San Francisco precipitation

 $Julien\ JACQUES \\ 2/19/2020$

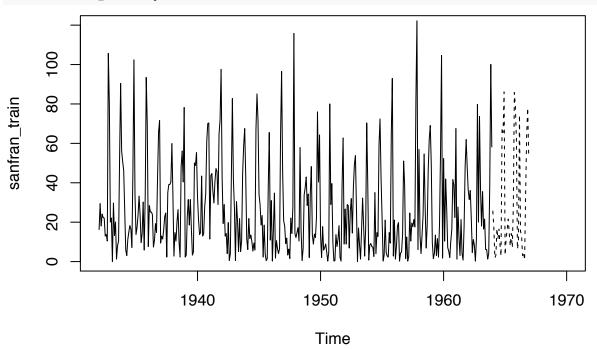
We extract training and test set

```
data=scan(file="http://eric.univ-lyon2.fr/~jjacques/Download/DataSet/sanfran.dat",skip=1)
sanfran<-ts(data,start=c(1932,1),end=c(1966,12),freq=12)
library(forecast)</pre>
```

```
## Warning: package 'forecast' was built under R version 3.5.2
sanfran_train=window(sanfran,start=c(1932,1),end=c(1963,12))
sanfran_test=window(sanfran,start=c(1964,1),end=c(1966,12))
```

We can plot both

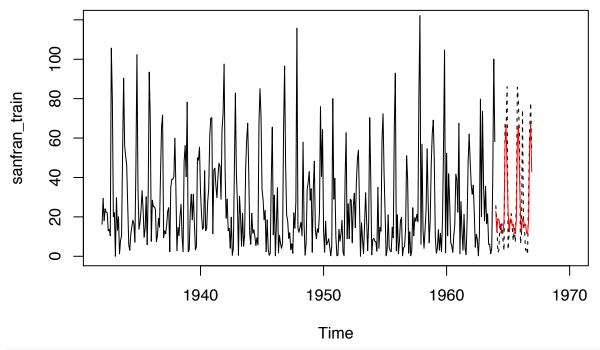
```
plot(sanfran_train,xlim=c(1932,1970),ylim=c(0,120))
lines(sanfran_test,lty=2)
```



Forecasting with exponential smoothing

We see a seasonal pattern, probably additive.

```
library(forecast)
h=hw(sanfran_train,seasonal='additive',damped=FALSE,h=36)
plot(sanfran_train,xlim=c(1932,1970),ylim=c(0,120))
lines(sanfran_test,lty=2)
lines(h$mean,col=2)
```

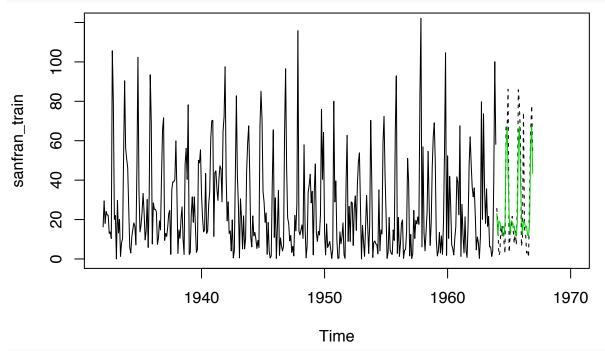


```
print(sqrt(mean((h$mean-sanfran_test)^2)))
```

[1] 15.86614

We can compare with a damped version, the result are slightly better

```
hd=hw(sanfran_train,seasonal='additive',damped=TRUE,h=36)
plot(sanfran_train,xlim=c(1932,1970),ylim=c(0,120))
lines(sanfran_test,lty=2)
lines(hd$mean,col=3)
```

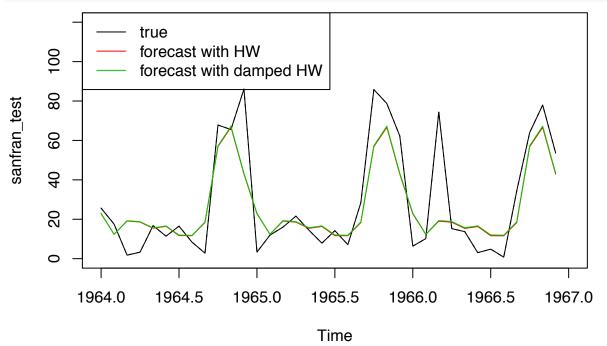


print(sqrt(mean((hd\$mean-sanfran_test)^2)))

[1] 15.77082

We can zoom on the prediction

```
plot(sanfran_test,xlim=c(1964,1967),ylim=c(0,120))
lines(sanfran_test,lty=2)
lines(h$mean,col=2)
lines(hd$mean,col=3)
legend('topleft',col=1:3,lty=1,legend=c('true','forecast with HW','forecast with damped HW'))
```



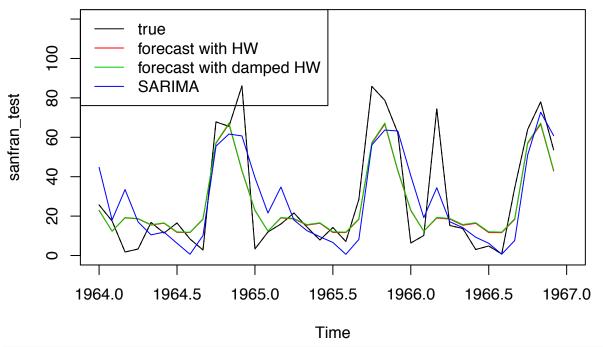
The difference is almost null between HW and its damped version. Indeed, if we have a look to the ϕ parameter, it is very close to 1 ($\phi = 0.9725$): the damping effect is almost null.

Forecasting with SARIMA

Simple solution: with auto.arima function:

```
fit=auto.arima(sanfran_train)
summary(fit)
## Series: sanfran_train
```

```
## Series: sanfran_train
   ARIMA(0,0,1)(2,1,0)[12] with drift
##
##
   Coefficients:
                                        drift
##
                               sar2
             ma1
                      sar1
##
         -0.0108
                   -0.6204
                            -0.2710
                                      -0.0061
          0.0510
                    0.0508
                             0.0521
                                       0.0415
##
   s.e.
##
## sigma^2 estimated as 327.7:
                                 log likelihood=-1605.73
## AIC=3221.46
                 AICc=3221.62
                                 BIC=3241.05
##
## Training set error measures:
##
                          ME
                                 RMSE
                                            MAE MPE MAPE
                                                                MASE
```



```
print(sqrt(mean((prev$mean-sanfran_test)^2)))
```

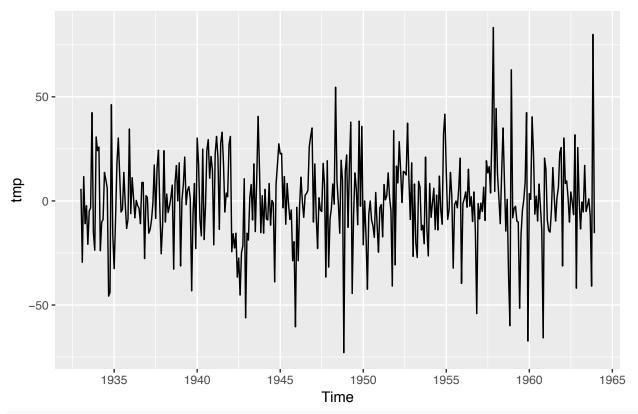
[1] 16.57343

The forecast is not better than with HW.

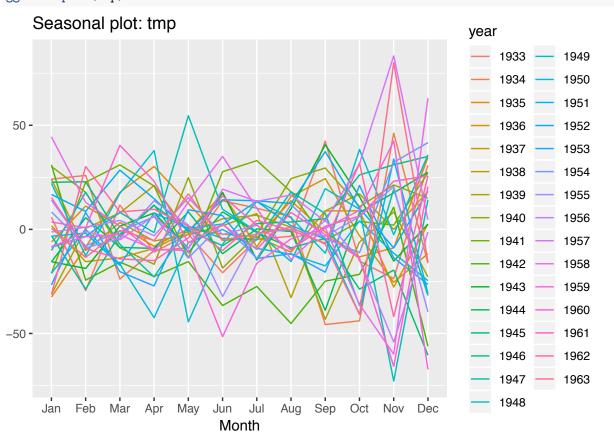
We can try to choose manually the ordre of the SARIMA model.

Let's start by differeciating the serie.

```
tmp=diff(sanfran_train,lag=12)
autoplot(tmp)
```

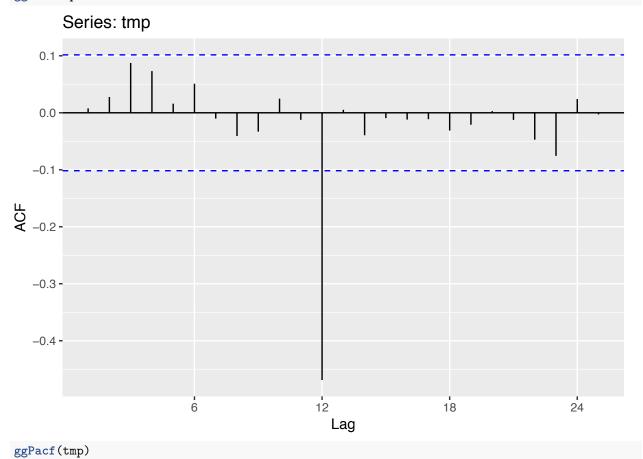


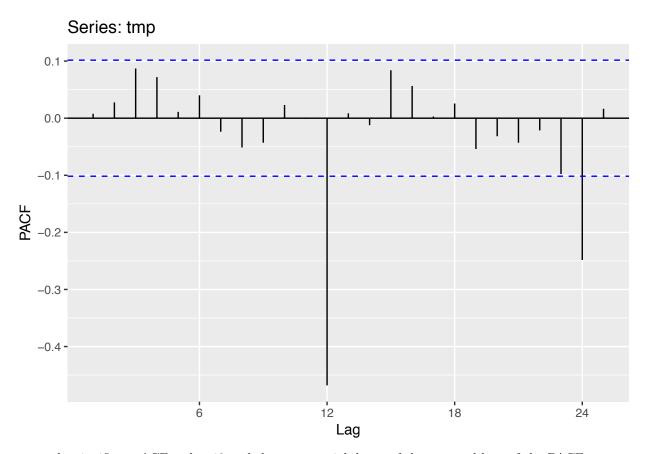
ggseasonplot(tmp)



It seems approximatively stationary. Let's look at the ACF and PACF $\,$

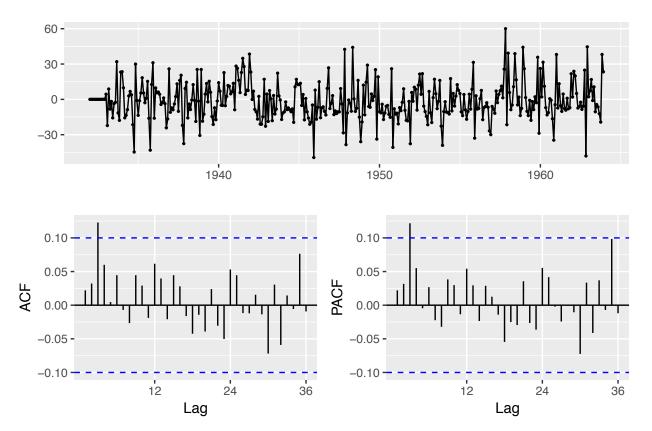
ggAcf(tmp)





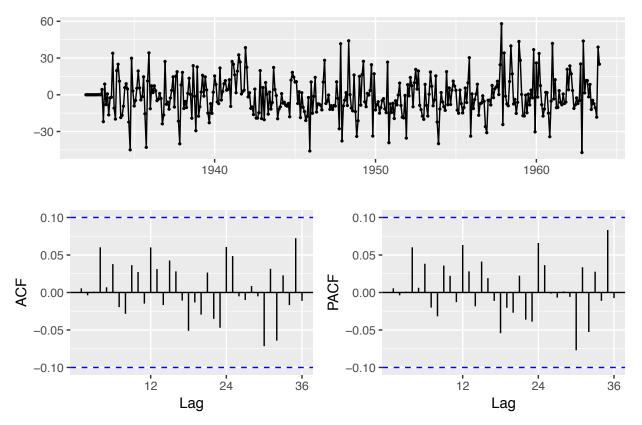
- the significant ACF at lag 12 and the exponential decay of the seasonal lags of the PACF suggest a seasonal MA_1

```
fit=Arima(sanfran_train, order=c(0,0,0), seasonal=c(0,1,1))
fit %>% residuals() %>% ggtsdisplay()
```



• There is still significant ACF and PACF at lag 3. We can add some additional non-seasonal terms, with an $SARIMA_{(0,0,3)(0,1,1)_{12}}$ (or $SARIMA_{(3,0,0)(0,1,1)_{12}}$)

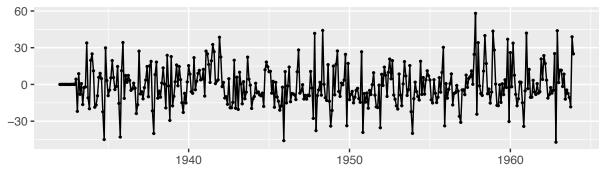
```
fit1=Arima(sanfran_train, order=c(0,0,3), seasonal=c(0,1,1))
fit1 %>% residuals() %>% ggtsdisplay()
```

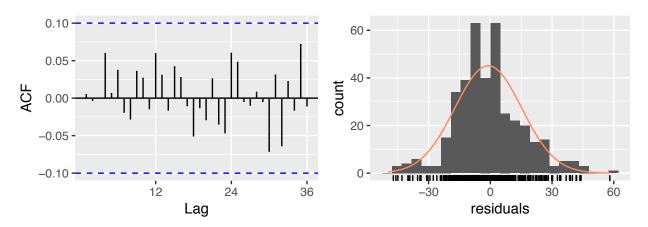


It seems that we have captured all auto-correlations

checkresiduals(fit1)



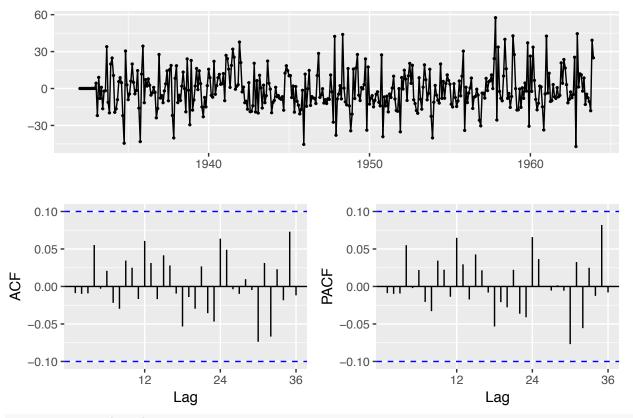




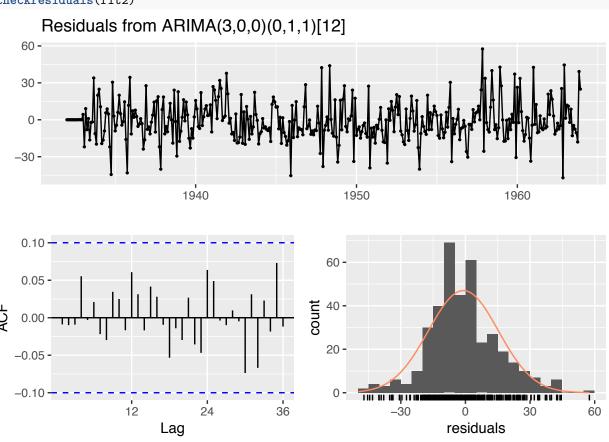
```
##
## Ljung-Box test
##
## data: Residuals from ARIMA(0,0,3)(0,1,1)[12]
## Q* = 11.139, df = 20, p-value = 0.9425
##
## Model df: 4. Total lags used: 24
```

We have the same result with an $SARIMA_{(3,0,0)(0,1,1)_{12}}$

```
fit2=Arima(sanfran_train, order=c(3,0,0), seasonal=c(0,1,1))
fit2 %>% residuals() %>% ggtsdisplay()
```







```
##
    Ljung-Box test
##
##
## data: Residuals from ARIMA(3,0,0)(0,1,1)[12]
## Q* = 10.834, df = 20, p-value = 0.9504
##
## Model df: 4.
                  Total lags used: 24
Both model are acceptable. We can compare the AICc:
cat('AICc for SARIMA_{(0,0,3)(0,1,1)_{12}} : ',fit1$aicc,'\n')
## AICc for SARIMA_{(0,0,3)(0,1,1)_{12}} : 3171.135
cat('AICc for SARIMA_{(3,0,0)(0,1,1)_{12}} : ',fit2$aicc,'\n')
## AICc for SARIMA_{(3,0,0)(0,1,1)_{12}} : 3170.473
The second one is better, we select it for forecasting.
We can forecast the next 30 values and compare the results
prev=forecast(fit2,h=36)
plot(sanfran_test,xlim=c(1964,1967),ylim=c(0,120))
lines(sanfran_test,lty=2)
lines(h$mean,col=2)
lines(hd$mean,col=3)
lines(prev$mean,col=4)
legend('topleft',col=1:4,lty=1,legend=c('true','forecast with HW','forecast with damped HW','SARIMA'))
                   true
                   forecast with HW
     100
                   forecast with damped HW
                   SARIMA
     80
sanfran_test
     9
     4
     0
          1964.0
                      1964.5
                                 1965.0
                                             1965.5
                                                        1966.0
                                                                    1966.5
                                                                                1967.0
                                              Time
```

[1] 17.49685

The results are not better with this model than with this one selected by auto.arima.

print(sqrt(mean((prev\$mean-sanfran_test)^2)))