 \cap Y < 3)$, (b) $P(X + Y < 3)$, (c) $P(X < 1 | 4 < 3)$

$$[\text{Ans. (a) } \frac{3}{8}, \quad \text{(b) } \frac{5}{24}, \quad \text{(c) } \frac{3}{5}]$$

5. Joint distribution of X and Y is given by

$$f(x, y) = 4xye^{-(x^2 + y^2)}, \quad x \geq 0, y \geq 0$$

Test whether X and Y are independent. Also find the conditional density of X given $Y = y$.

$$[\text{Ans. Yes, } f(X = x | Y = y) = 2xe^{-x^2}]$$

3.9. BAYE'S THEOREM

Statement. Let S be the sample space and let E_1, E_2, \dots, E_n be n mutually exclusive events associated with a random experiment. If A is any event which occurs with E_1 or E_2 or ... or E_n , then

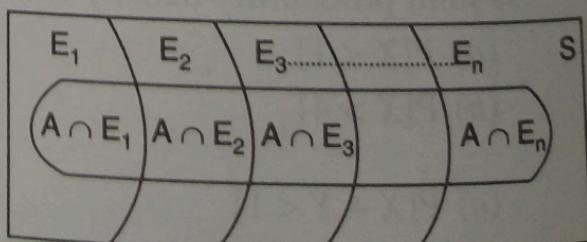


Fig. 3.1