In letters $E[IK] = I_{K} = \begin{cases} 1 \leftarrow Y_{n} \\ \infty & 1 - \frac{1}{n} \end{cases}$ $I_{K} = \begin{cases} 0 & \text{otherwise} \end{cases}$ Assignment-3 7) There were In envelops and n letters $R(I_k=1)=\frac{1}{n}$ X = H of reductly matched envelops, $X = f(I_K)$ k from 1 to note $X = f_n(I_K)$ Kfrom I to n. $E[X] = \sum_{k=1}^n E[I_k] = \sum_{k=1}^n Y_n = 1$ $=\sum_{k}T_{k}$ TO Var(x) = E(x2)-E(x)) -> momenté formulae. = E[(25x)2] - 12 (M) 112 - 5 - [. .] 1 1 1 $= E \left(\sum_{k=1}^{n} I_{k}^{2} + 2 \left(n_{c_{k}} \right) \left(\sum_{i \neq j} \sum_{i \neq j} I_{k}^{T_{i}} \right) \right) - 1$ ZE (Jr) +2 (C2) ZIE (JKJ) -1 = 271+2[n(1) 81 = m-1+2n(n-1)=1 (1,c 15) 0 (E +j) B(IK=1 \$Ij=1) = PV(Ij=1), BA(IK=1) = 1 - 1 $E\left[1^{k-1}j\right] = \sum_{k=1}^{\infty} \sum_{j=1}^{\infty} \left[1^{k-1}j\right] \cdot \frac{1}{n(n-1)}$ As no plancours) = Pr(Ix-1) ->0 $P_{\delta}(X=\delta) = \frac{e^{-\lambda}\lambda^{\delta}}{Y!}$

18: Ducess is a student recieving a grade that has not been recived before. ith \$ (iti) success. Let X; Lee the mond papers between Cm. $X = f(x;'s) = \sum x_i + l$ X; = Yit1 -Y; City Chiterra Xi is the wardom variable. $Y_{K} = \sum_{i=0}^{K-1} x_{i} \Rightarrow E[Y_{K}] = E[X_{i}] = E[X_{i}] + \underbrace{-PE(X_{i})}_{i=0} + \underbrace$ Comment = G (= + = + + + + + + + -) = 14,7 1. E[X|Y=:]= Zn Pxy (2/y) = 1x = + 2x = +3x = +1x = +2x0 +3x = x 1x0+2x = +3x = q GITT-C = 37/18. E[X/X=1] 1=1/5/3 = 37/18. 2. Equium probability of element being 1=> probability of element being 0=1-p Marie @ Expected value of the own stort with 1. & ends with 2 6 E(x1)= 1-(17)= 10. It starts with 0 of ends with 1)=) If we combine /E(x)=p. /++(1-12)(11-p)=== $E(x) = \frac{1}{1-p}$ run ends with o=1-p 1 runs ends with 1 = P. E(XV) = P(+p) + (-p) + = P+ (-p) = 2p-2p+1

$$= e^{-0.01} + \lambda e^{-0.01} + \lambda^{\frac{2}{2}} e^{-0.01} + \lambda^{\frac{3}{2}} e^{-0.01}$$

brobability of sucess in each trail is
$$P = \frac{1}{2}m$$
.

$$= \frac{P(x=kt) + x>k}{P(x>k)} = \frac{(1-p)^{k+j-1}P}{(1-p)^{k}}$$

$$= \frac{(1-p)^{j-1}}{P(x>k)}$$

E[X|X>K] = E(X) = 1/p -> : Geometric dictsi button.

E[X|X7k] = K+1=2m+K. When correct password founded

 $X = Y\cos(2\pi\theta) = Y\cos(\pi\theta)$ $0=0 \Rightarrow \chi=Y \Rightarrow Y=0$ (0)

$$Y = Y Sin \left(\frac{2\pi a}{\delta}\right) = Y sin \left(\frac{\pi a}{4}\right)$$

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X = 1/52 0=1_ -1/2 0=3

0=4

2=0. 0=6 6-27

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marginal 11 PMF of X

P(x=x) = 1/8, p(x=x/12) = 1/4, p(x=0) = 1/4, p(x=x/2) = 1/4

on arginal PMF&Y

P(Y=Y)=1/8, P(Y=YISL)= /4, P(Y=0)= /4, P(Y=-152)= /4

p(x=0)= Y4

P(Y= Y/1) = P(x/V)+P(Y=0)+P(Y=-WL)+P(Y=-1)

= Yyt Yy + Yy + 1/8 = 7/8.

p(x = x/12, x = x/12) = b(x=x/1222x= x/12)=1/8.

p(x(x/v2) = p(x=-1) = 1/8.

(2) P(x=0)= 0-5 P(Y=1 | Y=0)= 0.1, P(Y=0 | X=1)=0.2

$$P(Y=1|X=0)=1-P(BY=1|X=0)=1-0.9$$

$$P(X=1 & Y=0) = P(X=1) \cdot P(Y=0 | X=1) = 0$$

$$P(x=0 \notin Y=1) = P(x=0) \cdot P(Y=1|x=0) = 0.5 \times 0.1 = 0.05$$

$$P(Y=1|X=1) = 1 - P(Y=0|X=1) = 1 - 0.2 = 0.8$$

$$P(x=1, Y=1) = P(x=1) \cdot P(Y=1|x=1) = 05x08=09$$

(a) joint pMFIS
$$p(x,y) = \begin{cases} 0.45 & \alpha = 0, y = 0 \\ 0.10 & \gamma = 4, y = 0 \end{cases}$$

$$x = 0.45 & \gamma = 0.45 & \gamma$$

$$P_{X}(x) = \begin{cases} 0.5 & \lambda = 1 \\ 0.5 & \lambda = 0 \end{cases} P_{Y}(Y) = \begin{cases} 0.55 & y = 0 \\ 0.45 & y = 1 \end{cases}.$$

$$X = e^{-\lambda_1} \lambda_1^{K}$$

$$Y = e^{-\lambda_1} \lambda_1^{K}$$

$$X = e^{-\lambda_1} \lambda_1^{K}$$

$$P(z) = 2^{\frac{1}{2}} P(x-k) P(x-z-k) = \frac{2}{k-0} \frac{e^{\lambda_1} \lambda_1^{k}}{k!} \frac{e^{\lambda_1} k_2^{k-k}}{(z-k)!}$$

$$\int_{K=0}^{\infty} (Z) = e^{\lambda_1 + \lambda_2} \frac{\lambda_1}{(Z-K)!} \frac{\lambda_1}{(Z-K)!}$$

$$P_{X}(x) = \begin{cases} V_{3}^{2} X = 1 \\ (V_{2})^{2}(2x) & x = 4 \end{cases}$$

$$= (X_{3})^{2} (2x) + (X_{3})^{2}(2x) + (X$$

D Marginal PMF

$$P_{m}(m) = \sum_{n=0}^{\infty} P_{m,n}(m,n) = \sum_{n=0}^{\infty} \frac{1}{(k+1)^{2}} = \frac{(k+1)^{2}}{(k+1)^{2}} = \frac{1}{(k+1)^{2}}$$

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$$= e^{-\lambda} \frac{p' \lambda'}{i!} \left[\sum_{k=1}^{\infty} \frac{[k-\hat{D}]}{[k-\hat{D}]} \times \frac{\lambda^{k-1}}{[k-\hat{D}]} \left[(1-p)^{k-1} \right] \right]$$

$$= e^{-\lambda} (\lambda P)^{K} e^{-\lambda} e^{(1-P)\lambda} = (\lambda P)^{K} e^{-P\lambda}$$

$$= \frac{1}{K!} e^{-\lambda} (\lambda P)^{K} e^{-\lambda} e^{-\lambda}$$

$$Z = \eta - y = \chi - (\eta - y) = 2y - h = \frac{\eta + \lambda}{2}$$

$$P(Z=Z_0) = P(X=\frac{Z_0 th}{Z}) = \begin{pmatrix} y \\ z_0 th \end{pmatrix} = \begin{pmatrix} y \\ z_0 th \end{pmatrix} = \begin{pmatrix} y \\ z_0 th \end{pmatrix}$$

$$E(Z) = E[2x-n] = 2 E[x]-n = 2np-n$$

= $(2p-1) n$.

$$P(x=k) = \frac{e^{-x}n^{k}}{k!}$$
 Here $k=10, k=11$.

$$P(x=11) = e^{-10} \times 10^{11}$$

Rate for each booth
$$\lambda = \frac{30}{N}$$
.

$$[\lambda=6]$$

$$p(x_{4.5}) = e^{-5} \left(5^{\circ} + \frac{5!}{1!} + \frac{5^{2}}{2!} + \frac{5^{4}}{3!} + \frac{5^{4}}{5!} \right)$$

Constant of the last of the la

$$X_1 = 1,2$$
 $Y_1 = 1,2$

$$P(1,1) = 2K$$
, $P(1,2) = 4K$, $P(2,1) = 5K$, $P(2,1) = 6K$.

$$P(x=0P(Y=1) = P(x=1, Y=1) \Rightarrow \frac{3}{18} \times \frac{8}{18} \neq \frac{3}{18}$$

$$(2) P(x=i, Y=i) = (i)$$

$$P(x=x) = \sum_{y} P(x=y, Y=y) = \frac{e^{-2\lambda}}{2i}$$

$$= Z_{i} \left(\frac{1}{i} \right) e^{-2\lambda} \int_{1}^{1} Z_{i} e^{-2\lambda} Z_{i} \left(\frac{1}{i} \right) \frac{\lambda^{i}}{1!}$$

$$\frac{1}{2}\left(\frac{1}{2}\right) = \frac{1}{2}\left(\frac{1}{2}\right) = \frac{1}{2}\left(\frac{1}{2}\right)$$

$$E[x^{n}] = \int x^{n} f(x) dx = \int x^{n} dx = \frac{1}{n+1}.$$

$$Va_{1}(x^{n}) = E[x^{n}] - (E(x^{n}))^{2}$$

$$= \int x^{n} dx - \frac{1}{(n+1)^{2}} = \frac{1}{2n+1} - \frac{1}{(n+1)^{2}}.$$

$$\frac{b(X=k)}{b(X=k+1)} = \frac{u^{(K+1)}}{u^{(K+1)}} = \frac{u^{(K+1)}}{u^{(K+1)}} = \frac{u^{(K+1)}}{u^{(K+1)}} = \frac{u^{(K+1)}}{u^{(K+1)}}$$

$$\frac{b(X=k+1)}{b(X=k+1)} = \frac{u^{(K+1)}}{u^{(K+1)}} = \frac{u^{(K+1)}}{u^{(K+1)}} = \frac{u^{(K+1)}}{u^{(K+1)}}$$

If
$$E \leq (n+1)P-1) \Rightarrow P(x=E) \uparrow$$
 $E \Rightarrow (n+1)P-1) \Rightarrow P(x=E) \downarrow$

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$$P(X=K) = (1-P)^{K-1}P$$

$$P(X=n+K| X>n) = \frac{P(X=n+K| X>n)}{P(X>n)}$$

$$= \frac{P(X=n+K| X>n)}{P(X>n)} = \frac{(1-P)^{(n+K-1)}P}{(1-P)^{(n+K-1)}P}$$

$$= \frac{(1-P)^{(n+K-1)}P}{P(X>n)} = \frac{(1-P)^{(n+K-1)}P}{(1-P)^{(n+K-1)}P}$$

$$= \frac{(1-P)^{(n+K-1)}P}{P(X>n)} = \frac{(1-P)^{(n+K-1)}P}{(1-P)^{(n+K-1)}P}$$
The process has already taken more than n=torress trails without The brocess has already taken more than n=torress trails without The brocess disconst change the probability distribution of remaining

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