浙江工业大学 2010 - 2011 学年第二学期 概率论与数理统计试卷

姓名:	学号:	班级:	任课教师:	
/ -		<i>></i> /	に パラスノル・	

- 一. 填空题 (每空 2 分, 共 30 分)
 - 1. <u>0.8</u>
 - $2. \ \ \frac{32}{125}$, $\frac{48}{125}$
 - 3. a = 3
 - 4. $1, 1, \frac{8}{9}$

5. 1, -1,
$$\underbrace{\begin{cases} 2xe^{-x^2}, & x > 0\\ 0, & x \le 0 \end{cases}}$$

- 6. 0.6826
- 7. <u>9</u>, <u>1</u>
- $8. \quad (4.412, 5.588)$
- 9. $\frac{1}{\sqrt{80\pi}}e^{-\frac{(z-3)^2}{80}}$
- 二. 选择题(每题 2 分, 共 10 分) 1. A 2. B 3. B 4. A 5. C
- 三. 计算题 (第一题 8 分, 第四题 12 分, 其余每题 10 分, 共 60 分)
 - 1. **解:** 记 A_1, A_2, A_3 分别表示该产品为甲、乙、丙厂生产,B 表示该产品为次品。

$$P(B) = P(A_1)P(B|A_1) + P(A_2)P(B|A_2) + P(A_3)P(B|A_3)$$

= 0.4 \times 0.01 + 0.4 \times 0.02 + 0.2 \times 0.04 = 0.02

2)
$$P(A_3|B) = \frac{P(A_3)P(B|A_3)}{P(B)} = \frac{0.008}{0.02} = 0.4$$

2. **解**:记鞋子配成的双数为X,则

$$P(X = 0) = \frac{C_5^4 2^4}{C_{10}^4} = \frac{8}{21}$$

$$P(X = 1) = \frac{C_5^1 C_4^2 2^2}{C_{10}^4} = \frac{4}{7}$$

$$P(X = 2) = \frac{C_5^2}{C_{10}^4} = \frac{1}{21}$$

$$EX = \frac{4}{7} \times 1 + \frac{1}{21} \times 2 = \frac{2}{3}$$

3. 解:

1)
$$P(X>1)=e^{-\lambda 1}=e^{-\frac{1}{2}}\Rightarrow \lambda=\frac{1}{2}$$
,从而
$$P(X>2)=e^{-\lambda 2}=e^{-1}$$

2)
$$EX^2 = (EX)^2 + Var(X) = 4 + 4 = 8$$

3) 对 y > 0,

$$F_Y(y) = P(Y \le y) = P(X^2 \le y)$$

= $1 - P(X > \sqrt{y}) = 1 - e^{-\frac{1}{2}\sqrt{y}}$

从而密度函数为
$$f_Y(y) = F_Y'(y) = \begin{cases} \frac{1}{4\sqrt{y}} e^{-\frac{1}{2}\sqrt{y}}, & y > 0\\ 0, & y \le 0 \end{cases}$$

4. 解:

1)
$$1 = \int_0^2 \int_0^2 a(x+y) dx dy = a \int_0^2 (2+2y) dy = 8a$$
, 从而 $a = \frac{1}{8}$ 。 2)

$$P(X+Y \le 2) = \int_0^2 \int_0^{2-x} a(x+y)dydx$$
$$= a[\int_0^2 x(2-x) + \frac{1}{2}(2-x)^2 dx]$$
$$= a[4 - \frac{8}{3} + \frac{1}{2} \times \frac{8}{3}] = \frac{1}{3}$$

3)

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

$$= \begin{cases} \frac{1}{4} + \frac{1}{4}x, & 0 < x < 2\\ 0, & 其它 \end{cases}$$

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx$$

$$= \begin{cases} \frac{1}{4} + \frac{1}{4}y, & 0 < y < 2\\ 0, & 其它 \end{cases}$$

4)由对称性

$$EX = EY = \int_0^2 (\frac{1}{4} + \frac{1}{4}x)x dx = \frac{1}{2} + \frac{1}{4}\frac{8}{3} = \frac{7}{6}$$

$$E(XY) = \int_0^2 \int_0^2 \frac{1}{8}(x+y)xy dx dy$$

$$= \frac{1}{8} \int_0^2 (2x^2 + \frac{8}{3}x) dx = \frac{1}{8}(2 \times \frac{8}{3} + \frac{8}{3} \times 2) = \frac{4}{3}$$

$$Cov(X,Y) = E(XY) - EXEY = \frac{4}{3} - (\frac{7}{6})^2 = -\frac{1}{36}$$

$$EX^2 = EY^2 = \int_0^2 \frac{1}{4}(x+1)x^2 dx = \frac{1}{4}(4 + \frac{8}{3}) = \frac{5}{3}$$

$$Var(X) = Var(Y) = \frac{5}{3} - (\frac{7}{6})^2 = \frac{11}{36}$$

$$\rho(X,Y) = \frac{Cov(X,Y)}{\sqrt{Var(X)Var(Y)}} = -\frac{1}{11}$$

5. 解:

1) 矩估计:

$$EX = \int_0^a \frac{2x}{a^2} x dx = \frac{2}{3}a$$

从而
$$a=\frac{3}{2}EX$$
, $\hat{a}=\frac{3}{2}\overline{X}$

2) 极大似然估计: 似然函数为

$$L(a) = \begin{cases} \Pi_{i=1}^{n} \frac{2x_{i}}{a^{2}}, & 0 < x_{i} < a, i = 1, 2, \dots, n \\ 0, & \not\exists \dot{\Xi} \end{cases}$$
$$= \begin{cases} \Pi_{i=1}^{n} \frac{2x_{i}}{a^{2}}, & a \leq \max(x_{1}, x_{2}, \dots, x_{n}) \\ 0, & \not\exists \dot{\Xi} \end{cases}$$

该函数关于 a 是单调递减函数,其极大似然估计为 $\hat{a} = \max(x_1, x_2, \cdots, x_n)$ 。

6. 解:

- 1) $H_0: \mu = \mu_0 = 500, \ H_1: \mu \neq \mu_0$
- 2) $t = \frac{\overline{X} \mu_0}{s/\sqrt{n}} = \frac{488 500}{6.29/9} = 17.17$
- 3) 拒绝域为 $(-\infty, -2.306) \cup (2.306, \infty)$
- 4) 在拒绝域中,拒绝原假设,不能认为该包装机包装食盐的平均 重量是正常的。

浙江工业大学 2011 - 2012 学年第二学期 概率论与数理统计试卷

姓名: _	学号:	班级:	任课教师:	
一. 填空	题(每空 2 分,共 28 分)			
10).4			
21	$\frac{1}{2}$			
3. $\frac{1}{2}$	1 <u>9</u> 27			
41	2			
5. $\frac{2}{3}$	2			
66	$-2\sqrt{2}$			
7	$\geq \frac{3}{4}$			
80	.6826			
9	$\sqrt{2}$, 2 , 2			
10	(39.51, 40.49)			
二. 选择	题(每题3分,共12分)			
1. B				
2. B				
3. D				
4. B				

三. 解答题 (共60分)

1. (8分)

解:

- 1) 3本语文书放在一起的概率为 $p_1 = \frac{3!8!}{10!} = \frac{1}{15}$;
- 2) 4 本数学书放在一起的概率为 $p_3 = \frac{4!7!}{10!} = \frac{1}{30}$; 4 本数学书放在一起且 3 本物理书放在一起的概率为 $p_4 = \frac{4!3!5!}{10!} = \frac{1}{210}$, 则所求概率为 $p_2 = p_3 p_4 = \frac{1}{25}$ 。

2. (10分)

解: 用 A_1, A_2 分别表示被保险人为甲类和乙类,用 B 表示被保险人发生事故。

1) $P(B) = P(A_1)P(B|A_1) + P(A_2)P(B|A_2) = 0.3 \times 0.4 + 0.7 \times 0.2 = 0.26$:

2)
$$P(A_1|B) = \frac{P(A_1B)}{P(B)} = \frac{0.3 \times 0.4}{0.26} = \frac{6}{13} = 0.46$$

3. (10分)

解:

1)

$$F(x) = \begin{cases} 0, & x < 0\\ \int_0^x se^{-\frac{s^2}{2}} ds = 1 - e^{\frac{-s^2}{2}}, & x > 0 \end{cases}$$

2) $y=x^2$ 在 $(0,\infty)$ 上单调, $x=\sqrt{y}$,

$$f_Y(y) = f_X(\sqrt{y}) \frac{1}{2\sqrt{y}} = \begin{cases} \frac{1}{2}e^{-\frac{y}{2}}, & y > 0\\ 0, & y < 0 \end{cases}$$

4. (12分)

解:

$$\int_0^1 dy \int_0^y Ax dx$$
$$=A \int_0^1 \frac{1}{2} y^2 dy = \frac{A}{6}$$

$$3$$
) 对 $0 < x < 1$,

$$f_X(x) = \int_x^1 Ax dy = 6x(1-x)$$

对于 0 < y < 1,

$$f_Y(y) = \int_0^y Ax dx = 3y^2$$

3)

$$EX = \int_0^1 6x(1-x)xdx = 2 - \frac{3}{2} = \frac{1}{2}$$

$$EX^2 = \int_0^1 6x(1-x)x^2dx = \frac{3}{2} - \frac{6}{5} = \frac{3}{10}$$

$$EY = \int_0^1 3y^2ydy = \frac{3}{4}$$

$$EY^2 = \int_0^1 3y^2y^2dy = \frac{3}{5}$$

$$EXY = \int_0^1 \int_0^y 6xxydxdy = \int_0^1 2y^4dy = \frac{2}{5}$$

$$cov(X,Y) = \frac{2}{5} - \frac{1}{2} \times \frac{3}{4} = \frac{1}{40}$$

$$DX = \frac{3}{10} - \frac{1}{2^2} = \frac{1}{20}$$

$$DY = \frac{3}{5} - (\frac{3}{4})^2 = \frac{3}{80}$$

$$\rho(X,Y) = \frac{1}{\sqrt{3}}$$

5. (10分)

解:

$$\begin{split} EX &= \int_0^{+\infty} \lambda^2 x e^{-\lambda x} x dx = \frac{2}{\lambda} \\ \lambda &= \frac{2}{EX} \\ \hat{\lambda} &= \frac{2}{\overline{X}} \end{split}$$

2) 极大似然函数为:

$$L(\lambda) = \prod_{i=1}^{n} \lambda^{2} x_{i} e^{-\lambda x_{i}}$$
$$\frac{\partial \ln L}{\partial \lambda} = \sum_{i=1}^{n} \frac{2}{\lambda} - x_{i}$$
$$\hat{\lambda} = \frac{2}{\overline{X}}$$

6. $(10 \ \%)$ **M**: $H_0: \mu = \mu_0 = 72; \ H_1: \mu \neq \mu_0$

$$t = \frac{\overline{X} - \mu_0}{S/\sqrt{n}} \approx -2.453$$

拒绝域为 $(-\infty, -2.2622) \cup (2.2622, \infty)$,在拒绝域中,故拒绝原假设,铅中毒患者与正常人的脉搏有显著差异。

- \equiv DACCD
- 三、 1. 解: A_1 表示乘火车; A_2 表示乘轮船; A_3 表示乘汽车; A_4 表示乘飞机; B表示迟到;

则

$$P(A_1) = \frac{3}{10}, P(A_2) = \frac{1}{5}, P(A_3) = \frac{1}{10}, P(A_4) = \frac{2}{5}$$
$$P(B|A_1) = \frac{1}{4}, P(B|A_2) = \frac{1}{3}, P(B|A_3) = \frac{1}{2}, P(B|A_4) = 0$$

由全概率公式,

$$P(B) = \sum_{i=1}^{4} P(A_i) P(B|A_i) = \frac{23}{120}$$

由贝叶斯公式,

$$P(A_1|B) = \frac{P(A_1)P(B|A_1)}{P(B)} = \frac{9}{23}$$

2. 1)
$$P(x \ge a) = 1 - P(x < a) = P(x < a)$$

所以
$$P(x < a) = \frac{1}{2}$$

$$\mathbb{P}\left(x < a\right) = \int_{0}^{a} 2x dx = \frac{1}{2}$$

得
$$a = \frac{\sqrt{2}}{2}$$

如果
$$y < 0, F(y) = 0$$

如果
$$y > 1, F(y) = 1$$

当
$$0 \le y \le 1$$
, $F(y) = P(0 \le X \le \sqrt{y}) = \int_0^{\sqrt{y}} 2x dx = y$ 所以

 $f(y) = \begin{cases} 1, & 0 < y < 1 \\ 0, & else \end{cases}$

3. 1)
$$ext{in} \int_{0}^{+\infty} \int_{0}^{+\infty} Axe^{-(x+y)} dxdy = 1 \quad \text{if} \quad A = 1$$

2)
$$f_X(x) = \int_0^{+\infty} x e^{-(x+y)} dy = x e^{-x}$$

 $f_Y(y) = \int_0^{+\infty} x e^{-(x+y)} dx = e^{-y}$
 $f_Y(x) = f_X(x) = f_Y(x) = f_Y(y)$

所以相互独立

3)
$$P(X+Y<1) = \int_{0}^{1} dx \int_{0}^{1-x} xe^{-(x+y)} dy = 1 - \frac{5}{2}e^{-1}$$

4. 由于
$$X \sim P(2)$$
, 所以 $E(X) = 2$, $Var(X) = 2$

由于
$$Y \sim U(0,6)$$
, 所以 $E(Y) = 3$, $Var(Y) = 3$

则
$$E(Z) = 3E(X) - 2E(Y) = 0$$

$$Cov(X,Y) = \rho_{XY} \sqrt{Var(X)Var(Y)} = 1$$

$$Var(Z) = 9Var(X) + 4Var(Y) - 12Cov(X,Y)$$
$$= 18 + 12 - 12$$
$$= 18$$

5. 1)
$$\boxplus E(X) = \int_{1}^{+\infty} x\theta x^{-\theta-1} dx = \frac{\theta}{\theta-1} = \overline{X}$$

得
$$\hat{\theta} = \frac{\overline{X}}{\overline{X}-1}$$

2) 似然函数为

$$L(\theta) = \prod_{i=1}^{n} \theta x_i^{-\theta-1} = \theta^n \left(\prod_{i=1}^{n} x_i \right)^{\theta-1}$$

对数似然函数为

$$\ln L(\theta) = n \ln \theta - (\theta + 1) \sum_{i=1}^{n} \ln x_i$$

求导,并令导数等于0

$$\frac{\partial}{\partial \theta} \ln L(\theta) = \frac{n}{\theta} - \sum_{i=1}^{n} \ln x_i = 0$$

得
$$\hat{\theta} = n / \sum_{i=1}^{n} \ln x_i$$

6.
$$H_0: \mu = 72; H_1: \mu \neq 72$$

当 H_0 成立时, 统计量

$$T = \frac{\overline{X} - \mu_0}{s / \sqrt{n}} \sim t (n - 1)$$

计算,
$$\overline{X} = 66.44, s = 7.18$$
, 则

$$t = \frac{66.44 - 72}{7.18 / \sqrt{16}} = -3.098$$

由于 $|t| = 3.098 > t_{15} (0.025) = 2.1315$, 拒绝原假设。

浙江工业大学 2012 - 2013 学年第二学期 概率论与数理统计试卷

姓名:	学号:	班级:	任课教师:	
一. 填空题	(每空2分,共22分	分)		
1 0.6	_°			
20.12	. 0.37 .			
36	, <u>0.72</u> °			
40.1_	_°			
54	$, \underline{4}, \underline{>\frac{5}{9}} .$			
6. (9.10	0,15.56) 。			
7 0.68	<u>26</u> .			
二. 选择题	(每题3分,共18分	分)		
1. D				
2. B				
3. C				
4. B				
5. A				

6. C

三. 计算题 (共 60 分)

1. (12分)解:

1),2)

1),2)				
Y	0	1	2	
0	$\frac{1}{15}$	$\frac{4}{15}$	$\frac{1}{15}$	$\frac{6}{15}$
1	0	$\frac{4}{15}$	$\frac{4}{15}$	$\frac{8}{15}$
2	0	0	$\frac{1}{15}$	$\frac{1}{15}$
	$\frac{1}{15}$	$\frac{8}{15}$	$\frac{6}{15}$	

3)
$$E(X-1)(Y-1) = \frac{1}{15} + \frac{1}{15} + \frac{1}{15} \times (-1) = \frac{1}{15}$$
.

2. (16分)解:

$$1$$
) $\,A+B=1, -A+B=0 \Rightarrow A=\frac{1}{2}, B=\frac{1}{2};$

2)
$$f(x) = \begin{cases} \frac{1}{2}\cos x, & -\frac{\pi}{2} \le x \le \frac{\pi}{2} \\ 0, & 其它 \end{cases}$$

3)
$$P(-\frac{\pi}{6} < X < \frac{\pi}{6}) = \frac{3}{4} - \frac{1}{4} = \frac{1}{2};$$

4)
$$E\cos(X) = \int_{-\pi/2}^{\pi/2} \frac{1}{2}\cos^2 x dx = \frac{1}{4} \int_{-\pi/2}^{\pi/2} (\cos x + 1) dx = \frac{1}{4} (2 + \pi)$$

3. (12分)解:

1)

$$1 = \int_0^1 \int_0^x C(y+1)dydx$$
$$= C \int_0^1 \frac{1}{2}x^2 + xdx$$
$$= \frac{2}{3}C \Rightarrow C = \frac{3}{2}$$

2)

$$P(X < 2Y) = \int_0^1 \int_{x/2}^x \frac{3}{2} (y+1) dy dx$$
$$= \frac{3}{4} \int_0^1 \frac{3}{4} x^2 + x dx$$
$$= \frac{3}{4} (\frac{1}{4} + \frac{1}{2}) = \frac{9}{16}$$

3)

$$f_X(x) = \int_0^x \frac{3}{2}(y+1)dy = \frac{3}{4}x^2 + \frac{3}{2}x$$
$$f_Y(y) = \int_y^1 \frac{3}{2}(y+1)dx = \frac{3}{2}(1-y^2)$$

 $f(x,y) \neq f_X(x)f_Y(y)$, 故不独立。

- 4. (10分)解:
 - 1) 矩估计:

$$EX = E(X - 1) + 1 = \lambda + 1 \Rightarrow \lambda = EX - 1$$

矩估计为 $\hat{\lambda} = \overline{X} - 1$ 。

2) 极大似然估计:

$$L(\lambda) = \prod_{i=1}^{n} e^{-\lambda} \frac{\lambda^{x_i - 1}}{(x_i - 1)!}$$
$$\frac{\partial \ln L}{\partial \lambda} = \sum_{i=1}^{n} \left[-1 + \frac{x_i - 1}{\lambda} \right]$$

极大似然估计为 $\hat{\lambda} = \overline{X} - 1$ 。

- 5. (10分)解:
 - 1) $H_0: \mu = \mu_0 = 18$, $H_1: \mu \neq \mu_0$
 - 2) $u = \frac{\overline{X} \mu_0}{\sigma_0 / \sqrt{n}} = 0.6$
 - 3) 拒绝域为 (-∞, -1.96) ∪ (1.96, +∞)
 - 4) 不在拒绝域中,故可以认为该灌装机工作正常。

浙工大 2013/2014 第一学期概率统计试卷(A)参考答案

- 一、填空题(每空3分, 共30分)
- 1, 8/15; 2, $(1-p)^3$; 3, B(n, p); 4, 4; 5, 5/7;
- 6. $\lambda_1 + \lambda_2$; 7. 72; 8. np^2 ; 9. $\frac{a}{2} + \frac{3b}{4} = 1$; 10. F(1,1);
- 二、单选题(每小题 2 分,共 16 分)
- 1, A; 2, C; 3, B; 4, D; 5, C; 6, B; 7, B; 8, C;
- 三.解答下列各题(每小题6分,共18分)
- 1. 设事件 A, B 满足 $P(B|A) = P(\overline{B}|\overline{A}) = \frac{1}{3}$, $P(A) = \frac{1}{3}$, 求 P(B) 。

解:
$$P(B|\overline{A}) = 1 - \frac{1}{3} = \frac{2}{3}, P(\overline{A}) = \frac{2}{3},$$

 $P(B) = P(\overline{A})P(B|\overline{A}) + P(A)P(B|A) = \frac{2}{3} \cdot \frac{2}{3} + \frac{1}{3} \cdot \frac{1}{3} = \frac{5}{9}$

2. 解:因为 $Y=e^X$,故Y不取负值,从而,当y<0时,则 $f_y(y)=0$;当y>0时,

$$F_Y(y) = P\{Y \le y\} = P\{0 < Y \le y\} = P\{0 < e^X \le y\}$$

= $P\{-\infty < X \le \ln y\} = \Phi(\ln y)$.

从而,
$$y > 0$$
时, $f_Y(y) = F_Y'(y) = \Phi'(x)|_{x=\ln y} \cdot \frac{1}{y} = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(\ln y)^2} \cdot \frac{1}{y}.$

所以
$$f_{Y}(y) = \begin{cases} \frac{1}{\sqrt{2\pi y}} e^{-(\ln y)^{2}/2}, & y > 0, \\ 0, & 其它 \end{cases}$$

3.
$$\Re$$
: $\pm 1 = a + \frac{1}{8} + \frac{1}{4} + \frac{1}{8} + b + \frac{1}{4} = a + b = \frac{1}{4}$

曲
$$P(X = x_1, Y = y_3) = P(X = x_1)P(Y = y_3)$$
 得

$$\frac{1}{4} = (a + \frac{3}{8})\frac{1}{2}$$
, $\mathbb{E}[a = \frac{1}{8}]$, $\mathbb{E}[b = \frac{1}{8}]$

Y的分布律为

Y	y_1	y_2	y_3
P	1/4	1/4	1/2

四.解答下列各题(每小题 10 分,共 30 分)

1.#\textbf{F}:
$$E(X) = \frac{1}{8} \int_0^2 dx \int_0^2 x(x+y) dy = \frac{7}{6}, E(Y) = \frac{7}{6},$$

$$E(X^2) = \frac{1}{8} \int_0^2 dx \int_0^2 x^2 (x+y) dy = \frac{5}{3}, E(Y^2) = \frac{5}{3},$$

$$Var(X) = \frac{5}{3} - \frac{49}{36} = \frac{11}{36} = Var(Y).$$

$$E(XY) = \frac{1}{8} \int_0^2 dx \int_0^2 xy(x+y) dy = \frac{4}{3},$$

$$Cov(X,Y) = E(XY) - E(X)E(Y) = \frac{4}{3} - \frac{9}{36} = -\frac{1}{36},$$

$$\rho_{XY} = \frac{Cov(X,Y)}{\sqrt{Var(X)Var(Y)}} = -\frac{1}{36} \cdot \frac{36}{11} = -\frac{1}{11} \quad \circ$$

2. 解: (1) σ^2 的置信度 $1-\alpha$ 的置信区间为 $\left[\frac{(n-1)S^2}{\chi_{\frac{\alpha}{2}}^2(n-1)}, \frac{(n-1)S^2}{\chi_{\frac{1-\alpha}{2}}^2(n-1)}\right]$,

计算得 σ^2 的置信度 95%的置信区间为[707.22, 3104.4].

(2) $H_0: \mu \le 1800, H_1: \mu > 1800,$

选取统计量
$$t = \frac{\overline{X} - \mu_0}{S / \sqrt{n}} \sim t(n-1)$$

 $\alpha = 0.05$ 时,拒绝域为 $t > t_{0.05}(15) = 1.753$,

计算得

$$t = \frac{\overline{x - \mu_0}}{s / \sqrt{n}} = \frac{1832 - 1800}{36 / 4} \approx 3.556 > t_{0.05}(15) = 1.753,$$

所以拒绝 H_0 ,可以认为灯泡的平均寿命显著的大于 1800。

3. 解:似然函数为 $L(\alpha) = (\alpha + 1)^n (x_1 x_2 \cdots x_n)^{\alpha}$

$$\ln L(\alpha) = \ln[(\alpha + 1)^n (x_1 \ x_2 \cdots x_n)^{\alpha}] = n \ln (1 + \alpha) + \alpha \sum_{i=1}^n \ln x_i$$

所以参数
$$\alpha$$
 的最大似然估计量为 $\hat{\alpha} = -\frac{n}{\sum\limits_{i=1}^{n} \ln X_i} - 1.$

五、本大题两个小题任选一题作答(6分)

1. 设
$$X_1, X_2, X_3, X_4$$
为来自总体 $N(1, \sigma^2)$ 的样本,求统计量 $\frac{X_1 - X_2}{|X_3 + X_4 - 2|}$ 的分布。

解 因 X_1, X_2, X_3, X_4 是来自总体 $N(1, \sigma^2)$ 的样本,所以 X_1, X_2, X_3, X_4 相互独立且均 服 从 正 态 分 布 $N(1, \sigma^2)$ 。 由 正 态 分 布 的 叠 加 原 理 知 : $X_1 - X_2 \sim N(0, 2\sigma^2)$;

$$\frac{X_1 - X_2}{\sqrt{2}\sigma} \sim N(0,1), X_3 + X_4 - 2 \sim N(0,2\sigma^2); \frac{X_3 + X_4 - 2}{\sqrt{2}\sigma} \sim N(0,1),$$

故
$$(\frac{X_3 + X_4 - 2}{\sqrt{2}\sigma})^2 \sim \chi^2(1)$$
,且 $\frac{X_1 - X_2}{\sqrt{2}\sigma}$ 与 $(\frac{X_3 + X_4 - 2}{\sqrt{2}\sigma})^2$ 相互独立,

故由 t 分布的定义知:

$$\frac{X_1 - X_2}{\sqrt{2}\sigma} / \sqrt{\left(\frac{X_3 + X_4 - 2}{\sqrt{2}\sigma}\right)^2} = \frac{X_1 - X_2}{|X_3 + X_4 - 2|} \sim t(1).$$

2 解: 设 X_i ($i = 1, 2, \dots, 1200$) 为第 i 笔销售收入实际少收的金额,

则 $X_1, X_2, \cdots, X_{1200}$ 相互独立,服从同一分布U(0,0.1),

$$\coprod E(X_i) = 0.05, Var(X_i) = 1/1200$$
,

$$E(\sum_{i=1}^{1200} X_i) = \sum_{i=1}^{1200} E(X_i) = 60, Var(\sum_{i=1}^{1200} X_i) = \sum_{i=1}^{1200} Var(X_i) = 1,$$

由中心极限定理 $\sum_{i=1}^{1200} X_i \sim N(60,1)$

$$P(\sum_{i=1}^{1200} X_i \le 62) = P(\frac{\sum_{i=1}^{1200} X_i - 60}{1} \le \frac{62 - 60}{1}) \approx \Phi(2) = 0.9772$$

浙江工业大学 2013 - 2014 学年第二学期 概率论与数理统计参考答案

一. 填空题 (每空 3 分, 共 30 分)

- 1. $\frac{1}{3}$
- 2. $\frac{3}{7}$
- 3. $\frac{1}{4}$
- 4. __1__
- 5. <u>9</u>, <u>13</u>
- 6. $\frac{1}{2}$
- 7. <u>0.9544</u>
- 8. (2,1) , <u>1</u>

二. 选择题 (每题 2 分, 共 10 分)

- 1. C
- 2. D
- 3. A
- 4. B
- 5. B

三. 解答题 (共60分)

1. 解:

1)
$$P(X = \frac{\sqrt{3}}{4}) = \frac{6}{C_6^3} = 0.3;$$

3)
$$EX = \frac{\sqrt{3}}{4}[1 \times 0.3 + 2 \times 0.6 + 3 \times 0.1] = \frac{9\sqrt{3}}{20};$$

4)
$$EX^2 = \frac{3}{16}[1 \times 0.3 + 4 \times 0.6 + 9 \times 0.1] = \frac{27}{40};$$

$$Var(X) = EX^2 - (EX)^2 = \frac{27}{400}$$
°

2. 解:

1)
$$1 = \int_0^1 Cx(1-x)dx = C(\frac{1}{2} - \frac{1}{3}) \Rightarrow C = 6;$$

2)
$$x < 0, F(x) = 0$$
; $x > 1, F(x) = 1$; $0 \le x \le 1$, $F(x) = \int_0^x 6(s - s^2) ds = 3x^2 - 2x^3$;

3)
$$0 < y < 1$$
, $F_Y(y) = P(Y \le y) = P((2X - 1)^2 \le y) = P(\frac{1 - \sqrt{y}}{2} \le X \le \frac{1 + \sqrt{y}}{2}) = \frac{3}{2}\sqrt{y} - \frac{1}{2}y\sqrt{y}$,从而

$$f_Y(y) = \begin{cases} \frac{3}{4} \frac{1-y}{\sqrt{y}}, & 0 < y < 1 \\ 0, & \text{其它} \end{cases}$$

3. 解:

1)
$$1 = \int_0^1 \int_0^1 Cx(1-y)dxdy = \frac{C}{4} \Rightarrow C = 4;$$

2)

$$f_X(x) = \int_0^1 4x(1-y)dy = 2x, \quad 0 < x < 1$$

$$f_Y(y) = \int_0^1 4x(1-y)dx = 2(1-y), \quad 0 < y < 1$$

$$f(x,y) = f_X(x)f_Y(y)$$

因此,X,Y独立。

3)

$$P(X < Y) = \int_0^1 \int_0^y 4x(1-y)dxdy$$
$$= \int_0^1 2y^2(1-y)dy = 2(\frac{1}{3} - \frac{1}{4}) = \frac{1}{6}$$

4. 解:

矩估计:
$$EX = 0 \times (1-\theta) + 2 \times \theta - \theta^2) + 3 \times \theta^2 = \theta^2 + 2\theta$$
, $\theta = \sqrt{1+EX}-1$, $\hat{\theta} = \sqrt{1+\overline{X}}-1 = \sqrt{\frac{12}{5}}-1$;

极大似然估计: $L(\theta)=(1-\theta)^2(\theta-\theta^2)^2\theta^2=\theta^4(1-\theta)^4$,极大似然估计 $\tilde{\theta}$ 为最大值点 $\frac{1}{9}$ 。

5. 解:

 $H_0: \mu = \mu_0 = 20$, $H_1: \mu \neq \mu_0$

 $t=rac{\overline{x}-\mu_0}{S/\sqrt{n}}=2$;

拒绝域为 $(-\infty, -2.1315) \cup (2.1315, \infty)$;

t 的值不在拒绝域中,认为该机器生产的螺丝长度正常。

浙江工业大学 2014 - 2015 学年第一学期 概率论与数理统计试卷

姓名:	字号:	 	

- 一. **填空题** (每空2分, 共22分)
 - 1. __0.4__
 - 2. $\frac{5}{9}$
 - 3. __2__
 - 4. 3, $\frac{1}{2}$
 - 5. $\frac{1}{3}$
 - 6. <u>0.8186</u>
 - 7. <u>2</u>, <u>\sqrt{2}</u>, <u>2</u>
 - 8. $\overline{X} \frac{S}{\sqrt{n}} t_{\frac{\alpha}{2}}(n-1)$
- 二. 选择题 (每题 3 分, 共 18 分)
 - 1. B
 - 2. D
 - 3. B
 - 4. A
 - 5. C
 - 6. B

三. 解答题 (共60分)

1. (8分)解: A表示带菌, B表示阳性, 则

$$P(A|B) = \frac{P(A)P(B|A)}{P(A)P(B|A) + P(\bar{A})P(B|\bar{A})}$$

$$= \frac{0.1 \times 0.95}{0.1 \times 0.95 + 0.9 \times 0.01}$$

$$= \frac{95}{104}$$

$$\approx 0.9135$$

2. (8分)解:

$$EY^2 = 6^2 \times 0.4 + 0 = 14.4.$$

3. (12分,每小题4分)解:

$$1 = \int_{-\infty}^{\infty} f(x)dx$$
$$= \int_{0}^{1} cxdx + \int_{1}^{2} c(3-x)dx$$
$$= \frac{c}{2} + \frac{3c}{2} = 2c$$
$$\Rightarrow c = \frac{1}{2}$$

$$EX = \int_{-\infty}^{\infty} x f(x) dx$$

$$= \int_{0}^{1} cx^{2} dx + \int_{1}^{2} c(3 - x) x dx$$

$$= \frac{c}{3} + \frac{9}{2}c - \frac{7}{3}c = \frac{5}{4}$$

$$EX^{2} = \int_{0}^{1} cx^{3} dx + \int_{1}^{2} c(3 - x) x^{2} dx$$

$$= \frac{c}{4} + 7c - \frac{15}{4}c = \frac{7}{4}$$

$$Var(X) = EX^{2} - (EX)^{2} = \frac{3}{16}$$

3)
$$\forall 0 < y < 4, \ h(y) = \sqrt{y},$$

$$f_Y(y) = f_X(h(y))|h'(y)|$$

$$= \begin{cases} \frac{1}{4}, & 0 < y < 1\\ \frac{3}{4\sqrt{y}} - \frac{1}{4}, & 1 < y < 4 \end{cases}$$

4. (12分)每小题4分解:

1)

$$1 = \int_0^1 \int_0^y A(2x+y)dxdy$$
$$= \int_0^1 2Ay^2dy = \frac{2A}{3}$$
$$\Rightarrow A = \frac{3}{2}$$

2)

$$P(X+Y<1) = \int_0^{\frac{1}{2}} \int_x^{1-x} A(2x+y) dy dx$$
$$= \int_0^{\frac{1}{2}} 2Ax(1-2x) + \frac{A}{2} [(1-x)^2 - x^2] dx$$
$$= \int_0^{\frac{1}{2}} \frac{1}{2} A + Ax - 4Ax^2 dx = \frac{5}{24} A = \frac{5}{16}$$

$$\begin{split} f_X(x) &= \int_x^1 A(2x+y) dy = 3x(1-x) + \frac{3}{4}(1-x^2) = \frac{3}{4} + 3x - \frac{15}{4}x^2, \quad 0 < x < 1 \\ f_Y(y) &= \int_0^y A(2x+y) dx = 3y^2, \quad 0 < y < 1 \\ f(x,y) &\neq f_X(x) f_Y(y), 所以 X, Y 不独立。 \end{split}$$

- 5. (10分)解:
 - 1) 矩估计:

$$EX = \int_{2}^{\infty} \lambda x e^{-\lambda(x-2)} dx = \frac{1}{\lambda} + 2$$
$$\Rightarrow \lambda = \frac{1}{EX - 2}$$

故矩估计 $\hat{\lambda} = \frac{1}{\overline{X}-2}$ 。

2) 极大似然估计:

$$L(\lambda) = \prod_{i=1}^{n} \lambda e^{-\lambda(x_i - 2)}$$
$$\frac{\partial \ln L}{\partial \lambda} = \sum_{i=1}^{n} \frac{1}{\lambda} - (x_i - 2) = 0$$

得极大似然估计 $\hat{\lambda} = \frac{1}{\overline{X}-2}$ 。

6. $(10 \ \%)$ **M**: $H_0: \mu = (\leq)\mu_0 = 900$, $H_1: \mu > \mu_0$

$$t = \frac{\overline{X} - \mu_0}{S/\sqrt{n}} \approx 5.06$$

拒绝域为 (1.8331, ∞),

在拒绝域中,拒绝原假设,该肥料显著地提高了农作物的产量。

浙江工业大学 2014 - 2015 学年第二学期 概率论与数理统计参考答案

- 一. **填空题** (每空2分,共28分)
 - 1. _0.2_
 - $2. \frac{8}{15}$
 - 3. $\frac{-\frac{2}{3}}{3}$, $\frac{4}{3}$
 - 4. $\frac{1}{2}$, $-\frac{1}{2}$, $\frac{1}{2}\sin x$, $0 < x < \pi$ 0, 其它
 - 5. 103, $\frac{56}{5}$
 - 6. 3, $\frac{2}{3}$
 - 7. $\frac{5}{9}$
 - 8. $\bar{X} + \frac{S}{\sqrt{n}} t_{\alpha/2}(n-1)$
 - 9. $2\Phi(1) 1$

二. 选择题 (每题 3 分, 共 12 分)

- 1. B
- 2. B
- 3. D
- 4. C

三. 解答题 (共60分)

1. 解:

X	-1	0	1	
1	$\frac{1}{4}$	$\frac{1}{3}$	$\frac{1}{6}$	$\frac{3}{4}$
2	$\frac{1}{12}$	$\frac{1}{9}$	$\frac{1}{18}$	$\frac{1}{4}$
	$\frac{1}{3}$	$\frac{4}{9}$	$\frac{2}{9}$	

3)
$$P(X+Y>1) = \frac{1}{6} + \frac{1}{9} + \frac{1}{18} = \frac{1}{3}$$

2. 解:

1)
$$1 = \int_0^1 c(1-x^2)dx = c[1-\frac{1}{3}] \Rightarrow c = \frac{3}{2}$$
;

2)
$$EX = \int_0^1 xc(1-x^2)dx = c[\frac{1}{2} - \frac{1}{4}] = \frac{3}{8};$$

$$EX^2 = \int_0^1 x^2 c(1 - x^2) dx = c\left[\frac{1}{3} - \frac{1}{5}\right] = \frac{1}{5};$$

$$Var(X) = EX^2 - (EX)^2 = \frac{19}{320};$$

$$3)$$
 $X = \sqrt{Y}$,从而

$$f_Y(y) = \begin{cases} \frac{3}{4} \left[\frac{1}{\sqrt{y}} - \sqrt{y} \right], & 0 < y < 1 \\ 0, & \sharp : \Xi \end{cases}$$

3. 解:

1)
$$1 = \int_0^1 \int_0^1 c(1+y) dx dy = \frac{3}{2}c \Rightarrow c = \frac{2}{3};$$

2)
$$P(X < Y) = \int_0^1 \int_0^y c(1+y) dx dy = c \int_0^1 y(1+y) dy = \frac{5}{6}c = \frac{5}{9};$$

3)
$$EX = \int_0^1 \int_0^1 xc(1+y)dxdy = \frac{1}{2}$$
; $EY = \int_0^1 \int_0^1 yc(1+y)dy = \frac{5}{9}$;

$$EXY = \int_0^1 \int_0^1 xyc(1+y)dxdy = \frac{1}{2}c\int_0^1 y(1+y)dy = \frac{5}{18}$$

从而
$$Cov(X,Y) = EXY - EX EY = 0$$
,即 $\rho = 0$ 。

4. 解:

矩估计: $EX=\int_0^1 x\alpha x^{\alpha-1}dx=\frac{\alpha}{\alpha+1}$,从而 $\alpha=\frac{EX}{1-EX}$,即矩估计 为 $\tilde{\alpha}=\frac{\bar{X}}{1-\bar{X}}$;

极大似然估计: $L(\alpha) = \prod_{i=1}^{n} \alpha x_i^{\alpha-1}$,

$$\frac{\partial \ln L}{\partial \alpha} = \sum_{i=1}^{n} \left[\frac{1}{\alpha} + \ln x_i \right] = 0$$

可得极大似然估计为 $\hat{\alpha} = -\frac{n}{\sum_{i=1}^{n} \ln x_i}$ 。

5. **M**: $H_0: \mu = \mu_0 = 1000$, $H_1: \mu \neq \mu_0$;

 $t = \frac{\bar{x} = \mu_0}{s/\sqrt{n}} = -1.8$;

拒绝域为 $(-\infty, -2.1315) \cup (2.1315, \infty)$;

不在拒绝域中,可以认为这批鱼的平均重量为1000克。

浙江工业大学 2015-2016 学年第一学期概率统计试卷参考答案

一填空题

- 1. 0.2
- 2. 0.25
- 3. 0.1
- 4. 0.5
- 5. 10
- 6.

U	0	1
P	0. 16	0.84
V	0	1
P	0.64	0. 36

- 7. 9 3
- 8. [10.02,11.98]
- 二 选择题
- 1. C 2.D 3.B 4.D 5.C
- 三 解答题
- 1. 以 A 表示事件"产品出厂";以 B 表示事件"产品需进一步调试"由全概率公式,得

$$\alpha = P(A|B)P(B) + P(A|\overline{B})P(\overline{B}) = 0.8 \times 0.3 + 1 \times 0.7 = 0.94$$

根据二项分布,得

$$\beta = C_{100}^2 (1 - 0.94)^2 0.94^{98} = 0.04144168$$

2. 以题意可得 x 的分布律为

X	-1	1	2	
P	0.3	0.5	0.2	

得 $Y = X^2 + 1$ 的分布律

Y	2	2	5	
P	0.3	0.5	0.2	

合并相同的项有

Y	2	5
P	0.8	0.2

$$E(Y) = 2 \times 0.8 + 5 \times 0.2 = 2.6$$

$$E(Y^2) = 2^2 \times 0.8 + 5^2 \times 0.2 = 8.2$$

$$Var(Y) = E(Y^2) - (E(Y))^2 = 8.2 - 2.6^2 = 1.44$$

3. 1) X的边缘密度函数为

$$f_X(x) = \int f(x,y) dy$$

如果 0<x<1,有

$$f_X(x) = \int f(x,y)dy = \int_0^1 (2-x-y)dy = \frac{3}{2}-x$$

否则

$$f_{X}(x)=0$$

于是

$$f_X(x) = \begin{cases} \frac{3}{2} - x & 0 < x < 1 \\ 0 & else \end{cases}$$

类似地

$$f_{Y}(y) = \begin{cases} \frac{3}{2} - y & 0 < y < 1 \\ 0 & else \end{cases}$$

显然

$$f(x,y) \neq f_X(x) f_Y(y)$$

所以, X和Y不独立。

2)
$$P(X > 2Y) = \int_{0}^{1} dx \int_{0}^{\frac{1}{2}x} (2 - x - y) dy = \frac{7}{24}$$

4 由
$$EX = \overline{X}$$
,而

$$EX = \int_{0}^{\theta} \frac{x(6\theta x - 6x^{2})}{\theta^{3}} dx = \frac{\theta}{2}$$

得
$$\hat{\theta} = 2\overline{X}$$

5 似然函数为

$$L(\theta) = P(X=3)P(X=1)P(X=3)P(X=0)P(X=3)P(X=1)P(X=2)P(X=3)$$

= $4\theta^6 (1-\theta)^2 (1-2\theta)^4$

取对数得

$$\ln L(\theta) = 6 \ln \theta + 2 \ln (1 - \theta) + 4 \ln (1 - 2\theta) + \ln 4$$

求导,

$$\frac{dL(\theta)}{d\theta} = \frac{6}{\theta} - \frac{2}{1-\theta} - \frac{8}{1-2\theta} = \frac{2(12\theta^2 - 14\theta + 3)}{\theta(1-\theta)(1-2\theta)} = 0$$

注意到
$$0 < \theta < \frac{1}{2}$$
, 得 $\hat{\theta} = \frac{7 - \sqrt{13}}{12}$

6
$$H_0: \sigma_0^2 = 1.6^2 H_1: \sigma_0^2 \neq 1.6^2$$

当 H_0 成立时, 统计量

$$\chi^2 = \frac{(n-1)s^2}{\sigma_0^2} \sim \chi^2 (n-1)$$

计算,
$$\chi_0^2 = \frac{(9-1)\times1.1^2}{1.6^2} = 3.78 < \chi_{0.025}^2 (8) = 17.535$$
,

接受原假设。

7 记
$$X_i = \begin{cases} 1 & \text{第}i \land \text{的士发生事故} \\ 0 & \text{第}i \land \text{的士无事故} \end{cases}$$

出事故的的士数
$$X = \sum_{i=1}^{500} X_i \sim B(500, 0.004)$$

保险公司一年赚钱不小于 200000 元的事件为

$$\{500 \times 800 \ge 500 \times 800 - 50000 X \ge 200000 \}$$

即 事件{0≤X≤4}, 从而有

$$P(0 < X < 4) = P\left(0 < \sum_{i=1}^{500} X_i < 4\right)$$

$$= P\left(\frac{0 - 500 \times 0.004}{\sqrt{500 \times 0.004 \times (1 - 0.004)}} < \frac{\sum_{i=1}^{500} X_i - E\sum_{i=1}^{500} X_i}{\sqrt{Var \sum_{i=1}^{500} X_i}} < \frac{4 - 500 \times 0.004}{\sqrt{500 \times 0.004 \times (1 - 0.004)}}\right)$$

$$= \Phi\left(\frac{2}{\sqrt{1.992}}\right) - \Phi\left(\frac{-2}{\sqrt{1.992}}\right)$$

$$= 2\Phi\left(\frac{2}{\sqrt{1.992}}\right) - 1$$

$$= 2\Phi\left(1.42\right) - 1$$

$$= 2 \times 0.9222 - 1 = 0.8444$$

浙江工业大学 2015 - 2016 学年第二学期 概率论与数理统计试卷

一. 填空题, 每空3分。

- 1. _0.1_;
- 2. <u>0.5</u>;
- 3. <u>0.2</u>;
- 4. $\frac{1}{24}$;
- 5. _4_;
- 6. $\begin{cases} 1 3^{-x}, & x \ge 0 \\ 0, & \text{其它} \end{cases}$
- 7. 0.05;
- 8. $\frac{5}{9}$;
- 9. $\sqrt{\frac{3}{2}}$;
- 10. (9.46, 12.54) .

二. 选择题, 每题 3 分。

- 1. C
- 2. B
- 3. B
- 4. D

三. 解答题, 共58分。

1. (8 分) 解: A_1, A_2, A_3, A_4 分别表示"乘火车"、"乘轮船"、"乘汽车"、"乘飞机",B 表示迟到,则

$$P(A_1|B) = \frac{0.3 \times \frac{1}{4}}{0.3 \times \frac{1}{4} + 0.2 \times \frac{1}{3} + 0.1 \times \frac{1}{12}} = 0.5.$$

即乘火车的概率为0.5。

2. (10分)解:一只零件的寿命大于1500小时的概率为

$$\int_{1500}^{\infty} \frac{1000}{x^2} \, dx = \frac{2}{3},$$

则 5 只零件中至少有两只寿命大于 1500 小时的概率为

$$1 - (\frac{1}{3})^5 - C_5^1(\frac{2}{3})(\frac{1}{3})^4 = 1 - \frac{11}{3^5} = \frac{232}{243} \approx 0.955.$$

3. (10分)解:

(a)

$$1 = \int_0^1 \int_0^2 c(6 - x - y) \, dy dx = c \int_0^1 12 - 2x - 2 \, dx = c[10 - 1] = 9c$$

从而 $c = \frac{1}{9}$;

(b)

$$\begin{split} P(X \geq \frac{1}{2}, Y > 1) &= \int_{\frac{1}{2}}^{1} \int_{1}^{2} c(6 - x - y) \; dy dx \\ &= c \int_{\frac{1}{2}}^{1} 6 - x - \frac{3}{2} dx \\ &= c [\frac{9}{4} - \frac{3}{8}] = \frac{5}{24}. \end{split}$$

4. (10分)解:

$$a+b+c=\frac{1}{2}$$

$$c=2a$$

$$ab=\frac{1}{64}$$

解得

$$\begin{cases} a = \frac{1}{8} \\ b = \frac{1}{8} \\ c = \frac{1}{4} \end{cases}, \begin{cases} a = \frac{1}{24} \\ b = = \frac{3}{8} \\ c = \frac{1}{12}. \end{cases}$$

- 5. (10分)解:
 - (a) 矩估计:

$$EX = \int_{1}^{2} \theta x \, dx + \int_{2}^{3} (1 - \theta) x \, dx = \frac{5}{2} - \theta$$
$$\theta = \frac{5}{2} - EX$$

矩估计 $\hat{\theta} = \frac{5}{2} - \bar{X}$;

(b) 极大似然估计:

$$L(\theta) = \theta^{N} (1 - \theta)^{n-N}$$
$$\frac{\partial \ln L}{\partial \theta} = \frac{N}{\theta} - \frac{n-N}{1-\theta} = 0$$

解得极大似然估计 $\hat{\theta} = \frac{N}{n}$ 。

6. (10分)解:

(a) $H_0: \sigma^2 = \sigma_0^2 = 5000, H_1: \sigma^2 \neq \sigma_0^2;$

(b) $\chi^2 = \frac{1}{\sigma_0^2}(n-1)s^2 = 46$;

(c) 拒绝域为 (0,11.52) ∪ (44.313,∞);

(d) 在拒绝域中, 拒绝原假设, 可以认为这批电池寿命的波动性有显著变化。

2016/2017学年第2学期概率试卷答案

一、填空.
$$1. \quad \frac{1}{3} \qquad 2. \quad (1-r)^{10} \qquad 3. \quad 0.045$$

$$4. \quad 2 \qquad 5. \quad 3, 0.6 \qquad 6. \quad \frac{1}{\sqrt{2\pi\times97}}e^{-\frac{x^2}{194}}, \ -\infty < x < \infty$$

$$7. \quad e^{-2}, B(5, e^{-2}) \qquad 8. \quad \frac{1}{2n-2} \qquad 9. \quad \frac{1}{4} \qquad 10. \quad \frac{3}{4}$$

7.
$$e^{-2}$$
, $B(5, e^{-2})$ 8. $\frac{1}{2n-2}$ 9. $\frac{1}{4}$ 10. $\frac{3}{4}$

二、选择题.

三、计算题.

1. 解: (1)记X表示开锁需要的次数,则

$$P(X \le 3) = P(X = 1) + P(X = 2) + P(X = 3) = \frac{1}{8} + \frac{7}{8} \cdot \frac{1}{7} + \frac{7}{8} \cdot \frac{6}{7} \cdot \frac{1}{6} = \frac{3}{8}$$
(2)

$$\begin{split} &P\{\max X, Y \geq 0\} = 1 - P\{\max X, Y < 0\} \\ &= P\{X \geq 0, Y \geq 0\} + P\{X \geq 0, Y \leq 0\} + P\{X \leq 0, Y \geq 0\} \\ &= \frac{3}{7} + P\{X \geq 0\} - P\{X \geq 0, Y \geq 0\} + P\{Y \geq 0\} - P\{X \geq 0, Y \geq 0\} \\ &= \frac{3}{7} + \frac{1}{7} + \frac{1}{7} = \frac{5}{7} \end{split}$$

2. \mathbf{M} :(1) \mathbf{X} 的分布函数是 $\mathbf{F}(\mathbf{x}) = \int_{-\infty}^{\mathbf{x}} f(s) ds$

当x > 0时,

$$F(x) = \int_0^x \frac{2}{\pi(t^2 + 1)} dt = \frac{2}{\pi} \arctan x$$

当 $x < 0$ 时, $F(x) = 0$.

$$F(x) = \begin{cases} \frac{2}{\pi} \arctan x, & \exists x > 0 \text{时}, \\ 0, & \text{其它}. \end{cases}$$

(2)因为

$$F(y)=P(Y\leq y)=P(\ln X\leq y)=P(X\leq e^y)=\int_{-\infty}^{e^y}f(x)dx$$
关于y求导,得概率密度函数

$$f(y) = \frac{2e^y}{\pi(e^{2y}+1)}, -\infty < y < \infty.$$

3. 解: (1)X的边缘概率密度函数是

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy = \begin{cases} \int_x^1 8xy dy = 4x(1 - x^2), & \text{\pm dot} 0 \le x \le 1 \text{\pm b}, \\ 0, & \text{\pm kc}, \end{cases}$$

Y的边缘概率密度函数是

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dy = \begin{cases} \int_0^y 8xy dy = 4y^3, & \text{\pm j} 0 \le y \le 1 \text{ \pm j}, \\ 0, & \text{\pm j} \text{ \pm j}, \end{cases}$$

(2)因为 $f(x,y) \neq f_X(x)f_Y(y)$,故变量X和Y不是相互独立.

(3)

$$P(X+Y \ge 1) = \iint_{X+Y>1} f(x,y) \, dx \, dy = \int_{1/2}^{1} dy \int_{1-y}^{y} 8xy \, dx = \frac{5}{6}.$$

4. 解: (1)由 $EX = \int_1^{+\infty} \frac{x\beta}{x^{\beta+1}} dx = \frac{\beta}{\beta-1}$,得 $\beta = \frac{EX}{EX-1},$

所以 β 的矩估计量是 $\hat{\beta} = \frac{\bar{X}}{\bar{X}-1}$.

(2)令极大似然函数是

$$L = \prod_{i=1}^{n} \frac{\beta}{x_i^{\beta+1}} = \frac{\beta^n}{\prod_{i=1}^{n} x_i^{\beta+1}}$$

取对数,得

$$\ln L = n \ln(\beta) - (\beta + 1) \sum_{i=1}^{n} \ln(x_i),$$

关于 β 求导,得

$$\frac{d \ln L}{d\beta} = \frac{n}{\beta} - \sum_{i=1}^{n} \ln(x_i),$$

令上式等于0,解此方程,得其极大值点是

$$\beta = \frac{n}{\sum_{i=1}^{n} \ln(x_i)},$$

 β 的极大似然估计是 $\hat{\beta} = \frac{n}{\sum_{i=1}^{n} \ln(X_i)}$.

5.解: (1)因为
$$T = \frac{\bar{X} - \mu}{S/\sqrt{n}} \sim t(n-1),$$

所以μ的置信度为95%的置信区间是

$$(\bar{X} - \frac{S}{\sqrt{n}}t_{0.025}(n-1), \bar{X} + \frac{S}{\sqrt{n}}t_{0.025}(n-1)).$$

将 $\bar{x} = 6.97, s = 0.194, n = 9, t_{0.025}(8) = 2.31$ 代入上式,得 μ 的置信度为95%的置信区间是(6.8209, 7.1191).

(2)因为 $7 \in (6.8209, 7.1191)$,所以生产铁水含碳量正常 $(\alpha = 0.05)$.

四、解:由题意,得

$$Var(X) = Var(Y) = \sigma^2,$$

则

$$Var(U) = Var(aX + bY) = (a^{2} + b^{2})\sigma^{2},$$

$$Var(V) = Var(aX - bY) = (a^{2} + b^{2})\sigma^{2},$$

$$Var(U + V) = Var(2aX) = 4a^{2}\sigma^{2},$$

$$Cov(U, V) = \frac{Var(U + V) - Var(U) - Var(V)}{2} = (a^{2} - b^{2})\sigma^{2}$$

$$\rho_{UV} = \frac{Cov(U, V)}{\sqrt{Var(U)}\sqrt{Var(V)}} = \frac{a^{2} + b^{2}}{a^{2} - b^{2}}.$$